## Statistics Toolbox ${ }^{\text {TM }}$

User's Guide

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## Statistics Toolbox ${ }^{\text {TM }}$ User's Guide

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## Getting Started

## Product Overview

Statistics Toolbox ${ }^{\text {TM }}$ software extends MATLAB ${ }^{\circledR}$ to support a wide range of common statistical tasks. The toolbox contains two categories of tools:

- Building-block statistical functions for use in MATLAB programming
- Graphical user interfaces (GUIs) for interactive data analysis

Code for the building-block functions is open and extensible. Use the MATLAB Editor to review, copy, and edit code for any function. Extend the toolbox by copying code to new files or by writing files that call toolbox functions.

GUIs allow you to perform statistical visualization and analysis without writing code. You interact with the GUIs using sliders, input fields, buttons, etc. and the GUIs automatically call building-block functions.

## Organizing Data

- "Introduction to Data Types" on page 2-2
- "MATLAB Arrays" on page 2-4
- "Statistical Arrays" on page 2-11
- "Grouped Data" on page 2-34


## Introduction to Data Types

MATLAB data is placed into "data containers" in the form of workspace variables. All workspace variables organize data into some form of array. For statistical purposes, arrays are viewed as tables of values.

MATLAB variables use different structures to organize data:

- 2-D numerical arrays (matrices) organize observations and measured variables by rows and columns, respectively. (See "Other Data Structures" in the MATLAB documentation.)
- Multidimensional arrays organize multidimensional observations or experimental designs. (See "Multidimensional Arrays" in the MATLAB documentation.)
- Cell and structure arrays organize heterogeneous data of different types, sizes, units, etc. (See "Cell Arrays" and "Structures" in the MATLAB documentation.)

Data types determine the kind of data variables contain. (See "Classes (Data Types)" in the MATLAB documentation.)

These basic MATLAB container variables are reviewed, in a statistical context, in the section on "MATLAB Arrays" on page 2-4.

These variables are not specifically designed for statistical data, however. Statistical data generally involves observations of multiple variables, with measurements of heterogeneous type and size. Data may be numerical (of type single or double), categorical, or in the form of descriptive metadata. Fitting statistical data into basic MATLAB variables, and accessing it efficiently, can be cumbersome.

Statistics Toolbox software offers two additional types of container variables specifically designed for statistical data:

- "Categorical Arrays" on page 2-13 accommodate data in the form of discrete levels, together with its descriptive metadata.
- "Dataset Arrays" on page 2-23 encapsulate heterogeneous data and metadata, including categorical data, which is accessed and manipulated using familiar methods analogous to those for numerical matrices.

These statistical container variables are discussed in the section on "Statistical Arrays" on page 2-11.

## MATLAB Arrays

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"Numerical Data" on page 2-4
"Heterogeneous Data" on page 2-7
"Statistical Functions" on page 2-9

## Numerical Data

MATLAB two-dimensional numerical arrays (matrices) containing statistical data use rows to represent observations and columns to represent measured variables. For example,

```
load fisheriris % Fisher's iris data (1936)
```

loads the variables meas and species into the MATLAB workspace. The meas variable is a 150 -by- 4 numerical matrix, representing 150 observations of 4 different measured variables (by column: sepal length, sepal width, petal length, and petal width, respectively).


The observations in meas are of three different species of iris (setosa, versicolor, and virginica), which can be separated from one another using the 150-by-1 cell array of strings species:

```
setosa_indices = strcmp('setosa',species);
setosa = meas(setosa_indices,:);
```

The resulting setosa variable is 50 -by-4, representing 50 observations of the 4 measured variables for iris setosa.

To access and display the first five observations in the setosa data, use row, column parenthesis indexing:

| SetosaObs | $=$ setosa(1:5,:) |  |  |
| ---: | :--- | ---: | :--- |
| SetosaObs |  |  |  |
| 5.1000 | 3.5000 | 1.4000 | 0.2000 |
| 4.9000 | 3.0000 | 1.4000 | 0.2000 |
| 4.7000 | 3.2000 | 1.3000 | 0.2000 |
| 4.6000 | 3.1000 | 1.5000 | 0.2000 |
| 5.0000 | 3.6000 | 1.4000 | 0.2000 |

The data are organized into a table with implicit column headers "Sepal Length," "Sepal Width," "Petal Length," and "Petal Width." Implicit row headers are "Observation 1," "Observation 2," "Observation 3," etc.

Similarly, 50 observations for iris versicolor and iris virginica can be extracted from the meas container variable:

```
versicolor_indices = strcmp('versicolor',species);
versicolor = meas(versicolor_indices,:);
virginica_indices = strcmp('virginica',species);
virginica = meas(virginica_indices,:);
```

Because the data sets for the three species happen to be of the same size, they can be reorganized into a single 50-by-4-by- 3 multidimensional array:

```
iris = cat(3,setosa,versicolor,virginica);
```

The iris array is a three-layer table with the same implicit row and column headers as the setosa, versicolor, and virginica arrays. The implicit layer names, along the third dimension, are "Setosa," "Versicolor," and "Virginica." The utility of such a multidimensional organization depends on assigning meaningful properties of the data to each dimension.

To access and display data in a multidimensional array, use parenthesis indexing, as for 2-D arrays. The following gives the first five observations of sepal lengths in the setosa data:

```
SetosaSL = iris(1:5,1,1)
SetosaSL =
    5.1000
    4.9000
    4.7000
    4.6000
    5.0000
```

Multidimensional arrays provide a natural way to organize numerical data for which the observations, or experimental designs, have many dimensions. If, for example, data with the structure of iris are collected by multiple observers, in multiple locations, over multiple dates, the entirety of the data can be organized into a single higher dimensional array with dimensions for "Observer," "Location," and "Date." Likewise, an experimental design calling for $m$ observations of $n p$-dimensional variables could be stored in an $m$-by- $n$-by- $p$ array.

Numerical arrays have limitations when organizing more general statistical data. One limitation is the implicit nature of the metadata. Another is the requirement that multidimensional data be of commensurate size across all dimensions. If variables have different lengths, or the number of variables differs by layer, then multidimensional arrays must be artificially padded with NaNs to indicate "missing values." These limitations are addressed by dataset arrays (see "Dataset Arrays" on page 2-23), which are specifically designed for statistical data.

## Heterogeneous Data

MATLAB data types include two container variables-cell arrays and structure arrays-that allow you to combine metadata with variables of different types and sizes.

The data in the variables setosa, versicolor, and virginica created in "Numerical Data" on page 2-4 can be organized in a cell array, as follows:

```
iris1 = cell(51,5,3); % Container variable
obsnames = strcat({'Obs'},num2str((1:50)','%-d'));
iris1(2:end,1,:) = repmat(obsnames,[1 1 3]);
varnames = {'SepalLength','SepalWidth',...
    'PetalLength','PetalWidth'};
iris1(1,2:end,:) = repmat(varnames,[1 1 3]);
iris1(2:end,2:end,1) = num2cell(setosa);
iris1(2:end,2:end,2) = num2cell(versicolor);
iris1(2:end,2:end,3) = num2cell(virginica);
iris1{1,1,1} = 'Setosa';
iris1{1,1,2} = 'Versicolor';
iris1{1,1,3} = 'Virginica';
```

To access and display the cells, use parenthesis indexing. The following displays the first five observations in the setosa sepal data:

| SetosaSLSW $=$ | iris1 $(1: 6,1: 3,1)$ |  |  |
| ---: | ---: | ---: | ---: |
| SetosaSLSW $=$ |  |  |  |
| 'Setosa' | 'SepalLength' | 'SepalWidth ' |  |
| 'Obs1' | [ | $5.1000]$ | [ |
| 'Obs2' | $[$ | $4.5000]$ | $[$ |

Here, the row and column headers have been explicitly labeled with metadata.
To extract the data subset, use row, column curly brace indexing:

```
subset = reshape([iris1{2:6,2:3,1}],5,2)
subset =
\begin{tabular}{ll}
5.1000 & 3.5000 \\
4.9000 & 3.0000 \\
4.7000 & 3.2000 \\
4.6000 & 3.1000 \\
5.0000 & 3.6000
\end{tabular}
```

While cell arrays are useful for organizing heterogeneous data, they may be cumbersome when it comes to manipulating and analyzing the data. MATLAB and Statistics Toolbox statistical functions do not accept data in the form of cell arrays. For processing, data must be extracted from the cell array to a numerical container variable, as in the preceding example. The indexing can become complicated for large, heterogeneous data sets. This limitation of cell arrays is addressed by dataset arrays (see "Dataset Arrays" on page 2-23), which are designed to store general statistical data and provide easy access.

The data in the preceding example can also be organized in a structure array, as follows:

```
iris2.data = cat(3,setosa,versicolor,virginica);
iris2.varnames = {'SepalLength','SepalWidth',...
    'PetalLength','PetalWidth'};
iris2.obsnames = strcat({'Obs'},num2str((1:50)','%-d'));
iris2.species = {'setosa','versicolor','virginica'};
```

The data subset is then returned using a combination of dot and parenthesis indexing:

```
subset = iris2.data(1:5,1:2,1)
subset =
    5.1000 3.5000
    4.9000 3.0000
    4.7000 3.2000
    4.6000 3.1000
    5.0000 3.6000
```

For statistical data, structure arrays have many of the same limitations as cell arrays. Once again, dataset arrays (see "Dataset Arrays" on page 2-23), designed specifically for general statistical data, address these limitations.

## Statistical Functions

One of the advantages of working in the MATLAB language is that functions operate on entire arrays of data, not just on single scalar values. The functions are said to be vectorized. Vectorization allows for both efficient problem formulation, using array-based data, and efficient computation, using vectorized statistical functions.

When MATLAB and Statistics Toolbox statistical functions operate on a vector of numerical data (either a row vector or a column vector), they return a single computed statistic:

```
% Fisher's setosa data:
load fisheriris
setosa_indices = strcmp('setosa',species);
setosa = meas(setosa_indices,:);
% Single variable from the data:
setosa_sepal_length = setosa(:,1);
% Standard deviation of the variable:
std(setosa_sepal_length)
ans =
    0.3525
```

When statistical functions operate on a matrix of numerical data, they treat the columns independently, as separate measured variables, and return a vector of statistics-one for each variable:

```
std(setosa)
ans =
    0.3525 0.3791 0.1737 0.1054
```

The four standard deviations are for measurements of sepal length, sepal width, petal length, and petal width, respectively.

Compare this to

```
std(setosa(:))
ans =
    1.8483
```

which gives the standard deviation across the entire array (all measurements).
Compare the preceding statistical calculations to the more generic mathematical operation

```
sin(setosa)
```

This operation returns a 50 -by- 4 array the same size as setosa. The sin function is vectorized in a different way than the std function, computing one scalar value for each element in the array.

MATLAB and Statistics Toolbox statistical functions, like std, must be distinguished from general mathematical functions like sin. Both are vectorized, and both are useful for working with array-based data, but only statistical functions summarize data across observations (rows) while preserving variables (columns). This property of statistical functions may be explicit, as with std, or implicit, as with regress. To see how a particular function handles array-based data, consult its reference page.

MATLAB statistical functions expect data input arguments to be in the form of numerical arrays. If data is stored in a cell or structure array, it must be extracted to a numerical array, via indexing, for processing. Statistics Toolbox functions are more flexible. Many toolbox functions accept data input arguments in the form of both numerical arrays and dataset arrays (see "Dataset Arrays" on page 2-23), which are specifically designed for storing general statistical data.

## Statistical Arrays

In this section...<br>"Introduction to Statistical Arrays" on page 2-11<br>"Categorical Arrays" on page 2-13<br>"Dataset Arrays" on page 2-23

## Introduction to Statistical Arrays

As discussed in "MATLAB Arrays" on page 2-4, MATLAB data types include arrays for numerical, logical, and character data, as well as cell and structure arrays for heterogeneous collections of data.

Statistics Toolbox software offers two additional types of arrays specifically designed for statistical data:

- "Categorical Arrays" on page 2-13
- "Dataset Arrays" on page 2-23

Categorical arrays store data with values in a discrete set of levels. Each level is meant to capture a single, defining characteristic of an observation. If no ordering is encoded in the levels, the data and the array are nominal. If an ordering is encoded, the data and the array are ordinal.

Categorical arrays also store labels for the levels. Nominal labels typically suggest the type of an observation, while ordinal labels suggest the position or rank.

Dataset arrays collect heterogeneous statistical data and metadata, including categorical data, into a single container variable. Like the numerical matrices discussed in "Numerical Data" on page 2-4, dataset arrays can be viewed as tables of values, with rows representing different observations and columns representing different measured variables. Like the cell and structure arrays discussed in "Heterogeneous Data" on page 2-7, dataset arrays can accommodate variables of different types, sizes, units, etc.

Dataset arrays combine the organizational advantages of these basic MATLAB data types while addressing their shortcomings with respect to storing complex statistical data.

Both categorical and dataset arrays have associated methods for assembling, accessing, manipulating, and processing the collected data. Basic array operations parallel those for numerical, cell, and structure arrays.

## Categorical Arrays

- "Categorical Data" on page 2-13
- "Categorical Arrays" on page 2-14
- "Using Categorical Arrays" on page 2-16


## Categorical Data

Categorical data take on values from only a finite, discrete set of categories or levels. Levels may be determined before the data are collected, based on the application, or they may be determined by the distinct values in the data when converting them to categorical form. Predetermined levels, such as a set of states or numerical intervals, are independent of the data they contain. Any number of values in the data may attain a given level, or no data at all. Categorical data show which measured values share common levels, and which do not.

Levels may have associated labels. Labels typically express a defining characteristic of an observation, captured by its level.

If no ordering is encoded in the levels, the data are nominal. Nominal labels typically indicate the type of an observation. Examples of nominal labels are \{false, true\}, \{male, female\}, and \{Afghanistan, ..., Zimbabwe\}. For nominal data, the numeric or lexicographic order of the labels is irrelevant-Afghanistan is not considered to be less than, equal to, or greater than Zimbabwe.

If an ordering is encoded in the levels-for example, if levels labeled "red", "green", and "blue" represent wavelengths-the data are ordinal. Labels for ordinal levels typically indicate the position or rank of an observation. Examples of ordinal labels are $\{0,1\},\{\mathrm{mm}, \mathrm{cm}, \mathrm{m}, \mathrm{km}\}$, and \{poor, satisfactory, outstanding\}. The ordering of the levels may or may not correspond to the numeric or lexicographic order of the labels.

## Categorical Arrays

Categorical data can be represented using MATLAB integer arrays, but this method has a number of drawbacks. First, it removes all of the useful metadata that might be captured in labels for the levels. Labels must be stored separately, in character arrays or cell arrays of strings. Secondly, this method suggests that values stored in the integer array have their usual numeric meaning, which, for categorical data, they may not. Finally, integer types have a fixed set of levels (for example, -128:127 for all int8 arrays), which cannot be changed.

Categorical arrays, available in Statistics Toolbox software, are specifically designed for storing, manipulating, and processing categorical data and metadata. Unlike integer arrays, each categorical array has its own set of levels, which can be changed. Categorical arrays also accommodate labels for levels in a natural way. Like numerical arrays, categorical arrays take on different shapes and sizes, from scalars to $N$-D arrays.

Organizing data in a categorical array can be an end in itself. Often, however, categorical arrays are used for further statistical processing. They can be used to index into other variables, creating subsets of data based on the category of observation, or they can be used with statistical functions that accept categorical inputs. For examples, see "Grouped Data" on page 2-34.

Categorical arrays come in two types, depending on whether the collected data is understood to be nominal or ordinal. Nominal arrays are constructed with nominal; ordinal arrays are constructed with ordinal. For example,

```
load fisheriris
ndata = nominal(species,{'A','B','C'});
```

creates a nominal array with levels $\mathrm{A}, \mathrm{B}$, and C from the species data in fisheriris.mat, while

```
odata = ordinal(ndata,{},{'C','A','B'});
```

encodes an ordering of the levels with C < A < B. See "Using Categorical Arrays" on page 2-16, and the reference pages for nominal and ordinal, for further examples.

Categorical arrays are implemented as objects of the categorical class. The class is abstract, defining properties and methods common to both
the nominal class and ordinal class. Use the corresponding constructors, nominal or ordinal, to create categorical arrays. Methods of the classes are used to display, summarize, convert, concatenate, and access the collected data. Many of these methods are invoked using operations analogous to those for numerical arrays, and do not need to be called directly (for example, [ ] invokes horzcat). Other methods, such as reorderlevels, must be called directly.

## Using Categorical Arrays

This section provides an extended tutorial example demonstrating the use of categorical arrays with methods of the nominal class and ordinal class.

- "Constructing Categorical Arrays" on page 2-16
- "Accessing Categorical Arrays" on page 2-18
- "Combining Categorical Arrays" on page 2-19
- "Computing with Categorical Arrays" on page 2-20


## Constructing Categorical Arrays.

1 Load the 150-by-4 numerical array meas and the 150-by-1 cell array of strings species:

```
load fisheriris % Fisher's iris data (1936)
```

The data are 150 observations of four measured variables (by column number: sepal length, sepal width, petal length, and petal width, respectively) over three species of iris (setosa, versicolor, and virginica).

2 Use nominal to create a nominal array from species:

```
n1 = nominal(species);
```

3 Open species and n 1 side by side in the Variable Editor (see "Viewing and Editing Workspace Variables with the Variable Editor" in the MATLAB documentation). Note that the string information in species has been converted to categorical form, leaving only information on which data share the same values, indicated by the labels for the levels.

By default, levels are labeled with the distinct values in the data (in this case, the strings in species). Give alternate labels with additional input arguments to the nominal constructor:

```
n2 = nominal(species,{'species1','species2','species3'});
```

4 Open n2 in the Variable Editor, and compare it with species and n1. The levels have been relabeled.

5 Suppose that the data are considered to be ordinal. A characteristic of the data that is not reflected in the labels is the diploid chromosome count, which orders the levels corresponding to the three species as follows:

```
species1 < species3 < species2
```

Use ordinal to cast n2 as an ordinal array:

```
01 = ordinal(n2,{},{'species1','species3','species2'});
```

The second input argument to ordinal is the same as for nominal-a list of labels for the levels in the data. If it is unspecified, as above, the labels are inherited from the data, in this case n2. The third input argument of ordinal indicates the ordering of the levels, in ascending order.

6 When displayed side by side in the Variable Editor, 01 does not appear any different than n2. This is because the data in 01 have not been sorted. It is important to recognize the difference between the ordering of the levels in an ordinal array and sorting the actual data according to that ordering. Use sort to sort ordinal data in ascending order:

```
02 = sort(01);
```

When displayed in the Variable Editor, o2 shows the data sorted by diploid chromosome count.

7 To find which elements moved up in the sort, use the < operator for ordinal arrays:

```
moved_up = (01 < 02);
```

The operation returns a logical array moved_up, indicating which elements have moved up (the data for species3).

8 Use getlabels to display the labels for the levels in ascending order:

```
labels2 = getlabels(02)
labels2 =
    'species1' 'species3' 'species2'
```

9 The sort function reorders the display of the data, but not the order of the levels. To reorder the levels, use reorderlevels:

```
03 = reorderlevels(o2,labels2([11 3 2]));
labels3 = getlabels(o3)
labels3 =
    'species1' 'species2' 'species3'
o4 = sort(o3);
```

These operations return the levels in the data to their original ordering, by species number, and then sort the data for display purposes.

Accessing Categorical Arrays. Categorical arrays are accessed using parenthesis indexing, with syntax that parallels similar operations for numerical arrays (see "Numerical Data" on page 2-4).

Parenthesis indexing on the right-hand side of an assignment is used to extract the lowest 50 elements from the ordinal array 04:

```
low50 = 04(1:50);
```

Suppose you want to categorize the data in 04 with only two levels: low (the data in low50) and high (the rest of the data). One way to do this is to use an assignment with parenthesis indexing on the left-hand side:

```
05 = 04; % Copy 04
o5(1:50) = 'low';
Warning: Categorical level 'low' being added.
05(51:end) = 'high';
Warning: Categorical level 'high' being added.
```

Note the warnings: the assignments move data to new levels. The old levels, though empty, remain:

```
getlabels(05)
ans =
    'species1' 'species2' 'species3' 'low' 'high'
```

The old levels are removed using droplevels:

```
05 = droplevels(o5,{'species1','species2','species3'});
```

Another approach to creating two categories in 05 from the three categories in 04 is to merge levels, using mergelevels:

```
05 = mergelevels(04,{'species1'},'low');
05 = mergelevels(o5,{'species2','species3'},'high');
getlabels(05)
ans =
    'low' 'high'
```

The merged levels are removed and replaced with the new levels.
Combining Categorical Arrays. Categorical arrays are concatenated using square brackets. Again, the syntax parallels similar operations for numerical arrays (see "Numerical Data" on page 2-4). There are, however, restrictions:

- Only categorical arrays of the same type can be combined. You cannot concatenate a nominal array with an ordinal array.
- Only ordinal arrays with the same levels, in the same order, can be combined.
- Nominal arrays with different levels can be combined to produce a nominal array whose levels are the union of the levels in the component arrays.

First use ordinal to create ordinal arrays from the variables for sepal length and sepal width in meas. Categorize the data as short or long depending on whether they are below or above the median of the variable, respectively:

```
sl = meas(:,1); % Sepal length data
sw = meas(:,2); % Sepal width data
SL1 = ordinal(sl,{'short','long'},[],...
    [min(sl),median(sl),max(sl)]);
SW1 = ordinal(sw,{'short','long'},[],...
    [min(sw),median(sw),max(sw)]);
```

Because SL1 and SW1 are ordinal arrays with the same levels, in the same order, they can be concatenated:

```
S1 = [SL1,SW1];
S1(1:10,:)
ans =
    short long
    short long
    short long
```

| short | long |
| :--- | :--- |
| short | long |
| short | long |
| short | long |
| short | long |
| short | short |
| short | long |

The result is an ordinal array S 1 with two columns.
If, on the other hand, the measurements are cast as nominal, different levels can be used for the different variables, and the two nominal arrays can still be combined:

```
SL2 = nominal(sl,{'short','long'},[],...
    [min(sl),median(sl),max(sl)]);
SW2 = nominal(sw,{'skinny','wide'},[],...
    [min(sw),median(sw),max(sw)]);
S2 = [SL2,SW2];
getlabels(S2)
ans =
    'short' 'long' 'skinny' 'wide'
S2(1:10,:)
ans =
    short wide
    short wide
    short wide
    short wide
    short wide
    short wide
    short wide
    short wide
    short skinny
    short wide
```

Computing with Categorical Arrays. Categorical arrays are used to index into other variables, creating subsets of data based on the category of observation, and they are used with statistical functions that accept categorical inputs, such as those described in "Grouped Data" on page 2-34.

Use ismember to create logical variables based on the category of observation. For example, the following creates a logical index the same size as species that is true for observations of iris setosa and false elsewhere. Recall that n1 = nominal(species):

```
SetosaObs = ismember(n1,'setosa');
```

Since the code above compares elements of n 1 to a single value, the same operation is carried out by the equality operator:

```
SetosaObs = (n1 == 'setosa');
```

The SetosaObs variable is used to index into meas to extract only the setosa data:

```
SetosaData = meas(SetosaObs,:);
```

Categorical arrays are also used as grouping variables. The following plot summarizes the sepal length data in meas by category:

```
boxplot(sl,n1)
```



## Dataset Arrays

- "Statistical Data" on page 2-23
- "Dataset Arrays" on page 2-24
- "Using Dataset Arrays" on page 2-25


## Statistical Data

MATLAB data containers (variables) are suitable for completely homogeneous data (numeric, character, and logical arrays) and for completely heterogeneous data (cell and structure arrays). Statistical data, however, are often a mixture of homogeneous variables of heterogeneous types and sizes. Dataset arrays are suitable containers for this kind of data.

Dataset arrays can be viewed as tables of values, with rows representing different observations or cases and columns representing different measured variables. In this sense, dataset arrays are analogous to the numerical arrays for statistical data discussed in "Numerical Data" on page 2-4. Basic methods for creating and manipulating dataset arrays parallel the syntax of corresponding methods for numerical arrays.

While each column of a dataset array must be a variable of a single type, each row may contain an observation consisting of measurements of different types. In this sense, dataset arrays lie somewhere between variables that enforce complete homogeneity on the data and those that enforce nothing. Because of the potentially heterogeneous nature of the data, dataset arrays have indexing methods with syntax that parallels corresponding methods for cell and structure arrays (see "Heterogeneous Data" on page 2-7).

## Dataset Arrays

Dataset arrays are variables created with dataset. For example, the following creates a dataset array from observations that are a combination of categorical and numerical measurements:

```
load fisheriris
NumObs = size(meas,1);
NameObs = strcat({'Obs'},num2str((1:NumObs)','%-d'));
iris = dataset({nominal(species),'species'},...
    {meas,'SL','SW','PL','PW'},...
    'ObsNames',NameObs);
iris(1:5,:)
ans =
\begin{tabular}{llrrll} 
& species & SL & SW & PL & PW \\
Obs1 & setosa & 5.1 & 3.5 & 1.4 & 0.2 \\
Obs2 & setosa & 4.9 & 3 & 1.4 & 0.2 \\
Obs3 & setosa & 4.7 & 3.2 & 1.3 & 0.2 \\
Obs4 & setosa & 4.6 & 3.1 & 1.5 & 0.2 \\
Obs5 & setosa & 5 & 3.6 & 1.4 & 0.2
\end{tabular}
```

When creating a dataset array, variable names and observation names can be assigned together with the data. Other metadata associated with the array can be assigned with set and accessed with get:

```
iris = set(iris,'Description','Fisher''s Iris Data');
get(iris)
    Description: 'Fisher's Iris Data'
    Units: {}
    DimNames: {'Observations' 'Variables'}
    UserData: []
    ObsNames: {150x1 cell}
    VarNames: {'species' 'SL' 'SW' 'PL' 'PW'}
```

Dataset arrays are implemented as objects of the dataset class. Methods of the class are used to display, summarize, convert, concatenate, and access the collected data. Many of these methods are invoked using operations analogous to those for numerical arrays, and do not need to be called directly (for example, [] invokes horzcat). Other methods, such as sortrows, must be called directly.

## Using Dataset Arrays

This section provides an extended tutorial example demonstrating the use of dataset arrays with methods of the dataset class.

- "Constructing Dataset Arrays" on page 2-25
- "Accessing Dataset Arrays" on page 2-27
- "Combining Dataset Arrays" on page 2-29
- "Removing Observations from Dataset Arrays" on page 2-31
- "Computing with Dataset Arrays" on page 2-31

Constructing Dataset Arrays. Load the 150-by-4 numerical array meas and the 150 -by- 1 cell array of strings species:

```
load fisheriris % Fisher's iris data (1936)
```

The data are 150 observations of four measured variables (by column number: sepal length, sepal width, petal length, and petal width, respectively) over three species of iris (setosa, versicolor, and virginica).

Use dataset to create a dataset array iris from the data, assigning variable names species, SL, SW, PL, and PW and observation names Obs1, Obs2, Obs3, etc.:

```
NumObs = size(meas,1);
NameObs = strcat({'Obs'},num2str((1:NumObs)','%-d'));
iris = dataset({nominal(species),'species'},...
    {meas,'SL','SW','PL','PW'},...
    'ObsNames',NameObs);
iris(1:5,:)
ans =
\begin{tabular}{llllll} 
& species & \multicolumn{1}{l}{ SL } & \multicolumn{1}{l}{ SW } & PL & PW \\
Obs1 & setosa & 5.1 & 3.5 & 1.4 & 0.2 \\
Obs2 & setosa & 4.9 & 3 & 1.4 & 0.2 \\
Obs3 & setosa & 4.7 & 3.2 & 1.3 & 0.2 \\
Obs4 & setosa & 4.6 & 3.1 & 1.5 & 0.2 \\
Obs5 & setosa & 5 & 3.6 & 1.4 & 0.2
\end{tabular}
```

The cell array of strings species is first converted to a categorical array of type nominal before inclusion in the dataset array. For further information on categorical arrays, see "Categorical Arrays" on page 2-13.

Use set to set properties of the array:

```
desc = 'Fisher''s iris data (1936)';
units = [{''} repmat({'cm'},1,4)];
info = 'http://en.wikipedia.org/wiki/R.A._Fisher';
iris = set(iris,'Description',desc,...
    'Units',units,...
    'UserData',info);
```

Use get to view properties of the array:

```
get(iris)
    Description: 'Fisher's iris data (1936)'
            Units: {'' 'cm' 'cm' 'cm' 'cm'}
            DimNames: {'Observations' 'Variables'}
            UserData: 'http://en.wikipedia.org/wiki/R.A._Fisher'
            ObsNames: {150x1 cell}
            VarNames: {'species' 'SL' 'SW' 'PL' 'PW'}
get(iris(1:5,:),'ObsNames')
ans =
    'Obs1'
    'Obs2'
    'Obs3'
    'Obs4'
    'Obs5'
```

For a table of accessible properties of dataset arrays, with descriptions, see the reference on the dataset class.

Accessing Dataset Arrays. Dataset arrays support multiple types of indexing. Like the numerical matrices described in "Numerical Data" on page $2-4$, parenthesis () indexing is used to access data subsets. Like the cell and structure arrays described in "Heterogeneous Data" on page 2-7, dot . indexing is used to access data variables and curly brace \{\} indexing is used to access data elements.

Use parenthesis indexing to assign a subset of the data in iris to a new dataset array iris1:

```
iris1 = iris(1:5,2:3)
iris1 =
\begin{tabular}{lrr} 
& \multicolumn{1}{l}{ SL } & \multicolumn{1}{l}{ SW } \\
Obs1 & 5.1 & 3.5 \\
Obs2 & 4.9 & 3 \\
Obs3 & 4.7 & 3.2 \\
Obs4 & 4.6 & 3.1 \\
Obs5 & 5 & 3.6
\end{tabular}
```

Similarly, use parenthesis indexing to assign new data to the first variable in iris1:

```
iris1(:,1) = dataset([5.2;4.9;4.6;4.6;5])
iris1 =
    SL SW
    Obs1 5.2 3.5
    Obs2 4.9 3
    Obs3 4.6 3.2
    Obs4 4.6 3.1
    Obs5 5 3.6
```

Variable and observation names can also be used to access data:

```
SepalObs = iris1({'Obs1','Obs3','Obs5'},'SL')
SepalObs =
            SL
    Obs1 5.2
    Obs3 4.6
    Obs5 5
```

Dot indexing is used to access variables in a dataset array, and can be combined with other indexing methods. For example, apply zscore to the data in SepalObs as follows:

```
ScaledSepalObs = zscore(iris1.SL([1 3 5]))
ScaledSepalObs =
    0.8006
    -1.1209
    0.3203
```

The following code extracts the sepal lengths in iris1 corresponding to sepal widths greater than 3:

```
BigSWLengths = iris1.SL(iris1.SW > 3)
BigSWLengths =
    5.2000
    4.6000
    4.6000
    5.0000
```

Dot indexing also allows entire variables to be deleted from a dataset array:

```
iris1.SL = []
iris1 =
SW
Obs 1 3.5
Obs 2 3
Obs 3 3.2
Obs 4 3.1
Obs 5 3.6
```

Dynamic variable naming works for dataset arrays just as it does for structure arrays. For example, the units of the SW variable are changed in iris1 as follows:

```
varname = 'SW';
iris1.(varname) = iris1.(varname)*10
iris1 =
    SW
    Obs1 35
    Obs2 30
```

```
    Obs3 32
    Obs4 31
    Obs5 36
iris1 = set(iris1,'Units',{'mm'});
```

Curly brace indexing is used to access individual data elements. The following are equivalent:

```
iris1{1,1}
ans =
    3 5
iris1{'Obs1','SW'}
ans =
    35
```

Combining Dataset Arrays. Combine two dataset arrays into a single dataset array using square brackets:

```
SepalData = iris(:,{'SL','SW'});
PetalData = iris(:,{'PL','PW'});
newiris = [SepalData,PetalData];
size(newiris)
ans =
    150 4
```

For horizontal concatenation, as in the preceding example, the number of observations in the two dataset arrays must agree. Observations are matched up by name (if given), regardless of their order in the two data sets.

The following concatenates variables within a dataset array and then deletes the component variables:

```
newiris.SepalData = [newiris.SL,newiris.SW];
newiris.PetalData = [newiris.PL,newiris.PW];
newiris(:,{'SL','SW','PL','PW'}) = [];
size(newiris)
ans =
    150 2
size(newiris.SepalData)
ans =
```

newiris is now a 150-by-2 dataset array containing two 150-by-2 numerical arrays as variables.

Vertical concatenation is also handled in a manner analogous to numerical arrays:

```
newobs = dataset({[5.3 4.2; 5.0 4.1],'PetalData'},...
    {[5.5 2; 4.8 2.1],'SepalData'});
newiris = [newiris;newobs];
size(newiris)
ans =
    152 2
```

For vertical concatenation, as in the preceding example, the names of the variables in the two dataset arrays must agree. Variables are matched up by name, regardless of their order in the two data sets.

Expansion of variables is also accomplished using direct assignment to new rows:

```
newiris(153,:) = dataset({[5.1 4.0],'PetalData'},...
    {[5.1 4.2],'SepalData'});
```

A different type of concatenation is performed by join, which takes the data in one dataset array and assigns it to the rows of another dataset array, based on matching values in a common key variable. For example, the following creates a dataset array with diploid chromosome counts for each species of iris:

```
snames = nominal({'setosa';'versicolor';'virginica'});
CC = dataset({snames,'species'},{[38;108;70],'cc'})
CC =
    species cc
    setosa 38
    versicolor 108
    virginica 70
```

This data is broadcast to the rows of iris using join:

```
iris2 = join(iris,CC);
```

|  | species | SL | SW | PL | PW | CC |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Obs 1 | setosa | 5.1 | 3.5 | 1.4 | 0.2 | 38 |
| Obs2 | setosa | 4.9 | 3 | 1.4 | 0.2 | 38 |
| Obs51 | versicolor | 7 | 3.2 | 4.7 | 1.4 | 108 |
| Obs52 | versicolor | 6.4 | 3.2 | 4.5 | 1.5 | 108 |
| Obs101 | virginica | 6.3 | 3.3 | 6 | 2.5 | 70 |
| Obs102 | virginica | 5.8 | 2.7 | 5.1 | 1.9 | 70 |

Removing Observations from Dataset Arrays. Use one of the following commands to remove observations or variables from a dataset (ds):

- Remove a variable by name:
ds.var = [];
- Remove the $j$ th variable:

$$
\mathrm{ds}(:, j)=[] ;
$$

- Remove the $j$ th variable:

```
ds = ds(:,[1:(j-1) (j+1):end]);
```

- Remove the $i$ th observation:
ds(i,:) = [];
- Remove the $i$ th observation:

```
ds = ds([1:(i-1) (i+1):end],:);
```

- Remove the $j$ th variable and $i$ th observation:

```
ds = ds([1:(i-1) (i+1):end],[1:(j-1) (j+1):end]);
```

Computing with Dataset Arrays. Use summary to provide summary statistics for the component variables of a dataset array:

```
summary(newiris)
Fisher's iris data (1936)
SepalData: [153x2 double]
```

| min | 4.3000 | 2 |
| :---: | :---: | ---: |
| 1st Q | 5.1000 | 2.8000 |
| median | 5.8000 | 3 |
| 3rd Q | 6.4000 | 3.3250 |
| max | 7.9000 | 4.4000 |
| PetalData: [153x2 double] |  |  |
| min | 1 | 0.1000 |
| 1st Q | 1.6000 | 0.3000 |
| median | 4.4000 | 1.3000 |
| 3rd Q | 5.1000 | 1.8000 |
| max | 6.9000 | 4.2000 |

To apply other statistical functions, use dot indexing to access relevant variables:

```
SepalMeans = mean(newiris.SepalData)
SepalMeans =
    5.8294 3.0503
```

The same result is obtained with datasetfun, which applies functions to dataset array variables:

```
means = datasetfun(@mean,newiris,'UniformOutput',false)
means =
    [1x2 double] [1x2 double]
SepalMeans = means{1}
SepalMeans =
    5.8294 3.0503
```

An alternative approach is to cast data in a dataset array as double and apply statistical functions directly. Compare the following two methods for computing the covariance of the length and width of the SepalData in newiris:

```
covs = datasetfun(@cov,newiris,'UniformOutput',false)
covs =
    [2x2 double] [2x2 double]
SepalCovs = covs{1}
SepalCovs =
    0.6835 -0.0373
    -0.0373 0.2054
```

```
SepalCovs = cov(double(newiris(:,1)))
SepalCovs =
    0.6835 -0.0373
    -0.0373 0.2054
```


## Grouped Data

In this section...<br>"Grouping Variables" on page 2-34<br>"Group Definition" on page 2-35<br>"Functions for Grouped Data" on page 2-35<br>"Using Grouping Variables" on page 2-36

## Grouping Variables

Grouping variables are utility variables used to indicate which elements in a data set are to be considered together when computing statistics and creating visualizations. Typically, you use grouping variables for classification, where you give or try to infer the group of an observation. Grouping variables may be:

- Numeric vectors
- String arrays (also called character arrays)
- Cell arrays of strings
- Categorical arrays
- Logical vectors

Grouping variables have the same length as the variables (columns) in a data set. Observations (rows) $i$ and $j$ are considered to be in the same group if the values of the corresponding grouping variable are identical at those indices. Grouping variables with multiple columns are used to specify different groups within multiple variables.

For example, the following command loads the 150-by-4 numerical array meas and the 150 -by- 1 cell array of strings species into the workspace:

```
load fisheriris % Fisher's iris data (1936)
```

The data are 150 observations of four measured variables (by column number: sepal length, sepal width, petal length, and petal width, respectively) over three species of iris (setosa, versicolor, and virginica). To group the
observations by species, the following are all acceptable (and equivalent) grouping variables:

```
group1 = species; % Cell array of strings
group2 = grp2idx(species); % Numeric vector
group3 = char(species); % Character array
group4 = nominal(species); % Categorical array
```

These grouping variables can be supplied as input arguments to any of the functions described in "Functions for Grouped Data" on page 2-35. Examples are given in "Using Grouping Variables" on page 2-36.

## Group Definition

Each level of a grouping variable defines a group. The levels and the order of levels are decided as follows:

- For a numeric vector or a logical vector $G$, the set of group levels is the distinct values of $G$. The order is the sorted order of the unique values.
- For a cell array of strings or a character array G, the set of group levels is the distinct strings of $G$. The order for strings is the order of their first appearance in G.
- For a categorical vector G, the set of group levels and their order match the output of the getlabels (G) method.

Some functions, such as grpstats, can take a cell array of several grouping variables (such as \{G1 G2 G3\}) to group the observations in the data set by each combination of the grouping variable levels. The order is decided first by the order of the first grouping variables, then by the order of the second grouping variable, and so on.

## Functions for Grouped Data

The following table lists general Statistics Toolbox functions that accept a grouping variable group as an input argument. The grouping variable may be in the form of a numeric or logical vector, string array, cell array of strings, or categorical array, as described in "Grouping Variables" on page 2-34.

For a full description of the syntax of any particular function, and examples of its use, consult its reference page, linked from the table. "Using Grouping Variables" on page 2-36 also includes examples.

| Function | Basic Syntax for Grouped Data |
| :--- | :--- |
| gplotmatrix | gplotmatrix $(x, y$, group $)$ |
| grp2idx | $[G, G N]=\operatorname{grp} 2 i d x($ group $)$ |
| grpstats | means $=\operatorname{grpstats}(X$, group $)$ |
| gscatter | gscatter $(x, y$, group $)$ |

## Using Grouping Variables

This section provides an example demonstrating the use of grouping variables and associated functions. Grouping variables are introduced in "Grouping Variables" on page 2-34. A list of general functions accepting grouping variables appears in "Functions for Grouped Data" on page 2-35.

1 Load the 150 -by- 4 numerical array meas and the 150 -by- 1 cell array of strings species:

```
load fisheriris % Fisher's iris data (1936)
```

The data are 150 observations of four measured variables (by column number: sepal length, sepal width, petal length, and petal width, respectively) over three species of iris (setosa, versicolor, and virginica).

2 Compute some basic statistics for the data (median and interquartile range), by group, using the grpstats function:

```
[order,number,group_median,group_iqr] = ...
    grpstats(meas,species,{'gname','numel',@median,@iqr})
order =
        'setosa'
        'versicolor'
        'virginica'
number =
    50 50 50 50
```

```
    50 50 50 50
    50 50 50 50
group_median =
    5.0000 3.4000 1.5000 0.2000
    5.9000 2.8000 4.3500 1.3000
    6.5000 3.0000 5.5500 2.0000
group_iqr =
    0.4000 0.5000 0.2000 0.1000
    0.7000 0.5000 0.6000 0.3000
    0.7000 0.4000 0.8000 0.5000
```

The statistics appear in 3-by-4 arrays, corresponding to the 3 groups (categories) and 4 variables in the data. The order of the groups appears in the group names in order.

3 You can use grouping variables to create visualizations based on categories of observations. The following scatter plot, created with the gscatter function, shows the correlation between sepal length and sepal width in two species of iris. Use ismember to subset the two species from species:

```
subset = ismember(species,{'setosa','versicolor'});
scattergroup = species(subset);
gscatter(meas(subset,1),\ldots.
    meas(subset,2),...
    scattergroup)
xlabel('Sepal Length')
ylabel('Sepal Width')
```



## Descriptive Statistics

- "Introduction to Descriptive Statistics" on page 3-2
- "Measures of Central Tendency" on page 3-3
- "Measures of Dispersion" on page 3-5
- "Measures of Shape" on page 3-7
- "Resampling Statistics" on page 3-9
- "Data with Missing Values" on page 3-14


## Introduction to Descriptive Statistics

You may need to summarize large, complex data sets-both numerically and visually-to convey their essence to the data analyst and to allow for further processing. This chapter focuses on numerical summaries; Chapter 4, "Statistical Visualization" focuses on visual summaries.

## Measures of Central Tendency

Measures of central tendency locate a distribution of data along an appropriate scale.

The following table lists the functions that calculate the measures of central tendency.

| Function Name | Description |
| :--- | :--- |
| geomean | Geometric mean |
| harmmean | Harmonic mean |
| mean | Arithmetic average |
| median | 50 th percentile |
| mode | Most frequent value |
| trimmean | Trimmed mean |

The average is a simple and popular estimate of location. If the data sample comes from a normal distribution, then the sample mean is also optimal (MVUE of $\mu$ ).

Unfortunately, outliers, data entry errors, or glitches exist in almost all real data. The sample mean is sensitive to these problems. One bad data value can move the average away from the center of the rest of the data by an arbitrarily large distance.

The median and trimmed mean are two measures that are resistant (robust) to outliers. The median is the 50th percentile of the sample, which will only change slightly if you add a large perturbation to any value. The idea behind the trimmed mean is to ignore a small percentage of the highest and lowest values of a sample when determining the center of the sample.

The geometric mean and harmonic mean, like the average, are not robust to outliers. They are useful when the sample is distributed lognormal or heavily skewed.

The following example shows the behavior of the measures of location for a sample with one outlier.

```
x = [ones(1,6) 100]
X =
    1
locate = [geomean(x) harmmean(x) mean(x) median(x)...
    trimmean(x,25)]
locate =
    1.9307 1.1647 15.1429 1.0000 1.0000
```

You can see that the mean is far from any data value because of the influence of the outlier. The median and trimmed mean ignore the outlying value and describe the location of the rest of the data values.

## Measures of Dispersion

The purpose of measures of dispersion is to find out how spread out the data values are on the number line. Another term for these statistics is measures of spread.

The table gives the function names and descriptions.

| Function <br> Name | Description |
| :--- | :--- |
| iqr | Interquartile range |
| mad | Mean absolute deviation |
| moment | Central moment of all orders |
| range | Range |
| std | Standard deviation |
| var | Variance |

The range (the difference between the maximum and minimum values) is the simplest measure of spread. But if there is an outlier in the data, it will be the minimum or maximum value. Thus, the range is not robust to outliers.

The standard deviation and the variance are popular measures of spread that are optimal for normally distributed samples. The sample variance is the MVUE of the normal parameter $\sigma^{2}$. The standard deviation is the square root of the variance and has the desirable property of being in the same units as the data. That is, if the data is in meters, the standard deviation is in meters as well. The variance is in meters ${ }^{2}$, which is more difficult to interpret.

Neither the standard deviation nor the variance is robust to outliers. A data value that is separate from the body of the data can increase the value of the statistics by an arbitrarily large amount.

The mean absolute deviation (MAD) is also sensitive to outliers. But the MAD does not move quite as much as the standard deviation or variance in response to bad data.

The interquartile range (IQR) is the difference between the 75th and 25th percentile of the data. Since only the middle $50 \%$ of the data affects this measure, it is robust to outliers.

The following example shows the behavior of the measures of dispersion for a sample with one outlier.

```
x = [ones(1,6) 100]
x =
```



```
stats = [iqr(x) mad(x) range(x) std(x)]
stats =
    04.2449 99.0000 37.4185
```


## Measures of Shape

Quantiles and percentiles provide information about the shape of data as well as its location and spread.

The quantile of order $p(0 \leq p \leq 1)$ is the smallest $x$ value where the cumulative distribution function equals or exceeds $p$. The function quantile computes quantiles as follows:
$1 n$ sorted data points are the $0.5 / n, 1.5 / n, \ldots,(n-0.5) / n$ quantiles.
2 Linear interpolation is used to compute intermediate quantiles.
3 The data min or max are assigned to quantiles outside the range.
4 Missing values are treated as NaN , and removed from the data.
Percentiles, computed by the prctile function, are quantiles for a certain percentage of the data, specified for $0 \leq p \leq 100$.

The following example shows the result of looking at every quartile (quantiles with orders that are multiples of 0.25 ) of a sample containing a mixture of two distributions.

```
x = [normrnd(4,1,1,100) normrnd(6,0.5,1,200)];
p = 100*(0:0.25:1);
y = prctile(x,p);
z = [p;y]
z =
\begin{tabular}{rrrrr}
0 & 25.0000 & 50.0000 & 75.0000 & 100.0000 \\
1.8293 & 4.6728 & 5.6459 & 6.0766 & 7.1546
\end{tabular}
```

A box plot helps to visualize the statistics:

```
boxplot(x)
```



## Column Number

The long lower tail and plus signs show the lack of symmetry in the sample values. For more information on box plots, see "Box Plots" on page 4-6.

The shape of a data distribution is also measured by the Statistics Toolbox functions skewness, kurtosis, and, more generally, moment.

## Resampling Statistics

In this section...<br>"The Bootstrap" on page 3-9<br>"The Jackknife" on page 3-12<br>"Parallel Computing Support for Resampling Methods" on page 3-13

## The Bootstrap

The bootstrap procedure involves choosing random samples with replacement from a data set and analyzing each sample the same way. Sampling with replacement means that each observation is selected separately at random from the original dataset. So a particular data point from the original data set could appear multiple times in a given bootstrap sample. The number of elements in each bootstrap sample equals the number of elements in the original data set. The range of sample estimates you obtain enables you to establish the uncertainty of the quantity you are estimating.

This example from Efron and Tibshirani [33] compares Law School Admission Test (LSAT) scores and subsequent law school grade point average (GPA) for a sample of 15 law schools.

```
load lawdata
plot(lsat,gpa,'+')
lsline
```



The least-squares fit line indicates that higher LSAT scores go with higher law school GPAs. But how certain is this conclusion? The plot provides some intuition, but nothing quantitative.

You can calculate the correlation coefficient of the variables using the corr function.

```
rhohat = corr(lsat,gpa)
rhohat =
    0.7764
```

Now you have a number describing the positive connection between LSAT and GPA; though it may seem large, you still do not know if it is statistically significant.

Using the bootstrp function you can resample the lsat and gpa vectors as many times as you like and consider the variation in the resulting correlation coefficients.

Here is an example.

```
rhos1000 = bootstrp(1000,'corr',lsat,gpa);
```

This command resamples the lsat and gpa vectors 1000 times and computes the corr function on each sample. Here is a histogram of the result.

```
hist(rhos1000,30)
set(get(gca,'Children'),'FaceColor',[.8 .8 1])
```



Nearly all the estimates lie on the interval [0.4 1.0].

It is often desirable to construct a confidence interval for a parameter estimate in statistical inferences. Using the bootci function, you can use bootstrapping to obtain a confidence interval. The confidence interval for the lsat and gpa data is computed as:

```
ci = bootci(5000,@corr,lsat,gpa)
ci =
```

0.3313
0.9427

Therefore, a $95 \%$ confidence interval for the correlation coefficient between LSAT and GPA is [0.33 0.94]. This is strong quantitative evidence that LSAT and subsequent GPA are positively correlated. Moreover, this evidence does not require any strong assumptions about the probability distribution of the correlation coefficient.

Although the bootci function computes the Bias Corrected and accelerated ( BCa ) interval as the default type, it is also able to compute various other types of bootstrap confidence intervals, such as the studentized bootstrap confidence interval.

## The Jackknife

Similar to the bootstrap is the jackknife, which uses resampling to estimate the bias of a sample statistic. Sometimes it is also used to estimate standard error of the sample statistic. The jackknife is implemented by the Statistics Toolbox function jackknife.

The jackknife resamples systematically, rather than at random as the bootstrap does. For a sample with $n$ points, the jackknife computes sample statistics on $n$ separate samples of size $n-1$. Each sample is the original data with a single observation omitted.

In the previous bootstrap example you measured the uncertainty in estimating the correlation coefficient. You can use the jackknife to estimate the bias, which is the tendency of the sample correlation to over-estimate or under-estimate the true, unknown correlation. First compute the sample correlation on the data:

```
load lawdata
rhohat = corr(lsat,gpa)
rhohat =
    0.7764
```

Next compute the correlations for jackknife samples, and compute their mean:

```
jackrho = jackknife(@corr,lsat,gpa);
meanrho = mean(jackrho)
meanrho =
```

0.7759

Now compute an estimate of the bias:

```
n = length(lsat);
biasrho = (n-1) * (meanrho-rhohat)
biasrho =
```

    -0.0065
    The sample correlation probably underestimates the true correlation by about this amount.

## Parallel Computing Support for Resampling Methods

For information on computing resampling statistics in parallel, see Chapter 17, "Parallel Statistics".

## Data with Missing Values

Many data sets have one or more missing values. It is convenient to code missing values as NaN (Not a Number) to preserve the structure of data sets across multiple variables and observations.

For example:

```
X = magic(3);
X([1 5]) = [NaN NaN]
X =
    NaN 1 6
    3 NaN 7
    4 9 2
```

Normal MATLAB arithmetic operations yield NaN values when operands are NaN :

```
s1 = sum(X)
s1 =
    NaN NaN 15
```

Removing the NaN values would destroy the matrix structure. Removing the rows containing the NaN values would discard data. Statistics Toolbox functions in the following table remove NaN values only for the purposes of computation.

| Function | Description |
| :--- | :--- |
| nancov | Covariance matrix, ignoring NaN values |
| nanmax | Maximum, ignoring NaN values |
| nanmean | Mean, ignoring NaN values |
| nanmedian | Median, ignoring NaN values |
| nanmin | Minimum, ignoring NaN values |
| nanstd | Standard deviation, ignoring NaN values |
| nansum | Sum, ignoring NaN values |
| nanvar | Variance, ignoring NaN values |

For example:

```
s2 = nansum(X)
s2 =
    7 10 15
```

Other Statistics Toolbox functions also ignore NaN values. These include iqr, kurtosis, mad, prctile, range, skewness, and trimmean.

## Statistical Visualization

- "Introduction to Statistical Visualization" on page 4-2
- "Scatter Plots" on page 4-3
- "Box Plots" on page 4-6
- "Distribution Plots" on page 4-8


## Introduction to Statistical Visualization

Statistics Toolbox data visualization functions add to the extensive graphics capabilities already in MATLAB.

- Scatter plots are a basic visualization tool for multivariate data. They are used to identify relationships among variables. Grouped versions of these plots use different plotting symbols to indicate group membership. The gname function is used to label points on these plots with a text label or an observation number.
- Box plots display a five-number summary of a set of data: the median, the two ends of the interquartile range (the box), and two extreme values (the whiskers) above and below the box. Because they show less detail than histograms, box plots are most useful for side-by-side comparisons of two distributions.
- Distribution plots help you identify an appropriate distribution family for your data. They include normal and Weibull probability plots, quantile-quantile plots, and empirical cumulative distribution plots.

Advanced Statistics Toolbox visualization functions are available for specialized statistical analyses.

Note For information on creating visualizations of data by group, see "Grouped Data" on page 2-34.

## Scatter Plots

A scatter plot is a simple plot of one variable against another. The MATLAB functions plot and scatter produce scatter plots. The MATLAB function plotmatrix can produce a matrix of such plots showing the relationship between several pairs of variables.

Statistics Toolbox functions gscatter and gplotmatrix produce grouped versions of these plots. These are useful for determining whether the values of two variables or the relationship between those variables is the same in each group.

Suppose you want to examine the weight and mileage of cars from three different model years.

```
load carsmall
gscatter(Weight,MPG,Model_Year,'','xos')
```



This shows that not only is there a strong relationship between the weight of a car and its mileage, but also that newer cars tend to be lighter and have better gas mileage than older cars.

The default arguments for gscatter produce a scatter plot with the different groups shown with the same symbol but different colors. The last two arguments above request that all groups be shown in default colors and with different symbols.

The carsmall data set contains other variables that describe different aspects of cars. You can examine several of them in a single display by creating a grouped plot matrix.

```
xvars = [Weight Displacement Horsepower];
yvars = [MPG Acceleration];
```



The upper right subplot displays MPG against Horsepower, and shows that over the years the horsepower of the cars has decreased but the gas mileage has improved.

The gplotmatrix function can also graph all pairs from a single list of variables, along with histograms for each variable. See "MANOVA" on page 8-39.

## Box Plots

The graph below, created with the boxplot command, compares petal lengths in samples from two species of iris.

```
    load fisheriris
    s1 = meas(51:100,3);
    s2 = meas(101:150,3);
boxplot([s1 s2],'notch','on',...
    'labels',{'versicolor','virginica'})
```



This plot has the following features:

- The tops and bottoms of each "box" are the 25 th and 75 th percentiles of the samples, respectively. The distances between the tops and bottoms are the interquartile ranges.
- The line in the middle of each box is the sample median. If the median is not centered in the box, it shows sample skewness.
- The whiskers are lines extending above and below each box. Whiskers are drawn from the ends of the interquartile ranges to the furthest observations within the whisker length (the adjacent values).
- Observations beyond the whisker length are marked as outliers. By default, an outlier is a value that is more than 1.5 times the interquartile range away from the top or bottom of the box, but this value can be adjusted with additional input arguments. Outliers are displayed with a red + sign.
- Notches display the variability of the median between samples. The width of a notch is computed so that box plots whose notches do not overlap (as above) have different medians at the $5 \%$ significance level. The significance level is based on a normal distribution assumption, but comparisons of medians are reasonably robust for other distributions. Comparing box-plot medians is like a visual hypothesis test, analogous to the $t$ test used for means.


## Distribution Plots

In this section...<br>"Normal Probability Plots" on page 4-8<br>"Quantile-Quantile Plots" on page 4-10<br>"Cumulative Distribution Plots" on page 4-12<br>"Other Probability Plots" on page 4-14

## Normal Probability Plots

Normal probability plots are used to assess whether data comes from a normal distribution. Many statistical procedures make the assumption that an underlying distribution is normal, so normal probability plots can provide some assurance that the assumption is justified, or else provide a warning of problems with the assumption. An analysis of normality typically combines normal probability plots with hypothesis tests for normality, as described in Chapter 7, "Hypothesis Tests".

The following example shows a normal probability plot created with the normplot function.

```
x = normrnd(10,1,25,1);
normplot(x)
```



The plus signs plot the empirical probability versus the data value for each point in the data. A solid line connects the 25 th and 75 th percentiles in the data, and a dashed line extends it to the ends of the data. The $y$-axis values are probabilities from zero to one, but the scale is not linear. The distance between tick marks on the $y$-axis matches the distance between the quantiles of a normal distribution. The quantiles are close together near the median (probability $=0.5$ ) and stretch out symmetrically as you move away from the median.

In a normal probability plot, if all the data points fall near the line, an assumption of normality is reasonable. Otherwise, the points will curve away from the line, and an assumption of normality is not justified.

For example:

$$
x=\operatorname{exprnd}(10,100,1) ;
$$

normplot(x)


The plot is strong evidence that the underlying distribution is not normal.

## Quantile-Quantile Plots

Quantile-quantile plots are used to determine whether two samples come from the same distribution family. They are scatter plots of quantiles computed from each sample, with a line drawn between the first and third quartiles. If the data falls near the line, it is reasonable to assume that the two samples come from the same distribution. The method is robust with respect to changes in the location and scale of either distribution.

To create a quantile-quantile plot, use the qqplot function.

The following example shows a quantile-quantile plot of two samples from Poisson distributions.

```
x = poissrnd(10,50,1);
y = poissrnd(5,100,1);
qqplot(x,y);
```



Even though the parameters and sample sizes are different, the approximate linear relationship suggests that the two samples may come from the same distribution family. As with normal probability plots, hypothesis tests, as described in Chapter 7, "Hypothesis Tests", can provide additional justification for such an assumption. For statistical procedures that depend on the two samples coming from the same distribution, however, a linear quantile-quantile plot is often sufficient.

The following example shows what happens when the underlying distributions are not the same.

```
x = normrnd(5,1,100,1);
y = wblrnd(2,0.5,100,1);
qqplot(x,y);
```



These samples clearly are not from the same distribution family.

## Cumulative Distribution Plots

An empirical cumulative distribution function (cdf) plot shows the proportion of data less than each $x$ value, as a function of $x$. The scale on the $y$-axis is linear; in particular, it is not scaled to any particular distribution. Empirical cdf plots are used to compare data cdfs to cdfs for particular distributions.

To create an empirical cdf plot, use the cdfplot function (or ecdf and stairs).
The following example compares the empirical cdf for a sample from an extreme value distribution with a plot of the cdf for the sampling distribution. In practice, the sampling distribution would be unknown, and would be chosen to match the empirical cdf.

```
y = evrnd(0,3,100,1);
cdfplot(y)
hold on
x = -20:0.1:10;
f = evcdf(x,0,3);
plot(x,f,'m')
legend('Empirical','Theoretical','Location','NW')
```



## Other Probability Plots

A probability plot, like the normal probability plot, is just an empirical cdf plot scaled to a particular distribution. The $y$-axis values are probabilities from zero to one, but the scale is not linear. The distance between tick marks is the distance between quantiles of the distribution. In the plot, a line is drawn between the first and third quartiles in the data. If the data falls near the line, it is reasonable to choose the distribution as a model for the data.

To create probability plots for different distributions, use the probplot function.

For example, the following plot assesses two samples, one from a Weibull distribution and one from a Rayleigh distribution, to see if they may have come from a Weibull population.

```
x1 = wblrnd(3,3,100,1);
x2 = raylrnd(3,100,1);
probplot('weibull',[x1 x2])
legend('Weibull Sample','Rayleigh Sample','Location','NW')
```



The plot gives justification for modeling the first sample with a Weibull distribution; much less so for the second sample.

A distribution analysis typically combines probability plots with hypothesis tests for a particular distribution, as described in Chapter 7, "Hypothesis Tests".

## Probability Distributions

- "Using Probability Distributions" on page 5-2
- "Supported Distributions" on page 5-3
- "Working with Distributions Through GUIs" on page 5-9
- "Statistics Toolbox Distribution Functions" on page 5-52
- "Using Probability Distribution Objects" on page 5-84
- "Probability Distributions Used for Multivariate Modeling" on page 5-99


## Using Probability Distributions

Probability distributions are theoretical distributions based on assumptions about a source population. They assign probability to the event that a random variable takes on a specific, discrete value, or falls within a specified range of continuous values. There are two main types of models:

- Parametric Models-Choose a model based on a parametric family of probability distributions and then adjust the parameters to fit the data. For information on supported parametric distributions, see "Parametric Distributions" on page 5-4.
- Nonparametric Models-When data or statistics do not follow any standard probability distribution, nonparametric models may be appropriate. For information on supported nonparametric distributions, see "Nonparametric Distributions" on page 5-8.

The Statistics Toolbox provides several ways of working with both parametric and nonparametric probability distributions:

- Graphic User Interfaces (GUIs)—Interact with the distributions to visualize distributions, fit a distribution to your data, or generate random data using a specific distribution. For more information, see "Working with Distributions Through GUIs" on page 5-9.
- Command Line Functions-Use command-line functions to further explore the distributions, fit relevant models to your data, or generate random data. For more information on using functions, see "Statistics Toolbox Distribution Functions" on page 5-52.
- Distribution Objects-Use objects to explore and fit your data to a distribution, save the results to a single entity, and generate random data from the resulting parameters. For more information, see "Using Probability Distribution Objects" on page 5-84.


## Supported Distributions

In this section...<br>"Parametric Distributions" on page 5-4<br>"Nonparametric Distributions" on page 5-8

Probability distributions supported by the Statistics Toolbox are cross-referenced with their supporting functions and GUIs in the following tables. The tables use the following abbreviations for distribution functions:

- pdf - Probability density functions
- cdf - Cumulative distribution functions
- inv - Inverse cumulative distribution functions
- stat - Distribution statistics functions
- fit - Distribution fitting functions
- like - Negative log-likelihood functions
- rnd - Random number generators

For more detailed explanations of each supported distribution, see Appendix B, "Distribution Reference".

## Parametric Distributions

## Continuous Distributions (Data)

| Name | pdf | cdf | inv | stat | fit | like | rnd |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Beta | betapdf, pdf | betacdf, cdf | betainv, icdf | betastat | betafit, <br> fitdist, <br> mle, <br> dfittool | betalike | betarnd, random, randtool |
| BirnbaumSaunders |  |  |  |  | mle, <br> fitdist, <br> dfittool |  |  |
| Exponential | exppdf, pdf | expcdf, cdf | expinv, icdf | expstat | ```expfit, mle, fitdist, dfittool``` | explike | exprnd, random, randtool |
| Extreme value | evpdf, pdf | evcdf, cdf | evinv, icdf | evstat | evfit, mle, fitdist, dfittool | evlike | evrnd, random, randtool |
| Gamma | gampdf, pdf | gamcdf, cdf | gaminv, icdf | gamstat | ```gamfit, mle, fitdist, dfittool``` | gamlike | gamrnd, <br> randg, <br> random, randtool |
| Generalized extreme value | gevpdf, pdf | gevcdf, cdf | gevinv, icdf | gevstat | ```gevfit, mle, fitdist, dfittool``` | gevlike | gevrnd, random, randtool |
| Generalized <br> Pareto | gppdf, pdf | gpcdf, cdf | gpinv, icdf | gpstat | gpfit, mle, <br> fitdist, <br> dfittool | gplike | gprnd, random, randtool |
| Inverse Gaussian |  |  |  |  | mle, <br> fitdist, <br> dfittool |  |  |
| Johnson system |  |  |  |  | johnsrnd |  | johnsrnd |


| Name | pdf | cdf | inv | stat | fit | like | rnd |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Logistic |  |  |  |  | mle, <br> fitdist, <br> dfittool |  |  |
| Loglogistic |  |  |  |  | mle, <br> fitdist, <br> dfittool |  |  |
| Lognormal | lognpdf, <br> pdf | logncdf, <br> cdf | logninv, <br> icdf | lognstat | lognfit, <br> mle, <br> fitdist, <br> dfittool | lognlike | lognrnd, <br> random, <br> randtool |
| Nakagami |  |  |  |  | mle, <br> fitdist, <br> dfittool |  |  |
| Normal <br> (Gaussian) | normpdf, <br> pdf | normcdf, <br> cdf | norminv, <br> icdf | normstat | normfit, <br> mle, <br> fitdist, <br> dfittool | normlike | normrnd, <br> rand, <br> random, <br> randtool |
| Pearson <br> system |  |  |  |  | pearsrnd |  | pearsrnd |

## Continuous Distributions (Statistics)

| Name | pdf | cdf | inv | stat | fit | like | rnd |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Chi-square | chi2pdf, <br> pdf | chi2cdf, <br> cdf | chi2inv, <br> icdf | chi2stat |  |  | chi2rnd, <br> random, <br> randtool |
| $F$ | fpdf, pdf | fcdf, cdff | finv, <br> icdf | fstat |  |  | frnd, <br> random, <br> randtool |
| Noncentral <br> chi-square | ncx2pdf, <br> pdf | ncx2cdf, <br> cdf | ncx2inv, <br> icdf | ncx2stat |  | ncx2rnd, <br> random, <br> randtool |  |
| Noncentral <br> $F$ | ncfpdf, <br> pdf | ncfcdf, <br> cdf | ncfinv, <br> icdf | ncfstat |  | ncfrnd, <br> random, <br> randtool |  |
| Noncentral <br> $t$ | nctpdf, <br> pdf | nctcdf, <br> cdf | nctinv, <br> icdf | nctstat |  | nctrnd, <br> random, <br> randtool |  |
| Student's $t$ | tpdf, pdf | tcdf, cdf | tinv, <br> icdf | tstat |  | trnd, <br> random, <br> randtool |  |
| $t$ location- <br> scale |  |  |  | mle, <br> fitdist, <br> dfittool |  |  |  |

## Discrete Distributions

| Name | pdf | cdf | inv | stat | fit | like | rnd |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Binomial | binopdf, pdf | binocdf, cdf | binoinv, icdf | binostat | binofit, mle, fitdist, dfittool |  | binornd, random, randtool |
| Bernoulli |  |  |  |  | mle |  |  |
| Geometric | geopdf, pdf | geocdf, cdf | geoinv, icdf | geostat | mle |  | geornd, random, randtool |
| Hypergeomet | rłygepdf, pdf | hygecdf, cdf | hygeinv, icdf | hygestat |  |  | hygernd, random |
| Multinomial | mnpdf |  |  |  |  |  | mnrnd |
| Negative binomial | nbinpdf, pdf | nbincdf, cdf | $\begin{aligned} & \text { nbininv, } \\ & \text { icdf } \end{aligned}$ | nbinstat | nbinfit, mle, <br> fitdist, dfittool |  | nbinrnd, random, randtool |
| Poisson | $\begin{aligned} & \text { poisspdf, } \\ & \text { pdf } \end{aligned}$ | $\begin{aligned} & \text { poisscdf, } \\ & \text { cdf } \end{aligned}$ | $\begin{aligned} & \text { poissinv, } \\ & \text { icdf } \end{aligned}$ | poisstat | ```poissfit, mle, fitdist, dfittool``` |  | poissrnd, random, randtool |
| Uniform (discrete) | unidpdf, pdf | unidcdf, cdf | $\begin{aligned} & \text { unidinv, } \\ & \text { icdf } \end{aligned}$ | unidstat | mle |  | unidrnd, random, randtool |

## Multivariate Distributions

| Name | pdf | cdf | inv | stat | fit | like | rnd |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Gaussian <br> copula | copulapdf | copulacdf |  | copulastat | copulafit |  | copularnd |
| Gaussian <br> mixture | pdf | cdf |  |  | fit |  | random |
| $t$ copula | copulapdf | copulacdf |  | copulastat | copulafit |  | copularnd |
| Clayton <br> copula | copulapdf | copulacdf |  | copulastat | copulafit |  | copularnd |
| Frank <br> copula | copulapdf | copulacdf |  | copulastat | copulafit |  | copularnd |
| Gumbel <br> copula | copulapdf | copulacdf |  | copulastat | copulafit |  | copularnd |
| Inverse <br> Wishart |  |  |  |  | iwishrnd |  |  |
| Multivariate <br> normal | mvnpdf | mvncdf |  |  |  | mvnrnd |  |
| Multivariate <br> $t$ | mvtpdf | mvtcdf |  |  |  | mvtrnd |  |
| Wishart |  |  |  |  |  | wishrnd |  |

Nonparametric Distributions

| Name | pdf | cdf | inv | stat | fit | like | rnd |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Nonparametric | ksdensity | ksdensity | ksdensity |  | ksdensity, <br> fitdist, <br> dfittool |  |  |

# Working with Distributions Through GUls 

In this section...<br>"Exploring Distributions" on page 5-9<br>"Modeling Data Using the Distribution Fitting Tool" on page 5-11<br>"Visually Exploring Random Number Generation" on page 5-49

This section describes Statistics Toolbox GUIs that provide convenient, interactive access to the distribution functions described in "Statistics Toolbox Distribution Functions" on page 5-52.

## Exploring Distributions

To interactively see the influence of parameter changes on the shapes of the pdfs and cdfs of supported Statistics Toolbox distributions, use the Probability Distribution Function Tool.

Run the tool by typing disttool at the command line.


Start by selecting a distribution. Then choose the function type: probability density function (pdf) or cumulative distribution function (cdf).

After the plot appears, you can

- Calculate a new function value by
- Typing a new x value in the text box on the $x$-axis
- Dragging the vertical reference line.
- Clicking in the figure where you want the line to be.

The new function value appears in the text box to the left of the plot.

- For cdf plots, find critical values corresponding to a specific probability by typing the desired probability in the text box on the $y$-axis or by dragging the horizontal reference line.
- Use the controls at the bottom of the window to set parameter values for the distribution and to change their upper and lower bounds.


## Modeling Data Using the Distribution Fitting Tool

The Distribution Fitting Tool is a GUI for fitting univariate distributions to data. This section describes how to use the Distribution Fitting Tool this tool and covers the following topics:

- "Opening the Distribution Fitting Tool" on page 5-12
- "Creating and Managing Data Sets" on page 5-14
- "Creating a New Fit" on page 5-19
- "Displaying Results" on page 5-24
- "Managing Fits" on page 5-26
- "Evaluating Fits" on page 5-28
- "Excluding Data" on page 5-32
- "Saving and Loading Sessions" on page 5-38
- "Example: Fitting a Distribution" on page 5-39
- "Generating a File to Fit and Plot Distributions" on page 5-46
- "Using Custom Distributions" on page 5-47
- "Additional Distributions Available in the Distribution Fitting Tool" on page 5-49


## Opening the Distribution Fitting Tool

To open the Distribution Fitting Tool, enter the command

```
dfittool
```



Adjusting the Plot. Buttons at the top of the tool allow you to adjust the plot displayed in this window:

- $\quad$ - Toggle the legend on (default) or off.
- 䜹 — Toggle grid lines on or off (default).
-     - Restore default axes limits.

Displaying the Data. The Display type field specifies the type of plot displayed in the main window. Each type corresponds to a probability function, for example, a probability density function. The following display types are available:

- Density (PDF) — Display a probability density function (PDF) plot for the fitted distribution.
- Cumulative probability (CDF) - Display a cumulative probability plot of the data.
- Quantile (inverse CDF) - Display a quantile (inverse CDF) plot.
- Probability plot - Display a probability plot.
- Survivor function - Display a survivor function plot of the data.
- Cumulative hazard - Display a cumulative hazard plot of the data.

Inputting and Fitting Data. The task buttons enable you to perform the tasks necessary to fit distributions to data. Each button opens a new dialog box in which you perform the task. The buttons include:

- Data - Import and manage data sets. See "Creating and Managing Data Sets" on page 5-14.
- New Fit - Create new fits. See "Creating a New Fit" on page 5-19.
- Manage Fits - Manage existing fits. See "Managing Fits" on page 5-26.
- Evaluate - Evaluate fits at any points you choose. See "Evaluating Fits" on page 5-28.
- Exclude - Create rules specifying which values to exclude when fitting a distribution. See "Excluding Data" on page 5-32.

The display pane displays plots of the data sets and fits you create. Whenever you make changes in one of the dialog boxes, the results in the display pane update.

Saving and Customizing Distributions. The Distribution Fitting Tool menus contain items that enable you to do the following:

- Save and load sessions. See "Saving and Loading Sessions" on page 5-38.
- Generate a file with which you can fit distributions to data and plot the results independently of the Distribution Fitting Tool. See "Generating a File to Fit and Plot Distributions" on page 5-46.
- Define and import custom distributions. See "Using Custom Distributions" on page 5-47.


## Creating and Managing Data Sets

This section describes how to create and manage data sets.
To begin, click the Data button in the Distribution Fitting Tool to open the Data dialog box shown in the following figure.


Importing Data. The Import workspace vectors pane enables you to create a data set by importing a vector from the MATLAB workspace. The following sections describe the fields in this pane and give appropriate values for vectors imported from the MATLAB workspace:

- Data - The drop-down list in the Data field contains the names of all matrices and vectors, other than 1-by-1 matrices (scalars) in the MATLAB workspace. Select the array containing the data you want to fit. The actual data you import must be a vector. If you select a matrix in the Data field, the first column of the matrix is imported by default. To select a different column or row of the matrix, click Select Column or Row. This displays
the matrix in the Variable Editor, where you can select a row or column by highlighting it with the mouse.

Alternatively, you can enter any valid MATLAB expression in the Data field.

When you select a vector in the Data field, a histogram of the data appears in the Data preview pane.

- Censoring - If some of the points in the data set are censored, enter a Boolean vector, of the same size as the data vector, specifying the censored entries of the data. A 1 in the censoring vector specifies that the corresponding entry of the data vector is censored, while a 0 specifies that the entry is not censored. If you enter a matrix, you can select a column or row by clicking Select Column or Row. If you do not want to censor any data, leave the Censoring field blank.
- Frequency - Enter a vector of positive integers of the same size as the data vector to specify the frequency of the corresponding entries of the data vector. For example, a value of 7 in the 15 th entry of frequency vector specifies that there are 7 data points corresponding to the value in the 15 th entry of the data vector. If all entries of the data vector have frequency 1 , leave the Frequency field blank.
- Data set name - Enter a name for the data set you import from the workspace, such as My data.

After you have entered the information in the preceding fields, click Create Data Set to create the data set My data.

Managing Data Sets. The Manage data sets pane enables you to view and manage the data sets you create. When you create a data set, its name appears in the Data sets list. The following figure shows the Manage data sets pane after creating the data set My data.


For each data set in the Data sets list, you can:

- Select the Plot check box to display a plot of the data in the main Distribution Fitting Tool window. When you create a new data set, Plot is selected by default. Clearing the Plot check box removes the data from the plot in the main window. You can specify the type of plot displayed in the Display type field in the main window.
- If Plot is selected, you can also select Bounds to display confidence interval bounds for the plot in the main window. These bounds are pointwise confidence bounds around the empirical estimates of these functions. The bounds are only displayed when you set Display Type in the main window to one of the following:
- Cumulative probability (CDF)
- Survivor function
- Cumulative hazard

The Distribution Fitting Tool cannot display confidence bounds on density (PDF), quantile (inverse CDF), or probability plots. Clearing the Bounds check box removes the confidence bounds from the plot in the main window.

When you select a data set from the list, the following buttons are enabled:

- View - Display the data in a table in a new window.
- Set Bin Rules - Defines the histogram bins used in a density (PDF) plot.
- Rename - Rename the data set.
- Delete - Delete the data set.

Setting Bin Rules. To set bin rules for the histogram of a data set, click Set Bin Rules. This opens the Set Bin Width Rules dialog box.


You can select from the following rules:

- Freedman-Diaconis rule - Algorithm that chooses bin widths and locations automatically, based on the sample size and the spread of the data. This rule, which is the default, is suitable for many kinds of data.
- Scott rule - Algorithm intended for data that are approximately normal. The algorithm chooses bin widths and locations automatically.
- Number of bins - Enter the number of bins. All bins have equal widths.
- Bins centered on integers - Specifies bins centered on integers.
- Bin width — Enter the width of each bin. If you select this option, you can also select:
- Automatic bin placement - Place the edges of the bins at integer multiples of the Bin width.
- Bin boundary at - Enter a scalar to specify the boundaries of the bins. The boundary of each bin is equal to this scalar plus an integer multiple of the Bin width.

The Set Bin Width Rules dialog box also provides the following options:

- Apply to all existing data sets - Apply the rule to all data sets. Otherwise, the rule is only applied to the data set currently selected in the Data dialog box.
- Save as default - Apply the current rule to any new data sets that you create. You can also set default bin width rules by selecting Set Default Bin Rules from the Tools menu in the main window.


## Creating a New Fit

This section describes how to create a new fit. To begin, click the New Fit button at the top of the main window to open the New Fit dialog box. If you created the data set My data, it appears in the Data field.


| Field Name | Description |
| :--- | :--- |
| Distribution | Select the type of distribution to fit from the Distribution <br> drop-down list. See "Available Distributions" on page 5-22 <br> for a list of distributions supported by the Distribution <br> Fitting Tool. <br> Only the distributions that apply to the values of the <br> selected data set appear in the Distribution field. For <br> example, positive distributions are not displayed when the <br> data include values that are zero or negative. <br> You can specify either a parametric or a nonparametric <br> distribution. When you select a parametric distribution <br> from the drop-down list, a description of its parameters <br> appears in the Normal pane. The Distribution Fitting Tool <br> estimates these parameters to fit the distribution to the <br> data set. When you select Nonparametric fit, options for <br> the fit appear in the pane, as described in "Further Options <br> for Nonparametric Fits" on page 5-23. |
| Exclusion <br> rule | Specify a rule to exclude some data in the Exclusion rule <br> field. Create an exclusion rule by clicking Exclude in <br> the Distribution Fitting Tool. For more information, see <br> "Excluding Data" on page 5-32. |

Apply the New Fit. Click Apply to fit the distribution. For a parametric fit, the Results pane displays the values of the estimated parameters. For a nonparametric fit, the Results pane displays information about the fit.

When you click Apply, the Distribution Fitting Tool displays a plot of the distribution, along with the corresponding data.

Note When you click Apply, the title of the dialog box changes to Edit Fit. You can now make changes to the fit you just created and click Apply again to save them. After closing the Edit Fit dialog box, you can reopen it from the Fit Manager dialog box at any time to edit the fit.

After applying the fit, you can save the information to the workspace using probability distribution objects by clicking Save to workspace. See "Using Probability Distribution Objects" on page 5-84 for more information.

Available Distributions. This section lists the distributions available in the Distribution Fitting Tool.

Most, but not all, of the distributions available in the Distribution Fitting Tool are supported elsewhere in Statistics Toolbox software (see "Supported Distributions" on page 5-3), and have dedicated distribution fitting functions. These functions compute the majority of the fits in the Distribution Fitting Tool, and are referenced in the list below.

Other fits are computed using functions internal to the Distribution Fitting Tool. Distributions that do not have corresponding Statistics Toolbox fitting functions are described in "Additional Distributions Available in the Distribution Fitting Tool" on page 5-49.

Not all of the distributions listed below are available for all data sets. The Distribution Fitting Tool determines the extent of the data (nonnegative, unit interval, etc.) and displays appropriate distributions in the Distribution drop-down list. Distribution data ranges are given parenthetically in the list below.

- Beta (unit interval values) distribution, fit using the function betafit.
- Binomial (nonnegative values) distribution, fit using the function binopdf.
- Birnbaum-Saunders (positive values) distribution.
- Exponential (nonnegative values) distribution, fit using the function expfit.
- Extreme value (all values) distribution, fit using the function evfit.
- Gamma (positive values) distribution, fit using the function gamfit.
- Generalized extreme value (all values) distribution, fit using the function gevfit.
- Generalized Pareto (all values) distribution, fit using the function gpfit.
- Inverse Gaussian (positive values) distribution.
- Logistic (all values) distribution.
- Loglogistic (positive values) distribution.
- Lognormal (positive values) distribution, fit using the function lognfit.
- Nakagami (positive values) distribution.
- Negative binomial (nonnegative values) distribution, fit using the function nbinpdf.
- Nonparametric (all values) distribution, fit using the function ksdensity. See "Further Options for Nonparametric Fits" on page 5-23 for a description of available options.
- Normal (all values) distribution, fit using the function normfit.
- Poisson (nonnegative integer values) distribution, fit using the function poisspdf.
- Rayleigh (positive values) distribution using the function raylfit.
- Rician (positive values) distribution.
- $t$ location-scale (all values) distribution.
- Weibull (positive values) distribution using the function wblfit.

Further Options for Nonparametric Fits. When you select Non-parametric in the Distribution field, a set of options appears in the Non-parametric pane, as shown in the following figure.


The options for nonparametric distributions are:

- Kernel - Type of kernel function to use.
- Normal
- Box
- Triangle
- Epanechnikov
- Bandwidth - The bandwidth of the kernel smoothing window. Select auto for a default value that is optimal for estimating normal densities. This value appears in the Fit results pane after you click Apply. Select specify and enter a smaller value to reveal features such as multiple modes or a larger value to make the fit smoother.
- Domain - The allowed $x$-values for the density.
- unbounded - The density extends over the whole real line.
- positive - The density is restricted to positive values.
- specify - Enter lower and upper bounds for the domain of the density.

When you select positive or specify, the nonparametric fit has zero probability outside the specified domain.

## Displaying Results

This section explains the different ways to display results in the Distribution Fitting Tool window. This window displays plots of:

- The data sets for which you select Plot in the Data dialog box
- The fits for which you select Plot in the Fit Manager dialog box
- Confidence bounds for:
- Data sets for which you select Bounds in the Data dialog box
- Fits for which you select Bounds in the Fit Manager dialog box

The following fields are available.

Display Type. The Display Type field in the main window specifies the type of plot displayed. Each type corresponds to a probability function, for example, a probability density function. The following display types are available:

- Density (PDF) — Display a probability density function (PDF) plot for the fitted distribution. The main window displays data sets using a probability histogram, in which the height of each rectangle is the fraction of data points that lie in the bin divided by the width of the bin. This makes the sum of the areas of the rectangles equal to 1 .
- Cumulative probability (CDF) - Display a cumulative probability plot of the data. The main window displays data sets using a cumulative probability step function. The height of each step is the cumulative sum of the heights of the rectangles in the probability histogram.
- Quantile (inverse CDF) - Display a quantile (inverse CDF) plot.
- Probability plot - Display a probability plot of the data. You can specify the type of distribution used to construct the probability plot in the Distribution field, which is only available when you select Probability plot. The choices for the distribution are:
- Exponential
- Extreme value
- Logistic
- Log-Logistic
- Lognormal
- Normal
- Rayleigh
- Weibull

In addition to these choices, you can create a probability plot against a parametric fit that you create in the New Fit pane. These fits are added at the bottom of the Distribution drop-down list when you create them.

- Survivor function - Display survivor function plot of the data.
- Cumulative hazard - Display cumulative hazard plot of the data.

Note Some distributions are unavailable if the plotted data includes 0 or negative values.

Confidence Bounds. You can display confidence bounds for data sets and fits when you set Display Type to Cumulative probability (CDF), Survivor function, Cumulative hazard, or, for fits only, Quantile (inverse CDF).

- To display bounds for a data set, select Bounds next to the data set in the Data sets pane of the Data dialog box.
- To display bounds for a fit, select Bounds next to the fit in the Fit Manager dialog box. Confidence bounds are not available for all fit types.

To set the confidence level for the bounds, select Confidence Level from the View menu in the main window and choose from the options.

## Managing Fits

This section describes how to manage fits that you have created. To begin, click the Manage Fits button in the Distribution Fitting Tool. This opens the Fit Manager dialog box as shown in the following figure.


The Table of fits displays a list of the fits you create, with the following options:

- Plot - Select Plot to display a plot of the fit in the main window of the Distribution Fitting Tool. When you create a new fit, Plot is selected by default. Clearing the Plot check box removes the fit from the plot in the main window.
- Bounds - If Plot is selected, you can also select Bounds to display confidence bounds in the plot. The bounds are displayed when you set Display Type in the main window to one of the following:
- Cumulative probability (CDF)
- Quantile (inverse CDF)
- Survivor function
- Cumulative hazard

The Distribution Fitting Tool cannot display confidence bounds on density (PDF) or probability plots. In addition, bounds are not supported for nonparametric fits and some parametric fits.

Clearing the Bounds check box removes the confidence intervals from the plot in the main window.

When you select a fit in the Table of fits, the following buttons are enabled below the table:

- New Fit - Open a New Fit window.
- Copy - Create a copy of the selected fit.
- Edit - Open an Edit Fit dialog box, where you can edit the fit.

Note You can only edit the currently selected fit in the Edit Fit dialog box. To edit a different fit, select it in the Table of fits and click Edit to open another Edit Fit dialog box.

- Save to workspace - Save the selected fit as a distribution object. See "Using Probability Distribution Objects" on page 5-84 for more information.
- Delete - Delete the selected fit.


## Evaluating Fits

The Evaluate dialog box enables you to evaluate any fit at whatever points you choose. To open the dialog box, click the Evaluate button in the Distribution Fitting Tool. The following figure shows the Evaluate dialog box.


The Evaluate dialog box contains the following items:

- Fit pane - Display the names of existing fits. Select one or more fits that you want to evaluate. Using your platform specific functionality, you can select multiple fits.
- Function - Select the type of probability function you want to evaluate for the fit. The available functions are
- Density (PDF) - Computes a probability density function.
- Cumulative probability (CDF) - Computes a cumulative probability function.
- Quantile (inverse CDF) - Computes a quantile (inverse CDF) function.
- Survivor function - Computes a survivor function.
- Cumulative hazard - Computes a cumulative hazard function.
- Hazard rate - Computes the hazard rate.
- At $\mathbf{x}=-$ Enter a vector of points or the name of a workspace variable containing a vector of points at which you want to evaluate the distribution function. If you change Function to Quantile (inverse CDF), the field name changes to At $\mathbf{p}=$ and you enter a vector of probability values.
- Compute confidence bounds - Select this box to compute confidence bounds for the selected fits. The check box is only enabled if you set Function to one of the following:
- Cumulative probability (CDF)
- Quantile (inverse CDF)
- Survivor function
- Cumulative hazard

The Distribution Fitting Tool cannot compute confidence bounds for nonparametric fits and for some parametric fits. In these cases, the tool returns NaN for the bounds.

- Level - Set the level for the confidence bounds.
- Plot function - Select this box to display a plot of the distribution function, evaluated at the points you enter in the $\mathbf{A t} \mathbf{x}=$ field, in a new window.

Note The settings for Compute confidence bounds, Level, and Plot function do not affect the plots that are displayed in the main window of the Distribution Fitting Tool. The settings only apply to plots you create by clicking Plot function in the Evaluate window.

Click Apply to apply these settings to the selected fit. The following figure shows the results of evaluating the cumulative density function for the fit My fit, created in "Example: Fitting a Distribution" on page 5-39, at the points in the vector -3:0.5:3.


The window displays the following values in the columns of the table to the right of the Fit pane:

- X - The entries of the vector you enter in At $\mathbf{x}=$ field
- Y - The corresponding values of the CDF at the entries of X
- LB - The lower bounds for the confidence interval, if you select Compute confidence bounds
- UB - The upper bounds for the confidence interval, if you select Compute confidence bounds

To save the data displayed in the Evaluate window, click Export to Workspace. This saves the values in the table to a matrix in the MATLAB workspace.

## Excluding Data

To exclude values from fit, click the Exclude button in the main window of the Distribution Fitting Tool. This opens the Exclude window, in which you can create rules for excluding specified values. You can use these rules to exclude data when you create a new fit in the New Fit window. The following figure shows the Exclude window.


To create an exclusion rule:

1 Exclusion Rule Name-Enter a name for the exclusion rule in the Exclusion rule name field.

2 Exclude Sections-In the Exclude sections pane, you can specify bounds for the excluded data:

- In the Lower limit: exclude $\mathbf{Y}$ drop-down list, select <= or < from the drop-down list and enter a scalar in the field to the right. This excludes values that are either less than or equal to or less than that scalar, respectively.
- In the Upper limit: exclude $\mathbf{Y}$ drop-down list, select >= or > from the drop-down list and enter a scalar in the field to the right to exclude values that are either greater than or equal to or greater than the scalar, respectively.


## OR

Exclude Graphically-The Exclude Graphically button enables you to define the exclusion rule by displaying a plot of the values in a data set and selecting the bounds for the excluded data with the mouse. For example, if you created the data set My data, described in "Creating and Managing Data Sets" on page 5-14, select it from the drop-down list next to Exclude graphically and then click the Exclude graphically button. This displays the values in My data in a new window as shown in the following figure.


To set a lower limit for the boundary of the excluded region, click Add Lower Limit. This displays a vertical line on the left side of the plot window. Move the line with the mouse to the point you where you want the lower limit, as shown in the following figure.


Moving the vertical line changes the value displayed in the Lower limit: exclude data field in the Exclude window, as shown in the following figure.


The value displayed corresponds to the $x$-coordinate of the vertical line.
Similarly, you can set the upper limit for the boundary of the excluded region by clicking Add Upper Limit and moving the vertical line that appears at the right side of the plot window. After setting the lower and upper limits, click Close and return to the Exclude window.

3 Create Exclusion Rule-Once you have set the lower and upper limits for the boundary of the excluded data, click Create Exclusion Rule to create the new rule. The name of the new rule now appears in the Existing exclusion rules pane.

When you select an exclusion rule in the Existing exclusion rules pane, the following buttons are enabled:

- Copy - Creates a copy of the rule, which you can then modify. To save the modified rule under a different name, click Create Exclusion Rule.
- View - Opens a new window in which you can see which data points are excluded by the rule. The following figure shows a typical example.


The shaded areas in the plot graphically display which data points are excluded. The table to the right lists all data points. The shaded rows indicate excluded points:

- Rename - Renames the rule
- Delete - Deletes the rule

Once you define an exclusion rule, you can use it when you fit a distribution to your data. The rule does not exclude points from the display of the data set.

## Saving and Loading Sessions

This section explains how to save your work in the current Distribution Fitting Tool session and then load it in a subsequent session, so that you can continue working where you left off.

Saving a Session. To save the current session, select Save Session from the File menu in the main window. This opens a dialog box that prompts you to enter a filename, such as my_session.dfit, for the session. Clicking Save saves the following items created in the current session:

- Data sets
- Fits
- Exclusion rules
- Plot settings
- Bin width rules

Loading a Session. To load a previously saved session, select Load Session from the File menu in the main window and enter the name of a previously saved session. Clicking Open restores the information from the saved session to the current session of the Distribution Fitting Tool.

## Example: Fitting a Distribution

This section presents an example that illustrates how to use the Distribution Fitting Tool. The example involves the following steps:

- "Step 1: Generate Random Data" on page 5-39
- "Step 2: Import Data" on page 5-39
- "Step 3: Create a New Fit" on page 5-42

Step 1: Generate Random Data. To try the example, first generate some random data to which you will fit a distribution. The following command generates a vector data, of length 100, whose entries are random numbers from a normal distribution with mean. 36 and standard deviation 1.4.

```
data = normrnd(.36, 1.4, 100, 1);
```

Step 2: Import Data. Open the distribution fitting tool:

```
dfittool
```

To import the vector data into the Distribution Fitting Tool, click the Data button in main window. This opens the window shown in the following figure.


The Data field displays all numeric arrays in the MATLAB workspace. Select data from the drop-down list, as shown in the following figure.

| Data: | lnone) |
| :--- | :--- |
| Censoring: | (none) |
|  | data |

This displays a histogram of the data in the Data preview pane.
In the Data set name field, type a name for the data set, such as My data, and click Create Data Set to create the data set. The main window of the Distribution Fitting Tool now displays a larger version of the histogram in the Data preview pane, as shown in the following figure.


Note Because the example uses random data, you might see a slightly different histogram if you try this example for yourself.

Step 3: Create a New Fit. To fit a distribution to the data, click New Fit in the main window of the Distribution Fitting Tool. This opens the window shown in the following figure.


To fit a normal distribution, the default entry of the Distribution field, to My data:

1 Enter a name for the fit, such as My fit, in the Fit name field.
2 Select My data from the drop-down list in the Data field.

## 3 Click Apply.

The Results pane displays the mean and standard deviation of the normal distribution that best fits My data, as shown in the following figure.


The main window of the Distribution Fitting Tool displays a plot of the normal distribution with this mean and standard deviation, as shown in the following figure.


## Generating a File to Fit and Plot Distributions

The Generate Code option in the File menu enables you to create a file that

- Fits the distributions used in the current session to any data vector in the MATLAB workspace.
- Plots the data and the fits.

After you end the current session, you can use the file to create plots in a standard MATLAB figure window, without having to reopen the Distribution Fitting Tool.

As an example, assuming you created the fit described in "Creating a New Fit" on page $5-19$, do the following steps:

1 Select Generate Code from the File menu.
2 Choose File > Save as in the MATLAB Editor window. Save the file as normal_fit.m in a folder on the MATLAB path.

You can then apply the function normal_fit to any vector of data in the MATLAB workspace. For example, the following commands

```
new_data = normrnd(4.1, 12.5, 100, 1);
newfit = normal_fit(new_data)
legend('New Data', 'My fit')
```

generate newfit, a fitted normal distribution of the data, and generates a plot of the data and the fit.

```
newfit =
normal distribution
    mu = 3.19148
    sigma = 12.5631
```



Note By default, the file labels the data in the legend using the same name as the data set in the Distribution Fitting Tool. You can change the label using the legend command, as illustrated by the preceding example.

## Using Custom Distributions

This section explains how to use custom distributions with the Distribution Fitting Tool.

Defining Custom Distributions. To define a custom distribution, select Define Custom Distribution from the File menu. This opens a file template in the MATLAB editor. You then edit this file so that it computes the distribution you want.

The template includes example code that computes the Laplace distribution, beginning at the lines

```
%
% Remove the following return statement to define the
% Laplace distributon
%
return
```

To use this example, simply delete the command return and save the file. If you save the template in a folder on the MATLAB path, under its default name dfittooldists.m, the Distribution Fitting Tool reads it in automatically when you start the tool. You can also save the template under a different name, such as laplace.m, and then import the custom distribution as described in the following section.

Importing Custom Distributions. To import a custom distribution, select Import Custom Distributions from the File menu. This opens a dialog box in which you can select the file that defines the distribution. For example, if you created the file laplace.m, as described in the preceding section, you can enter laplace.m and select Open in the dialog box. The Distribution field of the New Fit window now contains the option Laplace.

## Additional Distributions Available in the Distribution Fitting Tool

The following distributions are available in the Distribution Fitting Tool, but do not have dedicated distribution functions as described in "Statistics Toolbox Distribution Functions" on page 5-52. The distributions can be used with the functions pdf, cdf, icdf, and mle in a limited capacity. See the reference pages for these functions for details on the limitations.

- "Birnbaum-Saunders Distribution" on page B-10
- "Inverse Gaussian Distribution" on page B-45
- "Loglogistic Distribution" on page B-50
- "Logistic Distribution" on page B-49
- "Nakagami Distribution" on page B-70
- "Rician Distribution" on page B-93
- "t Location-Scale Distribution" on page B-97

For a complete list of the distributions available for use with the Distribution Fitting Tool, see "Supported Distributions" on page 5-3. Distributions listing dfittool in the fit column of the tables in that section can be used with the Distribution Fitting Tool.

## Visually Exploring Random Number Generation

The Random Number Generation Tool is a graphical user interface that generates random samples from specified probability distributions and displays the samples as histograms. Use the tool to explore the effects of changing parameters and sample size on the distributions.

Run the tool by typing randtool at the command line.


Start by selecting a distribution, then enter the desired sample size.

## You can also

- Use the controls at the bottom of the window to set parameter values for the distribution and to change their upper and lower bounds.
- Draw another sample from the same distribution, with the same size and parameters.
- Export the current sample to your workspace. A dialog box enables you to provide a name for the sample.


## Statistics Toolbox Distribution Functions

In this section...<br>"Probability Density Functions" on page 5-52<br>"Cumulative Distribution Functions" on page 5-62<br>"Inverse Cumulative Distribution Functions" on page 5-66<br>"Distribution Statistics Functions" on page 5-68<br>"Distribution Fitting Functions" on page 5-70<br>"Negative Log-Likelihood Functions" on page 5-77<br>"Random Number Generators" on page 5-80

For each distribution supported by Statistics Toolbox software, a selection of the distribution functions described in this section is available for statistical programming. This section gives a general overview of the use of each type of function, independent of the particular distribution. For specific functions available for specific distributions, see "Supported Distributions" on page 5-3.

## Probability Density Functions

- "Estimating PDFs with Parameters" on page 5-52
- "Estimating PDFs without Parameters" on page 5-55


## Estimating PDFs with Parameters

Probability density functions (pdfs) for supported Statistics Toolbox distributions all end with pdf, as in binopdf or exppdf. For more information on specific function names for specific distributions see "Supported Distributions" on page 5-3.

Each function represents a parametric family of distributions. Input arguments are arrays of outcomes followed by a list of parameter values specifying a particular member of the distribution family.

For discrete distributions, the pdf assigns a probability to each outcome. In this context, the pdf is often called a probability mass function (pmf).

For example, the discrete binomial pdf

$$
f(k)=\binom{n}{k} p^{k}(1-p)^{n-k}
$$

assigns probability to the event of $k$ successes in $n$ trials of a Bernoulli process (such as coin flipping) with probability $p$ of success at each trial. Each of the integers $k=0,1,2, \ldots, n$ is assigned a positive probability, with the sum of the probabilities equal to 1 . Compute the probabilities with the binopdf function:

```
p = 0.2; % Probability of success for each trial
n = 10; % Number of trials
k = 0:n; % Outcomes
m = binopdf(k,n,p); % Probability mass vector
bar(k,m) % Visualize the probability distribution
set(get(gca,'Children'),'FaceColor',[.8 . 8 1])
grid on
```



For continuous distributions, the pdf assigns a probability density to each outcome. The probability of any single outcome is zero. The pdf must be integrated over a set of outcomes to compute the probability that an outcome falls within that set. The integral over the entire set of outcomes is 1 .

For example, the continuous exponential pdf

$$
f(t)=\lambda e^{-\lambda t}
$$

is used to model the probability that a process with constant failure rate $\lambda$ will have a failure within time $t$. Each time $t>0$ is assigned a positive probability density. Densities are computed with the exppdf function:

```
lambda = 2; % Failure rate
t = 0:0.01:3; % Outcomes
f = exppdf(t,1/lambda); % Probability density vector
plot(t,f) % Visualize the probability distribution
grid on
```



Probabilities for continuous pdfs can be computed with the quad function. In the example above, the probability of failure in the time interval $[0,1]$ is computed as follows:

```
f_lambda = @(t)exppdf(t,1/lambda); % Pdf with fixed lambda
P = quad(f_lambda,0,1) % Integrate from 0 to 1
P =
    0.8647
```

Alternatively, the cumulative distribution function (cdf) for the exponential function, expcdf, can be used:

```
P = expcdf(1,1/lambda) % Cumulative probability from 0 to 1
P =
    0.8647
```


## Estimating PDFs without Parameters

A distribution of data can be described graphically with a histogram:

```
cars = load('carsmall','MPG','Origin');
MPG = cars.MPG;
hist(MPG)
set(get(gca,'Children'),'FaceColor',[.8 .8 1])
```



You can also describe a data distribution by estimating its density.
The ksdensity function does this using a kernel smoothing method. A nonparametric density estimate of the previous data, using the default kernel and bandwidth, is given by:

```
[f,x] = ksdensity(MPG);
plot(x,f);
title('Density estimate for MPG')
```



Controlling Probability Density Curve Smoothness. The choice of kernel bandwidth controls the smoothness of the probability density curve. The following graph shows the density estimate for the same mileage data using different bandwidths. The default bandwidth is in blue and looks like the preceding graph. Estimates for smaller and larger bandwidths are in red and green.

The first call to ksdensity returns the default bandwidth, $u$, of the kernel smoothing function. Subsequent calls modify this bandwidth.

```
[f,x,u] = ksdensity(MPG);
plot(x,f)
title('Density estimate for MPG')
hold on
```

```
[f,x] = ksdensity(MPG,'width',u/3);
plot(x,f,'r');
[f,x] = ksdensity(MPG,'width',u*3);
plot(x,f,'g');
legend('default width','1/3 default','3*default')
hold off
```

Density estimate for MPG


The default bandwidth seems to be doing a good job-reasonably smooth, but not so smooth as to obscure features of the data. This bandwidth is the one that is theoretically optimal for estimating densities for the normal distribution.

The green curve shows a density with the kernel bandwidth set too high. This curve smooths out the data so much that the end result looks just like the kernel function. The red curve has a smaller bandwidth and is rougher looking than the blue curve. It may be too rough, but it does provide an indication that there might be two major peaks rather than the single peak of the blue curve. A reasonable choice of width might lead to a curve that is intermediate between the red and blue curves.

Specifying Kernel Smoothing Functions. You can also specify a kernel function by supplying either the function name or a function handle. The four preselected functions, 'normal', 'epanechnikov', 'box', and 'triangle', are all scaled to have standard deviation equal to 1 , so they perform a comparable degree of smoothing.

Using default bandwidths, you can now plot the same mileage data, using each of the available kernel functions.

```
hname = {'normal' 'epanechnikov' 'box' 'triangle'};
colors = {'r' 'b' 'g' 'm'};
for j=1:4
    [f,x] = ksdensity(MPG,'kernel',hname{j});
    plot(x,f,colors{j});
    hold on;
end
legend(hname{:});
hold off
```



The density estimates are roughly comparable, but the box kernel produces a density that is rougher than the others.

Comparing Density Estimates. While it is difficult to overlay two histograms to compare them, you can easily overlay smooth density estimates. For example, the following graph shows the MPG distributions for cars from different countries of origin:

```
Origin = cellstr(cars.Origin);
I = strcmp('USA',Origin);
J = strcmp('Japan',Origin);
K = ~(I|J);
MPG_USA = MPG(I);
MPG_Japan = MPG(J);
MPG_Europe = MPG(K);
```



For piecewise probability density estimation, using kernel smoothing in the center of the distribution and Pareto distributions in the tails, see "Fitting Piecewise Distributions" on page 5-72.

## Cumulative Distribution Functions

- "Estimating Parametric CDFs" on page 5-62
- "Estimating Empirical CDFs" on page 5-63


## Estimating Parametric CDFs

Cumulative distribution functions (cdfs) for supported Statistics Toolbox distributions all end with cdf, as in binocdf or expcdf. Specific function names for specific distributions can be found in "Supported Distributions" on page 5-3.

Each function represents a parametric family of distributions. Input arguments are arrays of outcomes followed by a list of parameter values specifying a particular member of the distribution family.

For discrete distributions, the $\operatorname{cdf} F$ is related to the $\operatorname{pdf} f$ by

$$
F(x)=\sum_{y \leq x} f(y)
$$

For continuous distributions, the $\operatorname{cdf} F$ is related to the $\operatorname{pdf} f$ by

$$
F(x)=\int_{-\infty}^{x} f(y) d y
$$

Cdfs are used to compute probabilities of events. In particular, if $F$ is a cdf and $x$ and $y$ are outcomes, then

- $P(y \leq x)=F(x)$
- $P(y>x)=1-F(x)$
- $P\left(x_{1}<y \leq x_{2}\right)=F\left(x_{2}\right)-F\left(x_{1}\right)$

For example, the $t$-statistic

$$
t=\frac{\bar{x}-\mu}{s / \sqrt{n}}
$$

follows a Student's $t$ distribution with $n-1$ degrees of freedom when computed from repeated random samples from a normal population with mean $\mu$. Here $\bar{x}$ is the sample mean, $s$ is the sample standard deviation, and $n$ is the sample size. The probability of observing a $t$-statistic greater than or equal to the value computed from a sample can be found with the tcdf function:

```
mu = 1; % Population mean
sigma = 2; % Population standard deviation
n = 100; % Sample size
x = normrnd(mu,sigma,n,1); % Random sample from population
xbar = mean(x); % Sample mean
s = std(x); % Sample standard deviation
t = (xbar-mu)/(s/sqrt(n)) % t-statistic
t =
    0.2489
p = 1-tcdf(t,n-1) % Probability of larger t-statistic
p =
    0.4020
```

This probability is the same as the $p$ value returned by a $t$-test of the null hypothesis that the sample comes from a normal population with mean $\mu$ :

```
[h,ptest] = ttest(x,mu,0.05,'right')
h =
    0
ptest =
    0.4020
```


## Estimating Empirical CDFs

The ksdensity function produces an empirical version of a probability density function (pdf). That is, instead of selecting a density with a particular parametric form and estimating the parameters, it produces a nonparametric density estimate that adapts itself to the data.

Similarly, it is possible to produce an empirical version of the cumulative distribution function (cdf). The ecdf function computes this empirical cdf. It
returns the values of a function $F$ such that $F(x)$ represents the proportion of observations in a sample less than or equal to $x$.

The idea behind the empirical cdf is simple. It is a function that assigns probability $1 / n$ to each of $n$ observations in a sample. Its graph has a stair-step appearance. If a sample comes from a distribution in a parametric family (such as a normal distribution), its empirical cdf is likely to resemble the parametric distribution. If not, its empirical distribution still gives an estimate of the cdf for the distribution that generated the data.

The following example generates 20 observations from a normal distribution with mean 10 and standard deviation 2 . You can use ecdf to calculate the empirical cdf and stairs to plot it. Then you overlay the normal distribution curve on the empirical function.

```
x = normrnd(10,2,20,1);
[f,xf] = ecdf(x);
stairs(xf,f)
hold on
xx=linspace(5,15,100);
yy = normcdf(xx,10,2);
plot(xx,yy,'r:')
hold off
legend('Empirical cdf','Normal cdf',2)
```



The empirical cdf is especially useful in survival analysis applications. In such applications the data may be censored, that is, not observed exactly. Some individuals may fail during a study, and you can observe their failure time exactly. Other individuals may drop out of the study, or may not fail until after the study is complete. The ecdf function has arguments for dealing with censored data. In addition, you can use the coxphfit function with individuals that have predictors that are not the same.

For piecewise probability density estimation, using the empirical cdf in the center of the distribution and Pareto distributions in the tails, see "Fitting Piecewise Distributions" on page 5-72.

## Inverse Cumulative Distribution Functions

Inverse cumulative distribution functions for supported Statistics Toolbox distributions all end with inv, as in binoinv or expinv. Specific function names for specific distributions can be found in "Supported Distributions" on page 5-3.

Each function represents a parametric family of distributions. Input arguments are arrays of cumulative probabilities from 0 to 1 followed by a list of parameter values specifying a particular member of the distribution family.

For continuous distributions, the inverse cdf returns the unique outcome whose cdf value is the input cumulative probability.

For example, the expinv function can be used to compute inverses of exponential cumulative probabilities:

```
x = 0.5:0.2:1.5 % Outcomes
x =
    0.5000}00.7000 0.9000 1.1000 1.3000 1.5000
p = expcdf(x,1) % Cumulative probabilities
p =
    0.3935 0.5034 0.5934 0.6671 0.7275 0.7769
expinv(p,1) % Return original outcomes
ans =
    0.5000 0.7000 0.9000 1.1000 1.3000 1.5000
```

For discrete distributions, there may be no outcome whose cdf value is the input cumulative probability. In these cases, the inverse cdf returns the first outcome whose cdf value equals or exceeds the input cumulative probability.

For example, the binoinv function can be used to compute inverses of binomial cumulative probabilities:

```
x = 0:5 % Some possible outcomes
p = binocdf(x,10,0.2) % Their cumulative probabilities
p =
    0.1074 0.3758 0.6778
q = [.1 .2 .3 .4] % New trial probabilities
```

```
q =
    0.1000 0.2000 0.3000 0.4000
binoinv(q,10,0.2) % Their corresponding outcomes
ans =
    0
```

The inverse cdf is useful in hypothesis testing, where critical outcomes of a test statistic are computed from cumulative significance probabilities. For example, norminv can be used to compute a $95 \%$ confidence interval under the assumption of normal variability:

```
p = [0.025 0.975]; % Interval containing 95% of [0,1]
x = norminv(p,0,1) % Assume standard normal variability
x =
    -1.9600 1.9600 % 95% confidence interval
n = 20; % Sample size
y = normrnd(8,1,n,1); % Random sample (assume mean is unknown)
ybar = mean(y);
ci = ybar + (1/sqrt(n))*x % Confidence interval for mean
ci =
    7.6779 8.5544
```


## Distribution Statistics Functions

Distribution statistics functions for supported Statistics Toolbox distributions all end with stat, as in binostat or expstat. Specific function names for specific distributions can be found in "Supported Distributions" on page 5-3.

Each function represents a parametric family of distributions. Input arguments are lists of parameter values specifying a particular member of the distribution family. Functions return the mean and variance of the distribution, as a function of the parameters.

For example, the wblstat function can be used to visualize the mean of the Weibull distribution as a function of its two distribution parameters:

```
a = 0.5:0.1:3;
b = 0.5:0.1:3;
[A,B] = meshgrid(a,b);
M = wblstat(A,B);
surfc(A,B,M)
```



## Distribution Fitting Functions

- "Fitting Regular Distributions" on page 5-70
- "Fitting Piecewise Distributions" on page 5-72


## Fitting Regular Distributions

Distribution fitting functions for supported Statistics Toolbox distributions all end with fit, as in binofit or expfit. Specific function names for specific distributions can be found in "Supported Distributions" on page 5-3.

Each function represents a parametric family of distributions. Input arguments are arrays of data, presumed to be samples from some member of the selected distribution family. Functions return maximum likelihood estimates (MLEs) of distribution parameters, that is, parameters for the distribution family member with the maximum likelihood of producing the data as a random sample.

The Statistics Toolbox function mle is a convenient front end to the individual distribution fitting functions, and more. The function computes MLEs for distributions beyond those for which Statistics Toolbox software provides specific pdf functions.

For some pdfs, MLEs can be given in closed form and computed directly. For other pdfs, a search for the maximum likelihood must be employed. The search can be controlled with an options input argument, created using the statset function. For efficient searches, it is important to choose a reasonable distribution model and set appropriate convergence tolerances.

MLEs can be heavily biased, especially for small samples. As sample size increases, however, MLEs become unbiased minimum variance estimators with approximate normal distributions. This is used to compute confidence bounds for the estimates.

For example, consider the following distribution of means from repeated random samples of an exponential distribution:

```
mu = 1; % Population parameter
n = 1e3; % Sample size
ns = 1e4; % Number of samples
```

```
samples = exprnd(mu,n,ns); % Population samples
means = mean(samples); % Sample means
```

The Central Limit Theorem says that the means will be approximately normally distributed, regardless of the distribution of the data in the samples. The normfit function can be used to find the normal distribution that best fits the means:

```
[muhat,sigmahat,muci,sigmaci] = normfit(means)
muhat =
    1.0003
sigmahat =
    0.0319
muci =
    0.9997
    1.0010
sigmaci =
    0.0314
    0.0323
```

The function returns MLEs for the mean and standard deviation and their $95 \%$ confidence intervals.

To visualize the distribution of sample means together with the fitted normal distribution, you must scale the fitted pdf, with area $=1$, to the area of the histogram being used to display the means:

```
numbins = 50;
hist(means,numbins)
set(get(gca,'Children'),'FaceColor',[.8 .8 1])
hold on
[bincounts,binpositions] = hist(means,numbins);
binwidth = binpositions(2) - binpositions(1);
histarea = binwidth*sum(bincounts);
x = binpositions(1):0.001:binpositions(end);
y = normpdf(x,muhat,sigmahat);
plot(x,histarea*y,'r','LineWidth',2)
```



## Fitting Piecewise Distributions

The parametric methods discussed in "Fitting Regular Distributions" on page 5-70 fit data samples with smooth distributions that have a relatively low-dimensional set of parameters controlling their shape. These methods work well in many cases, but there is no guarantee that a given sample will be described accurately by any of the supported Statistics Toolbox distributions.

The empirical distributions computed by ecdf and discussed in "Estimating Empirical CDFs" on page 5-63 assign equal probability to each observation in a sample, providing an exact match of the sample distribution. However, the distributions are not smooth, especially in the tails where data may be sparse.

The paretotails function fits a distribution by piecing together the empirical distribution in the center of the sample with smooth generalized Pareto distributions (GPDs) in the tails. The output is an object of the paretotails class, with associated methods to evaluate the cdf, inverse cdf, and other functions of the fitted distribution.

As an example, consider the following data, with about $20 \%$ outliers:

```
left_tail = - exprnd(1,10,1);
right_tail = exprnd(5,10,1);
center = randn(80,1);
data = [left_tail;center;right_tail];
```

Neither a normal distribution nor a $t$ distribution fits the tails very well:

```
probplot(data);
p = fitdist(data,'tlocationscale');
h = probplot(gca,p);
set(h,'color','r','linestyle','-')
title('{\bf Probability Plot}')
legend('Normal','Data','t','Location','NW')
```



On the other hand, the empirical distribution provides a perfect fit, but the outliers make the tails very discrete:

```
ecdf(data)
```



Random samples generated from this distribution by inversion might include, for example, values around 4.33 and 9.25 , but nothing in-between.

The paretotails function provides a single, well-fit model for the entire sample. The following uses generalized Pareto distributions (GPDs) for the lower and upper $10 \%$ of the data:

```
pfit = paretotails(data,0.1,0.9)
pfit =
Piecewise distribution with 3 segments
    -Inf < x < -1.30726 (0< p < 0.1)
        lower tail, GPD(-1.10167,1.12395)
    -1.30726<x<1.27213 (0.1< p < 0.9)
        interpolated empirical cdf
    1.27213<x< Inf (0.9<p<1)
```



Access information about the fit using the methods of the paretotails class. Options allow for nonparametric estimation of the center of the cdf.

## Negative Log-Likelihood Functions

Negative log-likelihood functions for supported Statistics Toolbox distributions all end with like, as in explike. Specific function names for specific distributions can be found in "Supported Distributions" on page 5-3.

Each function represents a parametric family of distributions. Input arguments are lists of parameter values specifying a particular member of the distribution family followed by an array of data. Functions return the negative log-likelihood of the parameters, given the data.

Negative log-likelihood functions are used as objective functions in search algorithms such as the one implemented by the MATLAB function fminsearch. Additional search algorithms are implemented by Optimization Toolbox ${ }^{\text {TM }}$ functions and Global Optimization Toolbox functions.

When used to compute maximum likelihood estimates (MLEs), negative log-likelihood functions allow you to choose a search algorithm and exercise low-level control over algorithm execution. By contrast, the functions discussed in "Distribution Fitting Functions" on page 5-70 use preset algorithms with options limited to those set by the statset function.

Likelihoods are conditional probability densities. A parametric family of distributions is specified by its $\operatorname{pdf} f(x, a)$, where $x$ and $a$ represent the variables and parameters, respectively. When $a$ is fixed, the pdf is used to compute the density at $x, f(x \mid a)$. When $x$ is fixed, the pdf is used to compute the likelihood of the parameters $a, f(a \mid x)$. The joint likelihood of the parameters over an independent random sample $X$ is

$$
L(a)=\prod_{x \in X} f(a \mid x)
$$

Given $X$, MLEs maximize $L(\alpha)$ over all possible $\alpha$.
In numerical algorithms, the log-likelihood function, $\log (L(a))$, is (equivalently) optimized. The logarithm transforms the product of potentially small likelihoods into a sum of logs, which is easier to distinguish from 0 in computation. For convenience, Statistics Toolbox negative log-likelihood functions return the negative of this sum, since the optimization algorithms to which the values are passed typically search for minima rather than maxima.

For example, use gamrnd to generate a random sample from a specific gamma distribution:

```
a = [1,2];
X = gamrnd(a(1),a(2),1e3,1);
```

Given X, the gamlike function can be used to visualize the likelihood surface in the neighborhood of a:

```
mesh = 50;
delta = 0.5;
a1 = linspace(a(1)-delta,a(1)+delta,mesh);
a2 = linspace(a(2)-delta,a(2)+delta,mesh);
logL = zeros(mesh); % Preallocate memory
for i = 1:mesh
        for j = 1:mesh
        logL(i,j) = gamlike([a1(i),a2(j)],X);
        end
end
[A1,A2] = meshgrid(a1,a2);
surfc(A1,A2,logL)
```



The MATLAB function fminsearch is used to search for the minimum of the likelihood surface:

```
LL = @(u)gamlike([u(1),u(2)],X); % Likelihood given X
MLES = fminsearch(LL,[1,2])
MLES =
    1.0231 1.9729
```

These can be compared to the MLEs returned by the gamfit function, which uses a combination search and solve algorithm:

```
ahat = gamfit(X)
ahat =
    1.0231 1.9728
```

The MLEs can be added to the surface plot (rotated to show the minimum):

```
    hold on
plot3(MLES(1),MLES(2),LL(MLES),...
    'ro','MarkerSize',5,...
        'MarkerFaceColor','r')
```



## Random Number Generators

The Statistics Toolbox supports the generation of random numbers from various distributions. Each RNG represents a parametric family of distributions. RNGs return random numbers from the specified distribution in an array of the specified dimensions. Specific RNG names for specific distributions are in "Supported Distributions" on page 5-3.

Other random number generation functions which do not support specific distributions include:

- cvpartition
- datasample
- hmmgenerate
- lhsdesign
- lhsnorm
- mhsample
- random
- randsample
- slicesample

RNGs in Statistics Toolbox software depend on MATLAB's default random number stream via the rand and randn functions, each RNG uses one of the techniques discussed in "Common Generation Methods" on page 6-5 to generate random numbers from a given distribution.

By controlling the default random number stream and its state, you can control how the RNGs in Statistics Toolbox software generate random values. For example, to reproduce the same sequence of values from an RNG, you can save and restore the default stream's state, or reset the default stream. For details on managing the default random number stream, see "Managing the Global Stream".

MATLAB initializes the default random number stream to the same state each time it starts up. Thus, RNGs in Statistics Toolbox software will generate the same sequence of values for each MATLAB session unless you
modify that state at startup. One simple way to do that is to add commands to startup.m such as
rng('shuffle');
that initialize MATLAB's default random number stream to a different state for each session.

## Dependencies of the Random Number Generators

The following table lists the dependencies of Statistics Toolbox RNGs on the MATLAB base RNGs rand, randi, and/or randn.

| RNG | MATLAB Base RNG |
| :---: | :---: |
| betarnd | rand, randn |
| binornd | rand |
| chi2rnd | rand, randn |
| evrnd | rand |
| exprnd | rand |
| datasample | rand, randi, randperm |
| frnd | rand, randn |
| gamrnd | rand, randn |
| geornd | rand |
| gevrnd | rand |
| gprnd | rand |
| hygernd | rand |
| iwishrnd | rand, randn |
| johnsrnd | randn |
| lhsdesign | rand |
| lhsnorm | rand |
| lognrnd | randn |
| mhsample | rand or randn, depending on the RNG given for the proposal distribution |
| mvnrnd | randn |
| mvtrnd | rand, randn |


|  |  |
| :--- | :--- |
| RNG | MATLAB Base RNG |
| nbinrnd | rand, randn |
| ncfrnd | rand, randn |
| nctrnd | rand, randn |
| ncx2rnd | randn |
| normrnd | randn |
| pearsrnd | distribution type |
| poissrnd | rand, randn or randn, depending on the |
| random | rand |
| randsample | randn |
| raylrnd | rand |
| slicesample | rand, randn |
| trnd | rand |
| unidrnd | rand |
| unifrnd | rand |
| wblrnd | rand, randn |
| wishrnd | randibution |

## Using Probability Distribution Objects

In this section...<br>"Using Distribution Objects" on page 5-84<br>"What are Objects?" on page 5-85<br>"Creating Distribution Objects" on page 5-88<br>"Object-Supported Distributions" on page 5-89<br>"Performing Calculations Using Distribution Objects" on page 5-90<br>"Capturing Results Using Distribution Objects" on page 5-97

## Using Distribution Objects

For many distributions supported by Statistics Toolbox software, objects are available for statistical analysis. This section gives a general overview of the uses of distribution objects, including sample work flows. For information on objects available for specific distributions, see "Object-Supported Distributions" on page 5-89.

Probability distribution objects allow you to easily fit, access, and store distribution information for a given data set. The following operations are easier to perform using distribution objects:

- Grouping a single dataset in a number of different ways using group names, and then fit a distribution to each group. For an example of how to fit distributions to grouped data, see "Example: Fitting Distributions to Grouped Data Within a Single Dataset" on page 5-91.
- Fitting different distributions to the same set of data. For an example of how objects make fitting multiple distribution types easier, see "Example: Fitting Multiple Distribution Types to a Single Dataset" on page 5-95.
- Sharing fitted distributions across workspaces. For an example of sharing information using probability distribution objects, see "Example: Saving and Sharing Distribution Fit Data" on page 5-97.


## Deciding to Use Distribution Objects

If you know the type of distribution you would like to use, objects provide a less complex interface than functions and a more efficient functionality than the dfittool GUI.

If you are a novice statistician who would like to explore how various distributions look without having to manipulate data, see "Working with Distributions Through GUIs" on page 5-9.

If you have no data to fit, but want to calculate a pdf, cdf, etc for various parameters, see "Statistics Toolbox Distribution Functions" on page 5-52.

## What are Objects?

Objects are, in short, a convenient way of storing data. They allow you to set rules for the types of data to store, while maintaining some flexibility for the actual values of the data. For example, in statistics groups of distributions have some general things in common:

- All distributions have a name (ex, Normal).
- Parametric distributions have parameters.
- Nonparametric distributions have kernel-smoothing functions.

Objects store all this information within properties. Classes of related objects (for example, all univariate parametric distributions) have the same properties with values and types relevant to a specified distribution. In addition to storing information within objects, you can perform certain actions (called methods) on objects.

Subclasses (for example, ProbDistParametric is a subclass of ProbDist) contain the same properties and methods as the original class, in addition to other properties relevant to that subclass. This concept is called inheritance. Inheritance means that subclasses of a class have all of its properties and methods. For example, parametric distributions, which are a subset (subclass) of probability distributions, have input data and a distribution name. The following diagram illustrates this point:


The left side of this diagram shows the inheritance line from all probability distributions down to univariate parametric probability distributions. The right side shows the lineage down to univariate kernel distributions. Here is how to interpret univariate parametric distribution lineage:

- ProbDist is a class of objects that includes all probability distributions. All probability distribution objects have at least these properties:
- DistName - the name of the distribution (for example Normal or Weibull)
- InputData - the data fit to the distribution

In addition, you can perform the following actions on these objects, using the following methods:

- cdf - Return the cumulative distribution function for a specified distribution.
- pdf - Return the probability density function for a specified distribution.
- random - Generate random numbers based on a specified distribution.
- ProbDistParametric is a class of objects that includes all parametric probability distributions. All parametric probability distribution objects have the properties and methods of a ProbDist object, in addition to at least the following properties:
- NLogL - Negative log likelihood for input data
- NumParams - Number of parameters for that distribution
- ParamCov - Covariance matrix of parameter estimates
- ParamDescription - Descriptions of parameters
- ParamNames - Names of parameters
- Params - Values of parameters

No additional unique methods apply to ProbDistParametric objects.

- ProbDistUnivParam is a class of objects that includes only univariate parametric probability distributions. In addition to the properties and methods of ProbDist and ProbDistParametric objects, these objects also have at least the following methods:
- icdf - Return the inverse cumulative distribution function for a specified distribution based on a given set of data.
- iqr - Return the interquartile range for a specified distribution based on a given set of data.
- mean - Return the mean for a specified distribution based on a given set of data.
- median - Return the median for a specified distribution based on a given set of data.
- paramci - Return the parameter confidence intervals for a specified distribution based on a given set of data.
- std - Return the standard deviation for a specified distribution based on a given set of data.
- var - Return the variance for a specified distribution based on a given set of data.
No additional unique properties apply to ProbDistUnivParam objects.

The univariate nonparametric lineage reads in a similar manner, with different properties and methods. For more information on nonparametric objects and their methods and properties, see ProbDistKernel and ProbDistUnivKernel.

For more detailed information on object-oriented programming in MATLAB, see Object-Oriented Programming.

## Creating Distribution Objects

There are two ways to create distribution objects:

- Use the fitdist function. See "Creating Distribution Objects Using fitdist" on page 5-88.
- Use the object constructor. See "Creating Distribution Objects Using Constructors" on page 5-88.


## Creating Distribution Objects Using fitdist

Using the fitdist function is the simplest way of creating distribution objects. Like the *fit functions, fitdist fits your data to a specified distribution and returns relevant distribution information. fitdist creates an object relevant to the type of distribution you specify: if you specify a parametric distribution, it returns a ProbDistUnivParam object. For examples of how to use fitdist to fit your data, see "Performing Calculations Using Distribution Objects" on page 5-90.

## Creating Distribution Objects Using Constructors

If you know the distribution you would like to use and would like to create a univariate parametric distribution with known parameters, you can use the ProbDistUnivParam constructor. For example, create a normal distribution with mean 100 and standard deviation 10 :

```
pd = ProbDistUnivParam('normal',[100 10])
```

For nonparametric distributions, you must have a dataset. Using fitdist is a simpler way to fit nonparametric data, but you can use the ProbDistUnivKernel constructor as well. For example, create a nonparametric distribution of the MPG data from carsmall.mat:

```
load carsmall
pd = ProbDistUnivKernel(MPG)
```


## Object-Supported Distributions

Object-oriented programming in the Statistics Toolbox supports the following distributions.

## Parametric Distributions

Use the following distribution to create ProbDistUnivParam objects using fitdist. For more information on the cumulative distribution function (cdf) and probability density function (pdf) methods, as well as other available methods, see the ProbDistUnivParam class reference page.

| Supported Distribution | Input to fitdist |
| :--- | :--- |
| "Beta Distribution" on page B-4 | 'beta' |
| "Binomial Distribution" on page B-7 | 'binomial' |
| "Birnbaum-Saunders Distribution" <br> on page B-10 | 'birnbaumsaunders ' |
| "Exponential Distribution" on page <br> B-16 | 'exponential' |
| "Extreme Value Distribution" on <br> page B-19 | 'extreme value' or 'ev' |
| "Gamma Distribution" on page B-27 | 'gamma' |
| "Generalized Extreme Value <br> Distribution" on page B-32 | 'generalized extreme value ' or |
| "Generalized Pareto Distribution" on <br> page B-37 | 'generalized pareto' or 'gp ' |
| "Inverse Gaussian Distribution" on <br> page B-45 | 'inversegaussian' |
| "Logistic Distribution" on page B-49 | 'logistic' |
| "Loglogistic Distribution" on page <br> B-50 | 'loglogistic' |


| Supported Distribution | Input to fitdist |
| :--- | :--- |
| "Lognormal Distribution" on page <br> B-51 | 'lognormal' |
| "Nakagami Distribution" on page <br> B-70 | 'nakagami' |
| "Negative Binomial Distribution" on <br> page B-72 | 'negative binomial' or 'nbin' |
| "Normal Distribution" on page B-83 | 'normal' |
| "Poisson Distribution" on page B-89 | 'poisson' |
| "Rayleigh Distribution" on page B-91 | 'rayleigh' |
| "Rician Distribution" on page B-93 | 'rician' |
| "t Location-Scale Distribution" on <br> page B-97 | 'tlocationscale' |
| "Weibull Distribution" on page B-103 | 'weibull' or 'wbl' |

## Nonparametric Distributions

Use the following distributions to create ProbDistUnivKernel objects. For more information on the cumulative distribution function (cdf) and probability density function (pdf) methods, as well as other available methods, see the ProbDistUnivKernel class reference page.

| Supported Distribution | Input to fitdist |
| :--- | :--- |
| "Nonparametric Distributions" on <br> page B-82 | 'kernel' |

## Performing Calculations Using Distribution Objects

Distribution objects make it easier for you to perform calculations on complex datasets. The following sample workflows show some of the functionality of these objects.

- "Example: Fitting a Single Distribution to a Single Dataset" on page 5-91
- "Example: Fitting Distributions to Grouped Data Within a Single Dataset" on page 5-91
- "Example: Fitting Multiple Distribution Types to a Single Dataset" on page 5-95


## Example: Fitting a Single Distribution to a Single Dataset

Fit a single Normal distribution to a dataset using fitdist:

```
load carsmall
NormDist = fitdist(MPG,'normal')
NormDist =
normal distribution
    mu = 23.7181
    sigma = 8.03573
```

The output MATLAB returns is a ProbDistUnivParam object with a DistName property of 'normal distribution'. The ParamNames property contains the strings mu and sigma, while the Params property contains the parameter values.

## Example: Fitting Distributions to Grouped Data Within a Single Dataset

Often, datasets are collections of data you can group in different ways. Using fitdist and the data from carsmall.mat, group the MPG data by country of origin, then fit a Weibull distribution each group:

```
load carsmall
[WeiByOrig, Country] = fitdist(MPG,'weibull','by',Origin)
Warning: Error while fitting group 'Italy':
Not enough data in X to fit this distribution.
> In fitdist at 171
WeiByOrig =
```

```
Columns 1 through 4
[1x1 ProbDistUnivParam] [1x1 ProbDistUnivParam] ...
[1x1 ProbDistUnivParam] [1x1 ProbDistUnivParam]
Columns 5 through 6
[1x1 ProbDistUnivParam] []
Country =
'USA'
'France'
'Japan'
'Germany '
'Sweden'
'Italy'
```

A warning appears informing you that, since the data only represents one Italian car, fitdist cannot fit a Weibull distribution to that group. Each one of the five other groups now has a distribution object associated with it, represented in the cell array wd. Each object contains properties that hold information about the data, the distribution, and the parameters. For more information on what properties exist and what information they contain, see ProbDistUnivParam or ProbDistUnivKernel.

Now access two of the objects and their properties:

```
% Get USA fit
distusa = WeiByOrig{1};
% Use the InputData property of ProbDistUnivParam objects to see
% the actual data used to fit the distribution:
dusa = distusa.InputData.data;
% Get Japan fit and data
distjapan = WeiByOrig{3};
djapan = distjapan.InputData.data;
```

Now you can easily compare PDFs using the pdf method of the ProbDistUnivParam class:

```
time = linspace(0,45);
pdfjapan = pdf(distjapan,time);
pdfusa = pdf(distusa,time);
hold on
plot(time,[pdfjapan;pdfusa])
l = legend('Japan','USA')
set(l,'Location','Best')
xlabel('MPG')
ylabel('Probability Density')
```



You could then further group the data and compare, for example, MPG by year for American cars:

```
load carsmall
[WeiByYearOrig, Names] = fitdist(MPG,'weibull','by',...
    {Origin Model_Year});
USA70 = WeiByYearOrig{1};
USA76 = WeiByYearOrig{2};
USA82 = WeiByYearOrig{3};
time = linspace(0,45);
pdf70 = pdf(USA70,time);
pdf76 = pdf(USA76,time);
pdf82 = pdf(USA82,time);
line(t,[pdf70;pdf76;pdf82])
l = legend('1970','1976','1982')
set(l,'Location','Best')
title('USA Car MPG by Year')
xlabel('MPG')
ylabel('Probability Density')
```



## Example: Fitting Multiple Distribution Types to a Single Dataset

Distribution objects make it easy to fit multiple distributions to the same dataset, while minimizing workspace clutter. For example, use fitdist to group the MPG data by country of origin, then fit Weibull, Normal, Logistic, and nonparametric distributions for each group:

```
load carsmall;
[WeiByOrig, Country] = fitdist(MPG,'weibull','by',Origin);
[NormByOrig, Country] = fitdist(MPG,'normal','by',Origin);
[LogByOrig, Country] = fitdist(MPG,'logistic','by',Origin);
[KerByOrig, Country] = fitdist(MPG,'kernel','by',Origin);
```

Extract the fits for American cars and compare the fits visually against a histogram of the original data:

```
WeiUSA = WeiByOrig{1};
NormUSA = NormByOrig{1};
LogUSA = LogByOrig{1};
KerUSA = KerByOrig{1};
% Since all three distributions use the same set of data,
% you can extract the data from any of them:
data = WeiUSA.InputData.data;
% Create a histogram of the data:
[n,y] = hist(data,10);
b = bar(y,n,'hist');
set(b,'FaceColor',[1,0.8,0])
% Scale the density by the histogram area, for easier display:
area = sum(n) * (y(2)-y(1));
time = linspace(0,45);
pdfWei = pdf(WeiUSA,time);
pdfNorm = pdf(NormUSA,time);
pdfLog = pdf(LogUSA,time);
pdfKer = pdf(KerUSA,time);
allpdf = [pdfWei;pdfNorm;pdfLog;pdfKer];
line(t,area * allpdf)
l = legend('Data','Weibull','Normal','Logistic','Kernel')
set(l,'Location','Best')
title('USA Car')
xlabel('MPG')
```



You can see that only the nonparametric kernel distribution, KerUSA, comes close to revealing the two modes in the data.

## Capturing Results Using Distribution Objects

Distribution objects allow you to share both your dataset and your analysis results simply by saving the information to a .mat file.

## Example: Saving and Sharing Distribution Fit Data

Using the premise from the previous set of examples, group the MPG data in carsmall.mat by country of origin and fit four different distributions to each of the six sets of data:

```
load carsmall;
[WeiByOrig, Country] = fitdist(MPG,'weibull','by',Origin);
[NormByOrig, Country] = fitdist(MPG,'normal','by',Origin);
[LogByOrig, Country] = fitdist(MPG,'logistic','by',Origin);
[KerByOrig, Country] = fitdist(MPG,'kernel','by',Origin);
```

Combine all four fits and the country labels into a single cell array, including "headers" to indicate which distributions correspond to which objects. Then, save the array to a .mat file:

```
AllFits = cell(['Country' Country'; 'Weibull' WeiByOrig;...
    'Normal' NormByOrig; 'Logistic' LogByOrig; 'Kernel',...
    KerByOrig]);
save('CarSmallFits.mat','AllFits');
```

To show that the data is both safely saved and easily restored, clear your workspace of relevant variables. This command clears only those variables associated with this example:

```
clear('Weight','Acceleration','AllFits','Country',...
'Cylinders','Displacement','Horsepower','KerByOrig' , ...
'LogByOrig','MPG','Model','Model_Year','NormByOrig' , ...
'Origin','WeiByOrig')
```

Now, load the data:
load CarSmallFits
AllFits

You can now access the distributions objects as in the previous examples.

# Probability Distributions Used for Multivariate Modeling 

In this section...<br>"Gaussian Mixture Models" on page 5-99<br>"Copulas" on page 5-107

## Gaussian Mixture Models

- "Creating Gaussian Mixture Models" on page 5-99
- "Simulating Gaussian Mixtures" on page 5-105

Gaussian mixture models are formed by combining multivariate normal density components. For information on individual multivariate normal densities, see "Multivariate Normal Distribution" on page B-58 and related distribution functions listed under "Multivariate Distributions" on page 5-8.

In Statistics Toolbox software, use the gmdistribution class to fit data using an expectation maximization (EM) algorithm, which assigns posterior probabilities to each component density with respect to each observation. The fitting method uses an iterative algorithm that converges to a local optimum. Clustering using Gaussian mixture models is sometimes considered a soft clustering method. The posterior probabilities for each point indicate that each data point has some probability of belonging to each cluster.

For more information on clustering with Gaussian mixture models, see "Gaussian Mixture Models" on page 11-28. This section describes their creation.

## Creating Gaussian Mixture Models

- "Specifying a Model" on page 5-100
- "Fitting a Model to Data" on page 5-102

Specifying a Model. Use the gmdistribution constructor to create Gaussian mixture models with specified means, covariances, and mixture proportions. The following creates an object of the gmdistribution class defining a two-component mixture of bivariate Gaussian distributions:

```
MU = [1 2;-3 -5]; % Means
SIGMA = cat(3,[2 0;0 .5],[1 0;0 1]); % Covariances
p = ones(1,2)/2; % Mixing proportions
obj = gmdistribution(MU,SIGMA,p);
```

Display properties of the object with the MATLAB function fieldnames:

```
properties = fieldnames(obj)
properties =
    'NDimensions'
    'DistName'
    'NComponents'
    'PComponents'
    'mu'
    'Sigma'
    'NlogL'
    'AIC'
    'BIC'
    'Converged'
    'Iters'
    'SharedCov'
    'CovType'
    'RegV'
```

The gmdistribution reference page describes these properties. To access the value of a property, use dot indexing:

```
dimension = obj.NDimensions
dimension =
    2
name = obj.DistName
name =
gaussian mixture distribution
```

Use the methods pdf and cdf to compute values and visualize the object:

```
ezsurf(@(x,y)pdf(obj,[x y]),[-10 10],[-10 10])
```

$\operatorname{pdf}($ obj, $[x, y])$


```
ezsurf(@(x,y)cdf(obj,[x y]),[-10 10],[-10 10])
```



Fitting a Model to Data. You can also create Gaussian mixture models by fitting a parametric model with a specified number of components to data. The fit method of the gmdistribution class uses the syntax obj = gmdistribution.fit ( $X, k$ ), where $X$ is a data matrix and $k$ is the specified number of components. Choosing a suitable number of components $k$ is essential for creating a useful model of the data-too few components fails to model the data accurately; too many components leads to an over-fit model with singular covariance matrices.

The following example illustrates this approach.
First, create some data from a mixture of two bivariate Gaussian distributions using the mvnrnd function:

```
MU1 = [1 2];
SIGMA1 = [2 0; 0 .5];
```

```
MU2 = [-3 -5];
SIGMA2 = [1 0; 0 1];
X = [mvnrnd(MU1,SIGMA1,1000);
mvnrnd(MU2,SIGMA2, 1000)];
scatter(X(:, 1),X(:, 2),10,'.')
```



Next, fit a two-component Gaussian mixture model:

```
options = statset('Display','final');
obj = gmdistribution.fit(X,2,'Options',options);
hold on
h = ezcontour(@(x,y)pdf(obj,[x y]),[-8 6],[-8 6]);
hold off
```



Among the properties of the fit are the parameter estimates:

```
ComponentMeans = obj.mu
ComponentMeans =
    0.9391 2.0322
    -2.9823 -4.9737
ComponentCovariances = obj.Sigma
ComponentCovariances(:,:,1) =
    1.7786 -0.0528
    -0.0528 0.5312
ComponentCovariances(:,:,2) =
    1.0491 -0.0150
```

```
    -0.0150 0.9816
MixtureProportions = obj.PComponents
MixtureProportions =
    0.5000 0.5000
```

The two-component model minimizes the Akaike information:

```
AIC = zeros(1,4);
obj = cell(1,4);
for k = 1:4
    obj{k} = gmdistribution.fit(X,k);
    AIC(k)= obj{k}.AIC;
end
[minAIC,numComponents] = min(AIC);
numComponents
numComponents =
    2
model = obj{2}
model =
Gaussian mixture distribution
with 2 components in 2 dimensions
Component 1:
Mixing proportion: 0.500000
Mean: 0.9391 2.0322
Component 2:
Mixing proportion: 0.500000
Mean: -2.9823 -4.9737
```

Both the Akaike and Bayes information are negative log-likelihoods for the data with penalty terms for the number of estimated parameters. You can use them to determine an appropriate number of components for a model when the number of components is unspecified.

## Simulating Gaussian Mixtures

Use the method random of the gmdistribution class to generate random data from a Gaussian mixture model created with gmdistribution or fit.

For example, the following specifies a gmdistribution object consisting of a two-component mixture of bivariate Gaussian distributions:

```
MU = [1 2;-3 -5];
SIGMA = cat(3,[2 0;0 .5],[1 0;0 1]);
p = ones(1,2)/2;
obj = gmdistribution(MU,SIGMA,p);
ezcontour(@(x,y)pdf(obj,[x y]),[-10 10],[-10 10])
hold on
```



Use random (gmdistribution) to generate 1000 random values:

```
Y = random(obj,1000);
scatter(Y(:,1),Y(:,2),10,'.')
```



## Copulas

- "Determining Dependence Between Simulation Inputs" on page 5-108
- "Constructing Dependent Bivariate Distributions" on page 5-112
- "Using Rank Correlation Coefficients" on page 5-116
- "Using Bivariate Copulas" on page 5-119
- "Higher Dimension Copulas" on page 5-126
- "Archimedean Copulas" on page 5-128
- "Simulating Dependent Multivariate Data Using Copulas" on page 5-130
- "Example: Fitting Copulas to Data" on page 5-135

Copulas are functions that describe dependencies among variables, and provide a way to create distributions that model correlated multivariate data. Using a copula, you can construct a multivariate distribution by specifying marginal univariate distributions, and then choose a copula to provide a correlation structure between variables. Bivariate distributions, as well as distributions in higher dimensions, are possible.

## Determining Dependence Between Simulation Inputs

One of the design decisions for a Monte Carlo simulation is a choice of probability distributions for the random inputs. Selecting a distribution for each individual variable is often straightforward, but deciding what dependencies should exist between the inputs may not be. Ideally, input data to a simulation should reflect what you know about dependence among the real quantities you are modeling. However, there may be little or no information on which to base any dependence in the simulation. In such cases, it is useful to experiment with different possibilities in order to determine the model's sensitivity.

It can be difficult to generate random inputs with dependence when they have distributions that are not from a standard multivariate distribution. Further, some of the standard multivariate distributions can model only limited types of dependence. It is always possible to make the inputs independent, and while that is a simple choice, it is not always sensible and can lead to the wrong conclusions.

For example, a Monte-Carlo simulation of financial risk could have two random inputs that represent different sources of insurance losses. You could model these inputs as lognormal random variables. A reasonable question to ask is how dependence between these two inputs affects the results of the simulation. Indeed, you might know from real data that the same random conditions affect both sources; ignoring that in the simulation could lead to the wrong conclusions.

## Example: Generate and Exponentiate Normal Random Variables.

The lognrnd function simulates independent lognormal random variables. In the following example, the mvnrnd function generates n pairs of independent normal random variables, and then exponentiates them. Notice that the covariance matrix used here is diagonal:

```
n = 1000;
sigma = .5;
SigmaInd = sigma.^2 .* [1 0; 0 1]
SigmaInd =
    0.25 0
    0 0.25
ZInd = mvnrnd([0 0],SigmaInd,n);
XInd = exp(ZInd);
```

```
plot(XInd(:,1),XInd(:, 2),'.')
```

plot(XInd(:,1),XInd(:, 2),'.')
axis([0 5 0 5])
axis([0 5 0 5])
axis equal
axis equal
xlabel('X1')
xlabel('X1')
ylabel('X2')

```
ylabel('X2')
```



Dependent bivariate lognormal random variables are also easy to generate using a covariance matrix with nonzero off-diagonal terms:

```
rho = .7;
SigmaDep = sigma.^2 .* [1 rho; rho 1]
SigmaDep =
    0.25 0.175
    0.175 0.25
```

ZDep = mvnrnd([0 0],SigmaDep,n);
XDep $=\exp (Z D e p) ;$

A second scatter plot demonstrates the difference between these two bivariate distributions:

```
plot(XDep(:,1),XDep(:,2),'.')
axis([0 5 0 5])
axis equal
xlabel('X1')
ylabel('X2')
```



It is clear that there is a tendency in the second data set for large values of X1 to be associated with large values of X2, and similarly for small values. The correlation parameter, $\rho$, of the underlying bivariate normal determines this dependence. The conclusions drawn from the simulation could well depend on whether you generate X1 and X2 with dependence. The bivariate lognormal distribution is a simple solution in this case; it easily generalizes to higher dimensions in cases where the marginal distributions are different lognormals.

Other multivariate distributions also exist. For example, the multivariate $t$ and the Dirichlet distributions simulate dependent $t$ and beta random variables, respectively. But the list of simple multivariate distributions is not long, and they only apply in cases where the marginals are all in the same family (or even the exact same distributions). This can be a serious limitation in many situations.

## Constructing Dependent Bivariate Distributions

Although the construction discussed in the previous section creates a bivariate lognormal that is simple, it serves to illustrate a method that is more generally applicable.

1 Generate pairs of values from a bivariate normal distribution. There is statistical dependence between these two variables, and each has a normal marginal distribution.

2 Apply a transformation (the exponential function) separately to each variable, changing the marginal distributions into lognormals. The transformed variables still have a statistical dependence.

If a suitable transformation can be found, this method can be generalized to create dependent bivariate random vectors with other marginal distributions. In fact, a general method of constructing such a transformation does exist, although it is not as simple as exponentiation alone.

By definition, applying the normal cumulative distribution function (cdf), denoted here by $\Phi$, to a standard normal random variable results in a random variable that is uniform on the interval $[0,1]$. To see this, if $Z$ has a standard normal distribution, then the cdf of $U=\Phi(Z)$ is
$\operatorname{Pr}\{U \leq u\}=\operatorname{Pr}\{\Phi(Z) \leq u\}=\operatorname{Pr}\left(Z \leq \Phi^{-1}(u)\right\}=u$
and that is the cdf of a $\operatorname{Unif}(0,1)$ random variable. Histograms of some simulated normal and transformed values demonstrate that fact:

```
n = 1000;
z = normrnd(0,1,n,1);
hist(z,-3.75:.5:3.75)
xlim([-4 4])
title('1000 Simulated N(0,1) Random Values')
xlabel('Z')
ylabel('Frequency')
set(get(gca,'Children'),'FaceColor',[.8 .8 1])
```




Borrowing from the theory of univariate random number generation, applying the inverse cdf of any distribution, $F$, to a $\operatorname{Unif}(0,1)$ random variable results in a random variable whose distribution is exactly $F$ (see "Inversion Methods" on page 6-7). The proof is essentially the opposite of the preceding proof for the forward case. Another histogram illustrates the transformation to a gamma distribution:

```
x = gaminv(u,2,1);
hist(x,.25:.5:9.75)
title('1000 Simulated N(0,1) Values Transformed to Gamma(2,1)')
xlabel('X')
ylabel('Frequency')
set(get(gca,'Children'),'FaceColor',[.8 .8 1])
```



You can apply this two-step transformation to each variable of a standard bivariate normal, creating dependent random variables with arbitrary marginal distributions. Because the transformation works on each component separately, the two resulting random variables need not even have the same marginal distributions. The transformation is defined as:

$$
\begin{aligned}
Z & =\left[Z_{1}, Z_{2}\right] \square N\left([0,0],\left[\begin{array}{ll}
1 & \rho \\
\rho & 1
\end{array}\right]\right) \\
U & =\left[\Phi\left(Z_{1}\right), \Phi\left(Z_{2}\right)\right] \\
X & =\left[G_{1}\left(U_{1}\right), G_{2}\left(U_{2}\right)\right]
\end{aligned}
$$

where $G_{1}$ and $G_{2}$ are inverse cdfs of two possibly different distributions. For example, the following generates random vectors from a bivariate distribution with $t_{5}$ and $\operatorname{Gamma}(2,1)$ marginals:

```
n = 1000; rho = .7;
Z = mvnrnd([0 0],[1 rho; rho 1],n);
U = normcdf(Z);
X = [gaminv(U(:,1),2,1) tinv(U(:, 2),5)];
scatterhist(X(:,1),X(:,2),'Direction','out')
```



This plot has histograms alongside a scatter plot to show both the marginal distributions, and the dependence.

## Using Rank Correlation Coefficients

The correlation parameter, $\rho$, of the underlying bivariate normal determines the dependence between X1 and X2 in this construction. However, the linear correlation of X1 and X2 is not $\rho$. For example, in the original lognormal case, a closed form for that correlation is:

$$
\operatorname{cor}(X 1, X 2)=\frac{e^{\rho \sigma^{2}}-1}{e^{\sigma^{2}}-1}
$$

which is strictly less than $\rho$, unless $\rho$ is exactly 1 . In more general cases such as the Gamma/t construction, the linear correlation between X1 and X2 is difficult or impossible to express in terms of $\rho$, but simulations show that the same effect happens.

That is because the linear correlation coefficient expresses the linear dependence between random variables, and when nonlinear transformations are applied to those random variables, linear correlation is not preserved. Instead, a rank correlation coefficient, such as Kendall's $\tau$ or Spearman's $\rho$, is more appropriate.

Roughly speaking, these rank correlations measure the degree to which large or small values of one random variable associate with large or small values of another. However, unlike the linear correlation coefficient, they measure the association only in terms of ranks. As a consequence, the rank correlation is preserved under any monotonic transformation. In particular, the transformation method just described preserves the rank correlation. Therefore, knowing the rank correlation of the bivariate normal $Z$ exactly determines the rank correlation of the final transformed random variables, $X$. While the linear correlation coefficient, $\rho$, is still needed to parameterize the underlying bivariate normal, Kendall's $\tau$ or Spearman's $\rho$ are more useful in describing the dependence between random variables, because they are invariant to the choice of marginal distribution.

For the bivariate normal, there is a simple one-to-one mapping between Kendall's $\tau$ or Spearman's $\rho$, and the linear correlation coefficient $\rho$ :

$$
\begin{aligned}
& \tau=\frac{2}{\pi} \arcsin (\rho) \quad \text { or } \quad \rho=\sin \left(\tau \frac{\pi}{2}\right) \\
& \rho_{\mathrm{s}}=\frac{6}{\pi} \arcsin \left(\frac{\rho}{2}\right) \text { or } \quad \rho=2 \sin \left(\rho_{\mathrm{s}} \frac{\pi}{6}\right)
\end{aligned}
$$

The following plot shows the relationship:

```
rho = -1:.01:1;
```

```
    tau = 2.*asin(rho)./pi;
    rho_s = 6.*asin(rho./2)./pi;
    plot(rho,tau,'b-','LineWidth', 2)
    hold on
    plot(rho,rho_s,'g-','LineWidth',2)
    plot([-1 1],[-1 1],'k:','LineWidth',2)
    axis([-1 1-1 -1 1])
    xlabel('rho')
    ylabel('Rank correlation coefficient')
    legend('Kendall''s {\it\tau}', ...
    'Spearman''s {\it\rho_s}', ...
    'location','NW')
```



Thus, it is easy to create the desired rank correlation between X1 and X2, regardless of their marginal distributions, by choosing the correct $\rho$ parameter value for the linear correlation between Z1 and Z2.

For the multivariate normal distribution, Spearman's rank correlation is almost identical to the linear correlation. However, this is not true once you transform to the final random variables.

## Using Bivariate Copulas

The first step of the construction described in the previous section defines what is known as a bivariate Gaussian copula. A copula is a multivariate probability distribution, where each random variable has a uniform marginal distribution on the unit interval $[0,1]$. These variables may be completely independent, deterministically related (e.g., U2 = U1), or anything in between. Because of the possibility for dependence among variables, you can use a copula to construct a new multivariate distribution for dependent variables. By transforming each of the variables in the copula separately using the inversion method, possibly using different cdfs, the resulting distribution can have arbitrary marginal distributions. Such multivariate distributions are often useful in simulations, when you know that the different random inputs are not independent of each other.

Statistics Toolbox functions compute:

- Probability density functions (copulapdf) and the cumulative distribution functions (copulacdf) for Gaussian copulas
- Rank correlations from linear correlations (copulastat) and vice versa (copulaparam)
- Random vectors (copularnd)
- Parameters for copulas fit to data (copulafit)

For example, use the copularnd function to create scatter plots of random values from a bivariate Gaussian copula for various levels of $\rho$, to illustrate the range of different dependence structures. The family of bivariate Gaussian copulas is parameterized by the linear correlation matrix:

$$
\mathrm{P}=\left(\begin{array}{ll}
1 & \rho \\
\rho & 1
\end{array}\right)
$$

U1 and U2 approach linear dependence as $\rho$ approaches $\pm 1$, and approach complete independence as $\rho$ approaches zero:

```
n = 500;
U = copularnd('Gaussian',[1 .8; . 8 1],n);
subplot(2,2,1)
plot(U(:,1),U(:,2),'.')
title('{\it\rho} = 0.8')
xlabel('U1')
ylabel('U2')
U = copularnd('Gaussian',[1 .1; .1 1],n);
subplot(2,2,2)
plot(U(:,1),U(:,2),'.')
title('{\it\rho} = 0.1')
xlabel('U1')
ylabel('U2')
U = copularnd('Gaussian',[1 -.1; -.1 1],n);
subplot(2,2,3)
plot(U(:,1),U(:,2),'.')
title('{\it\rho} = -0.1')
xlabel('U1')
ylabel('U2')
U = copularnd('Gaussian',[1 -.8; -. 8 1],n);
subplot(2,2,4)
plot(U(:,1),U(:,2),'.')
title('{\it\rho} = -0.8')
xlabel('U1')
ylabel('U2')
```



The dependence between U1 and U2 is completely separate from the marginal distributions of $\mathrm{X} 1=\mathrm{G}(\mathrm{U} 1)$ and $\mathrm{X} 2=\mathrm{G}(\mathrm{U} 2) . \mathrm{X} 1$ and X 2 can be given any marginal distributions, and still have the same rank correlation. This is one of the main appeals of copulas-they allow this separate specification of dependence and marginal distribution. You can also compute the pdf (copulapdf) and the cdf (copulacdf) for a copula. For example, these plots show the pdf and cdf for $\rho=.8$ :

```
u1 = linspace(1e-3,1-1e-3,50);
u2 = linspace(1e-3,1-1e-3,50);
[U1,U2] = meshgrid(u1,u2);
Rho = [1 . 8; .8 1];
f = copulapdf('t',[U1(:) U2(:)],Rho,5);
f = reshape(f,size(U1));
surf(u1,u2,log(f),'FaceColor','interp','EdgeColor','none')
```




A different family of copulas can be constructed by starting from a bivariate $t$ distribution and transforming using the corresponding $t$ cdf. The bivariate $t$ distribution is parameterized with $P$, the linear correlation matrix, and $v$, the degrees of freedom. Thus, for example, you can speak of a $t_{1}$ or a $t_{5}$ copula, based on the multivariate $t$ with one and five degrees of freedom, respectively.

Just as for Gaussian copulas, Statistics Toolbox functions for $t$ copulas compute:

- Probability density functions (copulapdf) and the cumulative distribution functions (copulacdf) for Gaussian copulas
- Rank correlations from linear correlations (copulastat) and vice versa (copulaparam)
- Random vectors (copularnd)
- Parameters for copulas fit to data (copulafit)

For example, use the copularnd function to create scatter plots of random values from a bivariate $t_{1}$ copula for various levels of $\rho$, to illustrate the range of different dependence structures:

```
n = 500;
nu = 1;
U = copularnd('t',[1 .8; .8 1],nu,n);
subplot(2,2,1)
plot(U(:,1),U(:,2),'.')
title('{\it\rho} = 0.8')
xlabel('U1')
ylabel('U2')
U = copularnd('t',[1 .1; .1 1],nu,n);
subplot(2,2,2)
plot(U(:,1),U(:,2),'.')
title('{\it\rho} = 0.1')
xlabel('U1')
ylabel('U2')
U = copularnd('t',[1 -.1; -.1 1],nu,n);
subplot(2,2,3)
plot(U(:,1),U(:,2),'.')
title('{\it\rho} = -0.1')
xlabel('U1')
ylabel('U2')
```

U = copularnd('t',[1-.8; -. 8 1],nu, n);
subplot (2,2,4)
plot(U(:,1),U(:,2),'.')
title('\{\it\rho\} = -0.8')
xlabel('U1')
ylabel('U2')


A $t$ copula has uniform marginal distributions for U1 and U2, just as a Gaussian copula does. The rank correlation $\tau$ or $\rho_{s}$ between components in a $t$ copula is also the same function of $\rho$ as for a Gaussian. However, as these plots demonstrate, a $t_{1}$ copula differs quite a bit from a Gaussian copula, even when their components have the same rank correlation. The difference is in their dependence structure. Not surprisingly, as the degrees of freedom parameter $v$ is made larger, a $t_{v}$ copula approaches the corresponding Gaussian copula.

As with a Gaussian copula, any marginal distributions can be imposed over a $t$ copula. For example, using a $t$ copula with 1 degree of freedom, you can again generate random vectors from a bivariate distribution with $\operatorname{Gamma}(2,1)$ and $t_{5}$ marginals using copularnd:

```
n = 1000;
rho = .7;
nu = 1;
```

    U = copularnd('t',[1 rho; rho 1],nu,n);
    U = copularnd('t',[1 rho; rho 1],nu,n);
    X = [gaminv(U(:,1),2,1) tinv(U(:,2),5)];
    X = [gaminv(U(:,1),2,1) tinv(U(:,2),5)];
    scatterhist(X(:,1),X(:,2),'Direction','out')
    scatterhist(X(:,1),X(:,2),'Direction','out')
    

Compared to the bivariate Gamma/t distribution constructed earlier, which was based on a Gaussian copula, the distribution constructed here, based on a $t_{1}$ copula, has the same marginal distributions and the same rank correlation between variables but a very different dependence structure. This illustrates the fact that multivariate distributions are not uniquely defined by their marginal distributions, or by their correlations. The choice of a particular copula in an application may be based on actual observed data, or different copulas may be used as a way of determining the sensitivity of simulation results to the input distribution.

## Higher Dimension Copulas

The Gaussian and $t$ copulas are known as elliptical copulas. It is easy to generalize elliptical copulas to a higher number of dimensions. For example, simulate data from a trivariate distribution with $\operatorname{Gamma}(2,1)$, Beta(2,2), and $t_{5}$ marginals using a Gaussian copula and copularnd, as follows:

```
    n = 1000;
    Rho = [1 .4 .2; . 4 1 -. 8; .2 -. 8 1];
    U = copularnd('Gaussian',Rho,n);
    X = [gaminv(U(:,1),2,1) betainv(U(:,2),2,2) tinv(U(:, 3),5)];
    subplot(1,1,1)
    plot3(X(:, 1),X(:,2),X(:,3),'.')
    grid on
    view([-55, 15])
    xlabel('X1')
    ylabel('X2')
zlabel('X3')
```



Notice that the relationship between the linear correlation parameter $\rho$ and, for example, Kendall's $\tau$, holds for each entry in the correlation matrix $P$ used here. You can verify that the sample rank correlations of the data are approximately equal to the theoretical values:

```
tauTheoretical = 2.*asin(Rho)./pi
tauTheoretical =
```

| 1 | 0.26198 | 0.12819 |
| ---: | ---: | ---: |
| 0.26198 | 1 | -0.59033 |
| 0.12819 | -0.59033 | 1 |
| tauSample $=$ corr (X, 'type ' , 'Kendall' $)$ |  |  |
| tauSample $=$ |  |  |
| 1 | 0.27254 | 0.12701 |
| 0.27254 | 1 | -0.58182 |
| 0.12701 | -0.58182 | 1 |

## Archimedean Copulas

Statistics Toolbox functions are available for three bivariate Archimedean copula families:

- Clayton copulas
- Frank copulas
- Gumbel copulas

These are one-parameter families that are defined directly in terms of their cdfs, rather than being defined constructively using a standard multivariate distribution.

To compare these three Archimedean copulas to the Gaussian and $t$ bivariate copulas, first use the copulastat function to find the rank correlation for a Gaussian or $t$ copula with linear correlation parameter of 0.8 , and then use the copulaparam function to find the Clayton copula parameter that corresponds to that rank correlation:

```
tau = copulastat('Gaussian',.8 ,'type','kendall')
tau =
    0.59033
alpha = copulaparam('Clayton',tau,'type','kendall')
alpha =
    2.882
```

Finally, plot a random sample from the Clayton copula with copularnd. Repeat the same procedure for the Frank and Gumbel copulas:

```
n = 500;
U = copularnd('Clayton',alpha,n);
subplot(3,1,1)
plot(U(:,1),U(:,2),'.');
title(['Clayton Copula, {\it\alpha} = ',sprintf('%0.2f',alpha)])
xlabel('U1')
ylabel('U2')
alpha = copulaparam('Frank',tau,'type','kendall');
U = copularnd('Frank',alpha,n);
subplot(3,1,2)
plot(U(:,1),U(:, 2),'.')
title(['Frank Copula, {\it\alpha} = ',sprintf('%0.2f',alpha)])
xlabel('U1')
ylabel('U2')
alpha = copulaparam('Gumbel',tau,'type','kendall');
U = copularnd('Gumbel',alpha,n);
subplot(3,1,3)
plot(U(:,1),U(:, 2),'.')
title(['Gumbel Copula, {\it\alpha} = ',sprintf('%0.2f',alpha)])
xlabel('U1')
ylabel('U2')
```



## Simulating Dependent Multivariate Data Using Copulas

To simulate dependent multivariate data using a copula, you must specify each of the following:

- The copula family (and any shape parameters)
- The rank correlations among variables
- Marginal distributions for each variable

Suppose you have return data for two stocks and want to run a Monte Carlo simulation with inputs that follow the same distributions as the data:

```
load stockreturns
nobs = size(stocks,1);
subplot(2,1,1)
hist(stocks(:, 1),10)
xlim([-3.5 3.5])
xlabel('X1')
ylabel('Frequency')
set(get(gca,'Children'),'FaceColor',[. . . 8 1])
subplot(2,1,2)
hist(stocks(:,2),10)
xlim([-3.5 3.5])
xlabel('X2')
ylabel('Frequency')
set(get(gca,'Children'),'FaceColor',[.8 . 8 1])
```



You could fit a parametric model separately to each dataset, and use those estimates as the marginal distributions. However, a parametric model may not be sufficiently flexible. Instead, you can use a nonparametric model to transform to the marginal distributions. All that is needed is a way to compute the inverse cdf for the nonparametric model.

The simplest nonparametric model is the empirical cdf, as computed by the ecdf function. For a discrete marginal distribution, this is appropriate. However, for a continuous distribution, use a model that is smoother than the step function computed by ecdf. One way to do that is to estimate the empirical cdf and interpolate between the midpoints of the steps with a piecewise linear function. Another way is to use kernel smoothing with ksdensity. For example, compare the empirical cdf to a kernel smoothed cdf estimate for the first variable:

```
    [Fi,xi] = ecdf(stocks(:,1));
    stairs(xi,Fi,'b','LineWidth', 2)
hold on
    Fi_sm = ksdensity(stocks(:,1),xi,'function','cdf','width',.15);
    plot(xi,Fi_sm,'r-','LineWidth',1.5)
    xlabel('X1')
    ylabel('Cumulative Probability')
    legend('Empirical','Smoothed','Location','NW')
grid on
```



For the simulation, experiment with different copulas and correlations. Here, you will use a bivariate $t$ copula with a fairly small degrees of freedom
parameter. For the correlation parameter, you can compute the rank correlation of the data, and then find the corresponding linear correlation parameter for the $t$ copula using copulaparam:

```
nu = 5;
tau = corr(stocks(:,1),stocks(:,2),'type','kendall')
tau =
    0.51798
rho = copulaparam('t', tau, nu, 'type','kendall')
rho =
    0.72679
```

Next, use copularnd to generate random values from the $t$ copula and transform using the nonparametric inverse cdfs. The ksdensity function allows you to make a kernel estimate of distribution and evaluate the inverse cdf at the copula points all in one step:

```
n = 1000;
U = copularnd('t',[1 rho; rho 1],nu,n);
X1 = ksdensity(stocks(:,1),U(:,1),\ldots
    'function','icdf','width',.15);
X2 = ksdensity(stocks(:,2),U(:,2),...
    'function','icdf','width',.15);
```

Alternatively, when you have a large amount of data or need to simulate more than one set of values, it may be more efficient to compute the inverse cdf over a grid of values in the interval $(0,1)$ and use interpolation to evaluate it at the copula points:

```
p = linspace(0.00001,0.99999,1000);
G1 = ksdensity(stocks(:,1),p,'function','icdf','width',0.15);
X1 = interp1(p,G1,U(:,1),'spline');
G2 = ksdensity(stocks(:,2),p,'function','icdf','width',0.15);
X2 = interp1(p,G2,U(:,2),'spline');
scatterhist(X1,X2,'Direction','out')
```



The marginal histograms of the simulated data are a smoothed version of the histograms for the original data. The amount of smoothing is controlled by the bandwidth input to ksdensity.

## Example: Fitting Copulas to Data

The copulafit function is used to calibrate copulas with data. To generate data Xsim with a distribution "just like" (in terms of marginal distributions and correlations) the distribution of data in the matrix $X$ :

1 Fit marginal distributions to the columns of $x$.
2 Use appropriate cdf functions to transform $X$ to $U$, so that $U$ has values between 0 and 1.

3 Use copulafit to fit a copula to $U$.
4 Generate new data Usim from the copula.

5 Use appropriate inverse cdf functions to transform Usim to Xsim.
The following example illustrates the procedure.
Load and plot simulated stock return data:
load stockreturns
x = stocks(:,1);
y = stocks(:,2);

```
scatterhist(x,y,'Direction','out')
```



Transform the data to the copula scale (unit square) using a kernel estimator of the cumulative distribution function:

```
u = ksdensity(x,x,'function','cdf');
v = ksdensity(y,y,'function','cdf');
```

```
scatterhist(u,v,'Direction','out')
xlabel('u')
ylabel('v')
```



Fit a $t$ copula:

```
[Rho,nu] = copulafit('t',[u v],'Method','ApproximateML')
Rho =
    1.0000 0.7220
    0.7220 1.0000
nu =
    3.2017e+006
```

Generate a random sample from the $t$ copula:

```
r = copularnd('t',Rho,nu,1000);
u1 = r(:,1);
v1 = r(:,2);
```

```
scatterhist(u1,v1,'Direction','out')
xlabel('u')
ylabel('v')
set(get(gca,'children'),'marker','.')
```



Transform the random sample back to the original scale of the data:

```
x1 = ksdensity(x,u1,'function','icdf');
y1 = ksdensity(y,v1,'function','icdf');
scatterhist(x1,y1,'Direction','out')
set(get(gca,'children'),'marker','.')
```



As the example illustrates, copulas integrate naturally with other distribution fitting functions.

# Random Number Generation 

- "Generating Random Data" on page 6-2
- "Random Number Generation Functions" on page 6-3
- "Common Generation Methods" on page 6-5
- "Representing Sampling Distributions Using Markov Chain Samplers" on page 6-13
- "Generating Quasi-Random Numbers" on page 6-15
- "Generating Data Using Flexible Families of Distributions" on page 6-25


## Generating Random Data

Pseudorandom numbers are generated by deterministic algorithms. They are "random" in the sense that, on average, they pass statistical tests regarding their distribution and correlation. They differ from true random numbers in that they are generated by an algorithm, rather than a truly random process.

Random number generators (RNGs) like those in MATLAB are algorithms for generating pseudorandom numbers with a specified distribution.

For more information on random number generators for supported distributions, see "Random Number Generators" on page 5-80.

For more information on the GUI for generating random numbers from supported distributions, see "Visually Exploring Random Number Generation" on page 5-49.

## Random Number Generation Functions

The following table lists the supported distributions and their respective random number generation functions. For more information on other functions for each distribution, see "Supported Distributions" on page 5-3. For more information on random number generators, see "Random Number Generators" on page 5-80.

| Distribution | Random Number Generation Function |
| :--- | :--- |
| Beta | betarnd, random, randtool |
| Binomial | binornd, random, randtool |
| Chi-square | chi2rnd, random, randtool |
| Clayton copula | copularnd |
| Exponential | exprnd, random, randtool |
| Extreme value | evrnd, random, randtool |
| $F$ | frnd, random, randtool |
| Frank copula | copularnd |
| Gamma | gamrnd, randg, random, randtool |
| Gaussian copula | copularnd |
| Gaussian mixture | random |
| Generalized extreme <br> value | gevrnd, random, randtool |
| Generalized Pareto | gprnd, random, randtool |
| Geometric | geornd, random, randtool |
| Gumbel copula | copularnd |
| Hypergeometric | hygernd, random |
| Inverse Wishart | iwishrnd |
| Johnson system | johnsrnd |
| Lognormal | lognrnd, random, randtool |
| Multinomial | mnrnd |


| Distribution | Random Number Generation Function |
| :--- | :--- |
| Multivariate normal | mvnrnd |
| Multivariate $t$ | mvtrnd |
| Negative binomial | nbinrnd, random, randtool |
| Noncentral chi-square | ncx2rnd, random, randtool |
| Noncentral $F$ | ncfrnd, random, randtool |
| Noncentral $t$ | nctrnd, random, randtool |
| Normal (Gaussian) | normrnd, randn, random, randtool |
| Pearson system | pearsrnd |
| Piecewise | random |
| Poisson | poissrnd, random, randtool |
| Rayleigh | raylrnd, random, randtool |
| Student's $t$ | trnd, random, randtool |
| $t$ copula | copularnd |
| Uniform (continuous) | unifrnd, rand, random |
| Uniform (discrete) | unidrnd, random, randtool |
| Weibull | wblrnd, random |
| Wishart | wishrnd |

## Common Generation Methods

```
In this section...
"Direct Methods" on page 6-5
"Inversion Methods" on page 6-7
"Acceptance-Rejection Methods" on page 6-9
```

Methods for generating pseudorandom numbers usually start with uniform random numbers, like the MATLAB rand function produces. The methods described in this section detail how to produce random numbers from other distributions.

## Direct Methods

Direct methods directly use the definition of the distribution.
For example, consider binomial random numbers. A binomial random number is the number of heads in $N$ tosses of a coin with probability $p$ of a heads on any single toss. If you generate $N$ uniform random numbers on the interval $(0,1)$ and count the number less than $p$, then the count is a binomial random number with parameters $N$ and $p$.

This function is a simple implementation of a binomial RNG using the direct approach:

```
function X = directbinornd(N,p,m,n)
X = zeros(m,n); % Preallocate memory
for i = 1:m*n
    u = rand(N,1);
    X(i) = sum(u < p);
end
```

For example:

```
X = directbinornd(100,0.3,1e4,1);
hist(X,101)
set(get(gca,'Children'),'FaceColor',[.8 .8 1])
```



The Statistics Toolbox function binornd uses a modified direct method, based on the definition of a binomial random variable as the sum of Bernoulli random variables.

You can easily convert the previous method to a random number generator for the Poisson distribution with parameter $\lambda$. The Poisson distribution is the limiting case of the binomial distribution as $N$ approaches infinity, $p$ approaches zero, and $N p$ is held fixed at $\lambda$. To generate Poisson random numbers, create a version of the previous generator that inputs $\lambda$ rather than $N$ and $p$, and internally sets $N$ to some large number and $p$ to $\lambda / N$.

The Statistics Toolbox function poissrnd actually uses two direct methods:

- A waiting time method for small values of $\lambda$
- A method due to Ahrens and Dieter for larger values of $\lambda$


## Inversion Methods

Inversion methods are based on the observation that continuous cumulative distribution functions (cdfs) range uniformly over the interval ( 0,1 ). If $u$ is a uniform random number on ( 0,1 ), then using $X=F^{-1}(U)$ generates a random number $X$ from a continuous distribution with specified $\operatorname{cdf} F$.

For example, the following code generates random numbers from a specific exponential distribution using the inverse cdf and the MATLAB uniform random number generator rand:

```
mu = 1;
X = expinv(rand(1e4,1),mu);
```

Compare the distribution of the generated random numbers to the pdf of the specified exponential by scaling the pdf to the area of the histogram used to display the distribution:

```
numbins = 50;
hist(X, numbins)
set(get(gca,'Children'),'FaceColor',[.8 .8 1])
hold on
[bincounts,binpositions] = hist(X,numbins);
binwidth = binpositions(2) - binpositions(1);
histarea = binwidth*sum(bincounts);
x = binpositions(1):0.001:binpositions(end);
y = exppdf(x,mu);
plot(x,histarea*y,'r','LineWidth',2)
```



Inversion methods also work for discrete distributions. To generate a random number $X$ from a discrete distribution with probability mass vector $P\left(X=x_{i}\right)=$ $p_{i}$ where $x_{0}<x_{1}<x_{2}<\ldots$, generate a uniform random number $u$ on $(0,1)$ and then set $X=x_{i}$ if $F\left(x_{i-1}\right)<u<F\left(x_{i}\right)$.

For example, the following function implements an inversion method for a discrete distribution with probability mass vector $p$ :

```
function X = discreteinvrnd(p,m,n)
X = zeros(m,n); % Preallocate memory
for i = 1:m*n
    u = rand;
    I = find(u < cumsum(p));
    X(i) = min(I);
end
```

Use the function to generate random numbers from any discrete distribution:

```
p = [0.1 0.2 0.3 0.2 0.1 0.1]; % Probability mass vector
X = discreteinvrnd(p,1e4,1);
[n,x] = hist(X,length(p));
bar(1:length(p),n)
set(get(gca,'Children'),'FaceColor',[.8 . 8 1])
```



## Acceptance-Rejection Methods

The functional form of some distributions makes it difficult or time-consuming to generate random numbers using direct or inversion methods.
Acceptance-rejection methods provide an alternative in these cases.
Acceptance-rejection methods begin with uniform random numbers, but require an additional random number generator. If your goal is to generate a random number from a continuous distribution with $\operatorname{pdf} f$, acceptance-rejection methods first generate a random number from a continuous distribution with pdf $g$ satisfying $f(x) \leq c g(x)$ for some $c$ and all $x$.


A continuous acceptance-rejection RNG proceeds as follows:
1 Chooses a density $g$.
2 Finds a constant $c$ such that $f(x) / g(x) \leq c$ for all $x$.
3 Generates a uniform random number $u$.
4 Generates a random number $v$ from $g$.
5 If $c u \leq f(v) / g(v)$, accepts and returns $v$.
6 Otherwise, rejects $v$ and goes to step 3 .
For efficiency, a "cheap" method is necessary for generating random numbers from $g$, and the scalar $c$ should be small. The expected number of iterations to produce a single random number is $c$.

The following function implements an acceptance-rejection method for generating random numbers from $\operatorname{pdf} f$, given $f$, $g$, the RNG grnd for $g$, and the constant $c$ :

```
function X = accrejrnd(f,g,grnd,c,m,n)
X = zeros(m,n); % Preallocate memory
for i = 1:m*n
    accept = false;
    while accept == false
        u = rand();
        v = grnd();
        if c*u <= f(v)/g(v)
            X(i) = v;
            accept = true;
        end
    end
end
```

For example, the function $f(x)=x \mathrm{e}^{-x_{2} / 2}$ satisfies the conditions for a pdf on $[0, \infty)$ (nonnegative and integrates to 1 ). The exponential pdf with mean $1, f(x)=\mathrm{e}^{-x}$, dominates $g$ for $c$ greater than about 2.2. Thus, you can use rand and exprnd to generate random numbers from $f$ :

```
f = @(x)x.*exp(-(x.^2)/2);
g = @(x)exp(-x);
grnd = @()exprnd(1);
X = accrejrnd(f,g,grnd,2.2,1e4,1);
```

The $\operatorname{pdf} f$ is actually a Rayleigh distribution with shape parameter 1 . This example compares the distribution of random numbers generated by the acceptance-rejection method with those generated by raylrnd:

```
Y = raylrnd(1,1e4,1);
hist([X Y])
h = get(gca,'Children');
set(h(1),'FaceColor',[.8 . 8 1])
legend('A-R RNG','Rayleigh RNG')
```



The Statistics Toolbox function raylrnd uses a transformation method, expressing a Rayleigh random variable in terms of a chi-square random variable, which you compute using randn.

Acceptance-rejection methods also work for discrete distributions. In this case, the goal is to generate random numbers from a distribution with probability mass $P_{p}(X=i)=p_{i}$, assuming that you have a method for generating random numbers from a distribution with probability mass $P_{q}(X=i)=q_{i}$. The RNG proceeds as follows:

1 Chooses a density $P_{q}$.
2 Finds a constant $c$ such that $p_{i} / q_{i} \leq c$ for all $i$.
3 Generates a uniform random number $u$.
4 Generates a random number $v$ from $P_{q}$.
5 If $c u \leq p_{v} / q_{v}$, accepts and returns $v$.
6 Otherwise, rejects $v$ and goes to step 3 .

## Representing Sampling Distributions Using Markov Chain Samplers

In this section...<br>"Using the Metropolis-Hastings Algorithm" on page 6-13<br>"Using Slice Sampling" on page 6-14

The methods in "Common Generation Methods" on page 6-5 might be inadequate when sampling distributions are difficult to represent in computations. Such distributions arise, for example, in Bayesian data analysis and in the large combinatorial problems of Markov chain Monte Carlo (MCMC) simulations. An alternative is to construct a Markov chain with a stationary distribution equal to the target sampling distribution, using the states of the chain to generate random numbers after an initial burn-in period in which the state distribution converges to the target.

## Using the Metropolis-Hastings Algorithm

The Metropolis-Hastings algorithm draws samples from a distribution that is only known up to a constant. Random numbers are generated from a distribution with a probability density function that is equal to or proportional to a proposal function.

To generate random numbers:
1 Assume an initial value $x(t)$.
2 Draw a sample, $y(t)$, from a proposal distribution $q(y \mid x(t))$.
3 Accept $y(t)$ as the next sample $x(t+1)$ with probability $r(x(t), y(t))$, and keep $x(t)$ as the next sample $x(t+1)$ with probability $1-r(x(t), y(t))$, where:

$$
r(x, y)=\min \left\{\frac{f(y)}{f(x)} \frac{q(x \mid y)}{q(y \mid x)}, 1\right\}
$$

4 Increment $t \rightarrow t+1$, and repeat steps 2 and 3 until you get the desired number of samples.

Generate random numbers using the Metropolis-Hastings method with the mhsample function. To produce quality samples efficiently with the Metropolis-Hastings algorithm, it is crucial to select a good proposal distribution. If it is difficult to find an efficient proposal distribution, use the slice sampling algorithm (slicesample) without explicitly specifying a proposal distribution.

## Using Slice Sampling

In instances where it is difficult to find an efficient Metropolis-Hastings proposal distribution, the slice sampling algorithm does not require an explicit specification. The slice sampling algorithm draws samples from the region under the density function using a sequence of vertical and horizontal steps. First, it selects a height at random from 0 to the density function $f(x)$. Then, it selects a new $x$ value at random by sampling from the horizontal "slice" of the density above the selected height. A similar slice sampling algorithm is used for a multivariate distribution.

If a function $f(x)$ proportional to the density function is given, then do the following to generate random numbers:

1 Assume an initial value $x(t)$ within the domain of $f(x)$.
2 Draw a real value $y$ uniformly from $(0, f(x(t))$ ), thereby defining a horizontal "slice" as $S=\{x: y<f(x)\}$.

3 Find an interval $I=(L, R)$ around $x(t)$ that contains all, or much of the "slice" $S$.

4 Draw the new point $x(t+1)$ within this interval.
5 Increment $t \rightarrow t+1$ and repeat steps 2 through 4 until you get the desired number of samples.

Slice sampling can generate random numbers from a distribution with an arbitrary form of the density function, provided that an efficient numerical procedure is available to find the interval $I=(L, R)$, which is the "slice" of the density.

Generate random numbers using the slice sampling method with the slicesample function.

## Generating Quasi-Random Numbers

In this section...<br>"Quasi-Random Sequences" on page 6-15<br>"Quasi-Random Point Sets" on page 6-16<br>"Quasi-Random Streams" on page 6-23

## Quasi-Random Sequences

Quasi-random number generators (QRNGs) produce highly uniform samples of the unit hypercube. QRNGs minimize the discrepancy between the distribution of generated points and a distribution with equal proportions of points in each sub-cube of a uniform partition of the hypercube. As a result, QRNGs systematically fill the "holes" in any initial segment of the generated quasi-random sequence.

Unlike the pseudorandom sequences described in "Common Generation Methods" on page 6-5, quasi-random sequences fail many statistical tests for randomness. Approximating true randomness, however, is not their goal. Quasi-random sequences seek to fill space uniformly, and to do so in such a way that initial segments approximate this behavior up to a specified density.

QRNG applications include:

- Quasi-Monte Carlo (QMC) integration. Monte Carlo techniques are often used to evaluate difficult, multi-dimensional integrals without a closed-form solution. QMC uses quasi-random sequences to improve the convergence properties of these techniques.
- Space-filling experimental designs. In many experimental settings, taking measurements at every factor setting is expensive or infeasible. Quasi-random sequences provide efficient, uniform sampling of the design space.
- Global optimization. Optimization algorithms typically find a local optimum in the neighborhood of an initial value. By using a quasi-random sequence of initial values, searches for global optima uniformly sample the basins of attraction of all local minima.


## Example: Using Scramble, Leap, and Skip

Imagine a simple 1-D sequence that produces the integers from 1 to 10 . This is the basic sequence and the first three points are [1,2,3]:

## 12345678910

Now look at how Scramble, Leap, and Skip work together:

- Scramble - Scrambling shuffles the points in one of several different ways. In this example, assume a scramble turns the sequence into $1,3,5,7,9,2,4,6,8,10$. The first three points are now [1,3,5]:


## 13579246810

- Skip - A Skip value specifies the number of initial points to ignore. In this example, set the Skip value to 2 . The sequence is now $5,7,9,2,4,6,8,10$ and the first three points are [5,7,9]:
$13 \lcm{579246810}$
- Leap - A Leap value specifies the number of points to ignore for each one you take. Continuing the example with the Skip set to 2 , if you set the Leap to 1 , the sequence uses every other point. In this example, the sequence is now $5,9,4,8$ and the first three points are [5,9,4]:



## Quasi-Random Point Sets

Statistics Toolbox functions support these quasi-random sequences:

- Halton sequences. Produced by the haltonset function. These sequences use different prime bases to form successively finer uniform partitions of the unit interval in each dimension.
- Sobol sequences. Produced by the sobolset function. These sequences use a base of 2 to form successively finer uniform partitions of the unit interval, and then reorder the coordinates in each dimension.
- Latin hypercube sequences. Produced by the lhsdesign function. Though not quasi-random in the sense of minimizing discrepancy, these sequences nevertheless produce sparse uniform samples useful in experimental designs.

Quasi-random sequences are functions from the positive integers to the unit hypercube. To be useful in application, an initial point set of a sequence must be generated. Point sets are matrices of size $n$-by- $d$, where $n$ is the number of points and $d$ is the dimension of the hypercube being sampled. The functions haltonset and sobolset construct point sets with properties of a specified quasi-random sequence. Initial segments of the point sets are generated by the net method of the qrandset class (parent class of the haltonset class and sobolset class), but points can be generated and accessed more generally using parenthesis indexing.

Because of the way in which quasi-random sequences are generated, they may contain undesirable correlations, especially in their initial segments, and especially in higher dimensions. To address this issue, quasi-random point sets often skip, leap over, or scramble values in a sequence. The haltonset and sobolset functions allow you to specify both a Skip and a Leap property of a quasi-random sequence, and the scramble method of the qrandset class allows you apply a variety of scrambling techniques. Scrambling reduces correlations while also improving uniformity.

## Example: Generate a Quasi-Random Point Set

This example uses haltonset to construct a 2-D Halton point set-an object, p, of the haltonset class-that skips the first 1000 values of the sequence and then retains every 101st point:

```
p = haltonset(2,'Skip',1e3,'Leap',1e2)
p =
    Halton point set in 2 dimensions (8.918019e+013 points)
    Properties:
                            Skip : 1000
        Leap : 100
    ScrambleMethod : none
```

The object $p$ encapsulates properties of the specified quasi-random sequence. The point set is finite, with a length determined by the Skip and Leap properties and by limits on the size of point set indices (maximum value of $2^{53}$ ).

Use scramble to apply reverse-radix scrambling:

```
p = scramble(p,'RR2')
p =
    Halton point set in 2 dimensions (8.918019e+013 points)
    Properties:
            Skip : }100
            Leap : 100
    ScrambleMethod : RR2
```

Use net to generate the first 500 points:

```
XO = net(p,500);
```

This is equivalent to:

```
XO = p(1:500,:);
```

Values of the point set X0 are not generated and stored in memory until you access $p$ using net or parenthesis indexing.

To appreciate the nature of quasi-random numbers, create a scatter of the two dimensions in XO :

```
scatter(XO(:,1),XO(:,2),5,'r')
axis square
title('{\bf Quasi-Random Scatter}')
```


## Quasi-Random Scatter



Compare this to a scatter of uniform pseudorandom numbers generated by the MATLAB rand function:

```
X = rand(500,2);
scatter(X(:,1),X(:,2),5,'b')
axis square
title('{\bf Uniform Random Scatter}')
```



The quasi-random scatter appears more uniform, avoiding the clumping in the pseudorandom scatter.

In a statistical sense, quasi-random numbers are too uniform to pass traditional tests of randomness. For example, a Kolmogorov-Smirnov test, performed by kstest, is used to assess whether or not a point set has a uniform random distribution. When performed repeatedly on uniform pseudorandom samples, such as those generated by rand, the test produces a uniform distribution of $p$-values:

```
nTests = 1e5;
sampSize = 50;
PVALS = zeros(nTests,1);
for test = 1:nTests
    x = rand(sampSize,1);
    [h,pval] = kstest(x,[x,x]);
```



The results are quite different when the test is performed repeatedly on uniform quasi-random samples:

```
p = haltonset(1,'Skip',1e3,'Leap',1e2);
p = scramble(p,'RR2');
nTests = 1e5;
sampSize = 50;
PVALS = zeros(nTests,1);
for test = 1:nTests
```



Small $p$-values call into question the null hypothesis that the data are uniformly distributed. If the hypothesis is true, about $5 \%$ of the $p$-values are expected to fall below 0.05 . The results are remarkably consistent in their failure to challenge the hypothesis.

## Quasi-Random Streams

Quasi-random streams, produced by the qrandstream function, are used to generate sequential quasi-random outputs, rather than point sets of a specific size. Streams are used like pseudoRNGS, such as rand, when client applications require a source of quasi-random numbers of indefinite size that can be accessed intermittently. Properties of a quasi-random stream, such as its type (Halton or Sobol), dimension, skip, leap, and scramble, are set when the stream is constructed.

In implementation, quasi-random streams are essentially very large quasi-random point sets, though they are accessed differently. The state of a quasi-random stream is the scalar index of the next point to be taken from the stream. Use the qrand method of the qrandstream class to generate points from the stream, starting from the current state. Use the reset method to reset the state to 1 . Unlike point sets, streams do not support parenthesis indexing.

## Example: Generate a Quasi-Random Stream

For example, the following code, taken from the example at the end of "Quasi-Random Point Sets" on page 6-16, uses haltonset to create a quasi-random point set $p$, and then repeatedly increments the index into the point set, test, to generate different samples:

```
p = haltonset(1,'Skip',1e3,'Leap',1e2);
p = scramble(p,'RR2');
nTests = 1e5;
sampSize = 50;
PVALS = zeros(nTests,1);
for test = 1:nTests
    x = p(test:test+(sampSize-1),:);
    [h,pval] = kstest(x,[x,x]);
    PVALS(test) = pval;
end
```

The same results are obtained by using qrandstream to construct a quasi-random stream $q$ based on the point set $p$ and letting the stream take care of increments to the index:

```
p = haltonset(1,'Skip',1e3,'Leap',1e2);
p = scramble(p,'RR2');
q = qrandstream(p)
nTests = 1e5;
sampSize = 50;
PVALS = zeros(nTests,1);
for test = 1:nTests
    X = qrand(q,sampSize);
    [h,pval] = kstest(X,[X,X]);
    PVALS(test) = pval;
end
```


# Generating Data Using Flexible Families of Distributions 

In this section...<br>"Pearson and Johnson Systems" on page 6-25<br>"Generating Data Using the Pearson System" on page 6-26<br>"Generating Data Using the Johnson System" on page 6-28

## Pearson and Johnson Systems

As described in "Using Probability Distributions" on page 5-2, choosing an appropriate parametric family of distributions to model your data can be based on a priori or a posteriori knowledge of the data-producing process, but the choice is often difficult. The Pearson and Johnson systems can make such a choice unnecessary. Each system is a flexible parametric family of distributions that includes a wide range of distribution shapes, and it is often possible to find a distribution within one of these two systems that provides a good match to your data.

## Data Input

The following parameters define each member of the Pearson and Johnson systems

- Mean - Estimated by mean
- Standard deviation - Estimated by std
- Skewness - Estimated by skewness
- Kurtosis - Estimated by kurtosis

These statistics can also be computed with the moment function. The Johnson system, while based on these four parameters, is more naturally described using quantiles, estimated by the quantile function.

The Statistics Toolbox functions pearsrnd and johnsrnd take input arguments defining a distribution (parameters or quantiles, respectively) and return the type and the coefficients of the distribution in the corresponding system. Both functions also generate random numbers from the specified distribution.

As an example, load the data in carbig.mat, which includes a variable MPG containing measurements of the gas mileage for each car.

```
load carbig
MPG = MPG(~isnan(MPG));
[n,x] = hist(MPG,15);
bar(x,n)
set(get(gca,'Children'),'FaceColor',[.8 . 8 1])
```



The following two sections model the distribution with members of the Pearson and Johnson systems, respectively.

## Generating Data Using the Pearson System

The statistician Karl Pearson devised a system, or family, of distributions that includes a unique distribution corresponding to every valid combination of mean, standard deviation, skewness, and kurtosis. If you compute sample
values for each of these moments from data, it is easy to find the distribution in the Pearson system that matches these four moments and to generate a random sample.

The Pearson system embeds seven basic types of distribution together in a single parametric framework. It includes common distributions such as the normal and $t$ distributions, simple transformations of standard distributions such as a shifted and scaled beta distribution and the inverse gamma distribution, and one distribution-the Type IV-that is not a simple transformation of any standard distribution.

For a given set of moments, there are distributions that are not in the system that also have those same first four moments, and the distribution in the Pearson system may not be a good match to your data, particularly if the data are multimodal. But the system does cover a wide range of distribution shapes, including both symmetric and skewed distributions.

To generate a sample from the Pearson distribution that closely matches the MPG data, simply compute the four sample moments and treat those as distribution parameters.

```
moments = {mean(MPG),std(MPG),skewness(MPG),kurtosis(MPG)};
[r,type] = pearsrnd(moments{:},10000,1);
```

The optional second output from pearsrnd indicates which type of distribution within the Pearson system matches the combination of moments.

```
type
type =
```

    1
    In this case, pearsrnd has determined that the data are best described with a Type I Pearson distribution, which is a shifted, scaled beta distribution.

Verify that the sample resembles the original data by overlaying the empirical cumulative distribution functions.

```
ecdf(MPG);
[Fi,xi] = ecdf(r);
hold on, stairs(xi,Fi,'r'); hold off
```



## Generating Data Using the Johnson System

Statistician Norman Johnson devised a different system of distributions that also includes a unique distribution for every valid combination of mean, standard deviation, skewness, and kurtosis. However, since it is more natural to describe distributions in the Johnson system using quantiles, working with this system is different than working with the Pearson system.

The Johnson system is based on three possible transformations of a normal random variable, plus the identity transformation. The three nontrivial cases are known as SL, SU, and SB, corresponding to exponential, logistic, and hyperbolic sine transformations. All three can be written as

$$
\mathrm{X}=\gamma+\delta \cdot \Gamma\left(\frac{(\mathrm{Z}-\xi)}{\lambda}\right)
$$

where Z is a standard normal random variable, $\Gamma$ is the transformation, and $\gamma, \delta, \xi$, and $\lambda$ are scale and location parameters. The fourth case, SN , is the identity transformation.

To generate a sample from the Johnson distribution that matches the MPG data, first define the four quantiles to which the four evenly spaced standard normal quantiles of $-1.5,-0.5,0.5$, and 1.5 should be transformed. That is, you compute the sample quantiles of the data for the cumulative probabilities of $0.067,0.309,0.691$, and 0.933 .

```
probs = normcdf([-1.5 -0.5 0.5 1.5])
probs =
    0.066807 0.30854 0.69146 0.93319
quantiles = quantile(MPG,probs)
quantiles =
    13.0000 18.0000 27.2000 36.0000
```

Then treat those quantiles as distribution parameters.

```
[r1,type] = johnsrnd(quantiles,10000,1);
```

The optional second output from johnsrnd indicates which type of distribution within the Johnson system matches the quantiles.

```
type
type =
SB
```

You can verify that the sample resembles the original data by overlaying the empirical cumulative distribution functions.

```
ecdf(MPG);
[Fi,xi] = ecdf(r1);
hold on, stairs(xi,Fi,'r'); hold off
```



In some applications, it may be important to match the quantiles better in some regions of the data than in others. To do that, specify four evenly spaced standard normal quantiles at which you want to match the data, instead of the default $-1.5,-0.5,0.5$, and 1.5 . For example, you might care more about matching the data in the right tail than in the left, and so you specify standard normal quantiles that emphasizes the right tail.

```
qnorm = [-.5 . 25 1 1.75];
probs = normcdf(qnorm);
qemp = quantile(MPG,probs);
r2 = johnsrnd([qnorm; qemp],10000,1);
```

However, while the new sample matches the original data better in the right tail, it matches much worse in the left tail.

```
[Fj,xj] = ecdf(r2);
hold on, stairs(xj,Fj,'g'); hold off
```



## Hypothesis Tests

- "Introduction to Hypothesis Tests" on page 7-2
- "Hypothesis Test Terminology" on page 7-3
- "Hypothesis Test Assumptions" on page 7-5
- "Example: Hypothesis Testing" on page 7-7
- "Available Hypothesis Tests" on page 7-13


## Introduction to Hypothesis Tests

Hypothesis testing is a common method of drawing inferences about a population based on statistical evidence from a sample.

As an example, suppose someone says that at a certain time in the state of Massachusetts the average price of a gallon of regular unleaded gas was $\$ 1.15$. How could you determine the truth of the statement? You could try to find prices at every gas station in the state at the time. That approach would be definitive, but it could be time-consuming, costly, or even impossible.

A simpler approach would be to find prices at a small number of randomly selected gas stations around the state, and then compute the sample average.

Sample averages differ from one another due to chance variability in the selection process. Suppose your sample average comes out to be $\$ 1.18$. Is the $\$ 0.03$ difference an artifact of random sampling or significant evidence that the average price of a gallon of gas was in fact greater than $\$ 1.15$ ? Hypothesis testing is a statistical method for making such decisions.

## Hypothesis Test Terminology

All hypothesis tests share the same basic terminology and structure.

- A null hypothesis is an assertion about a population that you would like to test. It is "null" in the sense that it often represents a status quo belief, such as the absence of a characteristic or the lack of an effect. It may be formalized by asserting that a population parameter, or a combination of population parameters, has a certain value. In the example given in the "Introduction to Hypothesis Tests" on page 7-2, the null hypothesis would be that the average price of gas across the state was $\$ 1.15$. This is written $H_{0}: \mu=1.15$.
- An alternative hypothesis is a contrasting assertion about the population that can be tested against the null hypothesis. In the example given in the "Introduction to Hypothesis Tests" on page 7-2, possible alternative hypotheses are:
$H_{1}: \mu \neq 1.15$ - State average was different from $\$ 1.15$ (two-tailed test)
$H_{1}: \mu>1.15$ - State average was greater than $\$ 1.15$ (right-tail test)
$H_{1}: \mu<1.15$ - State average was less than $\$ 1.15$ (left-tail test)
- To conduct a hypothesis test, a random sample from the population is collected and a relevant test statistic is computed to summarize the sample. This statistic varies with the type of test, but its distribution under the null hypothesis must be known (or assumed).
- The $p$ value of a test is the probability, under the null hypothesis, of obtaining a value of the test statistic as extreme or more extreme than the value computed from the sample.
- The significance level of a test is a threshold of probability $a$ agreed to before the test is conducted. A typical value of $\alpha$ is 0.05 . If the $p$ value of a test is less than $a$, the test rejects the null hypothesis. If the $p$ value is greater than $a$, there is insufficient evidence to reject the null hypothesis. Note that lack of evidence for rejecting the null hypothesis is not evidence for accepting the null hypothesis. Also note that substantive "significance" of an alternative cannot be inferred from the statistical significance of a test.
- The significance level $\alpha$ can be interpreted as the probability of rejecting the null hypothesis when it is actually true-a type I error. The distribution of the test statistic under the null hypothesis determines the probability
$\alpha$ of a type I error. Even if the null hypothesis is not rejected, it may still be false-a type II error. The distribution of the test statistic under the alternative hypothesis determines the probability $\beta$ of a type II error. Type II errors are often due to small sample sizes. The power of a test, $1-\beta$, is the probability of correctly rejecting a false null hypothesis.
- Results of hypothesis tests are often communicated with a confidence interval. A confidence interval is an estimated range of values with a specified probability of containing the true population value of a parameter. Upper and lower bounds for confidence intervals are computed from the sample estimate of the parameter and the known (or assumed) sampling distribution of the estimator. A typical assumption is that estimates will be normally distributed with repeated sampling (as dictated by the Central Limit Theorem). Wider confidence intervals correspond to poor estimates (smaller samples); narrow intervals correspond to better estimates (larger samples). If the null hypothesis asserts the value of a population parameter, the test rejects the null hypothesis when the hypothesized value lies outside the computed confidence interval for the parameter.


## Hypothesis Test Assumptions

Different hypothesis tests make different assumptions about the distribution of the random variable being sampled in the data. These assumptions must be considered when choosing a test and when interpreting the results.

For example, the $z$-test (ztest) and the $t$-test (ttest) both assume that the data are independently sampled from a normal distribution. Statistics Toolbox functions are available for testing this assumption, such as chi2gof, jbtest, lillietest, and normplot.

Both the $z$-test and the $t$-test are relatively robust with respect to departures from this assumption, so long as the sample size $n$ is large enough. Both tests compute a sample mean $\bar{x}$, which, by the Central Limit Theorem, has an approximately normal sampling distribution with mean equal to the population mean $\mu$, regardless of the population distribution being sampled.

The difference between the $z$-test and the $t$-test is in the assumption of the standard deviation $\sigma$ of the underlying normal distribution. A $z$-test assumes that $\sigma$ is known; a $t$-test does not. As a result, a $t$-test must compute an estimate $s$ of the standard deviation from the sample.

Test statistics for the $z$-test and the $t$-test are, respectively,

$$
\begin{aligned}
& z=\frac{\bar{x}-\mu}{\sigma / \sqrt{n}} \\
& t=\frac{\bar{x}-\mu}{s / \sqrt{n}}
\end{aligned}
$$

Under the null hypothesis that the population is distributed with mean $\mu$, the $z$-statistic has a standard normal distribution, $N(0,1)$. Under the same null hypothesis, the $t$-statistic has Student's $t$ distribution with $n-1$ degrees of freedom. For small sample sizes, Student's $t$ distribution is flatter and wider than $N(0,1)$, compensating for the decreased confidence in the estimate $s$. As sample size increases, however, Student's $t$ distribution approaches the standard normal distribution, and the two tests become essentially equivalent.

Knowing the distribution of the test statistic under the null hypothesis allows for accurate calculation of $p$-values. Interpreting $p$-values in the context of the test assumptions allows for critical analysis of test results.

Assumptions underlying Statistics Toolbox hypothesis tests are given in the reference pages for implementing functions.

## Example: Hypothesis Testing

This example uses the gas price data in the file gas.mat. The file contains two random samples of prices for a gallon of gas around the state of Massachusetts in 1993. The first sample, price1, contains 20 random observations around the state on a single day in January. The second sample, price2, contains 20 random observations around the state one month later.

```
load gas
prices = [price1 price2];
```

As a first step, you might want to test the assumption that the samples come from normal distributions.

A normal probability plot gives a quick idea.
normplot(prices)


Both scatters approximately follow straight lines through the first and third quartiles of the samples, indicating approximate normal distributions. The February sample (the right-hand line) shows a slight departure from normality in the lower tail. A shift in the mean from January to February is evident.

A hypothesis test is used to quantify the test of normality. Since each sample is relatively small, a Lilliefors test is recommended.

```
lillietest(price1)
ans =
    0
lillietest(price2)
ans =
    0
```

The default significance level of lillietest is $5 \%$. The logical 0 returned by each test indicates a failure to reject the null hypothesis that the samples are normally distributed. This failure may reflect normality in the population or it may reflect a lack of strong evidence against the null hypothesis due to the small sample size.

Now compute the sample means:

```
sample_means = mean(prices)
sample_means =
    115.1500 118.5000
```

You might want to test the null hypothesis that the mean price across the state on the day of the January sample was $\$ 1.15$. If you know that the standard deviation in prices across the state has historically, and consistently, been $\$ 0.04$, then a $z$-test is appropriate.

```
[h,pvalue,ci] = ztest(price1/100,1.15,0.04)
h =
    0
pvalue =
    0.8668
ci =
    1.1340
    1.1690
```

The logical output $\mathrm{h}=0$ indicates a failure to reject the null hypothesis at the default significance level of $5 \%$. This is a consequence of the high probability under the null hypothesis, indicated by the $p$ value, of observing a value as extreme or more extreme of the $z$-statistic computed from the sample. The $95 \%$ confidence interval on the mean [1.1340 1.1690] includes the hypothesized population mean of $\$ 1.15$.

Does the later sample offer stronger evidence for rejecting a null hypothesis of a state-wide average price of $\$ 1.15$ in February? The shift shown in the probability plot and the difference in the computed sample means suggest this. The shift might indicate a significant fluctuation in the market, raising questions about the validity of using the historical standard deviation. If a known standard deviation cannot be assumed, a $t$-test is more appropriate.

```
[h,pvalue,ci] = ttest(price2/100,1.15)
h =
    1
pvalue =
    4.9517e-004
ci =
    1.1675
    1.2025
```

The logical output $\mathrm{h}=1$ indicates a rejection of the null hypothesis at the default significance level of $5 \%$. In this case, the $95 \%$ confidence interval on the mean does not include the hypothesized population mean of $\$ 1.15$.

You might want to investigate the shift in prices a little more closely. The function ttest2 tests if two independent samples come from normal distributions with equal but unknown standard deviations and the same mean, against the alternative that the means are unequal.

```
[h,sig,ci] = ttest2(price1,price2)
h =
    1
sig =
    0.0083
ci =
    -5.7845
    -0.9155
```

The null hypothesis is rejected at the default $5 \%$ significance level, and the confidence interval on the difference of means does not include the hypothesized value of 0 .

A notched box plot is another way to visualize the shift.


The plot displays the distribution of the samples around their medians. The heights of the notches in each box are computed so that the side-by-side boxes have nonoverlapping notches when their medians are different at a default $5 \%$ significance level. The computation is based on an assumption of normality in the data, but the comparison is reasonably robust for other distributions. The side-by-side plots provide a kind of visual hypothesis test, comparing medians rather than means. The plot above appears to barely reject the null hypothesis of equal medians.

The nonparametric Wilcoxon rank sum test, implemented by the function ranksum, can be used to quantify the test of equal medians. It tests if two independent samples come from identical continuous (not necessarily normal) distributions with equal medians, against the alternative that they do not have equal medians.

```
[p,h] = ranksum(price1,price2)
\(p\) =
    0.0095
\(\mathrm{h}=\)
    1
```

The test rejects the null hypothesis of equal medians at the default $5 \%$ significance level.

## Available Hypothesis Tests

| Function | Description |
| :--- | :--- |
| ansaribradley | Ansari-Bradley test. Tests if two independent samples <br> come from the same distribution, against the alternative <br> that they come from distributions that have the same <br> median and shape but different variances. |
| chi2gof | Chi-square goodness-of-fit test. Tests if a sample comes <br> from a specified distribution, against the alternative <br> that it does not come from that distribution. |
| dwtest | Durbin-Watson test. Tests if the residuals from a linear <br> regression are uncorrelated, against the alternative <br> that there is autocorrelation among them. |
| jbtest | Jarque-Bera test. Tests if a sample comes from a <br> normal distribution with unknown mean and variance, <br> against the alternative that it does not come from a <br> normal distribution. |
| kstest | One-sample Kolmogorov-Smirnov test. Tests if a sample <br> comes from a continuous distribution with specified <br> parameters, against the alternative that it does not <br> come from that distribution. |
| kstest2 | Two-sample Kolmogorov-Smirnov test. Tests if two <br> samples come from the same continuous distribution, <br> against the alternative that they do not come from the <br> same distribution. |
| lillietest | Lilliefors test. Tests if a sample comes from a <br> distribution in the normal family, against the <br> alternative that it does not come from a normal <br> distribution. |
| linhyptest | Linear hypothesis test. Tests if H*b = c for parameter <br> estimates b with estimated covariance H and specified <br> c, against the alternative that H*b $\quad$ c. |


| Function | Description |
| :--- | :--- |
| ranksum | Wilcoxon rank sum test. Tests if two independent <br> samples come from identical continuous distributions <br> with equal medians, against the alternative that they <br> do not have equal medians. |
| runstest | Runs test. Tests if a sequence of values comes in <br> random order, against the alternative that the ordering <br> is not random. |
| signrank | One-sample or paired-sample Wilcoxon signed rank test. <br> Tests if a sample comes from a continuous distribution <br> symmetric about a specified median, against the <br> alternative that it does not have that median. |
| signtest | One-sample or paired-sample sign test. Tests if a <br> sample comes from an arbitrary continuous distribution <br> with a specified median, against the alternative that it <br> does not have that median. |
| ttest | One-sample or paired-sample $t$-test. Tests if a sample <br> comes from a normal distribution with unknown <br> variance and a specified mean, against the alternative <br> that it does not have that mean. |
| ttest2 | Two-sample $t$-test. Tests if two independent samples <br> come from normal distributions with unknown but <br> equal (or, optionally, unequal) variances and the same <br> mean, against the alternative that the means are <br> unequal. |
| vartest2 | One-sample chi-square variance test. Tests if a sample <br> comes from a normal distribution with specified <br> variance, against the alternative that it comes from a <br> normal distribution with a different variance. |
| vartest | Two-sample $F$-test for equal variances. Tests if two <br> independent samples come from normal distributions <br> with the same variance, against the alternative that <br> they come from normal distributions with different <br> variances. |


| Function | Description |
| :--- | :--- |
| vartestn | Bartlett multiple-sample test for equal variances. Tests <br> if multiple samples come from normal distributions <br> with the same variance, against the alternative that <br> they come from normal distributions with different <br> variances. |
| ztest | One-sample $z$-test. Tests if a sample comes from a <br> normal distribution with known variance and specified <br> mean, against the alternative that it does not have that <br> mean. |

Note In addition to the previous functions, Statistics Toolbox functions are available for analysis of variance (ANOVA), which perform hypothesis tests in the context of linear modeling. These functions are discussed in Chapter 8, "Analysis of Variance".

## Analysis of Variance

- "Introduction to Analysis of Variance" on page 8-2
- "ANOVA" on page 8-3
- "MANOVA" on page 8-39


## Introduction to Analysis of Variance

Analysis of variance (ANOVA) is a procedure for assigning sample variance to different sources and deciding whether the variation arises within or among different population groups. Samples are described in terms of variation around group means and variation of group means around an overall mean. If variations within groups are small relative to variations between groups, a difference in group means may be inferred. Chapter 7, "Hypothesis Tests" are used to quantify decisions.

This chapter treats ANOVA among groups, that is, among categorical predictors. ANOVA for regression, with continuous predictors, is discussed in "Tabulating Diagnostic Statistics" on page 9-13.

Multivariate analysis of variance (MANOVA), for data with multiple measured responses, is also discussed in this chapter.

## ANOVA

## In this section...

"One-Way ANOVA" on page 8-3
"Two-Way ANOVA" on page 8-9
"N-Way ANOVA" on page 8-12
"Other ANOVA Models" on page 8-26
"Analysis of Covariance" on page 8-27
"Nonparametric Methods" on page 8-35

## One-Way ANOVA

- "Introduction to One-Way ANOVA" on page 8-3
- "Example: One-Way ANOVA" on page 8-4
- "Multiple Comparisons" on page 8-6
- "Example: Multiple Comparisons" on page 8-7


## Introduction to One-Way ANOVA

The purpose of one-way ANOVA is to find out whether data from several groups have a common mean. That is, to determine whether the groups are actually different in the measured characteristic.

One-way ANOVA is a simple special case of the linear model. The one-way ANOVA form of the model is

$$
y_{i j}=\alpha_{. j}+\varepsilon_{i j}
$$

where:

- $y_{i j}$ is a matrix of observations in which each column represents a different group.
- $\alpha_{. j}$ is a matrix whose columns are the group means. (The "dot $j$ " notation means that $\alpha$ applies to all rows of column $j$. That is, the value $\alpha_{i j}$ is the same for all i.)
- $\varepsilon_{i j}$ is a matrix of random disturbances.

The model assumes that the columns of $y$ are a constant plus a random disturbance. You want to know if the constants are all the same.

## Example: One-Way ANOVA

The data below comes from a study by Hogg and Ledolter [48] of bacteria counts in shipments of milk. The columns of the matrix hogg represent different shipments. The rows are bacteria counts from cartons of milk chosen randomly from each shipment. Do some shipments have higher counts than others?

```
load hogg
hogg
hogg =
    24
    15
    21
    27
```




```
[p,tbl,stats] = anova1(hogg);
p
p =
    1.1971e-04
```

The standard ANOVA table has columns for the sums of squares, degrees of freedom, mean squares (SS/df), $F$ statistic, and $p$ value.


You can use the $F$ statistic to do a hypothesis test to find out if the bacteria counts are the same. anova1 returns the $p$ value from this hypothesis test.

In this case the $p$ value is about 0.0001 , a very small value. This is a strong indication that the bacteria counts from the different shipments are not the same. An $F$ statistic as extreme as the observed $F$ would occur by chance only once in 10,000 times if the counts were truly equal.

The $p$ value returned by anova1 depends on assumptions about the random disturbances $\varepsilon_{i j}$ in the model equation. For the $p$ value to be correct, these disturbances need to be independent, normally distributed, and have constant variance.

You can get some graphical assurance that the means are different by looking at the box plots in the second figure window displayed by anova1. Note, however, that the notches are used for a comparison of medians, not a comparison of means. For more information on this display, see "Box Plots" on page 4-6.


## Multiple Comparisons

Sometimes you need to determine not just whether there are any differences among the means, but specifically which pairs of means are significantly different. It is tempting to perform a series of $t$ tests, one for each pair of means, but this procedure has a pitfall.

In a $t$ test, you compute a $t$ statistic and compare it to a critical value. The critical value is chosen so that when the means are really the same (any apparent difference is due to random chance), the probability that the $t$ statistic will exceed the critical value is small, say $5 \%$. When the means are different, the probability that the statistic will exceed the critical value is larger.

In this example there are five means, so there are 10 pairs of means to compare. It stands to reason that if all the means are the same, and if there is a $5 \%$ chance of incorrectly concluding that there is a difference in one pair,
then the probability of making at least one incorrect conclusion among all 10 pairs is much larger than $5 \%$.

Fortunately, there are procedures known as multiple comparison procedures that are designed to compensate for multiple tests.

## Example: Multiple Comparisons

You can perform a multiple comparison test using the multcompare function and supplying it with the stats output from anova1.

```
load hogg
[p,tbl,stats] = anova1(hogg);
[c,m] = multcompare(stats)
c =
\begin{tabular}{rrrrr}
1.0000 & 2.0000 & 2.4953 & 10.5000 & 18.5047 \\
1.0000 & 3.0000 & 4.1619 & 12.1667 & 20.1714 \\
1.0000 & 4.0000 & 6.6619 & 14.6667 & 22.6714 \\
1.0000 & 5.0000 & -2.0047 & 6.0000 & 14.0047 \\
2.0000 & 3.0000 & -6.3381 & 1.6667 & 9.6714 \\
2.0000 & 4.0000 & -3.8381 & 4.1667 & 12.1714 \\
2.0000 & 5.0000 & -12.5047 & -4.5000 & 3.5047 \\
3.0000 & 4.0000 & -5.5047 & 2.5000 & 10.5047 \\
3.0000 & 5.0000 & -14.1714 & -6.1667 & 1.8381 \\
4.0000 & 5.0000 & -16.6714 & -8.6667 & -0.6619
\end{tabular}
m =
    23.8333 1.9273
    13.3333 1.9273
    11.6667 1.9273
    9.1667 1.9273
    17.8333 1.9273
```

The first output from multcompare has one row for each pair of groups, with an estimate of the difference in group means and a confidence interval for that group. For example, the second row has the values

$$
\begin{array}{lllll}
1.0000 & 3.0000 & 4.1619 & 12.1667 & 20.1714
\end{array}
$$

indicating that the mean of group 1 minus the mean of group 3 is estimated to be 12.1667 , and a $95 \%$ confidence interval for this difference is
[4.1619, 20.1714]. This interval does not contain 0 , so you can conclude that the means of groups 1 and 3 are different.

The second output contains the mean and its standard error for each group.
It is easier to visualize the difference between group means by looking at the graph that multcompare produces.

There are five groups. The graph instructs you to Click on the group you want to test. Three groups have slopes significantly different from group one.

Click on the group you want to test


The graph shows that group 1 is significantly different from groups 2,3 , and 4. By using the mouse to select group 4, you can determine that it is also significantly different from group 5. Other pairs are not significantly different.

## Two-Way ANOVA

- "Introduction to Two-Way ANOVA" on page 8-9
- "Example: Two-Way ANOVA" on page 8-10


## Introduction to Two-Way ANOVA

The purpose of two-way ANOVA is to find out whether data from several groups have a common mean. One-way ANOVA and two-way ANOVA differ in that the groups in two-way ANOVA have two categories of defining characteristics instead of one.

Suppose an automobile company has two factories, and each factory makes the same three models of car. It is reasonable to ask if the gas mileage in the cars varies from factory to factory as well as from model to model. There are two predictors, factory and model, to explain differences in mileage.

There could be an overall difference in mileage due to a difference in the production methods between factories. There is probably a difference in the mileage of the different models (irrespective of the factory) due to differences in design specifications. These effects are called additive.

Finally, a factory might make high mileage cars in one model (perhaps because of a superior production line), but not be different from the other factory for other models. This effect is called an interaction. It is impossible to detect an interaction unless there are duplicate observations for some combination of factory and car model.

Two-way ANOVA is a special case of the linear model. The two-way ANOVA form of the model is

$$
y_{i j k}=\mu+\alpha_{. j}+\beta_{i .}+\gamma_{i j}+\varepsilon_{i j k}
$$

where, with respect to the automobile example above:

- $y_{i j k}$ is a matrix of gas mileage observations (with row index $i$, column index $j$, and repetition index $k$ ).
- $\mu$ is a constant matrix of the overall mean gas mileage.
- $a_{j}$ is a matrix whose columns are the deviations of each car's gas mileage (from the mean gas mileage $\mu$ ) that are attributable to the car's model. All values in a given column of $\alpha_{. j}$ are identical, and the values in each row of $a_{. j}$ sum to 0 .
- $\beta_{i .}$ is a matrix whose rows are the deviations of each car's gas mileage (from the mean gas mileage $\mu$ ) that are attributable to the car's factory. All values in a given row of $\beta_{i}$ are identical, and the values in each column of $\beta_{i}$. sum to 0 .
- $\gamma_{i j}$ is a matrix of interactions. The values in each row of $\gamma_{i j}$ sum to 0 , and the values in each column of $\gamma_{i j}$ sum to 0 .
- $\varepsilon_{i j k}$ is a matrix of random disturbances.


## Example: Two-Way ANOVA

The purpose of the example is to determine the effect of car model and factory on the mileage rating of cars.

```
load mileage
mileage
mileage =
    33.3000 34.5000 37.4000
    33.4000 34.8000 36.8000
    32.9000 33.8000 37.6000
    32.6000 33.4000 36.6000
    32.5000 33.7000 37.0000
    33.0000 33.9000 36.7000
cars = 3;
[p,tbl,stats] = anova2(mileage,cars);
p
p =
    0.0000 0.0039 0.8411
```

There are three models of cars (columns) and two factories (rows). The reason there are six rows in mileage instead of two is that each factory provides
three cars of each model for the study. The data from the first factory is in the first three rows, and the data from the second factory is in the last three rows.

The standard ANOVA table has columns for the sums of squares, degrees-of-freedom, mean squares (SS/df), $F$ statistics, and $p$-values.

| - Figure No. 1: Two-way ANOVA |  |  |  | - $\square_{\text {a }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| File Edit Iools Window Help |  |  |  |  |  |
| ANOVA Table |  |  |  |  |  |
| Source | 55 | df | MS | F | Prob>F $\triangle$ |
| Columns | 53.3511 | 2 | 26.6756 | 234.22 | 0 |
| Rows | 1.445 | 1 | 1.445 | 12.69 | 0.0039 |
| Interaction | 0.04 | 2 | 0.02 | 0.18 | 0.8411 |
| Error | 1.3667 | 12 | 0.1139 |  |  |
| Total | 56.2028 | 17 |  |  | $\square$ |

You can use the $F$ statistics to do hypotheses tests to find out if the mileage is the same across models, factories, and model-factory pairs (after adjusting for the additive effects). anova2 returns the $p$ value from these tests.

The $p$ value for the model effect is zero to four decimal places. This is a strong indication that the mileage varies from one model to another. An $F$ statistic as extreme as the observed $F$ would occur by chance less than once in 10,000 times if the gas mileage were truly equal from model to model. If you used the multcompare function to perform a multiple comparison test, you would find that each pair of the three models is significantly different.

The $p$ value for the factory effect is 0.0039 , which is also highly significant. This indicates that one factory is out-performing the other in the gas mileage of the cars it produces. The observed $p$ value indicates that an $F$ statistic as extreme as the observed $F$ would occur by chance about four out of 1000 times if the gas mileage were truly equal from factory to factory.

There does not appear to be any interaction between factories and models. The $p$ value, 0.8411 , means that the observed result is quite likely ( 84 out 100 times) given that there is no interaction.

The $p$-values returned by anova2 depend on assumptions about the random disturbances $\varepsilon_{i j k}$ in the model equation. For the $p$-values to be correct these disturbances need to be independent, normally distributed, and have constant variance.

In addition, anova2 requires that data be balanced, which in this case means there must be the same number of cars for each combination of model and factory. The next section discusses a function that supports unbalanced data with any number of predictors.

## N-Way ANOVA

- "Introduction to N-Way ANOVA" on page 8-12
- "N-Way ANOVA with a Small Data Set" on page 8-13
- "N-Way ANOVA with a Large Data Set" on page 8-15
- "ANOVA with Random Effects" on page 8-19


## Introduction to N-Way ANOVA

You can use N -way ANOVA to determine if the means in a set of data differ when grouped by multiple factors. If they do differ, you can determine which factors or combinations of factors are associated with the difference.

N-way ANOVA is a generalization of two-way ANOVA. For three factors, the model can be written

$$
y_{i j k l}=\mu+\alpha_{. j .}+\beta_{i . .}+\gamma_{. . k}+(\alpha \beta)_{i j .}+(\alpha \gamma)_{i . k}+(\beta \gamma)_{. j k}+(\alpha \beta \gamma)_{i j k}+\varepsilon_{i j k l}
$$

In this notation parameters with two subscripts, such as $(\alpha \beta)_{i j}$, represent the interaction effect of two factors. The parameter $(\alpha \beta \gamma)_{i j k}$ represents the three-way interaction. An ANOVA model can have the full set of parameters or any subset, but conventionally it does not include complex interaction terms unless it also includes all simpler terms for those factors. For example, one would generally not include the three-way interaction without also including all two-way interactions.

The anovan function performs N-way ANOVA. Unlike the anova1 and anova2 functions, anovan does not expect data in a tabular form. Instead, it expects a vector of response measurements and a separate vector (or text array) containing the values corresponding to each factor. This input data format is more convenient than matrices when there are more than two factors or when the number of measurements per factor combination is not constant.

## N-Way ANOVA with a Small Data Set

Consider the following two-way example using anova2.

```
m = [23 15 20;27 17 63;43 3 55;41 9 90]
m =
    23 15 20
    27 17 63
    43 3 55
    41 9 90
anova2(m,2)
ans =
    0.0197 0.2234 0.2663
```

The factor information is implied by the shape of the matrix $m$ and the number of measurements at each factor combination (2). Although anova2 does not actually require arrays of factor values, for illustrative purposes you could create them as follows.

```
cfactor = repmat(1:3,4,1)
cfactor =
    1 2
    1 2 3
    1 2 3
    1 2
rfactor = [ones(2,3); 2*ones(2,3)]
rfactor =
```

    \(1 \quad 1 \quad 1\)
    | 1 | 1 | 1 |
| :--- | :--- | :--- |
| 2 | 2 | 2 |
| 2 | 2 | 2 |

The cfactor matrix shows that each column of $m$ represents a different level of the column factor. The rfactor matrix shows that the top two rows of $m$ represent one level of the row factor, and bottom two rows of $m$ represent a second level of the row factor. In other words, each value $m(i, j)$ represents an observation at column factor level cfactor(i,j) and row factor level rfactor(i,j).

To solve the above problem with anovan, you need to reshape the matrices $m$, cfactor, and rfactor to be vectors.

```
m = m(:);
cfactor = cfactor(:);
rfactor = rfactor(:);
[m cfactor rfactor]
ans =
    23 1 1
    27 1 1
    43 1 2
    41 1 2
    15 2 1
    17 2 1
    3 2 2
    9 2 2
    20 3 1
    63 3 1
    55 3 2
    90 3 2
anovan(m,{cfactor rfactor},2)
ans =
    0.0197
```

0.2234
0.2663

## N-Way ANOVA with a Large Data Set

The previous example used anova2 to study a small data set measuring car mileage. This example illustrates how to analyze a larger set of car data with mileage and other information on 406 cars made between 1970 and 1982. First, load the data set and look at the variable names.
load carbig
whos

| Name | Size | Bytes | Class |
| :--- | :--- | ---: | :--- |
|  |  |  |  |
| Acceleration | $406 \times 1$ | 3248 | double array |
| Cylinders | $406 \times 1$ | 3248 | double array |
| Displacement | $406 \times 1$ | 3248 | double array |
| Horsepower | $406 \times 1$ | 3248 | double array |
| MPG | $406 \times 1$ | 3248 | double array |
| Model | $406 \times 36$ | 29232 | char array |
| Model_Year | $406 \times 1$ | 3248 | double array |
| Origin | $406 \times 7$ | 5684 | char array |
| Weight | $406 \times 1$ | 3248 | double array |
| cyl4 | $406 \times 5$ | 4060 | char array |
| org | $406 \times 7$ | 5684 | char array |
| when | $406 \times 5$ | 4060 | char array |

The example focusses on four variables. MPG is the number of miles per gallon for each of 406 cars (though some have missing values coded as NaN). The other three variables are factors: cyl4 (four-cylinder car or not), org (car originated in Europe, Japan, or the USA), and when (car was built early in the period, in the middle of the period, or late in the period).

First, fit the full model, requesting up to three-way interactions and Type 3 sums-of-squares.

```
varnames = {'Origin';'4Cyl';'MfgDate'};
anovan(MPG,{org cyl4 when},3,3,varnames)
ans =
    0.0000
```

NaN
0
0.7032
0.0001
0.2072
0.6990


Constrained (Type III) sums of squares. Terms marked with \# are not full rank.

Note that many terms are marked by a \# symbol as not having full rank, and one of them has zero degrees of freedom and is missing a $p$ value. This can happen when there are missing factor combinations and the model has higher-order terms. In this case, the cross-tabulation below shows that there are no cars made in Europe during the early part of the period with other than four cylinders, as indicated by the 0 in table $(2,1,1)$.

```
[table, chi2, p, factorvals] = crosstab(org,when,cyl4)
table(:,:,1) =
\begin{tabular}{rrr}
82 & 75 & 25 \\
0 & 4 & 3 \\
3 & 3 & 4
\end{tabular}
```

```
table(:,:,2) =
    12 22 38
    23 26 17
    12 25 32
chi2 =
    207.7689
p =
    0
factorvals =
\begin{tabular}{llr} 
'USA' & 'Early' & 'Other' \\
'Europe' & 'Mid' & 'Four' \\
'Japan' & 'Late' & {[]}
\end{tabular}
```

Consequently it is impossible to estimate the three-way interaction effects, and including the three-way interaction term in the model makes the fit singular.

Using even the limited information available in the ANOVA table, you can see that the three-way interaction has a $p$ value of 0.699 , so it is not significant. So this time you examine only two-way interactions.

```
[p,tbl,stats,terms] = anovan(MPG,{org cyl4 when},2,3,varnames);
terms
\begin{tabular}{rll} 
terms \(=\) & & \\
1 & 0 & 0 \\
0 & 1 & 0 \\
0 & 0 & 1 \\
1 & 1 & 0 \\
1 & 0 & 1 \\
0 & 1 & 1
\end{tabular}
```

| - Figure No. 1: N-Way ANOVA |  |  |  | - |
| :---: | :---: | :---: | :---: | :---: |
| File Edit View Insert Iools Window Help Analysis of Variance |  |  |  |  |
|  |  |  |  |  |
| Source | Sum Sq. d.f | Mean Sq | F | Prob>F $\triangle$ |
| Origin | $532.6 \quad 2$ | 266.29 | 18.82 | 0 |
| 4 Cyl | 1769.81 | 1769.85 | 125.11 | 0 |
| MfgDate | 2887.12 | 1443.55 | 102.05 | 0 |
| Origin*4Cyl | 12.52 | 6.27 | 0.44 | 0.6422 |
| Origin*MfgDate | 350.44 | 87.59 | 6.19 | 0.0001 |
| $4 \mathrm{Cyl} * \mathrm{Mfg}$ Date | $31 \quad 2$ | 15.52 | 1.1 | 0.3348 |
| Error | 5432.1384 | 14.15 |  |  |
| Total | 24252.6397 |  |  |  |
| Constrained (Type III) sums of squares. |  |  |  |  |

Now all terms are estimable. The $p$-values for interaction term 4 (Origin*4Cyl) and interaction term 6 (4Cyl*MfgDate) are much larger than a typical cutoff value of 0.05 , indicating these terms are not significant. You could choose to omit these terms and pool their effects into the error term. The output terms variable returns a matrix of codes, each of which is a bit pattern representing a term. You can omit terms from the model by deleting their entries from terms and running anovan again, this time supplying the resulting vector as the model argument.

```
terms([4 6],:) = []
terms =
\begin{tabular}{lll}
1 & 0 & 0 \\
0 & 1 & 0 \\
0 & 0 & 1 \\
1 & 0 & 1
\end{tabular}
anovan(MPG,{org cyl4 when},terms,3,varnames)
ans =
    1.0e-003 *
```

0.0000

0
0
0.1140


Now you have a more parsimonious model indicating that the mileage of these cars seems to be related to all three factors, and that the effect of the manufacturing date depends on where the car was made.

## ANOVA with Random Effects

- "Introduction to ANOVA with Random Effects" on page 8-19
- "Setting Up the Model" on page 8-20
- "Fitting a Random Effects Model" on page 8-21
- "F Statistics for Models with Random Effects" on page 8-22
- "Variance Components" on page 8-24

Introduction to ANOVA with Random Effects. In an ordinary ANOVA model, each grouping variable represents a fixed factor. The levels of that factor are a fixed set of values. Your goal is to determine whether different factor levels lead to different response values. This section presents an example that shows how to use anovan to fit models where a factor's levels represent a random selection from a larger (infinite) set of possible levels.

Setting Up the Model. To set up the example, first load the data, which is stored in a 6 -by- 3 matrix, mileage.

```
load mileage
```

The anova2 function works only with balanced data, and it infers the values of the grouping variables from the row and column numbers of the input matrix. The anovan function, on the other hand, requires you to explicitly create vectors of grouping variable values. To create these vectors, do the following steps:

1 Create an array indicating the factory for each value in mileage. This array is 1 for the first column, 2 for the second, and 3 for the third.

```
factory = repmat(1:3,6,1);
```

2 Create an array indicating the car model for each mileage value. This array is 1 for the first three rows of mileage, and 2 for the remaining three rows.

```
carmod = [ones(3,3); 2*ones(3,3)];
```

3 Turn these matrices into vectors and display them.

```
mileage = mileage(:);
factory = factory(:);
carmod = carmod(:);
[mileage factory carmod]
ans =
```

| 33.3000 | 1.0000 | 1.0000 |
| :--- | :--- | :--- |
| 33.4000 | 1.0000 | 1.0000 |
| 32.9000 | 1.0000 | 1.0000 |
| 32.6000 | 1.0000 | 2.0000 |
| 32.5000 | 1.0000 | 2.0000 |
| 33.0000 | 1.0000 | 2.0000 |
| 34.5000 | 2.0000 | 1.0000 |
| 34.8000 | 2.0000 | 1.0000 |
| 33.8000 | 2.0000 | 1.0000 |
| 33.4000 | 2.0000 | 2.0000 |
| 33.7000 | 2.0000 | 2.0000 |


| 33.9000 | 2.0000 | 2.0000 |
| :--- | :--- | :--- |
| 37.4000 | 3.0000 | 1.0000 |
| 36.8000 | 3.0000 | 1.0000 |
| 37.6000 | 3.0000 | 1.0000 |
| 36.6000 | 3.0000 | 2.0000 |
| 37.0000 | 3.0000 | 2.0000 |
| 36.7000 | 3.0000 | 2.0000 |

Fitting a Random Effects Model. Continuing the example from the preceding section, suppose you are studying a few factories but you want information about what would happen if you build these same car models in a different factory-either one that you already have or another that you might construct. To get this information, fit the analysis of variance model, specifying a model that includes an interaction term and that the factory factor is random.

```
[pvals,tbl,stats] = anovan(mileage, {factory carmod}, ...
'model',2, 'random',1,'varnames',{'Factory' 'Car Model'});
```

Analysis of Variance

| Source | Sum Sq. | d.f. Mean Sq. | F | Prob>F |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Factory | 53.3511 | 2 | 26.6756 | 1333.78 | 0.0007 |
| Car Model | 1.445 | 1 | 1.445 | 72.25 | 0.0136 |
| Factory*Car Model | 0.04 | 2 | 0.02 | 0.18 | 0.8411 |
| Error | 1.3667 | 12 | 0.1139 |  |  |
| Total | 56.2028 | 17 |  |  |  |
|  |  |  |  |  |  |

In the fixed effects version of this fit, which you get by omitting the inputs ' random ' , 1 in the preceding code, the effect of car model is significant, with a $p$ value of 0.0039 . But in this example, which takes into account the random variation of the effect of the variable 'Car Model' from one factory to another, the effect is still significant, but with a higher $p$ value of 0.0136 .

F Statistics for Models with Random Effects. The $F$ statistic in a model having random effects is defined differently than in a model having all fixed effects. In the fixed effects model, you compute the $F$ statistic for any term by taking the ratio of the mean square for that term with the mean square for error. In a random effects model, however, some $F$ statistics use a different mean square in the denominator.

In the example described in "Setting Up the Model" on page 8-20, the effect of the variable 'Factory' could vary across car models. In this case, the interaction mean square takes the place of the error mean square in the $F$ statistic. The $F$ statistic for factory is:

```
F = 1.445 / 0.02
F =
```

72.2500

The degrees of freedom for the statistic are the degrees of freedom for the numerator (1) and denominator (2) mean squares. Therefore the $p$ value for the statistic is:

```
pval = 1 - fcdf(F,1,2)
pval =
    0.0136
```

With random effects, the expected value of each mean square depends not only on the variance of the error term, but also on the variances contributed by the random effects. You can see these dependencies by writing the expected values as linear combinations of contributions from the various model terms. To find the coefficients of these linear combinations, enter stats.ems, which returns the ems field of the stats structure:

```
stats.ems
ans =
    6.0000 0.0000 3.0000 1.0000
    0.0000 9.0000 3.0000 1.0000
```

| 0.0000 | 0.0000 | 3.0000 | 1.0000 |
| ---: | ---: | ---: | ---: |
| 0 | 0 | 0 | 1.0000 |

To see text representations of the linear combinations, enter

```
stats.txtems
ans =
    '6*V(Factory)+3*V(Factory*Car Model)+V(Error)'
    '9*Q(Car Model)+3*V(Factory*Car Model)+V(Error)'
    '3*V(Factory*Car Model)+V(Error)'
    'V(Error)'
```

The expected value for the mean square due to car model (second term) includes contributions from a quadratic function of the car model effects, plus three times the variance of the interaction term's effect, plus the variance of the error term. Notice that if the car model effects were all zero, the expression would reduce to the expected mean square for the third term (the interaction term). That is why the $F$ statistic for the car model effect uses the interaction mean square in the denominator.

In some cases there is no single term whose expected value matches the one required for the denominator of the $F$ statistic. In that case, the denominator is a linear combination of mean squares. The stats structure contains fields giving the definitions of the denominators for each $F$ statistic. The txtdenom field, stats.txtdenom, gives a text representation, and the denom field gives a matrix that defines a linear combination of the variances of terms in the model. For balanced models like this one, the denom matrix, stats.denom, contains zeros and ones, because the denominator is just a single term's mean square:

```
stats.txtdenom
ans =
    'MS(Factory*Car Model)'
    'MS(Factory*Car Model)'
    'MS(Error)'
stats.denom
```

```
ans =
\begin{tabular}{rrr}
-0.0000 & 1.0000 & 0.0000 \\
0.0000 & 1.0000 & -0.0000 \\
0.0000 & 0 & 1.0000
\end{tabular}
```

Variance Components. For the model described in "Setting Up the Model" on page 8-20, consider the mileage for a particular car of a particular model made at a random factory. The variance of that car is the sum of components, or contributions, one from each of the random terms.

```
stats.rtnames
ans =
    'Factory'
    'Factory*Car Model'
    'Error
```

You do not know those variances, but you can estimate them from the data. Recall that the ems field of the stats structure expresses the expected value of each term's mean square as a linear combination of unknown variances for random terms, and unknown quadratic forms for fixed terms. If you take the expected mean square expressions for the random terms, and equate those expected values to the computed mean squares, you get a system of equations that you can solve for the unknown variances. These solutions are the variance component estimates. The varest field contains a variance component estimate for each term. The rtnames field contains the names of the random terms.

```
stats.varest
ans =
    4.4426
    -0.0313
    0.1139
```

Under some conditions, the variability attributed to a term is unusually low, and that term's variance component estimate is negative. In those cases it
is common to set the estimate to zero, which you might do, for example, to create a bar graph of the components.

```
bar(max(0,stats.varest))
set(gca,'xtick',1:3,'xticklabel',stats.rtnames)
```



You can also compute confidence bounds for the variance estimate. The anovan function does this by computing confidence bounds for the variance expected mean squares, and finding lower and upper limits on each variance component containing all of these bounds. This procedure leads to a set of bounds that is conservative for balanced data. (That is, $95 \%$ confidence bounds will have a probability of at least $95 \%$ of containing the true variances if the number of observations for each combination of grouping variables is the same.) For unbalanced data, these are approximations that are not guaranteed to be conservative.

```
[{'Term' 'Estimate' 'Lower' 'Upper'};
    stats.rtnames, num2cell([stats.varest stats.varci])]
ans =
    'Term' 'Estimate' 'Lower' 'Upper'
```

| 'Factory' | 4.4426] | [1.0736] | [175.6038] |
| :---: | :---: | :---: | :---: |
| 'Factory*Car Model' | [ -0.0313] | NaN ] | NaN ] |
| Error' | 0.1139] | [0.0586] | $0.3103]$ |

## Other ANOVA Models

The anovan function also has arguments that enable you to specify two other types of model terms. First, the 'nested' argument specifies a matrix that indicates which factors are nested within other factors. A nested factor is one that takes different values within each level its nested factor.

For example, the mileage data from the previous section assumed that the two car models produced in each factory were the same. Suppose instead, each factory produced two distinct car models for a total of six car models, and we numbered them 1 and 2 for each factory for convenience. Then, the car model is nested in factory. A more accurate and less ambiguous numbering of car model would be as follows:


However, it is common with nested models to number the nested factor the same way in each nested factor.

Second, the 'continuous ' argument specifies that some factors are to be treated as continuous variables. The remaining factors are categorical variables. Although the anovan function can fit models with multiple continuous and categorical predictors, the simplest model that combines one predictor of each type is known as an analysis of covariance model. The next section describes a specialized tool for fitting this model.

## Analysis of Covariance

- "Introduction to Analysis of Covariance" on page 8-27
- "Analysis of Covariance Tool" on page 8-27
- "Confidence Bounds" on page 8-32
- "Multiple Comparisons" on page 8-34


## Introduction to Analysis of Covariance

Analysis of covariance is a technique for analyzing grouped data having a response ( $y$, the variable to be predicted) and a predictor ( $x$, the variable used to do the prediction). Using analysis of covariance, you can model $y$ as a linear function of $x$, with the coefficients of the line possibly varying from group to group.

## Analysis of Covariance Tool

The aoctool function opens an interactive graphical environment for fitting and prediction with analysis of covariance (ANOCOVA) models. It fits the following models for the $i$ th group:

| Same mean | $y=a+\varepsilon$ |
| :--- | :--- |
| Separate means | $y=\left(a+a_{i}\right)+\varepsilon$ |
| Same line | $y=\alpha+\beta x+\varepsilon$ |
| Parallel lines | $y=\left(a+a_{i}\right)+\beta x+\varepsilon$ |
| Separate lines | $y=\left(\alpha+a_{i}\right)+\left(\beta+\beta_{i}\right) x+\varepsilon$ |

For example, in the parallel lines model the intercept varies from one group to the next, but the slope is the same for each group. In the same mean model, there is a common intercept and no slope. In order to make the group coefficients well determined, the tool imposes the constraints

$$
\sum \alpha_{j}=\sum \beta_{j}=0
$$

The following steps describe the use of aoctool.

1 Load the data．The Statistics Toolbox data set carsmall．mat contains information on cars from the years 1970，1976，and 1982．This example studies the relationship between the weight of a car and its mileage， and whether this relationship has changed over the years．To start the demonstration，load the data set．

```
load carsmall
```

The Workspace Browser shows the variables in the data set．

| －）Workspace |  |  | －$\square$ 미 $x$ |
| :---: | :---: | :---: | :---: |
| File Edit Yiew Web Window Help |  |  |  |
|  | Stack：Base |  |  |
| Name | Size | Bytes | Class |
| 囲MPG <br> 囲Model＿Year <br> 囲Weight <br> 囲Horsepower <br> 囲Cylinders <br> 囲Acceleration <br> 囲Displacement <br> abc Model <br> abo Origin | $\begin{aligned} & 100 \times 1 \\ & 100 \times 1 \\ & 100 \times 1 \\ & 100 \times 1 \\ & 100 \times 1 \\ & 100 \times 1 \\ & 100 \times 1 \\ & 100 \times 36 \\ & 100 \times 7 \end{aligned}$ | 800 <br> 800 <br> 800 <br> 800 <br> 800 <br> 800 <br> 800 <br> 7200 <br> 1400 | double array double array double array double array double array double array double array char array char array |

You can also use aoctool with your own data．
2 Start the tool．The following command calls aoctool to fit a separate line to the column vectors Weight and MPG for each of the three model group defined in Model＿Year．The initial fit models the $y$ variable，MPG，as a linear function of the $x$ variable，Weight．
[h,atab,ctab,stats] = aoctool(Weight,MPG,Model_Year);

See the aoctool function reference page for detailed information about calling aoctool.

3 Examine the output. The graphical output consists of a main window with a plot, a table of coefficient estimates, and an analysis of variance table. In the plot, each Model_Year group has a separate line. The data points for each group are coded with the same color and symbol, and the fit for each group has the same color as the data points.


The coefficients of the three lines appear in the figure titled ANOCOVA Coefficients. You can see that the slopes are roughly -0.0078 , with a small deviation for each group:

- Model year 1970: $y=(45.9798-8.5805)+(-0.0078+0.002) x+\varepsilon$
- Model year 1976: $y=(45.9798-3.8902)+(-0.0078+0.0011) x+\varepsilon$
- Model year 1982: $y=(45.9798+12.4707)+(-0.0078-0.0031) x+\varepsilon$

| -) Figure No. 3: ANOCOVA Coefficients |  |  |  | - - 미 |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| File Edit Yiew Insert Iools Window Help |  |  |  |  |  |
| Coefficient Estimates |  |  |  |  |  |
| Term | Estimate | Std. Err. | T | Prob>\|T| | $\triangle$ |
| Intercept | 45.9798 | 1.52085 | 30.23 | 0 |  |
| 70 | -8.5805 | 1. 96186 | -4.37 |  |  |
| 76 | -3.8902 | 1. 86864 | -2.08 | 0.0403 |  |
| 82 | 12.4707 | 2.5568 | 4.88 | 0 |  |
| Slope | -0.0078 | 0.00056 | -14 | 0 |  |
| 70 | 0.002 | 0.00066 | 2.96 | 0.0039 |  |
| 76 | 0.0011 | 0.00065 | 1.74 | 0.0849 |  |
| 82 | -0.0031 | 0.001 | -3.1 | 0.0026 |  |

Because the three fitted lines have slopes that are roughly similar, you may wonder if they really are the same. The Model_Year*Weight interaction expresses the difference in slopes, and the ANOVA table shows a test for the significance of this term. With an $F$ statistic of 5.23 and a $p$ value of 0.0072 , the slopes are significantly different.

| -) Figure No. 2: AnOCOVA Test Results |  |  |  | - $\mid$ ㅁ] $\times$ |
| :---: | :---: | :---: | :---: | :---: |
| File Edit view Insert I | Iools Window |  |  |  |
| ANOVA Table |  |  |  |  |
| Source d. | d.f Sum Sq | Mean Sq | F | Prob $>\mathrm{F} \triangle$ |
| Model_Year | 2807.69 | 403.84 | 51.98 | 0 |
| Weight | 12050.2 | 2050.2 | 263.87 |  |
| Model_Year*Weigh | 281.22 | 40.61 | 5.23 | 0.0072 |
| Error 8 | $88 \quad 683.74$ | 7.77 |  | -1 |

4 Constrain the slopes to be the same. To examine the fits when the slopes are constrained to be the same, return to the ANOCOVA Prediction Plot window and use the Model pop-up menu to select a Parallel Lines model. The window updates to show the following graph.


Though this fit looks reasonable, it is significantly worse than the Separate Lines model. Use the Model pop-up menu again to return to the original model.

## Confidence Bounds

The example in "Analysis of Covariance Tool" on page 8-27 provides estimates of the relationship between MPG and Weight for each Model_Year, but how accurate are these estimates? To find out, you can superimpose confidence bounds on the fits by examining them one group at a time.

1 In the Model_Year menu at the lower right of the figure, change the setting from All Groups to 82 . The data and fits for the other groups are dimmed, and confidence bounds appear around the 82 fit.


The dashed lines form an envelope around the fitted line for model year 82 . Under the assumption that the true relationship is linear, these bounds provide a $95 \%$ confidence region for the true line. Note that the fits for the other model years are well outside these confidence bounds for Weight values between 2000 and 3000 .

2 Sometimes it is more valuable to be able to predict the response value for a new observation, not just estimate the average response value. Use the aoctool function Bounds menu to change the definition of the confidence bounds from Line to Observation. The resulting wider intervals reflect the uncertainty in the parameter estimates as well as the randomness of a new observation.


Like the polytool function, the aoctool function has cross hairs that you can use to manipulate the Weight and watch the estimate and confidence bounds along the $y$-axis update. These values appear only when a single group is selected, not when All Groups is selected.

## Multiple Comparisons

You can perform a multiple comparison test by using the stats output structure from aoctool as input to the multcompare function. The multcompare function can test either slopes, intercepts, or population marginal means (the predicted MPG of the mean weight for each group). The example in "Analysis of Covariance Tool" on page $8-27$ shows that the slopes are not all the same, but could it be that two are the same and only the other one is different? You can test that hypothesis.
multcompare(stats,0.05,'on', '', 's')
ans =
1.0000
1.0000

This matrix shows that the estimated difference between the intercepts of groups 1 and 2 (1970 and 1976) is 0.0008 , and a confidence interval for the difference is $[-0.0012,0.0029]$. There is no significant difference between the two. There are significant differences, however, between the intercept for 1982 and each of the other two. The graph shows the same information.


Note that the stats structure was created in the initial call to the aoctool function, so it is based on the initial model fit (typically a separate-lines model). If you change the model interactively and want to base your multiple comparisons on the new model, you need to run aoctool again to get another stats structure, this time specifying your new model as the initial model.

## Nonparametric Methods

- "Introduction to Nonparametric Methods" on page 8-36
- "Kruskal-Wallis Test" on page 8-36
- "Friedman's Test" on page 8-37


## Introduction to Nonparametric Methods

Statistics Toolbox functions include nonparametric versions of one-way and two-way analysis of variance. Unlike classical tests, nonparametric tests make only mild assumptions about the data, and are appropriate when the distribution of the data is non-normal. On the other hand, they are less powerful than classical methods for normally distributed data.

Both of the nonparametric functions described here will return a stats structure that can be used as an input to the multcompare function for multiple comparisons.

## Kruskal-Wallis Test

The example "Example: One-Way ANOVA" on page 8-4 uses one-way analysis of variance to determine if the bacteria counts of milk varied from shipment to shipment. The one-way analysis rests on the assumption that the measurements are independent, and that each has a normal distribution with a common variance and with a mean that was constant in each column. You can conclude that the column means were not all the same. The following example repeats that analysis using a nonparametric procedure.

The Kruskal-Wallis test is a nonparametric version of one-way analysis of variance. The assumption behind this test is that the measurements come from a continuous distribution, but not necessarily a normal distribution. The test is based on an analysis of variance using the ranks of the data values, not the data values themselves. Output includes a table similar to an ANOVA table, and a box plot.

You can run this test as follows:
load hogg
p = kruskalwallis(hogg)
$\mathrm{p}=$
0.0020

The low $p$ value means the Kruskal-Wallis test results agree with the one-way analysis of variance results.

## Friedman's Test

"Example: Two-Way ANOVA" on page 8-10 uses two-way analysis of variance to study the effect of car model and factory on car mileage. The example tests whether either of these factors has a significant effect on mileage, and whether there is an interaction between these factors. The conclusion of the example is there is no interaction, but that each individual factor has a significant effect. The next example examines whether a nonparametric analysis leads to the same conclusion.

Friedman's test is a nonparametric test for data having a two-way layout (data grouped by two categorical factors). Unlike two-way analysis of variance, Friedman's test does not treat the two factors symmetrically and it does not test for an interaction between them. Instead, it is a test for whether the columns are different after adjusting for possible row differences. The test is based on an analysis of variance using the ranks of the data across categories of the row factor. Output includes a table similar to an ANOVA table.

You can run Friedman's test as follows.

```
load mileage
p = friedman(mileage,3)
p =
    7.4659e-004
```

Recall the classical analysis of variance gave a $p$ value to test column effects, row effects, and interaction effects. This $p$ value is for column effects. Using either this $p$ value or the $p$ value from ANOVA ( $p<0.0001$ ), you conclude that there are significant column effects.

In order to test for row effects, you need to rearrange the data to swap the roles of the rows in columns. For a data matrix $x$ with no replications, you could simply transpose the data and type
p = friedman(x')

With replicated data it is slightly more complicated. A simple way is to transform the matrix into a three-dimensional array with the first dimension
representing the replicates, swapping the other two dimensions, and restoring the two-dimensional shape.

```
x = reshape(mileage, [3 2 3]);
x = permute(x,[\begin{array}{lll}{1 3 2]);}\end{array}]
x = reshape(x,[9 2])
x =
    33.3000 32.6000
    33.4000 32.5000
    32.9000 33.0000
    34.5000 33.4000
    34.8000 33.7000
    33.8000 33.9000
    37.4000 36.6000
    36.8000 37.0000
    37.6000 36.7000
friedman(x,3)
ans =
    0.0082
```

Again, the conclusion is similar to that of the classical analysis of variance. Both this $p$ value and the one from ANOVA ( $p=0.0039$ ) lead you to conclude that there are significant row effects.

You cannot use Friedman's test to test for interactions between the row and column factors.

## MANOVA

## In this section...

"Introduction to MANOVA" on page 8-39
"ANOVA with Multiple Responses" on page 8-39

## Introduction to MANOVA

The analysis of variance technique in "Example: One-Way ANOVA" on page 8-4 takes a set of grouped data and determine whether the mean of a variable differs significantly among groups. Often there are multiple response variables, and you are interested in determining whether the entire set of means is different from one group to the next. There is a multivariate version of analysis of variance that can address the problem.

## ANOVA with Multiple Responses

The carsmall data set has measurements on a variety of car models from the years 1970, 1976, and 1982. Suppose you are interested in whether the characteristics of the cars have changed over time.

First, load the data.

| load carsmall |  |  |  |
| :--- | ---: | ---: | :--- |
| whos |  |  |  |
| Name | Size | Bytes | Class |
| Acceleration | $100 \times 1$ | 800 | double array |
| Cylinders | $100 \times 1$ | 800 | double array |
| Displacement | $100 \times 1$ | 800 | double array |
| Horsepower | $100 \times 1$ | 800 | double array |
| MPG | $100 \times 1$ | 800 | double array |
| Model | $100 \times 36$ | 7200 | char array |
| Model_Year | $100 \times 1$ | 800 | double array |
| Origin | $100 \times 7$ | 1400 | char array |
| Weight | $100 \times 1$ | 800 | double array |

Four of these variables (Acceleration, Displacement, Horsepower, and MPG) are continuous measurements on individual car models. The variable

Model_Year indicates the year in which the car was made. You can create a grouped plot matrix of these variables using the gplotmatrix function.

(When the second argument of gplotmatrix is empty, the function graphs the columns of the $x$ argument against each other, and places histograms along the diagonals. The empty fourth argument produces a graph with the default colors. The fifth argument controls the symbols used to distinguish between groups.)

It appears the cars do differ from year to year. The upper right plot, for example, is a graph of MPG versus Weight. The 1982 cars appear to have higher mileage than the older cars, and they appear to weigh less on average. But as a group, are the three years significantly different from one another? The manova1 function can answer that question.

```
[d,p,stats] = manova1(x,Model_Year)
```

```
d =
        2
p =
    1.0e-006 *
            0
        0.1141
stats =
                    W: [4\times4 double]
                    B: [4x4 double]
                    T: [4x4 double]
                dfW: 90
                dfB: 2
                dfT: 92
        lambda: [2x1 double]
            chisq: [2x1 double]
        chisqdf: [2x1 double]
        eigenval: [4x1 double]
        eigenvec: [4x4 double]
            canon: [100x4 double]
            mdist: [100x1 double]
            gmdist: [3x3 double]
```

The manova1 function produces three outputs:

- The first output, d , is an estimate of the dimension of the group means. If the means were all the same, the dimension would be 0 , indicating that the means are at the same point. If the means differed but fell along a line, the dimension would be 1 . In the example the dimension is 2 , indicating that the group means fall in a plane but not along a line. This is the largest possible dimension for the means of three groups.
- The second output, p , is a vector of $p$-values for a sequence of tests. The first $p$ value tests whether the dimension is 0 , the next whether the dimension is 1 , and so on. In this case both $p$-values are small. That's why the estimated dimension is 2 .
- The third output, stats, is a structure containing several fields, described in the following section.


## The Fields of the stats Structure

The W, B, and T fields are matrix analogs to the within, between, and total sums of squares in ordinary one-way analysis of variance. The next three fields are the degrees of freedom for these matrices. Fields lambda, chisq, and chisqdf are the ingredients of the test for the dimensionality of the group means. (The $p$-values for these tests are the first output argument of manova1.)

The next three fields are used to do a canonical analysis. Recall that in principal components analysis ("Principal Component Analysis (PCA)" on page 10-31) you look for the combination of the original variables that has the largest possible variation. In multivariate analysis of variance, you instead look for the linear combination of the original variables that has the largest separation between groups. It is the single variable that would give the most significant result in a univariate one-way analysis of variance. Having found that combination, you next look for the combination with the second highest separation, and so on.

The eigenvec field is a matrix that defines the coefficients of the linear combinations of the original variables. The eigenval field is a vector measuring the ratio of the between-group variance to the within-group variance for the corresponding linear combination. The canon field is a matrix of the canonical variable values. Each column is a linear combination of the mean-centered original variables, using coefficients from the eigenvec matrix.

A grouped scatter plot of the first two canonical variables shows more separation between groups then a grouped scatter plot of any pair of original variables. In this example it shows three clouds of points, overlapping but with distinct centers. One point in the bottom right sits apart from the others. By using the gname function, you can see that this is the 20th point.

```
c1 = stats.canon(:,1);
c2 = stats.canon(:,2);
gscatter(c2,c1,Model_Year,[],'oxs')
gname
```



Roughly speaking, the first canonical variable, c1, separates the 1982 cars (which have high values of c1) from the older cars. The second canonical variable, c2, reveals some separation between the 1970 and 1976 cars.

The final two fields of the stats structure are Mahalanobis distances. The mdist field measures the distance from each point to its group mean. Points with large values may be outliers. In this data set, the largest outlier is the one in the scatter plot, the Buick Estate station wagon. (Note that you could have supplied the model name to the gname function above if you wanted to label the point with its model name rather than its row number.)

```
max(stats.mdist)
ans =
    31.5273
find(stats.mdist == ans)
ans =
```

```
    20
Model(20,:)
ans =
    buick_estate_wagon_(sw)
```

The gmdist field measures the distances between each pair of group means. The following commands examine the group means and their distances:

```
grpstats(x, Model_Year)
ans =
    1.0e+003 *
        0.0177 0.1489 0.2869 3.4413
        0.0216 0.1011 0.1978 3.0787
        0.0317 0.0815 0.1289 2.4535
stats.gmdist
ans =
\begin{tabular}{rrr}
0 & 3.8277 & 11.1106 \\
3.8277 & 0 & 6.1374 \\
11.1106 & 6.1374 & 0
\end{tabular}
```

As might be expected, the multivariate distance between the extreme years 1970 and 1982 (11.1) is larger than the difference between more closely spaced years (3.8 and 6.1). This is consistent with the scatter plots, where the points seem to follow a progression as the year changes from 1970 through 1976 to 1982. If you had more groups, you might find it instructive to use the manovacluster function to draw a diagram that presents clusters of the groups, formed using the distances between their means.

# Parametric Regression Analysis 

- "Introduction to Parametric Regression Analysis" on page 9-2
- "Linear Regression" on page 9-3
- "Nonlinear Regression" on page 9-72


## Introduction to Parametric Regression Analysis

Regression is the process of fitting models to data. The process depends on the model. If a model is parametric, regression estimates the parameters from the data. If a model is linear in the parameters, estimation is based on methods from linear algebra that minimize the norm of a residual vector. If a model is nonlinear in the parameters, estimation is based on search methods from optimization that minimize the norm of a residual vector. Nonparametric models, like "Classification Trees and Regression Trees" on page 13-27, use methods all their own.

This chapter considers data and models with continuous predictors and responses. Categorical predictors are the subject of Chapter 8, "Analysis of Variance". Categorical responses are the subject of Chapter 12, "Parametric Classification" and Chapter 13, "Nonparametric Supervised Learning".

## Linear Regression

## In this section...

"Linear Regression Models" on page 9-3
"Multiple Linear Regression" on page 9-8
"Robust Regression" on page 9-14
"Stepwise Regression" on page 9-19
"Ridge Regression" on page 9-29
"Lasso and Elastic Net" on page 9-32
"Partial Least Squares" on page 9-46
"Polynomial Models" on page 9-51
"Response Surface Models" on page 9-59
"Generalized Linear Models" on page 9-66
"Multivariate Regression" on page 9-71

## Linear Regression Models

In statistics, linear regression models often take the form of something like this:

$$
y=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{1} x_{2}+\beta_{4} x_{1}^{2}+\beta_{5} x_{2}^{2}+\varepsilon
$$

Here a response variable $y$ is modeled as a combination of constant, linear, interaction, and quadratic terms formed from two predictor variables $x_{1}$ and $x_{2}$. Uncontrolled factors and experimental errors are modeled by $\varepsilon$. Given data on $x_{1}, x_{2}$, and $y$, regression estimates the model parameters $\beta_{j}(j=1, \ldots, 5)$.

More general linear regression models represent the relationship between a continuous response $y$ and a continuous or categorical predictor $\boldsymbol{x}$ in the form:

$$
y=\beta_{1} f_{1}(\boldsymbol{x})+\ldots+\beta_{p} f_{p}(\boldsymbol{x})+\varepsilon
$$

The response is modeled as a linear combination of (not necessarily linear) functions of the predictor, plus a random error $\varepsilon$. The expressions $f_{j}(\boldsymbol{x})(j=1$, $\ldots, p)$ are the terms of the model. The $\beta_{j}(j=1, \ldots, p)$ are the coefficients. Errors $\varepsilon$ are assumed to be uncorrelated and distributed with mean 0 and constant (but unknown) variance.

Examples of linear regression models with a scalar predictor variable $x$ include:

- Linear additive (straight-line) models - Terms are $f_{1}(x)=1$ and $f_{2}(x)=x$.
- Polynomial models - Terms are $f_{1}(x)=1, f_{2}(x)=x, \ldots, f_{p}(x)=x^{p-1}$.
- Chebyshev orthogonal polynomial models - Terms are $f_{1}(x)=1, f_{2}(x)=x$, $\ldots, f_{p}(x)=2 x f_{p-1}(x)-f_{p-2}(x)$.
- Fourier trigonometric polynomial models - Terms are $f_{1}(x)=1 / 2$ and sines and cosines of different frequencies.

Examples of linear regression models with a vector of predictor variables $\boldsymbol{x}$ $=\left(x_{1}, \ldots, x_{N}\right)$ include:

- Linear additive (hyperplane) models - Terms are $f_{1}(\boldsymbol{x})=1$ and $f_{k+1}(\boldsymbol{x})=$ $x_{k}(k=1, \ldots, N)$.
- Pairwise interaction models - Terms are linear additive terms plus $g_{k_{1} k_{2}}(\boldsymbol{x})$ $=x_{k_{1}} x_{k_{2}}\left(k_{1}, k_{2}=1, \ldots, N, k_{1} \neq k_{2}\right)$.
- Quadratic models - Terms are pairwise interaction terms plus $h_{k}(\boldsymbol{x})=$ $x_{k}{ }^{2}(k=1, \ldots, N)$.
- Pure quadratic models - Terms are quadratic terms minus the $g_{k_{1} k_{2}}(\boldsymbol{x})$ terms.

Whether or not the predictor $\boldsymbol{x}$ is a vector of predictor variables, multivariate regression refers to the case where the response $\boldsymbol{y}=\left(y_{1}, \ldots, y_{M}\right)$ is a vector of $M$ response variables. See "Multivariate Regression" on page 9-71 for more on multivariate regression models.

Given $n$ independent observations $\left(\boldsymbol{x}_{1}, y_{1}\right), \ldots,\left(\boldsymbol{x}_{n}, y_{n}\right)$ of the predictor $\boldsymbol{x}$ and the response $y$, the linear regression model becomes an $n$-by- $p$ system of equations:

$$
\underbrace{\left(\begin{array}{c}
y_{1} \\
\vdots \\
y_{n}
\end{array}\right)}_{y}=\underbrace{\left(\begin{array}{ccc}
f_{1}\left(\boldsymbol{x}_{1}\right) & \cdots & f_{p}\left(\boldsymbol{x}_{1}\right) \\
\vdots & \ddots & \vdots \\
f_{1}\left(\boldsymbol{x}_{n}\right) & \ldots & f_{p}\left(\boldsymbol{x}_{n}\right)
\end{array}\right)}_{X} \underbrace{\left(\beta_{1}\right.}_{\beta} \begin{array}{c}
\vdots \\
\beta_{p}
\end{array})+\underbrace{\left(\begin{array}{c}
\varepsilon_{1} \\
\vdots \\
\varepsilon_{n}
\end{array}\right)}_{\varepsilon}
$$

$X$ is the design matrix of the system. The columns of $X$ are the terms of the model evaluated at the predictors. To fit the model to the data, the system must be solved for the $p$ coefficient values in $\beta=\left(\beta_{1}, \ldots, \beta_{p}\right)^{T}$.

The MATLAB backslash operator <br>(mldivide) solves systems of linear equations. Ignoring the unknown error $\varepsilon$, MATLAB estimates model coefficients in $y=X \beta$ using

```
betahat = X\y
```

where $X$ is the design matrix and $y$ is the vector of observed responses. MATLAB returns the least-squares solution to the system; betahat minimizes the norm of the residual vector $\mathrm{y}-\mathrm{X}$ *beta over all beta. If the system is consistent, the norm is 0 and the solution is exact. In this case, the regression model interpolates the data. In more typical regression cases where $n>p$ and the system is overdetermined, the least-squares solution estimates model coefficients obscured by the error $\varepsilon$.

The least-squares estimator betahat has several important statistical properties. First, it is unbiased, with expected value $\beta$. Second, by the Gauss-Markov theorem, it has minimum variance among all unbiased estimators formed from linear combinations of the response data. Under the additional assumption that $\varepsilon$ is normally distributed, betahat is a maximum likelihood estimator. The assumption also implies that the estimates themselves are normally distributed, which is useful for computing confidence intervals. Even without the assumption, by the Central Limit theorem, the estimates have an approximate normal distribution if the sample size is large enough.

Visualize the least-squares estimator as follows.


For betahat to minimize norm(y-X*beta), y-X*betahat must be perpendicular to the column space of $X$, which contains all linear combinations of the model terms. This requirement is summarized in the normal equations, which express vanishing inner products between $y-X *$ betahat and the columns of $X$ :

$$
X^{T}(y-X \hat{\beta})=0
$$

or

$$
X^{T} X \hat{\beta}=X^{T} y
$$

If X is $n$-by- $p$, the normal equations are a $p$-by- $p$ square system with solution betahat $=\operatorname{inv}\left(X^{\prime} * X\right) * X^{\prime} * y$, where inv is the MATLAB inverse operator. The matrix $\operatorname{inv}\left(X^{\prime} * X\right) * X^{\prime}$ is the pseudoinverse of $X$, computed by the MATLAB function pinv.

The normal equations are often badly conditioned relative to the original system $y=X \beta$ (the coefficient estimates are much more sensitive to the model error $\varepsilon$ ), so the MATLAB backslash operator avoids solving them directly.

Instead, a $Q R$ (orthogonal, triangular) decomposition of X is used to create a simpler, more stable triangular system:

$$
\begin{aligned}
X^{T} X \hat{\beta} & =X^{T} y \\
(Q R)^{T}(Q R) \hat{\beta} & =(Q R)^{T} y \\
R^{T} Q^{T} Q R \hat{\beta} & =R^{T} Q^{T} y \\
R^{T} R \hat{\beta} & =R^{T} Q^{T} y \\
R \hat{\beta} & =Q^{T} y
\end{aligned}
$$

Statistics Toolbox functions like regress and regstats call the MATLAB backslash operator to perform linear regression. The $Q R$ decomposition is also used for efficient computation of confidence intervals.

Once betahat is computed, the model can be evaluated at the predictor data:

$$
\text { yhat }=X * \text { betahat }
$$

or

$$
\text { yhat }=X * \operatorname{inv}\left(X^{\prime} * X\right) * X^{\prime *} y
$$

$H=X * \operatorname{inv}\left(X^{\prime} * X\right) * X^{\prime}$ is the hat matrix. It is a square, symmetric $n$-by- $n$ matrix determined by the predictor data. The diagonal elements $\mathrm{H}(\mathrm{i}, \mathrm{i})$ $(i=1, \ldots, n)$ give the leverage of the $i$ th observation. Since yhat $=H^{*} y$, leverage values determine the influence of the observed response $y(i)$ on the predicted response yhat(i). For leverage values near 1, the predicted response approximates the observed response. The Statistics Toolbox function leverage computes leverage values from a $Q R$ decomposition of x .

Component residual values in $y$-yhat are useful for detecting failures in model assumptions. Like the errors in $\varepsilon$, residuals have an expected value of 0 . Unlike the errors, however, residuals are correlated, with nonconstant variance. Residuals may be "Studentized" (scaled by an estimate of their standard deviation) for comparison. Studentized residuals are used by Statistics Toolbox functions like regress and robustfit to identify outliers in the data.

## Multiple Linear Regression

- "Introduction to Multiple Linear Regression" on page 9-8
- "Programmatic Multiple Linear Regression" on page 9-9
- "Interactive Multiple Linear Regression" on page 9-11
- "Tabulating Diagnostic Statistics" on page 9-13


## Introduction to Multiple Linear Regression

The system of linear equations

$$
\underbrace{\left(\begin{array}{c}
y_{1} \\
\vdots \\
y_{n}
\end{array}\right)}_{y}=\underbrace{\left(\begin{array}{ccc}
f_{1}\left(\boldsymbol{x}_{1}\right) & \ldots & f_{p}\left(\boldsymbol{x}_{1}\right) \\
\vdots & \ddots & \vdots \\
f_{1}\left(\boldsymbol{x}_{n}\right) & \ldots & f_{p}\left(\boldsymbol{x}_{n}\right)
\end{array}\right)}_{X} \underbrace{\left(\begin{array}{c}
\beta_{1} \\
\vdots \\
\beta_{p}
\end{array}\right)}_{\beta}+\underbrace{\left(\begin{array}{c}
\varepsilon_{1} \\
\vdots \\
\varepsilon_{n}
\end{array}\right)}_{\varepsilon}
$$

in "Linear Regression Models" on page 9-3 expresses a response $y$ as a linear combination of model terms $f_{j}(\boldsymbol{x})(j=1, \ldots, p)$ at each of the observations ( $\boldsymbol{x}_{1}$, $\left.y_{1}\right), \ldots,\left(\boldsymbol{x}_{n}, y_{n}\right)$.

If the predictor $\boldsymbol{x}$ is multidimensional, so are the functions $f_{j}$ that form the terms of the model. For example, if the predictor is $\boldsymbol{x}=\left(x_{1}, x_{2}\right)$, terms for the model might include $f_{1}(\boldsymbol{x})=x_{1}$ (a linear term), $f_{2}(\boldsymbol{x})=x_{1}{ }^{2}$ (a quadratic term), and $f_{3}(\boldsymbol{x})=x_{1} x_{2}$ (a pairwise interaction term). Typically, the function $f(\boldsymbol{x})=1$ is included among the $f_{j}$, so that the design matrix $X$ contains a column of 1 s and the model contains a constant ( $y$-intercept) term.

Multiple linear regression models are useful for:

- Understanding which terms $f_{j}(\boldsymbol{x})$ have greatest effect on the response (coefficients $\beta_{j}$ with greatest magnitude)
- Finding the direction of the effects (signs of the $\beta_{j}$ )
- Predicting unobserved values of the response $(y(\boldsymbol{x})$ for new $\boldsymbol{x})$

The Statistics Toolbox functions regress and regstats are used for multiple linear regression analysis.

## Programmatic Multiple Linear Regression

For example, the file moore.mat contains the 20 -by- 6 data matrix moore. The first five columns are measurements of biochemical oxygen demand on five predictor variables. The final column contains the observed responses. Use regress to find coefficient estimates betahat for a linear additive model as follows. Before using regress give the design matrix X1 a first column of 1 s to include a constant term in the model, betahat(1).

```
load moore
X1 = [ones(size(moore,1),1) moore(:,1:5)];
y = moore(:,6);
betahat = regress(y,X1)
betahat =
    -2.1561
    -0.0000
        0.0013
        0.0001
        0.0079
        0.0001
```

The MATLAB backslash (mldivide) operator, which regress calls, obtains the same result:

```
betahat = X1\y
betahat =
    -2.1561
    -0.0000
        0.0013
        0.0001
        0.0079
        0.0001
```

The advantage of working with regress is that it allows for additional inputs and outputs relevant to statistical analysis of the regression. For example:

```
alpha = 0.05;
[betahat,Ibeta,res,Ires,stats] = regress(y,X1,alpha);
```

returns not only the coefficient estimates in betahat, but also

- Ibeta - A p-by-2 matrix of $95 \%$ confidence intervals for the coefficient estimates, using a 100*(1-alpha) \% confidence level. The first column contains lower confidence bounds for each of the $p$ coefficient estimates; the second column contains upper confidence bounds.
- res - An $n$-by- 1 vector of residuals.
- Ires - An n-by-2 matrix of intervals that can be used to diagnose outliers. If the interval Ires (i, :) for observation i does not contain zero, the corresponding residual is larger than expected in 100* (1-alpha) \% of new observations, suggesting an outlier.
- stats - A 1-by-4 vector that contains, in order, the $R^{2}$ statistic, the $F$ statistic and its $p$ value, and an estimate of the error variance. The statistics are computed assuming the model contains a constant term, and are incorrect otherwise.

Visualize the residuals, in case (row number) order, with the rcoplot function:

```
rcoplot(res,Ires)
```



The interval around the first residual, shown in red when plotted, does not contain zero. This indicates that the residual is larger than expected in $95 \%$ of new observations, and suggests the data point is an outlier.

Outliers in regression appear for a variety of reasons:
1 If there is sufficient data, $5 \%$ of the residuals, by the definition of rint, are too big.

2 If there is a systematic error in the model (that is, if the model is not appropriate for generating the data under model assumptions), the mean of the residuals is not zero.

3 If the errors in the model are not normally distributed, the distributions of the residuals may be skewed or leptokurtic (with heavy tails and more outliers).

When errors are normally distributed, Ires(i,:) is a confidence interval for the mean of res(i) and checking if the interval contains zero is a test of the null hypothesis that the residual has zero mean.

## Interactive Multiple Linear Regression

The function regstats also performs multiple linear regression, but computes more statistics than regress. By default, regstats automatically adds a first column of 1 s to the design matrix (necessary for computing the $F$ statistic and its $p$ value), so a constant term should not be included explicitly as for regress. For example:

```
X2 = moore(:,1:5);
stats = regstats(y,X2);
```

creates a structure stats with fields containing regression statistics. An optional input argument allows you to specify which statistics are computed.

To interactively specify the computed statistics, call regstats without an output argument. For example:

```
regstats(y,X2)
```

opens the following interface.


Select the check boxes corresponding to the statistics you want to compute and click OK. Selected statistics are returned to the MATLAB workspace. Names
of container variables for the statistics appear on the right-hand side of the interface, where they can be changed to any valid MATLAB variable name.

## Tabulating Diagnostic Statistics

The regstats function computes statistics that are typically used in regression diagnostics. Statistics can be formatted into standard tabular displays in a variety of ways. For example, the tstat field of the stats output structure of regstats is itself a structure containing statistics related to the estimated coefficients of the regression. Dataset arrays (see "Dataset Arrays" on page 2-23) provide a natural tabular format for the information:

```
t = stats.tstat;
CoeffTable = dataset({t.beta,'Coef'},{t.se,'StdErr'}, ...
    {t.t,'tStat'},{t.pval,'pVal'})
CoeffTable =
    Coef StdErr tStat pVal
    -2.1561 0.91349 -2.3603 0.0333
    -9.0116e-006 0.00051835 -0.017385 0.98637
        0.0013159 0.0012635 1.0415 0.31531
        0.0001278 7.6902e-005 1.6618 0.11876
        0.0078989 0.014 0.56421 0.58154
        0.00014165 7.3749e-005 1.9208 0.075365
```

The MATLAB function fprintf gives you control over tabular formatting. For example, the fstat field of the stats output structure of regstats is a structure with statistics related to the analysis of variance (ANOVA) of the regression. The following commands produce a standard regression ANOVA table:

```
f = stats.fstat;
fprintf('\n')
fprintf('Regression ANOVA');
fprintf('\n\n')
fprintf('%6s','Source');
fprintf('%10s','df','SS','MS','F','P');
fprintf('\n')
fprintf('%6s','Regr');
```

```
fprintf('%10.4f',f.dfr,f.ssr,f.ssr/f.dfr,f.f,f.pval);
fprintf('\n')
fprintf('%6s','Resid');
fprintf('%10.4f',f.dfe,f.sse,f.sse/f.dfe);
fprintf('\n')
fprintf('%6s','Total');
fprintf('%10.4f',f.dfe+f.dfr,f.sse+f.ssr);
fprintf('\n')
```

The result looks like this:
Regression ANOVA

| Source | df | SS | MS | $F$ | $P$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| Regr | 5.0000 | 4.1084 | 0.8217 | 11.9886 | 0.0001 |
| Resid | 14.0000 | 0.9595 | 0.0685 |  |  |
| Total | 19.0000 | 5.0679 |  |  |  |

## Robust Regression

- "Introduction to Robust Regression" on page 9-14
- "Programmatic Robust Regression" on page 9-15
- "Interactive Robust Regression" on page 9-16


## Introduction to Robust Regression

The models described in "Linear Regression Models" on page 9-3 are based on certain assumptions, such as a normal distribution of errors in the observed responses. If the distribution of errors is asymmetric or prone to outliers, model assumptions are invalidated, and parameter estimates, confidence intervals, and other computed statistics become unreliable. The Statistics Toolbox function robustfit is useful in these cases. The function implements a robust fitting method that is less sensitive than ordinary least squares to large changes in small parts of the data.

Robust regression works by assigning a weight to each data point. Weighting is done automatically and iteratively using a process called iteratively
reweighted least squares. In the first iteration, each point is assigned equal weight and model coefficients are estimated using ordinary least squares. At subsequent iterations, weights are recomputed so that points farther from model predictions in the previous iteration are given lower weight. Model coefficients are then recomputed using weighted least squares. The process continues until the values of the coefficient estimates converge within a specified tolerance.

## Programmatic Robust Regression

The example in "Multiple Linear Regression" on page 9-8 shows an outlier when ordinary least squares is used to model the response variable as a linear combination of the five predictor variables. To determine the influence of the outlier, compare the coefficient estimates computed by regress:

```
load moore
X1 = [ones(size(moore,1),1) moore(:,1:5)];
y = moore(:,6);
betahat = regress(y,X1)
betahat =
```

    -2.1561
    -0.0000
        0.0013
        0.0001
        0.0079
        0.0001
    to those computed by robustfit:

```
X2 = moore(:,1:5);
robustbeta = robustfit(X2,y)
robustbeta =
    -1.7516
    0.0000
    0.0009
    0.0002
    0.0060
    0.0001
```

By default, robustfit automatically adds a first column of 1 s to the design matrix, so a constant term does not have to be included explicitly as for
regress. In addition, the order of inputs is reversed for robustfit and regress.

To understand the difference in the coefficient estimates, look at the final weights given to the data points in the robust fit:

```
[robustbeta,stats] = robustfit(X2,y);
stats.w'
ans =
    Columns 1 through 5
        0.0246 0.9986 0.9763 0.9323 0.9704
    Columns 6 through 10
        0.8597 0.9180 0.9992 0.9590}00.964
        Columns }11\mathrm{ through 15
        0.9769 0.9868 0.9999 0.9976 0.8122
    Columns 16 through 20
        0.9733 0.9892 0.9988 0.8974 0.6774
```

The first data point has a very low weight compared to the other data points, and so is effectively ignored in the robust regression.

## Interactive Robust Regression

The robustdemo function shows the difference between ordinary least squares and robust fitting for data with a single predictor. You can use data provided with the demo, or you can supply your own data. The following steps show you how to use robustdemo.

1 Start the demo. To begin using robustdemo with the built-in data, simply enter the function name at the command line:

```
robustdemo
```

Use left mouse button to select and drag points
Use right mouse button to query point properties


Least squares:
Robust:
$Y=-0.188327+1.10351 * X$
$Y=-1.77278+1.50415 * X$

RMS error $=2.21375$
RMS error $=1.43663$

The resulting figure shows a scatter plot with two fitted lines. The red line is the fit using ordinary least-squares regression. The green line is the fit using robust regression. At the bottom of the figure are the equations for the fitted lines, together with the estimated root mean squared errors for each fit.

2 View leverages and robust weights. Right-click on any data point to see its least-squares leverage and robust weight.


In the built-in data, the rightmost point has a relatively high leverage of 0.35 . The point exerts a large influence on the least-squares fit, but its small robust weight shows that it is effectively excluded from the robust fit.

3 See how changes in the data affect the fits. With the left mouse button, click and hold on any data point and drag it to a new location. When you release the mouse button, the displays update.


Bringing the rightmost data point closer to the least-squares line makes the two fitted lines nearly identical. The adjusted rightmost data point has significant weight in the robust fit.

## Stepwise Regression

- "Introduction to Stepwise Regression" on page 9-20
- "Programmatic Stepwise Regression" on page 9-21
- "Interactive Stepwise Regression" on page 9-27


## Introduction to Stepwise Regression

Multiple linear regression models, as described in "Multiple Linear Regression" on page 9-8, are built from a potentially large number of predictive terms. The number of interaction terms, for example, increases exponentially with the number of predictor variables. If there is no theoretical basis for choosing the form of a model, and no assessment of correlations among terms, it is possible to include redundant terms in a model that confuse the identification of significant effects.

Stepwise regression is a systematic method for adding and removing terms from a multilinear model based on their statistical significance in a regression. The method begins with an initial model and then compares the explanatory power of incrementally larger and smaller models. At each step, the $p$ value of an $F$-statistic is computed to test models with and without a potential term. If a term is not currently in the model, the null hypothesis is that the term would have a zero coefficient if added to the model. If there is sufficient evidence to reject the null hypothesis, the term is added to the model. Conversely, if a term is currently in the model, the null hypothesis is that the term has a zero coefficient. If there is insufficient evidence to reject the null hypothesis, the term is removed from the model. The method proceeds as follows:

1 Fit the initial model.
2 If any terms not in the model have $p$-values less than an entrance tolerance (that is, if it is unlikely that they would have zero coefficient if added to the model), add the one with the smallest $p$ value and repeat this step; otherwise, go to step 3 .

3 If any terms in the model have $p$-values greater than an exit tolerance (that is, if it is unlikely that the hypothesis of a zero coefficient can be rejected), remove the one with the largest $p$ value and go to step 2; otherwise, end.

Depending on the terms included in the initial model and the order in which terms are moved in and out, the method may build different models from the same set of potential terms. The method terminates when no single step improves the model. There is no guarantee, however, that a different initial model or a different sequence of steps will not lead to a better fit. In this sense, stepwise models are locally optimal, but may not be globally optimal.

Statistics Toolbox functions for stepwise regression are:

- stepwisefit - A function that proceeds automatically from a specified initial model and entrance/exit tolerances
- stepwise - An interactive tool that allows you to explore individual steps in the regression


## Programmatic Stepwise Regression

For example, load the data in hald.mat, which contains observations of the heat of reaction of various cement mixtures:

| load hald <br> whos |  |  |  |  |
| :--- | :--- | ---: | :--- | :--- |
| $\quad$ Name | Size | Bytes | Class | Attributes |
|  |  |  |  |  |
| $\quad$Description $22 \times 58$ 2552 char <br> hald $13 \times 5$ 520 double <br> heat $13 \times 1$ 104 double <br> ingredients $13 \times 4$ 416 double. |  |  |  |  |

The response (heat) depends on the quantities of the four predictors (the columns of ingredients).

Use stepwisefit to carry out the stepwise regression algorithm, beginning with no terms in the model and using entrance/exit tolerances of 0.05/0.10 on the $p$-values:

```
stepwisefit(ingredients,heat,...
    'penter',0.05,'premove',0.10);
Initial columns included: none
Step 1, added column 4, p=0.000576232
Step 2, added column 1, p=1.10528e-006
Final columns included: 1 4
    'Coeff' 'Std.Err.' 'Status' 'P'
    [ 1.4400] [ 0.1384] 'In' [1.1053e-006]
    [ 0.4161] [ 0.1856] 'Out' [ 0.0517]
    [-0.4100] [ 0.1992] 'Out' [ 0.0697]
    [-0.6140] [ 0.0486] 'In' [1.8149e-007]
```

stepwisefit automatically includes an intercept term in the model, so you do not add it explicitly to ingredients as you would for regress. For terms not
in the model, coefficient estimates and their standard errors are those that result if the term is added.

The inmodel parameter is used to specify terms in an initial model:

```
initialModel = ...
    [false true false false]; % Force in 2nd term
stepwisefit(ingredients,heat,...
    'inmodel',initialModel,...
    'penter',.05,'premove',0.10);
Initial columns included: 2
Step 1, added column 1, p=2.69221e-007
Final columns included: 1 2
    'Coeff' 'Std.Err.' 'Status' 'P'
    [ 1.4683] [ 0.1213] 'In' [2.6922e-007]
    [ 0.6623] [ 0.0459] 'In' [5.0290e-008]
    [ 0.2500] [ 0.1847] 'Out' [ 0.2089]
    [-0.2365] [ 0.1733] 'Out' [ 0.2054]
```

The preceding two models, built from different initial models, use different subsets of the predictive terms. Terms 2 and 4, swapped in the two models, are highly correlated:

```
term2 = ingredients(:,2);
term4 = ingredients(:,4);
R = corrcoef(term2,term4)
R =
    1.0000 -0.9730
    -0.9730 1.0000
```

To compare the models, use the stats output of stepwisefit:

```
[betahat1,se1,pval1,inmodel1,stats1] = ...
    stepwisefit(ingredients,heat,...
    'penter',.05,'premove',0.10,\ldots.
    'display','off');
[betahat2,se2,pval2,inmodel2,stats2] = ...
    stepwisefit(ingredients,heat,...
                                    'inmodel',initialModel,...
                                    'penter',.05,'premove',0.10,...
                                    'display','off');
```

```
RMSE1 = stats1.rmse
RMSE1 =
    2.7343
RMSE2 = stats2.rmse
RMSE2 =
    2.4063
```

The second model has a lower Root Mean Square Error (RMSE).
An added variable plot is used to determine the unique effect of adding a new term to a model. The plot shows the relationship between the part of the response unexplained by terms already in the model and the part of the new term unexplained by terms already in the model. The "unexplained" parts are measured by the residuals of the respective regressions. A scatter of the residuals from the two regressions forms the added variable plot.

For example, suppose you want to add term2 to a model that already contains the single term term1. First, consider the ability of term2 alone to explain the response:

```
load hald
term2 = ingredients(:,2);
[b2,Ib2,res2] = regress(heat,[ones(size(term2)) term2]);
scatter(term2,heat)
xlabel('Term 2')
ylabel('Heat')
hold on
x2 = 20:80;
y2 = b2(1) + b2(2)*x2;
plot(x2,y2,'r')
title('{\bf Response Explained by Term 2: Ignoring Term 1}')
```



Next, consider the following regressions involving the model term term1:

```
term1 = ingredients(:,1);
[b1,Ib1,res1] = regress(heat,[ones(size(term1)) term1]);
[b21,Ib21,res21] = regress(term2,[ones(size(term1)) term1]);
bres = regress(res1,[ones(size(res21)) res21]);
```

A scatter of the residuals res1 vs. the residuals res12 forms the added variable plot:

```
figure
scatter(res21,res1)
xlabel('Residuals: Term 2 on Term 1')
ylabel('Residuals: Heat on Term 1')
hold on
```

```
xres = -30:30;
yres = bres(1) + bres(2)*xres;
plot(xres,yres,'r')
title('{\bf Response Explained by Term 2: Adjusted for Term 1}')
```

Response Explained by Term 2: Adjusted for Term 1


Since the plot adjusted for term1 shows a stronger relationship (less variation along the fitted line) than the plot ignoring term1, the two terms act jointly to explain extra variation. In this case, adding term2 to a model consisting of term1 would reduce the RMSE.

The Statistics Toolbox function addedvarplot produces added variable plots. The previous plot is essentially the one produced by the following:

```
figure
```

```
inmodel = [true false false false];
addedvarplot(ingredients,heat,2,inmodel)
Added variable plot for X2 Adjusted for X1
```



In addition to the scatter of residuals, the plot shows $95 \%$ confidence intervals on predictions from the fitted line. The fitted line has intercept zero because, under the assumptions outlined in "Linear Regression Models" on page 9-3, both of the plotted variables have mean zero. The slope of the fitted line is the coefficient that term2 would have if it was added to the model with term1.

The addevarplot function is useful for considering the unique effect of adding a new term to an existing model with any number of terms.

## Interactive Stepwise Regression

The stepwise interface provides interactive features that allow you to investigate individual steps in a stepwise regression, and to build models from arbitrary subsets of the predictive terms. To open the interface with data from hald.mat:

```
load hald
stepwise(ingredients,heat)
```



The upper left of the interface displays estimates of the coefficients for all potential terms, with horizontal bars indicating $90 \%$ (colored) and $95 \%$ (grey) confidence intervals. The red color indicates that, initially, the terms are not in the model. Values displayed in the table are those that would result if the terms were added to the model.

The middle portion of the interface displays summary statistics for the entire model. These statistics are updated with each step.

The lower portion of the interface, Model History, displays the RMSE for the model. The plot tracks the RMSE from step to step, so you can compare the optimality of different models. Hover over the blue dots in the history to see which terms were in the model at a particular step. Click on a blue dot in the history to open a copy of the interface initialized with the terms in the model at that step.

Initial models, as well as entrance/exit tolerances for the $p$-values of $F$-statistics, are specified using additional input arguments to stepwise. Defaults are an initial model with no terms, an entrance tolerance of 0.05 , and an exit tolerance of 0.10 .

To center and scale the input data (compute $z$-scores) to improve conditioning of the underlying least-squares problem, select Scale Inputs from the Stepwise menu.

You proceed through a stepwise regression in one of two ways:
1 Click Next Step to select the recommended next step. The recommended next step either adds the most significant term or removes the least significant term. When the regression reaches a local minimum of RMSE, the recommended next step is "Move no terms." You can perform all of the recommended steps at once by clicking All Steps.

2 Click a line in the plot or in the table to toggle the state of the corresponding term. Clicking a red line, corresponding to a term not currently in the model, adds the term to the model and changes the line to blue. Clicking a blue line, corresponding to a term currently in the model, removes the term from the model and changes the line to red.

To call addedvarplot and produce an added variable plot from the stepwise interface, select Added Variable Plot from the Stepwise menu. A list of terms is displayed. Select the term you want to add, and then click OK.

Click Export to display a dialog box that allows you to select information from the interface to save to the MATLAB workspace. Check the information you want to export and, optionally, change the names of the workspace variables to be created. Click OK to export the information.

## Ridge Regression

- "Introduction to Ridge Regression" on page 9-29
- "Example: Ridge Regression" on page 9-30


## Introduction to Ridge Regression

Coefficient estimates for the models described in "Multiple Linear Regression" on page $9-8$ rely on the independence of the model terms. When terms are correlated and the columns of the design matrix $X$ have an approximate linear dependence, the matrix $\left(X^{T} X\right)^{-1}$ becomes close to singular. As a result, the least-squares estimate

$$
\hat{\beta}=\left(X^{T} X\right)^{-1} X^{T} y
$$

becomes highly sensitive to random errors in the observed response $y$, producing a large variance. This situation of multicollinearity can arise, for example, when data are collected without an experimental design.

Ridge regression addresses the problem by estimating regression coefficients using

$$
\hat{\beta}=\left(X^{T} X+k I\right)^{-1} X^{T} y
$$

where $k$ is the ridge parameter and $I$ is the identity matrix. Small positive values of $k$ improve the conditioning of the problem and reduce the variance of the estimates. While biased, the reduced variance of ridge estimates often result in a smaller mean square error when compared to least-squares estimates.

The Statistics Toolbox function ridge carries out ridge regression.

## Example: Ridge Regression

For example, load the data in acetylene.mat, with observations of the predictor variables $\mathrm{x} 1, \mathrm{x} 2, \mathrm{x} 3$, and the response variable y :

```
load acetylene
```

Plot the predictor variables against each other:

```
subplot(1,3,1)
plot(x1,x2,'.')
xlabel('x1'); ylabel('x2'); grid on; axis square
subplot(1,3,2)
plot(x1,x3,'.')
xlabel('x1'); ylabel('x3'); grid on; axis square
subplot(1,3,3)
plot(x2,x3,'.')
xlabel('x2'); ylabel('x3'); grid on; axis square
```



Note the correlation between $\times 1$ and the other two predictor variables.
Use ridge and x2fx to compute coefficient estimates for a multilinear model with interaction terms, for a range of ridge parameters:

```
X = [x1 x2 x3];
D = x2fx(X,'interaction');
```

```
D(:,1) = []; % No constant term
k = 0:1e-5:5e-3;
betahat = ridge(y,D,k);
```

Plot the ridge trace:

```
figure
plot(k,betahat,'LineWidth',2)
ylim([-100 100])
grid on
xlabel('Ridge Parameter')
ylabel('Standardized Coefficient')
title('{\bf Ridge Trace}')
legend('x1','x2','x3','x1x2','x1x3','x2x3')
```



The estimates stabilize to the right of the plot. Note that the coefficient of the $\times 2 \times 3$ interaction term changes sign at a value of the ridge parameter $\approx$ $5 \times 10^{-4}$.

## Lasso and Elastic Net

Lasso is a regularization technique. Use lasso to:

- Reduce the number of predictors in a regression model.
- Identify important predictors.
- Select among redundant predictors.
- Produce shrinkage estimates with potentially lower predictive errors than ordinary least squares.

Elastic net is a related technique. Use elastic net when you have several highly correlated variables. lasso provides elastic net regularization when you set the Alpha name-value pair to a number strictly between 0 and 1 .

See "Lasso and Elastic Net Details" on page 9-44.

## Example: Lasso Regularization

To see how lasso identifies and discards unnecessary predictors:
1 Generate 200 samples of five-dimensional artificial data $X$ from exponential distributions with various means:

```
rng(3,'twister') % for reproducibility
X = zeros(200,5);
for ii = 1:5
    X(:,ii) = exprnd(ii,200,1);
end
```

2 Generate response data $Y=X * r+e p s$ where $r$ has just two nonzero components, and the noise eps is normal with standard deviation 0.1:

```
r = [0;2;0;-3;0];
Y = X*r + randn(200,1)*.1;
```

3 Fit a cross-validated sequence of models with lasso, and plot the result:

```
[b fitinfo] = lasso(X,Y,'CV',10);
lassoPlot(b,fitinfo,'PlotType','Lambda','XScale','log');
```


## Trace Plot of coefficients fit by Lasso



The plot shows the nonzero coefficients in the regression for various values of the Lambda regularization parameter. Larger values of Lambda appear on the left side of the graph, meaning more regularization, resulting in fewer nonzero regression coefficients.

The dashed vertical lines represent the Lambda value with minimal mean squared error (on the right), and the Lambda value with minimal mean squared error plus one standard deviation. This latter value is a recommended setting for Lambda. These lines appear only when you perform cross validation. Cross validate by setting the 'CV' name-value pair. This example uses 10 -fold cross validation.

The upper part of the plot shows the degrees of freedom (df), meaning the number of nonzero coefficients in the regression, as a function of Lambda. On the left, the large value of Lambda causes all but one coefficient to be

0 . On the right all five coefficients are nonzero, though the plot shows only two clearly. The other three coefficients are so small that you cannot visually distinguish them from 0 .

For small values of Lambda (toward the right in the plot), the coefficient values are close to the least-squares estimate. See step 5 on page 9-35.

4 Find the Lambda value of the minimal cross-validated mean squared error plus one standard deviation. Examine the MSE and coefficients of the fit at that Lambda:

```
lam = fitinfo.Index1SE;
fitinfo.MSE(lam)
ans =
    0.0101
b(:,lam)
ans =
    0
        1.9850
        0
        -2.9946
        0
```

lasso did a good job finding the coefficient vector $r$.
5 For comparison, find the least-squares estimate of $r$ :

```
rhat = X\Y
rhat =
    -0.0038
    1.9952
    0.0014
    -2.9993
    0.0031
```

The estimate $\mathrm{b}(:, \mathrm{lam})$ has slightly more mean squared error than the mean squared error of rhat:

```
res = X*rhat - Y; % calculate residuals
MSEmin = res'*res/200 % b(:,lam) value is 0.0101
MSEmin =
    0.0088
```

But b(:, lam) has only two nonzero components, and therefore can provide better predictive estimates on new data.

## Example: Lasso and Elastic Net Regularization with Cross Validation

Consider predicting the mileage (MPG) of a car based on its weight, displacement, horsepower, and acceleration. The carbig data contains these measurements. The data seem likely to be correlated, making elastic net an attractive choice.

1 Load the data:

```
    load carbig
```

2 Extract the continuous (noncategorical) predictors (lasso does not handle categorical predictors):

```
X = [Acceleration Displacement Horsepower Weight];
```

3 Perform a lasso fit with 10 -fold cross validation:
[b fitinfo] = lasso(X,MPG, 'CV',10);
4 Plot the result:

```
lassoPlot(b,fitinfo,'PlotType','Lambda','XScale','log');
```

Trace Plot of coefficients fit by Lasso
df


5 Calculate the correlation of the predictors:

```
% Eliminate NaNs so corr runs
nonan = ~any(isnan([X MPG]),2);
Xnonan = X(nonan,:);
MPGnonan = MPG(nonan,:);
corr(Xnonan)
ans =
    1.0000 -0.5438 -0.6892 -0.4168
    -0.5438 1.0000 0.8973 0.9330
    -0.6892 0.8973 1.0000 0.8645
    -0.4168 0.9330 0.8645 1.0000
```

6 Because some predictors are highly correlated, perform elastic net fitting. Use Alpha = 0.5:

```
[ba fitinfoa] = lasso(X,MPG,'CV',10,'Alpha',.5);
```

7 Plot the result. Name each predictor so you can tell which curve is which:

```
pnames = {'Acceleration','Displacement',...
    'Horsepower','Weight'};
lassoPlot(ba,fitinfoa,'PlotType','Lambda',...
    'XScale','log','PredictorNames',pnames);
```

                    Trace Plot of coefficients fit by Elastic Net (Alpha \(=0.5\) )
                                    df
    

When you activate the data cursor
and click the plot, you see the name of the predictor, the coefficient, the value of Lambda, and the index of that point, meaning the column in $b$ associated with that fit.

Here, the elastic net and lasso results are not very similar. Also, the elastic net plot reflects a notable qualitative property of the elastic net technique. The elastic net retains three nonzero coefficients as Lambda increases (toward the left of the plot), and these three coefficients reach 0 at about the same Lambda value. In contrast, the lasso plot shows two of the three coefficients becoming 0 at the same value of Lambda, while another coefficient remains nonzero for higher values of Lambda.

This behavior exemplifies a general pattern. In general, elastic net tends to retain or drop groups of highly correlated predictors as Lambda increases. In contrast, lasso tends to drop smaller groups, or even individual predictors.

## Example: Wide Data via Lasso and Parallel Computing

Lasso and elastic net are especially well suited to wide data, meaning data with more predictors than observations. Obviously, there are redundant predictors in this type of data. Use lasso along with cross validation to identify important predictors.

Cross validation can be slow. If you have a Parallel Computing Toolbox ${ }^{\text {TM }}$ license, speed the computation using parallel computing.

1 Load the spectra data:

```
load spectra
Description
Description =
== Spectral and octane data of gasoline ==
NIR spectra and octane numbers of 60 gasoline samples
NIR: NIR spectra, measured in 2 nm intervals from 900 nm to 1700 nm
octane: octane numbers
spectra: a dataset array containing variables for NIR and octane
```

Reference:
Kalivas, John H., "Two Data Sets of Near Infrared Spectra," Chemometrics and Intelligent Laboratory Systems, v. 37 (1997) pp. 255259

2 Compute the default lasso fit:
[b fitinfo] = lasso(NIR,octane);

3 Plot the number of predictors in the fitted lasso regularization as a function of Lambda, using a logarithmic $x$-axis:
lassoPlot(b,fitinfo,'PlotType','Lambda','XScale', 'log');

Trace Plot of coefficients fit by Lasso


4 It is difficult to tell which value of Lambda is appropriate. To determine a good value, try fitting with cross validation:
tic

```
[b fitinfo] = lasso(NIR,octane,'CV',10);
% A time-consuming operation
toc
Elapsed time is 161.933311 seconds.
```

5 Plot the result:

```
lassoPlot(b,fitinfo,'PlotType','Lambda','XScale','log');
```

Trace Plot of coefficients fit by Lasso


You can see the suggested value of Lambda is over $1 \mathrm{e}-2$, and the Lambda with minimal MSE is under 1e-2. These values are in the fitinfo structure:

```
fitinfo.LambdaMinMSE
ans =
    0.0057
```

```
fitinfo.Lambda1SE
ans =
    0.0190
```

6 Examine the quality of the fit for the suggested value of Lambda:

```
lambdaindex = fitinfo.Index1SE;
fitinfo.MSE(lambdaindex)
ans =
    0.0532
fitinfo.DF(lambdaindex)
ans =
    1 1
```

The fit uses just 11 of the 401 predictors, and achieves a cross-validated MSE of 0.0532 .

7 Examine the plot of cross-validated MSE:

```
lassoPlot(b,fitinfo,'PlotType','CV');
% Use a log scale for MSE to see small MSE values better
set(gca,'YScale','log');
```



As Lambda increases (toward the left), MSE increases rapidly. The coefficients are reduced too much and they do not adequately fit the responses.

As Lambda decreases, the models are larger (have more nonzero coefficients). The increasing MSE suggests that the models are overfitted.

The default set of Lambda values does not include values small enough to include all predictors. In this case, there does not appear to be a reason to look at smaller values. However, if you want smaller values than the default, use the LambdaRatio parameter, or supply a sequence of Lambda values using the Lambda parameter. For details, see the lasso reference page.

8 To compute the cross-validated lasso estimate faster, use parallel computing (available with a Parallel Computing Toolbox license):

```
matlabpool open
Starting matlabpool using the 'local' configuration ...
    connected to 4 labs.
opts = statset('UseParallel','always');
tic;
[b fitinfo] = lasso(NIR,octane,'CV',10,'Options',opts);
toc
```

Elapsed time is 77.910882 seconds.

Computing in parallel is more than twice as fast on this problem using a quad-core processor.

## Lasso and Elastic Net Details

Overview of Lasso and Elastic Net. Lasso is a regularization technique for performing linear regression. Lasso includes a penalty term that constrains the size of the estimated coefficients. Therefore, it resembles ridge regression. Lasso is a shrinkage estimator: it generates coefficient estimates that are biased to be small. Nevertheless, a lasso estimator can have smaller mean squared error than an ordinary least-squares estimator when you apply it to new data.

Unlike ridge regression, as the penalty term increases, lasso sets more coefficients to zero. This means that the lasso estimator is a smaller model, with fewer predictors. As such, lasso is an alternative to stepwise regression and other model selection and dimensionality reduction techniques.

Elastic net is a related technique. Elastic net is a hybrid of ridge regression and lasso regularization. Like lasso, elastic net can generate reduced models by generating zero-valued coefficients. Empirical studies have suggested that the elastic net technique can outperform lasso on data with highly intercorrelated predictors.

Definition of Lasso. The lasso technique solves this regularization problem. For a given value of $\lambda$, a nonnegative parameter, lasso solves the problem

$$
\min _{\beta_{0}, \beta}\left(\frac{1}{2 N} \sum_{i=1}^{N}\left(y_{i}-\beta_{0}-x_{i}^{T} \beta\right)+\lambda \sum_{j=1}^{p}\left|\beta_{j}\right|\right)
$$

where

- $N$ is the number of observations.
- $y_{i}$ is the response at observation $i$.
- $x_{i}$ is data, a vector of $p$ values at observation $i$.
- $\lambda$ is a positive regularization parameter corresponding to one value of Lambda.
- The parameters $\beta_{0}$ and $\beta$ are scalar and $p$-vector respectively.

As $\lambda$ increases, the number of nonzero components of $\beta$ decreases.

The lasso problem involves the $L^{1}$ norm of $\beta$, as contrasted with the elastic net algorithm.

Definition of Elastic Net. The elastic net technique solves this regularization problem. For an $a$ strictly between 0 and 1, and a nonnegative $\lambda$, elastic net solves the problem

$$
\min _{\beta_{0}, \beta}\left(\frac{1}{2 N} \sum_{i=1}^{N}\left(y_{i}-\beta_{0}-x_{i}^{T} \beta\right)+\lambda P_{\alpha}(\beta)\right)
$$

where

$$
P_{\alpha}(\beta)=\frac{(1-\alpha)}{2}\|\beta\|_{2}^{2}+\alpha\|\beta\|_{1}=\sum_{j=1}^{p}\left(\frac{(1-\alpha)}{2} \beta_{j}^{2}+\alpha\left|\beta_{j}\right|\right) .
$$

Elastic net is the same as lasso when $\alpha=1$. As $a$ shrinks toward 0 , elastic net approaches ridge regression. For other values of $\alpha$, the penalty term $P_{a}(\beta)$ interpolates between the $L^{1}$ norm of $\beta$ and the squared $L^{2}$ norm of $\beta$.

## References

[1] Tibshirani, R. Regression shrinkage and selection via the lasso. Journal of the Royal Statistical Society, Series B, Vol 58, No. 1, pp. 267-288, 1996.
[2] Zou, H. and T. Hastie. Regularization and variable selection via the elastic net. Journal of the Royal Statistical Society, Series B, Vol. 67, No. 2, pp. 301-320, 2005.
[3] Friedman, J., R. Tibshirani, and T. Hastie. Regularization paths for generalized linear models via coordinate descent. Journal of Statistical Software, Vol 33, No. 1, 2010. http://www.jstatsoft.org/v33/i01
[4] Hastie, T., R. Tibshirani, and J. Friedman. The Elements of Statistical Learning, 2nd edition. Springer, New York, 2008.

## Partial Least Squares

- "Introduction to Partial Least Squares" on page 9-46
- "Example: Partial Least Squares" on page 9-47


## Introduction to Partial Least Squares

Partial least-squares (PLS) regression is a technique used with data that contain correlated predictor variables. This technique constructs new predictor variables, known as components, as linear combinations of the original predictor variables. PLS constructs these components while considering the observed response values, leading to a parsimonious model with reliable predictive power.

The technique is something of a cross between multiple linear regression and principal component analysis:

- Multiple linear regression finds a combination of the predictors that best fit a response.
- Principal component analysis finds combinations of the predictors with large variance, reducing correlations. The technique makes no use of response values.
- PLS finds combinations of the predictors that have a large covariance with the response values.

PLS therefore combines information about the variances of both the predictors and the responses, while also considering the correlations among them.

PLS shares characteristics with other regression and feature transformation techniques. It is similar to ridge regression in that it is used in situations with correlated predictors. It is similar to stepwise regression (or more general feature selection techniques) in that it can be used to select a smaller set of model terms. PLS differs from these methods, however, by transforming the original predictor space into the new component space.

The Statistics Toolbox function plsregress carries out PLS regression.

## Example: Partial Least Squares

For example, consider the data on biochemical oxygen demand in moore.mat, padded with noisy versions of the predictors to introduce correlations:

```
load moore
y = moore(:,6); % Response
X0 = moore(:,1:5); % Original predictors
X1 = X0+10*randn(size(X0)); % Correlated predictors
X = [X0,X1];
```

Use plsregress to perform PLS regression with the same number of components as predictors, then plot the percentage variance explained in the response as a function of the number of components:

```
[XL,yl,XS,YS,beta,PCTVAR] = plsregress(X,y,10);
plot(1:10,cumsum(100*PCTVAR(2,:)),'-bo');
xlabel('Number of PLS components');
ylabel('Percent Variance Explained in y');
```



Choosing the number of components in a PLS model is a critical step. The plot gives a rough indication, showing nearly $80 \%$ of the variance in y explained by the first component, with as many as five additional components making significant contributions.

The following computes the six-component model:

```
[XL,yl,XS,YS,beta,PCTVAR,MSE,stats] = plsregress(X,y,6);
yfit = [ones(size(X,1),1) X]*beta;
plot(y,yfit,'o')
```



The scatter shows a reasonable correlation between fitted and observed responses, and this is confirmed by the $R^{2}$ statistic:

```
TSS = sum((y-mean(y)).^2);
RSS = sum((y-yfit).^2);
Rsquared = 1 - RSS/TSS
Rsquared =
    0.8421
```

A plot of the weights of the ten predictors in each of the six components shows that two of the components (the last two computed) explain the majority of the variance in $X$ :

```
plot(1:10,stats.W, 'o-');
legend({'c1','c2','c3','c4','c5','c6'},'Location','NW')
xlabel('Predictor');
ylabel('Weight');
```



A plot of the mean-squared errors suggests that as few as two components may provide an adequate model:

```
[axes,h1,h2] = plotyy(0:6,MSE(1,:),0:6,MSE(2,:));
set(h1,'Marker','o')
set(h2,'Marker','o')
legend('MSE Predictors','MSE Response')
xlabel('Number of Components')
```



The calculation of mean-squared errors by plsregress is controlled by optional parameter name/value pairs specifying cross-validation type and the number of Monte Carlo repetitions.

## Polynomial Models

- "Introduction to Polynomial Models" on page 9-51
- "Programmatic Polynomial Regression" on page 9-52
- "Interactive Polynomial Regression" on page 9-57


## Introduction to Polynomial Models

Polynomial models are a special case of the linear models discussed in "Linear Regression Models" on page 9-3. Polynomial models have the advantages of being simple, familiar in their properties, and reasonably flexible for following
data trends. They are also robust with respect to changes in the location and scale of the data (see "Conditioning Polynomial Fits" on page 9-55). However, polynomial models may be poor predictors of new values. They oscillate between data points, especially as the degree is increased to improve the fit. Asymptotically, they follow power functions, leading to inaccuracies when extrapolating other long-term trends. Choosing a polynomial model is often a trade-off between a simple description of overall data trends and the accuracy of predictions made from the model.

## Programmatic Polynomial Regression

- "Functions for Polynomial Fitting" on page 9-52
- "Displaying Polynomial Fits" on page 9-54
- "Conditioning Polynomial Fits" on page 9-55

Functions for Polynomial Fitting. To fit polynomials to data, MATLAB and Statistics Toolbox software offer a number of dedicated functions. The MATLAB function polyfit computes least-squares coefficient estimates for polynomials of arbitrary degree. For example:

```
x = 0:5; % x data
y = [2 1 4 4 3 2]; % y data
p = polyfit(x,y,3) % Degree 3 fit
p =
    -0.1296 0.6865 -0.1759 1.6746
```

Polynomial coefficients in p are listed from highest to lowest degree, so $p(x)$ $\approx-0.13 x^{3}+0.69 x^{2}-0.18 x+1.67$. For convenience, polyfit sets up the Vandermonde design matrix (vander) and calls backslash (mldivide) to perform the fit.

Once the coefficients of a polynomial are collected in a vector $p$, use the MATLAB function polyval to evaluate the polynomial at arbitrary inputs. For example, the following plots the data and the fit over a range of inputs:

```
plot(x,y,'ro','LineWidth',2) % Plot data
hold on
xfit = -1:0.01:6;
yfit = polyval(p,xfit);
```

```
plot(xfit,yfit,'LineWidth',2) % Plot fit
ylim([0,5])
grid on
```



Use the MATLAB function roots to find the roots of $p$ :

```
r = roots(p)
r=
    5.4786
    -0.0913 + 1.5328i
    -0.0913 - 1.5328i
```

The MATLAB function poly solves the inverse problem, finding a polynomial with specified roots. poly is the inverse of roots up to ordering, scaling, and round-off error.

An optional output from polyfit is passed to polyval or to the Statistics Toolbox function polyconf to compute prediction intervals for the fit. For example, the following computes $95 \%$ prediction intervals for new observations at each value of the predictor x :

```
[p,S] = polyfit(x,y,3);
[yhat,delta] = polyconf(p,x,S);
PI = [yhat-delta;yhat+delta]'
PI =
    -5.3022 8.6514
    -4.2068 8.3179
    -2.9899 9.0534
    -2.1963 9.8471
    -2.6036 9.9211
    -5.2229 8.7308
```

Optional input arguments to polyconf allow you to compute prediction intervals for estimated values (yhat) as well as new observations, and to compute the bounds simultaneously for all x instead of nonsimultaneously (the default). The confidence level for the intervals can also be set.

Displaying Polynomial Fits. The documentation example function polydemo combines the functions polyfit, polyval, roots, and polyconf to produce a formatted display of data with a polynomial fit.

Note Statistics Toolbox documentation example files are located in the \help\toolbox\stats \examples subdirectory of your MATLAB root folder (matlabroot). This subdirectory is not on the MATLAB path at installation. To use the files in this subdirectory, either add the subdirectory to the MATLAB path (addpath) or make the subdirectory your current working folder (cd).

For example, the following uses polydemo to produce a display of simulated data with a quadratic trend, a fitted polynomial, and $95 \%$ prediction intervals for new observations:

```
x = -5:5;
y = x.^2 - 5*x - 3 + 5*randn(size(x));
p = polydemo(x,y,2,0.05)
```


polydemo calls the documentation example function polystr to convert the coefficient vector $p$ into a string for the polynomial expression displayed in the figure title.

Conditioning Polynomial Fits. If $x$ and $y$ data are on very different scales, polynomial fits may be badly conditioned, in the sense that coefficient estimates are very sensitive to random errors in the data. For example, using polyfit to estimate coefficients of a cubic fit to the U.S. census data in census.mat produces the following warning:

```
load census
x = cdate;
```

```
y = pop;
p = polyfit(x,y,3);
Warning: Polynomial is badly conditioned.
    Add points with distinct X values,
    reduce the degree of the polynomial,
    or try centering and scaling as
    described in HELP POLYFIT.
```

The following implements the suggested centering and scaling, and demonstrates the robustness of polynomial fits under these transformations:

```
plot(x,y,'ro') % Plot data
hold on
z = zscore(x); % Compute z-scores of x data
zfit = linspace(z(1),z(end),100);
pz = polyfit(z,y,3); % Compute conditioned fit
yfit = polyval(pz,zfit);
xfit = linspace(x(1),x(end),100);
plot(xfit,yfit,'b-') % Plot conditioned fit vs. x data
grid on
```



## Interactive Polynomial Regression

The functions polyfit, polyval, and polyconf are interactively applied to data using two graphical interfaces for polynomial fitting:

- "The Basic Fitting Tool" on page 9-57
- "The Polynomial Fitting Tool" on page 9-58

The Basic Fitting Tool. The Basic Fitting Tool is a MATLAB interface, discussed in "Interactive Fitting" in the MATLAB documentation. The tool allows you to:

- Fit interpolants and polynomials of degree $\leq 10$
- Plot residuals and compute their norm
- Interpolate or extrapolate values from the fit
- Save results to the MATLAB workspace

The Polynomial Fitting Tool. The Statistics Toolbox function polytool opens the Polynomial Fitting Tool. For example, the following opens the interface using simulated data with a quadratic trend and displays a fitted polynomial with $95 \%$ prediction intervals for new observations:

```
x = -5:5;
y = x.^2 - 5*x - 3 + 5*randn(size(x));
polytool(x,y,2,0.05)
```



The tool allows you to:

- Interactively change the degree of the fit. Change the value in the Degree text box at the top of the figure.
- Evaluate the fit and the bounds using a movable crosshair. Click, hold, and drag the crosshair to change its position.
- Export estimated coefficients, predicted values, prediction intervals, and residuals to the MATLAB workspace. Click Export to a open a dialog box with choices for exporting the data.

Options for the displayed bounds and the fitting method are available through menu options at the top of the figure:

- The Bounds menu lets you choose between bounds on new observations (the default) and bounds on estimated values. It also lets you choose between nonsimultaneous (the default) and simultaneous bounds. See polyconf for a description of these options.
- The Method menu lets you choose between ordinary least-squares regression and robust regression, as described in "Robust Regression" on page 9-14.


## Response Surface Models

- "Introduction to Response Surface Models" on page 9-59
- "Programmatic Response Surface Methodology" on page 9-60
- "Interactive Response Surface Methodology" on page 9-65


## Introduction to Response Surface Models

Polynomial models are generalized to any number of predictor variables $x_{i}$ ( $i$ $=1, \ldots, N$ ) as follows:

$$
y(x)=a_{0}+\sum_{i=0}^{N} a_{i} x_{i}+\sum_{i<j}^{N} a_{i j} x_{i} x_{j}+\sum_{i=0}^{N} a_{i i} x_{i}^{2}+\ldots
$$

The model includes, from left to right, an intercept, linear terms, quadratic interaction terms, and squared terms. Higher order terms would follow, as necessary.

Response surface models are multivariate polynomial models. They typically arise in the design of experiments (see Chapter 15, "Design of Experiments"), where they are used to determine a set of design variables that optimize a response. Linear terms alone produce models with response surfaces that are hyperplanes. The addition of interaction terms allows for warping of the hyperplane. Squared terms produce the simplest models in which the response surface has a maximum or minimum, and so an optimal response.

Response surface methodology (RSM) is the process of adjusting predictor variables to move the response in a desired direction and, iteratively, to an optimum. The method generally involves a combination of both computation and visualization. The use of quadratic response surface models makes the method much simpler than standard nonlinear techniques for determining optimal designs.

## Programmatic Response Surface Methodology

The file reaction.mat contains simulated data on the rate of a chemical reaction:
load reaction

The variables include:

- rate - A 13-by-1 vector of observed reaction rates
- reactants - A 13-by-3 matrix of reactant concentrations
- xn - The names of the three reactants
- yn - The name of the response

In "Nonlinear Regression" on page 9-72, the nonlinear Hougen-Watson model is fit to the data using nlinfit. However, there may be no theoretical basis for choosing a particular model to fit the data. A quadratic response surface model provides a simple way to determine combinations of reactants that lead to high reaction rates.

As described in "Multiple Linear Regression" on page 9-8, the regress and regstats functions fit linear models-including response surface models-to data using a design matrix of model terms evaluated at predictor data. The
x2fx function converts predictor data to design matrices for quadratic models. The regstats function calls x2fx when instructed to do so.

For example, the following fits a quadratic response surface model to the data in reaction.mat:

```
stats = regstats(rate,reactants,'quadratic','beta');
b = stats.beta; % Model coefficients
```

The 10 -by- 1 vector b contains, in order, a constant term and then the coefficients for the model terms $x_{1}, x_{2}, x_{3}, x_{1} x_{2}, x_{1} x_{3}, x_{2} x_{3}, x_{1}^{2}, x_{2}{ }^{2}$, and $x_{3}{ }^{2}$, where $x_{1}, x_{2}$, and $x_{3}$ are the three columns of reactants. The order of coefficients for quadratic models is described in the reference page for x 2 fx .

Since the model involves only three predictors, it is possible to visualize the entire response surface using a color dimension for the reaction rate:

```
x1 = reactants(:,1);
x2 = reactants(:,2);
x3 = reactants(:,3);
xx1 = linspace(min(x1),max(x1),25);
xx2 = linspace(min(x2),max(x2),25);
xx3 = linspace(min(x3),max (x3),25);
[X1,X2,X3] = meshgrid(xx1,xx2,xx3);
RATE = b(1) + b(2)*X1 + b(3)*X2 + b(4)*X3 + ...
    b(5)*X1.*X2 + b(6)*X1.*X3 + b(7)*X2.*X3 + ...
    b(8)*X1.^2 + b(9)*X2.^2 + b(10)*X3.^2;
hmodel = scatter3(X1(:),X2(:),X3(:),5,RATE(:),'filled');
hold on
hdata = scatter3(x1,x2,x3,'ko','filled');
axis tight
xlabel(xn(1,:))
ylabel(xn(2,:))
zlabel(xn(3,:))
hbar = colorbar;
ylabel(hbar,yn);
title('{\bf Quadratic Response Surface Model}')
```



The plot show a general increase in model response, within the space of the observed data, as the concentration of $n$-pentane increases and the concentrations of hydrogen and isopentane decrease.

Before trying to determine optimal values of the predictors, perhaps by collecting more data in the direction of increased reaction rate indicated by the plot, it is helpful to evaluate the geometry of the response surface. If $x$ $=\left(x_{1}, x_{2}, x_{3}\right)^{T}$ is the vector of predictors, and $H$ is the matrix such that $x^{T} H x$ gives the quadratic terms of the model, the model has a unique optimum if and only if $H$ is positive definite. For the data in this example, the model does not have a unique optimum:

```
H = [b(8),b(5)/2,b(6)/2; ...
    b(5)/2,b(9),b(7)/2; ...
    b(6)/2,b(7)/2,b(10)];
lambda = eig(H)
lambda =
1.0e-003 *
    -0.1303
    0.0412
    0.4292
```

The negative eigenvalue shows a lack of positive definiteness. The saddle in the model is visible if the range of the predictors in the plot ( $\mathrm{xx} 1, \mathrm{xx} 2$, and xx 3 ) is expanded:


When the number of predictors makes it impossible to visualize the entire response surface, 3 -, 2 -, and 1 -dimensional slices provide local views. The MATLAB function slice displays 2 -dimensional contours of the data at fixed values of the predictors:


One-dimensional contours are displayed by the Response Surface Tool, rstool, described in the next section.

## Interactive Response Surface Methodology

The Statistics Toolbox function rstool opens a GUI for interactively investigating simultaneous one-dimensional contours of multidimensional response surface models. For example, the following opens the interface with a quadratic response surface fit to the data in reaction.mat:

```
load reaction
alpha = 0.01; % Significance level
rstool(reactants,rate,'quadratic',alpha,xn,yn)
```



A sequence of plots is displayed, each showing a contour of the response surface against a single predictor, with all other predictors held fixed. Confidence intervals for new observations are shown as dashed red curves above and below the response. Predictor values are displayed in the text boxes on the horizontal axis and are marked by vertical dashed blue lines
in the plots. Predictor values are changed by editing the text boxes or by dragging the dashed blue lines. When you change the value of a predictor, all plots update to show the new point in predictor space.

Note The Statistics Toolbox demonstration function rsmdemo generates simulated data for experimental settings specified by either the user or by a $D$-optimal design generated by cordexch. It uses the rstool interface to visualize response surface models fit to the data, and it uses the nlintool interface to visualize a nonlinear model fit to the data.

## Generalized Linear Models

- "Introduction to Generalized Linear Models" on page 9-66
- "Example: Generalized Linear Models" on page 9-67


## Introduction to Generalized Linear Models

Linear regression models describe a linear relationship between a response and one or more predictive terms. Many times, however, a nonlinear relationship exists. "Nonlinear Regression" on page 9-72 describes general nonlinear models. A special class of nonlinear models, known as generalized linear models, makes use of linear methods.

Recall that linear models have the following characteristics:

- At each set of values for the predictors, the response has a normal distribution with mean $\mu$.
- A coefficient vector $b$ defines a linear combination $X b$ of the predictors $X$.
- The model is $\mu=X b$.

In generalized linear models, these characteristics are generalized as follows:

- At each set of values for the predictors, the response has a distribution that may be normal, binomial, Poisson, gamma, or inverse Gaussian, with parameters including a mean $\mu$.
- A coefficient vector $b$ defines a linear combination $X b$ of the predictors $X$.
- A link function $f$ defines the model as $f(\mu)=X b$.


## Example: Generalized Linear Models

The following data are derived from carbig.mat, which contains measurements of large cars of various weights. Each weight in whas a corresponding number of cars in total and a corresponding number of poor-mileage cars in poor:

```
w = [2100 2300 2500 2700 2900 3100 ...
    3300 3500 3700 3900 4100 4300]';
total = [48 42 31 34 31 21 23 23 21 16 17 21]';
poor = [1 2 0 3 8 8 14 17 19 15 17 21]';
```

A plot shows that the proportion of poor-mileage cars follows an S-shaped sigmoid:

```
plot(w,poor./total,'x','LineWidth',2)
grid on
xlabel('Weight')
ylabel('Proportion of Poor-Mileage Cars')
```



The logistic model is useful for proportion data. It defines the relationship between the proportion $p$ and the weight $w$ by:

$$
\log [p /(1-p)]=b_{1}+b_{2} w
$$

Some of the proportions in the data are 0 and 1, making the left-hand side of this equation undefined. To keep the proportions within range, add relatively small perturbations to the poor and total values. A semi-log plot then shows a nearly linear relationship, as predicted by the model:

```
p_adjusted = (poor+.5)./(total+1);
semilogy(w,p_adjusted./(1-p_adjusted),'x','LineWidth',2)
grid on
xlabel('Weight')
ylabel('Adjusted p / (1 - p)')
```



It is reasonable to assume that the values of poor follow binomial distributions, with the number of trials given by total and the percentage of successes depending on $w$. This distribution can be accounted for in the context of a logistic model by using a generalized linear model with link function $\log (\mu /(1-\mu))=X b$.

Use the glmfit function to carry out the associated regression:

```
b = glmfit(w,[poor total],'binomial','link','logit')
b =
    -13.3801
    0.0042
```

To use the coefficients in $b$ to compute fitted proportions, invert the logistic relationship:

$$
p=1 /\left(1+\exp \left(-b_{1}-b_{2} w\right)\right)
$$

Use the glmval function to compute the fitted values:

```
    x = 2100:100:4500;
    y = glmval(b,x,'logit');
    plot(w,poor./total,'x','LineWidth',2)
    hold on
    plot(x,y,'r-','LineWidth',2)
    grid on
    xlabel('Weight')
    ylabel('Proportion of Poor-Mileage Cars')
```



The previous is an example of logistic regression. For an example of a kind of stepwise logistic regression, analogous to stepwise regression for linear models, see "Sequential Feature Selection" on page 10-23.

## Multivariate Regression

Whether or not the predictor $\boldsymbol{x}$ is a vector of predictor variables, multivariate regression refers to the case where the response $\boldsymbol{y}=\left(y_{1}, \ldots, y_{M}\right)$ is a vector of $M$ response variables.

The Statistics Toolbox functions mvregress and mvregresslike are used for multivariate regression analysis.

## Nonlinear Regression

In this section...<br>"Nonlinear Regression Models" on page 9-72<br>"Parametric Models" on page 9-73<br>"Mixed-Effects Models" on page 9-78

## Nonlinear Regression Models

The models described in "Linear Regression Models" on page 9-3 are often called empirical models, because they are based solely on observed data. Model parameters typically have no relationship to any mechanism producing the data. To increase the accuracy of a linear model within the range of observations, the number of terms is simply increased.

Nonlinear models, on the other hand, typically involve parameters with specific physical interpretations. While they require a priori assumptions about the data-producing process, they are often more parsimonious than linear models, and more accurate outside the range of observed data.

Parametric nonlinear models represent the relationship between a continuous response variable and one or more predictor variables (either continuous or categorical) in the form $y=f(X, \beta)+\varepsilon$, where

- $y$ is an $n$-by- 1 vector of observations of the response variable.
- $X$ is an $n$-by- $p$ design matrix determined by the predictors.
- $\beta$ is a $p$-by- 1 vector of unknown parameters to be estimated.
- $f$ is any function of $X$ and $\beta$.
- $\varepsilon$ is an $n$-by- 1 vector of independent, identically distributed random disturbances.

Nonparametric models do not attempt to characterize the relationship between predictors and response with model parameters. Descriptions are often graphical, as in the case of "Classification Trees and Regression Trees" on page 13-27.

## Parametric Models

- "A Parametric Nonlinear Model" on page 9-73
- "Confidence Intervals for Parameter Estimates" on page 9-75
- "Confidence Intervals for Predicted Responses" on page 9-75
- "Interactive Nonlinear Parametric Regression" on page 9-76


## A Parametric Nonlinear Model

The Hougen-Watson model (Bates and Watts, [2], pp. 271-272) for reaction kinetics is an example of a parametric nonlinear model. The form of the model is

$$
\text { rate }=\frac{\beta_{1} x_{2}-x_{3} / \beta_{5}}{1+\beta_{2} x_{1}+\beta_{3} x_{2}+\beta_{4} x_{3}}
$$

where rate is the reaction rate, $x_{1}, x_{2}$, and $x_{3}$ are concentrations of hydrogen, $n$-pentane, and isopentane, respectively, and $\beta_{1}, \beta_{2}, \ldots, \beta_{5}$ are the unknown parameters.

The file reaction.mat contains simulated reaction data:
load reaction
The variables are:

- rate - A 13-by-1 vector of observed reaction rates
- reactants - A 13-by-3 matrix of reactant concentrations
- beta - A 5-by-1 vector of initial parameter estimates
- model - The name of a function file for the model
- $x n$ - The names of the reactants
- yn - The name of the response

The function for the model is hougen, which looks like this:
type hougen

```
function yhat = hougen(beta,x)
%HOUGEN Hougen-Watson model for reaction kinetics.
% YHAT = HOUGEN(BETA,X) gives the predicted values of the
% reaction rate, YHAT, as a function of the vector of
% parameters, BETA, and the matrix of data, X.
% BETA must have five elements and X must have three
% columns.
%
% The model form is:
% y = (b1*x2 - x3/b5)./(1+b2*x1+b3*x2+b4*x3)
b1 = beta(1);
b2 = beta(2);
b3 = beta(3);
b4 = beta(4);
b5 = beta(5);
x1 = x(:,1);
x2 = x(:,2);
x3 = x(:,3);
yhat = (b1*x2 - x3/b5)./(1+b2*x1+b3*x2+b4*x3);
```

The function nlinfit is used to find least-squares parameter estimates for nonlinear models. It uses the Gauss-Newton algorithm with Levenberg-Marquardt modifications for global convergence.
nlinfit requires the predictor data, the responses, and an initial guess of the unknown parameters. It also requires a function handle to a function that takes the predictor data and parameter estimates and returns the responses predicted by the model.

To fit the reaction data, call nlinfit using the following syntax:

```
load reaction
betahat = nlinfit(reactants,rate,@hougen,beta)
betahat =
    1.2526
    0.0628
```

$$
\begin{aligned}
& 0.0400 \\
& 0.1124 \\
& 1.1914
\end{aligned}
$$

The output vector betahat contains the parameter estimates.
The function nlinfit has robust options, similar to those for robustfit, for fitting nonlinear models to data with outliers.

## Confidence Intervals for Parameter Estimates

To compute confidence intervals for the parameter estimates, use the function nlparci, together with additional outputs from nlinfit:

```
[betahat,resid,J] = nlinfit(reactants,rate,@hougen,beta);
betaci = nlparci(betahat,resid,J)
betaci =
\begin{tabular}{ll}
-0.7467 & 3.2519 \\
-0.0377 & 0.1632 \\
-0.0312 & 0.1113 \\
-0.0609 & 0.2857 \\
-0.7381 & 3.1208
\end{tabular}
```

The columns of the output betaci contain the lower and upper bounds, respectively, of the (default) $95 \%$ confidence intervals for each parameter.

## Confidence Intervals for Predicted Responses

The function nlpredci is used to compute confidence intervals for predicted responses:

```
[yhat,delta] = nlpredci(@hougen,reactants,betahat,resid,J);
opd = [rate yhat delta]
opd =
\begin{tabular}{rrr}
8.5500 & 8.4179 & 0.2805 \\
3.7900 & 3.9542 & 0.2474 \\
4.8200 & 4.9109 & 0.1766 \\
0.0200 & -0.0110 & 0.1875 \\
2.7500 & 2.6358 & 0.1578 \\
14.3900 & 14.3402 & 0.4236 \\
2.5400 & 2.5662 & 0.2425
\end{tabular}
```

| 4.3500 | 4.0385 | 0.1638 |
| ---: | ---: | ---: |
| 13.0000 | 13.0292 | 0.3426 |
| 8.5000 | 8.3904 | 0.3281 |
| 0.0500 | -0.0216 | 0.3699 |
| 11.3200 | 11.4701 | 0.3237 |
| 3.1300 | 3.4326 | 0.1749 |

The output opd contains the observed rates in the first column and the predicted rates in the second column. The (default) $95 \%$ simultaneous confidence intervals on the predictions are the values in the second column $\pm$ the values in the third column. These are not intervals for new observations at the predictors, even though most of the confidence intervals do contain the original observations.

## Interactive Nonlinear Parametric Regression

Calling nlintool opens a graphical user interface (GUI) for interactive exploration of multidimensional nonlinear functions, and for fitting parametric nonlinear models. The GUI calls nlinfit, and requires the same inputs. The interface is analogous to polytool and rstool for polynomial models.

Open nlintool with the reaction data and the hougen model by typing

```
load reaction
nlintool(reactants,rate,@hougen, beta, 0.01,xn,yn)
```



You see three plots. The response variable for all plots is the reaction rate, plotted in green. The red lines show confidence intervals on predicted responses. The first plot shows hydrogen as the predictor, the second shows $n$-pentane, and the third shows isopentane.

Each plot displays the fitted relationship of the reaction rate to one predictor at a fixed value of the other two predictors. The fixed values are in the text boxes below each predictor axis. Change the fixed values by typing in a new value or by dragging the vertical lines in the plots to new positions. When you change the value of a predictor, all plots update to display the model at the new point in predictor space.

While this example uses only three predictors, nlintool can accommodate any number of predictors.

Note The Statistics Toolbox demonstration function rsmdemo generates simulated data for experimental settings specified by either the user or by a $D$-optimal design generated by cordexch. It uses the rstool interface to visualize response surface models fit to the data, and it uses the nlintool interface to visualize a nonlinear model fit to the data.

## Mixed-Effects Models

- "Introduction to Mixed-Effects Models" on page 9-78
- "Mixed-Effects Model Hierarchy" on page 9-79
- "Specifying Mixed-Effects Models" on page 9-81
- "Specifying Covariate Models" on page 9-84
- "Choosing nlmefit or nlmefitsa" on page 9-85
- "Using Output Functions with Mixed-Effects Models" on page 9-88
- "Example: Mixed-Effects Models Using nlmefit and nlmefitsa" on page 9-93
- "Example: Examining Residuals for Model Verification" on page 9-108


## Introduction to Mixed-Effects Models

In statistics, an effect is anything that influences the value of a response variable at a particular setting of the predictor variables. Effects are translated into model parameters. In linear models, effects become coefficients, representing the proportional contributions of model terms. In nonlinear models, effects often have specific physical interpretations, and appear in more general nonlinear combinations.

Fixed effects represent population parameters, assumed to be the same each time data is collected. Estimating fixed effects is the traditional domain of regression modeling. Random effects, by comparison, are sample-dependent random variables. In modeling, random effects act like additional error terms, and their distributions and covariances must be specified.

For example, consider a model of the elimination of a drug from the bloodstream. The model uses time $t$ as a predictor and the concentration of the drug $C$ as the response. The nonlinear model term $C_{0} e^{-r t}$ combines parameters $C_{0}$ and $r$, representing, respectively, an initial concentration and an elimination rate. If data is collected across multiple individuals, it is reasonable to assume that the elimination rate is a random variable $r_{i}$ depending on individual $i$, varying around a population mean $\bar{r}$. The term $C_{0} e^{-r t}$ becomes

$$
C_{0} e^{-\left[\bar{r}+\left(r_{i}-\bar{r}\right)\right] t}=C_{0} e^{-\left(\beta+b_{i}\right) t},
$$

where $\beta=\bar{r}$ is a fixed effect and $b_{i}=r_{i}-\bar{r}$ is a random effect.
Random effects are useful when data falls into natural groups. In the drug elimination model, the groups are simply the individuals under study. More sophisticated models might group data by an individual's age, weight, diet, etc. Although the groups are not the focus of the study, adding random effects to a model extends the reliability of inferences beyond the specific sample of individuals.

Mixed-effects models account for both fixed and random effects. As with all regression models, their purpose is to describe a response variable as a function of the predictor variables. Mixed-effects models, however, recognize correlations within sample subgroups. In this way, they provide a compromise between ignoring data groups entirely and fitting each group with a separate model.

## Mixed-Effects Model Hierarchy

Suppose data for a nonlinear regression model falls into one of $m$ distinct groups $i=1, \ldots, m$. To account for the groups in a model, write response $j$ in group $i$ as:

$$
y_{i j}=f\left(\varphi, x_{i j}\right)+\varepsilon_{i j}
$$

$y_{i j}$ is the response, $x_{i j}$ is a vector of predictors, $\varphi$ is a vector of model parameters, and $\varepsilon_{i j}$ is the measurement or process error. The index $j$ ranges from 1 to $n_{i}$, where $n_{i}$ is the number of observations in group $i$. The function $f$ specifies the form of the model. Often, $x_{i j}$ is simply an observation time $t_{i j}$.

The errors are usually assumed to be independent and identically, normally distributed, with constant variance.

Estimates of the parameters in $\varphi$ describe the population, assuming those estimates are the same for all groups. If, however, the estimates vary by group, the model becomes

$$
y_{i j}=f\left(\varphi_{i}, x_{i j}\right)+\varepsilon_{i j}
$$

In a mixed-effects model, $\varphi_{i}$ may be a combination of a fixed and a random effect:

$$
\varphi_{i}=\beta+b_{i}
$$

The random effects $b_{i}$ are usually described as multivariate normally distributed, with mean zero and covariance $\Psi$. Estimating the fixed effects $\beta$ and the covariance of the random effects $\Psi$ provides a description of the population that does not assume the parameters $\varphi_{i}$ are the same across groups. Estimating the random effects $b_{i}$ also gives a description of specific groups within the data.

Model parameters do not have to be identified with individual effects. In general, design matrices $A$ and $B$ are used to identify parameters with linear combinations of fixed and random effects:

$$
\varphi_{i}=A \beta+B b_{i}
$$

If the design matrices differ among groups, the model becomes

$$
\varphi_{i}=A_{i} \beta+B_{i} b_{i}
$$

If the design matrices also differ among observations, the model becomes

$$
\begin{aligned}
\varphi_{i j} & =A_{i j} \beta+B_{i j} b_{i} \\
y_{i j} & =f\left(\varphi_{i j}, x_{i j}\right)+\varepsilon_{i j}
\end{aligned}
$$

Some of the group-specific predictors in $x_{i j}$ may not change with observation $j$. Calling those $v_{i}$, the model becomes

$$
y_{i j}=f\left(\varphi_{i j}, x_{i j}, v_{i}\right)+\varepsilon_{i j}
$$

## Specifying Mixed-Effects Models

Suppose data for a nonlinear regression model falls into one of $m$ distinct groups $i=1, \ldots, m$. (Specifically, suppose that the groups are not nested.) To specify a general nonlinear mixed-effects model for this data:

1 Define group-specific model parameters $\varphi_{i}$ as linear combinations of fixed effects $\beta$ and random effects $b_{i}$.

2 Define response values $y_{i}$ as a nonlinear function $f$ of the parameters and group-specific predictor variables $X_{i}$.

The model is:

$$
\begin{aligned}
\varphi_{i} & =A_{i} \beta+B_{i} b_{i} \\
y_{i} & =f\left(\varphi_{i}, X_{i}\right)+\varepsilon_{i} \\
b_{i} & \square N(0, \Psi) \\
\varepsilon_{i} & \square N\left(0, \sigma^{2}\right)
\end{aligned}
$$

This formulation of the nonlinear mixed-effects model uses the following notation:
$\varphi_{i} \quad$ A vector of group-specific model parameters
$\beta \quad$ A vector of fixed effects, modeling population parameters
$b_{i} \quad$ A vector of multivariate normally distributed group-specific random effects
$A_{i} \quad$ A group-specific design matrix for combining fixed effects
$B_{i} \quad$ A group-specific design matrix for combining random effects
$X_{i} \quad$ A data matrix of group-specific predictor values
$y_{i} \quad$ A data vector of group-specific response values

| $f$ | A general, real-valued function of $\varphi_{i}$ and $X_{i}$ |
| :--- | :--- |
| $\varepsilon_{i}$ | A vector of group-specific errors, assumed to be independent, <br> identically, normally distributed, and independent of $b_{i}$ |
| $\Psi$ | A covariance matrix for the random effects |
| $\sigma^{2}$ | The error variance, assumed to be constant across observations |

For example, consider a model of the elimination of a drug from the bloodstream. The model incorporates two overlapping phases:

- An initial phase $p$ during which drug concentrations reach equilibrium with surrounding tissues
- A second phase $q$ during which the drug is eliminated from the bloodstream

For data on multiple individuals $i$, the model is

$$
y_{i j}=C_{p i} e^{-r_{p i} t_{i j}}+C_{q i} e^{-r_{q i} t_{i j}}+\varepsilon_{i j},
$$

where $y_{i j}$ is the observed concentration in individual $i$ at time $t_{i j}$. The model allows for different sampling times and different numbers of observations for different individuals.

The elimination rates $r_{p i}$ and $r_{q i}$ must be positive to be physically meaningful. Enforce this by introducing the $\log$ rates $R_{p i}=\log \left(r_{p i}\right)$ and $R_{q i}=\log \left(r_{q i}\right)$ and reparametrizing the model:

$$
y_{i j}=C_{p i} e^{-\exp \left(R_{p i}\right) t_{i j}}+C_{q i} e^{-\exp \left(R_{q i}\right) t_{i j}}+\varepsilon_{i j}
$$

Choosing which parameters to model with random effects is an important consideration when building a mixed-effects model. One technique is to add random effects to all parameters, and use estimates of their variances to determine their significance in the model. An alternative is to fit the model separately to each group, without random effects, and look at the variation of the parameter estimates. If an estimate varies widely across groups, or if confidence intervals for each group have minimal overlap, the parameter is a good candidate for a random effect.

To introduce fixed effects $\beta$ and random effects $b_{i}$ for all model parameters, reexpress the model as follows:

$$
\begin{aligned}
y_{i j}= & {\left[\bar{C}_{p}+\left(C_{p i}-\bar{C}_{p}\right)\right] e^{-\exp \left[\bar{R}_{p}+\left(R_{p i}-\bar{R}_{p}\right)\right] t_{i j}}+} \\
& {\left[\bar{C}_{q}+\left(C_{q i}-\bar{C}_{q}\right)\right] e^{-\exp \left[\bar{R}_{q}+\left(R_{q i}-\bar{R}_{q}\right)\right] t_{i j}}+\varepsilon_{i j} } \\
= & \left(\beta_{1}+b_{1 i}\right) e^{-\exp \left(\beta_{2}+b_{2 i}\right) t_{i j}}+ \\
& \left(\beta_{3}+b_{3 i}\right) e^{-\exp \left(\beta_{4}+b_{4 i}\right) t_{i j}}+\varepsilon_{i j}
\end{aligned}
$$

In the notation of the general model:

$$
\beta=\left(\begin{array}{c}
\beta_{1} \\
\vdots \\
\beta_{4}
\end{array}\right), b_{i}=\left(\begin{array}{c}
b_{i 1} \\
\vdots \\
b_{i 4}
\end{array}\right), y_{i}=\left(\begin{array}{c}
y_{i 1} \\
\vdots \\
y_{i n_{i}}
\end{array}\right), X_{i}=\left(\begin{array}{c}
t_{i 1} \\
\vdots \\
t_{i n_{i}}
\end{array}\right),
$$

where $n_{i}$ is the number of observations of individual $i$. In this case, the design matrices $A_{i}$ and $B_{i}$ are, at least initially, 4-by-4 identity matrices. Design matrices may be altered, as necessary, to introduce weighting of individual effects, or time dependency.

Fitting the model and estimating the covariance matrix $\Psi$ often leads to further refinements. A relatively small estimate for the variance of a random effect suggests that it can be removed from the model. Likewise, relatively small estimates for covariances among certain random effects suggests that a full covariance matrix is unnecessary. Since random effects are unobserved, $\Psi$ must be estimated indirectly. Specifying a diagonal or block-diagonal covariance pattern for $\Psi$ can improve convergence and efficiency of the fitting algorithm.

Statistics Toolbox functions nlmefit and nlmefitsa fit the general nonlinear mixed-effects model to data, estimating the fixed and random effects. The functions also estimate the covariance matrix $\Psi$ for the random effects. Additional diagnostic outputs allow you to assess tradeoffs between the number of model parameters and the goodness of fit.

## Specifying Covariate Models

If the model in "Specifying Mixed-Effects Models" on page 9-81 assumes a group-dependent covariate such as weight ( $w$ ) the model becomes:

$$
\left(\begin{array}{l}
\varphi_{1} \\
\varphi_{2} \\
\varphi_{3}
\end{array}\right)=\left(\begin{array}{llll}
1 & 0 & 0 & w_{i} \\
0 & 1 & 0 & 0 \\
0 & 0 & 1 & 0
\end{array}\right)\left(\begin{array}{l}
\beta_{1} \\
\beta_{2} \\
\beta_{3} \\
\beta_{4}
\end{array}\right)+\left(\begin{array}{lll}
1 & 0 & 0 \\
0 & 1 & 0 \\
0 & 0 & 1
\end{array}\right)\left(\begin{array}{l}
b_{1} \\
b_{2} \\
b_{3}
\end{array}\right)
$$

Thus, the parameter $\varphi_{\mathrm{i}}$ for any individual in the ith group is:

$$
\left(\begin{array}{c}
\varphi_{1_{i}} \\
\varphi_{2_{i}} \\
\varphi_{3_{i}}
\end{array}\right)=\left(\begin{array}{c}
\beta_{1}+\beta_{4} * w_{i} \\
\beta_{2} \\
\beta_{3}
\end{array}\right)+\left(\begin{array}{l}
b_{1_{i}} \\
b_{2_{i}} \\
b_{3_{i}}
\end{array}\right)
$$

To specify a covariate model, use the 'FEGroupDesign' option.
'FEGroupDesign' is a $p-b y-q-b y-m$ array specifying a different $p-b y-q$ fixed-effects design matrix for each of the $m$ groups. Using the previous example, the array resembles the following:


1 Create the array.

```
% Number of parameters in the model (Phi)
num_params = 3;
```

```
% Number of covariates
num_cov = 1;
% Assuming number of groups in the data set is 7
num_groups = 7;
% Array of covariate values
covariates = [75; 52; 66; 55; 70; 58; 62 ];
A = repmat(eye(num_params, num_params+num_cov),...
[1,1,num_groups]);
A(1,num_params+1,1:num_groups) = covariates(:,1)
```

2 Create a struct with the specified design matrix.

```
options.FEGroupDesign = A;
```

3 Specify the arguments for nlmefit (or nlmefitsa) as shown in "Example: Mixed-Effects Models Using nlmefit and nlmefitsa" on page 9-93.

## Choosing nlmefit or nlmefitsa

Statistics Toolbox provides two functions, nlmefit and nlmefitsa for fitting non-linear mixed-effects models. Each function provides different capabilities, which may help you decide which to use.

- "Approximation Methods" on page 9-85
- "Parameters Specific to nlmefitsa" on page 9-86
- "Model and Data Requirements" on page 9-87

Approximation Methods. nlmefit provides the following four approximation methods for fitting non-linear mixed-effects models:

- 'LME' - Use the likelihood for the linear mixed-effects model at the current conditional estimates of beta and B. This is the default.
- 'RELME ' - Use the restricted likelihood for the linear mixed-effects model at the current conditional estimates of beta and $B$.
- 'FO' - First-order Laplacian approximation without random effects.
- 'FOCE ' - First-order Laplacian approximation at the conditional estimates of B.
nlmefitsa provides an additional approximation method, Stochastic Approximation Expectation-Maximization (SAEM) [24] with three steps :

1 Simulation: Generate simulated values of the random effects $b$ from the posterior density $p(b \mid \Sigma)$ given the current parameter estimates.

2 Stochastic approximation: Update the expected value of the log likelihood function by taking its value from the previous step, and moving part way toward the average value of the log likelihood calculated from the simulated random effects.

3 Maximization step: Choose new parameter estimates to maximize the log likelihood function given the simulated values of the random effects.

Both nlmefit and nlmefitsa attempt to find parameter estimates to maximize a likelihood function, which is difficult to compute. nlmefit deals with the problem by approximating the likelihood function in various ways, and maximizing the approximate function. It uses traditional optimization techniques that depend on things like convergence criteria and iteration limits.
nlmefitsa, on the other hand, simulates random values of the parameters in such a way that in the long run they converge to the values that maximize the exact likelihood function. The results are random, and traditional convergence tests don't apply. Therefore nlmefitsa provides options to plot the results as the simulation progresses, and to re-start the simulation multiple times. You can use these features to judge whether the results have converged to the accuracy you desire.

Parameters Specific to nlmefitsa. The following parameters are specific to nlmefitsa. Most control the stochastic algorithm.

- Cov0 - Initial value for the covariance matrix PSI. Must be an $r$-by- $r$ positive definite matrix. If empty, the default value depends on the values of BETAO.
- ComputeStdErrors - true to compute standard errors for the coefficient estimates and store them in the output STATS structure, or false (default) to omit this computation.
- LogLikMethod - Specifies the method for approximating the log likelihood.
- NBurnIn - Number of initial burn-in iterations during which the parameter estimates are not recomputed. Default is 5 .
- NIterations - Controls how many iterations are performed for each of three phases of the algorithm.
- NMCMCIterations - Number of Markov Chain Monte Carlo (MCMC) iterations.

Model and Data Requirements. There are some differences in the capabilities of nlmefit and nlmefitsa. Therefore some data and models are usable with either function, but some may require you to choose just one of them.

- Error models - nlmefitsa supports a variety of error models. For example, the standard deviation of the response can be constant, proportional to the function value, or a combination of the two. nlmefit fits models under the assumption that the standard deviation of the response is constant. One of the error models, 'exponential', specifies that the log of the response has a constant standard deviation. You can fit such models using nlmefit by providing the log response as input, and by re-writing the model function to produce the log of the nonlinear function value.
- Random effects - Both functions fit data to a nonlinear function with parameters, and the parameters may be simple scalar values or linear functions of covariates. nlmefit allows any coefficients of the linear functions to have both fixed and random effects. nlmefitsa supports random effects only for the constant (intercept) coefficient of the linear functions, but not for slope coefficients. So in the example in "Specifying Covariate Models" on page 9-84, nlmefitsa can treat only the first three beta values as random effects.
- Model form - nlmefit supports a very general model specification, with few restrictions on the design matrices that relate the fixed coefficients and the random effects to the model parameters. nlmefitsa is more restrictive:
- The fixed effect design must be constant in every group (for every individual), so an observation-dependent design is not supported.
- The random effect design must be constant for the entire data set, so neither an observation-dependent design nor a group-dependent design is supported.
- As mentioned under Random Effects, the random effect design must not specify random effects for slope coefficients. This implies that the design must consist of zeros and ones.
- The random effect design must not use the same random effect for multiple coefficients, and cannot use more than one random effect for any single coefficient.
- The fixed effect design must not use the same coefficient for multiple parameters. This implies that it can have at most one non-zero value in each column.
If you want to use nlmefitsa for data in which the covariate effects are random, include the covariates directly in the nonlinear model expression. Don't include the covariates in the fixed or random effect design matrices.
- Convergence - As described in the Model form, nlmefit and nlmefitsa have different approaches to measuring convergence. nlmefit uses traditional optimization measures, and nlmefitsa provides diagnostics to help you judge the convergence of a random simulation.

In practice, nlmefitsa tends to be more robust, and less likely to fail on difficult problems. However, nlmefit may converge faster on problems where it converges at all. Some problems may benefit from a combined strategy, for example by running nlmefitsa for a while to get reasonable parameter estimates, and using those as a starting point for additional iterations using nlmefit.

## Using Output Functions with Mixed-Effects Models

The Outputfen field of the options structure specifies one or more functions that the solver calls after each iteration. Typically, you might use an output function to plot points at each iteration or to display optimization quantities from the algorithm. To set up an output function:

1 Write the output function as a MATLAB file function or subfunction.
2 Use statset to set the value of Outputfen to be a function handle, that is, the name of the function preceded by the @ sign. For example, if the output function is outfun.m, the command

```
options = statset('OutputFcn', @outfun);
```

specifies OutputFen to be the handle to outfun. To specify multiple output functions, use the syntax:

```
options = statset('OutputFcn',{@outfun, @outfun2});
```

3 Call the optimization function with options as an input argument.
For an example of an output function, see "Sample Output Function" on page 9-93.

Structure of the Output Function. The function definition line of the output function has the following form:

```
stop = outfun(beta,status,state)
```

where

- beta is the current fixed effects.
- status is a structure containing data from the current iteration. "Fields in status" on page 9-89 describes the structure in detail.
- state is the current state of the algorithm. "States of the Algorithm" on page 9-90 lists the possible values.
- stop is a flag that is true or false depending on whether the optimization routine should quit or continue. See "Stop Flag" on page 9-91 for more information.

The solver passes the values of the input arguments to outfun at each iteration.

Fields in status. The following table lists the fields of the status structure:

| Field | Description |
| :--- | :--- |
| procedure | • 'ALT' - alternating algorithm for the optimization of <br> the linear mixed effects or restricted linear mixed effects <br> approximations |
|  | - 'LAP' - optimization of the Laplacian approximation for <br> first order or first order conditional estimation |
| iteration | An integer starting from 0. |


| Field | Description |
| :---: | :---: |
| inner | A structure describing the status of the inner iterations within the ALT and LAP procedures, with the fields: <br> - procedure - When procedure is 'ALT': <br> - 'PNLS' (penalized non-linear least squares) <br> - 'LME' (linear mixed-effects estimation) <br> - 'none' <br> When procedure is 'LAP', <br> - 'PNLS' (penalized non-linear least squares) <br> - 'PLM' (profiled likelihood maximization) <br> - 'none' <br> - state - one of the following: <br> - 'init' <br> - 'iter' <br> - 'done' <br> - 'none' <br> - iteration - an integer starting from 0, or NaN. For nlmefitsa with burn-in iterations, the output function is called after each of those iterations with a negative value for STATUS.iteration. |
| fval | The current log likelihood |
| Psi | The current random-effects covariance matrix |
| theta | The current parameterization of Psi |
| mse | The current error variance |

States of the Algorithm. The following table lists the possible values for state:

| state | Description |
| :--- | :--- |
| 'init' | The algorithm is in the initial state before the first <br> iteration. |
| 'iter' | The algorithm is at the end of an iteration. |
| 'done' | The algorithm is in the final state after the last iteration. |

The following code illustrates how the output function might use the value of state to decide which tasks to perform at the current iteration:

```
switch state
    case 'iter'
        % Make updates to plot or guis as needed
    case 'init'
        % Setup for plots or guis
    case 'done'
        % Cleanup of plots, guis, or final plot
otherwise
end
```

Stop Flag. The output argument stop is a flag that is true or false. The flag tells the solver whether it should quit or continue. The following examples show typical ways to use the stop flag.

## Stopping an Optimization Based on Intermediate Results

The output function can stop the estimation at any iteration based on the values of arguments passed into it. For example, the following code sets stop to true based on the value of the log likelihood stored in the 'fval 'field of the status structure:

```
stop = outfun(beta,status,state)
stop = false;
% Check if loglikelihood is more than 132.
if status.fval > -132
    stop = true;
end
```


## Stopping an Iteration Based on GUI Input

If you design a GUI to perform nlmefit iterations, you can make the output function stop when a user clicks a Stop button on the GUI. For example, the following code implements a dialog to cancel calculations:

```
function retval = stop_outfcn(beta,str,status)
persistent h stop;
if isequal(str.inner.state,'none')
    switch(status)
        case 'init'
            % Initialize dialog
            stop = false;
            h = msgbox('Press STOP to cancel calculations.',...
                    'NLMEFIT: Iteration O ');
            button = findobj(h,'type','uicontrol');
            set(button,'String','STOP','Callback',@stopper)
            pos = get(h,'Position');
            pos(3) = 1.1 * pos(3);
            set(h,'Position',pos)
            drawnow
            case 'iter'
            % Display iteration number in the dialog title
            set(h,'Name',sprintf('NLMEFIT: Iteration %d',...
                    str.iteration))
            drawnow;
        case 'done'
            % Delete dialog
            delete(h);
        end
end
if stop
        % Stop if the dialog button has been pressed
        delete(h)
end
retval = stop;
        function stopper(varargin)
        % Set flag to stop when button is pressed
        stop = true;
```

```
    disp('Calculation stopped.')
    end
end
```

Sample Output Function. nmlefitoutputfon is the sample Statistics Toolbox output function for nlmefit and nlmefitsa. It initializes or updates a plot with the fixed-effects (BETA) and variance of the random effects (diag(STATUS.Psi)). For nlmefit, the plot also includes the log-likelihood (STATUS.fval).
nlmefitoutputfon is the default output function for nlmefitsa. To use it with nlmefit, specify a function handle for it in the options structure:

```
opt = statset('OutputFcn', @nlmefitoutputfcn, )
beta = nlmefit( , 'Options', opt, )
```

To prevent nlmefitsa from using of this function, specify an empty value for the output function:

```
opt = statset('OutputFcn', [], )
beta = nlmefitsa( , 'Options', opt, )
```

nlmefitoutputfon stops nlmefit or nlmefitsa if you close the figure that it produces.

## Example: Mixed-Effects Models Using nlmefit and nlmefitsa

The following example also works with nlmefitsa in place of nlmefit.
The data in indomethacin.mat records concentrations of the drug indomethacin in the bloodstream of six subjects over eight hours:

```
load indomethacin
gscatter(time,concentration,subject)
xlabel('Time (hours)')
ylabel('Concentration (mcg/ml)')
title('{\bf Indomethacin Elimination}')
hold on
```

Indomethacin Elimination

"Specifying Mixed-Effects Models" on page 9-81 discusses a useful model for this type of data. Construct the model via an anonymous function as follows:

```
model = @(phi,t)(phi(1)*exp(- exp(phi(2))*t) + ...
    phi(3)*exp(- exp(phi(4))*t));
```

Use the nlinfit function to fit the model to all of the data, ignoring subject-specific effects:

```
phiO = [1 1 2 1 1];
[phi,res] = nlinfit(time,concentration,model,phi0);
numObs = length(time);
numParams = 4;
df = numObs-numParams;
mse = (res'*res)/df
```

```
    mse =
        0.0304
    tplot = 0:0.01:8;
plot(tplot,model(phi,tplot),'k','LineWidth',2)
hold off
Indomethacin Elimination
```



A box plot of residuals by subject shows that the boxes are mostly above or below zero, indicating that the model has failed to account for subject-specific effects:

```
colors = 'rygcbm';
h = boxplot(res,subject,'colors',colors,'symbol','o');
set(h(~isnan(h)),'LineWidth', 2)
hold on
boxplot(res,subject,'colors','k','symbol','ko')
```

```
grid on
xlabel('Subject')
ylabel('Residual')
hold off
```



To account for subject-specific effects, fit the model separately to the data for each subject:

```
phiO = [1 2 1 1];
PHI = zeros(4,6);
RES = zeros(11,6);
for I = 1:6
    tI = time(subject == I);
    cI = concentration(subject == I);
    [PHI(:,I),RES(:,I)] = nlinfit(tI,cI,model,phiO);
end
```

```
PHI
PHI =
    2.0293 2.8277 5.4683 2.1981 3.5661 3.0023
    0.5794 0.8013 1.7498 0.2423 1.0408 1.0882
    0.1915 0.4989 1.6757 0.2545 0.2915 0.9685
    -1.7878 -1.6354 -0.4122 -1.6026 -1.5069 -0.8731
numParams = 24;
df = numObs-numParams;
mse = (RES(:)'*RES(:))/df
mse =
    0.0057
gscatter(time,concentration, subject)
xlabel('Time (hours)')
ylabel('Concentration (mcg/ml)')
title('{\bf Indomethacin Elimination}')
hold on
for I = 1:6
    plot(tplot,model(PHI(:, I),tplot),'Color',colors(I))
end
axis([0}08003.5]
hold off
```


## Indomethacin Elimination



PHI gives estimates of the four model parameters for each of the six subjects. The estimates vary considerably, but taken as a 24 -parameter model of the data, the mean-squared error of 0.0057 is a significant reduction from 0.0304 in the original four-parameter model.

A box plot of residuals by subject shows that the larger model accounts for most of the subject-specific effects:

```
h = boxplot(RES,'colors',colors,'symbol','o');
set(h(~isnan(h)),'LineWidth',2)
hold on
boxplot(RES,'colors','k','symbol','ko')
grid on
xlabel('Subject')
ylabel('Residual')
```

The spread of the residuals (the vertical scale of the box plot) is much smaller than in the previous box plot, and the boxes are now mostly centered on zero.

While the 24-parameter model successfully accounts for variations due to the specific subjects in the study, it does not consider the subjects as representatives of a larger population. The sampling distribution from which the subjects are drawn is likely more interesting than the sample itself. The purpose of mixed-effects models is to account for subject-specific variations more broadly, as random effects varying around population means.

Use the nlmefit function to fit a mixed-effects model to the data.
The following anonymous function, nlme_model, adapts the four-parameter model used by nlinfit to the calling syntax of nlmefit by allowing separate
parameters for each individual. By default, nlmefit assigns random effects to all the model parameters. Also by default, nlmefit assumes a diagonal covariance matrix (no covariance among the random effects) to avoid overparametrization and related convergence issues.

```
nlme_model = @(PHI,t)(PHI(:,1).*exp(-exp(PHI(:,2)).*t) + ...
    PHI(:, 3).* exp(-exp(PHI(:,4)).*t));
phiO = [1 2 1 1];
[phi,PSI,stats] = nlmefit(time,concentration,subject, ...
                                    [],nlme_model,phiO)
phi =
    2.8277
    0.7729
    0.4606
    -1.3459
PSI =
    0.3264 0 0 0
            0
            0 0
            0 0 0
stats =
            dfe: 57
            logl: 54.5882
            mse: 0.0066
            rmse: 0.0787
        errorparam: 0.0815
            aic: -91.1765
            bic: -93.0506
            covb: [4x4 double]
            sebeta: [0.2558 0.1066 0.1092 0.2244]
            ires: [66x1 double]
            pres: [66x1 double]
            iwres: [66x1 double]
            pwres: [66x1 double]
            cwres: [66x1 double]
```

The mean-squared error of 0.0066 is comparable to the 0.0057 of the 24 -parameter model without random effects, and significantly better than the 0.0304 of the four-parameter model without random effects.

The estimated covariance matrix PSI shows that the variance of the fourth random effect is essentially zero, suggesting that you can remove it to simplify the model. To do this, use the REParamsSelect parameter to specify the indices of the parameters to be modeled with random effects in nlmefit:

```
[phi,PSI,stats] = nlmefit(time,concentration,subject, ...
    [],nlme_model,phiO, ...
    'REParamsSelect',[1 2 3])
phi =
    2.8277
    0.7728
    0.4605
    -1.3460
PSI =
    0.3270 0 0
        0 0.0250 0
        0 0}00.012
stats =
            dfe: 58
            logl: 54.5875
            mse: 0.0066
            rmse: 0.0780
        errorparam: 0.0815
            aic: -93.1750
            bic: -94.8410
            covb: [4x4 double]
            sebeta: [0.2560 0.1066 0.1092 0.2244]
            ires: [66x1 double]
            pres: [66x1 double]
            iwres: [66x1 double]
            pwres: [66x1 double]
            cwres: [66x1 double]
```

The log-likelihood logl is almost identical to what it was with random effects for all of the parameters, the Akaike information criterion aic is reduced from - 91.1765 to -93.1750 , and the Bayesian information criterion bic is reduced from -93.0506 to -94.8410. These measures support the decision to drop the fourth random effect.

Refitting the simplified model with a full covariance matrix allows for identification of correlations among the random effects. To do this, use the CovPattern parameter to specify the pattern of nonzero elements in the covariance matrix:

```
[phi,PSI,stats] = nlmefit(time,concentration,subject, ...
    [],nlme_model,phi0, ...
    'REParamsSelect',[1 2 3], ...
    'CovPattern',ones(3))
phi =
    2.8148
    0.8293
    0.5613
    -1.1407
PSI =
    0.4767 0.1152 0.0499
    0.1152 0.0321 0.0032
    0.0499 0.0032 0.0236
stats =
        dfe: 55
        logl: 58.4731
            mse: 0.0061
            rmse: 0.0782
        errorparam: 0.0781
            aic: -94.9462
            bic: -97.2369
            covb: [4x4 double]
            sebeta: [0.3028 0.1103 0.1179 0.1662]
            ires: [66x1 double]
            pres: [66x1 double]
            iwres: [66x1 double]
            pwres: [66x1 double]
```

```
cwres: [66x1 double]
```

The estimated covariance matrix PSI shows that the random effects on the first two parameters have a relatively strong correlation, and both have a relatively weak correlation with the last random effect. This structure in the covariance matrix is more apparent if you convert PSI to a correlation matrix using corrcov:

```
RHO = corrcov(PSI)
RHO =
    1.0000 0.9316 0.4706
    0.9316 1.0000 0.1178
    0.4706 0.1178 1.0000
clf; imagesc(RHO)
set(gca,'XTick',[1 2 3],'YTick',[1 2 3])
title('{\bf Random Effect Correlation}')
h = colorbar;
set(get(h,'YLabel'),'String','Correlation');
```

Random Effect Correlation


Incorporate this structure into the model by changing the specification of the covariance pattern to block-diagonal:

```
P = [1 1 0;1 1 0;0 0 1] % Covariance pattern
P =
\begin{tabular}{lll}
1 & 1 & 0 \\
1 & 1 & 0 \\
0 & 0 & 1
\end{tabular}
[phi,PSI,stats,b] = nlmefit(time,concentration,subject, ...
    [],nlme_model,phi0, ...
    'REParamsSelect',[1 2 3], ...
    'CovPattern',P)
```

```
phi =
    2.7830
    0.8981
    0.6581
    -1.0000
PSI =
    0.5180 0.1069 0
    0.1069 0.0221 0
            0 0 0.0454
stats =
            dfe: 57
            logl: 58.0804
            mse: 0.0061
            rmse: 0.0768
            errorparam: 0.0782
                    aic: -98.1608
                    bic: -100.0350
                    covb: [4x4 double]
            sebeta: [[0.3171 0.1073 0.1384 0.1453]
                    ires: [66x1 double]
                    pres: [66x1 double]
            iwres: [66x1 double]
            pwres: [66x1 double]
            cwres: [66x1 double]
b =
\begin{tabular}{rrrrrr}
-0.8507 & -0.1563 & 1.0427 & -0.7559 & 0.5652 & 0.1550 \\
-0.1756 & -0.0323 & 0.2152 & -0.1560 & 0.1167 & 0.0320 \\
-0.2756 & 0.0519 & 0.2620 & 0.1064 & -0.2835 & 0.1389
\end{tabular}
```

The block-diagonal covariance structure reduces aic from -94.9462 to -98.1608 and bic from -97.2368 to -100.0350 without significantly affecting the log-likelihood. These measures support the covariance structure used in the final model.

The output $b$ gives predictions of the three random effects for each of the six subjects. These are combined with the estimates of the fixed effects in phi to produce the mixed-effects model.

Use the following commands to plot the mixed-effects model for each of the six subjects. For comparison, the model without random effects is also shown.

```
PHI = repmat(phi,1,6) + ... % Fixed effects
    [b(1,:);b(2,:);b(3,:);zeros(1,6)]; % Random effects
RES = zeros(11,6); % Residuals
colors = 'rygcbm';
for I = 1:6
    fitted_model = @(t)(PHI(1,I)*exp(-\operatorname{exp}(\operatorname{PHI}(2,I))*t) + ...
                                    PHI (3,I)*exp(- exp(PHI (4,I)) *t));
    tI = time(subject == I);
    cI = concentration(subject == I);
    RES(:,I) = cI - fitted_model(tI);
    subplot(2,3,I)
    scatter(tI,cI,20,colors(I),'filled')
    hold on
    plot(tplot,fitted_model(tplot),'Color',colors(I))
    plot(tplot,model(phi,tplot),'k')
    axis([0 8 0 3.5])
    xlabel('Time (hours)')
    ylabel('Concentration (mcg/ml)')
    legend(num2str(I),'Subject','Fixed')
end
```



If obvious outliers in the data (visible in previous box plots) are ignored, a normal probability plot of the residuals shows reasonable agreement with model assumptions on the errors:

```
clf; normplot(RES(:))
```



## Example: Examining Residuals for Model Verification

You can examine the stats structure, which is returned by both nlmefit and nlmefitsa, to determine the quality of your model. The stats structure contains fields with conditional weighted residuals (cwres field) and individual weighted residuals (iwres field). Since the model assumes that residuals are normally distributed, you can examine the residuals to see how well this assumption holds.

This example generates synthetic data using normal distributions. It shows how the fit statistics look:

- Good when testing against the same type of model as generates the data
- Poor when tested against incorrect data models

1 Initialize a 2-D model with 100 individuals:

```
nGroups = 100; % 100 Individuals
nlmefun = @(PHI,t)(PHI(:,1)*5 + PHI(:,2)^2.*t); % Regression fon
REParamSelect = [1 2]; % Both Parameters have random effect
errorParam = .03;
beta0 = [ 1.5 5]; % Parameter means
psi = [ 0.35 0; ... % Covariance Matrix
    0 0.51 ];
time =[0.25;0.5;0.75;1;1.25;2;3;4;5;6];
nParameters = 2;
rng(0,'twister') % for reproducibility
```

2 Generate the data for fitting with a proportional error model:

```
b_i = mvnrnd(zeros(1, numel(REParamSelect)), psi, nGroups);
individualParameters = zeros(nGroups,nParameters);
individualParameters(:, REParamSelect) = ...
    bsxfun(@plus,beta0(REParamSelect), b_i);
groups = repmat(1:nGroups,numel(time),1);
groups = vertcat(groups(:));
y = zeros(numel(time)*nGroups,1);
x = zeros(numel(time)*nGroups,1);
for i = 1:nGroups
    idx = groups == i;
    f = nlmefun(individualParameters(i,:), time);
    % Make a proportional error model for y:
    y(idx) = f + errorParam*f.*randn(numel(f),1);
    x(idx) = time;
end
P = [ 1 0 ; 0 1 ];
```

3 Fit the data using the same regression function and error model as the model generator:

```
[~,~,stats] = nlmefit(x,y,groups, ...
    [],nlmefun,[1 1],'REParamsSelect',REParamSelect,...
    'ErrorModel','Proportional','CovPattern',P);
```

4 Create a plotting routine by copying the following function definition, and creating a file plotResiduals.m on your MATLAB path:

```
function plotResiduals(stats)
pwres = stats.pwres;
iwres = stats.iwres;
cwres = stats.cwres;
figure
subplot(2,3,1);
normplot(pwres); title('PWRES')
subplot(2,3,4);
createhistplot(pwres);
subplot(2,3,2);
normplot(cwres); title('CWRES')
subplot(2,3,5);
createhistplot(cwres);
subplot(2,3,3);
normplot(iwres); title('IWRES')
subplot(2,3,6);
createhistplot(iwres); title('IWRES')
function createhistplot(pwres)
[x, n] = hist(pwres);
d = n(2)- n(1);
x = x/sum(x*d);
bar(n,x);
ylim([0 max(x)*1.05]);
hold on;
x2 = -4:0.1:4;
f2 = normpdf(x2,0,1);
plot(x2,f2,'r');
end
end
```

5 Plot the residuals using the plotResiduals function:
plotResiduals(stats);


The upper probability plots look straight, meaning the residuals are normally distributed. The bottom histogram plots match the superimposed normal density plot. So you can conclude that the error model matches the data.

6 For comparison, fit the data using a constant error model, instead of the proportional model that created the data:

```
[~,~,stats] = nlmefit(x,y,groups, ...
    [],nlmefun,[0 0],'REParamsSelect',REParamSelect,...
    'ErrorModel','Constant','CovPattern',P);
plotResiduals(stats);
```



The upper probability plots are not straight, indicating the residuals are not normally distributed. The bottom histogram plots are fairly close to the superimposed normal density plots.

7 For another comparison, fit the data to a different structural model than created the data:

```
nlmefun2 = @(PHI,t)(PHI(:,1)*5 + PHI(:,2).*t.^4);
[~,~,stats] = nlmefit(x,y,groups, ...
            [],nlmefun2,[0 0],'REParamsSelect',REParamSelect,...
            'ErrorModel','constant', 'CovPattern',P);
plotResiduals(stats);
```



Not only are the upper probability plots not straight, but the histogram plot is quite skewed compared to the superimposed normal density. These residuals are not normally distributed, and do not match the model.

## Multivariate Methods

- "Introduction to Multivariate Methods" on page 10-2
- "Multidimensional Scaling" on page 10-3
- "Procrustes Analysis" on page 10-14
- "Feature Selection" on page 10-23
- "Feature Transformation" on page 10-28


## Introduction to Multivariate Methods

Large, high-dimensional data sets are common in the modern era of computer-based instrumentation and electronic data storage. High-dimensional data present many challenges for statistical visualization, analysis, and modeling.

Data visualization, of course, is impossible beyond a few dimensions. As a result, pattern recognition, data preprocessing, and model selection must rely heavily on numerical methods.

A fundamental challenge in high-dimensional data analysis is the so-called curse of dimensionality. Observations in a high-dimensional space are necessarily sparser and less representative than those in a low-dimensional space. In higher dimensions, data over-represent the edges of a sampling distribution, because regions of higher-dimensional space contain the majority of their volume near the surface. (A d-dimensional spherical shell has a volume, relative to the total volume of the sphere, that approaches 1 as $d$ approaches infinity.) In high dimensions, typical data points at the interior of a distribution are sampled less frequently.

Often, many of the dimensions in a data set-the measured features-are not useful in producing a model. Features may be irrelevant or redundant. Regression and classification algorithms may require large amounts of storage and computation time to process raw data, and even if the algorithms are successful the resulting models may contain an incomprehensible number of terms.

Because of these challenges, multivariate statistical methods often begin with some type of dimension reduction, in which data are approximated by points in a lower-dimensional space. Dimension reduction is the goal of the methods presented in this chapter. Dimension reduction often leads to simpler models and fewer measured variables, with consequent benefits when measurements are expensive and visualization is important.

## Multidimensional Scaling

In this section...<br>"Introduction to Multidimensional Scaling" on page 10-3<br>"Classical Multidimensional Scaling" on page 10-3<br>"Nonclassical Multidimensional Scaling" on page 10-8<br>"Nonmetric Multidimensional Scaling" on page 10-10

## Introduction to Multidimensional Scaling

One of the most important goals in visualizing data is to get a sense of how near or far points are from each other. Often, you can do this with a scatter plot. However, for some analyses, the data that you have might not be in the form of points at all, but rather in the form of pairwise similarities or dissimilarities between cases, observations, or subjects. There are no points to plot.

Even if your data are in the form of points rather than pairwise distances, a scatter plot of those data might not be useful. For some kinds of data, the relevant way to measure how near two points are might not be their Euclidean distance. While scatter plots of the raw data make it easy to compare Euclidean distances, they are not always useful when comparing other kinds of inter-point distances, city block distance for example, or even more general dissimilarities. Also, with a large number of variables, it is very difficult to visualize distances unless the data can be represented in a small number of dimensions. Some sort of dimension reduction is usually necessary.

Multidimensional scaling (MDS) is a set of methods that address all these problems. MDS allows you to visualize how near points are to each other for many kinds of distance or dissimilarity metrics and can produce a representation of your data in a small number of dimensions. MDS does not require raw data, but only a matrix of pairwise distances or dissimilarities.

## Classical Multidimensional Scaling

- "Introduction to Classical Multidimensional Scaling" on page 10-4
- "Example: Multidimensional Scaling" on page 10-6


## Introduction to Classical Multidimensional Scaling

The function cmdscale performs classical (metric) multidimensional scaling, also known as principal coordinates analysis. cmdscale takes as an input a matrix of inter-point distances and creates a configuration of points. Ideally, those points are in two or three dimensions, and the Euclidean distances between them reproduce the original distance matrix. Thus, a scatter plot of the points created by cmdscale provides a visual representation of the original distances.

As a very simple example, you can reconstruct a set of points from only their inter-point distances. First, create some four dimensional points with a small component in their fourth coordinate, and reduce them to distances.

```
X = [ normrnd(0,1,10,3), normrnd(0,.1,10,1) ];
D = pdist(X,'euclidean');
```

Next, use cmdscale to find a configuration with those inter-point distances. cmdscale accepts distances as either a square matrix, or, as in this example, in the vector upper-triangular form produced by pdist.
[Y,eigvals] = cmdscale(D);
cmdscale produces two outputs. The first output, Y , is a matrix containing the reconstructed points. The second output, eigvals, is a vector containing the sorted eigenvalues of what is often referred to as the "scalar product matrix," which, in the simplest case, is equal to $Y^{*} Y^{\prime}$. The relative magnitudes of those eigenvalues indicate the relative contribution of the corresponding columns of $Y$ in reproducing the original distance matrix $D$ with the reconstructed points.

```
format short g
[eigvals eigvals/max(abs(eigvals))]
ans =
    12.623 1
    4.3699 0.34618
    1.9307 0.15295
    0.025884 0.0020505
    1.7192e-015 1.3619e-016
    6.8727e-016 5.4445e-017
```

$$
\begin{array}{rr}
4.4367 e-017 & 3.5147 e-018 \\
-9.2731 e-016 & -7.3461 e-017 \\
-1.327 e-015 & -1.0513 e-016 \\
-1.9232 e-015 & -1.5236 e-016
\end{array}
$$

If eigvals contains only positive and zero (within round-off error) eigenvalues, the columns of $Y$ corresponding to the positive eigenvalues provide an exact reconstruction of $D$, in the sense that their inter-point Euclidean distances, computed using pdist, for example, are identical (within round-off) to the values in $D$.

```
maxerr4 = max(abs(D - pdist(Y))) % exact reconstruction
maxerr4 =
    2.6645e-015
```

If two or three of the eigenvalues in eigvals are much larger than the rest, then the distance matrix based on the corresponding columns of $Y$ nearly reproduces the original distance matrix $D$. In this sense, those columns form a lower-dimensional representation that adequately describes the data. However it is not always possible to find a good low-dimensional reconstruction.

```
% good reconstruction in 3D
maxerr3 = max(abs(D - pdist(Y(:,1:3))))
maxerr3 =
    0.029728
% poor reconstruction in 2D
maxerr2 = max(abs(D - pdist(Y(:,1:2))))
maxerr2 =
    0.91641
```

The reconstruction in three dimensions reproduces D very well, but the reconstruction in two dimensions has errors that are of the same order of magnitude as the largest values in D.

```
max(max(D))
ans =
    3.4686
```

Often, eigvals contains some negative eigenvalues, indicating that the distances in D cannot be reproduced exactly. That is, there might not be any configuration of points whose inter-point Euclidean distances are given by D. If the largest negative eigenvalue is small in magnitude relative to the largest positive eigenvalues, then the configuration returned by cmdscale might still reproduce $D$ well.

## Example: Multidimensional Scaling

Given only the distances between 10 US cities, cmdscale can construct a map of those cities. First, create the distance matrix and pass it to cmdscale. In this example, $D$ is a full distance matrix: it is square and symmetric, has positive entries off the diagonal, and has zeros on the diagonal.

```
cities = ...
{'Atl','Chi','Den','Hou','LA','Mia','NYC','SF','Sea','WDC'};
D = [ 0 587 1212 701 1936 604 748 2139 2182 543;
    587 0 920 940 1745 1188 713 1858 1737 597;
    1212 920 0 879 831 1726 1631 949 1021 1494;
        701 940 879 0 1374 968 1420 1645 1891 1220;
        1936 1745 831 1374 0 2339 2451 347 959 2300;
        604 1188 1726 968 2339 0 1092 2594 2734 923;
        748 713 1631 1420 2451 1092 0 2571 2408 205;
        2139 1858 949 1645 347 2594 2571 0 678 2442;
        2182 1737 1021 1891 959 2734 2408 678 0, 2329;
[Y,eigvals] = cmdscale(D);
```

Next, look at the eigenvalues returned by cmdscale. Some of these are negative, indicating that the original distances are not Euclidean. This is because of the curvature of the earth.

```
format short g
[eigvals eigvals/max(abs(eigvals))]
ans =
    9.5821e+006 1
    1.6868e+006 0.17604
        8157.3 0.0008513
        1432.9 0.00014954
        508.67 5.3085e-005
        25.143 2.624e-006
```

```
5.3394e-010 5.5722e-017
    -897.7 -9.3685e-005
    -5467.6 -0.0005706
    -35479 -0.0037026
```

However, in this case, the two largest positive eigenvalues are much larger in magnitude than the remaining eigenvalues. So, despite the negative eigenvalues, the first two coordinates of $Y$ are sufficient for a reasonable reproduction of $D$.

```
Dtriu = D(find(tril(ones(10),-1)))';
maxrelerr = max(abs(Dtriu-pdist(Y(:,1:2))))./max(Dtriu)
maxrelerr =
    0.0075371
```

Here is a plot of the reconstructed city locations as a map. The orientation of the reconstruction is arbitrary. In this case, it happens to be close to, although not exactly, the correct orientation.

```
plot(Y(:,1),Y(:,2),'.')
text(Y(:,1)+25,Y(:,2),cities)
xlabel('Miles')
ylabel('Miles')
```



## Nonclassical Multidimensional Scaling

The function mdscale performs nonclassical multidimensional scaling. As with cmdcale, you use mdscale either to visualize dissimilarity data for which no "locations" exist, or to visualize high-dimensional data by reducing its dimensionality. Both functions take a matrix of dissimilarities as an input and produce a configuration of points. However, mdscale offers a choice of different criteria to construct the configuration, and allows missing data and weights.

For example, the cereal data include measurements on 10 variables describing breakfast cereals. You can use mdscale to visualize these data in two dimensions. First, load the data. For clarity, this example code selects a subset of 22 of the observations.

```
load cereal.mat
X = [Calories Protein Fat Sodium Fiber ...
    Carbo Sugars Shelf Potass Vitamins];
```

```
% Take a subset from a single manufacturer
mfg1 = strcmp('G',cellstr(Mfg));
X = X(mfg1,:);
size(X)
ans =
    22 10
```

Then use pdist to transform the 10 -dimensional data into dissimilarities. The output from pdist is a symmetric dissimilarity matrix, stored as a vector containing only the ( $23 * 22 / 2$ ) elements in its upper triangle.

```
dissimilarities = pdist(zscore(X),'cityblock');
size(dissimilarities)
ans =
    1 231
```

This example code first standardizes the cereal data, and then uses city block distance as a dissimilarity. The choice of transformation to dissimilarities is application-dependent, and the choice here is only for simplicity. In some applications, the original data are already in the form of dissimilarities.

Next, use mdscale to perform metric MDS. Unlike cmdscale, you must specify the desired number of dimensions, and the method to use to construct the output configuration. For this example, use two dimensions. The metric STRESS criterion is a common method for computing the output; for other choices, see the mdscale reference page in the online documentation. The second output from mdscale is the value of that criterion evaluated for the output configuration. It measures the how well the inter-point distances of the output configuration approximate the original input dissimilarities:

```
[Y,stress] =...
mdscale(dissimilarities,2,'criterion','metricstress');
stress
stress =
    0.1856
```

A scatterplot of the output from mdscale represents the original 10 -dimensional data in two dimensions, and you can use the gname function to label selected points:

```
plot(Y(:,1),Y(:, 2),'o','LineWidth', 2);
```

gname (Name (mfg1))


## Nonmetric Multidimensional Scaling

Metric multidimensional scaling creates a configuration of points whose inter-point distances approximate the given dissimilarities. This is sometimes too strict a requirement, and non-metric scaling is designed to relax it a bit. Instead of trying to approximate the dissimilarities themselves, non-metric scaling approximates a nonlinear, but monotonic, transformation of them. Because of the monotonicity, larger or smaller distances on a plot of the output will correspond to larger or smaller dissimilarities, respectively. However, the nonlinearity implies that mdscale only attempts to preserve the
ordering of dissimilarities. Thus, there may be contractions or expansions of distances at different scales.

You use mdscale to perform nonmetric MDS in much the same way as for metric scaling. The nonmetric STRESS criterion is a common method for computing the output; for more choices, see the mdscale reference page in the online documentation. As with metric scaling, the second output from mdscale is the value of that criterion evaluated for the output configuration. For nonmetric scaling, however, it measures the how well the inter-point distances of the output configuration approximate the disparities. The disparities are returned in the third output. They are the transformed values of the original dissimilarities:

```
[Y,stress,disparities] = ...
mdscale(dissimilarities,2,'criterion','stress');
stress
stress =
    0.1562
```

To check the fit of the output configuration to the dissimilarities, and to understand the disparities, it helps to make a Shepard plot:

```
distances = pdist(Y);
[dum,ord] = sortrows([disparities(:) dissimilarities(:)]);
plot(dissimilarities,distances,'bo', ...
    dissimilarities(ord),disparities(ord),'r.-', ...
    [0 25],[0 25],'k-')
xlabel('Dissimilarities')
ylabel('Distances/Disparities')
legend({'Distances' 'Disparities' '1:1 Line'},...
    'Location', 'NorthWest');
```



This plot shows that mdscale has found a configuration of points in two dimensions whose inter-point distances approximates the disparities, which in turn are a nonlinear transformation of the original dissimilarities. The concave shape of the disparities as a function of the dissimilarities indicates that fit tends to contract small distances relative to the corresponding dissimilarities. This may be perfectly acceptable in practice.
mdscale uses an iterative algorithm to find the output configuration, and the results can often depend on the starting point. By default, mdscale uses cmdscale to construct an initial configuration, and this choice often leads to a globally best solution. However, it is possible for mdscale to stop at a configuration that is a local minimum of the criterion. Such
cases can be diagnosed and often overcome by running mdscale multiple times with different starting points. You can do this using the 'start' and 'replicates' parameters. The following code runs five replicates of MDS, each starting at a different randomly-chosen initial configuration. The criterion value is printed out for each replication; mdscale returns the configuration with the best fit.

```
opts = statset('Display','final');
[Y,stress] =...
mdscale(dissimilarities,2,'criterion','stress',...
'start','random','replicates',5,'Options',opts);
35 iterations, Final stress criterion = 0.156209
3 1 ~ i t e r a t i o n s , ~ F i n a l ~ s t r e s s ~ c r i t e r i o n ~ = ~ 0 . 1 5 6 2 0 9 ~
4 8 \text { iterations, Final stress criterion = 0.171209}
33 iterations, Final stress criterion = 0.175341
3 2 ~ i t e r a t i o n s , ~ F i n a l ~ s t r e s s ~ c r i t e r i o n ~ = ~ 0 . 1 8 5 8 8 1
```

Notice that mdscale finds several different local solutions, some of which do not have as low a stress value as the solution found with the cmdscale starting point.

## Procrustes Analysis

In this section...<br>"Comparing Landmark Data" on page 10-14<br>"Data Input" on page 10-14<br>"Preprocessing Data for Accurate Results" on page 10-15<br>"Example: Comparing Handwritten Shapes" on page 10-16

## Comparing Landmark Data

The procrustes function analyzes the distribution of a set of shapes using Procrustes analysis. This analysis method matches landmark data (geometric locations representing significant features in a given shape) to calculate the best shape-preserving Euclidian transformations. These transformations minimize the differences in location between compared landmark data.

Procrustes analysis is also useful in conjunction with multidimensional scaling. In "Example: Multidimensional Scaling" on page 10-6 there is an observation that the orientation of the reconstructed points is arbitrary. Two different applications of multidimensional scaling could produce reconstructed points that are very similar in principle, but that look different because they have different orientations. The procrustes function transforms one set of points to make them more comparable to the other.

## Data Input

The procrustes function takes two matrices as input:

- The target shape matrix $X$ has dimension $n \times p$, where $n$ is the number of landmarks in the shape and $p$ is the number of measurements per landmark.
- The comparison shape matrix $Y$ has dimension $n \times q$ with $q \leq p$. If there are fewer measurements per landmark for the comparison shape than the target shape ( $q<p$ ), the function adds columns of zeros to $Y$, yielding an $n \times p$ matrix.

The equation to obtain the transformed shape, $Z$, is

$$
\begin{equation*}
Z=b Y T+c \tag{10-1}
\end{equation*}
$$

where:

- $b$ is a scaling factor that stretches $(b>1)$ or shrinks $(b<1)$ the points.
- $T$ is the orthogonal rotation and reflection matrix.
- $c$ is a matrix with constant values in each column, used to shift the points.

The procrustes function chooses $b, T$, and $c$ to minimize the distance between the target shape $X$ and the transformed shape $Z$ as measured by the least squares criterion:

$$
\sum_{i=1}^{n} \sum_{j=1}^{p}\left(X_{i j}-Z_{i j}\right)^{2}
$$

## Preprocessing Data for Accurate Results

Procrustes analysis is appropriate when all $p$ measurement dimensions have similar scales. The analysis would be inaccurate, for example, if the columns of $Z$ had different scales:

- The first column is measured in milliliters ranging from 2,000 to 6,000 .
- The second column is measured in degrees Celsius ranging from 10 to 25 .
- The third column is measured in kilograms ranging from 50 to 230 .

In such cases, standardize your variables by:
1 Subtracting the sample mean from each variable.
2 Dividing each resultant variable by its sample standard deviation.
Use the zscore function to perform this standardization.

## Example: Comparing Handwritten Shapes

In this example, use Procrustes analysis to compare two handwritten number threes. Visually and analytically explore the effects of forcing size and reflection changes as follows:

- "Step 1: Load and Display the Original Data" on page 10-16
- "Step 2: Calculate the Best Transformation" on page 10-17
- "Step 3: Examine the Similarity of the Two Shapes" on page 10-18
- "Step 4: Restrict the Form of the Transformations" on page 10-20


## Step 1: Load and Display the Original Data

Input landmark data for two handwritten number threes:

```
A = [11 39;17 42;25 42;25 40;23 36;19 35;30 34;35 29;...
30 20;18 19];
B = [15 31;20 37;30 40;29 35;25 29;29 31;31 31;35 20;...
29 10;25 18];
```

Create $X$ and $Y$ from $A$ and $B$, moving $B$ to the side to make each shape more visible:

```
X = A;
Y = B + repmat([25 0], 10,1);
```

Plot the shapes, using letters to designate the landmark points. Lines in the figure join the points to indicate the drawing path of each shape.

```
plot(X(:,1), X(:,2),'r-', Y(:,1), Y(:,2),'b-');
text(X(:,1), X(:,2),('abcdefghij')')
text(Y(:,1), Y(:,2),('abcdefghij')')
legend('X = Target','Y = Comparison','location','SE')
set(gca,'YLim',[0 55],'XLim',[0 65]);
```



## Step 2: Calculate the Best Transformation

Use Procrustes analysis to find the transformation that minimizes distances between landmark data points.

Call procrustes as follows:

```
[d, Z, tr] = procrustes(X,Y);
```

The outputs of the function are:

- d - A standardized dissimilarity measure.)
- Z - A matrix of the transformed landmarks.
- tr - A structure array of the computed transformation with fields $\mathrm{T}, \mathrm{b}$, and c which correspond to the transformation equation, Equation 10-1.

Visualize the transformed shape, Z, using a dashed blue line:

```
plot(X(:,1), X(:,2),'r-', Y(:,1), Y(:,2),'b-',...
Z(:,1),Z(:,2),'b:');
text(X(:,1), X(:,2),('abcdefghij')')
text(Y(:,1), Y(:,2),('abcdefghij')')
text(Z(:,1), Z(:,2),('abcdefghij')')
legend('X = Target','Y = Comparison',...
'Z = Transformed','location','SW')
set(gca,'YLim',[0 55],'XLim',[0 65]);
```



## Step 3: Examine the Similarity of the Two Shapes

Use two different numerical values to assess the similarity of the target shape and the transformed shape.

Dissimilarity Measure d. The dissimilarity measure d gives a number between 0 and 1 describing the difference between the target shape and the transformed shape. Values near 0 imply more similar shapes, while values near 1 imply dissimilarity. For this example:
$\mathrm{d}=$
0.1502

The small value of $d$ in this case shows that the two shapes are similar.
procrustes calculates $d$ by comparing the sum of squared deviations between the set of points with the sum of squared deviations of the original points from their column means:

```
numerator = sum(sum((X-Z).^2))
numerator =
    166.5321
denominator = sum(sum(bsxfun(@minus,X,mean(X)).^2))
denominator =
    1.1085e+003
ratio = numerator/denominator
ratio =
    0.1502
```

Note The resulting measure $d$ is independent of the scale of the size of the shapes and takes into account only the similarity of landmark data. "Examining the Scaling Measure b" on page 10-19 shows how to examine the size similarity of the shapes.

Examining the Scaling Measure b. The target and comparison threes in the previous figure visually show that the two numbers are of a similar size. The closeness of calculated value of the scaling factor b to 1 supports this observation as well:

```
tr.b
ans =
0.9291
```

The sizes of the target and comparison shapes appear similar. This visual impression is reinforced by the value of $b=0.93$, which implies that the best transformation results in shrinking the comparison shape by a factor . 93 (only 7\%).

## Step 4: Restrict the Form of the Transformations

Explore the effects of manually adjusting the scaling and reflection coefficients.

Fixing the Scaling Factor $\mathbf{b}=\mathbf{1}$. Force $b$ to equal 1 (set 'Scaling' to false) to examine the amount of dissimilarity in size of the target and transformed figures:

```
ds = procrustes(X,Y,'Scaling',false)
ds =
    0.1552
```

In this case, setting 'Scaling' to false increases the calculated value of d only 0.0049 , which further supports the similarity in the size of the two number threes. A larger increase in d would have indicated a greater size discrepancy.

Forcing a Reflection in the Transformation. This example requires only a rotation, not a reflection, to align the shapes. You can show this by observing that the determinant of the matrix T is 1 in this analysis:

```
det(tr.T)
ans =
    1.0000
```

If you need a reflection in the transformation, the determinant of Tis-1. You can force a reflection into the transformation as follows:

```
[dr,Zr,trr] = procrustes(X,Y,'Reflection',true);
dr
dr =
```

0.8130

The d value increases dramatically, indicating that a forced reflection leads to a poor transformation of the landmark points. A plot of the transformed shape shows a similar result:

- The landmark data points are now further away from their target counterparts.
- The transformed three is now an undesirable mirror image of the target three.

```
plot(X(:,1), X(:,2),'r-', Y(:,1), Y(:,2),'b-',...
Zr(:,1),Zr(:,2),'b:');
text(X(:,1), X(:,2),('abcdefghij')')
text(Y(:,1), Y(:,2),('abcdefghij')')
text(Zr(:,1), Zr(:,2),('abcdefghij')')
legend('X = Target','Y = Comparison',...
'Z = Transformed','location','SW')
set(gca,'YLim',[0 55],'XLim',[0 65]);
```



It appears that the shapes might be better matched if you flipped the transformed shape upside down. Flipping the shapes would make the transformation even worse, however, because the landmark data points would be further away from their target counterparts. From this example, it is clear that manually adjusting the scaling and reflection parameters is generally not optimal.

## Feature Selection

In this section...<br>"Introduction to Feature Selection" on page 10-23<br>"Sequential Feature Selection" on page 10-23

## Introduction to Feature Selection

Feature selection reduces the dimensionality of data by selecting only a subset of measured features (predictor variables) to create a model. Selection criteria usually involve the minimization of a specific measure of predictive error for models fit to different subsets. Algorithms search for a subset of predictors that optimally model measured responses, subject to constraints such as required or excluded features and the size of the subset.

Feature selection is preferable to feature transformation when the original units and meaning of features are important and the modeling goal is to identify an influential subset. When categorical features are present, and numerical transformations are inappropriate, feature selection becomes the primary means of dimension reduction.

## Sequential Feature Selection

- "Introduction to Sequential Feature Selection" on page 10-23
- "Example: Sequential Feature Selection" on page 10-24


## Introduction to Sequential Feature Selection

A common method of feature selection is sequential feature selection. This method has two components:

- An objective function, called the criterion, which the method seeks to minimize over all feasible feature subsets. Common criteria are mean squared error (for regression models) and misclassification rate (for classification models).
- A sequential search algorithm, which adds or removes features from a candidate subset while evaluating the criterion. Since an exhaustive
comparison of the criterion value at all $2^{n}$ subsets of an $n$-feature data set is typically infeasible (depending on the size of $n$ and the cost of objective calls), sequential searches move in only one direction, always growing or always shrinking the candidate set.

The method has two variants:

- Sequential forward selection (SFS), in which features are sequentially added to an empty candidate set until the addition of further features does not decrease the criterion.
- Sequential backward selection (SBS), in which features are sequentially removed from a full candidate set until the removal of further features increase the criterion.

Stepwise regression is a sequential feature selection technique designed specifically for least-squares fitting. The functions stepwise and stepwisefit make use of optimizations that are only possible with least-squares criteria. Unlike generalized sequential feature selection, stepwise regression may remove features that have been added or add features that have been removed.

The Statistics Toolbox function sequentialfs carries out sequential feature selection. Input arguments include predictor and response data and a function handle to a file implementing the criterion function. Optional inputs allow you to specify SFS or SBS, required or excluded features, and the size of the feature subset. The function calls cvpartition and crossval to evaluate the criterion at different candidate sets.

## Example: Sequential Feature Selection

For example, consider a data set with 100 observations of 10 predictors. As described in "Example: Generalized Linear Models" on page 9-67, the following generates random data from a logistic model, with a binomial distribution of responses at each set of values for the predictors. Some coefficients are set to zero so that not all of the predictors affect the response:

```
n = 100;
m = 10;
X = rand(n,m);
b = [1 0 0 2 .5 0 0 0.1 0 1];
Xb = X*b';
```

```
p = 1./(1+exp(-Xb));
N = 50;
y = binornd(N,p);
```

The glmfit function fits a logistic model to the data:

```
Y = [y N*ones(size(y))];
[b0,dev0,stats0] = glmfit(X,Y,'binomial');
% Display coefficient estimates and their standard errors:
modelO = [bO statsO.se]
model0 =
    0.3115 0.2596
    0.9614 0.1656
    -0.1100 0.1651
    -0.2165 0.1683
    1.9519 0.1809
    0.5683 0.2018
    -0.0062 0.1740
    0.0651 0.1641
    -0.1034 0.1685
    0.0017 0.1815
    0.7979 0.1806
% Display the deviance of the fit:
dev0
dev0 =
    101.2594
```

This is the full model, using all of the features (and an initial constant term). Sequential feature selection searches for a subset of the features in the full model with comparative predictive power.

First, you must specify a criterion for selecting the features. The following function, which calls glmfit and returns the deviance of the fit (a generalization of the residual sum of squares) is a useful criterion in this case:

```
function dev = critfun(X,Y)
[b,dev] = glmfit(X,Y,'binomial');
```

You should create this function as a file on the MATLAB path.
The function sequentialfs performs feature selection, calling the criterion function via a function handle:

```
maxdev = chi2inv(.95,1);
opt = statset('display','iter',...
    'TolFun',maxdev,...
    'TolTypeFun','abs');
inmodel = sequentialfs(@critfun,X,Y,...
    'cv','none',...
    'nullmodel',true,...
    'options',opt,...
    'direction','forward');
Start forward sequential feature selection:
Initial columns included: none
Columns that can not be included: none
Step 1, used initial columns, criterion value 309.118
Step 2, added column 4, criterion value 180.732
Step 3, added column 1, criterion value 138.862
Step 4, added column 10, criterion value 114.238
Step 5, added column 5, criterion value 103.503
Final columns included: 14510
```

The iterative display shows a decrease in the criterion value as each new feature is added to the model. The final result is a reduced model with only four of the original ten features: columns $1,4,5$, and 10 of $X$. These features are indicated in the logical vector inmodel returned by sequentialfs.

The deviance of the reduced model is higher than for the full model, but the addition of any other single feature would not decrease the criterion by more than the absolute tolerance, maxdev, set in the options structure. Adding a feature with no effect reduces the deviance by an amount that has a chi-square distribution with one degree of freedom. Adding a significant feature results in a larger change. By setting maxdev to chi2inv(.95,1), you instruct sequentialfs to continue adding features so long as the change in deviance is more than would be expected by random chance.

The reduced model (also with an initial constant term) is:

```
[b,dev,stats] = glmfit(X(:,inmodel),Y,'binomial');
% Display coefficient estimates and their standard errors:
model = [b stats.se]
model =
    0.0784 0.1642
    1.0040 0.1592
    1.9459 0.1789
    0.6134 0.1872
    0.8245 0.1730
```


## Feature Transformation

In this section...<br>"Introduction to Feature Transformation" on page 10-28<br>"Nonnegative Matrix Factorization" on page 10-28<br>"Principal Component Analysis (PCA)" on page 10-31<br>"Factor Analysis" on page 10-45

## Introduction to Feature Transformation

Feature transformation is a group of methods that create new features (predictor variables). The methods are useful for dimension reduction when the transformed features have a descriptive power that is more easily ordered than the original features. In this case, less descriptive features can be dropped from consideration when building models.

Feature transformation methods are contrasted with the methods presented in "Feature Selection" on page 10-23, where dimension reduction is achieved by computing an optimal subset of predictive features measured in the original data.

The methods presented in this section share some common methodology. Their goals, however, are essentially different:

- Nonnegative matrix factorization is used when model terms must represent nonnegative quantities, such as physical quantities.
- Principal component analysis is used to summarize data in fewer dimensions, for example, to visualize it.
- Factor analysis is used to build explanatory models of data correlations.


## Nonnegative Matrix Factorization

- "Introduction to Nonnegative Matrix Factorization" on page 10-29
- "Example: Nonnegative Matrix Factorization" on page 10-29


## Introduction to Nonnegative Matrix Factorization

Nonnegative matrix factorization (NMF) is a dimension-reduction technique based on a low-rank approximation of the feature space. Besides providing a reduction in the number of features, NMF guarantees that the features are nonnegative, producing additive models that respect, for example, the nonnegativity of physical quantities.

Given a nonnegative $m$-by- $n$ matrix $X$ and a positive integer $k<\min (m, n)$, NMF finds nonnegative $m$-by- $k$ and $k$-by- $n$ matrices $W$ and $H$, respectively, that minimize the norm of the difference $X-W H$. $W$ and $H$ are thus approximate nonnegative factors of $X$.

The $k$ columns of $W$ represent transformations of the variables in $X$; the $k$ rows of $H$ represent the coefficients of the linear combinations of the original $n$ variables in $X$ that produce the transformed variables in $W$. Since $k$ is generally smaller than the rank of $X$, the product $W H$ provides a compressed approximation of the data in $X$. A range of possible values for $k$ is often suggested by the modeling context.

The Statistics Toolbox function nnmf carries out nonnegative matrix factorization. nnmf uses one of two iterative algorithms that begin with random initial values for $W$ and $H$. Because the norm of the residual $X$ - WH may have local minima, repeated calls to nnmf may yield different factorizations. Sometimes the algorithm converges to a solution of lower rank than $k$, which may indicate that the result is not optimal.

## Example: Nonnegative Matrix Factorization

For example, consider the five predictors of biochemical oxygen demand in the data set moore.mat:

```
load moore
X = moore(:,1:5);
```

The following uses nnmf to compute a rank-two approximation of $X$ with a multiplicative update algorithm that begins from five random initial values for W and H :

```
opt = statset('MaxIter',10,'Display','final');
[WO,HO] = nnmf(X,2,'replicates',5,...
    'options',opt,...
```

|  |  | 'algorithm', 'mult'); <br> rep <br> iteration resid | \|delta x| |
| ---: | :---: | :---: | :---: |
| 1 | 10 | 358.296 | 0.00190554 |
| 2 | 10 | 78.3556 | 0.000351747 |
| 3 | 10 | 230.962 | 0.0172839 |
| 4 | 10 | 326.347 | 0.00739552 |
| 5 | 10 | 361.547 | 0.00705539 |
| Final root mean square residual $=78.3556$ |  |  |  |

The 'mult' algorithm is sensitive to initial values, which makes it a good choice when using 'replicates' to find W and H from multiple random starting values.

Now perform the factorization using an alternating least-squares algorithm, which converges faster and more consistently. Run 100 times more iterations, beginning from the initial WO and HO identified above:

```
opt = statset('Maxiter',1000,'Display','final');
[W,H] = nnmf(X,2,'wO',WO,'hO',HO,...
    'options',opt,...
    'algorithm','als');
    rep iteration rms resid |delta x|
    1 3 77.5315 3.52673e-005
Final root mean square residual = 77.5315
```

The two columns of W are the transformed predictors. The two rows of H give the relative contributions of each of the five predictors in $X$ to the predictors in w :

```
H
H =
\begin{tabular}{lllll}
0.0835 & 0.0190 & 0.1782 & 0.0072 & 0.9802 \\
0.0558 & 0.0250 & 0.9969 & 0.0085 & 0.0497
\end{tabular}
```

The fifth predictor in $X$ (weight 0.9802 ) strongly influences the first predictor in W. The third predictor in $X$ (weight 0.9969 ) strongly influences the second predictor in $W$.

Visualize the relative contributions of the predictors in X with a biplot, showing the data and original variables in the column space of w :

```
    biplot(H','scores',W,'varlabels',{'','','v3','','v5'});
    axis([0 1.1 0 1.1])
    xlabel('Column 1')
    ylabel('Column 2')
```



## Principal Component Analysis (PCA)

- "Introduction to Principal Component Analysis (PCA)" on page 10-31
- "Example: Principal Component Analysis" on page 10-33


## Introduction to Principal Component Analysis (PCA)

One of the difficulties inherent in multivariate statistics is the problem of visualizing data that has many variables. The MATLAB function plot displays a graph of the relationship between two variables. The plot3 and surf commands display different three-dimensional views. But when
there are more than three variables, it is more difficult to visualize their relationships.

Fortunately, in data sets with many variables, groups of variables often move together. One reason for this is that more than one variable might be measuring the same driving principle governing the behavior of the system. In many systems there are only a few such driving forces. But an abundance of instrumentation enables you to measure dozens of system variables. When this happens, you can take advantage of this redundancy of information. You can simplify the problem by replacing a group of variables with a single new variable.

Principal component analysis is a quantitatively rigorous method for achieving this simplification. The method generates a new set of variables, called principal components. Each principal component is a linear combination of the original variables. All the principal components are orthogonal to each other, so there is no redundant information. The principal components as a whole form an orthogonal basis for the space of the data.

There are an infinite number of ways to construct an orthogonal basis for several columns of data. What is so special about the principal component basis?

The first principal component is a single axis in space. When you project each observation on that axis, the resulting values form a new variable. And the variance of this variable is the maximum among all possible choices of the first axis.

The second principal component is another axis in space, perpendicular to the first. Projecting the observations on this axis generates another new variable. The variance of this variable is the maximum among all possible choices of this second axis.

The full set of principal components is as large as the original set of variables. But it is commonplace for the sum of the variances of the first few principal components to exceed $80 \%$ of the total variance of the original data. By examining plots of these few new variables, researchers often develop a deeper understanding of the driving forces that generated the original data.

You can use the function princomp to find the principal components. To use princomp, you need to have the actual measured data you want to analyze. However, if you lack the actual data, but have the sample covariance or correlation matrix for the data, you can still use the function pcacov to perform a principal components analysis. See the reference page for pcacov for a description of its inputs and outputs.

## Example: Principal Component Analysis

- "Computing Components" on page 10-33
- "Component Coefficients" on page 10-36
- "Component Scores" on page 10-36
- "Component Variances" on page 10-40
- "Hotelling's T2" on page 10-42
- "Visualizing the Results" on page 10-42

Computing Components. Consider a sample application that uses nine different indices of the quality of life in 329 U.S. cities. These are climate, housing, health, crime, transportation, education, arts, recreation, and economics. For each index, higher is better. For example, a higher index for crime means a lower crime rate.

Start by loading the data in cities.mat.

| load cities |  |  |  |
| :--- | ---: | ---: | :--- |
| whos | Size | Bytes | Class |
| Name | $9 \times 14$ | 252 | char array |
| categories | $329 \times 43$ | 28294 | char array |
| names | $329 \times 9$ | 23688 | double array |

The whos command generates a table of information about all the variables in the workspace.

The cities data set contains three variables:

- categories, a string matrix containing the names of the indices
- names, a string matrix containing the 329 city names
- ratings, the data matrix with 329 rows and 9 columns

The categories variable has the following values:

```
categories
categories =
    climate
    housing
    health
    crime
    transportation
    education
    arts
    recreation
    economics
```

The first five rows of names are

```
first5 = names(1:5,:)
first5 =
    Abilene, TX
    Akron, OH
    Albany, GA
    Albany-Troy, NY
    Albuquerque, NM
```

To get a quick impression of the ratings data, make a box plot.

```
boxplot(ratings,'orientation','horizontal','labels',categories)
```

This command generates the plot below. Note that there is substantially more variability in the ratings of the arts and housing than in the ratings of crime and climate.


Ordinarily you might also graph pairs of the original variables, but there are 36two-variable plots. Perhaps principal components analysis can reduce the number of variables you need to consider.

Sometimes it makes sense to compute principal components for raw data. This is appropriate when all the variables are in the same units. Standardizing the data is often preferable when the variables are in different units or when the variance of the different columns is substantial (as in this case).

You can standardize the data by dividing each column by its standard deviation.

```
stdr = std(ratings);
sr = ratings./repmat(stdr,329,1);
```

Now you are ready to find the principal components.

```
[coefs,scores,variances,t2] = princomp(sr);
```

The following sections explain the four outputs from princomp.
Component Coefficients. The first output of the princomp function, coefs, contains the coefficients of the linear combinations of the original variables that generate the principal components. The coefficients are also known as loadings.

The first three principal component coefficient vectors are:

| 0.2064 | 0.2178 | -0.6900 |
| :---: | :---: | :---: |
| 0.3565 | 0.2506 | -0.2082 |
| 0.4602 | -0.2995 | -0.0073 |
| 0.2813 | 0.3553 | 0.1851 |
| 0.3512 | -0.1796 | 0.1464 |
| 0.2753 | -0.4834 | 0.2297 |
| 0.4631 | -0.1948 | -0.0265 |
| 0.3279 | 0.3845 | -0.0509 |
| 0.1354 | 0.4713 | 0.6073 |

The largest coefficients in the first column (first principal component) are the third and seventh elements, corresponding to the variables health and arts. All the coefficients of the first principal component have the same sign, making it a weighted average of all the original variables.

The principal components are unit length and orthogonal:

```
I = c3'*c3
I =
    \(1.0000-0.0000-0.0000\)
    \(-0.0000 \quad 1.0000 \quad-0.0000\)
    \(-0.0000 \quad-0.0000 \quad 1.0000\)
```

Component Scores. The second output, scores, contains the coordinates of the original data in the new coordinate system defined by the principal components. This output is the same size as the input data matrix.

A plot of the first two columns of scores shows the ratings data projected onto the first two principal components. princomp computes the scores to have mean zero.

```
plot(scores(:,1),scores(:,2),'+')
xlabel('1st Principal Component')
ylabel('2nd Principal Component')
```



Note the outlying points in the right half of the plot.
While it is possible to create a three-dimensional plot using three columns of scores, the examples in this section create two-dimensional plots, which are easier to describe.

The function gname is useful for graphically identifying a few points in a plot like this. You can call gname with a string matrix containing as many case
labels as points in the plot. The string matrix names works for labeling points with the city names.

```
gname(names)
```

Move your cursor over the plot and click once near each point in the right half. As you click each point, it is labeled with the proper row from the names string matrix. Here is the plot after a few clicks:


When you are finished labeling points, press the Return key.
The labeled cities are some of the biggest population centers in the United States. They are definitely different from the remainder of the data, so perhaps they should be considered separately. To remove the labeled cities from the data, first identify their corresponding row numbers as follows:

1 Close the plot window.
2 Redraw the plot by entering

```
plot(scores(:,1),scores(:,2),'+')
xlabel('1st Principal Component');
ylabel('2nd Principal Component');
```

3 Enter gname without any arguments.
4 Click near the points you labeled in the preceding figure. This labels the points by their row numbers, as shown in the following figure.


Then you can create an index variable containing the row numbers of all the metropolitan areas you choose.

```
metro = [43 65 179 213 234 270 314];
names(metro,:)
ans =
    Boston, MA
    Chicago, IL
    Los Angeles, Long Beach, CA
    New York, NY
    Philadelphia, PA-NJ
    San Francisco, CA
    Washington, DC-MD-VA
```

To remove these rows from the ratings matrix, enter the following.

```
rsubset = ratings;
nsubset = names;
nsubset(metro,:) = [];
rsubset(metro,:) = [];
size(rsubset)
ans =
    322 9
```

Component Variances. The third output, variances, is a vector containing the variance explained by the corresponding principal component. Each column of scores has a sample variance equal to the corresponding element of variances.

```
variances
variances =
    3.4083
    1.2140
    1.1415
    0.9209
    0.7533
    0.6306
    0.4930
    0.3180
    0.1204
```

You can easily calculate the percent of the total variability explained by each principal component.

```
percent_explained = 100*variances/sum(variances)
percent explained =
    37.8699
    13.4886
    12.6831
    10.2324
        8.3698
        7.0062
        5.4783
        3.5338
        1.3378
```

Use the pareto function to make a scree plot of the percent variability explained by each principal component.

```
pareto(percent_explained)
xlabel('Principal Component')
ylabel('Variance Explained (%)')
```



The preceding figure shows that the only clear break in the amount of variance accounted for by each component is between the first and second components. However, that component by itself explains less than $40 \%$ of the variance, so more components are probably needed. You can see that the first three principal components explain roughly two-thirds of the total variability in the standardized ratings, so that might be a reasonable way to reduce the dimensions in order to visualize the data.

Hotelling's T2. The last output of the princomp function, t2, is Hotelling'sT ${ }^{2}$, a statistical measure of the multivariate distance of each observation from the center of the data set. This is an analytical way to find the most extreme points in the data.

```
[st2, index] = sort(t2,'descend'); % Sort in descending order.
extreme = index(1)
extreme =
    213
names(extreme,:)
ans =
    New York, NY
```

It is not surprising that the ratings for New York are the furthest from the average U.S. town.

Visualizing the Results. Use the biplot function to help visualize both the principal component coefficients for each variable and the principal component scores for each observation in a single plot. For example, the following command plots the results from the principal components analysis on the cities and labels each of the variables.

```
biplot(coefs(:,1:2), 'scores',scores(:,1:2),...
'varlabels',categories);
axis([-. 26 1- -.51 .51]);
```



Each of the nine variables is represented in this plot by a vector, and the direction and length of the vector indicates how each variable contributes to the two principal components in the plot. For example, you have seen that the first principal component, represented in this biplot by the horizontal axis, has positive coefficients for all nine variables. That corresponds to the nine vectors directed into the right half of the plot. You have also seen that the second principal component, represented by the vertical axis, has positive coefficients for the variables education, health, arts, and transportation, and negative coefficients for the remaining five variables. That corresponds to vectors directed into the top and bottom halves of the plot, respectively. This indicates that this component distinguishes between cities that have high values for the first set of variables and low for the second, and cities that have the opposite.

The variable labels in this figure are somewhat crowded. You could either leave out the VarLabels parameter when making the plot, or simply select and drag some of the labels to better positions using the Edit Plot tool from the figure window toolbar.

Each of the 329 observations is represented in this plot by a point, and their locations indicate the score of each observation for the two principal components in the plot. For example, points near the left edge of this plot have the lowest scores for the first principal component. The points are scaled to fit within the unit square, so only their relative locations may be determined from the plot.

You can use the Data Cursor, in the Tools menu in the figure window, to identify the items in this plot. By clicking on a variable (vector), you can read off that variable's coefficients for each principal component. By clicking on an observation (point), you can read off that observation's scores for each principal component.

You can also make a biplot in three dimensions. This can be useful if the first two principal coordinates do not explain enough of the variance in your data. Selecting Rotate 3D in the Tools menu enables you to rotate the figure to see it from different angles.

```
biplot(coefs(:,1:3), 'scores',scores(:,1:3),...
'obslabels',names);
axis([-.26 1 -. 51 . 51 -.61 . 81]);
view([30 40]);
```



## Factor Analysis

- "Introduction to Factor Analysis" on page 10-45
- "Example: Factor Analysis" on page 10-46


## Introduction to Factor Analysis

Multivariate data often includes a large number of measured variables, and sometimes those variables overlap, in the sense that groups of them might be dependent. For example, in a decathlon, each athlete competes in 10 events, but several of them can be thought of as speed events, while others can be thought of as strength events, etc. Thus, you can think of a competitor's 10 event scores as largely dependent on a smaller set of three or four types of athletic ability.

Factor analysis is a way to fit a model to multivariate data to estimate just this sort of interdependence. In a factor analysis model, the measured variables depend on a smaller number of unobserved (latent) factors. Because each factor might affect several variables in common, they are known as common factors. Each variable is assumed to be dependent on a linear combination of the common factors, and the coefficients are known as loadings. Each measured variable also includes a component due to independent random variability, known as specific variance because it is specific to one variable.

Specifically, factor analysis assumes that the covariance matrix of your data is of the form

$$
\sum_{x}=\Lambda \Lambda^{\mathrm{T}}+\Psi
$$

where $\Lambda$ is the matrix of loadings, and the elements of the diagonal matrix $\Psi$ are the specific variances. The function factoran fits the Factor Analysis model using maximum likelihood.

## Example: Factor Analysis

- "Factor Loadings" on page 10-46
- "Factor Rotation" on page 10-48
- "Factor Scores" on page 10-50
- "Visualizing the Results" on page 10-52

Factor Loadings. Over the course of 100 weeks, the percent change in stock prices for ten companies has been recorded. Of the ten companies, the first four can be classified as primarily technology, the next three as financial, and the last three as retail. It seems reasonable that the stock prices for companies that are in the same sector might vary together as economic conditions change. Factor Analysis can provide quantitative evidence that companies within each sector do experience similar week-to-week changes in stock price.

In this example, you first load the data, and then call factoran, specifying a model fit with three common factors. By default, factoran computes rotated estimates of the loadings to try and make their interpretation simpler. But in this example, you specify an unrotated solution.

## load stockreturns

```
[Loadings,specificVar,T,stats] = ...
    factoran(stocks,3,'rotate','none');
```

The first two factoran return arguments are the estimated loadings and the estimated specific variances. Each row of the loadings matrix represents one of the ten stocks, and each column corresponds to a common factor. With unrotated estimates, interpretation of the factors in this fit is difficult because most of the stocks contain fairly large coefficients for two or more factors.

```
Loadings
Loadings =
    0.8885 0.2367 -0.2354
    0.7126 0.3862 0.0034
    0.3351 0.2784 -0.0211
    0.3088 0.1113 -0.1905
    0.6277 -0.6643 0.1478
    0.4726 -0.6383 0.0133
    0.1133 -0.5416 0.0322
    0.6403 0.1669 0.4960
    0.2363 0.5293 0.5770
    0.1105 0.1680 0.5524
```

Note "Factor Rotation" on page 10-48 helps to simplify the structure in the Loadings matrix, to make it easier to assign meaningful interpretations to the factors.

From the estimated specific variances, you can see that the model indicates that a particular stock price varies quite a lot beyond the variation due to the common factors.

```
specificVar
specificVar =
    0.0991
    0.3431
    0.8097
    0.8559
    0.1429
```

0.3691
0.6928
0.3162
0.3311

$$
0.6544
$$

A specific variance of 1 would indicate that there is no common factor component in that variable, while a specific variance of 0 would indicate that the variable is entirely determined by common factors. These data seem to fall somewhere in between.

The $p$ value returned in the stats structure fails to reject the null hypothesis of three common factors, suggesting that this model provides a satisfactory explanation of the covariation in these data.

```
stats.p
ans =
    0.8144
```

To determine whether fewer than three factors can provide an acceptable fit, you can try a model with two common factors. The $p$ value for this second fit is highly significant, and rejects the hypothesis of two factors, indicating that the simpler model is not sufficient to explain the pattern in these data.

```
[Loadings2,specificVar2,T2,stats2] = ...
    factoran(stocks, 2,'rotate','none');
stats2.p
ans =
    3.5610e-006
```

Factor Rotation. As the results illustrate, the estimated loadings from an unrotated factor analysis fit can have a complicated structure. The goal of factor rotation is to find a parameterization in which each variable has only a small number of large loadings. That is, each variable is affected by a small number of factors, preferably only one. This can often make it easier to interpret what the factors represent.

If you think of each row of the loadings matrix as coordinates of a point in $M$-dimensional space, then each factor corresponds to a coordinate axis. Factor rotation is equivalent to rotating those axes and computing new
loadings in the rotated coordinate system. There are various ways to do this. Some methods leave the axes orthogonal, while others are oblique methods that change the angles between them. For this example, you can rotate the estimated loadings by using the promax criterion, a common oblique method.

```
[LoadingsPM,specVarPM] = factoran(stocks,3,'rotate','promax');
LoadingsPM
LoadingsPM =
\begin{tabular}{rrr}
0.9452 & 0.1214 & -0.0617 \\
0.7064 & -0.0178 & 0.2058 \\
0.3885 & -0.0994 & 0.0975 \\
0.4162 & -0.0148 & -0.1298 \\
0.1021 & 0.9019 & 0.0768 \\
0.0873 & 0.7709 & -0.0821 \\
-0.1616 & 0.5320 & -0.0888 \\
0.2169 & 0.2844 & 0.6635 \\
0.0016 & -0.1881 & 0.7849 \\
-0.2289 & 0.0636 & 0.6475
\end{tabular}
```

Promax rotation creates a simpler structure in the loadings, one in which most of the stocks have a large loading on only one factor. To see this structure more clearly, you can use the biplot function to plot each stock using its factor loadings as coordinates.

```
biplot(LoadingsPM,'varlabels',num2str((1:10)'));
axis square
view(155,27);
```



This plot shows that promax has rotated the factor loadings to a simpler structure. Each stock depends primarily on only one factor, and it is possible to describe each factor in terms of the stocks that it affects. Based on which companies are near which axes, you could reasonably conclude that the first factor axis represents the financial sector, the second retail, and the third technology. The original conjecture, that stocks vary primarily within sector, is apparently supported by the data.

Factor Scores. Sometimes, it is useful to be able to classify an observation based on its factor scores. For example, if you accepted the three-factor model and the interpretation of the rotated factors, you might want to categorize each week in terms of how favorable it was for each of the three stock sectors, based on the data from the 10 observed stocks.

Because the data in this example are the raw stock price changes, and not just their correlation matrix, you can have factoran return estimates of the
value of each of the three rotated common factors for each week. You can then plot the estimated scores to see how the different stock sectors were affected during each week.

```
[LoadingsPM,specVarPM,TPM,stats,F] = ...
    factoran(stocks, 3,'rotate','promax');
```

```
plot3(F(:,1),F(:,2),F(:,3),'b.')
line([-4 4 NaN O O NaN O 0], [0 O NaN -4 4 NaN O 0],...
    [O O NaN O O NaN -4 4], 'Color','black')
xlabel('Financial Sector')
ylabel('Retail Sector')
zlabel('Technology Sector')
grid on
axis square
view(-22.5, 8)
```



Retail Sector
Financial Sector

Oblique rotation often creates factors that are correlated. This plot shows some evidence of correlation between the first and third factors, and you can investigate further by computing the estimated factor correlation matrix.

```
inv(TPM'*TPM)
ans =
\begin{tabular}{rrr}
1.0000 & 0.1559 & 0.4082 \\
0.1559 & 1.0000 & -0.0559 \\
0.4082 & -0.0559 & 1.0000
\end{tabular}
```

Visualizing the Results. You can use the biplot function to help visualize both the factor loadings for each variable and the factor scores for each observation in a single plot. For example, the following command plots the results from the factor analysis on the stock data and labels each of the 10 stocks.

```
biplot(LoadingsPM,'scores',F,'varlabels',num2str((1:10)'))
xlabel('Financial Sector')
ylabel('Retail Sector')
zlabel('Technology Sector')
axis square
view(155,27)
```



In this case, the factor analysis includes three factors, and so the biplot is three-dimensional. Each of the 10 stocks is represented in this plot by a vector, and the direction and length of the vector indicates how each stock depends on the underlying factors. For example, you have seen that after promax rotation, the first four stocks have positive loadings on the first factor, and unimportant loadings on the other two factors. That first factor, interpreted as a financial sector effect, is represented in this biplot as one of the horizontal axes. The dependence of those four stocks on that factor corresponds to the four vectors directed approximately along that axis. Similarly, the dependence of stocks 5,6 , and 7 primarily on the second factor, interpreted as a retail sector effect, is represented by vectors directed approximately along that axis.

Each of the 100 observations is represented in this plot by a point, and their locations indicate the score of each observation for the three factors. For example, points near the top of this plot have the highest scores for the
technology sector factor. The points are scaled to fit within the unit square, so only their relative locations can be determined from the plot.

You can use the Data Cursor tool from the Tools menu in the figure window to identify the items in this plot. By clicking a stock (vector), you can read off that stock's loadings for each factor. By clicking an observation (point), you can read off that observation's scores for each factor.

## Cluster Analysis

- "Introduction to Cluster Analysis" on page 11-2
- "Hierarchical Clustering" on page 11-3
- "K-Means Clustering" on page 11-21
- "Gaussian Mixture Models" on page 11-28


## Introduction to Cluster Analysis

Cluster analysis, also called segmentation analysis or taxonomy analysis, creates groups, or clusters, of data. Clusters are formed in such a way that objects in the same cluster are very similar and objects in different clusters are very distinct. Measures of similarity depend on the application.
"Hierarchical Clustering" on page 11-3 groups data over a variety of scales by creating a cluster tree or dendrogram. The tree is not a single set of clusters, but rather a multilevel hierarchy, where clusters at one level are joined as clusters at the next level. This allows you to decide the level or scale of clustering that is most appropriate for your application. The Statistics Toolbox function clusterdata performs all of the necessary steps for you. It incorporates the pdist, linkage, and cluster functions, which may be used separately for more detailed analysis. The dendrogram function plots the cluster tree.
"K-Means Clustering" on page 11-21 is a partitioning method. The function kmeans partitions data into $k$ mutually exclusive clusters, and returns the index of the cluster to which it has assigned each observation. Unlike hierarchical clustering, $k$-means clustering operates on actual observations (rather than the larger set of dissimilarity measures), and creates a single level of clusters. The distinctions mean that $k$-means clustering is often more suitable than hierarchical clustering for large amounts of data.
"Gaussian Mixture Models" on page 11-28 form clusters by representing the probability density function of observed variables as a mixture of multivariate normal densities. Mixture models of the gmdistribution class use an expectation maximization (EM) algorithm to fit data, which assigns posterior probabilities to each component density with respect to each observation. Clusters are assigned by selecting the component that maximizes the posterior probability. Clustering using Gaussian mixture models is sometimes considered a soft clustering method. The posterior probabilities for each point indicate that each data point has some probability of belonging to each cluster. Like $k$-means clustering, Gaussian mixture modeling uses an iterative algorithm that converges to a local optimum. Gaussian mixture modeling may be more appropriate than $k$-means clustering when clusters have different sizes and correlation within them.

## Hierarchical Clustering

In this section...<br>"Introduction to Hierarchical Clustering" on page 11-3<br>"Algorithm Description" on page 11-3<br>"Similarity Measures" on page 11-4<br>"Linkages" on page 11-6<br>"Dendrograms" on page 11-8<br>"Verifying the Cluster Tree" on page 11-10<br>"Creating Clusters" on page 11-16

## Introduction to Hierarchical Clustering

Hierarchical clustering groups data over a variety of scales by creating a cluster tree or dendrogram. The tree is not a single set of clusters, but rather a multilevel hierarchy, where clusters at one level are joined as clusters at the next level. This allows you to decide the level or scale of clustering that is most appropriate for your application. The Statistics Toolbox function clusterdata supports agglomerative clustering and performs all of the necessary steps for you. It incorporates the pdist, linkage, and cluster functions, which you can use separately for more detailed analysis. The dendrogram function plots the cluster tree.

## Algorithm Description

To perform agglomerative hierarchical cluster analysis on a data set using Statistics Toolbox functions, follow this procedure:

1 Find the similarity or dissimilarity between every pair of objects in the data set. In this step, you calculate the distance between objects using the pdist function. The pdist function supports many different ways to compute this measurement. See "Similarity Measures" on page $11-4$ for more information.

2 Group the objects into a binary, hierarchical cluster tree. In this step, you link pairs of objects that are in close proximity using the linkage
function. The linkage function uses the distance information generated in step 1 to determine the proximity of objects to each other. As objects are paired into binary clusters, the newly formed clusters are grouped into larger clusters until a hierarchical tree is formed. See "Linkages" on page 11-6 for more information.

3 Determine where to cut the hierarchical tree into clusters. In this step, you use the cluster function to prune branches off the bottom of the hierarchical tree, and assign all the objects below each cut to a single cluster. This creates a partition of the data. The cluster function can create these clusters by detecting natural groupings in the hierarchical tree or by cutting off the hierarchical tree at an arbitrary point.

The following sections provide more information about each of these steps.

Note The Statistics Toolbox function clusterdata performs all of the necessary steps for you. You do not need to execute the pdist, linkage, or cluster functions separately.

## Similarity Measures

You use the pdist function to calculate the distance between every pair of objects in a data set. For a data set made up of $m$ objects, there are $m^{*}(m-$ $1) / 2$ pairs in the data set. The result of this computation is commonly known as a distance or dissimilarity matrix.

There are many ways to calculate this distance information. By default, the pdist function calculates the Euclidean distance between objects; however, you can specify one of several other options. See pdist for more information.

Note You can optionally normalize the values in the data set before calculating the distance information. In a real world data set, variables can be measured against different scales. For example, one variable can measure Intelligence Quotient (IQ) test scores and another variable can measure head circumference. These discrepancies can distort the proximity calculations. Using the zscore function, you can convert all the values in the data set to use the same proportional scale. See zscore for more information.

For example, consider a data set, X , made up of five objects where each object is a set of $x, y$ coordinates.

- Object 1: 1, 2
- Object 2: 2.5, 4.5
- Object 3: 2, 2
- Object 4: 4, 1.5
- Object 5: 4, 2.5

You can define this data set as a matrix

$$
X=[12 ; 2.54 .5 ; 22 ; 41.5 ; 42.5]
$$

and pass it to pdist. The pdist function calculates the distance between object 1 and object 2 , object 1 and object 3 , and so on until the distances between all the pairs have been calculated. The following figure plots these objects in a graph. The Euclidean distance between object 2 and object 3 is shown to illustrate one interpretation of distance.


## Distance Information

The pdist function returns this distance information in a vector, Y , where each element contains the distance between a pair of objects.

```
Y = pdist(X)
Y =
    Columns 1 through 5
        2.9155 1.0000 3.0414 3.0414 2.5495
    Columns 6 through }1
        3.3541 2.5000 2.0616 2.0616 1.0000
```

To make it easier to see the relationship between the distance information generated by pdist and the objects in the original data set, you can reformat the distance vector into a matrix using the squareform function. In this matrix, element $i, j$ corresponds to the distance between object $i$ and object $j$ in the original data set. In the following example, element 1,1 represents the distance between object 1 and itself (which is zero). Element 1,2 represents the distance between object 1 and object 2 , and so on.

| squareform $(Y)$ <br> ans $=$ | 0 | 2.9155 | 1.0000 | 3.0414 |
| ---: | ---: | ---: | ---: | ---: |
| 2.9155 | 0 | 2.5495 | 3.3541 | 2.0414 |
| 1.0000 | 2.5495 | 0 | 2.0616 | 2.0616 |
| 3.0414 | 3.3541 | 2.0616 | 0 | 1.0000 |
| 3.0414 | 2.5000 | 2.0616 | 1.0000 | 0 |

## Linkages

Once the proximity between objects in the data set has been computed, you can determine how objects in the data set should be grouped into clusters, using the linkage function. The linkage function takes the distance information generated by pdist and links pairs of objects that are close together into binary clusters (clusters made up of two objects). The linkage function then links these newly formed clusters to each other and to other objects to create bigger clusters until all the objects in the original data set are linked together in a hierarchical tree.

For example, given the distance vector $Y$ generated by pdist from the sample data set of $x$ - and $y$-coordinates, the linkage function generates a hierarchical cluster tree, returning the linkage information in a matrix, $z$.

```
Z = linkage(Y)
Z =
    4.0000 5.0000 1.0000
```

| 1.0000 | 3.0000 | 1.0000 |
| :--- | :--- | :--- |
| 6.0000 | 7.0000 | 2.0616 |
| 2.0000 | 8.0000 | 2.5000 |

In this output, each row identifies a link between objects or clusters. The first two columns identify the objects that have been linked. The third column contains the distance between these objects. For the sample data set of $x$ and $y$-coordinates, the linkage function begins by grouping objects 4 and 5, which have the closest proximity (distance value $=1.0000$ ). The linkage function continues by grouping objects 1 and 3 , which also have a distance value of 1.0000 .

The third row indicates that the linkage function grouped objects 6 and 7. If the original sample data set contained only five objects, what are objects 6 and 7 ? Object 6 is the newly formed binary cluster created by the grouping of objects 4 and 5 . When the linkage function groups two objects into a new cluster, it must assign the cluster a unique index value, starting with the value $m+1$, where $m$ is the number of objects in the original data set. (Values 1 through $m$ are already used by the original data set.) Similarly, object 7 is the cluster formed by grouping objects 1 and 3 .
linkage uses distances to determine the order in which it clusters objects. The distance vector $Y$ contains the distances between the original objects 1 through 5. But linkage must also be able to determine distances involving clusters that it creates, such as objects 6 and 7 . By default, linkage uses a method known as single linkage. However, there are a number of different methods available. See the linkage reference page for more information.

As the final cluster, the linkage function grouped object 8, the newly formed cluster made up of objects 6 and 7 , with object 2 from the original data set. The following figure graphically illustrates the way linkage groups the objects into a hierarchy of clusters.


## Dendrograms

The hierarchical, binary cluster tree created by the linkage function is most easily understood when viewed graphically. The Statistics Toolbox function dendrogram plots the tree, as follows:
dendrogram(Z)


In the figure, the numbers along the horizontal axis represent the indices of the objects in the original data set. The links between objects are represented as upside-down U -shaped lines. The height of the U indicates the distance between the objects. For example, the link representing the cluster containing objects 1 and 3 has a height of 1 . The link representing the cluster that groups object 2 together with objects 1, 3, 4, and 5, (which are already clustered as object 8 ) has a height of 2.5 . The height represents the distance linkage computes between objects 2 and 8 . For more information about creating a dendrogram diagram, see the dendrogram reference page.

## Verifying the Cluster Tree

After linking the objects in a data set into a hierarchical cluster tree, you might want to verify that the distances (that is, heights) in the tree reflect the original distances accurately. In addition, you might want to investigate natural divisions that exist among links between objects. Statistics Toolbox functions are available for both of these tasks, as described in the following sections:

- "Verifying Dissimilarity" on page 11-10
- "Verifying Consistency" on page 11-11


## Verifying Dissimilarity

In a hierarchical cluster tree, any two objects in the original data set are eventually linked together at some level. The height of the link represents the distance between the two clusters that contain those two objects. This height is known as the cophenetic distance between the two objects. One way to measure how well the cluster tree generated by the linkage function reflects your data is to compare the cophenetic distances with the original distance data generated by the pdist function. If the clustering is valid, the linking of objects in the cluster tree should have a strong correlation with the distances between objects in the distance vector. The cophenet function compares these two sets of values and computes their correlation, returning a value called the cophenetic correlation coefficient. The closer the value of the cophenetic correlation coefficient is to 1 , the more accurately the clustering solution reflects your data.

You can use the cophenetic correlation coefficient to compare the results of clustering the same data set using different distance calculation methods or clustering algorithms. For example, you can use the cophenet function to evaluate the clusters created for the sample data set

```
\(c=\operatorname{cophenet}(Z, Y)\)
c \(=\)
    0.8615
```

where $Z$ is the matrix output by the linkage function and $Y$ is the distance vector output by the pdist function.

Execute pdist again on the same data set, this time specifying the city block metric. After running the linkage function on this new pdist output using the average linkage method, call cophenet to evaluate the clustering solution.

```
Y = pdist(X,'cityblock');
Z = linkage(Y,'average');
c = cophenet(Z,Y)
c =
    0.9047
```

The cophenetic correlation coefficient shows that using a different distance and linkage method creates a tree that represents the original distances slightly better.

## Verifying Consistency

One way to determine the natural cluster divisions in a data set is to compare the height of each link in a cluster tree with the heights of neighboring links below it in the tree.

A link that is approximately the same height as the links below it indicates that there are no distinct divisions between the objects joined at this level of the hierarchy. These links are said to exhibit a high level of consistency, because the distance between the objects being joined is approximately the same as the distances between the objects they contain.

On the other hand, a link whose height differs noticeably from the height of the links below it indicates that the objects joined at this level in the cluster tree are much farther apart from each other than their components were when they were joined. This link is said to be inconsistent with the links below it.

In cluster analysis, inconsistent links can indicate the border of a natural division in a data set. The cluster function uses a quantitative measure of inconsistency to determine where to partition your data set into clusters.

The following dendrogram illustrates inconsistent links. Note how the objects in the dendrogram fall into two groups that are connected by links at a much higher level in the tree. These links are inconsistent when compared with the links below them in the hierarchy.

These links show inconsistency when compared to the links below them.


These links show consistency.

The relative consistency of each link in a hierarchical cluster tree can be quantified and expressed as the inconsistency coefficient. This value compares the height of a link in a cluster hierarchy with the average height of links below it. Links that join distinct clusters have a high inconsistency coefficient; links that join indistinct clusters have a low inconsistency coefficient.

To generate a listing of the inconsistency coefficient for each link in the cluster tree, use the inconsistent function. By default, the inconsistent
function compares each link in the cluster hierarchy with adjacent links that are less than two levels below it in the cluster hierarchy. This is called the depth of the comparison. You can also specify other depths. The objects at the bottom of the cluster tree, called leaf nodes, that have no further objects below them, have an inconsistency coefficient of zero. Clusters that join two leaves also have a zero inconsistency coefficient.

For example, you can use the inconsistent function to calculate the inconsistency values for the links created by the linkage function in "Linkages" on page 11-6.

| $I=$ | inconsistent $(Z)$ |  |  |
| ---: | :--- | ---: | ---: |
| $I=$ |  |  |  |
|  | 1.0000 | 0 | 1.0000 |

The inconsistent function returns data about the links in an ( $m-1$ )-by- 4 matrix, whose columns are described in the following table.

| Column | Description |
| :--- | :--- |
| 1 | Mean of the heights of all the links included in the calculation |
| 2 | Standard deviation of all the links included in the calculation |
| 3 | Number of links included in the calculation |
| 4 | Inconsistency coefficient |

In the sample output, the first row represents the link between objects 4 and 5 . This cluster is assigned the index 6 by the linkage function. Because both 4 and 5 are leaf nodes, the inconsistency coefficient for the cluster is zero. The second row represents the link between objects 1 and 3 , both of which are also leaf nodes. This cluster is assigned the index 7 by the linkage function.

The third row evaluates the link that connects these two clusters, objects 6 and 7. (This new cluster is assigned index 8 in the linkage output). Column 3 indicates that three links are considered in the calculation: the link itself and the two links directly below it in the hierarchy. Column 1 represents the mean of the heights of these links. The inconsistent function uses the height
information output by the linkage function to calculate the mean. Column 2 represents the standard deviation between the links. The last column contains the inconsistency value for these links, 1.1547. It is the difference between the current link height and the mean, normalized by the standard deviation:

```
(2.0616 - 1.3539) / . }612
ans =
    1.1547
```

The following figure illustrates the links and heights included in this calculation.


Note In the preceding figure, the lower limit on the $y$-axis is set to 0 to show the heights of the links. To set the lower limit to 0 , select Axes Properties from the Edit menu, click the $\mathbf{Y}$ Axis tab, and enter 0 in the field immediately to the right of $\mathbf{Y}$ Limits.

Row 4 in the output matrix describes the link between object 8 and object 2 . Column 3 indicates that two links are included in this calculation: the link itself and the link directly below it in the hierarchy. The inconsistency coefficient for this link is 0.7071 .

The following figure illustrates the links and heights included in this calculation.


## Creating Clusters

After you create the hierarchical tree of binary clusters, you can prune the tree to partition your data into clusters using the cluster function. The cluster function lets you create clusters in two ways, as discussed in the following sections:

- "Finding Natural Divisions in Data" on page 11-17
- "Specifying Arbitrary Clusters" on page 11-18


## Finding Natural Divisions in Data

The hierarchical cluster tree may naturally divide the data into distinct, well-separated clusters. This can be particularly evident in a dendrogram diagram created from data where groups of objects are densely packed in certain areas and not in others. The inconsistency coefficient of the links in the cluster tree can identify these divisions where the similarities between objects change abruptly. (See "Verifying the Cluster Tree" on page 11-10 for more information about the inconsistency coefficient.) You can use this value to determine where the cluster function creates cluster boundaries.

For example, if you use the cluster function to group the sample data set into clusters, specifying an inconsistency coefficient threshold of 1.2 as the value of the cutoff argument, the cluster function groups all the objects in the sample data set into one cluster. In this case, none of the links in the cluster hierarchy had an inconsistency coefficient greater than 1.2.

```
T = cluster(Z,'cutoff',1.2)
T =
```

    1
    1
    1
    1
    1
    The cluster function outputs a vector, T , that is the same size as the original data set. Each element in this vector contains the number of the cluster into which the corresponding object from the original data set was placed.

If you lower the inconsistency coefficient threshold to 0.8 , the cluster function divides the sample data set into three separate clusters.

```
T = cluster(Z,'cutoff',0.8)
T =
    3
    2
    3
    1
    1
```

This output indicates that objects 1 and 3 were placed in cluster 1, objects 4 and 5 were placed in cluster 2 , and object 2 was placed in cluster 3 .

When clusters are formed in this way, the cutoff value is applied to the inconsistency coefficient. These clusters may, but do not necessarily, correspond to a horizontal slice across the dendrogram at a certain height. If you want clusters corresponding to a horizontal slice of the dendrogram, you can either use the criterion option to specify that the cutoff should be based on distance rather than inconsistency, or you can specify the number of clusters directly as described in the following section.

## Specifying Arbitrary Clusters

Instead of letting the cluster function create clusters determined by the natural divisions in the data set, you can specify the number of clusters you want created.

For example, you can specify that you want the cluster function to partition the sample data set into two clusters. In this case, the cluster function creates one cluster containing objects $1,3,4$, and 5 and another cluster containing object 2 .

```
T = cluster(Z,'maxclust',2)
T =
    2
    1
    2
    2
    2
```

To help you visualize how the cluster function determines these clusters, the following figure shows the dendrogram of the hierarchical cluster tree. The horizontal dashed line intersects two lines of the dendrogram, corresponding to setting 'maxclust' to 2 . These two lines partition the objects into two clusters: the objects below the left-hand line, namely $1,3,4$, and 5 , belong to one cluster, while the object below the right-hand line, namely 2 , belongs to the other cluster.


On the other hand, if you set 'maxclust' to 3 , the cluster function groups objects 4 and 5 in one cluster, objects 1 and 3 in a second cluster, and object 2 in a third cluster. The following command illustrates this.

```
T = cluster(Z,'maxclust',3)
T =
    1
    3
    1
    2
    2
```

This time, the cluster function cuts off the hierarchy at a lower point, corresponding to the horizontal line that intersects three lines of the dendrogram in the following figure.


## K-Means Clustering

In this section...<br>"Introduction to K-Means Clustering" on page 11-21<br>"Creating Clusters and Determining Separation" on page 11-22<br>"Determining the Correct Number of Clusters" on page 11-23<br>"Avoiding Local Minima" on page 11-26

## Introduction to K-Means Clustering

$K$-means clustering is a partitioning method. The function kmeans partitions data into $k$ mutually exclusive clusters, and returns the index of the cluster to which it has assigned each observation. Unlike hierarchical clustering, $k$-means clustering operates on actual observations (rather than the larger set of dissimilarity measures), and creates a single level of clusters. The distinctions mean that $k$-means clustering is often more suitable than hierarchical clustering for large amounts of data.
kmeans treats each observation in your data as an object having a location in space. It finds a partition in which objects within each cluster are as close to each other as possible, and as far from objects in other clusters as possible. You can choose from five different distance measures, depending on the kind of data you are clustering.

Each cluster in the partition is defined by its member objects and by its centroid, or center. The centroid for each cluster is the point to which the sum of distances from all objects in that cluster is minimized. kmeans computes cluster centroids differently for each distance measure, to minimize the sum with respect to the measure that you specify.
kmeans uses an iterative algorithm that minimizes the sum of distances from each object to its cluster centroid, over all clusters. This algorithm moves objects between clusters until the sum cannot be decreased further. The result is a set of clusters that are as compact and well-separated as possible. You can control the details of the minimization using several optional input parameters to kmeans, including ones for the initial values of the cluster centroids, and for the maximum number of iterations.

## Creating Clusters and Determining Separation

The following example explores possible clustering in four-dimensional data by analyzing the results of partitioning the points into three, four, and five clusters.

Note Because each part of this example generates random numbers sequentially, i.e., without setting a new state, you must perform all steps in sequence to duplicate the results shown. If you perform the steps out of sequence, the answers will be essentially the same, but the intermediate results, number of iterations, or ordering of the silhouette plots may differ.

First, load some data:

```
load kmeansdata;
size(X)
ans =
    560 4
```

Even though these data are four-dimensional, and cannot be easily visualized, kmeans enables you to investigate whether a group structure exists in them. Call kmeans with k, the desired number of clusters, equal to 3. For this example, specify the city block distance measure, and use the default starting method of initializing centroids from randomly selected data points:

```
idx3 = kmeans(X,3,'distance','city');
```

To get an idea of how well-separated the resulting clusters are, you can make a silhouette plot using the cluster indices output from kmeans. The silhouette plot displays a measure of how close each point in one cluster is to points in the neighboring clusters. This measure ranges from +1 , indicating points that are very distant from neighboring clusters, through 0 , indicating points that are not distinctly in one cluster or another, to -1 , indicating points that are probably assigned to the wrong cluster. silhouette returns these values in its first output:

```
[silh3,h] = silhouette(X,idx3,'city');
set(get(gca,'Children'),'FaceColor',[.8 .8 1])
xlabel('Silhouette Value')
ylabel('Cluster')
```



From the silhouette plot, you can see that most points in the third cluster have a large silhouette value, greater than 0.6 , indicating that the cluster is somewhat separated from neighboring clusters. However, the first cluster contains many points with low silhouette values, and the second contains a few points with negative values, indicating that those two clusters are not well separated.

## Determining the Correct Number of Clusters

Increase the number of clusters to see if kmeans can find a better grouping of the data. This time, use the optional 'display ' parameter to print information about each iteration:

```
idx4 = kmeans(X,4, 'dist','city', 'display','iter');
    iter phase num sum
        1 1 560 2897.56
```

| 2 | 1 | 53 | 2736.67 |
| :---: | :---: | :---: | :---: |
| 3 | 1 | 50 | 2476.78 |
| 4 | 1 | 102 | 1779.68 |
| 5 | 1 | 5 | 1771.1 |
| 6 | 2 | 0 | 1771.1 |
| 6 iterations, total sum of distances $=1771.1$ |  |  |  |

Notice that the total sum of distances decreases at each iteration as kmeans reassigns points between clusters and recomputes cluster centroids. In this case, the second phase of the algorithm did not make any reassignments, indicating that the first phase reached a minimum after five iterations. In some problems, the first phase might not reach a minimum, but the second phase always will.

A silhouette plot for this solution indicates that these four clusters are better separated than the three in the previous solution:

```
[silh4,h] = silhouette(X,idx4,'city');
set(get(gca,'Children'),'FaceColor',[.8 . 8 1])
xlabel('Silhouette Value')
ylabel('Cluster')
```



A more quantitative way to compare the two solutions is to look at the average silhouette values for the two cases:

```
mean(silh3)
ans =
    0.52594
mean(silh4)
ans =
    0.63997
```

Finally, try clustering the data using five clusters:

```
idx5 = kmeans(X,5,'dist','city','replicates',5);
[silh5,h] = silhouette(X,idx5,'city');
set(get(gca,'Children'),'FaceColor',[.8 .8 1])
xlabel('Silhouette Value')
```



This silhouette plot indicates that this is probably not the right number of clusters, since two of the clusters contain points with mostly low silhouette values. Without some knowledge of how many clusters are really in the data, it is a good idea to experiment with a range of values for k .

## Avoiding Local Minima

Like many other types of numerical minimizations, the solution that kmeans reaches often depends on the starting points. It is possible for kmeans to reach a local minimum, where reassigning any one point to a new cluster would increase the total sum of point-to-centroid distances, but where a
better solution does exist. However, you can use the optional 'replicates' parameter to overcome that problem.

For four clusters, specify five replicates, and use the 'display ' parameter to print out the final sum of distances for each of the solutions.

```
[idx4,cent4,sumdist] = kmeans(X,4,'dist','city',...
    'display','final','replicates',5);
1 7 \text { iterations, total sum of distances = 2303.36}
5 iterations, total sum of distances = 1771.1
6 iterations, total sum of distances = 1771.1
5 iterations, total sum of distances = 1771.1
8 iterations, total sum of distances = 2303.36
```

The output shows that, even for this relatively simple problem, non-global minima do exist. Each of these five replicates began from a different randomly selected set of initial centroids, and kmeans found two different local minima. However, the final solution that kmeans returns is the one with the lowest total sum of distances, over all replicates.

```
sum(sumdist)
ans =
    1771.1
```


## Gaussian Mixture Models

In this section...<br>"Introduction to Gaussian Mixture Models" on page 11-28<br>"Clustering with Gaussian Mixtures" on page 11-28

## Introduction to Gaussian Mixture Models

Gaussian mixture models are formed by combining multivariate normal density components. For information on individual multivariate normal densities, see "Multivariate Normal Distribution" on page B-58 and related distribution functions listed under "Multivariate Distributions" on page 5-8.

In Statistics Toolbox software, use the gmdistribution class to fit data using an expectation maximization (EM) algorithm, which assigns posterior probabilities to each component density with respect to each observation.

Gaussian mixture models are often used for data clustering. Clusters are assigned by selecting the component that maximizes the posterior probability. Like $k$-means clustering, Gaussian mixture modeling uses an iterative algorithm that converges to a local optimum. Gaussian mixture modeling may be more appropriate than $k$-means clustering when clusters have different sizes and correlation within them. Clustering using Gaussian mixture models is sometimes considered a soft clustering method. The posterior probabilities for each point indicate that each data point has some probability of belonging to each cluster.

Creation of Gaussian mixture models is described in the "Gaussian Mixture Models" on page 5-99 section of Chapter 5, "Probability Distributions". This section describes their application in cluster analysis.

## Clustering with Gaussian Mixtures

Gaussian mixture distributions can be used for clustering data, by realizing that the multivariate normal components of the fitted model can represent clusters.

1 To demonstrate the process, first generate some simulated data from a mixture of two bivariate Gaussian distributions using the mvnrnd function:

```
mu1 = [1 2];
sigma1 = [3 .2; .2 2];
mu2 = [-1 -2];
sigma2 = [2 0; 0 1];
X = [mvnrnd(mu1,sigma1,200);mvnrnd(mu2,sigma2,100)];
scatter(X(:,1),X(:,2),10,'ko')
```



2 Fit a two-component Gaussian mixture distribution. Here, you know the correct number of components to use. In practice, with real data, this decision would require comparing models with different numbers of components.

```
options = statset('Display','final');
gm = gmdistribution.fit(X,2,'Options',options);
```

This displays

```
49 iterations, log-likelihood = -1207.91
```

3 Plot the estimated probability density contours for the two-component mixture distribution. The two bivariate normal components overlap, but their peaks are distinct. This suggests that the data could reasonably be divided into two clusters:

```
hold on
ezcontour(@(x,y)pdf(gm,[x y]),[-8 6],[-8 6]);
hold off
```



4 Partition the data into clusters using the cluster method for the fitted mixture distribution. The cluster method assigns each point to one of the two components in the mixture distribution.

```
idx = cluster(gm,X);
cluster1 = (idx == 1);
cluster2 = (idx == 2);
scatter(X(cluster1,1),X(cluster1,2),10,'r+');
hold on
scatter(X(cluster2,1),X(cluster2,2),10,'bo');
hold off
legend('Cluster 1','Cluster 2','Location','NW')
```



Each cluster corresponds to one of the bivariate normal components in the mixture distribution. cluster assigns points to clusters based on the estimated posterior probability that a point came from a component; each point is assigned to the cluster corresponding to the highest posterior probability. The posterior method returns those posterior probabilities.

For example, plot the posterior probability of the first component for each point:

```
P = posterior(gm,X);
scatter(X(cluster1,1),X(cluster1,2),10,P(cluster1,1),'+')
hold on
scatter(X(cluster2,1),X(cluster2,2),10,P(cluster2,1), 'o')
hold off
legend('Cluster 1','Cluster 2','Location','NW')
clrmap = jet(80); colormap(clrmap(9:72,:))
```

```
ylabel(colorbar,'Component 1 Posterior Probability')
```



## Soft Clustering Using Gaussian Mixture Distributions

An alternative to the previous example is to use the posterior probabilities for "soft clustering". Each point is assigned a membership score to each cluster. Membership scores are simply the posterior probabilities, and describe how similar each point is to each cluster's archetype, i.e., the mean of the corresponding component. The points can be ranked by their membership score in a given cluster:

```
[~,order] = sort(P(:,1));
plot(1:size(X,1),P(order,1),'r-',1:size(X,1),P(order,2),'b-');
legend({'Cluster 1 Score' 'Cluster 2 Score'},'location','NW');
ylabel('Cluster Membership Score');
xlabel('Point Ranking');
```



Although a clear separation of the data is hard to see in a scatter plot of the data, plotting the membership scores indicates that the fitted distribution does a good job of separating the data into groups. Very few points have scores close to 0.5 .

Soft clustering using a Gaussian mixture distribution is similar to fuzzy K-means clustering, which also assigns each point to each cluster with a membership score. The fuzzy K-means algorithm assumes that clusters are roughly spherical in shape, and all of roughly equal size. This is comparable to a Gaussian mixture distribution with a single covariance matrix that is shared across all components, and is a multiple of the identity matrix. In contrast, gmdistribution allows you to specify different covariance options. The default is to estimate a separate, unconstrained covariance matrix for
each component. A more restricted option, closer to K-means, would be to estimate a shared, diagonal covariance matrix:

```
gm2 = gmdistribution.fit(X,2,'CovType','Diagonal',...
    'SharedCov',true);
```

This covariance option is similar to fuzzy K-means clustering, but provides more flexibility by allowing unequal variances for different variables.

You can compute the soft cluster membership scores without computing hard cluster assignments, using posterior, or as part of hard clustering, as the second output from cluster:

```
P2 = posterior(gm2,X); % equivalently [idx,P2] = cluster(gm2,X)
[~,order] = sort(P2(:,1));
plot(1:size(X,1),P2(order,1),'r-',1:size(X,1),P2(order,2),'b-');
legend({'Cluster 1 Score' 'Cluster 2 Score'},'location','NW');
ylabel('Cluster Membership Score');
xlabel('Point Ranking');
```



## Assigning New Data to Clusters

In the previous example, fitting the mixture distribution to data using fit, and clustering those data using cluster, are separate steps. However, the same data are used in both steps. You can also use the cluster method to assign new data points to the clusters (mixture components) found in the original data.

1 Given a data set X , first fit a Gaussian mixture distribution. The previous code has already done that.
gm
gm =
Gaussian mixture distribution with 2 components in 2 dimensions

```
Component 1:
Mixing proportion: 0.312592
Mean: -0.9082 -2.1109
Component 2:
Mixing proportion: 0.687408
Mean: 0.9532 1.8940
```

2 You can then use cluster to assign each point in a new data set, $Y$, to one of the clusters defined for the original data:

```
Y = [mvnrnd(mu1,sigma1,50);mvnrnd(mu2,sigma2,25)];
idx = cluster(gm,Y);
cluster1 = (idx == 1);
cluster2 = (idx == 2);
scatter(Y(cluster1,1),Y(cluster1,2),10,'r+');
hold on
scatter(Y(cluster2,1),Y(cluster2,2),10,'bo');
hold off
legend('Class 1','Class 2','Location','NW')
```



As with the previous example, the posterior probabilities for each point can be treated as membership scores rather than determining "hard" cluster assignments.

For cluster to provide meaningful results with new data, Y should come from the same population as X , the original data used to create the mixture distribution. In particular, the estimated mixing probabilities for the Gaussian mixture distribution fitted to X are used when computing the posterior probabilities for Y .

## Parametric Classification

- "Introduction to Parametric Classification" on page 12-2
- "Discriminant Analysis" on page 12-3
- "Naive Bayes Classification" on page 12-29
- "Performance Curves" on page 12-32


## Introduction to Parametric Classification

Models of data with a categorical response are called classifiers. A classifier is built from training data, for which classifications are known. The classifier assigns new test data to one of the categorical levels of the response.

Parametric methods, like "Discriminant Analysis" on page 12-3, fit a parametric model to the training data and interpolate to classify test data.

Nonparametric methods, like "Classification Trees and Regression Trees" on page 13-27, use other means to determine classifications. In this sense, classification methods are analogous to the methods discussed in "Nonlinear Regression" on page 9-72.

## Discriminant Analysis

```
In this section...
"About Discriminant Analysis" on page 12-3
"Example: Create Discriminant Analysis Classifiers" on page 12-3
"How the ClassificationDiscriminant.fit Method Creates a Classifier" on
page 12-4
"How the predict Method Classifies" on page 12-6
"Example: Creating and Visualizing a Discriminant Analysis Classifier"
on page 12-9
"Improving a Discriminant Analysis Classifier" on page 12-14
"Examining the Gaussian Mixture Assumption" on page 12-22
"Bibliography" on page 12-28
```


## About Discriminant Analysis

Discriminant analysis is a classification method. It assumes that different classes generate data based on different Gaussian distributions.

- To train (create) a classifier, the fitting function estimates the parameters of a Gaussian distribution for each class (see "How the ClassificationDiscriminant.fit Method Creates a Classifier" on page 12-4).
- To predict the classes of new data, the trained classifier by default finds the class with the largest posterior probability (see "How the predict Method Classifies" on page 12-6).

To learn how to prepare your data for discriminant analysis and create a classifier, see "Steps in Supervised Learning (Machine Learning)" on page 13-2.

## Example: Create Discriminant Analysis Classifiers

To create the basic types of discriminant analysis classifiers for the Fisher iris data:

1 Load the data:

```
load fisheriris;
```

2 Create a default (linear) discriminant analysis classifier:

```
linclass = ClassificationDiscriminant.fit(meas,species);
```

To visualize the classification boundaries of a 2-D linear classification of the data, see Linear Discriminant Classification - Fisher Training Data on page 12-12.

3 Classify an iris with average measurements:

```
meanmeas = mean(meas);
meanclass = predict(linclass,meanmeas)
meanclass =
    'versicolor'
```

4 Create a quadratic classifier:

```
quadclass = ClassificationDiscriminant.fit(meas,species,...
    'discrimType','quadratic');
```

To visualize the classification boundaries of a 2-D quadratic classification of the data, see Quadratic Discriminant Classification - Fisher Training Data on page 12-14.

5 Classify an iris with average measurements using the quadratic classifier:

```
meanclass2 = predict(quadclass,meanmeas)
meanclass2 =
    'versicolor'
```


## How the ClassificationDiscriminant.fit Method Creates a Classifier

The model for discriminant analysis is:

- Each class $(Y)$ generates data $(X)$ using a multivariate normal distribution. In other words, the model assumes $X$ has a Gaussian mixture distribution (gmdistribution).
- For linear discriminant analysis, the model has the same covariance matrix for each class; only the means vary.
- For quadratic discriminant analysis, both means and covariances of each class vary.

Under this modeling assumption, ClassificationDiscriminant.fit infers the mean and covariance parameters of each class.

- For linear discriminant analysis, it computes the sample mean of each class. Then it computes the sample covariance by first subtracting the sample mean of each class from the observations of that class, and taking the empirical covariance matrix of the result.
- For quadratic discriminant analysis, it computes the sample mean of each class. Then it computes the sample covariances by first subtracting the sample mean of each class from the observations of that class, and taking the empirical covariance matrix of each class.

The fit method does not use prior probabilities or costs for fitting.

## Weighted Observations

The fit method constructs weighted classifiers using the following scheme. Suppose $M$ is an $N$-by- $K$ class membership matrix:
$M_{n k}=1$ if observation $n$ is from class $k$
$M_{n k}=0$ otherwise.
The estimate of the class mean for unweighted data is

$$
\hat{\mu}_{k}=\frac{\sum_{n=1}^{N} M_{n k} x_{n}}{\sum_{n=1}^{N} M_{n k}} .
$$

For weighted data with positive weights $w_{n}$, the natural generalization is

$$
\hat{\mu}_{k}=\frac{\sum_{n=1}^{N} M_{n k} w_{n} x_{n}}{\sum_{n=1}^{N} M_{n k} w_{n}} .
$$

The unbiased estimate of the pooled-in covariance matrix for unweighted data is

$$
\hat{\Sigma}=\frac{\sum_{n=1}^{N} \sum_{k=1}^{K} M_{n k}\left(x_{n}-\hat{\mu}_{k}\right)\left(x_{n}-\hat{\mu}_{k}\right)^{T}}{N-K} .
$$

For quadratic discriminant analysis, the fit method uses $K=1$.
For weighted data, assuming the weights sum to 1 , the unbiased estimate of the pooled-in covariance matrix is

$$
\hat{\Sigma}=\frac{\sum_{n=1}^{N} \sum_{k=1}^{K} M_{n k} w_{n}\left(x_{n}-\hat{\mu}_{k}\right)\left(x_{n}-\hat{\mu}_{k}\right)^{T}}{1-\sum_{k=1}^{K} \frac{W_{k}^{(2)}}{W_{k}}}
$$

where

- $W_{k}=\sum_{n=1}^{N} M_{n k} w_{n}$ is the sum of the weights for class $k$.
- $W_{k}^{(2)}=\sum_{n=1}^{N} M_{n k} w_{n}^{2}$ is the sum of squared weights per class.


## How the predict Method Classifies

There are three elements in the predict classification algorithm:

- "Posterior Probability" on page 12-7
- "Prior Probability" on page 12-8
- "Cost" on page 12-8
predict classifies so as to minimize the expected classification cost:

$$
\hat{y}=\underset{y=1, \ldots, K}{\arg \min } \sum_{k=1}^{K} \hat{P}(k \mid x) C(y \mid k),
$$

where

- $\hat{y}$ is the predicted classification.
- $K$ is the number of classes.
- $\hat{P}(k \mid x)$ is the posterior probability of class $k$ for observation $x$.
- $C(y \mid k)$ is the cost of classifying an observation as $y$ when its true class is $k$.

The space of $X$ values divides into regions where a classification $Y$ is a particular value. The regions are separated by straight lines for linear discriminant analysis, and by conic sections (ellipses, hyperbolas, or parabolas) for quadratic discriminant analysis. For a visualization of these regions, see "Example: Creating and Visualizing a Discriminant Analysis Classifier" on page 12-9.

## Posterior Probability

The posterior probability that a point $x$ belongs to class $k$ is the product of the prior probability and the multivariate normal density. The density function of the multivariate normal with mean $\mu_{k}$ and covariance $\Sigma_{k}$ at a point $x$ is

$$
P(x \mid k)=\frac{1}{\left(2 \pi\left|\Sigma_{k}\right|\right)^{1 / 2}} \exp \left(-\frac{1}{2}\left(x-\mu_{k}\right)^{T} \Sigma_{k}^{-1}\left(x-\mu_{k}\right)\right),
$$

where $\left|\Sigma_{k}\right|$ is the determinant of $\Sigma_{k}$, and $\Sigma_{k}^{-1}$ is the inverse matrix.
Let $P(k)$ represent the prior probability of class $k$. Then the posterior probability that an observation $x$ is of class $k$ is

$$
\hat{P}(k \mid x)=\frac{P(x \mid k) P(k)}{P(x)}
$$

where $P(x)$ is a normalization constant, namely, the sum over $k$ of $P(x \mid k) P(k)$.

## Prior Probability

The prior probability is one of three choices:

- 'uniform' - The prior probability of class k is 1 over the total number of classes.
- 'empirical' - The prior probability of class $k$ is the number of training samples of class k divided by the total number of training samples.
- A numeric vector - The prior probability of class $k$ is the $j$ th element of the prior vector. See ClassificationDiscriminant.fit.

After creating a classifier obj, you can set the prior using dot addressing:

```
obj.Prior = v;
```

where $v$ is a vector of positive elements representing the frequency with which each element occurs. You do not need to retrain the classifier when you set a new prior.

## Cost

There are two costs associated with discriminant analysis classification: the true misclassification cost per class, and the expected misclassification cost per observation.

True Misclassification Cost per Class. Cost (i, j ) is the cost of classifying an observation into class $j$ if its true class is $i$. By default, $\operatorname{Cost}(i, j)=1$ if $i \sim=j$, and $\operatorname{Cost}(i, j)=0$ if $i=j$. In other words, the cost is 0 for correct classification, and 1 for incorrect classification.

You can set any cost matrix you like when creating a classifier. Pass the cost matrix in the Cost name-value pair in ClassificationDiscriminant.fit.

After you create a classifier obj, you can set a custom cost using dot addressing:
obj.Cost = B;
$B$ is a square matrix of size K -by- K when there are K classes. You do not need to retrain the classifier when you set a new cost.

Expected Misclassification Cost per Observation. Suppose you have Nobs observations that you want to classify with a trained discriminant analysis classifier obj. Suppose you have K classes. You place the observations into a matrix Xnew with one observation per row. The command

```
[label,score,cost] = predict(obj,Xnew)
```

returns, among other outputs, a cost matrix of size Nobs-by-K. Each row of the cost matrix contains the expected (average) cost of classifying the observation into each of the $K$ classes. cost ( $n, k$ ) is

$$
\sum_{i=1}^{K} \hat{P}(i \mid X n e w(n)) C(k \mid i)
$$

where

- $K$ is the number of classes.
- $\hat{P}(i \mid X n e w(n))$ is the posterior probability of class $i$ for observation Xnew(n).
- $C(k \mid i)$ is the cost of classifying an observation as $k$ when its true class is $i$.


## Example: Creating and Visualizing a Discriminant Analysis Classifier

This example shows both linear and quadratic classification of the Fisher iris data. The example uses only two of the four predictors to enable simple plotting.

1 Load the data:

```
load fisheriris
```

2 Use the petal length (PL) and petal width (PW) measurements:

```
PL = meas(:,3);
PW = meas(:,4);
```

3 Plot the data, showing the classification:

```
h1 = gscatter(PL,PW,species,'krb','ov^',[],'off');
set(h1,'LineWidth',2)
legend('Setosa','Versicolor','Virginica',...
                            'Location','best')
```



4 Create a linear classifier:

```
X = [PL,PW];
cls = ClassificationDiscriminant.fit(X,species);
```

5 Plot the classification boundaries:

```
hold on
K = cls.Coeffs(2,3).Const;
L = cls.Coeffs(2,3).Linear;
% Plot the curve K + [x,y]*L = 0:
f = @(x1,x2) K + L(1)*x1 + L(2)*x2;
h2 = ezplot(f,[.9 7.1 0 2.5]);
set(h2,'Color','r','LineWidth',2)
K = cls.Coeffs(1,2).Const;
L = cls.Coeffs(1,2).Linear;
% Plot the curve K + [x1,x2]*L = 0:
f = @(x1,x2) K + L(1)*x1 + L(2)*x2;
h3 = ezplot(f,[.9 7.1 0 2.5]);
set(h3,'Color','k','LineWidth',2)
axis([.9 7.1 0 2.5])
xlabel('Petal Length')
ylabel('Petal Width')
title('{\bf Linear Classification with Fisher Training Data}')
```



## Linear Discriminant Classification - Fisher Training Data

6 Create a quadratic discriminant classifier:

```
cqs = ClassificationDiscriminant.fit(X,species,...
    'DiscrimType', 'quadratic');
```

7 Plot the classification boundaries:

```
delete(h2); delete(h3) % remove the linear plots
K = cqs.Coeffs(2,3).Const;
L = cqs.Coeffs(2,3).Linear;
Q = cqs.Coeffs(2,3).Quadratic;
% Plot the curve K + [x1,x2]*L + [x1,x2]*Q*[x1,x2]'=0:
```

```
f = @(x1,x2) K + L(1)*x1 + L(2)*x2 + Q(1,1)*x1.^2 + ...
    (Q(1,2)+Q(2,1))*x1.*x2 + Q(2,2)*x2.^2;
h2 = ezplot(f,[.9 7.1 0 2.5]);
set(h2,'Color','r','LineWidth',2)
K = cqs.Coeffs(1,2).Const;
L = cqs.Coeffs(1,2).Linear;
Q = cqs.Coeffs(1,2).Quadratic;
% Plot the curve K + [x1,x2]*L + [x1,x2]*Q*[x1,x2]'=0:
f = @(x1,x2) K + L(1)*x1 + L(2)*x2 + Q(1,1)*x1.^2 + ...
    (Q(1,2)+Q(2,1))*x1.*x2 + Q(2,2)*x2.^2;
h3 = ezplot(f,[.9 7.1 0 1.02]); % plot the relevant
                                    % portion of the curve
set(h3,'Color','k','LineWidth',2)
axis([.9 7.1 0 2.5])
xlabel('Petal Length')
ylabel('Petal Width')
title('{\bf Quadratic Classification with Fisher Training Data}')
hold off
```



## Quadratic Discriminant Classification - Fisher Training Data

## Improving a Discriminant Analysis Classifier

- "Deal with Singular Data" on page 12-15
- "Choose a Discriminant Type" on page 12-16
- "Examine the Resubstitution Error and Confusion Matrix" on page 12-17
- "Cross Validation" on page 12-18
- "Change Costs and Priors" on page 12-19


## Deal with Singular Data

Discriminant analysis needs data sufficient to fit Gaussian models with invertible covariance matrices. If your data is not sufficient to fit such a model uniquely, ClassificationDiscriminant.fit fails. This section shows methods for handling failures.

Tip To obtain a discriminant analysis classifier without failure, set the DiscrimType name-value pair to 'pseudoLinear' or 'pseudoQuadratic' in ClassificationDiscriminant.fit.
"Pseudo" discriminants never fail, because they use the pseudoinverse of the covariance matrix $\Sigma_{k}$ (see pinv).

Example: Singular Covariance Matrix. When the covariance matrix of the fitted classifier is singular, ClassificationDiscriminant.fit fails:

```
load popcorn
X = popcorn(:,[1 2]);
Y = popcorn(:,3);
ppcrn = ClassificationDiscriminant.fit(X,Y);
Error using
ClassificationDiscriminant>ClassificationDiscriminant.ClassificationDiscriminant
You cannot obtain discriminant coefficients because the pooled-in matrix is singular.
Set 'discrimType' to 'pseudoLinear' or 'diagLinear'.
Error in FitTemplate>FitTemplate.fit at 212
    obj = this.MakeFitObject(X,Y,W,this.ModelParams,fitArgs{:});
Error in ClassificationDiscriminant>ClassificationDiscriminant.fit at 43
    this = fit(temp,X,Y);
```

To proceed with linear discriminant analysis, use a pseudoLinear or diagLinear discriminant type:

```
ppcrn = ClassificationDiscriminant.fit(X,Y,...
    'discrimType','pseudoLinear');
meanpredict = predict(ppcrn,mean(X))
```

```
meanpredict =
    4
```


## Choose a Discriminant Type

There are six types of discriminant analysis classifiers: linear and quadratic, with diagonal and pseudo variants of each type.

Tip To see if your covariance matrix is singular, set discrimType to 'linear ' or 'quadratic'. If it is singular, the ClassificationDiscriminant.fit method fails.

To obtain a classifier even when your covariance matrix is singular, set discrimType to 'pseudoQuadratic', 'pseudoLinear', or either diagonal type.

```
obj = ClassificationDiscriminant.fit(X,Y,...
    'discrimType','pseudoQuadratic') % or 'pseudoLinear'
```

Choose a classifier type by setting the discrimType name-value pair to one of:

- 'linear' (default) - Estimate one covariance matrix for all classes.
- 'quadratic' - Estimate one covariance matrix for each class.
- 'diagLinear' - Use the diagonal of the 'linear' covariance matrix, and use its pseudoinverse if necessary.
- 'diagQuadratic' - Use the diagonals of the 'quadratic' covariance matrices, and use their pseudoinverses if necessary.
- 'pseudoLinear' - Use the pseudoinverse of the 'linear' covariance matrix if necessary.
- 'pseudoQuadratic' - Use the pseudoinverses of the 'quadratic' covariance matrices if necessary.

ClassificationDiscriminant.fit can fail for the 'linear' and 'quadratic' classifiers. When it fails, it returns an explanation, as shown in "Deal with Singular Data" on page 12-15.

ClassificationDiscriminant.fit always succeeds with the diagonal and pseudo variants. For information about pseudoinverses, see pinv.

You can set the discriminant type using dot addressing after constructing a classifier:

```
obj.DiscrimType = 'discrimType'
```

You can change between linear types or between quadratic types, but cannot change between a linear and a quadratic type.

## Examine the Resubstitution Error and Confusion Matrix

The resubstitution error is the difference between the response training data and the predictions the classifier makes of the response based on the input training data. If the resubstitution error is high, you cannot expect the predictions of the classifier to be good. However, having low resubstitution error does not guarantee good predictions for new data. Resubstitution error is often an overly optimistic estimate of the predictive error on new data.

The confusion matrix shows how many errors, and which types, arise in resubstitution. When there are K classes, the confusion matrix R is a K-by-K matrix with
$R(i, j)=$ the number of observations of class $i$ that the classifier predicts to be of class $j$.

## Example: Resubstitution Error of a Discriminant Analysis Classifier.

Examine the resubstitution error of the default discriminant analysis classifier for the Fisher iris data:

```
load fisheriris
obj = ClassificationDiscriminant.fit(meas,species);
resuberror = resubLoss(obj)
resuberror =
    0.0200
```

The resubstitution error is very low, meaning obj classifies nearly all the Fisher iris data correctly. The total number of misclassifications is:

```
resuberror * obj.NObservations
ans =
    3.0000
```

To see the details of the three misclassifications, examine the confusion matrix:

```
R = confusionmat(obj.Y,resubPredict(obj))
R =
    50 0 0
        0 48 2
        0 1 49
obj.ClassNames
ans =
    'setosa'
    'versicolor'
    'virginica'
```

- $R(1,:)=\left[\begin{array}{lll}50 & 0 & 0\end{array}\right]$ means obj classifies all 50 setosa irises correctly.
- $R(2,:)=\left[\begin{array}{ll}0 & 48 \\ 2\end{array}\right]$ means obj classifies 48 versicolor irises correctly, and misclassifies two versicolor irises as virginica.
- $R(3,:)=\left[\begin{array}{lll}0 & 1 & 49\end{array}\right]$ means obj classifies 49 virginica irises correctly, and misclassifies one virginica iris as versicolor.


## Cross Validation

Typically, discriminant analysis classifiers are robust and do not exhibit overtraining when the number of predictors is much less than the number of observations. Nevertheless, it is good practice to cross validate your classifier to ensure its stability.

Example: Cross Validating a Discriminant Analysis Classifier. Try five-fold cross validation of a quadratic discriminant analysis classifier:

1 Load the Fisher iris data:

```
load fisheriris
```

2 Create a quadratic discriminant analysis classifier for the data:

```
quadisc = ClassificationDiscriminant.fit(meas,species,...
    'DiscrimType','quadratic');
```

3 Find the resubstitution error of the classifier:

```
qerror = resubLoss(quadisc)
qerror =
    0.0200
```

The classifier does an excellent job. Nevertheless, resubstitution error can be an optimistic estimate of the error when classifying new data. So proceed to cross validation.

4 Create a cross-validation model:

```
cvmodel = crossval(quadisc,'kfold',5);
```

5 Find the cross-validation loss for the model, meaning the error of the out-of-fold observations:

```
cverror = kfoldLoss(cvmodel)
cverror =
    0.0333
```

The cross-validated loss is nearly as low as the original resubstitution loss. Therefore, you can have confidence that the classifier is reasonably accurate.

## Change Costs and Priors

Sometimes you want to avoid certain misclassification errors more than others. For example, it might be better to have oversensitive cancer detection instead of undersensitive cancer detection. Oversensitive detection gives more false positives (unnecessary testing or treatment). Undersensitive detection gives more false negatives (preventable illnesses or deaths). The
consequences of underdetection can be high. Therefore, you might want to set costs to reflect the consequences.

Similarly, the training data $Y$ can have a distribution of classes that does not represent their true frequency. If you have a better estimate of the true frequency, you can include this knowledge in the classification prior property.

Example: Setting Custom Misclassification Costs. Consider the Fisher iris data. Suppose that the cost of classifying a versicolor iris as virginica is 10 times as large as making any other classification error. Create a classifier from the data, then incorporate this cost and then view the resulting classifier.

1 Load the Fisher iris data and create a default (linear) classifier as in "Example: Resubstitution Error of a Discriminant Analysis Classifier" on page 12-17:

```
load fisheriris
obj = ClassificationDiscriminant.fit(meas,species);
resuberror = resubLoss(obj)
resuberror =
    0.0200
R = confusionmat(obj.Y,resubPredict(obj))
R =
    50 0 0
    0 48 2
    0 1 49
obj.ClassNames
ans =
    'setosa'
    'versicolor'
    'virginica'
```

$R(2,:)=\left[\begin{array}{lll}0 & 48 & 2\end{array}\right]$ means obj classifies 48 versicolor irises correctly, and
misclassifies two versicolor irises as virginica.

2 Change the cost matrix to make fewer mistakes in classifying versicolor irises as virginica:

```
obj.Cost(2,3) = 10;
R2 = confusionmat(obj.Y,resubPredict(obj))
R2 =
\begin{tabular}{rrr}
50 & 0 & 0 \\
0 & 50 & 0 \\
0 & 7 & 43
\end{tabular}
```

obj now classifies all versicolor irises correctly, at the expense of increasing the number of misclassifications of virginica irises from 1 to 7 .

Example: Setting Alternative Priors. Consider the Fisher iris data. There are 50 irises of each kind in the data. Suppose that, in a particular region, you have historical data that shows virginica are five times as prevalent as the other kinds. Create a classifier that incorporates this information.

1 Load the Fisher iris data and make a default (linear) classifier as in "Example: Resubstitution Error of a Discriminant Analysis Classifier" on page 12-17:

```
load fisheriris
obj = ClassificationDiscriminant.fit(meas,species);
resuberror = resubLoss(obj)
resuberror =
    0.0200
R = confusionmat(obj.Y,resubPredict(obj))
R =
    50 0 0
    0 48 2
    0 1 49
obj.ClassNames
ans =
    'setosa'
    'versicolor'
```


## 'virginica'

$R(3,:)=\left[\begin{array}{lll}0 & 1 & 49\end{array}\right]$ means obj classifies 49 virginica irises correctly, and misclassifies one virginica iris as versicolor.

2 Change the prior to match your historical data, and examine the confusion matrix of the new classifier:

```
obj.Prior = [1 1 5];
R2 = confusionmat(obj.Y,resubPredict(obj))
R2 =
\begin{tabular}{rrr}
50 & 0 & 0 \\
0 & 46 & 4 \\
0 & 0 & 50
\end{tabular}
```

The new classifier classifies all virginica irises correctly, at the expense of increasing the number of misclassifications of versicolor irises from 2 to 4.

## Examining the Gaussian Mixture Assumption

Discriminant analysis assumes that the data comes from a Gaussian mixture model (see "How the ClassificationDiscriminant.fit Method Creates a Classifier" on page 12-4). If the data appears to come from a Gaussian mixture model, you can expect discriminant analysis to be a good classifier. Furthermore, the default linear discriminant analysis assumes that all class covariance matrices are equal. This section shows methods to check these assumptions:

- "Bartlett Test of Equal Covariance Matrices for Linear Discriminant Analysis" on page 12-22
- "Q-Q Plot" on page 12-24
- "Mardia Kurtosis Test of Multivariate Normality" on page 12-27


## Bartlett Test of Equal Covariance Matrices for Linear Discriminant Analysis

The Bartlett test (see Box [1]) checks equality of the covariance matrices of the various classes. If the covariance matrices are equal, the test indicates that linear discriminant analysis is appropriate. If not, consider using
quadratic discriminant analysis, setting the DiscrimType name-value pair to 'quadratic' in ClassificationDiscriminant.fit.

The Bartlett test assumes normal (Gaussian) samples, where neither the means nor covariance matrices are known. To determine whether the covariances are equal, compute the following quantities:

- Sample covariance matrices per class $\sigma_{i}, 1 \leq i \leq k$, where $k$ is the number of classes.
- Pooled-in covariance matrix $\sigma$.
- Test statistic $V$ :

$$
V=(n-k) \log (|\Sigma|)-\sum_{i=1}^{k}\left(n_{i}-1\right) \log \left(\left|\Sigma_{i}\right|\right)
$$

where $n$ is the total number of observations, and $n_{i}$ is the number of observations in class $i$, and $|\Sigma|$ means the determinant of the matrix $\Sigma$.

- Asymptotically, as the number of observations in each class $n_{i}$ become large, $V$ is distributed approximately $x^{2}$ with $k d(d+1) / 2$ degrees of freedom, where $d$ is the number of predictors (number of dimensions in the data).

The Bartlett test is to check whether $V$ exceeds a given percentile of the $x^{2}$ distribution with $k d(d+1) / 2$ degrees of freedom. If it does, then reject the hypothesis that the covariances are equal.

Example: Bartlett Test for Equal Covariance Matrices. Check whether the Fisher iris data is well modeled by a single Gaussian covariance, or whether it would be better to model it as a Gaussian mixture.

```
load fisheriris;
prednames = {'SepalLength','SepalWidth',...
    'PetalLength','PetalWidth'};
L = ClassificationDiscriminant.fit(meas,species,...
    'PredictorNames',prednames);
Q = ClassificationDiscriminant.fit(meas,species,...
    'PredictorNames',prednames,'DiscrimType','quadratic');
D = 4; % Number of dimensions of X
Nclass = [50 50 50];
```

```
N = L.NObservations;
K = numel(L.ClassNames);
SigmaQ = Q.Sigma;
SigmaL = L.Sigma;
logV = (N-K)*log(det(SigmaL));
for k=1:K
    logV = logV - (Nclass(k)-1)*log(det(SigmaQ(:,:,k)));
end
nu = (K-1)*D*(D+1)/2;
pval = 1-chi2cdf(logV,nu)
pval =
    0
```

The Bartlett test emphatically rejects the hypothesis of equal covariance matrices. If pval had been greater than 0.05 , the test would not have rejected the hypothesis. The result indicates to use quadratic discriminant analysis, as opposed to linear discriminant analysis.

## Q-Q Plot

A Q-Q plot graphically shows whether an empirical distribution is close to a theoretical distribution. If the two are equal, the Q-Q plot lies on a $45^{\circ}$ line. If not, the Q-Q plot strays from the $45^{\circ}$ line.

Check Q-Q Plots for Linear and Quadratic Discriminants. For linear discriminant analysis, use a single covariance matrix for all classes.

```
load fisheriris;
prednames = {'SepalLength','SepalWidth',...
    'PetalLength','PetalWidth'};
L = ClassificationDiscriminant.fit(meas,species,...
    'PredictorNames',prednames);
N = L.NObservations;
K = numel(L.ClassNames);
mahL = mahal(L,L.X,'ClassLabels',L.Y);
D = 4;
expQ = chi2inv(((1:N)-0.5)/N,D); % expected quantiles
[mahL,sorted] = sort(mahL); % sorted obbserved quantiles
figure;
gscatter(expQ,mahL,L.Y(sorted),'bgr',[],[],'off');
```

```
legend('virginica','versicolor','setosa','Location','NW');
xlabel('Expected quantile');
ylabel('Observed quantile');
line([0 20],[0 20],'color','k');
```



Overall, the agreement between the expected and observed quantiles is good. Look at the right half of the plot. The deviation of the plot from the $45^{\circ}$ line upward indicates that the data has tails heavier than a normal distribution. There are three possible outliers on the right: two observations from class 'setosa' and one observation from class 'virginica'.

As shown in "Bartlett Test of Equal Covariance Matrices for Linear Discriminant Analysis" on page 12-22, the data does not match a single covariance matrix. Redo the calculations for a quadratic discriminant.

```
load fisheriris;
```

```
prednames = {'SepalLength','SepalWidth',...
    'PetalLength','PetalWidth'};
Q = ClassificationDiscriminant.fit(meas,species,...
    'PredictorNames',prednames,'DiscrimType', 'quadratic');
Nclass = [50 50 50];
N = L.NObservations;
K = numel(L.ClassNames);
mahQ = mahal(Q,Q.X,'ClassLabels',Q.Y);
expQ = chi2inv(((1:N)-0.5)/N,D);
[mahQ,sorted] = sort(mahQ);
figure;
gscatter(expQ,mahQ,Q.Y(sorted),'bgr',[],[],'off');
legend('virginica','versicolor','setosa','Location', 'NW');
xlabel('Expected quantile');
ylabel('Observed quantile for QDA');
line([0 20],[0 20],'color','k');
```



The Q-Q plot shows a better agreement between the observed and expected quantiles. There is only one outlier candidate, from class 'setosa'.

## Mardia Kurtosis Test of Multivariate Normality

The Mardia kurtosis test (see Mardia [2]) is an alternative to examining a Q-Q plot. It gives a numeric approach to deciding if data matches a Gaussian mixture model.

In the Mardia kurtosis test you compute $M$, the mean of the fourth power of the Mahalanobis distance of the data from the class means. If the data is normally distributed with constant covariance matrix (and is thus suitable for linear discriminant analysis), $M$ is asymptotically distributed as normal with mean $d(d+2)$ and variance $8 d(d+2) / n$, where

- $d$ is the number of predictors (number of dimensions in the data).
- $n$ is the total number of observations.

The Mardia test is two sided: check whether $M$ is close enough to $d(d+2)$ with respect to a normal distribution of variance $8 d(d+2) / n$.

Example: Mardia Kurtosis Test for Linear and Quadratic
Discriminants. Check whether the Fisher iris data is approximately normally distributed for both linear and quadratic discriminant analysis. According to "Bartlett Test of Equal Covariance Matrices for Linear Discriminant Analysis" on page 12-22, the data is not normal for linear discriminant analysis (the covariance matrices are different). "Check Q-Q Plots for Linear and Quadratic Discriminants" on page 12-24 indicates that the data is well modeled by a Gaussian mixture model with different covariances per class. Check these conclusions with the Mardia kurtosis test:

```
load fisheriris;
prednames = {'SepalLength','SepalWidth',...
    'PetalLength','PetalWidth'};
L = ClassificationDiscriminant.fit(meas,species,...
    'PredictorNames',prednames);
mahL = mahal(L,L.X,'ClassLabels',L.Y);
D = 4;
N = L.NObservations;
obsKurt = mean(mahL.^2);
```

```
expKurt = D*(D+2);
varKurt = 8*D*(D+2)/N;
[~,pval] = ztest(obsKurt,expKurt,sqrt(varKurt))
pval =
    0.0208
```

The Mardia test indicates to reject the hypothesis that the data is normally distributed.

Continuing the example with quadratic discriminant analysis:

```
Q = ClassificationDiscriminant.fit(meas,species,...
    'PredictorNames',prednames,'DiscrimType','quadratic');
mahQ = mahal(Q,Q.X,'ClassLabels',Q.Y);
obsKurt = mean(mahQ.^2);
[~,pval] = ztest(obsKurt,expKurt,sqrt(varKurt))
pval =
    0.7230
```

Because pval is high, you conclude the data are consistent with the multivariate normal distribution.

## Bibliography

[1] Box, G. E. P. A General Distribution Theory for a Class of Likelihood Criteria. Biometrika 36(3), pp. 317-346, 1949.
[2] Mardia, K. V. Measures of multivariate skewness and kurtosis with applications. Biometrika 57 (3), pp. 519-530, 1970.

## Naive Bayes Classification

The Naive Bayes classifier is designed for use when features are independent of one another within each class, but it appears to work well in practice even when that independence assumption is not valid. It classifies data in two steps:

1 Training step: Using the training samples, the method estimates the parameters of a probability distribution, assuming features are conditionally independent given the class.

2 Prediction step: For any unseen test sample, the method computes the posterior probability of that sample belonging to each class. The method then classifies the test sample according the largest posterior probability.

The class-conditional independence assumption greatly simplifies the training step since you can estimate the one-dimensional class-conditional density for each feature individually. While the class-conditional independence between features is not true in general, research shows that this optimistic assumption works well in practice. This assumption of class independence allows the Naive Bayes classifier to better estimate the parameters required for accurate classification while using less training data than many other classifiers. This makes it particularly effective for datasets containing many predictors or features.

To learn how to prepare your data for Naive Bayes classification and create a classifier, see "Steps in Supervised Learning (Machine Learning)" on page 13-2.

## Supported Distributions

Naive Bayes classification is based on estimating $P(X \mid Y)$, the probability or probability density of features $X$ given class $Y$. The Naive Bayes classification object NaiveBayes provides support for normal (Gaussian), kernel, multinomial, and multivariate multinomial distributions. It is possible to use different distributions for different features.

## Normal (Gaussian) Distribution

The 'normal' distribution is appropriate for features that have normal distributions in each class. For each feature you model with a normal distribution, the Naive Bayes classifier estimates a separate normal distribution for each class by computing the mean and standard deviation of the training data in that class. For more information on normal distributions, see "Normal Distribution" on page B-83.

## Kernel Distribution

The 'kernel' distribution is appropriate for features that have a continuous distribution. It does not require a strong assumption such as a normal distribution and you can use it in cases where the distribution of a feature may be skewed or have multiple peaks or modes. It requires more computing time and more memory than the normal distribution. For each feature you model with a kernel distribution, the Naive Bayes classifier computes a separate kernel density estimate for each class based on the training data for that class. By default the kernel is the normal kernel, and the classifier selects a width automatically for each class and feature. It is possible to specify different kernels for each feature, and different widths for each feature or class.

## Multinomial Distribution

The multinomial distribution (specify with the 'mn' keyword) is appropriate when all features represent counts of a set of words or tokens. This is sometimes called the "bag of words" model. For example, an email spam classifier might be based on features that count the number of occurrences of various tokens in an email. One feature might count the number of exclamation points, another might count the number of times the word "money" appears, and another might count the number of times the recipient's name appears. This is a Naive Bayes model under the further assumption that the total number of tokens (or the total document length) is independent of response class.

For the multinomial option, each feature represents the count of one token. The classifier counts the set of relative token probabilities separately for each class. The classifier defines the multinomial distribution for each row by the vector of probabilities for the corresponding class, and by $N$, the total token count for that row.

Classification is based on the relative frequencies of the tokens. For a row in which no token appears, N is 0 and no classification is possible. This classifier is not appropriate when the total number of tokens provides information about the response class.

## Multivariate Multinomial Distribution

The multivariate multinomial distribution (specify with the 'mvmn' keyword) is appropriate for categorical features. For example, you could fit a feature describing the weather in categories such as rain/sun/snow/clouds using the multivariate multinomial model. The feature categories are sometimes called the feature levels, and differ from the class levels for the response variable.

For each feature you model with a multivariate multinomial distribution, the Naive Bayes classifier computes a separate set of probabilities for the set of feature levels for each class.

## Performance Curves

In this section...<br>"Introduction to Performance Curves" on page 12-32<br>"What are ROC Curves?" on page 12-32<br>"Evaluating Classifier Performance Using perfcurve" on page 12-32

## Introduction to Performance Curves

After a classification algorithm such as NaiveBayes or TreeBagger has trained on data, you may want to examine the performance of the algorithm on a specific test dataset. One common way of doing this would be to compute a gross measure of performance such as quadratic loss or accuracy, averaged over the entire test dataset.

## What are ROC Curves?

You may want to inspect the classifier performance more closely, for example, by plotting a Receiver Operating Characteristic (ROC) curve. By definition, a ROC curve [1,2] shows true positive rate versus false positive rate (equivalently, sensitivity versus 1 -specificity) for different thresholds of the classifier output. You can use it, for example, to find the threshold that maximizes the classification accuracy or to assess, in more broad terms, how the classifier performs in the regions of high sensitivity and high specificity.

## Evaluating Classifier Performance Using perfcurve

perfcurve computes measures for a plot of classifier performance. You can use this utility to evaluate classifier performance on test data after you train the classifier. Various measures such as mean squared error, classification error, or exponential loss can summarize the predictive power of a classifier in a single number. However, a performance curve offers more information as it lets you explore the classifier performance across a range of thresholds on its output.

You can use perfcurve with any classifier or, more broadly, with any method that returns a numeric score for an instance of input data. By convention adopted here,

- A high score returned by a classifier for any given instance signifies that the instance is likely from the positive class.
- A low score signifies that the instance is likely from the negative classes.

For some classifiers, you can interpret the score as the posterior probability of observing an instance of the positive class at point $X$. An example of such a score is the fraction of positive observations in a leaf of a decision tree. In this case, scores fall into the range from 0 to 1 and scores from positive and negative classes add up to unity. Other methods can return scores ranging between minus and plus infinity, without any obvious mapping from the score to the posterior class probability.
perfcurve does not impose any requirements on the input score range. Because of this lack of normalization, you can use perfcurve to process scores returned by any classification, regression, or fit method. perfcurve does not make any assumptions about the nature of input scores or relationships between the scores for different classes. As an example, consider a problem with three classes, A, B, and C, and assume that the scores returned by some classifier for two instances are as follows:

|  | A | B | C |
| :--- | :--- | :--- | :--- |
| instance 1 | 0.4 | 0.5 | 0.1 |
| instance 2 | 0.4 | 0.1 | 0.5 |

If you want to compute a performance curve for separation of classes $A$ and $B$, with $C$ ignored, you need to address the ambiguity in selecting $A$ over $B$. You could opt to use the score ratio, $s(A) / s(B)$, or score difference, $s(A)-s(B)$; this choice could depend on the nature of these scores and their normalization. perfcurve always takes one score per instance. If you only supply scores for class A, perfcurve does not distinguish between observations 1 and 2 . The performance curve in this case may not be optimal.
perfcurve is intended for use with classifiers that return scores, not those that return only predicted classes. As a counter-example, consider a decision tree that returns only hard classification labels, 0 or 1 , for data with two classes. In this case, the performance curve reduces to a single point because classified instances can be split into positive and negative categories in one way only.

For input, perfcurve takes true class labels for some data and scores assigned by a classifier to these data. By default, this utility computes a Receiver Operating Characteristic (ROC) curve and returns values of 1 -specificity, or false positive rate, for $X$ and sensitivity, or true positive rate, for $Y$. You can choose other criteria for $X$ and $Y$ by selecting one out of several provided criteria or specifying an arbitrary criterion through an anonymous function. You can display the computed performance curve using plot ( $\mathrm{X}, \mathrm{Y}$ ).
perfcurve can compute values for various criteria to plot either on the $x$ - or the $y$-axis. All such criteria are described by a 2 -by- 2 confusion matrix, a 2 -by- 2 cost matrix, and a 2 -by- 1 vector of scales applied to class counts.

The confusion matrix, $C$, is defined as

$$
\left(\begin{array}{ll}
T P & F N \\
F P & T N
\end{array}\right)
$$

where

- $P$ stands for "positive".
- $N$ stands for "negative".
- $T$ stands for "true".
- $F$ stands for "false".

For example, the first row of the confusion matrix defines how the classifier identifies instances of the positive class: $C(1,1)$ is the count of correctly identified positive instances and $C(1,2)$ is the count of positive instances misidentified as negative.

The cost matrix defines the cost of misclassification for each category:

$$
\left(\begin{array}{ll}
\operatorname{Cost}(P \mid P) & \operatorname{Cost}(N \mid P) \\
\operatorname{Cost}(P \mid N) & \operatorname{Cost}(N \mid N)
\end{array}\right)
$$

where Cost (I $\mid \mathrm{J})$ is the cost of assigning an instance of class $J$ to class I. Usually Cost $(I \mid J)=0$ for $I=J$. For flexibility, perfcurve allows you to specify nonzero costs for correct classification as well.

The two scales include prior information about class probabilities. perfcurve computes these scales by taking scale ( P ) = prior ( P ) *N and scale $(N)=\operatorname{prior}(N) * P$ and normalizing the sum scale (P)+scale(N) to 1. $P=T P+F N$ and $N=T N+F P$ are the total instance counts in the positive and negative class, respectively. The function then applies the scales as multiplicative factors to the counts from the corresponding class: perfcurve multiplies counts from the positive class by scale (P) and counts from the negative class by scale(N). Consider, for example, computation of positive predictive value, PPV = TP/ (TP+FP). TP counts come from the positive class and FP counts come from the negative class. Therefore, you need to scale TP by scale (P) and FP by scale(N), and the modified formula for PPV with prior probabilities taken into account is now:

$$
P P V=\frac{\operatorname{scale}(P) * T P}{\operatorname{scale}(P) * T P+\operatorname{scale}(N) * F P}
$$

If all scores in the data are above a certain threshold, perfcurve classifies all instances as 'positive'. This means that TP is the total number of instances in the positive class and FP is the total number of instances in the negative class. In this case, PPV is simply given by the prior:

$$
P P V=\frac{\operatorname{prior}(P)}{\operatorname{prior}(P)+\operatorname{prior}(N)}
$$

The perfcurve function returns two vectors, X and Y , of performance measures. Each measure is some function of confusion, cost, and scale values. You can request specific measures by name or provide a function handle to compute a custom measure. The function you provide should take confusion, cost, and scale as its three inputs and return a vector of output values.

The criterion for X must be a monotone function of the positive classification count, or equivalently, threshold for the supplied scores. If perfcurve cannot perform a one-to-one mapping between values of the X criterion and score thresholds, it exits with an error message.

By default, perfcurve computes values of the $X$ and $Y$ criteria for all possible score thresholds. Alternatively, it can compute a reduced number of specific $X$ values supplied as an input argument. In either case, for $M$ requested values, perfcurve computes $M+1$ values for $X$ and $Y$. The first value out of these $M+1$ values is special. perfcurve computes it by setting the TP instance count
to zero and setting TN to the total count in the negative class. This value corresponds to the 'reject all' threshold. On a standard ROC curve, this translates into an extra point placed at $(0,0)$.

If there are NaN values among input scores, perfcurve can process them in either of two ways:

- It can discard rows with NaN scores.
- It can add them to false classification counts in the respective class.

That is, for any threshold, instances with NaN scores from the positive class are counted as false negative (FN), and instances with NaN scores from the negative class are counted as false positive (FP). In this case, the first value of $X$ or $Y$ is computed by setting TP to zero and setting TN to the total count minus the NaN count in the negative class. For illustration, consider an example with two rows in the positive and two rows in the negative class, each pair having a NaN score:

Class
Negative
Negative
Positive
Positive

## Score

0.2 NaN
0.7

NaN

If you discard rows with NaN scores, then as the score cutoff varies, perfcurve computes performance measures as in the following table. For example, a cutoff of 0.5 corresponds to the middle row where rows 1 and 3 are classified correctly, and rows 2 and 4 are omitted.

| TP | FN | FP | TN |
| :--- | :--- | :--- | :--- |
| 0 | 1 | 0 | 1 |
| 1 | 0 | 0 | 1 |
| 1 | 0 | 1 | 0 |

If you add rows with NaN scores to the false category in their respective classes, perfcurve computes performance measures as in the following table. For example, a cutoff of 0.5 corresponds to the middle row where now rows

2 and 4 are counted as incorrectly classified. Notice that only the FN and FP columns differ between these two tables.

| TP | FN | FP | TN |
| :--- | :--- | :--- | :--- |
| 0 | 2 | 1 | 1 |
| 1 | 1 | 1 | 1 |
| 1 | 1 | 2 | 0 |

For data with three or more classes, perfcurve takes one positive class and a list of negative classes for input. The function computes the $X$ and $Y$ values using counts in the positive class to estimate TP and FN, and using counts in all negative classes to estimate TN and FP. perfcurve can optionally compute $Y$ values for each negative class separately and, in addition to $Y$, return a matrix of size $M$-by- $C$, where $M$ is the number of elements in $X$ or $Y$ and $C$ is the number of negative classes. You can use this functionality to monitor components of the negative class contribution. For example, you can plot TP counts on the X -axis and FP counts on the Y -axis. In this case, the returned matrix shows how the FP component is split across negative classes.

You can also use perfcurve to estimate confidence intervals. perfcurve computes confidence bounds using either cross-validation or bootstrap. If you supply cell arrays for labels and scores, perfcurve uses cross-validation and treats elements in the cell arrays as cross-validation folds. If you set input parameter NBoot to a positive integer, perfcurve generates nboot bootstrap replicas to compute pointwise confidence bounds.
perfcurve estimates the confidence bounds using one of two methods:

- Vertical averaging (VA) - estimate confidence bounds on $Y$ and $T$ at fixed values of $X$. Use the XVals input parameter to use this method for computing confidence bounds.
- Threshold averaging (TA) - estimate confidence bounds for $X$ and $Y$ at fixed thresholds for the positive class score. Use the TVals input parameter to use this method for computing confidence bounds.

To use observation weights instead of observation counts, you can use the 'Weights' parameter in your call to perfcurve. When you use this parameter, to compute $\mathrm{X}, \mathrm{Y}$ and T or to compute confidence bounds by cross-validation, perfcurve uses your supplied observation weights instead of
observation counts. To compute confidence bounds by bootstrap, perfcurve samples $N$ out of $N$ with replacement using your weights as multinomial sampling probabilities.

## Nonparametric Supervised Learning

- "Supervised Learning (Machine Learning) Workflow and Algorithms" on page 13-2
- "Classification Using Nearest Neighbors" on page 13-9
- "Classification Trees and Regression Trees" on page 13-27
- "Ensemble Methods" on page 13-51
- "Bibliography" on page 13-130


# Supervised Learning (Machine Learning) Workflow and Algorithms 

In this section...<br>"Steps in Supervised Learning (Machine Learning)" on page 13-2<br>"Characteristics of Algorithms" on page 13-7

## Steps in Supervised Learning (Machine Learning)

Supervised learning (machine learning) takes a known set of input data and known responses to the data, and seeks to build a predictor model that generates reasonable predictions for the response to new data.


2


Suppose you want to predict if someone will have a heart attack within a year. You have a set of data on previous people, including age, weight, height, blood pressure, etc. You know if the previous people had heart attacks within a year of their data measurements. So the problem is combining all the existing data into a model that can predict whether a new person will have a heart attack within a year.

Supervised learning splits into two broad categories:

- Classification for responses that can have just a few known values, such as 'true' or 'false'. Classification algorithms apply to nominal, not ordinal response values.
- Regression for responses that are a real number, such as miles per gallon for a particular car.

You can have trouble deciding whether you have a classification problem or a regression problem. In that case, create a regression model first, because they are often more computationally efficient.

While there are many Statistics Toolbox algorithms for supervised learning, most use the same basic workflow for obtaining a predictor model. (Detailed instruction on the steps for ensemble learning is in "Framework for Ensemble Learning" on page 13-51.) The steps for supervised learning are:

1 "Prepare Data" on page 13-3
2 "Choose an Algorithm" on page 13-4
3 "Fit a Model" on page 13-4
4 "Choose a Validation Method" on page 13-5
5 "Examine Fit and Update Until Satisfied" on page 13-6
6 "Use Fitted Model for Predictions" on page 13-6

## Prepare Data

All supervised learning methods start with an input data matrix, usually called $X$ here. Each row of $X$ represents one observation. Each column of X represents one variable, or predictor. Represent missing entries with NaN values in X. Statistics Toolbox supervised learning algorithms can handle NaN values, either by ignoring them or by ignoring any row with a NaN value.

You can use various data types for response data $Y$. Each element in $Y$ represents the response to the corresponding row of $X$. Observations with missing $Y$ data are ignored.

- For regression, Y must be a numeric vector with the same number of elements as the number of rows of $X$.
- For classification, Y can be any of these data types. This table also contains the method of including missing entries.

| Data Type | Missing Entry |
| :--- | :--- |
| Numeric vector | NaN |
| Categorical vector | <undefined> |
| Character array | Row of spaces |
| Cell array of strings | '' |
| Logical vector | (Cannot represent) |

## Choose an Algorithm

There are tradeoffs between several characteristics of algorithms, such as:

- Speed of training
- Memory usage
- Predictive accuracy on new data
- Transparency or interpretability, meaning how easily you can understand the reasons an algorithm makes its predictions

Details of the algorithms appear in "Characteristics of Algorithms" on page 13-7. More detail about ensemble algorithms is in "Choose an Applicable Ensemble Method" on page 13-54.

## Fit a Model

The fitting function you use depends on the algorithm you choose.

| Algorithm | Fitting Function |
| :--- | :--- |
| Classification Trees | ClassificationTree.fit |
| Regression Trees | RegressionTree.fit |
| Discriminant Analysis <br> (classification) | ClassificationDiscriminant.fit |
| Naive Bayes (classification) | NaiveBayes.fit |


| Algorithm | Fitting Function |
| :--- | :--- |
| Classification or Regression <br> Ensembles | fitensemble |
| Classification or Regression <br> Ensembles in Parallel | TreeBagger |

## Choose a Validation Method

The three main methods to examine the accuracy of the resulting fitted model are:

- Examine the resubstitution error. For examples, see:
- "Example: Resubstitution Error of a Classification Tree" on page 13-34
- "Example: Cross Validating a Regression Tree" on page 13-35
- "Example: Test Ensemble Quality" on page 13-60
- "Example: Resubstitution Error of a Discriminant Analysis Classifier" on page 12-17
- Examine the cross-validation error. For examples, see:
- "Example: Cross Validating a Regression Tree" on page 13-35
- "Example: Test Ensemble Quality" on page 13-60
- "Example: Classification with Many Categorical Levels" on page 13-72
- "Example: Cross Validating a Discriminant Analysis Classifier" on page 12-18
- Examine the out-of-bag error for bagged decision trees. For examples, see:
- "Example: Test Ensemble Quality" on page 13-60
- "Workflow Example: Regression of Insurance Risk Rating for Car Imports with TreeBagger" on page 13-97
- "Workflow Example: Classifying Radar Returns for Ionosphere Data with TreeBagger" on page 13-106


## Examine Fit and Update Until Satisfied

After validating the model, you might want to change it for better accuracy, better speed, or to use less memory.

- Change fitting parameters to try to get a more accurate model. For examples, see:
- "Example: Tuning RobustBoost" on page 13-92
- "Example: Unequal Classification Costs" on page 13-67
- "Improving a Discriminant Analysis Classifier" on page 12-14
- Change fitting parameters to try to get a smaller model. This sometimes gives a model with more accuracy. For examples, see:
- "Example: Selecting Appropriate Tree Depth" on page 13-36
- "Example: Pruning a Classification Tree" on page 13-39
- "Example: Surrogate Splits" on page 13-77
- "Example: Regularizing a Regression Ensemble" on page 13-82
- "Workflow Example: Regression of Insurance Risk Rating for Car Imports with TreeBagger" on page 13-97
- "Workflow Example: Classifying Radar Returns for Ionosphere Data with TreeBagger" on page 13-106
- Try a different algorithm. For applicable choices, see:
- "Characteristics of Algorithms" on page 13-7
- "Choose an Applicable Ensemble Method" on page 13-54

When satisfied with a model of some types, you can trim it using the appropriate compact method (compact for classification trees, compact for classification ensembles, compact for regression trees, compact for regression ensembles, compact for discriminant analysis). compact removes training data and pruning information, so the model uses less memory.

## Use Fitted Model for Predictions

To predict classification or regression response for most fitted models, use the predict method:

```
Ypredicted = predict(obj,Xnew)
```

- obj is the fitted model object.
- Xnew is the new input data.
- Ypredicted is the predicted response, either classification or regression.

For classregtree, use the eval method instead of predict.

## Characteristics of Algorithms

This table shows typical characteristics of the various supervised learning algorithms. The characteristics in any particular case can vary from the listed ones. Use the table as a guide for your initial choice of algorithms, but be aware that the table can be inaccurate for some problems.

## Characteristics of Supervised Learning Algorithms

| Algorithm | Predictive <br> Accuracy | Fitting <br> Speed | Prediction <br> Speed | Memory <br> Usage | Easy to <br> Interpret | Handles <br> Categorical <br> Predictors |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Trees | Low | Fast | Fast | Low | Yes | Yes |
| Boosted Trees | High | Medium | Medium | Medium | No | Yes |
| Bagged Trees | High | Slow | Slow | High | No | Yes |
| SVM in <br> Bioinformatics <br> Toolbox | High | Medium | $*$ | $*$ | $*$ | No |
| Naive Bayes | Low | $* *$ | $* *$ | $* *$ | Yes | Yes |
| Nearest <br> Neighbor | $* * *$ | Fast*** | Medium | High | No | Yes*** |
| Discriminant <br> Analysis | $* * * *$ | Fast | Fast | Low | Yes | No |

*     - SVM prediction speed and memory usage are good if there are few support vectors, but can be poor if there are many support vectors. When you use a kernel function, it can be difficult to interpret how SVM classifies data,
though the default linear scheme is easy to interpret. SVM is available if you have a Bioinformatics Toolbox license.
** - Naive Bayes speed and memory usage are good for simple distributions, but can be poor for kernel distributions and large data sets.
*** - Nearest Neighbor usually has good predictions in low dimensions, but can have poor predictions in high dimensions. For linear search, Nearest Neighbor does not perform any fitting. For $k$ d-trees, Nearest Neighbor does perform fitting. Nearest Neighbor can have either continuous or categorical predictors, but not both.
**** - Discriminant Analysis is accurate when the modeling assumptions are satisfied (multivariate normal by class). Otherwise, the predictive accuracy varies.


# Classification Using Nearest Neighbors 

In this section...<br>"Pairwise Distance" on page 13-9<br>" $k$-Nearest Neighbor Search and Radius Search" on page 13-12

## Pairwise Distance

Categorizing query points based on their distance to points in a training dataset can be a simple yet effective way of classifying new points. You can use various metrics to determine the distance, described next. Use pdist2 to find the distance between a sets of data and query points.

## Distance Metrics

Given an $m x$-by- $n$ data matrix X , which is treated as $m x$ (1-by- $n$ ) row vectors $\mathrm{x}_{1}, \mathrm{x}_{2}, \ldots, \mathrm{x}_{m \mathrm{x}}$, and $m y$-by- $n$ data matrix Y , which is treated as $m y$ (1-by- $n$ ) row vectors $\mathrm{y}_{1}, \mathrm{y}_{2}, \ldots, \mathrm{y}_{m \mathrm{y}}$, the various distances between the vector $\mathrm{x}_{s}$ and $\mathrm{y}_{t}$ are defined as follows:

- Euclidean distance

$$
d_{s t}^{2}=\left(x_{s}-y_{t}\right)\left(x_{s}-y_{t}\right)^{\prime}
$$

The Euclidean distance is a special case of the Minkowski metric, where p $=2$.

- Standardized Euclidean distance

$$
d_{s t}^{2}=\left(x_{s}-y_{t}\right) V^{-1}\left(x_{s}-y_{t}\right)^{\prime}
$$

where V is the $n$-by- $n$ diagonal matrix whose $j$ th diagonal element is $\mathrm{S}(j)^{2}$, where $S$ is the vector containing the inverse weights.

- Mahalanobis distance

$$
d_{s t}^{2}=\left(x_{s}-y_{t}\right) C^{-1}\left(x_{s}-y_{t}\right)^{\prime}
$$

where $C$ is the covariance matrix.

- City block metric

$$
d_{s t}=\sum_{j=1}^{n}\left|x_{s j}-y_{t j}\right|
$$

The city block distance is a special case of the Minkowski metric, where p $=1$.

- Minkowski metric

$$
d_{s t}=\sqrt[p]{\sum_{j=1}^{n}\left|x_{s j}-y_{t j}\right|^{p}}
$$

For the special case of $p=1$, the Minkowski metric gives the city block metric, for the special case of $p=2$, the Minkowski metric gives the Euclidean distance, and for the special case of $p=\infty$, the Minkowski metric gives the Chebychev distance.

- Chebychev distance

$$
d_{s t}=\max _{j}\left\{\left|x_{s j}-y_{t j}\right|\right\}
$$

The Chebychev distance is a special case of the Minkowski metric, where $p$ $=\infty$.

- Cosine distance

$$
d_{s t}=\left(1-\frac{x_{s} y_{t}^{\prime}}{\sqrt{\left(x_{s} x_{s}^{\prime}\right)\left(y_{t} y_{t}^{\prime}\right)}}\right)
$$

- Correlation distance

$$
d_{s t}=1-\frac{\left(x_{s}-\bar{x}_{s}\right)\left(y_{t}-\bar{y}_{t}\right)^{\prime}}{\sqrt{\left(x_{s}-\bar{x}_{s}\right)\left(x_{s}-\bar{x}_{s}\right)^{\prime}} \sqrt{\left(y_{t}-\bar{y}_{t}\right)\left(y_{t}-\bar{y}_{t}\right)^{\prime}}}
$$

where

$$
\begin{aligned}
& \bar{x}_{s}=\frac{1}{n} \sum_{j} x_{s j} \\
& \mathrm{~d} \\
& \bar{y}_{t}=\frac{1}{n} \sum_{j} y_{t j}
\end{aligned}
$$

- Hamming distance

$$
d_{s t}=\left(\#\left(x_{s j} \neq y_{t j}\right) / n\right)
$$

- Jaccard distance

$$
d_{s t}=\frac{\#\left[\left(x_{s j} \neq y_{t j}\right) \cap\left(\left(x_{s j} \neq 0\right) \cup\left(y_{t j} \neq 0\right)\right)\right]}{\#\left[\left(x_{s j} \neq 0\right) \cup\left(y_{t j} \neq 0\right)\right]}
$$

- Spearman distance

$$
d_{s t}=1-\frac{\left(r_{s}-\bar{r}_{s}\right)\left(r_{t}-\bar{r}_{t}\right)^{\prime}}{\sqrt{\left(r_{s}-\bar{r}_{s}\right)\left(r_{s}-\bar{r}_{s}\right)^{\prime}} \sqrt{\left(r_{t}-\bar{r}_{t}\right)\left(r_{t}-\bar{r}_{t}\right)^{\prime}}}
$$

where

- $r_{s j}$ is the rank of $x_{s j}$ taken over $x_{1 j}, x_{2 j}, \ldots x_{m x, j}$, as computed by tiedrank.
- $r_{t j}$ is the rank of $y_{t j}$ taken over $y_{1 j}, y_{2 j}, \ldots y_{m y, j}$, as computed by tiedrank.
- $r_{s}$ and $r_{t}$ are the coordinate-wise rank vectors of $x_{s}$ and $y_{t}$, i.e., $r_{s}=\left(r_{s 1}\right.$, $\left.r_{s 2}, \ldots r_{s n}\right)$ and $r_{t}=\left(r_{t 1}, r_{t 2}, \ldots r_{t n}\right)$.
- $\bar{r}_{s}=\frac{1}{n} \sum_{j} r_{s j}=\frac{(n+1)}{2}$.
- $\bar{r}_{t}=\frac{1}{n} \sum_{j} r_{t j}=\frac{(n+1)}{2}$.


## k-Nearest Neighbor Search and Radius Search

Given a set $X$ of $n$ points and a distance function, $k$-nearest neighbor ( $k \mathrm{NN}$ ) search lets you find the $k$ closest points in $X$ to a query point or set of points Y . The $k \mathrm{NN}$ search technique and $k \mathrm{NN}$-based algorithms are widely used as benchmark learning rules. The relative simplicity of the $k \mathrm{NN}$ search technique makes it easy to compare the results from other classification techniques to $k N N$ results. The technique has been used in various areas such as:

- bioinformatics
- image processing and data compression
- document retrieval
- computer vision
- multimedia database
- marketing data analysis

You can use $k \mathrm{NN}$ search for other machine learning algorithms, such as:

- $k N N$ classification
- local weighted regression
- missing data imputation and interpolation
- density estimation

You can also use $k N N$ search with many distance-based learning functions, such as K-means clustering.

In contrast, for a positive real value $r$, rangesearch finds all points in $X$ that are within a distance $r$ of each point in $Y$. This fixed-radius search is closely related to $k \mathrm{NN}$ search, as it supports the same distance metrics and search classes, and uses the same search algorithms.

## $\boldsymbol{k}$-Nearest Neighbor Search Using Exhaustive Search

When your input data meets any of the following criteria, knnsearch uses the exhaustive search method by default to find the $k$-nearest neighbors:

- The number of columns of $X$ is more than 10 .
- X is sparse.
- The distance measure is either:
- 'seuclidean'
- 'mahalanobis'
- 'cosine'
- 'correlation'
- 'spearman'
- 'hamming'
- 'jaccard'
- A custom distance function
knnsearch also uses the exhaustive search method if your search object is an ExhaustiveSearcher object. The exhaustive search method finds the distance from each query point to every point in $X$, ranks them in ascending order, and returns the $k$ points with the smallest distances. For example, this diagram shows the $k=3$ nearest neighbors.



## k-Nearest Neighbor Search Using a kd-Tree

When your input data meets all of the following criteria, knnsearch creates a $k$ d-tree by default to find the $k$-nearest neighbors:

- The number of columns of $X$ is less than 10 .
- X is not sparse.
- The distance measure is either:
- 'euclidean' (default)
- 'cityblock'
- 'minkowski'
- 'chebychev'
knnsearch also uses a $k$ d-tree if your search object is a KDTreeSearcher object.
$k$ d-trees divide your data into nodes with at most BucketSize (default is 50) points per node, based on coordinates (as opposed to categories). The following diagrams illustrate this concept using patch objects to color code the different "buckets."


When you want to find the $k$-nearest neighbors to a given query point, knnsearch does the following:

1 Determines the node to which the query point belongs. In the following example, the query point $(32,90)$ belongs to Node 4.

2 Finds the closest $k$ points within that node and its distance to the query point. In the following example, the points in red circles are equidistant from the query point, and are the closest points to the query point within Node 4.

3 Chooses all other nodes having any area that is within the same distance, in any direction, from the query point to the $k$ th closest point. In this example, only Node 3 overlaps the solid black circle centered at the query point with radius equal to the distance to the closest points within Node 4.

4 Searches nodes within that range for any points closer to the query point. In the following example, the point in a red square is slightly closer to the query point than those within Node 4.


| $\times$ | Query |
| :---: | :---: |
| $\square$ | Neighbors |
|  | Cuts |
| $\cdots \cdots \cdots$ | Neighborhood |
| $\square$ | Node 1 |
| $\square$ | Node 2 |
| $\square$ | Node 3 |
| $\square$ | Node 4 |
| $\square$ | Node 5 |
| $\square$ | Node 6 |
| $\square$ | Node 7 |
| $\square$ | Node 8 |

Using a $k$ d-tree for large datasets with fewer than 10 dimensions (columns) can be much more efficient than using the exhaustive search method, as knnsearch needs to calculate only a subset of the distances. To maximize the efficiency of $k$ d-trees, use a KDTreeSearcher object.

## What Are Search Objects?

Basically, objects are a convenient way of storing information. Classes of related objects (for example, all search objects) have the same properties with values and types relevant to a specified search method. In addition to storing information within objects, you can perform certain actions (called methods) on objects.

All search objects have a knnsearch method specific to that class. This lets you efficiently perform a $k$-nearest neighbors search on your object for that specific object type. In addition, there is a generic knnsearch function that searches without creating or using an object.

To determine which type of object and search method is best for your data, consider the following:

- Does your data have many columns, say more than 10 ? The ExhaustiveSearcher object may perform better.
- Is your data sparse? Use the ExhaustiveSearcher object.
- Do you want to use one of these distance measures to find the nearest neighbors? Use the ExhaustiveSearcher object.
- 'seuclidean'
- 'mahalanobis'
- 'cosine'
- 'correlation'
- 'spearman'
- 'hamming'
- 'jaccard'
- A custom distance function
- Is your dataset huge (but with fewer than 10 columns)? Use the KDTreeSearcher object.
- Are you searching for the nearest neighbors for a large number of query points? Use the KDTreeSearcher object.

For more detailed information on object-oriented programming in MATLAB, see Object-Oriented Programming.

## Example: Classifying Query Data Using knnsearch

1 Classify a new point based on the last two columns of the Fisher iris data. Using only the last two columns makes it easier to plot:

```
load fisheriris
x = meas(:,3:4);
gscatter(x(:,1),x(:,2),species)
set(legend,'location','best')
```



2 Plot the new point:
newpoint = [5 1.45];
line(newpoint(1), newpoint(2),'marker','x','color','k',... 'markersize',10,'linewidth' , 2)


3 Find the 10 sample points closest to the new point:

```
[n,d] = knnsearch(x,newpoint,'k',10)
line(x(n,1),x(n,2),'color',[.5 .5 .5],'marker','o',...
    'linestyle','none','markersize',10)
```



4 It appears that knnsearch has found only the nearest eight neighbors. In fact, this particular dataset contains duplicate values:

```
x(n,:)
ans =
    5.0000 1.5000
    4.9000 1.5000
    4.9000 1.5000
    5.1000 1.5000
    5.1000 1.6000
    4.8000 1.4000
    5.0000 1.7000
    4.7000 1.4000
    4.7000 1.4000
```

$$
4.7000 \quad 1.5000
$$

5 To make duplicate values visible on the plot, use the following code:

```
% jitter to make repeated points visible
xj = x + .05*(rand(150,2)-.5);
gscatter(xj(:,1),xj(:,2),species)
```

The jittered points do not affect any analysis of the data, only the visualization. This example does not jitter the points.

6 Make the axes equal so the calculated distances correspond to the apparent distances on the plot axis equal and zoom in to see the neighbors better:


7 Find the species of the 10 neighbors:

```
tabulate(species(n))
\begin{tabular}{rrr} 
Value & Count & Percent \\
virginica & 2 & \(20.00 \%\) \\
versicolor & 8 & \(80.00 \%\)
\end{tabular}
```

Using a rule based on the majority vote of the 10 nearest neighbors, you can classify this new point as a versicolor.

8 Visually identify the neighbors by drawing a circle around the group of them:

```
% Define the center and diameter of a circle, based on the
% location of the new point:
ctr = newpoint - d(end);
diameter = 2*d(end);
% Draw a circle around the 10 nearest neighbors:
h = rectangle('position',[ctr,diameter,diameter],...
    'curvature',[1 1]);
set(h,'linestyle',':')
```



9 Using the same dataset, find the 10 nearest neighbors to three new points:

```
figure
newpoint2 = [5 1.45;6 2;2.75 .75];
gscatter(x(:,1),x(:,2),species)
legend('location','best')
[n2,d2] = knnsearch(x,newpoint2,'k',10);
line(x(n2,1),x(n2,2),'color',[.5 .5 .5],'marker','o',...
    'linestyle','none','markersize',10)
line(newpoint2(:,1),newpoint2(:,2),'marker','x','color','k',...
    'markersize',10,'linewidth',2,'linestyle','none')
```



10 Find the species of the 10 nearest neighbors for each new point:

```
tabulate(species(n2(1,:)))
            Value Count Percent
            virginica 2 20.00%
                versicolor 8 80.00%
tabulate(species(n2(2,:)))
            Value Count Percent
    virginica 10 100.00%
tabulate(species(n2(3,:)))
            Value Count Percent
        versicolor 7 70.00%
        setosa 3 30.00%
```

For further examples using knnsearch methods and function, see the individual reference pages.

## Classification Trees and Regression Trees

In this section...<br>"What Are Classification Trees and Regression Trees?" on page 13-27<br>"Example: Creating a Classification Tree" on page 13-28<br>"Example: Creating a Regression Tree" on page 13-28<br>"Viewing a Tree" on page 13-29<br>"How the Fit Methods Create Trees" on page 13-31<br>"Predicting Responses With Classification Trees and Regression Trees" on page 13-33<br>"Improving Classification Trees and Regression Trees" on page 13-34<br>"Alternative: classregtree" on page 13-43

## What Are Classification Trees and Regression Trees?

Classification trees and regression trees predict responses to data. To predict a response, follow the decisions in the tree from the root (beginning) node down to a leaf node. The leaf node contains the response. Classification trees give responses that are nominal, such as 'true' or 'false'. Regression trees give numeric responses.

Statistics Toolbox trees are binary. Each step in a prediction involves checking the value of one predictor (variable). For example, here is a simple classification tree:


This tree predicts classifications based on two predictors, x 1 and x 2 . To predict, start at the top node, represented by a triangle ( $\Delta$ ). The first decision is whether $x 1$ is smaller than 0.5 . If so, follow the left branch, and see that the tree classifies the data as type 0 .

If, however, $x 1$ exceeds 0.5 , then follow the right branch to the lower-right triangle node. Here the tree asks if $x 2$ is smaller than 0.5 . If so, then follow the left branch to see that the tree classifies the data as type 0 . If not, then follow the right branch to see that the that the tree classifies the data as type 1 .

To learn how to prepare your data for classification or regression using decision trees, see "Steps in Supervised Learning (Machine Learning)" on page 13-2.

## Example: Creating a Classification Tree

To create a classification tree for the ionosphere data:

```
load ionosphere % contains X and Y variables
ctree = ClassificationTree.fit(X,Y)
ctree =
ClassificationTree:
            PredictorNames: {1x34 cell}
    CategoricalPredictors: []
            ResponseName: 'Y'
            ClassNames: {'b' 'g'}
            ScoreTransform: 'none'
            NObservations: 351
```


## Example: Creating a Regression Tree

To create a regression tree for the carsmall data based on the Horsepower and Weight vectors for data, and MPG vector for response:

```
load carsmall % contains Horsepower, Weight, MPG
X = [Horsepower Weight];
rtree = RegressionTree.fit(X,MPG)
```

```
rtree =
RegressionTree:
    PredictorNames: {'x1' 'x2'}
CategoricalPredictors: []
ResponseName: 'Y'
ResponseTransform: 'none'
NObservations: 94
```


## Viewing a Tree

There are two ways to view a tree:

- view(tree) returns a text description of the tree.
- view(tree,'mode', 'graph') returns a graphic description of the tree.
"Example: Creating a Classification Tree" on page 13-28 has the following two views:

```
load fisheriris
ctree = ClassificationTree.fit(meas,species);
view(ctree)
Decision tree for classification
1 if x3<2.45 then node 2 elseif x3>=2.45 then node 3 else setosa
2 class = setosa
3 if x4<1.75 then node 4 elseif x4>=1.75 then node 5 else versicolor
4 if x3<4.95 then node 6 elseif x3>=4.95 then node 7 else versicolor
5 class = virginica
6 if x4<1.65 then node 8 elseif x4>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
view(ctree,'mode','graph')
```



Similarly, "Example: Creating a Regression Tree" on page 13-28 has the following two views:

```
load carsmall % contains Horsepower, Weight, MPG
X = [Horsepower Weight];
rtree = RegressionTree.fit(X,MPG,'MinParent',30);
view(rtree)
Decision tree for regression
1 if x2<3085.5 then node 2 elseif x2>=3085.5 then node 3 else 23.7181
2 if x1<89 then node 4 elseif x1>=89 then node 5 else 28.7931
3 if x1<115 then node 6 elseif x1>=115 then node 7 else 15.5417
4 if x2<2162 then node 8 elseif x2>=2162 then node 9 else 30.9375
5 fit = 24.0882
6 fit = 19.625
7 fit = 14.375
```

```
8 fit = 33.3056
9 fit = 29
```

view(rtree,'mode', 'graph')


## How the Fit Methods Create Trees

The ClassificationTree.fit and RegressionTree.fit methods perform the following steps to create decision trees:

1 Start with all input data, and examine all possible binary splits on every predictor.

2 Select a split with best optimization criterion.

- If the split leads to a child node having too few observations (less than the MinLeaf parameter), select a split with the best optimization criterion subject to the MinLeaf constraint.

3 Impose the split.
4 Repeat recursively for the two child nodes.
The explanation requires two more items: description of the optimization criterion, and stopping rule.

Stopping rule: Stop splitting when any of the following hold:

- The node is pure.
- For classification, a node is pure if it contains only observations of one class.
- For regression, a node is pure if the mean squared error (MSE) for the observed response in this node drops below the MSE for the observed response in the entire data multiplied by the tolerance on quadratic error per node (qetoler parameter).
- There are fewer than MinParent observations in this node.
- Any split imposed on this node would produce children with fewer than MinLeaf observations.


## Optimization criterion:

- Regression: mean-squared error (MSE). Choose a split to minimize the MSE of predictions compared to the training data.
- Classification: One of three measures, depending on the setting of the SplitCriterion name-value pair:
- 'gdi' (Gini's diversity index, the default)
- 'twoing'
- 'deviance'

For details, see ClassificationTree "Definitions" on page 20-210.

For a continuous predictor, a tree can split halfway between any two adjacent unique values found for this predictor. For a categorical predictor with $L$ levels, a classification tree needs to consider $2^{L-1}-1$ splits. To obtain this formula, observe that you can assign $L$ distinct values to the left and right nodes in $2^{L}$ ways. Two out of these $2^{L}$ configurations would leave either left or right node empty, and therefore should be discarded. Now divide by 2 because left and right can be swapped. A classification tree can thus process only categorical predictors with a moderate number of levels. A regression tree employs a computational shortcut: it sorts the levels by the observed mean response, and considers only the $L-1$ splits between the sorted levels.

## Predicting Responses With Classification Trees and Regression Trees

After creating a tree, you can easily predict responses for new data. Suppose Xnew is new data that has the same number of columns as the original data x . To predict the classification or regression based on the tree and the new data, enter

```
Ynew = predict(tree,Xnew);
```

For each row of data in Xnew, predict runs through the decisions in tree and gives the resulting prediction in the corresponding element of Ynew. For more information for classification, see the classification predict reference page; for regression, see the regression predict reference page.

For example, to find the predicted classification of a point at the mean of the ionosphere data:

```
load ionosphere % contains X and Y variables
ctree = ClassificationTree.fit(X,Y);
Ynew = predict(ctree,mean(X))
Ynew =
    'g'
```

To find the predicted MPG of a point at the mean of the carsmall data:

```
load carsmall % contains Horsepower, Weight, MPG
X = [Horsepower Weight];
rtree = RegressionTree.fit(X,MPG);
```

```
Ynew = predict(rtree,mean(X))
Ynew =
```

    28.7931
    
## Improving Classification Trees and Regression Trees

You can tune trees by setting name-value pairs in ClassificationTree.fit and RegressionTree.fit. The remainder of this section describes how to determine the quality of a tree, how to decide which name-value pairs to set, and how to control the size of a tree:

- "Examining Resubstitution Error" on page 13-34
- "Cross Validation" on page 13-35
- "Control Depth or "Leafiness'"" on page 13-35
- "Pruning" on page 13-39


## Examining Resubstitution Error

Resubstitution error is the difference between the response training data and the predictions the tree makes of the response based on the input training data. If the resubstitution error is high, you cannot expect the predictions of the tree to be good. However, having low resubstitution error does not guarantee good predictions for new data. Resubstitution error is often an overly optimistic estimate of the predictive error on new data.

Example: Resubstitution Error of a Classification Tree. Examine the resubstitution error of a default classification tree for the Fisher iris data:

```
load fisheriris
ctree = ClassificationTree.fit(meas,species);
resuberror = resubLoss(ctree)
resuberror =
    0.0200
```

The tree classifies nearly all the Fisher iris data correctly.

## Cross Validation

To get a better sense of the predictive accuracy of your tree for new data, cross validate the tree. By default, cross validation splits the training data into 10 parts at random. It trains 10 new trees, each one on nine parts of the data. It then examines the predictive accuracy of each new tree on the data not included in training that tree. This method gives a good estimate of the predictive accuracy of the resulting tree, since it tests the new trees on new data.

Example: Cross Validating a Regression Tree. Examine the resubstitution and cross-validation accuracy of a regression tree for predicting mileage based on the carsmall data:

```
load carsmall
X = [Acceleration Displacement Horsepower Weight];
rtree = RegressionTree.fit(X,MPG);
resuberror = resubLoss(rtree)
resuberror =
    4 . 7 1 8 8
```

The resubstitution loss for a regression tree is the mean-squared error. The resulting value indicates that a typical predictive error for the tree is about the square root of 4.7 , or a bit over 2 .

Now calculate the error by cross validating the tree:

```
cvrtree = crossval(rtree);
cvloss = kfoldLoss(cvrtree)
cvloss =
    23.4808
```

The cross-validated loss is almost 25, meaning a typical predictive error for the tree on new data is about 5 . This demonstrates that cross-validated loss is usually higher than simple resubstitution loss.

## Control Depth or "Leafiness"

When you grow a decision tree, consider its simplicity and predictive power. A deep tree with many leaves is usually highly accurate on the training data.

However, the tree is not guaranteed to show a comparable accuracy on an independent test set. A leafy tree tends to overtrain, and its test accuracy is often far less than its training (resubstitution) accuracy. In contrast, a shallow tree does not attain high training accuracy. But a shallow tree can be more robust - its training accuracy could be close to that of a representative test set. Also, a shallow tree is easy to interpret.

If you do not have enough data for training and test, estimate tree accuracy by cross validation.

For an alternative method of controlling the tree depth, see "Pruning" on page 13-39.

Example: Selecting Appropriate Tree Depth. This example shows how to control the depth of a decision tree, and how to choose an appropriate depth.

1 Load the ionosphere data:

```
load ionosphere
```

2 Generate minimum leaf occupancies for classification trees from 10 to 100, spaced exponentially apart:

```
leafs = logspace(1,2,10);
```

3 Create cross validated classification trees for the ionosphere data with minimum leaf occupancies from leafs:

```
N = numel(leafs);
err = zeros(N,1);
for n=1:N
    t = ClassificationTree.fit(X,Y,'crossval','on',...
        'minleaf',leafs(n));
    err(n) = kfoldLoss(t);
end
plot(leafs,err);
xlabel('Min Leaf Size');
ylabel('cross-validated error');
```



The best leaf size is between about 20 and 50 observations per leaf.
4 Compare the near-optimal tree with at least 40 observations per leaf with the default tree, which uses 10 observations per parent node and 1 observation per leaf.

```
DefaultTree = ClassificationTree.fit(X,Y);
view(DefaultTree,'mode','graph')
```



OptimalTree = ClassificationTree.fit(X,Y,'minleaf', 40); view(OptimalTree,'mode', 'graph')


```
resubOpt = resubLoss(OptimalTree);
lossOpt = kfoldLoss(crossval(OptimalTree));
resubDefault = resubLoss(DefaultTree);
lossDefault = kfoldLoss(crossval(DefaultTree));
resubOpt,resubDefault,lossOpt,lossDefault
resubOpt =
    0.0883
```

```
resubDefault =
    0.0114
lossOpt =
    0.1054
lossDefault =
    0.1026
```

The near-optimal tree is much smaller and gives a much higher resubstitution error. Yet it gives similar accuracy for cross-validated data.

## Pruning

Pruning optimizes tree depth (leafiness) is by merging leaves on the same tree branch. "Control Depth or "Leafiness"" on page 13-35 describes one method for selecting the optimal depth for a tree. Unlike in that section, you do not need to grow a new tree for every node size. Instead, grow a deep tree, and prune it to the level you choose.

Prune a tree at the command line using the prune method (classification) or prune method (regression). Alternatively, prune a tree interactively with the tree viewer:

```
view(tree,'mode','graph')
```

To prune a tree, the tree must contain a pruning sequence. By default, both ClassificationTree.fit and RegressionTree.fit calculate a pruning sequence for a tree during construction. If you construct a tree with the 'Prune ' name-value pair set to 'off', or if you prune a tree to a smaller level, the tree does not contain the full pruning sequence. Generate the full pruning sequence with the prune method (classification) or prune method (regression).

Example: Pruning a Classification Tree. This example creates a classification tree for the ionosphere data, and prunes it to a good level.

1 Load the ionosphere data:
load ionosphere

2 Construct a default classification tree for the data:

```
tree = ClassificationTree.fit(X,Y);
```

3 View the tree in the interactive viewer:

```
view(tree,'mode','graph')
```



4 Find the optimal pruning level by minimizing cross-validated loss:

```
[~,~,~,bestlevel] = cvLoss(tree,...
    'subtrees','all','treesize','min')
bestlevel =
    6
```

5 Prune the tree to level 6 in the interactive viewer:


The pruned tree is the same as the near-optimal tree in "Example: Selecting Appropriate Tree Depth" on page 13-36.

6 Set 'treesize' to 'se' (default) to find the maximal pruning level for which the tree error does not exceed the error from the best level plus one standard deviation:

```
[~,~,~,bestlevel] = cvLoss(tree,'subtrees','all')
bestlevel =
    6
```

In this case the level is the same for either setting of 'treesize'.
7 Prune the tree to use it for other purposes:

```
tree = prune(tree,'Level',6);
view(tree,'mode','graph')
```



## Alternative: classregtree

The ClassificationTree and RegressionTree classes are new in MATLAB R2011a. Previously, you represented both classification trees and regression trees with a classregtree object. The new classes provide all the functionality of the classregtree class, and are more convenient when used in conjunction with "Ensemble Methods" on page 13-51.

Before the classregtree class, there were treefit, treedisp, treeval, treeprune, and treetest functions. Statistics Toolbox software maintains these only for backward compatibility.

## Example: Creating Classification Trees Using classregtree

This example uses Fisher's iris data in fisheriris.mat to create a classification tree for predicting species using measurements of sepal length, sepal width, petal length, and petal width as predictors. Here, the predictors are continuous and the response is categorical.

1 Load the data and use the classregtree constructor of the classregtree class to create the classification tree:

```
load fisheriris
t = classregtree(meas,species,...
    'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
```

t is a classregtree object and can be operated on with any class method.
2 Use the type method of the classregtree class to show the type of the tree:

```
treetype = type(t)
treetype =
classification
```

classregtree creates a classification tree because species is a cell array of strings, and the response is assumed to be categorical.

3 To view the tree, use the view method of the classregtree class:

```
view(t)
```




The tree predicts the response values at the circular leaf nodes based on a series of questions about the iris at the triangular branching nodes. A true answer to any question follows the branch to the left. A false follows the branch to the right.

4 The tree does not use sepal measurements for predicting species. These can go unmeasured in new data, and you can enter them as NaN values for predictions. For example, to use the tree to predict the species of an iris with petal length 4.8 and petal width 1.6 , type:

```
predicted = t([NaN NaN 4.8 1.6])
predicted =
    'versicolor'
```

The object allows for functional evaluation, of the form $t(X)$. This is a shorthand way of calling the eval method of the classregtree class. The predicted species is the left leaf node at the bottom of the tree in the previous view.

5 You can use a variety of methods of the classregtree class, such as cutvar and cuttype to get more information about the split at node 6 that makes the final distinction between versicolor and virginica:

```
var6 = cutvar(t,6) % What variable determines the split?
var6 =
    'PW'
type6 = cuttype(t,6) % What type of split is it?
type6 =
    'continuous'
```

6 Classification trees fit the original (training) data well, but can do a poor job of classifying new values. Lower branches, especially, can be strongly affected by outliers. A simpler tree often avoids overfitting. You can use the prune method of the classregtree class to find the next largest tree from an optimal pruning sequence:

```
pruned = prune(t,'level',1)
pruned =
Decision tree for classification
    if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
    class = setosa
    if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
    if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
    class = virginica
    class = versicolor
    class = virginica
view(pruned)
```

|  |  |  |  |  |  |  |  | - |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Click to display: | Identity | $\square$ | Magnification: | 100\% | $\checkmark$ | Pruning level: | 0 of 3 | - |



To find the best classification tree, employing the techniques of resubstitution and cross validation, use the test method of the classregtree class.

## Example: Creating Regression Trees Using classregtree

This example uses the data on cars in carsmall.mat to create a regression tree for predicting mileage using measurements of weight and the number of cylinders as predictors. Here, one predictor (weight) is continuous and the other (cylinders) is categorical. The response (mileage) is continuous.

1 Load the data and use the classregtree constructor of the classregtree class to create the regression tree:

```
load carsmall
t = classregtree([Weight, Cylinders],MPG,...
    'cat',2,'splitmin',20,...
    'names',{'W','C'})
t =
Decision tree for regression
    if W<3085.5 then node 2 elseif W>=3085.5 then node 3 else 23.7181
    if W<2371 then node 4 elseif W>=2371 then node 5 else 28.7931
    if C=8 then node 6 elseif C in {4 6} then node 7 else 15.5417
    if W<2162 then node 8 elseif W>=2162 then node 9 else 32.0741
    if C=6 then node 10 elseif C=4 then node 11 else 25.9355
    if W<4381 then node 12 elseif W>=4381 then node 13 else 14.2963
    fit = 19.2778
    fit = 33.3056
    fit = 29.6111
    fit = 23.25
    if W<2827.5 then node 14 elseif W>=2827.5 then node 15 else 27.2143
    if W<3533.5 then node 16 elseif W>=3533.5 then node 17 else 14.8696
    fit = 11
    fit = 27.6389
    fit = 24.6667
    fit = 16.6
    fit = 14.3889
```

$t$ is a classregtree object and can be operated on with any of the methods of the class.

2 Use the type method of the classregtree class to show the type of the tree:

```
treetype = type(t)
treetype =
regression
```

classregtree creates a regression tree because MPG is a numerical vector, and the response is assumed to be continuous.

3 To view the tree, use the view method of the classregtree class:

```
view(t)
```




The tree predicts the response values at the circular leaf nodes based on a series of questions about the car at the triangular branching nodes. A true answer to any question follows the branch to the left; a false follows the branch to the right.

4 Use the tree to predict the mileage for a 2000 -pound car with either 4,6 , or 8 cylinders:

```
mileage2K = t([2000 4; 2000 6; 2000 8])
mileage2K =
    33.3056
    33.3056
    33.3056
```

The object allows for functional evaluation, of the form $t(X)$. This is a shorthand way of calling the eval method of the classregtree class.

5 The predicted responses computed above are all the same. This is because they follow a series of splits in the tree that depend only on weight, terminating at the left-most leaf node in the view above. A 4000-pound car, following the right branch from the top of the tree, leads to different predicted responses:

```
mileage4K = t([4000 4; 4000 6; 4000 8])
```

```
mileage4K =
    19.2778
    19.2778
    14.3889
```

6 You can use a variety of other methods of the classregtree class, such as cutvar, cuttype, and cutcategories, to get more information about the split at node 3 that distinguishes the 8 -cylinder car:

```
var3 = cutvar(t,3) % What variable determines the split?
var3 =
    'C'
type3 = cuttype(t,3) % What type of split is it?
type3 =
    'categorical'
c = cutcategories(t,3) % Which classes are sent to the left
    % child node, and which to the right?
c =
    [8] [1x2 double]
c{1}
ans =
    8
c{2}
ans =
    4 6
```

Regression trees fit the original (training) data well, but may do a poor job of predicting new values. Lower branches, especially, may be strongly affected by outliers. A simpler tree often avoids over-fitting. To find the best regression tree, employing the techniques of resubstitution and cross validation, use the test method of the classregtree class.

## Ensemble Methods

In this section...<br>"Framework for Ensemble Learning" on page 13-51<br>"Basic Ensemble Examples" on page 13-58<br>"Test Ensemble Quality" on page 13-60<br>"Classification: Imbalanced Data or Unequal Misclassification Costs" on page 13-65<br>"Example: Classification with Many Categorical Levels" on page 13-72<br>"Example: Surrogate Splits" on page 13-77<br>"Ensemble Regularization" on page 13-81<br>"Example: Tuning RobustBoost" on page 13-92<br>"TreeBagger Examples" on page 13-96<br>"Ensemble Algorithms" on page 13-118

## Framework for Ensemble Learning

You have several methods for melding results from many weak learners into one high-quality ensemble predictor. These methods follow, as closely as possible, the same syntax, so you can try different methods with only minor changes in your commands.

Create an ensemble with the fitensemble function. The syntax of fitensemble is

```
ens = fitensemble(X,Y,model,numberens,learners)
```

- X is the matrix of data, each row containing one observation, each column contains one predictor variable.
- Y is the responses, with the same number of observations as rows in X .
- model is a string naming the type of ensemble.
- numberens is the number of weak learners in ens from each element of learners. So the number of elements in ens is numberens times the number of elements in learners.
- learners is a string naming a weak learner, is a weak learner template, or is a cell array of such templates.

Pictorially, here is the information you need to create an ensemble:


For all classification or nonlinear regression problems, follow these steps to create an ensemble:

1 "Put Predictor Data in a Matrix" on page 13-52
2 "Prepare Response Data" on page 13-53
3 "Choose an Applicable Ensemble Method" on page 13-54
4 "Set the Number of Ensemble Members" on page 13-55
5 "Prepare the Weak Learners" on page 13-55
6 "Call fitensemble" on page 13-56

## Put Predictor Data in a Matrix

All supervised learning methods start with a data matrix, usually called $X$ in this documentation. Each row of $X$ represents one observation. Each column of $X$ represents one variable, or predictor.

Currently, you can use only decision trees as learners for ensembles. Decision trees can handle NaN values in X. Such values are called "missing." If you have
some missing values in a row of $X$, a decision tree finds optimal splits using nonmissing values only. If an entire row consists of NaN, fitensemble ignores that row. If you have data with a large fraction of missing values in X, use surrogate decision splits. For examples of surrogate splits, see "Example: Unequal Classification Costs" on page 13-67 and "Example: Surrogate Splits" on page 13-77.

## Prepare Response Data

You can use a wide variety of data types for response data.

- For regression ensembles, Y must be a numeric vector with the same number of elements as the number of rows of $X$.
- For classification ensembles, Y can be any of the following data types. The table also contains the method of including missing entries.

| Data Type | Missing Entry |
| :--- | :--- |
| Numeric vector | NaN |
| Categorical vector | <undefined> |
| Character array | Row of spaces |
| Cell array of strings | '' |
| Logical vector | (not possible to represent) |

fitensemble ignores missing values in $Y$ when creating an ensemble.
For example, suppose your response data consists of three observations in the following order: true, false, true. You could express $Y$ as:

- [1;0;1] (numeric vector)
- nominal(\{'true','false','true'\}) (categorical vector)
- [true;false;true] (logical vector)
- ['true ';'false';'true '] (character array, padded with spaces so each row has the same length)
- \{'true', 'false', 'true'\} (cell array of strings)

Use whichever data type is most convenient. Since you cannot represent missing values with logical entries, do not use logical entries when you have missing values in Y .

## Choose an Applicable Ensemble Method

fitensemble uses one of these algorithms to create an ensemble.

- For classification with two classes:
- 'AdaBoostM1'
- 'LogitBoost'
- 'GentleBoost'
- 'RobustBoost'
- 'Bag'
- For classification with three or more classes:
- 'AdaBoostM2'
- 'Bag'
- For regression:
- 'LSBoost'
- 'Bag'

Since 'Bag' applies to all methods, indicate whether you want a classifier or regressor with the type name-value pair set to 'classification' or 'regression'.

For descriptions of the various algorithms, and aid in choosing which applies to your data, see "Ensemble Algorithms" on page 13-118. The following table gives characteristics of the various algorithms. In the table titles:

- Regress. - Regression
- Classif. - Classification
- Preds. - Predictors
- Estim. - Estimate
- Gen. - Generalization
- Pred. - Prediction
- Mem. - Memory usage

| Algorithm | Regress.Binary <br> Classif. | Binary <br> Classif. <br> Multi- <br> Level <br> Preds. | Classif. <br> 3+ <br> Classes | Auto <br> Estim. <br> Gen. <br> Error | Fast <br> Train | Fast <br> Pred. | Low <br> Mem. |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Bag | $\times$ | $\times$ |  | $\times$ | $\times$ |  |  |  |
| AdaBoostM1 |  | $\times$ |  |  |  | $\times$ | $\times$ | $\times$ |
| AdaBoostM2 |  |  |  | $\times$ |  | $\times$ | $\times$ | $\times$ |
| LogitBoost |  | $\times$ | $\times$ |  |  | $\times$ | $\times$ | $\times$ |
| GentleBoost | $\times$ | $\times$ |  |  | $\times$ | $\times$ | $\times$ |  |
| RobustBoost |  | $\times$ |  |  |  |  | $\times$ | $\times$ |
| LSBoost | $\times$ |  |  |  |  | $\times$ | $\times$ | $\times$ |

## Set the Number of Ensemble Members

Choosing the size of an ensemble involves balancing speed and accuracy.

- Larger ensembles take longer to train and to generate predictions.
- Some ensemble algorithms can become overtrained (inaccurate) when too large.

To set an appropriate size, consider starting with several dozen to several hundred members in an ensemble, training the ensemble, and then checking the ensemble quality, as in "Example: Test Ensemble Quality" on page 13-60. If it appears that you need more members, add them using the resume method (classification) or the resume method (regression). Repeat until adding more members does not improve ensemble quality.

## Prepare the Weak Learners

Currently there is one built-in weak learner type: 'Tree'. To create an ensemble with the default tree options, pass in 'Tree' as the weak learner.

To set a nondefault classification tree learner, create a classification tree template with the ClassificationTree.template method.

Similarly, to set a nondefault regression tree learner, create a regression tree template with the RegressionTree.template method.

While you can give fitensemble a cell array of learner templates, the most common usage is to give just one weak learner template.

For examples using a template, see "Example: Unequal Classification Costs" on page 13-67 and "Example: Surrogate Splits" on page 13-77.

## Common Settings for Weak Learners.

- The depth of the weak learner tree makes a difference for training time, memory usage, and predictive accuracy. You control the depth with two parameters:
- MinLeaf - Each leaf has at least MinLeaf observations. Set small values of MinLeaf to get a deep tree.
- MinParent - Each branch node in the tree has at least MinParent observations. Set small values of MinParent to get a deep tree.

If you supply both MinParent and MinLeaf, the learner uses the setting that gives larger leaves:
MinParent = max(MinParent,2*MinLeaf)

- Surrogate - Grow decision trees with surrogate splits when Surrogate is ' on '. Use surrogate splits when your data has missing values.

Note Surrogate splits cause training to be slower and use more memory.

## Call fitensemble

The syntax of fitensemble is

```
ens = fitensemble(X,Y,model,numberens,learners)
```

- $X$ is the matrix of data. Each row contains one observation, and each column contains one predictor variable.
- Y is the responses, with the same number of observations as rows in X .
- model is a string naming the type of ensemble.
- numberens is the number of weak learners in ens from each element of learners. So the number of elements in ens is numberens times the number of elements in learners.
- learners is a string naming a weak learner, a weak learner template, or a cell array of such strings and templates.

The result of fitensemble is an ensemble object, suitable for making predictions on new data. For a basic example of creating a classification ensemble, see "Creating a Classification Ensemble" on page 13-58. For a basic example of creating a regression ensemble, see "Creating a Regression Ensemble" on page 13-59.

Where to Set Name-Value Pairs. There are several name-value pairs you can pass to fitensemble, and several that apply to the weak learners (ClassificationTree.template and RegressionTree.template). To determine which option (name-value pair) is appropriate, the ensemble or the weak learner:

- Use template name-value pairs to control the characteristics of the weak learners.
- Use fitensemble name-value pairs to control the ensemble as a whole, either for algorithms or for structure.

For example, to have an ensemble of boosted classification trees with each tree deeper than the default, set the ClassificationTree.template name-value pairs (MinLeaf and MinParent) to smaller values than the defaults. This causes the trees to be leafier (deeper).

To name the predictors in the ensemble (part of the structure of the ensemble), use the PredictorNames name-value pair in fitensemble.

## Basic Ensemble Examples

## Creating a Classification Ensemble

Create a classification ensemble for the Fisher iris data, and use it to predict the classification of a flower with average measurements.

1 Load the data:

```
load fisheriris
```

2 The predictor data $X$ is the meas matrix.
3 The response data $Y$ is the species cell array.
4 The only boosted classification ensemble for three or more classes is 'AdaBoostM2'.

5 For this example, arbitrarily take an ensemble of 100 trees.
6 Use a default tree template.
7 Create the ensemble:

```
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree')
ens =
classreg.learning.classif.ClassificationEnsemble:
    PredictorNames: {'x1' 'x2' 'x3' 'x4'}
    CategoricalPredictors: []
            ResponseName: 'Y'
                ClassNames: {'setosa' 'versicolor' 'virginica'}
            ScoreTransform: 'none'
            NObservations: 150
                    NTrained: 100
                            Method: 'AdaBoostM2'
            LearnerNames: {'Tree'}
        ReasonForTermination: [1x77 char]
            FitInfo: [100x1 double]
            FitInfoDescription: [2x83 char]
```

8 Predict the classification of a flower with average measurements:

```
flower = predict(ens,mean(meas))
flower =
    'versicolor'
```


## Creating a Regression Ensemble

Create a regression ensemble to predict mileage of cars based on their horsepower and weight, trained on the carsmall data. Use the resulting ensemble to predict the mileage of a car with 150 horsepower weighing 2750 lbs.

1 Load the data:

```
load carsmall
```

2 Prepare the input data.

```
X = [Horsepower Weight];
```

3 The response data $Y$ is MPG.
4 The only boosted regression ensemble type is 'LSBoost'.
5 For this example, arbitrarily take an ensemble of 100 trees.
6 Use a default tree template.
7 Create the ensemble:

```
ens = fitensemble(X,MPG,'LSBoost',100,'Tree')
ens =
classreg.learning.regr.RegressionEnsemble:
            PredictorNames: {'x1' 'x2'}
    CategoricalPredictors: []
            ResponseName: 'Y'
            ResponseTransform: 'none'
            NObservations: 94
            NTrained: 100
                Method: 'LSBoost'
            LearnerNames: {'Tree'}
```

```
ReasonForTermination: [1x77 char]
            FitInfo: [100x1 double]
FitInfoDescription: [2x83 char]
    Regularization: []
```

8 Predict the mileage of a car with 150 horsepower weighing 2750 lbs:

```
mileage = ens.predict([150 2750])
mileage =
    22.6735
```


## Test Ensemble Quality

Usually you cannot evaluate the predictive quality of an ensemble based on its performance on training data. Ensembles tend to "overtrain," meaning they produce overly optimistic estimates of their predictive power. This means the result of resubLoss for classification (resubLoss for regression) usually indicates lower error than you get on new data.

To obtain a better idea of the quality of an ensemble, use one of these methods:

- Evaluate the ensemble on an independent test set (useful when you have a lot of training data).
- Evaluate the ensemble by cross validation (useful when you don't have a lot of training data).
- Evaluate the ensemble on out-of-bag data (useful when you create a bagged ensemble with fitensemble).


## Example: Test Ensemble Quality

This example uses a bagged ensemble so it can use all three methods of evaluating ensemble quality.

1 Generate an artificial dataset with 20 predictors. Each entry is a random number from 0 to 1 . The initial classification:

$$
\begin{aligned}
& Y=1 \text { when } X(1)+X(2)+X(3)+X(4)+X(5)>2.5 \\
& Y=0 \text { otherwise. }
\end{aligned}
$$

```
rng(1,'twister') % for reproducibility
X = rand(2000,20);
Y = sum(X(:,1:5),2) > 2.5;
```

In addition, to add noise to the results, randomly switch $10 \%$ of the classifications:

```
idx = randsample(2000,200);
Y(idx) = ~Y(idx);
```


## 2 Independent Test Set

Create independent training and test sets of data. Use 70\% of the data for a training set by calling cvpartition with the holdout option:

```
cvpart = cvpartition(Y,'holdout',0.3);
Xtrain = X(training(cvpart),:);
Ytrain = Y(training(cvpart),:);
Xtest = X(test(cvpart),:);
Ytest = Y(test(cvpart),:);
```

3 Create a bagged classification ensemble of 200 trees from the training data:

```
bag = fitensemble(Xtrain,Ytrain,'Bag',200,'Tree',...
    'type','classification')
bag =
classreg.learning.classif.ClassificationBaggedEnsemble:
            PredictorNames: {1x20 cell}
    CategoricalPredictors: []
            ResponseName: 'Y'
            ClassNames: [0 1]
            ScoreTransform: 'none'
                NObservations: 1400
                        NTrained: 200
                                    Method: 'Bag'
            LearnerNames: {'Tree'}
            ReasonForTermination: [1x77 char]
                    FitInfo: []
            FitInfoDescription: 'None'
```

```
    FResample: 1
    Replace: 1
UseObsForLearner: [1400x200 logical]
```

4 Plot the loss (misclassification) of the test data as a function of the number of trained trees in the ensemble:

```
figure;
plot(loss(bag,Xtest,Ytest,'mode','cumulative'));
xlabel('Number of trees');
ylabel('Test classification error');
```



## 5 Cross validation

Generate a five-fold cross-validated bagged ensemble:

```
cv = fitensemble(X,Y,'Bag',200,'Tree',...
    'type','classification','kfold',5)
```

```
cv =
classreg.learning.partition.ClassificationPartitionedEnsemble:
    CrossValidatedModel: 'Bag'
            PredictorNames: {1x20 cell}
    CategoricalPredictors: []
            ResponseName: 'Y'
            NObservations: 2000
            KFold: 5
            Partition: [1x1 cvpartition]
                NTrainedPerFold: [200 200 200 200 200]
            ClassNames: [0 1]
            ScoreTransform: 'none'
```

6 Examine the cross validation loss as a function of the number of trees in the ensemble:

```
figure;
plot(loss(bag,Xtest,Ytest,'mode','cumulative'));
hold
plot(kfoldLoss(cv,'mode','cumulative'),'r.');
hold off;
xlabel('Number of trees');
ylabel('Classification error');
legend('Test','Cross-validation','Location','NE');
```



Cross validating gives comparable estimates to those of the independent set.

## 7 Out-of-Bag Estimates

Generate the loss curve for out-of-bag estimates, and plot it along with the other curves:

```
figure;
plot(loss(bag,Xtest,Ytest,'mode','cumulative'));
hold
plot(kfoldLoss(cv,'mode','cumulative'),'r.');
plot(oobLoss(bag,'mode','cumulative'),'k--');
hold off;
xlabel('Number of trees');
ylabel('Classification error');
legend('Test','Cross-validation','Out of bag','Location','NE');
```



The out-of-bag estimates are again comparable to those of the other methods.

## Classification: Imbalanced Data or Unequal Misclassification Costs

In many real-world applications, you might prefer to treat classes in your data asymmetrically. For example, you might have data with many more observations of one class than of any other. Or you might work on a problem in which misclassifying observations of one class has more severe consequences than misclassifying observations of another class. In such situations, you can use two optional parameters for fitensemble: prior and cost.

By using prior, you set prior class probabilities (that is, class probabilities used for training). Use this option if some classes are under- or overrepresented in your training set. For example, you might obtain your training data by simulation. Because simulating class A is more expensive than class B, you opt to generate fewer observations of class A and more
observations of class B. You expect, however, that class A and class B are mixed in a different proportion in the real world. In this case, set prior probabilities for class A and B approximately to the values you expect to observe in the real world. fitensemble normalizes prior probabilities to make them add up to 1; multiplying all prior probabilities by the same positive factor does not affect the result of classification.

If classes are adequately represented in the training data but you want to treat them asymmetrically, use the cost parameter. Suppose you want to classify benign and malignant tumors in cancer patients. Failure to identify a malignant tumor (false negative) has far more severe consequences than misidentifying benign as malignant (false positive). You should assign high cost to misidentifying malignant as benign and low cost to misidentifying benign as malignant.

You must pass misclassification costs as a square matrix with nonnegative elements. Element C(i,j) of this matrix is the cost of classifying an observation into class $j$ if the true class is $i$. The diagonal elements $C(i, i)$ of the cost matrix must be 0 . For the example above, you can choose malignant tumor to be class 1 and benign tumor to be class 2 . Then you can set the cost matrix to

$$
\left[\begin{array}{ll}
0 & c \\
1 & 0
\end{array}\right]
$$

where $c>1$ is the cost of misidentifying a malignant tumor as benign. Costs are relative-multiplying all costs by the same positive factor does not affect the result of classification.

If you have only two classes, fitensemble adjusts their prior probabilities using $\tilde{P}_{i}=C_{i j} P_{i}$ for class $i=1,2$ and $j \neq i . P_{i}$ are prior probabilities either passed into fitensemble or computed from class frequencies in the training
data, and $\tilde{P}_{i}$ are adjusted prior probabilities. Then fitensemble uses the default cost matrix

$$
\left[\begin{array}{ll}
0 & 1 \\
1 & 0
\end{array}\right]
$$

and these adjusted probabilities for training its weak learners. Manipulating the cost matrix is thus equivalent to manipulating the prior probabilities.

If you have three or more classes, fitensemble also converts input costs into adjusted prior probabilities. This conversion is more complex. First, fitensemble attempts to solve a matrix equation described in Zhou and Liu [12]. If it fails to find a solution, fitensemble applies the "average cost" adjustment described in Breiman et al. [4]. For more information, see .Zadrozny, Langford, and Abe [11]

## Example: Unequal Classification Costs

This example uses data on patients with hepatitis to see if they live or die as a result of the disease. The data is described at http://archive.ics.uci.edu/ml/datasets/Hepatitis.

1 Load the data into a file named hepatitis.txt:

```
s = urlread(['http://archive.ics.uci.edu/ml/' ...
    'machine-learning-databases/hepatitis/hepatitis.data']);
fid = fopen('hepatitis.txt','w');
fwrite(fid,s);
fclose(fid);
```

2 Load the data hepatitis.txt into a dataset, with variable names describing the fields in the data:

```
VarNames = {'die_or_live' 'age' 'sex' 'steroid' 'antivirals' 'fatigue' ...
    'malaise' 'anorexia' 'liver_big' 'liver_firm' 'spleen_palpable' ...
    'spiders' 'ascites' 'varices' 'bilirubin' 'alk_phosphate' 'sgot' ...
    'albumin' 'protime' 'histology'};
ds = dataset('file','hepatitis.txt','VarNames',VarNames,...
    'Delimiter',',','ReadVarNames',false,'TreatAsEmpty','?',...
    'Format', '%f%f%f%f%f%f%f%f%f%f%f%f%f%f%f%f%f%f%f%f');
```

ds is a dataset with 155 observations and 20 variables:

```
size(ds)
```

ans $=$

3 Convert the data in the dataset to the format for ensembles: a numeric matrix of predictors, and a cell array with outcome names: 'Die' or 'Live'. The first field in the dataset has the outcomes.

```
X = double(ds(:,2:end));
ClassNames = {'Die' 'Live'};
Y = ClassNames(ds.die_or_live);
```

4 Inspect the data for missing values:

```
figure;
bar(sum(isnan(X),1)/size(X,1));
xlabel('Predictor');
ylabel('Fraction of missing values');
```



Most predictors have missing values, and one has nearly $45 \%$ of missing values. Therefore, use decision trees with surrogate splits for better accuracy. Because the dataset is small, training time with surrogate splits should be tolerable.

5 Create a classification tree template that uses surrogate splits:

```
rng(0,'twister') % for reproducibility
t = ClassificationTree.template('surrogate','on');
```

6 Examine the data or the description of the data to see which predictors are categorical:

```
X(1:5,:)
ans =
```

Columns 1 through 6

| 30.0000 | 2.0000 | 1.0000 | 2.0000 | 2.0000 | 2.0000 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 50.0000 | 1.0000 | 1.0000 | 2.0000 | 1.0000 | 2.0000 |
| 78.0000 | 1.0000 | 2.0000 | 2.0000 | 1.0000 | 2.0000 |
| 31.0000 | 1.0000 | NaN | 1.0000 | 2.0000 | 2.0000 |
| 34.0000 | 1.0000 | 2.0000 | 2.0000 | 2.0000 | 2.0000 |

Columns 7 through 12

| 2.0000 | 1.0000 | 2.0000 | 2.0000 | 2.0000 | 2.0000 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 2.0000 | 1.0000 | 2.0000 | 2.0000 | 2.0000 | 2.0000 |
| 2.0000 | 2.0000 | 2.0000 | 2.0000 | 2.0000 | 2.0000 |
| 2.0000 | 2.0000 | 2.0000 | 2.0000 | 2.0000 | 2.0000 |
| 2.0000 | 2.0000 | 2.0000 | 2.0000 | 2.0000 | 2.0000 |

Columns 13 through 18

| 2.0000 | 1.0000 | 85.0000 | 18.0000 | 4.0000 | NaN |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 2.0000 | 0.9000 | 135.0000 | 42.0000 | 3.5000 | NaN |
| 2.0000 | 0.7000 | 96.0000 | 32.0000 | 4.0000 | NaN |
| 2.0000 | 0.7000 | 46.0000 | 52.0000 | 4.0000 | 80.0000 |
| 2.0000 | 1.0000 | NaN | 200.0000 | 4.0000 | NaN |

```
Column 19
    1.0000
    1.0000
    1.0000
    1.0000
    1.0000
```

It appears that predictors 2 through 13 are categorical, as well as predictor 19. You can confirm this inference with the dataset description at http://archive.ics.uci.edu/ml/datasets/Hepatitis.

7 List the categorical variables:

```
ncat = [2:13,19];
```

8 Create a cross-validated ensemble using 200 learners and the GentleBoost algorithm:

```
a = fitensemble(X,Y,'GentleBoost',200,t,...
    'PredictorNames',VarNames(2:end),'LearnRate',0.1,...
    'CategoricalPredictors',ncat,'kfold',5);
figure;
plot(kfoldLoss(a,'mode','cumulative','lossfun','exponential'));
xlabel('Number of trees');
ylabel('Cross-validated exponential loss');
```



9 Inspect the confusion matrix to see which people the ensemble predicts correctly:

```
[Yfit,Sfit] = kfoldPredict(a); %
confusionmat(Y,Yfit,'order',ClassNames)
ans =
    16 16
    10 113
```

Of the 123 people who live, the ensemble predicts correctly that 113 will live. But for the 32 people who die of hepatitis, the ensemble only predicts correctly that half will die of hepatitis.

10 There are two types of error in the predictions of the ensemble:

- Predicting that the patient lives, but the patient dies
- Predicting that the patient dies, but the patient lives

Suppose you believe that the first error is five times worse than the second. Make a new classification cost matrix that reflects this belief:

```
cost.ClassNames = ClassNames;
cost.ClassificationCosts = [0 5; 1 0];
```

11 Create a new cross-validated ensemble using cost as misclassification cost, and inspect the resulting confusion matrix:

```
aC = fitensemble(X,Y,'GentleBoost',200,t,...
    'PredictorNames',VarNames(2:end),'LearnRate',0.1,...
    'CategoricalPredictors',ncat,'kfold',5,...
    'cost',cost);
[YfitC,SfitC] = kfoldPredict(aC);
confusionmat(Y,YfitC,'order',ClassNames)
ans =
    19 13
    9 114
```

As expected, the new ensemble does a better job classifying the people who die. Somewhat surprisingly, the new ensemble also does a better job classifying the people who live, though the result is not statistically significantly better. The results of the cross validation are random, so this result is simply a statistical fluctuation. The result seems to indicate that the classification of people who live is not very sensitive to the cost.

## Example: Classification with Many Categorical Levels

Generally, you cannot use classification with more than 31 levels in any categorical predictor. However, two boosting algorithms can classify data with many categorical levels: LogitBoost and GentleBoost. For details, see "LogitBoost" on page 13-125 and "GentleBoost" on page 13-126.

This example uses demographic data from the U.S. Census, available at http://archive.ics.uci.edu/ml/machine-learning-databases/adult/. The objective of the researchers who posted the data is predicting whether an individual makes more than $\$ 50,000 /$ year, based on a set of characteristics.

You can see details of the data, including predictor names, in the adult. names file at the site.

1 Load the 'adult.data' file from the UCI Machine Learning Repository:

```
s = urlread(['http://archive.ics.uci.edu/ml/' ...
    'machine-learning-databases/adult/adult.data']);
```

2 'adult.data' represents missing data as '?'. Replace instances of missing data with the blank string ' ':

```
s = strrep(s,'?','');
```

3 Put the data into a MATLAB dataset array:

```
fid = fopen('adult.txt','w');
fwrite(fid,s);
fclose(fid);
clear s;
VarNames = {'age' 'workclass' 'fnlwgt' 'education' 'education_num' ...
    'marital_status' 'occupation' 'relationship' 'race' ...
    'sex' 'capital_gain' 'capital_loss' ...
    'hours_per_week' 'native_country' 'income'};
ds = dataset('file','adult.txt','VarNames',VarNames,...
    'Delimiter',',','ReadVarNames',false,'Format',...
    '%u%s%u%s%u%S%S%s%s%s%u%u%u%s%s');
cat = ~datasetfun(@isnumeric,ds(:,1:end-1)); % Logical indices
% of categorical variables
catcol = find(cat); % indices of categorical variables
```

4 Many predictors in the data are categorical. Convert those fields in the dataset array to nominal:

```
ds.workclass = nominal(ds.workclass);
ds.education = nominal(ds.education);
ds.marital_status = nominal(ds.marital_status);
ds.occupation = nominal(ds.occupation);
ds.relationship = nominal(ds.relationship);
ds.race = nominal(ds.race);
ds.sex = nominal(ds.sex);
```

```
ds.native_country = nominal(ds.native_country);
ds.income = nominal(ds.income);
```

5 Convert the dataset array into numerical variables for fitensemble:

```
X = double(ds(:,1:end-1));
Y = ds.income;
```

6 Some variables have many levels. Plot the number of levels of each predictor:

```
ncat = zeros(1,numel(catcol));
for c=1:numel(catcol)
    [~,gn] = grp2idx(X(:,catcol(c)));
    ncat(c) = numel(gn);
end
figure;
bar(catcol,ncat);
xlabel('Predictor');
ylabel('Number of categories');
```



Predictor 14 ('native_country') has more than 40 categorical levels. This is too many levels for any method except LogitBoost and GentleBoost.

7 Create classification ensembles using both LogitBoost and GentleBoost:

```
lb = fitensemble(X,Y,'LogitBoost',300,'Tree','CategoricalPredictors',cat,...
    'PredictorNames',VarNames(1:end-1),'ResponseName','income');
gb = fitensemble(X,Y,'GentleBoost',300,'Tree','CategoricalPredictors',cat,...
    'PredictorNames',VarNames(1:end-1),'ResponseName','income');
```

8 Examine the resubstitution error for the two ensembles:

```
figure;
plot(resubLoss(lb,'mode','cumulative'));
hold on
plot(resubLoss(gb,'mode','cumulative'),'r--');
hold off
xlabel('Number of trees');
ylabel('Resubstitution error');
legend('LogitBoost','GentleBoost','Location','NE');
```



The algorithms have similar resubstitution error.
9 Estimate the generalization error for the two algorithms by cross validation.

```
lbcv = crossval(lb,'kfold',5);
gbcv = crossval(gb,'kfold',5);
figure;
plot(kfoldLoss(lbcv,'mode','cumulative'));
hold on
plot(kfoldLoss(gbcv,'mode','cumulative'),'r--');
hold off
xlabel('Number of trees');
ylabel('Cross-validated error');
legend('LogitBoost','GentleBoost','Location','NE');
```



The cross-validated loss is nearly the same as the resubstitution error.

## Example: Surrogate Splits

When you have missing data, trees and ensembles of trees give better predictions when they include surrogate splits. Furthermore, estimates of predictor importance are often different with surrogate splits. Eliminating unimportant predictors can save time and memory for predictions, and can make predictions easier to understand.

This example shows the effects of surrogate splits for predictions for data containing missing entries in both training and test sets. There is a redundant predictor in the data, which the surrogate split uses to infer missing values. While the example is artificial, it shows the value of surrogate splits with missing data.

1 Generate and plot two different normally-distributed populations, one with 5000 members, one with 10,000 members:

```
rng(1,'twister') % for reproducibility
N = 5000;
N1 = 2*N; % number in population 1
N2 = N; % number in population 2
mu1 = [-1 -1]/2; % mean of population 1
mu2 = [1 1]/2; % mean of population 2
S1 = [3 -2.5;...
    -2.5 3]; % variance of population 1
S2 = [3 2.5;...
    2.5 3]; % variance of population 2
X1 = mvnrnd(mu1,S1,N1); % population 1
X2 = mvnrnd(mu2,S2,N2); % population 2
X = [X1; X2]; % total population
Y = ones(N1+N2,1); % label population 1
Y(N1+1:end) = 2; % label population 2
figure
plot(X1(:,1),X1(:,2),'k.','MarkerSize',2)
hold on
plot(X2(:,1),X2(:,2),'rx','MarkerSize',3);
hold off
axis square
```



There is a good deal of overlap between the data points. You cannot expect perfect classification of this data.

2 Make a third predictor that is the same as the first component of $X$ :

$$
X=[X X(:, 1)] ;
$$

3 Remove half the values of predictor 1 at random:

$$
X(\operatorname{rand}(\operatorname{size}(X(:, 1)))<0.5,1)=\operatorname{NaN} ;
$$

4 Partition the data into a training set and a test set:

```
cv = cvpartition(Y,'holdout',0.3); % 30% test data
```

```
Xtrain = X(training(cv),:);
Ytrain = Y(training(cv));
Xtest = X(test(cv),:);
Ytest = Y(test(cv));
```

5 Create two Bag ensembles: one with surrogate splits, one without. First create the template for surrogate splits, then train both ensembles:

```
templS = ClassificationTree.template('surrogate','on');
bag = fitensemble(Xtrain,Ytrain,'Bag',50,'Tree',...
    'type','class','nprint',10);
Training Bag...
Grown weak learners: 10
Grown weak learners: 20
Grown weak learners: 30
Grown weak learners: 40
Grown weak learners: 50
bagS = fitensemble(Xtrain,Ytrain,'Bag',50,templS,...
    'type','class','nprint',10);
Training Bag...
Grown weak learners: 10
Grown weak learners: 20
Grown weak learners: 30
Grown weak learners: 40
Grown weak learners: 50
```

6 Examine the accuracy of the two ensembles for predicting the test data:

```
figure
plot(loss(bag,Xtest,Ytest,'mode','cumulative'));
hold on
plot(loss(bagS,Xtest,Ytest,'mode','cumulative'),'r--');
hold off;
legend('Without surrogate splits','With surrogate splits');
xlabel('Number of trees');
ylabel('Test classification error');
```



The ensemble with surrogate splits is obviously more accurate than the ensemble without surrogate splits.

7 Check the statistical significance of the difference in results with the McNemar test:

```
Yfit = predict(bag,Xtest);
YfitS = predict(bagS,Xtest);
N10 = sum(Yfit==Ytest & YfitS~=Ytest);
N01 = sum(Yfit~=Ytest & YfitS==Ytest);
mcnemar = (abs(N1O-NO1) - 1)^2/(N10+NO1);
pval = 1 - chi2cdf(mcnemar,1)
pval =
    0
```

The extremely low $p$-value indicates that the ensemble with surrogate splits is better in a statistically significant manner.

## Ensemble Regularization

Regularization is a process of choosing fewer weak learners for an ensemble in a way that does not diminish predictive performance. Currently you can regularize regression ensembles.

The regularize method finds an optimal set of learner weights $a_{t}$ that minimize

$$
\sum_{n=1}^{N} w_{n} g\left(\left(\sum_{t=1}^{T} \alpha_{t} h_{t}\left(x_{n}\right)\right), y_{n}\right)+\lambda \sum_{t=1}^{T}\left|\alpha_{t}\right| .
$$

Here

- $\lambda \geq 0$ is a parameter you provide, called the lasso parameter.
- $h_{t}$ is a weak learner in the ensemble trained on $N$ observations with predictors $x_{n}$, responses $y_{n}$, and weights $w_{n}$.
- $g(f, y)=(f-y)^{2}$ is the squared error.

The ensemble is regularized on the same $\left(x_{n}, y_{n}, w_{n}\right)$ data used for training, so

$$
\sum_{n=1}^{N} w_{n} g\left(\left(\sum_{t=1}^{T} \alpha_{t} h_{t}\left(x_{n}\right)\right), y_{n}\right)
$$

is the ensemble resubstitution error (MSE).
If you use $\lambda=0$, regularize finds the weak learner weights by minimizing the resubstitution MSE. Ensembles tend to overtrain. In other words, the resubstitution error is typically smaller than the true generalization error. By making the resubstitution error even smaller, you are likely to make the ensemble accuracy worse instead of improving it. On the other hand, positive values of $\lambda$ push the magnitude of the $\alpha_{t}$ coefficients to 0 . This often improves the generalization error. Of course, if you choose $\lambda$ too large, all the optimal coefficients are 0 , and the ensemble does not have any accuracy. Usually you can find an optimal range for $\lambda$ in which the accuracy of the regularized ensemble is better or comparable to that of the full ensemble without regularization.

A nice feature of lasso regularization is its ability to drive the optimized coefficients precisely to zero. If a learner's weight $\alpha_{t}$ is 0 , this learner can be excluded from the regularized ensemble. In the end, you get an ensemble with improved accuracy and fewer learners.

## Example: Regularizing a Regression Ensemble

This example uses data for predicting the insurance risk of a car based on its many attributes.

1 Load the imports-85 data into the MATLAB workspace:

```
load imports-85;
```

2 Look at a description of the data to find the categorical variables and predictor names:

```
Description
Description =
1 9 8 5 \text { Auto Imports Database from the UCI repository}
http://archive.ics.uci.edu/ml/machine-learning-databases/autos/imports-85.names
Variables have been reordered to place variables with numeric values (referred
to as "continuous" on the UCI site) to the left and categorical values to the
right. Specifically, variables 1:16 are: symboling, normalized-losses,
wheel-base, length, width, height, curb-weight, engine-size, bore, stroke,
compression-ratio, horsepower, peak-rpm, city-mpg, highway-mpg, and price.
Variables 17:26 are: make, fuel-type, aspiration, num-of-doors, body-style,
drive-wheels, engine-location, engine-type, num-of-cylinders, and fuel-system.
```

The objective of this process is to predict the "symboling," the first variable in the data, from the other predictors. "symboling" is an integer from - 3 (good insurance risk) to 3 (poor insurance risk). You could use a classification ensemble to predict this risk instead of a regression ensemble. As stated in "Steps in Supervised Learning (Machine Learning)" on page 13-2, when you have a choice between regression and classification, you should try regression first. Furthermore, this example is to show regularization, which currently works only for regression.

3 Prepare the data for ensemble fitting:

```
Y = X(:,1);
X(:,1) = [];
VarNames = {'normalized-losses' 'wheel-base' 'length' 'width' 'height' ...
    'curb-weight' 'engine-size' 'bore' 'stroke' 'compression-ratio' ...
    'horsepower' 'peak-rpm' 'city-mpg' 'highway-mpg' 'price' 'make' ...
    'fuel-type' 'aspiration' 'num-of-doors' 'body-style' 'drive-wheels' ...
    'engine-location' 'engine-type' 'num-of-cylinders' 'fuel-system'};
catidx = 16:25; % indices of categorical predictors
```

4 Create a regression ensemble from the data using 300 default trees:

```
ls = fitensemble(X,Y,'LSBoost',300,'Tree','LearnRate',0.1,\ldots
    'PredictorNames',VarNames,'ResponseName','symboling',...
    'CategoricalPredictors',catidx)
ls =
classreg.learning.regr.RegressionEnsemble:
            PredictorNames: {1x25 cell}
    CategoricalPredictors: [16 17 18 19 20 21 22 23 24 25]
            ResponseName: 'symboling'
            ResponseTransform: 'none'
            NObservations: 205
                NTrained: 300
                    Method: 'LSBoost'
            LearnerNames: {'Tree'}
        ReasonForTermination: [1x77 char]
            FitInfo: [300x1 double]
        FitInfoDescription: [2x83 char]
            Regularization: []
```

The final line, Regularization, is empty ([ ]). To regularize the ensemble, you have to use the regularize method.

5 Cross validate the ensemble, and inspect its loss curve.

```
cv = crossval(ls,'kfold',5);
figure;
plot(kfoldLoss(cv,'mode','cumulative'));
```

```
xlabel('Number of trees');
ylabel('Cross-validated MSE');
```



It appears you might obtain satisfactory performance from a smaller ensemble, perhaps one containing from 50 to 100 trees.

6 Call the regularize method to try to find trees that you can remove from the ensemble. By default, regularize examines 10 values of the lasso (Lambda) parameter spaced exponentially.

```
ls = regularize(ls)
ls =
classreg.learning.regr.RegressionEnsemble:
            PredictorNames: {1x25 cell}
    CategoricalPredictors: [16 17 18 19 20 21 22 23 24 25]
```

```
    ResponseName: 'symboling'
    ResponseTransform: 'none'
    NObservations: 205
            NTrained: 300
            Method: 'LSBoost'
    LearnerNames: {'Tree'}
ReasonForTermination: [1x77 char]
            FitInfo: [300x1 double]
FitInfoDescription: [2x83 char]
    Regularization: [1x1 struct]
```

The Regularization property is no longer empty.
7 Plot the resubstitution mean-squared error (MSE) and number of learners with nonzero weights against the lasso parameter. Separately plot the value at Lambda=0. Use a logarithmic scale since the values of Lambda are exponentially spaced.

```
figure;
semilogx(ls.Regularization.Lambda,ls.Regularization.ResubstitutionMSE);
line([1e-3 1e-3],[ls.Regularization.ResubstitutionMSE(1) ...
    ls.Regularization.ResubstitutionMSE(1)],...
    'marker','x','markersize',12,'color','b');
rO = resubLoss(ls);
line([ls.Regularization.Lambda(2) ls.Regularization.Lambda(end)],...
    [rO r0],'color','r','LineStyle','--');
xlabel('Lambda');
ylabel('Resubstitution MSE');
annotation('textbox',[0.5 0.22 0.5 0.05],'String','unregularized ensemble',...
    'color','r','FontSize',14,'LineStyle','none');
figure;
loglog(ls.Regularization.Lambda,sum(ls.Regularization.TrainedWeights>0,1));
line([1e-3 1e-3],...
    [sum(ls.Regularization.TrainedWeights(:,1)>0) ...
    sum(ls.Regularization.TrainedWeights(:,1)>0)],...
    'marker','x','markersize',12,'color','b');
line([ls.Regularization.Lambda(2) ls.Regularization.Lambda(end)],...
    [ls.NTrained ls.NTrained],...
```

```
    'color','r','LineStyle','--');
xlabel('Lambda');
ylabel('Number of learners');
annotation('textbox',[0.3 0.8 0.5 0.05],'String','unregularized ensemble',...
    'color','r','FontSize',14,'LineStyle','none');
```




8 The resubstitution MSE values are likely to be overly optimistic. To obtain more reliable estimates of the error associated with various values of Lambda, cross validate the ensemble using cvshrink. Plot the resulting cross validation loss (MSE) and number of learners against Lambda.

```
rng(0,'Twister') % for reproducibility
[mse,nlearn] = cvshrink(ls,'lambda',ls.Regularization.Lambda,'kfold',5);
figure;
semilogx(ls.Regularization.Lambda,ls.Regularization.ResubstitutionMSE);
hold;
semilogx(ls.Regularization.Lambda,mse,'r--');
hold off;
xlabel('Lambda');
ylabel('Mean squared error');
legend('resubstitution','cross-validation','Location','NW');
```

```
line([1e-3 1e-3],[ls.Regularization.ResubstitutionMSE(1) ...
        ls.Regularization.ResubstitutionMSE(1)],...
    'marker','x','markersize',12,'color','b');
line([1e-3 1e-3],[mse(1) mse(1)],'marker','o',...
    'markersize',12,'color','r','LineStyle','--');
figure;
loglog(ls.Regularization.Lambda,sum(ls.Regularization.TrainedWeights>0,1));
hold;
loglog(ls.Regularization.Lambda,nlearn,'r--');
hold off;
xlabel('Lambda');
ylabel('Number of learners');
legend('resubstitution','cross-validation','Location','NE');
line([1e-3 1e-3],...
    [sum(ls.Regularization.TrainedWeights(:,1)>0) ...
    sum(ls.Regularization.TrainedWeights(:,1)>0)],...
    'marker','x','markersize',12,'color','b');
line([1e-3 1e-3],[nlearn(1) nlearn(1)],'marker','o',...
```

'markersize', 12, 'color','r','LineStyle','--');



Examining the cross-validated error shows that the cross-validation MSE is almost flat for Lambda up to a bit over 1e-2.

9 Examine ls.Regularization. Lambda to find the highest value that gives MSE in the flat region (up to a bit over 1e-2):

```
jj = 1:length(ls.Regularization.Lambda);
[jj;ls.Regularization.Lambda]
ans =
```

Columns 1 through 6
1.0000
2.0000
3.0000
4.0000
5.0000
6.0000
$0 \quad 0.0014$
0.0033
0.0077
0.0183
0.0435

Columns 7 through 10

| 7.0000 | 8.0000 | 9.0000 | 10.0000 |
| :--- | :--- | :--- | ---: |
| 0.1031 | 0.2446 | 0.5800 | 1.3754 |

Element 5 of ls.Regularization.Lambda has value 0.0183 , the largest in the flat range.

10 Reduce the ensemble size using the shrink method. shrink returns a compact ensemble with no training data. The generalization error for the new compact ensemble was already estimated by cross validation in mse (5).

```
cmp = shrink(ls,'weightcolumn',5)
cmp =
classreg.learning.regr.CompactRegressionEnsemble:
    PredictorNames: {1x25 cell}
    CategoricalPredictors: [16 17 18 19 20 21 22 23 24 25]
            ResponseName: 'symboling'
        ResponseTransform: 'none'
        NTrained: 18
```

There are only 18 trees in the new ensemble, notably reduced from the 300 in ls.

11 Compare the sizes of the ensembles:

```
sz(1) = whos('cmp'); sz(2) = whos('ls');
[sz(1).bytes sz(2).bytes]
ans =
    162270 2791024
```

The reduced ensemble is about $6 \%$ the size of the original.
12 Compare the MSE of the reduced ensemble to that of the original ensemble:

```
figure;
plot(kfoldLoss(cv,'mode','cumulative'));
hold on
plot(cmp.NTrained,mse(5),'ro','MarkerSize',12);
xlabel('Number of trees');
```

```
ylabel('Cross-validated MSE');
legend('unregularized ensemble','regularized ensemble',...
    'Location', 'NE');
hold off
```



The reduced ensemble gives low loss while using many fewer trees.

## Example: Tuning RobustBoost

The RobustBoost algorithm can make good classification predictions even when the training data has noise. However, the default RobustBoost parameters can produce an ensemble that does not predict well. This example shows one way of tuning the parameters for better predictive accuracy.

1 Generate data with label noise. This example has twenty uniform random numbers per observation, and classifies the observation as 1 if the sum of the first five numbers exceeds 2.5 (so is larger than average), and 0 otherwise:

```
rng(0,'twister') % for reproducibility
Xtrain = rand(2000,20);
Ytrain = sum(Xtrain(:,1:5),2) > 2.5;
```

2 To add noise, randomly switch $10 \%$ of the classifications:

```
idx = randsample(2000,200);
Ytrain(idx) = ~Ytrain(idx);
```

3 Create an ensemble with AdaBoostM1 for comparison purposes:

```
ada = fitensemble(Xtrain,Ytrain,'AdaBoostM1',...
    300,'Tree','LearnRate',0.1);
```

4 Create an ensemble with RobustBoost. Since the data has $10 \%$ incorrect classification, perhaps an error goal of $15 \%$ is reasonable.

```
rb1 = fitensemble(Xtrain,Ytrain,'RobustBoost',300,...
    'Tree', 'RobustErrorGoal',0.15,'RobustMaxMargin',1);
```

5 Try setting a high value of the error goal, 0.6. You get an error:
rb2 = fitensemble(Xtrain, Ytrain, 'RobustBoost', 300,'Tree', 'RobustErrorGoal', 0.6)

Error using RobustBoost>RobustBoost.RobustBoost
For the chosen values of 'RobustMaxMargin' and 'RobustMarginSigma', you must set 'RobustErrorGoal' to a value between 0 and 0.5.

6 Create an ensemble with an error goal in the allowed range, 0.4:

```
rb2 = fitensemble(Xtrain,Ytrain,'RobustBoost',300,...
    'Tree', 'RobustErrorGoal', 0.4);
```

7 Create an ensemble with very optimistic error goal, 0.01:

```
rb3 = fitensemble(Xtrain,Ytrain,'RobustBoost',300,...
    'Tree','RobustErrorGoal',0.01);
```

8 Compare the resubstitution error of the four ensembles:
figure

```
plot(resubLoss(rb1,'mode','cumulative'));
hold on
plot(resubLoss(rb2,'mode','cumulative'),'r--');
plot(resubLoss(rb3,'mode','cumulative'),'k-.');
plot(resubLoss(ada,'mode','cumulative'),'g.');
hold off;
xlabel('Number of trees');
ylabel('Resubstitution error');
legend('ErrorGoal=0.15','ErrorGoal=0.4','ErrorGoal=0.01' ,...
'AdaBoostM1','Location', 'NE');
```



All the RobustBoost curves show lower resubstitution error than the AdaBoostM1 curve. The error goal of 0.15 curve shows the lowest
resubstitution error over most of the range. However, its error is rising in the latter half of the plot, while the other curves are still descending.

9 Generate test data to see the predictive power of the ensembles. Test the four ensembles:

```
Xtest = rand(2000,20);
Ytest = sum(Xtest(:,1:5),2) > 2.5;
idx = randsample(2000,200);
Ytest(idx) = ~Ytest(idx);
figure;
plot(loss(rb1,Xtest,Ytest,'mode','cumulative'));
hold on
plot(loss(rb2,Xtest,Ytest,'mode','cumulative'),'r--');
plot(loss(rb3,Xtest,Ytest,'mode','cumulative'),'k-.');
plot(loss(ada,Xtest,Ytest,'mode','cumulative'),'g.');
hold off;
xlabel('Number of trees');
ylabel('Test error');
legend('ErrorGoal=0.15','ErrorGoal=0.4','ErrorGoal=0.01',...
    'AdaBoostM1','Location','NE');
```



The error curve for error goal 0.15 is lowest (best) in the plotted range. The curve for error goal 0.4 seems to be converging to a similar value for a large number of trees, but more slowly. AdaBoostM1 has higher error than the curve for error goal 0.15 . The curve for the too-optimistic error goal 0.01 remains substantially higher (worse) than the other algorithms for most of the plotted range.

## TreeBagger Examples

TreeBagger ensembles have more functionality than those constructed with fitensemble; see TreeBagger Features Not in fitensemble on page 13-120. Also, some property and method names differ from their fitensemble
counterparts. This section contains examples of workflow for regression and classification that use this extra TreeBagger functionality.

## Workflow Example: Regression of Insurance Risk Rating for Car Imports with TreeBagger

In this example, use a database of 1985 car imports with 205 observations, 25 input variables, and one response variable, insurance risk rating, or "symboling." The first 15 variables are numeric and the last 10 are categorical. The symboling index takes integer values from -3 to 3 .

1 Load the dataset and split it into predictor and response arrays:

```
load imports-85;
Y = X(:,1);
X = X(:,2:end);
```

2 Because bagging uses randomized data drawings, its exact outcome depends on the initial random seed. To reproduce the exact results in this example, use the random stream settings:

```
rng(1945,'twister')
```

Finding the Optimal Leaf Size. For regression, the general rule is to set leaf size to 5 and select one third of input features for decision splits at random. In the following step, verify the optimal leaf size by comparing mean-squared errors obtained by regression for various leaf sizes. oobError computes MSE versus the number of grown trees. You must set oobpred to ' on' to obtain out-of-bag predictions later.

```
leaf = [1 5 10 20 50 100];
col = 'rgbcmy';
figure(1);
for i=1:length(leaf)
    b = TreeBagger(50,X,Y,'method','r','oobpred','on',\ldots.
    'cat',16:25,'minleaf',leaf(i));
    plot(oobError(b),col(i));
    hold on;
end
xlabel('Number of Grown Trees');
ylabel('Mean Squared Error');
```

```
legend({'1' '5' '10' '20' '50' '100'},'Location','NorthEast');
hold off;
```



The red (leaf size 1) curve gives the lowest MSE values.

## Estimating Feature Importance.

1 In practical applications, you typically grow ensembles with hundreds of trees. Only 50 trees were used in "Finding the Optimal Leaf Size" on page 13-97 for faster processing. Now that you have estimated the optimal leaf size, grow a larger ensemble with 100 trees and use it for estimation of feature importance:

```
b = TreeBagger(100,X,Y,'method','r','oobvarimp','on',...
'cat',16:25,'minleaf',1);
```

2 Inspect the error curve again to make sure nothing went wrong during training:

```
figure(2);
plot(oobError(b));
xlabel('Number of Grown Trees');
ylabel('Out-of-Bag Mean Squared Error');
```



Prediction ability should depend more on important features and less on unimportant features. You can use this idea to measure feature importance.

For each feature, you can permute the values of this feature across all of the observations in the data set and measure how much worse the mean-squared error (MSE) becomes after the permutation. You can repeat this for each feature.

1 Using the following code, plot the increase in MSE due to permuting out-of-bag observations across each input variable. The OOBPermutedVarDeltaError array stores the increase in MSE averaged over all trees in the ensemble and divided by the standard deviation taken over the trees, for each variable. The larger this value, the more important the variable. Imposing an arbitrary cutoff at 0.65 , you can select the five most important features.

```
figure(3);
bar(b.00BPermutedVarDeltaError);
xlabel('Feature Number');
ylabel('Out-Of-Bag Feature Importance');
idxvar = find(b.O0BPermutedVarDeltaError>0.65)
idxvar =
```



2 The OOBIndices property of TreeBagger tracks which observations are out of bag for what trees. Using this property, you can monitor the fraction of observations in the training data that are in bag for all trees. The curve starts at approximately $2 / 3$, the fraction of unique observations selected by one bootstrap replica, and goes down to 0 at approximately 10 trees.

```
finbag = zeros(1,b.NTrees);
for t=1:b.NTrees
    finbag(t) = sum(all(~b.00BIndices(:,1:t),2));
end
finbag = finbag / size(X,1);
```

```
figure(4);
plot(finbag);
xlabel('Number of Grown Trees');
ylabel('Fraction of in-Bag Observations');
```



Growing Trees on a Reduced Set of Features. Using just the five most powerful features selected in "Estimating Feature Importance" on page 13-98, determine if it is possible to obtain a similar predictive power. To begin, grow 100 trees on these features only. The first three of the five selected features are numeric and the last two are categorical.

```
b5v = TreeBagger(100,X(:,idxvar),Y,'method','r',...
'oobvarimp','on','cat',4:5,'minleaf',1);
```

```
figure(5);
plot(oobError(b5v));
xlabel('Number of Grown Trees');
ylabel('Out-of-Bag Mean Squared Error');
figure(6);
bar(b5v.00BPermutedVarDeltaError);
xlabel('Feature Index');
ylabel('Out-of-Bag Feature Importance');
```




These five most powerful features give the same MSE as the full set, and the ensemble trained on the reduced set ranks these features similarly to each other. Features 1 and 2 from the reduced set perhaps could be removed without a significant loss in the predictive power.

Finding Outliers. To find outliers in the training data, compute the proximity matrix using fillProximities:

```
b5v = fillProximities(b5v);
```

The method normalizes this measure by subtracting the mean outlier measure for the entire sample, taking the magnitude of this difference and dividing the result by the median absolute deviation for the entire sample:

```
figure(7);
hist(b5v.OutlierMeasure);
xlabel('Outlier Measure');
ylabel('Number of Observations');
```



Discovering Clusters in the Data. By applying multidimensional scaling to the computed matrix of proximities, you can inspect the structure of the input data and look for possible clusters of observations. The mdsProx method returns scaled coordinates and eigenvalues for the computed proximity matrix. If run with the colors option, this method makes a scatter plot of two scaled coordinates, first and second by default.

```
figure(8);
[~,e] = mdsProx(b5v,'colors','k');
xlabel('1st Scaled Coordinate');
ylabel('2nd Scaled Coordinate');
```



Assess the relative importance of the scaled axes by plotting the first 20 eigenvalues:

```
figure(9);
bar(e(1:20));
xlabel('Scaled Coordinate Index');
ylabel('Eigenvalue');
```



Saving the Ensemble Configuration for Future Use. To use the trained ensemble for predicting the response on unseen data, store the ensemble to disk and retrieve it later. If you do not want to compute predictions for out-of-bag data or reuse training data in any other way, there is no need to store the ensemble object itself. Saving the compact version of the ensemble would be enough in this case. Extract the compact object from the ensemble:

```
c = compact(b5v)
C =
Ensemble with 100 decision trees:
Method: regression
Nvars:
    5
```

This object can be now saved in a *.mat file as usual.

## Workflow Example: Classifying Radar Returns for Ionosphere Data with TreeBagger

You can also use ensembles of decision trees for classification. For this example, use ionosphere data with 351 observations and 34 real-valued predictors. The response variable is categorical with two levels:

- 'g' for good radar returns
- 'b' for bad radar returns

The goal is to predict good or bad returns using a set of 34 measurements. The workflow resembles that for "Workflow Example: Regression of Insurance Risk Rating for Car Imports with TreeBagger" on page 13-97.

1 Fix the initial random seed, grow 50 trees, inspect how the ensemble error changes with accumulation of trees, and estimate feature importance. For classification, it is best to set the minimal leaf size to 1 and select the square root of the total number of features for each decision split at random. These are the default settings for a TreeBagger used for classification.

```
load ionosphere;
rng(1945,'twister')
b = TreeBagger(50,X,Y,'oobvarimp','on');
```

```
figure(10);
plot(oobError(b));
xlabel('Number of Grown Trees');
ylabel('Out-of-Bag Classification Error');
```



2 The method trains ensembles with few trees on observations that are in bag for all trees. For such observations, it is impossible to compute the true out-of-bag prediction, and TreeBagger returns the most probable class for classification and the sample mean for regression. You can change the default value returned for in-bag observations using the DefaultYfit property. If you set the default value to an empty string for classification, the method excludes in-bag observations from computation of the out-of-bag error. In this case, the curve is more variable when the number of trees is small, either because some observations are never out of bag (and are therefore excluded) or because their predictions are based on few trees.

```
b.DefaultYfit = '';
figure(11);
plot(oobError(b));
xlabel('Number of Grown Trees');
ylabel('Out-of-Bag Error Excluding in-Bag Observations');
```



3 The OOBIndices property of TreeBagger tracks which observations are out of bag for what trees. Using this property, you can monitor the fraction of observations in the training data that are in bag for all trees. The curve starts at approximately $2 / 3$, the fraction of unique observations selected by one bootstrap replica, and goes down to 0 at approximately 10 trees.

```
finbag = zeros(1,b.NTrees);
for t=1:b.NTrees
    finbag(t) = sum(all(~b.OOBIndices(:,1:t),2));
end
finbag = finbag / size(X,1);
figure(12);
plot(finbag);
xlabel('Number of Grown Trees');
ylabel('Fraction of in-Bag Observations');
```



4 Estimate feature importance:
figure(13);
bar(b.00BPermutedVarDeltaError);
xlabel('Feature Index');
ylabel('Out-of-Bag Feature Importance'); idxvar $=$ find(b.OOBPermutedVarDeltaError>0.8)
idxvar =
$\begin{array}{lllll}3 & 4 & 5 & 7 & 8\end{array}$


5 Having selected the five most important features, grow a larger ensemble on the reduced feature set. Save time by not permuting out-of-bag observations to obtain new estimates of feature importance for the reduced feature set (set oobvarimp to 'off'). You would still be interested in obtaining out-of-bag estimates of classification error (set oobpred to 'on ').

```
b5v = TreeBagger(100,X(:,idxvar),Y,'oobpred','on');
figure(14);
plot(oobError(b5v));
xlabel('Number of Grown Trees');
ylabel('Out-of-Bag Classification Error');
```



6 For classification ensembles, in addition to classification error (fraction of misclassified observations), you can also monitor the average classification margin. For each observation, the margin is defined as the difference between the score for the true class and the maximal score for other classes predicted by this tree. The cumulative classification margin uses the scores averaged over all trees and the mean cumulative classification margin is the cumulative margin averaged over all observations. The oobMeanMargin method with the 'mode' argument set to 'cumulative' (default) shows how the mean cumulative margin changes as the ensemble grows: every new element in the returned array represents the cumulative margin obtained by including a new tree in the ensemble. If training is successful, you would expect to see a gradual increase in the mean classification margin.

For decision trees, a classification score is the probability of observing an instance of this class in this tree leaf. For example, if the leaf of a grown decision tree has five 'good' and three 'bad ' training observations in it, the scores returned by this decision tree for any observation fallen on this leaf are $5 / 8$ for the 'good' class and $3 / 8$ for the 'bad' class. These probabilities are called 'scores' for consistency with other classifiers that might not have an obvious interpretation for numeric values of returned predictions.
figure(15);
plot(oobMeanMargin(b5v));

```
xlabel('Number of Grown Trees');
ylabel('Out-of-Bag Mean Classification Margin');
```



7 Compute the matrix of proximities and look at the distribution of outlier measures. Unlike regression, outlier measures for classification ensembles are computed within each class separately.

```
b5v = fillProximities(b5v);
figure(16);
hist(b5v.OutlierMeasure);
xlabel('Outlier Measure');
ylabel('Number of Observations');
```



8 All extreme outliers for this dataset come from the 'good ' class:

```
b5v.Y(b5v.OutlierMeasure>40)
ans =
    'g'
    'g'
    'g'
    'g'
    'g''
```

9 As for regression, you can plot scaled coordinates, displaying the two classes in different colors using the colors argument of mdsProx. This argument takes a string in which every character represents a color. To find the order of classes used by the ensemble, look at the ClassNames property:

```
b5v.ClassNames
ans =
    'g'
    'b'
```

The 'good' class is first and the 'bad' class is second. Display scaled coordinates using red for 'good ' and blue for 'bad' observations:

```
figure(17);
[s,e] = mdsProx(b5v,'colors','rb');
xlabel('1st Scaled Coordinate');
ylabel('2nd Scaled Coordinate');
```



10 Plot the first 20 eigenvalues obtained by scaling. The first eigenvalue in this case clearly dominates and the first scaled coordinate is most important.

```
figure(18);
bar(e(1:20));
xlabel('Scaled Coordinate Index');
ylabel('Eigenvalue');
```



Plotting a Classification Performance Curve. Another way of exploring the performance of a classification ensemble is to plot its Receiver Operating Characteristic (ROC) curve or another performance curve suitable for the current problem. First, obtain predictions for out-of-bag observations. For a classification ensemble, the oobPredict method returns a cell array of classification labels (' $g$ ' or ' $b$ ' for ionosphere data) as the first output argument and a numeric array of scores as the second output argument. The returned array of scores has two columns, one for each class. In this case, the first column is for the 'good ' class and the second column is for the 'bad ' class. One column in the score matrix is redundant because the scores represent class probabilities in tree leaves and by definition add up to 1 .

```
[Yfit,Sfit] = oobPredict(b5v);
```

Use the perfcurve utility (see "Performance Curves" on page 12-32) to compute a performance curve. By default, perfcurve returns the standard ROC curve, which is the true positive rate versus the false positive rate. perfcurve requires true class labels, scores, and the positive class label for input. In this case, choose the 'good' class as positive. The scores for this class are in the first column of Sfit.

```
[fpr,tpr] = perfcurve(b5v.Y,Sfit(:,1),'g');
figure(19);
plot(fpr,tpr);
```

```
xlabel('False Positive Rate');
ylabel('True Positive Rate');
```



Instead of the standard ROC curve, you might want to plot, for example, ensemble accuracy versus threshold on the score for the 'good' class. The ycrit input argument of perfcurve lets you specify the criterion for the $y$-axis, and the third output argument of perfcurve returns an array of thresholds for the positive class score. Accuracy is the fraction of correctly classified observations, or equivalently, 1 minus the classification error.

```
[fpr,accu,thre] = perfcurve(b5v.Y,Sfit(:,1),'g','ycrit','accu');
figure(20);
plot(thre,accu);
xlabel('Threshold for ''good'' Returns');
ylabel('Classification Accuracy');
```



The curve shows a flat region indicating that any threshold from 0.2 to 0.6 is a reasonable choice. By default, the function assigns classification labels using 0.5 as the boundary between the two classes. You can find exactly what accuracy this corresponds to:

```
i50 = find(accu>=0.50,1,'first')
accu(abs(thre-0.5)<eps)
```


## returns

$i 50=$
2
ans =
0.9430

The maximal accuracy is a little higher than the default one:

```
[maxaccu,iaccu] = max(accu)
```

returns

```
iaccu =
    91
```

The optimal threshold is therefore:

```
thre(iaccu)
ans =
    0.5056
```


## Ensemble Algorithms

- "Bagging" on page 13-118
- "AdaBoostM1" on page 13-122
- "AdaBoostM2" on page 13-124
- "LogitBoost" on page 13-125
- "GentleBoost" on page 13-126
- "RobustBoost" on page 13-127
- "LSBoost" on page 13-128


## Bagging

Bagging, which stands for "bootstrap aggregation", is a type of ensemble learning. To bag a weak learner such as a decision tree on a dataset, generate many bootstrap replicas of this dataset and grow decision trees on these replicas. Obtain each bootstrap replica by randomly selecting N observations out of $N$ with replacement, where $N$ is the dataset size. To find the predicted response of a trained ensemble, take an average over predictions from individual trees.

Bagged decision trees were introduced in MATLAB R2009a as TreeBagger. The fitensemble function lets you bag in a manner consistent with boosting. An ensemble of bagged trees, either ClassificationBaggedEnsemble or RegressionBaggedEnsemble, returned by fitensemble offers almost the same functionally as TreeBagger. Discrepancies between TreeBagger and
the new framework are described in detail in TreeBagger Features Not in fitensemble on page 13-120.

Bagging works by training learners on resampled versions of the data. This resampling is usually done by bootstrapping observations, that is, selecting $N$ out of $N$ observations with replacement for every new learner. In addition, every tree in the ensemble can randomly select predictors for decision splits-a technique known to improve the accuracy of bagged trees.

By default, the minimal leaf sizes for bagged trees are set to 1 for classification and 5 for regression. Trees grown with the default leaf size are usually very deep. These settings are close to optimal for the predictive power of an ensemble. Often you can grow trees with larger leaves without losing predictive power. Doing so reduces training and prediction time, as well as memory usage for the trained ensemble.

Another important parameter is the number of predictors selected at random for every decision split. This random selection is made for every split, and every deep tree involves many splits. By default, this parameter is set to a square root of the number of predictors for classification, and one third of predictors for regression.

Several features of bagged decision trees make them a unique algorithm. Drawing N out of N observations with replacement omits on average $37 \%$ of observations for each decision tree. These are "out-of-bag" observations. You can use them to estimate the predictive power and feature importance. For each observation, you can estimate the out-of-bag prediction by averaging over predictions from all trees in the ensemble for which this observation is out of bag. You can then compare the computed prediction against the observed response for this observation. By comparing the out-of-bag predicted responses against the observed responses for all observations used for training, you can estimate the average out-of-bag error. This out-of-bag average is an unbiased estimator of the true ensemble error. You can also obtain out-of-bag estimates of feature importance by randomly permuting out-of-bag data across one variable or column at a time and estimating the increase in the out-of-bag error due to this permutation. The larger the increase, the more important the feature. Thus, you need not supply test data for bagged ensembles because you obtain reliable estimates of the predictive power and feature importance in the process of training, which is an attractive feature of bagging.

Another attractive feature of bagged decision trees is the proximity matrix. Every time two observations land on the same leaf of a tree, their proximity increases by 1 . For normalization, sum these proximities over all trees in the ensemble and divide by the number of trees. The resulting matrix is symmetric with diagonal elements equal to 1 and off-diagonal elements ranging from 0 to 1 . You can use this matrix for finding outlier observations and discovering clusters in the data through multidimensional scaling.

For examples using bagging, see:

- "Example: Test Ensemble Quality" on page 13-60
- "Example: Surrogate Splits" on page 13-77
- "Workflow Example: Regression of Insurance Risk Rating for Car Imports with TreeBagger" on page 13-97
- "Workflow Example: Classifying Radar Returns for Ionosphere Data with TreeBagger" on page 13-106

For references related to bagging, see Breiman [1], [2], and [3].
Comparison of TreeBagger and Bagged Ensembles. fitensemble produces bagged ensembles that have most, but not all, of the functionality of TreeBagger objects. Additionally, some functionality has different names in the new bagged ensembles.

## TreeBagger Features Not in fitensemble

| Feature | TreeBagger Property | TreeBagger Method |
| :--- | :--- | :--- |
| Computation of proximity <br> matrix | Proximity | fillProximities, mdsProx |
| Computation of outliers | OutlierMeasure |  |
| Out-of-bag estimates of <br> predictor importance | OOBPermutedVarDeltaError, <br> OOBPermutedVarDeltaMeanMargin, <br> OOBPermutedVarCountRaiseMargin |  |

## TreeBagger Features Not in fitensemble (Continued)

| Feature | TreeBagger Property | TreeBagger Method |
| :--- | :--- | :--- |
| Merging two ensembles <br> trained separately |  | append |
| Parallel computation for <br> creating ensemble | Set the UseParallel name-value <br> pair to 'always'; see Chapter 17, <br> "Parallel Statistics" |  |

## Differing Names Between TreeBagger and Bagged Ensembles

| Feature | TreeBagger | Bagged Ensembles |
| :--- | :--- | :--- |
| Split criterion contributions <br> for each predictor | DeltaCritDecisionSplit <br> property | First output of <br> predictorImportance <br> (classification) or <br> predictorImportance <br> (regression) |
| Predictor associations | VarAssoc property | Second output of <br> predictorImportance <br> (classification) or <br> predictorImportance <br> (regression) |
| Error (misclassification <br> probability or mean-squared <br> error) | error and oobError methods | loss and oobLoss methods <br> (classification); loss and <br> oobLoss methods (regression) |
| Train additional trees and add <br> to ensemble | growTrees method | resume method (classification); <br> resume method (regression) |
| Mean classification margin <br> per tree | meanMargin and <br> oobMeanMargin methods | edge and oobEdge methods <br> (classification); |

In addition, two important changes were made to training and prediction for bagged classification ensembles:

- If you pass a misclassification cost matrix to TreeBagger, it passes this matrix along to the trees. If you pass a misclassification cost matrix to fitensemble, it uses this matrix to adjust the class prior probabilities.
fitensemble then passes the adjusted prior probabilities and the default cost matrix to the trees. The default cost matrix is ones (K) -eye (K) for K classes.
- Unlike the loss and edge methods in the new framework, the TreeBagger error and meanMargin methods do not normalize input observation weights of the prior probabilities in the respective class.


## AdaBoostM1

AdaBoostM1 is a very popular boosting algorithm for binary classification. The algorithm trains learners sequentially. For every learner with index $t$, AdaBoostM1 computes the weighted classification error

$$
\varepsilon_{t}=\sum_{n=1}^{N} d_{n}^{(t)} \mathrm{I}\left(y_{n}=h_{t}\left(x_{n}\right)\right)
$$

- $x_{n}$ is a vector of predictor values for observation $n$.
- $y_{n}$ is the true class label.
- $h_{t}$ is the prediction of learner (hypothesis) with index $t$.
- I is the indicator function.
- $d_{n}^{(t)}$ is the weight of observation $n$ at step $t$.

AdaBoostM1 then increases weights for observations misclassified by learner $t$ and reduces weights for observations correctly classified by learner $t$. The next learner $t+1$ is then trained on the data with updated weights $d_{n}^{(t+1)}$.

After training finishes, AdaBoostM1 computes prediction for new data using

$$
f(x)=\sum_{t=1}^{T} \alpha_{t} h_{t}(x)
$$

where

$$
\alpha_{t}=\frac{1}{2} \log \frac{1-\varepsilon_{t}}{\varepsilon_{t}}
$$

are weights of the weak hypotheses in the ensemble.
Training by AdaBoostM1 can be viewed as stagewise minimization of the exponential loss:

$$
\sum_{n=1}^{N} w_{n} \exp \left(-y_{n} f\left(x_{n}\right)\right)
$$

where

- $y_{n} \in\{-1,+1\}$ is the true class label.
- $w_{n}$ are observation weights normalized to add up to 1 .
- $f\left(x_{n}\right) \in(-\infty,+\infty)$ is the predicted classification score.

The observation weights $w_{n}$ are the original observation weights you passed to fitensemble.

The second output from the predict method of an AdaBoostM1 classification ensemble is an $N$-by- 2 matrix of classification scores for the two classes and $N$ observations. The second column in this matrix is always equal to minus the first column. predict returns two scores to be consistent with multiclass models, though this is redundant since the second column is always the negative of the first.

Most often AdaBoostM1 is used with decision stumps (default) or shallow trees. If boosted stumps give poor performance, try setting the minimal parent node size to one quarter of the training data.

By default, the learning rate for boosting algorithms is 1. If you set the learning rate to a lower number, the ensemble learns at a slower rate, but can converge to a better solution. 0.1 is a popular choice for the learning rate. Learning at a rate less than 1 is often called "shrinkage".

For examples using AdaBoostM1, see "Example: Tuning RobustBoost" on page 13-92.

For references related to AdaBoostM1, see Freund and Schapire [6], Schapire et al. [10], Friedman, Hastie, and Tibshirani [8], and Friedman [7].

## AdaBoostM2

AdaBoostM2 is an extension of AdaBoostM1 for multiple classes. Instead of weighted classification error, AdaBoostM2 uses weighted pseudo-loss for $N$ observations and $K$ classes:

$$
\varepsilon_{t}=\frac{1}{2} \sum_{n=1}^{N} \sum_{k \neq y_{n}} d_{n, k}^{(t)}\left(1-h_{t}\left(x_{n}, y_{n}\right)+h_{t}\left(x_{n}, k\right)\right)
$$

where

- $h_{t}\left(x_{n}, k\right)$ is the confidence of prediction by learner at step $t$ into class $k$ ranging from 0 (not at all confident) to 1 (highly confident).
- $d_{n, k}^{(t)}$ are observation weights at step $t$ for class $k$.
- $y_{n}$ is the true class label taking one of the $K$ values.
- The second sum is over all classes other than the true class $y_{n}$.

Interpreting the pseudo-loss is harder than classification error, but the idea is the same. Pseudo-loss can be used as a measure of the classification accuracy from any learner in an ensemble. Pseudo-loss typically exhibits the same behavior as a weighted classification error for AdaBoostM1: the first few learners in a boosted ensemble give low pseudo-loss values. After the first few training steps, the ensemble begins to learn at a slower pace, and the pseudo-loss value approaches 0.5 from below.

For examples using AdaBoostM2, see "Creating a Classification Ensemble" on page 13-58.

For references related to AdaBoostM2, see Freund and Schapire [6].

## LogitBoost

LogitBoost is another popular algorithm for binary classification.
LogitBoost works similarly to AdaBoostM1, except it minimizes binomial deviance

$$
\sum_{n=1}^{N} w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(x_{n}\right)\right)\right)
$$

where

- $y_{n} \epsilon\{-1,+1\}$ is the true class label.
- $w_{n}$ are observation weights normalized to add up to 1 .
- $f\left(x_{n}\right) \in(-\infty,+\infty)$ is the predicted classification score.

Binomial deviance assigns less weight to badly misclassified observations (observations with large negative values of $y_{n} f\left(x_{n}\right)$ ). LogitBoost can give better average accuracy than AdaBoostM1 for data with poorly separable classes.

Learner $t$ in a LogitBoost ensemble fits a regression model to response values

$$
\tilde{y}_{n}=\frac{y_{n}^{*}-p_{t}\left(x_{n}\right)}{p_{t}\left(x_{n}\right)\left(1-p_{t}\left(x_{n}\right)\right)}
$$

where

- $y^{*}{ }_{n} \epsilon\{0,+1\}$ are relabeled classes ( 0 instead of -1 ).
- $p_{t}\left(x_{n}\right)$ is the current ensemble estimate of the probability for observation $x_{n}$ to be of class 1 .

Fitting a regression model at each boosting step turns into a great computational advantage for data with multilevel categorical predictors. Take a categorical predictor with $L$ levels. To find the optimal decision split on such a predictor, classification tree needs to consider $2^{L-1}-1$ splits. A regression tree needs to consider only $L-1$ splits, so the processing time
can be much shorter. LogitBoost is recommended for categorical predictors with many levels.
fitensemble computes and stores the mean-squared error

$$
\sum_{n=1}^{N} d_{n}^{(t)}\left(\tilde{y}_{n}-h_{t}\left(x_{n}\right)\right)^{2}
$$

in the FitInfo property of the ensemble object. Here

- $d_{n}^{(t)}$ are observation weights at step $t$ (the weights add up to 1 ).
- $h_{t}\left(x_{n}\right)$ are predictions of the regression model $h_{t}$ fitted to response values $\tilde{y}_{n}$.

Values $y_{n}$ can range from $-\infty$ to $+\infty$, so the mean-squared error does not have well-defined bounds.

For examples using LogitBoost, see "Example: Classification with Many Categorical Levels" on page 13-72.

For references related to LogitBoost, see Friedman, Hastie, and Tibshirani [8].

## GentleBoost

GentleBoost (also known as Gentle AdaBoost) combines features of AdaBoostM1 and LogitBoost. Like AdaBoostM1, GentleBoost minimizes the exponential loss. But its numeric optimization is set up differently. Like LogitBoost, every weak learner fits a regression model to response values $y_{n} \in\{-1,+1\}$. This makes GentleBoost another good candidate for binary classification of data with multilevel categorical predictors.
fitensemble computes and stores the mean-squared error

$$
\sum_{n=1}^{N} d_{n}^{(t)}\left(\tilde{y}_{n}-h_{t}\left(x_{n}\right)\right)^{2}
$$

in the FitInfo property of the ensemble object, where

- $d_{n}^{(t)}$ are observation weights at step $t$ (the weights add up to 1 ).
- $h_{t}\left(x_{n}\right)$ are predictions of the regression model $h_{t}$ fitted to response values $y_{n}$.

As the strength of individual learners weakens, the weighted mean-squared error approaches 1 .

For examples using GentleBoost, see "Example: Unequal Classification Costs" on page 13-67, "Example: Classification with Many Categorical Levels" on page 13-72.

For references related to GentleBoost, see Friedman, Hastie, and Tibshirani [8].

## RobustBoost

Boosting algorithms such as AdaBoostM1 and LogitBoost increase weights for misclassified observations at every boosting step. These weights can become very large. If this happens, the boosting algorithm sometimes concentrates on a few misclassified observations and neglects the majority of training data. Consequently the average classification accuracy suffers.

In this situation, you can try using RobustBoost. This algorithm does not assign almost the entire data weight to badly misclassified observations. It can produce better average classification accuracy.

Unlike AdaBoostM1 and LogitBoost, RobustBoost does not minimize a specific loss function. Instead, it maximizes the number of observations with the classification margin above a certain threshold.

RobustBoost trains based on time evolution. The algorithm starts at $t=0$. At every step, RobustBoost solves an optimization problem to find a positive step in time $\Delta t$ and a corresponding positive change in the average margin for training data $\Delta m$. RobustBoost stops training and exits if at least one of these three conditions is true:

- Time $t$ reaches 1.
- RobustBoost cannot find a solution to the optimization problem with positive updates $\Delta t$ and $\Delta m$.
- RobustBoost grows as many learners as you requested.

Results from RobustBoost can be usable for any termination condition. Estimate the classification accuracy by cross validation or by using an independent test set.

To get better classification accuracy from RobustBoost, you can adjust three parameters in fitensemble: RobustErrorGoal, RobustMaxMargin, and RobustMarginSigma. Start by varying values for RobustErrorGoal from 0 to 1. The maximal allowed value for RobustErrorGoal depends on the two other parameters. If you pass a value that is too high, fitensemble produces an error message showing the allowed range for RobustErrorGoal.

For examples using RobustBoost, see "Example: Tuning RobustBoost" on page 13-92

For references related to RobustBoost, see Freund [5].

## LSBoost

You can use least squares boosting (LSBoost) to fit regression ensembles. At every step, the ensemble fits a new learner to the difference between the observed response and the aggregated prediction of all learners grown previously. The ensemble fits to minimize mean-squared error.

You can use LSBoost with shrinkage by passing in LearnRate parameter. By default this parameter is set to 1 , and the ensemble learns at the maximal speed. If you set LearnRate to a value from 0 to 1 , the ensemble fits every new learner to $y_{n}-\eta f\left(x_{n}\right)$, where

- $y_{n}$ is the observed response.
- $f\left(x_{n}\right)$ is the aggregated prediction from all weak learners grown so far for observation $x_{n}$.
- $\eta$ is the learning rate.

For examples using LSBoost, see "Creating a Regression Ensemble" on page 13-59, "Example: Regularizing a Regression Ensemble" on page 13-82

For references related to LSBoost, see Hastie, Tibshirani, and Friedman [9], Chapters 7 (Model Assessment and Selection) and 15 (Random Forests).

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## Markov Models

- "Introduction to Markov Models" on page 14-2
- "Markov Chains" on page 14-3
- "Hidden Markov Models (HMM)" on page 14-5


## Introduction to Markov Models

Markov processes are examples of stochastic processes-processes that generate random sequences of outcomes or states according to certain probabilities. Markov processes are distinguished by being memoryless-their next state depends only on their current state, not on the history that led them there. Models of Markov processes are used in a wide variety of applications, from daily stock prices to the positions of genes in a chromosome.

## Markov Chains

A Markov model is given visual representation with a state diagram, such as the one below.


## State Diagram for a Markov Model

The rectangles in the diagram represent the possible states of the process you are trying to model, and the arrows represent transitions between states. The label on each arrow represents the probability of that transition. At each step of the process, the model may generate an output, or emission, depending on which state it is in, and then make a transition to another state. An important characteristic of Markov models is that the next state depends only on the current state, and not on the history of transitions that lead to the current state.

For example, for a sequence of coin tosses the two states are heads and tails. The most recent coin toss determines the current state of the model and each subsequent toss determines the transition to the next state. If the coin is fair, the transition probabilities are all $1 / 2$. The emission might simply be the current state. In more complicated models, random processes at each state will generate emissions. You could, for example, roll a die to determine the emission at any step.

Markov chains are mathematical descriptions of Markov models with a discrete set of states. Markov chains are characterized by:

- A set of states $\{1,2, \ldots, M\}$
- An $M$-by- $M$ transition matrix $T$ whose $i, j$ entry is the probability of a transition from state $i$ to state $j$. The sum of the entries in each row of $T$ must be 1, because this is the sum of the probabilities of making a transition from a given state to each of the other states.
- A set of possible outputs, or emissions, $\left\{s_{1}, s_{2}, \ldots, s_{N}\right\}$. By default, the set of emissions is $\{1,2, \ldots, N\}$, where $N$ is the number of possible emissions, but you can choose a different set of numbers or symbols.
- An $M$-by- $N$ emission matrix $E$ whose $i, k$ entry gives the probability of emitting symbol $s_{\mathrm{k}}$ given that the model is in state $i$.

Markov chains begin in an initial state $i_{0}$ at step 0 . The chain then transitions to state $i_{1}$ with probability $T_{1 i_{1}}$, and emits an output $s_{k_{1}}$ with probability
$E_{i_{1} k_{1}}$. Consequently, the probability of observing the sequence of states $i_{1} i_{2} \ldots i_{r}$ and the sequence of emissions $s_{k_{1}} s_{k_{2}} \ldots s_{k_{r}}$ in the first $r$ steps, is

$$
T_{1 i_{1}} E_{i_{1} k_{1}} T_{i_{1} i_{2}} E_{i_{2} k_{2}} \ldots T_{i_{r-1} i_{r}} E_{i_{r} k}
$$

## Hidden Markov Models (HMM)

In this section...<br>"Introduction to Hidden Markov Models (HMM)" on page 14-5<br>"Analyzing Hidden Markov Models" on page 14-7

## Introduction to Hidden Markov Models (HMM)

A hidden Markov model (HMM) is one in which you observe a sequence of emissions, but do not know the sequence of states the model went through to generate the emissions. Analyses of hidden Markov models seek to recover the sequence of states from the observed data.

As an example, consider a Markov model with two states and six possible emissions. The model uses:

- A red die, having six sides, labeled 1 through 6.
- A green die, having twelve sides, five of which are labeled 2 through 6, while the remaining seven sides are labeled 1.
- A weighted red coin, for which the probability of heads is .9 and the probability of tails is .1.
- A weighted green coin, for which the probability of heads is .95 and the probability of tails is .05 .

The model creates a sequence of numbers from the set $\{1,2,3,4,5,6\}$ with the following rules:

- Begin by rolling the red die and writing down the number that comes up, which is the emission.
- Toss the red coin and do one of the following:
- If the result is heads, roll the red die and write down the result.
- If the result is tails, roll the green die and write down the result.
- At each subsequent step, you flip the coin that has the same color as the die you rolled in the previous step. If the coin comes up heads, roll the same die as in the previous step. If the coin comes up tails, switch to the other die.

The state diagram for this model has two states, red and green, as shown in the following figure.


You determine the emission from a state by rolling the die with the same color as the state. You determine the transition to the next state by flipping the coin with the same color as the state.

The transition matrix is:

$$
T=\left[\begin{array}{ll}
0.9 & 0.1 \\
0.05 & 0.95
\end{array}\right]
$$

The emissions matrix is:

$$
E=\left[\begin{array}{cccccc}
\frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} \\
\frac{7}{12} & \frac{1}{12} & \frac{1}{12} & \frac{1}{12} & \frac{1}{12} & \frac{1}{12}
\end{array}\right]
$$

The model is not hidden because you know the sequence of states from the colors of the coins and dice. Suppose, however, that someone else is generating
the emissions without showing you the dice or the coins. All you see is the sequence of emissions. If you start seeing more 1 s than other numbers, you might suspect that the model is in the green state, but you cannot be sure because you cannot see the color of the die being rolled.

Hidden Markov models raise the following questions:

- Given a sequence of emissions, what is the most likely state path?
- Given a sequence of emissions, how can you estimate transition and emission probabilities of the model?
- What is the forward probability that the model generates a given sequence?
- What is the posterior probability that the model is in a particular state at any point in the sequence?


## Analyzing Hidden Markov Models

- "Generating a Test Sequence" on page 14-8
- "Estimating the State Sequence" on page 14-8
- "Estimating Transition and Emission Matrices" on page 14-9
- "Estimating Posterior State Probabilities" on page 14-11
- "Changing the Initial State Distribution" on page 14-12

Statistics Toolbox functions related to hidden Markov models are:

- hmmgenerate - Generates a sequence of states and emissions from a Markov model
- hmmestimate - Calculates maximum likelihood estimates of transition and emission probabilities from a sequence of emissions and a known sequence of states
- hmmtrain - Calculates maximum likelihood estimates of transition and emission probabilities from a sequence of emissions
- hmmviterbi - Calculates the most probable state path for a hidden Markov model
- hmmdecode - Calculates the posterior state probabilities of a sequence of emissions

This section shows how to use these functions to analyze hidden Markov models.

## Generating a Test Sequence

The following commands create the transition and emission matrices for the model described in the "Introduction to Hidden Markov Models (HMM)" on page 14-5:

```
TRANS = [.9 .1; .05 .95;];
EMIS = [1/6, 1/6, 1/6, 1/6, 1/6, 1/6;...
7/12, 1/12, 1/12, 1/12, 1/12, 1/12];
```

To generate a random sequence of states and emissions from the model, use hmmgenerate:

```
[seq,states] = hmmgenerate(1000,TRANS,EMIS);
```

The output seq is the sequence of emissions and the output states is the sequence of states.
hmmgenerate begins in state 1 at step 0 , makes the transition to state $i_{1}$ at step 1 , and returns $i_{1}$ as the first entry in states. To change the initial state, see "Changing the Initial State Distribution" on page 14-12.

## Estimating the State Sequence

Given the transition and emission matrices TRANS and EMIS, the function hmmviterbi uses the Viterbi algorithm to compute the most likely sequence of states the model would go through to generate a given sequence seq of emissions:

```
likelystates = hmmviterbi(seq, TRANS, EMIS);
```

likelystates is a sequence the same length as seq.
To test the accuracy of hmmviterbi, compute the percentage of the actual sequence states that agrees with the sequence likelystates.

```
sum(states==likelystates)/1000
ans =
    0.8200
```

In this case, the most likely sequence of states agrees with the random sequence $82 \%$ of the time.

## Estimating Transition and Emission Matrices

- "Using hmmestimate" on page 14-9
- "Using hmmtrain" on page 14-10

The functions hmmestimate and hmmtrain estimate the transition and emission matrices TRANS and EMIS given a sequence seq of emissions.

Using hmmestimate. The function hmmestimate requires that you know the sequence of states states that the model went through to generate seq.

The following takes the emission and state sequences and returns estimates of the transition and emission matrices:

```
[TRANS_EST, EMIS_EST] = hmmestimate(seq, states)
TRANS_EST =
0.8989 0.1011
0.0585 0.9415
EMIS EST =
\begin{tabular}{llllll}
0.1721 & 0.1721 & 0.1749 & 0.1612 & 0.1803 & 0.1393 \\
0.5836 & 0.0741 & 0.0804 & 0.0789 & 0.0726 & 0.1104
\end{tabular}
```

You can compare the outputs with the original transition and emission matrices, TRANS and EMIS:

```
TRANS
```

TRANS =
$0.9000 \quad 0.1000$
$0.0500 \quad 0.9500$
EMIS

| EMIS $=$ |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 0.1667 | 0.1667 | 0.1667 | 0.1667 | 0.1667 | 0.1667 |
| 0.5833 | 0.0833 | 0.0833 | 0.0833 | 0.0833 | 0.0833 |

Using hmmtrain. If you do not know the sequence of states states, but you have initial guesses for TRANS and EMIS, you can still estimate TRANS and EMIS using hmmtrain.

Suppose you have the following initial guesses for TRANS and EMIS.

```
TRANS_GUESS = [.85 .15; .1 .9];
EMIS_GUESS = [.17 . 16 .17 .16 .17 .17;.6 .08 . 08 . 08 . 08 08];
```

You estimate TRANS and EMIS as follows:

hmmtrain uses an iterative algorithm that alters the matrices TRANS_GUESS and EMIS_GUESS so that at each step the adjusted matrices are more likely to generate the observed sequence, seq. The algorithm halts when the matrices in two successive iterations are within a small tolerance of each other.

If the algorithm fails to reach this tolerance within a maximum number of iterations, whose default value is 100 , the algorithm halts. In this case, hmmtrain returns the last values of TRANS_EST and EMIS_EST and issues a warning that the tolerance was not reached.

If the algorithm fails to reach the desired tolerance, increase the default value of the maximum number of iterations with the command:

```
hmmtrain(seq,TRANS_GUESS,EMIS_GUESS,'maxiterations',maxiter)
```

where maxiter is the maximum number of steps the algorithm executes.

Change the default value of the tolerance with the command:

```
hmmtrain(seq, TRANS_GUESS, EMIS_GUESS, 'tolerance', tol)
```

where tol is the desired value of the tolerance. Increasing the value of tol makes the algorithm halt sooner, but the results are less accurate.

Two factors reduce the reliability of the output matrices of hmmtrain:

- The algorithm converges to a local maximum that does not represent the true transition and emission matrices. If you suspect this, use different initial guesses for the matrices TRANS_EST and EMIS_EST.
- The sequence seq may be too short to properly train the matrices. If you suspect this, use a longer sequence for seq.


## Estimating Posterior State Probabilities

The posterior state probabilities of an emission sequence seq are the conditional probabilities that the model is in a particular state when it generates a symbol in seq, given that seq is emitted. You compute the posterior state probabilities with hmmdecode:

```
PSTATES = hmmdecode(seq,TRANS,EMIS)
```

The output PSTATES is an $M$-by- $L$ matrix, where $M$ is the number of states and $L$ is the length of seq. $\operatorname{PSTATES}(i, j)$ is the conditional probability that the model is in state $i$ when it generates the $j$ th symbol of seq, given that seq is emitted.
hmmdecode begins with the model in state 1 at step 0 , prior to the first emission. $\operatorname{PSTATES}(i, 1)$ is the probability that the model is in state $i$ at the following step 1. To change the initial state, see "Changing the Initial State Distribution" on page 14-12.

To return the logarithm of the probability of the sequence seq, use the second output argument of hmmdecode:

```
[PSTATES,logpseq] = hmmdecode(seq,TRANS,EMIS)
```

The probability of a sequence tends to 0 as the length of the sequence increases, and the probability of a sufficiently long sequence becomes less
than the smallest positive number your computer can represent. hmmdecode returns the logarithm of the probability to avoid this problem.

## Changing the Initial State Distribution

By default, Statistics Toolbox hidden Markov model functions begin in state 1. In other words, the distribution of initial states has all of its probability mass concentrated at state 1 . To assign a different distribution of probabilities, $p=$ [ $p_{1}, p_{2}, \ldots, p_{M}$ ], to the $M$ initial states, do the following:

1 Create an $M+1$-by- $M+1$ augmented transition matrix, $\hat{T}$ of the following form:

$$
\hat{T}=\left[\begin{array}{ll}
0 & p \\
0 & T
\end{array}\right]
$$

where $T$ is the true transition matrix. The first column of $\hat{T}$ contains $M+1$ zeros. $p$ must sum to 1 .

2 Create an $M+1$-by- $N$ augmented emission matrix, $\hat{E}$, that has the following form:

$$
\hat{E}=\left[\begin{array}{l}
0 \\
E
\end{array}\right]
$$

If the transition and emission matrices are TRANS and EMIS, respectively, you create the augmented matrices with the following commands:

```
TRANS_HAT = [0 p; zeros(size(TRANS,1),1) TRANS];
EMIS_HAT = [zeros(1,size(EMIS,2)); EMIS];
```


## Design of Experiments

- "Introduction to Design of Experiments" on page 15-2
- "Full Factorial Designs" on page 15-3
- "Fractional Factorial Designs" on page 15-5
- "Response Surface Designs" on page 15-9
- "D-Optimal Designs" on page 15-15


## Introduction to Design of Experiments

Passive data collection leads to a number of problems in statistical modeling. Observed changes in a response variable may be correlated with, but not caused by, observed changes in individual factors (process variables). Simultaneous changes in multiple factors may produce interactions that are difficult to separate into individual effects. Observations may be dependent, while a model of the data considers them to be independent.

Designed experiments address these problems. In a designed experiment, the data-producing process is actively manipulated to improve the quality of information and to eliminate redundant data. A common goal of all experimental designs is to collect data as parsimoniously as possible while providing sufficient information to accurately estimate model parameters.

For example, a simple model of a response $y$ in an experiment with two controlled factors $x_{1}$ and $x_{2}$ might look like this:

$$
y=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{1} x_{2}+\varepsilon
$$

Here $\varepsilon$ includes both experimental error and the effects of any uncontrolled factors in the experiment. The terms $\beta_{1} x_{1}$ and $\beta_{2} x_{2}$ are main effects and the term $\beta_{3} x_{1} x_{2}$ is a two-way interaction effect. A designed experiment would systematically manipulate $x_{1}$ and $x_{2}$ while measuring $y$, with the objective of accurately estimating $\beta_{0}, \beta_{1}, \beta_{2}$, and $\beta_{3}$.

## Full Factorial Designs

## In this section...

"Multilevel Designs" on page 15-3
"Two-Level Designs" on page 15-4

## Multilevel Designs

To systematically vary experimental factors, assign each factor a discrete set of levels. Full factorial designs measure response variables using every treatment (combination of the factor levels). A full factorial design for $n$ factors with $N_{1}, \ldots, N_{n}$ levels requires $N_{1} \times \ldots \times N_{n}$ experimental runs-one for each treatment. While advantageous for separating individual effects, full factorial designs can make large demands on data collection.

As an example, suppose a machine shop has three machines and four operators. If the same operator always uses the same machine, it is impossible to determine if a machine or an operator is the cause of variation in production. By allowing every operator to use every machine, effects are separated. A full factorial list of treatments is generated by the Statistics Toolbox function fullfact:

```
dFF = fullfact([3,4])
dFF =
    1
    2 1
    3 1
    1 2
    2 2
    3
    3
    2 3
    3
    4
    2 4
    3 4
```

Each of the $3 \times 4=12$ rows of dFF represent one machine/operator combination.

## Two-Level Designs

Many experiments can be conducted with two-level factors, using two-level designs. For example, suppose the machine shop in the previous example always keeps the same operator on the same machine, but wants to measure production effects that depend on the composition of the day and night shifts. The Statistics Toolbox function ff2n generates a full factorial list of treatments:

| dFF2 | $=f f 2 n(4)$ |  |  |
| ---: | :--- | ---: | :--- |
| dFF2 | $=$ |  |  |
| 0 | 0 | 0 | 0 |
| 0 | 0 | 0 | 1 |
| 0 | 0 | 1 | 0 |
| 0 | 0 | 1 | 1 |
| 0 | 1 | 0 | 0 |
| 0 | 1 | 0 | 1 |
| 0 | 1 | 1 | 0 |
| 0 | 1 | 1 | 1 |
| 1 | 0 | 0 | 0 |
| 1 | 0 | 0 | 1 |
| 1 | 0 | 1 | 0 |
| 1 | 0 | 1 | 1 |
| 1 | 1 | 0 | 0 |
| 1 | 1 | 0 | 1 |
| 1 | 1 | 1 | 0 |
| 1 | 1 | 1 | 1 |

Each of the $2^{4}=16$ rows of dFF2 represent one schedule of operators for the day (0) and night (1) shifts.

## Fractional Factorial Designs

In this section...<br>"Introduction to Fractional Factorial Designs" on page 15-5<br>"Plackett-Burman Designs" on page 15-5<br>"General Fractional Designs" on page 15-6

## Introduction to Fractional Factorial Designs

Two-level designs are sufficient for evaluating many production processes. Factor levels of $\pm 1$ can indicate categorical factors, normalized factor extremes, or simply "up" and "down" from current factor settings. Experimenters evaluating process changes are interested primarily in the factor directions that lead to process improvement.

For experiments with many factors, two-level full factorial designs can lead to large amounts of data. For example, a two-level full factorial design with 10 factors requires $2^{10}=1024$ runs. Often, however, individual factors or their interactions have no distinguishable effects on a response. This is especially true of higher order interactions. As a result, a well-designed experiment can use fewer runs for estimating model parameters.

Fractional factorial designs use a fraction of the runs required by full factorial designs. A subset of experimental treatments is selected based on an evaluation (or assumption) of which factors and interactions have the most significant effects. Once this selection is made, the experimental design must separate these effects. In particular, significant effects should not be confounded, that is, the measurement of one should not depend on the measurement of another.

## Plackett-Burman Designs

Plackett-Burman designs are used when only main effects are considered significant. Two-level Plackett-Burman designs require a number of experimental runs that are a multiple of 4 rather than a power of 2 . The MATLAB function hadamard generates these designs:

```
dPB = hadamard(8)
```

$\mathrm{dPB}=$|  |  |  |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 1 | -1 | 1 | -1 | 1 | -1 | 1 | -1 |
| 1 | 1 | -1 | -1 | 1 | 1 | -1 | -1 |
| 1 | -1 | -1 | 1 | 1 | -1 | -1 | 1 |
| 1 | 1 | 1 | 1 | -1 | -1 | -1 | -1 |
| 1 | -1 | 1 | -1 | -1 | 1 | -1 | 1 |
| 1 | 1 | -1 | -1 | -1 | -1 | 1 | 1 |
| 1 | -1 | -1 | 1 | -1 | 1 | 1 | -1 |

Binary factor levels are indicated by $\pm 1$. The design is for eight runs (the rows of dPB) manipulating seven two-level factors (the last seven columns of dPB). The number of runs is a fraction $8 / 2^{7}=0.0625$ of the runs required by a full factorial design. Economy is achieved at the expense of confounding main effects with any two-way interactions.

## General Fractional Designs

At the cost of a larger fractional design, you can specify which interactions you wish to consider significant. A design of resolution $R$ is one in which no $n$-factor interaction is confounded with any other effect containing less than $R-n$ factors. Thus, a resolution III design does not confound main effects with one another but may confound them with two-way interactions (as in "Plackett-Burman Designs" on page 15-5), while a resolution IV design does not confound either main effects or two-way interactions but may confound two-way interactions with each other.

Specify general fractional factorial designs using a full factorial design for a selected subset of basic factors and generators for the remaining factors. Generators are products of the basic factors, giving the levels for the remaining factors. Use the Statistics Toolbox function fracfact to generate these designs:

| -1 | -1 | -1 | -1 | -1 | -1 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| -1 | -1 | -1 | 1 | 1 | 1 |
| -1 | -1 | 1 | -1 | 1 | 1 |
| -1 | -1 | 1 | 1 | -1 | -1 |
| -1 | 1 | -1 | -1 | 1 | -1 |


| -1 | 1 | -1 | 1 | -1 | 1 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| -1 | 1 | 1 | -1 | -1 | 1 |
| -1 | 1 | 1 | 1 | 1 | -1 |
| 1 | -1 | -1 | -1 | -1 | 1 |
| 1 | -1 | -1 | 1 | 1 | -1 |
| 1 | -1 | 1 | -1 | 1 | -1 |
| 1 | -1 | 1 | 1 | -1 | 1 |
| 1 | 1 | -1 | -1 | 1 | 1 |
| 1 | 1 | -1 | 1 | -1 | -1 |
| 1 | 1 | 1 | -1 | -1 | -1 |
| 1 | 1 | 1 | 1 | 1 | 1 |

This is a six-factor design in which four two-level basic factors (a, b, c, and $d$ in the first four columns of dfF) are measured in every combination of levels, while the two remaining factors (in the last three columns of dfF) are measured only at levels defined by the generators bcd and acd, respectively. Levels in the generated columns are products of corresponding levels in the columns that make up the generator.

The challenge of creating a fractional factorial design is to choose basic factors and generators so that the design achieves a specified resolution in a specified number of runs. Use the Statistics Toolbox function fracfactgen to find appropriate generators:

```
generators = fracfactgen('a b c d e f',4,4)
generators =
    'a'
    'b'
'c'
'd'
'bcd'
'acd'
```

These are generators for a six-factor design with factors a through f, using $2^{4}$ $=16$ runs to achieve resolution IV. The fracfactgen function uses an efficient search algorithm to find generators that meet the requirements.

An optional output from fracfact displays the confounding pattern of the design:

```
[dfF,confounding] = fracfact(generators);
```

```
confounding
confounding =
    'Term' 'Generator' 'Confounding
    'X1' 'a' 'X1'
    'X2' 'b' 'X2'
    'X3' 'c' 'X3'
    'X4' 'd' 'X4'
    'X5' 'bcd' 'X5'
    'X6' 'acd' 'X6'
    'X1*X2' 'ab' 'X1*X2 + X5*X6'
    'X1*X3' 'ac' 'X1*X3 + X4*X6'
    'X1*X4' 'ad' 'X1*X4 + X3*X6'
    'X1*X5' 'abcd' 'X1*X5 + X2*X6'
    'X1*X6' 'cd' 'X1*X6 + X2*X5 + X3*X4'
    'X2*X3' 'bc' 'X2*X3 + X4*X5'
    'X2*X4' 'bd' 'X2*X4 + X3*X5'
    'X2*X5' 'cd' 'X1*X6 + X2*X5 + X3*X4'
    'X2*X6' 'abcd' 'X1*X5 + X2*X6'
    'X3*X4' 'cd' 'X1*X6 + X2*X5 + X3*X4'
    'X3*X5' 'bd' 'X2*X4 + X3*X5'
    'X3*X6' 'ad' 'X1*X4 + X3*X6'
    'X4*X5' 'bc' 'X2*X3 + X4*X5'
    'X4*X6' 'ac' 'X1*X3 + X4*X6'
    'X5*X6' 'ab' 'X1*X2 + X5*X6'
```

The confounding pattern shows that main effects are effectively separated by the design, but two-way interactions are confounded with various other two-way interactions.

## Response Surface Designs

In this section...<br>"Introduction to Response Surface Designs" on page 15-9<br>"Central Composite Designs" on page 15-9<br>"Box-Behnken Designs" on page 15-13

## Introduction to Response Surface Designs

As discussed in "Response Surface Models" on page 9-59, quadratic response surfaces are simple models that provide a maximum or minimum without making additional assumptions about the form of the response. Quadratic models can be calibrated using full factorial designs with three or more levels for each factor, but these designs generally require more runs than necessary to accurately estimate model parameters. This section discusses designs for calibrating quadratic models that are much more efficient, using three or five levels for each factor, but not using all combinations of levels.

## Central Composite Designs

Central composite designs (CCDs), also known as Box-Wilson designs, are appropriate for calibrating the full quadratic models described in "Response Surface Models" on page 9-59. There are three types of CCDs-circumscribed, inscribed, and faced-pictured below:



Each design consists of a factorial design (the corners of a cube) together with center and star points that allow for estimation of second-order effects. For a full quadratic model with $n$ factors, CCDs have enough design points to estimate the $(n+2)(n+1) / 2$ coefficients in a full quadratic model with $n$ factors.

The type of CCD used (the position of the factorial and star points) is determined by the number of factors and by the desired properties of the design. The following table summarizes some important properties. A design is rotatable if the prediction variance depends only on the distance of the design point from the center of the design.

| Design | Rotatable | Factor <br> Levels | Uses Points <br> Outside $\mathbf{1}$ | Accuracy of <br> Estimates |
| :--- | :--- | :--- | :--- | :--- |
| Circumscribed <br> (CCC) | Yes | 5 | Yes | Good over entire <br> design space |
| Inscribed <br> (CCI) | Yes | 5 | No | Good over central <br> subset of design space |
| Faced (CCF) | No | 3 | No | Fair over entire <br> design space; poor <br> for pure quadratic <br> coefficients |

Generate CCDs with the Statistics Toolbox function ccdesign:

```
dCC = ccdesign(3,'type','circumscribed')
dCC =
\begin{tabular}{rrr}
-1.0000 & -1.0000 & -1.0000 \\
-1.0000 & -1.0000 & 1.0000 \\
-1.0000 & 1.0000 & -1.0000 \\
-1.0000 & 1.0000 & 1.0000 \\
1.0000 & -1.0000 & -1.0000 \\
1.0000 & -1.0000 & 1.0000 \\
1.0000 & 1.0000 & -1.0000 \\
1.0000 & 1.0000 & 1.0000 \\
-1.6818 & 0 & 0 \\
1.6818 & 0 & 0
\end{tabular}
            0 -1.6818 0
            0 1.6818 0
            0 0 -1.6818
            0 0 1.6818
            0 0 0
            0
            0
            0
            0
            0 0
```

The repeated center point runs allow for a more uniform estimate of the prediction variance over the entire design space.

## Box-Behnken Designs

Like the designs described in "Central Composite Designs" on page 15-9, Box-Behnken designs are used to calibrate full quadratic models. Box-Behnken designs are rotatable and, for a small number of factors (four or less), require fewer runs than CCDs. By avoiding the corners of the design space, they allow experimenters to work around extreme factor combinations. Like an inscribed CCD, however, extremes are then poorly estimated.

The geometry of a Box-Behnken design is pictured in the following figure.


Design points are at the midpoints of edges of the design space and at the center, and do not contain an embedded factorial design.

Generate Box-Behnken designs with the Statistics Toolbox function bbdesign:

| dBB $=$ bbdesign (3) |  |  |
| ---: | :--- | ---: |
| dBB $=$ |  |  |
| -1 | -1 | 0 |
| -1 | 1 | 0 |
| 1 | -1 | 0 |
| 1 | 1 | 0 |
| -1 | 0 | -1 |
| -1 | 0 | 1 |
| 1 | 0 | -1 |
| 1 | 0 | 1 |
| 0 | -1 | -1 |
| 0 | -1 | 1 |
| 0 | 1 | -1 |
| 0 | 1 | 1 |
| 0 | 0 | 0 |
| 0 | 0 | 0 |
| 0 | 0 | 0 |

Again, the repeated center point runs allow for a more uniform estimate of the prediction variance over the entire design space.

## D-Optimal Designs

In this section...<br>"Introduction to D-Optimal Designs" on page 15-15<br>"Generating D-Optimal Designs" on page 15-16<br>"Augmenting D-Optimal Designs" on page 15-19<br>"Specifying Fixed Covariate Factors" on page 15-20<br>"Specifying Categorical Factors" on page 15-21<br>"Specifying Candidate Sets" on page 15-21

## Introduction to D-Optimal Designs

Traditional experimental designs ("Full Factorial Designs" on page 15-3, "Fractional Factorial Designs" on page 15-5, and "Response Surface Designs" on page $15-9$ ) are appropriate for calibrating linear models in experimental settings where factors are relatively unconstrained in the region of interest. In some cases, however, models are necessarily nonlinear. In other cases, certain treatments (combinations of factor levels) may be expensive or infeasible to measure. D-optimal designs are model-specific designs that address these limitations of traditional designs.

A $D$-optimal design is generated by an iterative search algorithm and seeks to minimize the covariance of the parameter estimates for a specified model. This is equivalent to maximizing the determinant $D=\left|X^{T} X\right|$, where $X$ is the design matrix of model terms (the columns) evaluated at specific treatments in the design space (the rows). Unlike traditional designs, $D$-optimal designs do not require orthogonal design matrices, and as a result, parameter estimates may be correlated. Parameter estimates may also be locally, but not globally, $D$-optimal.

There are several Statistics Toolbox functions for generating $D$-optimal designs:

| Function | Description |
| :--- | :--- |
| candexch | Uses a row-exchange algorithm to generate a $D$-optimal design <br> with a specified number of runs for a specified model and a <br> specified candidate set. This is the second component of the <br> algorithm used by rowexch. |
| candgen | Generates a candidate set for a specified model. This is the <br> first component of the algorithm used by rowexch. |
| cordexch | Uses a coordinate-exchange algorithm to generate a $D$-optimal <br> design with a specified number of runs for a specified model. |
| daugment | Uses a coordinate-exchange algorithm to augment an existing <br> $D$-optimal design with additional runs to estimate additional <br> model terms. |
| dcovary | Uses a coordinate-exchange algorithm to generate a $D$-optimal <br> design with fixed covariate factors. |
| rowexch | Uses a row-exchange algorithm to generate a $D$-optimal design <br> with a specified number of runs for a specified model. The <br> algorithm calls candgen and then candexch. (Call candexch <br> separately to specify a candidate set.) |

The following sections explain how to use these functions to generate $D$-optimal designs.

Note The Statistics Toolbox function rsmdemo generates simulated data for experimental settings specified by either the user or by a $D$-optimal design generated by cordexch. It uses the rstool interface to visualize response surface models fit to the data, and it uses the nlintool interface to visualize a nonlinear model fit to the data.

## Generating D-Optimal Designs

Two Statistics Toolbox algorithms generate $D$-optimal designs:

- The cordexch function uses a coordinate-exchange algorithm
- The rowexch function uses a row-exchange algorithm

Both cordexch and rowexch use iterative search algorithms. They operate by incrementally changing an initial design matrix $X$ to increase $D=\left|X^{T} X\right|$ at each step. In both algorithms, there is randomness built into the selection of the initial design and into the choice of the incremental changes. As a result, both algorithms may return locally, but not globally, $D$-optimal designs. Run each algorithm multiple times and select the best result for your final design. Both functions have a 'tries' parameter that automates this repetition and comparison.

At each step, the row-exchange algorithm exchanges an entire row of $X$ with a row from a design matrix $C$ evaluated at a candidate set of feasible treatments. The rowexch function automatically generates a $C$ appropriate for a specified model, operating in two steps by calling the candgen and candexch functions in sequence. Provide your own $C$ by calling candexch directly. In either case, if $C$ is large, its static presence in memory can affect computation.

The coordinate-exchange algorithm, by contrast, does not use a candidate set. (Or rather, the candidate set is the entire design space.) At each step, the coordinate-exchange algorithm exchanges a single element of $X$ with a new element evaluated at a neighboring point in design space. The absence of a candidate set reduces demands on memory, but the smaller scale of the search means that the coordinate-exchange algorithm is more likely to become trapped in a local minimum than the row-exchange algorithm.

For example, suppose you want a design to estimate the parameters in the following three-factor, seven-term interaction model:

$$
y=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+\beta_{12} x_{1} x_{2}+\beta_{13} x_{1} x_{3}+\beta_{23} x_{2} x_{3}+\varepsilon
$$

Use cordexch to generate a $D$-optimal design with seven runs:

```
nfactors = 3;
nruns = 7;
[dCE,X] = cordexch(nfactors,nruns,'interaction','tries',10)
dCE =
\begin{tabular}{rrr}
-1 & 1 & 1 \\
-1 & -1 & -1 \\
1 & 1 & 1 \\
-1 & 1 & -1 \\
1 & -1 & 1
\end{tabular}
```

$\mathrm{X}=$| 1 | -1 | -1 |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| -1 | -1 | 1 |  |  |  |  |
|  |  |  |  | 1 | -1 | -1 |
| 1 | -1 | 1 | 1 |  |  |  |
| 1 | -1 | -1 | -1 | 1 | 1 | 1 |
| 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 1 | -1 | 1 | -1 | -1 | 1 | -1 |
| 1 | 1 | -1 | 1 | -1 | 1 | -1 |
| 1 | 1 | -1 | -1 | -1 | -1 | 1 |
| 1 | -1 | -1 | 1 | 1 | -1 | -1 |

Columns of the design matrix $X$ are the model terms evaluated at each row of the design dCE. The terms appear in order from left to right:

1 Constant term
2 Linear terms (1, 2, 3)
3 Interaction terms (12, 13, 23)
Use $X$ to fit the model, as described in "Linear Regression" on page 9-3, to response data measured at the design points in dCE.

Use rowexch in a similar fashion to generate an equivalent design:

```
[dRE,X] = rowexch(nfactors, nruns,'interaction','tries', 10)
dRE =
\begin{tabular}{rrr}
-1 & -1 & 1 \\
1 & -1 & 1
\end{tabular}
\(1 \quad-1 \quad-1\)
            \(1 \quad 1 \quad 1\)
            \(\begin{array}{lll}-1 & -1 & -1\end{array}\)
            \(\begin{array}{lll}-1 & 1 & -1\end{array}\)
X =
\begin{tabular}{rrrrrrr}
1 & -1 & -1 & 1 & 1 & -1 & -1 \\
1 & 1 & -1 & 1 & -1 & 1 & -1 \\
1 & 1 & -1 & -1 & -1 & -1 & 1 \\
1 & 1 & 1 & 1 & 1 & 1 & 1 \\
1 & -1 & -1 & -1 & 1 & 1 & 1 \\
1 & -1 & 1 & -1 & -1 & 1 & -1
\end{tabular}
```


## Augmenting D-Optimal Designs

In practice, you may want to add runs to a completed experiment to learn more about a process and estimate additional model coefficients. The daugment function uses a coordinate-exchange algorithm to augment an existing $D$-optimal design.

For example, the following eight-run design is adequate for estimating main effects in a four-factor model:


To estimate the six interaction terms in the model, augment the design with eight additional runs:

```
dCEinteraction = daugment(dCEmain,8,'interaction')
dCEinteraction =
\begin{tabular}{rrrr}
1 & -1 & -1 & 1 \\
-1 & -1 & 1 & 1 \\
-1 & 1 & -1 & 1 \\
1 & 1 & 1 & -1 \\
1 & 1 & 1 & 1 \\
-1 & 1 & -1 & -1 \\
1 & -1 & -1 & -1 \\
-1 & -1 & 1 & -1 \\
-1 & 1 & 1 & 1 \\
-1 & -1 & -1 & -1 \\
1 & -1 & 1 & -1 \\
1 & 1 & -1 & 1 \\
-1 & 1 & 1 & -1
\end{tabular}
```

| 1 | 1 | -1 | -1 |
| ---: | ---: | ---: | ---: |
| 1 | -1 | 1 | 1 |
| 1 | 1 | 1 | -1 |

The augmented design is full factorial, with the original eight runs in the first eight rows.

The 'start' parameter of the candexch function provides the same functionality as daugment, but uses a row exchange algorithm rather than a coordinate-exchange algorithm.

## Specifying Fixed Covariate Factors

In many experimental settings, certain factors and their covariates are constrained to a fixed set of levels or combinations of levels. These cannot be varied when searching for an optimal design. The dcovary function allows you to specify fixed covariate factors in the coordinate exchange algorithm.

For example, suppose you want a design to estimate the parameters in a three-factor linear additive model, with eight runs that necessarily occur at different times. If the process experiences temporal linear drift, you may want to include the run time as a variable in the model. Produce the design as follows:

```
time = linspace(-1,1,8)';
[dCV,X] = dcovary(3,time,'linear')
dCV =
\begin{tabular}{rrrrr}
-1.0000 & 1.0000 & 1.0000 & -1.0000 & \\
1.0000 & -1.0000 & -1.0000 & -0.7143 & \\
-1.0000 & -1.0000 & -1.0000 & -0.4286 & \\
1.0000 & -1.0000 & 1.0000 & -0.1429 & \\
1.0000 & 1.0000 & -1.0000 & 0.1429 & \\
-1.0000 & 1.0000 & -1.0000 & 0.4286 & \\
1.0000 & 1.0000 & 1.0000 & 0.7143 & \\
-1.0000 & -1.0000 & 1.0000 & 1.0000 & \\
& & & & \\
1.0000 & -1.0000 & 1.0000 & 1.0000 & -1.0000 \\
1.0000 & 1.0000 & -1.0000 & -1.0000 & -0.7143 \\
1.0000 & -1.0000 & -1.0000 & -1.0000 & -0.4286 \\
1.0000 & 1.0000 & -1.0000 & 1.0000 & -0.1429
\end{tabular}
```

| 1.0000 | 1.0000 | 1.0000 | -1.0000 | 0.1429 |
| ---: | ---: | ---: | ---: | ---: |
| 1.0000 | -1.0000 | 1.0000 | -1.0000 | 0.4286 |
| 1.0000 | 1.0000 | 1.0000 | 1.0000 | 0.7143 |
| 1.0000 | -1.0000 | -1.0000 | 1.0000 | 1.0000 |

The column vector time is a fixed factor, normalized to values between $\pm 1$. The number of rows in the fixed factor specifies the number of runs in the design. The resulting design dCV gives factor settings for the three controlled model factors at each time.

## Specifying Categorical Factors

Categorical factors take values in a discrete set of levels. Both cordexch and rowexch have a 'categorical' parameter that allows you to specify the indices of categorical factors and a 'levels' parameter that allows you to specify a number of levels for each factor.

For example, the following eight-run design is for a linear additive model with five factors in which the final factor is categorical with three levels:
dCEcat $=$ cordexch(5,8,'linear', 'categorical', 5, 'levels' , 3)
dCEcat $=$

-1 |  |  |  |  |  |
| ---: | ---: | ---: | ---: | :--- |
| -1 | -1 | 1 | 1 | 2 |
| 1 | 1 | 1 | -1 | 3 |
| 1 | 1 | -1 | -1 | 3 |
| 1 | -1 | -1 | 1 | 3 |
| -1 | 1 | -1 | 1 | 1 |
| -1 | 1 | 1 | -1 | 3 |
| 1 | -1 | 1 | -1 | 1 |

## Specifying Candidate Sets

The row-exchange algorithm exchanges rows of an initial design matrix $X$ with rows from a design matrix $C$ evaluated at a candidate set of feasible treatments. The rowexch function automatically generates a $C$ appropriate for a specified model, operating in two steps by calling the candgen and candexch functions in sequence. Provide your own $C$ by calling candexch directly.

For example, the following uses rowexch to generate a five-run design for a two-factor pure quadratic model using a candidate set that is produced internally:

```
dRE1 = rowexch(2,5,'purequadratic','tries',10)
dRE1 =
    -1 1
    0
    1 -1
    0
    1 1
```

The same thing can be done using candgen and candexch in sequence:

```
[dC,C] = candgen(2,'purequadratic') % Candidate set, C
dC =
    -1 -1
        0 -1
        1
        0
        0
        -1 1
        0
        1
C =
    1
    1
    1
    1 0
    1
    1
    1
    1
treatments = candexch(C,5,'tries',10) % D-opt subset
treatments =
    2
    1
    7
    3
```

```
    4
dRE2 = dC(treatments,:) % Display design
dRE2 =
    0 -1
    -1 -1
    -1 1
    1 -1
    -1 0
```

You can replace $C$ in this example with a design matrix evaluated at your own candidate set. For example, suppose your experiment is constrained so that the two factors cannot have extreme settings simultaneously. The following produces a restricted candidate set:

```
constraint = sum(abs(dC),2) < 2; % Feasible treatments
my_dC = dC(constraint,:)
my_dC =
\begin{tabular}{rr}
0 & -1 \\
-1 & 0 \\
0 & 0 \\
1 & 0 \\
0 & 1
\end{tabular}
```

Use the x2fx function to convert the candidate set to a design matrix:

```
my_C = x2fx(my_dC,'purequadratic')
my_C =
\begin{tabular}{rrrrr}
1 & 0 & -1 & 0 & 1 \\
1 & -1 & 0 & 1 & 0 \\
1 & 0 & 0 & 0 & 0 \\
1 & 1 & 0 & 1 & 0 \\
1 & 0 & 1 & 0 & 1
\end{tabular}
```

Find the required design in the same manner:

```
my_treatments = candexch(my_C,5,'tries',10) % D-opt subset
my_treatments =
    2
    4
    5
    1
```

```
    3
my_dRE = my_dC(my_treatments,:) % Display design
my_dRE =
    -1 0
    1 0
    0 1
0 -1
0
```


## Statistical Process Control

- "Introduction to Statistical Process Control" on page 16-2
- "Control Charts" on page 16-3
- "Capability Studies" on page 16-6


## Introduction to Statistical Process Control

Statistical process control (SPC) refers to a number of different methods for monitoring and assessing the quality of manufactured goods. Combined with methods from the Chapter 15, "Design of Experiments", SPC is used in programs that define, measure, analyze, improve, and control development and production processes. These programs are often implemented using "Design for Six Sigma" methodologies.

## Control Charts

A control chart displays measurements of process samples over time. The measurements are plotted together with user-defined specification limits and process-defined control limits. The process can then be compared with its specifications-to see if it is in control or out of control.

The chart is just a monitoring tool. Control activity might occur if the chart indicates an undesirable, systematic change in the process. The control chart is used to discover the variation, so that the process can be adjusted to reduce it.

Control charts are created with the controlchart function. Any of the following chart types may be specified:

- Xbar or mean
- Standard deviation
- Range
- Exponentially weighted moving average
- Individual observation
- Moving range of individual observations
- Moving average of individual observations
- Proportion defective
- Number of defectives
- Defects per unit
- Count of defects

Control rules are specified with the controlrules function.
For example, the following commands create an xbar chart, using the "Western Electric 2" rule (2 of 3 points at least 2 standard errors above the center line) to mark out of control measurements:

```
load parts;
st = controlchart(runout,'rules','we2');
```

```
    x = st.mean;
    cl = st.mu;
    se = st.sigma./sqrt(st.n);
    hold on
    plot(cl+2*se,'m')
```

XBAR control chart


Measurements that violate the control rule can then be identified:

```
R = controlrules('we2',x,cl,se);
I = find(R)
```

I =
21
23

## Capability Studies

Before going into production, many manufacturers run a capability study to determine if their process will run within specifications enough of the time. Capability indices produced by such a study are used to estimate expected percentages of defective parts.

Capability studies are conducted with the capability function. The following capability indices are produced:

- mu - Sample mean
- sigma - Sample standard deviation
- P - Estimated probability of being within the lower (L) and upper (U) specification limits
- Pl - Estimated probability of being below L
- Pu - Estimated probability of being above $U$
- Cp - (U-L) /(6*sigma)
- Cpl-(mu-L)./(3.*sigma)
- Cpu - (U-mu). /(3.*sigma)
- Cpk - min(Cpl,Cpu)

As an example, simulate a sample from a process with a mean of 3 and a standard deviation of 0.005 :

```
data = normrnd(3,0.005,100,1);
```

Compute capability indices if the process has an upper specification limit of 3.01 and a lower specification limit of 2.99:

```
S = capability(data,[2.99 3.01])
S =
    mu: 3.0006
    sigma: 0.0047
            P: 0.9669
            Pl: 0.0116
            Pu: 0.0215
```

```
    Cp: 0.7156
Cpl: 0.7567
Cpu: 0.6744
Cpk: 0.6744
```

Visualize the specification and process widths:

```
capaplot(data,[2.99 3.01]);
grid on
```

Probability Between Limits $=0.96688$


## Parallel Statistics

- "Quick Start Parallel Computing for Statistics Toolbox" on page 17-2
- "Concepts of Parallel Computing in Statistics Toolbox" on page 17-7
- "When to Run Statistical Functions in Parallel" on page 17-8
- "Working with parfor" on page 17-10
- "Reproducibility in Parallel Statistical Computations" on page 17-13
- "Examples of Parallel Statistical Functions" on page 17-19


## Quick Start Parallel Computing for Statistics Toolbox

Note To use parallel computing as described in this chapter, you must have a Parallel Computing Toolbox license.

```
In this section...
"What Is Parallel Statistics Functionality?" on page 17-2
"How To Compute in Parallel" on page 17-3
"Example: Parallel Treebagger" on page 17-5
```


## What Is Parallel Statistics Functionality?

You can use any of the Statistics Toolbox functions with Parallel Computing Toolbox constructs such as parfor and spmd. However, some functions, such as those with interactive displays, can lose functionality in parallel. In particular, displays and interactive usage are not effective on workers (see "Vocabulary for Parallel Computation" on page 17-7).

Additionally, the following functions are enhanced to use parallel computing internally. These functions use parfor internally to parallelize calculations.

- bootci
- bootstrp
- candexch
- cordexch
- crossval
- daugment
- dcovary
- jackknife
- nnmf
- plsregress
- rowexch
- sequentialfs
- TreeBagger
- TreeBagger.growTrees

This chapter gives the simplest way to use these enhanced functions in parallel. For more advanced topics, including the issues of reproducibility and nested parfor loops, see the other sections in this chapter.

For information on parallel statistical computing at the command line, enter

```
help parallelstats
```


## How To Compute in Parallel

To have a function compute in parallel:
1 "Open matlabpool" on page 17-3
2 "Set the UseParallel Option to 'always" on page 17-5
3 "Call the Function Using the Options Structure" on page 17-5

## Open matlabpool

To run a statistical computation in parallel, first set up a parallel environment.

Note Setting up a parallel environment can take several seconds.

Multicore. For a multicore machine, enter the following at the MATLAB command line:
matlabpool open $n$
$n$ is the number of workers you want to use.

Network. If you have multiple processors on a network, use Parallel Computing Toolbox functions and MATLAB ${ }^{\circledR}$ Distributed Computing Server ${ }^{\text {TM }}$ software to establish parallel computation. Make sure that your system is configured properly for parallel computing. Check with your system administrator, or refer to the Parallel Computing Toolbox documentation, or the Administrator Guide documentation for MATLAB Distributed Computing Server.

Many parallel statistical functions call a function that can be one you define in a file. For example, jackknife calls a function (jackfun) that can be a built-in MATLAB function such as corr, but can also be a function you define. Built-in functions are available to all workers. However, you must take extra steps to enable workers to access a function file that you define.

To place a function file on the path of all workers, and check that it is accessible:

1 At the command line, enter

```
matlabpool open conf
matlabpool open conf \(n\)
```

or
where conf is your configuration, and $n$ is the number of processors you want to use.

2 If network_file_path is the network path to your function file, enter pctRunOnAll('addpath network_file_path')
so the worker processors can access your function file.
3 Check whether the file is on the path of every worker by entering:

```
pctRunOnAll('which filename')
```

If any worker does not have a path to the file, it reports:

```
filename not found.
```


## Set the UseParallel Option to 'always'

Create an options structure with the statset function. To run in parallel, set the UseParallel option to 'always':

```
paroptions = statset('UseParallel','always');
```


## Call the Function Using the Options Structure

Call your function with syntax that uses the options structure. For example:

```
% Run crossval in parallel
cvMse = crossval('mse',x,y,'predfun',regf,'Options',paroptions);
% Run bootstrp in parallel
sts = bootstrp(100,@(x)[mean(x) std(x)],y,'Options',paroptions);
% Run TreeBagger in parallel
b = TreeBagger(50,meas,spec,'O0BPred','on','Options',paroptions);
```

For more complete examples of parallel statistical functions, see "Example: Parallel Treebagger" on page 17-5 and "Examples of Parallel Statistical Functions" on page 17-19.

After you have finished computing in parallel, close the parallel environment:

```
matlabpool close
```

Tip To save time, keep the pool open if you expect to compute in parallel again soon.

## Example: Parallel Treebagger

To run the example "Workflow Example: Regression of Insurance Risk Rating for Car Imports with TreeBagger" on page 13-97 in parallel:

1 Set up the parallel environment to use two cores:

```
matlabpool open 2
```

Starting matlabpool using the 'local' configuration ...

```
connected to 2 labs.
```

2 Set the options to use parallel processing:

```
paroptions = statset('UseParallel','always');
```

3 Load the problem data and separate it into input and response:

```
load imports-85;
Y = X(:,1);
X = X(:,2:end);
```

4 Estimate feature importance using leaf size 1 and 1000 trees in parallel. Time the function for comparison purposes:

```
tic
b = TreeBagger(1000,X,Y,'Method','r','OOBVarImp','on',...
    'cat',16:25,'MinLeaf',1,'Options',paroptions);
toc
Elapsed time is 37.357930 seconds.
```

5 Perform the same computation in serial for timing comparison:

```
tic
b = TreeBagger(1000,X,Y,'Method','r','OOBVarImp','on',...
    'cat',16:25,'MinLeaf',1); % No options gives serial
toc
Elapsed time is 63.921864 seconds.
```

Computing in parallel took less than $60 \%$ of the time of computing serially.

# Concepts of Parallel Computing in Statistics Toolbox 

In this section...<br>"Subtleties in Parallel Computing" on page 17-7<br>"Vocabulary for Parallel Computation" on page 17-7

## Subtleties in Parallel Computing

There are two main subtleties in parallel computations:

- Nested parallel evaluations (see "No Nested parfor Loops" on page 17-11). Only the outermost parfor loop runs in parallel, the others run serially.
- Reproducible results when using random numbers (see "Reproducibility in Parallel Statistical Computations" on page 17-13). How can you get exactly the same results when repeatedly running a parallel computation that uses random numbers?


## Vocabulary for Parallel Computation

- worker - An independent MATLAB session that runs code distributed by the client.
- client - The MATLAB session with which you interact, and that distributes jobs to workers.
- parfor - A Parallel Computing Toolbox function that distributes independent code segments to workers (see "Working with parfor" on page 17-10).
- random stream - A pseudorandom number generator, and the sequence of values it generates. MATLAB implements random streams with the RandStream class.
- reproducible computation - A computation that can be exactly replicated, even in the presence of random numbers (see "Reproducibility in Parallel Statistical Computations" on page 17-13).


# When to Run Statistical Functions in Parallel 

In this section...<br>"Why Run in Parallel?" on page 17-8<br>"Factors Affecting Speed" on page 17-8<br>"Factors Affecting Results" on page 17-9

## Why Run in Parallel?

The main reason to run statistical computations in parallel is to gain speed, meaning to reduce the execution time of your program or functions. "Factors Affecting Speed" on page 17-8 discusses the main items affecting the speed of programs or functions. "Factors Affecting Results" on page 17-9 discusses details that can cause a parallel run to give different results than a serial run.

## Factors Affecting Speed

Some factors that can affect the speed of execution of parallel processing are:

- Parallel environment setup. It takes time to run matlabpool to begin computing in parallel. If your computation is fast, the setup time can exceed any time saved by computing in parallel.
- Parallel overhead. There is overhead in communication and coordination when running in parallel. If function evaluations are fast, this overhead could be an appreciable part of the total computation time. Thus, solving a problem in parallel can be slower than solving the problem serially. For an example, see Improving Optimization Performance with Parallel Computing in MATLAB Digest, March 2009.
- No nested parfor loops. This is described in "Working with parfor" on page 17-10. parfor does not work in parallel when called from within another parfor loop. If you have programmed your custom functions to take advantage of parallel processing, the limitation of no nested parfor loops can cause a parallel function to run slower than expected.
- When executing serially, parfor loops run slightly slower than for loops.
- Passing parameters. Parameters are automatically passed to worker sessions during the execution of parallel computations. If there are many
parameters, or they take a large amount of memory, passing parameters can slow the execution of your computation.
- Contention for resources: network and computing. If the pool of workers has low bandwidth or high latency, parallel computation can be slow.


## Factors Affecting Results

Some factors can affect results when using parallel processing. There are several caveats related to parfor listed in "Limitations" in the Parallel Computing Toolbox documentation. Some important factors are:

- Persistent or global variables. If any functions use persistent or global variables, these variables can take different values on different worker processors. Furthermore, they might not be cleared properly on the worker processors.
- Accessing external files. External files can be accessed unpredictably during a parallel computation. The order of computations is not guaranteed during parallel processing, so external files can be accessed in unpredictable order, leading to unpredictable results. Furthermore, if multiple processors try to read an external file simultaneously, the file can become locked, leading to a read error, and halting function execution.
- Noncomputational functions, such as input, plot, and keyboard, can behave badly when used in your custom functions. When called in a parfor loop, these functions are executed on worker machines. This can cause a worker to become nonresponsive, since it is waiting for input.
- parfor does not allow break or return statements.
- The random numbers you use can affect the results of your computations. See "Reproducibility in Parallel Statistical Computations" on page 17-13.


## Working with parfor

In this section...<br>"How Statistical Functions Use parfor" on page 17-10<br>"Characteristics of parfor" on page 17-11

## How Statistical Functions Use parfor

parfor is a Parallel Computing Toolbox function similar to a for loop. Parallel statistical functions call parfor internally. parfor distributes computations to worker processors.


## Characteristics of parfor

More caveats related to parfor appear in "Limitations" in the Parallel Computing Toolbox documentation.

## No Nested parfor Loops

parfor does not work in parallel when called from within another parfor loop, or from an spmd block. Parallelization occurs only at the outermost level.

Suppose, for example, you want to apply jackknife to your function userfen, which calls parfor, and you want to call jackknife in a loop. The following figure shows three cases:

1 The outermost loop is parfor. Only that loop runs in parallel.
2 The outermost parfor loop is in jackknife. Only jackknife runs in parallel.

3 The outermost parfor loop is in userfon. userfon uses parfor in parallel.

Bold indicates the function that runs in parallel


## When parfor Runs in Parallel

## Reproducibility in Parallel Statistical Computations

In this section...<br>"Issues and Considerations in Reproducing Parallel Computations" on page 17-13<br>"Running Reproducible Parallel Computations" on page 17-14<br>"Subtleties in Parallel Statistical Computation Using Random Numbers" on page 17-15

## Issues and Considerations in Reproducing Parallel Computations

A reproducible computation is one that gives the same results every time it runs. Reproducibility is important for:

- Debugging - To correct an anomalous result, you need to reproduce the result.
- Confidence - When you can reproduce results, you can investigate and understand them.
- Modifying existing code - When you change existing code, you want to ensure that you do not break anything.

Generally, you do not need to ensure reproducibility for your computation. Often, when you want reproducibility, the simplest technique is to run in serial instead of in parallel. In serial computation you can simply call the rng function as follows:

```
s = rng % Obtain the current state of the random stream
% run the statistical function
rng(s) % Reset the stream to the previous state
% run the statistical function again, obtain identical results
```

This section addresses the case when your function uses random numbers, and you want reproducible results in parallel. This section also addresses the case when you want the same results in parallel as in serial.

## Running Reproducible Parallel Computations

To run a Statistics Toolbox function reproducibly:
1 Set the UseSubstreams option to 'always'.
2 Set the Streams option to a type that supports substreams: 'mlfg6331_64' or 'mrg32k3a'. For information on these streams, see "Choosing a Random Number Generator" in the MATLAB Mathematics documentation.

3 To compute in parallel, set the UseParallel option to 'always'.
4 Call the function with the options structure.
5 To reproduce the computation, reset the stream, then call the function again.

To understand why this technique gives reproducibility, see "How Substreams Enable Reproducible Parallel Computations" on page 17-15.

For example, to use the 'mlfg6331_64' stream for reproducible computation:
1 Create an appropriate options structure:

```
s = RandStream('mlfg6331_64');
options = statset('UseParallel','always', ...
    'Streams',s,'UseSubstreams', 'always');
```

2 Run your parallel computation. For instructions, see "Quick Start Parallel Computing for Statistics Toolbox" on page 17-2.

3 Reset the random stream:

```
reset(s);
```

4 Rerun your parallel computation. You obtain identical results.
For an example of a parallel computation run this reproducible way, see "Reproducible Parallel Bootstrap" on page 17-23.

## Subtleties in Parallel Statistical Computation Using Random Numbers

## What Are Substreams?

A substream is a portion of a random stream that RandStream can access quickly. There is a number $M$ such that for any positive integer $k$, RandStream can go the kMth pseudorandom number in the stream. From that point, RandStream can generate the subsequent entries in the stream. Currently, RandStream has $M=2^{72}$, about 5e21, or more.

Beginning
of stream
M 2M
3M
$\mid$ Substream 1 Substream 2 Substream 3|

The entries in different substreams have good statistical properties, similar to the properties of entries in a single stream: independence, and lack of $k$-way correlation at various lags. The substreams are so long that you can view the substreams as being independent streams, as in the following picture.

| Substream 1 |
| :--- |
| Random Number 1 |
| Random Number 2 |
| Random Number 3 |
| $\ldots$ |
|  |



Substream 3
Random Number 1
Random Number 2 Random Number 3

Two RandStream stream types support substreams: 'mlfg6331_64' and 'mrg32k3a'.

## How Substreams Enable Reproducible Parallel Computations

When MATLAB performs computations in parallel with parfor, each worker receives loop iterations in an unpredictable order. Therefore, you cannot
predict which worker gets which iteration, so cannot determine the random numbers associated with each iteration.

Substreams allow MATLAB to tie each iteration to a particular sequence of random numbers. parfor gives each iteration an index. The iteration uses the index as the substream number. Since the random numbers are associated with the iterations, not with the workers, the entire computation is reproducible.

To obtain reproducible results, simply reset the stream, and all the substreams generate identical random numbers when called again. This method succeeds when all the workers use the same stream, and the stream supports substreams. This concludes the discussion of how the procedure in "Running Reproducible Parallel Computations" on page 17-14 gives reproducible parallel results.

## Random Numbers on the Client or Workers

A few functions generate random numbers on the client before distributing them to parallel workers. The workers do not use random numbers, so operate purely deterministically. For these functions, you can run a parallel computation reproducibly using any random stream type.

The functions that operate this way include:

- crossval
- plsregress
- sequentialfs

To obtain identical results, reset the random stream on the client, or the random stream you pass to the client. For example:

```
s = rng % Obtain the current state of the random stream
% run the statistical function
rng(s) % Reset the stream to the previous state
% run the statistical function again, obtain identical results
```

While this method enables you to run reproducibly in parallel, the results can differ from a serial computation. The reason for the difference is parfor loops run in reverse order from for loops. Therefore, a serial computation can
generate random numbers in a different order than a parallel computation. For unequivocal reproducibility, use the technique in "Running Reproducible Parallel Computations" on page 17-14.

## Distributing Streams Explicitly

For testing or comparison using particular random number algorithms, you must set the random number generators. How do you set these generators in parallel, or initialize streams on each worker in a particular way? Or you might want to run a computation using a different sequence of random numbers than any other you have run. How can you ensure the sequence you use is statistically independent?

Parallel Statistics Toolbox functions allow you to set random streams on each worker explicitly. For information on creating multiple streams, enter help RandStream/create at the command line. To create four independent streams using the 'mrg32k3a' generator:

```
s = RandStream.create('mrg32k3a','NumStreams',4,...
    'CellOutput',true);
```

Pass these streams to a statistical function using the Streams option. For example:

```
matlabpool open 4 % if you have at least 4 cores
s = RandStream.create('mrg32k3a','NumStreams',4,...
    'CellOutput',true); % create 4 independent streams
paroptions = statset('UseParallel','always',...
    'Streams',s); % set the 4 different streams
x = [randn(700,1); 4 + 2*randn(300,1)];
latt = -4:0.01:12;
myfun = @(X) ksdensity(X,latt);
pdfestimate = myfun(x);
B = bootstrp(200,myfun,x,'Options',paroptions);
```

See "Example: Parallel Bootstrap" on page 17-21 for a plot of the results of this computation.

This method of distributing streams gives each worker a different stream for the computation. However, it does not allow for a reproducible computation, because the workers perform the 200 bootstraps in an unpredictable order. If
you want to perform a reproducible computation, use substreams as described in "Running Reproducible Parallel Computations" on page 17-14.

If you set the UseSubstreams option to 'always', then set the Streams option to a single random stream of the type that supports substreams ('mlfg6331_64' or 'mrg32k3a'). This setting gives reproducible computations.

## Examples of Parallel Statistical Functions

In this section...<br>"Example: Parallel Jackknife" on page 17-19<br>"Example: Parallel Cross Validation" on page 17-20<br>"Example: Parallel Bootstrap" on page 17-21

## Example: Parallel Jackknife

This example is from the jackknife function reference page, but runs in parallel.

```
matlabpool open
opts = statset('UseParallel','always');
sigma = 5;
y = normrnd(0,sigma,100,1);
m = jackknife(@var, y,1,'Options',opts);
n = length(y);
bias = -sigma^2 / n % known bias formula
jbias = (n - 1)*(mean(m)-var(y,1)) % jackknife bias estimate
bias =
    -0.2500
jbias =
    -0.2698
```

This simple example is not a good candidate for parallel computation:

```
% How long to compute in serial?
tic;m = jackknife(@var,y,1);toc
Elapsed time is 0.023852 seconds.
% How long to compute in parallel?
tic;m = jackknife(@var,y,1,'Options',opts);toc
Elapsed time is 1.911936 seconds.
```

jackknife does not use random numbers, so gives the same results every time, whether run in parallel or serial.

## Example: Parallel Cross Validation

- "Simple Parallel Cross Validation" on page 17-20
- "Reproducible Parallel Cross Validation" on page 17-20


## Simple Parallel Cross Validation

This example is the same as the first in the crossval function reference page, but runs in parallel.

```
matlabpool open
opts = statset('UseParallel','always');
load('fisheriris');
y = meas(:,1);
X = [ones(size(y,1),1),meas(:,2:4)];
regf=@(XTRAIN,ytrain,XTEST)(XTEST*regress(ytrain,XTRAIN));
cvMse = crossval('mse',X,y,'Predfun',regf,'Options',opts)
cvMse =
    0.0999
```

This simple example is not a good candidate for parallel computation:

```
% How long to compute in serial?
tic;cvMse = crossval('mse',X,y,'Predfun',regf);toc
Elapsed time is 0.046005 seconds.
% How long to compute in parallel?
tic;cvMse = crossval('mse',X,y,'Predfun',regf,...
    'Options',opts);toc
Elapsed time is 1.333021 seconds.
```


## Reproducible Parallel Cross Validation

To run crossval in parallel in a reproducible fashion, set the options and reset the random stream appropriately (see "Running Reproducible Parallel Computations" on page 17-14).

```
matlabpool open
```

```
s = RandStream('mlfg6331_64');
options = statset('UseParallel','always',...
    'Streams',s,'UseSubstreams',''always');
load('fisheriris');
y = meas(:,1);
X = [ones(size(y,1),1),meas(:,2:4)];
regf=@(XTRAIN, ytrain,XTEST)(XTEST*regress(ytrain,XTRAIN));
cvMse = crossval('mse',X,y,'Predfun',regf,'Options',opts)
cvMse =
    0.1020
```

Reset the stream and the result is identical:

```
reset(s)
cvMse = crossval('mse',X,y,'Predfun',regf,'Options',opts)
cvMse =
    0.1020
```


## Example: Parallel Bootstrap

- "Bootstrap in Serial and Parallel" on page 17-21
- "Reproducible Parallel Bootstrap" on page 17-23


## Bootstrap in Serial and Parallel

Here is an example timing a bootstrap in parallel versus in serial. The example generates data from a mixture of two Gaussians, constructs a nonparametric estimate of the resulting data, and uses a bootstrap to get a sense of the sampling variability.

1 Generate the data:

```
% Generate a random sample of size 1000,
% from a mixture of two Gaussian distributions
x = [randn(700,1); 4 + 2*randn(300,1)];
```

2 Construct a nonparametric estimate of the density from the data:

```
latt = -4:0.01:12;
myfun = @(X) ksdensity(X,latt);
pdfestimate = myfun(x);
```

3 Bootstrap the estimate to get a sense of its sampling variability. Run the bootstrap in serial for timing comparison.

```
tic;B = bootstrp(200,myfun,x);toc
Elapsed time is 17.455586 seconds.
```

4 Run the bootstrap in parallel for timing comparison:

```
matlabpool open
Starting matlabpool using the 'local' configuration ...
connected to 2 labs.
opt = statset('UseParallel','always');
tic;B = bootstrp(200,myfun,x,'Options',opt);toc
Elapsed time is 9.984345 seconds.
```

Computing in parallel is nearly twice as fast as computing in serial for this example.

Overlay the ksdensity density estimate with the 200 bootstrapped estimates obtained in the parallel bootstrap. You can get a sense of how to assess the accuracy of the density estimate from this plot.

```
hold on
for i=1:size(B,1),
    plot(latt,B(i,:),'c:')
end
plot(latt,pdfestimate);
xlabel('x');ylabel('Density estimate')
```



## Reproducible Parallel Bootstrap

To run the example in parallel in a reproducible fashion, set the options appropriately (see "Running Reproducible Parallel Computations" on page 17-14). First set up the problem and parallel environment as in "Bootstrap in Serial and Parallel" on page 17-21. Then set the options to use substreams along with a stream that supports substreams.

```
s = RandStream('mlfg6331_64'); % has substreams
opts = statset('UseParallel','always',...
    'Streams',s,'UseSubstreams', 'always');
B2 = bootstrp(200,myfun, x,'Options',opts);
```

To rerun the bootstrap and get the same result:

```
reset(s) % set the stream to initial state
B3 = bootstrp(200,myfun,x,'Options',opts);
isequal(B2,B3) % check if same results
ans =
    1
```


## Function Reference

File I/O (p. 18-2)
Data Organization (p. 18-3)
Descriptive Statistics (p. 18-8)
Statistical Visualization (p. 18-11)
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Hypothesis Tests (p. 18-31)
Analysis of Variance (p. 18-32)
Parametric Regression Analysis (p. 18-33)

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Cluster Analysis (p. 18-38)
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Design of Experiments (p. 18-56)
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Data file input/output
Data arrays and groups
Data summaries
Data patterns and trends
Modeling data frequency
Inferences from data
Modeling data variance
Continuous data models

Visualization and reduction
Identifying data categories
Identifying data categories
Categorical data models
Classification and regression via trees, bagging, boosting, and more

Stochastic data models
Systematic data collection
Production monitoring
Interactive tools
General purpose

File I/O

| caseread | Read case names from file |
| :--- | :--- |
| casewrite | Write case names to file |
| tblread | Read tabular data from file |
| tblwrite | Write tabular data to file |
| tdfread | Read tab-delimited file |
| xptread | Create dataset array from data <br> stored in SAS XPORT format file |

## Data Organization

Categorical Arrays (p. 18-3)
Dataset Arrays (p. 18-6)
Grouped Data (p. 18-7)

## Categorical Arrays

| addlevels (categorical) | Add levels to categorical array |
| :--- | :--- |
| cat (categorical) | Concatenate categorical arrays |
| categorical | Create categorical array <br> Convert categorical array to cell <br> array of strings |
| char (categorical) | Convert categorical array to <br> character array |
| circshift (categorical) | Shift categorical array circularly |
| ctranspose (categorical) | Transpose categorical matrix <br> Convert categorical array to double <br> array |
| double (categorical) | Drop levels <br> droplevels (categorical) |
| end (categorical) | Last index in indexing expression for <br> categorical array |
| flipdim (categorical) | Flip categorical array along specified <br> dimension |
| fliplr (categorical) | Flip categorical matrix in left/right <br> direction |
| flipud (categorical) | Flip categorical matrix in up/down <br> direction |
| getlabels (categorical) | Access categorical array labels |
| getlevels (categorical) | Get categorical array levels |

Plot histogram of categorical data
Horizontal concatenation for categorical arrays

Convert categorical array to signed 16-bit integer array

Convert categorical array to signed 32-bit integer array

Convert categorical array to signed 64-bit integer array

Convert categorical array to signed 8 -bit integer array
Set intersection for categorical arrays

Inverse permute dimensions of categorical array

True for empty categorical array
True if categorical arrays are equal
Test for levels
True for elements of categorical array in set

Test for membership
True if categorical array is scalar
Test for undefined elements
True if categorical array is vector
Length of categorical array
Element counts by level
Merge levels
Number of dimensions of categorical array

Construct nominal categorical array

| numel (categorical) | Number of elements in categorical <br> array |
| :--- | :--- |
| ordinal | Construct ordinal categorical array |
| permute (categorical) | Permute dimensions of categorical <br> array |
| reorderlevels (categorical) | Reorder levels |
| repmat (categorical) | Replicate and tile categorical array |
| reshape (categorical) | Resize categorical array |
| rot90 (categorical) | Rotate categorical matrix 90 degrees |
| setdiff (categorical) | Set difference for categorical arrays |
| setlabels (categorical) | Label levels |
| setxor (categorical) | Set exclusive-or for categorical |
| arrays |  |
| shiftdim (categorical) | Shift dimensions of categorical array |
| single (categorical) | Convert categorical array to single |
| array |  |

uint64 (categorical)
uint8 (categorical)
union (categorical)
unique (categorical)
vertcat (categorical)

## Dataset Arrays

| cat (dataset) | Concatenate dataset arrays |
| :--- | :--- |
| cellstr (dataset) | Create cell array of strings from <br> dataset array |
| dataset | Construct dataset array <br> datasetfun (dataset) <br> variables function to dataset array |
| double (dataset) | Convert dataset variables to double <br> array |
| end (dataset) | Last index in indexing expression for <br> dataset array |
| export (dataset) | Write dataset array to file |
| get (dataset) | Access dataset array properties |
| grpstats (dataset) | Summary statistics by group for <br> dataset arrays |
| horzcat (dataset) | Horizontal concatenation for dataset <br> arrays |
| isempty (dataset) | True for empty dataset array |
| join (dataset) | Merge observations |
| length (dataset) | Length of dataset array |

```
ndims (dataset)
numel (dataset)
replacedata (dataset)
set (dataset)
single (dataset)
size (dataset)
sortrows (dataset)
stack (dataset)
summary (dataset)
unique (dataset)
unstack (dataset)
vertcat (dataset)
```


## Grouped Data

gplotmatrix
grp2idx
grpstats
gscatter

Number of dimensions of dataset array

Number of elements in dataset array
Replace dataset variables
Set and display properties
Convert dataset variables to single array

Size of dataset array
Sort rows of dataset array
Stack data from multiple variables into single variable

Print summary of dataset array
Unique observations in dataset array
Unstack data from single variable into multiple variables

Vertical concatenation for dataset arrays

Matrix of scatter plots by group
Create index vector from grouping variable

Summary statistics by group
Scatter plot by group

## Descriptive Statistics

Summaries (p. 18-8)<br>Measures of Central Tendency<br>Measures of Dispersion (p. 18-8)<br>Measures of Shape (p. 18-9)<br>Statistics Resampling (p. 18-9)<br>Data with Missing Values (p. 18-9)<br>Data Correlation (p. 18-10)

## Summaries

crosstab
grpstats
summary (categorical)
tabulate

Cross-tabulation
Summary statistics by group
Summary statistics for categorical array
Frequency table

## Measures of Central Tendency

geomean
harmmean
trimmean

Geometric mean
Harmonic mean
Mean excluding outliers

## Measures of Dispersion

Interquartile range
Mean or median absolute deviation
moment
range

## Measures of Shape

kurtosis
moment
prctile
quantile
skewness
zscore

## Statistics Resampling

bootci

bootstrp
jackknife

Kurtosis
Central moments
Calculate percentile values
Quantiles
Skewness
Standardized $z$-scores

## Data with Missing Values

| nancov | Covariance ignoring NaN values |
| :--- | :--- |
| nanmax | Maximum ignoring NaN values |
| nanmean | Mean ignoring NaN values |
| nanmedian | Median ignoring NaN values |
| nanmin | Minimum ignoring NaN values |
| nanstd | Standard deviation ignoring NaN <br> values |
| nansum | Sum ignoring NaN values |
| nanvar | Variance, ignoring NaN values |

## Data Correlation

| canoncorr | Canonical correlation |
| :--- | :--- |
| cholcov | Cholesky-like covariance <br> decomposition |
| cophenet | Cophenetic correlation coefficient <br> corr <br> corrcov |
| partialcorr | Linear or rank correlation <br> Convert covariance matrix to <br> correlation matrix |
| tiedrank | Linear or rank partial correlation <br> coefficients |
|  | Rank adjusted for ties |

## Statistical Visualization

Distribution Plots (p. 18-11)
Scatter Plots (p. 18-12)
ANOVA Plots (p. 18-12)
Regression Plots (p. 18-13)
Multivariate Plots (p. 18-13)
Cluster Plots (p. 18-13)
Classification Plots (p. 18-14)
DOE Plots (p. 18-14)
SPC Plots (p. 18-14)
Distribution Plots

| boxplot |  |
| :--- | :--- |
| cdfplot | Box plot |
| dfittool | Empirical cumulative distribution <br> function plot |
| disttool | Interactive distribution fitting <br> Interactive density and distribution <br> plots |
| ecdfhist | Empirical cumulative distribution <br> function histogram |
| fsurfht | Interactive contour plot |
| hist3 | Bivariate histogram |
| histfit | Histogram with normal fit <br> normplot |
| normspec | Normal probability plot |
| normal density plot between |  |
| pareto | specifications |
| probplot | Pareto chart |

qqplot<br>randtool<br>scatterhist<br>surfht<br>wblplot

## Scatter Plots

gline
gname
gplotmatrix
gscatter
lsline
refcurve
refline
scatterhist

Quantile-quantile plot
Interactive random number generation
Scatter plot with marginal histograms
Interactive contour plot
Weibull probability plot

Interactively add line to plot
Add case names to plot
Matrix of scatter plots by group
Scatter plot by group
Add least-squares line to scatter plot
Add reference curve to plot
Add reference line to plot
Scatter plot with marginal histograms

## ANOVA Plots

## anova1

aoctool
manovacluster
multcompare

One-way analysis of variance
Interactive analysis of covariance
Dendrogram of group mean clusters following MANOVA
Multiple comparison test

## Regression Plots

addedvarplot
gline
lassoPlot
lsline
polytool
rcoplot
refcurve
refline
robustdemo
rsmdemo
rstool
view (classregtree)

## Multivariate Plots

andrewsplot
biplot
glyphplot
parallelcoords

## Cluster Plots

dendrogram

manovacluster
silhouette

Added-variable plot
Interactively add line to plot
Trace plot of lasso fit
Add least-squares line to scatter plot
Interactive polynomial fitting
Residual case order plot
Add reference curve to plot
Add reference line to plot
Interactive robust regression
Interactive response surface demonstration
Interactive response surface modeling
Plot tree

Andrews plot
Biplot
Glyph plot
Parallel coordinates plot

Dendrogram plot
Dendrogram of group mean clusters following MANOVA
Silhouette plot

## Classification Plots

perfcurve
view (classregtree)

## DOE Plots

interactionplot
maineffectsplot
multivarichart
rsmdemo
rstool

## SPC Plots

capaplot
controlchart
histfit
normspec

Compute Receiver Operating Characteristic (ROC) curve or other performance curve for classifier output

Plot tree

Interaction plot for grouped data Main effects plot for grouped data Multivari chart for grouped data Interactive response surface demonstration

Interactive response surface modeling

Process capability plot
Shewhart control charts
Histogram with normal fit
Normal density plot between specifications

## Probability Distributions

Distribution Objects (p. 18-15)
Distribution Plots (p. 18-16)
Probability Density (p. 18-17)
Cumulative Distribution (p. 18-19)
Inverse Cumulative Distribution
(p. 18-21)

Distribution Statistics (p. 18-23)
Distribution Fitting (p. 18-24)
Negative Log-Likelihood (p. 18-26)
Random Number Generators
(p. 18-26)

Quasi-Random Numbers (p. 18-28)
Piecewise Distributions (p. 18-29)

## Distribution Objects

| cdf (ProbDist) | Return cumulative distribution <br> function (CDF) for ProbDist object |
| :--- | :--- |
| fitdist | Fit probability distribution to data |
| icdf (ProbDistUnivKernel) | Return inverse cumulative <br> distribution function (ICDF) for <br> ProbDistUnivKernel object |
| icdf (ProbDistUnivParam) | Return inverse cumulative <br> distribution function (ICDF) for <br> ProbDistUnivParam object |
| iqr (ProbDistUnivKernel) | Return interquartile range (IQR) for <br> ProbDistUnivKernel object |
| iqr (ProbDistUnivParam) | Return interquartile range (IQR) for <br> ProbDistUnivParam object |

mean (ProbDistUnivParam)<br>median (ProbDistUnivKernel)<br>median (ProbDistUnivParam)<br>paramci (ProbDistUnivParam)<br>pdf (ProbDist)<br>ProbDistUnivKernel<br>ProbDistUnivParam<br>random (ProbDist)<br>std (ProbDistUnivParam)<br>var (ProbDistUnivParam)

Return mean of ProbDistUnivParam object

Return median of ProbDistUnivKernel object

Return median of ProbDistUnivParam object

Return parameter confidence intervals of ProbDistUnivParam object

Return probability density function (PDF) for ProbDist object

Construct ProbDistUnivKernel object

Construct ProbDistUnivParam object

Generate random number drawn from ProbDist object

Return standard deviation of ProbDistUnivParam object

Return variance of ProbDistUnivParam object

## Distribution Plots

boxplot<br>cdfplot<br>dfittool<br>disttool<br>ecdfhist

Box plot
Empirical cumulative distribution function plot

Interactive distribution fitting
Interactive density and distribution plots
Empirical cumulative distribution function histogram
fsurfht
hist3
histfit
normplot
normspec
pareto
probplot
qqplot
randtool
scatterhist
surfht
wblplot

## Probability Density

betapdf<br>binopdf<br>chi2pdf<br>copulapdf<br>disttool<br>evpdf<br>exppdf

Interactive contour plot
Bivariate histogram
Histogram with normal fit
Normal probability plot
Normal density plot between specifications

Pareto chart
Probability plots
Quantile-quantile plot
Interactive random number generation

Scatter plot with marginal histograms

Interactive contour plot
Weibull probability plot

Beta probability density function
Binomial probability density function

Chi-square probability density function

Copula probability density function
Interactive density and distribution plots

Extreme value probability density function

Exponential probability density function

| fpdf | F probability density function |
| :--- | :--- |
| gampdf |  |
| geopdf | Gamma probability density function <br> Geometric probability density <br> function |
| gevpdf | Generalized extreme value <br> probability density function |
| gppdf | Generalized Pareto probability <br> density function |
| hygepdf | Hypergeometric probability density <br> function |
| ksdensity | Kernel smoothing density estimate |
| lognpdf | Lognormal probability density <br> function |
| mnpdf | Multinomial probability density <br> function |
| mvnpdf | Multivariate normal probability <br> density function |
| mvtpdf | Multivariate $t$ probability density <br> function |
| nbinpdf | Negative binomial probability <br> density function |
| ncfpdf | Noncentral $F$ probability density <br> function |
| nctpdf | Noncentral $t$ probability density <br> function |
| ncx2pdf | Noncentral chi-square probability <br> density function |
| normpdf | Normal probability density function |
| pdf | Probability density functions <br> Grobability density function for |
| Gaussian mixture distribution |  |

```
pdf (piecewisedistribution)
poisspdf
random (piecewisedistribution)
raylpdf
tpdf
unidpdf
unifpdf
wblpdf
```


## Cumulative Distribution

| betacdf | Beta cumulative distribution <br> function |
| :--- | :--- |
| binocdf | Binomial cumulative distribution <br> function |
| cdf | Cumulative distribution functions |
| cdf (gmdistribution) | Cumulative distribution function for <br> Gaussian mixture distribution |
| cdf (piecewisedistribution) | Cumulative distribution function for <br> piecewise distribution |
| cdfplot | Empirical cumulative distribution <br> function plot |
| chi2cdf | Chi-square cumulative distribution <br> function |
| copulacdf | Copula cumulative distribution <br> function |


| disttool | Interactive density and distribution <br> plots |
| :--- | :--- |
| ecdf | Empirical cumulative distribution <br> function |
| ecdfhist | Empirical cumulative distribution <br> function histogram |
| evcdf | Extreme value cumulative <br> distribution function |
| expcdf | Exponential cumulative distribution <br> function |
| fcdf | Fcumulative distribution function <br> gamma cumulative distribution |
| gamcdf | Gunction |
| geocdf | Geometric cumulative distribution <br> function |
| gevcdf | Generalized extreme value <br> cumulative distribution function |
| gpcdf | Generalized Pareto cumulative <br> distribution function |
| hygecdf | Hypergeometric cumulative <br> distribution function |
| logncdf | Lognormal cumulative distribution <br> function |
| mvncdf | Multivariate normal cumulative <br> distribution function |
| nvtcdf | Multivariate $t$ cumulative <br> distribution function |
|  | Noncentral $F$ cumulative <br> distribution function |
|  | Noncentral $t$ cumulative distribution <br> function |
|  |  |


| ncx2cdf | Noncentral chi-square cumulative <br> distribution function |
| :--- | :--- |
| normcdf | Normal cumulative distribution <br> function |
| poisscdf | Poisson cumulative distribution <br> function <br> raylcdf |
| tcdf | Rayleigh cumulative distribution <br> function |
| unidcdf | Student's $t$ cumulative distribution <br> function |
| unifcdf | Discrete uniform cumulative <br> distribution function |
| wblcdf | Continuous uniform cumulative <br> distribution function |
|  | Weibull cumulative distribution <br> function |

## Inverse Cumulative Distribution

| betainv | Beta inverse cumulative distribution <br> function |
| :--- | :--- |
| binoinv | Binomial inverse cumulative <br> distribution function |
| chi2inv | Chi-square inverse cumulative <br> distribution function |
| evinv | Extreme value inverse cumulative <br> distribution function |
| expinv | Exponential inverse cumulative <br> distribution function |
| finv | $F$ inverse cumulative distribution <br> function |


| gaminv | Gamma inverse cumulative <br> distribution function |
| :--- | :--- |
| geoinv | Geometric inverse cumulative <br> distribution function |
| gevinv | Generalized extreme value inverse <br> cumulative distribution function |
| gpinv | Generalized Pareto inverse <br> cumulative distribution function |
| hygeinv | Hypergeometric inverse cumulative <br> distribution function |
| icdf | Inverse cumulative distribution <br> functions |
| icdf (piecewisedistribution) | Inverse cumulative distribution <br> function for piecewise distribution |
| logninv | Lognormal inverse cumulative <br> distribution function |
| nbininv | Negative binomial inverse <br> cumulative distribution function |
| ncfinv | Noncentral $F$ inverse cumulative <br> distribution function |
| nctinv | Noncentral $t$ inverse cumulative <br> distribution function |
| ncx2inv | Noncentral chi-square inverse <br> cumulative distribution function |
| norminv | Normal inverse cumulative <br> distribution function |
| poissinv | Poisson inverse cumulative <br> distribution function |
| raylinv | Rayleigh inverse cumulative <br> distribution function |
| Student's $t$ inverse cumulative |  |
| distribution function |  |

unidinv<br>unifinv<br>wblinv

Discrete uniform inverse cumulative distribution function

Continuous uniform inverse cumulative distribution function

Weibull inverse cumulative distribution function

## Distribution Statistics

| betastat | Beta mean and variance |
| :--- | :--- |
| binostat | Binomial mean and variance |
| chi2stat | Chi-square mean and variance |
| copulastat | Copula rank correlation |
| evstat | Extreme value mean and variance |
| expstat | Exponential mean and variance |
| fstat | $F$ mean and variance |
| gamstat | Gamma mean and variance |
| geostat | Geometric mean and variance |
| gevstat | Generalized extreme value mean <br> and variance |
| gpstat | Generalized Pareto mean and <br>  <br> variance |
| hygestat | Hypergeometric mean and variance |
| lognstat | Lognormal mean and variance |
| nbinstat | Negative binomial mean and |
|  | variance |
| ncfstat | Noncentral $F$ mean and variance |
| nctstat | Noncentral $t$ mean and variance |
| ncx2stat | Noncentral chi-square mean and <br> variance |

normstat<br>poisstat<br>raylstat<br>tstat<br>unidstat<br>unifstat<br>wblstat

## Distribution Fitting

Supported Distributions (p. 18-24)
Piecewise Distributions (p. 18-25)

## Supported Distributions

| betafit | Beta parameter estimates |
| :--- | :--- |
| binofit | Binomial parameter estimates |
| copulafit | Fit copula to data |
| copulaparam | Copula parameters as function of <br> rank correlation |
| dfittool | Interactive distribution fitting |
| evfit | Extreme value parameter estimates |
| expfit | Exponential parameter estimates |
| fit (gmdistribution) | Gaussian mixture parameter <br> estimates |
| gamfit | Gamma parameter estimates <br> gevfit |
|  | Generalized extreme value <br> parameter estimates |

gpfit
histfit
johnsrnd
lognfit
mle
mlecov
nbinfit
normfit
normplot
pearsrnd
poissfit
raylfit
unifit
wblfit
wblplot

Generalized Pareto parameter estimates
Histogram with normal fit Johnson system random numbers Lognormal parameter estimates Maximum likelihood estimates Asymptotic covariance of maximum likelihood estimators

Negative binomial parameter estimates

Normal parameter estimates
Normal probability plot Pearson system random numbers Poisson parameter estimates Rayleigh parameter estimates Continuous uniform parameter estimates
Weibull parameter estimates
Weibull probability plot

## Piecewise Distributions

boundary (piecewisedistribution)
lowerparams (paretotails)
nsegments (piecewisedistribution)
paretotails
piecewisedistribution
segment (piecewisedistribution)
upperparams (paretotails)

Piecewise distribution boundaries
Lower Pareto tails parameters
Number of segments
Construct Pareto tails object
Create piecewise distribution object
Segments containing values
Upper Pareto tails parameters

## Negative Log-Likelihood

\(\left.$$
\begin{array}{ll}\text { betalike } & \text { Beta negative log-likelihood } \\
\text { evlike } & \begin{array}{l}\text { Extreme value negative } \\
\text { log-likelihood }\end{array} \\
\text { explike } & \text { Exponential negative log-likelihood } \\
\text { gamlike } & \begin{array}{l}\text { Gamma negative log-likelihood }\end{array} \\
\text { gevlike } & \begin{array}{l}\text { Generalized extreme value negative } \\
\text { log-likelihood }\end{array} \\
\text { gplike } & \begin{array}{l}\text { Generalized Pareto negative } \\
\text { log-likelihood }\end{array} \\
\text { lognlike } & \begin{array}{l}\text { Lognormal negative log-likelihood } \\
\text { mvregresslike }\end{array}
$$ <br>
Negative log-likelihood for <br>

multivariate regression\end{array}\right\}\)| Normal negative log-likelihood |
| :--- |
| wbllike |

## Random Number Generators

betarnd
binornd
chi2rnd
copularnd
datasample
evrnd
exprnd
frnd
gamrnd
geornd

Beta random numbers
Binomial random numbers
Chi-square random numbers
Copula random numbers
Randomly sample from data, with or without replacement

Extreme value random numbers
Exponential random numbers
$F$ random numbers
Gamma random numbers
Geometric random numbers

```
gevrnd
gprnd
hygernd
iwishrnd
johnsrnd
lhsdesign
lhsnorm
lognrnd
mhsample
mnrnd
mvnrnd
mvtrnd
nbinrnd
ncfrnd
nctrnd
ncx2rnd
normrnd
pearsrnd
poissrnd
randg
random
random (gmdistribution)
random (piecewisedistribution)
gevrnd
gprnd
hygernd
iwishrnd
johnsrnd
lhsdesign
lhsnorm
lognrnd
mhsample
mnrnd
mvnrnd
mvtrnd
nbinrnd
ncfrnd
nctrnd
ncx2rnd
normrnd
pearsrnd
poissrnd
randg
random
random (gmdistribution)
random (piecewisedistribution)
```

Generalized extreme value random numbers

Generalized Pareto random numbers
Hypergeometric random numbers
Inverse Wishart random numbers
Johnson system random numbers
Latin hypercube sample
Latin hypercube sample from normal distribution

Lognormal random numbers
Metropolis-Hastings sample
Multinomial random numbers
Multivariate normal random numbers

Multivariate $t$ random numbers Negative binomial random numbers Noncentral $F$ random numbers Noncentral $t$ random numbers Noncentral chi-square random numbers

Normal random numbers
Pearson system random numbers
Poisson random numbers
Gamma random numbers
Random numbers
Random numbers from Gaussian mixture distribution

Random numbers from piecewise distribution

| randsample | Random sample |
| :--- | :--- |
| randtool | Interactive random number <br> generation |
| raylrnd | Rayleigh random numbers |
| slicesample | Slice sampler |
| trnd | Student's $t$ random numbers |
| unidrnd | Discrete uniform random numbers |
| unifrnd | Continuous uniform random |
| wblrnd | numbers |
| wishrnd | Weibull random numbers |
|  | Wishart random numbers |

## Quasi-Random Numbers

| addlistener (qrandstream) | Add listener for event |
| :--- | :--- |
| delete (qrandstream) | Delete handle object |
| end (qrandset) | Last index in indexing expression for <br> point set |
| eq (qrandstream) | Test handle equality <br> findobj (qrandstream) |
| findprop (qrandstream) | Find objects matching specified <br> conditions |
| ge (qrandstream) | Find property of MATLAB handle <br> object |
| gt (qrandstream) | Greater than or equal relation for <br> handles |
| haltonset | Greater than relation for handles |
| isvalid (qrandstream) | Construct Halton quasi-random <br> point set |
|  | Test handle validity |

```
le (qrandstream)
length (qrandset)
lt (qrandstream)
ndims (qrandset)
ne (qrandstream)
net (qrandset)
notify (qrandstream)
qrand (qrandstream)
qrandset
qrandstream
rand (qrandstream)
reset (qrandstream)
scramble (qrandset)
size (qrandset)
sobolset
```


## Piecewise Distributions

boundary (piecewisedistribution)
cdf (piecewisedistribution)
icdf (piecewisedistribution)
lowerparams (paretotails)

Less than or equal relation for handles

Length of point set
Less than relation for handles
Number of dimensions in matrix
Not equal relation for handles
Generate quasi-random point set
Notify listeners of event
Generate quasi-random points from stream

Abstract quasi-random point set class

Construct quasi-random number stream

Generate quasi-random points from stream

Reset state
Scramble quasi-random point set Number of dimensions in matrix Construct Sobol quasi-random point set

Piecewise distribution boundaries
Cumulative distribution function for piecewise distribution

Inverse cumulative distribution function for piecewise distribution

Lower Pareto tails parameters

| nsegments (piecewisedistribution) | Number of segments |
| :--- | :--- |
| paretotails | Construct Pareto tails object |
| pdf (piecewisedistribution) | Probability density function for <br> piecewise distribution |
| piecewisedistribution | Create piecewise distribution object |
| random (piecewisedistribution) | Random numbers from piecewise <br> distribution |
| segment (piecewisedistribution) | Segments containing values |
| upperparams (paretotails) | Upper Pareto tails parameters |

## Hypothesis Tests

| ansaribradley | Ansari-Bradley test |
| :--- | :--- |
| barttest | Bartlett's test |
| canoncorr | Canonical correlation |
| chi2gof | Chi-square goodness-of-fit test |
| dwtest | Durbin-Watson test |
| friedman | Friedman's test |
| jbtest | Jarque-Bera test |
| kruskalwallis | Kruskal-Wallis test |
| kstest | One-sample Kolmogorov-Smirnov |
| kstest2 | test |
| lillietest | Two-sample Kolmogorov-Smirnov |
| linhyptest | test |
| ranksum | Lilliefors test |
| runstest | Linear hypothesis test |
| sampsizepwr | Wilcoxon rank sum test |
| signrank | Run test for randomness |
| signtest | Sample size and power of test |
| ttest | Wilcoxon signed rank test |
| ttest2 | Sign test |
| vartest | One-sample and paired-sample $t$-test |
| vartest2 | Two-sample $t$-test |
| vartestn | Chi-square variance test |
|  | Two-sample $F$-test for equal |
|  | variances |
| Bartlett multiple-sample test for |  |
| equal variances |  |

zscore
ztest

Standardized $z$-scores
$z$-test

## Analysis of Variance

ANOVA Plots (p. 18-32)
ANOVA Operations (p. 18-32)

## ANOVA Plots

anova 1
aoctool
manovacluster
multcompare

## ANOVA Operations

anova1
anova2
anovan
aoctool
dummyvar
friedman
kruskalwallis
manova1

One-way analysis of variance Interactive analysis of covariance

Dendrogram of group mean clusters following MANOVA

Multiple comparison test

One-way analysis of variance
Two-way analysis of variance
$N$-way analysis of variance
Interactive analysis of covariance
Create dummy variables
Friedman's test
Kruskal-Wallis test
One-way multivariate analysis of variance

## manovacluster <br> multcompare

## Parametric Regression Analysis

Regression Plots (p. 18-33)
Linear Regression (p. 18-34)
Nonlinear Regression (p. 18-35)

## Regression Plots

| addedvarplot | Added-variable plot |
| :--- | :--- |
| gline | Interactively add line to plot |
| lassoPlot | Trace plot of lasso fit |
| lsline | Add least-squares line to scatter plot |
| polytool | Interactive polynomial fitting |
| rcoplot | Residual case order plot |
| refcurve | Add reference curve to plot |
| refline | Add reference line to plot |
| robustdemo | Interactive robust regression |
| rsmdemo | Interactive response surface <br> demonstration |
| rstool | Interactive response surface <br> modeling |
| view (classregtree) | Plot tree |

## Linear Regression

| coxphfit | Cox proportional hazards regression |
| :--- | :--- |
| dummyvar | Create dummy variables |
| glmfit | Generalized linear model regression |
| glmval | Generalized linear model values |
| invpred | Inverse prediction |
| lasso | Regularized least-squares regression <br> using lasso or elastic net algorithms |
| leverage | Leverage |
| mnrfit | Multinomial logistic regression |
| mnrval | Multinomial logistic regression <br> mvregress |
| values |  |
| mvregresslike | Multivariate linear regression |
| plsregress | Negative log-likelihood for <br> multivariate regression |
| polyconf | Partial least-squares regression |
| polytool | Polynomial confidence intervals |
| regress | Interactive polynomial fitting |
| regstats | Multiple linear regression |
| ridge | Regression diagnostics |
| robustdemo | Ridge regression |
| robustfit | Interactive robust regression |
| rsmdemo | Robust regression |
| rstool | Interactive response surface |
| demonstration |  |
| stepwise | Interactive response surface |
| modeling |  |

stepwisefit
x2fx

## Nonlinear Regression

dummyvar<br>hougen<br>nlinfit<br>nlintool<br>nlmefit<br>nlmefitsa<br>nlparci<br>nlpredci

Stepwise regression
Convert predictor matrix to design matrix

Create dummy variables
Hougen-Watson model
Nonlinear regression
Interactive nonlinear regression
Nonlinear mixed-effects estimation
Fit nonlinear mixed effects model with stochastic EM algorithm

Nonlinear regression parameter confidence intervals

Nonlinear regression prediction confidence intervals

## Multivariate Methods

Multivariate Plots (p. 18-36)
Multidimensional Scaling (p. 18-36)
Procrustes Analysis (p. 18-36)
Feature Selection (p. 18-37)
Feature Transformation (p. 18-37)

## Multivariate Plots

andrewsplot
Andrews plot
biplot
glyphplot
parallelcoords
Biplot
Glyph plot
Parallel coordinates plot

## Multidimensional Scaling

cmdscale
mahal
mdscale
pdist
squareform

## Procrustes Analysis

Classical multidimensional scaling
Mahalanobis distance
Nonclassical multidimensional scaling

Pairwise distance between pairs of objects
Format distance matrix

Procrustes analysis

## Feature Selection

sequentialfs

Sequential feature selection

## Feature Transformation

Nonnegative Matrix Factorization (p. 18-37)

Principal Component Analysis
(p. 18-37)

Factor Analysis (p. 18-37)

## Nonnegative Matrix Factorization

```
nnmf Nonnegative matrix factorization
```


## Principal Component Analysis

barttest
pareto
pcacov
pcares
princomp

Bartlett's test
Pareto chart
Principal component analysis on covariance matrix

Residuals from principal component analysis
Principal component analysis (PCA) on data

Factor analysis

## Cluster Analysis

Cluster Plots (p. 18-38)<br>Hierarchical Clustering (p. 18-38)<br>K-Means Clustering (p. 18-39)<br>Gaussian Mixture Models (p. 18-39)

## Cluster Plots

| dendrogram | Dendrogram plot |
| :--- | :--- |
| manovacluster | Dendrogram of group mean clusters <br> following MANOVA |
| silhouette | Silhouette plot |

## Hierarchical Clustering

| cluster | Construct agglomerative clusters <br> from linkages |
| :--- | :--- |
| clusterdata | Agglomerative clusters from data |
| cophenet | Cophenetic correlation coefficient |
| inconsistent | Inconsistency coefficient |
| linkage | Agglomerative hierarchical cluster <br> tree |
| pdist | Pairwise distance between pairs of <br> objects |
| squareform | Format distance matrix |

## K-Means Clustering

| kmeans | $K$-means clustering |
| :--- | :--- |
| mahal | Mahalanobis distance |

## Gaussian Mixture Models

cdf (gmdistribution)
cluster (gmdistribution)
fit (gmdistribution)
gmdistribution
mahal (gmdistribution)
pdf (gmdistribution)
posterior (gmdistribution)
random (gmdistribution)

Cumulative distribution function for Gaussian mixture distribution

Construct clusters from Gaussian mixture distribution

Gaussian mixture parameter estimates

Construct Gaussian mixture distribution

Mahalanobis distance to component means

Probability density function for Gaussian mixture distribution Posterior probabilities of components
Random numbers from Gaussian mixture distribution

## Model Assessment

```
confusionmat
crossval
cvpartition
repartition (cvpartition)
```

Confusion matrix
Loss estimate using cross-validation
Create cross-validation partition for data

Repartition data for cross-validation

```
test (cvpartition)
training (cvpartition)
```


## Parametric Classification

Discriminant Analysis (p. 18-40)
Naive Bayes Classification (p. 18-41)
Classification Plots (p. 18-42)

## Discriminant Analysis

ClassificationDiscriminant
ClassificationPartitionedModel
classify
compact
(ClassificationDiscriminant)
CompactClassificationDiscriminant crossval
(ClassificationDiscriminant)
edge
(CompactClassificationDiscriminant)
fit (ClassificationDiscriminant)
kfoldEdge
(ClassificationPartitionedModel)
kfoldfun
(ClassificationPartitionedModel)
kfoldLoss
(ClassificationPartitionedModel)

Test indices for cross-validation
Training indices for cross-validation

Discriminant analysis classification Cross-validated classification model Discriminant analysis

Compact discriminant analysis classifier

Compact discriminant analysis class
Cross-validated discriminant analysis classifier

Classification edge

Fit discriminant analysis classifier
Classification edge for observations not used for training

Cross validate function

Classification loss for observations not used for training

| kfoldMargin <br> (ClassificationPartitionedModel) | Classification margins for <br> observations not used for training |
| :--- | :--- |
| kfoldPredict |  |
| (ClassificationPartitionedModel) | Predict response for observations not <br> used for training |
| loss | Classification error |
| (CompactClassificationDiscriminant) |  |
| mahal | Mahalanobis distance to class means |
| (CompactClassificationDiscriminant) |  |
| make (ClassificationDiscriminant) | Construct discriminant analysis <br> classifier from parameters |
| margin | Classification margins |
| (CompactClassificationDiscriminant) | Predict classification |
| predict |  |
| (CompactClassificationDiscriminant) | Classification edge by resubstitution |
| resubEdge |  |
| (ClassificationDiscriminant) | Classification error by resubstitution |
| resubLoss <br> (ClassificationDiscriminant) | Classification margins by <br> resubMargin <br> (ClassificationDiscriminant) |
| resubstitution |  |
| resubPredict | Predict resubstitution response of |
| classifier |  |

## Naive Bayes Classification

fit (NaiveBayes)<br>NaiveBayes<br>posterior (NaiveBayes)<br>predict (NaiveBayes)

Create Naive Bayes classifier object by fitting training data

Create NaiveBayes object
Compute posterior probability of each class for test data

Predict class label for test data

## Classification Plots

perfcurve
view (classregtree)

Compute Receiver Operating Characteristic (ROC) curve or other performance curve for classifier output

Plot tree

## Nonparametric Supervised Learning

## Distance Computation and Nearest Neighbor Search

| createns | Create object to use in $k$-nearest <br> neighbors search |
| :--- | :--- |
| knnsearch | Find $k$-nearest neighbors using data |
| knnsearch (ExhaustiveSearcher) | Find $k$-nearest neighbors using <br> ExhaustiveSearcher object |
| knnsearch (KDTreeSearcher) | Find $k$-nearest neighbors using <br> KDTreeSearcher object |
| pdist | Pairwise distance between pairs of <br> objects |
| pdist2 | Pairwise distance between two sets <br> of observations |
| rangesearch | Find all neighbors within specified <br> distance |
| rangesearch (ExhaustiveSearcher) | Find all neighbors within specified <br> distance using object |
| rangesearch (KDTreeSearcher) | Find all neighbors within specified <br> distance using object |
| relieff | Importance of attributes (predictors) <br> using ReliefF algorithm |

## Classification Trees

| catsplit (classregtree) | Categorical splits used for branches <br> in decision tree |
| :--- | :--- |
| children (classregtree) | Child nodes |
| classcount (classregtree) | Class counts |
| ClassificationPartitionedModel | Cross-validated classification model |


| ClassificationTree | Binary decision tree for classification |
| :---: | :---: |
| classname (classregtree) | Class names for classification decision tree |
| classprob (classregtree) | Class probabilities |
| classregtree | Construct classification and regression trees |
| classregtree | Classification and regression trees |
| compact (ClassificationTree) | Compact tree |
| CompactClassificationTree | Compact classification tree |
| crossval (ClassificationTree) | Cross-validated decision tree |
| cutcategories (classregtree) | Cut categories |
| cutpoint (classregtree) | Decision tree cut point values |
| cuttype (classregtree) | Cut types |
| cutvar (classregtree) | Cut variable names |
| cvloss (ClassificationTree) | Classification error by cross validation |
| edge (CompactClassificationTree) | Classification edge |
| eval (classregtree) | Predicted responses |
| fit (ClassificationTree) | Fit classification tree |
| isbranch (classregtree) | Test node for branch |
| kfoldEdge (ClassificationPartitionedModel) | Classification edge for observations not used for training |
| kfoldfun (ClassificationPartitionedModel) | Cross validate function |
| kfoldLoss <br> (ClassificationPartitionedModel) | Classification loss for observations not used for training |
| kfoldMargin (ClassificationPartitionedModel) | Classification margins for observations not used for training |
| kfoldPredict <br> (ClassificationPartitionedModel) | Predict response for observations not used for training |


| loss (CompactClassificationTree) | Classification error |
| :--- | :--- |
| margin (CompactClassificationTree) | Classification margins |
| meansurrvarassoc (classregtree) | Mean predictive measure of <br> association for surrogate splits in <br> decision tree |
| meanSurrVarAssoc | Mean predictive measure of <br> association for surrogate splits in <br> decision tree |
| (CompactClassificationTree) | Class values of nodes of classification |
| nodeclass (classregtree) | Return vector of node errors |


| surrcutcategories (classregtree) | Categories used for surrogate splits <br> in decision tree |
| :--- | :--- |
| surrcutflip (classregtree) | Numeric cutpoint assignments used <br> for surrogate splits in decision tree |
| surrcutpoint (classregtree) | Cutpoints used for surrogate splits <br> in decision tree |
| surrcuttype (classregtree) | Types of surrogate splits used at <br> branches in decision tree |
| surrcutvar (classregtree) | Variables used for surrogate splits <br> in decision tree |
| surrvarassoc (classregtree) | Predictive measure of association for <br> surrogate splits in decision tree |
| template (ClassificationTree) | Create classification template |
| test (classregtree) | Error rate |
| type (classregtree) | Tree type |
| varimportance (classregtree) | Compute embedded estimates of |
| view (classregtree) | input feature importance |

## Regression Trees

| catsplit (classregtree) | Categorical splits used for branches <br> in decision tree |
| :--- | :--- |
| children (classregtree) | Child nodes |
| classregtree | Construct classification and <br> regression trees |
| classregtree | Classification and regression trees |
| compact (RegressionTree) | Compact regression tree |
| CompactRegressionTree | Compact regression tree |


| crossval (RegressionTree) | Cross-validated decision tree |
| :--- | :--- |
| cutcategories (classregtree) | Cut categories |
| cutpoint (classregtree) | Decision tree cut point values |
| cuttype (classregtree) | Cut types |
| cutvar (classregtree) | Cut variable names |
| cvloss (RegressionTree) | Regression error by cross validation |
| eval (classregtree) | Predicted responses |
| fit (RegressionTree) | Binary decision tree for regression |
| isbranch (classregtree) | Test node for branch |
| kfoldfun | Cross validate function |
| (RegressionPartitionedModel) | Cross-validation loss of partitioned |
| kfoldLoss | regression model |
| (RegressionPartitionedModel) | Predict response for observations not |
| kfoldPredict | used for training. |
| (RegressionPartitionedModel) | Regression error |
| loss (CompactRegressionTree) | Mean predictive measure of |
| meansurrvarassoc (classregtree) | association for surrogate splits in <br> decision tree |
|  | Mean predictive measure of <br> meanSurrVarAssoc |
| association for surrogate splits in |  |
| (CompactRegressionTree) | decision tree |
| nodeerr (classregtree) | Return vector of node errors |
| nodemean (classregtree) | Mean values of nodes of regression |
| nodeprob (classregtree) | tree |
| nodesize (classregtree) | Node probabilities |
| numnodes (classregtree) | Return node size |
| parent (classregtree) | Number of nodes |
| predict (CompactRegressionTree) | Predict response of regression tree |
|  |  |
| Parent node |  |


| predictorImportance <br> (CompactRegressionTree) <br> prune (classregtree) | Estimates of predictor importance |
| :--- | :--- |
| prune (RegressionTree) | Prune tree |
| prunelist (classregtree) | Produce sequence of subtrees by <br> pruning |
| RegressionPartitionedModel | Pruning levels for decision tree <br> nodes |
| RegressionTree | Cross-validated regression model <br> Resression tree |
| resubLoss (RegressionTree) | Regression error by resubstitution |
| resubPredict (RegressionTree) | Predict resubstitution response of <br> tree |
| risk (classregtree) | Node risks <br> Categories used for surrogate splits |
| surrcutcategories (classregtree) | in decision tree |
| surrcutflip (classregtree) | Numeric cutpoint assignments used <br> for surrogate splits in decision tree |
| surrcutpoint (classregtree) | Cutpoints used for surrogate splits <br> in decision tree |
| surrcuttype (classregtree) | Types of surrogate splits used at <br> branches in decision tree |
| surrcutvar (classregtree) | Variables used for surrogate splits <br> in decision tree |
| surrvarassoc (classregtree) | Predictive measure of association for <br> surrogate splits in decision tree |
| template (RegressionTree) | Create regression template |
| test (classregtree) | Error rate |
| type (classregtree) | Tree type |
| varimportance (classregtree) | Compute embedded estimates of <br> input feature importance |


| view (classregtree) | Plot tree |
| :--- | :--- |
| view (CompactRegressionTree) | View tree |

## Ensemble Methods - Classification

| append (TreeBagger) | Append new trees to ensemble |
| :--- | :--- |
| ClassificationBaggedEnsemble | Classification ensemble grown by <br> resampling |
| ClassificationEnsemble | Ensemble classifier |
| ClassificationPartitionedEnsemble | Cross-validated classification <br> ensemble |
| combine (CompactTreeBagger) | Combine two ensembles |
| compact (ClassificationEnsemble) | Compact classification ensemble <br> Compact ensemble of decision trees |
| compact (TreeBagger) | Compact classification ensemble <br> Class |
| CompactClassificationEnsemble | Compact ensemble of decision trees <br> grown by bootstrap aggregation |
| CompactTreeBagger | Cross validate ensemble |
| crossval (ClassificationEnsemble) | Classification edge |
| edge | Error (misclassification probability <br> or MSE) |
| (CompactClassificationEnsemble) | Error (misclassification probability <br> or MSE) |
| error (CompactTreeBagger) | Proximity matrix for training data |
| error (TreeBagger) | Fitted ensemble for classification or <br> regression |
| fillProximities (TreeBagger) | Train additional trees and add to <br> ensemble |
| fitensemble | growTrees (TreeBagger) |


| kfoldEdge |  |
| :--- | :--- |
| (ClassificationPartitionedEnsemble) | Classification edge for observations |
| not used for training |  |
| kfoldfun |  |
| (ClassificationPartitionedModel) | Cross validate function |
| kfoldLoss |  |
| (ClassificationPartitionedEnsemble) | Classification loss for observations |
| not used for training |  |
| kfoldMargin | Classification margins for <br> (ClassificationPartitionedModel) |
| observations not used for training <br> kfoldPredict <br> (ClassificationPartitionedModel) | Predict response for observations not <br> used for training |
| loss |  |
| (CompactClassificationEnsemble) | Classification error |
| margin | Classification margins |
| (CompactClassificationEnsemble) | Classification margin |
| margin (CompactTreeBagger) | Classification margin |
| margin (TreeBagger) | Multidimensional scaling of |
| mdsProx (CompactTreeBagger) | proximity matrix |
| mdsProx (TreeBagger) | Multidimensional scaling of |
| proximity matrix |  |

```
oobPredict (ClassificationBaggedEnsemble)
oobPredict (TreeBagger)
```

outlierMeasure (CompactTreeBagger) predict (CompactClassificationEnsemble) predict (CompactTreeBagger) predict (TreeBagger)
predictorImportance (CompactClassificationEnsemble)
proximity (CompactTreeBagger) resubEdge (ClassificationEnsemble) resubLoss (ClassificationEnsemble) resubMargin (ClassificationEnsemble)
resubPredict
(ClassificationEnsemble)
resume (ClassificationEnsemble)
resume
(ClassificationPartitionedEnsemble)
SetDefaultYfit (CompactTreeBagger)

TreeBagger

Predict out-of-bag response of ensemble

Ensemble predictions for out-of-bag observations

Outlier measure for data

Predict classification

Predict response
Predict response
Estimates of predictor importance

Proximity matrix for data
Classification edge by resubstitution
Classification error by resubstitution
Classification margins by resubstitution
Predict ensemble response by resubstitution

Resume training ensemble
Resume training learners on cross-validation folds

Set default value for predict

Bootstrap aggregation for ensemble of decision trees

## Ensemble Methods - Regression

append (TreeBagger)
combine (CompactTreeBagger)

Append new trees to ensemble Combine two ensembles

| compact (RegressionEnsemble) | Create compact regression ensemble |
| :--- | :--- |
| compact (TreeBagger) | Compact ensemble of decision trees |
| CompactRegressionEnsemble | Compact regression ensemble class |
| CompactTreeBagger | Compact ensemble of decision trees <br> grown by bootstrap aggregation |
| crossval (RegressionEnsemble) | Cross validate ensemble |
| cvshrink (RegressionEnsemble) | Cross validate shrinking (pruning) <br> ensemble |
| error (CompactTreeBagger) | Error (misclassification probability <br> or MSE) |
| error (TreeBagger) | Error (misclassification probability <br> or MSE) |
| fillProximities (TreeBagger) | Proximity matrix for training data |
| fitensemble | Fitted ensemble for classification or <br> regression |
| growTrees (TreeBagger) | Train additional trees and add to <br> ensemble |
| kfoldfun | Cross validate function |
| (RegressionPartitionedModel) | Cross-validation loss of partitioned <br> kfoldLoss <br> regression ensemble |
| kfoldPredict | Predict response for observations not |
| (RegressionPartitionedModel) | used for training. |
| loss (CompactRegressionEnsemble) | Regression error |
| mdsProx (CompactTreeBagger) | Multidimensional scaling of <br> proximity matrix |
| mdsProx (TreeBagger) | Multidimensional scaling of <br> proximity matrix |
| meanMargin (CompactTreeBagger) | Mean classification margin <br> Out-of-bag error |
| oobError (TreeBagger) |  |


| oobLoss <br> (RegressionBaggedEnsemble) | Out-of-bag regression error |
| :--- | :--- |
| oobPredict |  |
| (RegressionBaggedEnsemble) |  |
| oobPredict (TreeBagger) |  |$\quad$| Predict out-of-bag response of |
| :--- |
| ensemble |
| Ensemble predictions for out-of-bag |
| observations |



## Hidden Markov Models

hmmdecode<br>hmmestimate<br>hmmgenerate<br>hmmtrain<br>hmmviterbi

Hidden Markov model posterior state probabilities

Hidden Markov model parameter estimates from emissions and states

Hidden Markov model states and emissions

Hidden Markov model parameter estimates from emissions

Hidden Markov model most probable state path

## Design of Experiments

DOE Plots (p. 18-56)<br>Full Factorial Designs (p. 18-56)<br>Fractional Factorial Designs<br>(p. 18-57)<br>Response Surface Designs (p. 18-57)<br>D-Optimal Designs (p. 18-57)<br>Latin Hypercube Designs (p. 18-57)<br>Quasi-Random Designs (p. 18-58)

## DOE Plots

interactionplot
maineffectsplot
multivarichart
rsmdemo
rstool

Interaction plot for grouped data
Main effects plot for grouped data
Multivari chart for grouped data
Interactive response surface demonstration

Interactive response surface modeling

Two-level full factorial design Full factorial design

## Fractional Factorial Designs

fracfact<br>fracfactgen<br>Fractional factorial design<br>Fractional factorial design generators

## Response Surface Designs

bbdesign
ccdesign

## D-Optimal Designs

candexch

candgen
cordexch
daugment
dcovary
rowexch
rsmdemo

Box-Behnken design
Central composite design
$D$-optimal design from candidate set using row exchanges

Candidate set generation
Coordinate exchange
$D$-optimal augmentation
$D$-optimal design with fixed covariates

Row exchange
Interactive response surface demonstration

## Latin Hypercube Designs

| lhsdesign | Latin hypercube sample |
| :--- | :--- |
| lhsnorm | Latin hypercube sample from normal <br> distribution |

## Quasi-Random Designs

| addlistener (qrandstream) | Add listener for event |
| :--- | :--- |
| delete (qrandstream) | Delete handle object |
| end (qrandset) | Last index in indexing expression for <br> point set |
| eq (qrandstream) | Test handle equality |
| findobj (qrandstream) | Find objects matching specified <br> conditions |
| findprop (qrandstream) | Find property of MATLAB handle <br> object |
| ge (qrandstream) | Greater than or equal relation for <br> handles |
| gt (qrandstream) | Greater than relation for handles |
| haltonset | Construct Halton quasi-random <br> point set |
| isvalid (qrandstream) | Test handle validity |
| le (qrandstream) | Less than or equal relation for <br> handles |
| length (qrandset) | Length of point set <br> lt (qrandstream) |
| ndims (qrandset) | Less than relation for handles <br> ne (qrandstream) |
| net (qrandset) | Number of dimensions in matrix |
| notify (qrandstream) | Generate quasi-random point set |
| qrand (qrandstream) | Notify listeners of event |
| qrandset | Generate quasi-random points from <br> stream |
| qrandstream | Abstract quasi-random point set <br> class |
| stream |  |

rand (qrandstream)<br>reset (qrandstream)<br>scramble (qrandset)<br>size (qrandset)<br>sobolset

Generate quasi-random points from stream

Reset state
Scramble quasi-random point set
Number of dimensions in matrix
Construct Sobol quasi-random point set

## Statistical Process Control

SPC Plots (p. 18-60)
SPC Functions (p. 18-60)

## SPC Plots

capaplot
controlchart
histfit
normspec

## SPC Functions

capability
controlrules
gagerr

Process capability plot
Shewhart control charts
Histogram with normal fit
Normal density plot between specifications

Process capability indices
Western Electric and Nelson control rules

Gage repeatability and reproducibility study

## GUIs

| aoctool | Interactive analysis of covariance |
| :--- | :--- |
| dfittool | Interactive distribution fitting |
| disttool | Interactive density and distribution <br> plots |
| fsurfht | Interactive contour plot |
| polytool | Interactive polynomial fitting |
| randtool | Interactive random number <br> generation |
| regstats | Regression diagnostics |
| robustdemo | Interactive robust regression |
| rsmdemo | Interactive response surface <br> demonstration |
| rstool | Interactive response surface |
| modeling |  |$\quad$| Interactive contour plot |
| :--- |

## Utilities

| combnk | Enumeration of combinations |
| :--- | :--- |
| perms | Enumeration of permutations |
| statget | Access values in statistics options |
| structure |  |
| statset | Create statistics options structure |
| zscore | Standardized $z$-scores |

## Class Reference

- "Data Organization" on page 19-2
- "Probability Distributions" on page 19-3
- "Gaussian Mixture Models" on page 19-4
- "Model Assessment" on page 19-4
- "Parametric Classification" on page 19-5
- "Supervised Learning" on page 19-6
- "Quasi-Random Design of Experiments" on page 19-8


## Data Organization

In this section...<br>"Categorical Arrays" on page 19-2<br>"Dataset Arrays" on page 19-2

## Categorical Arrays

| categorical | Arrays for categorical data |
| :--- | :--- |
| nominal | Arrays for nominal categorical data |
| ordinal | Arrays for ordinal categorical data |

## Dataset Arrays

dataset Arrays for statistical data

## Probability Distributions

In this section...<br>"Distribution Objects" on page 19-3<br>"Quasi-Random Numbers" on page 19-3<br>"Piecewise Distributions" on page 19-4

## Distribution Objects

| ProbDist | Object representing probability <br> distribution |
| :--- | :--- |
| ProbDistKernel | Object representing nonparametric <br> probability distribution defined by <br> kernel smoothing |
| ProbDistParametric | Object representing parametric <br> probability distribution |
| ProbDistUnivKernel | Object representing univariate <br> kernel probability distribution |
| ProbDistUnivParam | Object representing univariate <br> parametric probability distribution |

## Quasi-Random Numbers

haltonset
qrandset
qrandstream
sobolset

Halton quasi-random point sets
Quasi-random point sets
Quasi-random number streams
Sobol quasi-random point sets

## Piecewise Distributions

paretotails Empirical distributions with Pareto tails<br>piecewisedistribution<br>Piecewise-defined distributions

## Gaussian Mixture Models

gmdistribution

Gaussian mixture models

## Model Assessment

cvpartition
Data partitions for cross-validation

## Parametric Classification

In this section...<br>"Discriminant Analysis" on page 19-5<br>"Naive Bayes Classification" on page 19-5<br>"Distance Classifiers" on page 19-5

## Discriminant Analysis

| ClassificationDiscriminant | Discriminant analysis classification |
| :--- | :--- |
| ClassificationPartitionedModel | Cross-validated classification model |
| CompactClassificationDiscriminant | Compact discriminant analysis class |

## Naive Bayes Classification

NaiveBayes
Naive Bayes classifier

## Distance Classifiers

| ExhaustiveSearcher | Nearest neighbors search using <br> exhaustive search |
| :--- | :--- |
| KDTreeSearcher | Nearest neighbors search using <br> $k$ d-tree |
| NeighborSearcher | Nearest neighbor search object |

## Supervised Learning

In this section...<br>"Classification Trees" on page 19-6<br>"Classification Ensemble Classes" on page 19-6<br>"Regression Trees" on page 19-7<br>"Regression Ensemble Classes" on page 19-7

## Classification Trees

| ClassificationPartitionedModel | Cross-validated classification model |
| :--- | :--- |
| ClassificationTree | Binary decision tree for classification |
| classregtree | Classification and regression trees |
| CompactClassificationTree | Compact classification tree |

## Classification Ensemble Classes

| ClassificationBaggedEnsemble | Classification ensemble grown by <br> resampling |
| :--- | :--- |
| ClassificationEnsemble | Ensemble classifier |
| ClassificationPartitionedEnsemble | Cross-validated classification <br> ensemble |
| CompactClassificationEnsemble | Compact classification ensemble <br> class |
| CompactTreeBagger | Compact ensemble of decision trees <br> grown by bootstrap aggregation |
| TreeBagger | Bootstrap aggregation for ensemble <br> of decision trees |

## Regression Trees

| classregtree | Classification and regression trees |
| :--- | :--- |
| CompactRegressionTree | Compact regression tree |
| RegressionPartitionedModel | Cross-validated regression model |
| RegressionTree | Regression tree |

## Regression Ensemble Classes

CompactRegressionEnsemble
CompactTreeBagger

RegressionBaggedEnsemble

RegressionEnsemble
RegressionPartitionedEnsemble
TreeBagger

Compact regression ensemble class
Compact ensemble of decision trees grown by bootstrap aggregation

Regression ensemble grown by resampling

Ensemble regression
Cross-validated regression ensemble Bootstrap aggregation for ensemble of decision trees

## Quasi-Random Design of Experiments

haltonset

qrandset
qrandstream
sobolset

Halton quasi-random point sets
Quasi-random point sets
Quasi-random number streams
Sobol quasi-random point sets

Functions - Alphabetical
List

> Purpose Added-variable plot
> Syntax addedvarplot ( $\mathrm{X}, \mathrm{y}$, num, inmodel) addedvarplot( $\mathrm{X}, \mathrm{y}$, num, inmodel, stats)
> addedvarplot ( $\mathrm{X}, \mathrm{y}$, num, inmodel) displays an added variable plot using the predictive terms in X , the response values in y , the added term in column num of $X$, and the model with current terms specified by inmodel. X is an $n$-by- $p$ matrix of $n$ observations of $p$ predictive terms. y is vector of $n$ response values. num is a scalar index specifying the column of X with the term to be added. inmodel is a logical vector of $p$ elements specifying the columns of $X$ in the current model. By default, all elements of inmodel are false.

Note addedvarplot automatically includes a constant term in all models. Do not enter a column of 1 s directly into $X$.
addedvarplot ( $\mathrm{X}, \mathrm{y}$, num, inmodel, stats) uses the stats output from the stepwisefit function to improve the efficiency of repeated calls to addedvarplot. Otherwise, this syntax is equivalent to the previous syntax.

Added variable plots are used to determine the unique effect of adding a new term to a multilinear model. The plot shows the relationship between the part of the response unexplained by terms already in the model and the part of the new term unexplained by terms already in the model. The "unexplained" parts are measured by the residuals of the respective regressions. A scatter of the residuals from the two regressions forms the added variable plot.

In addition to the scatter of residuals, the plot produced by addedvarplot shows $95 \%$ confidence intervals on predictions from the fitted line. The fitted line has intercept zero because, under typical linear model assumptions, both of the plotted variables have mean zero. The slope of the fitted line is the coefficient that the new term would have if it were added to the model with terms inmodel.

Added variable plots are sometimes known as partial regression leverage plots.

Examples Load the data in hald.mat, which contains observations of the heat of reaction of various cement mixtures:

| load hald <br> whos <br> Name | Size | Bytes | Class |
| :--- | :--- | ---: | :--- | Attributes

Create an added variable plot to investigate the addition of the third column of ingredients to a model consisting of the first two columns:

```
inmodel = [true true false false];
addedvarplot(ingredients,heat,3,inmodel)
```

Added variable plot for X 3 Adjusted for $\mathrm{X} 1, \mathrm{X} 2$


The wide scatter and the low slope of the fitted line are evidence against the statistical significance of adding the third column to the model.

## Purpose Add levels to categorical array <br> Syntax $\quad$ B $=\operatorname{addlevels}(A$, newlevels $)$

Description $\quad B=\operatorname{addlevels}(A$, newlevels) adds new levels to the categorical array A. newlevels is a cell array of strings or a 2-D character matrix that specifies the levels to add. addlevels adds the new levels at the end of the list of possible categorical levels in A, but does not modify the value of any element. B does not contain elements at the new levels.

## Examples Example 1

Add levels for additional species in Fisher's iris data:

```
load fisheriris
species = nominal(species,...
    {'Species1','Species2','Species3'},...
    {'setosa','versicolor','virginica'});
species = addlevels(species,{'Species4','Species5'});
getlabels(species)
ans =
    'Species1' 'Species2' 'Species3' 'Species4' 'Species5'
```


## Example 2

1 Load patient data from the CSV file hospital. dat and store the information in a dataset array with observation names given by the first column in the data (patient identification):

```
patients = dataset('file','hospital.dat',...
    'delimiter',',',...
    'ReadObsNames',true);
```

2 Make the $\{0,1\}$-valued variable smoke nominal, and change the labels to 'No' and 'Yes':

```
patients.smoke = nominal(patients.smoke,{'No','Yes'});
```

3 Add new levels to smoke as placeholders for more detailed histories of smokers:

```
patients.smoke = addlevels(patients.smoke,...
    {'0-5 Years','5-10 Years','LongTerm'});
```

4 Assuming the nonsmokers have never smoked, relabel the 'No' level:

```
patients.smoke = setlabels(patients.smoke,'Never','No');
```

5 Drop the undifferentiated 'Yes' level from smoke:
patients.smoke = droplevels(patients.smoke,'Yes');
Warning: OLDLEVELS contains categorical levels that were present in A, caused some array elements to have undefined levels.

Note that smokers now have an undefined level.
6 Set each smoker to one of the new levels, by observation name:

```
patients.smoke('YPL-320') = '5-10 Years';
```

See Also
droplevels | getlabels | islevel | mergelevels | reorderlevels

## Purpose Add listener for event

```
Syntax
el = addlistener(hsource,'eventname',callback)
el = addlistener(hsource,property,'eventname',callback)
```

Description
el = addlistener(hsource,'eventname',callback) creates a listener for the event named eventname, the source of which is handle object hsource. If hsource is an array of source handles, the listener responds to the named event on any handle in the array. callback is a function handle that is invoked when the event is triggered.
el = addlistener(hsource,property,'eventname',callback) adds a listener for a property event. eventname must be one of the strings 'PreGet', 'PostGet', 'PreSet', and 'PostSet'. property must be either a property name or cell array of property names, or a meta.property or array of meta.property. The properties must belong to the class of hsource. If hsource is scalar, property can include dynamic properties.

For all forms, addlistener returns an event.listener. To remove a listener, delete the object returned by addlistener. For example, delete(el) calls the handle class delete method to remove the listener and delete it from the workspace.

[^0]
## gmdistribution.AIC property

Purpose Akaike Information Criterion
Description The Akaike Information Criterion: $2 * \mathrm{NlogL}+2 * \mathrm{~m}$, where m is the number of estimated parameters.

Note This property applies only to gmdistribution objects constructed with fit.

## Purpose <br> Syntax <br> Description

Andrews plot
andrewsplot (X)
andrewsplot(X,...,'Standardize',standopt)
andrewsplot(X,...,'Quantile', alpha)
andrewsplot( $\mathrm{X}, \ldots$, , $\mathrm{Group}^{\prime}$, group)
andrewsplot(X,...,'PropName', PropVal,....)
$\mathrm{h}=$ andrewsplot $(\mathrm{X}, \ldots$ )
andrewsplot (X) creates an Andrews plot of the multivariate data in the matrix $X$. The rows of $X$ correspond to observations, the columns to variables. Andrews plots represent each observation by a function $f(t)$ of a continuous dummy variable $t$ over the interval $[0,1] . f(t)$ is defined for the $i$ th observation in X as

$$
f(t)=X(i, 1) / \sqrt{2}+X(i, 2) \sin (2 \pi t)+X(i, 3) \cos (2 \pi t)+\ldots
$$

andrewsplot treats NaN values in X as missing values and ignores the corresponding rows.
andrewsplot (X,...,'Standardize',standopt) creates an Andrews plot where standopt is one of the following:

- 'on' - scales each column of $X$ to have mean 0 and standard deviation 1 before making the plot.
- 'PCA' - creates an Andrews plot from the principal component scores of $X$, in order of decreasing eigenvalue. (See princomp.)
- 'PCAStd ' - creates an Andrews plot using the standardized principal component scores. (See princomp.)
andrewsplot( $\mathrm{X}, \ldots$, '. Quantile', alpha) plots only the median and the alpha and ( $1-$ alpha) quantiles of $f(t)$ at each value of $t$. This is useful if $X$ contains many observations.
andrewsplot (X, .., 'Group ', group) plots the data in different groups with different colors. Groups are defined by group, a numeric array containing a group index for each observation. group can also be a


## andrewsplot

categorical array, character matrix, or cell array of strings containing a group name for each observation. (See "Grouped Data" on page 2-34.)
andrewsplot(X,...,'PropName',PropVal,...) sets optional lineseries object properties to the specified values for all lineseries objects created by andrewsplot. (See Lineseries Properties.)
$\mathrm{h}=$ andrewsplot $(\mathrm{X}, \ldots$ ) returns a column vector of handles to the lineseries objects created by andrewsplot, one handle per row of X. If you use the 'Quantile' input parameter, h contains one handle for each of the three lineseries objects created. If you use both the 'Quantile' and the 'Group' input parameters, $h$ contains three handles for each group.

## Examples Make a grouped plot of the Fisher iris data:

```
load fisheriris
andrewsplot(meas,'group',species)
```



Plot only the median and quartiles of each group:

```
andrewsplot(meas,'group',species,'quantile',.25)
```



| See Also | parallelcoords \| glyphplot |
| :--- | :--- |
| How To | - "Grouped Data" on page 2-34 |

One-way analysis of variance

```
\(p=\operatorname{anova1}(X)\)
\(p=\) anova1 ( \(X\), group)
p = anova1(X,group,displayopt)
[p,table] = anova1(...)
[p,table,stats] = anova1(...)
```

$\mathrm{p}=$ anova1(X) performs balanced one-way ANOVA for comparing the means of two or more columns of data in the matrix $X$, where each column represents an independent sample containing mutually independent observations. The function returns the $p$ value under the null hypothesis that all samples in X are drawn from populations with the same mean.

If $p$ is near zero, it casts doubt on the null hypothesis and suggests that at least one sample mean is significantly different than the other sample means. Common significance levels are 0.05 or 0.01 .

The anova1 function displays two figures, the standard ANOVA table and a box plot of the columns of $X$.

The standard ANOVA table divides the variability of the data into two parts:

- Variability due to the differences among the column means (variability between groups)
- Variability due to the differences between the data in each column and the column mean (variability within groups)

The standard ANOVA table has six columns:
1 The source of the variability.
2 The sum of squares (SS) due to each source.
3 The degrees of freedom (df) associated with each source.

4 The mean squares (MS) for each source, which is the ratio SS/df.
5 The $F$-statistic, which is the ratio of the mean squares.
6 The $p$ value, which is derived from the cdf of $F$.
The box plot of the columns of X suggests the size of the $F$-statistic and the $p$ value. Large differences in the center lines of the boxes correspond to large values of $F$ and correspondingly small values of $p$.
anova1 treats NaN values as missing, and disregards them.
$p$ = anova1(X,group) performs ANOVA by group. For more information on grouping variables, see "Grouped Data" on page 2-34.

If $X$ is a matrix, anova1 treats each column as a separate group, and evaluates whether the population means of the columns are equal. This form of anova1 is appropriate when each group has the same number of elements (balanced ANOVA). group can be a character array or a cell array of strings, with one row per column of $X$, containing group names. Enter an empty array ([]) or omit this argument if you do not want to specify group names.

If $X$ is a vector, group must be a categorical variable, vector, string array, or cell array of strings with one name for each element of $X$. $X$ values corresponding to the same value of group are placed in the same group. This form of anova1 is appropriate when groups have different numbers of elements (unbalanced ANOVA).

If group contains empty or NaN -valued cells or strings, the corresponding observations in X are disregarded.
$\mathrm{p}=$ anova1(X, group, displayopt) enables the ANOVA table and box plot displays when displayopt is 'on' (default) and suppresses the displays when displayopt is 'off'. Notches in the boxplot provide a test of group medians (see boxplot) different from the $F$ test for means in the ANOVA table.
[ $p$, table] = anova1(...) returns the ANOVA table (including column and row labels) in the cell array table. Copy a text version of
the ANOVA table to the clipboard using the Copy Text item on the Edit menu.
[p,table,stats] $=$ anova1(...) returns a structure stats used to perform a follow-up multiple comparison test. anova1 evaluates the hypothesis that the samples all have the same mean against the alternative that the means are not all the same. Sometimes it is preferable to perform a test to determine which pairs of means are significantly different, and which are not. Use the multcompare function to perform such tests by supplying the stats structure as input.

## Assumptions

The ANOVA test makes the following assumptions about the data in X :

- All sample populations are normally distributed.
- All sample populations have equal variance.
- All observations are mutually independent.

The ANOVA test is known to be robust with respect to modest violations of the first two assumptions.

## Examples Example 1

Create $X$ with columns that are constants plus random normal disturbances with mean zero and standard deviation one:

```
X = meshgrid(1:5)
X =
    1
    1 2 3 4 5
    1
    1
    1 2 3 4 5
X = X + normrnd(0,1,5,5)
X =
    1.3550 2.0662 2.4688 5.9447 5.4897
```

| 2.0693 | 1.7611 | 1.4864 | 4.8826 | 6.3222 |
| ---: | ---: | ---: | ---: | ---: |
| 2.1919 | 0.7276 | 3.1905 | 4.8768 | 4.6841 |
| 2.7620 | 1.8179 | 3.9506 | 4.4678 | 4.9291 |
| -0.3626 | 1.1685 | 3.5742 | 2.1945 | 5.9465 |

Perform one-way ANOVA:



The very small $p$ value indicates that differences between column means are highly significant. The probability of this outcome under the null hypothesis (that samples drawn from the same population would have means differing by the amounts seen in X ) is equal to the $p$ value.

## Example 2

The following example is from a study of the strength of structural beams in Hogg. The vector strength measures deflections of beams in thousandths of an inch under 3,000 pounds of force. The vector alloy identifies each beam as steel ('st'), alloy 1 ('al1'), or alloy 2 ('al2'). (Although alloy is sorted in this example, grouping variables do not need to be sorted.) The null hypothesis is that steel beams are equal in strength to beams made of the two more expensive alloys.

```
strength = [lllllll
```

```
    78 75 76 77 79 79 77 78 82 79];
alloy = {'st','st','st','st','st','st','st','st',...
    al1','al1','al1','al1','al1','al1',...
    'al2','al2','al2','al2','al2','al2'};
p = anova1(strength,alloy)
p =
    1.5264e-004
```

ANOVA Table

| Source | SS | df | MS | F | Prob $>$ F | 4 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Columns | 184.8 | 2 | 92.4 | 15.4 | 0.0002 |  |
| Error | 102 | 17 | 6 |  |  |  |
| Total | 286.8 | 19 |  |  |  | $\checkmark$ |



The $p$ value suggests rejection of the null hypothesis. The box plot shows that steel beams deflect more than beams made of the more expensive alloys.

References

How To

## See Also

[1] Hogg, R. V., and J. Ledolter. Engineering Statistics. New York: MacMillan, 1987.
anova2 | anovan | boxplot | manova1 | multcompare

- "Grouped Data" on page 2-34

Purpose Two-way analysis of variance

Syntax $\quad$| $p=\operatorname{anova2}(X$, reps $)$ |  |
| :--- | :--- |
|  | $p=\operatorname{anova2}(X$, reps, displayopt $)$ |
|  | $[p$, table $]=$ anova2 $(\ldots)$ |
|  | $[p$, table, stats $]=\operatorname{anova} 2(\ldots)$ |

## Description

p = anova2(X,reps) performs a balanced two-way ANOVA for comparing the means of two or more columns and two or more rows of the observations in X. The data in different columns represent changes in factor $A$. The data in different rows represent changes in factor $B$. If there is more than one observation for each combination of factors, input reps indicates the number of replicates in each position, which must be constant. (For unbalanced designs, use anovan.)

The matrix below shows the format for a set-up where column factor A has two levels, row factor B has three levels, and there are two replications (reps $=2$ ). The subscripts indicate row, column, and replicate, respectively.

| $A=1 \quad A=2$ |  |
| :---: | :---: |
| $\left.x_{111} \quad x_{121}\right]$ | $B=1$ |
| $x_{112} \quad x_{122}$ |  |
| $x_{211} \quad x_{221}$ | $B=2$ |
| $x_{212} \quad x_{222}$ |  |
| $x_{311} \quad x_{321}$ |  |
| $x_{312} \quad x_{322}$ | $B=3$ |

When reps is 1 (default), anova2 returns two $p$-values in vector p :
1 The $p$ value for the null hypothesis, $\mathrm{H}_{0 \mathrm{~A}}$, that all samples from factor A (i.e., all column-samples in X ) are drawn from the same population

2 The $p$ value for the null hypothesis, $\mathrm{H}_{0 \mathrm{~B}}$, that all samples from factor B (i.e., all row-samples in X ) are drawn from the same population

When reps is greater than 1 , anova2 returns a third $p$ value in vector p :

3 The $p$ value for the null hypothesis, $\mathrm{H}_{0 \mathrm{AB}}$, that the effects due to factors A and B are additive (i.e., that there is no interaction between factors A and B)

If any $p$ value is near zero, this casts doubt on the associated null hypothesis. A sufficiently small $p$ value for $\mathrm{H}_{0 \mathrm{~A}}$ suggests that at least one column-sample mean is significantly different that the other column-sample means; i.e., there is a main effect due to factor A. A sufficiently small $p$ value for $\mathrm{H}_{0 \mathrm{~B}}$ suggests that at least one row-sample mean is significantly different than the other row-sample means; i.e., there is a main effect due to factor B. A sufficiently small $p$ value for $H_{0 A B}$ suggests that there is an interaction between factors $A$ and $B$. The choice of a limit for the $p$ value to determine whether a result is "statistically significant" is left to the researcher. It is common to declare a result significant if the $p$ value is less than 0.05 or 0.01 .
anova2 also displays a figure showing the standard ANOVA table, which divides the variability of the data in $X$ into three or four parts depending on the value of reps:

- The variability due to the differences among the column means
- The variability due to the differences among the row means
- The variability due to the interaction between rows and columns (if reps is greater than its default value of one)
- The remaining variability not explained by any systematic source

The ANOVA table has five columns:

- The first shows the source of the variability.
- The second shows the Sum of Squares (SS) due to each source.
- The third shows the degrees of freedom (df) associated with each source.
- The fourth shows the Mean Squares (MS), which is the ratio SS/df.
- The fifth shows the $F$ statistics, which is the ratio of the mean squares.
p = anova2(X,reps,displayopt) enables the ANOVA table display when displayopt is 'on' (default) and suppresses the display when displayopt is 'off'.
[ $p$, table] = anova2(...) returns the ANOVA table (including column and row labels) in cell array table. (Copy a text version of the ANOVA table to the clipboard by using the Copy Text item on the Edit menu.)
[p,table,stats] = anova2(...) returns a stats structure that you can use to perform a follow-up multiple comparison test.

The anova2 test evaluates the hypothesis that the row, column, and interaction effects are all the same, against the alternative that they are not all the same. Sometimes it is preferable to perform a test to determine which pairs of effects are significantly different, and which are not. Use the multcompare function to perform such tests by supplying the stats structure as input.

## Examples

The data below come from a study of popcorn brands and popper type (Hogg 1987). The columns of the matrix popcorn are brands (Gourmet, National, and Generic). The rows are popper type (Oil and Air.) The study popped a batch of each brand three times with each popper. The values are the yield in cups of popped popcorn.
load popcorn

| popcorn <br> popcorn |  |  |
| :--- | :--- | :--- |
| 5.5000 | 4.5000 | 3.5000 |
| 5.5000 | 4.5000 | 4.0000 |
| 6.0000 | 4.0000 | 3.0000 |
| 6.5000 | 5.0000 | 4.0000 |
| 7.0000 | 5.5000 | 5.0000 |

```
        7.0000 5.0000 4.5000
    p = anova2(popcorn,3)
p =
    0.0000 0.0001 0.7462
```

| ANOVA Table |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Source | 55 | df | MS | F | Prob $>\mathrm{F}$ | - |
| Columns | 15.75 | 2 | 7.875 | 56.7 | 0 |  |
| Rows | 4.5 | 1 | 4.5 | 32.4 | 0.0001 |  |
| Interaction | 0.0833 | 2 | 0.04167 | 0.3 | 0.7462 |  |
| Error | 1.6667 | 12 | 0.13889 |  |  |  |
| Total | 22 | 17 |  |  |  |  |

The vector p shows the $p$-values for the three brands of popcorn, 0.0000 , the two popper types, 0.0001 , and the interaction between brand and popper type, 0.7462 . These values indicate that both popcorn brand and popper type affect the yield of popcorn, but there is no evidence of a synergistic (interaction) effect of the two.

The conclusion is that you can get the greatest yield using the Gourmet brand and an Air popper (the three values popcorn $(4: 6,1)$ ).

## References

See Also
[1] Hogg, R. V., and J. Ledolter. Engineering Statistics. New York: MacMillan, 1987.

```
anova1 | anovan
```

```
Purpose \(N\)-way analysis of variance
Syntax
p = anovan(y,group)
p = anovan(y,group,param,val)
[p,table] = anovan(y,group,param,val)
[p,table,stats] = anovan(y,group,param,val)
[p,table,stats,terms] = anovan(y,group, param, val)
```


## Description

$p$ = anovan(y, group) performs multiway (n-way) analysis of variance (ANOVA) for testing the effects of multiple factors on the mean of the vector y. (See "Grouped Data".) This test compares the variance explained by factors to the left over variance that cannot be explained. The factors and factor levels of the observations in $y$ are assigned by the cell array group. Each of the cells in the cell array group contains a list of factor levels identifying the observations in $y$ with respect to one of the factors. The list within each cell can be a categorical array, numeric vector, character matrix, or single-column cell array of strings, and must have the same number of elements as $y$. The fitted ANOVA model includes the main effects of each grouping variable. All grouping variables are treated as fixed effects by default. The result $p$ is a vector of $p$-values, one per term. For an example, see "Example of Three-Way ANOVA" on page 20-28.
$p$ = anovan(y, group, param, val) specifies one or more of the parameter name/value pairs described in the following table.

| Parameter | Value |
| :--- | :--- |
| 'alpha' | A number between 0 and 1 requesting 100(1- <br> alpha)\% confidence bounds (default 0.05 for 95\% <br> confidence) |
| 'continuous ' | A vector of indices indicating which grouping <br> variables should be treated as continuous predictors <br> rather than as categorical predictors. |
| 'display' | 'on' displays an ANOVA table (the default) <br> 'off' omits the display |


| Parameter | Value |
| :--- | :--- |
| 'model' | The type of model used. See "Model Type" on page <br> $20-26$ for a description of this parameter. |
| 'nested ' | A matrix M of 0's and 1's specifying the nesting <br> relationships among the grouping variables. M(i,j) is <br> 1 if variable i is nested in variable j. |
| 'random' | A vector of indices indicating which grouping <br> variables are random effects (all are fixed by default). <br> See "ANOVA with Random Effects" on page 8-19 for <br> an example of how to use 'random '. |
| 'sstype' | 1, 2, 3 (default), or h specifies the type of sum of <br> squares. See "Sum of Squares" on page 20-27 for a <br> description of this parameter. |
| 'varnames ' | A character matrix or a cell array of strings specifying <br> names of grouping variables, one per grouping <br> variable. When you do not specify 'varnames ', the <br> default labels 'X1' 'X2 ', 'X3', ..., 'XN' are used. <br> See "ANOVA with Random Effects" on page 8-19 for <br> an example of how to use 'varnames '. |

[p,table] = anovan(y,group, param, val) returns the ANOVA table (including factor labels) in cell array table. (Copy a text version of the ANOVA table to the clipboard by using the Copy Text item on the Edit menu.)
[p,table,stats] = anovan(y,group,param,val) returns a stats structure that you can use to perform a follow-up multiple comparison test with the multcompare function. See "The Stats Structure" on page 20-31The Stats Structure for more information.
[p,table,stats,terms] = anovan(y,group, param, val) returns the main and interaction terms used in the ANOVA computations. The terms are encoded in the output matrix terms using the same format described above for input 'model'. When you specify 'model' itself in this matrix format, the matrix returned in terms is identical.

## Model Type

This section explains how to use the argument 'model' with the syntax:

```
[...] = anovan(y,group,'model',modeltype)
```

The argument model type, which specifies the type of model the function uses, can be any one of the following:

- 'linear' - The default 'linear' model computes only the $p$-values for the null hypotheses on the $N$ main effects.
- 'interaction' - The 'interaction' model computes the $p$-values
for null hypotheses on the $N$ main effects and the $\binom{N}{2}$ two-factor
interactions.
- 'full' - The 'full' model computes the $p$-values for null hypotheses on the $N$ main effects and interactions at all levels.
- An integer - For an integer value of modeltype, $k(k \leq N)$, anovan computes all interaction levels through the kth level. For example, the value 3 means main effects plus two- and three-factor interactions. The values $\mathrm{k}=1$ and $\mathrm{k}=2$ are equivalent to the 'linear' and 'interaction' specifications, respectively, while the value $\mathrm{k}=\mathrm{N}$ is equivalent to the 'full' specification.
- A matrix of term definitions having the same form as the input to the x2fx function. All entries must be 0 or 1 (no higher powers).

For more precise control over the main and interaction terms that anovan computes, modeltype can specify a matrix containing one row for each main or interaction term to include in the ANOVA model. Each row defines one term using a vector of $N$ zeros and ones. The table below illustrates the coding for a 3 -factor ANOVA.

| Matrix Row | ANOVA Term |
| :--- | :--- |
| $\left[\begin{array}{lll}1 & 0 & 0\end{array}\right]$ | Main term A |
| $\left[\begin{array}{lll}0 & 1 & 0\end{array}\right]$ | Main term B |


| Matrix Row | ANOVA Term |
| :--- | :--- |
| $\left[\begin{array}{lll}0 & 0 & 1\end{array}\right]$ | Main term C |
| $\left[\begin{array}{lll}1 & 1 & 0\end{array}\right]$ | Interaction term AB |
| $\left[\begin{array}{lll}1 & 0 & 1\end{array}\right]$ | Interaction term AC |
| $\left[\begin{array}{lll}0 & 1 & 1\end{array}\right]$ | Interaction term BC |
| $\left[\begin{array}{lll}1 & 1 & 1\end{array}\right]$ | Interaction term ABC |

For example, if modeltype is the matrix $\left[\begin{array}{lllllll}0 & 1 & 0 ; 0 & 0 & 1 ; 0 & 1 & 1\end{array}\right]$, the output vector p contains the $p$-values for the null hypotheses on the main effects B and C and the interaction effect BC , in that order. A simple way to generate the modeltype matrix is to modify the terms output, which codes the terms in the current model using the format described above. If anovan returns [ 0 1 $0 ; 0$ 0 1; 0 111] for terms, for example, and there is no significant result for interaction BC , you can recompute the ANOVA on just the main effects B and C by specifying [0 1 0;0 0 1] for modeltype.

## Sum of Squares

This section explains how to use the argument 'sstype' with the syntax:
[...] = anovan(y,group,'sstype',type)
This syntax computes the ANOVA using the type of sum of squares specified by type, which can be $1,2,3$, or h . While the numbers $1-3$ designate Type 1 , Type 2, or Type 3 sum of squares, respectively, $h$ represents a hierarchical model similar to type 2, but with continuous as well as categorical factors used to determine the hierarchy of terms. The default value is 3 . For a model containing main effects but no interactions, the value of type only influences computations on unbalanced data.

The sum of squares for any term is determined by comparing two models. The Type 1 sum of squares for a term is the reduction in residual sum of squares obtained by adding that term to a fit that already includes the terms listed before it. The Type 2 sum of squares is
the reduction in residual sum of squares obtained by adding that term to a model consisting of all other terms that do not contain the term in question. The Type 3 sum of squares is the reduction in residual sum of squares obtained by adding that term to a model containing all other terms, but with their effects constrained to obey the usual "sigma restrictions" that make models estimable.

Suppose you are fitting a model with two factors and their interaction, and that the terms appear in the order $A, B, A B$. Let $R(\cdot)$ represent the residual sum of squares for a model, so for example $R(A, B, A B)$ is the residual sum of squares fitting the whole model, $R(A)$ is the residual sum of squares fitting just the main effect of $A$, and $R(1)$ is the residual sum of squares fitting just the mean. The three types of sums of squares are as follows:

| Term | Type 1 Sum of Squares | Type 2 Sum of <br> Squares | Type 3 Sum of <br> Squares |
| :--- | :--- | :--- | :--- |
| $A$ | $R(1)-R(A)$ | $R(B)-R(A, B)$ | $R(B, A B)-R(A, B, A B)$ |
| $B$ | $R(A)-R(A, B)$ | $R(A)-R(A, B)$ | $R(A, A B)-R(A, B, A B)$ |
| $A B$ | $R(A, B)-R(A, B, A B)$ | $R(A, B)-R(A, B, A B)$ | $R(A, B)-R(A, B, A B)$ |

The models for Type 3 sum of squares have sigma restrictions imposed. This means, for example, that in fitting $R(B, A B)$, the array of $A B$ effects is constrained to sum to 0 over $A$ for each value of $B$, and over $B$ for each value of $A$.

## Example of Three-Way ANOVA

As an example of three-way ANOVA, consider the vector y and group inputs below.

```
y = [52.7 57.5 45.9 44.5 53.0 57.0 45.9 44.0]';
g1 = [1 2 1 2 1 2 1 2];
g2 = {'hi';'hi';'lo';'lo';'hi';'hi';'lo';'lo'};
g3 = {'may';'may';'may';'may';'june';'june';'june';'june'};
```

This defines a three-way ANOVA with two levels of each factor. Every observation in $y$ is identified by a combination of factor levels. If the factors are A, B, and C, then observation y (1) is associated with

- Level 1 of factor A
- Level 'hi' of factor B
- Level 'may' of factor C

Similarly, observation y (6) is associated with

- Level 2 of factor A
- Level 'hi' of factor B
- Level 'june' of factor C

To compute the ANOVA, enter

```
p = anovan(y,{g1 g2 g3})
p =
    0.4174
    0.0028
    0.9140
```

Output vector p contains $p$-values for the null hypotheses on the N main effects. Element $p(1)$ contains the $p$ value for the null hypotheses, $\mathrm{H}_{0 \mathrm{~A}}$, that samples at all levels of factor A are drawn from the same population; element $\mathrm{p}(2)$ contains the $p$ value for the null hypotheses, $\mathrm{H}_{0 \mathrm{~B}}$, that samples at all levels of factor B are drawn from the same population; and so on.

If any $p$ value is near zero, this casts doubt on the associated null hypothesis. For example, a sufficiently small $p$ value for $\mathrm{H}_{0 \mathrm{~A}}$ suggests that at least one A-sample mean is significantly different from the other A-sample means; that is, there is a main effect due to factor A. You need to choose a bound for the $p$ value to determine whether a result is statistically significant. It is common to declare a result significant if the $p$ value is less than 0.05 or 0.01 .
anovan also displays a figure showing the standard ANOVA table, which by default divides the variability of the data in $x$ into

- The variability due to differences between the levels of each factor accounted for in the model (one row for each factor)
- The remaining variability not explained by any systematic source

The ANOVA table has six columns:

- The first shows the source of the variability.
- The second shows the sum of squares (SS) due to each source.
- The third shows the degrees of freedom (df) associated with each source.
- The fourth shows the mean squares (MS), which is the ratio $\mathrm{SS} / \mathrm{df}$.
- The fifth shows the $F$ statistics, which are the ratios of the mean squares.
- The sixth shows the $p$-values for the $F$ statistics.

The table is shown in the following figure:

| Analysis of Variance |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Source | Sum Sq. | d.f. | Mean Sq | F | Prob>F | - |
| X 1 | 3.781 | 1 | 3.781 | 0.82 | 0.4174 |  |
| X 2 | 199.001 | 1 | 199.001 | 42.95 | 0.0028 |  |
| X3 | 0.061 | 1 | 0.061 | 0.01 | 0.914 |  |
| Error | 18.535 | 4 | 4.634 |  |  |  |
| Total | 221.379 | 7 |  |  |  |  |

Constrained (Type III) sums of squares.

## Two-Factor Interactions

By default, anovan computes $p$-values just for the three main effects. To also compute $p$-values for the two-factor interactions, X 1 * $\mathrm{X} 2, \mathrm{X} 1 * \mathrm{X} 3$,
and X2*X3, add the name/value pair 'model', 'interaction' as input arguments.

```
p = anovan(y,{g1 g2 g3},'model','interaction')
p =
    0.0347
    0.0048
    0.2578
    0.0158
    0.1444
    0.5000
```

The first three entries of p are the $p$-values for the main effects. The last three entries are the $p$-values for the two-factor interactions. You can determine the order in which the two-factor interactions occur from the ANOVAN table shown in the following figure.

Analysis of Variance

| Source | Sum Sq. | d.f. | Mean Sq. | F | Prob $>F$ | $\Delta$ |
| :--- | ---: | :--- | ---: | ---: | ---: | ---: |
| X1 | 3.781 | 1 | 3.781 | 336.11 | 0.0347 |  |
| X2 | 199.001 | 1 | 199.001 | 17689 | 0.0048 |  |
| X3 | 0.061 | 1 | 0.061 | 5.44 | 0.2578 |  |
| X1*X2 | 18.301 | 1 | 18.301 | 1626.78 | 0.0158 |  |
| X1*X3 | 0.211 | 1 | 0.211 | 18.78 | 0.1444 |  |
| X2*X3 | 0.011 | 1 | 0.011 | 1 | 0.5 |  |
| Error | 0.011 | 1 | 0.011 |  |  |  |
| Total | 221.379 | 7 |  |  |  |  |

Constrained (Type III) sums of squares.

## The Stats Structure

The anovan test evaluates the hypothesis that the different levels of a factor (or more generally, a term) have the same effect, against the alternative that they do not all have the same effect. Sometimes it is preferable to perform a test to determine which pairs of levels are significantly different, and which are not. Use the multcompare function to perform such tests by supplying the stats structure as input.

The stats structure contains the fields listed below, in addition to a number of other fields required for doing multiple comparisons using the multcompare function:

| Field | Description |
| :--- | :--- |
| coeffs | Estimated coefficients |
| coeffnames | Name of term for each coefficient |
| vars | Matrix of grouping variable values for each term |
| resid | Residuals from the fitted model |

The stats structure also contains the following fields if there are random effects:

| Field | Description |
| :--- | :--- |
| ems | Expected mean squares |
| denom | Denominator definition |
| rtnames | Names of random terms |
| varest | Variance component estimates (one per random term) |
| varci | Confidence intervals for variance components |

"Example: Two-Way ANOVA" on page $8-10$ shows how to use anova2 to analyze the effects of two factors on a response in a balanced design. For a design that is not balanced, use anovan instead.

The data in carbig.mat gives measurements on 406 cars. Use anonvan to study how the mileage depends on where and when the cars were made:

```
load carbig
p = anovan(MPG,{org when},'model',2,'sstype',3,...
    'varnames',{'Origin';'Mfg date'})
p =
```

```
    0
    0
0.3059
```

| Analysis of Variance |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Source | Sum Sq. d.f. | Mean Sq | F | Prob>F $\triangle$ |
| Origin | 5727.22 | 2863.58 | 115.09 | 0 |
| Mfg date | 4710.32 | 2355.15 | 94.65 | 0 |
| Origin*Mfg date | 120.54 | 30.12 | 1.21 | 0.3059 |
| Error | 9679.1389 | 24.88 |  |  |
| Total | 24252.6397 |  |  |  |

Constrained (Type III) sums of squares.
The $p$ value for the interaction term is not small, indicating little evidence that the effect of the year or manufacture (when) depends on where the car was made (org). The linear effects of those two factors, however, are significant.

## References

See Also
How To
[1] Hogg, R. V., and J. Ledolter. Engineering Statistics. New York: MacMillan, 1987.
See Also anova1 | anova2 | multcompare

- "Grouped Data" on page 2-34


## ansaribradley

## Purpose Ansari-Bradley test

Syntax $\quad h=\operatorname{ansaribradley}(x, y)$
h = ansaribradley ( $x, y$, alpha)
h = ansaribradley( $\mathrm{x}, \mathrm{y}$, alpha,tail)
[h,p] = ansaribradley(...)
[h,p,stats] = ansaribradley(...)
[...] = ansaribradley(x,y,alpha,tail,exact)
$[. .]=$. ansaribradley (x,y,alpha,tail,exact,dim)

## Description

$\mathrm{h}=$ ansaribradley ( $\mathrm{x}, \mathrm{y}$ ) performs an Ansari-Bradley test of the hypothesis that two independent samples, in the vectors $x$ and $y$, come from the same distribution, against the alternative that they come from distributions that have the same median and shape but different dispersions (e.g. variances). The result is $\mathrm{h}=0$ if the null hypothesis of identical distributions cannot be rejected at the $5 \%$ significance level, or $h=1$ if the null hypothesis can be rejected at the $5 \%$ level. $x$ and $y$ can have different lengths.
x and y can also be matrices or N -dimensional arrays. For matrices, ansaribradley performs separate tests along each column, and returns a vector of results. $x$ and $y$ must have the same number of columns. For $N$-dimensional arrays, ansaribradley works along the first nonsingleton dimension. $x$ and $y$ must have the same size along all the remaining dimensions.
$\mathrm{h}=$ ansaribradley ( $\mathrm{x}, \mathrm{y}$, alpha) performs the test at the significance level (100*alpha), where alpha is a scalar.
$\mathrm{h}=$ ansaribradley (x,y,alpha, tail) performs the test against the alternative hypothesis specified by the string tail. tail is one of:

- 'both' - Two-tailed test (dispersion parameters are not equal)
- 'right' - Right-tailed test (dispersion of $X$ is greater than dispersion of Y )
- 'left' - Left-tailed test (dispersion of $X$ is less than dispersion of $Y$ )
[ $\mathrm{h}, \mathrm{p}$ ] = ansaribradley(...) returns the $p$ value, i.e., the probability of observing the given result, or one more extreme, by chance if the null hypothesis is true. Small values of $p$ cast doubt on the validity of the null hypothesis.
[h,p,stats] = ansaribradley(...) returns a structure stats with the following fields:
- W - Value of the test statistic W, which is the sum of the Ansari-Bradley ranks for the X sample
- Wstar - Approximate normal statistic W*
[...] = ansaribradley (x,y,alpha,tail,exact) computes p using an exact calculation of the distribution of $W$ with exact $=$ 'on'. This can be time-consuming for large samples. exact $=$ 'off' computes $p$ using a normal approximation for the distribution of $W^{*}$. The default if exact is empty is to use the exact calculation if $N$, the total number of rows in x and y , is 25 or less, and to use the normal approximation if $N$ $>25$. Pass in [] for alpha and tail to use their default values while specifying a value for exact. Note that $N$ is computed before any NaN values (representing missing data) are removed.
[...] = ansaribradley(x,y,alpha,tail,exact,dim) works along dimension dim of $x$ and $y$.
The Ansari-Bradley test is a nonparametric alternative to the two-sample $F$ test of equal variances. It does not require the assumption that $x$ and $y$ come from normal distributions. The dispersion of a distribution is generally measured by its variance or standard deviation, but the Ansari-Bradley test can be used with samples from distributions that do not have finite variances.

The theory behind the Ansari-Bradley test requires that the groups have equal medians. Under that assumption and if the distributions in each group are continuous and identical, the test does not depend on the distributions in each group. If the groups do not have the same medians, the results may be misleading. Ansari and Bradley recommend subtracting the median in that case, but the distribution of
the resulting test, under the null hypothesis, is no longer independent of the common distribution of $x$ and $y$. If you want to perform the tests with medians subtracted, you should subtract the medians from $x$ and y before calling ansaribradley.

## Examples

Is the dispersion significantly different for two model years?

```
load carsmall
[h,p,stats] = ansaribradley(MPG(Model_Year==82),MPG(Model_Year==76))
h =
            0
p =
            0.8426
stats =
            W: 526.9000
    Wstar: 0.1986
```


## See Also

vartest | vartestn | ttest2

## Purpose <br> Interactive analysis of covariance

## Syntax

## Description

```
aoctool(x,y,group)
aoctool(x,y,group,alpha)
aoctool(x,y,group,alpha, xname, yname,gname)
aoctool(x,y,group,alpha, xname, yname,gname,displayopt)
aoctool(x,y,group, alpha, xname, yname, gname,displayopt,model)
h = aoctool(...)
[h,atab,ctab] = aoctool(...)
[h,atab,ctab,stats] = aoctool(...)
```

aoctool ( $x, y$, group) fits a separate line to the column vectors, $x$ and $y$, for each group defined by the values in the array group. group may be a categorical variable, vector, character array, or cell array of strings. (See "Grouped Data" on page 2-34.) These types of models are known as one-way analysis of covariance (ANOCOVA) models. The output consists of three figures:

- An interactive graph of the data and prediction curves
- An ANOVA table
- A table of parameter estimates

You can use the figures to change models and to test different parts of the model. More information about interactive use of the aoctool function appears in "Analysis of Covariance Tool" on page 8-27.
aoctool( $x, y$, group, alpha) determines the confidence levels of the prediction intervals. The confidence level is $100(1$-alpha) $\%$. The default value of alpha is 0.05 .
aoctool(x,y,group, alpha, xname, yname, gname) specifies the name to use for the $x, y$, and $g$ variables in the graph and tables. If you enter simple variable names for the $x, y$, and $g$ arguments, the aoctool function uses those names. If you enter an expression for one of these arguments, you can specify a name to use in place of that expression by supplying these arguments. For example, if you enter $m(:, 2)$ as the $x$ argument, you might choose to enter 'Col 2' as the xname argument.
aoctool(x,y,group, alpha, xname, yname, gname, displayopt) enables the graph and table displays when displayopt is 'on' (default) and suppresses those displays when displayopt is 'off'.
aoctool(x,y,group, alpha, xname, yname, gname, displayopt, model) specifies the initial model to fit. The value of model can be any of the following:

- 'same mean' - Fit a single mean, ignoring grouping
- 'separate means' - Fit a separate mean to each group
- 'same line' - Fit a single line, ignoring grouping
- 'parallel lines' - Fit a separate line to each group, but constrain the lines to be parallel
- 'separate lines' - Fit a separate line to each group, with no constraints
$h=$ aoctool(...) returns a vector of handles to the line objects in the plot.
[h,atab,ctab] = aoctool(...) returns cell arrays containing the entries in ANOVA table (atab) and the table of coefficient estimates (ctab). (You can copy a text version of either table to the clipboard by using the Copy Text item on the Edit menu.)
[h,atab,ctab,stats] = aoctool(...) returns a stats structure that you can use to perform a follow-up multiple comparison test. The ANOVA table output includes tests of the hypotheses that the slopes or intercepts are all the same, against a general alternative that they are not all the same. Sometimes it is preferable to perform a test to determine which pairs of values are significantly different, and which are not. You can use the multcompare function to perform such tests by supplying the stats structure as input. You can test either the slopes, the intercepts, or population marginal means (the heights of the curves at the mean $x$ value).

Examples

This example illustrates how to fit different models non-interactively. After loading the smaller car data set and fitting a separate-slopes model, you can examine the coefficient estimates.

```
load carsmall
[h,a,c,s] = aoctool(Weight,MPG,Model_Year,0.05,...
                                    '','','','off','separate lines');
c(:,1:2)
ans =
    'Term' 'Estimate'
    'Intercept' [45.97983716833132]
    ' 70' [-8.58050531454973]
    ' 76' [-3.89017396094922]
    ' 82' [12.47067927549897]
    'Slope' [-0.00780212907455]
    ' 70' [ 0.00195840368824]
    ' 76' [ 0.00113831038418]
    ' 82' [-0.00309671407243]
```

Roughly speaking, the lines relating MPG to Weight have an intercept close to 45.98 and a slope close to -0.0078 . Each group's coefficients are offset from these values somewhat. For instance, the intercept for the cars made in 1970 is $45.98-8.58=37.40$.

Next, try a fit using parallel lines. (The ANOVA table shows that the parallel-lines fit is significantly worse than the separate-lines fit.)

```
[h,a,c,s] = aoctool(Weight,MPG,Model_Year,0.05,...
                                    '','','','off','parallel lines');
c(:,1:2)
ans =
    'Term' 'Estimate'
    'Intercept' [43.38984085130596]
    ' 70' [-3.27948192983761]
    ' 76' [-1.35036234809006]
```

```
' 82' [ 4.62984427792768]
'Slope' [-0.00664751826198]
```

Again, there are different intercepts for each group, but this time the slopes are constrained to be the same.

## See Also anova1 | multcompare | polytool

How To . "Grouped Data" on page 2-34

## TreeBagger.append

| Purpose | Append new trees to ensemble |
| :--- | :--- |
| Syntax | $B=$ append ( $B$, other) |
| Description | $B=$ append ( $B$, other $)$ appends the trees from the other ensemble to <br> those in $B$. This method checks for consistency of the $X$ and $Y$ properties <br> of the two ensembles, as well as consistency of their compact objects and <br> out-of-bag indices, before appending the trees. The output ensemble $B$ <br> takes training parameters such as FBoot, Prior Cost, and other from <br> the $B$ input. There is no attempt to check if these training parameters <br> are consistent between the two objects. |

See Also<br>CompactTreeBagger.combine

## ProbDistKernel.BandWidth property

| Purpose | Read-only value specifying bandwidth of kernel smoothing function <br> for ProbDistKernel object |
| :--- | :--- |
| Description $\quad$BandWidth is a read-only property of the ProbDistKernel class. <br> BandWidth is a value specifying the width of the kernel smoothing <br> function used to compute a nonparametric estimate of the probability <br> distribution when creating a ProbDistKernel object. |  |
| Values | For a distribution specified to cover only the positive numbers or only <br> a finite interval, the data are transformed before the kernel density is <br> applied, and the bandwidth is on the scale of the transformed data. |
| Use this information to view and compare the width of the kernel <br> smoothing function used to create distributions. |  |

## See Also ksdensity

## barttest

## Purpose Bartlett's test

```
Syntax ndim = barttest(X,alpha)
[ndim,prob,chisquare] = barttest(X,alpha)
```

Description

## Examples

See Also
ndim = barttest(X, alpha) returns the number of dimensions necessary to explain the nonrandom variation in the data matrix $X$, using the significance probability alpha. The dimension is determined by a series of hypothesis tests. The test for ndim=1 tests the hypothesis that the variances of the data values along each principal component are equal, the test for ndim=2 tests the hypothesis that the variances along the second through last components are equal, and so on.
[ndim, prob, chisquare] = barttest(X, alpha) returns the number of dimensions, the significance values for the hypothesis tests, and the $x^{2}$ values associated with the tests.

```
X = mvnrnd([0 0],[1 0.99; 0.99 1],20);
X(:,3:4) = mvnrnd([0 0],[1 0.99; 0.99 1],20);
X(:,5:6) = mvnrnd([0 0],[1 0.99; 0.99 1],20);
[ndim, prob] = barttest(X,0.05)
ndim =
        3
prob =
            0
            0
            0
        0.5081
        0.6618
```

princomp | pcacov | pcares

## bbdesign

Purpose Box-Behnken design
Syntax
dBB = bbdesign(n)
[dBB,blocks] = bbdesign(n)
[...] = bbdesign(n,param,val)
Description
$d B B=$ bbdesign( $n$ ) generates a Box-Behnken design for $n$ factors. $n$ must be an integer 3 or larger. The output matrix dBB is $m$-by-n, where $m$ is the number of runs in the design. Each row represents one run, with settings for all factors represented in the columns. Factor values are normalized so that the cube points take values between - 1 and 1 .
[dBB,blocks] = bbdesign(n) requests a blocked design. The output blocks is an $m$-by- 1 vector of block numbers for each run. Blocks indicate runs that are to be measured under similar conditions to minimize the effect of inter-block differences on the parameter estimates.
[...] = bbdesign( $n$, param, val) specifies one or more optional parameter/value pairs for the design. The following table lists valid parameter/value pairs.

| Parameter | Description | Values |
| :--- | :--- | :--- |
| 'center' | Number of <br> center points. | Integer. The default depends on <br> n. |
| 'blocksize' | Maximum <br> number of <br> points per block. | Integer. The default is Inf. |

## Examples The following creates a 3 -factor Box-Behnken design:

| dBB | $=\operatorname{bbdesign}(3)$ |  |
| ---: | :--- | ---: | :--- |
| dBB | $=$ |  |
| -1 | -1 | 0 |
| -1 | 1 | 0 |
| 1 | -1 | 0 |

## bbdesign

| 1 | 1 | 0 |
| ---: | ---: | ---: |
| -1 | 0 | -1 |
| -1 | 0 | 1 |
| 1 | 0 | -1 |
| 1 | 0 | 1 |
| 0 | -1 | -1 |
| 0 | -1 | 1 |
| 0 | 1 | -1 |
| 0 | 1 | 1 |
| 0 | 0 | 0 |
| 0 | 0 | 0 |
| 0 | 0 | 0 |

The center point is run 3 times to allow for a more uniform estimate of the prediction variance over the entire design space.
Visualize the design as follows:

```
plot3(dBB(:,1),dBB(:,2),dBB(:,3),'ro',...
            'MarkerFaceColor','b')
X = [ 1 - 1 -1 -1 1 1 -1 -1 -1 1 1 1 -1 -1; ...
    1 1 1 1 -1 1 1 1 1 1 -1 1 1 1 - 1 - 1];
Y = [-1 -1 1 1 -1 - -1 -1 1 1 -1 1 1 -1 1 - 1; ...
    1 -1 1 1 1 1 1 - - 1 1 1 1 1 - -1 1 - 1];
Z = [14 1 1 1 1 - - 1 -1 -1 - 1 - 1 -1 -1 -1; ...
    1 1 1 1 1 - - - - - - - -1 1 1 1 1 1];
line(X,Y,Z,'Color','b')
axis square equal
```


## bbdesign



## Purpose

Beta cumulative distribution function

## Syntax

p = betacdf( $\mathrm{X}, \mathrm{A}, \mathrm{B}$ )
$p=\operatorname{betacdf}(X, A, B)$ returns the beta cdf at each of the values in $X$ using the corresponding parameters in $A$ and $B . X, A$, and $B$ can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The parameters in A and B must all be positive, and the values in $X$ must lie on the interval [0,1].

The beta cdf for a given value $x$ and given pair of parameters $a$ and $b$ is

$$
p=F(x \mid a, b)=\frac{1}{B(a, b)} \int_{0}^{x} t^{a-1}(1-t)^{b-1} d t
$$

where $B(\cdot)$ is the Beta function.

## Examples

```
x = 0.1:0.2:0.9;
a = 2;
b = 2;
p = betacdf(x,a,b)
p =
```



```
a = [1 2 3 3];
p = betacdf(0.5,a,a)
p =
    0.5000 0.5000 0.5000
```

See Also cdf | betapdf | betainv | betastat | betalike | betarnd | betafit
How To . "Beta Distribution" on page B-4

## Purpose Beta parameter estimates

Syntax phat = betafit(data)
[phat,pci] = betafit(data,alpha)

Description

Examples
phat $=$ betafit(data) computes the maximum likelihood estimates of the beta distribution parameters $a$ and $b$ from the data in the vector data and returns a column vector containing the $a$ and $b$ estimates, where the beta cdf is given by

$$
F(x \mid a, b)=\frac{1}{B(a, b)} \int_{0}^{x} t^{a-1}(1-t)^{b-1} d t
$$

and $B(\cdot)$ is the Beta function. The elements of data must lie in the open interval $(0,1)$, where the beta distribution is defined. However, it is sometimes also necessary to fit a beta distribution to data that include exact zeros or ones. For such data, the beta likelihood function is unbounded, and standard maximum likelihood estimation is not possible. In that case, betafit maximizes a modified likelihood that incorporates the zeros or ones by treating them as if they were values that have been left-censored at sqrt (realmin) or right-censored at 1 -eps/2, respectively.
[phat,pci] = betafit(data,alpha) returns confidence intervals on the $a$ and $b$ parameters in the 2 -by- 2 matrix pci. The first column of the matrix contains the lower and upper confidence bounds for parameter $a$, and the second column contains the confidence bounds for parameter $b$. The optional input argument alpha is a value in the range $[0,1]$ specifying the width of the confidence intervals. By default, alpha is 0.05 , which corresponds to $95 \%$ confidence intervals. The confidence intervals are based on a normal approximation for the distribution of the logs of the parameter estimates.

This example generates 100 beta distributed observations. The true $a$ and $b$ parameters are 4 and 3 , respectively. Compare these to the
values returned in $p$ by the beta fit. Note that the columns of ci both bracket the true parameters.

```
data = betarnd(4,3,100,1);
[p,ci] = betafit(data,0.01)
p =
    5.5328 3.8097
ci =
    3.6538 2.6197
    8.3781 5.5402
```


## References

See Also
How To . "Beta Distribution" on page B-4

Purpose Beta inverse cumulative distribution function

## Syntax $\quad X=\operatorname{betainv}(P, A, B)$

Description

Algorithms

## Examples

$X=$ betainv( $P, A, B)$ computes the inverse of the beta cdf with parameters specified by $A$ and $B$ for the corresponding probabilities in $P$. $P, A$, and $B$ can be vectors, matrices, or multidimensional arrays that are all the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The parameters in $A$ and $B$ must all be positive, and the values in $P$ must lie on the interval $[0,1]$.

The inverse beta cdf for a given probability $p$ and a given pair of parameters $a$ and $b$ is

$$
x=F^{-1}(p \mid a, b)=\{x: F(x \mid a, b)=p\}
$$

where

$$
p=F(x \mid a, b)=\frac{1}{B(a, b)} \int_{0}^{x} t^{a-1}(1-t)^{b-1} d t
$$

and $B(\cdot)$ is the Beta function. Each element of output X is the value whose cumulative probability under the beta cdf defined by the corresponding parameters in $A$ and $B$ is specified by the corresponding value in $P$.

The betainv function uses Newton's method with modifications to constrain steps to the allowable range for $x$, i.e., [ $\left.\begin{array}{ll}1 & 1\end{array}\right]$.

$$
\begin{aligned}
& p=\left[\begin{array}{lll}
0.01 & 0.5 & 0.99
\end{array}\right] ; \\
& x=\operatorname{betainv}(p, 10,5) \\
& x= \\
& \\
& 0.3726 \quad 0.6742 \quad 0.8981
\end{aligned}
$$

According to this result, for a beta cdf with $a=10$ and $b=5$, a value less than or equal to 0.3726 occurs with probability 0.01 . Similarly,
values less than or equal to 0.6742 and 0.8981 occur with respective probabilities 0.5 and 0.99 .

See Also<br>icdf | betapdf | betafit | betainv | betastat | betalike | betarnd | betacdf<br>How To . "Beta Distribution" on page B-4

Purpose Beta negative log-likelihood

| Syntax | nlogL $=$ betalike $($ params, data $)$ |
| :--- | :--- |
| $[n \log L, A V A R]$ |  |$=$ betalike $($ params, data $)$.

Description nlogL = betalike(params, data) returns the negative of the beta log-likelihood function for the beta parameters $a$ and $b$ specified in vector params and the observations specified in the column vector data.
The elements of data must lie in the open interval $(0,1)$, where the beta distribution is defined. However, it is sometimes also necessary to fit a beta distribution to data that include exact zeros or ones. For such data, the beta likelihood function is unbounded, and standard maximum likelihood estimation is not possible. In that case, betalike computes a modified likelihood that incorporates the zeros or ones by treating them as if they were values that have been left-censored at sqrt (realmin) or right-censored at $1-\mathrm{eps} / 2$, respectively.
[nlogL,AVAR] = betalike(params,data) also returns AVAR, which is the asymptotic variance-covariance matrix of the parameter estimates if the values in params are the maximum likelihood estimates. AVAR is the inverse of Fisher's information matrix. The diagonal elements of AVAR are the asymptotic variances of their respective parameters.
betalike is a utility function for maximum likelihood estimation of the beta distribution. The likelihood assumes that all the elements in the data sample are mutually independent. Since betalike returns the negative beta log-likelihood function, minimizing betalike using fminsearch is the same as maximizing the likelihood.

Examples This example continues the betafit example, which calculates estimates of the beta parameters for some randomly generated beta distributed data.

```
r = betarnd(4,3,100,1);
[nlogl,AVAR] = betalike(betafit(r),r)
nlogl =
```

```
    -27.5996
AVAR =
    0.2783 0.1316
    0.1316 0.0867
```


## See Also

betapdf | betafit | betainv | betastat | betarnd | betacdf
How To - "Beta Distribution" on page B-4

Purpose Beta probability density function
Syntax $\quad Y=\operatorname{betapdf}(X, A, B)$

Description

## Examples

## See Also

$Y=$ betapdf( $X, A, B$ ) computes the beta pdf at each of the values in $X$ using the corresponding parameters in $A$ and $B . X, A$, and $B$ can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions of the other inputs. The parameters in A and B must all be positive, and the values in X must lie on the interval [ 0,1$]$.

The beta probability density function for a given value $x$ and given pair of parameters $a$ and $b$ is

$$
y=f(x \mid a, b)=\frac{1}{B(a, b)} x^{a-1}(1-x)^{b-1} I_{(0,1)}(x)
$$

where $B(\cdot)$ is the Beta function. The indicator function $I_{(0,1)}(x)$ ensures that only values of $x$ in the range ( 01 ) have nonzero probability. The uniform distribution on (01) is a degenerate case of the beta pdf where $a=1$ and $b=1$.

A likelihood function is the pdf viewed as a function of the parameters. Maximum likelihood estimators (MLEs) are the values of the parameters that maximize the likelihood function for a fixed value of $x$.

```
a = [0.5 1; 2 4]
a =
        0.5000 1.0000
        2.0000 4.0000
y = betapdf(0.5,a,a)
y =
        0.6366 1.0000
        1.5000 2.1875
```

pdf | betafit | betainv | betastat | betalike | betarnd | betacdf

How To

- "Beta Distribution" on page B-4


## betarnd

## Purpose Beta random numbers

Syntax $\quad$| $R$ | $=\operatorname{betarnd}(A, B)$ |
| ---: | :--- |
| $R$ | $=\operatorname{betarnd}(A, B, m, n, \ldots)$ |
| $R$ | $=\operatorname{betarnd}(A, B,[m, n, \ldots])$ |

Description

## Examples

$r=$

$$
\begin{array}{lllll}
0.5974 & 0.4777 & 0.5538 & 0.5465 & 0.6327
\end{array}
$$

$r=\operatorname{betarnd}(4,2,2,3)$
$r=$

| 0.3943 | 0.6101 | 0.5768 |
| :--- | :--- | :--- |
| 0.5990 | 0.2760 | 0.5474 |

See Also
random | betapdf | betafit | betainv | betastat | betalike | betacdf

How To . "Beta Distribution" on page B-4

## betastat

Purpose Beta mean and variance

## Syntax $\quad[M, V]=\operatorname{betastat}(A, B)$

Description $[M, V]=$ betastat $(A, B)$, with $A>0$ and $B>0$, returns the mean of and variance for the beta distribution with parameters specified by $A$ and B. A and B can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of M and V . A scalar input for $A$ or $B$ is expanded to a constant array with the same dimensions as the other input.

The mean of the beta distribution with parameters $a$ and $b$ is $a /(a+b)$ and the variance is

$$
\frac{a b}{(a+b+1)(a+b)^{2}}
$$

## Examples

If parameters $a$ and $b$ are equal, the mean is $1 / 2$.

```
a = 1:6;
[m,v] = betastat(a,a)
m =
    0.5000}00.5000 0.5000 0.5000 0.5000 0.5000
v =
    0.0833 0.0500
```


## See Also

betapdf | betafit | betainv | betalike | betarnd | betacdf
How To

- "Beta Distribution" on page B-4

Purpose Bayes Information Criterion
Description The Bayes Information Criterion: $2 * N \log L+m * \log (n)$, where $n$ is the number of observations and m is the number of estimated parameters.

## binocdf

## Purpose Binomial cumulative distribution function

## Syntax $\quad Y=\operatorname{binocdf}(X, N, P)$

Description

Examples
$Y=$ binocdf( $X, N, P$ ) computes a binomial cdf at each of the values in $X$ using the corresponding number of trials in $N$ and probability of success for each trial in P. X, N, and P can be vectors, matrices, or multidimensional arrays that are all the same size. A scalar input is expanded to a constant array with the same dimensions of the other inputs. The values in $N$ must all be positive integers, the values in $X$ must lie on the interval $[0, N]$, and the values in $P$ must lie on the interval [0, 1].
The binomial cdf for a given value $x$ and a given pair of parameters $n$ and $p$ is

$$
y=F(x \mid n, p)=\sum_{i=0}^{x}\binom{n}{i} p^{i}(1-p)^{(n-i)} I_{(0,1, \ldots, n)}(i) .
$$

The result, $y$, is the probability of observing up to $x$ successes in $n$ independent trials, where the probability of success in any given trial is $p$. The indicator function $I_{(0,1, \ldots, n)}(i)$ ensures that $x$ only adopts values of $0,1, \ldots, n$.

If a baseball team plays 162 games in a season and has a $50-50$ chance of winning any game, then the probability of that team winning more than 100 games in a season is:

```
1 - binocdf(100,162,0.5)
```

The result is 0.001 (i.e., 1-0.999). If a team wins 100 or more games in a season, this result suggests that it is likely that the team's true probability of winning any game is greater than 0.5 .

[^1]How To

- "Binomial Distribution" on page B-7


## binofit

Purpose Binomial parameter estimates
$\begin{array}{cl}\text { Syntax } \quad & \text { phat }=\operatorname{binofit}(x, n) \\ & {[\text { phat, pci }]=\operatorname{binofit}(x, n)} \\ & {[p h a t, p c i]=\operatorname{binofit}(x, n, a l p h a)}\end{array}$

## Description

phat $=$ binofit $(x, n)$ returns a maximum likelihood estimate of the probability of success in a given binomial trial based on the number of successes, $x$, observed in $n$ independent trials. If $x=(x(1), x(2)$, $\ldots x(k))$ is a vector, binofit returns a vector of the same size as $x$ whose ith entry is the parameter estimate for $x(i)$. All $k$ estimates are independent of each other. If $n=(n(1), n(2), \ldots, n(k))$ is a vector of the same size as $x$, the binomial fit, binofit, returns a vector whose ith entry is the parameter estimate based on the number of successes $x(i)$ in $n(i)$ independent trials. A scalar value for $x$ or $n$ is expanded to the same size as the other input.
[phat, pci] = binofit(x, n) returns the probability estimate, phat, and the $95 \%$ confidence intervals, pci. binofit uses the Clopper-Pearson method to calculate confidence intervals.
[phat,pci] = binofit(x,n,alpha) returns the 100(1-alpha)\% confidence intervals. For example, alpha $=0.01$ yields $99 \%$ confidence intervals.

Note binofit behaves differently than other Statistics Toolbox functions that compute parameter estimates, in that it returns independent estimates for each entry of $x$. By comparison, expfit returns a single parameter estimate based on all the entries of $x$.

Unlike most other distribution fitting functions, the binofit function treats its input $x$ vector as a collection of measurements from separate samples. If you want to treat $x$ as a single sample and compute a single parameter estimate for it, you can use binofit ( $\operatorname{sum}(x), \operatorname{sum}(n))$ when n is a vector, and binofit (sum $(\mathrm{X}), \mathrm{N}$ *length $(\mathrm{X})$ ) when n is a scalar.
This example generates a binomial sample of 100 elements, where the probability of success in a given trial is 0.6 , and then estimates this probability from the outcomes in the sample.
r = binornd(100,0.6);
r = binornd(100,0.6);
r = binornd(100,0.6);
[phat,pci] = binofit(r,100)
[phat,pci] = binofit(r,100)
[phat,pci] = binofit(r,100)
phat =
phat =
phat =
0.5800
0.5800
0.5800
pci =
pci =
pci =
0.4771 0.6780
0.4771 0.6780
0.4771 0.6780
The $95 \%$ confidence interval, pci, contains the true value, 0.6.

References<br>[1] Johnson, N. L., S. Kotz, and A. W. Kemp. Univariate Discrete Distributions. Hoboken, NJ: Wiley-Interscience, 1993.

See Also<br>mle | binopdf | binocdf | binoinv | binostat | binornd<br>How To . "Binomial Distribution" on page B-7

## binoinv

## Purpose Binomial inverse cumulative distribution function

$$
\text { Syntax } \quad X=\operatorname{binoinv}(Y, N, P)
$$

Description $\quad X=\operatorname{binoinv}(Y, N, P)$ returns the smallest integer $X$ such that the binomial cdf evaluated at $X$ is equal to or exceeds $Y$. You can think of $Y$ as the probability of observing $X$ successes in $N$ independent trials where $P$ is the probability of success in each trial. Each $X$ is a positive integer less than or equal to $N$.
$Y, N$, and $P$ can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The parameters in $N$ must be positive integers, and the values in both $P$ and $Y$ must lie on the interval [01].

## Examples If a baseball team has a 50-50 chance of winning any game, what is a reasonable range of games this team might win over a season of 162 games?

```
binoinv([0.05 0.95],162,0.5)
ans =
    71 91
```

This result means that in $90 \%$ of baseball seasons, a .500 team should win between 71 and 91 games.

| See Also | icdf \| binopdf | binocdf | binofit | binostat | binornd |
| :--- | :--- |
| How To | - "Binomial Distribution" on page B-7 |

## Purpose

Binomial probability density function

## Syntax

$Y=\operatorname{binopdf}(X, N, P)$

Examples A Quality Assurance inspector tests 200 circuit boards a day. If $2 \%$ of the boards have defects, what is the probability that the inspector will find no defective boards on any given day?

```
binopdf(0,200,0.02)
ans =
    0.0176
```

What is the most likely number of defective boards the inspector will find?

```
defects=0:200;
y = binopdf(defects,200,.02);
```


## binopdf

$$
\begin{aligned}
& {[x, i]=\max (y) ;} \\
& \text { defects }(i) \\
& \text { ans }= \\
& \quad 4
\end{aligned}
$$

See Also pdf | binoinv | binocdf | binofit | binostat | binornd
How To - "Binomial Distribution" on page B-7

## binornd

## Purpose Binomial random numbers

Syntax
R = binornd(N, P)
R = binornd(N, $\mathrm{P}, \mathrm{m}, \mathrm{n}, \ldots$ )
$R=$ binornd(N,P,[m,n,...])

## Algorithms

## Examples

```
n = 10:10:60;
    r1 = binornd(n,1./n)
r1 =
        2
r2 = binornd(n,1./n,[1 6])
r2 =
```



```
r3 = binornd(n,1./n,1,6)
r3 =
    0
```

See Also random | binoinv | binocdf | binofit | binostat | binopdf

## binornd

How To

- "Binomial Distribution" on page B-7


## Purpose

Binomial mean and variance

## Syntax

$[\mathrm{M}, \mathrm{V}]=$ binostat( $\mathrm{N}, \mathrm{P}$ )
$[M, V]=$ binostat ( $N, P$ ) returns the mean of and variance for the binomial distribution with parameters specified by the number of trials, N , and probability of success for each trial, P. N and P can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of M and V . A scalar input for N or P is expanded to a constant array with the same dimensions as the other input.

The mean of the binomial distribution with parameters $n$ and $p$ is $n p$. The variance is $n p q$, where $q=1-p$.

## Examples

```
n = logspace(1,5,5)
n =
    10}1001001000 10000 10000
[m,v] = binostat(n,1./n)
m =
    1 1 1 1 1 1 1
v =
```



```
[m,v] = binostat(n,1/2)
m =
    5 50 500 5000 50000
v =
    1.0e+04 *
    0.0003 0.0025 0.0250 0.2500 2.5000
```

See Also binoinv | binocdf | binofit | binornd | binopdf
How To . "Binomial Distribution" on page B-7

## biplot

## Purpose <br> Syntax <br> Description

Biplot

## Input Arguments

biplot(coefs) h = biplot(coefs,'Name', Value)
biplot (coefs) creates a biplot of the coefficients in the matrix coefs. The biplot is 2-D if coefs has two columns or 3-D if it has three columns. coefs usually contains principal component coefficients created with princomp, pcacov, or factor loadings estimated with factoran. The axes in the biplot represent the principal components or latent factors (columns of coefs), and the observed variables (rows of coefs) are represented as vectors.
A biplot allows you to visualize the magnitude and sign of each variable's contribution to the first two or three principal components, and how each observation is represented in terms of those components.
biplot imposes a sign convention, forcing the element with largest magnitude in each column of coefs to be positive. This flips some of the vectors in coefs to the opposite direction, but often makes the plot easier to read. Interpretation of the plot is unaffected, because changing the sign of a coefficient vector does not change its meaning.
h = biplot(coefs, 'Name', Value) specifies one or more name/value input pairs and returns a column vector of handles to the graphics objects created by biplot. The $h$ contains, in order, handles corresponding to variables (line handles, followed by marker handles, followed by text handles), to observations (if present, marker handles followed by text handles), and to the axis lines.

## Name-Value Pair Arguments

Scores
Plots both coefs and the scores in the matrix scores in the biplot. scores usually contains principal component scores created with princomp or factor scores estimated with factoran.

Each observation (row of scores) is represented as a point in the biplot.

## VarLabels

Labels each vector (variable) with the text in the character array or cell array varlabels.

## ObsLabels

Uses the text in the character array or cell array obslabels as observation names when displaying data cursors.

## Positive

- 'true' - restricts the biplot to the positive quadrant (in 2-D) or octant (in 3-D).
- 'false' - makes the biplot over the range $+/-\max (\operatorname{coefs}(:))$ for all coordinates.


## Default: false

## PropertyName

Specifies optional property name/value pairs for all line graphics objects created by biplot.

Perform a principal component analysis of the data in carsmall.mat:

```
load carsmall
x = [Acceleration Displacement Horsepower MPG Weight];
x = x(all(~isnan(x),2),:);
[coefs,score] = princomp(zscore(x));
```

View the data and the original variables in the space of the first three principal components:

```
vbls = \{'Accel','Disp','HP','MPG','Wgt'\};
biplot(coefs(:, 1:3), 'scores', score(:, 1:3), ...
        'varlabels', vbls);
```



Component 2
Component 1

See Also
factoran | nnmf | princomp | pcacov | rotatefactors
Purpose Bootstrap confidence interval
Syntax

ci = bootci(nboot,bootfun,...)
ci = bootci(nboot,\{bootfun,...\},'alpha',alpha)
ci = bootci(nboot,\{bootfun,...\},...,'type',type)
ci = bootci(nboot,\{bootfun,...\},....','type','student',
'nbootstd', nbootstd)
ci = bootci(nboot,\{bootfun,...\},...,'type','student','stderr',
stderr)
ci = bootci(nboot,\{bootfun,...\},...,'Weights', weights)
ci = bootci(nboot,\{bootfun,...\},...,'Options',options)
[ci,bootstat] = bootci(...)

## Description

ci $=$ bootci(nboot, bootfun, ...) computes the $95 \%$ bootstrap confidence interval of the statistic computed by the function bootfun. nboot is a positive integer indicating the number of bootstrap samples used in the computation. bootfun is a function handle specified with @, and must return a scalar. The third and later input arguments to bootci are data (scalars, column vectors, or matrices) that are used to create inputs to bootfun. bootci creates each bootstrap sample by sampling with replacement from the rows of the non-scalar data arguments (these must have the same number of rows). Scalar data are passed to bootfun unchanged.
If bootfun returns a scalar, ci is a vector containing the lower and upper bounds of the confidence interval. If bootfun returns a vector of length $m$, ci is an array of size 2 -by- $m$, where $\mathrm{ci}(1,:)$ are lower bounds and $\mathrm{ci}(2,:)$ are upper bounds. If bootfun returns an array of size $m$-by- $n$-by- $p$-by-..., ci is an array of size 2 -by- $m$-by- $n$-by- $p$-by-..., where $\operatorname{ci}(1,:,:,:, \ldots)$ is an array of lower bounds and $\operatorname{ci}(2,:,:,:, \ldots)$ is an array of upper bounds.
ci $=$ bootci(nboot, \{bootfun, ...\},'alpha', alpha) computes the 100* ( 1 -alpha) bootstrap confidence interval of the statistic defined by the function bootfun. bootfun and the data that bootci passes to it are contained in a single cell array. alpha is a scalar between 0 and 1 . The default value of alpha is 0.05 .
ci = bootci(nboot,\{bootfun,...\},...,'type',type) computes the bootstrap confidence interval of the statistic defined by the function bootfun. type is the confidence interval type, chosen from among the following strings:

- 'norm' or 'normal' - Normal approximated interval with bootstrapped bias and standard error.
- 'per' or 'percentile' - Basic percentile method.
- 'cper' or 'corrected percentile' - Bias corrected percentile method.
- 'bca' - Bias corrected and accelerated percentile method. This is the default.
- 'stud' or 'student' - Studentized confidence interval.
ci =
bootci(nboot,\{bootfun,...\},...,'type','student', 'nbootstd', nbootstd)
computes the studentized bootstrap confidence interval of the statistic defined by the function bootfun. The standard error of the bootstrap statistics is estimated using bootstrap, with nbootstd bootstrap data samples. nbootstd is a positive integer value. The default value of nbootstd is 100 .
ci =
bootci(nboot, \{bootfun, ...\},...,'type', 'student', 'stderr', stderr)
computes the studentized bootstrap confidence interval of statistics defined by the function bootfun. The standard error of the bootstrap statistics is evaluated by the function stderr. stderr is a function handle. stderr takes the same arguments as bootfun and returns the standard error of the statistic computed by bootfun.
ci = bootci(nboot, \{bootfun, ...\},..., 'Weights', weights) specifies observation weights. weights must be a vector of non-negative numbers with at least one positive element. The number of elements in weights must be equal to the number of rows in non-scalar input arguments to bootfun. To obtain one bootstrap replicate,
bootstrp samples $N$ out of $N$ with replacement using these weights as multinomial sampling probabilities.
ci = bootci(nboot,\{bootfun,...\},...,'Options',options) specifies options that govern the computation of bootstrap iterations. One option requests that bootci perform bootstrap iterations using multiple processors, if the Parallel Computing Toolbox is available. Two options specify the random number streams to be used in bootstrap resampling. This argument is a struct that you can create with a call to statset. You can retrieve values of the individual fields with a call to statget. Applicable statset parameters are:
- 'UseParallel' - If 'always' and if a matlabpool of the Parallel Computing Toolbox is open, compute bootstrap iterations in parallel. If the Parallel Computing Toolbox is not installed, or a matlabpool is not open, computation occurs in serial mode. Default is 'never', or serial computation.
- UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'.
- Streams - A RandStream object or cell array of such objects. If you do not specify Streams, bootci uses the default stream or streams. If you choose to specify Streams, use a single object except in the case
- You have an open MATLAB pool
- UseParallel is 'always'
- UseSubstreams is 'never'

In that case, use a cell array the same size as the MATLAB pool.
For more information on using parallel computing, see Chapter 17, "Parallel Statistics".
[ci,bootstat] = bootci(...) also returns the bootstrapped statistic computed for each of the nboot bootstrap replicate samples. Each row of bootstat contains the results of applying bootfun to one bootstrap

## bootci

sample. If bootfun returns a matrix or array, then this output is converted to a row vector for storage in bootstat.

## Examples

Compute the confidence interval for the capability index in statistical process control:

```
y = normrnd(1,1,30,1); % Simulated process data
LSL = -3; USL = 3; % Process specifications
capable = @(x)(USL-LSL)./(6* std(x)); % Process capability
ci = bootci(2000,capable,y) % BCa confidence interval
ci =
    0.8122
    1.2657
sci = bootci(2000,{capable,y},'type','student') % Studentized ci
sci =
    0.7739
    1.2707
```

See Also bootstrp | jackknife | statget | statset | randsample | parfor

| Purpose | Bootstrap sampling |
| :---: | :---: |
| Syntax | ```bootstat = bootstrp(nboot,bootfun,d1,...) [bootstat,bootsam] = bootstrp(...) bootstat = bootstrp(...,'Name',Value)``` |
| Description | bootstat = bootstrp(nboot,bootfun,d1,...) draws nboot bootstrap data samples, computes statistics on each sample using bootfun, and returns the results in the matrix bootstat. nboot must be a positive integer. bootfun is a function handle specified with a. Each row of bootstat contains the results of applying bootfun to one bootstrap sample. If bootfun returns a matrix or array, then this output is converted to a row vector for storage in bootstat. |
|  | The third and later input arguments ( $\mathrm{d} 1, \ldots$ ) are data (scalars, column vectors, or matrices) used to create inputs to bootfun. bootstrp creates each bootstrap sample by sampling with replacement from the rows of the non-scalar data arguments (these must have the same number of rows). bootfun accepts scalar data unchanged. |
|  | [bootstat, bootsam] = bootstrp(...) returns an n-by-nboot matrix of bootstrap indices, bootsam. Each column in bootsam contains indices of the values that were drawn from the original data sets to constitute the corresponding bootstrap sample. For example, if d1, .. each contain 16 values, and nboot $=4$, then bootsam is a 16 -by- 4 matrix. The first column contains the indices of the 16 values drawn from $\mathrm{d} 1, \ldots$, for the first of the four bootstrap samples, the second column contains the indices for the second of the four bootstrap samples, and so on. (The bootstrap indices are the same for all input data sets.) To get the output samples bootsam without applying a function, set bootfun to empty ([]). |
|  | bootstat $=$ bootstrp(..., 'Name', Value) uses additional arguments specified by one or more Name, Value pair arguments. The name-value pairs must appear after the data arguments. The available name-value pairs: |

- 'Weights' - Observation weights. The weights value must be a vector of nonnegative numbers with at least one positive element. The number of elements in weights must be equal to the number of rows in non-scalar input arguments to bootstrp. To obtain one bootstrap replicate, bootstrp samples $N$ out of $N$ with replacement using these weights as multinomial sampling probabilities.
- 'Options' - The value is a structure that contains options specifying whether to compute bootstrap iterations in parallel, and specifying how to use random numbers during the bootstrap sampling. Create the options structure with statset. Applicable statset parameters are:
- 'UseParallel' - If 'always' and if a matlabpool of the Parallel Computing Toolbox is open, compute bootstrap iterations in parallel. If the Parallel Computing Toolbox is not installed, or a matlabpool is not open, computation occurs in serial mode. Default is 'never', meaning serial computation.
- UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'.
- Streams - A RandStream object or cell array of such objects. If you do not specify Streams, bootstrp uses the default stream or streams. If you choose to specify Streams, use a single object except in the case
- You have an open MATLAB pool
- UseParallel is 'always'
- UseSubstreams is 'never'

In that case, use a cell array the same size as the MATLAB pool.
For more information on using parallel computing, see Chapter 17, "Parallel Statistics".

## Examples Bootstrapping a Correlation Coefficient Standard Error

Load a data set containing the LSAT scores and law-school GPA for 15 students. These 15 data points are resampled to create 1000 different data sets, and the correlation between the two variables is computed for each data set.

```
load lawdata
[bootstat,bootsam] = bootstrp(1000,@corr,lsat,gpa);
```

Display the first 5 bootstrapped correlation coefficients.

```
bootstat(1:5,:)
ans =
    0.6600
    0.7969
    0.5807
    0.8766
    0.9197
```

Display the indices of the data selected for the first 5 bootstrap samples.


## bootstrp

hist(bootstat)


The histogram shows the variation of the correlation coefficient across all the bootstrap samples. The sample minimum is positive, indicating that the relationship between LSAT score and GPA is not accidental.
Finally, compute a bootstrap standard of error for the estimated correlation coefficient.

```
se = std(bootstat)
se =
    0.1327
```


## Estimating the Density of Bootstrapped Statistic

Compute a sample of 100 bootstrapped means of random samples taken from the vector Y , and plot an estimate of the density of these bootstrapped means:

```
y = exprnd(5,100,1);
m = bootstrp(100,@mean,y);
[fi,xi] = ksdensity(m);
plot(xi,fi);
```



## bootstrp

## Bootstrapping More Than One Statistic

Compute a sample of 100 bootstrapped means and standard deviations of random samples taken from the vector Y , and plot the bootstrap estimate pairs:

```
y = exprnd(5,100,1);
stats = bootstrp(100,@(x)[mean(x) std(x)],y);
plot(stats(:,1),stats(:,2),'o')
```



## Bootstrapping a Regression Model

Estimate the standard errors for a coefficient vector in a linear regression by bootstrapping residuals:

```
load hald
x = [ones(size(heat)),ingredients];
y = heat;
b = regress(y,x);
yfit = x*b;
resid = y - yfit;
se = std(bootstrp(...
    1000,@(bootr)regress(yfit+bootr,x),resid));
```

See Also
hist | bootci | ksdensity | parfor | random | randsample | RandStream | statget | statset

## boxplot

Purpose Box plot
Syntax boxplot(X)
boxplot(X,G)
boxplot(axes, X,...)
boxplot(...,'Name', value)
Description
boxplot $(\mathrm{X})$ produces a box plot of the data in X . If X is a matrix, there is one box per column; if $X$ is a vector, there is just one box. On each box, the central mark is the median, the edges of the box are the 25th and 75 th percentiles, the whiskers extend to the most extreme data points not considered outliers, and outliers are plotted individually.
boxplot $(X, G)$ specifies one or more grouping variables $G$, producing a separate box for each set of $X$ values sharing the same $G$ value or values (see "Grouped Data" on page 2-34). Grouping variables must have one row per element of $X$, or one row per column of $X$. Specify a single grouping variable in $G$ using a vector, a character array, a cell array of strings, or a vector categorical array; specify multiple grouping variables in G using a cell array of these variable types, such as \{G1 G2 G3\}, or by using a matrix. If multiple grouping variables are used, they must all be the same length. Groups that contain a NaN value or an empty string in a grouping variable are omitted, and are not counted in the number of groups considered by other parameters.

By default, character and string grouping variables are sorted in the order they initially appear in the data, categorical grouping variables are sorted by the order of their levels, and numeric grouping variables are sorted in numeric order. To control the order of groups, do one of the following:

- Use categorical variables in $G$ and specify the order of their levels.
- Use the 'grouporder' parameter described below.
- Pre-sort your data.
boxplot(axes, $\mathrm{X}, \ldots$ ) creates the plot in the axes with handle axes.


## boxplot

boxplot(..., 'Name', value) specifies one or more optional parameter name/value pairs, as described in the following table. Specify Name in single quotes.
\(\left.$$
\begin{array}{l|l}\hline \text { Name } & \text { Value } \\
\hline \text { plotstyle } & \begin{array}{l}\text { - 'traditional' — Traditional box style. } \\
\text { This is the default. }\end{array}
$$ <br>
\hline bompact ' — Box style designed for plots <br>
with many groups. This style changes the <br>
defaults for some other parameters, as <br>

described in the following table.\end{array}\right\}\)| - 'outline' — Draws an unfilled box with |
| :--- |
| dashed whiskers. This is the default. |
| - 'filled ' — Draws a narrow filled box with |
| lines for whiskers. |

## boxplot

| Name | Value |
| :--- | :--- |
| datalim | A two-element vector containing lower and <br> upper limits, used by 'extrememode' to <br> determine which points are extreme. The <br> default is [ - Inf Inf ]. |
| extrememode | - 'clip' - Moves data outside the datalim <br> limits to the limit. This is the default. <br> - ' compress ' - Evenly distributes data <br> outside the datalim limits in a region just <br> outside the limit, retaining the relative <br> order of the points. |
| A dotted line marks the limit if any points |  |
| are outside it, and two gray lines mark |  |
| the compression region if any points are |  |
| compressed. Values at +/-Inf can be clipped |  |
| or compressed, but NaN values still do not |  |
| appear on the plot. Box notches are drawn to |  |
| scale and may extend beyond the bounds if the |  |
| median is inside the limit; they are not drawn |  |
| if the median is outside the limits. |  |

## boxplot

| Name | Value <br> fullfactors <br> - 'off' - This is the default. <br> - on' — Create a group for each possible <br> combination of group variable values, <br> including combinations that do not appear <br> in the data. |
| :--- | :--- |
| factorseparator | Specifies which factors should have their <br> values separated by a grid line. The value <br> may be 'auto ' or a vector of grouping variable <br> numbers. For example, [1 2] adds a separator <br> line when the first or second grouping variable <br> changes value. 'auto' is [ ] for one grouping <br> variable and [1] for two or more grouping <br> variables. The default is [ ]. |
| factorgap | Specifies an extra gap to leave between boxes <br> when the corresponding grouping factor <br> changes value, expressed as a percentage of <br> the width of the plot. For example, with [3 1], <br> the gap is 3\% of the width of the plot between <br> groups with different values of the first <br> grouping variable, and 1\% between groups <br> with the same value of the first grouping <br> variable but different values for the second. <br> 'auto' specifies that boxplot should choose a <br> gap automatically. The default is [ ]. |
| grouporder | Order of groups for plotting, specified as a <br> cell array of strings. With multiple grouping <br> variables, separate values within each string <br> with a comma. Using categorical arrays as <br> grouping variables is an easier way to control <br> the order of the boxes. The default is [ ], which <br> does not reorder the boxes. |

## boxplot



## boxplot

$\left.\begin{array}{l|l}\hline \text { Name } & \text { Value } \\ \hline \text { labelverbosity } & \begin{array}{l}\text { - 'all' - Displays every label. This is the } \\ \text { default. }\end{array} \\ \text { 'minor' — Displays a label for a factor only } \\ \text { when that factor has a different value from } \\ \text { the previous group. } \\ \text { - 'majorminor' — Displays a label for a } \\ \text { factor when that factor or any factor major } \\ \text { to it has a different value from the previous } \\ \text { group. }\end{array}\right\}$

## boxplot

| Name | Value |
| :--- | :--- |
| outliersize | Size of the marker used for outliers, in points. <br> The default is 6 (6/72 inch). |
| positions | Box positions specified as a numeric vector <br> with one entry per group or X value. The <br> default is 1: numGroups, where numGroups is <br> the number of groups. |
| symbol | Symbol and color to use for outliers, using <br> the same values as the LineSpec parameter <br> in plot. The default is 'r+'. If the symbol <br> is omitted then the outliers are invisible; if <br> the color is omitted then the outliers have the <br> same color as their corresponding box. |
| whisker | Maximum whisker length $w$. The default is a <br> $w$ of 1.5. Points are drawn as outliers if they |
| are larger than $q_{3}+w\left(q_{3}-q_{1}\right)$ or smaller than |  |
| $q_{1}-w\left(q_{3}-q_{1}\right)$, where $q_{1}$ and $q_{3}$ are the 25th |  |
| and 75th percentiles, respectively. The default |  |
| of 1.5 corresponds to approximately +/-2.7o |  |
| and 99.3 coverage if the data are normally |  |
| distributed. The plotted whisker extends to |  |
| the adjacent value, which is the most extreme |  |
| data value that is not an outlier. Set whisker |  |
| to 0 to give no whiskers and to make every |  |
| point outside of $q_{1}$ and $q_{3}$ an outlier. |  |

When the plotstyle parameter takes the value 'compact', the following default values for other parameters apply.

## boxplot

| Parameter | Default when plotstyle is 'compact' |
| :--- | :--- |
| boxstyle | 'filled' |
| factorseparator | 'auto' |
| factorgap | 'auto' |
| jitter | 0.5 |
| labelorientation | 'inline' |
| labelverbosity | 'majorminor' |
| medianstyle | 'target' |
| outliersize | 4 |
| symbol | 'o' |

You can see data values and group names using the data cursor in the figure window. The cursor shows the original values of any points affected by the datalim parameter. You can label the group to which an outlier belongs using the gname function.
To modify graphics properties of a box plot component, use findobj with the Tag property to find the component's handle. Tag values for box plot components depend on parameter settings, and are listed in the table below.

| Parameter <br> Settings | Tag Values |
| :--- | :--- |
| All settings | • 'Box' <br> • 'Outliers' |
| When plotstyle is <br> 'traditional' | • 'Median' <br> - 'Upper Whisker' <br>  <br> - 'Lower Whisker' <br>  <br>  <br>  <br> • 'Upper Adjacent Value' <br> • 'Lower Adjacent Value' |

## boxplot

| Parameter <br> Settings | Tag Values |
| :--- | :--- |
| When plotstyle is <br> 'compact' | • 'Whisker' <br> • 'MedianOuter' <br> • 'MedianInner' |
| When notch is <br> 'marker' | • 'NotchLo' <br> • 'NotchHi' |

## Examples

## Example 1

Create a box plot of car mileage, grouped by country:
load carsmall
boxplot(MPG,Origin)


## Example 2

Create notched box plots for two groups of sample data:

```
x1 = normrnd(5,1,100,1);
x2 = normrnd(6,1,100,1);
boxplot([x1,x2],'notch','on')
```


## boxplot



The difference between the medians of the two groups is approximately 1. Since the notches in the box plot do not overlap, you can conclude, with $95 \%$ confidence, that the true medians do differ.

The following figure shows the box plot for the same data with the length of the whiskers specified as 1.0 times the interquartile range. Points beyond the whiskers are displayed using + .

```
boxplot([x1,x2],'notch','on','whisker',1)
```



## Example 3

A plotstyle of 'compact' is useful for large numbers of groups:

```
X = randn(100,25);
subplot(2,1,1)
boxplot(X)
subplot(2,1,2)
boxplot(X,'plotstyle','compact')
```


## boxplot



## References

[1] McGill, R., J. W. Tukey, and W. A. Larsen. "Variations of Boxplots." The American Statistician. Vol. 32, No. 1, 1978, pp. 12-16.
[2] Velleman, P.F., and D.C. Hoaglin. Applications, Basics, and Computing of Exploratory Data Analysis. Pacific Grove, CA: Duxbury Press, 1981.
[3] Nelson, L. S. "Evaluating Overlapping Confidence Intervals." Journal of Quality Technology. Vol. 21, 1989, pp. 140-141.

[^2]
## piecewisedistribution.boundary

Purpose Piecewise distribution boundaries
Syntax $\quad[p, q]=$ boundary $(o b j)$,
Description $[p, q]=$ boundary $(o b j)$ returns the boundary points between segments of the piecewise distribution object, obj. p is a vector of cumulative probabilities at each boundary. $q$ is a vector of quantiles at each boundary.
$[p, q]=$ boundary (obj,i) returns $p$ and $q$ for the $i$ th boundary.
Examples
Fit Pareto tails to a $t$ distribution at cumulative probabilities 0.1 and 0.9:
$\mathrm{t}=\operatorname{trnd}(3,100,1)$;
obj = paretotails(t, 0.1,0.9);
[p,q] = boundary(obj)
$p=$
0.1000
0.9000
q $=$
-1.7766
1.8432
See Also paretotails | cdf | icdf | nsegments

| Purpose | $D$-optimal design from candidate set using row exchanges |
| :---: | :---: |
| Syntax | ```rlist = candexch(C,nrows) rlist = candexch(C,nrows,Name,Value)``` |
| Description | rlist $=$ candexch ( $C$, nrows $)$ uses a row-exchange algorithm to select a $D$-optimal design from the candidate set C . |
|  | rlist $=\operatorname{candexch}(C$, nrows, Name, Value) generates a $D$-optimal design with additional options specified by one or more Name, Value pair arguments. |
| Input Arguments | c |
|  | N -by- P matrix containing the values of P model terms at each of N points. |
|  | nrows |
|  | The desired number of rows in the design. |
|  | Name-Value Pair Arguments |
|  | Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN. |
|  | display |
|  | When 'on', displays iteration number. Disable the display by setting to 'off'. |
|  | Default: 'on', except when the UseParallel option is 'always ' |
|  | init |
|  | nrows-by-P matrix giving an initial design. |
|  | Default: A random subset of the rows of C |

maxiter
Maximum number of iterations, a positive integer.
Default: 10
options
A structure that specifies whether to run in parallel, and specifies the random stream or streams. Create the options structure with statset. Option fields:

- UseParallel - Set to 'always' to compute in parallel. Default is 'never'.
- UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'.
- Streams - A RandStream object or cell array of such objects. If you do not specify Streams, candexch uses the default stream or streams. If you choose to specify Streams, use a single object except in the case
- You have an open MATLAB pool
- UseParallel is 'always'
- UseSubstreams is 'never'

In that case, use a cell array the same size as the MATLAB pool.
For more information on using parallel computing, see Chapter 17, "Parallel Statistics".

## Default: []

## start

An nobs-by-p matrix of factor settings, specifying a set of nobs fixed design points to include in the design. candexch finds
nrows additional rows to add to the start design. The parameter provides the same functionality as the daugment function, using a row-exchange algorithm rather than a coordinate-exchange algorithm.

Default: []

## tries

Number of times to try to generate a design from a new starting point. The algorithm uses random points for each try, except possibly the first.

## Default: 1

## Output rlist

Arguments

## Examples

## Algorithms

## Alternatives

candexch selects a starting design $X$ at random, and uses a row-exchange algorithm to iteratively replace rows of $X$ by rows of $C$ in an attempt to improve the determinant of X '*X.

The rowexch function also generates $D$-optimal designs using a row-exchange algorithm, but it automatically generates a candidate
set that is appropriate for a specified model. The daugment function augments a set of fixed design points using a coordinate-exchange algorithm; the 'start' parameter provides the same functionality using the row exchange algorithm.
See Also candgen | rowexch | cordexch | daugment | x2fx
Tutorials- "Specifying Candidate Sets" on page 15-21
How To - "D-Optimal Designs" on page 15-15

## Purpose Candidate set generation

Syntax $\quad d C=$ candgen(nfactors,'model')
[dC,C] = candgen(nfactors,'model')
[...] = candgen(nfactors,'model','Name',value)

## Description

dC = candgen(nfactors,'model') generates a candidate set dC of treatments appropriate for estimating the parameters in the model with nfactors factors. dC has nfactors columns and one row for each candidate treatment. model is one of the following strings, specified inside single quotes:

- linear - Constant and linear terms. This is the default.
- interaction - Constant, linear, and interaction terms
- quadratic - Constant, linear, interaction, and squared terms
- purequadratic - Constant, linear, and squared terms

Alternatively, model can be a matrix specifying polynomial terms of arbitrary order. In this case, model should have one column for each factor and one row for each term in the model. The entries in any row of model are powers for the factors in the columns. For example, if a model has factors $\mathrm{X} 1, \mathrm{X} 2$, and X 3 , then a row [ $\left.\begin{array}{lll}0 & 1 & 2\end{array}\right]$ in model specifies the term (X1.^0).*(X2.^1) .* (X3.^2). A row of all zeros in model specifies a constant term, which can be omitted.
[dC,C] = candgen(nfactors,'model') also returns the design matrix C evaluated at the treatments in dC . The order of the columns of C for a full quadratic model with $n$ terms is:

1 The constant term
2 The linear terms in order $1,2, \ldots, n$
3 The interaction terms in order (1, 2), (1, 3), $\ldots,(1, n),(2,3), \ldots,(n-1, n)$
4 The squared terms in order $1,2, \ldots, n$

Other models use a subset of these terms, in the same order.
Pass C to candexch to generate a $D$-optimal design using a coordinate-exchange algorithm.
[...] = candgen(nfactors,'model','Name', value) specifies one or more optional name/value pairs for the design. Valid parameters and their values are listed in the following table. Specify Name inside single quotes.

| Name | Value |
| :--- | :--- |
| bounds | Lower and upper bounds for each factor, specified as <br> a 2-by-nfactors matrix. Alternatively, this value <br> can be a cell array containing nfactors elements, <br> each element specifying the vector of allowable <br> values for the corresponding factor. |
| categorical | Indices of categorical predictors. |
| levels | Vector of number of levels for each factor. |

Note The rowexch function automatically generates a candidate set using candgen, and then creates a $D$-optimal design from that candidate set using candexch. Call candexch separately to specify your own candidate set to the row-exchange algorithm.

## Examples

The following example uses rowexch to generate a five-run design for a two-factor pure quadratic model using a candidate set that is produced internally:

```
dRE1 = rowexch(2,5,'purequadratic','tries',10)
dRE1 =
    -1 1
    0
    1 -1
    0
```

The same thing can be done using candgen and candexch in sequence:

```
[dC,C] = candgen(2,'purequadratic') % Candidate set, C
dC =
    -1 -1
        0 -1
        1 -1
        -1 0
        0
        0
        -1 1
        0
        1
C =
\begin{tabular}{rrrrr}
1 & -1 & -1 & 1 & 1 \\
1 & 0 & -1 & 0 & 1 \\
1 & 1 & -1 & 1 & 1 \\
1 & -1 & 0 & 1 & 0 \\
1 & 0 & 0 & 0 & 0 \\
1 & 1 & 0 & 1 & 0 \\
1 & -1 & 1 & 1 & 1 \\
1 & 0 & 1 & 0 & 1 \\
1 & 1 & 1 & 1 & 1
\end{tabular}
treatments = candexch(C,5,'tries',10) % Find D-opt subset
treatments =
    2
        1
        7
        3
        4
dRE2 = dC(treatments,:) % Display design
dRE2 =
    0 -1
    -1 -1
    -1 1
```


## candgen



See Also
candexch | rowexch

## Purpose <br> Syntax <br> Description

Canonical correlation
[A,B] = canoncorr(X,Y)
[A,B,r] = canoncorr(X,Y)
[A,B,r,U,V] = canoncorr(X,Y)
[A,B,r,U,V,stats] = canoncorr(X,Y)
$[A, B]=$ canoncorr $(X, Y)$ computes the sample canonical coefficients for the $n$-by-d1 and $n$-by-d2 data matrices $X$ and $Y$. $X$ and $Y$ must have the same number of observations (rows) but can have different numbers of variables (columns). A and B are d1-by-d and d2-by-d matrices, where $d=\min (\operatorname{rank}(X), \operatorname{rank}(Y))$. The $j$ th columns of $A$ and $B$ contain the canonical coefficients, i.e., the linear combination of variables making up the $j$ th canonical variable for $X$ and $Y$, respectively. Columns of $A$ and $B$ are scaled to make the covariance matrices of the canonical variables the identity matrix (see $U$ and $V$ below). If $X$ or $Y$ is less than full rank, canoncorr gives a warning and returns zeros in the rows of $A$ or $B$ corresponding to dependent columns of $X$ or $Y$.
$[A, B, r]=$ canoncorr $(X, Y)$ also returns a 1-by-d vector containing the sample canonical correlations. The $j$ th element of $r$ is the correlation between the $j$ th columns of $U$ and $V$ (see below).
$[A, B, r, U, V]=$ canoncorr $(X, Y)$ also returns the canonical variables, scores. $U$ and $V$ are $n$-by- $d$ matrices computed as

$$
\begin{aligned}
& U=(X-\operatorname{repmat}(\operatorname{mean}(X), N, 1)) * A \\
& V=(Y-\operatorname{repmat}(\operatorname{mean}(Y), N, 1)) * B
\end{aligned}
$$

[A,B,r,U,V,stats] = canoncorr(X,Y) also returns a structure stats containing information relating to the sequence of $d$ null
hypotheses $H_{0}^{(k)}$, that the ( $\mathrm{k}+1$ )st through dth correlations are all zero, for $k=0:(d-1)$. stats contains seven fields, each a 1-by-d vector with elements corresponding to the values of $k$, as described in the following table:

| Field | Description |
| :--- | :--- |
| Wilks | Wilks' lambda (likelihood ratio) statistic |
| chisq | Bartlett's approximate chi-squared statistic for $H_{0}^{(k)}$ <br> with Lawley's modification |
| pChisq | Right-tail significance level for chisq |
| F | Rao's approximate $F$ statistic for $H_{0}^{(k)}$ |
| pF | Right-tail significance level for F |
| df1 | Degrees of freedom for the chi-squared statistic, and <br> the numerator degrees of freedom for the $F$ statistic |
| df2 | Denominator degrees of freedom for the $F$ statistic |

## Examples

```
load carbig;
X = [Displacement Horsepower Weight Acceleration MPG];
nans = sum(isnan(X),2) > 0;
[A B r U V] = canoncorr(X(~nans,1:3),X(~nans,4:5));
plot(U(:,1),V(:,1),'.')
xlabel('0.0025*Disp+0.020*HP-0.000025*Wgt')
ylabel('-0.17*Accel-0.092*MPG')
```



## References

[1] Krzanowski, W. J. Principles of Multivariate Analysis: A User's Perspective. New York: Oxford University Press, 1988.
[2] Seber, G. A. F. Multivariate Observations. Hoboken, NJ: John Wiley \& Sons, Inc., 1984.

See Also manova1 | princomp

Purpose Process capability indices
Syntax $\quad$ S = capability (data, specs)
Description
$S$ = capability(data, specs) estimates capability indices for measurements in data given the specifications in specs. data can be either a vector or a matrix of measurements. If data is a matrix, indices are computed for the columns. specs can be either a two-element vector of the form [ $\mathrm{L}, \mathrm{U}$ ] containing lower and upper specification limits, or (if data is a matrix) a two-row matrix with the same number of columns as data. If there is no lower bound, use - Inf as the first element of specs. If there is no upper bound, use Inf as the second element of specs.

The output S is a structure with the following fields:

- mu - Sample mean
- sigma - Sample standard deviation
- P - Estimated probability of being within limits
- Pl - Estimated probability of being below L
- Pu - Estimated probability of being above $U$
- Cp - (U-L)/(6*sigma)
- Cpl-(mu-L)./(3.*sigma)
- Cpu - (U-mu)./(3.*sigma)
- Cpk - min(Cpl, Cpu)

Indices are computed under the assumption that data values are independent samples from a normal population with constant mean and variance.

Indices divide a "specification width" (between specification limits) by a "process width" (between control limits). Higher ratios indicate a process with fewer measurements outside of specification.

Examples
Simulate a sample from a process with a mean of 3 and a standard deviation of 0.005 :

```
data = normrnd(3,0.005,100,1);
```

Compute capability indices if the process has an upper specification limit of 3.01 and a lower specification limit of 2.99:

```
S = capability(data,[2.99 3.01])
S =
```

    mu: 3.0006
    sigma: 0.0047
            P: 0.9669
            Pl: 0.0116
            Pu: 0.0215
            Cp: 0.7156
        Cpl: 0.7567
        Cpu: 0.6744
        Cpk: 0.6744
    Visualize the specification and process widths:
capaplot(data,[2.99 3.01]);
grid on


## References <br> [1] Montgomery, D. Introduction to Statistical Quality Control. Hoboken, NJ: John Wiley \& Sons, 1991, pp. 369-374.

See Also<br>capaplot | histfit

## Purpose Process capability plot

Syntax $\quad \begin{aligned} & p=\text { capaplot (data, specs) } \\ & {[p, h]=\operatorname{capaplot}(\text { data }, \text { specs })}\end{aligned}$
Description
$p$ = capaplot(data, specs) estimates the mean of and variance for the observations in input vector data, and plots the pdf of the resulting $T$ distribution. The observations in data are assumed to be normally distributed. The output, $p$, is the probability that a new observation from the estimated distribution will fall within the range specified by the two-element vector specs. The portion of the distribution between the lower and upper bounds specified in specs is shaded in the plot.
$[p, h]=$ capaplot(data, specs) additionally returns handles to the plot elements in $h$.
capaplot treats NaN values in data as missing, and ignores them.

## Examples

Simulate a sample from a process with a mean of 3 and a standard deviation of 0.005 :

```
data = normrnd(3,0.005,100,1);
```

Compute capability indices if the process has an upper specification limit of 3.01 and a lower specification limit of 2.99:

```
S = capability(data,[2.99 3.01])
S =
    mu: 3.0006
    sigma: 0.0047
            P: 0.9669
            Pl: 0.0116
            Pu: 0.0215
            Cp: 0.7156
            Cpl: 0.7567
            Cpu: 0.6744
            Cpk: 0.6744
```

Visualize the specification and process widths:

```
capaplot(data,[2.99 3.01]);
grid on
```

Probability Between Limits $=0.96688$


## See Also

capability | histfit

## Purpose Read case names from file

Syntax names = caseread('filename')
names $=$ caseread
Description names = caseread('filename') reads the contents of filename and returns a string matrix of names. filename is the name of a file in the current folder, or the complete path name of any file elsewhere. caseread treats each line as a separate case.
names = caseread displays the Select File to Open dialog box for interactive selection of the input file.

## Examples <br> Read the file months. dat created using the casewrite function.

```
type months.dat
January
February
March
April
May
names = caseread('months.dat')
names =
January
February
March
April
May
```

See Also
casewrite | gname | tdfread | tblread

Purpose Write case names to file

```
Syntax casewrite(strmat,'filename')
casewrite(strmat)
```

Description casewrite(strmat,'filename') writes the contents of string matrix strmat to filename. Each row of strmat represents one case name. filename is the name of a file in the current folder, or the complete path name of any file elsewhere. casewrite writes each name to a separate line in filename.
casewrite (strmat) displays the Select File to Write dialog box for interactive specification of the output file.

## Examples

```
strmat = char('January','February',...
    'March','April','May')
strmat =
January
February
March
April
May
casewrite(strmat,'months.dat')
type months.dat
January
February
March
April
May
```

See Also gname | caseread | tblwrite | tdfread

## Purpose Concatenate categorical arrays

Syntax
c = cat(dim, $A, B, \ldots)$

Description $\quad c=\operatorname{cat}(\operatorname{dim}, A, B, \ldots)$ concatenates the categorical arrays $A, B, \ldots$ along dimension dim. All inputs must have the same size except along dimension dim. The set of categorical levels for C is the sorted union of the sets of levels of the inputs, as determined by their labels.

See Also
cat | horzcat | vertcat

## Purpose Arrays for categorical data

Description
cat
cellstr
char
circshift
ctranspose
disp
display
double
droplevels
end

Construction
categorical
addlevels
Create categorical array

## Methods

Add levels to categorical array

Concatenate categorical arrays Convert categorical array to cell array of strings
Convert categorical array to character array
Shift categorical array circularly
Transpose categorical matrix
Display categorical array
Display categorical array
Convert categorical array to double array
Drop levels
Last index in indexing expression for categorical array

| flipdim | Flip categorical array along specified dimension |
| :---: | :---: |
| fliplr | Flip categorical matrix in left/right direction |
| flipud | Flip categorical matrix in up/down direction |
| getlabels | Access categorical array labels |
| getlevels | Get categorical array levels |
| hist | Plot histogram of categorical data |
| horzcat | Horizontal concatenation for categorical arrays |
| int16 | Convert categorical array to signed 16 -bit integer array |
| int32 | Convert categorical array to signed 32 -bit integer array |
| int64 | Convert categorical array to signed 64-bit integer array |
| int8 | Convert categorical array to signed 8-bit integer array |
| intersect | Set intersection for categorical arrays |
| ipermute | Inverse permute dimensions of categorical array |
| isempty | True for empty categorical array |
| isequal | True if categorical arrays are equal |
| islevel | Test for levels |
| ismember | True for elements of categorical array in set |


| isscalar | True if categorical array is scalar |
| :---: | :---: |
| isundefined | Test for undefined elements |
| isvector | True if categorical array is vector |
| length | Length of categorical array |
| levelcounts | Element counts by level |
| ndims | Number of dimensions of categorical array |
| numel | Number of elements in categorical array |
| permute | Permute dimensions of categorical array |
| reorderlevels | Reorder levels |
| repmat | Replicate and tile categorical array |
| reshape | Resize categorical array |
| rot90 | Rotate categorical matrix 90 degrees |
| setdiff | Set difference for categorical arrays |
| setlabels | Label levels |
| setxor | Set exclusive-or for categorical arrays |
| shiftdim | Shift dimensions of categorical array |
| single | Convert categorical array to single array |
| size | Size of categorical array |

squeeze
subsasgn
subsindex
subsref
summary
times
transpose
uint16
uint32
uint64
uint8
union
unique
vertcat

Squeeze singleton dimensions from categorical array
Subscripted assignment for categorical array

Subscript index for categorical
array
Subscripted reference for categorical array

Summary statistics for categorical array

Product of categorical arrays
Transpose categorical matrix
Convert categorical array to unsigned 16-bit integers

Convert categorical array to unsigned 32-bit integers

Convert categorical array to unsigned 64-bit integers

Convert categorical array to unsigned 8-bit integers

Set union for categorical arrays
Unique values in categorical array

Vertical concatenation for categorical arrays

Text labels for levels
Text label for undefined levels

Copy
Semantics

How To . "Categorical Arrays" on page 2-13
Purpose Create categorical array
Description categorical is an abstract class, and you cannot create instances of it directly. You must create nominal or ordinal arrays.
See Also ..... nominal | ordinal

Purpose Concatenate dataset arrays

## Syntax <br> ds = cat(dim, ds1, ds2, ...)

Description ds = cat (dim, ds1, ds2, ...) concatenates the dataset arrays ds1, ds2, ... along dimension dim by calling the dataset/horzcat or dataset/vertcat method. dim must be 1 or 2 .

See Also horzcat | vertcat

Purpose
Syntax
v=catsplit(t)
v=catsplit(t,j)

Description $\quad v=$ catsplit ( $t$ ) returns an $n$-by- 2 cell array $v$. Each row in $v$ gives left and right values for a categorical split. For each branch node $j$ based on a categorical predictor variable $z$, the left child is chosen if $z$ is in $v(j, 1)$ and the right child is chosen if $z$ is in $v(j, 2)$. The splits are in the same order as nodes of the tree. Nodes for these splits can be found by running cuttype and selecting 'categorical' cuts from top to bottom. $\mathrm{v}=$ catsplit $(\mathrm{t}, \mathrm{j})$ takes an array j of rows and returns the splits for the specified rows.

## See Also

classregtree

Purpose Cumulative distribution function for Gaussian mixture distribution

## Syntax <br> $y=\operatorname{cdf}(o b j, X)$

Description
$y=\operatorname{cdf}(o b j, X)$ returns a vector $y$ of length $n$ containing the values of the cumulative distribution function (cdf) for the gmdistribution object obj, evaluated at the $n$-by- $d$ data matrix $X$, where $n$ is the number of observations and $d$ is the dimension of the data. obj is an object created by gmdistribution or fit. $\mathrm{y}(\mathrm{I})$ is the cdf of observation I.

Examples Create a gmdistribution object defining a two-component mixture of bivariate Gaussian distributions:

```
MU = [1 2;-3 -5];
SIGMA = cat(3,[2 0;0 .5],[1 0;0 1]);
p = ones(1,2)/2;
obj = gmdistribution(MU,SIGMA,p);
ezsurf(@(x,y)cdf(obj,[x y]),[-10 10],[-10 10])
```



See Also
gmdistribution | fit | pdf | mvncdf

Purpose
Syntax

## Description

Central composite design

```
dCC \(=\) ccdesign( \(n\) )
[dCC,blocks] = ccdesign(n)
[...] \(=\operatorname{ccdesign(n,'Name',value)~}\)
dCC = ccdesign(n)
[...] = ccdesign(n,'Name',value)
```

$d C C=\operatorname{ccdesign}(n)$ generates a central composite design for $n$ factors. $n$ must be an integer 2 or larger. The output matrix dCC is $m$-by-n, where $m$ is the number of runs in the design. Each row represents one run, with settings for all factors represented in the columns. Factor values are normalized so that the cube points take values between - 1 and 1 .
[dCC,blocks] = ccdesign(n) requests a blocked design. The output blocks is an $m$-by- 1 vector of block numbers for each run. Blocks indicate runs that are to be measured under similar conditions to minimize the effect of inter-block differences on the parameter estimates.
[...] = ccdesign( n , 'Name', value) specifies one or more optional name/value pairs for the design. Valid parameters and their values are listed in the following table. Specify Name in single quotes.

| Parameter | Description | Values |
| :--- | :--- | :--- |
| center | Number of <br> center points. | • Integer - Number of center <br> points to include. <br> • 'uniform' — Select number of <br> center points to give uniform <br> precision. |
| - 'orthogonal' — Select |  |  |
| number of center points to give |  |  |
| an orthogonal design. This is |  |  |
| the default. |  |  |


| Parameter | Description | Values |
| :--- | :--- | :--- |
|  | as an exponent <br> of $1 / 2$. | $\bullet 2-1 / 4$ fraction. |
| type | Type of CCD. | $\bullet$ ' circumscribed ' - <br> Circumscribed (CCC). This is <br> the default. |
| blocksize | 'inscribed' — Inscribed <br> (CCI). <br> Maximum <br> number of <br> points per block. | Integer. The default is Inf. |

## Examples

The following creates a 2 -factor CCC:

```
dCC = ccdesign(2,'type','circumscribed')
dCC =
        -1.0000 -1.0000
        -1.0000 1.0000
        1.0000 -1.0000
        1.0000 1.0000
        -1.4142 0
        1.4142 0
            0 -1.4142
            0 1.4142
            0
            0
            0
            0 0
            0 0
            0 0
            0 0
            0
```

The center point is run 8 times to reduce the correlations among the coefficient estimates.

Visualize the design as follows:

```
plot(dCC(:,1),dCC(:,2),'ro','MarkerFaceColor','b')
X = [1 -1 -1 -1; 1 1 1 - 1];
Y = [-1 -1 1 -1; 1 -1 1 1];
line(X,Y,'Color','b')
axis square equal
```



## See Also

bbdesign

## Purpose Cumulative distribution functions

Syntax $\quad Y=\operatorname{cdf}\left(\right.$ ' $^{\prime}$ name $\left.{ }^{\prime}, X, A\right)$
$Y=c d f(' n a m e ', X, A, B)$
$Y=\operatorname{cdf}('$ name', $X, A, B, C)$

## Description

$Y=\operatorname{cdf}($ ' name', $X, A)$ computes the cumulative distribution function for the one-parameter family of distributions specified by name. A contains parameter values for the distribution. The cumulative distribution function is evaluated at the values in $X$ and its values are returned in Y .

If $X$ and $A$ are arrays, they must be the same size. If $X$ is a scalar, it is expanded to a constant matrix the same size as $A$. If A is a scalar, it is expanded to a constant matrix the same size as $X$.
$Y$ is the common size of $X$ and $A$ after any necessary scalar expansion.
$Y=\operatorname{cdf}($ ' name' $, X, A, B)$ computes the cumulative distribution function for two-parameter families of distributions, where parameter values are given in $A$ and $B$.

If $X, A$, and $B$ are arrays, they must be the same size. If $X$ is a scalar, it is expanded to a constant matrix the same size as $A$ and $B$. If either $A$ or $B$ are scalars, they are expanded to constant matrices the same size as $X$.
$Y$ is the common size of $X, A$, and $B$ after any necessary scalar expansion.
$Y=\operatorname{cdf}($ ' $n a m e ', X, A, B, C)$ computes the cumulative distribution function for three-parameter families of distributions, where parameter values are given in $A, B$, and $C$.

If $X, A, B$, and $C$ are arrays, they must be the same size. If $X$ is a scalar, it is expanded to a constant matrix the same size as $A, B$, and $C$. If any of $A, B$ or $C$ are scalars, they are expanded to constant matrices the same size as $X$.
$Y$ is the common size of $X, A, B$, and $C$ after any necessary scalar expansion.
Acceptable strings for name (specified in single quotes) are:

| name | Distribution | Input <br> Parameter A | Input <br> Parameter B | Input <br> Parameter C |
| :--- | :--- | :--- | :--- | :--- |
| beta or Beta | "Beta <br> Distribution" <br> on page B-4 | a | b | - |
| bino or <br> Binomial | "Binomial <br> Distribution" <br> on page B-7 | n: number of <br> trials | p: probability of <br> success for each <br> trial | - |
| chi2 or <br> Chisquare | "Chi-Square <br> Distribution" <br> on page B-12 | $v:$ degrees of <br> freedom | - | - |
| exp or <br> Exponential | "Exponential <br> Distribution" on <br> page B-16 | $\mu:$ mean | - | - |
| ev or Extreme <br> Value | "Extreme Value <br> Distribution" on <br> page B-19 | $\mu:$ location <br> parameter | $\sigma:$ scale <br> parameter | - |
| f or F | "F Distribution" <br> on page B-25 | $v 1:$ numerator <br> degrees of <br> freedom | $v 2:$ denominator <br> degrees of <br> freedom | - |
| gam or Gamma | "Gamma <br> Distribution" <br> on page B-27 | a: shape <br> parameter | b: scale <br> parameter | - |
| "Generalized <br> Extreme Value <br> Distribution" on <br> page B-32 | k: shape <br> parameter | $\sigma:$ scale <br> parameter | $\mu:$ location <br> parameter |  |
| "Generalized <br> Pareto <br> Distribution" <br> on page B-37 | k: tail <br> index (shape) <br> parameter | $\sigma:$ scale <br> parameter | $\mu:$ threshold <br> (location) <br> parameter |  |
| Extreme Value | - |  |  |  |
| generalized <br> Pareto | ored | - | - |  |


| name | Distribution | Input <br> Parameter A | Input <br> Parameter B | Input <br> Parameter C |
| :--- | :--- | :--- | :--- | :--- |
| geo or <br> Geometric | "Geometric <br> Distribution" <br> on page B-41 | p: probability <br> parameter | - | - |
| hyge or <br> Hypergeometric | "Hypergeometric <br> Distribution" on <br> page B-43 | M: size of the <br> population | K: number of <br> items with <br> the desired <br> characteristic in <br> the population | n: number of <br> samples drawn |
| logn or <br> Lognormal | "Lognormal <br> Distribution" <br> on page B-51 | $\mu$ | $\sigma$ | - |
| nbin or <br> Negative <br> Binomial | "Negative <br> Binomial <br> Distribution" <br> on page B-72 | r: number of <br> successes | p: probability <br> of success in a <br> single trial | - |
| ncf or <br> Noncentral F | "Noncentral F <br> Distribution" on <br> page B-78 | $v 1:$ numerator <br> degrees of <br> freedom | $v 2:$ denominator <br> degrees of <br> freedom | f: noncentrality <br> parameter |
| nct or <br> Noncentral t | "Noncentral t <br> Distribution" on <br> page B-80 | $v:$ degrees of <br> freedom | $\delta:$ noncentrality <br> parameter | - |
| ncx2 or <br> Noncentral <br> Chi-square | "Noncentral <br> Chi-Square <br> Distribution" <br> on page B-76 | $v:$ degrees of <br> freedom | $\delta:$ noncentrality <br> parameter | - |
| norm or Normal | "Normal <br> Distribution" <br> on page B-83 | $\mu:$ mean | $\sigma:$ standard <br> deviation | - |


| name | Distribution | Input <br> Parameter A | Input <br> Parameter B | Input <br> Parameter C |
| :--- | :--- | :--- | :--- | :--- |
| poiss or <br> Poisson | "Poisson <br> Distribution" <br> on page B-89 | $\lambda:$ mean | - | - |
| rayl or <br> Rayleigh | "Rayleigh <br> Distribution" <br> on page B-91 | b: scale <br> parameter | - | - |
| t or T | "Student's t <br> Distribution" <br> on page B-95 | $v:$ degrees of <br> freedom | - | - |
| unif or Uniform | "Uniform <br> Distribution <br> (Continuous)" <br> on page B-99 | a: lower <br> endpoint <br> (minimum) | end upper <br> (maximum) | - |
| unid or <br> Discrete <br> Uniform | "Uniform <br> Distribution <br> (Discrete)" on <br> page B-101 | N: maximum <br> observable <br> value | - | - |
| wbl or Weibull | "Weibull <br> Distribution" <br> on page B-103 | a: scale <br> parameter | b: shape <br> parameter | - |

Examples Compute the cdf of the normal distribution with mean 0 and standard deviation 1 at inputs $-2,-1,0,1,2$ :

```
p1 = cdf('Normal',-2:2,0,1)
p1 =
\[
\begin{array}{lllll}
0.0228 & 0.1587 & 0.5000 & 0.8413 & 0.9772
\end{array}
\]
```

The order of the parameters is the same as for normodf.
Compute the cdfs of Poisson distributions with rate parameters $0,1, \ldots$, 4 at inputs $1,2, \ldots, 5$, respectively:

```
p2 = cdf('Poisson',0:4,1:5)
p2 =
    0.3679 0.4060 0.4232 0.4335 0.4405
```

The order of the parameters is the same as for poisscdf.

## See Also <br> pdf | icdf

## piecewisedistribution.cdf

Purpose Cumulative distribution function for piecewise distribution

$$
\text { Syntax } \quad P=\operatorname{cdf}(o b j, x)
$$

Description $\quad P=c d f(o b j, X)$ returns an array $P$ of values of the cumulative distribution function for the piecewise distribution object obj, evaluated at the values in the array $X$.

Examples $\quad$ Fit Pareto tails to a $t$ distribution at cumulative probabilities 0.1 and 0.9:

```
    t = trnd(3,100,1);
    obj = paretotails(t,0.1,0.9);
    [p,q] = boundary(obj)
    p =
        0.1000
        0.9000
    q =
        -1.7766
        1.8432
    cdf(obj,q)
    ans =
        0.1000
        0.9000
```

See Also paretotails | pdf | icdf

## Purpose

Return cumulative distribution function (CDF) for ProbDist object
Syntax $\quad Y=\operatorname{cdf}(P D, X)$
Description
$Y=\operatorname{cdf}(P D, X)$ returns $Y$, an array containing the cumulative distribution function (CDF) for the ProbDist object PD, evaluated at values in $X$.

## Input <br> Arguments

$x$

Y
Output $\quad Y$
Arguments

An object of the class ProbDistUnivParam or ProbDistUnivKernel.

A numeric array of values where you want to evaluate the CDF.

An array containing the cumulative distribution function (CDF) for the ProbDist object $P D$.

See Also cdf
See Also cdf

Purpose Empirical cumulative distribution function plot

```
Syntax cdfplot(X)
h = cdfplot(X)
[h,stats] = cdfplot(X)
```


## Description

## Examples

The following example compares the empirical cdf for a sample from an extreme value distribution with a plot of the cdf for the sampling distribution. In practice, the sampling distribution would be unknown, and would be chosen to match the empirical cdf.

```
y = evrnd(0,3,100,1);
cdfplot(y)
hold on
x = -20:0.1:10;
f = evcdf(x,0,3);
plot(x,f,'m')
legend('Empirical','Theoretical','Location','NW')
```



See Also
ecdf

Purpose Convert categorical array to cell array of strings

## Syntax <br> $B=$ cellstr(A)

Description $B=$ cellstr $(A)$ converts the categorical array $A$ to a cell array of strings. Each element of B contains the categorical level label for the corresponding element of A.

See Also char | getlabels

## Purpose Create cell array of strings from dataset array

Syntax
$\mathrm{B}=\mathrm{cellstr}(\mathrm{A})$
$\mathrm{B}=\mathrm{cellstr}(\mathrm{A}, \mathrm{VARS})$

Description
$\mathrm{B}=$ cellstr(A) returns the contents of the dataset A , converted to a cell array of strings. The variables in the dataset must support the conversion and must have compatible sizes.
$B=$ cellstr(A,VARS) returns the contents of the dataset variables specified by VARS. VARS is a positive integer, a vector of positive integers, a variable name, a cell array containing one or more variable names, or a logical vector.

See Also<br>dataset.double | dataset.replacedata

Purpose $\quad$ Convert categorical array to character array

## Syntax <br> $B=\operatorname{char}(A)$

Description $\quad B=\operatorname{char}(A)$ converts the categorical array $A$ to a 2-D character matrix. char does not preserve the shape of $A$. $B$ contains numel ( $A$ ) rows, and each row of $B$ contains the categorical level label for the corresponding element of $A(:)$.

See Also cellstr | getlabels

## Purpose <br> Chi-square cumulative distribution function

## Syntax <br> $P=\operatorname{chi2cdf}(X, V)$

Description
$P=\operatorname{chi2cdf}(X, V)$ computes the chi-square cdf at each of the values in $X$ using the corresponding degrees of freedom in $V$. $X$ and $V$ can be vectors, matrices, or multidimensional arrays that have the same size. A scalar input is expanded to a constant array with the same dimensions as the other input.

The degrees of freedom parameters in $V$ must be positive integers, and the values in X must lie on the interval [ 0 Inf].

The $x^{2} \operatorname{cdf}$ for a given value $x$ and degrees-of-freedom $v$ is

$$
p=F(x \mid v)=\int_{0}^{x} \frac{t^{(v-2) / 2} e^{-t / 2}}{2^{v / 2} \Gamma(v / 2)} d t
$$

where $\Gamma(\cdot)$ is the Gamma function.
The chi-square density function with $v$ degrees-of-freedom is the same as the gamma density function with parameters $v / 2$ and 2 .

## Examples

## See Also

cdf | chi2pdf | chi2inv | chi2stat | chi2rnd
How To

- "Chi-Square Distribution" on page B-12


## Purpose Chi-square goodness-of-fit test

Syntax

```
h = chi2gof(x)
[h,p] = chi2gof(...)
[h,p,stats] = chi2gof(...)
[...] = chi2gof(X,'Name',value)
```


## Description

$h=c h i 2 g o f(x)$ performs a chi-square goodness-of-fit test of the default null hypothesis that the data in vector x are a random sample from a normal distribution with mean and variance estimated from $x$, against the alternative that the data are not normally distributed with the estimated mean and variance. The result $h$ is 1 if the null hypothesis can be rejected at the $5 \%$ significance level. The result h is 0 if the null hypothesis cannot be rejected at the $5 \%$ significance level.
The null distribution can be changed from a normal distribution to an arbitrary discrete or continuous distribution. See the syntax for specifying optional argument name/value pairs below.

The test is performed by grouping the data into bins, calculating the observed and expected counts for those bins, and computing the chi-square test statistic

$$
\chi^{2}=\sum_{i=1}^{N}\left(O_{i}-E_{i}\right)^{2} / E_{i}
$$

where $O_{i}$ are the observed counts and $E_{i}$ are the expected counts. The statistic has an approximate chi-square distribution when the counts are sufficiently large. Bins in either tail with an expected count less than 5 are pooled with neighboring bins until the count in each extreme bin is at least 5 . If bins remain in the interior with counts less than 5 , chi2gof displays a warning. In this case, you should use fewer bins, or provide bin centers or edges, to increase the expected counts in all bins. (See the syntax for specifying optional argument name/value pairs below.) chi2gof sets the number of bins, nbins, to 10 by default, and compares the test statistic to a chi-square distribution with nbins - 3 degrees of freedom to take into account the two estimated parameters.
$[\mathrm{h}, \mathrm{p}]=$ chi2gof(...) also returns the $p$ value of the test, p . The $p$ value is the probability, under assumption of the null hypothesis, of observing the given statistic or one more extreme.
[h,p,stats] = chi2gof(...) also returns a structure stats with the following fields:

- chi2stat - The chi-square statistic
- df - Degrees of freedom
- edges - Vector of bin edges after pooling
- 0 - Observed count in each bin
- E - Expected count in each bin
[...] = chi2gof(X, 'Name', value) specifies one or more optional argument name/value pairs chosen from the following lists. Argument names are case insensitive and partial matches are allowed. Specify Name in single quotes.
The following name/value pairs control the initial binning of the data before pooling. You should not specify more than one of these options.
- nbins - The number of bins to use. Default is 10 .
- ctrs - A vector of bin centers
- edges - A vector of bin edges

The following name/value pairs determine the null distribution for the test. Do not specify both cdf and expected.

- cdf - A fully specified cumulative distribution function. This can be a function name, a function handle, or a ProbDist object of the ProbDistUnivParam class or ProbDistUnivKernel class. When 'cdf' is a function name or handle, the distribution function must take $x$ as its only argument. Alternately, you can provide a cell array whose first element is a function name or handle, and whose later


## chi2gof

elements are parameter values, one per cell. The function must take $x$ as its first argument, and other parameters as later arguments.

- expected - A vector with one element per bin specifying the expected counts for each bin.
- nparams - The number of estimated parameters; used to adjust the degrees of freedom to be nbins - 1 - nparams, where nbins is the number of bins.

If your cdf or expected input depends on estimated parameters, you should use nparams to ensure that the degrees of freedom for the test is correct. If cdfis a cell array, the default value of nparams is the number of parameters in the array; otherwise the default is 0 .
The following name/value pairs control other aspects of the test.

- emin - The minimum allowed expected value for a bin; any bin in either tail having an expected value less than this amount is pooled with a neighboring bin. Use the value 0 to prevent pooling. The default is 5 .
- frequency - A vector the same length as $x$ containing the frequency of the corresponding xvalues
- alpha - Significance level for the test. The default is 0.05 .


## Examples Example 1

Equivalent ways to test against an unspecified normal distribution with estimated parameters:

```
x = normrnd(50,5,100,1);
[h,p] = chi2gof(x)
h =
    0
p =
    0.7532
```

```
[h,p] = chi2gof(x,'cdf',@(z)normcdf(z,mean(x),std(x)),'nparams',2)
h =
    0
p =
    0.7532
[h,p] = chi2gof(x,'cdf',{@normcdf,mean(x),std(x)})
h =
    0
p =
    0.7532
```


## Example 2

Test against the standard normal:

```
x = randn(100,1);
[h,p] = chi2gof(x,'cdf',@normcdf)
h =
    0
p =
    0.9443
```


## Example 3

Test against the standard uniform:

```
x = rand(100,1);
n = length(x);
edges = linspace(0,1,11);
expectedCounts = n * diff(edges);
[h,p,st] = chi2gof(x,'edges',edges,...
    'expected',expectedCounts)
h =
            0
p =
    0.3191
```

```
st =
    chi2stat: 10.4000
        df: 9
        edges: [1\times11 double]
            0: [[\begin{array}{llllllllllll}{6}&{11}&{4}&{12}&{15}&{8}&{14}&{9}&{11}&{10}\end{array}]
            E: [1\times10 double]
```


## Example 4

Test against the Poisson distribution by specifying observed and expected counts:

```
bins = 0:5;
obsCounts = [ll 16 10 12 4 2];
n = sum(obsCounts);
lambdaHat = sum(bins.*obsCounts)/n;
expCounts = n*poisspdf(bins,lambdaHat);
[h,p,st] = chi2gof(bins,'ctrs',bins,...
                                    'frequency',obsCounts, ...
                                    'expected', expCounts,...
                                    'nparams',1)
h =
            0
p =
    0.4654
st =
    chi2stat: 2.5550
            df: 3
            edges: [1\times6 double]
                0: [[\begin{array}{llllll}{6}&{16}&{10}&{12}&{6}\end{array}]
                E: [ll.0429 13.8041 13.5280 8.8383 6.0284]
```

crosstab | lillietest | kstest | chi2cdf | chi2pdf | chi2inv |
chi2stat | chi2rnd

How To

- "Chi-Square Distribution" on page B-12


## Purpose

Chi-square inverse cumulative distribution function

## Syntax

$X=\operatorname{chi2inv}(P, V)$
$X=$ chi2inv $(P, V)$ computes the inverse of the chi-square cdf with degrees of freedom specified by $V$ for the corresponding probabilities in P. P and V can be vectors, matrices, or multidimensional arrays that have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs.

The degrees of freedom parameters in $V$ must be positive integers, and the values in $P$ must lie in the interval [01].

The inverse chi-square cdf for a given probability $p$ and v degrees of freedom is

$$
x=F^{-1}(p \mid v)=\{x: F(x \mid v)=p\}
$$

where

$$
p=F(x \mid v)=\int_{0}^{x} \frac{t^{(v-2) / 2} e^{-t / 2}}{2^{v / 2} \Gamma(v / 2)} d t
$$

and $\Gamma(\cdot)$ is the Gamma function. Each element of output X is the value whose cumulative probability under the chi-square cdf defined by the corresponding degrees of freedom parameter in $V$ is specified by the corresponding value in P .

## Examples

Find a value that exceeds $95 \%$ of the samples from a chi-square distribution with 10 degrees of freedom.

```
x = chi2inv(0.95,10)
x =
```

    18.3070
    You would observe values greater than 18.3 only $5 \%$ of the time by chance.

See Also icdf | chi2cdf | chi2pdf | chi2stat | chi2rnd<br>How To . "Chi-Square Distribution" on page B-12

## Purpose

Chi-square probability density function

## Syntax

Description

## Examples

See Also
$Y=\operatorname{chi} 2 p d f(X, V)$ constant array with the same dimensions as the other input. the values in X must lie on the interval [ 0 Inf].

$$
y=f(x \mid v)=\frac{x^{(v-2) / 2} e^{-x / 2}}{2^{v / 2} \Gamma(v / 2)}
$$

where $\Gamma(\cdot)$ is the Gamma function. density function with parameters $v / 2$ and 2 ).

```
nu = 1:6;
\(x=n u ;\)
\(y=\) chi2pdf(x, nu)
\(y=\)
    \(\begin{array}{llllll}0.2420 & 0.1839 & 0.1542 & 0.1353 & 0.1220 & 0.1120\end{array}\)
``` density of the mean falls as nu increases.
```

pdf | chi2cdf | chi2inv | chi2stat | chi2rnd

```
\(Y=\) chi2pdf \((X, V)\) computes the chi-square pdf at each of the values in \(X\) using the corresponding degrees of freedom in \(V . X\) and \(V\) can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of the output Y . A scalar input is expanded to a

The degrees of freedom parameters in \(V\) must be positive integers, and

The chi-square pdf for a given value \(x\) and \(v\) degrees of freedom is

If \(x\) is standard normal, then \(x^{2}\) is distributed chi-square with one degree of freedom. If \(x_{1}, x_{2}, \ldots, x_{\mathrm{n}}\) are \(n\) independent standard normal observations, then the sum of the squares of the \(x\) 's is distributed chi-square with \(n\) degrees of freedom (and is equivalent to the gamma

The mean of the chi-square distribution is the value of the degrees of freedom parameter, nu. The above example shows that the probability

\author{
How To \\ - "Chi-Square Distribution" on page B-12
}

\section*{Purpose \\ Syntax \\ Description}

Chi-square random numbers
\(\mathrm{R}=\mathrm{chi2} \mathrm{rnd}(\mathrm{V})\)
R = chi2rnd(V,m,n,...)
\(R=\operatorname{chi2rnd}(V,[m, n, \ldots])\)

\section*{Examples}
```

r = chi2rnd(1:6)

```
\(r=\)
    \(\begin{array}{llllll}0.0037 & 3.0377 & 7.8142 & 0.9021 & 3.2019 & 9.0729\end{array}\)
\(r=\operatorname{chi} 2 r n d(6,[16])\)
\(r=\)
    \(\begin{array}{llllll}6.5249 & 2.6226 & 12.2497 & 3.0388 & 6.3133 & 5.0388\end{array}\)
\(r=\operatorname{chi2rnd}(1: 6,1,6)\)
\(r=\)
    \(\begin{array}{llllll}0.7638 & 6.0955 & 0.8273 & 3.2506 & 1.5469 & 10.9197\end{array}\)

\section*{See Also}
random | chi2cdf | chi2pdf | chi2inv | chi2stat
How To
- "Chi-Square Distribution" on page B-12

Purpose Chi-square mean and variance
Syntax \(\quad[M, V]=\) chi2stat (NU)
Description
\([\mathrm{M}, \mathrm{V}]=\) chi2stat(NU) returns the mean of and variance for the chi-square distribution with degrees of freedom parameters specified by NU.

The mean of the chi-square distribution is \(v\), the degrees of freedom parameter, and the variance is 2 v .

Examples


\author{
See Also chi2cdf | chi2pdf | chi2inv | chi2rnd \\ How To . "Chi-Square Distribution" on page B-12
}

\section*{classregtree.children}

Purpose Child nodes
Syntax
C = children(t)
C = children(t,nodes)

Description
\(\mathrm{C}=\) children(t) returns an \(n\)-by- 2 array C containing the numbers of the child nodes for each node in the tree \(t\), where \(n\) is the number of nodes. Leaf nodes have child node 0 .

C = children(t,nodes) takes a vector nodes of node numbers and returns the children for the specified nodes.

\section*{Examples \\ Create a classification tree for Fisher's iris data:}
```

load fisheriris;
t = classregtree(meas,species,...
'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
view(t)

```
\begin{tabular}{|c|c|c|c|c|c|c|c|c|}
\hline & & & & & & & & - \\
\hline Click to display: & Identity & \(\checkmark\) & Magnification: & 100\% & \(\checkmark\) & Pruning level: & 0 of 4 & \\
\hline
\end{tabular}

\[
\begin{array}{cc}
C= & \text { children }(t) \\
C= & \\
2 & 3 \\
0 & 0 \\
4 & 5 \\
6 & 7 \\
0 & 0 \\
8 & 9 \\
0 & 0 \\
0 & 0
\end{array}
\]

\section*{classregtree.children}
\(0 \quad 0\)

\author{
References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.
}

See Also
classregtree | numnodes | parent

\section*{Purpose Cholesky-like covariance decomposition}
Syntax \(\quad\)\begin{tabular}{ll}
\(T=\operatorname{chol} \operatorname{cov}(\) SIGMA \()\) \\
{\([T\), num \(]\)} & \(=\operatorname{cholcov}(\) SIGMA \()\) \\
{\([T\), num \(]\)} & \(=\operatorname{cholcov}(\) SIGMA, 0\()\)
\end{tabular}

\section*{Description}

\section*{Examples}

The following 4 -by- 4 covariance matrix is rank-deficient:


3

Use cholcov to factor C1:
```

T = cholcov(C1)
T =
-0.2113 0.7887 -0.5774 0
0.7887 -0.2113 -0.5774 0
1.1547 1.1547 1.1547 1.7321
C2 = T'*T
C2 =

| 2.0000 | 1.0000 | 1.0000 | 2.0000 |
| :--- | :--- | :--- | :--- |
| 1.0000 | 2.0000 | 1.0000 | 2.0000 |
| 1.0000 | 1.0000 | 2.0000 | 2.0000 |
| 2.0000 | 2.0000 | 2.0000 | 3.0000 |

```

Use T to generate random data with the specified covariance:
```

C3 = cov(randn(1e6,3)*T)
C3 =
1.9973 0.9982 0.9995 1.9975
0.9982 1.9962 0.9969 1.9956
0.9995 0.9969 1.9980 1.9972
1.9975 1.9956 1.9972 2.9951

```

See Also
chol | cov

\section*{Purpose Shift categorical array circularly}
Syntax
B = circshift(A,shiftsize)

Description
\(B=\operatorname{circshift}(A, s h i f t s i z e)\) circularly shifts the values in the categorical array A by shiftsize elements. shiftsize is a vector of integer scalars where the \(n\)-th element specifies the shift amount for the \(n\)-th dimension of array \(A\). If an element in shiftsize is positive, the values of A are shifted down (or to the right). If it is negative, the values of A are shifted up (or to the left).

See Also permute | shiftdim

\section*{NaiveBayes.ClsNonEmpty property}

Purpose Flag for non-empty classes
Description The CIsNonEmpty property is a logical vector of length NClasses specifying which classes are not empty. When the grouping variable is categorical, it may contain categorical levels that don't appear in the elements of the grouping variable. Those levels are empty and NaiveBayes ignores them for the purposes of training the classifier.

\section*{Purpose Class counts}
```

Syntax
P = classcount(t)
P = classcount(t,nodes)

```

Description
\(P=c l a s s c o u n t(t)\) returns an \(n\)-by- \(m\) array \(P\) of class counts for the nodes in the classification tree t , where \(n\) is the number of nodes and \(m\) is the number of classes. For any node number i, the class counts \(P(i,:)\) are counts of observations (from the data used in fitting the tree) from each class satisfying the conditions for node i.

P = classcount( t , nodes) takes a vector nodes of node numbers and returns the class counts for the specified nodes.

\section*{Examples}

Create a classification tree for Fisher's iris data:
```

load fisheriris;
t = classregtree(meas,species,...
'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
view(t)

```


\begin{tabular}{|c|c|c|}
\hline \multicolumn{3}{|l|}{\(\mathrm{P}=\) classcount(t)} \\
\hline \(\mathrm{P}=\) & & \\
\hline 50 & 50 & 50 \\
\hline 50 & 0 & 0 \\
\hline 0 & 50 & 50 \\
\hline 0 & 49 & 5 \\
\hline 0 & 1 & 45 \\
\hline 0 & 47 & 1 \\
\hline 0 & 2 & 4 \\
\hline 0 & 47 & 0 \\
\hline
\end{tabular}
\(\begin{array}{lll}0 & 0 & 1\end{array}\)
References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

\section*{See Also}

\section*{ClassificationBaggedEnsemble}
Purpose Classification ensemble grown by resampling
Description
ClassificationBaggedEnsemble combines a set of trained weak learner models and data on which these learners were trained. It can predict ensemble response for new data by aggregating predictions from its weak learners.

\section*{Construction}

\section*{ens =}
fitensemble(X,Y,'bag', nlearn,learners,'type','classification') creates a bagged classification ensemble. For syntax details, see the fitensemble reference page.

\section*{Properties}

\section*{CategoricalPredictors}
List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of X .
CombineWeights
String describing how ens combines weak learner weights, either 'WeightedSum' or 'WeightedAverage'.

\section*{FitInfo}
Numeric array of fit information. The FitInfoDescription property describes the content of this array.

\section*{FitInfoDescription}
String describing the meaning of the FitInfo array.

\section*{FResample}
Numeric scalar between 0 and 1. FResample is the fraction of training data fitensemble resampled at random for every weak learner when constructing the ensemble.

\section*{Method}
String describing the method that creates ens.

\footnotetext{
ModelParams
}

\section*{ClassificationBaggedEnsemble}

Parameters used in training ens.

\section*{NTrained}

Number of trained weak learners in ens, a scalar.

\section*{PredictorNames}

Cell array of names for the predictor variables, in the order in which they appear in \(X\).

\section*{ReasonForTermination}

String describing the reason fitensemble stopped adding weak learners to the ensemble.

\section*{Replace}

Logical value indicating if the ensemble was trained with replacement (true) or without replacement (false).

\section*{ResponseName}

String with the name of the response variable \(Y\).

\section*{ScoreTransform}

Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none' means @(x)x. For a list of built-in transformation functions and the syntax of custom transformation functions, see ClassificationTree.fit.

Add or change a ScoreTransform function by dot addressing:
ens.ScoreTransform = 'function'
or
ens.ScoreTransform = @function

\section*{Trained}

Trained learners, a cell array of compact classification models.

\section*{ClassificationBaggedEnsemble}

\section*{TrainedWeights}

Numeric vector of trained weights for the weak learners in ens. TrainedWeights has \(T\) elements, where \(T\) is the number of weak learners in learners.

\section*{UseObsForLearner}

Logical matrix of size N -by-NTrained, where N is the number of observations in the training data and NTrained is the number of trained weak learners. UseObsForLearner ( \(I, J\) ) is true if observation I was used for training learner \(J\), and is false otherwise.
w
Scaled weights, a vector with length n , the number of rows in X . The sum of the elements of \(W\) is 1 .

X
Matrix of predictor values that trained the ensemble. Each column of \(X\) represents one variable, and each row represents one observation.

Y
Numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. Each row of \(Y\) represents the classification of the corresponding row of \(X\).

\section*{Methods}
oobEdge
oobLoss
oobMargin
oobPredict

Out-of-bag classification edge
Out-of-bag classification error
Out-of-bag classification margins
Predict out-of-bag response of ensemble

\section*{ClassificationBaggedEnsemble}

\section*{Inherited Methods}
\begin{tabular}{ll} 
compact & Compact classification ensemble \\
crossval & \begin{tabular}{l} 
Cross validate ensemble \\
resubEdge \\
resssification edge by \\
resubstitution \\
Classification error by \\
resubstitution \\
Classification margins by \\
resubstitution
\end{tabular} \\
resubPredict & \begin{tabular}{l} 
Predict ensemble response by \\
resubstitution
\end{tabular} \\
resume & Resume training ensemble \\
edge & Classification edge \\
loss & Classification error \\
margin & Classification margins \\
predict & Predict classification \\
predictorImportance & Estimates of predictor importance
\end{tabular}

\section*{Copy Semantics \\ Value. To learn how value classes affect copy operations, see Copying Objects in the MATLAB Programming Fundamentals documentation.}

\section*{Examples}

Construct a bagged ensemble for the ionosphere data, and examine its resubstitution loss:
```

load ionosphere
rng(0,'twister') % for reproducibility
ens = fitensemble(X,Y,'bag',100,'Tree',...
'type','classification');
L = resubLoss(ens)

```

\section*{ClassificationBaggedEnsemble}
\(\mathrm{L}=0\)

The ensemble does a perfect job classifying its training data.

\section*{See Also ClassificationEnsemble | fitensemble}

How To . "Ensemble Methods" on page 13-51

\section*{ClassificationDiscriminant}

\section*{Superclasses CompactClassificationDiscriminant \\ Purpose Discriminant analysis classification \\ Description A ClassificationDiscriminant object encapsulates a discriminant analysis classifier, which is a Gaussian mixture model for data generation. A ClassificationDiscriminant object can predict responses for new data using the predict method. The object contains the data used for training, so can compute resubstitution predictions. \\ Construction \\ obj \(=\) ClassificationDiscriminant.fit \((X, Y)\) creates a discriminant classification object based on the input variables (also known as predictors, features, or attributes) \(X\) and output (response) Y. For syntax details, see ClassificationDiscriminant.fit. \\ obj = ClassificationDiscriminant.fit(X,Y,Name, Value) creates a classifier with additional options specified by one or more Name, Value pair arguments. If you use one of the following five options, obj is of class ClassificationPartitionedModel: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Otherwise, obj is of class ClassificationDiscriminant. \\ Input Arguments}

X
Matrix of numeric predictor values. Each column of \(X\) represents one variable, and each row represents one observation.

NaN values in X are considered missing values. Observations with missing values for \(X\) are not used in the fit.

Numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. Each row of \(Y\) represents the classification of the corresponding row of \(X\). NaN values in \(Y\) are considered missing values. Observations with missing values for \(Y\) are not used in the fit.

\section*{ClassificationDiscriminant}

\section*{Properties}

BetweenSigma
p -by-p matrix, the between-class covariance, where p is the number of predictors.

CategoricalPredictors
List of categorical predictors, always empty ([ ]) for discriminant analysis.

\section*{ClassNames}

List of the elements in the training data \(Y\) with duplicates removed. ClassNames can be a numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. ClassNames has the same data type as the data in the argument Y .

\section*{Coeffs}
\(k\)-by- k structure of coefficient matrices, where k is the number of classes. Coeffs(i,j) contains coefficients of the linear or quadratic boundaries between classes i and j. Fields in Coeffs(i,j):
- DiscrimType
- Class1 - ClassNames(i)
- Class2 - ClassNames(j)
- Const - A scalar
- Linear - A vector with \(p\) components, where \(p\) is the number of columns in \(X\)
- Quadratic - p-by-p matrix, exists for quadratic DiscrimType

The equation of the boundary between class \(i\) and class \(j\) is
\[
\text { Const + Linear * } x+x^{\prime} \text { * Quadratic * } x=0
\]
where \(x\) is a column vector of length \(p\).

\section*{ClassificationDiscriminant}

If ClassificationDiscriminant.fit had the FillCoeffs name-value pair set to 'off' when constructing the classifier, Coeffs is empty ([]).

\section*{Cost}

Square matrix, where \(\operatorname{Cost}(i, j)\) is the cost of classifying a point into class \(j\) if its true class is \(i\). Cost is \(K\)-by- \(K\), where \(K\) is the number of classes.

Change a Cost matrix using dot addressing:
```

obj.Cost = costMatrix

```

DiscrimType
String specifying the discriminant type. One of:
- 'linear'
- 'quadratic'
- 'diagLinear'
- 'diagQuadratic'
- 'pseudoLinear'
- 'pseudoQuadratic'

\section*{LogDetSigma}

Logarithm of the determinant of the within-class covariance matrix. The type of LogDetSigma depends on the discriminant type:
- Scalar for linear discriminant analysis
- Vector of length K for quadratic discriminant analysis, where K is the number of classes

\footnotetext{
ModelParams
}

\section*{ClassificationDiscriminant}

Parameters used in training obj.
Mu
Matrix of class means of size K-by-p, where \(K\) is the number of classes, and \(p\) is the number of predictors. Each row of Mu represents the mean of the multivariate normal distribution of the corresponding class. The class indices are in the ClassNames attribute.

\section*{NObservations}

Number of observations in the training data, a numeric scalar. NObservations can be less than the number of rows of input data \(X\) when there are missing values in \(X\) or response \(Y\).

\section*{PredictorNames}

Cell array of names for the predictor variables, in the order in which they appear in the training data \(X\).

Prior
Prior probabilities for each class. Prior is a numeric vector whose entries relate to the corresponding ClassNames property.
Add or change a Prior vector using dot addressing:
obj.Prior = priorVector

\section*{ResponseName}

String describing the response variable Y .
ScoreTransform
Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none' means @(x)x. For a list of built-in transformation functions and the syntax of custom transformation functions, see ClassificationDiscriminant.fit.

Add or change a ScoreTransform function by dot addressing:

\section*{ClassificationDiscriminant}
```

    cobj.ScoreTransform = 'function'
    or
cobj.ScoreTransform = @function

```

\section*{Sigma}

Within-class covariance matrix or matrices. The dimensions depend on DiscrimType:
- 'linear' (default) - Matrix of size \(p-b y-p\), where \(p\) is the number of predictors
- 'quadratic' - Array of size p-by-p-by-K, where K is the number of classes
- 'diagLinear' - Row vector of length \(p\)
- 'diagQuadratic' — Array of size 1-by-p-by-K
- 'pseudoLinear' - Matrix of size p-by-p
- 'pseudoQuadratic' - Array of size p-by-p-by-K

W
Scaled weights, a vector with length \(n\), the number of rows in \(X\).

X
Matrix of predictor values. Each column of X represents one predictor (variable), and each row represents one observation.

\section*{Xcentered}
\(X\) data with class means subtracted. If \(Y(i)\) is of class \(j\),
```

Xcentered(i,:) = X(i,:) - Mu(j,:),

```
where \(M u\) is the class mean property.

\section*{ClassificationDiscriminant}

Numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. Each row of \(Y\) represents the classification of the corresponding row of \(X\).

\section*{Methods}
\begin{tabular}{ll} 
compact & \begin{tabular}{l} 
Compact discriminant analysis \\
classifier
\end{tabular} \\
crossval & \begin{tabular}{l} 
Cross-validated discriminant \\
analysis classifier
\end{tabular} \\
fit & \begin{tabular}{l} 
Fit discriminant analysis \\
classifier
\end{tabular} \\
make & \begin{tabular}{l} 
Construct discriminant analysis \\
classifier from parameters
\end{tabular} \\
resubEdge & \begin{tabular}{l} 
Classification edge by \\
resubstitution \\
Classification error by \\
resubstitution
\end{tabular} \\
resubLoss & \begin{tabular}{l} 
Classification margins by \\
resubstitution
\end{tabular} \\
resubMargin & \begin{tabular}{l} 
Predict resubstitution response \\
classifier
\end{tabular} \\
resubPredict & \begin{tabular}{l} 
Classification edge
\end{tabular} \\
Inherited Methods & \begin{tabular}{l} 
Classification error \\
edge
\end{tabular} \\
loss & \begin{tabular}{l} 
Mahalanobis distance to class \\
means
\end{tabular} \\
mahal & Classification margins
\end{tabular}

\section*{ClassificationDiscriminant}

\section*{Definitions Discriminant Classification}

The model for discriminant analysis is:
- Each class (Y) generates data (X) using a multivariate normal distribution. That is, the model assumes X has a Gaussian mixture distribution (gmdistribution).
- For linear discriminant analysis, the model has the same covariance matrix for each class, only the means vary.
- For quadratic discriminant analysis, both means and covariances of each class vary.
predict classifies so as to minimize the expected classification cost:
\[
\hat{y}=\underset{y=1, \ldots, K}{\arg \min } \sum_{k=1}^{K} \hat{P}(k \mid x) C(y \mid k),
\]
where
- \(\hat{y}\) is the predicted classification.
- \(K\) is the number of classes.
- \(\hat{P}(k \mid x)\) is the posterior probability of class \(k\) for observation \(x\).
- \(C(y \mid k)\) is the cost of classifying an observation as \(y\) when its true class is \(k\).

For details, see "How the predict Method Classifies" on page 12-6.
Copy Value. To learn how value classes affect copy operations, see Copying Semantics Objects in the MATLAB Programming Fundamentals documentation.

\section*{ClassificationDiscriminant}
```

Examples Create a discriminant analysis classifier for the Fisher iris data:
load fisheriris
obj = ClassificationDiscriminant.fit(meas,species)
obj =
ClassificationDiscriminant:
PredictorNames: {'x1' 'x2' 'x3' 'x4'}
ResponseName: 'Y'
ClassNames: {'setosa' 'versicolor' 'virginica'}
ScoreTransform: 'none'
NObservations: 150
DiscrimType: 'linear'
Mu: [3\times4 double]
Coeffs: [3x3 struct]
See Also CompactClassificationDiscriminant |
ClassificationDiscriminant.fit
How To . "Discriminant Analysis" on page 12-3

```

\section*{ClassificationEnsemble}

\section*{Superclasses CompactClassificationEnsemble}

\section*{Purpose \\ Ensemble classifier}

Description

\section*{Construction}

\section*{Properties}

ClassificationEnsemble combines a set of trained weak learner models and data on which these learners were trained. It can predict ensemble response for new data by aggregating predictions from its weak learners. It also stores data used for training and can compute resubstitution predictions. It can resume training if desired.
ens = fitensemble( \(\mathrm{X}, \mathrm{Y}\), method, nlearn, learners) returns an ensemble model that can predict responses to data. The ensemble consists of models listed in learners. For more information on the syntax, see the fitensemble function reference page.
ens = fitensemble(X,Y,method,nlearn,learners,Name,Value) returns an ensemble model with additional options specified by one or more Name, Value pair arguments. For more information on the syntax, see the fitensemble function reference page.

\section*{CategoricalPredictors}

List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of \(X\).

\section*{ClassNames}

List of the elements in \(Y\) with duplicates removed. ClassNames can be a numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. ClassNames has the same data type as the data in the argument Y .

\section*{CombineWeights}

String describing how ens combines weak learner weights, either 'WeightedSum' or 'WeightedAverage'.

Cost

\section*{ClassificationEnsemble}

Square matrix where \(\operatorname{Cost}(i, j)\) is the cost of classifying a point into class j if its true class is i .

\section*{FitInfo}

Numeric array of fit information. The FitInfoDescription property describes the content of this array.

\section*{FitInfoDescription}

String describing the meaning of the FitInfo array.

\section*{LearnerNames}

Cell array of strings with names of weak learners in the ensemble. The name of each learner appears just once. For example, if you have an ensemble of 100 trees, LearnerNames is \{'Tree'\}.

\section*{Method}

String describing the method that creates ens.

\section*{ModelParams}

Parameters used in training ens.

\section*{NObservations}

Numeric scalar containing the number of observations in the training data.

\section*{NTrained}

Number of trained weak learners in ens, a scalar.

\section*{PredictorNames}

Cell array of names for the predictor variables, in the order in which they appear in \(X\).

Prior
Prior probabilities for each class. Prior is a numeric vector whose entries relate to the corresponding ClassNames property.

\footnotetext{
ReasonForTermination
}

String describing the reason fitensemble stopped adding weak learners to the ensemble.

\section*{ResponseName}

String with the name of the response variable Y .

\section*{ScoreTransform}

Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none' means @(x)x. For a list of built-in transformation functions and the syntax of custom transformation functions, see ClassificationTree.fit.

Add or change a ScoreTransform function by dot addressing:
```

ens.ScoreTransform = 'function'

```
or
```

ens.ScoreTransform = @function

```

Trained
Trained learners, a cell array of compact classification models.

\section*{TrainedWeights}

Numeric vector of trained weights for the weak learners in ens. TrainedWeights has T elements, where T is the number of weak learners in learners.

W
Scaled weights, a vector with length \(n\), the number of rows in \(X\). The sum of the elements of \(W\) is 1 .

X
Matrix of predictor values that trained the ensemble. Each column of \(X\) represents one variable, and each row represents one observation.

\section*{ClassificationEnsemble}

\section*{Y}

Numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. Each row of \(Y\) represents the classification of the corresponding row of \(X\).

\section*{Methods}
\begin{tabular}{ll} 
compact & Compact classification ensemble \\
crossval & Cross validate ensemble \\
resubEdge & \begin{tabular}{l} 
Classification edge by \\
resubstitution
\end{tabular} \\
resubLoss & \begin{tabular}{l} 
Classification error by \\
resubstitution \\
resubMargin
\end{tabular} \\
Classification margins by \\
resubstitution \\
resumeredict & \begin{tabular}{l} 
Predict ensemble response by \\
resubstitution \\
Resume training ensemble
\end{tabular}
\end{tabular}

\section*{Inherited Methods}
\begin{tabular}{ll} 
edge & Classification edge \\
loss & Classification error \\
margin & Classification margins \\
predict & Predict classification \\
predictorImportance & Estimates of predictor importance
\end{tabular}

Value. To learn how value classes affect copy operations, see Copying Objects in the MATLAB Programming Fundamentals documentation.
Examples Construct a boosted classification ensemble for the ionosphere data,using the AdaBoostM1 method:
```

load ionosphere
ens = fitensemble(X,Y,'AdaBoostM1',100,'Tree')
ens =
classreg.learning.classif.ClassificationEnsemble:
PredictorNames: {1x34 cell}
CategoricalPredictors: []
ResponseName: 'Response'
ClassNames: {'b' 'g'}
ScoreTransform: 'none'
NObservations: 351
NTrained: 100
Method: 'AdaBoostM1'
LearnerNames: {'Tree'}
ReasonForTermination: [1x77 char]
FitInfo: [100x1 double]
FitInfoDescription: [2x83 char]

```
Predict the classification of the mean of X :
```

ypredict = predict(ens,mean(X))
ypredict =
'g'

```

\section*{See Also ClassificationTree | fitensemble | RegressionEnsemble | CompactClassificationEnsemble}
How To
- Chapter 13, "Nonparametric Supervised Learning"

\section*{ClassificationPartitionedEnsemble}

\section*{Purpose Cross-validated classification ensemble \\ Construction \\ Properties \\ ClassificationPartitionedEnsemble is a set of classification ensembles trained on cross-validated folds. Estimate the quality of classification by cross validation using one or more "kfold" methods: kfoldPredict, kfoldLoss, kfoldMargin, kfoldEdge, and kfoldfun. \\ Every "kfold" method uses models trained on in-fold observations to predict response for out-of-fold observations. For example, suppose you cross validate using five folds. In this case, every training fold contains roughly \(4 / 5\) of the data and every test fold contains roughly \(1 / 5\) of the data. The first model stored in Trained\{1\} was trained on \(X\) and \(Y\) with the first \(1 / 5\) excluded, the second model stored in Trained\{2\} was trained on X and Y with the second \(1 / 5\) excluded, and so on. When you call kfoldPredict, it computes predictions for the first \(1 / 5\) of the data using the first model, for the second \(1 / 5\) of data using the second model, and so on. In short, response for every observation is computed by kfoldPredict using the model trained without this observation. \\ cvens = crossval(ens) creates a cross-validated ensemble from ens, a classification ensemble. For syntax details, see the crossval method reference page. \\ cvens = fitensemble(X,Y,method, nlearn, learners, name, value) creates a cross-validated ensemble when name is one of 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. For syntax details, see the fitensemble function reference page. \\ CategoricalPredictors \\ List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of X . \\ ClassNames}

List of the elements in \(Y\) with duplicates removed. ClassNames can be a numeric vector, vector of categorical variables (nominal

\section*{ClassificationPartitionedEnsemble}
or ordinal), logical vector, character array, or cell array of strings. ClassNames has the same data type as the data in the argument \(Y\).

\section*{Combiner}

Cell array of combiners across all folds.

\section*{Cost}

Square matrix, where \(\operatorname{Cost}(i, j)\) is the cost of classifying a point into class \(j\) if its true class is \(i\).

\section*{CrossValidatedModel}

Name of the cross-validated model, a string.

\section*{Kfold}

Number of folds used in a cross-validated ensemble, a positive integer.

\section*{ModelParams}

Object holding parameters of cvens.

\section*{NObservations}

Number of data points used in training the ensemble, a positive integer.

\section*{NTrainedPerFold}

Number of data points used in training each fold of the ensemble, a positive integer.

\section*{Partition}

Partition of class cvpartition used in creating the cross-validated ensemble.

\section*{PredictorNames}

Cell array of names for the predictor variables, in the order in which they appear in \(X\).
```

Prior

```

\section*{ClassificationPartitionedEnsemble}

Prior probabilities for each class. Prior is a numeric vector whose entries relate to the corresponding ClassNames property.

\section*{ResponseName}

Name of the response variable Y , a string.

\section*{ScoreTransform}

Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none' means @(x)x. For a list of built-in transformation functions and the syntax of custom transformation functions, see ClassificationTree.fit.

Add or change a ScoreTransform function by dot addressing:
```

ens.ScoreTransform = 'function'

```
or
```

ens.ScoreTransform = @function

```

\section*{Trainable}

Cell array of ensembles trained on cross-validation folds. Every ensemble is full, meaning it contains its training data and weights. Trained

Cell array of compact ensembles trained on cross-validation folds. W

Scaled weights, a vector with length n , the number of rows in X .
X
A matrix of predictor values. Each column of \(X\) represents one variable, and each row represents one observation.

\section*{Y}

A numeric column vector with the same number of rows as \(X\). Each entry in Y is the response to the data in the corresponding row of X .

\section*{ClassificationPartitionedEnsemble}

\section*{Methods}
\begin{tabular}{ll} 
kfoldEdge & \begin{tabular}{l} 
Classification edge for \\
observations not used for \\
training
\end{tabular} \\
kfoldLoss & \begin{tabular}{l} 
Classification loss for \\
observations not used for \\
training
\end{tabular} \\
resume & \begin{tabular}{l} 
Resume training learners on \\
cross-validation folds
\end{tabular} \\
Inherited Methods & \begin{tabular}{l} 
Classification edge for \\
observations not used for \\
training
\end{tabular} \\
kfoldEdge & \begin{tabular}{l} 
Cross validate function
\end{tabular} \\
kfoldfun & \begin{tabular}{l} 
Classification loss for \\
observations not used for \\
training
\end{tabular} \\
kfoldLoss & \begin{tabular}{l} 
Classification margins for \\
observations not used for training
\end{tabular} \\
kfoldMargin & \begin{tabular}{l} 
Predict response for observations \\
not used for training
\end{tabular} \\
kfoldPredict &
\end{tabular}

\section*{Copy Semantics}

Examples

Value. To learn how value classes affect copy operations, see Copying Objects in the MATLAB Programming Fundamentals documentation.

Evaluate the k-fold cross-validation error for a classification ensemble that models the Fisher iris data:
```

load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
cvens = crossval(ens);
L = kfoldLoss(cvens)

```

\section*{ClassificationPartitionedEnsemble}

\author{
L = \\ 0.0533 \\ \begin{tabular}{ll} 
See Also & \begin{tabular}{l} 
RegressionPartitionedEnsemble | \\
ClassificationPartitionedModel | ClassificationEnsemble
\end{tabular} \\
How To & - Chapter 13, "Nonparametric Supervised Learning"
\end{tabular}
}

\section*{ClassificationPartitionedModel}
Purpose Cross-validated classification model
Description

\section*{Construction}
ClassificationPartitionedModel is a set of classification models trained on cross-validated folds. Estimate the quality of classification by cross validation using one or more "kfold" methods: kfoldPredict, kfoldLoss, kfoldMargin, kfoldEdge, and kfoldfun.
Every "kfold" method uses models trained on in-fold observations to predict response for out-of-fold observations. For example, suppose you cross validate using five folds. In this case, every training fold contains roughly \(4 / 5\) of the data and every test fold contains roughly \(1 / 5\) of the data. The first model stored in Trained\{1\} was trained on \(X\) and \(Y\) with the first \(1 / 5\) excluded, the second model stored in Trained \(\{2\}\) was trained on \(X\) and \(Y\) with the second \(1 / 5\) excluded, and so on. When you call kfoldPredict, it computes predictions for the first \(1 / 5\) of the data using the first model, for the second \(1 / 5\) of data using the second model, and so on. In short, response for every observation is computed by kfoldPredict using the model trained without this observation.
cvmodel \(=\) crossval(obj) creates a cross-validated classification model from a classification model.
cvmodel \(=\) ClassificationTree.fit( \(X, Y\), name, value) or cvmodel \(=\) ClassificationDiscriminant.fit(X,Y, name, value) creates a cross-validated model when name is one of 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. For syntax details, see the ClassificationTree.fit or ClassificationDiscriminant.fit function reference pages.

\section*{Input Arguments}

\section*{obj}
A classification model. obj can be a classification tree constructed using ClassificationTree.fit, or a discriminant analysis classifier constructed using ClassificationDiscriminant.fit.

\section*{ClassificationPartitionedModel}

\section*{Properties}

CategoricalPredictors
List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of \(X\).
ClassNames
List of the elements in \(Y\) with duplicates removed. ClassNames can be a numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. ClassNames has the same data type as the data in the argument Y .

Cost
Square matrix, where \(\operatorname{Cost}(i, j)\) is the cost of classifying a point into class \(j\) if its true class is \(i\).

If cvmodel is a cross-validated ClassificationDiscriminant model, you can change a cost matrix by dot addressing:
```

cvmodel.Cost = costMatrix

```

\section*{CrossValidatedModel}

Name of the cross-validated model, a string.

\section*{Kfold}

Number of folds used in cross-validated model, a positive integer.

\section*{ModelParams}

Object holding parameters of cvmodel.

\section*{Partition}

The partition of class cvpartition used in creating the cross-validated model.

\section*{PredictorNames}

A cell array of names for the predictor variables, in the order in which they appear in \(X\).

\section*{ClassificationPartitionedModel}

Prior
Prior probabilities for each class. Prior is a numeric vector whose entries relate to the corresponding ClassNames property.

If cvmodel is a cross-validated ClassificationDiscriminant model, you can change a Prior vector by dot addressing:
```

cvmodel.Prior = priorVector

```

\section*{ResponseName}

Name of the response variable \(Y\), a string.

\section*{ScoreTransform}

Function handle for transforming scores, or string representing a built-in transformation function.
\begin{tabular}{l|l}
\hline String & Formula \\
\hline 'symmetric' & \(2 x-1\) \\
\hline 'invlogit' & \(\log (x /(1-x))\) \\
\hline 'ismax' & \begin{tabular}{l} 
Set score for the class with the \\
largest score to 1, and scores for all \\
other classes to 0.
\end{tabular} \\
\hline 'symmetricismax' & \begin{tabular}{l} 
Set score for the class with the \\
largest score to 1, and scores for \\
all other classes to -1.
\end{tabular} \\
\hline 'none' & \(x\) \\
\hline 'logit' & \(1 /\left(1+e^{-x}\right)\) \\
\hline 'doublelogit' & \(1 /\left(1+e^{-2 x}\right)\) \\
\hline 'symmetriclogit' & \(2 /\left(1+e^{-x}\right)-1\) \\
\hline 'sign' & \begin{tabular}{l}
-1 for \(x<0\) \\
0 for \(x=0\) \\
1 for \(x>0\)
\end{tabular} \\
\hline
\end{tabular}

\section*{ClassificationPartitionedModel}

You can include your own function handle for transforming scores. Your function should accept a matrix (the original scores) and return a matrix of the same size (the transformed scores).
Add or change a ScoreTransform function by dot addressing:
```

cvmodel.ScoreTransform = 'function'
Or
cvmodel.ScoreTransform = @function

```

\section*{Trained}

The trained learners, a cell array of compact classification models.

\section*{W}

The scaled weights, a vector with length n , the number of rows in \(X\).

X
A matrix of predictor values. Each column of X represents one variable, and each row represents one observation.

Y
A numeric column vector with the same number of rows as \(X\). Each entry in \(Y\) is the response to the data in the corresponding row of \(X\).

\section*{Methods}
\begin{tabular}{ll} 
kfoldEdge & \begin{tabular}{l} 
Classification edge for \\
observations not used for \\
training
\end{tabular} \\
kfoldfun & \begin{tabular}{l} 
Cross validate function
\end{tabular} \\
kfoldLoss & \begin{tabular}{l} 
Classification loss for \\
observations not used for \\
training
\end{tabular}
\end{tabular}

\section*{ClassificationPartitionedModel}


\section*{ClassificationTree}
\begin{tabular}{|c|c|}
\hline Superclasses & CompactClassificationTree \\
\hline Purpose & Binary decision tree for classification \\
\hline Description & A decision tree with binary splits for classification. An object of class ClassificationTree can predict responses for new data with the predict method. The object contains the data used for training, so can compute resubstitution predictions. \\
\hline \multirow[t]{5}{*}{Construction} & \begin{tabular}{l}
tree = ClassificationTree.fit( \(\mathrm{X}, \mathrm{Y}\) ) returns a classification tree based on the input variables (also known as predictors, features, or attributes) X and output (response) Y . tree is a binary tree, where each branching node is split based on the values of a column of \(X\). \\
tree \(=\) ClassificationTree.fit(X,Y,Name, Value) fits a tree with additional options specified by one or more Name, Value pair arguments. If you use one of the following five options, tree is of class ClassificationPartitionedModel: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Otherwise, tree is of class ClassificationTree.
\end{tabular} \\
\hline & Input Arguments \\
\hline & X \\
\hline & A matrix of numeric predictor values. Each column of X represents one variable, and each row represents one observation. \\
\hline & NaN values in X are taken to be missing values. Observations with all missing values for X are not used in the fit. Observations with some missing values for X are used to find splits on variables for which these observations have valid values. \\
\hline
\end{tabular}

\section*{Y}

A numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. Each row of \(Y\) represents the classification of the corresponding

\section*{ClassificationTree}

\begin{abstract}
row of \(X\). For numeric \(Y\), consider using RegressionTree.fit instead of ClassificationTree.fit.

NaN values in \(Y\) are taken to be missing values. Observations with missing values for Y are not used in the fit.
\end{abstract}

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{CategoricalPredictors}

List of categorical predictors. Pass CategoricalPredictors as one of:
- A numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of \(X\).
- A logical vector of length \(p\), where a true entry means that the corresponding column of X is a categorical variable.
- 'all', meaning all predictors are categorical.
- A cell array of strings, where each element in the array is the name of a predictor variable. The names must match entries in PredictorNames values.
- A character matrix, where each row of the matrix is a name of a predictor variable. The names must match entries in PredictorNames values. Pad the names with extra blanks so each row of the character matrix has the same length.

Default: []

\section*{ClassNames}

Array of class names. Use the data type that exists in Y .

\section*{ClassificationTree}

Use ClassNames to order the classes or to select a subset of classes for training.

Default: The class names that exist in \(Y\)

\section*{Cost}

Square matrix, where \(\operatorname{Cost}(i, j)\) is the cost of classifying a point into class \(j\) if its true class is i. Alternatively, Cost can be a structure S having two fields: S . ClassNames containing the group names as a variable of the same type as Y , and \(S . C l a s s i f i c a t i o n C o s t s ~ c o n t a i n i n g ~ t h e ~ c o s t ~ m a t r i x . ~\)

Default: \(\operatorname{Cost}(i, j)=1\) if \(i \sim=j\), and \(\operatorname{Cost}(i, j)=0\) if \(i=j\)

\section*{crossval}

If ' on ', grows a cross-validated decision tree with 10 folds. You can use 'kfold', 'holdout', 'leaveout', or 'cvpartition' parameters to override this cross-validation setting. You can only use one of these four parameters ('kfold', 'holdout', 'leaveout', or 'cvpartition') at a time when creating a cross-validated tree.

Alternatively, cross validate tree later using the crossval method.

Default: 'off'
cupartition
Partition created with cvpartition to use in a cross-validated tree. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
holdout
Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar

\section*{ClassificationTree}
from 0 to 1. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

\section*{kfold}

Number of folds to use in a cross-validated tree, a positive integer. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

Default: 10

\section*{leaveout}

Use leave-one-out cross validation by setting to 'on'. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

\section*{MergeLeaves}

When 'on', ClassificationTree.fit merges leaves that originate from the same parent node, and that give a sum of risk values greater or equal to the risk associated with the parent node. When 'off', ClassificationTree.fit does not merge leaves.

Default: 'on'
MinLeaf
Each leaf has at least MinLeaf observations per tree leaf. If you supply both MinParent and MinLeaf, ClassificationTree.fit uses the setting that gives larger leaves: MinParent=max (MinParent, 2*MinLeaf).

Default: 1
MinParent

\section*{ClassificationTree}

Each branch node in the tree has at least MinParent observations. If you supply both MinParent and MinLeaf, ClassificationTree.fit uses the setting that gives larger leaves: MinParent=max (MinParent,2*MinLeaf).

Default: 10

\section*{NVarToSample}

Number of predictors to select at random for each split. Can be a positive integer or 'all', which means use all available predictors.

Default: 'all'

\section*{PredictorNames}

A cell array of names for the predictor variables, in the order in which they appear in X .

Default: \{'x1', 'x2', ...\}
prior
Prior probabilities for each class. Specify as one of:
- A string:
- 'empirical' determines class probabilities from class frequencies in Y. If you pass observation weights, they are used to compute the class probabilities.
- 'uniform' sets all class probabilities equal.
- A vector (one scalar value for each class)
- A structure \(S\) with two fields:
- S.ClassNames containing the class names as a variable of the same type as \(Y\)

\section*{ClassificationTree}
- S.ClassProbs containing a vector of corresponding probabilities

If you set values for both weights and prior, the weights are renormalized to add up to the value of the prior probability in the respective class.

Default: 'empirical'

\section*{Prune}

When 'on', ClassificationTree.fit grows the classification tree, and computes the optimal sequence of pruned subtrees. When 'off' ClassificationTree.fit grows the classification tree without pruning.

Default: 'on'

\section*{PruneCriterion}

String with the pruning criterion, either 'error' or 'impurity'.
Default: 'error'

\section*{ResponseName}

Name of the response variable \(Y\), a string.
Default: 'Response'

\section*{ScoreTransform}

Function handle for transforming scores, or string representing a built-in transformation function.
\begin{tabular}{l|l}
\hline String & Formula \\
\hline 'symmetric' & \(2 x-1\) \\
\hline 'invlogit' & \(\log (x /(1-x))\) \\
\hline
\end{tabular}

\section*{ClassificationTree}
\begin{tabular}{l|l}
\hline String & Formula \\
\hline ' ismax' & \begin{tabular}{l} 
Set score for the class with the \\
largest score to 1, and scores for all \\
other classes to 0.
\end{tabular} \\
\hline 'symmetricismax' & \begin{tabular}{l} 
Set score for the class with the \\
largest score to 1, and scores for \\
all other classes to -1.
\end{tabular} \\
\hline 'none' & \(x\) \\
\hline 'logit' & \(1 /\left(1+e^{-x}\right)\) \\
\hline 'doublelogit' & \(1 /\left(1+e^{-2 x}\right)\) \\
\hline 'symmetriclogit' & \(2 /\left(1+e^{-x}\right)-1\) \\
\hline 'sign' & \begin{tabular}{l}
-1 for \(x<0\) \\
0 for \(x=0\) \\
1 for \(x>0\)
\end{tabular} \\
\hline
\end{tabular}

You can include your own function handle for transforming scores. Your function should accept a matrix (the original scores) and return a matrix of the same size (the transformed scores).

Default: 'none'

\section*{SplitCriterion}

Criterion for choosing a split. One of 'gdi' (Gini's diversity index), 'twoing ' for the twoing rule, or 'deviance' for maximum deviance reduction (also known as cross entropy).

Default: 'gdi'

\section*{Surrogate}

When 'on', ClassificationTree.fit finds surrogate splits at each branch node. This setting improves the accuracy of predictions for data with missing values. The setting also enables

\section*{ClassificationTree}
you to compute measures of predictive association between predictors. This setting can use much time and memory.

Default: 'off'

\section*{weights}

Vector of observation weights. The length of weights is the number of rows in X. ClassificationTree.fit normalizes the weights in each class to add up to the value of the prior probability of the class.

Default: ones(size(X,1),1)

\section*{Properties}

\section*{CategoricalPredictors}

List of categorical predictors, a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of \(X\).

\section*{CatSplit}

An \(n\)-by- 2 cell array, where \(n\) is the number of nodes in tree. Each row in CatSplit gives left and right values for a categorical split. For each branch node \(j\) based on a categorical predictor variable \(z\), the left child is chosen if \(z\) is in CatSplit ( \(j, 1\) ) and the right child is chosen if \(z\) is in CatSplit ( \(j, 2\) ). The splits are in the same order as nodes of the tree. Find the nodes for these splits by selecting 'categorical' cuts from top to bottom in the CutType property.
Children
An \(n\)-by- 2 array containing the numbers of the child nodes for each node in tree, where \(n\) is the number of nodes. Leaf nodes have child node 0.

\section*{ClassCount}

An \(n\)-by- \(k\) array of class counts for the nodes in tree, where \(n\) is the number of nodes and \(k\) is the number of classes. For any node number \(i\), the class counts ClassCount (i,:) are counts of

\section*{ClassificationTree}
observations (from the data used in fitting the tree) from each class satisfying the conditions for node i.

\section*{ClassNames}

List of the elements in \(Y\) with duplicates removed. ClassNames can be a numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. ClassNames has the same data type as the data in the argument Y .

\section*{ClassProb}

An \(n\)-by- \(k\) array of class probabilities for the nodes in tree, where \(n\) is the number of nodes and \(k\) is the number of classes. For any node number \(i\), the class probabilities ClassProb(i,:) are the estimated probabilities for each class for a point satisfying the conditions for node i.

\section*{Cost}

Square matrix, where \(\operatorname{Cost}(i, j)\) is the cost of classifying a point into class j if its true class is i .

\section*{CutCategories}

An \(n\)-by- 2 cell array of the categories used at branches in tree, where \(n\) is the number of nodes. For each branch node i based on a categorical predictor variable \(x\), the left child is chosen if \(x\) is among the categories listed in CutCategories \(\{\mathrm{i}, 1\}\), and the right child is chosen if \(x\) is among those listed in CutCategories \(\{i, 2\}\). Both columns of CutCategories are empty for branch nodes based on continuous predictors and for leaf nodes.

CutPoint contains the cut points for 'continuous' cuts, and CutCategories contains the set of categories.

\section*{CutPoint}

An \(n\)-element vector of the values used as cut points in tree, where \(n\) is the number of nodes. For each branch node i based on a continuous predictor variable \(x\), the left child is chosen if \(x<C u t P o i n t(i)\) and the right child is chosen if \(x>=C u t P o i n t(i)\).

\section*{ClassificationTree}

CutPoint is NaN for branch nodes based on categorical predictors and for leaf nodes.

CutPoint contains the cut points for 'continuous ' cuts, and CutCategories contains the set of categories.

\section*{CutType}

An \(n\)-element cell array indicating the type of cut at each node in tree, where \(n\) is the number of nodes. For each node \(i\), CutType\{i\} is:
- 'continuous' - If the cut is defined in the form \(x<v\) for a variable x and cut point v .
- 'categorical' - If the cut is defined by whether a variable \(x\) takes a value in a set of categories.
- ' ' - If i is a leaf node.

CutPoint contains the cut points for 'continuous ' cuts, and CutCategories contains the set of categories.

\section*{CutVar}

An \(n\)-element cell array of the names of the variables used for branching in each node in tree, where \(n\) is the number of nodes. These variables are sometimes known as cut variables. For leaf nodes, CutVar contains an empty string.

CutPoint contains the cut points for 'continuous ' cuts, and CutCategories contains the set of categories.

\section*{IsBranch}

An \(n\)-element logical vector that is true for each branch node and false for each leaf node of tree.

\section*{ModelParams}

Parameters used in training tree.

\section*{NObservations}

\section*{ClassificationTree}

Number of observations in the training data, a numeric scalar. NObservations can be less than the number of rows of input data \(X\) when there are missing values in \(X\) or response \(Y\).

\section*{NodeClass}

An \(n\)-element cell array with the names of the most probable classes in each node of tree, where \(n\) is the number of nodes in the tree. Every element of this array is a string equal to one of the class names in ClassNames.

\section*{NodeErr}

An \(n\)-element vector of the errors of the nodes in tree, where \(n\) is the number of nodes. NodeErr(i) is the misclassification probability for node i.

\section*{NodeProb}

An \(n\)-element vector of the probabilities of the nodes in tree, where \(n\) is the number of nodes. The probability of a node is computed as the proportion of observations from the original data that satisfy the conditions for the node. This proportion is adjusted for any prior probabilities assigned to each class.

NodeSize
An \(n\)-element vector of the sizes of the nodes in tree, where \(n\) is the number of nodes. The size of a node is defined as the number of observations from the data used to create the tree that satisfy the conditions for the node.

\section*{NumNodes}

The number of nodes in tree.

\section*{Parent}

An \(n\)-element vector containing the number of the parent node for each node in tree, where \(n\) is the number of nodes. The parent of the root node is 0 .
```

PredictorNames

```

\section*{ClassificationTree}

A cell array of names for the predictor variables, in the order in which they appear in \(X\).

\section*{Prior}

Prior probabilities for each class. Prior is a numeric vector whose entries relate to the corresponding ClassNames property.

\section*{PruneList}

An \(n\)-element numeric vector with the pruning levels in each node of tree, where \(n\) is the number of nodes.

\section*{ResponseName}

String describing the response variable \(Y\).

\section*{Risk}

An \(n\)-element vector of the risk of the nodes in tree, where \(n\) is the number of nodes.
- If tree was grown with SplitCriterion set to either 'gdi' (default) or 'deviance', then the risk at node \(i\) is the impurity of node i times the node probability. See "Definitions" on page 20-210
- Otherwise,
\[
\text { Risk(i) }=\operatorname{NodeErr(i)} \text { * NodeProb(i). }
\]

\section*{ScoreTransform}

Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none' means @(x)x.

Add or change a ScoreTransform function by dot addressing:
```

tree.ScoreTransform = 'function'
or
tree.ScoreTransform = @function

```

\section*{ClassificationTree}

\section*{SurrCutCategories}

An \(n\)-element cell array of the categories used for surrogate splits in tree, where \(n\) is the number of nodes in tree. For each node k, SurrCutCategories \(\{k\}\) is a cell array. The length of SurrCutCategories \(\{\mathrm{k}\}\) is equal to the number of surrogate predictors found at this node. Every element of SurrCutCategories \(\{\mathrm{k}\}\) is either an empty string for a continuous surrogate predictor, or is a two-element cell array with categories for a categorical surrogate predictor. The first element of this two-element cell array lists categories assigned to the left child by this surrogate split, and the second element of this two-element cell array lists categories assigned to the right child by this surrogate split. The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutCategories contains an empty cell.

\section*{SurrCutFlip}

An \(n\)-element cell array of the numeric cut assignments used for surrogate splits in tree, where \(n\) is the number of nodes in tree. For each node k, SurrCutFlip \(\{k\}\) is a numeric vector. The length of SurrCutFlip \(\{\mathrm{k}\}\) is equal to the number of surrogate predictors found at this node. Every element of SurrCutFlip\{k\} is either zero for a categorical surrogate predictor, or a numeric cut assignment for a continuous surrogate predictor. The numeric cut assignment can be either -1 or +1 . For every surrogate split with a numeric cut \(C\) based on a continuous predictor variable \(Z\), the left child is chosen if \(Z<C\) and the cut assignment for this surrogate split is +1 , or if \(Z \geq C\) and the cut assignment for this surrogate split is -1 . Similarly, the right child is chosen if \(Z \geq C\) and the cut assignment for this surrogate split is +1 , or if \(Z<C\) and the cut assignment for this surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutFlip contains an empty array.

\section*{ClassificationTree}

\section*{SurrCutPoint}

An \(n\)-element cell array of the numeric values used for surrogate splits in tree, where \(n\) is the number of nodes in tree. For each node \(k\), SurrCutPoint \(\{\mathrm{k}\}\) is a numeric vector. The length of SurrCutPoint \(\{\mathrm{k}\}\) is equal to the number of surrogate predictors found at this node. Every element of SurrCutPoint \{k\} is either NaN for a categorical surrogate predictor, or a numeric cut for a continuous surrogate predictor. For every surrogate split with a numeric cut \(C\) based on a continuous predictor variable \(Z\), the left child is chosen if \(Z<C\) and SurrCutFlip for this surrogate split is -1 . Similarly, the right child is chosen if \(Z \geq C\) and SurrCutFlip for this surrogate split is +1 , or if \(Z<C\) and SurrCutFlip for this surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables returned by SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutPoint contains an empty cell.

\section*{SurrCutType}

An \(n\)-element cell array indicating types of surrogate splits at each node in tree, where \(n\) is the number of nodes in tree. For each node k, SurrCutType \(\{k\}\) is a cell array with the types of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutType contains an empty cell. A surrogate split type can be either 'continuous' if the cut is defined in the form \(Z<V\) for a variable \(Z\) and cut point \(V\) or 'categorical' if the cut is defined by whether \(Z\) takes a value in a set of categories.

\footnotetext{
SurrCutVar
}

\section*{ClassificationTree}

An \(n\)-element cell array of the names of the variables used for surrogate splits in each node in tree, where \(n\) is the number of nodes in tree. Every element of SurrCutVar is a cell array with the names of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutVar contains an empty cell.

\section*{SurrVarAssoc}

An \(n\)-element cell array of the predictive measures of association for surrogate splits in tree, where \(n\) is the number of nodes in tree. For each node k, SurrVarAssoc \(\{k\}\) is a numeric vector. The length of SurrVarAssoc \(\{k\}\) is equal to the number of surrogate predictors found at this node. Every element of SurrVarAssoc \{k\} gives the predictive measure of association between the optimal split and this surrogate split. The order of the surrogate split variables at each node is the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrVarAssoc contains an empty cell.
w
The scaled weights, a vector with length \(n\), the number of rows in \(X\).

X
A matrix of predictor values. Each column of \(X\) represents one variable, and each row represents one observation.

Y
A numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. Each row of \(Y\) represents the classification of the corresponding row of \(X\).

\section*{ClassificationTree}

\section*{Methods}
\begin{tabular}{ll} 
compact & Compact tree \\
crossval & \begin{tabular}{l} 
Cross-validated decision tree \\
Classification error by cross \\
validation
\end{tabular} \\
fit & \begin{tabular}{l} 
Fit classification tree \\
Produce sequence of subtrees by \\
pruning
\end{tabular} \\
prune & \begin{tabular}{l} 
Classification edge by \\
resubstitution
\end{tabular} \\
resubEdge & \begin{tabular}{l} 
Classification error by \\
resubstitution
\end{tabular} \\
resubLoss & \begin{tabular}{l} 
Classification margins by \\
resubstitution
\end{tabular} \\
resubMargin & \begin{tabular}{l} 
Predict resubstitution response of \\
tree
\end{tabular} \\
resubPredict & Create classification template
\end{tabular}

\section*{Inherited Methods}

\author{
edge \\ loss \\ margin \\ meanSurrVarAssoc \\ predict \\ predictorImportance \\ view
}

Classification edge
Classification error
Classification margins
Mean predictive measure of association for surrogate splits in decision tree

Predict classification
Estimates of predictor importance
View tree

\section*{ClassificationTree}

\section*{Definitions}

\section*{Impurity and Node Error}

ClassificationTree splits nodes based on either impurity or node error. Impurity means one of several things, depending on your choice of the SplitCriterion name-value pair:
- Gini's Diversity Index (gdi) - The Gini index of a node is
\[
1-\sum_{i} p^{2}(i)
\]
where the sum is over the classes \(i\) at the node, and \(p(i)\) is the observed fraction of classes with class \(i\) that reach the node. A node with just one class (a pure node) has Gini index 0; otherwise the Gini index is positive. So the Gini index is a measure of node impurity.
- Deviance ('deviance') — With \(p(i)\) defined as for the Gini index, the deviance of a node is
\[
-\sum_{i} p(i) \log p(i) .
\]

A pure node has deviance 0 ; otherwise, the deviance is positive.
- Twoing rule ('twoing') — Twoing is not a purity measure of a node, but is a different measure for deciding how to split a node. Let \(L(i)\) denote the fraction of members of class \(i\) in the left child node after a split, and \(R(i)\) denote the fraction of members of class \(i\) in the right child node after a split. Choose the split criterion to maximize
\[
P(L) P(R)\left(\sum_{i}|L(i)-R(i)|\right)^{2}
\]
where \(P(L)\) and \(P(R)\) are the fractions of observations that split to the left and right respectively. If the expression is large, the split made each child node purer. Similarly, if the expression is small, the split made each child node similar to each other, and hence similar to the parent node, and so the split did not increase node purity.

\section*{ClassificationTree}
- Node error - The node error is the fraction of misclassified classes at a node. If \(j\) is the class with largest number of training samples at a node, the node error is
\[
1-p(j)
\]

\section*{Copy Semantics}

Value. To learn how value classes affect copy operations, see Copying Objects in the MATLAB Programming Fundamentals documentation.

Examples Construct a classification tree for the data in ionosphere.mat:
```

load ionosphere
tc = ClassificationTree.fit(X,Y)
tc =
ClassificationTree:
PredictorNames: {1\times34 cell}
CategoricalPredictors: []
ResponseName: 'Response'
ClassNames: {'b' 'g'}
Cost: [2x2 double]
Prior: [0.3590 0.6410]
ScoreTransform: 'none'
X: [351x34 double]
Y: {351x1 cell}
W: [351x1 double]
ModelParams: [1\times1 classreg.learning.modelparams.TreeParams]
See Also RegressionTree | ClassificationEnsemble |
. "Classification Trees and Regression Trees" on page 13-27

```
How To

\section*{Purpose \\ Discriminant analysis}

Syntax
```

class = classify(sample,training,group)
class = classify(sample,training,group,'type')
class = classify(sample,training,group,'type',prior)
[class,err] = classify(...)
[class,err,POSTERIOR] = classify(...)
[class,err,POSTERIOR,logp] = classify(...)
[class,err,POSTERIOR,logp,coeff] = classify(...)

```

\section*{Description}
class = classify(sample,training, group) classifies each row of the data in sample into one of the groups in training. (See "Grouped Data" on page 2-34.) sample and training must be matrices with the same number of columns. group is a grouping variable for training. Its unique values define groups; each element defines the group to which the corresponding row of training belongs. group can be a categorical variable, a numeric vector, a string array, or a cell array of strings. training and group must have the same number of rows. classify treats NaNs or empty strings in group as missing values, and ignores the corresponding rows of training. The output class indicates the group to which each row of sample has been assigned, and is of the same type as group.
class = classify(sample,training, group, 'type') allows you to specify the type of discriminant function. Specify type inside single quotes. type is one of:
- linear - Fits a multivariate normal density to each group, with a pooled estimate of covariance. This is the default.
- diaglinear - Similar to linear, but with a diagonal covariance matrix estimate (naive Bayes classifiers).
- quadratic - Fits multivariate normal densities with covariance estimates stratified by group.
- diagquadratic - Similar to quadratic, but with a diagonal covariance matrix estimate (naive Bayes classifiers).
- mahalanobis - Uses Mahalanobis distances with stratified covariance estimates.
class = classify(sample,training, group,'type',prior) allows you to specify prior probabilities for the groups. prior is one of:
- A numeric vector the same length as the number of unique values in group (or the number of levels defined for group, if group is categorical). If group is numeric or categorical, the order of prior must correspond to the ordered values in group, or, if group contains strings, to the order of first occurrence of the values in group.
- A 1-by-1 structure with fields:
- prob - A numeric vector.
- group - Of the same type as group, containing unique values indicating the groups to which the elements of prob correspond.

As a structure, prior can contain groups that do not appear in group. This can be useful if training is a subset a larger training set. classify ignores any groups that appear in the structure but not in the group array.
- The string 'empirical', indicating that group prior probabilities should be estimated from the group relative frequencies in training.
prior defaults to a numeric vector of equal probabilities, i.e., a uniform distribution. prior is not used for discrimination by Mahalanobis distance, except for error rate calculation.
[class,err] = classify(...) also returns an estimate err of the misclassification error rate based on the training data. classify returns the apparent error rate, i.e., the percentage of observations in training that are misclassified, weighted by the prior probabilities for the groups.
[class,err,POSTERIOR] = classify(...) also returns a matrix POSTERIOR of estimates of the posterior probabilities that the \(j\) th training group was the source of the \(i\) th sample observation, i.e.,

\section*{classify}
\(\operatorname{Pr}(\) group \(j \mid\) obs \(i)\). POSTERIOR is not computed for Mahalanobis discrimination.
[class,err,POSTERIOR,logp] = classify(...) also returns a vector logp containing estimates of the logarithms of the unconditional predictive probability density of the sample observations, \(p(o b s i)=\) \(\sum p(o b s i \mid g r o u p j) \operatorname{Pr}(\) group \(j\) ) over all groups. logp is not computed for Mahalanobis discrimination.
[class,err,POSTERIOR,logp,coeff] = classify(...) also returns a structure array coeff containing coefficients of the boundary curves between pairs of groups. Each element coeff (I, J) contains information for comparing group I to group \(J\) in the following fields:
- type - Type of discriminant function, from the type input.
- name1 - Name of the first group.
- name2 - Name of the second group.
- const - Constant term of the boundary equation (K)
- linear - Linear coefficients of the boundary equation (L)
- quadratic - Quadratic coefficient matrix of the boundary equation (Q)

For the linear and diaglinear types, the quadratic field is absent, and a row \(x\) from the sample array is classified into group I rather than group \(J\) if \(0<K+x * L\). For the other types, \(x\) is classified into group I if \(0<K+x^{*} L+x^{*} Q^{*} x^{\prime}\).

\section*{Examples}

For training data, use Fisher's sepal measurements for iris versicolor and virginica:
```

load fisheriris
SL = meas(51:end,1);
SW = meas(51:end,2);
group = species(51:end);
h1 = gscatter(SL,SW,group,'rb','v^',[],'off');

```


Classify a grid of measurements on the same scale:
```

[X,Y] = meshgrid(linspace(4.5,8),linspace(2,4));
X = X(:); Y = Y(:);
[C,err,P,logp,coeff] = classify([X Y],[SL SW],...
group,'quadratic');

```

Visualize the classification:
```

hold on;
gscatter(X,Y,C,'rb','.',1, 'off');
K = coeff(1,2).const;
L = coeff(1,2).linear;
Q = coeff(1,2).quadratic;
% Function to compute K + L*v + v'*Q*v for multiple vectors
% v=[x;y]. Accepts x and y as scalars or column vectors.
f = @(x,y) K + [x y]*L + sum(([x y]*Q) .* [x y], 2);
h2 = ezplot(f,[4.5 8 2 4]);
set(h2,'Color','m','LineWidth', 2)
axis([4.5 8 2 4])
xlabel('Sepal Length')
ylabel('Sepal Width')
title('{\bf Classification with Fisher Training Data}')

```


\author{
References \\ See Also \\ classregtree | mahal | NaiveBayes \\ How To - "Grouped Data" on page 2-34
}

\section*{classregtree.classname}

Purpose Class names for classification decision tree
Syntax
CNAMES = classname(T)
CNAMES = classname(T,J)

Description
CNAMES = classname \((\mathrm{T})\) returns a cell array of strings with class names for this classification decision tree.

CNAMES = classname ( \(\mathrm{T}, \mathrm{J}\) ) takes an array J of class numbers and returns the class names for the specified numbers.

See Also classregtree

\section*{TreeBagger.ClassNames property}

Purpose Names of classes
Description The ClassNames property is a cell array containing the class names for the response variable Y . This property is empty for regression trees.

\section*{classregtree.classprob}

Purpose Class probabilities
```

Syntax
P = classprob(t)
P = classprob(t,nodes)

```
\(\mathrm{P}=\) classprob( t ) returns an \(n\)-by- \(m\) array P of class probabilities for the nodes in the classification tree t , where \(n\) is the number of nodes and \(m\) is the number of classes. For any node number \(i\), the class probabilities \(\mathrm{P}(\mathrm{i},:\) ) are the estimated probabilities for each class for a point satisfying the conditions for node i.
\(P=\) classprob(t, nodes) takes a vector nodes of node numbers and returns the class probabilities for the specified nodes.

\section*{Examples}

Create a classification tree for Fisher's iris data:
```

load fisheriris;
t = classregtree(meas,species,...
'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica

```
view(t)

```

P = classprob(t)
P =

| 0.3333 | 0.3333 | 0.3333 |
| ---: | ---: | ---: |
| 1.0000 | 0 | 0 |
| 0 | 0.5000 | 0.5000 |
| 0 | 0.9074 | 0.0926 |
| 0 | 0.0217 | 0.9783 |
| 0 | 0.9792 | 0.0208 |
| 0 | 0.3333 | 0.6667 |
| 0 | 1.0000 | 0 |

```

\section*{classregtree.classprob}

\author{
\(0 \quad 0 \quad 1.0000\) \\ References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification \\ See Also classregtree | numnodes
}

\section*{Purpose Classification and regression trees}

\section*{Construction classregtree}

Construct classification and regression trees

Categorical splits used for branches in decision tree

Child nodes
Class counts
Class names for classification decision tree

Class probabilities
Cut categories
Decision tree cut point values
Cut types
Cut variable names
Display classregtree object
Display classregtree object
Predicted responses
Test node for branch
Mean predictive measure of association for surrogate splits in decision tree

Class values of nodes of classification tree

Return vector of node errors
\begin{tabular}{ll} 
nodemean & \begin{tabular}{l} 
Mean values of nodes of \\
regression tree
\end{tabular} \\
nodeprob & Node probabilities \\
nodesize & Return node size \\
numnodes & Number of nodes \\
parent & Parent node \\
prune & Prune tree \\
prunelist & \begin{tabular}{l} 
Pruning levels for decision tree \\
nodes
\end{tabular} \\
risk & \begin{tabular}{l} 
Node risks \\
subsasgn
\end{tabular} \\
\begin{tabular}{l} 
Subscripted reference for \\
classregtree object
\end{tabular} \\
subsref & \begin{tabular}{l} 
Subscripted reference for \\
classregtree object
\end{tabular} \\
surrcutcategories & \begin{tabular}{l} 
Categories used for surrogate \\
splits in decision tree
\end{tabular} \\
surrcutflip & \begin{tabular}{l} 
Numeric cutpoint assignments \\
used for surrogate splits in
\end{tabular} \\
& \begin{tabular}{l} 
decision tree
\end{tabular} \\
surrcutpoint & \begin{tabular}{l} 
Cutpoints used for surrogate \\
splits in decision tree
\end{tabular} \\
surrcuttype & \begin{tabular}{l} 
Types of surrogate splits used at \\
branches in decision tree
\end{tabular} \\
surrcutvar & \begin{tabular}{l} 
Variables used for surrogate \\
splits in decision tree
\end{tabular} \\
surrvarassoc & \begin{tabular}{l} 
Predictive measure of association \\
for surrogate splits in decision \\
tree
\end{tabular} \\
test & \begin{tabular}{lr} 
Error rate
\end{tabular} \\
\hline
\end{tabular}
\begin{tabular}{ll} 
type & Tree type \\
varimportance & \begin{tabular}{l} 
Compute embedded estimates of \\
input feature importance
\end{tabular} \\
view & Plot tree
\end{tabular}

\section*{Properties}

\section*{Copy Semantics}

How To

Objects of the classregtree class have no properties accessible by dot indexing, get methods, or set methods. To obtain information about a classregtree object, use the appropriate method.

Value. To learn how this affects your use of the class, see Comparing Handle and Value Classes in the MATLAB Object-Oriented Programming documentation.
- "Ensemble Methods" on page 13-51
- "Classification Trees and Regression Trees" on page 13-27
- "Grouped Data" on page 2-34

Purpose Construct classification and regression trees
```

Syntax
$\mathrm{t}=\mathrm{classregtree}(\mathrm{X}, \mathrm{y})$
t = classregtree(X,y,'Name',value)

```
\(t=\) classregtree \((X, y)\) creates a decision tree \(t\) for predicting the response \(y\) as a function of the predictors in the columns of \(X . X\) is an \(n\)-by- \(m\) matrix of predictor values. If y is a vector of \(n\) response values, classregtree performs regression. If \(y\) is a categorical variable, character array, or cell array of strings, classregtree performs classification. Either way, t is a binary tree where each branching node is split based on the values of a column of X . NaN values in X or y are taken to be missing values. Observations with all missing values for \(X\) or missing values for \(y\) are not used in the fit. Observations with some missing values for X are used to find splits on variables for which these observations have valid values.
\(\mathrm{t}=\) classregtree (X, y, 'Name', value) specifies one or more optional parameter name/value pairs. Specify Name in single quotes. The following options are available:

For all trees:
- categorical - Vector of indices of the columns of \(X\) that are to be treated as unordered categorical variables
- method - Either 'classification' (default if y is text or a categorical variable) or 'regression' (default if \(y\) is numeric).
- names - A cell array of names for the predictor variables, in the order in which they appear in the \(X\) from which the tree was created.
- prune - 'on' (default) to compute the full tree and the optimal sequence of pruned subtrees, or 'off' for the full tree without pruning.
- minparent - A number \(k\) such that impure nodes must have \(k\) or more observations to be split (default is 10).
- minleaf - A minimal number of observations per tree leaf (default is 1). If you supply both 'minparent' and 'minleaf', classregtree uses the setting which results in larger leaves: minparent \(=\) max(minparent, 2*minleaf)
- mergeleaves - ' on ' (default) to merge leaves that originate from the same parent node and give the sum of risk values greater or equal to the risk associated with the parent node. If 'off', classregtree does not merge leaves.
- nvartosample - Number of predictor variables randomly selected for each split. By default all variables are considered for each decision split.
- stream - Random number stream. Default is the MATLAB default random number stream.
- surrogate - ' on' to find surrogate splits at each branch node. Default is 'off'. If you set this parameter to 'on',classregtree can run significantly slower and consume significantly more memory.
- weights - Vector of observation weights. By default the weight of every observation is 1 . The length of this vector must be equal to the number of rows in \(X\).

For regression trees only:
- qetoler - Defines tolerance on quadratic error per node for regression trees. Splitting nodes stops when quadratic error per node drops below qetoler*qed, where qed is the quadratic error for the entire data computed before the decision tree is grown: qed \(=\) norm( \(y\)-ybar) with ybar estimated as the average of the input array Y. Default value is 1e-6.

For classification trees only:
- cost - Square matrix \(C\), where \(C(i, j)\) is the cost of classifying a point into class \(j\) if its true class is \(i\) (default has \(C(i, j)=1\) if \(i \sim=j\), and \(C(i, j)=0\) if \(i=j)\). Alternatively, this value can be a structure \(S\) having two fields: S.group containing the group names as a categorical variable, character array, or cell array of strings; and S.cost containing the cost matrix C.
- splitcriterion - Criterion for choosing a split. One of 'gdi' (default) or Gini's diversity index, 'twoing ' for the twoing rule, or 'deviance' for maximum deviance reduction.
- priorprob - Prior probabilities for each class, specified as a string ('empirical' or 'equal') or as a vector (one value for each distinct group name) or as a structure \(S\) with two fields:
- S.group containing the group names as a categorical variable, character array, or cell array of strings
- S.prob containing a vector of corresponding probabilities.

If the input value is 'empirical' (default), class probabilities are determined from class frequencies in Y . If the input value is 'equal ', all class probabilities are set equal. If both observation weights and class prior probabilities are supplied, the weights are renormalized to add up to the value of the prior probability in the respective class.

\section*{Examples Create a classification tree for Fisher's iris data:}
```

load fisheriris;

```
```

t = classregtree(meas,species,...
'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica

```
view(t)

\section*{classregtree}
\begin{tabular}{|c|c|c|c|c|c|c|c|}
\hline Click to display: |dentity & \(\checkmark\) & Magnification: & 100\% & \(\checkmark\) & Pruning level: & 0 of 4 & - \\
\hline
\end{tabular}


\footnotetext{
References and Regression Trees. Boca Raton, FL: CRC Press, 1984.

See Also eval \| prune \| test | view
How To
- Grouped Data
- "Ensemble Methods" on page 13-51
}
[1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification

Purpose Class levels
Description The CLevels property is a vector of the same type as the grouping variable, containing the unique levels of the grouping variable.

Purpose
Syntax

Description

Examples

Construct agglomerative clusters from linkages

T = cluster(Z,'cutoff', c )
T = cluster(Z,'cutoff', c,'depth',d)
T = cluster(Z,'cutoff',c,'criterion',criterion)
T = cluster(Z,'maxclust', \(n\) )

T = cluster (Z, 'cutoff' c ) constructs clusters from the agglomerative hierarchical cluster tree, \(Z\), as generated by the linkage function. \(Z\) is a matrix of size ( \(m-1\) )-by- 3 , where \(m\) is the number of observations in the original data. c is a threshold for cutting Z into clusters. Clusters are formed when a node and all of its subnodes have inconsistent value less than c. All leaves at or below the node are grouped into a cluster. t is a vector of size m containing the cluster assignments of each observation.
If \(c\) is a vector, \(T\) is a matrix of cluster assignments with one column per cutoff value.

T = cluster(Z,'cutoff', c, 'depth', d) evaluates inconsistent values by looking to a depth \(d\) below each node. The default depth is 2 .

T = cluster(Z,'cutoff', c, 'criterion', criterion) uses the specified criterion for forming clusters, where criterion is one of the strings 'inconsistent' (default) or 'distance'. The 'distance' criterion uses the distance between the two subnodes merged at a node to measure node height. All leaves at or below a node with height less than c are grouped into a cluster.

T = cluster( Z, 'maxclust', n ) constructs a maximum of n clusters using the 'distance' criterion. cluster finds the smallest height at which a horizontal cut through the tree leaves n or fewer clusters.

If n is a vector, T is a matrix of cluster assignments with one column per maximum value.

Compare clusters from Fisher iris data with species:

\author{
load fisheriris
}
```

d = pdist(meas);
Z = linkage(d);
c = cluster(Z,'maxclust',3:5);
crosstab(c(:,1),species)
ans =
0 0 2
0 50 48
50 0 0
crosstab(c(:,2),species)
ans =
0 0 1
0 50 47
0 0
50 0 0
crosstab(c(:,3),species)
ans =

| 0 | 4 | 0 |
| ---: | ---: | ---: |
| 0 | 46 | 47 |
| 0 | 0 | 1 |
| 0 | 0 | 2 |
| 50 | 0 | 0 |

```

See Also
clusterdata | cophenet | inconsistent | linkage | pdist
\begin{tabular}{|c|c|}
\hline Purpose & Construct clusters from Gaussian mixture distribution \\
\hline \multirow[t]{5}{*}{Syntax} & idx = cluster(obj, X) \\
\hline & [idx, nlogl] = cluster(obj, X ) \\
\hline & [idx, nlogl, P] = cluster(obj, X ) \\
\hline & [idx, nlogl, P, logpdf] = cluster(obj, X) \\
\hline & [idx, nlogl, P, logpdf, M] = cluster(obj, X) \\
\hline \multirow[t]{6}{*}{Description} & idx \(=\) cluster \((\mathrm{obj}, \mathrm{X})\) partitions data in the \(n\)-by- \(d\) matrix X , where \(n\) is the number of observations and \(d\) is the dimension of the data, into \\
\hline & \(k\) clusters determined by the \(k\) components of the Gaussian mixture \\
\hline & distribution defined by obj. obj is an object created by gmdistribution or fit. idx is an \(n\)-by- 1 vector, where \(i d x\) (I) is the cluster index of \\
\hline & observation I. The cluster index gives the component with the largest \\
\hline & posterior probability for the observation, weighted by the component \\
\hline & \\
\hline
\end{tabular}

Note The data in X is typically the same as the data used to create the Gaussian mixture distribution defined by obj. Clustering with cluster is treated as a separate step, apart from density estimation. For cluster to provide meaningful clustering with new data, \(X\) should come from the same population as the data used to create obj.
cluster treats NaN values as missing data. Rows of X with NaN values are excluded from the partition.
[idx, nlogl] = cluster(obj, X) also returns nlogl, the negative log-likelihood of the data.
[idx, nlogl, P] = cluster(obj,X) also returns the posterior probabilities of each component for each observation in the \(n\)-by- \(k\) matrix \(P . P(I, J)\) is the probability of component \(J\) given observation \(I\).
[idx, nlogl, P, logpdf] = cluster(obj, X) also returns the \(n\)-by- 1 vector logpdf containing the logarithm of the estimated probability density function for each observation. The density estimate for
observation I is a sum over all components of the component density at I times the component probability.
[idx, nlogl, P, logpdf, M ] = cluster(obj, X) also returns an \(n\)-by- \(k\) matrix \(M\) containing Mahalanobis distances in squared units. \(M(I, J)\) is the Mahalanobis distance of observation I from the mean of component \(J\) J.

\section*{Examples}

Generate data from a mixture of two bivariate Gaussian distributions using the mvnrnd function:
```

MU1 = [1 2];
SIGMA1 = [2 0; 0 .5];
MU2 = [-3 -5];
SIGMA2 = [1 0; 0 1];
X = [mvnrnd(MU1,SIGMA1,1000);mvnrnd(MU2,SIGMA2,1000)];
scatter(X(:,1),X(:,2),10,'.')
hold on

```


Fit a two-component Gaussian mixture model:
```

obj = gmdistribution.fit(X,2);
h = ezcontour(@(x,y)pdf(obj,[x y]),[-8 6],[-8 6]);

```


Use the fit to cluster the data:
```

idx = cluster(obj,X);
cluster1 = X(idx == 1,:);
cluster2 = X(idx == 2,:);
delete(h)
h1 = scatter(cluster1(:,1),cluster1(:,2),10,'r.');
h2 = scatter(cluster2(:,1),cluster2(:,2),10,'g.');
legend([h1 h2],'Cluster 1','Cluster 2','Location','NW')

```


\author{
See Also \\ fit | gmdistribution | mahal | posterior
}

\section*{Purpose Agglomerative clusters from data}
\[
\begin{array}{ll}
\text { Syntax } & \mathrm{T}=\text { clusterdata }(\mathrm{X}, \text { cutoff }) \\
& \mathrm{T}=\text { clusterdata }(\mathrm{X}, \text { Name, Value })
\end{array}
\]

\section*{Description}
\(\mathrm{T}=\) clusterdata(X, cutoff)
T = clusterdata(X,Name, Value) clusters with additional options specified by one or more Name, Value pair arguments.

Tips
- The centroid and median methods can produce a cluster tree that is not monotonic. This occurs when the distance from the union of two clusters, \(r\) and \(s\), to a third cluster is less than the distance between \(r\) and \(s\). In this case, in a dendrogram drawn with the default orientation, the path from a leaf to the root node takes some downward steps. To avoid this, use another method. The following image shows a nonmonotonic cluster tree.


In this case, cluster 1 and cluster 3 are joined into a new cluster, while the distance between this new cluster and cluster 2 is less than the distance between cluster 1 and cluster 3 . This leads to a nonmonotonic tree.

\section*{clusterdata}
- You can provide the output \(T\) to other functions including dendrogram to display the tree, cluster to assign points to clusters, inconsistent to compute inconsistent measures, and cophenet to compute the cophenetic correlation coefficient.

\section*{Input Arguments \\ X}

Matrix with two or more rows. The rows represent observations, the columns represent categories or dimensions.

\section*{cutoff}

When \(0<\) cutoff < 2, clusterdata forms clusters when inconsistent values are greater than cutoff (see inconsistent). When cutoff is an integer \(\geq 2\), clusterdata interprets cutoff as the maximum number of clusters to keep in the hierarchical tree generated by linkage.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
```

criterion

```

Either 'inconsistent' or 'distance'. cutoff

Cutoff for inconsistent or distance measure, a positive scalar. When 0 < cutoff < 2, clusterdata forms clusters when inconsistent values are greater than cutoff (see inconsistent). When cutoff is an integer \(\geq 2\), clusterdata interprets cutoff as the maximum number of clusters to keep in the hierarchical tree generated by linkage. depth

Depth for computing inconsistent values, a positive integer.

\section*{distance}

Any of the distance metric names allowed by pdist (follow the 'minkowski' option by the value of the exponent \(p\) ):
\begin{tabular}{l|l}
\hline Metric & Description \\
\hline 'euclidean' & Euclidean distance (default). \\
\hline 'seuclidean' & \begin{tabular}{l} 
Standardized Euclidean distance. Each \\
coordinate difference between rows in X \\
is scaled by dividing by the corresponding \\
element of the standard deviation \\
S=nanstd (X). To specify another value for S, \\
use D=pdist (X, 'seuclidean' , S).
\end{tabular} \\
\hline 'cityblock' & \begin{tabular}{l} 
City block metric.
\end{tabular} \\
\hline 'minkowski' & \begin{tabular}{l} 
Minkowski distance. The default exponent \\
is 2. To specify a different exponent, use D \\
= pdist (X, 'minkowski' ,P), where P is a
\end{tabular} \\
\hline 'chebychev' & \begin{tabular}{l} 
Chebychev distance (maximum coordinate \\
difference).
\end{tabular} \\
\hline 'mahalanobis' & \begin{tabular}{l} 
Mahalanobis distance, using the \\
sample covariance of X as computed \\
by nancov. To compute the distance \\
with a different covariance, use D \(=\) \\
pdist (X, 'mahalanobis ', C), where the \\
matrix C is symmetric and positive definite.
\end{tabular} \\
\hline 'cosine' & \begin{tabular}{l} 
One minus the cosine of the included angle \\
between points (treated as vectors).
\end{tabular} \\
\hline 'correlation' & \begin{tabular}{l} 
One minus the sample correlation between \\
points (treated as sequences of values).
\end{tabular} \\
\hline 'spearman' & \begin{tabular}{l} 
One minus the sample Spearman's rank \\
correlation between observations (treated as \\
sequences of values).
\end{tabular} \\
\hline
\end{tabular}

\section*{clusterdata}
\(\left.\begin{array}{l|l}\hline \text { Metric } & \text { Description } \\
\hline \text { 'hamming' } & \begin{array}{l}\text { Hamming distance, which is the percentage } \\
\text { of coordinates that differ. }\end{array} \\
\hline \text { ' jaccard' } & \begin{array}{l}\text { One minus the Jaccard coefficient, which is } \\
\text { the percentage of nonzero coordinates that } \\
\text { differ. }\end{array} \\
\hline \begin{array}{l}\text { custom distance } \\
\text { function }\end{array} & \begin{array}{l}\text { A distance function specified using @: } \\
\text { D = pdist (X, @distfun) } \\
\text { A distance function must be of form }\end{array} \\
\text { d2 = distfun(XI , XJ) }\end{array}\right\}\)\begin{tabular}{l} 
taking as arguments a 1-by- \(n\) vector XI, \\
corresponding to a single row of X, and \\
an m2-by-n matrix XJ, corresponding to \\
multiple rows of X. distfun must accept a \\
matrix XJ with an arbitrary number of rows. \\
distfun must return an \(m 2\)-by-1 vector \\
of distances d2, whose \(k\) th element is the \\
distance between XI and XJ ( \(k,:\) ).
\end{tabular}
linkage
Any of the linkage methods allowed by the linkage function:
- 'average'
- 'centroid'
- 'complete'
- 'median'
- 'single'
- 'ward'
- 'weighted'

For details, see the definitions in the linkage function reference page.
```

maxclust

```

Maximum number of clusters to form, a positive integer.

\section*{savememory}

A string, either 'on' or 'off'. When applicable, the 'on' setting causes clusterdata to construct clusters without computing the distance matrix. savememory is applicable when:
- linkage is 'centroid', 'median', or 'ward'
- distance is 'euclidean' (default)

When savememory is 'on', linkage run time is proportional to the number of dimensions (number of columns of \(X\) ). When savememory is 'off', linkage memory requirement is proportional to \(\mathrm{N}^{2}\), where N is the number of observations. So choosing the best (least-time) setting for savememory depends on the problem dimensions, number of observations, and available memory. The default savememory setting is a rough approximation of an optimal setting.

Default: 'on' when X has 20 columns or fewer, or the computer does not have enough memory to store the distance matrix; otherwise 'off'

\section*{Output}

Arguments
\(T\) is a vector of size \(m\) containing a cluster number for each observation.
- When \(0<\) cutoff \(<2\), \(T\) = clusterdata(X, cutoff) is equivalent to:
```

Y = pdist(X,'euclid');

```

\section*{clusterdata}
```

Z = linkage(Y,'single');
T = cluster(Z,'cutoff',cutoff);

```
- When cutoff is an integer \(\geq 2, T=\) clusterdata( \(X\), cutoff) is equivalent to:
```

Y = pdist(X,'euclid');
Z = linkage(Y,'single');
T = cluster(Z,'maxclust',cutoff);

```

\section*{Examples The example first creates a sample data set of random numbers. It then} uses clusterdata to compute the distances between items in the data set and create a hierarchical cluster tree from the data set. Finally, the clusterdata function groups the items in the data set into three clusters. The example uses the find function to list all the items in cluster 2 , and the scatter3 function to plot the data with each cluster shown in a different color.
```

X = [gallery('uniformdata',[10 3],12);...
gallery('uniformdata',[10 3],13)+1.2;...
gallery('uniformdata',[10 3],14)+2.5];
T = clusterdata(X,'maxclust',3);
find(T==2)
ans =
11
12
13
14
15
16
1 7
18
19
20
scatter3(X(:,1),X(:,2),X(:,3),100,T,'filled')

```


Create a hierarchical cluster tree for a data with 20000 observations using Ward's linkage. If you set savememory to 'off', you can get an out-of-memory error if your machine doesn't have enough memory to hold the distance matrix.
```

X = rand(20000,3);
c = clusterdata(X,'linkage','ward','savememory','on',...
'maxclust',4);
scatter3(X(:,1),X(:,2),X(:,3),10,c)

```

\section*{clusterdata}


\section*{See Also}
cluster | inconsistent | kmeans | linkage | pdist
\begin{tabular}{|c|c|}
\hline Purpose & Classical multidimensional scaling \\
\hline Syntax & \[
\begin{aligned}
& Y=\text { cmdscale(D) } \\
& {[Y, e]=\text { cmdscale(D) }}
\end{aligned}
\] \\
\hline \multirow[t]{3}{*}{Description} & \begin{tabular}{l}
\(Y=\) cmdscale( \(D\) ) takes an \(n\)-by-n distance matrix \(D\), and returns an \(n\)-by- \(p\) configuration matrix \(Y\). Rows of \(Y\) are the coordinates of \(n\) points in \(p\)-dimensional space for some \(p<n\). When \(D\) is a Euclidean distance matrix, the distances between those points are given by D. \(p\) is the dimension of the smallest space in which the n points whose inter-point distances are given by D can be embedded. \\
\([Y, e]=\) cmdscale ( \(D\) ) also returns the eigenvalues of \(Y * Y\). When \(D\) is Euclidean, the first \(p\) elements of \(e\) are positive, the rest zero. If the first \(k\) elements of \(e\) are much larger than the remaining ( \(n-k\) ), then you can use the first \(k\) columns of \(Y\) as \(k\)-dimensional points whose inter-point distances approximate \(D\). This can provide a useful dimension reduction for visualization, e.g., for \(\mathrm{k}=2\).
\end{tabular} \\
\hline & D need not be a Euclidean distance matrix. If it is non-Euclidean or a more general dissimilarity matrix, then some elements of e are negative, and cmdscale chooses \(p\) as the number of positive eigenvalues. In this case, the reduction to \(p\) or fewer dimensions provides a reasonable approximation to \(D\) only if the negative elements of e are small in magnitude. \\
\hline & You can specify D as either a full dissimilarity matrix, or in upper triangle vector form such as is output by pdist. A full dissimilarity matrix must be real and symmetric, and have zeros along the diagonal and positive elements everywhere else. A dissimilarity matrix in upper triangle form must have real, positive entries. You can also specify D as a full similarity matrix, with ones along the diagonal and all other elements less than one. cmdscale transforms a similarity matrix to a dissimilarity matrix in such a way that distances between the points returned in Y equal or approximate sqrt(1-D). To use a different transformation, you must transform the similarities prior to calling cmdscale. \\
\hline
\end{tabular}
```

Examples
Generate some points in 4-D space, but close to 3-D space, then reduce them to distances only.

```
```

X = [normrnd(0,1,10,3) normrnd(0,.1,10,1)];

```
X = [normrnd(0,1,10,3) normrnd(0,.1,10,1)];
D = pdist(X,'euclidean');
D = pdist(X,'euclidean');
Find a configuration with those inter-point distances.
```

```
[Y,e] = cmdscale(D);
```

[Y,e] = cmdscale(D);
% Four, but fourth one small
% Four, but fourth one small
dim = sum(e > eps^(3/4))
dim = sum(e > eps^(3/4))
% Poor reconstruction
% Poor reconstruction
maxerr2 = max(abs(pdist(X)-pdist(Y(:,1:2))))
maxerr2 = max(abs(pdist(X)-pdist(Y(:,1:2))))
% Good reconstruction
% Good reconstruction
maxerr3 = max(abs(pdist(X)-pdist(Y(:,1:3))))
maxerr3 = max(abs(pdist(X)-pdist(Y(:,1:3))))
% Exact reconstruction
% Exact reconstruction
maxerr4 = max(abs(pdist(X)-pdist(Y)))
maxerr4 = max(abs(pdist(X)-pdist(Y)))
% D is now non-Euclidean
% D is now non-Euclidean
D = pdist(X,'cityblock');
D = pdist(X,'cityblock');
[Y,e] = cmdscale(D);
[Y,e] = cmdscale(D);
% One is large negative
% One is large negative
min(e)
min(e)
% Poor reconstruction
% Poor reconstruction
maxerr = max(abs(pdist(X)-pdist(Y)))

```
maxerr = max(abs(pdist(X)-pdist(Y)))
```


## References <br> [1] Seber, G. A. F. Multivariate Observations. Hoboken, NJ: John Wiley \& Sons, Inc., 1984.

## See Also mdscale | pdist | procrustes

## Purpose Class names

Description The CNames property is an NClasses-by- 1 cell array containing the group names, where NClasses number of groups in the grouping variable used to create the Naive Bayes classifier.

## CompactTreeBagger.combine

Purpose Combine two ensembles
Syntax $\quad B 1=\operatorname{combine}(B 1, B 2)$
Description $\quad B 1=\operatorname{combine}(B 1, B 2)$ appends decision trees from ensemble $B 2$ to those stored in B1 and returns ensemble B1. This method requires that the class and variable names be identical in both ensembles.

See Also TreeBagger.append

## Purpose Enumeration of combinations

## Syntax <br> C = combnk (v,k)

Description $\quad C=\operatorname{combnk}(v, k)$ returns all combinations of the $n$ elements in $v$ taken k at a time.
$\mathrm{C}=$ combnk( $\mathrm{v}, \mathrm{k})$ produces a matrix C with k columns and $n!/ k!(n-k)$ ! rows, where each row contains k of the elements in the vector v .

It is not practical to use this function if v has more than about 15 elements.

## Examples Combinations of characters from a string.

```
C = combnk('tendril',4);
last5 = C(31:35,:)
last5 =
tedr
tenl
teni
tenr
tend
```

Combinations of elements from a numeric vector.
c $=$ combnk $(1: 4,2)$
c $=$
34
24
23
14
13
12

See Also perms

## ClassificationDiscriminant.compact

Purpose Compact discriminant analysis classifier
Syntax ..... cobj = compact(obj)
Description ..... cobj $=$ compact (obj) creates a compact version of obj.
Input ..... objArgumentsDiscriminant analysis classifier created usingClassificationDiscriminant.fit.
Output ..... cobj
Arguments
Examples Compare the size of the discriminant analysis classifier for Fisher's iris data to the compact version of the classifier:

```
load fisheriris
fullobj = ClassificationDiscriminant.fit(meas,species);
cobj = compact(fullobj);
b = whos('fullobj'); % b.bytes = size of fullobj
c = whos('cobj'); % c.bytes = size of cobj
[b.bytes c.bytes] % shows cobj uses 60% of the memory
ans =
    18578 11498
```


## See Also

ClassificationDiscriminant
How To - "Discriminant Analysis" on page 12-3

## ClassificationEnsemble.compact

Purpose Compact classification ensemble
Syntax cens = compact(ens)
Description cens = compact(ens) creates a compact version of ens. You canpredict classifications using cens exactly as you can using ens.However, since cens does not contain training data, you cannot performsome actions, such as cross validation.
Input ..... ensArguments
Output Arguments ..... cens
A compact classification ensemble. cens has class CompactClassificationEnsemble.
Examples Compare the size of a classification ensemble for Fisher's iris data to the compact version of the ensemble:

```
load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
cens = compact(ens);
b = whos('ens'); % b.bytes = size of ens
c = whos('cens'); % c.bytes = size of ens
[b.bytes c.bytes] % shows cens uses less memory
ans =
    571727 532476
```

See Also ClassificationTree | fitensemble
How To - "Ensemble Methods" on page 13-51

## ClassificationTree.compact

Purpose Compact tree
Syntax ctree = compact(tree)
Description ctree $=$ compact(tree) creates a compact version of tree.
Input ..... tree
Arguments A classification tree created using ClassificationTree.fit.
Output ctreeArgumentsA compact decision tree. ctree has classCompactClassificationTree. You can predict classificationsusing ctree exactly as you can using tree. However, since ctreedoes not contain training data, you cannot perform some actions,such as cross validation.
ExamplesCompare the size of the classification tree for Fisher's iris data to thecompact version of the tree:

```
load fisheriris
fulltree = ClassificationTree.fit(meas,species);
ctree = compact(fulltree);
b = whos('fulltree'); % b.bytes = size of fulltree
c = whos('ctree'); % c.bytes = size of ctree
[b.bytes c.bytes] % shows ctree uses half the memory
ans =
    13913 6818
```

See Also CompactClassificationTree | ClassificationTree | predict
How To - Chapter 13, "Nonparametric Supervised Learning"

## RegressionEnsemble.compact

Purpose Create compact regression ensemble
Syntax cens = compact(ens)
Description cens = compact(ens) creates a compact version of ens. You canpredict regressions using cens exactly as you can using ens. However,since cens does not contain training data, you cannot perform someactions, such as cross validation.
Input ensArguments
A regression ensemble created with fitensemble.
Output Arguments ..... cens
A compact regression ensemble. cens is of class CompactRegressionEnsemble.
Examples Compare the size of a regression ensemble for the carsmall data to the compact version of the ensemble:

```
load carsmall
X = [Acceleration Cylinders Displacement Horsepower Weight];
ens = fitensemble(X,MPG,'LSBoost',100,'Tree');
cens = compact(ens);
b = whos('ens'); % b.bytes = size of ens
c = whos('cens'); % c.bytes = size of cens
[b.bytes c.bytes] % shows ctree uses less memory
ans =
    311789 287368
```

See Also RegressionEnsemble | CompactRegressionEnsemble
How To - Chapter 13, "Nonparametric Supervised Learning"

## RegressionTree.compact

Purpose Compact regression tree
Syntax

ctree = compact(tree)Descriptionctree $=$ compact(tree) creates a compact version of tree.
Input tree
Arguments A regression tree created using RegressionTree.fit.
Output ctreeArgumentsA compact regression tree. ctree has classCompactRegressionTree. You can predict regressionsusing ctree exactly as you can using tree. However, since ctreedoes not contain training data, you cannot perform some actions,such as cross validation.
ExamplesCompare the size of a regression tree for the carsmall data to thecompact version of the tree:

```
load carsmall
X = [Acceleration Cylinders Displacement Horsepower Weight];
fulltree = RegressionTree.fit(X,MPG);
ctree = compact(fulltree);
b = whos('fulltree'); % b.bytes = size of fulltree
c = whos('ctree'); % c.bytes = size of ctree
[b.bytes c.bytes] % shows ctree uses 2/3 the memory
ans =
    15715 10258
```

See Also CompactRegressionTree | RegressionTree | predict
How To - Chapter 13, "Nonparametric Supervised Learning"

| Purpose | Compact ensemble of decision trees |
| :--- | :--- |
| Description | Return an object of class CompactTreeBagger holding the structure <br> of the trained ensemble. The class is more compact than the full <br> TreeBagger class because it does not contain information for growing <br> more trees for the ensemble. In particular, it does not contain $X$ and |
| $Y$ used for training. |  |

See Also

CompactTreeBagger

## CompactClassificationDiscriminant

## Purpose Compact discriminant analysis class

Description

Construction

A CompactClassificationDiscriminant object is a compact version of a discriminant analysis classifier. The compact version does not include the data for training the classifier. Therefore, you cannot perform some tasks with a compact classifier, such as cross validation. Use a compact classifier for making predictions (classifications) of new data.
cobj $=$ compact (obj) constructs a compact classifier from a full classifier.
cobj = ClassificationDiscriminant.make(Mu,Sigma) constructs a compact discriminant analysis classifier from the class means Mu and covariance matrix Sigma. For syntax details, see ClassificationDiscriminant.make.

## Input Arguments

obj
Discriminant analysis classifier, created with ClassificationDiscriminant.fit.

BetweenSigma
p -by-p matrix, the between-class covariance, where p is the number of predictors.

CategoricalPredictors
List of categorical predictors, always empty ([ ]) for discriminant analysis.

## ClassNames

List of the elements in the training data $Y$ with duplicates removed. ClassNames can be a numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. ClassNames has the same data type as the data in the argument $Y$.

## CompactClassificationDiscriminant

Coeffs
k -by-k structure of coefficient matrices, where k is the number of classes. Coeffs (i,j) contains coefficients of the linear or quadratic boundaries between classes $i$ and $j$. Fields in Coeffs(i,j):

- DiscrimType
- Class1 - ClassNames(i)
- Class2 - ClassNames(j)
- Const - A scalar
- Linear - A vector with $p$ components, where $p$ is the number of columns in $X$
- Quadratic - p-by-p matrix, exists for quadratic DiscrimType

The equation of the boundary between class $i$ and class $j$ is

$$
\text { Const }+ \text { Linear * } x+x^{\prime} * \text { Quadratic } * x=0
$$

where $x$ is a column vector of length $p$.
If ClassificationDiscriminant.fit had the FillCoeffs name-value pair set to 'off' when constructing the classifier, Coeffs is empty ([]).

Cost
Square matrix, where $\operatorname{Cost}(i, j)$ is the cost of classifying a point into class $j$ if its true class is $i$. Cost is K-by-K, where $K$ is the number of classes.

Change a Cost matrix using dot addressing:

```
obj.Cost = costMatrix
```

DiscrimType
String specifying the discriminant type. One of:

## CompactClassificationDiscriminant

- 'linear'
- 'quadratic'
- 'diagLinear'
- 'diagQuadratic'
- 'pseudoLinear'
- 'pseudoQuadratic'


## LogDetSigma

Logarithm of the determinant of the within-class covariance matrix. The type of LogDetSigma depends on the discriminant type:

- Scalar for linear discriminant analysis
- Vector of length $K$ for quadratic discriminant analysis, where $K$ is the number of classes


## Mu

Matrix of class means of size K-by-p, where $K$ is the number of classes, and $p$ is the number of predictors. Each row of Mu represents the mean of the multivariate normal distribution of the corresponding class. The class indices are in the ClassNames attribute.

## PredictorNames

Cell array of names for the predictor variables, in the order in which they appear in the training data $X$.

Prior
Prior probabilities for each class. Prior is a numeric vector whose entries relate to the corresponding ClassNames property.

Add or change a Prior vector using dot addressing:
obj.Prior = priorVector

## CompactClassificationDiscriminant

## ResponseName

String describing the response variable Y .

## ScoreTransform

Function handle for transforming scores, or string representing a built-in transformation function. ' none' means no transformation; equivalently, 'none' means @(x)x. For a list of built-in transformation functions and the syntax of custom transformation functions, see ClassificationDiscriminant.fit.

Add or change a ScoreTransform function by dot addressing:

```
cobj.ScoreTransform = 'function'
```

or
cobj.ScoreTransform = @function

## Sigma

Within-class covariance matrix or matrices. The dimensions depend on DiscrimType:

- 'linear' (default) - Matrix of size p -by- p , where p is the number of predictors
- 'quadratic' - Array of size p-by-p-by-K, where $K$ is the number of classes
- 'diagLinear' - Row vector of length $p$
- 'diagQuadratic' - Array of size 1-by-p-by-K
- 'pseudoLinear' - Matrix of size p-by-p
- 'pseudoQuadratic' - Array of size p-by-p-by-K


## Methods

| edge | Classification edge |
| :--- | :--- |
| loss | Classification error |

## CompactClassificationDiscriminant

| mahal | Mahalanobis distance to class <br> means |
| :--- | :--- |
| margin | Classification margins |
| predict | Predict classification |

## Definitions Discriminant Classification

The model for discriminant analysis is:

- Each class (Y) generates data (X) using a multivariate normal distribution. That is, the model assumes X has a Gaussian mixture distribution (gmdistribution).
- For linear discriminant analysis, the model has the same covariance matrix for each class, only the means vary.
- For quadratic discriminant analysis, both means and covariances of each class vary.
predict classifies so as to minimize the expected classification cost:

$$
\hat{y}=\underset{y=1, \ldots, K}{\arg \min } \sum_{k=1}^{K} \hat{P}(k \mid x) C(y \mid k),
$$

where

- $\hat{y}$ is the predicted classification.
- $K$ is the number of classes.
- $\hat{P}(k \mid x)$ is the posterior probability of class $k$ for observation $x$.
- $C(y \mid k)$ is the cost of classifying an observation as $y$ when its true class is $k$.

For details, see "How the predict Method Classifies" on page 12-6.

## CompactClassificationDiscriminant

## Copy <br> Semantics

Value. To learn how value classes affect copy operations, see Copying Objects in the MATLAB Programming Fundamentals documentation.

## Examples

Construct a compact discriminant analysis classifier for the Fisher iris data, and compare its size to that of the full classifier:

```
load fisheriris
fullobj = ClassificationDiscriminant.fit(meas,species);
cobj = compact(fullobj);
b = whos('fullobj'); % b.bytes = size of fullobj
c = whos('cobj'); % c.bytes = size of cobj
[b.bytes c.bytes] % shows cobj uses 60% of the memory
ans =
    1 8 5 7 8
    1 1 4 9 8
```

Construct a compact discriminant analysis classifier from the means and covariances of the Fisher iris data:

```
load fisheriris
mu(1,:) = mean(meas(1:50,:));
mu(2,:) = mean(meas(51:100,:));
mu(3,:) = mean(meas(101:150,:));
mm1 = repmat(mu(1,:),50,1);
mm2 = repmat(mu(2,:),50,1);
mm3 = repmat(mu(3,:),50,1);
CC = meas;
cc(1:50,:) = cc(1:50,:) - mm1;
cc(51:100,:) = cc(51:100,:) - mm2;
cc(101:150,:) = cc(101:150,:) - mm3;
sigstar = cc' * cc / 147;
cpct = ClassificationDiscriminant.make(mu,sigstar,...
    'ClassNames',{'setosa','versicolor','virginica'});
```


## CompactClassificationDiscriminant

See Also ClassificationDiscriminant | compact | ClassificationDiscriminant.make | predict<br>How To . "Discriminant Analysis" on page 12-3

## CompactClassificationEnsemble

Purpose Compact classification ensemble class
Description Compact version of a classification ensemble (of classClassificationEnsemble). The compact version does not include thedata for training the classification ensemble. Therefore, you cannotperform some tasks with a compact classification ensemble, such ascross validation. Use a compact classification ensemble for makingpredictions (classifications) of new data.
Construction cens = compact(ens) constructs a compact decision ensemble from a full decision ensemble.
Input Arguments
ens
A classification ensemble created by fitensemble.
Properties
CategoricalPredictors
List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to $p$, where $p$ is the number of columns of $X$.

## ClassNames

List of the elements in $Y$ with duplicates removed. ClassNames can be a numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. ClassNames has the same data type as the data in the argument Y .

## CombineWeights

String describing how ens combines weak learner weights, either 'WeightedSum' or 'WeightedAverage'.

## Cost

Square matrix where $\operatorname{Cost}(i, j)$ is the cost of classifying a point into class j if its true class is i .
NTrained

## CompactClassificationEnsemble

Number of trained weak learners in cens, a scalar.

## PredictorNames

A cell array of names for the predictor variables, in the order in which they appear in $X$.

## Prior

Prior probabilities for each class. Prior is a numeric vector whose entries relate to the corresponding ClassNames property.

## ResponseName

String with the name of the response variable $Y$.

## ScoreTransform

Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none' means @(x)x. For a list of built-in transformation functions and the syntax of custom transformation functions, see ClassificationTree.fit.

Add or change a ScoreTransform function by dot addressing:

```
cens.ScoreTransform = 'function'
```

or

```
cens.ScoreTransform = @function
```


## Trained

Trained learners, a cell array of compact classification models.

## TrainedWeights

Numeric vector of trained weights for the weak learners in ens. TrainedWeights has T elements, where $T$ is the number of weak learners in learners.

## CompactClassificationEnsemble

## Methods

| edge | Classification edge |
| :--- | :--- |
| loss | Classification error |
| margin | Classification margins |
| predict | Predict classification |
| predictorImportance | Estimates of predictor importance |

Copy Semantics

Value. To learn how value classes affect copy operations, see Copying
Objects in the MATLAB Programming Fundamentals documentation.

## Examples

Create a compact classification ensemble for the ionosphere data:

```
load ionosphere
ens = fitensemble(X,Y,'AdaBoostM1',100,'Tree');
cens = compact(ens)
cens =
classreg.learning.classif.CompactClassificationEnsemble:
    PredictorNames: {1x34 cell}
        CategoricalPredictors: []
            ResponseName: 'Y'
                ClassNames: {'b' 'g'}
            ScoreTransform: 'none'
                NTrained: 100
```

See Also
fitensemble | ClassificationEnsemble | predict | compact
How To

- Chapter 13, "Nonparametric Supervised Learning"


## CompactClassificationTree

## Purpose Compact classification tree

Description

Construction
ctree $=$ compact(tree) constructs a compact decision tree from a full decision tree.

## Input Arguments

## tree

A decision tree constructed by ClassificationTree.fit.

## Properties

Compact version of a classification tree (of class ClassificationTree). The compact version does not include the data for training the classification tree. Therefore, you cannot perform some tasks with a compact classification tree, such as cross validation. Use a compact classification tree for making predictions (classifications) of new data.

## CategoricalPredictors

List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to $p$, where $p$ is the number of columns of X .

CatSplit
An $n$-by- 2 cell array, where $n$ is the number of nodes in tree. Each row in CatSplit gives left and right values for a categorical split. For each branch node $j$ based on a categorical predictor variable $z$, the left child is chosen if $z$ is in CatSplit ( $j, 1$ ) and the right child is chosen if $z$ is in CatSplit ( $j, 2$ ). The splits are in the same order as nodes of the tree. Find the nodes for these splits by selecting 'categorical' cuts from top to bottom in the CutType property.

## Children

An $n$-by- 2 array containing the numbers of the child nodes for each node in tree, where $n$ is the number of nodes. Leaf nodes have child node 0 .

ClassCount

## CompactClassificationTree

An $n$-by- $k$ array of class counts for the nodes in tree, where $n$ is the number of nodes and $k$ is the number of classes. For any node number i, the class counts ClassCount (i,:) are counts of observations (from the data used in fitting the tree) from each class satisfying the conditions for node i.

## ClassNames

List of the elements in Y with duplicates removed. ClassNames can be a numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. ClassNames has the same data type as the data in the argument Y .

## ClassProb

An $n$-by- $k$ array of class probabilities for the nodes in tree, where $n$ is the number of nodes and $k$ is the number of classes. For any node number $i$, the class probabilities ClassProb(i,:) are the estimated probabilities for each class for a point satisfying the conditions for node i.

## Cost

Square matrix, where $\operatorname{Cost}(i, j)$ is the cost of classifying a point into class $j$ if its true class is $i$.

## CutCategories

An $n$-by- 2 cell array of the categories used at branches in tree, where $n$ is the number of nodes. For each branch node $i$ based on a categorical predictor variable x , the left child is chosen if x is among the categories listed in CutCategories $\{i, 1\}$, and the right child is chosen if $x$ is among those listed in CutCategories $\{i, 2\}$. Both columns of CutCategories are empty for branch nodes based on continuous predictors and for leaf nodes.

CutPoint contains the cut points for 'continuous ' cuts, and CutCategories contains the set of categories.

[^3]
## CompactClassificationTree

An $n$-element vector of the values used as cut points in tree, where $n$ is the number of nodes. For each branch node i based on a continuous predictor variable $x$, the left child is chosen if $x<C u t P o i n t$ (i) and the right child is chosen if $x>=$ CutPoint (i). CutPoint is NaN for branch nodes based on categorical predictors and for leaf nodes.

CutPoint contains the cut points for 'continuous ' cuts, and CutCategories contains the set of categories.

## CutType

An $n$-element cell array indicating the type of cut at each node in tree, where $n$ is the number of nodes. For each node i, CutType\{i\} is:

- 'continuous' - If the cut is defined in the form $\mathrm{x}<\mathrm{v}$ for a variable x and cut point v .
- 'categorical' - If the cut is defined by whether a variable $x$ takes a value in a set of categories.
- ' ' - If i is a leaf node.

CutPoint contains the cut points for 'continuous' cuts, and CutCategories contains the set of categories.

## CutVar

An $n$-element cell array of the names of the variables used for branching in each node in tree, where $n$ is the number of nodes. These variables are sometimes known as cut variables. For leaf nodes, CutVar contains an empty string.

CutPoint contains the cut points for 'continuous' cuts, and CutCategories contains the set of categories.

## IsBranch

An $n$-element logical vector that is true for each branch node and false for each leaf node of tree.

## NodeClass

An $n$-element cell array with the names of the most probable classes in each node of tree, where $n$ is the number of nodes in the tree. Every element of this array is a string equal to one of the class names in ClassNames.

## NodeErr

An $n$-element vector of the errors of the nodes in tree, where $n$ is the number of nodes. NodeErr(i) is the misclassification probability for node i.

## NodeProb

An $n$-element vector of the probabilities of the nodes in tree, where $n$ is the number of nodes. The probability of a node is computed as the proportion of observations from the original data that satisfy the conditions for the node. This proportion is adjusted for any prior probabilities assigned to each class.

## NodeSize

An $n$-element vector of the sizes of the nodes in tree, where $n$ is the number of nodes. The size of a node is defined as the number of observations from the data used to create the tree that satisfy the conditions for the node.

## NumNodes

The number of nodes in tree.

## Parent

An $n$-element vector containing the number of the parent node for each node in tree, where $n$ is the number of nodes. The parent of the root node is 0 .

## PredictorNames

A cell array of names for the predictor variables, in the order in which they appear in $X$.

Prior

## CompactClassificationTree

Prior probabilities for each class. Prior is a numeric vector whose entries relate to the corresponding ClassNames property.

## PruneList

An $n$-element numeric vector with the pruning levels in each node of tree, where $n$ is the number of nodes.

## ResponseName

String describing the response variable Y .
Risk
An $n$-element vector of the risk of the nodes in tree, where $n$ is the number of nodes. For each node i,

$$
\text { Risk(i) }=\operatorname{NodeErr(i)~*~NodeProb(i).~}
$$

## ScoreTransform

Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none' means @(x)x. For a list of built-in transformation functions and the syntax of custom transformation functions, see ClassificationTree.fit.

Add or change a ScoreTransform function by dot addressing:

```
ctree.ScoreTransform = 'function'
or
ctree.ScoreTransform = @function
```


## SurrCutCategories

An $n$-element cell array of the categories used for surrogate splits in tree, where $n$ is the number of nodes in tree. For each node k, SurrCutCategories $\{\mathrm{k}\}$ is a cell array. The length of SurrCutCategories $\{\mathrm{k}\}$ is equal to the number of surrogate predictors found at this node. Every element of SurrCutCategories $\{\mathrm{k}\}$ is either an empty string for a continuous surrogate predictor, or is a two-element cell array with categories

## CompactClassificationTree

for a categorical surrogate predictor. The first element of this two-element cell array lists categories assigned to the left child by this surrogate split and the second element of this two-element cell array lists categories assigned to the right child by this surrogate split. The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutCategories contains an empty cell.

## SurrCutFlip

An $n$-element cell array of the numeric cut assignments used for surrogate splits in tree, where $n$ is the number of nodes in tree. For each node k, SurrCutFlip $\{k\}$ is a numeric vector. The length of SurrCutFlip\{k\} is equal to the number of surrogate predictors found at this node. Every element of SurrCutFlip\{k\} is either zero for a categorical surrogate predictor, or a numeric cut assignment for a continuous surrogate predictor. The numeric cut assignment can be either -1 or +1 . For every surrogate split with a numeric cut $C$ based on a continuous predictor variable $Z$, the left child is chosen if $Z<C$ and the cut assignment for this surrogate split is +1 , or if $Z \geq C$ and the cut assignment for this surrogate split is -1 . Similarly, the right child is chosen if $Z \geq C$ and the cut assignment for this surrogate split is +1 , or if $Z<C$ and the cut assignment for this surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutFlip contains an empty array.

## SurrCutPoint

An $n$-element cell array of the numeric values used for surrogate splits in tree, where $n$ is the number of nodes in tree. For each node $k$, SurrCutPoint $\{k\}$ is a numeric vector. The length of SurrCutPoint $\{\mathrm{k}\}$ is equal to the number of surrogate predictors found at this node. Every element of SurrCutPoint \{k\} is either NaN for a categorical surrogate predictor, or a numeric cut for a continuous surrogate predictor. For every surrogate split with a

## CompactClassificationTree

numeric cut $C$ based on a continuous predictor variable $Z$, the left child is chosen if $Z<C$ and SurrCutFlip for this surrogate split is -1 . Similarly, the right child is chosen if $Z \geq C$ and SurrCutFlip for this surrogate split is +1 , or if $Z<C$ and SurrCutFlip for this surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables returned by SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutPoint contains an empty cell.

## SurrCutType

An $n$-element cell array indicating types of surrogate splits at each node in tree, where $n$ is the number of nodes in tree. For each node $k$, SurrCutType $\{k\}$ is a cell array with the types of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutType contains an empty cell. A surrogate split type can be either 'continuous' if the cut is defined in the form $Z<V$ for a variable $Z$ and cut point V or 'categorical' if the cut is defined by whether $Z$ takes a value in a set of categories.

## SurrCutVar

An $n$-element cell array of the names of the variables used for surrogate splits in each node in tree, where $n$ is the number of nodes in tree. Every element of SurrCutVar is a cell array with the names of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutVar contains an empty cell.

## CompactClassificationTree

## SurrVarAssoc

An $n$-element cell array of the predictive measures of association for surrogate splits in tree, where $n$ is the number of nodes in tree. For each node k, SurrVarAssoc $\{k\}$ is a numeric vector. The length of SurrVarAssoc $\{k\}$ is equal to the number of surrogate predictors found at this node. Every element of SurrVarAssoc $\{k\}$ gives the predictive measure of association between the optimal split and this surrogate split. The order of the surrogate split variables at each node is the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrVarAssoc contains an empty cell.

## Methods

| edge | Classification edge |
| :--- | :--- |
| loss | Classification error |
| margin | Classification margins |
| meanSurrVarAssoc | Mean predictive measure of <br> association for surrogate splits in <br> decision tree |
| predict | Predict classification |
| predictorImportance | Estimates of predictor importance |
| view | View tree |

## Definitions Impurity and Node Error

ClassificationTree splits nodes based on either impurity or node error. Impurity means one of several things, depending on your choice of the SplitCriterion name-value pair:

- Gini's Diversity Index (gdi) - The Gini index of a node is

$$
1-\sum_{i} p^{2}(i),
$$

## CompactClassificationTree

where the sum is over the classes $i$ at the node, and $p(i)$ is the observed fraction of classes with class $i$ that reach the node. A node with just one class (a pure node) has Gini index 0; otherwise the Gini index is positive. So the Gini index is a measure of node impurity.

- Deviance ('deviance') - With $p(i)$ defined as for the Gini index, the deviance of a node is

$$
-\sum_{i} p(i) \log p(i) .
$$

A pure node has deviance 0 ; otherwise, the deviance is positive.

- Twoing rule ('twoing') — Twoing is not a purity measure of a node, but is a different measure for deciding how to split a node. Let $L(i)$ denote the fraction of members of class $i$ in the left child node after a split, and $R(i)$ denote the fraction of members of class $i$ in the right child node after a split. Choose the split criterion to maximize

$$
P(L) P(R)\left(\sum_{i}|L(i)-R(i)|\right)^{2}
$$

where $P(L)$ and $P(R)$ are the fractions of observations that split to the left and right respectively. If the expression is large, the split made each child node purer. Similarly, if the expression is small, the split made each child node similar to each other, and hence similar to the parent node, and so the split did not increase node purity.

- Node error - The node error is the fraction of misclassified classes at a node. If $j$ is the class with largest number of training samples at a node, the node error is

$$
1-p(j) .
$$

Value. To learn how value classes affect copy operations, see Copying Semantics

## CompactClassificationTree

| Examples | Construct a compact classification tree for the Fisher iris data, and compare the size of the resulting tree to that of the original tree: ```load fisheriris tree = ClassificationTree.fit(meas,species); ctree = compact(tree); t = whos('tree'); % t.bytes = size of tree in bytes c = whos('ctree'); % c.bytes = size of ctree in bytes [c.bytes t.bytes] ans = 6 8 1 8 1 3 9 1 3``` |
| :---: | :---: |
| See Also | ClassificationTree.fit \| ClassificationTree | compact |
| How To | - Chapter 13, "Nonparametric Supervised Learning" |

## CompactRegressionEnsemble

## Purpose Compact regression ensemble class

Description

Properties

Compact version of a regression ensemble (of class RegressionEnsemble). The compact version does not include the data for training the regression ensemble. Therefore, you cannot perform some tasks with a compact regression ensemble, such as cross validation. Use a compact regression ensemble for making predictions (regressions) of new data.

## Construction <br> cens = compact(ens) constructs a compact decision ensemble from a full decision ensemble. <br> Input Arguments

## ens

A regression ensemble created by fitensemble.
CategoricalPredictors
List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to $p$, where $p$ is the number of columns of $X$.

CombineWeights
A string describing how the ensemble combines learner predictions.

## NTrained

Number of trained learners in the ensemble, a positive scalar.

## PredictorNames

A cell array of names for the predictor variables, in the order in which they appear in $X$.

## ResponseName

A string with the name of the response variable $Y$.
ResponseTransform

## CompactRegressionEnsemble

Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none' means @(x)x.

Add or change a ResponseTransform function by dot addressing:

```
cens.ResponseTransform = @function
```


## Trained

The trained learners, a cell array of compact regression models.

## TrainedWeights

A numeric vector of weights the ensemble assigns to its learners. The ensemble computes predicted response by aggregating weighted predictions from its learners.

## Methods

| loss | Regression error |
| :--- | :--- |
| predict | Predict response of ensemble |
| predictorImportance | Estimates of predictor importance |

$\begin{array}{ll}\text { Copy } & \text { Value. To learn how value classes affect copy operations, see Copying } \\ \text { Semantics } & \text { Objects in the MATLAB Programming Fundamentals documentation. }\end{array}$
Examples Construct a regression ensemble for the carsmall data. Make a compact version of the ensemble, and compare its size to that of the full ensemble:

```
load carsmall
learner = RegressionTree.template('MinParent',20);
ens = fitensemble([Weight, Cylinders],MPG,...
    'LSBoost',100,learner,'PredictorNames',{'W','C'},...
    'categoricalpredictors',2);
cens = compact(ens);
ee = whos('ens'); % ee.bytes = size of ensemble in bytes
cee = whos('cens');
```


## CompactRegressionEnsemble

[ee.bytes cee.bytes]
ans =
606903587096
See Also fitensemble | RegressionEnsemble | predict | compact
How To • Chapter 13, "Nonparametric Supervised Learning"

## CompactRegressionTree

Purpose Compact regression tree
Description
Constructionctree = compact(tree) constructs a compact decision tree from afull decision tree.
Input Arguments
tree
A decision tree constructed by RegressionTree.fit.

## Properties

## CategoricalPredictors

List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to $p$, where $p$ is the number of columns of $X$.

## CatSplit

An $n$-by- 2 cell array, where $n$ is the number of nodes in tree. Each row in CatSplit gives left and right values for a categorical split. For each branch node $j$ based on a categorical predictor variable $z$, the left child is chosen if $z$ is in CatSplit ( $j, 1$ ) and the right child is chosen if $z$ is in CatSplit ( $\mathrm{j}, 2$ ). The splits are in the same order as nodes of the tree. Nodes for these splits can be found by running cuttype and selecting 'categorical' cuts from top to bottom.

## Children

An $n$-by- 2 array containing the numbers of the child nodes for each node in tree, where $n$ is the number of nodes. Leaf nodes have child node 0 .

[^4]
## CompactRegressionTree

An $n$-by- 2 cell array of the categories used at branches in tree, where $n$ is the number of nodes. For each branch node i based on a categorical predictor variable $x$, the left child is chosen if $x$ is among the categories listed in CutCategories $\{i, 1\}$, and the right child is chosen if $x$ is among those listed in CutCategories $\{i, 2\}$. Both columns of CutCategories are empty for branch nodes based on continuous predictors and for leaf nodes.

CutVar contains the cut points for 'continuous' cuts, and CutCategories contains the set of categories.

## CutPoint

An $n$-element vector of the values used as cut points in tree, where $n$ is the number of nodes. For each branch node i based on a continuous predictor variable $x$, the left child is chosen if CutPoint<v(i) and the right child is chosen if $x>=$ CutPoint(i). CutPoint is NaN for branch nodes based on categorical predictors and for leaf nodes.

## CutType

An $n$-element cell array indicating the type of cut at each node in tree, where $n$ is the number of nodes. For each node i, CutType\{i\} is:

- 'continuous' - If the cut is defined in the form $x<v$ for a variable x and cut point v .
- 'categorical' - If the cut is defined by whether a variable $x$ takes a value in a set of categories.
- ' ' - If i is a leaf node.

CutVar contains the cut points for 'continuous' cuts, and CutCategories contains the set of categories.

## CutVar

An $n$-element cell array of the names of the variables used for branching in each node in tree, where $n$ is the number of nodes.

## CompactRegressionTree

These variables are sometimes known as cut variables. For leaf nodes, CutVar contains an empty string.

CutVar contains the cut points for 'continuous' cuts, and CutCategories contains the set of categories.

## IsBranch

An $n$-element logical vector ib that is true for each branch node and false for each leaf node of tree.

## NodeErr

An $n$-element vector e of the errors of the nodes in tree, where $n$ is the number of nodes. $e(i)$ is the misclassification probability for node i.

## NodeMean

An $n$-element numeric array with mean values in each node of tree, where $n$ is the number of nodes in the tree. Every element in NodeMean is the average of the true $Y$ values over all observations in the node.

## NodeProb

An $n$-element vector p of the probabilities of the nodes in tree, where $n$ is the number of nodes. The probability of a node is computed as the proportion of observations from the original data that satisfy the conditions for the node. This proportion is adjusted for any prior probabilities assigned to each class.

## NodeSize

An $n$-element vector sizes of the sizes of the nodes in tree, where $n$ is the number of nodes. The size of a node is defined as the number of observations from the data used to create the tree that satisfy the conditions for the node.

## NumNodes

The number of nodes n in tree.
Parent

## CompactRegressionTree

An $n$-element vector $p$ containing the number of the parent node for each node in tree, where $n$ is the number of nodes. The parent of the root node is 0 .

## PredictorNames

A cell array of names for the predictor variables, in the order in which they appear in X .

## PruneList

An $n$-element numeric vector with the pruning levels in each node of tree, where $n$ is the number of nodes.

## ResponseName

Name of the response variable Y , a string.

## ResponseTransform

Function handle for transforming the raw response values (mean squared error). The function handle should accept a matrix of response values and return a matrix of the same size. The default string 'none' means $@(x) x$, or no transformation.

Add or change a ResponseTransform function by dot addressing:
ctree.ResponseTransform = @function

Risk
An $n$-element vector $r$ of the risk of the nodes in tree, where $n$ is the number of nodes. The risk $r(i)$ for node $i$ is the node error NodeErr(i) multiplied by the node probability NodeProb(i).

## SurrCutCategories

An $n$-element cell array of the categories used for surrogate splits in tree, where $n$ is the number of nodes in tree. For each node $k$, SurrCutCategories $\{\mathrm{k}\}$ is a cell array. The length of SurrCutCategories $\{\mathrm{k}\}$ is equal to the number of surrogate predictors found at this node. Every element of SurrCutCategories $\{k\}$ is either an empty string for a continuous

## CompactRegressionTree

surrogate predictor, or is a two-element cell array with categories for a categorical surrogate predictor. The first element of this two-element cell array lists categories assigned to the left child by this surrogate split, and the second element of this two-element cell array lists categories assigned to the right child by this surrogate split. The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutCategories contains an empty cell.

## SurrCutFlip

An $n$-element cell array of the numeric cut assignments used for surrogate splits in tree, where $n$ is the number of nodes in tree. For each node k, SurrCutFlip $\{k\}$ is a numeric vector. The length of SurrCutFlip\{k\} is equal to the number of surrogate predictors found at this node. Every element of SurrCutFlip\{k\} is either zero for a categorical surrogate predictor, or a numeric cut assignment for a continuous surrogate predictor. The numeric cut assignment can be either -1 or +1 . For every surrogate split with a numeric cut $C$ based on a continuous predictor variable $Z$, the left child is chosen if $Z<C$ and the cut assignment for this surrogate split is +1 , or if $Z \geq C$ and the cut assignment for this surrogate split is -1 . Similarly, the right child is chosen if $Z \geq C$ and the cut assignment for this surrogate split is +1 , or if $Z<C$ and the cut assignment for this surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutFlip contains an empty array.

## SurrCutPoint

An $n$-element cell array of the numeric values used for surrogate splits in tree, where $n$ is the number of nodes in tree. For each node k, SurrCutPoint $\{\mathrm{k}\}$ is a numeric vector. The length of SurrCutPoint $\{\mathrm{k}\}$ is equal to the number of surrogate predictors found at this node. Every element of SurrCutPoint \{k\} is either NaN for a categorical surrogate predictor, or a numeric cut for a

## CompactRegressionTree

continuous surrogate predictor. For every surrogate split with a numeric cut $C$ based on a continuous predictor variable $Z$, the left child is chosen if $Z<C$ and SurrCutFlip for this surrogate split is -1 . Similarly, the right child is chosen if $Z \geq C$ and SurrCutFlip for this surrogate split is +1 , or if $Z<C$ and SurrCutFlip for this surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables returned by SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutPoint contains an empty cell.

## SurrCutType

An $n$-element cell array indicating types of surrogate splits at each node in tree, where $n$ is the number of nodes in tree. For each node $k$, SurrCutType $\{k\}$ is a cell array with the types of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutType contains an empty cell. A surrogate split type can be either 'continuous' if the cut is defined in the form $Z<V$ for a variable $Z$ and cut point $V$ or 'categorical' if the cut is defined by whether $Z$ takes a value in a set of categories.

## SurrCutVar

An $n$-element cell array of the names of the variables used for surrogate splits in each node in tree, where $n$ is the number of nodes in tree. Every element of SurrCutVar is a cell array with the names of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutVar contains an empty cell.

## CompactRegressionTree

## SurrVarAssoc

An $n$-element cell array of the predictive measures of association for surrogate splits in tree, where $n$ is the number of nodes in tree. For each node k, SurrVarAssoc $\{k\}$ is a numeric vector. The length of SurrVarAssoc $\{k\}$ is equal to the number of surrogate predictors found at this node. Every element of SurrVarAssoc \{k\} gives the predictive measure of association between the optimal split and this surrogate split. The order of the surrogate split variables at each node is the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrVarAssoc contains an empty cell.

## Methods

| loss | Regression error <br> meanSurrVarAssoc |
| :--- | :--- |
| Mean predictive measure of <br> association for surrogate splits in <br> decision tree |  |
| predict | Predict response of regression <br> tree |
| predictorImportance | Estimates of predictor importance <br> view |
| View tree |  |

## Copy Value. To learn how value classes affect copy operations, see Copying Semantics Objects in the MATLAB Programming Fundamentals documentation.

## Examples Construct a regression tree for the carsmall data. Make a compact

 version of the tree, and compare its size to that of the full tree:```
load carsmall
tree = RegressionTree.fit([Weight, Cylinders],MPG,...
    'MinParent',20,...
    'PredictorNames',{'W','C'});
ctree = compact(tree);
t = whos('tree'); % t.bytes = size of tree in bytes
```


## CompactRegressionTree

```
c = whos('ctree'); % c.bytes = size of ctree in bytes
[c.bytes t.bytes]
ans =
    4972 8173
```


## See Also RegressionTree.fit | RegressionTree | compact

How To • Chapter 13, "Nonparametric Supervised Learning"

## CompactTreeBagger

## Purpose <br> Description

Construction
CompactTreeBagger

combine<br>error<br>margin<br>mdsProx<br>meanMargin<br>outlierMeasure<br>predict

Create CompactTreeBagger object

Combine two ensembles
Error (misclassification probability or MSE)

Classification margin
Multidimensional scaling of proximity matrix

Mean classification margin
Outlier measure for data
Predict response

## CompactTreeBagger

proximity

SetDefaultYfit

Proximity matrix for data
Set default value for predict

## Properties

## ClassNames

The ClassNames property is a cell array containing the class names for the response variable $Y$ supplied to TreeBagger. This property is empty for regression trees.

## DeltaCritDecisionSplit

The DeltaCritDecisionSplit property is a numeric array of size 1-by-Nvars of changes in the split criterion summed over splits on each variable, averaged across the entire ensemble of grown trees.

See also TreeBagger.DeltaCritDecisionSplit, classregtree.varimportance

## DefaultYfit

The DefaultYfit property controls what predicted value CompactTreeBagger returns when no prediction is possible, for example when the predict method needs to predict for an observation which has only false values in the matrix supplied through 'useifort' argument.

For classification, you can set this property to either ' ' or 'MostPopular'. If you choose 'MostPopular' (default), the property value becomes the name of the most probable class in the training data.

For regression, you can set this property to any numeric scalar. The default is the mean of the response for the training data.

See also predict, setDefaultYfit, TreeBagger.DefaultYfit. Method

The Method property is 'classification' for classification ensembles and 'regression' for regression ensembles.

## CompactTreeBagger

## NTrees

The NTrees property is a scalar equal to the number of decision trees in the ensemble.

## NVarSplit

The NVarSplit property is a numeric array of size 1-by-Nvars, where every element gives a number of splits on this predictor summed over all trees.

## Trees

The Trees property is a cell array of size NTrees-by- 1 containing the trees in the ensemble.

## VarAssoc

The VarAssoc property is a matrix of size Nuars-by-Nvars with predictive measures of variable association, averaged across the entire ensemble of grown trees. If you grew the ensemble setting 'surrogate' to 'on', this matrix for each tree is filled with predictive measures of association averaged over the surrogate splits. If you grew the ensemble setting 'surrogate' to 'off' (default), VarAssoc is diagonal.

See also classregtree.MeanSurrVarAssoc

## VarNames

The VarNames property is a cell array containing the names of the predictor variables (features). These names are taken from the optional 'names' parameter that supplied to TreeBagger. The default names are 'x1', 'x2', etc.

## Copy Semantics

## See Also

## CompactTreeBagger

## How To

- "Ensemble Methods" on page 13-51
- "Classification Trees and Regression Trees" on page 13-27
- "Grouped Data" on page 2-34


## CompactTreeBagger

Purpose Create CompactTreeBagger object
Description When you use the TreeBagger constructor to grow trees, it creates aCompactTreeBagger object. You can obtain the compact object from thefull TreeBagger object using the TreeBagger/compact method. You donot create an instance of CompactTreeBagger directly.
See Also TreeBagger
How To - Grouped Data- "Ensemble Methods" on page 13-51

## TreeBagger.ComputeOOBPrediction property

## Purpose Flag to compute out-of-bag predictions

## Description The Compute00BPrediction property is a logical flag specifying

 whether out-of-bag predictions for training observations should be computed. The default is false.If this flag is true, the following properties are available:

- OOBIndices
- OOBInstanceWeight

If this flag is true, the following methods can be called:

- oobError
- oobMargin
- oobMeanMargin

See Also oobError | OOBIndices | OOBInstanceWeight | oobMargin | oobMeanMargin

## TreeBagger.ComputeOOBVarImp property

| Purpose | Flag to compute out-of-bag variable importance |
| :--- | :--- |
| Description | The Compute00BVarImp property is a logical flag specifying whether <br> TreeBagger should compute out-of-bag estimates of variable <br> importance. The default is false. |
|  | If this flag is true, the following properties are available: |
| - OOBPermutedVarDeltaError |  |
| - OOBPermutedVarDeltaMeanMargin |  |

Purpose Confusion matrix

| Syntax | $C=$ confusionmat(group, grouphat) |
| :---: | :--- |
| $C=$ confusionmat(group, grouphat, ${ }^{\prime}$ order', grouporder) |  |
| $[C$, order $]=$ confusionmat $(\ldots)$ |  |

## Description

## Examples

C = confusionmat (group, grouphat) returns the confusion matrix C determined by the known and predicted groups in group and grouphat, respectively. group and grouphat are grouping variables with the same number of observations, as described in "Grouped Data" on page $2-34$. Input vectors must be of the same type. C is a square matrix with size equal to the total number of distinct elements in group and grouphat. $C(i, j)$ is a count of observations known to be in group i but predicted to be in group j. Group indices and their order are the same for the rows and columns of C , computed by grp2idx using grp2idx(group; grouphat). NaN, empty, or 'undefined' groups are not counted.

C = confusionmat(group,grouphat,'order',grouporder) uses grouporder to order the rows and columns of C. grouporder is a grouping variable containing all of the distinct elements in group and grouphat. If grouporder contains elements that are not in group or grouphat, the corresponding entries in $C$ will be 0 .
[C,order] = confusionmat (...) also returns the order of the rows and columns of $C$ in a variable order the same type as group and grouphat.

## Example 1

Display the confusion matrix for data with two misclassifications and one missing classification:

```
g1 = [llllllll
g2 = [1 1 1 2 3 4 NaN]'; % Predicted groups
[C,order] = confusionmat(g1,g2)
C =
```

| 2 | 0 | 0 | 0 |
| :---: | :---: | :---: | :---: |
| 0 | 1 | 1 | 0 |
| 0 | 0 | 0 | 1 |
| 0 | 0 | 0 | 0 |
| order $=$ |  |  |  |
| 1 |  |  |  |
| 2 |  |  |  |
| 3 |  |  |  |
| 4 |  |  |  |

## Example 2

Randomize the measurements and groups in Fisher's iris data:

```
load fisheriris
numObs = length(species);
p = randperm(numObs);
meas = meas(p,:);
species = species(p);
```

Use classify to classify measurements in the second half of the data, using the first half of the data for training:

```
half = floor(numObs/2);
training = meas(1:half,:);
trainingSpecies = species(1:half);
sample = meas(half+1:end,:);
grouphat = classify(sample,training,trainingSpecies);
```

Display the confusion matrix for the resulting classification:

```
group = species(half+1:end);
[C,order] = confusionmat(group,grouphat)
C =
    22 0 0
    22 0
    0 0 29
order =
    virginica'
```


# 'versicolor' <br> 'setosa' 

See Also crosstab | grp2idx
How To . "Grouped Data" on page 2-34

## Purpose <br> Syntax <br> Description

Shewhart control charts

```
controlchart(X)
controlchart(x,group)
controlchart(X,group)
[stats,plotdata] = controlchart(x,[group])
controlchart(x,group,'name',value)
```

controlchart (X) produces an xbar chart of the measurements in matrix $X$. Each row of $X$ is considered to be a subgroup of measurements containing replicate observations taken at the same time. The rows should be in time order. If X is a time series object, the time samples should contain replicate observations.

The chart plots the means of the subgroups in time order, a center line (CL) at the average of the means, and upper and lower control limits (UCL, LCL) at three standard errors from the center line. The standard error is the estimated process standard deviation divided by the square root of the subgroup size. Process standard deviation is estimated from the average of the subgroup standard deviations. Out of control measurements are marked as violations and drawn with a red circle. Data cursor mode is enabled, so clicking any data point displays information about that point.
controlchart (x, group) accepts a grouping variable group for a vector of measurements x. (See "Grouped Data" on page 2-34.) group is a categorical variable, vector, string array, or cell array of strings the same length as $x$. Consecutive measurements $x(n)$ sharing the same value of group( $n$ ) for $1 \leq n \leq$ length ( $x$ ) are defined to be a subgroup. Subgroups can have different numbers of observations.
controlchart (X, group) accepts a grouping variable group for a matrix of measurements in $X$. In this case, group is only used to label the time axis; it does not change the default grouping by rows.
[stats, plotdata] = controlchart(x,[group]) returns a structure stats of subgroup statistics and parameter estimates, and a structure plotdata of plotted values. plotdata contains one record for each chart.

The fields in stats and plotdata depend on the chart type.
The fields in stats are selected from the following:

- mean - Subgroup means
- std - Subgroup standard deviations
- range - Subgroup ranges
- n - Subgroup size, or total inspection size or area
- i - Individual data values
- ma - Moving averages
- mr - Moving ranges
- count - Count of defects or defective items
- mu - Estimated process mean
- sigma - Estimated process standard deviation
- p - Estimated proportion defective
- m - Estimated mean defects per unit

The fields in plotdata are the following:

- pts — Plotted point values
- cl - Center line
- lcl - Lower control limit
- ucl - Upper control limit
- se - Standard error of plotted point
- n - Subgroup size
- ooc - Logical that is true for points that are out of control
controlchart(x,group, 'name', value) specifies one or more of the following optional parameter name/value pairs, with name in single quotes:
- charttype - The name of a chart type chosen from among the following:
- 'xbar' - Xbar or mean
- 's' - Standard deviation
- 'r' - Range
- 'ewma' - Exponentially weighted moving average
- 'i' - Individual observation
- 'mr' - Moving range of individual observations
- 'ma' - Moving average of individual observations
- 'p' - Proportion defective
- 'np' - Number of defectives
- 'u' - Defects per unit
- 'c' - Count of defects

Alternatively, a parameter can be a cell array listing multiple compatible chart types. There are four sets of compatible types:

- 'xbar', 's', 'r', and 'ewma'
- 'i', 'mr', and 'ma'
- 'p' and 'np'
- 'u' and 'c'
- display - Either 'on ' (default) to display the control chart, or 'off' to omit the display
- label - A string array or cell array of strings, one per subgroup. This label is displayed as part of the data cursor for a point on the plot.
- lambda - A parameter between 0 and 1 controlling how much the current prediction is influenced by past observations in an EWMA plot. Higher values of 'lambda' give less weight to past observations and more weight to the current observation. The default is 0.4.
- limits' - A three-element vector specifying the values of the lower control limit, center line, and upper control limits. Default is to estimate the center line and to compute control limits based on the estimated value of sigma. Not permitted if there are multiple chart types.
- mean - Value for the process mean, or an empty value (default) to estimate the mean from $X$. This is the $p$ parameter for $p$ and $n p$ charts, the mean defects per unit for $u$ and $c$ charts, and the normal mu parameter for other charts.
- nsigma - The number of sigma multiples from the center line to a control limit. Default is 3 .
- parent - The handle of the axes to receive the control chart plot. Default is to create axes in a new figure. Not permitted if there are multiple chart types.
- rules - The name of a control rule, or a cell array containing multiple control rule names. These rules, together with the control limits, determine if a point is marked as out of control. The default is to apply no control rules, and to use only the control limits to decide if a point is out of control. See controlrules for more information. Control rules are applied to charts that measure the process level (xbar, i, c, u, p, and np) rather than the variability ( $r, s$ ), and they are not applied to charts based on moving statistics (ma, mr, ewma).
- sigma - Either a value for sigma, or a method of estimating sigma chosen from among 'std' (the default) to use the average within-subgroup standard deviation, 'range' to use the average subgroup range, and 'variance' to use the square root of the pooled variance. When creating $i, m r$, or ma charts for data not in subgroups, the estimate is always based on a moving range.
- specs - A vector specifying specification limits. Typically this is a two-element vector of lower and upper specification limits. Since specification limits typically apply to individual measurements, this parameter is primarily suitable for i charts. These limits are not plotted on $r$, $s$, or $m r$ charts.
- unit - The total number of inspected items for $p$ and $n p$ charts, and the size of the inspected unit for $u$ and $c$ charts. In both cases $X$ must be the count of the number of defects or defectives found. Default is 1 for $u$ and $c$ charts. This argument is required (no default) for $p$ and $n p$ charts.
- width - The width of the window used for computing the moving ranges and averages in mr and ma charts, and for computing the sigma estimate in $\mathrm{i}, \mathrm{mr}$, and ma charts. Default is 5 .


## Examples Create xbar and $r$ control charts for the data in parts.mat:

```
load parts
st = controlchart(runout,'chart',{'xbar' 'r'});
```

Control charts


Display the process mean and standard deviation:

```
fprintf('Parameter estimates: mu = %g, sigma = %g\n',st.mu,st.sigma);
Parameter estimates: mu = -0.0863889, sigma = 0.130215
```

controlrules

- "Grouped Data" on page 2-34


## Purpose Western Electric and Nelson control rules

```
Syntax
R = controlrules('rules',x,cl,se)
[R,RULES] = controlrules('rules',x,cl,se)
```


## Description

$\mathrm{R}=$ controlrules('rules', $\mathrm{x}, \mathrm{cl}, \mathrm{se})$ determines which points in the vector x violate the control rules in rules. cl is a vector of center-line values. se is a vector of standard errors. (Typically, control limits on a control chart are at the values $\mathrm{cl}-3^{*}$ se and $\mathrm{cl}+3^{*} \mathrm{se}$.) rules is the name of a control rule, or a cell array containing multiple control rule names, from the list below. If $x$ has $n$ values and rules contains $m$ rules, then $R$ is an $n$-by- $m$ logical array, with $R(i, j)$ assigned the value 1 if point i violates rule $\mathrm{j}, 0$ if it does not.
The following are accepted values for rules (specified inside single quotes):

- we1-1 point above cl + 3*se
- we2 - 2 of 3 above cl + 2*se
- we3 - 4 of 5 above cl + se
- we4 - 8 of 8 above cl
- we5 - 1 below cl 3*se
- we6 - 2 of 3 below cl $2 *$ se
- we $7-4$ of 5 below cl se
- we8 - 8 of 8 below cl
- we9 - 15 of 15 between cl se and cl + se
- we $10-8$ of 8 below cl se or above cl + se
- $\mathrm{n} 1-1$ point below $\mathrm{cl} 3^{*}$ se or above $\mathrm{cl}+3^{*}$ se
- n2 - 9 of 9 on the same side of cl
- n3 - 6 of 6 increasing or decreasing
- n4 - 14 alternating up/down
- n5-2 of 3 below cl 2*se or above cl + 2*se, same side
- n6-4 of 5 below cl se or above cl + se, same side
- n7-15 of 15 between cl se and cl+se
- n8 - 8 of 8 below cl se or above $\mathrm{cl}+\mathrm{se}$, either side
- we - All Western Electric rules
- n - All Nelson rules

For multi-point rules, a rule violation at point i indicates that the set of points ending at point $i$ triggered the rule. Point $i$ is considered to have violated the rule only if it is one of the points violating the rule's condition.

Any points with NaN as their x, cl, or se values are not considered to have violated rules, and are not counted in the rules for other points.

Control rules can be specified in the controlchart function as values for the 'rules' parameter.
[R,RULES] = controlrules('rules', $x, c l, s e)$ returns a cell array of text strings RULES listing the rules applied.

## Examples

Create an xbar chart using the we2 rule to mark out of control measurements:

```
load parts;
st = controlchart(runout,'rules','we2');
x = st.mean;
cl = st.mu;
se = st.sigma./sqrt(st.n);
hold on
plot(cl+2*se,'m')
```

XBAR control chart


Use controlrules to identify the measurements that violate the control rule:

```
R = controlrules('we2',x,cl,se);
I = find(R)
I =
    21
    23
```


## controlrules

24
25
26

See Also
controlchart

Purpose Determine if algorithm converged
Description Logical true if the algorithm has converged; logical false if the algorithm has not converged.

Note This property applies only to gmdistribution objects constructed with fit.

Purpose Cophenetic correlation coefficient
$\begin{array}{ll}\text { Syntax } & c=\operatorname{cophenet}(Z, Y) \\ {[c, d]=\operatorname{cophenet}(Z, Y)}\end{array}$
Description $\quad c=$ cophenet $(Z, Y)$ computes the cophenetic correlation coefficient for the hierarchical cluster tree represented by $Z . Z$ is the output of the linkage function. $Y$ contains the distances or dissimilarities used to construct $Z$, as output by the pdist function. $Z$ is a matrix of size ( $m-1$ )-by-3, with distance information in the third column. Y is a vector of size $m *(m-1) / 2$.
$[c, d]=$ cophenet $(Z, Y)$ returns the cophenetic distances $d$ in the same lower triangular distance vector format as Y .

The cophenetic correlation for a cluster tree is defined as the linear correlation coefficient between the cophenetic distances obtained from the tree, and the original distances (or dissimilarities) used to construct the tree. Thus, it is a measure of how faithfully the tree represents the dissimilarities among observations.

The cophenetic distance between two observations is represented in a dendrogram by the height of the link at which those two observations are first joined. That height is the distance between the two subclusters that are merged by that link.

The output value, c , is the cophenetic correlation coefficient. The magnitude of this value should be very close to 1 for a high-quality solution. This measure can be used to compare alternative cluster solutions obtained using different algorithms.
The cophenetic correlation between $Z(:, 3)$ and $Y$ is defined as

$$
c=\frac{\sum_{i<j}\left(Y_{i j}-y\right)\left(Z_{i j}-z\right)}{\sqrt{\sum_{i<j}\left(Y_{i j}-y\right)^{2} \sum_{i<j}\left(Z_{i j}-z\right)^{2}}}
$$

where:

- $Y_{i j}$ is the distance between objects $i$ and $j$ in $Y$.
- $Z_{i j}$ is the cophenetic distance between objects $i$ and $j$, from $Z(:, 3)$.
- $y$ and $z$ are the average of $Y$ and $Z(:, 3)$, respectively.


## Examples

```
X = [rand(10,3); rand(10,3)+1; rand(10,3)+2];
Y = pdist(X);
Z = linkage(Y,'average');
% Compute Spearman's rank correlation between the
% dissimilarities and the cophenetic distances
[c,D] = cophenet(Z,Y);
r = corr(Y',D','type','spearman')
r =
    0.8279
```

See Also

cluster | dendrogram | inconsistent | linkage | pdist | squareform

## Purpose Copula cumulative distribution function

```
Syntax
Y = copulacdf('Gaussian', U,rho)
Y = copulacdf('t',U,rho,NU)
Y = copulacdf('family', U,alpha)
```


## Description

## Examples

```
u = linspace(0,1,10);
[U1,U2] = meshgrid(u,u);
F = copulacdf('Clayton',[U1(:) U2(:)],1);
surf(U1,U2,reshape(F,10,10))
xlabel('u1')
ylabel('u2')
```



See Also
copulapdf | copularnd | copulastat | copulaparam

Purpose Fit copula to data

```
Syntax RHOHAT = copulafit('Gaussian',U)
[RHOHAT,nuhat] = copulafit('t',U)
[RHOHAT,nuhat,nuci] = copulafit('t',U)
paramhat = copulafit('family',U)
[paramhat,paramci] = copulafit('family',U)
[...] = copulafit(...,'alpha',alpha)
[...] = copulafit('t',U,'Method','ApproximateML')
[...] = copulafit(...,'Options',options)
```


## Description

RHOHAT = copulafit('Gaussian', U) returns an estimate RHOHAT of the matrix of linear correlation parameters for a Gaussian copula, given data in $\mathrm{U} . \mathrm{U}$ is an $n$-by- $p$ matrix of values in the open interval $(0,1)$ representing $n$ points in the $p$-dimensional unit hypercube.
[RHOHAT, nuhat] = copulafit('t', U ) returns an estimate RHOHAT of the matrix of linear correlation parameters for a $t$ copula and an estimate nuhat of the degrees of freedom parameter, given data in $U$. U is an $n$-by- $p$ matrix of values in the open interval $(0,1)$ representing $n$ points in the $p$-dimensional unit hypercube.
[RHOHAT, nuhat, nuci] = copulafit('t', U) also returns an approximate $95 \%$ confidence interval nuci for the degrees of freedom parameter estimated in nuhat.
paramhat = copulafit('family', U) returns an estimate paramhat of the copula parameter for an Archimedean copula specified by family, given data in $U$. $U$ is an $n$-by- 2 matrix of values in the open interval $(0,1)$ representing $n$ points in the unit square. family is one of Clayton, Frank, or Gumbel.
[paramhat,paramci] = copulafit('family',U) also returns an approximate $95 \%$ confidence interval paramci for the copula parameter estimated in paramhat.
[...] = copulafit(...,'alpha', alpha) returns approximate 100*(1-alpha) \% confidence intervals in nuci or paramci.

Note By default, copulafit uses maximum likelihood to fit a copula to $U$. When $U$ contains data transformed to the unit hypercube by parametric estimates of their marginal cumulative distribution functions, this is known as the Inference Functions for Margins (IFM) method. When U contains data transformed by the empirical cdf (see ecdf), this is known as Canonical Maximum Likelihood (CML).
[...] = copulafit('t',U,'Method','ApproximateML') fits a $t$ copula for large samples $U$ by maximizing an objective function that approximates the profile log-likelihood for the degrees of freedom parameter (see [1]). This method can be significantly faster than maximum likelihood, but the estimates and confidence limits may not be accurate for small to moderate sample sizes.
[...] = copulafit(...,'Options',options) specifies control parameters for the iterative parameter estimation algorithm using an options structure options as created by statset. Type statset('copulafit') at the command prompt for fields and default values used by copulafit. This argument is not applicable to the 'Gaussian' family.

## References

[1] Bouye, E., V. Durrleman, A. Nikeghbali, G. Riboulet, and T. Roncalli. "Copulas for Finance: A Reading Guide and Some Applications." Working Paper. Groupe de Recherche Operationnelle, Credit Lyonnais, 2000.

## Examples Load and plot simulated stock return data:

```
load stockreturns
x = stocks(:,1);
y = stocks(:,2);
scatterhist(x,y)
```



Transform the data to the copula scale (unit square) using a kernel estimator of the cumulative distribution function:

```
u = ksdensity(x,x,'function','cdf');
v = ksdensity(y,y,'function','cdf');
scatterhist(u,v)
xlabel('u')
ylabel('v')
```



Fit a $t$ copula:

```
[Rho,nu] = copulafit('t',[u v],'Method','ApproximateML')
Rho =
    1.0000 0.7220
    0.7220 1.0000
nu =
    2.8934e+006
```

Generate a random sample from the $t$ copula:

```
r = copularnd('t',Rho,nu,1000);
```

```
u1 = r(:,1);
v1 = r(:,2);
scatterhist(u1,v1)
xlabel('u')
ylabel('v')
set(get(gca,'children'),'marker','.')
```




Transform the random sample back to the original scale of the data:
x1 = ksdensity(x,u1,'function','icdf');

```
y1 = ksdensity(y,v1,'function','icdf');
scatterhist(x1,y1)
set(get(gca,'children'),'marker','.')
```



See Also
ecdf | copulacdf | copulaparam | copulapdf | copularnd | copulastat

# Purpose <br> Copula parameters as function of rank correlation 

Syntax
Description

## Examples

```
rho = copulaparam('Gaussian',R)
rho = copulaparam('t',R,NU)
alpha = copulaparam('family',R)
[...] = copulaparam(...,'type',type)
```

rho = copulaparam('Gaussian', R) returns the linear correlation parameters rho corresponding to a Gaussian copula having Kendall's rank correlation $R$. If $R$ is a scalar correlation coefficient, rho is a scalar correlation coefficient corresponding to a bivariate copula. If $R$ is a p -by-p correlation matrix, rho is a p -by-p correlation matrix.
rho = copulaparam('t',R,NU) returns the linear correlation parameters rho corresponding to a $t$ copula having Kendall's rank correlation R and degrees of freedom NU. If R is a scalar correlation coefficient, rho is a scalar correlation coefficient corresponding to a bivariate copula. If $R$ is a $p$-by-p correlation matrix, rho is a $p-b y-p$ correlation matrix.
alpha = copulaparam('family',R) returns the copula parameter alpha corresponding to a bivariate Archimedean copula having Kendall's rank correlation R. R is a scalar. family is one of Clayton, Frank, or Gumbel.
[...] = copulaparam(...,'type',type) assumes R is the specified type of rank correlation. type is 'Kendall' for Kendall's tau or 'Spearman' for Spearman's rho.
copulaparam uses an approximation to Spearman's rank correlation for some copula families when no analytic formula exists. The approximation is based on a smooth fit to values computed using Monte Carlo simulations.

Get the linear correlation coefficient corresponding to a bivariate Gaussian copula having a rank correlation of -0.5.

```
tau = -0.5
rho = copulaparam('gaussian',tau)
```

```
rho =
    -0.7071
% Generate dependent beta random values using that copula
u = copularnd('gaussian',rho,100);
b = betainv(u,2,2);
% Verify that the sample has a rank correlation
% approximately equal to tau
tau_sample = corr(b,'type','k')
tau_sample =
    1.0000 -0.4638
    -0.4638 1.0000
```

See Also
copulacdf | copulapdf | copularnd | copulastat

Purpose Copula probability density function


## Description

$Y=$ copulapdf('Gaussian', U, rho) returns the probability density of the Gaussian copula with linear correlation parameters rho, evaluated at the points in $U . U$ is an $n$-by-p matrix of values in [ 0,1 ], representing $n$ points in the $p$-dimensional unit hypercube. rho is a $p$-by-p correlation matrix. If $U$ is an $n$-by- 2 matrix, rho may also be a scalar correlation coefficient.
$Y=$ copulapdf('t', $U, r h o, N U)$ returns the probability density of the $t$ copula with linear correlation parameters rho and degrees of freedom parameter NU , evaluated at the points in $\mathrm{U} . \mathrm{U}$ is an $n$-by-p matrix of values in $[0,1]$. rho is a $p$-by-p correlation matrix. If $U$ is an $n$-by- 2 matrix, rho may also be a scalar correlation coefficient.
$\mathrm{Y}=$ copulapdf('family', $\mathrm{U}, \mathrm{alpha})$ returns the probability density of the bivariate Archimedean copula determined by family, with scalar parameter alpha, evaluated at the points in U. family is Clayton, Frank, or Gumbel. $U$ is an $n$-by- 2 matrix of values in $[0,1]$.

## Examples

```
u = linspace(0,1,10);
[U1,U2] = meshgrid(u,u);
F = copulapdf('Clayton',[U1(:) U2(:)],1);
surf(U1,U2,reshape(F,10,10))
xlabel('u1')
ylabel('u2')
```



See Also
copulacdf | copulaparam | copularnd | copulastat
Purpose Copula rank correlation

| Syntax | $\mathrm{R}=$ copulastat('Gaussian', rho) |
| :---: | :---: |
|  | $\mathrm{R}=$ copulastat('t', rho, NU) |
|  | R = copulastat('family',alpha) |
|  | R = copulastat(...,'type','type') |

## Description

## Examples

R = copulastat('Gaussian', rho) returns the Kendall's rank correlation R that corresponds to a Gaussian copula having linear correlation parameters rho. If rho is a scalar correlation coefficient, $R$ is a scalar correlation coefficient corresponding to a bivariate copula. If $r$ ho is a $p-b y-p$ correlation matrix, $R$ is a $p-b y-p$ correlation matrix.
$R=$ copulastat('t', rho, NU) returns the Kendall's rank correlation $R$ that corresponds to a $t$ copula having linear correlation parameters rho and degrees of freedom NU. If rho is a scalar correlation coefficient, R is a scalar correlation coefficient corresponding to a bivariate copula. If $r$ ho is a $p$-by- $p$ correlation matrix, $R$ is a $p-b y-p$ correlation matrix.

R = copulastat('family',alpha) returns the Kendall's rank correlation R that corresponds to a bivariate Archimedean copula with scalar parameter alpha. family is one of Clayton, Frank, or Gumbel.

R = copulastat(...,'type','type') returns the specified type of rank correlation. type is Kendall to compute Kendall's tau, or Spearman to compute Spearman's rho.
copulastat uses an approximation to Spearman's rank correlation for some copula families when no analytic formula exists. The approximation is based on a smooth fit to values computed using Monte-Carlo simulations.

Get the theoretical rank correlation coefficient for a bivariate.

```
%Gaussian copula with linear correlation parameter rho
rho = -.7071;
tau = copulastat('gaussian',rho)
tau =
    -0.5000
```

```
    % Generate dependent beta random values using that copula
    u = copularnd('gaussian',rho,100);
    b = betainv(u,2,2);
    % Verify that the sample has a rank correlation
    % approximately equal to tau
    tau_sample = corr(b,'type','k')
    tau_sample =
        1.0000 -0.5265
        -0.5265 1.0000
```

See Also copulacdf | copulaparam | copulapdf | copularnd

Purpose Copula random numbers

Syntax $\quad$| $U$ | $=$ copularnd ('Gaussian', rho, $N)$ |
| ---: | :--- |
| $U$ | $=$ copularnd $(' t ', r h o, N U, N)$ |
| $U$ | $=$ copularnd $(' f a m i l y ', ~ a l p h a, N)$ |

## Description

## Examples

Determine the linear correlation parameter corresponding to a bivariate Gaussian copula having a rank correlation of -0.5.

```
tau = -0.5
rho = copulaparam('gaussian',tau)
    rho =
        -0.7071
% Generate dependent beta random values using that copula
u = copularnd('gaussian',rho,100);
b = betainv(u,2,2);
```

```
% Verify that the sample has a rank correlation
% approximately equal to tau
tau_sample = corr(b,'type','kendall')
tau_sample =
        1.0000 -0.4537
    -0.4537 1.0000
```


## See Also

copulacdf | copulaparam | copulapdf | copulastat

## Purpose <br> Coordinate exchange

## Syntax

```
dCE = cordexch(nfactors,nruns)
[dCE,X] = cordexch(nfactors,nruns)
[dCE,X] = cordexch(nfactors,nruns,'model')
[dCE,X] = cordexch(...,'name',value)
```


## Description

dCE = cordexch(nfactors, nruns) uses a coordinate-exchange algorithm to generate a $D$-optimal design dCE with nruns runs (the rows of dCE) for a linear additive model with nfactors factors (the columns of dCE). The model includes a constant term.
[dCE, X] = cordexch(nfactors, nruns) also returns the associated design matrix $X$, whose columns are the model terms evaluated at each treatment (row) of dCE.
[dCE, X] = cordexch(nfactors, nruns,'model') uses the linear regression model specified in model. model is one of the following strings, specified inside single quotes:

- linear - Constant and linear terms. This is the default.
- interaction - Constant, linear, and interaction terms
- quadratic - Constant, linear, interaction, and squared terms
- purequadratic - Constant, linear, and squared terms

The order of the columns of $X$ for a full quadratic model with $n$ terms is:
1 The constant term
2 The linear terms in order $1,2, \ldots, n$
3 The interaction terms in order (1, 2), (1, 3), $\ldots,(1, n),(2,3), \ldots,(n-1, n)$
4 The squared terms in order $1,2, \ldots, n$
Other models use a subset of these terms, in the same order.

Alternatively, model can be a matrix specifying polynomial terms of arbitrary order. In this case, model should have one column for each factor and one row for each term in the model. The entries in any row of model are powers for the factors in the columns. For example, if a model has factors X1, X2, and X3, then a row [ $\left.\begin{array}{lll}0 & 1 & 2\end{array}\right]$ in model specifies the term (X1.^0).*(X2.^1).*(X3.^2). A row of all zeros in model specifies a constant term, which can be omitted.
[dCE,X] = cordexch(...,'name', value) specifies one or more optional name/value pairs for the design. Valid parameters and their values are listed in the following table. Specify name inside single quotes.

| name | Value |
| :--- | :--- |
| bounds | Lower and upper bounds for each factor, specified as <br> a 2-by-nfactors matrix. Alternatively, this value <br> can be a cell array containing nfactors elements, <br> each element specifying the vector of allowable <br> values for the corresponding factor. |
| categorical | Indices of categorical predictors. |
| display | Either 'on ' or ' off' to control display of the <br> iteration counter. The default is 'on '. |
| excludefun | Handle to a function that excludes undesirable <br> runs. If the function is $f$, it must support the syntax <br> $b=f(S)$, where $S$ is a matrix of treatments with <br> nfactors columns and $b$ is a vector of Boolean <br> values with the same number of rows as $S . b(i)$ is <br> true if the method should exclude $i$ th row $S$. |
| init | Initial design as a nruns-by-nfactors matrix. The <br> default is a randomly selected set of points. |
| levels | Vector of number of levels for each factor. |
| maxiter | Maximum number of iterations. The default is 10. |


| name | Value |
| :---: | :---: |
| tries | Number of times to try to generate a design from a new starting point. The algorithm uses random points for each try, except possibly the first. The default is 1 . |
| options | A structure that specifies whether to run in parallel, and specifies the random stream or streams. Create the options structure with statset. Option fields: <br> - UseParallel - Set to 'always' to compute in parallel. Default is 'never'. <br> - UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'. <br> - Streams - A RandStream object or cell array of such objects. If you do not specify Streams, cordexch uses the default stream or streams. If you choose to specify Streams, use a single object except in the case <br> - You have an open MATLAB pool <br> - UseParallel is 'always' <br> - UseSubstreams is 'never' <br> In that case, use a cell array the same size as the MATLAB pool. <br> For more information on using parallel computing, see Chapter 17, "Parallel Statistics". |

Both cordexch and rowexch use iterative search algorithms. They operate by incrementally changing an initial design matrix $X$ to increase $D=\left|X^{T} X\right|$ at each step. In both algorithms, there is randomness
built into the selection of the initial design and into the choice of the incremental changes. As a result, both algorithms may return locally, but not globally, $D$-optimal designs. Run each algorithm multiple times and select the best result for your final design. Both functions have a 'tries ' parameter that automates this repetition and comparison.

Unlike the row-exchange algorithm used by rowexch, cordexch does not use a candidate set. (Or rather, the candidate set is the entire design space.) At each step, the coordinate-exchange algorithm exchanges a single element of $X$ with a new element evaluated at a neighboring point in design space. The absence of a candidate set reduces demands on memory, but the smaller scale of the search means that the coordinate-exchange algorithm is more likely to become trapped in a local minimum.

## Examples

Suppose you want a design to estimate the parameters in the following three-factor, seven-term interaction model:

$$
y=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+\beta_{12} x_{1} x_{2}+\beta_{13} x_{1} x_{3}+\beta_{23} x_{2} x_{3}+\varepsilon
$$

Use cordexch to generate a $D$-optimal design with seven runs:

```
nfactors = 3;
nruns = 7;
[dCE,X] = cordexch(nfactors,nruns,'interaction','tries',10)
dCE =
\(X=\)\begin{tabular}{rrrrrrr}
-1 & 1 & 1 & & & & \\
-1 & -1 & -1 & & & & \\
1 & 1 & 1 & & & & \\
-1 & 1 & -1 & & & & \\
1 & -1 & 1 & & & & \\
1 & -1 & -1 & & & & \\
-1 & -1 & 1 & & & -1 & 1 \\
& & & 1 & -1 & 1 & 1 \\
1 & -1 & 1 & -1 & 1 & 1 & 1 \\
1 & -1 & -1 & -1 & 1 & 1 & 1 \\
1 & 1 & 1 & 1 & 1 \\
1 & -1 & 1 & -1 & -1 & 1 & -1
\end{tabular}
```

| 1 | 1 | -1 | 1 | -1 | 1 | -1 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 1 | -1 | -1 | -1 | -1 | 1 |
| 1 | -1 | -1 | 1 | 1 | -1 | -1 |

Columns of the design matrix $X$ are the model terms evaluated at each row of the design dCE. The terms appear in order from left to right: constant term, linear terms (1, 2, 3), interaction terms (12, 13, 23). Use $X$ to fit the model, as described in "Linear Regression" on page 9-3, to response data measured at the design points in dCE.

See Also rowexch | daugment | dcovary

## Purpose Linear or rank correlation

Syntax

```
RHO = corr(X)
RHO = corr(X,Y)
[RHO,PVAL] = corr(X,Y)
[RHO,PVAL] = corr(X,Y,'name',value)
```

Description
RHO $=\operatorname{corr}(\mathrm{X})$ returns a $p$-by- $p$ matrix containing the pairwise linear correlation coefficient between each pair of columns in the $n$-by- $p$ matrix $X$.

RHO $=\operatorname{corr}(\mathrm{X}, \mathrm{Y})$ returns a $p 1$-by- $p 2$ matrix containing the pairwise correlation coefficient between each pair of columns in the $n$-by- $p 1$ and $n$-by- $p 2$ matrices X and Y .
[RHO, PVAL] $=\operatorname{corr}(X, Y)$ also returns PVAL, a matrix of $p$-values for testing the hypothesis of no correlation against the alternative that there is a nonzero correlation. Each element of PVAL is the $p$ value for the corresponding element of RHO. If PVAL ( $i, j$ ) is small, say less than 0.05 , then the correlation $\mathrm{RHO}(\mathrm{i}, \mathrm{j})$ is significantly different from zero.
[RHO,PVAL] $=\operatorname{corr}(X, Y$, 'name', value) specifies one or more optional name/value pairs. Specify name inside single quotes. The following table lists valid parameters and their values.

| Parameter | Values |
| :--- | :--- |
| type | • 'Pearson' (the default) computes Pearson's <br> linear correlation coefficient |
|  | • 'Kendall' computes Kendall's tau |
| - 'Spearman' computes Spearman's rho |  |


| Parameter | Values |
| :--- | :--- |
|  | - 'pairwise 'computes RHO ( $i, j$ ) using rows <br> with no missing values in column i or $j$ |
| tail - The <br> alternative <br> hypothesis <br> against which to <br> compute $p$-values <br> for testing the <br> hypothesis of no <br> correlation | • 'both' - Correlation is not zero (the default) |

Using the 'pairwise' option for the rows parameter may return a matrix that is not positive definite. The 'complete' option always returns a positive definite matrix, but in general the estimates are based on fewer observations.
corr computes $p$-values for Pearson's correlation using a Student's $t$ distribution for a transformation of the correlation. This correlation is exact when X and Y are normal. corr computes $p$-values for Kendall's tau and Spearman's rho using either the exact permutation distributions (for small sample sizes), or large-sample approximations.
corr computes $p$-values for the two-tailed test by doubling the more significant of the two one-tailed $p$-values.

## References

[1] Gibbons, J.D. (1985) Nonparametric Statistical Inference, 2nd ed., M. Dekker.
[2] Hollander, M. and D.A. Wolfe (1973) Nonparametric Statistical Methods, Wiley.
[3] Kendall, M.G. (1970) Rank Correlation Methods, Griffin.
[4] Best, D.J. and D.E. Roberts (1975) "Algorithm AS 89: The Upper Tail Probabilities of Spearman's rho", Applied Statistics, 24:377-379.

See Also corrcoef | partialcorr | corrcov | tiedrank

## Purpose Convert covariance matrix to correlation matrix

```
Syntax
\(\mathrm{R}=\operatorname{corrcov}(\mathrm{C})\)
[R,sigma] = corrcov(C)
```

Description

Examples

See Also
$R=\operatorname{corrcov}(C)$ computes the correlation matrix $R$ corresponding to the covariance matrix C. C must be square, symmetric, and positive semi-definite.
[R,sigma] = corrcov(C) also computes the vector of standard deviations sigma.

Use cov and corrcoef to compute covariances and correlations, respectively, for sample data on weight and blood pressure (systolic, diastolic) in hospital.mat:

```
load hospital
X = [hospital.Weight hospital.BloodPressure];
C = cov(X)
C =
    706.0404 27.7879 41.0202
    27.7879 45.0622 23.8194
    41.0202 23.8194 48.0590
R = corrcoef(X)
R =
\begin{tabular}{lll}
1.0000 & 0.1558 & 0.2227 \\
0.1558 & 1.0000 & 0.5118 \\
0.2227 & 0.5118 & 1.0000
\end{tabular}
```

Compare R with the correlation matrix computed from C by corrcov:

```
corrcov(C)
```

ans $=$

| 1.0000 | 0.1558 | 0.2227 |
| :--- | :--- | :--- |
| 0.1558 | 1.0000 | 0.5118 |
| 0.2227 | 0.5118 | 1.0000 |

cov | corrcoef | corr | cholcov

## TreeBagger.Cost property

Purpose Misclassification costs
Description The cost property is a matrix with misclassification costs. This property is empty for ensembles of regression trees.
See Also classregtree

## gmdistribution.CovType property

## Purpose Type of covariance matrices

Description The string 'diagonal' if the covariance matrices are restricted to be diagonal; the string 'full' otherwise.

Purpose
Syntax
b $=\operatorname{coxphfit}(\mathrm{X}, \mathrm{y})$
b = coxphfit (X,y,'name', value)
[b,logl,H,stats] = coxphfit(...)
Cox proportional hazards regression
[b, igl, H stats] (

## Description

$\mathrm{b}=$ coxphfit $(\mathrm{X}, \mathrm{y})$ returns a $p$-by- 1 vector b of coefficient estimates for a Cox proportional hazards regression of the responses in $y$ on the predictors in X . X is an $n$-by- $p$ matrix of $p$ predictors at each of $n$ observations. y is an $n$-by- 1 vector of observed responses.

The phrase $h(t) * \exp (X * b)$ models the hazard rate for the distribution of $y$, where $h(t)$ is a common baseline hazard function. The model does not include a constant term, and $X$ cannot contain a column of 1 s .
b = coxphfit( $\mathrm{X}, \mathrm{y}$, 'name', value) specifies one or more optional parameter name/value pairs chosen from the following table. Specify name inside single quotes.

| Name | Value |
| :--- | :--- |
| baseline | The $X$ values at which to compute the baseline <br> hazard. Default is mean $(X)$, so the hazard at $X$ is <br> $h(t) * \exp ((X-m e a n(X)) * b)$. Enter 0 to compute <br> the baseline relative to 0 , so the hazard at $X$ is <br> $h(t) * \exp (X * b)$. |
| censoring | A Boolean array of the same size as $y$ that is 1 <br> for observations that are right-censored and 0 for <br> observations that are observed exactly. Default is all <br> observations observed exactly. |
| frequency | An array of the same size as $y$ containing nonnegative <br> integer counts. The $j^{\text {th }}$ element of this vector gives <br> the number of times the method observes the $j^{t h}$ <br> element of $y$ and the $j^{\text {th }}$ row of $X$. Default is one <br> observation per row of $X$ and $y$. |


| Name | Value |
| :--- | :--- |
| init | A vector containing initial values for the estimated <br> coefficients b. |
| options | A structure specifying control parameters for <br> the iterative algorithm used to estimate b. A <br> call to statset can create this argument. For <br> parameter names and default values, type <br> statset('coxphfit'). |

[b,logl, $\mathrm{H}, \mathrm{stats}]=$ coxphfit(...) returns additional results. logl is the log likelihood. H is a two-column matrix containing y values in the first column and the estimated baseline cumulative hazard evaluated at those values in the second column. stats is a structure that contains the fields:

- beta - Coefficient estimates (same as b)
- se - Standard errors of coefficient estimates b
- $z-z$ statistics for $b$ (b divided by standard error)
- $\mathrm{p}-p$-values for b
- covb - Estimated covariance matrix for $b$


## Examples Generate Weibull data depending on predictor x :

```
x = 4*rand(100,1);
A = 50*exp(-0.5*x); B = 2;
y = wblrnd(A,B);
```

Fit a Cox model :

```
[b,logL,H,stats] = coxphfit(x,y);
```

Show the Cox estimate of the baseline survivor function together with the known Weibull function:

```
stairs(H(:,1),exp(-H(:,2)))
```

```
    xx = linspace(0,100);
    line(xx,1-wblcdf(xx,50*exp(-0.5*mean(x)),B), 'color', 'r')
    xlim([0,50])
    legend('Survivor Function','Weibull Function')
```



## References

[1] Cox, D.R., and D. Oakes. Analysis of Survival Data. London: Chapman \& Hall, 1984.
[2] Lawless, J. F. Statistical Models and Methods for Lifetime Data. Hoboken, NJ: Wiley-Interscience, 2002.

## See Also

ecdf | statset | wblfit

## NaiveBayes.CPrior property

## Purpose Class priors

Description The CPrior property is a vector of length NClasses containing the class priors. The priors for empty classes are zero.
Purpose Create object to use in $k$-nearest neighbors search
Syntax

NS = createns(X)
NS = createns(X,'Name', Value)
$N S=$ createns $(X)$ uses the data observations in an $m x$-by- $n$ matrix $X$ to create an object NS. Rows of $X$ correspond to observations and columns correspond to variables. NS is either an ExhaustiveSearcher or a KDTreeSearcher object which you can use to find nearest neighbors in $X$ for desired query points. When NS is an ExhaustiveSearcher object, knnsearch uses the exhaustive search algorithm to find nearest neighbors. When NS is a KDTreeSearcher, createns creates and saves a $k$ d-tree based on $X$ in NS. knnsearch uses the $k d$-tree to find nearest neighbors. For information on these search methods, see " $k$-Nearest Neighbor Search and Radius Search" on page 13-12.
NS $=$ createns (X,'Name', Value) accepts one or more optional name/value pairs. Specify Name inside single quotes. Specify NSMethod to determine which type of object to create. The object's properties save the information when you specify other arguments. For more information on the objects' properties, see ExhaustiveSearcher and KDTreeSearcher.

## Input

Name-Value Pair Arguments

## NSMethod

Nearest neighbors search method, used to define the type of object created. Value is either:

- 'kdtree' - Create a KDTreeSearcher object. If you do not specify NSMethod, this is the default value when the number of columns of X is less than $10, \mathrm{X}$ is not sparse, and the distance measure is one of the following measures:
- 'euclidean' (default)
- 'cityblock'
- 'minkowski'
- 'chebychev'
- 'exhaustive' - Create an ExhaustiveSearcher object. If you do not specify NSMethod, this is the default value when the default criteria for 'kdtree' do not apply.


## Distance

A string or a function handle specifying the default distance metric used when you call the knnsearch method to find nearest neighbors for future query points. If you specify a distance metric but not an NSMethod, this input determines the type of object createns creates, according to the default values described in NSMethod.

For both KDTreeSearcher and ExhaustiveSearcher objects, the following options apply:

- 'euclidean' (default) - Euclidean distance.
- 'cityblock' - City block distance.
- 'chebychev' - Chebychev distance (maximum coordinate difference).
- 'minkowski' - Minkowski distance.

The following options apply only to ExhaustiveSearcher objects:

- 'seuclidean' - Standardized Euclidean distance. Each coordinate difference between rows in X and the query matrix is scaled by dividing by the corresponding element of the standard deviation computed from $X, S=$ nanstd ( $X$ ). To specify another value for $S$, use the Scale argument.
- 'mahalanobis' - Mahalanobis distance, which is computed using a positive definite covariance matrix $C$. The default value of $C$ is the sample covariance matrix of $X$, as computed by nancov (X). To change the value of $C$, use the Cov parameter.
- 'cosine' - One minus the cosine of the included angle between observations (treated as vectors).
- 'correlation' - One minus the sample linear correlation between observations (treated as sequences of values).
- 'spearman' - One minus the sample Spearman's rank correlation between observations (treated as sequences of values).
- 'hamming' - Hamming distance, which is percentage of coordinates that differ.
- 'jaccard' - One minus the Jaccard coefficient, which is the percentage of nonzero coordinates that differ.
- custom distance function - A distance function specified using @ (for example, @distfun). A distance function must be of the form function D2 = distfun(ZI, ZJ), taking as arguments a 1-by- $n$ vector ZI containing a single row from $X$ or from the query points Y , and an $m 2$-by- $n$ matrix ZJ containing multiple rows of X or Y , and returning an $m 2$-by- 1 vector of distances $d 2$, whose $j$ th element is the distance between the observations ZI and ZJ(j,:).

A positive scalar, $p$, indicating the exponent of the Minkowski distance. This parameter is only valid when Distance is 'minkowski'. Default is 2.

Cov
A positive definite matrix indicating the covariance matrix when computing the Mahalanobis distance. This parameter is only valid when Distance is 'mahalanobis'. Default is nancov ( X ).

Scale
A vector $S$ with the length equal to the number of columns in X. Each coordinate of $X$ and each query point is scaled by the corresponding element of $S$ when computing the standardized

Euclidean distance. This parameter is only valid when Distance is 'seuclidean'. Default is nanstd ( X ).

## BucketSize

A positive integer, indicating the maximum number of data points in each leaf node of the $k d$-tree. This argument is only meaningful when using the $k$ d-tree search method. Default is 50 .

## Examples <br> Create a $k$ d-tree with a Minkowski distance metric and a $P$ value of 5:

```
load fisheriris
x = meas(:,3:4);
% Since x has only two columns and the Distance is Minkowski,
% createns creates a KDTreeSearcher object by default:
knnobj = createns(x,'Distance','minkowski','P',5)
knnobj =
    KDTreeSearcher
    Properties:
        BucketSize: 50
            X: [150x2 double]
            Distance: 'minkowski'
    DistParameter: 5
```


## See Also ExhaustiveSearcher.knnsearch | KDTreeSearcher.knnsearch | ExhaustiveSearcher | KDTreeSearcher

How To

- " $k$-Nearest Neighbor Search and Radius Search" on page 13-12


## Purpose Cross-tabulation

```
Syntax table \(=\) crosstab \((x 1, x 2)\)
table \(=\) crosstab(x1, x2,..., xn)
[table,chi2,p] = crosstab(x1,..., xn)
[table,chi2, p,labels] = crosstab(x1,....,xn)
```


## Description

## Examples

## Example 1

Cross-tabulate two vectors with three and four distinct values, respectively:

$$
x=\left[\begin{array}{lllll}
1 & 1 & 2 & 3 & 1
\end{array}\right] ; y=\left[\begin{array}{llll}
1 & 2 & 5 & 3
\end{array}\right] \text {; }
$$

```
table = crosstab(x,y)
table =
\begin{tabular}{llll}
2 & 1 & 0 & 0 \\
0 & 0 & 0 & 1 \\
0 & 0 & 1 & 0
\end{tabular}
```


## Example 2

Generate two independent vectors, each containing 50 discrete uniform random numbers in the range $1: 3$ :

```
x1 = unidrnd(3,50,1);
x2 = unidrnd(3,50,1);
[table,chi2,p] = crosstab(x1,x2)
table =
    1 6
        5 5 2
        11 7 6
chi2 =
    7.5449
p =
    0.1097
```

At the $95 \%$ confidence level, the $p$ value fails to reject the null hypothesis that table is independent in each dimension.

## Example 3

The file carbig.mat contains measurements of large model cars during the years 1970-1982:

```
load carbig
[table,chi2,p,labels] = crosstab(cyl4,when,org)
table(:,:,1) =
    82 75 25
    12 22 38
table(:,:,2) =
            0 4
    23 26 17
table(:,:,3) =
```

```
        3 3 4
        12 25 32
chi2 =
    207.7689
p =
            0
label =
        'Other' 'Early' 'USA'
        'Four' 'Mid' 'Europe'
        [] 'Late' 'Japan'
```

table and label together show that the number of four-cylinder cars made in the USA during the late period of the data was $\operatorname{table}(2,3,1)$ or 38 cars.

## See Also <br> grp2idx | tabulate

How To

- "Grouped Data" on page 2-34

Purpose Loss estimate using cross-validation

Syntax<br>\section*{Description}

vals = crossval(fun, X)
vals = crossval(fun, X,Y,...)
mse = crossval('mse',X,y,'Predfun',predfun)
mcr = crossval('mcr',X,y,'Predfun',predfun)
val = crossval(criterion, $\mathrm{X} 1, \mathrm{X} 2, \ldots, \mathrm{y}$, 'Predfun', predfun)
vals $=$ crossval(...,'name', value)
vals $=$ crossval(fun, X ) performs 10 -fold cross-validation for the function fun, applied to the data in $X$.
fun is a function handle to a function with two inputs, the training subset of $X$, XTRAIN, and the test subset of X , XTEST, as follows:

```
testval = fun(XTRAIN,XTEST)
```

Each time it is called, fun should use XTRAIN to fit a model, then return some criterion testval computed on XTEST using that fitted model.
$X$ can be a column vector or a matrix. Rows of $X$ correspond to observations; columns correspond to variables or features. Each row of vals contains the result of applying fun to one test set. If testval is a non-scalar value, crossval converts it to a row vector using linear indexing and stored in one row of vals.
vals $=$ crossval(fun, $X, Y, \ldots$ ) is used when data are stored in separate variables X, Y, ... . All variables (column vectors, matrices, or arrays) must have the same number of rows. fun is called with the training subsets of $\mathrm{X}, \mathrm{Y}, \ldots$, followed by the test subsets of $\mathrm{X}, \mathrm{Y}, \ldots$, as follows:

```
testvals = fun(XTRAIN,YTRAIN,...,XTEST,YTEST,...)
```

mse = crossval('mse', X,y,'Predfun', predfun) returns mse, a scalar containing a 10 -fold cross-validation estimate of mean-squared error for the function predfun. X can be a column vector, matrix, or array of predictors. y is a column vector of response values. X and y must have the same number of rows.
predfun is a function handle called with the training subset of $X$, the training subset of $y$, and the test subset of $X$ as follows:
yfit = predfun(XTRAIN,ytrain, XTEST)

Each time it is called, predfun should use XTRAIN and ytrain to fit a regression model and then return fitted values in a column vector yfit. Each row of yfit contains the predicted values for the corresponding row of XTEST. crossval computes the squared errors between yfit and the corresponding response test set, and returns the overall mean across all test sets.
mcr = crossval('mcr',X,y,'Predfun', predfun) returns mcr, a scalar containing a 10 -fold cross-validation estimate of misclassification rate (the proportion of misclassified samples) for the function predfun. The matrix $X$ contains predictor values and the vector y contains class labels. predfun should use XTRAIN and YTRAIN to fit a classification model and return yfit as the predicted class labels for XTEST. crossval computes the number of misclassifications between yfit and the corresponding response test set, and returns the overall misclassification rate across all test sets.
val = crossval(criterion, $\mathrm{X} 1, \mathrm{X} 2, \ldots, \mathrm{y}$, 'Predfun', predfun), where criterion is 'mse' or 'mor', returns a cross-validation estimate of mean-squared error (for a regression model) or misclassification rate (for a classification model) with predictor values in $\mathrm{X} 1, \mathrm{X} 2, \ldots$ and, respectively, response values or class labels in y . $\mathrm{X} 1, \mathrm{X} 2, \ldots$ and y must have the same number of rows. predfun is a function handle called with the training subsets of $\mathrm{X} 1, \mathrm{X} 2, \ldots$, the training subset of y , and the test subsets of $\mathrm{X} 1, \mathrm{X} 2, \ldots$, as follows:
yfit=predfun(X1TRAIN, X2TRAIN, . . . , ytrain, X1TEST, X2TEST, . . . )
yfit should be a column vector containing the fitted values.
vals = crossval(..., 'name', value) specifies one or more optional parameter name/value pairs from the following table. Specify name inside single quotes.

| Name | Value |
| :--- | :--- |
| holdout | A scalar specifying the ratio or the number <br> of observations p for holdout cross-validation. <br> When $0<p<1$, approximately p*n observations <br> for the test set are randomly selected. When p <br> is an integer, p observations for the test set are <br> randomly selected. |
| kfold | A scalar specifying the number of folds k for <br> k-fold cross-validation. |
| leaveout | Specifies leave-one-out cross-validation. The <br> value must be 1. |
| mcreps | A positive integer specifying the number of <br> Monte-Carlo repetitions for validation. Ifthe <br> first input of crossval is 'mse ' or 'mcr ' <br> crossval returns the mean of mean-squared <br> error or misclassification rate across all of the <br> Monte-Carlo repetitions. Otherwise, crossval <br> concatenates the values vals from all of <br> the Monte-Carlo repetitions along the first <br> dimension. |
| partition | An object c of the cvpartition class, specifying <br> the cross-validation type and partition. |
| stratify | A column vector group specifying groups for <br> stratification. Both training and test sets have <br> roughly the same class proportions as in group. <br> NaNs or empty strings in group are treated as <br> missing values, and the corresponding rows of <br> the data are ignored. |
| options | A structure that specifies whether to run in <br> parallel, and specifies the random stream or <br> streams. Create the options structure with <br> statset. Option fields: |


| Name | Value |
| :---: | :---: |
|  | - UseParallel - Set to 'always' to compute in parallel. Default is 'never'. <br> - UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'. <br> - Streams - A RandStream object or cell array consisting of one such object. If you do not specify Streams, crossval uses the default stream. <br> For more information on using parallel computing, see Chapter 17, "Parallel Statistics". |

Only one of kfold, holdout, leaveout, or partition can be specified, and partition cannot be specified with stratify. If both partition and mcreps are specified, the first Monte-Carlo repetition uses the partition information in the cvpartition object, and the repartition method is called to generate new partitions for each of the remaining repetitions. If no cross-validation type is specified, the default is 10 -fold cross-validation.

Note When using cross-validation with classification algorithms, stratification is preferred. Otherwise, some test sets may not include observations from all classes.

## Examples Example 1

Compute mean-squared error for regression using 10 -fold cross-validation:

```
load('fisheriris');
y = meas(:,1);
X = [ones(size(y,1),1),meas(:,2:4)];
regf=@(XTRAIN,ytrain,XTEST)(XTEST*regress(ytrain,XTRAIN));
cvMse = crossval('mse',X,y,'predfun',regf)
cvMse =
    0.1015
```


## Example 2

Compute misclassification rate using stratified 10 -fold cross-validation:

```
load('fisheriris');
y = species;
X = meas;
cp = cvpartition(y,'k',10); % Stratified cross-validation
classf = @(XTRAIN, ytrain,XTEST)(classify(XTEST,XTRAIN,...
ytrain));
cvMCR = crossval('mcr',X,y,'predfun',classf,'partition',cp)
cvMCR =
    0.0200
```


## Example 3

Compute the confusion matrix using stratified 10 -fold cross-validation:

```
load('fisheriris');
y = species;
X = meas;
order = unique(y); % Order of the group labels
cp = cvpartition(y,'k',10); % Stratified cross-validation
f = @(xtr,ytr,xte,yte)confusionmat(yte,...
classify(xte,xtr,ytr),'order',order);
```

```
cfMat = crossval(f,X,y,'partition',cp);
cfMat = reshape(sum(cfMat),3,3)
cfMat =
\begin{tabular}{rrr}
50 & 0 & 0 \\
0 & 48 & 2 \\
0 & 1 & 49
\end{tabular}
```

cfMat is the summation of 10 confusion matrices from 10 test sets.

## References

[1] Hastie, T., R. Tibshirani, and J. Friedman. The Elements of Statistical Learning. New York: Springer, 2001.
See Also

cvpartition

How To - "Grouped Data" on page 2-34

## ClassificationDiscriminant.crossval

## Purpose Cross-validated discriminant analysis classifier

Syntax<br>Description

Tips

Input
Arguments
cvmodel = crossval(obj)
cvmodel = crossval(obj,Name,Value)
cvmodel $=$ crossval(obj) creates a partitioned model from obj, a fitted discriminant analysis classifier. By default, crossval uses 10 -fold cross validation on the training data to create cvmodel.
cvmodel = crossval(obj, Name, Value) creates a partitioned model with additional options specified by one or more Name, Value pair arguments.

- Assess the predictive performance of obj on cross-validated data using the "kfold" methods and properties of cvmodel, such as kfoldLoss.
obj
Discriminant analysis classifier, produced using ClassificationDiscriminant.fit.


## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
cvpartition
Object of class cvpartition, created by the cvpartition function. crossval splits the data into subsets with cupartition.
Use only one of these options at a time: 'cvpartition', 'holdout', 'kfold', or 'leaveout'.

Default: []

## ClassificationDiscriminant.crossval

## holdout

Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1. Use only one of these options at a time: 'cvpartition', 'holdout', 'kfold', or 'leaveout'.
kfold
Number of folds to use in a cross-validated classifier, a positive integer.

Use only one of these options at a time: 'cvpartition', 'holdout', 'kfold', or 'leaveout'.

Default: 10

## leaveout

Set to 'on' for leave-one-out cross validation.
Use only one of these options at a time: 'cvpartition', 'holdout', 'kfold', or 'leaveout'.

## Examples

Create a classification model for the Fisher iris data, and then create a cross-validation model. Evaluate the quality the model using kfoldLoss.

```
load fisheriris
obj = ClassificationDiscriminant.fit(meas,species);
cvmodel = crossval(obj);
L = kfoldLoss(cvmodel)
L =
    0.0200
```


## Alternatives

You can create a cross-validation classifier directly from the data, instead of creating a discriminant analysis classifier followed by a cross-validation classifier. To do so, include one of these options in

## ClassificationDiscriminant.crossval

> ClassificationDiscriminant.fit: 'crossval', 'cvpartition', 'holdout', 'kfold', or 'leaveout'.

## See Also ClassificationDiscriminant.fit | crossval | kfoldEdge | kfoldfun | kfoldLoss | kfoldMargin | kfoldPredict

How To

- "Discriminant Analysis" on page 12-3


## ClassificationEnsemble.crossval

## Purpose

Cross validate ensemble
Syntax

Description

## Input Arguments

```
cvens = crossval(ens)
cvens = crossval(ens,Name,Value)
```


## ens

cvens = crossval(ens) creates a cross-validated ensemble from ens, a classification ensemble. Default is 10 -fold cross validation.
cvens $=$ crossval(ens,Name, Value) creates a cross-validated ensemble with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

A classification ensemble created with fitensemble.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## cvpartition

A partition of class cvpartition. Sets the partition for cross validation.

Use no more than one of the name-value pairs cvpartition, holdout, kfold, or leaveout.
holdout
Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
kfold

## ClassificationEnsemble.crossval

Number of folds for cross validation, a numeric positive scalar.
Use no more than one of the name-value pairs 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
leaveout
If 'on ', use leave-one-out cross validation.
Use no more than one of the name-value pairs 'kfold', 'holdout', 'leaveout', or 'cvpartition'. nprint

Printout frequency, a positive integer scalar. Use this parameter to observe the training of cross-validation folds.

Default: 'off', meaning no printout

## Output Arguments

## Alternatives

You can create a cross-validation ensemble directly from the data, instead of creating an ensemble followed by a cross-validation ensemble. To do so, include one of these five options in fitensemble: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

## Examples

cvens
A cross-validated classification ensemble of class ClassificationPartitionedEnsemble.

Create a cross-validated classification model for the Fisher iris data, and assess its quality using the kfoldLoss method.

```
load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
cvens = crossval(ens);
L = kfoldLoss(cvens)
L =
    0.0467
```


# ClassificationEnsemble.crossval 

See Also ClassificationPartitionedEnsemble | cvpartition<br>How To • Chapter 13, "Nonparametric Supervised Learning"

## ClassificationTree.crossval

| Purpose | Cross-validated decision tree |
| :--- | :--- |
| Syntax | cvmodel$=\operatorname{crossval(model)}$ |
|  | cvmodel $=\operatorname{crossval}($ model, Name, Value $)$ |

Description cvmodel = crossval(model) creates a partitioned model from model, a fitted classification tree. By default, crossval uses 10 -fold cross validation on the training data to create cvmodel.
cvmodel = crossval(model, Name, Value) creates a partitioned model with additional options specified by one or more Name, Value pair arguments.

Tips

Input
Arguments

- Assess the predictive performance of model on cross-validated data using the "kfold" methods and properties of cvmodel, such as kfoldLoss.
model
A classification model, produced using ClassificationTree.fit.


## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
cvpartition
Object of class cvpartition, created by the cvpartition function. crossval splits the data into subsets with cupartition.
Use only one of these four options at a time: 'kfold', 'holdout ', 'leaveout', or 'cvpartition'.

Default: []
holdout

## ClassificationTree.crossval

Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
kfold
Number of folds to use in a cross-validated tree, a positive integer.
Use only one of these four options at a time: 'kfold', 'holdout',
'leaveout', or 'cvpartition'.
Default: 10

## leaveout

Set to 'on' for leave-one-out cross validation.

## Output Arguments

## Examples

cvmodel
A partitioned model of class ClassificationPartitionedModel.
Create a classification model for the ionosphere data, then create a cross-validation model. Evaluate the quality the model using kfoldLoss.
load ionosphere
tree $=$ ClassificationTree.fit(X,Y);
cvmodel = crossval(tree);
L = kfoldLoss(cvmodel)
L =
0.1168

## Alternatives

You can create a cross-validation tree directly from the data, instead of creating a decision tree followed by a cross-validation tree. To do so, include one of these five options in ClassificationTree.fit: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

## ClassificationTree.crossval

See Also ClassificationTree.fit | crossval<br>How To • Chapter 13, "Nonparametric Supervised Learning"

## RegressionEnsemble.crossval

## Purpose <br> Cross validate ensemble

Syntax $\quad \begin{aligned} \text { cvens } & =\operatorname{crossval(ens)} \\ \text { cvens } & =\operatorname{crossval}(\text { ens, Name, Value })\end{aligned}$

Description

## Input Arguments

cvens = crossval(ens) creates a cross-validated ensemble from ens, a regression ensemble. Default is 10 -fold cross validation.
cvens $=$ crossval(ens,Name, Value) creates a cross-validated ensemble with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## ens

A regression ensemble created with fitensemble.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## cvpartition

A partition of class cupartition. Sets the partition for cross validation.

Use no more than one of the name-value pairs cvpartition, holdout, kfold, and leaveout.
holdout
Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
kfold

## RegressionEnsemble.crossval

Number of folds for cross validation, a numeric positive scalar.
Use no more than one of the name-value pairs 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
leaveout
If 'on ', use leave-one-out cross-validation.
Use no more than one of the name-value pairs 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
nprint
Printout frequency, a positive integer scalar. Use this parameter to observe the training of cross-validation folds.

Default: 'off', meaning no printout

## Output Arguments

## Alternatives

## Examples

cvens
A cross-validated classification ensemble of class RegressionPartitionedEnsemble.

You can create a cross-validation ensemble directly from the data, instead of creating an ensemble followed by a cross-validation ensemble. To do so, include one of these five options in fitensemble: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

Create a cross-validated classification model for the carsmall data, and assess its quality using the kfoldLoss method:

```
X = [Acceleration Displacement Horsepower Weight];
rens = fitensemble(X,MPG,'LSBoost',100,'Tree');
cvens = crossval(rens);
L = kfoldLoss(cvens)
L =
    21.9868
```

See Also RegressionPartitionedEnsemble | cvpartition<br>How To • Chapter 13, "Nonparametric Supervised Learning"

## RegressionTree.crossval

| Purpose | Cross-validated decision tree |
| :---: | :---: |
| Syntax | ```cvmodel = crossval(model) cvmodel = crossval(model,Name,Value)``` |
| Description | cvmodel $=$ crossval(model) creates a partitioned model from model, a fitted regression tree. By default, crossval uses 10 -fold cross validation on the training data to create cvmodel. <br> cvmodel $=$ crossval(model, Name, Value) creates a partitioned model with additional options specified by one or more Name, Value pair arguments. |
| Tips | - Assess the predictive performance of model on cross-validated data using the "kfold" methods and properties of cvmodel, such as kfoldLoss. |
| Input <br> Arguments | model A regression model, produced using RegressionTree.fit. |
|  | Name-Value Pair Arguments |
|  | Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN. <br> cvpartition |
|  | Object of class cvpartition, created by the cvpartition function. crossval splits the data into subsets with cvpartition. |
|  | Use only one of these four options at a time: 'kfold', 'holdout', 'leaveout', or 'cvpartition'. |
|  | Default: [] |
|  | holdout |

holdout

## RegressionTree.crossval

Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
kfold
Number of folds to use in a cross-validated tree, a positive integer.
Use only one of these four options at a time: 'kfold', 'holdout',
'leaveout', or 'cvpartition'.
Default: 10

## leaveout

Set to 'on' for leave-one-out cross-validation.

## Output Arguments

## Examples

cvmodel
A partitioned model of class RegressionPartitionedModel.

Create a regression model of the carsmall data, and assess its accuracy with kfoldLoss:

```
load carsmall
X = [Acceleration Displacement Horsepower Weight];
tree = RegressionTree.fit(X,MPG);
cvtree = crossval(tree);
L = kfoldLoss(cvtree)
L =
    25.2432
```


## Alternatives

You can create a cross-validation tree directly from the data, instead of creating a decision tree followed by a cross-validation tree. To do so, include one of these five options in RegressionTree.fit: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

## RegressionTree.crossval

See Also RegressionTree.fit | crossval
How To - Chapter 13, "Nonparametric Supervised Learning"
Purpose Transpose categorical matrix
Syntax

$B=$ ctranspose(A)
$\begin{array}{ll}\text { Description } & B=\text { ctranspose (A) returns the transpose of the 2-D categorical matrix } \\ \text { A. Note that ctranspose is identical to transpose for categorical arrays. }\end{array}$
See Also transpose | permute

## classregtree.cutcategories

## Purpose Cut categories

```
Syntax C = cutcategories(t)
C = cutcategories(t,nodes)
```


## Description

$\mathrm{C}=$ cutcategories( t ) returns an $n$-by- 2 cell array C of the categories used at branches in the decision tree $t$, where $n$ is the number of nodes. For each branch node i based on a categorical predictor variable x, the left child is chosen if $x$ is among the categories listed in C $\{i, 1\}$, and the right child is chosen if $x$ is among those listed in $C\{i, 2\}$. Both columns of C are empty for branch nodes based on continuous predictors and for leaf nodes.
$C=$ cutcategories(t, nodes) takes a vector nodes of node numbers and returns the categories for the specified nodes.

## Examples Create a classification tree for car data:

```
load carsmall
t = classregtree([MPG Cylinders],Origin,...
    'names',{'MPG' 'Cyl'},'cat',2)
t =
Decision tree for classification
    if Cyl=4 then node 2 elseif Cyl in {6 8} then node 3 else USA
    if MPG<31.5 then node 4 elseif MPG>=31.5 then node 5 else USA
    if Cyl=6 then node 6 elseif Cyl=8 then node 7 else USA
    if MPG<21.5 then node 8 elseif MPG>=21.5 then node 9 else USA
    if MPG<41 then node 10 elseif MPG>=41 then node 11 else Japan
    if MPG<17 then node 12 elseif MPG>=17 then node 13 else USA
    class = USA
    class = France
    class = USA
    class = Japan
    class = Germany
    class = Germany
    class = USA
```



$$
\begin{aligned}
& C=\text { cutcategories }(t) \\
& C=
\end{aligned}
$$

[4] [1x2 double]
[] []
[6] [ 8]
[] []
[] []
[] []
[] []
[] []

## classregtree.cutcategories

```
    [] []
    [] []
    [] []
    [] []
    [] []
C{1,2}
ans =
    6
```

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

## See Also classregtree | cutpoint | cuttype | cutvar

Purpose
Decision tree cut point values

## Syntax

$v=$ cutpoint (t)
v = cutpoint(t,nodes)

Description
$v=$ cutpoint( $t$ ) returns an $n$-element vector $v$ of the values used as cut points in the decision tree t , where $n$ is the number of nodes. For each branch node $i$ based on a continuous predictor variable $x$, the left child is chosen if $x<v(i)$ and the right child is chosen if $x>=$ $v(i)$. $v$ is NaN for branch nodes based on categorical predictors and for leaf nodes.
$v=$ cutpoint( $t$, nodes) takes a vector nodes of node numbers and returns the cut points for the specified nodes.

## Examples Create a classification tree for car data:

```
load carsmall
t = classregtree([MPG Cylinders],Origin,...
            'names',{'MPG' 'Cyl'},'cat',2)
t =
Decision tree for classification
    if Cyl=4 then node 2 elseif Cyl in {6 8} then node 3 else USA
    if MPG<31.5 then node 4 elseif MPG>=31.5 then node 5 else USA
    if Cyl=6 then node 6 elseif Cyl=8 then node 7 else USA
    if MPG<21.5 then node 8 elseif MPG>=21.5 then node 9 else USA
    if MPG<41 then node 10 elseif MPG>=41 then node 11 else Japan
    if MPG<17 then node 12 elseif MPG>=17 then node 13 else USA
    class = USA
    class = France
    class = USA
    class = Japan
    class = Germany
    class = Germany
    class = USA
```


## classregtree.cutpoint

## view(t)

|  |  |  |  |  |  |  | 4 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Click to display: \|dentity |  | Magnification: | 100\% | $\checkmark$ | Pruning level: | 0 of 3 | $\forall$ |



$$
\begin{array}{r}
v=\text { cutpoint (t) } \\
v= \\
\text { NaN } \\
31.5000 \\
\mathrm{NaN} \\
21.5000 \\
41.0000 \\
17.0000 \\
\mathrm{NaN} \\
\mathrm{NaN} \\
\mathrm{NaN} \\
\mathrm{NaN}
\end{array}
$$

## classregtree.cutpoint

NaN
NaN
NaN

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.<br>See Also classregtree | cutcategories | cuttype | cutvar

## classregtree.cuttype

## Purpose <br> Cut types

Syntax
c = cuttype(t)
c = cuttype(t, nodes)

## Description

$c=$ cuttype( t ) returns an $n$-element cell array c indicating the type of cut at each node in the tree $t$, where $n$ is the number of nodes. For each node i, c\{i\} is:

- 'continuous ' - If the cut is defined in the form $\mathrm{x}<\mathrm{v}$ for a variable $x$ and cut point $v$.
- 'categorical' - If the cut is defined by whether a variable $x$ takes a value in a set of categories.
- ' ' - If $i$ is a leaf node.
cutvar returns the cut points for 'continuous' cuts, and cutcategories returns the set of categories.
c = cuttype(t, nodes) takes a vector nodes of node numbers and returns the cut types for the specified nodes.


## Examples <br> Create a classification tree for car data:

```
load carsmall
t = classregtree([MPG Cylinders],Origin,...
    'names',{'MPG' 'Cyl'},'cat',2)
t =
Decision tree for classification
    if Cyl=4 then node 2 elseif Cyl in {6 8} then node 3 else USA
    2 if MPG<31.5 then node 4 elseif MPG>=31.5 then node 5 else USA
    3 if Cyl=6 then node 6 elseif Cyl=8 then node 7 else USA
    4 if MPG<21.5 then node 8 elseif MPG>=21.5 then node 9 else USA
    5 if MPG<41 then node 10 elseif MPG>=41 then node 11 else Japan
    6 if MPG<17 then node 12 elseif MPG>=17 then node 13 else USA
    7 class = USA
```

| 8 | class = France |
| :---: | :---: |
| 9 | class = USA |
| 10 | class = Japan |
| 11 | class = Germany |
|  | class = Germany |
|  | class = USA |
|  | (t) |


| Click to display: | $\mid$ Identity | $\square$ | Magnification: | $100 \%$ | $\square$ |
| :--- | :--- | :--- | :--- | :--- | :--- | Pruning level: | 0 of 3 |  |
| :--- | :--- | :--- | :--- |



$$
\begin{aligned}
c= & \text { cuttype(t) } \\
c= & \\
& \text { 'categorical' } \\
& \text { 'continuous' } \\
& \text { 'contegorical' }
\end{aligned}
$$

## classregtree.cuttype

```
    'continuous'
    'continuous'
|
''
''
''
''
''
''
```

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.
See Also classregtree | numnodes | cutvar | cutcategories

## Purpose

Cut variable names

## Syntax

v = cutvar(t)
$v$ = cutvar(t, nodes)
[v,num] = cutvar(...)
$\mathrm{v}=$ cutvar( t$)$ returns an $n$-element cell array v of the names of the variables used for branching in each node of the tree $t$, where $n$ is the number of nodes. These variables are sometimes known as cut variables. For leaf nodes, v contains an empty string.
$\mathrm{v}=$ cutvar(t, nodes) takes a vector nodes of node numbers and returns the cut variables for the specified nodes.
[ v , num] $=$ cutvar(...) also returns a vector num containing the number of each variable.

Create a classification tree for car data:

```
load carsmall
t = classregtree([MPG Cylinders],Origin,...
    'names',{'MPG' 'Cyl'},'cat',2)
t =
Decision tree for classification
    if Cyl=4 then node 2 elseif Cyl in {6 8} then node 3 else USA
    if MPG<31.5 then node 4 elseif MPG>=31.5 then node 5 else USA
        if Cyl=6 then node 6 elseif Cyl=8 then node 7 else USA
        if MPG<21.5 then node 8 elseif MPG>=21.5 then node 9 else USA
        if MPG<41 then node 10 elseif MPG>=41 then node 11 else Japan
        if MPG<17 then node 12 elseif MPG>=17 then node 13 else USA
        class = USA
        class = France
        class = USA
        class = Japan
        class = Germany
        class = Germany
        lass = USA
```


## classregtree.cutvar

## view(t)

| Click to display: | $\mid$ Identity | $\square$ | Magnification: | $100 \%$ | $\square$ | Pruning level: |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 0 of 3 |  |  |  |  |  |  |



$$
[v, \text { num }]=\text { cutvar }(t)
$$

v =
'Cyl'
'MPG'
'Cyl'
'MPG '
'MPG '
'MPG '
1,
' '

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.
See Also classregtree | numnodes | children

## ClassificationTree.cvLoss

Purpose Classification error by cross validation
Syntax $\quad E=$ cvLoss(tree)
[E,SE] = cvLoss(tree)
[E,SE,Nleaf] = cvLoss(tree) [E,SE,Nleaf,BestLevel] = cvLoss(tree) [ $\mathrm{E}, \ldots$...] = cvLoss(tree, Name, Value)

## Description

Input Arguments
$\mathrm{E}=\mathrm{cvLoss}($ tree $)$ returns the cross-validated classification error (loss) for tree, a classification tree.
[ $E, S E]=$ cvLoss(tree) returns the standard error of $E$.
[ $\mathrm{E}, \mathrm{SE}, \mathrm{Nleaf}]=$ cvLoss(tree) returns the number of leaves of tree.
[E,SE,Nleaf, BestLevel] = cvLoss(tree) returns the optimal pruning level for tree.
[ $\mathrm{E}, \ldots$. $]$ = cvLoss(tree, Name, Value) cross validates with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
tree
A classification tree produced by ClassificationTree.fit.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## subtrees

Vector of pruning levels, with 0 representing the full, unpruned tree. tree must include a pruning sequence as created either by ClassificationTree.fit with 'prune' set to 'on', or by the prune method. The returned E, SE, and Nleaf are vectors of

## ClassificationTree.cvLoss

the same length as subtrees; BestLevel is a scalar. If you set subtrees to 'all', cvLoss uses the entire pruning sequence.

Default: 0

## treesize

One of the following strings:

- 'se' - cvLoss uses the smallest tree whose cost is within one standard error of the minimum cost.
- 'min' - cvLoss uses the minimal cost tree.

Default: 'se'

## kfold

Number of cross-validation samples, a positive integer.
Default: 10

Output
Arguments

E
The cross-validation classification error (loss). A vector or scalar depending on the setting of the subtrees name-value pair.

SE
The standard error of E . A vector or scalar depending on the setting of the subtrees name-value pair.

## Nleaf

Number of leaf nodes in tree. Leaf nodes are terminal nodes, which give classifications, not splits. A vector or scalar depending on the setting of the subtrees name-value pair.

## BestLevel

## ClassificationTree.cvLoss

By default, a scalar representing the largest pruning level that achieves a value of $E$ within $S E$ of the minimum error. If you set treesize to 'min', BestLevel is the smallest value in subtrees.

## Examples Compute the cross-validation error for the default classification tree

 for the ionosphere data:```
load ionosphere
tree = ClassificationTree.fit(X,Y);
[E,SE,Nleaf,BestLevel] = cvLoss(tree)
E =
    0.1282
SE =
    0.0178
Nleaf =
    1 9
BestLevel =
    0
```

Find the best level by using the subtrees name-value pair:

```
[~,~,~,BestLevel] = cvLoss(tree,'subtrees','all')
```

BestLevel =
6

## Alternatives

You can construct a cross-validated tree model with crossval, and call kfoldLoss instead of cvLoss. The alternative can save time if you are going to examine the cross-validated tree more than once.

However, unlike cvLoss, kfoldLoss does not return SE,Nleaf, or BestLevel. kfoldLoss also does not allow you to examine any error other than classification error.

# ClassificationTree.cvLoss 

See Also ClassificationTree.fit | crossval | loss | kfoldLoss<br>How To • Chapter 13, "Nonparametric Supervised Learning"

## RegressionTree.cvLoss

## Purpose Regression error by cross validation

Syntax
Description

Input Arguments
$\mathrm{E}=\mathrm{cvLoss}(\mathrm{tree})$
[E,SE] = cvLoss(tree)
[E,SE,Nleaf] = cvLoss(tree)
[E,SE,Nleaf,BestLevel] = cvLoss(tree)
[E,...] = cvLoss(tree,Name,Value)
$\mathrm{E}=\mathrm{cvLoss}($ tree $)$ returns the cross-validated regression error (loss) for tree, a regression tree.
[ $\mathrm{E}, \mathrm{SE}]=$ cvLoss(tree) returns the standard error of E .
[E,SE,Nleaf] = cvLoss(tree) returns the number of leaves (terminal nodes) in tree.
[E,SE,Nleaf, BestLevel] = cvLoss(tree) returns the optimal pruning level for tree.
[ $\mathrm{E}, \ldots$...] = cvLoss(tree, Name, Value) cross validates with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.
tree
A regression tree produced by RegressionTree.fit.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## subtrees

Vector of pruning levels, with 0 representing the full, unpruned tree. tree must include a pruning sequence as created either by ClassificationTree.fit with 'prune' set to 'on', or by
the prune method. The returned E, SE, and Nleaf are vectors of the same length as subtrees; BestLevel is a scalar. If you set subtrees to 'all', cvLoss uses the entire pruning sequence.

Default: 0

## treesize

One of the following strings:

- 'se' - cvLoss uses the smallest tree whose cost is within one standard error of the minimum cost.
- 'min' - cvLoss uses the minimal cost tree.

Default: 'se'
kfold
Number of cross-validation samples, a positive integer.
Default: 10

Output E
Arguments

The cross-validation mean squared error (loss). A vector or scalar depending on the setting of the subtrees name-value pair.

SE
The standard error of $E$. A vector or scalar depending on the setting of the subtrees name-value pair.

## Nleaf

Number of leaf nodes in tree. Leaf nodes are terminal nodes, which give responses, not splits. A vector or scalar depending on the setting of the subtrees name-value pair.

## BestLevel

## RegressionTree.cvLoss

By default, a scalar representing the largest pruning level that achieves a value of $E$ within $S E$ of the minimum error. If you set treesize to 'min', BestLevel is the smallest value in subtrees.

## Examples Compute the cross-validation error for the default classification tree for

 the carsmall data:```
load carsmall
X = [Displacement Horsepower Weight];
tree = RegressionTree.fit(X,MPG);
[E,SE,Nleaf,BestLevel] = cvLoss(tree)
E =
    30.7558
SE =
    6.0651
Nleaf =
    1 9
BestLevel =
    O
```

Find the best level by using the subtrees name-value pair:

```
[~,~,~,BestLevel] = cvLoss(tree,'subtrees','all')
BestLevel =
    15
```

Alternatives You can construct a cross-validated tree model with crossval, and call kfoldLoss instead of cvLoss. The alternative can save time if you are going to examine the cross-validated tree more than once.

However, unlike cvLoss, kfoldLoss does not return SE,Nleaf, or BestLevel.

## RegressionTree.cvLoss

See Also crossval | kfoldLoss | RegressionTree.fit | loss
How To $\quad$ Chapter 13, "Nonparametric Supervised Learning"

## Purpose Data partitions for cross-validation

Description
An object of the cupartition class defines a random partition on a set of data of a specified size. Use this partition to define test and training sets for validating a statistical model using cross-validation.

## Construction

Create cross-validation partition
for data

Display cvpartition object
Display cvpartition object
Repartition data for cross-validation

Test indices for cross-validation
Training indices for cross-validation

Number of observations (including observations with missing group values)
Number of test sets
Size of each test set
Size of each training set
Type of partition

Value. To learn how this affects your use of the class, see Comparing Handle and Value Classes in the MATLAB Object-Oriented Programming documentation.

```
Examples
Use a 10 -fold stratified cross-validation to compute the misclassification error for classify on iris data.
```

```
load('fisheriris');
```

load('fisheriris');
CVO = cvpartition(species,'k',10);
CVO = cvpartition(species,'k',10);
err = zeros(CVO.NumTestSets,1);
err = zeros(CVO.NumTestSets,1);
for i = 1:CVO.NumTestSets
for i = 1:CVO.NumTestSets
trIdx = CVO.training(i);
trIdx = CVO.training(i);
teIdx = CVO.test(i);
teIdx = CVO.test(i);
ytest = classify(meas(teIdx,:),meas(trIdx,:),...
ytest = classify(meas(teIdx,:),meas(trIdx,:),...
species(trIdx,:));
species(trIdx,:));
err(i) = sum(~strcmp(ytest,species(teIdx)));
err(i) = sum(~strcmp(ytest,species(teIdx)));
end
end
cvErr = sum(err)/sum(CVO.TestSize);

```
cvErr = sum(err)/sum(CVO.TestSize);
```

See Also crossval
How To

- "Grouped Data" on page 2-34

Purpose Create cross-validation partition for data

| Syntax | $c=$ cvpartition( $\mathrm{n}, \mathrm{\prime}$ kfold', k ) |
| :---: | :---: |
|  | c = cvpartition(group, 'kfold',k) |
|  | $\mathrm{c}=$ cvpartition( $\mathrm{n}, \mathrm{h}$ holdout', p ) |
|  | $\mathrm{c}=$ cvpartition(group, 'holdout', p ) |
|  | $c=$ cvpartition( $\mathrm{n}, \mathrm{\prime}$ leaveout') |
|  | $\mathrm{c}=$ cvpartition( $\mathrm{n}, \mathrm{'r}$ (esubstitution') |

## Description

c = cvpartition( $\mathrm{n}, \mathrm{'}^{\prime} \mathrm{kfold}{ }^{\prime}$, k ) constructs an object c of the cvpartition class defining a random partition for $k$-fold cross-validation on n observations. The partition divides the observations into k disjoint subsamples (or folds), chosen randomly but with roughly equal size. The default value of $k$ is 10 .
c = cvpartition(group, 'kfold', k) creates a random partition for a stratified $k$-fold cross-validation. group is a numeric vector, categorical array, string array, or cell array of strings indicating the class of each observation (see "Grouped Data" on page 2-34). Each subsample has roughly equal size and roughly the same class proportions as in group. cupartition treats NaNs or empty strings in group as missing values.
c = cvpartition( $n$, 'holdout', p ) creates a random partition for holdout validation on $n$ observations. This partition divides the observations into a training set and a test (or holdout) set. The parameter $p$ must be a scalar. When $0<p<1$, cvpartition randomly selects approximately $p * n$ observations for the test set. When $p$ is an integer, cvpartition randomly selects p observations for the test set. The default value of $p$ is $1 / 10$.
c = cvpartition(group, 'holdout', p) randomly partitions observations into a training set and a test set with stratification, using the class information in group; that is, both training and test sets have roughly the same class proportions as in group.
$\mathrm{c}=$ cvpartition( n, ' leaveout') creates a random partition for leave-one-out cross-validation on $n$ observations. Leave-one-out is a special case of ' $k f o l d$ ', in which the number of folds equals the number of observations.
$c=$ cvpartition(n, 'resubstitution') creates an object $c$ that does not partition the data. Both the training set and the test set contain all of the original $n$ observations.

## Examples

Use stratified 10 -fold cross-validation to compute misclassification rate:

```
load fisheriris;
y = species;
c = cvpartition(y,'k',10);
fun = @(xT,yT,xt,yt)(sum(~strcmp(yt,classify(xt,xT,yT))));
rate = sum(crossval(fun,meas,y,'partition',c))...
    /sum(c.TestSize)
rate =
    0.0200
```

See Also
crossval | repartition
How To • Grouped Data

## RegressionEnsemble.cvshrink

Purpose Cross validate shrinking (pruning) ensemble
Syntax
vals = cvshrink(ens)
[vals,nlearn] = cvshrink(ens)
[vals,nlearn] = cvshrink(ens,Name,Value)

## Description

Input
Arguments
vals $=$ cvshrink(ens) returns an L-by-T matrix with cross-validated values of the mean squared error. $L$ is the number of lambda values in the ens.Regularization structure. T is the number of threshold values on weak learner weights. If ens does not have a Regularization property filled in by the regularize method, pass a lambda name-value pair.
[vals, nlearn] = cvshrink(ens) returns an L-by-T matrix of the mean number of learners in the cross-validated ensemble.
[vals, nlearn] = cvshrink(ens,Name, Value) cross validates with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.
ens
A regression ensemble, created with fitensemble.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
cupartition
A partition created with cvpartition to use in a cross-validated tree. You can only use one of these four options at a time:
'kfold', 'holdout', 'leaveout', or 'cvpartition'.
holdout

Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

## kfold

Number of folds to use in a cross-validated tree, a positive integer. If you do not supply a cross-validation method, cvshrink uses 10 -fold cross validation. You can only use one of these four options at a time: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

Default: 10

## lambda

Vector of nonnegative regularization parameter values for lasso. If empty, cvshrink does not perform cross validation.

Default: []
leaveout
Use leave-one-out cross validation by setting to 'on '. You can only use one of these four options at a time: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
threshold
Numeric vector with lower cutoffs on weights for weak learners. cvshrink discards learners with weights below threshold in its cross-validation calculation.

Default: 0

## Output vals

Arguments
L-by-T matrix with cross-validated values of the mean squared error. $L$ is the number of values of the regularization parameter

## RegressionEnsemble.cvshrink

'lambda', and T is the number of 'threshold' values on weak learner weights.

## nlearn

L-by-T matrix with cross-validated values of the mean number of learners in the cross-validated ensemble. L is the number of values of the regularization parameter ' lambda', and $T$ is the number of 'threshold ' values on weak learner weights.

## Examples

Create a regression ensemble for predicting mileage from the carsmall data. Cross validate the ensemble for three values each of lambda and threshold.

```
load carsmall
X = [Displacement Horsepower Weight];
ens = fitensemble(X,MPG,'bag',100,'Tree',...
    'type','regression');
[vals nlearn] = cvshrink(ens,'lambda',[.01 .1 1],...
    'threshold',[0 .01 .1])
vals =
    20.0949 19.9007 131.6316
    20.0924 19.8431 128.0989
    19.9759 19.7987 119.5574
nlearn =
    13.3000 11.6000 3.5000
    13.2000 11.5000 3.6000
    13.4000 11.4000 3.9000
```

Clearly, setting a threshold of 0.1 leads to unacceptable errors, while a threshold of 0.01 gives similar errors to a threshold of 0 . The mean number of learners with a threshold of 0.1 is about 11.5 , whereas the mean number is about 13.2 when the threshold is 0 .

## See Also

```
regularize | shrink
```

How To • Chapter 13, "Nonparametric Supervised Learning"

Purpose
Syntax

## Description

Tips

Input
Arguments

```
```

y = datasample(data,k)

```
```

y = datasample(data,k)
y = datasample(data,k,dim)
y = datasample(data,k,dim)
[y,idx] = datasample(data,k,...)
[y,idx] = datasample(data,k,...)
[y,...] = datasample(s,data,k,...)
[y,...] = datasample(s,data,k,...)
[y,...] = datasample(data,k,Name, Value)
[y,...] = datasample(data,k,Name, Value)
[y,...] = datasample(data, $k$, dim,Name, Value)

```
```

[y,...] = datasample(data, $k$, dim,Name, Value)

```
```

Randomly sample from data, with or without replacement
$\mathrm{y}=$ datasample(data, k$)$ returns k observations sampled uniformly at random, with replacement, from the data in data.
$\mathrm{y}=$ datasample(data, $\mathrm{k}, \mathrm{dim})$ returns a sample taken along dimension dim of data.
[y,idx] = datasample(data,k,...) returns an index vector indicating which values datasample sampled from data.
$[y, \ldots]=$ datasample(s,data, $k, \ldots$ ) uses the random number stream $s$ to generate random numbers.
[y,...] = datasample(data,k,Name,Value) or [y,...] = datasample(data, k, dim,Name, Value) samples with additional options specified by one or more Name, Value pair arguments.

- To sample random integers with replacement from a range, use randi.
- To sample random integers without replacement, use randperm or datasample.
- To randomly sample from data, with or without replacement, use datasample.


## data

Vector, matrix, $N$-dimensional array, or dataset array representing the data from which to sample. By default, datasample regards the rows of a data matrix, or the first
nonsingleton dimension of a data array, as data elements. Change this behavior with the dim argument.

```
k
```

Positive integer, the number of samples.

```
dim
```

Integer specifying the dimension on which to take samples. For example, if data is a matrix and dim is 2 , $y$ contains a selection of columns in data. If data is a dataset array and dim is 2 , $y$ contains a selection of variables in data. Use dim to ensure sampling along a specific dimension regardless of whether data is a vector, matrix or N -dimensional array.

Default: 1

S
Random number stream. Create s using rng or RandStream.
Default: The global random number stream

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## Replace

Select the sample with replacement if Replace is true, or without replacement if Replace is false. If Replace is false, $k$ must not be larger than the number of data elements in data.

Default: true
Weights

Vector with the same number of elements as data elements in data, and with nonnegative elements. Sample with probability proportional to the elements of Weights.

Default: ones(datasize, 1), where datasize is the number of data elements in data

## Output Arguments

- If data is a vector, y is a vector containing k elements selected from data.
- If data is a matrix, y is a matrix containing k rows selected from data. Or, if dim $=2$, y is a matrix containing k columns selected from data
- If data is an $N$-dimensional array, datasample samples along its first non-singleton dimension. Or, if you give a dim name-value pair, datasample samples along the dimension dim.

When the sample is taken with replacement (default), y can contain repeated observations from data. Set the Replace name-value pair to false to sample without replacement.

Vector of indices indicating which elements datasample chose from data to create y. For example:

- If data is a vector, $y=$ data(idx).
- If data is a matrix, $y=\operatorname{data}(i d x,:)$.


## Examples Draw five unique values from the integers 1:10.

```
y = datasample(1:10,5,'Replace',false)
y =
    6
```

Generate a random sequence of the characters ACGT, with replacement, according to specified probabilities.

```
seq = datasample('ACGT',48,'Weights',[0.15 0.35 0.35 0.15])
seq =
CTTCGACTGTGAGTGGGCGCGACAAGGCTACCGGCCCGGGCGGCACTC
```

Select a random subset of columns from a data matrix.

```
X = randn(10,1000);
Y = datasample(X,5,2,'Replace',false)
\begin{tabular}{|c|c|c|c|c|}
\hline 0.7007 & 0.3382 & 2.1298 & -0.1891 & 0.5026 \\
\hline 0.6520 & -0.6693 & -0.1961 & -0.9915 & 1.9107 \\
\hline 0.1785 & 0.6640 & 2.3247 & -1.1735 & -1.0020 \\
\hline 1.6760 & 2.6102 & -0.8902 & -0.7735 & 1.8676 \\
\hline -0.3251 & -0.6415 & -0.2572 & -0.1629 & -1.0523 \\
\hline 0.1011 & 0.9323 & -1.3088 & -0.4477 & 0.8036 \\
\hline -0.5767 & -0.5778 & -0.8556 & 0.8672 & -0.0727 \\
\hline -0.0615 & -0.9084 & 0.9020 & -0.4185 & -1.9520 \\
\hline 0.7256 & -1.1228 & 0.7558 & 1.2691 & 2.4997 \\
\hline -1.2273 & 0.5754 & -0.8755 & -0.8224 & -1.2066 \\
\hline
\end{tabular}
```

Resample observations from a dataset array to create a bootstrap replicate dataset.
load hospital
y = datasample(hospital,size(hospital,1));

Use the second output to sample "in parallel" from two data vectors.

## datasample

```
x1 = randn(100,1);
x2 = randn(100,1);
[y1,idx] = datasample(x1,10);
y2 = x2(idx);
```

Algorithms
datasample uses randperm, rand, or randi to generate random values. Therefore, datasample changes the state of the MATLAB global random number generator. Control the random number generator using rng.

For selecting weighted samples without replacement, datasample uses the algorithm of Wong and Easton [1].

## References

Alternatives You can use randi or randperm to generate indices for random sampling with or without replacement, respectively. However, datasample can be more convenient because it samples directly from your data. datasample also allows weighted sampling.

See Also rand | randi | randperm | RandStream | rng

## Purpose Arrays for statistical data

## Construction

Use the dataset constructor to create a dataset array from variables in the MATLAB workspace. You can also create a dataset array by reading data from a text or spreadsheet file. You can access each variable in a dataset array much like fields in a structure, using dot subscripting. See the following section for a list of operations available for dataset arrays.

$$
\text { dataset } \quad \text { Construct dataset array }
$$

## Methods

cat
cellstr
datasetfun

Concatenate dataset arrays
Create cell array of strings from dataset array

Apply function to dataset array variables

## dataset

| disp | Display dataset array |
| :--- | :--- |
| display | Display dataset array |
| double | Convert dataset variables to <br> double array |
| end | Last index in indexing expression <br> for dataset array |
| export | Write dataset array to file |
| get | Access dataset array properties |
| grpstats | Summary statistics by group for <br> dataset arrays |
| horzcat | Horizontal concatenation for <br> dataset arrays |
| isempty | True for empty dataset array |
| join | Merge observations |
| length | Length of dataset array |
| ndims | Number of dimensions of dataset <br> array |
| numel | Number of elements in dataset <br> array |
| replacedata | Replace dataset variables |
| set | Set and display properties |
| single | Convert dataset variables to <br> single array |
| size | Size of dataset array <br> sortrows |
| Sort rows of dataset array |  |


| subsasgn | Subscripted assignment to <br> dataset array <br> Subscripted reference for dataset <br> array |
| :--- | :--- |
| summary | Print summary of dataset array |
| unique | Unique observations in dataset <br> array |
| unstack | Unstack data from single variable <br> into multiple variables |
| vertcat | Vertical concatenation for dataset <br> arrays |

## Properties A dataset array D has properties that store metadata (information about your data). Access or assign to a property using $\mathrm{P}=$ D.Properties.PropName or D.Properties.PropName = P, where PropName is one of the following:

| Description | String describing data set <br> Two-element cell array of strings <br> giving names of dimensions of <br> data set |
| :--- | :--- |
| ObsNames | Cell array of nonempty, <br> distinct strings giving names <br> of observations in data set |
| Units | Units of variables in data set |
| UserData | Variable containing additional <br> information associated with data <br> set |


| VarDescription | Cell array of strings giving <br> descriptions of variables in data <br> set |
| :--- | :--- |
| VarNames | Cell array giving names of <br> variables in data set |

## Examples

```
load hospital
h1 = hospital(1:10,:)
h2 = hospital(:,{'LastName' 'Age' 'Sex' 'Smoker'})
% Access and modify metadata
hospital.Properties.Description
hospital.Properties.VarNames{4} = 'Wgt'
% Create a new dataset variable from an existing one
hospital.AtRisk = hospital.Smoker | (hospital.Age > 40)
% Use individual variables to explore the data
boxplot(hospital.Age,hospital.Sex)
h3 = hospital(hospital.Age<30,...
    {'LastName' 'Age' 'Sex' 'Smoker'})
% Sort the observations based on two variables
h4 = sortrows(hospital,{'Sex','Age'})
```


## See Also

How To

- "Dataset Arrays" on page 2-23

| Purpose | Construct dataset array |
| :---: | :---: |
| Syntax | A = dataset(varspec, 'ParamName', Value) |
|  | A = dataset('File',filename,'ParamName', Value) |
|  | A = dataset('XLSFile', filename,'ParamName', Value) |
|  | A = dataset('XPTFile', xptfilename,'ParamName',Val |

Description
A = dataset(varspec,'ParamName', Value) creates dataset array A using one or more of the following workspace variable input methods and one or more optional name/value pairs (see Parameter Name/Value Pairs):

- VAR - a workspace variable. dataset uses the workspace name for the variable name in A. Variables can be arrays of any size, but all variables must have the same number of rows. VAR may also be an expression. In this case, dataset creates a default name automatically.
- \{VAR,name\} - a workspace variable, VAR and a variable name. dataset uses name as the variable name.
- \{VAR,name_1,...,name_m\} - an $m$-columned workspace variable, VAR . dataset uses the names name_1, . . ., name_m as variable names. You must include a name for every column in VAR. Each column becomes a separate variable in A.

You can combine these input methods to include as many variables and names as needed. Names must be valid, unique MATLAB identifier strings. For example input combinations, see Examples. For optional name/value pairs see Inputs.

Note Dataset arrays may contain built-in types or array objects as variables. Array objects must implement each of the following:

- Standard MATLAB parenthesis indexing of the form var(i,...), where $i$ is a numeric or logical vector corresponding to rows of the variable
- A size method with a dim argument
- A vertcat method

A = dataset('File',filename,'ParamName', Value) creates dataset array A from column-oriented data in the text file specified by the string filename. Variables in A are of type double if data in the corresponding column of the file, following the column header, are entirely numeric; otherwise the variables in A are cell arrays of strings. dataset converts empty fields to either NaN (for a numeric variable) or the empty string (for a string-valued variable). dataset ignores insignificant white space in the file. You cannot specify both a file and workspace variables as input. See Name/Value Pairs for more information.

A = dataset('XLSFile',filename,'ParamName', Value) creates dataset array A from column-oriented data in the Excel ${ }^{\circledR}$ spreadsheet specified by the string filename. Variables in A are of type double if data in the corresponding column of the spreadsheet, following the column header, are entirely numeric; otherwise the variables in A are cell arrays of strings. See Name/Value Pairs for more information.

A = dataset('XPTFile', xptfilename,'ParamName', Value) creates a dataset array from a SAS ${ }^{\circledR}$ XPORT format file. Variable names from the XPORT format file are preserved. Numeric data types in the XPORT format file are preserved but all other data types are converted to cell arrays of strings. The XPORT format allows for 28 missing data types. dataset represents these in the file by an upper case letter, '. ' or ' $\quad$ '. dataset converts all missing data to NaN values in A. However, if you
need the specific missing types you can use the xptread function to recover the information. See Name/Value Pairs for more information.

## Parameter Name/Value Pairs <br> Specify one or more of the following name/value pairs when constructing a dataset:

## VarNames

A cell array \{name_1, ..., name_m naming the $m$ variables in A with the specified variable names. Names must be valid, unique MATLAB identifier strings. The number of names must equal the number of variables in A. You cannot use the VarNames parameter if you provide names for individual variables using \{VAR, name\} pairs. To specify VarNames when using a file as input, set ReadVarNames to false.

## ObsNames

A cell array \{name_1, ..., name_n\} naming the $n$ observations in A with the specified observation names. The names need not be valid MATLAB identifier strings, but must be unique. The number of names must equal the number of observations (rows) in A. To specify ObsNames when using a file as input, set ReadObsNames to false.

## Name/value pairs available when using text files as inputs:

## Delimiter

A string indicating the character separating columns in the file. Values are

- ' $\backslash t$ ' (tab, the default when no format is specified)
- ' ' (space, the default when a format is specified)
- ', ' (comma)
- ';' (semicolon)
- ' $\mid$ ' (bar)

Format
A format string, as accepted by textscan. dataset reads the file using textscan, and creates variables in A according to the conversion specifiers in the format string. You may also provide any name/value pairs accepted by textscan. Using the Format parameter is much faster for large files. If ReadObsNames is true, the format string should include a format specifier for the first column of the file.

## HeaderLines

Numeric value indicating the number of lines to skip at the beginning of a file.

## Default: 0

## TreatAsEmpty

Specifies strings to treat as the empty string in a numeric column. Values may be a character string or a cell array of strings. The parameter applies only to numeric columns in the file; dataset does not accept numeric literals such as '-99'.

## Name/value pairs available when using text files or Excel spreadsheets as inputs:

## ReadVarNames

A logical value indicating whether (true) or not (false) to read variable names from the first row of the file. The default is true. If ReadVarNames is true, variable names in the column headers of the file or range (if using an Excel spreadsheet) cannot be empty.

## ReadObsNames

A logical value indicating whether (true) or not (false) to read observation names from the first column of the file or range (if using an Excel spreadsheet). The default is false. If ReadObsNames and ReadVarNames are both true, dataset saves
the header of the first column in the file or range as the name of the first dimension in A. Properties.DimNames.

When reading from an XPT format file, the ReadObsNames parameter name/value pair determines whether or not to try to use the first variable in the file as observation names. Specify as a logical value (default false). If the contents of the first variable are not valid observation names then dataset reads the variable into a variable of the dataset array and does not set the observation names.

## Name/value pairs available when using Excel spreadsheets as input:

## Sheet

A positive scalar value of type double indicating the sheet number, or a quoted string indicating the sheet name.

## Range

A string of the form 'C1:C2' where C1 and C2 are the names of cells at opposing corners of a rectangular region to be read, as for xlsread. By default, the rectangular region extends to the right-most column containing data. If the spreadsheet contains empty columns between columns of data, or if the spreadsheet contains figures or other non-tabular information, specify a range that contains only data.

Create a dataset array from workspace variables, including observation names:

```
load cereal
cereal = dataset(Calories,Protein,Fat,Sodium,Fiber,Carbo,...
    Sugars,'ObsNames',Name)
cereal.Properties.VarDescription = Variables(4:10,2);
```


## dataset

Create a dataset array from a single, multi-columned workspace variable, designating variable names for each column:

```
load cities
categories = cellstr(categories);
cities = dataset({ratings,categories{:}},...
    'ObsNames',cellstr(names))
```

Load data from a text or spreadsheet file

```
patients = dataset('File','hospital.dat',...
    'Delimiter',',','ReadObsNames',true)
patients2 = dataset('XLSFile','hospital.xls',...
    'ReadObsNames',true)
```

1 Load patient data from the CSV file hospital.dat and store the information in a dataset array with observation names given by the first column in the data (patient identification):

```
patients = dataset('file','hospital.dat', ...
    'format','%S%S%S%f%f%f%f%f%f%f%f%f', ...
    'Delimiter',',','ReadObsNames',true);
```

You can also load the data without specifying a format string. dataset will automatically create dataset variables that are either double arrays or cell arrays of strings, depending on the contents of the file:

```
patients = dataset('file','hospital.dat',...
    'delimiter',',',...
    'ReadObsNames',true);
```

2 Make the $\{0,1\}$-valued variable smoke nominal, and change the labels to 'No' and 'Yes':

```
patients.smoke = nominal(patients.smoke,{'No','Yes'});
```

3 Add new levels to smoke as placeholders for more detailed histories of smokers:

```
patients.smoke = addlevels(patients.smoke,...
    {'0-5 Years','5-10 Years','LongTerm'});
```

4 Assuming the nonsmokers have never smoked, relabel the 'No ' level:

```
patients.smoke = setlabels(patients.smoke,'Never','No');
```

5 Drop the undifferentiated 'Yes' level from smoke:

```
patients.smoke = droplevels(patients.smoke,'Yes');
Warning: OLDLEVELS contains categorical levels that
were present in A, caused some array elements to have undefined levels.
```

Note that smokers now have an undefined level.
6 Set each smoker to one of the new levels, by observation name:

```
patients.smoke('YPL-320') = '5-10 Years';
```


## See Also

tdfread | textscan | xlsread

## dataset.datasetfun

Purpose Apply function to dataset array variables
Syntax
b = datasetfun(fun, A)
[b,c,...] = datasetfun(fun,A)
[b,...] = datasetfun(fun,A,...,'UniformOutput',false)
$[b, \ldots]=$ datasetfun(fun, $A, \ldots$, '...DatasetOutput',true)
[b,...] = datasetfun(fun,A,...,'DataVars', vars)
[b,...] = datasetfun(fun,A,...,'ObsNames',obsnames)
[b,...] = datasetfun(fun,A,...,'ErrorHandler',efun)

## Description

b = datasetfun(fun, A) applies the function specified by fun to each variable of the dataset array $A$, and returns the results in the vector $b$. The $i$ th element of b is equal to fun applied to the $i$ th dataset variable of A. fun is a function handle to a function that takes one input argument and returns a scalar value. fun must return values of the same class each time it is called, and datasetfun concatenates them into the vector b. The outputs from fun must be one of the following types: numeric, logical, character, structure, or cell.

To apply functions that return results that are nonscalar or of different sizes and types, use the 'UniformOutput' or 'DatasetOutput' parameters described below.

Do not rely on the order in which datasetfun computes the elements of $b$, which is unspecified.

If fun is bound to more than one built-in function or file, (that is, if it represents a set of overloaded functions), datasetfun follows MATLAB dispatching rules in calling the function. (See "Determining Which Function Gets Called".)
[b,c,...] = datasetfun(fun, A), where fun is a function handle to a function that returns multiple outputs, returns vectors $\mathrm{b}, \mathrm{c}, \ldots$, each corresponding to one of the output arguments of fun. datasetfun calls fun each time with as many outputs as there are in the call to datasetfun. fun may return output arguments having different classes, but the class of each output must be the same each time fun is called.
[b,...] = datasetfun(fun,A,...,'UniformOutput',false) allows you to specify a function fun that returns values of different sizes or types. datasetfun returns a cell array (or multiple cell arrays), where the $i$ th cell contains the value of fun applied to the $i$ th dataset variable of A. Setting 'Uniform0utput' to true is equivalent to the default behavior.
[b, ...] = datasetfun(fun, A, ...,'DatasetOutput', true) specifies that the output(s) of fun are returned as variables in a dataset array (or multiple dataset arrays). fun must return values with the same number of rows each time it is called, but it may return values of any type. The variables in the output dataset array(s) have the same names as the variables in the input. Setting 'DatasetOutput' to false (the default) specifies that the type of the output(s) from datasetfun is determined by 'Uniform0utput'.
[b, ...] = datasetfun(fun, A, ...,'DataVars', vars) allows you to apply fun only to the dataset variables in A specified by vars. vars is a positive integer, a vector of positive integers, a variable name, a cell array containing one or more variable names, or a logical vector.
[b,...] = datasetfun(fun,A,..., 'ObsNames', obsnames) specifies observation names for the dataset output when 'DatasetOutput' is true.
[b,...] = datasetfun(fun,A,...,'ErrorHandler', efun), where efun is a function handle, specifies the MATLAB function to call if the call to fun fails. The error-handling function is called with the following input arguments:

- A structure with the fields identifier, message, and index, respectively containing the identifier of the error that occurred, the text of the error message, and the linear index into the input array(s) at which the error occurred
- The set of input arguments at which the call to the function failed

The error-handling function should either re-throw an error, or return the same number of outputs as fun. These outputs are then returned as

## dataset.datasetfun

the outputs of datasetfun. If 'UniformOutput' is true, the outputs of the error handler must also be scalars of the same type as the outputs of fun. For example, the following code could be saved in a file as the error-handling function:

```
function [A,B] = errorFunc(S,varargin)
warning(S.identifier,S.message);
A = NaN;
B = NaN;
```

If an error-handling function is not specified, the error from the call to fun is rethrown.

## Examples

Compute statistics on selected variables in the hospital dataset array:

```
load hospital
stats = ...
    datasetfun(@mean,hospital,...
                            'DataVars',{'Weight','BloodPressure'},...
                    'UniformOutput',false)
stats =
    [154] [1x2 double]
stats{2}
ans =
    122.7800 82.9600
```

Display the blood pressure variable:

```
datasetfun(@hist,hospital,...
    'DataVars','BloodPressure',...
    'UniformOutput',false);
title('{\bf Blood Pressure}')
legend('Systolic','Diastolic','Location','N')
```

Blood Pressure


See Also
grpstats

Purpose $\quad D$-optimal augmentation

```
Syntax
dCE2 = daugment(dCE,mruns)
[dCE2,X] = daugment(dCE,mruns)
[dCE2,X] = daugment(dCE,mruns,model)
[dCE2,X] = daugment(...,param1,val1,param2,val2,...)
```


## Description

dCE2 = daugment(dCE, mruns) uses a coordinate-exchange algorithm to $D$-optimally add mruns runs to an existing experimental design dCE for a linear additive model.
[dCE2, X] = daugment (dCE, mruns) also returns the design matrix $X$ associated with the augmented design.
[dCE2, X] = daugment(dCE,mruns,model) uses the linear regression model specified in model. model is one of the following strings:

- 'linear' - Constant and linear terms. This is the default.
- 'interaction' - Constant, linear, and interaction terms
- 'quadratic ' - Constant, linear, interaction, and squared terms
- 'purequadratic' - Constant, linear, and squared terms

The order of the columns of $X$ for a full quadratic model with $n$ terms is:

## 1 The constant term

2 The linear terms in order $1,2, \ldots, n$
3 The interaction terms in order (1, 2), (1, 3), $\ldots,(1, n),(2,3), \ldots,(n-1, n)$
4 The squared terms in order $1,2, \ldots, n$
Other models use a subset of these terms, in the same order.
Alternatively, model can be a matrix specifying polynomial terms of arbitrary order. In this case, model should have one column for each factor and one row for each term in the model. The entries in any row
of model are powers for the factors in the columns. For example, if a model has factors $\mathrm{X} 1, \mathrm{X} 2$, and X 3 , then a row [ $\left.\begin{array}{lll}0 & 1 & 2\end{array}\right]$ in model specifies the term (X1.^0).*(X2.^1) .* $\mathrm{X} 3 . \wedge^{\wedge} 2$ ). A row of all zeros in model specifies a constant term, which can be omitted.
[dCE2,X] = daugment(..., param1,val1,param2,val2,...) specifies additional parameter/value pairs for the design. Valid parameters and their values are listed in the following table.

| Parameter | Value |
| :--- | :--- |
| 'bounds' | Lower and upper bounds for each factor, specified <br> as a 2-by-nfactors matrix, where nfactors is the <br> number of factors. Alternatively, this value can be <br> a cell array containing nfactors elements, each <br> element specifying the vector of allowable values for <br> the corresponding factor. |
| 'categorical' | Indices of categorical predictors. |
| 'display' | Either 'on ' or ' off ' to control display of the <br> iteration counter. The default is 'on '. |
| 'excludefun' | Handle to a function that excludes undesirable <br> runs. If the function is $f$, it must support the syntax <br> $b=f(S)$, where $S$ is a matrix of treatments with <br> nfactors columns, where nfactors is the number <br> of factors, and $b$ is a vector of Boolean values with <br> the same number of rows as $S . b(i)$ is true if the $i$ th <br> row $S$ should be excluded. |
| 'init' | Initial design as an mruns-by-nfactors matrix, <br> where nfactors is the number of factors. The <br> default is a randomly selected set of points. |
| 'levels' | Vector of number of levels for each factor. |
| 'maxiter' | Maximum number of iterations. The default is 10. |


| Parameter | Value |
| :---: | :---: |
| 'options' | The value is a structure that contains options specifying whether to compute multiple tries in parallel, and specifying how to use random numbers when generating the starting points for the tries. Create the options structure with statset. Applicable statset parameters are: <br> - 'UseParallel' - If 'always' and if a matlabpool of the Parallel Computing Toolbox is open, compute in parallel. If the Parallel Computing Toolbox is not installed, or a matlabpool is not open, computation occurs in serial mode. Default is 'never', meaning serial computation. <br> - UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'. <br> - Streams - A RandStream object or cell array of such objects. If you do not specify Streams, daugment uses the default stream or streams. If you choose to specify Streams, use a single object except in the case <br> - You have an open MATLAB pool <br> - UseParallel is 'always' <br> - UseSubstreams is 'never' <br> In that case, use a cell array the same size as the MATLAB pool. |


| Parameter | Value |
| :--- | :--- |
|  | For more information on using parallel computing, <br> see Chapter 17, "Parallel Statistics". |
| 'tries ' | Number of times to try to generate a design from <br> a new starting point. The algorithm uses random <br> points for each try, except possibly the first. The <br> default is 1. |

Note The daugment function augments an existing design using a coordinate-exchange algorithm; the 'start' parameter of the candexch function provides the same functionality using a row-exchange algorithm.

## Examples

The following eight-run design is adequate for estimating main effects in a four-factor model:

| ```dCEmain = cordexch(4,8) dCEmain =``` |  |  |  |
| :---: | :---: | :---: | :---: |
| 1 | -1 | -1 | 1 |
| -1 | -1 | 1 | 1 |
| -1 | 1 | -1 | 1 |
| 1 | 1 | 1 | -1 |
| 1 | 1 | 1 | 1 |
| -1 | 1 | -1 | -1 |
| 1 | -1 | -1 | -1 |
| -1 | -1 | 1 | -1 |

To estimate the six interaction terms in the model, augment the design with eight additional runs:

```
dCEinteraction = daugment(dCEmain,8,'interaction')
dCEinteraction =
    1 
```

| -1 | -1 | 1 | 1 |
| ---: | ---: | ---: | ---: |
| -1 | 1 | -1 | 1 |
| 1 | 1 | 1 | -1 |
| 1 | 1 | 1 | 1 |
| -1 | 1 | -1 | -1 |
| 1 | -1 | -1 | -1 |
| -1 | -1 | 1 | -1 |
| -1 | 1 | 1 | 1 |
| -1 | -1 | -1 | -1 |
| 1 | -1 | 1 | -1 |
| 1 | 1 | -1 | 1 |
| -1 | 1 | 1 | -1 |
| 1 | 1 | -1 | -1 |
| 1 | -1 | 1 | 1 |
| 1 | 1 | 1 | -1 |

The augmented design is full factorial, with the original eight runs in the first eight rows.

See Also dcovary | cordexch | candexch

## Purpose

$D$-optimal design with fixed covariates
Syntax
dCV = dcovary(nfactors,fixed)
[dCV,X] = dcovary(nfactors,fixed)
[dCV,X] = dcovary(nfactors,fixed,model)
[dCV, X] = daugment(...,param1,val1,param2,val2,...)
Description
dCV = dcovary(nfactors,fixed) uses a coordinate-exchange algorithm to generate a $D$-optimal design for a linear additive model with nfactors factors, subject to the constraint that the model include the fixed covariate factors in fixed. The number of runs in the design is the number of rows in fixed. The design dCV augments fixed with initial columns for treatments of the model terms.
$[d C V, X]=$ dcovary (nfactors, fixed) also returns the design matrix X associated with the design.
[dCV, X] = dcovary(nfactors,fixed,model) uses the linear regression model specified in model. model is one of the following strings:

- 'linear' - Constant and linear terms. This is the default.
- 'interaction ' - Constant, linear, and interaction terms
- 'quadratic ' - Constant, linear, interaction, and squared terms
- 'purequadratic' - Constant, linear, and squared terms

The order of the columns of $X$ for a full quadratic model with $n$ terms is:
1 The constant term
2 The linear terms in order $1,2, \ldots, n$
3 The interaction terms in order $(1,2),(1,3), \ldots,(1, n),(2,3), \ldots,(n-1, n)$
4 The squared terms in order $1,2, \ldots, n$
Other models use a subset of these terms, in the same order.

Alternatively, model can be a matrix specifying polynomial terms of arbitrary order. In this case, model should have one column for each factor and one row for each term in the model. The entries in any row of model are powers for the factors in the columns. For example, if a model has factors $\mathrm{X} 1, \mathrm{X} 2$, and X 3 , then a row [ 012 1 2 ] in model specifies the term (X1.^0).*(X2.^1).*(X3.^2). A row of all zeros in model specifies a constant term, which can be omitted.
[dCV,X] = daugment(...,param1,val1,param2,val2,...) specifies additional parameter/value pairs for the design. Valid parameters and their values are listed in the following table.

| Parameter | Value |
| :---: | :---: |
| 'bounds ' | Lower and upper bounds for each factor, specified as a 2-by-nfactors matrix. Alternatively, this value can be a cell array containing nfactors elements, each element specifying the vector of allowable values for the corresponding factor. |
| 'categorical' | Indices of categorical predictors. |
| 'display' | Either 'on' or 'off' to control display of the iteration counter. The default is 'on'. |
| 'excludefun' | Handle to a function that excludes undesirable runs. If the function is $f$, it must support the syntax $b=f(S)$, where $S$ is a matrix of treatments with nfactors columns and $b$ is a vector of Boolean values with the same number of rows as $S . b(i)$ is true if the $i$ th row $S$ should be excluded. |
| 'init' | Initial design as an mruns-by-nfactors matrix. The default is a randomly selected set of points. |
| 'levels' | Vector of number of levels for each factor. |
| 'maxiter ${ }^{\text {' }}$ | Maximum number of iterations. The default is 10. |


| Parameter | Value |
| :---: | :---: |
| 'options' | The value is a structure that contains options specifying whether to compute multiple tries in parallel, and specifying how to use random numbers when generating the starting points for the tries. Create the options structure with statset. Applicable statset parameters are: <br> - 'UseParallel' - If 'always' and if a matlabpool of the Parallel Computing Toolbox is open, compute in parallel. If the Parallel Computing Toolbox is not installed, or a matlabpool is not open, computation occurs in serial mode. Default is 'never', meaning serial computation. <br> - UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'. <br> - Streams - A RandStream object or cell array of such objects. If you do not specify Streams, dcovary uses the default stream or streams. If you choose to specify Streams, use a single object except in the case <br> - You have an open MATLAB pool <br> - UseParallel is 'always' <br> - UseSubstreams is 'never' <br> In that case, use a cell array the same size as the MATLAB pool. |


| Parameter | Value |
| :--- | :--- |
|  | For more information on using parallel computing, <br> see Chapter 17, "Parallel Statistics". |
| 'tries ' | Number of times to try to generate a design from <br> a new starting point. The algorithm uses random <br> points for each try, except possibly the first. The <br> default is 1. |

## Examples

## Example 1

Suppose you want a design to estimate the parameters in a three-factor linear additive model, with eight runs that necessarily occur at different times. If the process experiences temporal linear drift, you may want to include the run time as a variable in the model. Produce the design as follows:

```
time = linspace(-1,1,8)';
[dCV1,X] = dcovary(3,time,'linear')
dCV1 =
\begin{tabular}{rrrrr}
-1.0000 & 1.0000 & 1.0000 & -1.0000 & \\
1.0000 & -1.0000 & -1.0000 & -0.7143 & \\
-1.0000 & -1.0000 & -1.0000 & -0.4286 & \\
1.0000 & -1.0000 & 1.0000 & -0.1429 & \\
1.0000 & 1.0000 & -1.0000 & 0.1429 & \\
-1.0000 & 1.0000 & -1.0000 & 0.4286 & \\
1.0000 & 1.0000 & 1.0000 & 0.7143 & \\
-1.0000 & -1.0000 & 1.0000 & 1.0000 & \\
& & & & \\
1.0000 & -1.0000 & 1.0000 & 1.0000 & -1.0000 \\
1.0000 & 1.0000 & -1.0000 & -1.0000 & -0.7143 \\
1.0000 & -1.0000 & -1.0000 & -1.0000 & -0.4286 \\
1.0000 & 1.0000 & -1.0000 & 1.0000 & -0.1429 \\
1.0000 & 1.0000 & 1.0000 & -1.0000 & 0.1429 \\
1.0000 & -1.0000 & 1.0000 & -1.0000 & 0.4286 \\
1.0000 & 1.0000 & 1.0000 & 1.0000 & 0.7143
\end{tabular}
```

$$
\begin{array}{lllll}
1.0000 & -1.0000 & -1.0000 & 1.0000 & 1.0000
\end{array}
$$

The column vector time is a fixed factor, normalized to values between $\pm 1$. The number of rows in the fixed factor specifies the number of runs in the design. The resulting design dCV gives factor settings for the three controlled model factors at each time.

## Example 2

The following example uses the dummyvar function to block an eight-run experiment into 4 blocks of size 2 for estimating a linear additive model with two factors:

```
fixed = dummyvar([1 1 1 2 2 3 3 4 4]);
dCV2 = dcovary(2,fixed(:,1:3),'linear')
dCV2 =
    1 1 1 1 0 0
    -1
    -1}10100010
    1 -1 0
    1 1 0 0 0
    -1
    -1 1
    1
```

The first two columns of dCV2 contain the settings for the two factors; the last three columns are dummy variable codings for the four blocks.

[^5]
## TreeBagger.DefaultYfit property

Purpose Default value returned by predict and oobPredict
Description
The DefaultYfit property controls what predicted value TreeBagger returns when no prediction is possible, for example when the oobPredict method needs to predict for an observation that is in-bag for all trees in the ensemble.

For classification, you can set this property to either ' ' or 'MostPopular'. If you choose 'MostPopular' (default), the property value becomes the name of the most probable class in the training data.
For regression, you can set this property to any numeric scalar. The default is the mean of the response for the training data.

If you set this property to ' ' for classification or NaN for regression, TreeBagger excludes the in-bag observations from computation of the out-of-bagerror and margin.

See Also<br>oobPredict | Predict | OOBIndices

Purpose Delete handle object
Syntax

delete(h)

Description delete $(\mathrm{h})$ deletes the handle object h , where h is a scalar handle. The delete method deletes a handle object but does not clear the handle from the workspace. A deleted handle is no longer valid.

See Also clear | isvalid | qrandstream

## TreeBagger.DeltaCritDecisionSplit property

Purpose Split criterion contributions for each predictor
Description $\quad \begin{aligned} & \text { The DeltaCritDecisionSplit property is a numeric array of size } \\ & \text { 1-by-Nvars of changes in the split criterion summed over splits on each } \\ & \text { variable, averaged across the entire ensemble of grown trees. }\end{aligned}$

See Also classregtree.varimportance


[ $\mathrm{H}, \mathrm{T}$, perm] $=$ dendrogram(...) generates a dendrogram and returns the permutation vector of the node labels of the leaves of the dendrogram. perm is ordered from left to right on a horizontal dendrogram and bottom to top for a vertical dendrogram.
[...] = dendrogram(...,'colorthreshold',t) assigns a unique color to each group of nodes in the dendrogram where the linkage is less than the threshold t . t is a value in the interval $[0, \max (\mathrm{Z}(:, 3))]$. Setting $t$ to the string 'default' is the same as $t=.7(\max (Z(:, 3)))$. 0 is the same as not specifying 'colorthreshold'. The value $\max (Z(:, 3))$ treats the entire tree as one group and colors it all one color.
[...] = dendrogram(...,'orientation','orient') orients the dendrogram within the figure window. Acceptable values for 'orient' are:

| Value | Description |
| :--- | :--- |
| 'top' | Top to bottom (default) |
| 'bottom' | Bottom to top |
| 'left' | Left to right |
| 'right' | Right to left |

[...] = dendrogram(...,'labels', S) accepts a character array or cell array of strings $S$ with one label for each observation. Any leaves in the tree containing a single observation are labeled with that observation's label.

## Examples

```
X = rand(100,2);
Y = pdist(X,'cityblock');
Z = linkage(Y,'average');
[H,T] = dendrogram(Z,'colorthreshold','default');
set(H,'LineWidth',2)
```



This output indicates that leaf node 20 in the dendrogram contains the original data points $20,49,62,65,73$, and 96 .

See Also
cluster | clusterdata | cophenet | inconsistent | linkage | silhouette

## dataset.Description property

Purpose String describing data set
Description Description is a string describing the data set. The default is an empty string

| Purpose | Interactive distribution fitting |
| :--- | :--- |
| Syntax | dfittool <br>  <br>  <br>  <br>  <br>  <br>  <br>  <br>  <br>  <br>  <br>  <br>  <br>  <br>  dfitittooltool $(y)$ |
|  | $y$, cens $)$ |

Description
dfittool opens a graphical user interface for displaying fit distributions to data. To fit distributions to your data and display them over plots of the empirical distributions, you can import data from the workspace.
dfittool(y) displays the Distribution Fitting Tool and creates a data set with data specified by the vector $y$.
dfittool (y, cens) uses the vector cens to specify whether the observation $\mathrm{y}(\mathrm{j})$ is censored, ( cens $(\mathrm{j})==1$ ) and/or observed, exactly (cens $(\mathrm{j})==0$ ). If cens is omitted or empty, no y values are censored.
dfittool (y, cens, freq) uses the vector freq to specify the frequency of each element of $y$. If freq is omitted or empty, all $y$ values have a frequency of 1 .
dfittool(y, cens,freq, dsname) creates a data set with the name dsname using the data vector $y$, censoring indicator cens, and frequency vector freq.
For more information, see "Modeling Data Using the Distribution Fitting Tool" on page 5-11.

See Also mle | randtool | disttool

## qrandset.Dimensions property

Purpose Number of dimensions
Description Number of dimensions in the point set. The Dimensions property of a point set contains a positive integer that indicates the number of dimensions for which the points have values. For example, a point set with Dimensions=5 produces points that each have five values.

Set this property by specifying the number of dimensions when constructing a new point set. After construction, you cannot change the value. The default number of dimensions is 2 .

## dataset.DimNames property

Purpose Two-element cell array of strings giving names of dimensions of data set
Description A two-element cell array of strings giving the names of the two dimensions of the data set. The default is \{'Observations' 'Variables'\}.

Purpose Display categorical array

## Syntax $\quad \operatorname{disp}(A)$

Description disp(A) prints the categorical array A without printing the array name. In all other ways it's the same as leaving the semicolon off an expression, except that empty arrays don't display.

See Also categorical | display

Purpose Display classregtree object
Syntax display(t)
Description display(t) prints the classregtree object $t$.
See Also ..... classregtree | view

## cvpartition.disp

Purpose Display cvpartition object

## Syntax disp(c)

Description disp(c) prints the cvpartition object c.
See Also cvpartition

| Purpose | Display dataset array |
| :--- | :--- |
| Syntax | disp(ds) |
| Description | disp(ds) prints the dataset array ds, including variable names and <br> observation names (if present), without printing the dataset name. In <br> all other ways it's the same as leaving the semicolon off an expression. |
| For numeric or categorical variables that are 2-D and have three or <br> fewer columns, disp prints the actual data using either short g, long <br> g, or bank format, depending on the current command line setting. |  |
| Otherwise, disp prints the size and type of each dataset element. |  |
| For character variables that are 2-D and 10 or fewer characters wide, |  |
| disp prints quoted strings. Otherwise, disp prints the size and type of |  |
| each dataset element. |  |
| For cell variables that are 2-D and have three or fewer columns, |  |
| disp prints the contents of each cell (or its size and type if too large). |  |
| Otherwise, disp prints the size of each dataset element. |  |

## gmdistribution.disp

Purpose Display Gaussian mixture distribution object

## Syntax disp(obj)

Description disp(obj) prints a text representation of the gmdistribution object, obj, without printing the object name. In all other ways it's the same as leaving the semicolon off an expression.

See Also gmdistribution | display
Purpose Display NaiveBayes classifier object
Syntax disp(nb)
Description disp(nb) prints a text representation of the NaiveBayes object nb,without printing the object name. In all other ways it's the same asleaving the semicolon off an expression.
See Also NaiveBayes | display

## piecewisedistribution.disp

Purpose Display piecewisedistribution object

## Syntax $\quad \operatorname{disp}(A)$

Description disp(A) prints a text representation of the piecewisedistribution object A, without printing the object name. In all other ways it's the same as leaving the semicolon off an expression.

See Also piecewisedistribution

## Purpose Display qrandset object

## Syntax disp(p)

Description disp(p) displays the properties of the quasi-random point set $s$, without printing the variable name. disp prints out the number of dimensions and points in the point-set, and follows this with the list of all property values for the object.

See Also qrandset

## qrandstream.disp

Purpose Display qrandstream object

## Syntax disp(q)

Description disp(q) displays the quasi-random stream $q$, without printing the variable name. disp prints the type and number of dimensions in the stream, and follows it with the list of point set properties.

See Also qrandstream
Purpose Display categorical array
Syntax display(A)
Description display(A) prints the categorical array A. categorical callsdisplay when a you do not use a semicolon to terminate a statement.
See Also ..... categorical | disp

## classregtree.display

Purpose Display classregtree object
Syntax display(t) ..... display(A)
Description display(t) prints the classregtree object t. classregtree callsdisplay when a you do not use a semicolon to terminate a statement.display(A) prints the categorical array A. categorical callsdisplaywhen a you do not use a semicolon to terminate a statement.
See Also ..... classregtree | eval | prune | test

Purpose Display cvpartition object

## Syntax display(c)

Description display (c) prints the cvpartition object c. cvpartition callsdisplay when a you do not use a semicolon to terminate a statement.

See Also cupartition

## dataset.display

Purpose Display dataset array

## Syntax display(ds)

Description display (ds) prints the dataset array ds, including variable names and observation names (if present). dataset callsdisplay when a you do not use a semicolon to terminate a statement

For numeric or categorical variables that are 2-D and have three or fewer columns, display prints the actual data. Otherwise, display prints the size and type of each dataset element.

For character variables that are 2-D and 10 or fewer characters wide, display prints quoted strings. Otherwise, display prints the size and type of each dataset element.

For cell variables that are 2-D and have three or fewer columns, display prints the contents of each cell (or its size and type if too large). Otherwise, display prints the size of each dataset element.
For time series variables, display prints columns for both the time and the data. If the variable is $2-\mathrm{D}$ and has three or fewer columns, display prints the actual data. Otherwise, display prints the size and type of each dataset element.

For other types of variables, display prints the size and type of each dataset element.

See Also dataset | display | format

Purpose Display Gaussian mixture distribution object

## Syntax display (obj)

Description display(obj) prints a text representation of the gmdistribution object obj. gmdistribution callsdisplay when a you do not use a semicolon to terminate a statement.

See Also gmdistribution | disp

## NaiveBayes.display

Purpose Display NaiveBayes classifier object
Syntax display(nb)
Description display (nb) prints a text representation of the NaiveBayes object nb. NaiveBayes callsdisplay when a you do not use a semicolon to terminate a statement.
See Also NaiveBayes | display
Purpose Display piecewisedistribution object
Syntax

display(A)

Description display (A) prints a text representation of the piecewisedistribution object A, without printing the object name. piecewisedistribution callsdisplay when a you do not use a semicolon to terminate a statement.

See Also<br>piecewisedistribution

## ProbDist.DistName property

```
Purpose Read-only string containing probability distribution name of ProbDist object
```


## Description <br> DistName is a read-only property of the ProbDist class. DistName is a

``` string containing the type of distribution used to create the object.
```


## Values <br> Possible values are:

```
- 'kernel'
- 'beta'
- 'binomial'
- 'birnbaumsaunders'
- 'exponential'
- 'extreme value'
- 'gamma'
- 'generalized extreme value'
- 'generalized pareto'
- 'inversegaussian'
- 'logistic'
- 'loglogistic'
- 'lognormal'
- 'nakagami'
- 'negative binomial'
- 'normal'
- 'poisson'
- 'rayleigh'
- 'rician'
```


## ProbDist.DistName property

- 'tlocationscale'
- 'weibull'

Use this information to view and compare the type of distribution used to create distribution objects.

## NaiveBayes.Dist property

## Purpose Distribution names

Description The Dist property is a string or a 1-by-NDims cell array of strings indicating the types of distributions for all the features. If all the features use the same type of distribution, Dist is a single string. Otherwise Dist(j) indicates the distribution type used for the j th feature.

The valid strings for this property are the following:

| 'normal' | Normal distribution. |
| :--- | :--- |
| 'kernel' | Kernel smoothing density <br> estimate. |
| 'mvmn' | Multivariate multinomial <br> distribution. |
| 'mn' | Multinomial bag-of-tokens model. |

## gmdistribution.DistName property

Purpose Type of distribution<br>Description The string 'gaussian mixture distribution'.

## disttool

Purpose Interactive density and distribution plots

## Syntax disttool

Description disttool is a graphical interface for exploring the effects of changing parameters on the plot of a cdf or pdf.

See Also randtool | dfittool

Purpose Convert categorical array to double array
Syntax $\quad B=\operatorname{double}(A)$
Description $\quad B=$ double $(A)$ converts the categorical array $A$ to a double array. Each element of B contains the internal categorical level code for the corresponding element of $A$.

See Also single

Purpose Convert dataset variables to double array
Syntax
b = double(A)
b = double(a, vars)

Description
$b=\operatorname{double}(A)$ returns the contents of the dataset $A$, converted to one double array. The classes of the variables in the dataset must support the conversion.
b = double(a,vars) returns the contents of the dataset variables specified by vars. vars is a positive integer, a vector of positive integers, a variable name, a cell array containing one or more variable names, or a logical vector.

## See Also

dataset | single | replacedata

## Purpose Drop levels

Syntax $\quad B=\operatorname{droplevels}(A)$
B = droplevels(A,oldlevels)
Description $\quad B=\operatorname{droplevels}(A)$ removes unused levels from the categorical array
A. B is a categorical array with the same size and values as A, but with a list of potential levels that includes only those present in some element of A .

B = droplevels(A, oldlevels) removes specified levels from the categorical array A. oldlevels is a cell array of strings or a 2-D character matrix specifying the levels to be removed.
droplevels removes levels, but does not remove elements. Elements of $B$ that correspond to elements of A having levels in oldlevels all have an undefined level.

## Examples

## Example 1

Drop unused age levels from the data in hospital.mat:

```
load hospital
edges = 0:10:100;
labels = strcat(num2str((0:10:90)','%d'),{'s'});
AgeGroup = ordinal(hospital.Age,labels,[],edges);
AgeGroup = droplevels(AgeGroup);
getlabels(AgeGroup)
ans =
    '20s' '30s' '40s' '50s'
```


## Example 2

1 Load patient data from the CSV file hospital. dat and store the information in a dataset array with observation names given by the first column in the data (patient identification):

```
patients = dataset('file','hospital.dat',...
    'delimiter',',',...
```

```
'ReadObsNames',true);
```

2 Make the $\{0,1\}$-valued variable smoke nominal, and change the labels to 'No' and 'Yes':

```
patients.smoke = nominal(patients.smoke,{'No','Yes'});
```

3 Add new levels to smoke as placeholders for more detailed histories of smokers:

```
patients.smoke = addlevels(patients.smoke,...
    {'0-5 Years','5-10 Years','LongTerm'});
```

4 Assuming the nonsmokers have never smoked, relabel the 'No ' level:

```
patients.smoke = setlabels(patients.smoke,'Never','No');
```

5 Drop the undifferentiated 'Yes' level from smoke:
patients.smoke = droplevels(patients.smoke,'Yes');
Warning: OLDLEVELS contains categorical levels that were present in A, caused some array elements to have undefined levels.

Note that smokers now have an undefined level.
6 Set each smoker to one of the new levels, by observation name:

```
patients.smoke('YPL-320') = '5-10 Years';
```

See Also
addlevels | getlabels | islevel | mergelevels | reorderlevels

Purpose
Create dummy variables
Syntax
Description
D = dummyvar (group)
$D=$ dummyvar(group) returns a matrix $D$ containing zeros and ones, whose columns are dummy variables for the grouping variable group. Columns of group represent categorical predictor variables, with values indicating categorical levels. Rows of group represent observations across variables.
group can be a numeric vector or categorical column vector representing levels within a single variable, a cell array containing one or more grouping variables, or a numeric matrix or cell array of categorical column vectors representing levels within multiple variables. If group is a numeric vector or matrix, values in any column must be positive integers in the range from 1 to the number of levels for the corresponding variable. In this case, dummyvars treats each column as a separate numeric grouping variable. With multiple grouping variables, the sets of dummy variable columns are in the same order as the grouping variables in group.

The order of the dummy variable columns in D matches the order of the groups defined by group. When group is a categorical vector, the groups and their order match the output of the getlabels (group) method. When group is a numeric vector, dummyvar assumes that the groups and their order are 1: max (group). In this respect, dummyvars treats a numeric grouping variable differently than grp2idx.
If group is $n$-by- $p, \mathrm{D}$ is $n$-by- $S$, where $S$ is the sum of the number of levels in each of the columns of group. The number of levels $s$ in any column of group is the maximum positive integer in the column or the number of categorical levels. Levels are considered distinct if they appear in different columns of group, even if they have the same value. Columns of $D$ are, from left to right, dummy variables created from the first column of group, followed by dummy variables created from the second column of group, etc.
dummyvar treats NaN values or undefined categorical levels in group as missing data and returns NaN values in D.

Dummy variables are used in regression analysis and ANOVA to indicate values of categorical predictors．

Note If a column of 1 s is introduced in the matrix D ，the resulting matrix $X=$［ones（size（ $\mathrm{D}, 1$ ），1）D］will be rank deficient．The matrix $D$ itself will be rank deficient if group has multiple columns． This is because dummy variables produced from any column of group always sum to a column of 1s．Regression and ANOVA calculations often address this issue by eliminating one dummy variable（implicitly setting the coefficients for dropped columns to zero）from each group of dummy variables produced by a column of group．

## Examples

Suppose you are studying the effects of two machines and three operators on a process．Use group to organize predictor data on machine－operator combinations：

```
machine = [\begin{array}{llllllll}{1}&{1}&{1}&{1}&{2}&{2}&{2}&{2]';}\end{array};⿱一⿻上丨
operator = [1 2 3 1 2 3 1 2]';
group = [machine operator]
group =
    1
    1 2
    1 3
    1
    2 2
    2 3
    2 1
    2
```

Use dummyvar to create dummy variables for a regression or ANOVA calculation：

```
D = dummyvar(group)
D =
    1 0
```

| 1 | 0 | 0 | 1 | 0 |
| :--- | :--- | :--- | :--- | :--- |
| 1 | 0 | 0 | 0 | 1 |
| 1 | 0 | 1 | 0 | 0 |
| 0 | 1 | 0 | 1 | 0 |
| 0 | 1 | 0 | 0 | 1 |
| 0 | 1 | 1 | 0 | 0 |
| 0 | 1 | 0 | 1 | 0 |

The first two columns of D represent observations of machine 1 and machine 2 , respectively; the remaining columns represent observations of the three operators.

## See Also regress | anova1

How To . "Grouped Data" on page 2-34

Purpose Durbin-Watson test
Syntax $\quad[P, D W]=\operatorname{dwtest}(R, X)$
[...] = dwtest(R,X,method)
[...] = dwtest(R,X,method,tail)

## Description

$[P, D W]=\operatorname{dwtest}(R, X)$ performs a Durbin-Watson test on the vector $R$ of residuals from a linear regression, where $X$ is the design matrix from that linear regression. $P$ is the computed $p$ value for the test, and DW is the Durbin-Watson statistic. The Durbin-Watson test is used to test if the residuals are uncorrelated, against the alternative that there is autocorrelation among them. Small values of P indicate that the residuals are correlated.
[...] = dwtest( $R, X$, method) specifies the method to be used in computing the p value. method can be either of the following:

- 'exact' - Calculates an exact $p$ value using the PAN algorithm (the default if the sample size is less than 400).
- 'approximate' - Calculates the p value using a normal approximation (the default if the sample size is 400 or larger).
[...] = dwtest(R,X,method,tail) performs the test against one of the following alternative hypotheses, specified by tail:

| Tail | Alternative Hypothesis |
| :--- | :--- |
| 'both' | Serial correlation is not 0. |
| 'right' | Serial correlation is greater than 0 (right-tailed test). |
| 'left' | Serial correlation is less than 0 (left-tailed test). |

## Examples

Fit a straight line to the census data, plot the residuals, and note their autocorrelation:

```
load census
```

$\mathrm{n}=$ length(cdate);

```
X = [ones(n,1),cdate];
[b,bint,r1] = regress(pop,X);
p1 = dwtest(r1,X)
plot(cdate,r1,'b-',cdate,zeros(n,1),'k:')
p1 =
    0
```

The test strongly rejects the hypothesis of no correlation in the residuals.


Adding a squared term gives less correlation in the residuals, but the resulting $p$ value is still significantly small:

```
X = [ones(n,1),cdate,cdate.^2];
[b,bint,r2] = regress(pop,X);
p2 = dwtest(r2,X)
```

line(cdate, r2,'color', 'r')
p2 =
0.0041


Purpose
Empirical cumulative distribution function
Syntax

```
\([f, x]=\operatorname{ecdf}(y)\)
[f,x,flo,fup] = ecdf(y)
ecdf(...)
ecdf(ax,...)
[...] = ecdf(y,param1,val1,param2,val2,...)
```

Description
$[f, x]=\operatorname{ecdf}(y)$ calculates the Kaplan-Meier estimate of the cumulative distribution function (cdf), also known as the empirical cdf. $y$ is a vector of data values. $f$ is a vector of values of the empirical cdf evaluated at $x$.
$[f, x, f l o, f u p]=\operatorname{ecdf}(y)$ also returns lower and upper confidence bounds for the cdf. These bounds are calculated using Greenwood's formula, and are not simultaneous confidence bounds.
$\operatorname{ecdf}(\ldots)$ without output arguments produces a plot of the empirical cdf.
$\operatorname{ecdf}(a x, \ldots)$ plots into axes ax instead of gca.
[...] = ecdf(y,param1,val1,param2,val2,...) specifies additional parameter/value pairs chosen from the following:

| Parameter | Value |
| :--- | :--- |
| 'censoring' | Boolean vector of the same size as $x$. Elements are <br> 1 for observations that are right-censored and 0 for <br> observations that are observed exactly. Default is all <br> observations observed exactly. |
| 'frequency' | Vector of the same size as $x$ containing nonnegative <br> integer counts. The jth element of this vector <br> gives the number of times the jth element of $x$ was <br> observed. Default is 1 observation per element of $x$. |
| 'alpha' | Value between 0 and 1 for a confidence level of <br> $100(1-a l p h a) \% . ~ D e f a u l t ~ i s ~ a l p h a=0.05 ~ f o r ~ 95 \% ~$ |
| confidence. |  |.


| Parameter | Value |
| :--- | :--- |
| 'function' | Type of function returned as the f output argument, <br> chosen from 'cdf' (default), 'survivor', or <br> 'cumulative hazard'. |
| 'bounds ' | Either 'on' to include bounds, or 'off' (the default) <br> to omit them. Used only for plotting. |

## Examples

Generate random failure times and random censoring times, and compare the empirical cdf with the known true cdf:

```
y = exprnd(10,50,1); % Random failure times exponential(10)
d = exprnd(20,50,1); % Drop-out times exponential(20)
t = min(y,d); % Observe the minimum of these times
censored = (y>d); % Observe whether the subject failed
% Calculate and plot empirical cdf and confidence bounds
[f,x,flo,fup] = ecdf(t,'censoring',censored);
stairs(x,f,'LineWidth',2)
hold on
stairs(x,flo,'r:','LineWidth', 2)
stairs(x,fup,'r:','LineWidth',2)
% Superimpose a plot of the known population cdf
xx = 0:.1:max(t);
yy = 1-exp(-xx/10);
plot(xx,yy,'g-','LineWidth',2)
legend('Empirical','LCB','UCB','Population',...
    'Location','SE')
hold off
```



## References

See Also cdfplot | ecdfhist

Purpose Empirical cumulative distribution function histogram

Syntax<br>\section*{Description}

```
n = ecdfhist(f,x)
n = ecdfhist(f,x,m)
n = ecdfhist(f,x,c)
[n,c] = ecdfhist(...)
ecdfhist(...)
```


## Examples

$n=$ ecdfhist $(f, x)$ takes a vector $f$ of empirical cumulative distribution function (cdf) values and a vector x of evaluation points, and returns a vector n containing the heights of histogram bars for 10 equally spaced bins. The function computes the bar heights from the increases in the empirical cdf, and normalizes them so that the area of the histogram is equal to 1 . In contrast, hist produces bars whose heights represent bin counts.
$n=\operatorname{ecdfhist}(f, x, m)$, where $m$ is a scalar, uses $m$ bins.
$n=$ ecdfhist $(f, x, c)$, where $c$ is a vector, uses bins with centers specified by c.
[ $n, c$ ] = ecdfhist(...) also returns the position of the bin centers in c .
ecdfhist(...) without output arguments produces a histogram bar plot of the results.

The following code generates random failure times and random
censoring times, and compares the empirical pdf with the known true pdf.

```
y = exprnd(10,50,1); % Random failure times
d = exprnd(20,50,1); % Drop-out times
t = min(y,d); % Observe the minimum of these times
censored = (y>d); % Observe whether the subject failed
% Calculate the empirical cdf and plot a histogram from it
[f,x] = ecdf(t,'censoring',censored);
ecdfhist(f,x)
```

```
set(get(gca,'Children'),'FaceColor',[.8 . 8 1])
hold on
```

\% Superimpose a plot of the known population pdf
xx = 0:.1:max(t);
yy = exp(-xx/10)/10;
plot(xx,yy,'r-','LineWidth',2)
hold off


See Also
ecdf | hist | histc

## CompactClassificationDiscriminant.edge

> Purpose Classification edge
> Syntax $\quad E=\operatorname{edge}(o b j, X, Y)$ E = edge(obj, X,Y,Name, Value)
> Description $E=\operatorname{edge}(o b j, X, Y)$ returns the classification edge for obj with data $X$ and classification $Y$.
> $\mathrm{E}=$ edge (obj, $\mathrm{X}, \mathrm{Y}$, Name, Value) computes the edge with additional options specified by one or more Name, Value pair arguments.
> Discriminant analysis classifier of class ClassificationDiscriminant or CompactClassificationDiscriminant, typically constructed with ClassificationDiscriminant.fit.
> x
> Matrix where each row represents an observation, and each column represents a predictor. The number of columns in X must equal the number of predictors in obj.
> Class labels, with the same data type as exists in obj. The number of elements of $Y$ must equal the number of rows of $X$.
> Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
> Observation weights, a numeric vector of length size ( $X, 1$ ). If you supply weights, edge computes the weighted classification edge.

## CompactClassificationDiscriminant.edge

## Output <br> Arguments

## Definitions

## Edge

The edge is the weighted mean value of the classification margin. The weights are the class probabilities in obj. Prior. If you supply weights in the weights name-value pair, those weights are normalized to sum to the prior probabilities in the respective classes, and are then used to compute the weighted average.

## Margin

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix X . A high value of margin indicates a more reliable prediction than a low value.

## Score (discriminant analysis)

For discriminant analysis, the score of a classification is the posterior probability of the classification. For the definition of posterior probability in discriminant analysis, see "Posterior Probability" on page 12-7.

Examples Compute the classification edge and margin for the Fisher iris data, trained on its first two columns of data, and view the last 10 entries:

```
load fisheriris
X = meas(:,1:2);
obj = ClassificationDiscriminant.fit(X,species);
E = edge(obj,X,species)
E =
    0.4980
```


## CompactClassificationDiscriminant.edge

```
M = margin(obj,X,species);
M(end-10:end)
ans =
    0 . 6 5 5 1
    0.4838
    0.6551
    -0.5127
    0.5659
    0.4611
    0.4949
    0.1024
    0.2787
    -0.1439
    -0.4444
```

The classifier trained on all the data is better:

```
obj = ClassificationDiscriminant.fit(meas,species);
E = edge(obj,meas,species)
E =
    0.9454
M = margin(obj,meas,species);
M(end-10:end)
ans =
    0.9983
    1.0000
    0.9991
    0.9978
    1.0000
    1.0000
    0.9999
    0.9882
```


# CompactClassificationDiscriminant.edge 

0.9937
1.0000
0.9649

See Also ClassificationDiscriminant | loss | margin | predict
How To . "Discriminant Analysis" on page 12-3

## CompactClassificationEnsemble.edge

Purpose Classification edge
$\begin{aligned} \text { Syntax } & E=\text { edge(ens, } X, Y) \\ E & =\text { edge(ens, } X, Y, \text { Name, value) }\end{aligned}$
Description $E=$ edge (ens $, X, Y)$ returns the classification edge for ens with data $X$ and classification $Y$.
$\mathrm{E}=$ edge(ens, $\mathrm{X}, \mathrm{Y}$, Name, Value) computes the edge with additional options specified by one or more Name, Value pair arguments.

## Input

Arguments
A classification ensemble constructed with fitensemble, or a compact classification ensemble constructed with compact.

X
A matrix where each row represents an observation, and each column represents a predictor. The number of columns in X must equal the number of predictors in ens.

Y
Class labels, with the same data type as exists in ens. The number of elements of $Y$ must equal the number of rows of $X$.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## learners

Indices of weak learners in the ensemble ranging from 1 to ens. NTrained. edge uses only these learners for calculating loss.

Default: 1:NTrained

## CompactClassificationEnsemble.edge

mode
String representing the meaning of the output E :

- 'ensemble' - E is a scalar value, the edge for the entire ensemble.
- 'individual' - E is a vector with one element per trained learner.
- 'cumulative' - E is a vector in which element J is obtained by using learners $1: \mathrm{J}$ from the input list of learners.

Default: 'ensemble'

## UseObsForLearner

A logical matrix of size N -by-T, where:

- $N$ is the number of rows of $X$.
- T is the number of weak learners in ens.

When UseObsForLearner( $\mathrm{i}, \mathrm{j}$ ) is true, learner j is used in predicting the class of row $i$ of $X$.

Default: true (N,T)

## weights

Observation weights, a numeric vector of length size ( $\mathrm{X}, 1$ ). If you supply weights, edge computes weighted classification edge.

Default: ones(size(X,1))

## Output E

Arguments
The classification edge, a vector or scalar depending on the setting of the mode name-value pair. Classification edge is weighted average classification margin.

## CompactClassificationEnsemble.edge

## Definitions

## Examples

Make a boosted ensemble classifier for the ionosphere data, and find the classification edge for the last few rows:

```
load ionosphere
ens = fitensemble(X,Y,'AdaboostM1',100,'Tree');
E = edge(ens,X(end-10:end,:),Y(end-10:end))
E =
    8.3310
```

See Also margin \| edge

## CompactClassificationEnsemble.edge

How To - Chapter 13, "Nonparametric Supervised Learning"

## CompactClassificationTree.edge

Purpose Classification edge
Syntax $\quad \begin{aligned} & E=\operatorname{edge}(\text { tree }, X, Y) \\ & E=\text { edge(tree }, X, Y, \text { Name, Value })\end{aligned}$
Description $E=$ edge(tree $, X, Y$ ) returns the classification edge for tree with data $X$ and classification $Y$.
$\mathrm{E}=$ edge(tree, $\mathrm{X}, \mathrm{Y}$, Name, Value) computes the edge with additional options specified by one or more Name, Value pair arguments.

## Input <br> Arguments

tree
A classification tree created by ClassificationTree.fit, or a compact classification tree created by compact.
x
A matrix where each row represents an observation, and each column represents a predictor. The number of columns in X must equal the number of predictors in tree.
Y
Class labels, with the same data type as exists in tree. The number of elements of $Y$ must equal the number of rows of $X$.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## weights

Observation weights, a numeric vector of length size ( $\mathrm{X}, 1$ ). If you supply weights, edge computes weighted classification edge.

Default: ones(size(X,1))

## CompactClassificationTree.edge

## Output Arguments <br> E

## Definitions

The edge, a scalar representing the weighted average value of the margin.

## Margin

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as the matrix X .

## Score (tree)

For trees, the score of a classification of a leaf node is the posterior probability of the classification at that node. The posterior probability of the classification at a node is the number of training sequences that lead to that node with the classification, divided by the number of training sequences that lead to that node.

For example, consider classifying a predictor X as true when $\mathrm{X}<0.15$ or $X>0.95$, and $X$ is false otherwise.

1
Generate 100 random points and classify them:

```
rng(0,'twister') % for reproducibility
X = rand(100,1);
Y = (abs(X - . 55) > .4);
tree = ClassificationTree.fit(X,Y);
view(tree,'mode','graph')
```


## CompactClassificationTree.edge



2
Prune the tree:

```
tree1 = prune(tree,'level',1);
view(tree1,'mode','graph')
```


## CompactClassificationTree.edge



The pruned tree correctly classifies observations that are less than 0.15 as true. It also correctly classifies observations from .15 to .94 as false. However, it incorrectly classifies observations that are greater than . 94 as false. Therefore, the score for observations that are greater than . 15 should be about $.05 / .85=.06$ for true, and about $.8 / .85=.94$ for false.

## 3

Compute the prediction scores for the first 10 rows of X :

```
[~,score] = predict(tree1,X(1:10));
[score X(1:10,:)]
ans =
\begin{tabular}{rrr}
0.9059 & 0.0941 & 0.8147 \\
0.9059 & 0.0941 & 0.9058 \\
0 & 1.0000 & 0.1270 \\
0.9059 & 0.0941 & 0.9134 \\
0.9059 & 0.0941 & 0.6324 \\
0 & 1.0000 & 0.0975 \\
0.9059 & 0.0941 & 0.2785
\end{tabular}
```


## CompactClassificationTree.edge

| 0.9059 | 0.0941 | 0.5469 |
| :--- | :--- | :--- |
| 0.9059 | 0.0941 | 0.9575 |
| 0.9059 | 0.0941 | 0.9649 |

Indeed, every value of $X$ (the rightmost column) that is less than 0.15 has associated scores (the left and center columns) of 0 and 1, while the other values of $X$ have associated scores of 0.91 and 0.09 . The difference (score 0.09 instead of the expected .06) is due to a statistical fluctuation: there are 8 observations in $X$ in the range (.95,1) instead of the expected 5 observations.

## Edge

The edge is the weighted mean value of the classification margin. The weights are the class probabilities in tree. Prior. If you supply weights in the weights name-value pair, those weights are normalized to sum to the prior probabilities in the respective classes, and are then used to compute the weighted average.

## Examples Compute the classification margin and edge for the Fisher iris data,

 trained on its first two columns of data, and view the last 10 entries:```
load fisheriris
X = meas(:,1:2);
tree = ClassificationTree.fit(X,species);
E = edge(tree,X,species)
E =
    0.6299
M = margin(tree,X,species);
M(end-10:end)
ans =
    0.1111
    0.1111
    0.1111
    -0.2857
```


## CompactClassificationTree.edge

0.6364
0.6364
0.1111
0.7500
1.0000
0.6364
0.2000

The classification tree trained on all the data is better:

```
tree = ClassificationTree.fit(meas,species);
E = edge(tree,meas,species)
E =
    0.9384
M = margin(tree,meas,species);
M(end-10:end)
ans =
    0.9565
    0.9565
    0.9565
    0.9565
    0.9565
    0.9565
    0.9565
    0.9565
    0.9565
    0.9565
    0.9565
```

See Also
margin | loss | predict

How To - Chapter 13, "Nonparametric Supervised Learning"

Purpose Last index in indexing expression for categorical array

## Syntax end $(A, k, n)$

Description end ( $A, k, n$ ) indexes expressions involving the categorical array $A$ when end is part of the $k$-th index out of $n$ indices. For example, the expression $A(e n d-1,:)$ calls A's end method with end (A, 1,2).

See Also single

Purpose
Last index in indexing expression for dataset array

## Syntax <br> end ( $A, k, n$ )

Description end $(A, k, n)$ is called for indexing expressions involving the dataset $A$ when end is part of the $k$-th index out of $n$ indices. For example, the expression A(end-1,:) calls A's end method with end (A, 1,2).

## qrandset.end

Purpose Last index in indexing expression for point set

## Syntax <br> end ( $\mathrm{p}, \mathrm{k}, \mathrm{n}$ )

Description end $(p, k, n)$ is called for indexing expressions involving the point set $p$ when end is part of the $k$-th index out of $n$ indices. For example, the expression $p$ (end-1,:) calls p's end method with end $(p, 1,2)$.

See Also qrandset

## Purpose <br> Extreme value cumulative distribution function

## Syntax

P = evcdf(X,mu,sigma)
[P,PLO,PUP] = evcdf(X,mu,sigma,pcov,alpha)
Description
$P=\operatorname{evcdf}(X, m u$, sigma) computes the cumulative distribution function
(cdf) for the type 1 extreme value distribution, with location parameter mu and scale parameter sigma, at each of the values in $\mathrm{X} . \mathrm{X}, \mathrm{mu}$, and sigma can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array of the same size as the other inputs. The default values for mu and sigma are 0 and 1, respectively.
[P,PLO,PUP] = evcdf(X,mu,sigma,pcov,alpha) produces confidence bounds for $P$ when the input parameters mu and sigma are estimates. pcov is a 2 -by- 2 covariance matrix of the estimated parameters. alpha has a default value of 0.05 , and specifies $100(1$ - alpha) $\%$ confidence bounds. PLO and PUP are arrays of the same size as P, containing the lower and upper confidence bounds.

The function evcdf computes confidence bounds for P using a normal approximation to the distribution of the estimate

$$
\frac{X-\hat{\mu}}{\hat{\sigma}}
$$

and then transforming those bounds to the scale of the output P. The computed bounds give approximately the desired confidence level when you estimate mu, sigma, and pcov from large samples, but in smaller samples other methods of computing the confidence bounds might be more accurate.

The type 1 extreme value distribution is also known as the Gumbel distribution. The version used here is suitable for modeling minima; the mirror image of this distribution can be used to model maxima by negating X. See "Extreme Value Distribution" on page B-19 for more details. If $x$ has a Weibull distribution, then $X=\log (x)$ has the type 1 extreme value distribution.

See Also cdf | evpdf | evinv \| evstat | evfit | evlike | evrnd<br>How To . "Extreme Value Distribution" on page B-19

Purpose
Syntax
Description
Extreme value parameter estimates

```
parmhat = evfit(data)
[parmhat,parmci] = evfit(data)
[parmhat,parmci] = evfit(data,alpha)
[...] = evfit(data,alpha,censoring)
[...] = evfit(data,alpha,censoring,freq)
[...] = evfit(data,alpha,censoring,freq,options)
```

parmhat = evfit(data) returns maximum likelihood estimates of the parameters of the type 1 extreme value distribution given the data in the vector data. parmhat (1) is the location parameter, , and parmhat(2) is the scale parameter, $\sigma$.
[parmhat, parmci] = evfit(data) returns $95 \%$ confidence intervals for the parameter estimates on the and $\sigma$ parameters in the 2 -by- 2 matrix parmci. The first column of the matrix of the extreme value fit contains the lower and upper confidence bounds for the parameter , and the second column contains the confidence bounds for the parameter $\sigma$.
[parmhat,parmci] = evfit(data,alpha) returns 100(1-alpha)\% confidence intervals for the parameter estimates, where alpha is a value in the range [ $\left.\begin{array}{ll}0 & 1\end{array}\right]$ specifying the width of the confidence intervals. By default, alpha is 0.05 , which corresponds to $95 \%$ confidence intervals.
[...] = evfit(data,alpha, censoring) accepts a Boolean vector, censoring, of the same size as data, which is 1 for observations that are right-censored and 0 for observations that are observed exactly.
[...] = evfit(data,alpha,censoring,freq) accepts a frequency vector, freq of the same size as data. Typically, freq contains integer frequencies for the corresponding elements in data, but can contain any nonnegative values. Pass in [] for alpha, censoring, or freq to use their default values.
[...] = evfit(data,alpha, censoring,freq,options) accepts a structure, options, that specifies control parameters for the iterative algorithm the function uses to compute maximum likelihood
estimates. You can create options using the function statset. Enter statset('evfit') to see the names and default values of the parameters that evfit accepts in the options structure. See the reference page for statset for more information about these options.
The type 1 extreme value distribution is also known as the Gumbel distribution. The version used here is suitable for modeling minima; the mirror image of this distribution can be used to model maxima by negating X. See "Extreme Value Distribution" on page B-19 for more details. If $x$ has a Weibull distribution, then $X=\log (x)$ has the type 1 extreme value distribution.

## See Also

mle | evlike | evpdf | evcdf | evinv | evstat | evrnd

## How To <br> - "Extreme Value Distribution" on page B-19

## Purpose <br> Extreme value inverse cumulative distribution function

Syntax

Description
$X=\operatorname{evinv(P,mu,sigma)~}$
[X,XLO,XUP] = evinv(P,mu,sigma,pcov,alpha)
$X=\operatorname{evinv}(P, m u$, sigma) returns the inverse cumulative distribution function (cdf) for a type 1 extreme value distribution with location parameter mu and scale parameter sigma, evaluated at the values in $P$. $P$, mu, and sigma can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array of the same size as the other inputs. The default values for mu and sigma are 0 and 1, respectively.
[X,XLO,XUP] = evinv(P,mu,sigma,pcov,alpha) produces confidence bounds for $X$ when the input parameters mu and sigma are estimates. pcov is the covariance matrix of the estimated parameters. alpha is a scalar that specifies $100(1-a l p h a) \%$ confidence bounds for the estimated parameters, and has a default value of 0.05 . XLO and XUP are arrays of the same size as $X$ containing the lower and upper confidence bounds.

The function evinv computes confidence bounds for P using a normal approximation to the distribution of the estimate

$$
\hat{\mu}+\hat{\sigma} q
$$

where $q$ is the Pth quantile from an extreme value distribution with parameters $\mu=0$ and $\sigma=1$. The computed bounds give approximately the desired confidence level when you estimate mu, sigma, and pcov from large samples, but in smaller samples other methods of computing the confidence bounds might be more accurate.

The type 1 extreme value distribution is also known as the Gumbel distribution. The version used here is suitable for modeling minima; the mirror image of this distribution can be used to model maxima by negating X. See "Extreme Value Distribution" on page B-19 for more details. If $x$ has a Weibull distribution, then $X=\log (x)$ has the type 1 extreme value distribution.

## evinv

See Also icdf | evcdf | evpdf | evstat | evfit | evlike | evrnd

## Purpose Test handle equality

Syntax
h1 == h2
tf = eq(h1, h2)

Description $\quad \mathrm{h} 1==\mathrm{h} 2$ performs element-wise comparisons between handle arrays h 1 and h2. h 1 and h2 must be of the same dimensions unless one is a scalar. The result is a logical array of the same dimensions, where each element is an element-wise equality result. If one of h1 or h2 is scalar, scalar expansion is performed and the result will match the dimensions of the array that is not scalar.
$\mathrm{tf}=\mathrm{eq}(\mathrm{h} 1, \mathrm{~h} 2)$ stores the result in a logical array of the same dimensions.

See Also qrandstream | ge | gt | le | lt | ne

## CompactTreeBagger.error

| Purpose | Error (misclassification probability or MSE) |
| :---: | :---: |
| Syntax | $\operatorname{err}=\operatorname{error}(\mathrm{B}, \mathrm{X}, \mathrm{Y})$ |
|  | err = error(B,X,Y,'param1', val1,'param2', val2,...) |

err $=\operatorname{error}(B, X, Y)$ computes the misclassification probability (for classification trees) or mean squared error (MSE, for regression trees) for each tree, for predictors $X$ given true response $Y$. For classification, $Y$ can be either a numeric vector, character matrix, cell array of strings, categorical vector or logical vector. For regression, Y must be a numeric vector. err is a vector with one error measure for each of the NTrees trees in the ensemble B.
err = error(B,X,Y,'param1',val1,'param2', val2,...) specifies optional parameter name/value pairs:
'mode ' String indicating how the method computes errors. If set to 'cumulative' (default), error computes cumulative errors and err is a vector of length NTrees, where the first element gives error from trees(1), second element gives error fromtrees(1:2) etc, up to trees(1:NTrees). If set to 'individual', err is a vector of length NTrees, where each element is an error from each tree in the ensemble. If set to 'ensemble', err is a scalar showing the cumulative error for the entire ensemble.
'trees ' Vector of indices indicating what trees to include in this calculation. By default, this argument is set to 'all' and the method uses all trees. If 'trees' is a numeric vector, the method returns a vector of length NTrees for 'cumulative' and 'individual' modes, where NTrees is the number of elements in the input vector, and a scalar for 'ensemble' mode. For example, in the 'cumulative ' mode, the first element gives error from trees (1), the second element gives error from trees(1:2) etc.

## CompactTreeBagger.error

> 'treeweights ' Vector of tree weights. This vector must have the same length as the 'trees ' vector. The method uses these weights to combine output from the specified trees by taking a weighted average instead of the simple non-weighted majority vote. You cannot use this argument in the 'individual' mode.

See Also TreeBagger.error

## TreeBagger.error



> 'treeweights ' Vector of tree weights. This vector must have the same length as the 'trees' vector. The method uses these weights to combine output from the specified trees by taking a weighted average instead of the simple non-weighted majority vote. You cannot use this argument in the 'individual' mode.

See Also CompactTreeBagger.error

Purpose Predicted responses
Syntax $\quad$ yfit $=\operatorname{eval}(t, x)$
yfit = eval(t, X,s)
[yfit, nodes] = eval(...)
[yfit,nodes,cnums] = eval(...)
[...] = t(X)
$[\ldots]=t(X, s)$

## Description

yfit $=$ eval( $\mathrm{t}, \mathrm{X}$ ) takes a classification or regression tree t and a matrix $X$ of predictors, and produces a vector yfit of predicted response values. For a regression tree, yfit(i) is the fitted response value for a point having the predictor values $\mathrm{X}(\mathrm{i},: \mathbf{)}$. For a classification tree, yfit(i) is the class into which the tree assigns the point with data X(i,:).
yfit $=$ eval( $\mathrm{t}, \mathrm{x}, \mathrm{s})$ takes an additional vector s of pruning levels, with 0 representing the full, unpruned tree. $t$ must include a pruning sequence as created by classregtree or by prune. If s has $k$ elements and X has $n$ rows, the output yfit is an $n$-by- $k$ matrix, with the j th column containing the fitted values produced by the $s(j)$ subtree. $s$ must be sorted in ascending order.

To compute fitted values for a tree that is not part of the optimal pruning sequence, first use prune to prune the tree.
[yfit, nodes] = eval(...) also returns a vector nodes the same size as yfit containing the node number assigned to each row of $X$. Use view to display the node numbers for any node you select.
[yfit,nodes,cnums] = eval(...) is valid only for classification trees. It returns a vector cnum containing the predicted class numbers.

NaN values in X are treated as missing. If eval encounters a missing value when it attempts to evaluate the split rule at a branch node, it cannot determine whether to proceed to the left or right child node. Instead, it sets the corresponding fitted value equal to the fitted value assigned to the branch node.

$$
[\ldots]=t(X) \text { or }[\ldots]=t(X, s) \text { also invoke eval. }
$$

Examples Create a classification tree for Fisher's iris data:

```
load fisheriris;
t = classregtree(meas,species,...
    'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
view(t)
```

|  |  |  |  |  |  |  |  | - |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Click to display: | Identity | $\square$ | Magnification: | 100\% | $\checkmark$ | Pruning level: | 0 of 4 | - |



Find assigned class names:

```
sfit = eval(t,meas);
```

Compute that proportion is correctly classified:

```
pct = mean(strcmp(sfit,species))
pct =
    0.9800
```

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.<br>See Also classregtree | prune | test | view

| Purpose | Extreme value negative log-likelihood |
| :---: | :---: |
| Syntax | nlogL = evlike(params,data) |
|  | [nlogL,AVAR] = evlike(params,data) |
|  | [...] = evlike(params,data, censoring) |
|  | [...] = evlike(params,data, censoring,freq) |

How To - "Extreme Value Distribution" on page B-19

## Description

nlogL = evlike(params, data) returns the negative of the log-likelihood for the type 1 extreme value distribution. params (1) is the tail location parameter, mu, and params (2) is the scale parameter, sigma. nlogL is a scalar.
[nlogL,AVAR] = evlike(params, data) returns the inverse of Fisher's information matrix, AVAR. If the input parameter values in params are the maximum likelihood estimates, the diagonal elements of AVAR are their asymptotic variances. AVAR is based on the observed Fisher's information, not the expected information.
[...] = evlike(params,data, censoring) accepts a Boolean vector of the same size as data, which is 1 for observations that are right-censored and 0 for observations that are observed exactly.
[...] = evlike(params,data, censoring,freq) accepts a frequency vector of the same size as data. freq typically contains integer frequencies for the corresponding elements in data, but can contain any nonnegative values. Pass in [] for censoring to use its default value.
The type 1 extreme value distribution is also known as the Gumbel distribution. The version used here is suitable for modeling minima; the mirror image of this distribution can be used to model maxima by negating data. See "Extreme Value Distribution" on page B-19 for more details. If $x$ has a Weibull distribution, then $X=\log (x)$ has the type 1 extreme value distribution.

## See Also evfit | evpdf \| evcdf | evinv \| evstat | evrnd

| Purpose | Extreme value probability density function |
| :---: | :---: |
| Syntax | $Y=\operatorname{evpdf}(\mathrm{X}, \mathrm{mu}$, sigma) |
| Description | $Y=\operatorname{evpdf}(X, m u$, sigma $)$ returns the pdf of the type 1 extreme value distribution with location parameter mu and scale parameter sigma, evaluated at the values in $X . X$, mu, and sigma can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array of the same size as the other inputs. The default values for mu and sigma are 0 and 1 , respectively. |
|  | The type 1 extreme value distribution is also known as the Gumbel distribution. The version used here is suitable for modeling minima; the mirror image of this distribution can be used to model maxima by negating X. See "Extreme Value Distribution" on page B-19 for more details. If $x$ has a Weibull distribution, then $X=\log (x)$ has the type 1 extreme value distribution. |
| See Also | pdf \| evcdf | evinv | evstat | evfit | evlike | evrnd |
| How To | . "Extreme Value Distribution" on page B-19 |

Purpose Extreme value random numbers

```
Syntax \(\quad R=\operatorname{evrnd}(m u\), sigma)
R = evrnd(mu,sigma,m,n,...)
\(R\) = evrnd(mu,sigma, [m,n,...])
```


## Description

## See Also

random | evpdf | evcdf | evinv | evstat | evfit \| evlike
How To

| Purpose | Extreme value mean and variance |
| :--- | :--- |
| Syntax | $[M, \mathrm{~V}]=$ evstat (mu, sigma) |
| Description | $[M, \mathrm{~V}]=$ evstat (mu, sigma) returns the mean of and variance for <br> the type 1 extreme value distribution with location parameter mu and <br> scale parameter sigma. mu and sigma can be vectors, matrices, or <br> multidimensional arrays that all have the same size. A scalar input is <br> expanded to a constant array of the same size as the other input. The <br> default values for mu and sigma are 0 and 1, respectively. |
| See Also | The type 1 extreme value distribution is also known as the Gumbel <br> distribution. The version used here is suitable for modeling minima; <br> the mirror image of this distribution can be used to model maxima. |
| See "Extreme Value Distribution" on page B-19 for more details. If |  |
| xhas a Weibull distribution, then $X=\log (x)$ has the type 1 extreme |  |
| value distribution. |  |

## Purpose Exponential cumulative distribution function

```
Syntax \(\quad P=\operatorname{expcdf}(X, m u)\)
[P,PLO,PUP] = expcdf(X,mu,pcov,alpha)
```


## Description

## Examples

$P=\operatorname{expcdf}(X, m u)$ computes the exponential cdf at each of the values in $X$ using the corresponding mean parameter mu. $X$ and mu can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other input. The parameters in mu must be positive.

The exponential cdf is

$$
p=F(x \mid u)=\int_{0}^{x} \frac{1}{\mu} e^{\frac{-t}{\mu}} d t=1-e^{\frac{-x}{\mu}}
$$

The result, $p$, is the probability that a single observation from an exponential distribution will fall in the interval [ $0 x$ ].
[P,PLO,PUP] = expcdf(X,mu,pcov,alpha) produces confidence bounds for $P$ when the input mean parameter mu is an estimate. pcov is the variance of the estimated mu. alpha specifies $100(1-a l p h a) \%$ confidence bounds. The default value of alpha is 0.05 . PLO and PUP are arrays of the same size as P containing the lower and upper confidence bounds. The bounds are based on a normal approximation for the distribution of the log of the estimate of mu. If you estimate mu from a set of data, you can get a more accurate set of bounds by applying expfit to the data to get a confidence interval for mu, and then evaluating expinv at the lower and upper endpoints of that interval.

The following code shows that the median of the exponential distribution is $* \log (2)$.

```
mu = 10:10:60;
p = expcdf(log(2)*mu,mu)
p =
    0.5000}00.5000 0.5000 0.5000 0.5000 0.5000
```

What is the probability that an exponential random variable is less than or equal to the mean, $\mu$ ?

```
mu = 1:6;
\(x=m u ;\)
\(p=\operatorname{expcdf}(x, m u)\)
p \(=\)
    \(\begin{array}{llllll}0.6321 & 0.6321 & 0.6321 & 0.6321 & 0.6321 & 0.6321\end{array}\)
```


## See Also

cdf | exppdf | expinv | expstat | expfit | explike | exprnd
How To

- "Exponential Distribution" on page B-16


## Purpose <br> Exponential parameter estimates

Syntax

```
muhat = expfit(data)
[muhat,muci] = expfit(data)
[muhat,muci] = expfit(data,alpha)
[...] = expfit(data,alpha,censoring)
[...] = expfit(data,alpha,censoring,freq)
```


## Description

## Examples

muhat $=$ expfit(data) estimates the mean of an exponentially distributed sample data. Each entry of muhat corresponds to the data in a column of data.
[muhat,muci] = expfit(data) returns $95 \%$ confidence intervals for the mean parameter estimates in matrix muci. The first row of muci contains the lower bounds of the confidence intervals, and the second row contains the upper bounds.
[muhat,muci] = expfit(data,alpha) returns 100(1-alpha)\% confidence intervals for the parameter estimates, where alpha is a value in the range [ 0 1] specifying the width of the confidence intervals. By default, alpha is 0.05 , which corresponds to $95 \%$ confidence intervals.
[...] = expfit(data, alpha, censoring) accepts a Boolean vector, censoring, of the same size as data, which is 1 for observations that are right-censored and 0 for observations that are observed exactly. data must be a vector in order to pass in the argument censoring.
[...] = expfit(data,alpha, censoring,freq) accepts a frequency vector, freq of the same size as data. Typically, freq contains integer frequencies for the corresponding elements in data, but can contain any nonnegative values. Pass in [] for alpha, censoring, or freq to use their default values.

The following estimates the mean mu of exponentially distributed data, and returns a $95 \%$ confidence interval for the estimate:

```
mu = 3;
data = exprnd(mu,100,1); % Simulated data
```

```
[muhat,muci] = expfit(data)
muhat =
    2.7511
muci =
    2.2826
    3.3813
```


## See Also

mle | explike | exppdf | expcdf | expinv | expstat | exprnd

## ExhaustiveSearcher

Superclasses NeighborSearcher
Purpose Nearest neighbors search using exhaustive search
Description

An ExhaustiveSearcher object performs $k \mathrm{NN}$ ( $k$-nearest neighbor) search using exhaustive search. Search objects store information about the data used, and the distance metric and parameters. The search performance for this object, compared with the KDTreeSearcher object, tends to be better for larger dimensions ( 10 or more) and worse for smaller dimensions. For more information on search objects, see "What Are Search Objects?" on page 13-17.

## Construction NS = ExhaustiveSearcher(X, 'Name', Value) constructs an

 ExhaustiveSearcher object based on $X$, where rows of $X$ correspond to observations and columns correspond to variables, using one or more optional name/value pairs. You can then use this object to find neighbors in $X$ nearest to the query points.NS = createns(X,'NSMethod','exhaustive','Name', Value) creates an ExhaustiveSearcher object based on $X$ using createns, where rows of $X$ correspond to observations and columns correspond to variables, using one or more optional name/value pairs. You can use this object to find neighbors in X nearest to the query points.

## Name-Value Pair Arguments

Both the ExhaustiveSearcher and the createns functions accept one or more of the following optional name/value pairs as input:

## Distance

A string or function handle specifying the default distance metric used when you call the knnsearch method.

- 'euclidean' — Euclidean distance (default).
- 'seuclidean' - Standardized Euclidean distance. Each coordinate difference between rows in $X$ and the query matrix is scaled by dividing by the corresponding element of the standard
deviation computed from X, S=nanstd (X). To specify another value for $S$, use the Scale argument.
- 'cityblock' - City block distance.
- 'chebychev' - Chebychev distance (maximum coordinate difference).
- 'minkowski' - Minkowski distance.
- 'mahalanobis' - Mahalanobis distance, computed using a positive definite covariance matrix $C$. The default value of $C$ is nancov ( X ). To change the value of C , use the Cov parameter.
- 'cosine' - One minus the cosine of the included angle between observations (treated as vectors).
- 'correlation' - One minus the sample linear correlation between observations (treated as sequences of values).
- 'spearman' - One minus the sample Spearman's rank correlation between observations (treated as sequences of values).
- 'hamming' - Hamming distance, percentage of coordinates that differ.
- 'jaccard' - One minus the Jaccard coefficient, the percentage of nonzero coordinates that differ.
- custom distance function - A distance function specified using @ (for example, @distfun). A distance function must be of the form function D2 = distfun(ZI, ZJ), taking as arguments a 1-by- $n$ vector ZI containing a single row from X or from the query points Y , an $m 2$-by- $n$ matrix ZJ containing multiple rows of $X$ or $Y$, and returning an $m 2$-by- 1 vector of distances D2, whose $j$ th element is the distance between the observations ZI and $\mathrm{ZJ}(j,:)$.

For more information on these distance metrics, see "Distance Metrics" on page 13-9.

## ExhaustiveSearcher

P
A positive scalar indicating the exponent of Minkowski distance. This parameter is only valid when Distance is 'minkowski'. Default is 2 .

Cov
A positive definite matrix indicating the covariance matrix when computing the Mahalanobis distance. This parameter is only valid when Distance is 'mahalanobis'. Default is nancov ( X ).

Scale
A vector $S$ with the length equal to the number of columns in X. Each coordinate of X and each query point is scaled by the corresponding element of $S$ when computing the standardized Euclidean distance. This parameter is only valid when Distance is 'seuclidean'. Default is nanstd ( X ).

## Properties <br> X

A matrix used to create the object

## Distance

A string specifying a built-in distance metric or a function handle that you provide when you create the object. This property is the default distance metric used when you call the knnsearch method to find nearest neighbors for future query points.

## DistParameter

Specifies the additional parameter for the chosen distance metric. The value is:

- If 'Distance' is 'minkowski': A positive scalar indicating the exponent of the Minkowski distance.
- If 'Distance' is 'mahalanobis': A positive definite matrix representing the covariance matrix used for computing the Mahalanobis distance.
- If 'Distance' is 'seuclidean': A vector representing the scale value of the data when computing the 'seuclidean' distance.
- Otherwise: Empty.

```
Methods
knnsearch
rangesearch
Find k-nearest neighbors using
ExhaustiveSearcher object
Find all neighbors within
specified distance using object
Examples Create an ExhaustiveSearcher object using the constructor:
```

```
load fisheriris
```

load fisheriris
x = meas(:,3:4);
x = meas(:,3:4);
NS = ExhaustiveSearcher(x,'distance','minkowski')
NS = ExhaustiveSearcher(x,'distance','minkowski')
NS =
NS =
ExhaustiveSearcher
ExhaustiveSearcher
Properties:
Properties:
X: [150x2 double]
X: [150x2 double]
Distance: 'minkowski'
Distance: 'minkowski'
DistParameter: 2

```
        DistParameter: 2
```

Create an ExhaustiveSearcher object using createns:

```
load fisheriris
x = meas(:,3:4);
NS = createns(x,'NsMethod','exhaustive',...
    'distance','minkowski')
```


## ExhaustiveSearcher

```
NS =
    ExhaustiveSearcher
    Properties:
                X: [150x2 double]
            Distance: 'minkowski'
        DistParameter: 2
```

For more in-depth examples using the knnsearch method, see the method reference page or see "Example: Classifying Query Data Using knnsearch" on page 13-18.

References<br>[1] Friedman, J. H., Bentely, J. and Finkel, R. A. (1977). An Algorithm for Finding Best Matches in Logarithmic Expected Time, ACM Transactions on Mathematical Software 3, 209.

See Also createns | KDTreeSearcher | NeighborSearcher
How To

- " $k$-Nearest Neighbor Search and Radius Search" on page 13-12
- "Distance Metrics" on page 13-9


## Purpose

Exponential inverse cumulative distribution function
Syntax
$\mathrm{x}=\operatorname{expinv}(\mathrm{P}, \mathrm{mu})$
[X,XLO,XUP] = expinv(X,mu,pcov,alpha)
$X=$ expinv $(P, m u)$ computes the inverse of the exponential cdf with parameters specified by mean parameter mu for the corresponding probabilities in P. P and mu can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other input. The parameters in mu must be positive and the values in $P$ must lie on the interval [01].
$[\mathrm{X}, \mathrm{XLO}, \mathrm{XUP}]=\operatorname{expinv}(\mathrm{X}, \mathrm{mu}, \mathrm{pcov}, \mathrm{alpha})$ produces confidence bounds for $X$ when the input mean parameter $m u$ is an estimate. $p c o v$ is the variance of the estimated mu. alpha specifies $100(1-$ alpha $) \%$ confidence bounds. The default value of alpha is 0.05 . XLO and XUP are arrays of the same size as $X$ containing the lower and upper confidence bounds. The bounds are based on a normal approximation for the distribution of the log of the estimate of mu. If you estimate mu from a set of data, you can get a more accurate set of bounds by applying expfit to the data to get a confidence interval for mu, and then evaluating expinv at the lower and upper end points of that interval.
The inverse of the exponential cdf is

$$
x=F^{-1}(p \mid \mu)=-\mu \ln (1-p)
$$

The result, $x$, is the value such that an observation from an exponential distribution with parameter $\mu$ will fall in the range $[0 x]$ with probability $p$.

## Examples

Let the lifetime of light bulbs be exponentially distributed with $\mu=700$ hours. What is the median lifetime of a bulb?

```
expinv(0.50,700)
ans =
```

Suppose you buy a box of " 700 hour" light bulbs. If 700 hours is the mean life of the bulbs, half of them will burn out in less than 500 hours.

## See Also <br> icdf | expcdf | exppdf | expstat | expfit | explike | exprnd

How To

- "Exponential Distribution" on page B-16

| Purpose | Exponential negative log-likelihood |
| :---: | :---: |
| Syntax | ```nlogL = explike(param,data) [nlogL,avar] = explike(param,data) [...] = explike(param,data,censoring) [...] = explike(param,data,censoring,freq)``` |
| Description | nlogL = explike(param,data) returns the negative of the log-likelihood for the exponential distribution. param is the mean parameter, mu. nlogL is a scalar. |
|  | [nlogL,avar] = explike(param,data) returns the inverse of Fisher's information, avar, a scalar. If the input parameter value in param is the maximum likelihood estimate, avar is its asymptotic variance. avar is based on the observed Fisher's information, not the expected information. |
|  | [...] = explike(param,data, censoring) accepts a Boolean vector, censoring, of the same size as data, which is 1 for observations that are right-censored and 0 for observations that are observed exactly. |
|  | [...] = explike(param,data, censoring,freq) accepts a frequency vector, freq, of the same size as data. The vector freq typically contains integer frequencies for the corresponding elements in data, but can contain any nonnegative values. Pass in [] for censoring to use its default value. |
| See Also | expcdf \| exppdf | expstat | expfit | expinv | exprnd |
| How To | - "Exponential Distribution" on page B-16 |

Purpose Write dataset array to file

```
Syntax export(DS,'file',filename)
export(DS)
export(DS,'file',filename,'Delimiter',delim)
export(DS,'XLSfile',filename)
export(DS,'XPTFile',filename)
export(DS,...,'WriteVarNames',false)
export(DS,...,'WriteObsNames',false)
```


## Description

export(DS,'file',filename) writes the dataset array DS to a tab-delimited text file, including variable names and observation names, if present. If the observation names exist, the name in the first column of the first line of the file is the first dimension name for the dataset (by default, 'Observations'). export overwrites any existing file named filename.
export(DS) writes to a text file whose default name is the name of the dataset array DS appended by '.txt'. If export cannot construct the file name from the dataset array input, it writes to the file 'dataset.txt'. export overwrites any existing file.
export(DS,'file',filename,'Delimiter', delim) writes the dataset array DS to a text file using the delimiter delim. delim must be one of the following:

- ' ' or 'space'
- ' $\backslash t$ ' or 'tab'
- ', ' or 'comma'
- ';' or 'semi'
- '|' or 'bar'
export(DS,'XLSfile',filename) writes the dataset array DS to a Microsoft ${ }^{\circledR}$ Excel ${ }^{\circledR}$ spreadsheet file, including variable names and observation names (if present). You can specify the 'Sheet' and
'Range' parameter name/value pairs, with parameter values as accepted by the xlsread function.
export(DS, 'XPTFile', filename) writes the dataset array DS to a SAS XPORT format file. When writing to an XPORT format file, variables must be scalar valued. export saves observation names to a variable called obsnames, unless the WriteObsNames parameter described below is false. The XPORT format restricts the length of variable names to eight characters; longer variable names are truncated.
export(DS,..., 'WriteVarNames', false) does not write the variable names to the text file. export(DS,..., 'WriteVarNames', true) is the default, writing the names as column headings in the first line of the file.
export(DS,...,'WriteObsNames',false) does not write the observation names to the text file. export(DS,...,'WriteObsNames', true) is the default, writing the names as the first column of the file.

In some cases, export creates a text file that does not represent A exactly, as described below. If you use dataset to read the file back into MATLAB, the new dataset array may not have exactly the same contents as the original dataset array. Save A as a MAT-file if you need to import it again as a dataset array.
export writes out numeric variables using long g format, and categorical or character variables as unquoted strings.

For non-character variables with more than one column, export writes out multiple delimiter-separated fields on each line, and constructs suitable column headings for the first line of the file.
export writes out variables that have more than two dimensions as a single empty field in each line of the file.
For cell-valued variables, export writes out the contents of each cell only when the cell contains a single row, and writes out a single empty field otherwise.

## dataset.export

In some cases, export creates a file that cannot be read back into MATLAB using dataset. Writing a dataset array that contains a cell-valued variable whose cell contents are not scalars results in a mismatch in the file between the number of fields on each line and the number of column headings on the first line. Writing a dataset array that contains a cell-valued variable whose cell contents are not all the same length results in a different number of fields on each line in the file. Therefore, if you might need to import a dataset array again, save it as a .mat file.

## Examples Move data between external text files and dataset arrays in the

 MATLAB workspace:```
A = dataset('file','sat2.dat','delimiter',',')
A =
    Test Gender Score
    'Verbal' 'Male' 470
    'Verbal' 'Female' 530
    'Quantitative' 'Male' 520
    'Quantitative' 'Female' 480
export(A(A.Score > 500,:),'file','HighScores.txt')
B = dataset('file','HighScores.txt','delimiter','\t')
B =
    Test Gender Score
    'Verbal' 'Female' 530
    'Quantitative' 'Male' 520
```

See Also dataset

## Purpose

Exponential probability density function

## Syntax <br> $Y=\operatorname{exppdf}(X, m u)$

Description

## Examples

See Also
How To
The exponential pdf is

$$
y=f(x \mid \mu)=\frac{1}{\mu} e^{\frac{-x}{\mu}}
$$

```
y = exppdf(5,1:5)
y =
        0.0067
y = exppdf(1:5,1:5)
y =
    0.3679
```

pdf | expcdf | expinv | expstat | expfit | explike | exprnd

- "Exponential Distribution" on page B-16
$Y=\operatorname{exppdf}(X, \mathrm{mu})$ returns the pdf of the exponential distribution with mean parameter mu, evaluated at the values in $X . X$ and mu can be vectors, matrices, or multidimensional arrays that have the same size. A scalar input is expanded to a constant array with the same dimensions as the other input. The parameters in mu must be positive.

The exponential pdf is the gamma pdf with its first parameter equal to 1 .
The exponential distribution is appropriate for modeling waiting times when the probability of waiting an additional period of time is independent of how long you have already waited. For example, the probability that a light bulb will burn out in its next minute of use is relatively independent of how many minutes it has already burned.

Purpose Exponential random numbers

```
Syntax \(\quad R=\operatorname{exprnd}(m u)\)
\(R=\operatorname{exprnd}(m u, m, n, \ldots)\)
\(R=\operatorname{exprnd}(m u,[m, n, \ldots])\)
```


## Description

## Examples

```
n1 = exprnd(5:10)
n1 =
    7.5943 18.3400 2.7113 3.0936 0.6078
n2 = exprnd(5:10,[1 6])
n2 =
    3.2752 1.1110 23.5530 23.4303 5.7190 3.9876
```

n3 $=\operatorname{exprnd}(5,2,3)$
n3 =
$24.3339 \quad 13.5271 \quad 1.8788$
$4.7932 \quad 4.3675 \quad 2.6468$
See Also random | expcdf | exppdf | expstat | expfit | explike | expinv
How To . "Exponential Distribution" on page B-16
Purpose Exponential mean and variance
Syntax [m,v] = expstat(mu)
Description $[m, v]=$ expstat $(\mathrm{mu})$ returns the mean of and variance for theexponential distribution with parameters mu. mu can be a vectors,matrix, or multidimensional array. The mean of the exponentialdistribution is $\mu$, and the variance is $\mu^{2}$.
Examples
[m,v] = expstat([10 10 100 1000])

m =

            \(\begin{array}{llll}1 & 10 & 100 & 1000\end{array}\)
    v =

    1100100001000000
    See Also expinv | expcdf | exppdf | expstat | expfit | explike | exprnd
How To - "Exponential Distribution" on page B-16

Purpose Factor analysis

```
Syntax lambda = factoran(X,m)
[lambda,psi] = factoran(X,m)
[lambda,psi,T] = factoran(X,m)
[lambda,psi,T,stats] = factoran(X,m)
[lambda,psi,T,stats,F] = factoran(X,m)
[...] = factoran(...,param1,val1,param2,val2,...)
```


## Definitions

## Description

factoran computes the maximum likelihood estimate (MLE) of the factor loadings matrix $\Lambda$ in the factor analysis model

$$
x=\mu+\Lambda f+e
$$

where $x$ is a vector of observed variables, $\mu$ is a constant vector of means, $\Lambda$ is a constant d-by-m matrix of factor loadings, $f$ is a vector of independent, standardized common factors, and $e$ is a vector of independent specific factors. $x, \mu$, and $e$ are of length $\mathrm{d} . f$ is of length m .
Alternatively, the factor analysis model can be specified as

$$
\operatorname{cov}(x)=\Lambda \Lambda^{T}+\Psi
$$

where $\Psi=\operatorname{cov}(e)$ is a d-by-d diagonal matrix of specific variances.
lambda $=$ factoran $(X, m)$ returns the maximum likelihood estimate, lambda, of the factor loadings matrix, in a common factor analysis model with $m$ common factors. $X$ is an $n$-by- $d$ matrix where each row is an observation of $d$ variables. The ( $i, j$ ) th element of the d-by-m matrix lambda is the coefficient, or loading, of the $j$ th factor for the ith variable. By default, factoran calls the function rotatefactors to rotate the estimated factor loadings using the 'varimax' option.
[lambda, psi] = factoran(X,m) also returns maximum likelihood estimates of the specific variances as a column vector $p s i$ of length $d$.
[lambda, psi,T] = factoran(X,m) also returns the m-by-m factor loadings rotation matrix $T$.
[lambda, psi, T,stats] = factoran(X,m) also returns a structure stats containing information relating to the null hypothesis, $\mathrm{H}_{0}$, that the number of common factors is m . stats includes the following fields:

| Field | Description |
| :--- | :--- |
| loglike | Maximized log-likelihood value |
| $d f e$ | Error degrees of freedom $=\left((d-m)^{\wedge} 2-(d+m)\right) / 2$ |
| chisq | Approximate chi-squared statistic for the null hypothesis |
| $p$ | Right-tail significance level for the null hypothesis |

factoran does not compute the chisq and p fields unless dfe is positive and all the specific variance estimates in psi are positive (see "Heywood Case" on page 20-537 below). If $X$ is a covariance matrix, then you must also specify the 'nobs ' parameter if you want factoran to compute the chisq and p fields.
[lambda, psi, T, stats,F] = factoran(X,m) also returns, in F, predictions of the common factors, known as factor scores. $F$ is an $n$-by-m matrix where each row is a prediction of $m$ common factors. If $X$ is a covariance matrix, factoran cannot compute $F$. factoran rotates $F$ using the same criterion as for lambda.
[...] = factoran(..., param1,val1,param2,val2,...) enables you to specify optional parameter name/value pairs to control the model fit and the outputs. The following are the valid parameter/value pairs.

| Parameter | Value |  |
| :--- | :--- | :--- |
| 'xtype' | Type of input in the matrix X. 'xtype ' can be one <br> of: |  |
|  | 'data' | Raw data (default) |
|  | 'covariance | ' Positive definite covariance or <br> correlation matrix |


| Parameter | Value |  |
| :---: | :---: | :---: |
| 'scores' | Method for predicting factor scores. 'scores' is ignored if X is not raw data. |  |
|  | 'wls' <br> 'Bartlett' | Synonyms for a weighted least-squares estimate that treats F as fixed (default) |
|  | 'regression <br> 'Thomson' | Synonyms for a minimum mean squared error prediction that is equivalent to a ridge regression |
| 'start' | Starting point for the specific variances psi in the maximum likelihood optimization. Can be specified as: |  |
|  | 'random' | Chooses d uniformly distributed values on the interval $[0,1]$. |
|  | 'Rsquared' | Chooses the starting vector as a scale factor times diag (inv (corrcoef(X))) (default). For examples, see Jöreskog [2]. |
|  | Positive integer | Performs the given number of maximum likelihood fits, each initialized as with 'random'. factoran returns the fit with the highest likelihood. |
|  | Matrix | Performs one maximum likelihood fit for each column of the specified matrix. The ith optimization is initialized with the values from the ith column. The matrix must have d rows. |


| Parameter | Value |  |
| :---: | :---: | :---: |
| 'rotate' | Method used to rotate factor loadings and scores. 'rotate' can have the same values as the 'Method' parameter of rotatefactors. See the reference page for rotatefactors for a full description of the available methods. |  |
|  | 'none ' | Performs no rotation. |
|  | 'equamax' | Special case of the orthomax rotation. Use the 'normalize', 'reltol', and 'maxit' parameters to control the details of the rotation. |
|  | 'orthomax' | Orthogonal rotation that maximizes a criterion based on the variance of the loadings. <br> Use the 'coeff', 'normalize', reltol', and 'maxit' parameters to control the details of the rotation. |
|  | 'parsimax' | Special case of the orthomax rotation (default). Use the 'normalize', 'reltol', and 'maxit' parameters to control the details of the rotation. |
|  | 'pattern' | Performs either an oblique rotation (the default) or an orthogonal rotation to best match a specified pattern matrix. Use the 'type' parameter to choose the type of rotation. Use the 'target' parameter to specify the pattern matrix. |


| Parameter | Value |  |
| :---: | :---: | :---: |
|  | 'procrustes | 'Performs either an oblique (the default) or an orthogonal rotation to best match a specified target matrix in the least squares sense. <br> Use the 'type ' parameter to choose the type of rotation. Use 'target' to specify the target matrix. |
|  | promax ${ }^{\prime}$ | Performs an oblique procrustes rotation to a target matrix determined by factoran as a function of an orthomax solution. <br> Use the 'power' parameter to specify the exponent for creating the target matrix. Because 'promax' uses 'orthomax' internally, you can also specify the parameters that apply to 'orthomax'. |
|  | 'quartimax' | Special case of the orthomax rotation (default). Use the 'normalize', 'reltol', and 'maxit' parameters to control the details of the rotation. |
|  | 'varimax' | Special case of the orthomax rotation (default). Use the 'normalize', 'reltol', and 'maxit' parameters to control the details of the rotation. |


| Parameter | Value |
| :---: | :---: |
|  | Function <br> Function handle to rotation function of the form <br> $[\mathrm{B}, \mathrm{T}]=$ <br> myrotation(A, ...) <br> where $A$ is a d-by-m matrix of unrotated factor loadings, $B$ is a $d$-by-m matrix of rotated loadings, and $T$ is the corresponding m -by-m rotation matrix. <br> Use the factoran parameter 'userargs' to pass additional arguments to this rotation function. See "Example 4" on page 20-542. |
| 'coeff' | Coefficient, often denoted as $\gamma$, defining the specific 'orthomax' criterion. Must be from 0 to 1. The value 0 corresponds to quartimax, and 1 corresponds to varimax. Default is 1 . |
| 'normalize' | Flag indicating whether the loading matrix should be row-normalized (1) or left unnormalized (0) for orthomax' or 'varimax' rotation. Default is 1 . |
| 'reltol' | Relative convergence tolerance for 'orthomax' or 'varimax' rotation. Default is sqrt(eps). |
| 'maxit' | Iteration limit for 'orthomax' or 'varimax' rotation. Default is 250 . |
| 'target' | Target factor loading matrix for 'procrustes' rotation. Required for 'procrustes ' rotation. No default value. |
| 'type ' | Type of 'procrustes ' rotation. Can be 'oblique' (default) or 'orthogonal'. |

## factoran

| Parameter | Value |
| :--- | :--- |
| 'power' | Exponent for creating the target matrix in the <br> 'promax ' rotation. Must be $\geq$ 1. Default is 4. |
| 'userargs ' | Denotes the beginning of additional input values <br> for a user-defined rotation function. factoran <br> appends all subsequent values, in order and <br> without processing, to the rotation function <br> argument list, following the unrotated factor <br> loadings matrix A. See "Example 4" on page 20-542. |
| 'nobs' | If X is a covariance or correlation matrix, indicates <br> the number of observations that were used in its <br> estimation. This allows calculation of significance <br> for the null hypothesis even when the original data <br> are not available. There is no default. 'nobs ' is <br> ignored if X is raw data. |
| 'delta' | Lower bound for the specific variances psi during <br> the maximum likelihood optimization. Default is <br> 0.005. |
| 'optimopts' | Structure that specifies control parameters for <br> the iterative algorithm the function uses to <br> compute maximum likelihood estimates. Create <br> this structure with the function statset. Enter <br> statset ('factoran' ) to see the names and <br> default values of the parameters that factoran <br> accepts in the options structure. See the reference <br> page for statset for more information about these <br> options. |

## Observed Data Variables

The variables in the observed data matrix $X$ must be linearly independent, i.e., $\operatorname{cov}(\mathrm{X})$ must have full rank, for maximum likelihood estimation to succeed. factoran reduces both raw data and a covariance matrix to a correlation matrix before performing the fit.
factoran standardizes the observed data $X$ to zero mean and unit variance before estimating the loadings lambda. This does not affect the model fit, because MLEs in this model are invariant to scale. However, lambda and psi are returned in terms of the standardized variables, i.e., lambda*lambda'+diag(psi) is an estimate of the correlation matrix of the original data $X$ (although not after an oblique rotation). See "Example 1" on page 20-537 and "Example 3" on page 20-539.

## Heywood Case

If elements of psi are equal to the value of the 'delta' parameter (i.e., they are essentially zero), the fit is known as a Heywood case, and interpretation of the resulting estimates is problematic. In particular, there can be multiple local maxima of the likelihood, each with different estimates of the loadings and the specific variances. Heywood cases can indicate overfitting (i.e., $m$ is too large), but can also be the result of underfitting.

## Rotation of Factor Loadings and Scores

Unless you explicitly specify no rotation using the 'rotate ' parameter, factoran rotates the estimated factor loadings, lambda, and the factor scores, F. The output matrix T is used to rotate the loadings, i.e., lambda $=$ lambda0*T, where lambda0 is the initial (unrotated) MLE of the loadings. T is an orthogonal matrix for orthogonal rotations, and the identity matrix for no rotation. The inverse of $T$ is known as the primary axis rotation matrix, while $T$ itself is related to the reference axis rotation matrix. For orthogonal rotations, the two are identical.
factoran computes factor scores that have been rotated by inv ( $\mathrm{T}^{\prime}$ ), i.e., $F=F 0$ * $\operatorname{inv}\left(T^{\prime}\right)$, where $F 0$ contains the unrotated predictions. The estimated covariance of $F$ is $\operatorname{inv}\left(T^{\prime *} T\right)$, which, for orthogonal or no rotation, is the identity matrix. Rotation of factor loadings and scores is an attempt to create a more easily interpretable structure in the loadings matrix after maximum likelihood estimation.

## Examples Example 1

Load the carbig data, and fit the default model with two factors.

## factoran

```
load carbig
X = [Acceleration Displacement Horsepower MPG Weight];
X = X(all(~isnan(X),2),:);
[Lambda,Psi,T,stats,F] = factoran(X,2,...
                            'scores','regression');
inv(T'*T) % Estimated correlation matrix of F, == eye(2)
Lambda*Lambda'+diag(Psi) % Estimated correlation matrix
Lambda*inv(T) % Unrotate the loadings
F*T' % Unrotate the factor scores
biplot(Lambda,... % Create biplot of two factors
    'LineWidth',2,...
    'MarkerSize', 20)
```



## Example 2

Although the estimates are the same, the use of a covariance matrix rather than raw data doesn't let you request scores or significance level:

```
[Lambda,Psi,T] = factoran(cov(X),2,'xtype','cov')
[Lambda,Psi,T] = factoran(corrcoef(X),2,'xtype','cov')
```


## Example 3

Use promax rotation:

```
[Lambda,Psi,T,stats,F] = factoran(X,2,'rotate','promax',...
    'powerpm',4);
```


## factoran

| $\operatorname{inv}\left(T^{\prime *} T\right)$ | \% Est'd corr of $F$, |
| :--- | :--- |
|  | \% no longer eye (2) |
| Lambda*inv ( $\left.T^{\prime *} T\right) *$ Lambda'+diag(Psi) | \% Est'd corr of $X$ |

Plot the unrotated variables with oblique axes superimposed.

```
invT = inv(T)
LambdaO = Lambda*invT
line([-invT(1,1) invT(1,1) NaN -invT(2,1) invT(2,1)], ...
    [-invT(1,2) invT(1,2) NaN -invT(2,2) invT(2,2)], ...
    'Color','r','linewidth',2)
hold on
biplot(Lambda0,...
    'LineWidth',2,...
    'MarkerSize',20)
xlabel('Loadings for unrotated Factor 1')
ylabel('Loadings for unrotated Factor 2')
```



Plot the rotated variables against the oblique axes:
biplot(Lambda, 'LineWidth', 2, 'MarkerSize', 20)

## factoran



## Example 4

Syntax for passing additional arguments to a user-defined rotation function:

```
[Lambda,Psi,T] = ...
```

factoran(X,2,'rotate', @myrotation,'userargs',1,'two');
References [1] Harman, H. H. Modern Factor Analysis. 3rd Ed. Chicago: University of Chicago Press, 1976.
[2] Jöreskog, K. G. "Some Contributions to Maximum Likelihood Factor Analysis." Psychometrika. Vol. 32, Issue 4, 1967, pp. 443-482.
[3] Lawley, D. N., and A. E. Maxwell. Factor Analysis as a Statistical Method. 2nd Ed. New York: American Elsevier Publishing Co., 1971.

## See Also

biplot | princomp | procrustes | pcacov | rotatefactors | statset

## TreeBagger.FBoot property

## Purpose Fraction of in-bag observations

Description The FBoot property is the fraction of observations to be randomly selected with replacement for each bootstrap replica. The size of each replica is given by $n$ *FBoot, where $n$ is the number of observations in the training set. The default value is 1 .

## Purpose $\quad F$ cumulative distribution function

$$
\text { Syntax } \quad P=\operatorname{fcdf}(X, V 1, V 2)
$$

Description $\quad P=\operatorname{fcdf}(X, V 1, V 2)$ computes the $F$ cdf at each of the values in $X$ using the corresponding numerator degrees of freedom V1 and denominator degrees of freedom V2. X, V1, and V2 can be vectors, matrices, or multidimensional arrays that are all the same size. A scalar input is expanded to a constant matrix with the same dimensions as the other inputs. V1 and V2 parameters must contain real positive values.

The $F$ cdf is

$$
p=F\left(x \mid v_{1}, v_{2}\right)=\int_{0}^{x} \frac{\Gamma\left[\frac{\left(v_{1}+v_{2}\right)}{2}\right]}{\Gamma\left(\frac{v_{1}}{2}\right) \Gamma\left(\frac{v_{2}}{2}\right)}\left(\frac{v_{1}}{v_{2}}\right)^{\frac{v_{1}}{2}} \frac{t^{\frac{v_{1}-2}{2}}}{\left[1+\left(\frac{v_{1}}{v_{2}}\right) t\right]^{\frac{v_{1}+v_{2}}{2}}} d t
$$

The result, $p$, is the probability that a single observation from an $F$ distribution with parameters $v_{1}$ and $v_{2}$ will fall in the interval [ $\left.0 x\right]$.

## Examples

The following illustrates a useful mathematical identity for the $F$ distribution:

```
nu1 = 1:5;
nu2 = 6:10;
x = 2:6;
F1 = fcdf(x,nu1,nu2)
F1 =
    0.7930
F2 = 1 - fcdf(1./x,nu2,nu1)
F2 =
    0.7930
```

See Also cdf | fpdf | finv | fstat | frnd

$$
\text { How To } \quad \text { "F Distribution" on page B-25 }
$$

## Purpose Two-level full factorial design

## Syntax <br> dFF2 = ff2n(n)

Description dFF2 $=f f 2 n(n)$ gives factor settings dFF2 for a two-level full factorial design with n factors. dFF2 is $m$-by- n , where $m$ is the number of treatments in the full-factorial design. Each row of dFF2 corresponds to a single treatment. Each column contains the settings for a single factor, with values of 0 and 1 for the two levels.

## Examples <br> dFF2 $=\mathrm{ff} 2 \mathrm{n}(3)$ <br> dFF2 = <br> $0 \quad 0 \quad 0$ <br> $0 \quad 0 \quad 1$ <br> $0 \quad 1 \quad 0$ <br> $0 \quad 1 \quad 1$ <br> 100 <br> 101 <br> $1 \quad 10$ <br> $1 \quad 1 \quad 1$

See Also fullfact

## TreeBagger.fillProximities

Purpose Proximity matrix for training data

```
Syntax
B = fillProximities(B)
B = fillProximities(B,'param1',val1,'param2',val2,...
```

Description
$B=$ fillProximities $(B)$ computes a proximity matrix for the training data and stores it in the Properties field of B.

B = fillProximities(B,'param1',val1,'param2', val2,...) specifies optional parameter name/value pairs:

| 'trees' | Either 'all' or a vector of indices of the trees in <br> the ensemble to be used in computing the proximity <br> matrix. Default is 'all'. |
| :--- | :--- |
| 'nprint' | Number of training cycles (grown trees) after which <br> TreeBagger displays a diagnostic message showing <br> training progress. Default is no diagnostic messages. |
|  |  |

See Also CompactTreeBagger.outlierMeasure | CompactTreeBagger.proximity

## Purpose Find objects matching specified conditions

Syntax $\quad h m=$ findobj(h, 'conditions')

Description The findobj method of the handle class follows the same syntax as the MATLAB findobj command, except that the first argument must be an array of handles to objects.
hm = findobj(h, 'conditions') searches the handle object array $h$ and returns an array of handle objects matching the specified conditions. Only the public members of the objects of $h$ are considered when evaluating the conditions.

See Also<br>findobj | qrandstream

## qrandstream.findprop

Purpose Find property of MATLAB handle object

## Syntax $\quad p=$ findprop(h,'propname')

Description $\quad p=$ findprop( $h$, 'propname') finds and returns the meta.property object associated with property name propname of scalar handle object h. propname must be a string. It can be the name of a property defined by the class of $h$ or a dynamic property added to scalar object $h$.

If no property named propname exists for object $h$, an empty meta.property array is returned.

See Also dynamicprops | findobj | meta.property | qrandstream

## Purpose $\quad F$ inverse cumulative distribution function

$$
\text { Syntax } \quad x=\operatorname{finv}(P, V 1, V 2)
$$

Description $\quad \mathrm{X}=\mathrm{finv}(\mathrm{P}, \mathrm{V} 1, \mathrm{~V} 2)$ computes the inverse of the $F$ cdf with numerator degrees of freedom V1 and denominator degrees of freedom V2 for the corresponding probabilities in P. P, V1, and V2 can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs.

V1 and V2 parameters must contain real positive values, and the values in $P$ must lie on the interval [01].

The $F$ inverse function is defined in terms of the $F$ cdf as

$$
x=F^{-1}\left(p \mid v_{1}, v_{2}\right)=\left\{x: F\left(x \mid v_{1}, v_{2}\right)=p\right\}
$$

where

$$
p=F\left(x \mid v_{1}, v_{2}\right)=\int_{0}^{x} \frac{\Gamma\left[\frac{\left(v_{1}+v_{2}\right)}{2}\right]}{\Gamma\left(\frac{v_{1}}{2}\right) \Gamma\left(\frac{v_{2}}{2}\right)}\left(\frac{v_{1}}{v_{2}}\right)^{\frac{v_{1}}{2}} \frac{t^{\frac{v_{1}-2}{2}}}{\left[1+\left(\frac{v_{1}}{v_{2}}\right) t^{\frac{v_{1}+v_{2}}{2}}\right.} d t
$$

Examples Find a value that should exceed $95 \%$ of the samples from an $F$ distribution with 5 degrees of freedom in the numerator and 10 degrees of freedom in the denominator.

$$
\begin{aligned}
& x=\operatorname{finv}(0.95,5,10) \\
& x= \\
& 3.3258
\end{aligned}
$$

You would observe values greater than 3.3258 only $5 \%$ of the time by chance.

See Also icdf | fcdf | fpdf | fstat | frnd
How To . "F Distribution" on page B-25

## ClassificationDiscriminant.fit

## Purpose

Fit discriminant analysis classifier
obj = ClassificationDiscriminant.fit(X,Y)
obj = ClassificationDiscriminant.fit(X,Y,Name, Value)

## Input

Arguments
obj $=$ ClassificationDiscriminant.fit( $\mathrm{X}, \mathrm{Y}$ ) returns a discriminant analysis classifier based on the input variables (also known as predictors, features, or attributes) $X$ and output (response) $Y$.
obj = ClassificationDiscriminant.fit(X,Y,Name, Value) fits a classifier with additional options specified by one or more Name, Value pair arguments. If you use one of the following five options, obj is of class ClassificationPartitionedModel: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Otherwise, obj is of class ClassificationDiscriminant.

X
Matrix of numeric predictor values. Each column of X represents one variable, and each row represents one observation.

ClassificationDiscriminant.fit considers NaN values in $X$ as missing values. ClassificationDiscriminant.fit does not use observations with missing values for $X$ in the fit.

## Y

Numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. Each row of $Y$ represents the classification of the corresponding row of $X$.

ClassificationDiscriminant.fit considers NaN values in $Y$ to be missing values. ClassificationDiscriminant.fit does not use observations with missing values for $Y$ in the fit.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must

## ClassificationDiscriminant.fit

appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

ClassNames
Array of class names. Use the data type that exists in $Y$.
Use ClassNames to order the classes or to select a subset of classes for training.

Default: Class names that exist in $Y$

## Cost

Square matrix, where $\operatorname{Cost}(i, j)$ is the cost of classifying a point into class $j$ if its true class is $i$. Alternatively, Cost can be a structure S having two fields: S . ClassNames containing the group names as a variable of the same type as Y , and S.ClassificationCosts containing the cost matrix.

Default: $\operatorname{Cost}(i, j)=1$ if $i \sim=j$, and $\operatorname{Cost}(i, j)=0$ if $i=j$
crossval
If 'on', creates a cross-validated classifier with 10 folds. You can use 'kfold', 'holdout', 'leaveout', or 'cvpartition' parameters to override this cross-validation setting. You can only use one of these options at a time for creating a cross-validated classifier: 'cvpartition', 'holdout', 'kfold', or 'leaveout'.

Alternatively, cross validate obj later using the crossval method.
Default: 'off'
cvpartition
Partition created with cvpartition to use in a cross-validated classifier. You can only use one of these options at a time for creating a cross-validated classifier: 'cvpartition', 'holdout', 'kfold', or 'leaveout'.

## ClassificationDiscriminant.fit

discrimType
String specifying the discriminant type. Case-insensitive. One of:

- 'linear'
- 'quadratic'
- 'diagLinear'
- 'diagQuadratic'
- 'pseudoLinear'
- 'pseudoQuadratic'

Default: 'linear'

## fillCoeffs

String, either 'on' or 'off', specifying whether to populate the Coeffs property in the classifier object. Setting to 'on' can be computationally intensive, especially when cross validating.

Default: 'on', except 'off' when cross validating

## holdout

Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1 . You can only use one of these options at a time for creating a cross-validated classifier: 'cvpartition', 'holdout', 'kfold', or 'leaveout'.
kfold
Number of folds to use in a cross-validated classifier, a positive integer. You can only use one of these options at a time for creating a cross-validated classifier: 'cvpartition', 'holdout', 'kfold', or 'leaveout'.

Default: 10

## ClassificationDiscriminant.fit

## leaveout

Use leave-one-out cross validation by setting to 'on'. You can only use one of these options at a time for creating a cross-validated classifier: 'cvpartition', 'holdout', 'kfold', or 'leaveout'.

## PredictorNames

Cell array of names for the predictor variables, in the order in which they appear in $X$.

Default: \{'x1', 'x2', ...\}
prior
Prior probabilities for each class. Specify as one of:

- A string:
- 'empirical' determines class probabilities from class frequencies in Y. If you pass observation weights, they are used to compute the class probabilities.
- 'uniform' sets all class probabilities equal.
- A vector (one scalar value for each class)
- A structure S with two fields:
- S.ClassNames containing the class names as a variable of the same type as Y
- S.ClassProbs containing a vector of corresponding probabilities

Default: 'empirical'

## ResponseName

Name of the response variable Y , a string.
Default: ' $Y^{\prime}$

## ClassificationDiscriminant.fit

## ScoreTransform

Function handle for transforming scores, or string representing a built-in transformation function.

| String | Formula |
| :--- | :--- |
| 'symmetric' | $2 x-1$ |
| 'invlogit' | $\log (x /(1-x))$ |
| 'ismax' | Set score for the class with the <br> largest score to 1, and scores for all <br> other classes to 0. |
| 'symmetricismax' | Set score for the class with the <br> largest score to 1, and scores for <br> all other classes to -1. |
| 'none' | $x$ |
| 'logit' | $1 /\left(1+e^{-x}\right)$ |
| 'doublelogit' | $1 /\left(1+e^{-2 x}\right)$ |
| 'symmetriclogit' | $2 /\left(1+e^{-x}\right)-1$ |
| 'sign' | -1 for $x<0$ <br> 0 for $x=0$ <br> 1 for $x>0$ |

You can include your own function handle for transforming scores. Your function should accept a matrix (the original scores) and return a matrix of the same size (the transformed scores).

Default: 'none'
weights
Vector of observation weights. The length of weights is the number of rows in X. ClassificationDiscriminant.fit normalizes the weights to sum to 1 .

## ClassificationDiscriminant.fit

Default: ones(size(X,1),1)

## Output

Arguments

## Definitions

obj
Discriminant analysis classifier. You can use obj to predict the response of new data using the predict method.

## Discriminant Classification

The model for discriminant analysis is:

- Each class (Y) generates data (X) using a multivariate normal distribution. That is, the model assumes X has a Gaussian mixture distribution (gmdistribution).
- For linear discriminant analysis, the model has the same covariance matrix for each class, only the means vary.
- For quadratic discriminant analysis, both means and covariances of each class vary.
predict classifies so as to minimize the expected classification cost:

$$
\hat{y}=\underset{y=1, \ldots, K}{\arg \min } \sum_{k=1}^{K} \hat{P}(k \mid x) C(y \mid k),
$$

where

- $\hat{y}$ is the predicted classification.
- $K$ is the number of classes.
- $\hat{P}(k \mid x)$ is the posterior probability of class $k$ for observation $x$.
- $C(y \mid k)$ is the cost of classifying an observation as $y$ when its true class is $k$.

For details, see "How the predict Method Classifies" on page 12-6.

## ClassificationDiscriminant.fit

```
Examples Construct a discriminant analysis classifier for the Fisher iris data:
```

```
load fisheriris
```

load fisheriris
obj = ClassificationDiscriminant.fit(meas,species)
obj = ClassificationDiscriminant.fit(meas,species)
obj =
obj =
ClassificationDiscriminant:
ClassificationDiscriminant:
PredictorNames: {'x1' 'x2' 'x3' 'x4'}
PredictorNames: {'x1' 'x2' 'x3' 'x4'}
ResponseName: 'Y'
ResponseName: 'Y'
ClassNames: {'setosa' 'versicolor' 'virginica'}
ClassNames: {'setosa' 'versicolor' 'virginica'}
ScoreTransform: 'none'
ScoreTransform: 'none'
NObservations: 150
NObservations: 150
DiscrimType: 'linear'
DiscrimType: 'linear'
Mu: [3\times4 double]
Mu: [3\times4 double]
Coeffs: [3x3 struct]

```
                Coeffs: [3x3 struct]
```

Alternatives The classify function also performs discriminant analysis. classify is usually more awkward to use:

- classify requires you to fit the classifier every time you make a new prediction.
- classify does not perform cross validation.
- classify requires you to fit the classifier when changing prior probabilities.


## See Also <br> ClassificationDiscriminant | ClassificationTree.fit <br> How To <br> - "Discriminant Analysis" on page 12-3

## ClassificationTree.fit

Purpose Fit classification tree
Syntax $\quad$ tree $=$ ClassificationTree.fit $(X, Y)$
tree $=$ ClassificationTree.fit $(X, Y$, Name, Value $)$
Description

Input x
Arguments
tree $=$ ClassificationTree.fit $(X, Y)$ returns a classification tree based on the input variables (also known as predictors, features, or attributes) X and output (response) Y . tree is a binary tree, where each branching node is split based on the values of a column of $X$.
tree $=$ ClassificationTree.fit(X,Y,Name, Value) fits a tree with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN. If you use one of the following five options, tree is of class ClassificationPartitionedModel: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Otherwise, tree is of class ClassificationTree.

A matrix of numeric predictor values. Each column of X represents one variable, and each row represents one observation.
ClassificationTree.fit considers NaN values in X as missing values. ClassificationTree.fit does not use observations with all missing values for X in the fit. ClassificationTree.fit uses observations with some missing values for X to find splits on variables for which these observations have valid values.

## Y

A numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings. Each row of $Y$ represents the classification of the corresponding row of $X$. For numeric $Y$, consider using RegressionTree.fit instead of ClassificationTree.fit.

## ClassificationTree.fit

ClassificationTree.fit considers NaN values in $Y$ to be missing values. ClassificationTree.fit does not use observations with missing values for $Y$ in the fit.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

## CategoricalPredictors

List of categorical predictors. Pass CategoricalPredictors as one of:

- A numeric vector with indices from 1 to $p$, where $p$ is the number of columns of $X$.
- A logical vector of length $p$, where a true entry means that the corresponding column of X is a categorical variable.
- 'all', meaning all predictors are categorical.
- A cell array of strings, where each element in the array is the name of a predictor variable. The names must match entries in PredictorNames values.
- A character matrix, where each row of the matrix is a name of a predictor variable. The names must match entries in PredictorNames values. Pad the names with extra blanks so each row of the character matrix has the same length.

Default: []

## ClassNames

Array of class names. Use the data type that exists in Y.
Use ClassNames to order the classes or to select a subset of classes for training.

## ClassificationTree.fit

Default: The class names that exist in $Y$

## Cost

Square matrix, where $\operatorname{Cost}(i, j)$ is the cost of classifying a point into class $j$ if its true class is i. Alternatively, Cost can be a structure S having two fields: S . ClassNames containing the group names as a variable of the same type as Y , and S.ClassificationCosts containing the cost matrix.

Default: $\operatorname{Cost}(i, j)=1$ if $i \sim=j$, and $\operatorname{Cost}(i, j)=0$ if $i=j$

## crossval

If ' on ', grows a cross-validated decision tree with 10 folds. You can use 'kfold', 'holdout', 'leaveout', or 'cvpartition' parameters to override this cross-validation setting. You can only use one of these four parameters ('kfold', 'holdout', 'leaveout', or 'cvpartition') at a time when creating a cross-validated tree.

Alternatively, cross validate tree later using the crossval method.

Default: 'off'
cvpartition
Partition created with cvpartition to use in a cross-validated tree. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
holdout
Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1 . You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

## ClassificationTree.fit

kfold
Number of folds to use in a cross-validated tree, a positive integer. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

Default: 10

## leaveout

Use leave-one-out cross validation by setting to 'on'. You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

## MergeLeaves

When 'on', ClassificationTree.fit merges leaves that originate from the same parent node, and that give a sum of risk values greater or equal to the risk associated with the parent node. When 'off', ClassificationTree.fit does not merge leaves.

Default: 'on'

## MinLeaf

Each leaf has at least MinLeaf observations per tree leaf. If you supply both MinParent and MinLeaf, ClassificationTree.fit uses the setting that gives larger leaves: MinParent=max(MinParent,2*MinLeaf).

Default: 1

## MinParent

Each branch node in the tree has at least MinParent observations. If you supply both MinParent and MinLeaf, ClassificationTree.fit uses the setting that gives larger leaves: MinParent=max(MinParent,2*MinLeaf).

## ClassificationTree.fit

Default: 10

## NVarToSample

Number of predictors to select at random for each split. Can be a positive integer or 'all', which means use all available predictors.

Default: 'all'

## PredictorNames

A cell array of names for the predictor variables, in the order in which they appear in $X$.

Default: \{'x1','x2', ...\} prior

Prior probabilities for each class. Specify as one of:

- A string:
- 'empirical' determines class probabilities from class frequencies in Y. If you pass observation weights, they are used to compute the class probabilities.
- 'uniform' sets all class probabilities equal.
- A vector (one scalar value for each class)
- A structure S with two fields:
- S.ClassNames containing the class names as a variable of the same type as $Y$
- S.ClassProbs containing a vector of corresponding probabilities


## ClassificationTree.fit

If you set values for both weights and prior, the weights are renormalized to add up to the value of the prior probability in the respective class.

Default: 'empirical'

## Prune

When 'on', ClassificationTree.fit grows the classification tree, and computes the optimal sequence of pruned subtrees. When 'off' ClassificationTree.fit grows the classification tree without pruning.

Default: 'on'

## PruneCriterion

String with the pruning criterion, either 'error' or 'impurity'.
Default: 'error'

## ResponseName

Name of the response variable $Y$, a string.
Default: 'Response'

## ScoreTransform

Function handle for transforming scores, or string representing a built-in transformation function.

| String | Formula |
| :--- | :--- |
| 'symmetric' | $2 x-1$ |
| 'invlogit' | $\log (x /(1-x))$ |

## ClassificationTree.fit

| String | Formula |
| :--- | :--- |
| 'ismax' | Set score for the class with the <br> largest score to 1, and scores for all <br> other classes to 0. |
| 'symmetricismax' | Set score for the class with the <br> largest score to 1, and scores for <br> all other classes to -1. |
| 'none' | $x$ |
| 'logit' | $1 /\left(1+e^{-x}\right)$ |
| 'doublelogit' | $1 /\left(1+e^{-2 x}\right)$ |
| 'symmetriclogit' | $2 /\left(1+e^{-x}\right)-1$ |
| 'sign' | -1 for $x<0$ <br> 0 for $x=0$ <br> 1 for $x>0$ |

You can include your own function handle for transforming scores. Your function should accept a matrix (the original scores) and return a matrix of the same size (the transformed scores).

Default: 'none'

## SplitCriterion

Criterion for choosing a split. One of 'gdi' (Gini's diversity index), 'twoing' for the twoing rule, or 'deviance' for maximum deviance reduction (also known as cross entropy).

Default: 'gdi'

## Surrogate

When 'on', ClassificationTree.fit finds surrogate splits at each branch node. This setting improves the accuracy of predictions for data with missing values. The setting also enables

## ClassificationTree.fit

you to compute measures of predictive association between predictors. This setting can use much time and memory.

Default: 'off'

## weights

Vector of observation weights. The length of weights is the number of rows in X. ClassificationTree.fit normalizes the weights in each class to add up to the value of the prior probability of the class.

Default: ones(size (X,1), 1)

## Output tree

Arguments
A classification tree object. You can use tree to predict the response of new data with the predict method.

## Definitions

## Impurity and Node Error

ClassificationTree splits nodes based on either impurity or node error. Impurity means one of several things, depending on your choice of the SplitCriterion name-value pair:

- Gini's Diversity Index (gdi) - The Gini index of a node is

$$
1-\sum_{i} p^{2}(i),
$$

where the sum is over the classes $i$ at the node, and $p(i)$ is the observed fraction of classes with class $i$ that reach the node. A node with just one class (a pure node) has Gini index 0 ; otherwise the Gini index is positive. So the Gini index is a measure of node impurity.

- Deviance ('deviance') — With $p(i)$ defined as for the Gini index, the deviance of a node is


## ClassificationTree.fit

$$
-\sum_{i} p(i) \log p(i) .
$$

A pure node has deviance 0 ; otherwise, the deviance is positive.

- Twoing rule ('twoing') — Twoing is not a purity measure of a node, but is a different measure for deciding how to split a node. Let $L(i)$ denote the fraction of members of class $i$ in the left child node after a split, and $R(i)$ denote the fraction of members of class $i$ in the right child node after a split. Choose the split criterion to maximize

$$
P(L) P(R)\left(\sum_{i}|L(i)-R(i)|\right)^{2}
$$

where $P(L)$ and $P(R)$ are the fractions of observations that split to the left and right respectively. If the expression is large, the split made each child node purer. Similarly, if the expression is small, the split made each child node similar to each other, and hence similar to the parent node, and so the split did not increase node purity.

- Node error - The node error is the fraction of misclassified classes at a node. If $j$ is the class with largest number of training samples at a node, the node error is

$$
1-p(j)
$$

Examples Construct a classification tree for the data in ionosphere.mat:

## ClassificationTree.fit

```
load ionosphere
tc = ClassificationTree.fit(X,Y)
tc =
ClassificationTree:
            VarNames: {1x34 cell}
    CategoricalVars: []
        ClassNames: {'b' 'g'}
            Cost: [2x2 double]
            Prior: [0.3590 0.6410]
        TransformScore: 'none'
            X: [351x34 double]
            Y: {351x1 cell}
            W: [351x1 double]
            ModelParams: [1x1 classreg.learning.modelparams.TreeParams]
```

See Also predict | ClassificationTree

```How To • Chapter 13, "Nonparametric Supervised Learning"
```

Purpose
Gaussian mixture parameter estimates

## Syntax

Description
obj = gmdistribution.fit( $\mathrm{X}, \mathrm{k}$ )
obj = gmdistribution.fit(...,param1,val1,param2,val2,...)
obj $=$ gmdistribution.fit ( $\mathrm{X}, \mathrm{k}$ ) uses an Expectation Maximization (EM) algorithm to construct an object obj of the gmdistribution class containing maximum likelihood estimates of the parameters in a Gaussian mixture model with k components for data in the $n$-by- $d$ matrix X , where $n$ is the number of observations and $d$ is the dimension of the data.
gmdistribution treats NaN values as missing data. Rows of X with NaN values are excluded from the fit.
obj = gmdistribution.fit(...,param1,val1,param2,val2,...) provides control over the iterative EM algorithm. Parameters and values are listed below.

| Parameter | Value |
| :--- | :--- |
| 'Start' | Method used to choose initial component parameters. <br> One of the following: |
|  | - 'randSample ' - To select k observations from X at <br> random as initial component means. The mixing <br> proportions are uniform. The initial covariance <br> matrices for all components are diagonal, where <br> the element j on the diagonal is the variance of <br> $\mathrm{X}(:, \mathrm{j})$. This is the default. |
|  | - $\mathrm{s}-\mathrm{A}$ structure array with fields mu, Sigma, <br> and PComponents. See gmdistribution for <br> descriptions of values. <br> - s-A vector of length $n$ containing an initial guess <br> of the component index for each point. |
| 'Replicates' | A positive integer giving the number of times to <br> repeat the EM algorithm, each time with a new set of |


| Parameter | Value |
| :--- | :--- |
|  | parameters. The solution with the largest likelihood <br> is returned. A value larger than 1 requires the <br> 'randSample' start method. The default is 1. |
| 'CovType' | 'diagonal' if the covariance matrices are restricted <br> to be diagonal; ' full' otherwise. The default is <br> 'full'. |
| 'SharedCov' | Logical true if all the covariance matrices are <br> restricted to be the same (pooled estimate); logical <br> false otherwise. |
| 'Regularize' | A nonnegative regularization number added to <br> the diagonal of covariance matrices to make them <br> positive-definite. The default is 0. |
| 'Options' | Options structure for the iterative EM algorithm, as <br> created by statset. gmdistribution.fit uses the <br> parameters 'Display' with a default value of 'off ', <br> 'MaxIter' with a default value of 100, and 'TolFun' <br> with a default value of 1e6. |

In some cases, gmdistribution may converge to a solution where one or more of the components has an ill-conditioned or singular covariance matrix.

The following issues may result in an ill-conditioned covariance matrix:

- The number of dimension of your data is relatively high and there are not enough observations.
- Some of the features (variables) of your data are highly correlated.
- Some or all the features are discrete.
- You tried to fit the data to too many components.

In general, you can avoid getting ill-conditioned covariance matrices by using one of the following precautions:

## gmdistribution.fit

- Pre-process your data to remove correlated features.
- Set 'SharedCov' to true to use an equal covariance matrix for every component.
- Set 'CovType' to 'diagonal'.
- Use 'Regularize' to add a very small positive number to the diagonal of every covariance matrix.
- Try another set of initial values.

In other cases gmdistribution may pass through an intermediate step where one or more of the components has an ill-conditioned covariance matrix. Trying another set of initial values may avoid this issue without altering your data or model.

## Examples Generate data from a mixture of two bivariate Gaussian distributions using the mvnrnd function:

```
MU1 = [1 2];
SIGMA1 = [2 0; 0 .5];
MU2 = [-3 -5];
SIGMA2 = [1 0; 0 1];
X = [mvnrnd(MU1,SIGMA1,1000);mvnrnd(MU2,SIGMA2,1000)];
scatter(X(:,1),X(:,2),10,'.')
hold on
```



Next, fit a two-component Gaussian mixture model:

```
options = statset('Display','final');
obj = gmdistribution.fit(X,2,'Options',options);
10 iterations, log-likelihood = -7046.78
h = ezcontour(@(x,y)pdf(obj,[x y]),[-8 6],[-8 6]);
```



Among the properties of the fit are the parameter estimates:

```
ComponentMeans = obj.mu
ComponentMeans =
    0.9391 2.0322
    -2.9823 -4.9737
ComponentCovariances = obj.Sigma
ComponentCovariances(:,:,1) =
    1.7786 
```

```
ComponentCovariances(:,:,2) =
    1.0491 -0.0150
    -0.0150 0.9816
MixtureProportions = obj.PComponents
MixtureProportions =
    0.5000 0.5000
```

The Akaike information is minimized by the two-component model:

```
AIC = zeros(1,4);
obj = cell(1,4);
for k = 1:4
    obj{k} = gmdistribution.fit(X,k);
    AIC(k)= obj{k}.AIC;
end
[minAIC,numComponents] = min(AIC);
numComponents
numComponents =
        2
model = obj{2}
model =
Gaussian mixture distribution
with 2 components in 2 dimensions
Component 1:
Mixing proportion: 0.500000
Mean: 0.9391 2.0322
Component 2:
Mixing proportion: 0.500000
Mean: -2.9823 -4.9737
```

Both the Akaike and Bayes information are negative log-likelihoods for the data with penalty terms for the number of estimated parameters. They are often used to determine an appropriate number of components for a model when the number of components is unspecified.

## gmdistribution.fit

References [1] McLachlan, G., and D. Peel. Finite Mixture Models. Hoboken, NJ: John Wiley \& Sons, Inc., 2000.<br>See Also gmdistribution | cluster

## Purpose

Create Naive Bayes classifier object by fitting training data
Syntax
nb = NaiveBayes.fit(training, class)
nb = NaiveBayes.fit(..., 'param1', val1, 'param2',val2, ...)
Description
nb = NaiveBayes.fit(training, class) builds a NaiveBayes classifier object nb. training is an $N$-by-D numeric matrix of training data. Rows of training correspond to observations; columns correspond to features. class is a classing variable for training (see "Grouped Data" on page 2-34) taking K distinct levels. Each element of class defines which class the corresponding row of training belongs to. training and class must have the same number of rows.
nb = NaiveBayes.fit(..., 'param1', val1, 'param2', val2, ...) specifies one or more of the following name/value pairs:

- 'Distribution' - a string or a 1-by-D cell vector of strings, specifying which distributions fit uses to model the data. If the value is a string, fit models all the features using one type of distribution. fit can also model different features using different types of distributions. If the value is a cell vector, its $j$ th element specifies the distribution fit uses for the jth feature. The available types of distributions are:

| 'normal' | Normal (Gaussian) distribution. |
| :--- | :--- |
| (default) |  |
| 'kernel' | Kernel smoothing density estimate. |

## NaiveBayes.fit

'mvmn' Multivariate multinomial distribution for discrete data. fit assumes each individual feature follows a multinomial model within a class. The parameters for a feature include the probabilities of all possible values that the corresponding feature can take.
'mn' Multinomial distribution for classifying the count-based data such as the bag-of-tokens model. In the bag-of-tokens model, the value of the $j$ th feature is the number of occurrences of the jth token in this observation, so it must be a non-negative integer. When 'mn' is used, fit considers each observation as multiple trials of a multinomial distribution, and considers each occurrence of a token as one trial. The number of categories (bins) in this multinomial model is the number of distinct tokens, i.e., the number of columns of training.

- 'Prior' - The prior probabilities for the classes, specified as one of the following:

| 'empirical' |  |
| :--- | :--- |
| (default) | fit estimates the prior probabilities from the <br> relative frequencies of the classes in training. |
| 'uniform' | The prior probabilities are equal for all classes. |

> vector
> structure
> A numeric vector of length $K$ specifying the prior probabilities in the class order of class.
> A structure $S$ containing class levels and their prior probabilities. S must have two fields:
> - S.prob: A numeric vector of prior probabilities.
> - S.class: A vector of the same type as class, containing unique class levels indicating the class for the corresponding element of prob. S.class must contain all the K levels in class. It can also contain classes that do not appear in class. This can be useful if training is a subset of a larger training set. fit ignores any classes that appear in S.class but not in class.

If the prior probabilities don't sum to one, fit will normalize them.

- 'KSWidth' - The bandwidth of the kernel smoothing window. The default is to select a default bandwidth automatically for each combination of feature and class, using a value that is optimal for a Gaussian distribution. You can specify the value as one of the following:

| scalar | Width for all features in all classes. |
| :--- | :--- |
| row | 1 -by-D vector where the $j$ th element is the bandwidth for <br> vector |
| the $j$ th feature in all classes. |  |
| column | K-by-1 vector where the ith element specifies the |
| vector | bandwidth for all features in the ith class. $K$ represents <br> the number of class levels. |

## NaiveBayes.fit

matrix K-by-D matrix $M$ where $M(i, j)$ specifies the bandwidth for the j th feature in the ith class.
structure A structure $S$ containing class levels and their bandwidths. S must have two fields:

- S.width - A numeric array of bandwidths specified as a row vector, or a matrix with $D$ columns.
- S.class - A vector of the same type as class, containing unique class levels indicating the class for the corresponding row of width.
- 'KSSupport' - The regions where the density can be applied. It can be a string, a two-element vector as shown below, or a 1-by-D cell array of these values:

| 'unbounded ' <br> (default) | The density can extend over the whole real <br> line. |
| :--- | :--- |
| 'positive' | The density is restricted to positive values. |
| $[L, U]$ | A two-element vector specifying the finite <br> lower bound L and upper bound U for the <br> support of the density. |

- 'KSType ' - The type of kernel smoother to use. It can be a string or a 1-by-D cell array of strings. Each string can be 'normal' (default), 'box', 'triangle', or 'epanechnikov'.

How To

- "Naive Bayes Classification" on page 12-29
- "Grouped Data" on page 2-34


## Purpose

Binary decision tree for regression
Syntax

Description

## Input Arguments

tree $=$ RegressionTree.fit $(X, Y)$
tree $=$ RegressionTree.fit(X,Y,Name,Value) node is split based on the values of a column of $X$. is of class RegressionTree.

X
tree $=$ RegressionTree.fit $(X, Y)$ returns a regression tree based on the input variables (also known as predictors, features, or attributes) $X$ and output (response) Y. tree is a binary tree where each branching
tree $=$ RegressionTree.fit(X,Y,Name, Value) fits a tree with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN. If you use one of the following five options, tree is of class RegressionPartitionedModel: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Otherwise, tree

A matrix of predictor values. Each column of $X$ represents one variable, and each row represents one observation.
RegressionTree.fit considers NaN values in X as missing values. RegressionTree.fit does not use observations with all missing values for X the fit. RegressionTree.fit uses observations with some missing values for X to find splits on variables for which these observations have valid values.

## Y

A numeric column vector with the same number of rows as $X$. Each entry in $Y$ is the response to the data in the corresponding row of $X$.
RegressionTree.fit considers NaN values in $Y$ to be missing values. RegressionTree.fit does not use observations with missing values for $Y$ in the fit.

## RegressionTree.fit

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## CategoricalPredictors

List of categorical predictors. Pass CategoricalPredictors as one of:

- A numeric vector with indices from 1 to $p$, where $p$ is the number of columns of $X$.
- A logical vector of length $p$, where a true entry means that the corresponding column of X is a categorical variable.
- 'all', meaning all predictors are categorical.
- A cell array of strings, where each element in the array is the name of a predictor variable. The names must match entries in the PredictorNames property.
- A character matrix, where each row of the matrix is a name of a predictor variable. Pad the names with extra blanks so each row of the character matrix has the same length.


## Default: []

## crossval

If 'on', grows a cross-validated decision tree with 10 folds. You can use 'kfold', 'holdout', 'leaveout', or 'cvpartition' parameters to override this cross-validation setting. You can only use one of these four parameters ('kfold', 'holdout', 'leaveout', or 'cvpartition') at a time when creating a cross-validated tree.

Alternatively, cross-validate tree later using the crossval method.
Default: 'off'

## cvpartition

A partition created with cupartition to use in cross-validated tree. You can only use one of these four options at a time:
'kfold', 'holdout', 'leaveout' and 'cvpartition'.
holdout
Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1 . You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout' and 'cvpartition'.

## kfold

Number of folds to use in a cross-validated tree, a positive integer. You can only use one of these four options at a time: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
Default: 10

## leaveout

Use leave-one-out cross validation by setting to 'on'. You can only use one of these four options at a time: 'kfold', 'holdout ', 'leaveout', or 'cvpartition'.

## MergeLeaves

When 'on', RegressionTree merges leaves that originate from the same parent node, and that give a sum of risk values greater or equal to the risk associated with the parent node. When 'off', RegressionTree does not merge leaves.
Default: 'on'
MinLeaf

## RegressionTree.fit

Each leaf has at least MinLeaf observations per tree leaf. If you supply both MinParent and MinLeaf, RegressionTree uses the setting that gives larger leaves: MinParent=max(MinParent, 2*MinLeaf).

Default: 1

## MinParent

Each branch node in the tree has at least MinParent observations. If you supply both MinParent and MinLeaf, RegressionTree uses the setting that gives larger leaves: MinParent=max(MinParent, 2*MinLeaf).

Default: 10

## NVarToSample

Number of predictors to select at random for each split. Can be a positive integer or 'all', which means use all available predictors.

Default: 'all'

## PredictorNames

A cell array of names for the predictor variables, in the order in which they appear in X .

Default: \{'x1','x2',...\}
Prune
When 'on', RegressionTree computes the full tree and the optimal sequence of pruned subtrees. When 'off' RegressionTree computes the full tree without pruning.

Default: 'on'

## PruneCriterion

String with the pruning criterion, always 'error'.
Default: 'error'
ResponseName
Name of the response variable Y , a string.
Default: ' $Y$ '

## ResponseTransform

Function handle for transforming the raw response values (mean squared error). The function handle should accept a matrix of response values and return a matrix of the same size. The default string ' none' means @(x)x, or no transformation.

Add or change a ResponseTransform function by dot addressing:
tree.ResponseTransform = @function
Default: 'none'

## SplitCriterion

Criterion for choosing a split, always the string 'MSE', meaning mean squared error.

Default: 'MSE'
Weights
Vector of observation weights. The length of weights is the number of rows in $X$.

Default: ones(size(X,1),1)

## RegressionTree.fit

```
Output tree
Arguments
Examples Load the data in carsmall.mat, and make a regression tree to predict the mileage of cars based on their weights and numbers of cylinders:
```

```
load carsmall
```

load carsmall
tree = RegressionTree.fit([Weight, Cylinders],MPG,...
tree = RegressionTree.fit([Weight, Cylinders],MPG,...
'MinParent',20,...
'MinParent',20,...
'PredictorNames',{'W','C'})
'PredictorNames',{'W','C'})
tree =
tree =
RegressionTree:
RegressionTree:
PredictorNames: {'W' 'C'}
PredictorNames: {'W' 'C'}
CategoricalPredictors: []
CategoricalPredictors: []
ResponseName: 'Y'
ResponseName: 'Y'
ResponseTransform: 'none'
ResponseTransform: 'none'
NObservations: 94

```
            NObservations: 94
```

Predict the mileage of a car that weighs 2200 lbs and has four cylinders:

```
predict(tree,[2200,4])
ans =
    29.6111
```

See Also
predict
How To • Chapter 13, "Nonparametric Supervised Learning"

Purpose Fit probability distribution to data
Syntax
$P D=$ fitdist (X, DistName)
[PDCA, GN, GL] = fitdist(X, DistName, 'By', GroupVar) ... = fitdist(..., param1, val1, param2, val2, ...)

Input Arguments
$P D=$ fitdist ( $X$, DistName) fits the probability distribution specified by DistName to the data in the column vector $X$, and returns $P D$, an object representing the fitted distribution.
[PDCA, GN, GL] = fitdist(X, DistName, 'By', GroupVar) takes a grouping variable, GroupVar, fits the specified distribution to the data in $X$ from each group, and returns PDCA, a cell array of the fitted probability distribution objects. GroupVar can also be a cell array of multiple grouping variables. $G N$ is a cell array of group labels. $G L$ is a cell array of grouping variable levels, with one column for each grouping variable. See "Grouped Data" on page 2-34 for more information.
... = fitdist(..., param1, val1, param2, val2, ...) specifies optional parameter name/value pairs, as described in the Parameter/Values table. Parameter and value names are case insensitive.
$x$ $\begin{array}{ll}\begin{array}{l}\text { Note Any NaN values in } X \text { are ignored by the } \\ \text { fitting calculations. Additionally, any NaN } \\ \text { values in the censoring vector or frequency } \\ \text { vector will cause the corresponding values in } \\ X \text { to be ignored by the fitting calculations. }\end{array} \\ \text { DistName } & \text { A string specifying a distribution. Choices are: }\end{array}$

- 'kernel' - To fit a nonparametric kernel-smoothing distribution.
- Any of the following to fit a parametric distribution:
- 'beta'
- 'binomial'
- 'birnbaumsaunders'
- 'exponential'
- 'extreme value' or 'ev'
- 'gamma'
- 'generalized extreme value' or 'gev'
- 'generalized pareto' or 'gp'
- 'inversegaussian'
- 'logistic'
- 'loglogistic'
- 'lognormal'
- 'nakagami'
- 'negative binomial' or 'nbin'
- 'normal'
- 'poisson'
- 'rayleigh'
- 'rician'
- 'tlocationscale'
- 'weibull' or 'wbl'

For more information on these parametric distributions, see Appendix B, "Distribution Reference".

GroupVar
A grouping variable or a cell array of multiple grouping variables. For more information on grouping variables, see "Grouped Data" on page 2-34.

| Parameter | Values |
| :---: | :---: |
| 'censoring | 'A Boolean vector the same size as $X$, containing 1 s when the corresponding elements in $X$ are right-censored observations and 0s when the corresponding elements are exact observations. Default is a vector of 0s. |
|  | Note Any NaN values in this censoring vector are ignored by the fitting calculations. Additionally, any NaN values in $X$ or the frequency vector will cause the corresponding values in the censoring vector to be ignored by the fitting calculations. |
| 'frequency | 'A vector the same size as $X$, containing nonnegative integers specifying the frequencies for the corresponding elements in $X$. Default is a vector of 1 s . |
|  | Note Any NaN values in this frequency vector are ignored by the fitting calculations. Additionally, any NaN values in $X$ or the censoring vector will cause the corresponding values in the frequency vector to be ignored by the fitting calculations. |

## fitdist

| Parameter Values |  |
| :--- | :--- |
| 'options' | A structure created by the statset function to specify <br> control parameters for the iterative fitting algorithm. |
| 'n' | For 'binomial' distributions only, a positive integer <br> specifying the N parameter (number of trials). |
| 'theta' | For 'generalized pareto' distributions only, value <br> specifying the theta (threshold) parameter for the <br> generalized Pareto distribution. Default is 0. |
| 'kernel' | For 'kernel' distributions only, a string specifying the <br> type of kernel smoother to use. Choices are: |
| - 'normal' (default) |  |
| - 'box' |  |
| - 'triangle' |  |
| - 'epanechnikov' |  |

## Output $P D$ <br> Arguments <br> PDCA GN GL

An object in either the ProbDistUnivKernel class or the ProbDistUnivParam class, which are derived from the ProbDist class.

A cell array of the fitted probability distribution objects.

A cell array of group labels.
A cell array of grouping variable levels, with one column for each grouping variable.

## Examples Creating a ProbDistUnivKernel Object

1 Load a MAT-file, included with the Statistics Toolbox software, which contains MPG, a column vector of data.
load carsmall

2 Create a ProbDistUnivKernel object by fitting a nonparametric kernel-smoothing distribution to the data:

```
ksd = fitdist(MPG,'kernel')
ksd =
kernel distribution
```

Kernel = normal
Bandwidth $=4.11428$
Support = unbounded

## Creating a ProbDistUnivParam Object

1 Load a MAT-file, included with the Statistics Toolbox software, which contains MPG, a column vector of data, and Origin, a cell array of seven grouping variables.

## fitdist

```
load carsmall
```

2 Create a cell array of ProbDistUnivParam objects by fitting a parametric distribution, namely a Weibull distribution, to the data, and also grouping the data. Since there is only one car from Italy, fitdist will return an error, since you cannot fit a distribution to a single observation.

```
wd = fitdist(MPG,'weibull','by',Origin)
```


## Algorithms

## References

Alternatives

See Also

The fitdist function fits most distributions using maximum likelihood. Two exceptions are the normal and lognormal distributions with uncensored data. For the uncensored normal distribution, the estimated value of the sigma parameter is the square root of the unbiased estimate of the variance. For the uncensored lognormal distribution, the estimated value of the sigma parameter is the square root of the unbiased estimate of the variance of the log of the data.
[1] Johnson, N. L., S. Kotz, and N. Balakrishnan. Continuous Univariate Distributions. Vol. 1, Hoboken, NJ: Wiley-Interscience, 1993.
[2] Johnson, N. L., S. Kotz, and N. Balakrishnan. Continuous Univariate Distributions. Vol. 2, Hoboken, NJ: Wiley-Interscience, 1994.
[3] Bowman, A. W., and A. Azzalini. Applied Smoothing Techniques for Data Analysis. New York: Oxford University Press, 1997.
dfittool - Opens a graphical user interface for displaying fit distributions to data, or for fitting distributions to your data and displaying them over plots of the empirical distributions, by importing data from the workspace.
disttool | randtool | statset | ProbDist | ProbDistUnivKernel | ProbDistUnivParam

## How To

- Appendix B, "Distribution Reference"
- "Grouped Data" on page 2-34


## fitensemble

Purpose Fitted ensemble for classification or regression
Syntax

ens = fitensemble(X,Y,method,nlearn,learners)

ens = fitensemble(X,Y,method,nlearn,learners,Name,Value)
ens = fitensemble( $X, Y$, method, nlearn, learners) creates an ensemble model that can predict responses to data. The ensemble consists of models listed in learners.
ens = fitensemble(X,Y,method, nlearn,learners,Name,Value) creates an ensemble model with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

- If $X$ has categorical predictors with many levels, use 'GentleBoost' or 'LogitBoost' for binary classification. Using other algorithms in this case would likely cause long training times and can exhaust memory.

X
A matrix of predictor values. Each column of $X$ represents one variable, and each row represents one observation.

Y
For classification, Y is a categorical variable, character array, or cell array of strings. Each row of $Y$ represents the classification of the corresponding row of $X$.
For regression, Y is a numeric column vector with the same number of rows as $X$. Each entry in $Y$ is the response to the data in the corresponding row of $X$. method

A case-insensitive string consisting of one of the following.

- For classification with two classes:


## fitensemble

- 'AdaBoostM1'
- 'LogitBoost'
- 'GentleBoost'
- 'RobustBoost'
- 'Bag'
- For classification with three or more classes:
- 'AdaBoostM2'
- 'Bag'
- For regression:
- 'LSBoost'
- 'Bag'

Since 'Bag' applies to all methods, indicate whether you want a classifier or regressor with the type name-value pair set to 'classification' or 'regression'.
nlearn
Number of ensemble learning cycles, a positive integer. At every training cycle, fitensemble loops over all learner templates in learners and trains one weak learner for every template. The total number of trained learners in ens is nlearn*numel(learners).
nlearn for ensembles can vary from a few dozen to a few thousand. Usually, an ensemble with a good predictive power needs between a few hundred and a few thousand weak learners. You do not have to train an ensemble for that many cycles at once. You can start by growing a few dozen learners, inspect the ensemble performance and, if necessary, train more weak learners using the resume method of the ensemble.

[^6]
## fitensemble

One of the following:

- A string with the name of a weak learner, such as 'tree'
- A single weak learner template you create with ClassificationTree.template or RegressionTree.template
- A cell array of weak learner templates

Create weak learner templates with ClassificationTree.template or RegressionTree.template. Usually you should supply only one weak learner template.

Ensemble performance depends on the parameters of the weak learners, and you can get poor performance using weak learners with default parameters. Specify the parameters for the weak learners in the template. Specify parameters for the ensemble in the fitensemble name-value pairs.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## All Ensembles:

## CategoricalPredictors

List of categorical predictors. Pass CategoricalPredictors as one of:

- A numeric vector with indices from 1 to $p$, where $p$ is the number of columns of $X$.
- A logical vector of length $p$, where a true entry means that the corresponding column of X is a categorical variable.
- 'all', meaning all predictors are categorical.


## fitensemble

- A cell array of strings, where each element in the array is the name of a predictor variable. The names must match entries in the PredictorNames property.
- A character matrix, where each row of the matrix is a name of a predictor variable. The names must match entries in the PredictorNames property. Pad the names with extra blanks so each row of the character matrix has the same length.


## Default: []

## crossval

If 'on', grows a cross-validated decision tree with 10 folds. You can use 'kfold', 'holdout', 'leaveout', or 'cvpartition' parameters to override this cross-validation setting. You can only use one of these four parameters ('kfold', 'holdout', 'leaveout', or 'cvpartition') at a time when creating a cross-validated tree.

Default: 'off'

## cvpartition

A partition created with cupartition to use in cross-validated tree. You can only use one of these four options at a time:
'kfold', 'holdout', 'leaveout', or 'cvpartition'.
fresample
Fraction of the training set to be selected by resampling for every weak learner. A numeric scalar between 0 and 1. This parameter has no effect unless you grow an ensemble by bagging or set 'resample' to 'on'. The default setting is the one used most often for an ensemble grown by resampling.

## Default: 1

holdout

## fitensemble

Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1 . You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
kfold
Number of folds to use in a cross-validated tree, a positive integer.
You can only use one of these four options at a time: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

Default: 10
leaveout
Use leave-one-out cross validation by setting to 'on '. You can only use one of these four options at a time: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
nprint
Printout frequency, a positive integer scalar. Set to 'off' for no printout. Use this parameter to track of how many weak learners have been trained so far. This is useful when you train ensembles with many learners on large datasets. If you use one of the cross-validation options, this parameter defines the printout frequency per number of cross-validation folds.

Default: 'off'
PredictorNames
A cell array of names for the predictor variables, in the order in which they appear in X .

Default: \{'x1','x2',...\}
replace

## fitensemble

'on' or 'off'. If 'on', sample with replacement; if 'off', sample without replacement. This parameter has no effect unless you grow an ensemble by bagging or set resample to 'on'. If you set resample to 'on' and replace to 'off', fitensemble samples training observations assuming uniform weights, and boosts by reweighting observations.

Default: 'on'

## resample

'on' or 'off'. If 'on', grows an ensemble by resampling, with the resampling fraction given by fresample, and sampling with or without replacement given by replace.

- Boosting - When 'off', the boosting algorithm reweights observations at every learning iteration. When 'on', the algorithm samples training observations using updated weights as the multinomial sampling probabilities.
- Bagging - You can use only the default value of this parameter ('on').

Default: 'off' for boosting, 'on' for bagging

## ResponseName

Name of the response variable Y , a string.
Default: ' $Y$ '
type
A string, either 'classification' or 'regression'. Specify type when the method is 'bag'.

## weights

Vector of observation weights. The length of weights is the number of rows in $X$.

## fitensemble

Default: ones(size(X,1),1)

## Classification Ensembles:

classnames
Array of class names. Specify a data type the same as exists in Y .
Default: The class names that exist in $Y$
cost
Square matrix $C$, where $C(i, j)$ is the cost of classifying a point into class $j$ if its true class is i. Alternatively, cost can be a structure S having two fields: S.ClassNames containing the group names as a categorical variable, character array, or cell array of strings; and S.ClassificationCosts containing the cost matrix C.

Default: $C(i, j)=1$ if $i \sim=j$, and $C(i, j)=0$ if $i=j$
prior
Prior probabilities for each class. Specify as one of:

- A string:
- 'empirical' determines class probabilities from class frequencies in Y. If you pass observation weights, they are used to compute the class probabilities.
- 'uniform' sets all class probabilities equal.
- A vector (one scalar value for each class)
- A structure S with two fields:
- S.ClassNames containing the class names as a categorical variable, character array, or cell array of strings


## fitensemble

- S.ClassProbs containing a vector of corresponding probabilities

If you set values for both weights and prior, the weights are renormalized to add up to the value of the prior probability in the respective class.

Default: 'empirical'

## AdaBoostM1, AdaBoostM2, LogitBoost, GentleBoost, and LSBoost:

LearnRate
Learning rate for shrinkage, a numeric scalar between 0 and 1. If you set the learning rate to a smaller value than 1 , the ensemble requires more learning iterations but often achieves a better accuracy. 0.1 is a popular choice for ensemble grown with shrinkage.

Default: 1

## RobustBoost:

## RobustErrorGoal

Target classification error for RobustBoost, a numeric scalar from 0 to 1 . Usually there is an optimal range for this parameter for your training data. If you set the error goal too low or too high, RobustBoost can produce a model with poor classification accuracy.

Default: 0.1
RobustMaxMargin
Maximal classification margin for RobustBoost in the training set, a nonnegative numeric scalar. RobustBoost minimizes the

## fitensemble

number of observations in the training set with classification margins below RobustMaxMargin.

Default: 0

## RobustMarginSigma

Spread of the distribution of classification margins over the training set for RobustBoost, a numeric positive scalar. You should consult literature on RobustBoost before setting this parameter

Default: 0.1

## Output <br> ens

 ArgumentsEnsemble object for predicting characteristics. The class of ens depends on settings. In the following table, cross-validation names are crossval, kfold, holdout, leaveout, and cvpartition.

| Settings | Class |
| :--- | :--- |
| resample name-value pair is 'off', and you <br> don't set a cross-validation name-value pair. | ClassificationEnsemble |
| resample name-value pair is 'off ', and you <br> don't set a cross-validation name-value pair. | RegressionEnsemble |
| resample name-value pair is 'on ', type <br> is 'classification', and you don't set a <br> cross-validation name-value pair. | ClassificationBaggedEnsemble |
| resample name-value pair is 'on ', type <br> is 'regression', and you don't set a <br> cross-validation name-value pair. | RegressionBaggedEnsemble |
| method is a classification method, and you <br> set a cross-validation name-value pair. | ClassificationPartitionedEnsemble |
| method is a regression method, and you set <br> a cross-validation name-value pair. | RegressionPartitionedEnsemble |

## fitensemble

Examples Train a boosting ensemble, and inspect the resubstitution loss:

```
load ionosphere;
ada = fitensemble(X,Y,'AdaBoostM1',100,'tree');
plot(resubLoss(ada,'mode','cumulative'));
xlabel('Number of decision trees');
ylabel('Resubstitution error');
```



## fitensemble

> Train a regression ensemble to predict car mileage based on the number of cylinders, engine displacement, horsepower, and weight. Predict the mileage for a four-cylinder car with a 200 cubic inch displacement, 150 horsepower, weighing 3000 lbs.

```
load carsmall
X = [Cylinders Displacement Horsepower Weight];
xnames = {'Cylinders' 'Displacement' 'Horsepower' 'Weight'};
t = RegressionTree.template('Surrogate','on');
rens = fitensemble(X,MPG,'LSBoost',100,t,'PredictorNames',xnames)
rens =
classreg.learning.regr.RegressionEnsemble:
            PredictorNames: {'Cylinders' 'Displacement' 'Horsepower' 'Weight'}
    CategoricalPredictors: []
            ResponseName: 'Y'
        ResponseTransform: 'none'
            NObservations: 94
            NTrained: 100
                    Method: 'LSBoost'
            LearnerNames: {'Tree'}
    ReasonForTermination: [1x77 char]
            FitInfo: [100x1 double]
        FitInfoDescription: [2x83 char]
            Regularization: []
predict(rens,[4 200 150 3000])
ans =
    20.4982
```

Train and estimate generalization error on a holdout sample:

```
load ionosphere;
ada = fitensemble(X,Y,'AdaBoostM1',100,'Tree',...
```

```
    'holdout',0.5);
plot(kfoldLoss(ada,'mode','cumulative'));
xlabel('Number of decision trees');
ylabel('Holdout error');
```



See Also ClassificationEnsemble | RegressionEnsemble |
ClassificationBaggedEnsemble | RegressionBaggedEnsemble
| ClassificationPartitionedEnsemble |
RegressionPartitionedEnsemble | ClassificationTree.template |
RegressionTree.template

How To - Chapter 13, "Nonparametric Supervised Learning"

Purpose Flip categorical array along specified dimension

## Syntax <br> $B=$ flipdim(A,dim)

Description $\quad B=$ flipdim (A, dim) returns the categorical array A with dimension dim flipped.

See Also fliplr | flipud | permute | rot90
Purpose Flip categorical matrix in left/right direction
Syntax

$$
B=f l i p l r(A)
$$

Description $B=f l i p l r(A)$ returns the $2-D$ categorical matrix $A$ with rows preserved and columns flipped in the left/right direction.
See Also flipdim | flipud | permute | rot90

## categorical.flipud

Purpose Flip categorical matrix in up/down direction

## Syntax <br> B = flipud(A)

Description $\quad B=$ flipud $(A)$ returns the 2-D categorical matrix $A$ with rows preserved and columns flipped in the up/down direction.

See Also flipdim | fliplr | permute | rot90

## Purpose

$F$ probability density function

## Syntax

$Y=f p d f(X, V 1, V 2)$
$Y=\operatorname{fpdf}(X, V 1, V 2)$ computes the $F$ pdf at each of the values in $X$ using the corresponding numerator degrees of freedom V1 and denominator degrees of freedom V2. X, V1, and V2 can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. V1 and V2 parameters must contain real positive values, and the values in X must lie on the interval $[0 \infty$ ).

The probability density function for the $F$ distribution is

$$
y=f\left(x \mid v_{1}, v_{2}\right)=\frac{\Gamma\left[\frac{\left(v_{1}+v_{2}\right)}{2}\right]}{\Gamma\left(\frac{v_{1}}{2}\right) \Gamma\left(\frac{v_{2}}{2}\right)}\left(\frac{v_{1}}{v_{2}}\right)^{\frac{v_{1}}{2}} \frac{x^{\frac{v_{1}-2}{2}}}{\left[1+\left(\frac{v_{1}}{v_{2}}\right) x\right]^{\frac{v_{1}+v_{2}}{2}}}
$$

## Examples

```
\(y=\operatorname{fpdf}(1: 6,2,2)\)
\(y=\)
    \(\begin{array}{llllll}0.2500 & 0.1111 & 0.0625 & 0.0400 & 0.0278 & 0.0204\end{array}\)
\(z=f p d f(3,5: 10,5: 10)\)
z =
    \(\begin{array}{llllll}0.0689 & 0.0659 & 0.0620 & 0.0577 & 0.0532 & 0.0487\end{array}\)
```

See Also
pdf | fcdf | finv | fstat | frnd

How To

- "F Distribution" on page B-25

Purpose Fractional factorial design
Syntax $\quad X=$ fracfact $($ gen $)$
[X,conf] = fracfact(gen)
[X,conf] = fracfact(gen, Name, Value)

Description

Input
Arguments
$X=$ fracfact (gen) creates the two-level fractional factorial design defined by the generator string gen.
[X, conf] = fracfact(gen) returns a cell array of strings containing the confounding pattern for the design.
[X,conf] = fracfact(gen, Name, Value) creates a fractional factorial designs with additional options specified by one or more Name, Value pair arguments.
gen
Either a cell array of strings where each cell contains one "word," or a string consisting of "words" separated by spaces. "Words" consist of case-sensitive letters or groups of letters, where 'a' represents string 1, ' $b$ ' represents string $2, \ldots$, ' $A$ ' represents string $27, \ldots$, 'Z' represents string 52.
Each word defines how the corresponding factor's levels are defined as products of generators from a $2^{\wedge} K$ full-factorial design. K is the number of letters of the alphabet in gen.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## FactorNames

Cell array specifying the name for each factor.
Default: \{'X1','X2',...\}

MaxInt
Positive integer setting the maximum level of interaction to include in the confounding output.

## Default: 2

## Output x

Arguments
The two-level fractional factorial design. X is a matrix of size N -by-P, where

- $N=2^{\wedge} K$, where $K$ is the number of letters of the alphabet in gen.
- $P$ is the number of words in gen.

Because $X$ is a two-level design, the components of $X$ are -1 . For the meaning of $X$, see "Fractional Factorial Designs" on page 15-5. conf

Cell array of strings containing the confounding pattern for the design.

## Examples Generate a fractional factorial design for four variables, where the fourth variable is the product of the first three:

| $\mathrm{x}=$ |  |  |  |
| :---: | :---: | :---: | :---: |
| -1 | -1 | -1 | -1 |
| -1 | -1 | 1 | 1 |
| -1 | 1 | -1 | 1 |
| -1 | 1 | 1 | -1 |
| 1 | -1 | -1 | 1 |
| 1 | -1 | 1 | -1 |
| 1 | 1 | -1 | -1 |
| 1 | 1 | 1 | 1 |

Find generators for a six-factor design that uses four factors and achieves resolution IV using fracfactgen. Use the result to specify the design:

```
generators = fracfactgen('a b c d e f',4, ... % 4 factors
    4) % resolution 4
generators =
    'a'
    'b'
    'c'
    'd'
    'bcd'
    ' acd '
x = fracfact(generators)
x =
\begin{tabular}{rrrrrr}
-1 & -1 & -1 & -1 & -1 & -1 \\
-1 & -1 & -1 & 1 & 1 & 1 \\
-1 & -1 & 1 & -1 & 1 & 1 \\
-1 & -1 & 1 & 1 & -1 & -1 \\
-1 & 1 & -1 & -1 & 1 & -1 \\
-1 & 1 & -1 & 1 & -1 & 1 \\
-1 & 1 & 1 & -1 & -1 & 1 \\
-1 & 1 & 1 & 1 & 1 & -1 \\
1 & -1 & -1 & -1 & -1 & 1 \\
1 & -1 & -1 & 1 & 1 & -1 \\
1 & -1 & 1 & -1 & 1 & -1 \\
1 & -1 & 1 & 1 & -1 & 1 \\
1 & 1 & -1 & -1 & 1 & 1 \\
1 & 1 & -1 & 1 & -1 & -1 \\
1 & 1 & 1 & -1 & -1 & -1 \\
1 & 1 & 1 & 1 & 1 & 1
\end{tabular}
```

References [1] Box, G. E. P., W. G. Hunter, and J. S. Hunter. Statistics for Experimenters. Hoboken, NJ: Wiley-Interscience, 1978.
See Also ff2n | fracfactgen | fullfact | hadamard
How To - "Fractional Factorial Designs" on page 15-5

Fractional factorial design generators

```
generators = fracfactgen(terms)
generators = fracfactgen(terms,k)
generators = fracfactgen(terms,k,R)
generators = fracfactgen(terms,k,R,basic)
```

generators $=$ fracfactgen(terms) uses the Franklin-Bailey algorithm to find generators for the smallest two-level fractional-factorial design for estimating linear model terms specified by terms. terms is a string consisting of words formed from the 52 case-sensitive letters a-Z, separated by spaces. Use 'a'-'z' for the first 26 factors, and, if necessary, 'A' - 'Z' for the remaining factors. For example, terms $=$ 'a b c ab ac'. Single-character words indicate main effects to be estimated; multiple-character words indicate interactions. Alternatively, terms is an $m$-by- $n$ matrix of 0 s and 1 s where $m$ is the number of model terms to be estimated and $n$ is the number of factors. For example, if terms contains rows [ $\left.\begin{array}{llll}0 & 1 & 0 & 0\end{array}\right]$ and [ $\left.\begin{array}{llll}1 & 0 & 0 & 1\end{array}\right]$, then the factor $b$ and the interaction between factors $a$ and d are included in the model. generators is a cell array of strings with one generator per cell. Pass generators to fracfact to produce the fractional-factorial design and corresponding confounding pattern.
generators = fracfactgen(terms,k) returns generators for a two-level fractional-factorial design with $2^{k}$-runs, if possible. If $k$ is [ ], fracfactgen finds the smallest design.
generators $=$ fracfactgen(terms, $k, R$ ) finds a design with resolution $R$, if possible. The default resolution is 3 .

A design of resolution $R$ is one in which no $n$-factor interaction is confounded with any other effect containing less than $R-n$ factors. Thus a resolution III design does not confound main effects with one another but may confound them with two-way interactions, while a resolution IV design does not confound either main effects or two-way interactions but may confound two-way interactions with each other.

If fracfactgen is unable to find a design at the requested resolution, it tries to find a lower-resolution design sufficient to calibrate the model.

If it is successful, it returns the generators for the lower-resolution design along with a warning. If it fails, it returns an error.
generators $=$ fracfactgen(terms, $k, R$, basic) also accepts a vector basic specifying the indices of factors that are to be treated as basic. These factors receive full-factorial treatments in the design. The default includes factors that are part of the highest-order interaction in terms.

## Examples

Suppose you wish to determine the effects of four two-level factors, for which there may be two-way interactions. A full-factorial design would require $2^{4}=16$ runs. The fracfactgen function finds generators for a resolution IV (separating main effects) fractional-factorial design that requires only $2^{3}=8$ runs:

```
generators = fracfactgen('a b c d',3,4)
generators =
    'a'
    'b'
    'c'
    'abc'
```

The more economical design and the corresponding confounding pattern are returned by fracfact:

```
[dfF,confounding] = fracfact(generators)
dfF =
```



## fracfactgen

| ' X3 ' | 'c' | ' X3 ' |
| :---: | :---: | :---: |
| ' $\mathrm{X} 4{ }^{\prime}$ | 'abc' | ' X4' |
| ' X1* 22 $^{\prime}$ | 'ab' | ${ }^{\prime} \mathrm{X} 1 * \mathrm{X} 2+\mathrm{X} 3 * \mathrm{X} 4{ }^{\prime}$ |
| ' X 1 * $\times 3{ }^{\prime}$ | 'ac' | ${ }^{\prime} \mathrm{X} 1 * \mathrm{X} 3+\mathrm{X} 2 * \mathrm{X} 4{ }^{\prime}$ |
| ' X 1 * $\mathrm{X} 4{ }^{\prime}$ | 'bc' | ${ }^{\prime} \mathrm{X} 1 * \mathrm{X} 4+\mathrm{X} 2 * \mathrm{X} 3$ ' |
| ' X 2 * $\times 3$ ' | 'bc ${ }^{\prime}$ | ${ }^{\prime} \mathrm{X} 1 * \mathrm{X} 4+\mathrm{X} 2 * \mathrm{X} 3$ ' |
| ' X 2 * $\mathrm{X} 4{ }^{\prime}$ | 'ac' | ${ }^{\prime} \mathrm{X} 1 * \mathrm{X} 3+\mathrm{X} 2 * \mathrm{X} 4{ }^{\prime}$ |
| 'X3*X4' | 'ab' | ' $\mathrm{X} 1 * \mathrm{X} 2+\mathrm{X} 3 * \mathrm{X} 4{ }^{\prime}$ |

The confounding pattern shows, for example, that the two-way interaction between X1 and X2 is confounded by the two-way interaction between X3 and X4.

References [1] Box, G. E. P., W. G. Hunter, and J. S. Hunter. Statistics for Experimenters. Hoboken, NJ: Wiley-Interscience, 1978.

See Also fracfact | hadamard
How To . "Fractional Factorial Designs" on page 15-5

## Purpose Friedman's test

Syntax $\quad$| $p=$ friedman $(X$, reps $)$ |
| :--- |
|  |
| $p=$ friedman $(X$, reps, displayopt $)$ |
|  |
|  |
|  |
| $[p$, table $]=$ friedman $(\ldots)$ |

## Description

$p$ = friedman(X, reps) performs the nonparametric Friedman's test to compare column effects in a two-way layout. Friedman's test is similar to classical balanced two-way ANOVA, but it tests only for column effects after adjusting for possible row effects. It does not test for row effects or interaction effects. Friedman's test is appropriate when columns represent treatments that are under study, and rows represent nuisance effects (blocks) that need to be taken into account but are not of any interest.

The different columns of $X$ represent changes in a factor $A$. The different rows represent changes in a blocking factor $B$. If there is more than one observation for each combination of factors, input reps indicates the number of replicates in each "cell," which must be constant.

The matrix below illustrates the format for a set-up where column factor A has three levels, row factor B has two levels, and there are two replicates (reps=2). The subscripts indicate row, column, and replicate, respectively.
$\left[\begin{array}{lll}x_{111} & x_{121} & x_{131} \\ x_{112} & x_{122} & x_{132} \\ x_{211} & x_{221} & x_{231} \\ x_{212} & x_{222} & x_{232}\end{array}\right]$

Friedman's test assumes a model of the form

$$
x_{i j k}=\mu+\alpha_{i}+\beta_{j}+\varepsilon_{i j k}
$$

## friedman

where $\mu$ is an overall location parameter, $\alpha_{i}$ represents the column effect, $\beta_{j}$ represents the row effect, and $\varepsilon_{i j k}$ represents the error. This test ranks the data within each level of B, and tests for a difference across levels of A. The p that friedman returns is the $p$ value for the null hypothesis that $\alpha_{i}=0$. If the $p$ value is near zero, this casts doubt on the null hypothesis. A sufficiently small $p$ value suggests that at least one column-sample median is significantly different than the others; i.e., there is a main effect due to factor A. The choice of a critical $p$ value to determine whether a result is "statistically significant" is left to the researcher. It is common to declare a result significant if the $p$ value is less than 0.05 or 0.01 .
friedman also displays a figure showing an ANOVA table, which divides the variability of the ranks into two or three parts:

- The variability due to the differences among the column effects
- The variability due to the interaction between rows and columns (if reps is greater than its default value of 1 )
- The remaining variability not explained by any systematic source

The ANOVA table has six columns:

- The first shows the source of the variability.
- The second shows the Sum of Squares (SS) due to each source.
- The third shows the degrees of freedom (df) associated with each source.
- The fourth shows the Mean Squares (MS), which is the ratio SS/df.
- The fifth shows Friedman's chi-square statistic.
- The sixth shows the $p$ value for the chi-square statistic.
$\mathrm{p}=\mathrm{friedman}(\mathrm{X}$, reps,displayopt) enables the ANOVA table display when displayopt is 'on' (default) and suppresses the display when displayopt is 'off'.
[ $p$,table] = friedman(...) returns the ANOVA table (including column and row labels) in cell array table. (You can copy a text version of the ANOVA table to the clipboard by selecting Copy Text from the Edit menu.
[p,table,stats] = friedman(...) returns a stats structure that you can use to perform a follow-up multiple comparison test. The friedman test evaluates the hypothesis that the column effects are all the same against the alternative that they are not all the same. Sometimes it is preferable to perform a test to determine which pairs of column effects are significantly different, and which are not. You can use the multcompare function to perform such tests by supplying the stats structure as input.


## Assumptions

Friedman's test makes the following assumptions about the data in X :

- All data come from populations having the same continuous distribution, apart from possibly different locations due to column and row effects.
- All observations are mutually independent.

The classical two-way ANOVA replaces the first assumption with the stronger assumption that data come from normal distributions.

Examples Let's repeat the example from the anova2 function, this time applying Friedman's test. Recall that the data below come from a study of popcorn brands and popper type (Hogg 1987). The columns of the matrix popcorn are brands (Gourmet, National, and Generic). The rows are popper type (Oil and Air). The study popped a batch of each brand three times with each popper. The values are the yield in cups of popped popcorn.

## friedman

```
load popcorn
popcorn
popcorn =
\begin{tabular}{lll}
5.5000 & 4.5000 & 3.5000 \\
5.5000 & 4.5000 & 4.0000 \\
6.0000 & 4.0000 & 3.0000 \\
6.5000 & 5.0000 & 4.0000 \\
7.0000 & 5.5000 & 5.0000 \\
7.0000 & 5.0000 & 4.5000
\end{tabular}
p = friedman(popcorn,3)
\(p=\)
    0.0010
```

| Friedman's ANOVA Table |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Source | 55 | df | MS | Chi-sq | Prob>Chi-sq | $\triangle$ |
| Columns <br> Interaction | $\begin{aligned} & 99.75 \\ & 0.0833 \end{aligned}$ | 2 | $\begin{gathered} 49.875 \\ 0.0417 \end{gathered}$ | 13.76 | 0.001 |  |
| Error | 16.1667 | 12 | 1.3472 |  |  |  |
| Total | 116 | 17 |  |  |  |  |

Test for column effects after row effects are removed
The small $p$ value of 0.001 indicates the popcorn brand affects the yield of popcorn. This is consistent with the results from anova2.

## References

[1] Hogg, R. V., and J. Ledolter. Engineering Statistics. New York: MacMillan, 1987.
[2] Hollander, M., and D. A. Wolfe. Nonparametric Statistical Methods. Hoboken, NJ: John Wiley \& Sons, Inc., 1999.

## See Also anova2

How To • multcompare

- kruskalwallis

Purpose $\quad F$ random numbers

Syntax $\quad$| $R$ | $=f r n d(V 1, V 2)$ |
| ---: | :--- |
| $R$ | $=f r n d(V 1, V 2, m, n, \ldots)$ |
| $R$ | $=f r n d(V 1, V 2,[m, n, \ldots])$ |

Description
$\mathrm{R}=\mathrm{frnd}(\mathrm{V} 1, \mathrm{~V} 2)$ generates random numbers from the $F$ distribution with numerator degrees of freedom V1 and denominator degrees of freedom V2. V1 and V2 can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input for V1 or V2 is expanded to a constant array with the same dimensions as the other input.
$R=f r n d(V 1, V 2, m, n, \ldots)$ or $R=f r n d(V 1, V 2,[m, n, \ldots])$ generates an $m$-by-n-by-... array containing random numbers from the $F$ distribution with parameters V1 and V2. V1 and V2 can each be scalars or arrays of the same size as $R$.

## Examples

```
n1 = frnd(1:6,1:6)
n1 =
    0.0022 0.3121 3.0528
n2 = frnd(2,2,[2 3])
n2 =
    0.3186 0.9727 3.0268
    0.2052 148.5816 0.2191
n3 = frnd([1 2 3;4 5 6],1,2,3)
n3 =
    0.6233 0.2322 31.5458
    2.5848 0.2121 4.4955
```

See Also random \| fpdf | fcdf | finv | fstat
How To . "F Distribution" on page B-25

## Purpose $\quad F$ mean and variance

## Syntax <br> $[M, V]=$ fstat $(V 1, V 2)$

Description
$[\mathrm{M}, \mathrm{V}]=\mathrm{fstat}(\mathrm{V} 1, \mathrm{~V} 2)$ returns the mean of and variance for the $F$ distribution with numerator degrees of freedom V1 and denominator degrees of freedom V2. V1 and V2 can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of M and V . A scalar input for V1 or V2 is expanded to a constant arrays with the same dimensions as the other input.

The mean of the $F$ distribution for values of $v_{2}$ greater than 2 is

$$
\frac{v_{2}}{v_{2}-2}
$$

The variance of the $F$ distribution for values of $v_{2}$ greater than 4 is

$$
\frac{2 v_{2}^{2}\left(v_{1}+v_{2}-2\right)}{v_{1}\left(v_{2}-2\right)^{2}\left(v_{2}-4\right)}
$$

The mean of the $F$ distribution is undefined if $v_{2}$ is less than 3 . The variance is undefined for $v_{2}$ less than 5.

## Examples fstat returns $N a N$ when the mean and variance are undefined.

```
[m,v] = fstat(1:5,1:5)
m =
        NaN NaN 3.0000 2.0000 1.6667
    v =
        NaN NaN NaN NaN 8.8889
```

See Also fpdf | fcdf | finv | frnd
How To . "F Distribution" on page B-25

Purpose Interactive contour plot

```
Syntax fsurfht(fun,xlims,ylims)
fsurfht(fun,xlims,ylims,p1,p2,p3,p4,p5)
```


## Description

Examples Plot the Gaussian likelihood function for the gas.mat data.

```
load gas
```

Create a function containing the following commands, and name it gauslike.m.

```
function z = gauslike(mu,sigma,p1)
n = length(p1);
z = ones(size(mu));
for i = 1:n
z = z .* (normpdf(p1(i),mu,sigma));
end
```

The gauslike function calls normpdf, treating the data sample as fixed and the parameters $\mu$ and $\sigma$ as variables. Assume that the gas prices are normally distributed, and plot the likelihood surface of the sample.

```
fsurfht('gauslike',[112 118],[3 5],price1)
```



The sample mean is the $x$ value at the maximum, but the sample standard deviation is not the $y$ value at the maximum.

```
mumax = mean(price1)
mumax =
    115.1500
sigmamax = std(price1)*sqrt(19/20)
sigmamax =
    3.7719
```


## fsurfht

Purpose Full factorial design
Syntax dFF = fullfact(levels)
Description dFF = fullfact(levels) gives factor settings dFF for a full factorialdesign with $n$ factors, where the number of levels for each factor is givenby the vector levels of length $n$. dFF is $m$-by- $n$, where $m$ is the numberof treatments in the full-factorial design. Each row of dFF correspondsto a single treatment. Each column contains the settings for a singlefactor, with integer values from one to the number of levels.
Examples The following generates an eight-run full-factorial design with twolevels in the first factor and four levels in the second factor:

```
dFF = fullfact([2 4])
dFF =
        1
        2 1
        1 2
        2
        1 3
        2 3
        1 4
        2 4
```

See Also
ff2n

Purpose Gage repeatability and reproducibility study

Syntax
gagerr(y, \{part, operator\})
gagerr(y, GROUP)
gagerr(y, part)
gagerr(..., param1, val1,param2,val2,...)
[TABLE, stats] = gagerr(...)

## Description

gagerr(y, \{part,operator\}) performs a gage repeatability and
reproducibility study on measurements in y collected by operator on part. y is a column vector containing the measurements on different parts. part and operator are categorical variables, numeric vectors, character matrices, or cell arrays of strings. The number of elements in part and operator should be the same as in $y$.
gagerr prints a table in the command window in which the decomposition of variance, standard deviation, study var ( 5.15 x standard deviation) are listed with respective percentages for different sources. Summary statistics are printed below the table giving the number of distinct categories (NDC) and the percentage of Gage R\&R of total variations (PRR).
gagerr also plots a bar graph showing the percentage of different components of variations. Gage $R \& R$, repeatability, reproducibility, and part-to-part variations are plotted as four vertical bars. Variance and study var are plotted as two groups.
To determine the capability of a measurement system using NDC, use the following guidelines:

- If NDC $>5$, the measurement system is capable.
- If $\mathrm{NDC}<2$, the measurement system is not capable.
- Otherwise, the measurement system may be acceptable.

To determine the capability of a measurement system using PRR, use the following guidelines:

- If $\mathrm{PRR}<10 \%$, the measurement system is capable.
- If PRR $>30 \%$, the measurement system is not capable.
- Otherwise, the measurement system may be acceptable.
gagerr(y, GROUP) performs a gage $R \& R$ study on measurements in $y$ with part and operator represented in GROUP. GROUP is a numeric matrix whose first and second columns specify different parts and operators, respectively. The number of rows in GROUP should be the same as the number of elements in y. (See "Grouped Data" on page 2-34.)
gagerr ( $y$, part) performs a gage $R \& R$ study on measurements in $y$ without operator information. The assumption is that all variability is contributed by part.
gagerr(..., param1,val1,param2,val2,...) performs a gage R\&R study using one or more of the following parameter name/value pairs:
- 'spec ' - A two-element vector that defines the lower and upper limit of the process, respectively. In this case, summary statistics printed in the command window include Precision-to-Tolerance Ratio (PTR). Also, the bar graph includes an additional group, the percentage of tolerance.
To determine the capability of a measurement system using PTR, use the following guidelines:
- If PTR $<0.1$, the measurement system is capable.
- If PTR $>0.3$, the measurement system is not capable.
- Otherwise, the measurement system may be acceptable.
- 'printtable' - A string with a value 'on' or 'off' that indicates whether the tabular output should be printed in the command window or not. The default value is 'on'.
- 'printgraph' - A string with a value 'on' or 'off' that indicates whether the bar graph should be plotted or not. The default value is 'on'.
- 'randomoperator' - A logical value, true or false, that indicates whether the effect of operator is random or not. The default value is true.
- 'model' - The model to use, specified by one of:
- 'linear' - Main effects only (default)
- 'interaction' - Main effects plus two-factor interactions
- 'nested' - Nest operator in part

The default value is 'linear'.
[TABLE, stats] $=$ gagerr(...) returns a 6-by-5 matrix TABLE and a structure stats. The columns of TABLE, from left to right, represent variance, percentage of variance, standard deviations, study var, and percentage of study var. The rows of TABLE, from top to bottom, represent different sources of variations: gage $R \& R$, repeatability, reproducibility, operator, operator and part interactions, and part. stats is a structure containing summary statistics for the performance of the measurement system. The fields of stats are:

- ndc - Number of distinct categories
- prr - Percentage of gage $R \& R$ of total variations
- ptr - Precision-to-tolerance ratio. The value is NaN if the parameter 'spec ' is not given.

Examples
Conduct a gage $R \& R$ study for a simulated measurement system using a mixed ANOVA model without interactions:

```
y = randn(100,1); % measurements
part = ceil(3*rand(100,1)); % parts
operator = ceil(4*rand(100,1)); % operators
gagerr(y,{part, operator},'randomoperator',true) % analysis
```

Source Variance \% Variance sigma 5.15*sigma \% 5.15*sigma

| Gage R\&R | 0.77 | 100.00 | 0.88 | 4.51 | 100.00 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Repeatability | 0.76 | 99.08 | 0.87 | 4.49 | 99.54 |
| Reproducibility | 0.01 | 0.92 | 0.08 | 0.43 | 9.61 |
| Operator | 0.01 | 0.92 | 0.08 | 0.43 | 9.61 |
| Part | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Total | 0.77 | 100.00 | 0.88 | 4.51 |  |

Number of distinct categories (NDC): 0
\% of Gage R\&R of total variations (PRR): 100.00

Note: The last column of the above table does not have to sum to $100 \%$


[^7]
## Purpose

Gamma cumulative distribution function

## Syntax

```
gamcdf(X,A,B)
[P,PLO,PUP] = gamcdf(X,A,B,pcov,alpha)
```

gamcdf $(X, A, B)$ computes the gamma cdf at each of the values in $X$ using the corresponding shape parameters in $A$ and scale parameters in B. X, A, and B can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The parameters in $A$ and $B$ must be positive.

The gamma cdf is

$$
p=F(x \mid a, b)=\frac{1}{b^{a} \Gamma(a)} \int_{0}^{x} t^{a-1} e^{\frac{-t}{b}} d t
$$

The result, $p$, is the probability that a single observation from a gamma distribution with parameters $a$ and $b$ will fall in the interval $[0 x]$.
[P,PLO,PUP] = gamcdf(X,A, B, pcov,alpha) produces confidence bounds for $P$ when the input parameters $A$ and $B$ are estimates. pcov is a 2 -by- 2 matrix containing the covariance matrix of the estimated parameters. alpha has a default value of 0.05 , and specifies $100(1$-alpha) $\%$ confidence bounds. PLO and PUP are arrays of the same size as P containing the lower and upper confidence bounds.
gammainc is the gamma distribution with $b$ fixed at 1 .

## Examples

```
a = 1:6;
b = 5:10;
prob = gamcdf(a.*b,a,b)
prob =
    0.6321 0.5940}00.5768 0.5665 0.5595 0.5543
```

The mean of the gamma distribution is the product of the parameters, $a b$. In this example, the mean approaches the median as it increases (i.e., the distribution becomes more symmetric).

See Also
cdf | gampdf | gaminv | gamstat | gamfit | gamlike | gamrnd | gamma
How To

- "Gamma Distribution" on page B-27


## Purpose

Gamma parameter estimates
Syntax
phat = gamfit(data)
[phat,pci] = gamfit(data)
[phat,pci] = gamfit(data,alpha)
[...] = gamfit(data,alpha, censoring,freq,options)

## Examples

phat $=$ gamfit(data) returns the maximum likelihood estimates (MLEs) for the parameters of the gamma distribution given the data in vector data.
[phat,pci] = gamfit(data) returns MLEs and $95 \%$ percent confidence intervals. The first row of pci is the lower bound of the confidence intervals; the last row is the upper bound.
[phat,pci] = gamfit(data,alpha) returns 100(1-alpha)\% confidence intervals. For example, alpha $=0.01$ yields $99 \%$ confidence intervals.
[...] = gamfit(data, alpha, censoring) accepts a Boolean vector of the same size as data that is 1 for observations that are right-censored and 0 for observations that are observed exactly.
[...] = gamfit(data,alpha, censoring,freq) accepts a frequency vector of the same size as data. freq typically contains integer frequencies for the corresponding elements in data, but may contain any nonnegative values.
[...] = gamfit(data,alpha, censoring,freq,options) accepts a structure, options, that specifies control parameters for the iterative algorithm the function uses to compute maximum likelihood estimates. The gamma fit function accepts an options structure which can be created using the function statset. Enter statset('gamfit') to see the names and default values of the parameters that gamfit accepts in the options structure.

Fit a gamma distribution to random data generated from a specified gamma distribution:

```
a = 2; b = 4;
data = gamrnd(a,b,100,1);
[p,ci] = gamfit(data)
p =
    2.1990 3.7426
ci =
    1.6840 2.8298
    2.7141 4.6554
```

References [1] Hahn, Gerald J., and S. S. Shapiro. Statistical Models in Engineering. Hoboken, NJ: John Wiley \& Sons, Inc., 1994, p. 88.
See Also mle | gamlike | gampdf | gamcdf | gaminv | gamstat | gamrnd
How To - "Gamma Distribution" on page B-27

## Purpose <br> Gamma inverse cumulative distribution function

Syntax
$X=\operatorname{gaminv}(P, A, B)$
[X,XLO,XUP] = gamcdf(P,A,B,pcov,alpha)

## Algorithms

Examples
$X=$ gaminv $(P, A, B)$ computes the inverse of the gamma cdf with shape parameters in $A$ and scale parameters in $B$ for the corresponding probabilities in P. P, A, and B can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The parameters in A and B must all be positive, and the values in $P$ must lie on the interval [01].

The gamma inverse function in terms of the gamma cdf is

$$
x=F^{-1}(p \mid a, b)=\{x: F(x \mid a, b)=p\}
$$

where

$$
p=F(x \mid a, b)=\frac{1}{b^{a} \Gamma(a)} \int_{0}^{x} t^{a-1} e^{\frac{-t}{b}} d t
$$

[X,XLO,XUP] = gamcdf(P,A,B,pcov,alpha) produces confidence bounds for $P$ when the input parameters $A$ and $B$ are estimates. pcov is a 2 -by- 2 matrix containing the covariance matrix of the estimated parameters. alpha has a default value of 0.05 , and specifies $100(1$-alpha) \% confidence bounds. PLO and PUP are arrays of the same size as P containing the lower and upper confidence bounds.

There is no known analytical solution to the integral equation above. gaminv uses an iterative approach (Newton's method) to converge on the solution.

This example shows the relationship between the gamma cdf and its inverse function.

```
a = 1:5;
b = 6:10;
x = gaminv(gamcdf(1:5,a,b),a,b)
x =
    1.0000 2.0000 3.0000 4.0000 5.0000
```


## See Also

icdf | gamcdf | gampdf | gamstat | gamfit | gamlike | gamrnd
How To . "Gamma Distribution" on page B-27

## Purpose Gamma negative log-likelihood

Syntax

Description

## Examples

```
nlogL = gamlike(params,data)
[nlogL,AVAR] = gamlike(params,data)
```

nlogL = gamlike(params, data) returns the negative of the gamma log-likelihood of the parameters, params, given data. params (1) =A, shape parameters, and params (2) $=B$, scale parameters.
[nlogL,AVAR] = gamlike(params,data) also returns AVAR, which is the asymptotic variance-covariance matrix of the parameter estimates when the values in params are the maximum likelihood estimates. AVAR is the inverse of Fisher's information matrix. The diagonal elements of AVAR are the asymptotic variances of their respective parameters.
[...] = gamlike(params,data, censoring) accepts a Boolean vector of the same size as data that is 1 for observations that are right-censored and 0 for observations that are observed exactly.
[...] = gamfit(params,data, censoring,freq) accepts a frequency vector of the same size as data. freq typically contains integer frequencies for the corresponding elements in data, but may contain any non-negative values.
gamlike is a utility function for maximum likelihood estimation of the gamma distribution. Since gamlike returns the negative gamma log-likelihood function, minimizing gamlike using fminsearch is the same as maximizing the likelihood.

Compute the negative log-likelihood of parameter estimates computed by the gamfit function:

```
a = 2; b = 3;
r = gamrnd(a,b,100,1);
[nlogL,AVAR] = gamlike(gamfit(r),r)
nlogL =
    267.5648
AVAR =
```

$$
\begin{array}{cl}
0.0788 & -0.1104 \\
-0.1104 & 0.1955
\end{array}
$$

See Also gamfit | gampdf | gamcdf | gaminv | gamstat | gamrnd
How To - "Gamma Distribution" on page B-27

## Purpose <br> Gamma probability density function

## Syntax <br> $Y=\operatorname{gampdf}(X, A, B)$

Description

## Examples

```
\(m u=1: 5 ;\)
y = gampdf(1,1,mu)
\(\mathrm{y}=\)
    \(\begin{array}{lllll}0.3679 & 0.3033 & 0.2388 & 0.1947 & 0.1637\end{array}\)
\(\mathrm{y} 1=\operatorname{exppdf}(1, \mathrm{mu})\)
y1 =
    \(\begin{array}{lllll}0.3679 & 0.3033 & 0.2388 & 0.1947 & 0.1637\end{array}\)
```

See Also pdf | gamcdf | gaminv | gamstat | gamfit | gamlike | gamrnd

How To<br>- "Gamma Distribution" on page B-27

Purpose Gamma random numbers
Syntax $R=\operatorname{gamrnd}(A, B)$
$R=\operatorname{gamrnd}(A, B, m, n, \ldots)$
$R=\operatorname{gamrnd}(A, B,[m, n, \ldots])$
Description
Examples
See Alsorandg | random | gampdf | gamcdf | gaminv | gamstat | gamfit |gamlike

- "Gamma Distribution" on page B-27

Purpose Gamma mean and variance

## Syntax <br> [M, V] = gamstat(A, B)

$[M, V]=\operatorname{gamstat}(A, B)$ returns the mean of and variance for the gamma distribution with shape parameters in $A$ and scale parameters in $B$. A and B can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of $M$ and $V$. A scalar input for $A$ or $B$ is expanded to a constant array with the same dimensions as the other input.

The mean of the gamma distribution with parameters $a$ and $b$ is $a b$. The variance is $a b^{2}$.

## Examples

```
[m,v] = gamstat(1:5,1:5)
m =
    1
v =
    1
[m,v] = gamstat(1:5,1./(1:5))
m =
    1 1 1 1 1 1 1 1
v =
    1.0000}00.5000 0.3333 0.2500 0.2000 
```


## See Also

gampdf | gamcdf | gaminv | gamfit | gamlike | gamrnd
How To - "Gamma Distribution" on page B-27

## Purpose Greater than or equal relation for handles

## Syntax <br> h1 >= h2

Description h1 >= h2 performs element-wise comparisons between handle arrays h 1 and h2. h 1 and h2 must be of the same dimensions unless one is a scalar. The result is a logical array of the same dimensions, where each element is an element-wise $>=$ result.

If one of h1 or h2 is scalar, scalar expansion is performed and the result will match the dimensions of the array that is not scalar.
$\mathrm{tf}=\mathrm{ge}(\mathrm{h} 1, \mathrm{~h} 2)$ stores the result in a logical array of the same dimensions.

See Also qrandstream | eq | gt | le | lt | ne

Purpose Geometric cumulative distribution function

## Syntax $\quad Y=\operatorname{geocdf}(X, P)$

Description

## Examples

See Also
How To
cdf | geopdf | geoinv | geostat | geornd | mle

- "Geometric Distribution" on page B-41
Purpose Geometric inverse cumulative distribution function
Syntax $X=\operatorname{geoinv}(Y, P)$
Description $X=$ geoinv $(Y, P)$ returns the smallest positive integer $X$ such that thegeometric cdf evaluated at $X$ is equal to or exceeds $Y$. You can think of$Y$ as the probability of observing $X$ successes in a row in independenttrials where $P$ is the probability of success in each trial.
$Y$ and $P$ can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input for P or Y is expanded to a constant array with the same dimensions as the other input. The values in $P$ and $Y$ must lie on the interval [01].


## Examples

## See Also icdf | geocdf | geopdf | geostat | geornd

How To

- "Geometric Distribution" on page B-41


## Purpose Geometric mean

## Syntax <br> $\mathrm{m}=$ geomean $(\mathrm{x})$ <br> geomean(X,dim)

## Description

$m=$ geomean $(x)$ calculates the geometric mean of a sample. For vectors, geomean $(x)$ is the geometric mean of the elements in $x$. For matrices, geomean $(X)$ is a row vector containing the geometric means of each column. For N-dimensional arrays, geomean operates along the first nonsingleton dimension of $X$.
geomean (X, dim) takes the geometric mean along the dimension dim of $X$.

The geometric mean is

$$
m=\left[\prod_{i=1}^{n} x_{i}\right]^{\frac{1}{n}}
$$

Examples The arithmetic mean is greater than or equal to the geometric mean.

```
x = exprnd(1,10,6);
geometric = geomean(x)
geometric =
    0.7466
average = mean(x)
average =
    1.3509 1.1583 0.9741 0.5319 1.0088
```


## See Also mean | median | harmmean | trimmean

How To . "Geometric Distribution" on page B-41

## Purpose

Geometric probability density function

## Syntax

Description

## Examples

$Y=\operatorname{geopdf}(X, P)$

The geometric pdf is

$$
y=f(x \mid p)=p q^{x} I_{(0,1, \ldots)}(x)
$$

where $q=1-p$.
$Y=$ geopdf $(X, P)$ computes the geometric pdf at each of the values in $X$ using the corresponding probabilities in $P$. $X$ and $P$ can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other input. The parameters in P must lie on the interval [01].

Suppose you toss a fair coin repeatedly. If the coin lands face up (heads), that is a success. What is the probability of observing exactly three tails before getting a heads?

```
p = geopdf(3,0.5)
p =
    0.0625
```

See Also pdf | geocdf | geoinv | geostat | geornd
How To . "Geometric Distribution" on page B-41

Purpose Geometric random numbers

Syntax

$R=$ geornd( $P$ )
R = geornd( $\mathrm{P}, \mathrm{m}, \mathrm{n}, \ldots$ )
$R=\operatorname{geornd}(P,[m, n, \ldots])$
$R=$ geornd $(P)$ generates geometric random numbers with probability parameter P. P can be a vector, a matrix, or a multidimensional array. The size of R is the size of P. The geometric distribution is useful when you want to model the number of successive failures preceding a success, where the probability of success in any given trial is the constant P. The parameters in P must lie in the interval [01].
$R=$ geornd( $P, m, n, \ldots$ ) or $R=$ geornd $(P,[m, n, \ldots])$ generates an $m$-by-n-by-... array containing random numbers from the geometric distribution with probability parameter P. P can be a scalar or an array of the same size as R .

```
Examples \(\quad r 1=\operatorname{geornd}\left(1 . / 2 .^{\wedge}(1: 6)\right)\)
\(r 1=\)
    \(\begin{array}{llllll}2 & 10 & 2 & 5 & 2 & 60\end{array}\)
r2 \(=\) geornd (0.01,[15])
r2 =
    \(\begin{array}{lllll}65 & 18 & 334 & 291 & 63\end{array}\)
r3 \(=\) geornd \((0.5,1,6)\)
r3 =
    \(\begin{array}{llllll}0 & 7 & 1 & 3 & 1 & 0\end{array}\)
```

See Also random | geopdf | geocdf | geoinv | geostat
How To . "Geometric Distribution" on page B-41

## Purpose

Geometric mean and variance

## Syntax

Description
$[M, V]=\operatorname{geostat}(P)$
$[M, V]=$ geostat $(P)$ returns the mean of and variance for the geometric distribution with corresponding probabilities in $P$.
The mean of the geometric distribution with parameter $p$ is $q / p$, where $q$ $=1-p$. The variance is $q / p^{2}$.

## Examples



See Also

How To
geopdf | geocdf | geoinv | geornd

- "Geometric Distribution" on page B-41

Purpose Access dataset array properties
Syntax
get (A)
$\mathrm{s}=\operatorname{get}(\mathrm{A})$
p = get(A,PropertyName)
p $=\operatorname{get}(\mathrm{A},\{$ PropertyName1,PropertyName2,...\})

## Description

get (A) displays a list of property/value pairs for the dataset array A.
$s=\operatorname{get}(A)$ returns the values in a scalar structure $s$ with field names given by the properties.
$\mathrm{p}=\operatorname{get}(\mathrm{A}$, PropertyName) returns the value of the property specified by the string PropertyName.
p $=\operatorname{get}(\mathrm{A},\{$ PropertyName1,PropertyName2, ...\}) allows multiple property names to be specified and returns their values in a cell array.

## Examples Create a dataset array from Fisher's iris data and access the information:

```
load fisheriris
NumObs = size(meas,1);
NameObs = strcat({'Obs'},num2str((1:NumObs)','%-d'));
iris = dataset({nominal(species),'species'},...
    {meas,'SL','SW','PL','PW'},...
    'ObsNames',NameObs);
get(iris)
    Description: ''
    Units: {}
    DimNames: {'Observations' 'Variables'}
    UserData: []
    ObsNames: {150x1 cell}
    VarNames: {'species' 'SL' 'SW' 'PL' 'PW'}
ON = get(iris,'ObsNames');
ON(1:3)
```

ans $=$
'Obs1'
'Obs2'
'Obs3'
See Also
set | summary

Purpose Access categorical array labels
Syntax labels = getlabels(A)
Description labels = getlabels $(A)$ returns the labels of the levels in the categorical array A as a cell array of strings labels. For ordinal A, the labels are returned in the order of the levels.

## Examples

## Example 1

Display levels in a nominal and an ordinal array:

```
standings = nominal({'Leafs','Canadiens','Bruins'});
getlabels(standings)
ans =
    'Bruins' 'Canadiens' 'Leafs'
standings = ordinal(1:3,{'Leafs','Canadiens','Bruins'});
getlabels(standings)
ans =
    'Leafs' 'Canadiens' 'Bruins'
```


## Example 2

Display age groups containing data in hospital.mat:

```
load hospital
edges = 0:10:100;
labels = strcat(num2str((0:10:90)','%d'),{'s'});
AgeGroup = ordinal(hospital.Age,labels,[],edges);
AgeGroup = droplevels(AgeGroup);
getlabels(AgeGroup)
ans =
    '20s' '30s' '40s' '50s'
```

See Also getlevels | setlabels
Purpose Get categorical array levels
Syntax S = getlevels(A)
Description $S=$ getlevels(A) returns the levels for the categorical array A. $S$ is a vector with the same type as $A$.
See Also ..... getlabels

| Purpose | Generalized extreme value cumulative distribution function |
| :---: | :---: |
| Syntax | $P=\operatorname{gevcdf}(\mathrm{X}, \mathrm{k}$, sigma, mu $)$ |
| Description | $P=\operatorname{gevcdf}(X, k$, sigma, mu) returns the cdf of the generalized extreme value (GEV) distribution with shape parameter k, scale parameter sigma, and location parameter, mu, evaluated at the values in X . The size of $P$ is the common size of the input arguments. A scalar input functions as a constant matrix of the same size as the other inputs. |
|  | Default values for $k$, sigma, and mu are 0,1 , and 0 , respectively. <br> When $\mathrm{k}<0$, the GEV is the type III extreme value distribution. When k > 0, the GEV distribution is the type II, or Frechet, extreme value distribution. If $w$ has a Weibull distribution as computed by the wblcdf function, then -w has a type III extreme value distribution and $1 / w$ has a type II extreme value distribution. In the limit as $k$ approaches 0 , the GEV is the mirror image of the type I extreme value distribution as computed by the evcdf function. |
|  | The mean of the GEV distribution is not finite when $k \geq 1$, and the variance is not finite when $k \geq 1 / 2$. The GEV distribution has positive density only for values of $X$ such that $k^{*}(X-m u) /$ sigma $>-1$. |
| References | [1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997. |
|  | [2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000. |
| See Also | cdf \| gevpdf | gevinv | gevstat | gevfit | gevlike | gevrnd |
| How To | . "Generalized Extreme Value Distribution" on page B-32 |

## Purpose

Generalized extreme value parameter estimates

## Syntax

```
parmhat = gevfit(X)
[parmhat,parmci] = gevfit(X)
[parmhat,parmci] = gevfit(X,alpha)
[...] = gevfit(X,alpha,options)
```

Description

## References

parmhat $=$ gevfit $(X)$ returns maximum likelihood estimates of the parameters for the generalized extreme value (GEV) distribution given the data in X. parmhat (1) is the shape parameter, $k$, parmhat (2) is the scale parameter, sigma, and parmhat(3) is the location parameter, mu.
[parmhat, parmci] = gevfit(X) returns $95 \%$ confidence intervals for the parameter estimates.
[parmhat, parmci] = gevfit(X,alpha) returns 100(1-alpha)\% confidence intervals for the parameter estimates.
[...] = gevfit(X,alpha,options) specifies control parameters for the iterative algorithm used to compute ML estimates. This argument can be created by a call to statset. See statset('gevfit') for parameter names and default values. Pass in [] for alpha to use the default values.

When $\mathrm{k}<0$, the GEV is the type III extreme value distribution. When k > 0, the GEV distribution is the type II, or Frechet, extreme value distribution. If $w$ has a Weibull distribution as computed by the wblfit function, then -w has a type III extreme value distribution and $1 / w$ has a type II extreme value distribution. In the limit as $k$ approaches 0 , the GEV is the mirror image of the type I extreme value distribution as computed by the evfit function.

The mean of the GEV distribution is not finite when $k \geq 1$, and the variance is not finite when $k \geq 1 / 2$. The GEV distribution is defined for $k *(X-m u) /$ sigma > -1 .
[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.

See Also<br>How To<br>mle | gevlike | gevpdf | gevcdf | gevinv | gevstat | gevrnd<br>- "Generalized Extreme Value Distribution" on page B-32

## Purpose

Syntax
Description

## References

See Also
How To

Generalized extreme value inverse cumulative distribution function
X = gevinv( $\mathrm{P}, \mathrm{k}$, sigma,mu)
$X=\operatorname{gevinv}(P, k$, sigma, mu $)$ returns the inverse cdf of the generalized extreme value (GEV) distribution with shape parameter $k$, scale parameter sigma, and location parameter mu, evaluated at the values in $P$. The size of $X$ is the common size of the input arguments. A scalar input functions as a constant matrix of the same size as the other inputs.
Default values for $k$, sigma, and mu are 0,1 , and 0 , respectively.
When $\mathrm{k}<0$, the GEV is the type III extreme value distribution. When k > 0, the GEV distribution is the type II, or Frechet, extreme value distribution. If whas a Weibull distribution as computed by the wblinv function, then -w has a type III extreme value distribution and $1 / w$ has a type II extreme value distribution. In the limit as $k$ approaches 0 , the GEV is the mirror image of the type I extreme value distribution as computed by the evinv function.
The mean of the GEV distribution is not finite when $k \geq 1$, and the variance is not finite when $k \geq 1 / 2$. The GEV distribution has positive density only for values of X such that $\mathrm{k} *(\mathrm{X}-\mathrm{mu}) /$ sigma $>-1$.
[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.
icdf | gevcdf | gevpdf | gevstat | gevfit | gevlike | gevrnd

- "Generalized Extreme Value Distribution" on page B-32

| Purpose | Generalized extreme value negative log-likelihood |
| :---: | :---: |
| Syntax | ```nlogL = gevlike(params,data) [nlogL,ACOV] = gevlike(params,data)``` |
| Description | nlogL = gevlike(params, data) returns the negative of the log-likelihood nlogL for the generalized extreme value (GEV) distribution, evaluated at parameters params. params (1) is the shape parameter, $k$, params (2) is the scale parameter, sigma, and params (3) is the location parameter, mu. |
|  | [nlogL, ACOV] = gevlike(params, data) returns the inverse of Fisher's information matrix, ACOV. If the input parameter values in params are the maximum likelihood estimates, the diagonal elements of ACOV are their asymptotic variances. ACOV is based on the observed Fisher's information, not the expected information. |
|  | When $\mathrm{k}<0$, the GEV is the type III extreme value distribution. When k > 0, the GEV distribution is the type II, or Frechet, extreme value distribution. If $w$ has a Weibull distribution as computed by the wbllike function, then -w has a type III extreme value distribution and $1 / w$ has a type II extreme value distribution. In the limit as $k$ approaches 0 , the GEV is the mirror image of the type I extreme value distribution as computed by the evlike function. |
|  | The mean of the GEV distribution is not finite when $k \geq 1$, and the variance is not finite when $k \geq 1 / 2$. The GEV distribution has positive density only for values of $X$ such that $k^{*}(X-m u) /$ sigma $>-1$. |
| References | [1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997. |
|  | [2] Kotz, S., and S. Nadarajah.Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000. |
| See Also | gevfit \| gevpdf | gevcdf | gevinv | gevstat | gevrnd |
| How To | - "Generalized Extreme Value Distribution" on page B-32 |

## Purpose

## Syntax

Description

## References

See Also
How To

Generalized extreme value probability density function
$Y=\operatorname{gevpdf}(X, k$, sigma,mu)
$Y=\operatorname{gevpdf}(X, k$, sigma, mu $)$ returns the pdf of the generalized extreme value (GEV) distribution with shape parameter $k$, scale parameter sigma, and location parameter, mu, evaluated at the values in $X$. The size of $Y$ is the common size of the input arguments. A scalar input functions as a constant matrix of the same size as the other inputs.

Default values for k , sigma, and mu are 0,1 , and 0 , respectively.
When $\mathrm{k}<0$, the GEV is the type III extreme value distribution. When k > 0, the GEV distribution is the type II, or Frechet, extreme value distribution. If $w$ has a Weibull distribution as computed by the wblpdf function, then -w has a type III extreme value distribution and $1 / w$ has a type II extreme value distribution. In the limit as $k$ approaches 0 , the GEV is the mirror image of the type I extreme value distribution as computed by the evcdf function.

The mean of the GEV distribution is not finite when $k \geq 1$, and the variance is not finite when $k \geq 1 / 2$. The GEV distribution has positive density only for values of X such that $\mathrm{k}^{*}(\mathrm{X}-\mathrm{mu}) /$ sigma $>-1$.
[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.
pdf | gevcdf | gevinv | gevstat | gevfit | gevlike | gevrnd

- "Generalized Extreme Value Distribution" on page B-32


## Purpose <br> Generalized extreme value random numbers

Syntax
$R=$ gevrnd(k,sigma,mu)
R = gevrnd(k,sigma,mu,m,n,...)
R = gevrnd(k,sigma,mu,[m,n,...])

## Description

## References

See Also
$R=$ gevrnd( $k$, sigma, mu) returns an array of random numbers chosen from the generalized extreme value (GEV) distribution with shape parameter k , scale parameter sigma, and location parameter, mu. The size of $R$ is the common size of the input arguments if all are arrays. If any parameter is a scalar, the size of $R$ is the size of the other parameters.
$R=$ gevrnd(k,sigma,mu,m,n,...) or $R=$ gevrnd(k,sigma, mu, [m, n, ...]) generates an m-by-n-by-... array containing random numbers from the GEV distribution with parameters k, sigma, and mu. The k, sigma, mu parameters can each be scalars or arrays of the same size as $R$.

When $\mathrm{k}<0$, the GEV is the type III extreme value distribution. When k > 0, the GEV distribution is the type II, or Frechet, extreme value distribution. If $w$ has a Weibull distribution as computed by the wblrnd function, then -w has a type III extreme value distribution and $1 / w$ has a type II extreme value distribution. In the limit as $k$ approaches 0 , the GEV is the mirror image of the type I extreme value distribution as computed by the evrnd function.
The mean of the GEV distribution is not finite when $k \geq 1$, and the variance is not finite when $k \geq 1 / 2$. The GEV distribution has positive density only for values of X such that $\mathrm{k}^{*}(\mathrm{X}-\mathrm{mu}) /$ sigma $>-1$.
[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.

```
random | gevpdf | gevcdf | gevinv | gevstat | gevfit | gevlike
```

How To

- "Generalized Extreme Value Distribution" on page B-32

Purpose Generalized extreme value mean and variance
Syntax $\quad[M, V]=$ gevstat $(k$, sigma, mu $)$

## References

See Also
How To
$[M, V]=$ gevstat ( $k$, sigma, mu) returns the mean of and variance for the generalized extreme value (GEV) distribution with shape parameter $k$, scale parameter sigma, and location parameter, mu. The sizes of $M$ and V are the common size of the input arguments. A scalar input functions as a constant matrix of the same size as the other inputs.
Default values for k , sigma, and mu are 0,1 , and 0 , respectively.
When $\mathrm{k}<0$, the GEV is the type III extreme value distribution. When k > 0 , the GEV distribution is the type II, or Frechet, extreme value distribution. If $w$ has a Weibull distribution as computed by the wblstat function, then -w has a type III extreme value distribution and $1 / w$ has a type II extreme value distribution. In the limit as $k$ approaches 0 , the GEV is the mirror image of the type I extreme value distribution as computed by the evstat function.

The mean of the GEV distribution is not finite when $k \geq 1$, and the variance is not finite when $k \geq 1 / 2$. The GEV distribution has positive density only for values of $X$ such that $\mathrm{k}^{*}(\mathrm{X}-\mathrm{mu}) /$ sigma $>-1$.
[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.
gevpdf | gevcdf | gevinv | gevfit | gevlike | gevrnd

- "Generalized Extreme Value Distribution" on page B-32


## Purpose Interactively add line to plot

Syntax
gline(h)
gline
hline = gline(...)
Description
gline ( h ) allows you to draw a line segment in the figure with handle $h$ by clicking the pointer at the two endpoints. A rubber-band line tracks the pointer movement.
gline with no input arguments defaults to $\mathrm{h}=\mathrm{gcf}$ and draws in the current figure.
hline $=$ gline(...) returns the handle hline to the line.

## Examples

Use gline to connect two points in a plot:

```
x = 1:10;
y = x + randn(1,10);
scatter(x,y,25,'b','*')
lsline
mu = mean(y);
hold on
plot([1 10],[mu mu],'ro')
hline = gline; % Connect circles
set(hline,'Color','r')
```



## See Also

refline | refcurve | lsline

## Purpose

Generalized linear model regression
Syntax
b = glmfit(X,y,distr)
b = glmfit(X,y,distr,param1,val1,param2,val2,...)
[b,dev] = glmfit(...)
[b,dev,stats] = glmfit(...)
$\mathrm{b}=\operatorname{glmfit}(\mathrm{X}, \mathrm{y}$, distr) returns a $p$-by- 1 vector b of coefficient estimates for a generalized linear regression of the responses in y on the predictors in X , using the distribution distr. X is an $n$-by- $p$ matrix of $p$ predictors at each of $n$ observations. distr can be any of the following strings: 'binomial', 'gamma', 'inverse gaussian', 'normal' (the default), and 'poisson'.

In most cases, $y$ is an $n$-by- 1 vector of observed responses. For the binomial distribution, y can be a binary vector indicating success or failure at each observation, or a two column matrix with the first column indicating the number of successes for each observation and the second column indicating the number of trials for each observation.

This syntax uses the canonical link (see below) to relate the distribution to the predictors.

Note By default, glmfit adds a first column of 1 s to X , corresponding to a constant term in the model. Do not enter a column of 1 s directly into $X$. You can change the default behavior of glmfit using the 'constant' parameter, below.
glmfit treats NaNs in either X or y as missing values, and ignores them. b = glmfit(X,y,distr,param1,val1,param2,val2,...) additionally allows you to specify optional parameter name/value pairs to control the model fit. Acceptable parameters are as follows:

| Parameter | Value | Description |
| :---: | :---: | :---: |
| 'link' | 'identity', default for the distribution 'normal' | $\mu=X b$ |
|  | ' $\log$ ', default for the distribution 'poisson' | $\log (\mu)=X b$ |
|  | 'logit', default for the distribution 'binomial' | $\log (\mu /(1-\mu))=X b$ |
|  | 'probit' | $\operatorname{norminv}(\mu)=X b$ |
|  | 'comploglog' | $\log (-\log (1-\mu))=X b$ |
|  | 'reciprocal' | $1 / \mu=X b$ |
|  | 'loglog', default for the distribution 'gamma' | $\log (-\log (\mu))=X b$ |
|  | p (a number), default for the distribution 'inverse gaussian' (with $p=-2$ ) | $\mu^{p}=X b$ |
|  | cell array of the form \{FL FD FI\}, containing three function handles, created using @, that define the link (FL), the derivative of the link (FD), and the inverse link (FI). | User-specified link function |


| Parameter | Value | Description |
| :--- | :--- | :--- |
| 'estdisp' | 'on' | Estimates a dispersion <br> parameter for the binomial <br> or Poisson distribution |
|  | 'off' (Default for <br> binomial or Poisson <br> distribution) | Uses the theoretical value of <br> 1.0 for those distributions |
| 'offset' | Vector | Used as an additional <br> predictor variable, but with a <br> coefficient value fixed at 1.0 |
| 'weights' | Vector of prior <br> weights, such as <br> the inverses of the <br> relative variance of <br> each observation | 'on' (default) |
| 'constant' | Includes a constant term in <br> the model. The coefficient of <br> the constant term is the first <br> element of b. |  |
|  | 'off' | Omit the constant term |

[b, dev] = glmfit(...)returns dev, the deviance of the fit at the solution vector. The deviance is a generalization of the residual sum of squares. It is possible to perform an analysis of deviance to compare several models, each a subset of the other, and to test whether the model with more terms is significantly better than the model with fewer terms.
[b,dev,stats] = glmfit(...) returns dev and stats.
stats is a structure with the following fields:

- beta - Coefficient estimates b
- dfe - Degrees of freedom for error
- s-Theoretical or estimated dispersion parameter


## glmfit

- sfit - Estimated dispersion parameter
- se - Vector of standard errors of the coefficient estimates $b$
- coeffcorr - Correlation matrix for b
- covb - Estimated covariance matrix for B
- $\mathrm{t}-t$ statistics for b
- $\mathrm{p}-p$-values for b
- resid - Vector of residuals
- residp - Vector of Pearson residuals
- residd - Vector of deviance residuals
- resida - Vector of Anscombe residuals

If you estimate a dispersion parameter for the binomial or Poisson distribution, then stats.s is set equal to stats.sfit. Also, the elements of stats.se differ by the factor stats.s from their theoretical values.

## Examples

Fit a probit regression model for y on x . Each $\mathrm{y}(\mathrm{i})$ is the number of successes in $n$ (i) trials.

```
x = [2100 2300 2500 2700 2900 3100 ...
    3300 3500 3700 3900 4100 4300]';
n = [48 42 31 34 31 21 23 23 21 16 17 21]';
y = [1 2 2 0 3 8 8 14 17 19 15 17 21]';
b = glmfit(x,[y n],'binomial','link','probit');
yfit = glmval(b, x,'probit','size', n);
plot(x, y./n,'o',x,yfit./n,'-','LineWidth',2)
```



## References

[1] Dobson, A. J. An Introduction to Generalized Linear Models. New York: Chapman \& Hall, 1990.
[2] McCullagh, P., and J. A. Nelder. Generalized Linear Models. New York: Chapman \& Hall, 1990.
[3] Collett, D. Modeling Binary Data. New York: Chapman \& Hall, 2002.

See Also
glmval | regress | regstats

Purpose Generalized linear model values

```
Syntax \(\quad\) yhat \(=\operatorname{glmval}(b, x\), link \()\)
[yhat,dylo,dyhi] = glmval(b,X,link,stats)
[...] = glmval(..., param1,val1,param2,val2,...)
```


## Description

yhat $=$ glmval $(b, x, l i n k)$ computes predicted values for the generalized linear model with link function link and predictors $X$. Distinct predictor variables should appear in different columns of X. b is a vector of coefficient estimates as returned by the glmfit function. link can be any of the strings used as values for the link parameter in the glmfit function.

Note By default, glmval adds a first column of 1 s to $X$, corresponding to a constant term in the model. Do not enter a column of 1 s directly into $X$. You can change the default behavior of glmval using the 'constant' parameter, below.
[yhat,dylo,dyhi] = glmval(b,X,link,stats) also computes 95\% confidence bounds for the predicted values. When the stats structure output of the glmfit function is specified, dylo and dyhi are also returned. dylo and dyhi define a lower confidence bound of yhat-dylo, and an upper confidence bound of yhat+dyhi. Confidence bounds are nonsimultaneous, and apply to the fitted curve, not to a new observation.
[...] = glmval(...,param1,val1,param2,val2,...) specifies optional parameter name/value pairs to control the predicted values. Acceptable parameters are:

| Parameter | Value |
| :--- | :--- |
| 'confidence ' — the confidence <br> level for the confidence bounds | A scalar between 0 and 1 |
| 'size' - the size parameter (N) <br> for a binomial model | A scalar, or a vector with one <br> value for each row of X |
| 'offset ' - used as an additional <br> predictor variable, but with a <br> coefficient value fixed at 1.0 | A vector |
| 'constant' | - 'on' — Includes a constant <br> term in the model. The <br> coefficient of the constant term <br> is the first element of b. <br> - 'off' - Omit the constant <br> term |

Fit a probit regression model for $y$ on $x$. Each $y(i)$ is the number of successes in $n(i)$ trials.

```
x = [2100 2300 2500 2700 2900 3100 ...
    3300 3500 3700 3900 4100 4300]';
n = [48 42 31 34 31 21 23 23 21 16 17 21]';
y = [1 2 2 0 3 8 8 14 17 19 15 17 21]';
b = glmfit(x,[y n],'binomial','link','probit');
yfit = glmval(b,x,'probit','size',n);
plot(x, y./n,'o',x,yfit./n,'-','LineWidth',2)
```



[^8]Purpose
Glyph plot
Syntax
glyphplot(X)
glyphplot(X,'glyph','face')
glyphplot(X,'glyph','face','features',f)
glyphplot(X,...,'grid', [rows, cols])
glyphplot(X,...,'grid',[rows, cols],'page', p)
glyphplot(X,...,'centers', C)
glyphplot,...,'centers', C,'radius', r)
glyphplot(X,...,'obslabels',labels)
glyphplot(X,...,'standardize', method)
glyphplot(X,...,prop1,val1,...)
h = glyphplot(X,...)
glyphplot(X) creates a star plot from the multivariate data in the $n$-by- $p$ matrix X . Rows of X correspond to observations, columns to variables. A star plot represents each observation as a "star" whose $i$ th spoke is proportional in length to the $i$ th coordinate of that observation. glyphplot standardizes X by shifting and scaling each column separately onto the interval [ 0,1 ] before making the plot, and centers the glyphs on a rectangular grid that is as close to square as possible. glyphplot treats NaNs in X as missing values, and does not plot the corresponding rows of X . glyphplot(X,'glyph', 'star') is a synonym for glyphplot(X).
glyphplot(X, 'glyph', 'face') creates a face plot from X. A face plot represents each observation as a "face," whose $i$ th facial feature is drawn with a characteristic proportional to the $i$ th coordinate of that observation. The features are described in "Face Features" on page 20-677Face Features.
glyphplot(X,'glyph','face','features',f) creates a face plot where the $i$ th element of the index vector $f$ defines which facial feature will represent the $i$ th column of X . f must contain integers from 0 to 17, where 0 indicate that the corresponding column of $X$ should not be plotted. See "Face Features" on page 20-677 for more information.

## glyphplot

glyphplot(X,...,'grid', [rows, cols]) organizes the glyphs into a rows-by-cols grid.
glyphplot(X, ...,'grid',[rows, cols],'page', p) organizes the glyph into one or more pages of a rows-by-cols grid, and displays the page $p$. If $p$ is a vector, glyphplot displays multiple pages in succession. If $p$ is 'all', glyphplot displays all pages. If $p$ is 'scroll', glyphplot displays a single plot with a scrollbar.
glyphplot(X,...,'centers',C) creates a plot with each glyph centered at the locations in the $n$-by- 2 matrix C .
glyphplot, ...,'centers', C, 'radius', r) creates a plot with glyphs positioned using $C$, and scale the glyphs so the largest has radius $r$.
glyphplot (X, ..., 'obslabels', labels) labels each glyph with the text in the character array or cell array of strings labels. By default, the glyphs are labelled 1:N. Use ' ' for blank labels.
glyphplot(X,...,'standardize', method) standardizes X before making the plot. Choices for method are

- 'column' - Maps each column of $X$ separately onto the interval [0,1]. This is the default.
- 'matrix' - Maps the entire matrix $X$ onto the interval $[0,1]$.
- 'PCA' - Transforms $X$ to its principal component scores, in order of decreasing eigenvalue, and maps each one onto the interval $[0,1]$.
- 'off' - No standardization. Negative values in X may make a star plot uninterpretable.
glyphplot (X, ...,prop1, val1,...) sets properties to the specified property values for all line graphics objects created by glyphplot.
$\mathrm{h}=$ glyphplot $(\mathrm{X}, \ldots)$ returns a matrix of handles to the graphics objects created by glyphplot. For a star plot, $h(:, 1)$ and $h(:, 2)$ contain handles to the line objects for each star's perimeter and spokes, respectively. For a face plot, h(: 1 ) and h(: 2 ) contain object handles
to the lines making up each face and to the pupils, respectively. $\mathrm{h}(:, 3)$ contains handles to the text objects for the labels, if present.


## Face Features

The following table describes the correspondence between the columns of the vector $f$, the value of the 'Features' input parameter, and the facial features of the glyph plot. If $X$ has fewer than 17 columns, unused features are displayed at their default value.

| Column | Facial Feature |
| :--- | :--- |
| 1 | Size of face |
| 2 | Forehead/jaw relative arc length |
| 3 | Shape of forehead |
| 4 | Shape of jaw |
| 5 | Width between eyes |
| 6 | Vertical position of eyes |
| 7 | Height of eyes |
| 8 | Width of eyes (this also affects eyebrow width) |
| 9 | Angle of eyes (this also affects eyebrow angle) |
| 10 | Vertical position of eyebrows |
| 11 | Width of eyebrows (relative to eyes) |
| 12 | Angle of eyebrows (relative to eyes) |
| 13 | Direction of pupils |
| 14 | Length of nose |
| 15 | Vertical position of mouth |
| 16 | Shape of mouth |
| 17 | Mouth arc length |

## glyphplot

## Examples

load carsmall
X = [Acceleration Displacement Horsepower MPG Weight];
glyphplot(X,'standardize','column',... 'obslabels', Model,...
'grid',[2 2],...
'page', 'scroll');

chevrolet chevelle malibu

plvmouth satellite

```
glyphplot(X,'glyph','face',...
    'obslabels',Model,...
    'grid',[2 3],...
    'page',9);
```


pontiac ventura $\mathrm{sj} \quad$ amc pacer $\mathrm{d} / \mathrm{l}$ volkswagen rabbit

datsun b-210

toyota corolla

ford pinto

## gmdistribution

## Purpose Gaussian mixture models

Description An object of the gmdistribution class defines a Gaussian mixture distribution, which is a multivariate distribution that consists of a mixture of one or more multivariate Gaussian distribution components. The number of components for a given gmdistribution object is fixed. Each multivariate Gaussian component is defined by its mean and covariance, and the mixture is defined by a vector of mixing proportions.

## Construction

To create a Gaussian mixture distribution by specifying the distribution parameters, use the gmdistribution constructor. To fit a Gaussian mixture distribution model to data, use gmdistribution.fit.
fit
gmdistribution

Methods
cdf
cluster
disp
display
fit
mahal

Gaussian mixture parameter estimates

Construct Gaussian mixture distribution

Cumulative distribution function for Gaussian mixture distribution

Construct clusters from Gaussian mixture distribution

Display Gaussian mixture distribution object

Display Gaussian mixture distribution object

Gaussian mixture parameter estimates

Mahalanobis distance to component means

| pdf | Probability density function for <br> Gaussian mixture distribution |
| :--- | :--- |
| posterior | Posterior probabilities of <br> components |
| random | Random numbers from Gaussian <br> mixture distribution |
| subsasgn | Subscripted reference for <br> Gaussian mixture distribution <br> object |
| subsref | Subscripted reference for <br> Gaussian mixture distribution <br> object |

## Properties

All objects of the class have the properties listed in the following table.

| CovType | Type of covariance matrices |
| :--- | :--- |
| DistName | Type of distribution |
| Mu | Input matrix of means MU |
| NComponents | Number $k$ of mixture components |
| NDimensions | Dimension $d$ of multivariate <br> Gaussian distributions |
| PComponents | Input vector of mixing proportions <br> true if all covariance matrices <br> are restricted to be the same |
| SharedCov | Input array of covariances |

Objects constructed with fit have the additional properties listed in the following table.

## gmdistribution

\(\left.\begin{array}{lll} \& AIC \& Akaike Information Criterion <br>

\& BIC \& Bayes Information Criterion\end{array}\right]\)| Determine if algorithm converged |
| :--- |

## Purpose

Construct Gaussian mixture distribution

## Syntax

Description

## Examples

obj = gmdistribution(mu,sigma, p) gmdistribution class defining a Gaussian mixture distribution. $k$ components. In this case, sigma(,$:, \mathrm{I})$ is the covariance of component I. component I. sigma is the pooled estimate of covariance. elements of the pooled estimate of covariance. default is equal proportions.
obj = gmdistribution(mu, sigma, $p$ ) constructs an object obj of the
mu is a $k$-by- $d$ matrix specifying the $d$-dimensional mean of each of the
sigma specifies the covariance of each component. The size of sigma is:

- $d$-by- $d$-by- $k$ if there are no restrictions on the form of the covariance.
- 1-by- $d$-by- $k$ if the covariance matrices are restricted to be diagonal, but not restricted to be same across components. In this case, sigma(:,: I) contains the diagonal elements of the covariance of
- $d$-by- $d$ matrix if the covariance matrices are restricted to be the same across components, but not restricted to be diagonal. In this case,
- 1-by- $d$ if the covariance matrices are restricted to be diagonal and the same across components. In this case, sigma contains the diagonal
$p$ is an optional 1-by- $k$ vector specifying the mixing proportions of each component. If $p$ does not sum to 1 , gmdistribution normalizes it. The

Create a gmdistribution object defining a two-component mixture of bivariate Gaussian distributions:

```
mu = [1 2;-3 -5];
sigma = cat(3,[2 0;0 .5],[1 0;0 1]);
p = ones(1,2)/2;
obj = gmdistribution(mu,sigma,p);
```



References [1] McLachlan, G., and D. Peel. Finite Mixture Models. Hoboken, NJ: John Wiley \& Sons, Inc., 2000.

See Also

fit | pdf | cdf | random | cluster | posterior | mahal

| Purpose | Add case names to plot |
| :--- | :--- |
| Syntax | gname(cases) |
|  | gname |
|  | $h=$ gname(cases, line_handle) |

## Description

## Examples

gname (cases) displays a figure window and waits for you to press a mouse button or a keyboard key. The input argument cases is a character array or a cell array of strings, in which each row of the character array or each element of the cell array contains the case name of a point. Moving the mouse over the graph displays a pair of cross-hairs. If you position the cross-hairs near a point with the mouse and click once, the graph displays the label corresponding to that point. Alternatively, you can click and drag the mouse to create a rectangle around several points. When you release the mouse button, the graph displays the labels for all points in the rectangle. Right-click a point to remove its label. When you are done labelling points, press the Enter or Escape key to stop labeling.
gname with no arguments labels each case with its case number.
cases typically contains unique case names for each point, and is a cell array of strings or a character matrix with each row representing a name. cases can also be any grouping variable, which gname converts to labels.
h = gname(cases, line_handle) returns a vector of handles to the text objects on the plot. Use the scalar line_handle to identify the correct line if there is more than one line object on the plot.

You can use gname to label plots created by the plot, scatter, gscatter, plotmatrix, and gplotmatrix functions.

This example uses the city ratings data sets to find out which cities are the best and worst for education and the arts.

```
load cities
education = ratings(:,6);
arts = ratings(:,7);
```

```
plot(education,arts,'+')
gname(names)
```



Click the point at the top of the graph to display its label, "New York."

## See Also

gtext | gscatter | gplotmatrix

## Purpose

## Syntax

Description

## References

How To

See Also

Generalized Pareto cumulative distribution function

P = gpcdf(X,K, sigma, theta)
$P=\operatorname{gpcdf}(X, K$, sigma, theta) returns the cdf of the generalized Pareto (GP) distribution with the tail index (shape) parameter K, scale parameter sigma, and threshold (location) parameter, theta, evaluated at the values in $X$. The size of $P$ is the common size of the input arguments. A scalar input functions as a constant matrix of the same size as the other inputs.

Default values for K, sigma, and theta are 0,1 , and 0 , respectively.
When $\mathrm{K}=0$ and theta $=0$, the GP is equivalent to the exponential distribution. When $K>0$ and theta $=$ sigma/K, the GP is equivalent to the Pareto distribution. The mean of the GP is not finite when $K \geq$ 1 , and the variance is not finite when $K \geq 1 / 2$. When $K \geq 0$, the GP has positive density for
$X>$ theta, or, when
$\mathrm{K}<0,0 \leq \frac{X-\theta}{\sigma} \leq-\frac{1}{K}$.
[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.
cdf | gppdf | gpinv | gpstat | gpfit | gplike | gprnd

- "Generalized Pareto Distribution" on page B-37

Purpose Generalized Pareto parameter estimates

```
Syntax parmhat = gpfit(X)
[parmhat,parmci] = gpfit(X)
[parmhat,parmci] = gpfit(X,alpha)
[...] = gpfit(X,alpha,options)
```


## Description

parmhat $=$ gpfit $(X)$ returns maximum likelihood estimates of the parameters for the two-parameter generalized Pareto (GP) distribution given the data in X. parmhat (1) is the tail index (shape) parameter, K and parmhat(2) is the scale parameter, sigma. gpfit does not fit a threshold (location) parameter.
[parmhat, parmci] = gpfit(X) returns $95 \%$ confidence intervals for the parameter estimates.
[parmhat, parmci] = gpfit(X,alpha) returns 100(1-alpha)\% confidence intervals for the parameter estimates.
[...] = gpfit(X, alpha,options) specifies control parameters for the iterative algorithm used to compute ML estimates. This argument can be created by a call to statset. See statset('gpfit') for parameter names and default values.

Other functions for the generalized Pareto, such as gpcdf allow a threshold parameter, theta. However, gpfit does not estimate theta. It is assumed to be known, and subtracted from $X$ before calling gpfit.
When $K=0$ and theta $=0$, the GP is equivalent to the exponential distribution. When $K>0$ and theta $=$ sigma/K, the GP is equivalent to the Pareto distribution. The mean of the GP is not finite when $\mathrm{K} \geq$ 1 , and the variance is not finite when $K \geq 1 / 2$. When $K \geq 0$, the GP has positive density for

X > theta, or, when

$$
0 \leq \frac{X-\theta}{\sigma} \leq-\frac{1}{K}
$$

References [1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
See Also mle | gplike | gppdf | gpcdf | gpinv | gpstat | gprnd
How To - "Generalized Pareto Distribution" on page B-37
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.

Purpose Generalized Pareto inverse cumulative distribution function
Syntax $\quad X=\operatorname{gpinv}(P, K$, sigma, theta)
Description
$X=$ gpinv( $P, K$, sigma, theta) returns the inverse cdf for a generalized Pareto (GP) distribution with tail index (shape) parameter K, scale parameter sigma, and threshold (location) parameter theta, evaluated at the values in $P$. The size of $X$ is the common size of the input arguments. A scalar input functions as a constant matrix of the same size as the other inputs.

Default values for K, sigma, and theta are 0,1 , and 0 , respectively.
When $K=0$ and theta $=0$, the GP is equivalent to the exponential distribution. When $K>0$ and theta $=$ sigma/K, the GP is equivalent to the Pareto distribution. The mean of the GP is not finite when $K \geq$ 1 , and the variance is not finite when $K \geq 1 / 2$. When $K \geq 0$, the GP has positive density for
$X>$ theta, or, when
$\mathrm{K}<0,0 \leq \frac{X-\theta}{\sigma} \leq-\frac{1}{K}$.

## References

[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.

See Also icdf | gpcdf | gppdf | gpstat | gpfit | gplike | gprnd
How To . "Generalized Pareto Distribution" on page B-37

## Purpose

Generalized Pareto negative log-likelihood
Syntax

Description

References

See Also gpfit | gppdf | gpcdf | gpinv | gpstat | gprnd
How To . "Generalized Pareto Distribution" on page B-37
nlogL = gplike(params,data)
[nlogL,ACOV] = gplike(params,data)
nlogL = gplike(params,data) returns the negative of the log-likelihood nlogL for the two-parameter generalized Pareto (GP) distribution, evaluated at parameters params. params(1) is the tail and params (3) is the threshold (location) parameter, mu. their asymptotic variances. ACOV is based on the observed Fisher's information, not the expected information. positive density for

X > theta, or, when
$\mathrm{K}<0,0 \leq \frac{X-\theta}{\sigma} \leq-\frac{1}{K}$. Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000. index (shape) parameter, K, params (2) is the scale parameter, sigma,
[nlogL,ACOV] = gplike(params,data) returns the inverse of Fisher's information matrix, ACOV. If the input parameter values in params are the maximum likelihood estimates, the diagonal elements of ACOV are

When $\mathrm{K}=0$ and theta $=0$, the GP is equivalent to the exponential distribution. When $K>0$ and theta $=$ sigma/ $K$, the GP is equivalent to the Pareto distribution. The mean of the GP is not finite when $K \geq$ 1 , and the variance is not finite when $K \geq 1 / 2$. When $K \geq 0$, the GP has
[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal

Purpose Generalized Pareto probability density function
Syntax $\quad P=\operatorname{gppdf}(X, K$, sigma, theta $)$
$P=\operatorname{gppdf}(X, K$, sigma, theta) returns the pdf of the generalized Pareto (GP) distribution with the tail index (shape) parameter K, scale parameter sigma, and threshold (location) parameter, theta, evaluated at the values in $X$. The size of $P$ is the common size of the input arguments. A scalar input functions as a constant matrix of the same size as the other inputs.

Default values for K , sigma, and theta are 0,1 , and 0 , respectively.
When $\mathrm{K}=0$ and theta $=0$, the GP is equivalent to the exponential distribution. When $K>0$ and theta $=$ sigma/K, the GP is equivalent to the Pareto distribution. The mean of the GP is not finite when $\mathrm{K} \geq$ 1 , and the variance is not finite when $K \geq 1 / 2$. When $K \geq 0$, the GP has positive density for

X > theta, or, when
$\mathrm{K}<0,0 \leq \frac{X-\theta}{\sigma} \leq-\frac{1}{K}$.

## References

[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.

## See Also <br> pdf | gpcdf | gpinv | gpstat | gpfit | gplike | gprnd

How To . "Generalized Pareto Distribution" on page B-37

| Purpose | Matrix of scatter plots by group |
| :---: | :---: |
| Syntax | ```gplotmatrix(x,y,group) gplotmatrix(x,y,group,clr,sym,siz) gplotmatrix(x,y,group, clr, sym, siz,doleg) gplotmatrix(x,y,group,clr,sym,siz,doleg,dispopt) gplotmatrix(x,y,group,clr,sym,siz,doleg,dispopt,xnam, ynam) [h,ax,bigax] = gplotmatrix(...)``` |
| Description | gplotmatrix ( $x, y$, group) creates a matrix of scatter plots. Each individual set of axes in the resulting figure contains a scatter plot of a column of $x$ against a column of $y$. All plots are grouped by the grouping variable group. (See "Grouped Data" on page 2-34.) |
|  | x and y are matrices with the same number of rows. If x has $p$ columns and $y$ has $q$ columns, the figure contains a $p$-by- $q$ matrix of scatter plots. If you omit $y$ or specify it as the empty matrix, [ ], gplotmatrix creates a square matrix of scatter plots of columns of $x$ against each other. <br> group is a grouping variable that can be a categorical variable, vector, string array, or cell array of strings. group must have the same number of rows as $x$ and $y$. Points with the same value of group are placed in the same group, and appear on the graph with the same marker and color. Alternatively, group can be a cell array containing several grouping variables (such as $\left\{\begin{array}{ll}\mathrm{g} 1 \mathrm{~g} 2 \mathrm{~g} 3\}\end{array}\right.$ ); in that case, observations are in the same group if they have common values of all grouping variables. |
|  | gplotmatrix ( $x, y$, group, clr, sym, siz) specifies the color, marker type, and size for each group. clr is a string array of colors recognized by the plot function. The default for clr is 'bgrcmyk'. sym is a string array of symbols recognized by the plot command, with the default value '. '. siz is a vector of sizes, with the default determined by the DefaultLineMarkerSize property. If you do not specify enough values for all groups, gplotmatrix cycles through the specified values as needed. |
|  | gplotmatrix ( $\mathrm{x}, \mathrm{y}$, group, clr, sym, siz, doleg) controls whether a legend is displayed on the graph (doleg is 'on', the default) or not (doleg is 'off'). |

## gplotmatrix

gplotmatrix(x,y,group, clr, sym, siz, doleg, dispopt) controls what appears along the diagonal of a plot matrix of $y$ versus x. Allowable values are 'none', to leave the diagonals blank, 'hist', to plot histograms, or 'variable', to write the variable names. gplotmatrix displays histograms along the diagonal only when there is only one variable (i.e., gplotmatrix(x,[],[],[],[],[],'hist').
gplotmatrix(x,y,group, clr,sym,siz,doleg,dispopt,xnam, ynam) specifies the names of the columns in the $x$ and $y$ arrays. These names are used to label the $x$ - and $y$-axes. xnam and ynam must be character arrays or cell arrays of strings, with one name for each column of $x$ and $y$, respectively.
[h,ax, bigax] = gplotmatrix(...) returns three arrays of handles. $h$ is an array of handles to the lines on the graphs. The array's third dimension corresponds to groups in G. ax is a matrix of handles to the axes of the individual plots. If dispopt is 'hist', ax contains one extra row of handles to invisible axes in which the histograms are plotted. bigax is a handle to big (invisible) axes framing the entire plot matrix. bigax is fixed to point to the current axes, so a subsequent title, xlabel, or ylabel command will produce labels that are centered with respect to the entire plot matrix.

## Examples

Load the cities data. The ratings array has ratings of the cities in nine categories (category names are in the array categories). group is a code whose value is 2 for the largest cities. You can make scatter plots of the first three categories against the other four, grouped by the city size code:

```
load discrim
gplotmatrix(ratings(:,1:2),ratings(:,[4 7]),group)
```

The output figure (not shown) has an array of graphs with each city group represented by a different color. The graphs are a little easier to read if you specify colors and plotting symbols, label the axes with the rating categories, and move the legend off the graphs:

```
gplotmatrix(ratings(:,1:2),ratings(:,[4 7]),group,...
```

```
'br','.o',[],'on','',categories(1:2,:),...
categories([4 7],:))
```



See Also grpstats | gscatter | plotmatrix

How To - "Grouped Data" on page 2-34

| Purpose | Generalized Pareto random numbers |
| :--- | :--- |
| Syntax | $R=\operatorname{gprnd}(K$, sigma, theta $)$ |
|  | $R=\operatorname{gprnd}(K$, sigma, theta, $m, n, \ldots)$ |
|  | $R=\operatorname{gprnd}(K$, sigma, theta, $[m, n, \ldots])$ |

## Description

## References

See Also
$R=$ gprnd(K, sigma, theta) returns an array of random numbers chosen from the generalized Pareto (GP) distribution with tail index (shape) parameter K, scale parameter sigma, and threshold (location) parameter, theta. The size of $R$ is the common size of the input arguments if all are arrays. If any parameter is a scalar, the size of $R$ is the size of the other parameters.
$R=$ gprnd(K, sigma, theta, $m, n, \ldots$ ) or $R=$ gprnd(K, sigma, theta, [m,n,...]) generates an m-by-n-by-... array. The K, sigma, theta parameters can each be scalars or arrays of the same size as $R$.

When $K=0$ and theta $=0$, the GP is equivalent to the exponential distribution. When $\mathrm{K}>0$ and theta $=$ sigma/K, the GP is equivalent to the Pareto distribution. The mean of the GP is not finite when $\mathrm{K} \geq$ 1 , and the variance is not finite when $K \geq 1 / 2$. When $K \geq 0$, the GP has positive density for

X > theta, or, when

$$
0 \leq \frac{X-\theta}{\sigma} \leq-\frac{1}{K}
$$

[1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997.
[2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.

random | gppdf | gpcdf | gpinv | gpstat | gpfit | gplike

## Purpose

Generalized Pareto mean and variance

## Syntax

[M, V] = gpstat(K, sigma, theta)
$[M, V]=$ gpstat (K, sigma, theta) returns the mean of and variance for the generalized Pareto (GP) distribution with the tail index (shape) parameter K, scale parameter sigma, and threshold (location) parameter, theta.
The default value for theta is 0 .
When $\mathrm{K}=0$ and theta $=0$, the GP is equivalent to the exponential distribution. When $K>0$ and theta $=$ sigma/K, the GP is equivalent to the Pareto distribution. The mean of the GP is not finite when $K \geq$ 1 , and the variance is not finite when $K \geq 1 / 2$. When $K \geq 0$, the GP has positive density for $X>$ theta, or when
$\mathrm{K}<0,0 \leq \frac{X-\theta}{\sigma} \leq-\frac{1}{K}$.

## References [1] Embrechts, P., C. Klüppelberg, and T. Mikosch. Modelling Extremal Events for Insurance and Finance. New York: Springer, 1997. <br> [2] Kotz, S., and S. Nadarajah. Extreme Value Distributions: Theory and Applications. London: Imperial College Press, 2000.

See Also
gppdf | gpcdf | gpinv | gpfit | gplike | gprnd

## TreeBagger.growTrees

Purpose Train additional trees and add to ensemble

```
Syntax
\(B=\) growTrees(B,ntrees)
B = growTrees(B,ntrees,'param1',val1,'param2',val2,...)
```

Description $\quad B=$ growTrees ( $B$, ntrees) grows ntrees new trees and appends them to those trees already stored in the ensemble B.

B = growTrees(B,ntrees,'param1',val1,'param2',val2,...) pecifies optional parameter name/value pairs:

| 'nprint' | Specifies that a diagnostic message showing <br> training progress should display after every <br> value training cycles (grown trees). Default is <br> no diagnostic messages. |
| :--- | :--- |
| 'options ' | A struct that specifies options that govern <br> computation when growing the ensemble of <br> decision trees. One option requests that the <br> computation of decision trees on multiple <br> bootstrap replicates uses multiple processors, <br> if the Parallel Computing Toolbox is available. <br> Two options specify the random number streams <br> to use in selecting bootstrap replicates. You can <br> create this argument with a call to statset. You <br> can retrieve values of the individual fields with a <br> call to statget. Applicable statset parameters <br> are: |

- 'UseParallel' - If 'always' and if a matlabpool of the Parallel Computing Toolbox is open, compute decision trees drawn on separate boostrap replicates in parallel. If the Parallel Computing Toolbox is not installed, or a matlabpool is not open, computation occurs in serial mode. Default is 'never ', or serial computation.
- UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'.
- Streams - A RandStream object or cell array of such objects. If you do not specify Streams, growTrees uses the default stream or streams. If you choose to specify Streams, use a single object except in the case
- You have an open MATLAB pool
- UseParallel is 'always'
- UseSubstreams is 'never'

In that case, use a cell array the same size as the MATLAB pool.

For more information on using parallel computing, see Chapter 17, "Parallel Statistics".

## See Also

## Purpose Create index vector from grouping variable

| Syntax | $[G, G N]=g r p 2 i d x(S)$ |
| :--- | :--- |
|  | $[G, G N, G L]=\operatorname{grp} 2 i d x(S)$ |

## Description

## Examples

[G,GN]=grp2idx(S) creates an index vector $G$ from the grouping variable S. S can be a categorical, numeric, or logical vector; a cell vector of strings; or a character matrix with each row representing a group label. The result G is a vector taking integer values from 1 up to the number K of distinct groups. GN is a cell array of strings representing group labels. $G N(G)$ reproduces $S$ (aside from any differences in type).

The order of GN depends on the grouping variable:

- For numeric and logical grouping variables, the order is the sorted order of S .
- For categorical grouping variables, the order is the order of getlabels(S).
- For string grouping variables, the order is the order of first appearance in $S$.
[G,GN,GL] = grp2idx(S) returns a column vector GL representing the group levels. The set of groups and their order in GL and GN are the same, except that GL has the same type as S . If S is a character matrix, $\mathrm{GL}(\mathrm{G},:$ ) reproduces S , otherwise $\mathrm{GL}(\mathrm{G})$ reproduces S .
grp2idx treats NaNs (numeric or logical), empty strings (char or cell array of strings), or <undefined> values (categorical) in S as missing values and returns NaNs in the corresponding rows of G. GN and GL don't include entries for missing values.

Load the data in hospital.mat and create a categorical grouping variable:

```
load hospital
edges = 0:10:100;
labels = strcat(num2str((0:10:90)','%d'),{'s'});
```

```
AgeGroup = ordinal(hospital.Age,labels,[],edges);
ages = hospital.Age(1:5)
ages =
    38
    4 3
    38
    4 0
    4 9
group = AgeGroup(1:5)
group =
    30s
    40s
    30s
    40s
    40s
indices = grp2idx(group)
indices =
    4
    5
    4
    5
    5
```

See Also
gscatter | grpstats | crosstab | getlabels
How To - "Grouped Data" on page 2-34

Purpose Summary statistics by group
Syntax $\quad$ means $=$ grpstats $(X)$
means = grpstats(X, group)
grpstats(X, group, alpha)
dsstats = grpstats(ds,groupvars) [A,B,...] = grpstats(X,group,whichstats)
[...] = grpstats(..., whichstats,'Param1',VAL1,'Param2',VAL2, ...)

## Description

means $=$ grpstats $(X)$ computes the mean of the entire sample without grouping, where X is a matrix of observations.
means $=$ grpstats ( X, group $)$ returns the means of each column of X by group. The array, group defines the grouping such that two elements of $X$ are in the same group if their corresponding group values are the same. (See "Grouped Data" on page 2-34.) The grouping variable group can be a categorical variable, vector, string array, or cell array of strings. It can also be a cell array containing several grouping variables (such as $\{g 1 \mathrm{~g} 2 \mathrm{~g} 3\}$ ) to group the values in X by each unique combination of grouping variable values.
grpstats ( X, group, alpha) displays a plot of the means versus index with 100(1-alpha)\% confidence intervals around each mean.
dsstats = grpstats(ds,groupvars), when ds is a dataset array, returns a dataset dsstats that contains the mean, computed by group, for variables in ds. groupvars specifies the grouping variables in ds that define the groups, and is a positive integer, a vector of positive integers, the name of a dataset variable, a cell array containing one or more dataset variable names, or a logical vector. A grouping variable may be a vector of categorical, logical, or numeric values, a character array of strings, or a cell vector of strings. dsstats contains those grouping variables, plus one variable giving the number of observations in ds for each group, as well as one variable for each of the remaining dataset variables in ds. These variables must be numeric or logical. dsstats contains one observation for each group of observations in ds.
groupvars can be [] or omitted to compute the mean of each variable across the entire dataset without grouping.
grpstats treats NaNs as missing values, and removes them.
grpstats ignores empty group names.
$[A, B, \ldots]=$ grpstats (X, group, whichstats) returns the statistics specified in whichstats. The input whichstats can be a single function handle or name, or a cell array containing multiple function handles or names. The number of outputs ( $\mathrm{A}, \mathrm{B}, \ldots$ ) must match the number function handles and names in whichstats. Acceptable names are as follows:

- 'mean' - mean
- 'sem' - standard error of the mean
- 'numel' - count, or number of non-NaN elements
- 'gname' - group name
- 'std' - standard deviation
- 'var' - variance
- 'min' - minimum
- 'max' - maximum
- 'range' - maximum - minimum
- 'meanci' - $95 \%$ confidence interval for the mean
- 'predci' - $95 \%$ prediction interval for a new observation

Each function included in whichstats must accept a column vector of data and compute a descriptive statistic for it. For example, @median and @skewness are suitable functions to apply to a numeric input. A function must return the same size output each time grpstats calls it, even if the input for some groups is empty. The function typically returns a scalar value, but may return an nvals-by- 1 column vector if the descriptive statistic is not a scalar (a confidence interval, for example). The size of each output A, B, ... is ngroups-by-ncols-by-nvals,
where ngroups is the number of groups, ncols is the number of columns in the data X , and nuals is the number of values returned by the function for data from a single group in one column of $X$. If $X$ is a vector of data, then the size of each output $\mathrm{A}, \mathrm{B}, \ldots$. is ngroups-by-nvals.

A function included in whichstats may also be written to accept a matrix of data and compute a descriptive statistic for each column. The function should return either a row vector, or an nuals-by-ncols matrix if the descriptive statistic is not a scalar.

For the case when data are contained in a numeric matrix X , a function specified in whichstats may also be written to accept a matrix of data and ompute a descriptive statistic for each column. The function should return either a row vector, or an nvals-by-ncols matrix if the descriptive statistic is not a scalar.

```
[...] =
grpstats(...,whichstats,'Param1',VAL1,'Param2',VAL2,....)
specifies additional parameter name/value pairs chosen
from the following:
```

| 'Alpha' | A value from 0 to 1 that <br> specifies the confidence level as <br> $100(1-a l p h a) \%$ for the 'meanci' |
| :--- | :--- |
| and 'predci' options. Default is |  |
| 0.05. |  |$\quad$| The names of the variables in |
| :--- |
| ds to which the functions in |
| whichstats should be applied. |
| dsstats contains one summary |
| statistic variable for each of |
| these data variables. datavars |
| is a positive integer, a vector |
| of positive integers, a variable |
| name, a cell array containing |

one or more variable names, or a logical vector.
'VarNames ' The names of the variables in dsstats. By default, grpstats uses the names from ds for the grouping variable names, and constructs names for the summary statistic variables based on the function name and the data variable names from ds.
dsstats contains ngroupvars + 1 + ndatavars*nfuns variables, where ngroupvars is the number of variables specified in groupvars, ndatavars is the number of variables specified in datavars, and nfuns is the number of summary statistics specified in whichstats.

## Examples

```
load carsmall
[m,p,g] = grpstats(Weight,Model_Year,...
    {'mean','predci','gname'})
n = length(m)
errorbar((1:n)',m,p(:,2)-m)
set(gca,'xtick',1:n,'xticklabel',g)
title('95% prediction intervals for mean weight by year')
```



## See Also <br> gscatter | grp2idx | dataset.grpstats

How To - "Grouped Data" on page 2-34
Purpose Summary statistics by group for dataset arrays
Syntax

B = grpstats(A,groupvars)
B = grpstats(A,groupvars,whichstats)
B = grpstats(A,groupvars,whichstats,...,'DataVars', vars)
B = grpstats(A,groupvars,whichstats,...,'VarNames', names)
$B$ = grpstats(A,groupvars) returns a dataset array B that contains the means, computed by group, for variables in the dataset array $A$. The optional input groupvars specifies the variables in A that define the groups. groupvars can be a positive integer, a vector of positive integers, a variable name, a cell array containing one or more variable names, or a logical vector. groupvars can also be [] or omitted to compute the means of the variables in A without grouping. Grouping variables can be vectors of categorical, logical, or numeric values, a character array of strings, or a cell vector of strings. (See "Grouped Data" on page 2-34.)
B contains the grouping variables, plus a variable giving the number of observations in A for each group, plus a variable for each of the remaining variables in $A$. $B$ contains one observation for each group of observations in $A$.
grpstats treats NaNs as missing values, and removes them.
B = grpstats(A,groupvars, whichstats) returns a dataset array B with variables for each of the statistics specified in whichstats, applied to each of the nongrouping variables in A. whichstats can be a single function handle or name, or a cell array containing multiple function handles or names. The names can be chosen from among the following:

- 'mean' - mean
- 'sem' - standard error of the mean
- 'numel' - count, or number of non-NaN elements
- 'gname' - group name
- 'std' - standard deviation


## dataset.grpstats

- 'var' - variance
- 'meanci' - $95 \%$ confidence interval for the mean
- 'predci' - $95 \%$ prediction interval for a new observation

Each function included in whichstats must accept a subset of the rows of a dataset variable, and compute column-wise descriptive statistics for it. A function should typically return a value that has one row but is otherwise the same size as its input data. For example, @median and @skewness are suitable functions to apply to a numeric dataset variable.

A summary statistic function may also return values with more than one row, provided the return values have the same number of rows each time grpstats applies the function to different subsets of data from a given dataset variable. For a dataset variable that is nobs-by-m-by-... if a summary statistic function returns values that are nvals-by-m-by-... then the corresponding summary statistic variable in $B$ is ngroups-by-m-by-...-by-nvals, where ngroups is the number of groups in $A$.

B = grpstats(A,groupvars,whichstats,...,'DataVars', vars) specifies the variables in A to which the functions in whichstats should be applied. The output dataset arrays contain one summary statistic variable for each of the specified variables. vars is a positive integer, a vector of positive integers, a variable name, a cell array containing one or more variable names, or a logical vector.

B = grpstats(A,groupvars,whichstats,...,'VarNames',names) specifies the names of the variables in B. By default, grpstats uses the names from $A$ for the grouping variables, and constructs names for the summary statistic variables based on the function name and the data variable names. The number of variables in B is ngroupvars +1 + ndatavars*nfuns, where ngroupvars is the number of variables specified in groupvars, ndatavars is the number of variables specified in vars, and nfuns is the number of summary statistics specified in whichstats.

Examples
Compute blood pressure statistics for the data in hospital.mat, by sex and smoker status:

```
load hospital
grpstats(hospital,...
    {'Sex','Smoker'},...
            {@median,@iqr},...
            'DataVars','BloodPressure')
ans =
Sex Smoker GroupCount
Female_0 Female false 40
Female_1 Female true 13
Male_0 Male false 26
Male_1 Male true 21
            median_BloodPressure
Female_0 119.5 - 79
Female_1 129 91
Male_0 119 79
Male_1 129 92
            iqr_BloodPressure
Female_0 6.5 5.5
Female_1 8 5.5
Male_0- 7 6
Male_1 10.5 4.5
```

See Also
grpstats | summary

## Purpose Scatter plot by group

```
Syntax gscatter(x,y,group)
gscatter(x,y,group,clr,sym,siz)
gscatter(x,y,group,clr,sym,siz,doleg)
gscatter(x,y,group,clr,sym,siz,doleg,xnam, ynam)
h = gscatter(...)
```


## Description

gscatter ( $x, y$, group) creates a scatter plot of $x$ and $y$, grouped by group. $x$ and $y$ are vectors of the same size. group is a grouping variable in the form of a categorical variable, vector, string array, or cell array of strings. (See "Grouped Data" on page 2-34.) Alternatively, group can be a cell array containing several grouping variables (such as $\left\{\begin{array}{ll}g 1 & \mathrm{~g} 2 \mathrm{~g} 3\end{array}\right\}$ ), in which case observations are in the same group if they have common values of all grouping variables. Points in the same group and appear on the graph with the same marker and color.
gscatter ( $\mathrm{x}, \mathrm{y}, \mathrm{group}, \mathrm{clr}$, sym, siz) specifies the color, marker type, and size for each group. clr is a string array of colors recognized by the plot function. The default for clr is 'bgrcmyk'. sym is a string array of symbols recognized by the plot command, with the default value '.'. siz is a vector of sizes, with the default determined by the 'DefaultLineMarkerSize' property. If you do not specify enough values for all groups, gscatter cycles through the specified values as needed.
gscatter( $\mathrm{x}, \mathrm{y}$, group, clr,sym,siz,doleg) controls whether a legend is displayed on the graph (doleg is 'on', the default) or not (doleg is 'off').
gscatter( $\mathrm{x}, \mathrm{y}$, group, clr, sym, siz,doleg, xnam, ynam) specifies the name to use for the $x$-axis and $y$-axis labels. If the x and y inputs are simple variable names and xnam and ynam are omitted, gscatter labels the axes with the variable names.
$\mathrm{h}=$ gscatter (...) returns an array of handles to the lines on the graph.

Examples
Load the cities data and look at the relationship between the ratings for climate (first column) and housing (second column) grouped by city size. We'll also specify the colors and plotting symbols.

```
load discrim
gscatter(ratings(:,1),ratings(:,2),group,'br','xo')
```



See Also
gplotmatrix | grpstats | scatter

## gscatter

How To

- "Grouped Data" on page 2-34


## Purpose Greater than relation for handles

## Syntax <br> h1 > h2

Description $\quad \mathrm{h} 1>\mathrm{h} 2$ performs element-wise comparisons between handle arrays h 1 and h2. h1 and h2 must be of the same dimensions unless one is a scalar. The result is a logical array of the same dimensions, where each element is an element-wise > result.

If one of h 1 or h2 is scalar, scalar expansion is performed and the result will match the dimensions of the array that is not scalar.
$\mathrm{tf}=\mathrm{gt}(\mathrm{h} 1, \mathrm{~h} 2)$ stores the result in a logical array of the same dimensions.

See Also<br>qrandstream \| eq | ge | le | lt | ne

## haltonset

Superclasses qrandset
Purpose Halton quasi-random point sets
Description haltonset is a quasi-random point set class that produces points from the Halton sequence.

## Construction

haltonset

> Construct Halton quasi-random point set

## Methods Inherited Methods

Methods in the following table are inherited from qrandset.

| disp | Display qrandset object |
| :--- | :--- |
| end | Last index in indexing expression <br> for point set |
| length | Length of point set |
| ndims | Number of dimensions in matrix |
| net | Generate quasi-random point set |
| scramble | Scramble quasi-random point set |
| size | Number of dimensions in matrix <br> subsref |
|  | Subscripted reference for <br> qrandset |

## Properties Inherited Properties

Properties in the following table are inherited from qrandset.

| Dimensions | Number of dimensions |
| :--- | :--- |
| Leap | Interval between points |


| ScrambleMethod | Settings that control scrambling |
| :--- | :--- |
| Skip | Number of initial points to omit <br> from sequence |
| Type | Name of sequence on which point <br> set $P$ is based |

Handle. To learn how this affects your use of the class, see Comparing Handle and Value Classes in the MATLAB Object-Oriented Programming documentation.

Copy
Semantics
References

See Also sobolset
How To
[1] Kocis, L., and W. J. Whiten, "Computational Investigations of Low-Discrepancy Sequences," ACM Transactions on Mathematical Software, Vol. 23, No. 2, pp. 266-294, 1997.

- "Quasi-Random Point Sets" on page 6-16


## haltonset

## Examples

## Purpose Construct Halton quasi-random point set

```
Syntax
p = haltonset(d)
p = haltonset(d,prop1,val1,prop2,val2,...)
p = haltonset(d,prop1,val1,prop2,val2,...)
```


## Description <br> Syntax

$p=$ haltonset $(d)$ constructs a d-dimensional point set $p$ of the haltonset class, with default property settings.
p = haltonset(d,prop1,val1,prop2,val2,...) specifies property name/value pairs used to construct $p$.

The object $p$ returned by haltonset encapsulates properties of a specified quasi-random sequence. The point set is finite, with a length determined by the Skip and Leap properties and by limits on the size of point set indices (maximum value of $2^{53}$ ). Values of the point set are not generated and stored in memory until you access p using net or parenthesis indexing.

Generate a 3-D Halton point set, skip the first 1000 values, and then retain every 101st point:

```
p = haltonset(3,'Skip',1e3,'Leap',1e2)
p =
    Halton point set in 3 dimensions (8.918019e+013 points)
    Properties:
        Skip : 1000
        Leap : 100
    ScrambleMethod : none
```

Use scramble to apply reverse-radix scrambling:

```
p = scramble(p,'RR2')
p =
    Halton point set in 3 dimensions (8.918019e+013 points)
    Properties:
            Skip : 1000
            Leap : 100
    ScrambleMethod : RR2
```

Use net to generate the first four points:

```
XO = net(p,4)
XO =
    0.0928 0.6950 0.0029
    0.6958 0.2958 0.8269
    0.3013 0.6497 0.4141
    0.9087 0.7883 0.2166
```

Use parenthesis indexing to generate every third point, up to the 11th point:

```
X = p(1:3:11,:)
X =
    0.0928 0.6950 0.0029
    0.9087 0.7883 0.2166
    0.3843 0.9840 0.9878
    0.6831 0.7357 0.7923
```


## References

[1] Kocis, L., and W. J. Whiten. "Computational Investigations of Low-Discrepancy Sequences." ACM Transactions on Mathematical Software. Vol. 23, No. 2, 1997, pp. 266-294.

See Also
net | scramble | sobolset

## harmmean

Purpose Harmonic mean

## Syntax <br> $\mathrm{m}=$ harmmean $(\mathrm{X})$ harmmean(X,dim)

Description
$m=$ harmmean $(X)$ calculates the harmonic mean of a sample. For vectors, harmmean $(x)$ is the harmonic mean of the elements in $x$. For matrices, harmmean $(X)$ is a row vector containing the harmonic means of each column. For $N$-dimensional arrays, harmmean operates along the first nonsingleton dimension of X .
harmmean ( $\mathrm{X}, \mathrm{dim}$ ) takes the harmonic mean along dimension dim of X .
The harmonic mean is

$$
m=\frac{n}{\sum_{i=1}^{n} \frac{1}{x_{i}}}
$$

Examples The arithmetic mean is greater than or equal to the harmonic mean.

```
x = exprnd(1,10,6);
harmonic = harmmean(x)
harmonic =
    0.3382 0.3200 0.3710}00.0540 0.4936 0.0907
average = mean(x)
average =
    1.3509
```


## See Also mean | median | geomean \| trimmean

Purpose
Syntax

Description
Bivariate histogram

```
hist3(X)
hist3(X,nbins)
hist3(X,ctrs)
hist3(X,'Edges',edges)
N = hist3(X,...)
[N,C] = hist3(X,...)
hist3(...,param1,val1,param2,val2,...)
```

hist3(X) bins the elements of the $m$-by- 2 matrix X into a 10 -by- 10 grid of equally spaced containers, and plots a histogram. Each column of $\mathbf{X}$ corresponds to one dimension in the bin grid.
hist3(X, nbins) plots a histogram using an nbins(1)-by-nbins (2) grid of bins. hist3( $X$, 'Nbins', nbins) is equivalent to hist3( $X$, nbins).
hist3 ( $\mathrm{X}, \mathrm{ctrs}$ ), where ctrs is a two-element cell array of numeric vectors with monotonically non-decreasing values, uses a 2 -D grid of bins centered on $\operatorname{ctrs}\{1\}$ in the first dimension and on $\operatorname{ctrs}\{2\}$ in the second. hist3 assigns rows of $X$ falling outside the range of that grid to the bins along the outer edges of the grid, and ignores rows of $X$ containing NaNs. hist3 ( $\mathrm{X}, \mathrm{C}^{\prime} \mathrm{Ctrs}$ ', ctrs) is equivalent to hist3(X, ctrs).
hist3( X, 'Edges', edges), where edges is a two-element cell array of numeric vectors with monotonically non-decreasing values, uses a 2-D grid of bins with edges at edges $\{1\}$ in the first dimension and at edges $\{2\}$ in the second. The $(i, j)$ th bin includes the value $X(k,:)$ if

```
edges{1}(i) <= X(k,1) < edges{1}(i+1)
edges{2}(j) <= X(k,2) < edges{2}(j+1)
```

Rows of $X$ that fall on the upper edges of the grid, edges $\{1\}$ (end) or edges $\{2\}$ (end), are counted in the ( $I, j$ )th or ( $i, J$ ) th bins, where I and $J$ are the lengths of edges $\{1\}$ and edges $\{2\}$. hist 3 does not count rows of $X$ falling outside the range of the grid. Use - Inf and Inf in edges to include all non- NaN values.
$N=$ hist3(X,...) returns a matrix containing the number of elements of $X$ that fall in each bin of the grid, and does not plot the histogram.
$[\mathrm{N}, \mathrm{C}]=$ hist3( $\mathrm{X}, \ldots$ ) returns the positions of the bin centers in a 1 -by- 2 cell array of numeric vectors, and does not plot the histogram. hist3(ax, $x, \ldots$ ) plots onto an axes with handle ax instead of the current axes. See the axes reference page for more information about handles to plots.
hist3(...,param1,val1,param2,val2,...) allows you to specify graphics parameter name/value pairs to fine-tune the plot.

## Examples

## Example 1

Make a 3-D figure using a histogram with a density plot underneath:

```
load seamount
dat = [-y,x]; % Grid corrected for negative y-values
hold on
hist3(dat) % Draw histogram in 2D
n = hist3(dat); % Extract histogram data;
        % default to 10x10 bins
n1 = n';
n1( size(n,1) + 1 ,size(n,2) + 1 ) = 0;
```

Generate grid for 2-D projected view of intensities:

```
xb = linspace(min(dat(:,1)),max(dat(:,1)),size(n,1)+1);
yb = linspace(min(dat(:,2)),max(dat(:,2)),size(n,1)+1);
```

Make a pseudocolor plot:

```
h = pcolor(xb,yb,n1);
```

Set the z-level and colormap of the displayed grid:

```
set(h, 'zdata', ones(size(n1)) * -max(max(n)))
colormap(hot) % heat map
title('Seamount: ...
```

```
Data Point Density Histogram and Intensity Map'); grid on
```

Display the default 3-D perspective view:

```
view(3);
```

Seamount: Data Point Density Histogram and Intensity Map


## Example 2

Use the car data to make a histogram on a 7 -by- 7 grid of bins.

```
load carbig
X = [MPG,Weight];
hist3(X,[7 7]);
xlabel('MPG'); ylabel('Weight');
```

Make a histogram with semi-transparent bars:

```
hist3(X,[7 7],'FaceAlpha',.65);
xlabel('MPG'); ylabel('Weight');
set(gcf,'renderer','opengl');
```



Specify bin centers, different in each direction; get back counts, but don't make the plot.

```
cnt = hist3(X, {0:10:50 2000:500:5000});
```


## Example 3

Make a histogram with bars colored according to height.

```
load carbig
X = [MPG,Weight];
hist3(X,[7 7]);
xlabel('MPG'); ylabel('Weight');
set(gcf,'renderer','opengl');
set(get(gca,'child'),'FaceColor','interp','CDataMode',...
'auto');
```



See Also accumarray | bar | bar3 | hist | histc

## Purpose Plot histogram of categorical data

Syntax

```
hist(Y)
hist(Y,X)
hist(ax,...)
N = hist(...)
[N,X] = hist(...)
```

Description
hist ( Y ) plots a histogram bar plot of the counts for each level of the categorical vector $Y$. If $Y$ is an $m$-by- $n$ categorical matrix, hist computes counts for each column of $Y$, and plots a group of $n$ bars for each categorical level.
hist $(\mathrm{Y}, \mathrm{X})$ plots bars only for the levels specified in $\mathrm{X} . \mathrm{X}$ is a categorical vector or a cell array of level names as strings.
hist ( $\mathrm{ax}, \ldots$ ) plots into the axes with handle ax instead of gca.
$N=$ hist (...) returns the counts for each categorical level. If $Y$ is a matrix, hist works down the columns of $Y$ and returns a matrix of counts with one column for each coluimn of $Y$ and one row for each cetegorical level.
$[\mathrm{N}, \mathrm{X}]=$ hist (...) returns the categorical levels corresponding to each count in $N$, or corresponding to each column of $N$ if $Y$ is a matrix.

## Examples

Create a histogram of age groups from the hospital.mat dataset:

```
load hospital
edges = 0:10:100;
labels = strcat(num2str((0:10:90)','%d'),{'s'});
AgeGroup = ordinal(hospital.Age,labels,[],edges);
AgeGroup = droplevels(AgeGroup);
hist(AgeGroup)
```



See Also
hist | categorical.levelcounts | categorical.getlevels

## Purpose <br> Histogram with normal fit

Syntax

```
histfit(data)
histfit(data,nbins)
histfit(data,nbins,dist)
h = histfit(...)
```

Description
histfit(data) plots a histogram of the values in the vector data using the number of bins equal to the square root of the number of elements in data, then superimposes a fitted normal distribution.
histfit(data, nbins) uses nbins bins for the histogram.
histfit(data, nbins, dist) plots a histogram with a density from the distribution specified by dist, one of the following strings:

- 'beta'
- 'birnbaumsaunders'
- 'exponential'
- 'extreme value' or ev'
- 'gamma'
- 'generalized extreme value' or 'gev'
- 'generalized pareto' or 'gp'
- 'inversegaussian'
- 'logistic'
- 'loglogistic'
- 'lognormal'
- 'nakagami'
- 'negative binomial' or 'nbin'
- 'normal' (default)
- 'poisson'
- 'rayleigh'
- 'rician'
- 'tlocationscale'
- 'weibull' or 'wbl'
$h=h i s t f i t(\ldots)$ returns a vector of handles $h$, where $h(1)$ is the handle to the histogram and $h(2)$ is the handle to the normal curve.


## Examples

```
r = normrnd(10,1,100,1);
histfit(r)
h = get(gca,'Children');
set(h(2),'FaceColor',[.8 . 8 1])
```



See Also
hist | normfit

## hmmdecode

Purpose<br>\section*{Syntax}<br>\section*{Description}

Hidden Markov model posterior state probabilities

PSTATES = hmmdecode(seq,TRANS, EMIS)
[PSTATES,logpseq] = hmmdecode(...)
[PSTATES,logpseq,FORWARD,BACKWARD,S] = hmmdecode(...)
hmmdecode(..., 'Symbols', SYMBOLS)
PSTATES = hmmdecode(seq, TRANS, EMIS) calculates the posterior state probabilities, PSTATES, of the sequence seq, from a hidden Markov model. The posterior state probabilities are the conditional probabilities of being at state $k$ at step $i$, given the observed sequence of symbols, sym. You specify the model by a transition probability matrix, TRANS, and an emissions probability matrix, EMIS. TRANS $(i, j)$ is the probability of transition from state $i$ to state $j$. EMIS ( $k$, seq) is the probability that symbol seq is emitted from state $k$.

PSTATES is an array with the same length as seq and one row for each state in the model. The $(i, j)$ th element of PSTATES gives the probability that the model is in state $i$ at the $j$ th step, given the sequence seq.

Note The function hmmdecode begins with the model in state 1 at step 0 , prior to the first emission. hmmdecode computes the probabilities in PSTATES based on the fact that the model begins in state 1 .
[PSTATES,logpseq] = hmmdecode(...) returns logpseq, the logarithm of the probability of sequence seq, given transition matrix TRANS and emission matrix EMIS.
[PSTATES, logpseq,FORWARD,BACKWARD,S] = hmmdecode(...) returns the forward and backward probabilities of the sequence scaled by $S$.
hmmdecode(..., 'Symbols', SYMBOLS) specifies the symbols that are emitted. SYMBOLS can be a numeric array or a cell array of the names of the symbols. The default symbols are integers 1 through $N$, where $N$ is the number of possible emissions.

## References

## Examples

See Also

## hmmestimate

Purpose
Syntax

Description

Hidden Markov model parameter estimates from emissions and states
[TRANS,EMIS] = hmmestimate(seq,states)
hmmestimate(...,'Symbols',SYMBOLS)
hmmestimate(...,'Statenames',STATENAMES)
hmmestimate(...,'Pseudoemissions', PSEUDOE)
hmmestimate(...,'Pseudotransitions', PSEUDOTR)
[TRANS,EMIS] = hmmestimate(seq,states) calculates the maximum likelihood estimate of the transition, TRANS, and emission, EMIS, probabilities of a hidden Markov model for sequence, seq, with known states, states.
hmmestimate(...,'Symbols', SYMBOLS) specifies the symbols that are emitted. SYMBOLS can be a numeric array or a cell array of the names of the symbols. The default symbols are integers 1 through N , where N is the number of possible emissions.
hmmestimate(...,'Statenames', STATENAMES) specifies the names of the states. STATENAMES can be a numeric array or a cell array of the names of the states. The default state names are 1 through $M$, where M is the number of states.
hmmestimate(...,'Pseudoemissions', PSEUDOE) specifies pseudocount emission values in the matrix PSEUDO. Use this argument to avoid zero probability estimates for emissions with very low probability that might not be represented in the sample sequence. PSEUDOE should be a matrix of size $m$-by- $n$, where $m$ is the number of states in the hidden Markov model and $n$ is the number of possible emissions. If the $i \rightarrow k$ emission does not occur in seq, you can set $\operatorname{PSEUDOE}(i, k)$ to be a positive number representing an estimate of the expected number of such emissions in the sequence seq.
hmmestimate(...,'Pseudotransitions',PSEUDOTR) specifies pseudocount transition values. You can use this argument to avoid zero probability estimates for transitions with very low probability that might not be represented in the sample sequence. PSEUDOTR should be a matrix of size $m$-by- $m$, where $m$ is the number of states in the hidden

Markov model. If the $i \rightarrow j$ transition does not occur in states, you can set PSEUDOTR ( $i, j$ ) to be a positive number representing an estimate of the expected number of such transitions in the sequence states.

## Pseudotransitions and Pseudoemissions

If the probability of a specific transition or emission is very low, the transition might never occur in the sequence states, or the emission might never occur in the sequence seq. In either case, the algorithm returns a probability of 0 for the given transition or emission in TRANS or EMIS. You can compensate for the absence of transition with the 'Pseudotransitions' and 'Pseudoemissions' arguments. The simplest way to do this is to set the corresponding entry of PSEUDO or
PSEUDOTR to 1. For example, if the transition $i \rightarrow j$ does not occur in states, set PSEUOTR $(i, j)=1$. This forces $\operatorname{TRANS}(i, j)$ to be positive.

If you have an estimate for the expected number of transitions $i \rightarrow j$ in a sequence of the same length as states, and the actual number of transitions $i \rightarrow j$ that occur in seq is substantially less than what you expect, you can set $\operatorname{PSEUOTR}(i, j)$ to the expected number. This increases the value of $\operatorname{TRANS}(i, j)$. For transitions that do occur in states with the frequency you expect, set the corresponding entry of PSEUDOTR to 0 , which does not increase the corresponding entry of TRANS.
If you do not know the sequence of states, use hmmtrain to estimate the model parameters.

## References

Examples
[1] Durbin, R., S. Eddy, A. Krogh, and G. Mitchison. Biological Sequence Analysis. Cambridge, UK: Cambridge University Press, 1998.

```
trans = [0.95,0.05; 0.10,0.90];
emis = [1/6 1/6 1/6 1/6 1/6 1/6;
    1/10 1/10 1/10 1/10 1/10 1/2];
[seq,states] = hmmgenerate(1000,trans,emis);
[estimateTR,estimateE] = hmmestimate(seq,states);
```


## hmmestimate

See Also hmmgenerate | hmmdecode | hmmviterbi | hmmtrain

## hmmgenerate

## Purpose Hidden Markov model states and emissions

Syntax

Description
[seq,states] = hmmgenerate(len,TRANS,EMIS)
hmmgenerate(...,'Symbols',SYMBOLS)
hmmgenerate(...,'Statenames', STATENAMES)
[seq, states] = hmmgenerate(len, TRANS,EMIS) takes a known Markov model, specified by transition probability matrix TRANS and emission probability matrix EMIS, and uses it to generate

- A random sequence seq of emission symbols
- A random sequence states of states

The length of both seq and states is len. $\operatorname{TRANS}(i, j)$ is the probability of transition from state $i$ to state $j$. EMIS $(k, l)$ is the probability that symbol $l$ is emitted from state $k$.

Note The function hmmgenerate begins with the model in state 1 at step 0, prior to the first emission. The model then makes a transition to state $\mathrm{i}_{1}$, with probability $T_{1 i}$, and generates an emission $a_{k_{1}}$ with probability $E_{i, k_{l} .}$. hmmgenerate returns $\mathrm{i}_{1}$ as the first entry of states, and $a_{k_{1}}$ as the first entry of seq.
hmmgenerate(...,'Symbols', SYMBOLS) specifies the symbols that are emitted. SYMBOLS can be a numeric array or a cell array of the names of the symbols. The default symbols are integers 1 through $N$, where $N$ is the number of possible emissions.
hmmgenerate(...,'Statenames', STATENAMES) specifies the names of the states. STATENAMES can be a numeric array or a cell array of the names of the states. The default state names are 1 through $M$, where $M$ is the number of states.

Since the model always begins at state 1 , whose transition probabilities are in the first row of TRANS, in the following example, the first entry of

## hmmgenerate

the output states is be 1 with probability 0.95 and 2 with probability 0.05 .

## Examples

```
trans = [0.95,0.05;
    0.10,0.90];
    emis = [1/6 1/6 1/6 1/6 1/6 1/6;
        1/10 1/10 1/10 1/10 1/10 1/2];
[seq,states] = hmmgenerate(100,trans,emis)
[seq,states] = hmmgenerate(100,trans,emis,...
    'Symbols',{'one','two','three','four','five','six'},...
    'Statenames',{'fair';'loaded'})
```

See Also hmmviterbi \| hmmdecode \| hmmestimate \| hmmtrain

Purpose
Hidden Markov model parameter estimates from emissions

## Syntax

## Description

```
[ESTTR,ESTEMIT] = hmmtrain(seq,TRGUESS,EMITGUESS)
hmmtrain(...,'Algorithm',algorithm)
hmmtrain(...,'Symbols',SYMBOLS)
hmmtrain(...,'Tolerance',tol)
hmmtrain(...,'Maxiterations',maxiter)
hmmtrain(...,'Verbose', true)
hmmtrain(...,'Pseudoemissions',PSEUDOE)
hmmtrain(...,'Pseudotransitions',PSEUDOTR)
```

[ESTTR,ESTEMIT] = hmmtrain(seq,TRGUESS,EMITGUESS) estimates the transition and emission probabilities for a hidden Markov model using the Baum-Welch algorithm. seq can be a row vector containing a single sequence, a matrix with one row per sequence, or a cell array with each cell containing a sequence. TRGUESS and EMITGUESS are initial estimates of the transition and emission probability matrices. $\operatorname{TRGUESS}(i, j)$ is the estimated probability of transition from state $i$ to state $j$. EMITGUESS ( $i, k$ ) is the estimated probability that symbol $k$ is emitted from state i.
hmmtrain(...,'Algorithm', algorithm) specifies the training algorithm. algorithm can be either 'BaumWelch' or 'Viterbi'. The default algorithm is 'BaumWelch'.
hmmtrain(...,'Symbols', SYMBOLS) specifies the symbols that are emitted. SYMBOLS can be a numeric array or a cell array of the names of the symbols. The default symbols are integers 1 through $N$, where $N$ is the number of possible emissions.
hmmtrain(..., 'Tolerance', tol) specifies the tolerance used for testing convergence of the iterative estimation process. The default tolerance is 1e-4.
hmmtrain(...,'Maxiterations', maxiter) specifies the maximum number of iterations for the estimation process. The default maximum is 100 .

## hmmtrain

hmmtrain(..., 'Verbose', true) returns the status of the algorithm at each iteration.
hmmtrain(..., 'Pseudoemissions', PSEUDOE) specifies pseudocount emission values for the Viterbi training algorithm. Use this argument to avoid zero probability estimates for emissions with very low probability that might not be represented in the sample sequence. PSEUDOE should be a matrix of size $m$-by- $n$, where $m$ is the number of states in the hidden Markov model and $n$ is the number of possible emissions. If the $i \rightarrow k$ emission does not occur in seq, you can set PSEUDOE ( $\mathrm{i}, \mathrm{k}$ ) to be a positive number representing an estimate of the expected number of such emissions in the sequence seq.
hmmtrain(..., 'Pseudotransitions', PSEUDOTR) specifies pseudocount transition values for the Viterbi training algorithm. Use this argument to avoid zero probability estimates for transitions with very low probability that might not be represented in the sample sequence. PSEUDOTR should be a matrix of size $m$-by- $m$, where $m$ is the number of states in the hidden Markov model. If the $i \rightarrow j$ transition does not occur in states, you can set PSEUDOTR ( $i, j$ ) to be a positive number representing an estimate of the expected number of such transitions in the sequence states.

If you know the states corresponding to the sequences, use hmmestimate to estimate the model parameters.

## Tolerance

The input argument 'tolerance ' controls how many steps the hmmtrain algorithm executes before the function returns an answer. The algorithm terminates when all of the following three quantities are less than the value that you specify for tolerance:

- The log likelihood that the input sequence seq is generated by the currently estimated values of the transition and emission matrices
- The change in the norm of the transition matrix, normalized by the size of the matrix
- The change in the norm of the emission matrix, normalized by the size of the matrix

The default value of 'tolerance' is .0001. Increasing the tolerance decreases the number of steps the hmmtrain algorithm executes before it terminates.

## maxiterations

The maximum number of iterations, 'maxiterations', controls the maximum number of steps the algorithm executes before it terminates. If the algorithm executes maxiter iterations before reaching the specified tolerance, the algorithm terminates and the function returns a warning. If this occurs, you can increase the value of 'maxiterations' to make the algorithm reach the desired tolerance before terminating.

## References

## Examples

See Also
[1] Durbin, R., S. Eddy, A. Krogh, and G. Mitchison. Biological Sequence Analysis. Cambridge, UK: Cambridge University Press, 1998.

```
trans = [0.95,0.05;
    0.10,0.90];
emis = [1/6, 1/6, 1/6, 1/6, 1/6, 1/6;
    1/10, 1/10, 1/10, 1/10, 1/10, 1/2];
seq1 = hmmgenerate(100,trans,emis);
seq2 = hmmgenerate(200,trans,emis);
seqs = {seq1,seq2};
[estTR,estE] = hmmtrain(seqs,trans,emis);
```

hmmgenerate | hmmdecode | hmmestimate | hmmviterbi

## hmmviterbi

| Purpose | Hidden Markov model most probable state path |
| :---: | :---: |
| Syntax | $\begin{aligned} & \text { STATES }=\text { hmmviterbi(seq, TRANS,EMIS) } \\ & \text { hmmviterbi(...,'Symbols',SYMBOLS) } \\ & \text { hmmviterbi(...,'Statenames', STATENAMES) } \end{aligned}$ |
| Description | STATES = hmmviterbi(seq,TRANS,EMIS) given a sequence, seq, calculates the most likely path through the hidden Markov model specified by transition probability matrix, TRANS, and emission probability matrix EMIS. TRANS $(i, j)$ is the probability of transition from state $i$ to state $j$. EMIS $(i, k)$ is the probability that symbol $k$ is emitted from state i. |

Note The function hmmviterbi begins with the model in state 1 at step 0 , prior to the first emission. hmmviterbi computes the most likely path based on the fact that the model begins in state 1 .
hmmviterbi(...,'Symbols',SYMBOLS) specifies the symbols that are emitted. SYMBOLS can be a numeric array or a cell array of the names of the symbols. The default symbols are integers 1 through $N$, where $N$ is the number of possible emissions.
hmmviterbi(..., 'Statenames', STATENAMES) specifies the names of the states. STATENAMES can be a numeric array or a cell array of the names of the states. The default state names are 1 through $M$, where $M$ is the number of states.

## Examples

```
trans = [0.95,0.05;
            0.10,0.90];
emis = [1/6 1/6 1/6 1/6 1/6 1/6;
    1/10 1/10 1/10 1/10 1/10 1/2];
[seq,states] = hmmgenerate(100,trans,emis);
estimatedStates = hmmviterbi(seq,trans,emis);
```

```
[seq,states] = ...
    hmmgenerate(100,trans,emis,...
                            'Statenames',{'fair';'loaded'});
estimatesStates = ...
    hmmviterbi(seq,trans,emis,...
        'Statenames',{'fair';'loaded'});
```


## References

See Also
[1] Durbin, R., S. Eddy, A. Krogh, and G. Mitchison. Biological Sequence Analysis. Cambridge, UK: Cambridge University Press, 1998.
hmmgenerate | hmmdecode | hmmestimate | hmmtrain

Purpose Horizontal concatenation for categorical arrays

```
Syntax \(\quad C=\operatorname{horzcat}(\operatorname{dim}, A, B, \ldots)\)
    \(C=\operatorname{horzcat}(A, B)\)
```

Description $\quad C=$ horzcat (dim, $A, B, \ldots$ ) horizontally concatenates the categorical arrays A, B, .. . For matrices, all inputs must have the same number of rows. For n-D arrays, all inputs must have the same sizes except in the second dimension. The set of categorical levels for C is the sorted union of the sets of levels of the inputs, as determined by their labels. $C=$ horzcat $(A, B)$ is called for the syntax [A B].

See Also<br>cat | vertcat

## Purpose Horizontal concatenation for dataset arrays

Syntax ds = horzcat(ds1, ds2, ...)
Description ds = horzcat(ds1, ds2, ...) horizontally concatenates the dataset arrays ds1, ds2, ... . You may concatenate dataset arrays that have duplicate variable names, however, the variables must contain identical data, and horzcat includes only one copy of the variable in the output dataset.

Observation names for all dataset arrays that have them must be identical except for order. horzcat concatenates by matching observation names when present, or by position for datasets that do not have observation names.

## See Also

cat | vertcat

Purpose Hougen-Watson model

## Syntax $\quad$ yhat $=$ hougen $($ beta,$x)$

Description $\quad y$ hat $=$ hougen $(b e t a, x)$ returns the predicted values of the reaction rate, yhat, as a function of the vector of parameters, beta, and the matrix of data, $X$. beta must have 5 elements and $X$ must have three columns.
hougen is a utility function for rsmdemo.
The model form is:

$$
\hat{y}=\frac{\beta_{1} x_{2}-x_{3} / \beta_{5}}{1+\beta_{2} x_{1}+\beta_{3} x_{2}+\beta_{4} x_{3}}
$$

References<br>[1] Bates, D. M., and D. G. Watts. Nonlinear Regression Analysis and Its Applications. Hoboken, NJ: John Wiley \& Sons, Inc., 1988.

## See Also <br> rsmdemo

## Purpose

Hypergeometric cumulative distribution function

## Syntax

hygecdf( $\mathrm{X}, \mathrm{M}, \mathrm{K}, \mathrm{N}$ )
Description

## Examples

Suppose you have a lot of 100 floppy disks and you know that 20 of them are defective. What is the probability of drawing zero to two defective floppies if you select 10 at random?

```
p = hygecdf(2,100,20,10)
p =
    0.6812
```

See Also cdf | hygepdf | hygeinv | hygestat | hygernd

## hygeinv

Purpose Hypergeometric inverse cumulative distribution function

## Syntax hygeinv ( $\mathrm{P}, \mathrm{M}, \mathrm{K}, \mathrm{N}$ )

Description

Examples
hygeinv $(P, M, K, N)$ returns the smallest integer $X$ such that the hypergeometric cdf evaluated at $X$ equals or exceeds $P$. You can think of $P$ as the probability of observing $X$ defective items in $N$ drawings without replacement from a group of $M$ items where $K$ are defective.

Suppose you are the Quality Assurance manager for a floppy disk manufacturer. The production line turns out floppy disks in batches of 1,000 . You want to sample 50 disks from each batch to see if they have defects. You want to accept $99 \%$ of the batches if there are no more than 10 defective disks in the batch. What is the maximum number of defective disks should you allow in your sample of 50 ?

```
x = hygeinv(0.99,1000,10,50)
x =
    3
```

What is the median number of defective floppy disks in samples of 50 disks from batches with 10 defective disks?

```
\(x=h y g e i n v(0.50,1000,10,50)\)
\(\mathrm{x}=\)
    0
```


## See Also icdf | hygecdf | hygepdf | hygestat | hygernd

## Purpose

Hypergeometric probability density function

## Syntax <br> $Y=\operatorname{hygepdf}(X, M, K, N)$

Description

## Examples

$Y=$ hygepdf( $X, M, K, N$ ) computes the hypergeometric pdf at each of the values in $X$ using the corresponding size of the population, $M$, number of items with the desired characteristic in the population, K , or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs.

The parameters in $M, K$, and $N$ must all be positive integers, with $N \leq M$. The values in $X$ must be less than or equal to all the parameter values.

The hypergeometric pdf is

$$
y=f(x \mid M, K, N)=\frac{\binom{K}{x}\binom{M-K}{N-x}}{\binom{M}{N}}
$$

The result, $y$, is the probability of drawing exactly $x$ of a possible $K$ items in $n$ drawings without replacement from a group of $M$ objects.
and number of samples drawn, $N . X, M, K$, and $N$ can be vectors, matrices,

Suppose you have a lot of 100 floppy disks and you know that 20 of them are defective. What is the probability of drawing 0 through 5 defective floppy disks if you select 10 at random?

```
p = hygepdf(0:5,100,20,10)
p =
    0.0951 0.2679
```

pdf | hygecdf | hygeinv | hygestat | hygernd

Purpose Hypergeometric random numbers

Syntax $\quad$| $R$ | $=\operatorname{hygernd}(M, K, N)$ |
| ---: | :--- |
| $R$ | $=\operatorname{hygernd}(M, K, N, m, n, \ldots)$ |
| $R$ | $=\operatorname{hygernd}(M, K, N,[m, n, \ldots])$ |

## Description

## Examples

See Also
$R=$ hygernd $(M, K, N)$ generates random numbers from the hypergeometric distribution with corresponding size of the population, M , number of items with the desired characteristic in the population, K , and number of samples drawn, N. M, K, and N can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of R. A scalar input for $M, K$, or $N$ is expanded to a constant array with the same dimensions as the other inputs.
$R=$ hygernd ( $M, K, N, m, n, \ldots$ ) or $R=$ hygernd ( $M, K, N,[m, n, \ldots]$ ) generates an $m$-by- $n$-by-... array. The $M, K, N$ parameters can each be scalars or arrays of the same size as R.

```
numbers = hygernd(1000,40,50)
numbers =
    1
```

[^9]
## Purpose

Hypergeometric mean and variance

## Syntax

[MN, V] = hygestat(M,K,N)
[MN, V] = hygestat ( $M, K, N$ ) returns the mean of and variance for the hypergeometric distribution with corresponding size of the population, M, number of items with the desired characteristic in the population, $K$, and number of samples drawn, $N$. Vector or matrix inputs for M, K, and N must have the same size, which is also the size of MN and V. A scalar input for $\mathrm{M}, \mathrm{K}$, or N is expanded to a constant matrix with the same dimensions as the other inputs.

The mean of the hypergeometric distribution with parameters $\mathrm{M}, \mathrm{K}$, and $N$ is $N K / M$, and the variance is $N K(M-K)(M-N) /\left[M^{\wedge} 2(M-1)\right]$.

## Examples

The hypergeometric distribution approaches the binomial distribution, where $p=K / M$, as $M$ goes to infinity.

```
[m,v] = hygestat(10.^(1:4),10.^(0:3),9)
m =
    \(0.9000 \quad 0.9000 \quad 0.9000 \quad 0.9000\)
v \(=\)
    \(0.0900 \quad 0.7445 \quad 0.8035 \quad 0.8094\)
[m, v] = binostat (9, 0.1)
m =
    0.9000
v \(=\)
    0.8100
```


## See Also

hygepdf | hygecdf | hygeinv | hygernd

## Purpose Inverse cumulative distribution functions

Syntax $\quad$| $Y$ | $=\operatorname{icdf}($ name $, X, A)$ |
| ---: | :--- |
| $Y$ | $=\operatorname{icdf}($ name $, X, A, B)$ |
| $Y$ | $=\operatorname{icdf}(\operatorname{name}, X, A, B, C)$ |

## Description

$Y=\operatorname{icdf}($ name $, X, A)$ computes the inverse cumulative distribution function for the one-parameter family of distributions specified by name. Parameter values for the distribution are given in A. The inverse cumulative distribution function is evaluated at the values in $X$ and its values are returned in Y .

If $X$ and $A$ are arrays, they must be the same size. If $X$ is a scalar, it is expanded to a constant matrix the same size as $A$. If A is a scalar, it is expanded to a constant matrix the same size as $X$.
$Y$ is the common size of $X$ and $A$ after any necessary scalar expansion.
$Y=\operatorname{icdf}($ name $, X, A, B)$ computes the inverse cumulative distribution function for two-parameter families of distributions, where parameter values are given in $A$ and $B$.

If $X, A$, and $B$ are arrays, they must be the same size. If $X$ is a scalar, it is expanded to a constant matrix the same size as A and B. If either A or B are scalars, they are expanded to constant matrices the same size as $X$.
$Y$ is the common size of $X, A$, and $B$ after any necessary scalar expansion.
$Y=\operatorname{icdf}($ name $, X, A, B, C)$ computes the inverse cumulative distribution function for three-parameter families of distributions, where parameter values are given in $A, B$, and $C$.

If $X, A, B$, and $C$ are arrays, they must be the same size. If $X$ is a scalar, it is expanded to a constant matrix the same size as $\mathrm{A}, \mathrm{B}$, and C . If any of A, B or C are scalars, they are expanded to constant matrices the same size as $X$.
$Y$ is the common size of $X, A, B$ and $C$ after any necessary scalar expansion.
Acceptable strings for name are:

| name | Distribution | Input Parameter A | Input Parameter B | Input <br> Parameter <br> C |
| :---: | :---: | :---: | :---: | :---: |
| 'beta' or 'Beta' | "Beta Distribution" on page B-4 | a | b | - |
| 'bino' or 'Binomial' | "Binomial Distribution" on page B-7 | n : number of trials | p : probability of success for each trial | - |
| 'chi2' or 'Chisquare' | "Chi-Square Distribution" on page B-12 | $v$ : degrees of freedom | - | - |
| $\begin{aligned} & \text { 'exp' or } \\ & \text { 'Exponential' } \end{aligned}$ | "Exponential Distribution" on page B-16 | $\mu$ : mean | - | - |
| 'ev' or 'Extreme Value' | "Extreme Value Distribution" on page B-19 | $\mu$ : location parameter | $\sigma$ : scale parameter | - |
| 'f' or 'F' | "F Distribution" on page B-25 | $v 1$ : numerator degrees of freedom | $v 2$ : denominator degrees of freedom | - |
| 'gam' or 'Gamma' | "Gamma Distribution" on page B-27 | a: shape parameter | b: scale parameter | - |
| 'gev' or 'Generalized Extreme Value' | "Generalized Extreme Value Distribution" on page B-32 | k: shape parameter | $\sigma$ : scale parameter | $\mu$ : location parameter |


| name | Distribution | Input <br> Parameter A | Input <br> Parameter B | Input <br> Parameter <br> C |
| :--- | :--- | :--- | :--- | :--- |
| 'gp' or <br> 'Generalized | "Generalized <br> Pareto <br> Distribution" <br> on page B-37 | k: tail <br> index (shape) <br> parameter | $\sigma:$ scale <br> parameter | $\mu:$ threshold <br> (location) <br> parameter |
| 'geo' or <br> 'Geometric' | "Geometric <br> Distribution" <br> on page B-41 | p: probability <br> parameter | - | - |
| 'hyge' or <br> 'Hypergeometric ' | "Hypergeometric <br> Distribution" on <br> page B-43 | M: size of the <br> population | K: number of <br> items with <br> the desired <br> characteristic in <br> the population | n: number <br> of samples <br> drawn |
| 'logn' or <br> 'Lognormal' | "Lognormal <br> Distribution" <br> on page B-51 | $\mu$ | $\sigma$ | - |
| 'nbin' or <br> 'Negative | "Negative <br> Binomial <br> Distribution" <br> Binomial' | r: number of <br> successes | p: probability <br> of success in a <br> single trial | - |
| 'ncf' or <br> 'Noncentral F' | "Noncentral F <br> Distribution" on <br> page B-78 | $v 1:$ numerator <br> degrees of <br> freedom | v2: <br> denominator <br> degrees of <br> freedom | $\delta:$ <br> noncentrality <br> parameter |
| 'nct' or <br> 'Noncentral t' | "Noncentral t <br> Distribution" on <br> page B-80 | $v:$ degrees of <br> freedom | $\delta:$ noncentrality <br> parameter | - |
| 'ncx2' or <br> 'Noncentral <br> Chi-square' | "Noncentral <br> Chi-Square <br> Distribution" <br> on page B-76 | $v:$ degrees of <br> freedom | $\delta:$ noncentrality <br> parameter | - |


| name | Distribution | Input <br> Parameter A | Input <br> Parameter B | Input <br> Parameter <br> C |
| :--- | :--- | :--- | :--- | :--- |
| 'norm' or <br> 'Normal' | "Normal <br> Distribution" <br> on page B-83 | $\mu:$ mean | $\sigma:$ standard <br> deviation | - |
| 'poiss' or <br> 'Poisson' | "Poisson <br> Distribution" <br> on page B-89 | $\lambda:$ mean | - | - |
| 'rayl' or <br> 'Rayleigh' | "Rayleigh <br> Distribution" <br> on page B-91 | b: scale <br> parameter | - | - |
| 't' or 'T' | "Student's t <br> Distribution" <br> on page B-95 | $v:$ degrees of <br> freedom | - | - |
| 'unif' or <br> 'Uniform' | "Uniform <br> Distribution <br> (Continuous)" <br> on page B-99 | a: lower <br> endpoint <br> (minimum) | b: upper <br> endpoint <br> (maximum) | - |
| 'unid' or <br> 'Discrete <br> Uniform' | "Uniform <br> Distribution <br> (Discrete)" on <br> page B-101 | N: maximum <br> observable <br> value | - | - |
| 'wbl' or <br> 'Weibull' | "Weibull <br> Distribution" <br> on page B-103 | a: scale <br> parameter | b: shape <br> parameter | - |

Examples
Compute the icdf of the normal distribution with mean 0 and standard deviation 1 at inputs $0.1,0.3, \ldots, 0.9$ :

```
x1 = icdf('Normal',0.1:0.2:0.9,0,1)
x1 =
    -1.2816 -0.5244 0
```

The order of the parameters is the same as for norminv.
Compute the icdfs of Poisson distributions with rate parameters $0,1, \ldots$, 4 at inputs $0.1,0.3, \ldots, 0.9$, respectively:

```
x2 = icdf('Poisson',0.1:0.2:0.9,0:4)
x2 =
    NaN 
```

The order of the parameters is the same as for poissinv.

```
See Also cdf | mle | pdf | random
```


## ProbDistUnivKernel.icdf

| Purpose | Return inverse cumulative distribution function (ICDF) for <br> ProbDistUnivKernel object |
| :--- | :--- |
| Syntax | $Y=$ icdf $(P D, P)$ |
| Description | $Y=$ icdf $(P D, P)$ returns $Y$, an array containing the inverse cumulative |
| distribution function (ICDF) for the ProbDistUnivKernel object $P D$, |  |
| evaluated at values in $P$. |  |

## ProbDistUnivParam.icdf

| Purpose | Return inverse cumulative distribution function (ICDF) for ProbDistUnivParam object |
| :---: | :---: |
| Syntax | $Y=\operatorname{icdf}(P D, P)$ |
| Description | $Y=\operatorname{icdf}(P D, P)$ returns $Y$, an array containing the inverse cumulative distribution function (ICDF) for the ProbDistUnivParam object PD, evaluated at values in $P$. |
| Input | PD An object of the class ProbDistUnivParam. |
| Argum | A numeric array of values from 0 to 1 where you want to evaluate the ICDF. |
| Output Arguments | An array containing the inverse cumulative distribution function (ICDF) for the ProbDistUnivParam object PD. |
| See Also | icdf |

See Also ..... icdf

```
Purpose Inverse cumulative distribution function for piecewise distribution
Syntax \(\quad X=\operatorname{icdf}(o b j, P)\)
Description \(\quad X=i c d f(o b j, P)\) returns an array \(X\) of values of the inverse cumulative
distribution function for the piecewise distribution object obj, evaluated
at the values in the array \(P\).
Examples \(\quad\) Fit Pareto tails to a \(t\) distribution at cumulative probabilities 0.1 and
0.9:
\(\mathrm{t}=\mathrm{trnd}(3,100,1)\);
obj = paretotails(t,0.1,0.9);
\([p, q]=\) boundary (obj)
\(p=\)
    0.1000
    0.9000
q \(=\)
    -1.7766
    1.8432
icdf(obj, p)
ans =
    -1.7766
    1.8432
```

See Also paretotails | cdf

## inconsistent

Purpose Inconsistency coefficient
Syntax $\quad \begin{aligned} & Y=\text { inconsistent }(Z) \\ & \\ & Y=\text { inconsistent }(Z, d)\end{aligned}$
Description $\quad Y=$ inconsistent $(Z)$ computes the inconsistency coefficient for each link of the hierarchical cluster tree $Z$, where $\mathbf{Z}$ is an ( $m-1$ )-by-3 matrix generated by the linkage function. The inconsistency coefficient characterizes each link in a cluster tree by comparing its height with the average height of other links at the same level of the hierarchy. The higher the value of this coefficient, the less similar the objects connected by the link.
$Y=$ inconsistent $(Z, d)$ computes the inconsistency coefficient for each link in the hierarchical cluster tree $Z$ to depth $d$, where $d$ is an integer denoting the number of levels of the cluster tree that are included in the calculation. By default, $\mathrm{d}=2$.

The output, Y , is an (m-1)-by-4 matrix formatted as follows.

| Column | Description |
| :--- | :--- |
| 1 | Mean of the heights of all the links included in the <br> calculation. |
| 2 | Standard deviation of the heights of all the links included <br> in the calculation. |
| 3 | Number of links included in the calculation. |
| 4 | Inconsistency coefficient. |

For each link, $k$, the inconsistency coefficient is calculated as:

$$
Y(k, 4)=(z(k, 3)-Y(k, 1)) / Y(k, 2)
$$

For leaf nodes, nodes that have no further nodes under them, the inconsistency coefficient is set to 0 .

Examples

```
    X = gallery('uniformdata',[10 2],12);
    Y = pdist(X);
    Z = linkage(Y,'single');
    dendrogram(Z)
```



| 0.1313 | 0 | 1.0000 | 0 |
| :---: | :---: | :---: | :---: |
| 0.1386 | 0 | 1.0000 | 0 |
| 0.1463 | 0.0109 | 2.0000 | 0.7071 |
| 0.2391 | 0 | 1.0000 | 0 |
| 0.1951 | 0.0568 | 4.0000 | 0.9425 |
| 0.2308 | 0.0543 | 4.0000 | 0.9320 |
| 0.2395 | 0.0748 | 4.0000 | 0.7636 |
| 0.2654 | 0.0945 | 4.0000 | 0.9203 |
| 0.3769 | 0.0950 | 3.0000 | 1.1040 |

## inconsistent

| References | [1] Jain, A., and R. Dubes. Algorithms for Clustering Data. Upper |
| :--- | :--- |
| Saddle River, NJ: Prentice-Hall, 1988. |  |$\quad$| [2] Zahn, C. T. "Graph-theoretical methods for detecting and describing |
| :--- |
| Gestalt clusters." IEEE Transactions on Computers. Vol. C-20, Issue |
| 1, 1971, pp. 68-86. |

## ProbDist.InputData property

## Purpose <br> Description

Read-only structure containing information about input data to ProbDist object

InputData is a read-only property of the ProbDist class. InputData is a structure containing information about input data to a ProbDist object. It includes the following fields:

- data
- cens
- freq


## Values

Possible values for the three fields in the structure are any data supplied to the fitdist function:

- data - Data passed to the fitdist function when creating the ProbDist object. This field is empty if the ProbDist object was created without fitting to data, that is by using the ProbDistUnivParam. ProbDistUnivParam constructor.
- cens - The vector supplied with the 'censoring' parameter when creating the ProbDist object using the fitdist function. This field is empty if the ProbDist object was created without fitting to data, that is by using the ProbDistUnivParam. ProbDistUnivParam constructor.
- freq - The vector supplied with the 'frequency ' parameter when creating the ProbDist object using the fitdist function. This field is empty if the ProbDist object was created without fitting to data, that is by using the ProbDistUnivParam. ProbDistUnivParam constructor.

Use this information to view and compare the data supplied to create distributions.

Purpose $\quad$ Convert categorical array to signed 8-bit integer array

$$
\text { Syntax } \quad B=\operatorname{int8}(A)
$$

Description $\quad B=$ int8 $(A)$ converts the categorical array $A$ to a signed 8-bit integer array. Each element of $B$ contains the internal categorical level code for the corresponding element of $A$.
Undefined elements of A are assigned the value 0 in B. If A contains more than intmax ('int8') levels, the internal codes will saturate to intmax('int8') when cast to int8.

See Also double | uint8
How To . "Integers"
Purpose Convert categorical array to signed 16-bit integer array
Syntax ..... B = int16(A)
Description $B=$ int16(A) converts the categorical array A to a signed 16-bit integerarray. Each element of B contains the internal categorical level code forthe corresponding element of A.Undefined elements of A are assigned the value 0 in B.
See Also ..... double | uint16
How To - "Integers"

Purpose Convert categorical array to signed 32-bit integer array

## Syntax <br> $B=\operatorname{int} 32(A)$

Description $\quad B=$ int32 $(A)$ converts the categorical array A to a signed 32-bit integer array. Each element of $B$ contains the internal categorical level code for the corresponding element of A.
Undefined elements of A are assigned the value 0 in B.
See Also double | uint32
How To . "Integers"
Purpose Convert categorical array to signed 64-bit integer array
Syntax B = int64(A)
Description $B=$ int64(A) converts the categorical array A to a signed 64-bit integerarray. Each element of B contains the internal categorical level code forthe corresponding element of A.Undefined elements of A are assigned the value 0 in B.
See Also ..... double | uint64
How To - "Integers"

## interactionplot

| Purpose | Interaction plot for grouped data |
| :--- | :--- |
| Syntax | interactionplot(Y,GROUP) <br> interactionplot (Y,GROUP, 'varnames ' , VARNAMES) <br> $[h, A X, b i g a x] ~$ |
|  | interactionplot $(\ldots)$ |

interactionplot(Y, GROUP) displays the two-factor interaction plot for the group means of matrix $Y$ with groups defined by entries in the cell array GROUP. $Y$ is a numeric matrix or vector. If $Y$ is a matrix, the rows represent different observations and the columns represent replications of each observation. If $Y$ is a vector, the rows give the means of each entry in the cell array GROUP. Each cell of GROUP must contain a grouping variable that can be a categorical variable, numeric vector, character matrix, or a single-column cell array of strings. (See "Grouped Data" on page 2-34.) GROUP can also be a matrix whose columns represent different grouping variables. Each grouping variable must have the same number of rows as Y . The number of grouping variables must be greater than 1.

The interaction plot is a matrix plot, with the number of rows and columns both equal to the number of grouping variables. The grouping variable names are printed on the diagonal of the plot matrix. The plot at off-diagonal position $(i, j)$ is the interaction of the two variables whose names are given at row diagonal $(i, i)$ and column diagonal $(j, j)$, respectively.
interactionplot(Y, GROUP, 'varnames', VARNAMES) displays the interaction plot with user-specified grouping variable names VARNAMES. VARNAMES is a character matrix or a cell array of strings, one per grouping variable. Default names are 'X1', 'X2', ... .
[h,AX,bigax] = interactionplot(...) returns a handle $h$ to the figure window, a matrix AX of handles to the subplot axes, and a handle bigax to the big (invisible) axes framing the subplots.

[^10]```
y = randn(1000,1); % response
group = ceil(3*rand(1000,4)); % four 3-level factors
interactionplot(y,group,'varnames',{'A','B','C','D'})
```



See Also
maineffectsplot | multivarichart
How To - "Grouped Data" on page 2-34

Purpose Set intersection for categorical arrays

## Syntax <br> $C=$ intersect $(A, B)$

Description $C=\operatorname{intersect}(A, B)$ when $A$ and $B$ are categorical arrays returns a categorical vector $C$ containing the values common to both $A$ and $B$. The result $C$ is sorted. The set of categorical levels for $C$ is the sorted union of the sets of levels of the inputs, as determined by their labels.
$[C, I A, I B]=\operatorname{UNION}(A, B)$ also returns index vectors IA and IB such that $C=A(I A)$ and $C=B(I B)$.

See Also ismember | setdiff | setxor | union | unique

## Purpose Inverse prediction

Syntax
XO = invpred(X,Y,YO)
[XO,DXLO,DXUP] = invpred(X,Y,YO)
[XO,DXLO, DXUP] = invpred(X,Y,YO,name1,val1,name2,val2,...)

## Description

$X 0=$ invpred $(X, Y, Y O)$ accepts vectors $X$ and $Y$ of the same length, fits a simple regression, and returns the estimated value $X 0$ for which the height of the line is equal to YO . The output, XO , has the same size as Y 0 , and $Y 0$ can be an array of any size.
[XO,DXLO,DXUP] = invpred (X,Y,YO) also computes $95 \%$ inverse prediction intervals. DXLO and DXUP define intervals with lower bound XO DXLO and upper bound XO+DXUP. Both DXLO and DXUP have the same size as YO.

The intervals are not simultaneous and are not necessarily finite. Some intervals may extend from a finite value to - Inf or +Inf, and some may extend over the entire real line.
[XO,DXLO,DXUP] = invpred(X,Y,YO,name1,val1,name2,val2,...) specifies optional argument name/value pairs chosen from the following list. Argument names are case insensitive and partial matches are allowed.

| Name | Value |
| :--- | :--- |
| 'alpha' | A value between 0 and 1 specifying a <br> confidence level of 100*(1-alpha)\%. Default <br> is alpha=0.05 for 95\% confidence. |
| 'predopt' | Either 'observation ', the default value to <br> compute the intervals for X0 at which a new <br> observation could equal Y0, or 'curve' to <br> compute intervals for the XO value at which <br> the curve is equal to YO. |

Examples $\quad$| $x=4^{*} \operatorname{rand}(25,1) ;$ |
| :--- |
| $y=10+5^{*} x+\operatorname{randn}(\operatorname{size}(x)) ;$ |
|  |
| scatter $(x, y)$ |
|  |
| $x 0=\operatorname{invpred}(x, y, 20)$ |

See Also polyfit | polyval | polyconf | polytool

# Purpose Inverse permute dimensions of categorical array 

$$
\text { Syntax } \quad A=\text { ipermute }(B, \text { order })
$$

Description
$A=$ ipermute $(B$, order $)$ is the inverse of permute. ipermute rearranges the dimensions of the categorical array B so that permute (A, order) will produce B. The array produced has the same values of A but the order of the subscripts needed to access any particular element are rearranged as specified by order. The elements of order must be a rearrangement of the numbers from 1 to n .

See Also permute

Purpose Interquartile range
Syntax

$\mathrm{y}=\mathrm{iqr}(\mathrm{X})$

iqr(X, dim)
Description
Tips
ExamplesThis Monte Carlo simulation shows the relative efficiency of the IQR tothe sample standard deviation for normal data.

```
x = normrnd(0,1,100,100);
s = std(x);
s_IQR = 0.7413*iqr(x);
efficiency = (norm(s-1)./norm(s_IQR-1)).^2
efficiency =
    0.3297
```

See Also std \| mad \| range

## ProbDistUnivKernel.iqr

## Purpose

Return interquartile range (IQR) for ProbDistUnivKernel object

## Syntax

Description
$Y=$ iqr ( $P D$ ) returns $Y$, the interquartile range for the ProbDistUnivKernel object PD. The interquartile range is the distance between the 75 th and 25 th percentiles.

## Input <br> Arguments <br> Output $Y$ Arguments

See Also iqr | ProbDistUnivKernel.icdf

## ProbDistUnivParam.iqr

Purpose Return interquartile range (IQR) for ProbDistUnivParam object
Syntax
$Y=i q r(P D)$

Description
$Y=$ iqr ( $P D$ ) returns $Y$, the interquartile range for the ProbDistUnivParam object $P D$. The interquartile range is the distance between the 75 th and 25 th percentiles.

## Input <br> $P D$ <br> Arguments <br> Output $\quad$ Arguments

An object of the class ProbDistUnivParam.

The value of the interquartile range for the ProbDistUnivParam object PD.

See Also iqr | ProbDistUnivParam.icdf

## Purpose Test node for branch

```
Syntax
ib = isbranch(t)
ib = isbranch(t,nodes)
```

Description $\quad i b=i s b r a n c h(t)$ returns an $n$-element logical vector $i b$ that is true for each branch node and false for each leaf node.
ib = isbranch(t, nodes) takes a vector nodes of node numbers and returns a vector of logical values for the specified nodes.

Examples Create a classification tree for Fisher's iris data:

```
load fisheriris;
t = classregtree(meas,species,...
    'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
```

view(t)

## classregtree.isbranch

|  |  |  |  |  |  |  |  | - |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Click to display: | Identity | $\checkmark$ | Magnification: | 100\% | $\square$ | Pruning level: | 0 of 4 | - |



[^11]References<br>[1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

## See Also

classregtree | cutvar | numnodes

## categorical.isempty

Purpose True for empty categorical array

## Syntax $\quad$ TF $=$ isempty $(A)$

Description TF = isempty (A) returns true (1) if A is an empty categorical array and false (0) otherwise. An empty array has no elements, that is numel $(A)=0$.

See Also numel | size
Purpose True for empty dataset array
Syntax tf = isempty(A)
Description tf = isempty (A) returns true (1) if A is an empty dataset and false ( 0 ) otherwise. An empty array has no elements, that is $\operatorname{prod}(\operatorname{size}(A))==0$.
See Also ..... size

Purpose True if categorical arrays are equal
Syntax $\quad \begin{aligned} T F & =\operatorname{isequal}(A, B) \\ T F & =\operatorname{isequal}(A, B, C, \ldots)\end{aligned}$
Description
$T F=$ isequal $(A, B)$ is true (1) if the categorical arrays $A$ and $B$ are the same class, have the same size and the same sets of levels, and contain the same values, and false (0) otherwise.

TF = isequal ( $A, B, C, \ldots$ ) is true (1) if all the input arguments are equal.

Elements with undefined levels are not considered equal to each other.

## See Also

getlabels

## Purpose Test for levels

> Syntax

Description
$I=$ islevel(levels, A) returns a logical array $I$ the same size as the string, cell array of strings, or $2-\mathrm{D}$ character matrix levels. I is true (1) where the corresponding element of levels is the label of a level in the categorical array A, even if the level contains no elements. $I$ is false (0) otherwise.

Examples Display age levels in the data in hospitl.mat, before and after dropping occupied levels:

```
load hospital
edges = 0:10:100;
labels = strcat(num2str((0:10:90)','%d'),{'s'});
disp(labels')
    'Os' '10s' '20s' '30s' '40s' '50s' '60s' '70s' '80s' '90s'
AgeGroup = ordinal(hospital.Age,labels,[],edges);
I = islevel(labels,AgeGroup);
disp(I')
    1
AgeGroup = droplevels(AgeGroup);
I = islevel(labels,AgeGroup);
disp(I')
    0
```


## See Also

Purpose Test for membership
$\begin{array}{ll}\text { Syntax } & I=\text { ismember }(A, \text { levels }) \\ {[I, I D X]=i s m e m b e r(A, l e v e l s)}\end{array}$
Description

Examples

## Example 1

$I=$ ismember (A, levels) returns a logical array $I$ the same size as the categorical array A. I is true (1) where the corresponding element of A is one of the levels specified by the labels in the categorical array, cell array of strings, or 2-D character array levels. I is false ( 0 ) otherwise.
[I, IDX] = ismember(A, levels) also returns an array of indices IDX containing the highest absolute index in levels for each element in A whose level is a member of levels, and 0 if there is no such index.

For nominal data:

```
load hospital
sex = hospital.Sex; % Nominal
smokers = hospital.Smoker; % Logical
I = ismember(sex(smokers),'Female');
I (1:5)
ans =
    0
    1
    0
    0
    0
```

The use of ismember above is equivalent to:

```
I = (sex(smokers) == 'Female');
```


## Example 2

For ordinal data:
load hospital
edges = 0:10:100;

```
labels = strcat(num2str((0:10:90)','%d'),{'s'});
AgeGroup = ordinal(hospital.Age,labels,[],edges);
I = ismember(AgeGroup(1:5),{'20s','30s'})
I =
    1
    0
    1
    0
    0
```

See Also islevel | isundefined

Purpose True for elements of categorical array in set

```
Syntax TF = ismember(A,levels)
[TF,LOC] = ismember(A,levels)
```

Description $\quad T F=$ ismember ( A, levels) returns a logical array the same size as the categorical array A, containing true (1) where the level of the corresponding element of $A$ is equal to one of the levels specified in levels, and false (0) otherwise. levels is a categorical array, or a cell array of strings or 2-D character array containing level labels.
[TF,LOC] = ismember(A,levels) also returns an index array LOC containing the highest absolute index in levels for each element in A whose level is a member of levels, and 0 if there is no such index.

## See Also

intersect | islevel | setdiff | setxor | union | unique
Purpose True if categorical array is scalar
Syntax TF = isscalar(A)
Description TF = isscalar(A) returns true (1) if the categorical array A is a 1-by-1 matrix, and false (0) otherwise.
See Also ..... isempty | isvector | size

Purpose Test for undefined elements
Syntax I = isundefined (A)
Description I = isundefined (A) returns a logical array I the same size as the categorical array A. I is true (1) where the corresponding element of A is not assigned to any level. I is false (0) where the corresponding element of A is assigned to a level.

Examples Create and display undefined levels in an ordinal array:

```
A = ordinal([ll 2 3 2 1],{'lo','med','hi'})
A =
    lo med hi med lo
A = droplevels(A,{'med','hi'})
Warning: OLDLEVELS contains categorical levels that
were present in A, caused some array elements to
have undefined levels.
A =
    lo <undefined> <undefined> <undefined> lo
I = isundefined(A)
I =
    0
```

See Also ..... islevel | ismember

## Purpose Test handle validity

## Syntax tf = isvalid(h)

Description $\quad \mathrm{tf}=$ isvalid( h ) performs an element-wise check for validity on the handle elements of $h$. The result is a logical array of the same dimensions as h , where each element is the element-wise validity result.

A handle is invalid if it has been deleted or if it is an element of a handle array and has not yet been initialized.

See Also delete | qrandstream

Purpose True if categorical array is vector

## Syntax <br> TF = isvector(A)

Description TF = isvector (A) returns true (1) if the categorical array A is a 1-by-n or $n$-by- 1 vector, where $n>=0$, and false ( 0 ) otherwise.

See Also isempty | isscalar | size

Purpose Number of iterations
Description The number of iterations of the algorithm.

Note This property applies only to gmdistribution objects constructed with fit.

## iwishrnd

Purpose Inverse Wishart random numbers

| Syntax | W $=$ iwishrnd $(T a u, d f)$ |
| :--- | :--- |
|  | W $=$ iwishrnd $(T a u, d f, D I)$ |
| $[W, D I]=\operatorname{iwishrnd}(T a u, d f)$ |  |

Description

See Also
wishrnd
How To . "Inverse Wishart Distribution" on page B-46

Jackknife sampling

```
jackstat = jackknife(jackfun,X)
jackstat = jackknife(jackfun,X,Y,...)
jackstat = jackknife(jackfun,...,'Options',option)
```


## Description

jackstat = jackknife(jackfun, X) draws jackknife data samples
from the n-by-p data array $X$, computes statistics on each sample using the function jackfun, and returns the results in the matrix jackstat. jackknife regards each row of $X$ as one data sample, so there are $n$ data samples. Each of the $n$ rows of jackstat contains the results of applying jackfun to one jackknife sample. jackfun is a function handle specified with @. Row i of jackstat contains the results for the sample consisting of $X$ with the ith row omitted:

```
s = x;
s(i,:) = [];
jackstat(i,:) = jackfun(s);
```

If jackfun returns a matrix or array, then this output is converted to a row vector for storage in jackstat. If X is a row vector, it is converted to a column vector.
jackstat $=$ jackknife(jackfun, X,Y,...) accepts additional arguments to be supplied as inputs to jackfun. They may be scalars, column vectors, or matrices. jackknife creates each jackknife sample by sampling with replacement from the rows of the non-scalar data arguments (these must have the same number of rows). Scalar data are passed to jackfun unchanged. Non-scalar arguments must have the same number of rows, and each jackknife sample omits the same row from these arguments.
jackstat = jackknife(jackfun, ...,'Options',option) provides an option to perform jackknife iterations in parallel, if the Parallel Computing Toolbox is available. Set 'Options ' as a structure you create with statset. jackknife uses the following field in the structure:

# 'UseParallel' If 'always' and if a matlabpool of the Parallel Computing Toolbox is open, use multiple processors to compute jackknife iterations. If the Parallel Computing Toolbox is not installed, or a matlabpool is not open, computation occurs in serial mode. Default is 'never', or serial computation. 

For more information on using parallel computing, see Chapter 17, "Parallel Statistics".

```
Examples Estimate the bias of the MLE variance estimator of random samples taken from the vector y using jackknife. The bias has a known formula in this problem, so you can compare the jackknife value to this formula.
```

```
sigma = 5;
```

sigma = 5;
y = normrnd(0,sigma,100,1);
y = normrnd(0,sigma,100,1);
m = jackknife(@var, y, 1);
m = jackknife(@var, y, 1);
n = length(y);
n = length(y);
bias = -sigma^2 / n % known bias formula
bias = -sigma^2 / n % known bias formula
jbias = (n - 1)*(mean(m)-var(y,1)) % jackknife bias estimate
jbias = (n - 1)*(mean(m)-var(y,1)) % jackknife bias estimate
bias =
bias =
-0.2500
-0.2500
jbias =
jbias =
-0.3378

```
    -0.3378
```

See Also bootstrp | random | randsample | hist | ksdensity
Tutorials

- "The Jackknife" on page 3-12

Purpose
Jarque-Bera test
Syntax
h = jbtest(x)
h = jbtest(x,alpha)
[h,p] = jbtest(...)
[h,p,jbstat] = jbtest(...)
[h, p,jbstat,critval] = jbtest(...)
[h,p,...] = jbtest(x,alpha,mctol)

## Description

$\mathrm{h}=\mathrm{jbtest}(\mathrm{x})$ performs a Jarque-Bera test of the null hypothesis that the sample in vector $x$ comes from a normal distribution with unknown mean and variance, against the alternative that it does not come from a normal distribution. The test is specifically designed for alternatives in "Generating Data Using the Pearson System" on page $6-26$ of distributions. The test returns the logical value $h=1$ if it rejects the null hypothesis at the $5 \%$ significance level, and $\mathrm{h}=0$ if it cannot. The test treats NaN values in x as missing values, and ignores them.

The Jarque-Bera test is a two-sided goodness-of-fit test suitable when a fully-specified null distribution is unknown and its parameters must be estimated. The test statistic is

$$
J B=\frac{n}{6}\left(s^{2}+\frac{(k-3)^{2}}{4}\right)
$$

where $n$ is the sample size, $s$ is the sample skewness, and $k$ is the sample kurtosis. For large sample sizes, the test statistic has a chi-square distribution with two degrees of freedom.

Jarque-Bera tests often use the chi-square distribution to estimate critical values for large samples, deferring to the Lilliefors test (see lillietest) for small samples. jbtest, by contrast, uses a table of critical values computed using Monte-Carlo simulation for sample sizes less than 2000 and significance levels between 0.001 and 0.50 . Critical values for a test are computed by interpolating into the table, using the analytic chi-square approximation only when extrapolating for larger sample sizes.
$\mathrm{h}=\mathrm{jbtest}(\mathrm{x}, \mathrm{alpha})$ performs the test at significance level alpha. alpha is a scalar in the range [0.001, 0.50]. To perform the test at a significance level outside of this range, use the motol input argument.
$[\mathrm{h}, \mathrm{p}]=\mathrm{jbtest}(\ldots)$ returns the $p$ value p , computed using inverse interpolation into the table of critical values. Small values of $p$ cast doubt on the validity of the null hypothesis. jbtest warns when $p$ is not found within the tabulated range of [0.001, 0.50], and returns either the smallest or largest tabulated value. In this case, you can use the mctol input argument to compute a more accurate $p$ value.
[h, p,jbstat] $=$ jbtest(...) returns the test statistic jbstat.
[h,p,jbstat,critval] = jbtest(...) returns the critical value critval for the test. When jbstat > critval, the null hypothesis is rejected at significance level alpha.
[h,p,...] = jbtest(x,alpha,mctol) computes a Monte-Carlo approximation for $p$ directly, rather than interpolating into the table of pre-computed values. This is useful when alpha or $p$ lie outside the range of the table. jbtest chooses the number of Monte Carlo replications, mcreps, large enough to make the Monte Carlo standard error for $p$, $\operatorname{sqrt}\left(p^{*}(1-p) / m c r e p s\right)$, less than mctol.

## Examples

Use jbtest to determine if car mileage, in miles per gallon (MPG), follows a normal distribution across different makes of cars:

```
load carbig
[h,p] = jbtest(MPG)
h =
    1
p =
    0.0022
```

The $p$ value is below the default significance level of $5 \%$, and the test rejects the null hypothesis that the distribution is normal.

With a log transformation, the distribution becomes closer to normal, but the $p$ value is still well below $5 \%$ :

```
[h,p] = jbtest(log(MPG))
\(\mathrm{h}=\)
    1
\(\mathrm{p}=\)
    0.0078
```

Decreasing the significance level makes it harder to reject the null hypothesis:

```
[h,p] = jbtest(log(MPG),0.0075)
h =
    0
\(p=\)
    0.0078
```


## References

[1] Jarque, C. M., and A. K. Bera. "A test for normality of observations and regression residuals." International Statistical Review. Vol. 55, No. 2, 1987, pp. 163-172.
[2] Deb, P., and M. Sefton. "The Distribution of a Lagrange Multiplier Test of Normality." Economics Letters. Vol. 51, 1996, pp. 123-130. This paper proposed a Monte Carlo simulation for determining the distribution of the test statistic. The results of this function are based on an independent Monte Carlo simulation, not the results in this paper.

Purpose Johnson system random numbers
Syntax $\quad \begin{aligned} & r=\text { johnsrnd(quantiles, } m, n) \\ & r=\text { johnsrnd(quantiles) } \\ & \\ & {[r, \text { type }]=j \text { johnsrnd }(\ldots)} \\ & \\ & {[r, \text { type }, \text { coefs }]=\text { johnsrnd }(\ldots)}\end{aligned}$

## Description

$r=j o h n s r n d(q u a n t i l e s, m, n)$ returns an $m-b y-n$ matrix of random numbers drawn from the distribution in the Johnson system that satisfies the quantile specification given by quantiles. quantiles is a four-element vector of quantiles for the desired distribution that correspond to the standard normal quantiles [ $-1.5-0.50 .51 .5]$. In other words, you specify a distribution from which to draw random values by designating quantiles that correspond to the cumulative probabilities [0.067 0.309 0.691 0.933]. quantiles may also be a 2-by-4 matrix whose first row contains four standard normal quantiles, and whose second row contains the corresponding quantiles of the desired distribution. The standard normal quantiles must be spaced evenly.

Note Because $r$ is a random sample, its sample quantiles typically differ somewhat from the specified distribution quantiles.
$r=$ johnsrnd(quantiles) returns a scalar value.
$r=$ johnsrnd(quantiles, $m, n, \ldots$ ) or $r=$
johnsrnd(quantiles, $m, n, \ldots]$ ) returns an $m$-by-n-by-... array.
[ $r$, type] $=$ johnsrnd (...) returns the type of the specified distribution within the Johnson system. type is 'SN', 'SL', 'SB', or 'SU'. Set $m$ and $n$ to zero to identify the distribution type without generating any random values.

The four distribution types in the Johnson system correspond to the following transformations of a normal random variate:

- 'SN ' - Identity transformation (normal distribution)
- 'SL' - Exponential transformation (lognormal distribution)
- 'SB' - Logistic transformation (bounded)
- 'SU' - Hyperbolic sine transformation (unbounded)
[r,type,coefs] = johnsrnd(...) returns coefficients coefs of the transformation that defines the distribution. coefs is [gamma, eta, epsilon, lambda]. If $z$ is a standard normal random variable and $h$ is one of the transformations defined above, $r=$ lambda*h( $z$-gamma)/eta) +epsilon is a random variate from the distribution type corresponding to $h$.


## Examples

Generate random values with longer tails than a standard normal:

```
r = johnsrnd([-1.7 -.5 .5 1.7],1000,1);
qqplot(r);
```



Generate random values skewed to the right:

```
r = johnsrnd([-1.3 -.5 .5 1.7],1000,1);
qqplot(r);
```



Generate random values that match some sample data well in the right-hand tail:

```
load carbig;
qnorm = [.5 1 1.5 2];
q = quantile(Acceleration, normcdf(qnorm));
r = johnsrnd([qnorm;q],1000,1);
[q;quantile(r,normcdf(qnorm))]
ans =
\begin{tabular}{llll}
16.7000 & 18.2086 & 19.5376 & 21.7263 \\
16.8190 & 18.2474 & 19.4492 & 22.4156
\end{tabular}
```

Determine the distribution type and the coefficients:

```
[r,type,coefs] = johnsrnd([qnorm;q],0)
r =
    []
type =
    SU
coefs =
    1.0920 0.5829 18.4382 1.4494
```


## See Also random | pearsrnd

How To . "Johnson System" on page B-48

| Purpose |
| :--- |
| Syntax |
| Description |

Merge observations

```
\(C=j 0 i n(A, B)\)
\(C=j o i n(A, B, k e y)\)
C = join(A,B,param1,val1,param2,val2,...)
[C,IB] = join(...)
C = join(A, B, 'Type', TYPE,...)
C = join(A, B,'Type',TYPE,'MergeKeys',true,...)
[C,IA,IB] = join(A,B,'Type',TYPE,...)
```

$C=j \operatorname{in}(A, B)$ creates a dataset array $C$ by merging observations from the two dataset arrays $A$ and $B$. join performs the merge by first finding key variables, that is, a pair of dataset variables, one in A and one in $B$, that share the same name. The key from B must contain unique values, and must contain all the values that are present in the key from $A$. join then uses these key variables to define a many-to-one correspondence between observations in A and those in B. join uses this correspondence to replicate the observations in $B$ and combine them with the observations in $A$ to create $C$.
$C=j o i n(A, B, k e y)$ performs the merge using the variable specified by key as the key variable in both A and B. key is a positive integer, a variable name, a cell array containing a variable name, or a logical vector with one true entry.

C contains one observation for each observation in A. Variables in C include all of the variables from $A$, as well as one variable corresponding to each variable in B (except for the key from B). If A and B contain variables with identical names, join adds the suffix '_left' and '_right' to the corresponding variables in C.
$C=j o i n(A, B, p a r a m 1, v a l 1, p a r a m 2, v a l 2, \ldots)$ specifies optional parameter name/value pairs to control how the dataset variables in A and $B$ are used in the merge. Parameters are:

- 'Keys ' - Specifies the variable to use as a keys in both A and B.
- 'LeftKeys' - Specifies the variable to use as a keys in A.
- 'RightKeys ' - Specifies the variable to use as a keys in B.

You may provide either the 'Keys ' parameter, or both the 'LeftKeys' and 'RightKeys' parameters. The value for these parameters is a positive integer, a variable name, a cell array containing a variable name, or a logical vector with one true entry. 'LeftKeys' or 'RightKeys' must both specify the same number of key variables, and join pairs the left and right keys are paired in the order specified.

- 'LeftVars' - Specifies the variables from A to include in C. By default, join includes all variables from A.
- 'RightVars' - Specifies the variables from B to include in C. By default, join includes all variables from $B$ except the key variable.

You can use 'LeftVars' or 'RightVars' to include or exclude key variables as well as data variables. The value for these parameters is a positive integer, a vector of positive integers, a variable name, a cell array containing one or more variable names, or a logical vector.
[C,IB] = join(...) returns an index vector IB, where join constructs C by horizontally concatenating A(:,LeftVars) and B(IB, RightVars). join can also perform more complicated inner and outer join operations that allow a many-to-many correspondence between A and B, and allow unmatched observations in either A or B.

C = join(A, B,'Type',TYPE,...) performs the join operation specified by TYPE. TYPE is one of 'inner', 'leftouter', 'rightouter', 'fullouter', or 'outer' (which is a synonym for 'fullouter'). For an inner join, C only contains observations corresponding to a combination of key values that occurred in both A and B. For a left (or right) outer join, $C$ also contains observations corresponding to keys in $A($ or $B)$ that did not match any in B (or A). Variables in C taken from A (or B) contain null values in those observations. A full outer join is equivalent to a left and right outer join. C contains variables corresponding to the key variables from both $A$ and $B$, and join sorts the observations in $C$ by the key values.

For inner and outer joins, C contains variables corresponding to the key variables from both A and B by default, as well as all the remaining variables. join sorts the observations in the result $C$ by the key values.

C = join(A, B,'Type', TYPE,'MergeKeys',true,...) includes a single variable in $C$ for each key variable pair from $A$ and $B$, rather than including two separate variables. For outer joins, join creates the single variable by merging the key values from A and B, taking values from $A$ where a corresponding observation exists in $A$, and from $B$ otherwise. Setting the 'MergeKeys' parameter to true overrides inclusion or exclusion of any key variables specified via the 'LeftVars' or 'RightVars' parameter. Setting the 'MergeKeys' parameter to false is equivalent to not passing in the 'MergeKeys' parameter.
[C,IA,IB] = join(A, B,'Type',TYPE,...) returns index vectors IA and IB indicating the correspondence between observations in $C$ and those in A and B. For an inner join, join constructs $C$ by horizontally concatenating A(IA,LeftVars) and B(IB,RightVars). For an outer join, IA or IB may also contain zeros, indicating the observations in C that do not correspond to observations in A or B , respectively.

## Examples

Create a dataset array from Fisher's iris data:

```
load fisheriris
NumObs = size(meas,1);
NameObs = strcat({'Obs'},num2str((1:NumObs)','%-d'));
iris = dataset({nominal(species),'species'},...
    {meas,'SL','SW','PL','PW'},...
    'ObsNames',NameObs);
```

Create a separate dataset array with the diploid chromosome counts for each species of iris:

```
snames = nominal({'setosa';'versicolor';'virginica'});
CC = dataset({snames,'species'},{[38;108;70],'cc'})
CC =
    species cc
```

```
setosa38
```

versicolor ..... 108
virginica ..... 70

Broadcast the data in CC to the rows of iris using the key variable species in each dataset:

```
iris2 = join(iris,CC);
iris2([1 2 51 52 101 102],:)
ans =
\begin{tabular}{llllllr} 
& Species & SL & SW & PL & PW & \multicolumn{1}{c}{ CC } \\
Obs1 & setosa & 5.1 & 3.5 & 1.4 & 0.2 & 38 \\
Obs2 & setosa & 4.9 & 3 & 1.4 & 0.2 & 38 \\
Obs51 & versicolor & 7 & 3.2 & 4.7 & 1.4 & 108 \\
Obs52 & versicolor & 6.4 & 3.2 & 4.5 & 1.5 & 108 \\
Obs101 & virginica & 6.3 & 3.3 & 6 & 2.5 & 70 \\
Obs102 & virginica & 5.8 & 2.7 & 5.1 & 1.9 & 70
\end{tabular}
```

Create two datasets and join them using the 'MergeKeys' flag:
\% Create two data sets that both contain the key variable
\% 'Key1'. The two arrays contain observations with common
\% values of Key1, but each array also contains observations
\% with values of Key1 not present in the other.
a = dataset(\{'a' 'b' 'c' 'e' 'h'\}',[1 2311 17]',...
'VarNames',\{'Key1' 'Var1'\})
b = dataset(\{'a' 'b' 'd' 'e'\}',[4 5 6 7]',...
'VarNames',\{'Key1' 'Var2'\})
\% Combine a and b with an outer join, which matches up \% observations with common key values, but also retains \% observations whose key values don't have a match.
\% Keep the key values as separate variables in the result. couter = join(a,b,'key','Key1','Type','outer')

```
% Join a and b, merging the key values as a single variable
% in the result.
coutermerge = join(a,b,'key','Key1','Type','outer',...
    'MergeKeys',true)
% Join a and b, retaining only observations whose key
% values match.
cinner = join(a,b,'key','Key1','Type','inner',...
    'MergeKeys',true)
a =
\begin{tabular}{cr} 
Key1 & Var1 \\
'a' & 1 \\
'b' & 2 \\
'c' & 3 \\
'e' & 11 \\
'h' & 17
\end{tabular}
b =
\begin{tabular}{ll} 
Key1 & Var2 \\
'a' & 4 \\
'b' & 5 \\
'd' & 6 \\
'e' & 7
\end{tabular}
```


## couter =

```
\begin{tabular}{lclr} 
Key1_left & Var1 & Key1_right & Var2 \\
'a' & 1 & 'a' \(^{\prime}\) & 4 \\
'b' & 2 & 'b' & 5 \\
'c' & 3 & \('\) '' & NaN \\
'' & NaN & 'd' & 6 \\
'e' & 11 & 'e' & 7
\end{tabular}
```


## dataset.join



See Also
sortrows

## KDTreeSearcher

Superclasses NeighborSearcher
Purpose Nearest neighbors search using $k d$-tree
Description A KDTreeSearcher object represents $k \mathrm{NN}$ ( $k$-nearest neighbor) searchusing a $k$ d-tree. Search objects store information about the data used,the distance metric and parameters, and the maximal number of datapoints in each leaf node. You cannot create this object for sparseinput data. The search performance for this object, compared with theExhaustiveSearcher object, tends to be better for smaller dimensions(10 or fewer) and worse for larger dimensions. For more information onsearch objects, see "What Are Search Objects?" on page 13-17.

## Construction

NS = KDTreeSearcher(X, 'Name', Value) constructs a $k$ d-tree based on $X$ and saves the information in a KDTreeSearcher object where rows of X correspond to observations and columns correspond to variables. You can use this tree to find neighbors in $X$ nearest to the query points. See the following section for optional name/value pairs.
NS = createns(X,'NSMethod','kdtree', 'Name', Value) creates a $k$ d-tree based on $X$ using createns and saves the information in a KDTreeSearcher object where rows of $X$ correspond to observations and columns correspond to variables. You can use this tree to find neighbors in $X$ nearest to the query points. See the following section for optional name/value pairs.

## Name-Value Pair Arguments

KDTreeSearcher and createns accept one or more of the following optional name/value pairs as input:

## Distance

A string specifying the default distance metric used when you call the knnsearch method.

- 'euclidean' - Euclidean distance (default).
- 'cityblock' - City block distance.


## KDTreeSearcher

- 'chebychev' - Chebychev distance (maximum coordinate difference).
- 'minkowski' - Minkowski distance.

For more information on these distance metrics, see "Distance Metrics" on page 13-9.

P
A positive scalar indicating the exponent of the Minkowski distance. This parameter is only valid when Distance is 'minkowski'. Default is 2.

## BucketSize

A positive integer, indicating the maximum number of data points in each leaf node of the $k$ d-tree. Default is 50 .

## Properties <br> X

A matrix used to create the object

## Distance

A string specifying a built-in distance metric that you provide when you create the object. This property is the default distance metric used when you call the knnsearch method to find nearest neighbors for future query points.

## DistParameter

Specifies the additional parameter for the chosen distance metric. The value is:

- If 'Distance' is 'minkowski': A positive scalar indicating the exponent of the Minkowski distance.
- Otherwise: Empty.

```
Methods
knnsearch
rangesearch
Find k-nearest neighbors using KDTreeSearcher object
Find all neighbors within specified distance using object
Examples Create a KDTreeSearcher object using the constructor:
```

```
load fisheriris
```

load fisheriris
x = meas(:,3:4);
x = meas(:,3:4);
kdtreeobj = kdtreesearcher(x,'distance','minkowski')
kdtreeobj = kdtreesearcher(x,'distance','minkowski')
kdtreeobj =
kdtreeobj =
KDTreeSearcher
Properties:
BucketSize: 50
X: [150x2 double]
Distance: 'minkowski'
DistParameter: 2

```

Create a KDTreeSearcher object using createns:
```

load fisheriris
x = meas(:,3:4);
kdtreeobj = createns(x,'NsMethod','kdtree',...
distance','minkowski')
kdtreeobj =
KDTreeSearcher
Properties:
BucketSize: 50

```

\section*{KDTreeSearcher}
```

X: [150x2 double]
Distance: 'minkowski'
DistParameter: 2

```

For more in-depth examples using the knnsearch method, see the method reference page or see "Example: Classifying Query Data Using knnsearch" on page 13-18.

\author{
References \\ [1] Friedman, J. H., Bentely, J., and Finkel, R. A. (1977). An Algorithm for Finding Best Matches in Logarithmic Expected Time, ACM Transactions on Mathematical Software 3, 209.
}

\section*{See Also \\ createns | ExhaustiveSearcher | NeighborSearcher}

How To . " \(k\)-Nearest Neighbor Search and Radius Search" on page 13-12

\section*{ProbDistKernel.Kernel property}
\begin{tabular}{ll} 
Purpose & \begin{tabular}{l} 
Read-only string specifying name of kernel smoothing function for \\
ProbDistKernel object
\end{tabular} \\
Description & \begin{tabular}{l} 
Kernel is a read-only property of the ProbDistKernel class. Kernel is \\
a string specifying the name of the kernel smoothing function used \\
to create a ProbDistKernel object.
\end{tabular} \\
Values & \begin{tabular}{l} 
' normal' \\
'box' \\
'triangle' \\
'epanechnikov'
\end{tabular} \\
See Also & \begin{tabular}{l} 
Use this information to view and compare the kernel smoothing \\
function used to create distributions.
\end{tabular} \\
ksdensity
\end{tabular}

\section*{ClassificationPartitionedEnsemble.kfoldEdge}
\begin{tabular}{|c|c|}
\hline Purpose & Classification edge for observations not used for training \\
\hline Syntax & ```
E = kfoldEdge(obj)
E = kfoldEdge(obj,Name,Value)
``` \\
\hline Description & \begin{tabular}{l}
E = kfoldEdge(obj) returns classification edge (average classification margin) obtained by cross-validated classification ensemble obj. For every fold, this method computes classification edge for in-fold observations using an ensemble trained on out-of-fold observations. \\
E = kfoldEdge(obj, Name, Value) calculates edge with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
\end{tabular} \\
\hline \begin{tabular}{l}
Input \\
Arguments
\end{tabular} & \begin{tabular}{l}
ens \\
Object of class ClassificationPartitionedEnsemble. Create ens with fitensemble along with one of the cross-validation options: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Alternatively, create ens from a classification ensemble with crossval.
\end{tabular} \\
\hline & Name-Value Pair Arguments \\
\hline & \begin{tabular}{l}
Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN. \\
folds
\end{tabular} \\
\hline & Indices of folds ranging from 1 to ens. KFold. Use only these folds for predictions. \\
\hline & Default: 1: ens.KFold \\
\hline & mode \\
\hline
\end{tabular}

\section*{ClassificationPartitionedEnsemble.kfoldEdge}

String representing the meaning of the output edge:
- 'average ' - edge is a scalar value, the average over all folds.
- 'individual' - edge is a vector of length ens.KFold with one element per fold.
- 'cumulative' - edge is a vector of length min(ens.NTrainedPerFold) in which element \(J\) is obtained by averaging values across all folds for weak learners \(1: J\) in each fold.

Default: 'average'

\section*{Output E \\ Arguments}

\section*{Definitions}

\section*{Edge}

The edge is the weighted mean value of the classification margin. The weights are the class probabilities in obj. Prior.

\section*{Margin}

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix obj. x .

\section*{Score (ensemble)}

For ensembles, a classification score represents the confidence of a classification into a class. The higher the score, the higher the confidence.

Different ensemble algorithms have different definitions for their scores. Furthermore, the range of scores depends on ensemble type. For example:

\section*{ClassificationPartitionedEnsemble.kfoldEdge}
- AdaBoostM1 scores range from \(-\infty\) to \(\infty\).
- Bag scores range from 0 to 1 .
```

Examples Compute the k-fold edge for an ensemble trained on the Fisher iris data:
load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
cvens = crossval(ens);
E = kfoldEdge(cvens)
E =
3.2078

```
See Also kfoldPredict | kfoldLoss | kfoldMargin | kfoldfun | crossval
How To
- Chapter 13, "Nonparametric Supervised Learning"

\section*{ClassificationPartitionedModel.kfoldEdge}
Purpose Classification edge for observations not used for training
Syntax

E = kfoldEdge(obj)

E = kfoldEdge(obj,Name,Value)
Description
InputArguments
\(\mathrm{E}=\mathrm{kfoldEdge}(\mathrm{obj})\) returns classification edge (average classification margin) obtained by cross-validated classification model obj. For every fold, this method computes classification edge for in-fold observations using an ensemble trained on out-of-fold observations.
E = kfoldEdge(obj, Name, Value) calculates edge with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
obj
Object of class ClassificationPartitionedModel.

\section*{Name-Value Pair Arguments}
Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{folds}
Indices of folds ranging from 1 to obj. KFold. Use only these folds for predictions.
Default: 1:obj.KFold
mode
String representing the meaning of the output edge:
- 'average ' - edge is a scalar value, the average over all folds.

\section*{ClassificationPartitionedModel.kfoldEdge}
- 'individual' - edge is a vector of length obj. KFold with one element per fold.

Default: 'average'

\section*{Output E}

\section*{Definitions}

\section*{Edge}

The edge is the weighted mean value of the classification margin. The weights are the class probabilities in obj. Prior. If you supply weights in the weights name-value pair, those weights are normalized to sum to the prior probabilities in the respective classes, and are then used to compute the weighted average.

\section*{Margin}

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix obj. x . A high value of margin indicates a more reliable prediction than a low value.

\section*{Score (discriminant analysis)}

For discriminant analysis, the score of a classification is the posterior probability of the classification. For the definition of posterior probability in discriminant analysis, see "Posterior Probability" on page 12-7.

\section*{Score (tree)}

For trees, the score of a classification of a leaf node is the posterior probability of the classification at that node. The posterior probability of the classification at a node is the number of training sequences that lead to that node with the classification, divided by the number of training sequences that lead to that node.

\section*{ClassificationPartitionedModel.kfoldEdge}

For example, consider classifying a predictor X as true when \(\mathrm{X}<0.15\) or \(X>0.95\), and \(X\) is false otherwise.

1
Generate 100 random points and classify them:
```

rng(0,'twister') % for reproducibility
X = rand(100,1);
Y = (abs(X - .55) > .4);
tree = ClassificationTree.fit(X,Y);
view(tree,'mode','graph')

```


2
Prune the tree:
```

tree1 = prune(tree,'level',1);
view(tree1,'mode','graph')

```

\section*{ClassificationPartitionedModel.kfoldEdge}


The pruned tree correctly classifies observations that are less than 0.15 as true. It also correctly classifies observations from .15 to .94 as false. However, it incorrectly classifies observations that are greater than . 94 as false. Therefore, the score for observations that are greater than .15 should be about \(.05 / .85=.06\) for true, and about \(.8 / .85=.94\) for false.

3
Compute the prediction scores for the first 10 rows of X :
```

[~,score] = predict(tree1,X(1:10));
[score X(1:10,:)]
ans =

| 0.9059 | 0.0941 | 0.8147 |
| ---: | ---: | ---: |
| 0.9059 | 0.0941 | 0.9058 |
| 0 | 1.0000 | 0.1270 |
| 0.9059 | 0.0941 | 0.9134 |
| 0.9059 | 0.0941 | 0.6324 |
| 0 | 1.0000 | 0.0975 |
| 0.9059 | 0.0941 | 0.2785 |

```

\section*{ClassificationPartitionedModel.kfoldEdge}
\begin{tabular}{lll}
0.9059 & 0.0941 & 0.5469 \\
0.9059 & 0.0941 & 0.9575 \\
0.9059 & 0.0941 & 0.9649
\end{tabular}

Indeed, every value of \(X\) (the rightmost column) that is less than 0.15 has associated scores (the left and center columns) of 0 and 1 , while the other values of \(X\) have associated scores of 0.91 and 0.09 . The difference (score 0.09 instead of the expected .06) is due to a statistical fluctuation: there are 8 observations in \(X\) in the range (.95,1) instead of the expected 5 observations.

\section*{Examples}

Compute the k-fold edge for a model trained on the Fisher iris data:
```

load fisheriris
tree = ClassificationTree.fit(meas,species);
cvtree = crossval(tree);
E = kfoldEdge(cvtree)
E =
0.8711

```
See Also \begin{tabular}{l} 
kfoldPredict | kfoldMargin | kfoldLoss | kfoldfun \\
| crossval | ClassificationPartitionedEnsemble | \\
ClassificationPartitionedModel
\end{tabular}

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\section*{ClassificationPartitionedModel.kfoldfun}

Purpose Cross validate function
Syntax \(\quad\) vals \(=\) kfoldfun (obj, fun \()\)

Description

Input
Arguments

\section*{Output}

Arguments
vals
The arrays of testvals output, concatenated vertically over all folds. For example, if testvals from every fold is a numeric vector of length N , kfoldfun returns a KFold-by-N numeric matrix with one row per fold.
```

Examples Cross validate a classification tree, and obtain the classification error
(see kfoldLoss):
load fisheriris
t = ClassificationTree.fit(meas,species);
rng(0,'twister') % for reproducibility
cv = crossval(t);
L = kfoldLoss(cv)
L =
0.0467

```

Examine the result when the error of misclassifying a flower as 'versicolor' is 10, and any other error is 1 :

1 Write a function file that gives a cost of 1 for misclassification, but 10 for misclassifying a flower as versicolor.
```

function averageCost = noversicolor(CMP,Xtrain,Ytrain,Wtrain,Xtest,Ytest,Wtest)
Ypredict = predict(CMP,Xtest);
misclassified = not(strcmp(Ypredict,Ytest)); % different result
classifiedAsVersicolor = strcmp(Ypredict,'versicolor'); % index of bad decisions
cost = sum(misclassified) + ...
9*sum(misclassified \& classifiedAsVersicolor); % total differences
averageCost = cost/numel(Ytest); % average error

```

2 Save the file as noversicolor.m on your MATLAB path.
3 Compute the mean misclassification error with the noversicolor cost:
```

mean(kfoldfun(cv,@noversicolor))
ans =
0.1667

```

\section*{ClassificationPartitionedModel.kfoldfun}

\author{
See Also kfoldPredict | kfoldEdge | kfoldMargin | kfoldLoss | crossval \\ How To - Chapter 13, "Nonparametric Supervised Learning"
}

\section*{RegressionPartitionedModel.kfoldfun}
Purpose Cross validate function

Cross validate function
Syntax vals = kfoldfun(obj,fun)

vals = kfoldfun(obj,fun)Description
Input
Arguments
vals \(=k f o l d f u n(o b j, f u n)\) cross validates the function fun by applying fun to the data stored in the cross-validated model obj. You must pass fun as a function handle.
obj
Object of class RegressionPartitionedModel or RegressionPartitionedEnsemble. Create obj with RegressionTree.fit or fitensemble along with one of the cross-validation options: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Alternatively, create obj from a regression tree or regression ensemble with crossval.
fun
A function handle for a cross-validation function. fun has the syntax
```

testvals = fun(CMP,Xtrain,Ytrain,Wtrain,Xtest,Ytest,Wtest)

```
- CMP is a compact model stored in one element of the obj.Trained property.
- Xtrain is the training matrix of predictor values.
- Ytrain is the training array of response values.
- Wtrain are the training weights for observations.
- Xtest and Ytest are the test data, with associated weights Wtest.
- The returned value testvals must have the same size across all folds.

\section*{RegressionPartitionedModel.kfoldfun}
OutputArguments
Examples Cross validate a regression tree, and obtain the mean squared error (see kfoldLoss):
```

load imports-85
t = RegressionTree.fit(X(:,[4 5]),X(:,16),...
'predictornames',{'length' 'width'},...
'responsename','price');
cv = crossval(t);
L = kfoldLoss(cv)
L =
1.5489e+007

```
Examine the result of simple averaging of responses instead of using predictions:
```

f = @(cmp,Xtrain,Ytrain,Wtrain,Xtest,Ytest,Wtest)...
mean((Ytest-mean(Ytrain)).^2)
mean(kfoldfun(cv,f))
ans =
6.3497e+007

```
See Also RegressionPartitionedEnsemble | kfoldPredict | kfoldLoss | RegressionPartitionedModel | crossval

\section*{How To}

\section*{ClassificationPartitionedEnsemble.kfoldLoss}
Purpose Classification loss for observations not used for training

Classification loss for observations not used for training

Arguments

L = kfoldLoss(ens)
L = kfoldLoss(ens,Name, Value)
\(\mathrm{L}=\) kfoldLoss(ens) returns loss obtained by cross-validated classification model ens. For every fold, this method computes classification loss for in-fold observations using a model trained on out-of-fold observations.

L = kfoldLoss(ens,Name, Value) calculates loss with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
ens
Object of class ClassificationPartitionedEnsemble. Create ens with fitensemble along with one of the cross-validation options: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Alternatively, create ens from a classification ensemble with crossval.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.
folds
Indices of folds ranging from 1 to ens.KFold. Use only these folds for predictions.

Default: 1:ens.KFold

\footnotetext{
lossfun
}

\section*{ClassificationPartitionedEnsemble.kfoldLoss}

Function handle or string representing a loss function. Built-in loss functions:
- 'binodeviance' - See "Loss Functions" on page 20-826
- 'classiferror' - Fraction of misclassified data
- 'exponential' - See "Loss Functions" on page 20-826

You can write your own loss function in the syntax described in "Loss Functions" on page 20-826.

Default: 'classiferror'
mode
A string for determining the output of kfoldLoss:
- 'average ' - L is a scalar, the loss averaged over all folds.
- 'individual' - L is a vector of length ens.KFold, where each entry is the loss for a fold.
- 'cumulative' - \(L\) is a vector in which element \(J\) is obtained by using learners \(1: J\) from the input list of learners.

Default: 'average'

\section*{Output \\ Arguments}

\section*{Definitions}

Loss, by default the fraction of misclassified data. L can be a vector, and can mean different things, depending on the name-value pair settings.

\section*{Loss Functions}

The built-in loss functions are:

\section*{ClassificationPartitionedEnsemble.kfoldLoss}
- 'binodeviance' - For binary classification, assume the classes \(y_{n}\) are -1 and 1 . With weight vector \(w\) normalized to have sum 1 , and predictions of row \(n\) of data \(X\) as \(f\left(X_{n}\right)\), the binomial deviance is
\[
\sum w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(X_{n}\right)\right)\right)
\]
- 'classiferror' - Fraction of misclassified data, weighted by \(w\).
- 'exponential' - With the same definitions as for 'binodeviance', the exponential loss is
\[
\sum w_{n} \exp \left(-y_{n} f\left(X_{n}\right)\right)
\]

To write your own loss function, create a function file of the form
```

function loss = lossfun(C,S,W,COST)

```
- \(N\) is the number of rows of ens. \(X\).
- K is the number of classes in ens, represented in ens.ClassNames.
- C is an \(N\)-by-K logical matrix, with one true per row for the true class. The index for each class is its position in tree.ClassNames.
- \(S\) is an \(N\)-by-K numeric matrix. \(S\) is a matrix of posterior probabilities for classes with one row per observation, similar to the score output from predict.
- W is a numeric vector with \(N\) elements, the observation weights.
- COST is a K-by-K numeric matrix of misclassification costs. The default 'classiferror' gives a cost of 0 for correct classification, and 1 for misclassification.
- The output loss should be a scalar.

Pass the function handle @lossfun as the value of the lossfun name-value pair.

\section*{ClassificationPartitionedEnsemble.kfoldLoss}
```

Examples Find the average cross-validated classification error for an ensemble
model of the ionosphere data:
load ionosphere
ens = fitensemble(X,Y,'AdaBoostM1',100,'Tree');
cvens = crossval(ens);
L = kfoldLoss(cvens)
L =
0.0826
See Also kfoldPredict | kfoldEdge | kfoldMargin | kfoldfun | crossval
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```

\section*{ClassificationPartitionedModel.kfoldLoss}
Description
InputArguments
Purpose Classification loss for observations not used for training

Classification loss for observations not used for training
Syntax

L = kfoldLoss(obj)

L = kfoldLoss(obj,Name,Value)

L = kfoldLoss(obj)

L = kfoldLoss(obj,Name,Value)
\(\mathrm{L}=\mathrm{kfoldLoss}(\mathrm{obj})\) returns loss obtained by cross-validated classification model obj. For every fold, this method computes classification loss for in-fold observations using a model trained on out-of-fold observations.

L = kfoldLoss(obj,Name, Value) calculates loss with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
obj
Object of class ClassificationPartitionedModel.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{folds}

Indices of folds ranging from 1 to obj. KFold. Use only these folds for predictions.

Default: 1:obj.KFold

\section*{lossfun}

Function handle or string representing a loss function. Built-in loss functions:
- 'binodeviance ' - See "Loss Functions" on page 20-830

\section*{ClassificationPartitionedModel.kfoldLoss}
- 'classiferror' - Fraction of misclassified data
- 'exponential' - See "Loss Functions" on page 20-830

You can write your own loss function in the syntax described in "Loss Functions" on page 20-830.

Default: 'classiferror'
mode
A string for determining the output of kfoldLoss:
- 'average ' - L is a scalar, the loss averaged over all folds.
- 'individual' - L is a vector of length obj.KFold, where each entry is the loss for a fold.

Default: 'average'

\section*{Output} Arguments

\section*{Definitions}

\section*{Classification Error}

The default classification error is the fraction of the data \(X\) that obj misclassifies, where \(Y\) are the true classifications.

Weighted classification error is the sum of weight \(i\) times the Boolean value that is 1 when obj misclassifies the \(i\) th row of \(X\), divided by the sum of the weights.

\section*{Loss Functions}

The built-in loss functions are:

\section*{ClassificationPartitionedModel.kfoldLoss}
- 'binodeviance' - For binary classification, assume the classes \(y_{n}\) are - 1 and 1 . With weight vector \(w\) normalized to have sum 1, and predictions of row \(n\) of data \(X\) as \(f\left(X_{n}\right)\), the binomial deviance is
\[
\sum w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(X_{n}\right)\right)\right)
\]
- 'classiferror' - Fraction of misclassified data, weighted by \(w\).
- 'exponential' - With the same definitions as for 'binodeviance', the exponential loss is
\[
\sum w_{n} \exp \left(-y_{n} f\left(X_{n}\right)\right)
\]

To write your own loss function, create a function file of the form
```

function loss = lossfun(C,S,W,COST)

```
- \(N\) is the number of rows of obj. \(X\).
- K is the number of classes in obj, represented in obj.ClassNames.
- C is an N-by-K logical matrix, with one true per row for the true class. The index for each class is its position in obj. ClassNames.
- \(S\) is an \(N\)-by-K numeric matrix. \(S\) is a matrix of posterior probabilities for classes with one row per observation, similar to the score output from predict.
- W is a numeric vector with \(N\) elements, the observation weights.
- COST is a K-by-K numeric matrix of misclassification costs. The default 'classiferror' gives a cost of 0 for correct classification, and 1 for misclassification.
- The output loss should be a scalar.

Pass the function handle @lossfun as the value of the lossfun name-value pair.

\section*{ClassificationPartitionedModel.kfoldLoss}
```

Examples Find the average cross-validated classification error for a model of the
ionosphere data:

```
```

load ionosphere

```
load ionosphere
tree = ClassificationTree.fit(X,Y);
tree = ClassificationTree.fit(X,Y);
cvtree = crossval(tree);
cvtree = crossval(tree);
L = kfoldLoss(cvtree)
L = kfoldLoss(cvtree)
L =
L =
    0.1197
    0.1197
See Also kfoldPredict | kfoldEdge | kfoldMargin | kfoldfun | crossval
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```


## RegressionPartitionedEnsemble.kfoldLoss

## Purpose

Cross-validation loss of partitioned regression ensemble
Syntax

L = kfoldLoss(cvens)
L = kfoldLoss(cvens,Name, Value)

Description

Input Arguments
$\mathrm{L}=\mathrm{kfold}$ Loss(cvens) returns the cross-validation loss of cvens.
L = kfoldLoss(cvens, Name, Value) returns cross-validation loss with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

## cvens

Object of class RegressionPartitionedEnsemble. Create
obj with fitensemble along with one of the cross-validation options: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Alternatively, create obj from a regression ensemble with crossval.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## folds

Indices of folds ranging from 1 to cvens. KFold. Use only these folds for predictions.

Default: 1:cvens.KFold

## lossfun

Function handle for loss function, or the string 'mse', meaning mean squared error. If you pass a function handle fun, loss calls it as

## RegressionPartitionedEnsemble.kfoldLoss

> fun(Y,Yfit,W)
where Y , Yfit, and W are numeric vectors of the same length.

- Y is the observed response.
- Yfit is the predicted response.
- $W$ is the observation weights.

The returned value fun(Y,Yfit, W) should be a scalar.
Default: 'mse'

## mode

String representing the meaning of the output L:

- 'ensemble' - L is a scalar value, the loss for the entire ensemble.
- 'individual' - L is a vector with one element per trained learner.
- 'cumulative' - $L$ is a vector in which element $J$ is obtained by using learners $1: J$ from the input list of learners.

Default: 'ensemble'

## Output

Arguments

## Examples

Find the cross-validation loss for a regression ensemble of the carsmall data:

[^12]
## RegressionPartitionedEnsemble.kfoldLoss

```
X = [Displacement Horsepower Weight];
rens = fitensemble(X,MPG,'LSboost',100,'Tree');
cvrens = crossval(rens);
L = kfoldLoss(cvrens)
L =
    25.6935
```


## See Also <br> RegressionPartitionedEnsemble | loss | kfoldPredict

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## RegressionPartitionedModel.kfoldLoss

| Purpose | Cross-validation loss of partitioned regression model |
| :--- | :--- |
| Syntax | $\mathrm{L}=$ kfoldLoss (cvmodel) <br> $\mathrm{L}=$ |
| = kfoldLoss(cvmodel, Name, Value) |  |

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
folds
Indices of folds ranging from 1 to obj. KFold. Use only these folds for predictions.

Default: 1:obj.KFold

## lossfun

Function handle for loss function, or the string 'mse', meaning mean squared error. If you pass a function handle fun, kfoldLoss calls it as

## RegressionPartitionedModel.kfoldLoss

|  | fun(Y,Yfit, W) |
| :---: | :---: |
|  | where Y, Yfit, and W are numeric vectors of the same length. |
|  | - Y is the observed response. |
|  | - Yfit is the predicted response. |
|  | - W is the observation weights. |
|  | The returned value fun(Y,Yfit, W) should be a scalar. |
|  | Default: 'mse' |
|  | mode |
|  | One of the following strings: |
|  | - 'average ' - L is the average loss over all folds. |
|  | - 'individual' - L is a vector of the individual losses of in-fold observations trained on out-of-fold data. |
|  | Default: 'average' |
| Output | L |
| Arguments | The loss (mean squared error) between the observations in a fold when compared against predictions made with a tree trained on the out-of-fold data. If mode is 'individual', $L$ is a vector of the losses. If mode is 'average ', L is the average loss. |
| Examples | Construct a partitioned regression model, and examine the cross-validation losses for the folds: |
|  | load carsmall |
|  | XX = [Cylinders Displacement Horsepower Weight]; $Y Y=M P G$; |
|  | cvmodel $=$ RegressionTree.fit(XX,YY,'crossval', 'on'); |

## RegressionPartitionedModel.kfoldLoss

$$
\begin{aligned}
& \text { L = } k f o l d L o s s(c v m o d e l, ' m o d e ~ ', ~ ' i n d i v i d u a l ') ~ \\
& L= \\
& 44.9635 \\
& 11.8525 \\
& 18.2046 \\
& 9.2965 \\
& 29.4329 \\
& 54.8659 \\
& 24.6446 \\
& 8.2085 \\
& 19.7593 \\
& 16.7394
\end{aligned}
$$

# Alternatives You can avoid constructing a cross-validated tree model by calling cvLoss instead of kfoldLoss. The cross-validated tree can save time if you are going to examine it more than once. 

## See Also loss | kfoldPredict

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## ClassificationPartitionedModel.kfoldMargin

## Purpose

Classification margins for observations not used for training

## Syntax

Description

## Input <br> Arguments

## Output <br> M

Arguments

## Definitions

M = kfoldMargin(obj)
M = kfoldMargin(obj) returns classification margins obtained by cross-validated classification model obj. For every fold, this method computes classification margins for in-fold observations using a model trained on out-of-fold observations.
obj
A partitioned classification model of type ClassificationPartitionedModel or ClassificationPartitionedEnsemble.

The classification margin.

## Margin

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix obj. x . A high value of margin indicates a more reliable prediction than a low value.

## Score (discriminant analysis)

For discriminant analysis, the score of a classification is the posterior probability of the classification. For the definition of posterior probability in discriminant analysis, see "Posterior Probability" on page 12-7.

## Score (ensemble)

For ensembles, a classification score represents the confidence of a classification into a class. The higher the score, the higher the confidence.

## ClassificationPartitionedModel.kfoldMargin

Different ensemble algorithms have different definitions for their scores. Furthermore, the range of scores depends on ensemble type. For example:

- AdaBoostM1 scores range from $-\infty$ to $\infty$.
- Bag scores range from 0 to 1 .


## Score (tree)

For trees, the score of a classification of a leaf node is the posterior probability of the classification at that node. The posterior probability of the classification at a node is the number of training sequences that lead to that node with the classification, divided by the number of training sequences that lead to that node.

For example, consider classifying a predictor $X$ as true when $X<0.15$ or $X>0.95$, and $X$ is false otherwise.

Generate 100 random points and classify them:

```
rng(0,'twister') % for reproducibility
X = rand(100,1);
Y = (abs(X - .55) > .4);
tree = ClassificationTree.fit(X,Y);
view(tree,'mode','graph')
```


## ClassificationPartitionedModel.kfoldMargin



2
Prune the tree:

```
tree1 = prune(tree,'level',1);
view(tree1,'mode','graph')
```


## ClassificationPartitionedModel.kfoldMargin



The pruned tree correctly classifies observations that are less than 0.15 as true. It also correctly classifies observations from .15 to .94 as false. However, it incorrectly classifies observations that are greater than . 94 as false. Therefore, the score for observations that are greater than .15 should be about $.05 / .85=.06$ for true, and about $.8 / .85=.94$ for false.

3
Compute the prediction scores for the first 10 rows of X :

```
[~,score] = predict(tree1,X(1:10));
[score X(1:10,:)]
ans =
\begin{tabular}{rrr}
0.9059 & 0.0941 & 0.8147 \\
0.9059 & 0.0941 & 0.9058 \\
0 & 1.0000 & 0.1270 \\
0.9059 & 0.0941 & 0.9134 \\
0.9059 & 0.0941 & 0.6324 \\
0 & 1.0000 & 0.0975 \\
0.9059 & 0.0941 & 0.2785
\end{tabular}
```


## ClassificationPartitionedModel.kfoldMargin

| 0.9059 | 0.0941 | 0.5469 |
| :--- | :--- | :--- |
| 0.9059 | 0.0941 | 0.9575 |
| 0.9059 | 0.0941 | 0.9649 |

Indeed, every value of $X$ (the rightmost column) that is less than 0.15 has associated scores (the left and center columns) of 0 and 1, while the other values of $X$ have associated scores of 0.91 and 0.09 . The difference (score 0.09 instead of the expected .06) is due to a statistical fluctuation: there are 8 observations in $X$ in the range (.95,1) instead of the expected 5 observations.

## Examples <br> Find the k-fold margins for an ensemble that classifies the ionosphere data:

```
load ionosphere
ens = fitensemble(X,Y,'AdaBoostM1',100,'Tree');
cvens = crossval(ens);
M = kfoldMargin(cvens);
[min(M) mean(M) max(M)]
ans =
    -7.6394 7.3469 22.4833
kfoldPredict | kfoldEdge | kfoldLoss | kfoldfun | crossval
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```

See Also

## ClassificationPartitionedModel.kfoldPredict

| Purpose | Predict response for observations not used for training |
| :--- | :--- |
| Syntax | $\begin{array}{l}\text { label }=\text { kfoldPredict }(o b j) \\ \\ \\ \\ \\ \\ \end{array}$ label, score $]=$ kfoldPredict $(o b j)$ |

Description

Input
Arguments

Output
Arguments
label $=$ kfoldPredict (obj) returns class labels predicted by obj, a cross-validated classification. For every fold, kfoldPredict predicts class labels for in-fold observations using a model trained on out-of-fold observations.
[label,score] = kfoldPredict(obj) returns the predicted classification scores for in-fold observations using a model trained on out-of-fold observations.
[label, score, cost] = kfoldPredict(obj) returns misclassification costs when obj is a cross-validated discriminant analysis classifier.
obj
Object of class ClassificationPartitionedModel or ClassificationPartitionedEnsemble.
label
Vector of class labels of the same type as the response data used in training obj. Each entry of label corresponds to a predicted class label for the corresponding row of $X$.

## score

Numeric matrix of size N -by-K, where N is the number of observations (rows) in obj. X , and K is the number of classes (in obj.ClassNames). score(i,j) represents the confidence that row i of obj. X is of class j . For details, see "Definitions" on page 20-845.

## ClassificationPartitionedModel.kfoldPredict

Numeric matrix of misclassification costs of size N-by-K. cost (i, j) is the average misclassification cost of predicting that row i of obj. X is of class j . Applies only to discriminant analysis classification.

## Definitions Cost (discriminant analysis)

The average misclassification cost is the mean misclassification cost for predictions made by the cross-validated classifiers trained on out-of-fold observations. The matrix of expected costs per observation is defined in "Cost" on page 12-8.

## Score (discriminant analysis)

For discriminant analysis, the score of a classification is the posterior probability of the classification. For the definition of posterior probability in discriminant analysis, see "Posterior Probability" on page 12-7.

## Score (ensemble)

For ensembles, a classification score represents the confidence of a classification into a class. The higher the score, the higher the confidence.

Different ensemble algorithms have different definitions for their scores. Furthermore, the range of scores depends on ensemble type. For example:

- AdaBoostM1 scores range from $-\infty$ to $\infty$.
- Bag scores range from 0 to 1.


## Score (tree)

For trees, the score of a classification of a leaf node is the posterior probability of the classification at that node. The posterior probability of the classification at a node is the number of training sequences that lead to that node with the classification, divided by the number of training sequences that lead to that node.

## ClassificationPartitionedModel.kfoldPredict

For example, consider classifying a predictor $X$ as true when $X<0.15$ or $X>0.95$, and $X$ is false otherwise.

1
Generate 100 random points and classify them:

```
rng(0,'twister') % for reproducibility
X = rand(100,1);
Y = (abs(X - . 55) > .4);
tree = ClassificationTree.fit(X,Y);
view(tree,'mode','graph')
```

| Click to display: | Identity | - | Magnification: | 100\% | $\checkmark$ | Pruning level: | 0 of 2 | 븐 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |



2
Prune the tree:

```
tree1 = prune(tree,'level',1);
view(tree1,'mode','graph')
```


## ClassificationPartitionedModel.kfoldPredict



The pruned tree correctly classifies observations that are less than 0.15 as true. It also correctly classifies observations from .15 to .94 as false. However, it incorrectly classifies observations that are greater than . 94 as false. Therefore, the score for observations that are greater than . 15 should be about $.05 / .85=.06$ for true, and about $.8 / .85=.94$ for false.

3
Compute the prediction scores for the first 10 rows of X :

```
[~,score] = predict(tree1,X(1:10));
[score X(1:10,:)]
ans =
\begin{tabular}{rrr}
0.9059 & 0.0941 & 0.8147 \\
0.9059 & 0.0941 & 0.9058 \\
0 & 1.0000 & 0.1270 \\
0.9059 & 0.0941 & 0.9134 \\
0.9059 & 0.0941 & 0.6324 \\
0 & 1.0000 & 0.0975 \\
0.9059 & 0.0941 & 0.2785
\end{tabular}
```


## ClassificationPartitionedModel.kfoldPredict

| 0.9059 | 0.0941 | 0.5469 |
| :--- | :--- | :--- |
| 0.9059 | 0.0941 | 0.9575 |
| 0.9059 | 0.0941 | 0.9649 |

Indeed, every value of $X$ (the rightmost column) that is less than 0.15 has associated scores (the left and center columns) of 0 and 1 , while the other values of $X$ have associated scores of 0.91 and 0.09 . The difference (score 0.09 instead of the expected .06) is due to a statistical fluctuation: there are 8 observations in $X$ in the range (.95,1) instead of the expected 5 observations.

```
Examples Find the cross-validation predictions for a model based on the Fisher
iris data:
```

```
load fisheriris
```

load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
cvens = crossval(ens);
cvens = crossval(ens);
[elabel escore] = kfoldPredict(cvens);
[elabel escore] = kfoldPredict(cvens);
max(escore)
max(escore)
ans =
ans =
9.3634 8.5624 9.3981
9.3634 8.5624 9.3981
min(escore)
min(escore)
ans =
ans =
0.0017 3.7518 0.8911
0.0017 3.7518 0.8911
See Also kfoldEdge | kfoldMargin | kfoldLoss | crossval
How To

- Chapter 13, "Nonparametric Supervised Learning"

```

\section*{RegressionPartitionedModel.kfoldPredict}

Purpose
Predict response for observations not used for training.

Syntax
Description

\section*{Input \\ Arguments}
yfit = kfoldPredict(obj)
yfit \(=k f o l d P r e d i c t(o b j)\) returns the predicted values for the responses of the training data based on obj, an object trained on out-of-fold observations.

\section*{obj}

Object of class RegressionPartitionedModel. Create obj with RegressionTree.fit or fitensemble along with one of the cross-validation options: 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. Alternatively, create obj from a regression tree or regression ensemble with crossval.

\section*{Output Arguments}

\section*{Examples}
yfit
A vector of predicted values for the response data based on a model trained on out-of-fold observations.

Construct a partitioned regression model, and examine the cross-validation loss. The cross-validation loss is the mean squared error between yfit and the true response data:
```

load carsmall
XX = [Cylinders Displacement Horsepower Weight];
YY = MPG;
tree = RegressionTree.fit(XX,YY);
cvmodel = crossval(tree);
L = kfoldLoss(cvmodel)
L =
26.5271
yfit = kfoldPredict(cvmodel);
mean( (yfit - tree.Y).^2 )

```

\section*{RegressionPartitionedModel.kfoldPredict}

\section*{ans =}
26.5271

See Also kfoldLoss
How To - Chapter 13, "Nonparametric Supervised Learning"
```

Purpose K
Syntax IDX = kmeans(X,k)
[IDX,C] = kmeans(X,k)
[IDX,C,sumd] = kmeans(X,k)
[IDX,C,sumd,D] = kmeans(X,k)
[...] = kmeans(...,param1,val1,param2,val2,...)

```

Description
IDX \(=\) kmeans \((X, k)\) partitions the points in the \(n\)-by- \(p\) data matrix \(X\) into k clusters. This iterative partitioning minimizes the sum, over all clusters, of the within-cluster sums of point-to-cluster-centroid distances. Rows of \(X\) correspond to points, columns correspond to variables. kmeans returns an \(n\)-by- 1 vector IDX containing the cluster indices of each point. By default, kmeans uses squared Euclidean distances. When X is a vector, kmeans treats it as an \(n\)-by- 1 data matrix, regardless of its orientation.
[IDX,C] = kmeans ( \(\mathrm{X}, \mathrm{k}\) ) returns the k cluster centroid locations in the k -by-p matrix C .
[IDX,C,sumd] = kmeans(X,k) returns the within-cluster sums of point-to-centroid distances in the 1 -by-k vector sumd.
[IDX, C, sumd, D] = kmeans ( \(\mathrm{X}, \mathrm{k}\) ) returns distances from each point to every centroid in the \(n\)-by- \(k\) matrix \(D\).
[...] = kmeans(...,param1,val1,param2,val2,...) enables you to specify optional parameter/value pairs to control the iterative algorithm used by kmeans. Valid parameter strings are listed in the following table.
\begin{tabular}{l|l|l}
\hline Parameter & Value \\
\hline 'distance' & \begin{tabular}{l} 
Distance measure, in p-dimensional space. kmeans \\
minimizes with respect to this parameter. kmeans \\
computes centroid clusters differently for the \\
different supported distance measures.
\end{tabular} \\
\hline & 'sqEuclidean' & \begin{tabular}{l} 
Squared Euclidean distance \\
(default). Each centroid is the \\
mean of the points in that cluster.
\end{tabular} \\
\hline & 'cityblock' & \begin{tabular}{l} 
Sum of absolute differences, i.e., \\
the L1 distance. Each centroid \\
is the component-wise median of \\
the points in that cluster.
\end{tabular} \\
\hline & 'cosine' & \begin{tabular}{l} 
One minus the cosine of the \\
included angle between points \\
(treated as vectors). Each \\
centroid is the mean of the points \\
in that cluster, after normalizing \\
those points to unit Euclidean \\
length.
\end{tabular} \\
\hline & 'correlation' & \begin{tabular}{l} 
One minus the sample correlation \\
between points (treated as \\
sequences of values). Each \\
centroid is the component-wise \\
mean of the points in that cluster, \\
after centering and normalizing \\
those points to zero mean and \\
unit standard deviation.
\end{tabular} \\
\hline & 'Hamming' & \begin{tabular}{l} 
Percentage of bits that differ (only \\
suitable for binary data). Each \\
centroid is the component-wise \\
median of points in that cluster.
\end{tabular} \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|}
\hline Parameter & \multicolumn{2}{|l|}{Value} \\
\hline \multirow[t]{4}{*}{'emptyaction'} & \multicolumn{2}{|l|}{Action to take if a cluster loses all its member observations.} \\
\hline & 'error' & Treat an empty cluster as an error (default). \\
\hline & 'drop ' & Remove any clusters that become empty. kmeans sets the corresponding return values in C and D to NaN. \\
\hline & 'singleton' & Create a new cluster consisting of the one point furthest from its centroid. \\
\hline 'onlinephase' & \multicolumn{2}{|l|}{Flag indicating whether kmeans should perform an online update phase in addition to a batch update phase. The online phase can be time consuming for large data sets, but guarantees a solution that is a local minimum of the distance criterion, that is, a partition of the data where moving any single point to a different cluster increases the total sum of distances.} \\
\hline & 'on' & Perform online update (default). \\
\hline & 'off ' & Do not perform online update. \\
\hline 'options' & \multicolumn{2}{|l|}{Options for the iterative algorithm used to minimize the fitting criterion, as created by statset.} \\
\hline 'replicates' & \multicolumn{2}{|l|}{Number of times to repeat the clustering, each with a new set of initial cluster centroid positions. kmeans returns the solution with the lowest value for sumd. You can supply 'replicates' implicitly by supplying a 3 D array as the value for the 'start' parameter.} \\
\hline
\end{tabular}
\begin{tabular}{l|l|l}
\hline Parameter & Value \\
\hline \multirow{6}{*}{ 'start' } & \begin{tabular}{l} 
Method used to choose the initial cluster centroid \\
positions, sometimes known as seeds.
\end{tabular} \\
\cline { 2 - 4 } & 'sample' & \begin{tabular}{l} 
Select k observations from X at \\
random (default).
\end{tabular} \\
\cline { 2 - 4 } & 'uniform' & \begin{tabular}{l} 
Select k points uniformly at \\
random from the range of X. Not \\
valid with Hamming distance.
\end{tabular} \\
\cline { 2 - 4 } & 'cluster' & \begin{tabular}{l} 
Perform a preliminary clustering \\
phase on a random 10\% \\
subsample of X. This preliminary \\
phase is itself initialized using
\end{tabular} \\
'sample'.
\end{tabular}

\section*{Algorithms}
kmeans uses a two-phase iterative algorithm to minimize the sum of point-to-centroid distances, summed over all \(k\) clusters:

1 The first phase uses batch updates, where each iteration consists of reassigning points to their nearest cluster centroid, all at once, followed by recalculation of cluster centroids. This phase occasionally does not converge to solution that is a local minimum, that is, a partition of the data where moving any single point to a different cluster increases the total sum of distances. This is more likely for small data sets. The batch phase is fast, but potentially only approximates a solution as a starting point for the second phase.

2 The second phase uses online updates, where points are individually reassigned if doing so will reduce the sum of distances, and cluster centroids are recomputed after each reassignment. Each iteration during the second phase consists of one pass though all the points. The second phase will converge to a local minimum, although there may be other local minima with lower total sum of distances. The problem of finding the global minimum can only be solved in general by an exhaustive (or clever, or lucky) choice of starting points, but using several replicates with random starting points typically results in a solution that is a global minimum.

\section*{References}

\section*{Examples}
[1] Seber, G. A. F. Multivariate Observations. Hoboken, NJ: John Wiley \& Sons, Inc., 1984.
[2] Spath, H. Cluster Dissection and Analysis: Theory, FORTRAN Programs, Examples. Translated by J. Goldschmidt. New York: Halsted Press, 1985.

The following creates two clusters from separated random data:
```

X = [randn(100,2)+ones(100,2);...
randn(100,2)-ones(100,2)];
opts = statset('Display','final');
[idx,ctrs] = kmeans(X,2,...
'Distance','city',...
'Replicates',5,...
'Options',opts);
5 iterations, total sum of distances = 284.671
4 iterations, total sum of distances = 284.671
4 iterations, total sum of distances = 284.671
3 iterations, total sum of distances = 284.671
3 iterations, total sum of distances = 284.671

```
```

plot(X(idx==1,1),X(idx==1,2),'r.','MarkerSize',12)

```
plot(X(idx==1,1),X(idx==1,2),'r.','MarkerSize',12)
hold on
hold on
plot(X(idx==2,1),X(idx==2,2),'b.','MarkerSize',12)
```

plot(X(idx==2,1),X(idx==2,2),'b.','MarkerSize',12)

```
```

plot(ctrs(:,1), ctrs(:,2),'kx',...
'MarkerSize', 12, 'LineWidth', 2)
plot(ctrs(:,1),ctrs(:,2),'ko',...
'MarkerSize',12, 'LineWidth', 2)
legend('Cluster 1','Cluster 2','Centroids',...
'Location','NW')

```


See Also linkage | clusterdata | silhouette

\section*{ExhaustiveSearcher.knnsearch}

\section*{Purpose}

Find \(k\)-nearest neighbors using ExhaustiveSearcher object

\section*{Syntax}

IDX = knnsearch(NS, Y)
[IDX, D] = knnsearch(NS,Y)
[IDX,D] = knnsearch(NS,Y,'Name', Value)

Input Arguments

IDX \(=\) knnsearch ( \(N S, Y\) ) finds the nearest neighbor (closest point) in NS. X for each point in Y. Rows of \(Y\) correspond to observations and columns correspond to features. Y must have the same number of columns as NS.X. IDX is a column vector with \(n y\) rows, where \(n y\) is the number of rows in Y. Each row in IDX contains the index of observation in NS.X which has the smallest distance to the corresponding observation in \(Y\).
[IDX, D] = knnsearch(NS,Y) returns a column vector D containing the distances between each observation in Y and the corresponding closest observation in NS. X. That is, \(D(i)\) is the distance between NS. X(IDX(i),:) and Y(i,:).
[IDX, D] = knnsearch(NS, Y, 'Name', Value) accepts one or more comma-separated argument name/value pairs. Specify Name inside single quotes.

\section*{Name-Value Pair Arguments}

K
A positive integer, \(k\), specifying the number of nearest neighbors in NS.X for each point in Y. Default is 1. IDX and D are \(n y\)-by- \(k\) matrices. D sorts the distances in each row in ascending order. Each row in IDX contains the indices of the \(k\) closest neighbors in NS. \(X\) corresponding to the \(k\) smallest distances in \(D\).

\section*{IncludeTies}

A logical value indicating whether knnsearch includes all the neighbors whose distance values are equal to the Kth smallest distance. If IncludeTies is true, knnsearch includes all these neighbors. In this case, IDX and D are ny-by- 1 cell arrays. Each

\section*{ExhaustiveSearcher.knnsearch}
row in IDX and \(D\) contains a vector with at least \(K\) numeric numbers. \(D\) sorts the distances in each vector in ascending order. Each row in IDX contains the indices of the closest neighbors corresponding to these smallest distances in \(D\).

Default: false

\section*{Distance}
- 'euclidean' - Euclidean distance (default).
- 'seuclidean' - Standardized Euclidean distance. Each coordinate difference between \(X\) and each query point is scaled by dividing by a scale value \(S\). The default value of S is NS.DistParameter if NS.Distance is 'seuclidean', otherwise the default is the standard deviation computed from X, S=nanstd (X). To specify another value for S, use the 'Scale' argument.
- 'cityblock' - City block distance.
- 'chebychev' - Chebychev distance (maximum coordinate difference).
- 'minkowski' - Minkowski distance.
- 'mahalanobis' - Mahalanobis distance, which is computed using a positive definite covariance matrix C . The default value of \(C\) is nancov \((X)\). To change the value of \(C\), use the Cov parameter.
- 'cosine' - One minus the cosine of the included angle between observations (treated as vectors).
- 'correlation' - One minus the sample linear correlation between observations (treated as sequences of values).
- 'spearman' - One minus the sample Spearman's rank correlation between observations (treated as sequences of values).
- 'hamming ' - Hamming distance, which is the percentage of coordinates that differ.
- 'jaccard' - One minus the Jaccard coefficient, which is the percentage of nonzero coordinates that differ.
- custom distance function - A distance function specified using @ (for example, @distfun). A distance function must be of the form function D2 = distfun(ZI, ZJ), taking as arguments a 1-by- \(n\) vector ZI containing a single row of from X or from the query points Y , an \(m 2\)-by- \(n\) matrix ZJ containing multiple rows of \(X\) or \(Y\), and returning an \(m 2\)-by- 1 vector of distances D2, whose \(j\) th element is the distance between the observations ZI and \(\mathrm{ZJ}(j,:)\).

Default is NS.Distance. For more information on these distance metrics, see "Distance Metrics" on page 13-9.

A positive scalar, \(p\), indicating the exponent of the Minkowski distance. This parameter is only valid if knnsearch uses the 'minkowski' distance metric. Default is NS. DistParameter if NS. Distance is 'minkowski' and 2 otherwise.

\section*{Cov}

A positive definite matrix indicating the covariance matrix when computing the Mahalanobis distance. This parameter is only valid when knnsearch uses the 'mahalanobis' distance metric. Default is NS.DistParameter if NS.Distance is 'mahalanobis', or nancov (X) otherwise.

Scale
A vector \(S\) with the length equal to the number of columns in X. Each coordinate of \(X\) and each query point is scaled by the corresponding element of \(S\) when computing the standardized Euclidean distance. This parameter is only valid when Distance is 'seuclidean'. Default is nanstd ( X ).

\section*{ExhaustiveSearcher.knnsearch}

Examples
Create an ExhaustiveSearcher object specifying 'cosine' as the distance metric. Perform a \(k\)-nearest neighbors search on the object using the mahalanobis metric and compare the results:
```

load fisheriris
x = meas(:,3:4);
exhaustiveobj = ExhaustiveSearcher(x,'Distance','cosine')
exhaustiveobj =
ExhaustiveSearcher
Properties:
X: [150x2 double]
Distance: 'cosine'
DistParameter: []
% Perform a knnsearch between x and a query point, using
% first cosine then mahalanobis distance metrics:
newpoint = [5 1.45];
[n,d]=knnsearch(exhaustiveobj, newpoint, 'k', 10);
[nmah,dmah] = knnsearch(exhaustiveobj,newpoint,'k',10,···
'distance','mahalanobis');
% Visualize the results of the two different nearest
% neighbors searches:
% First plot the training data:
gscatter(x(:,1),x(:,2),species)
% Plot an X for the query point:
line(newpoint(1),newpoint(2),'marker','x','color','k', ...
'markersize',10,'linewidth',2,'linestyle','none')
% Use circles to denote the cosine nearest neighbors:
line(x(n, 1), x(n,2),'color',[.5 .5 .5],'marker','o',...
'linestyle','none','markersize',10)
% Use pentagrams to denote the mahalanobis nearest neighbors:
line(x(nmah,1),x(nmah,2),'color',[.5 .5 .5],'marker','p',...

```
```

        'linestyle','none','markersize',10)
    legend('setosa','versicolor','virginica','query point',...
    'cosine','mahalanobis')
    set(legend,'location','best')
    ```


\section*{Algorithms}

See Also

How To

For information on a specific search algorithm, see "Distance Metrics" on page 13-9.
createns | ExhaustiveSearcher | KDTreeSearcher.knnsearch | knnsearch | rangesearch
- " \(k\)-Nearest Neighbor Search and Radius Search" on page 13-12

\section*{ExhaustiveSearcher.knnsearch}
- "Distance Metrics" on page 13-9

\section*{Purpose}

Find \(k\)-nearest neighbors using KDTreeSearcher object

IDX = knnsearch(NS, Y)
[IDX, D] = knnsearch(NS,Y)
[IDX,D] = knnsearch(NS,Y,'Name', Value)

Description

\section*{Input Arguments}

IDX = knnsearch(NS, \(Y\) ) finds the nearest neighbor (closest point) in NS. X for each point in Y. Rows of \(Y\) correspond to observations and columns correspond to features. Y must have the same number of columns as NS.X. IDX is a column vector with ny rows, where \(n y\) is the number of rows in Y. Each row in IDX contains the index of observation in NS. X which has the smallest distance to the corresponding observation in Y .
[IDX, D] = knnsearch(NS,Y) returns a column vector D containing the distances between each observation in \(Y\) and the corresponding closest observation in NS.X. That is, \(D(i)\) is the distance between NS. X(IDX(i),:) and Y(i,:).
[IDX, D] = knnsearch(NS, Y, 'Name', Value) accepts one or more comma-separated name/value pairs. Specify Name inside single quotes.

\section*{Name-Value Pair Arguments}

\section*{K}

A positive integer, \(k\), specifying the number of nearest neighbors in NS.X for each point in Y. Default is 1. IDX and D are \(n y\)-by- \(k\) matrices. D sorts the distances in each row in ascending order. Each row in IDX contains the indices of the \(k\) closest neighbors in NS. X corresponding to the \(k\) smallest distances in D.

\section*{Distance}

Select one of the following distance algorithms.
- 'euclidean' - Euclidean distance (default).
- 'cityblock' - City block distance.

\section*{KDTreeSearcher.knnsearch}
- 'chebychev' - Chebychev distance (maximum coordinate difference).
- 'minkowski' - Minkowski distance.

Default is NS.Distance. For more information on these distance metrics, see "Distance Metrics" on page 13-9.

\section*{IncludeTies}

A logical value indicating whether knnsearch includes all the neighbors whose distance values are equal to the Kth smallest distance. If IncludeTies is true, knnsearch includes all these neighbors. In this case, IDX and D are \(n y\)-by- 1 cell arrays. Each row in IDX and \(D\) contains a vector with at least \(K\) numeric numbers. \(D\) sorts the distances in each vector in ascending order. Each row in IDX contains the indices of the closest neighbors corresponding to these smallest distances in \(D\).

Default: false
P
A positive scalar, \(p\), indicating the exponent of the Minkowski distance. This parameter is only valid whenthe Distance is 'minkowski'. Default is NS.DistParameter if NS.Distance is 'minkowski' and 2 otherwise.

Create a KDTreeSearcher object specifying 'minkowski' as the distance metric with an exponent of 5 . Perform a \(k\)-nearest neighbors search on the object using the chebychev metric and compare the results:
```

load fisheriris
x = meas(:,3:4);
kdtreeNS = KDTreeSearcher(x,'Distance','minkowski','P',5)
kdtreeNS =
KDTreeSearcher

```
```

Properties:
BucketSize: 50
X: [150x2 double]
Distance: 'minkowski'
DistParameter: 5

```
\% Perform a knnsearch between \(X\) and a query point, using
\% first Minkowski then Chebychev distance metrics:
newpoint \(=\left[\begin{array}{ll}5 & 1.45\end{array}\right]\);
[ \(\mathrm{n}, \mathrm{d}\) ]=knnsearch(kdtreeNS, newpoint, 'k', 10);
[ncb, dcb] = knnsearch(kdtreeNS, newpoint, 'k', 10, ...
    'distance', 'chebychev');
\% Visualize the results of the two different nearest
\% neighbors searches:
\% First plot the training data:
gscatter(x(:,1), x(:,2), species)
\% Zoom in on the points of interest:
set(gca,'xlim',[4.5 5.5],'ylim',[12]); axis square
\% Plot an \(X\) for the query point:
line (newpoint (1), newpoint (2), 'marker', 'x', color', 'k', ...
    'markersize', 10,'linewidth', 2,'linestyle', 'none')
\% Use circles to denote the Minkowski nearest neighbors:

    'linestyle', 'none', 'markersize', 10)
\% Use pentagrams to denote the Chebychev nearest neighbors:

    'linestyle', 'none', 'markersize', 10)
legend('setosa', 'versicolor','virginica', 'query point',...
    'minkowski', 'chebychev')
set(legend, 'location', 'best')

\section*{KDTreeSearcher.knnsearch}


\section*{Algorithms}

How To

References [1] Friedman, J. H., Bentely, J., and Finkel, R. A. (1977). An Algorithm for Finding Best Matches in Logarithmic Expected Time, ACM Transactions on Mathematical Software 3, 209.
See Also

createns | ExhaustiveSearcher | ExhaustiveSearcher.knnsearch |
 knnsearch | rangesearch

For information on a specific search algorithm, see "Distance Metrics" on page 13-9.
- " \(k\)-Nearest Neighbor Search and Radius Search" on page 13-12
- "Distance Metrics" on page 13-9

Purpose
Find \(k\)-nearest neighbors using data

\section*{Syntax}

IDX = knnsearch (X,Y)
[IDX,D] = knnsearch(X,Y)
[IDX, D] = knnsearch(X,Y,'Name', Value)

\section*{Input Arguments}

IDX \(=\) knnsearch \((X, Y)\) finds the nearest neighbor in \(X\) for each point in \(Y\). \(X\) is an \(m x\)-by- \(n\) matrix and \(Y\) is an \(m y\)-by- \(n\) matrix. Rows of \(X\) and Y correspond to observations and columns correspond to variables. IDX is a column vector with \(m y\) rows. Each row in IDX contains the index of nearest neighbor in \(X\) for the corresponding row in \(Y\).
[IDX, D] = knnsearch ( \(\mathrm{X}, \mathrm{Y}\) ) returns an \(m y\)-by-1 vector D containing the distances between each observation in \(Y\) and the corresponding closest observation in \(X\). That is, \(D(i)\) is the distance between \(X(\operatorname{IDX}(i),:)\) and \(Y(i,:)\).
[IDX, D] = knnsearch(X,Y, 'Name', Value) accepts one or more optional comma-separated name/value pairs. Specify Name inside single quotes.
knnsearch does not save a search object. To create a search object, use createns.
- For a fixed positive integer K, knnsearch finds the \(K\) points in \(X\) that are nearest each point in \(Y\). In contrast, for a fixed positive real value \(r\), rangesearch finds all the points in \(X\) that are within a distance \(r\) of each point in \(Y\).

\section*{Name-Value Pair Arguments}

\section*{K}

Positive integer specifying the number of nearest neighbors in X for each point in Y. Default is 1. IDX and D are \(m y\)-by-K matrices. D sorts the distances in each row in ascending order. Each row in IDX contains the indices of the \(K\) closest neighbors in \(X\) corresponding to the K smallest distances in D .

\section*{IncludeTies}

A logical value indicating whether knnsearch includes all the neighbors whose distance values are equal to the Kth smallest distance. If IncludeTies is true, knnsearch includes all these neighbors. In this case, IDX and D are \(m y\)-by- 1 cell arrays. Each row in IDX and D contains a vector with at least K numeric numbers. D sorts the distances in each vector in ascending order. Each row in IDX contains the indices of the closest neighbors corresponding to these smallest distances in D.

Default: false

\section*{NSMethod}

Nearest neighbors search method. Value is either:
- 'kdtree' - Creates and uses a Kd-tree to find nearest neighbors. This is the default value when the number of columns of X is less than \(10, \mathrm{X}\) is not sparse, and the distance measure is one of the following measures. 'kdtree' is only valid when the distance measure is one of the following:
- 'euclidean'
- 'cityblock'
- 'minkowski'
- 'chebychev'
- 'exhaustive' - Uses the exhaustive search algorithm by computing the distance values from all the points in \(X\) to each point in \(Y\) to find nearest neighbors.

\section*{Distance}

A string or a function handle specifying the distance metric. The value can be one of the following:
- 'euclidean' - Euclidean distance (default).
- 'seuclidean' - Standardized Euclidean distance. Each coordinate difference between rows in X and the query matrix is scaled by dividing by the corresponding element of the standard deviation computed from \(X, S=\) nanstd ( \(X\) ). To specify another value for \(S\), use the Scale argument.
- 'cityblock' - City block distance.
- 'chebychev' - Chebychev distance (maximum coordinate difference).
- 'minkowski' - Minkowski distance. The default exponent is 2. To specify a different exponent, use the ' \(P\) ' argument.
- 'mahalanobis' - Mahalanobis distance, computed using a positive definite covariance matrix \(C\). The default value of \(C\) is nancov (X). To change the value of C , use the Cov parameter.
- 'cosine' - 1 minus the cosine of the included angle between observations (treated as vectors).
- 'correlation' - One minus the sample linear correlation between observations (treated as sequences of values).
- 'spearman' - One minus the sample Spearman's rank correlation between observations (treated as sequences of values).
- 'hamming' - Hamming distance, which is the percentage of coordinates that differ.
- 'jaccard' - One minus the Jaccard coefficient, which is the percentage of nonzero coordinates that differ.
- custom distance function - A distance function specified using @ (for example, @distfun). A distance function must be of the form function D2 \(=\) distfun(ZI, ZJ), taking as arguments a 1-by- \(n\) vector ZI containing a single row of X or Y , an \(m 2\)-by- \(n\) matrix ZJ containing multiple rows of X or Y , and returning an \(m 2\)-by- 1 vector of distances D2, whose \(j\) th element is the distance between the observations ZI and \(\mathrm{ZJ}(j,:)\).

For more information on these distance metrics, see "Distance Metrics" on page 13-9.

P
A positive scalar, \(p\), indicating the exponent of the Minkowski distance. This parameter is only valid if the Distance is 'minkowski'. Default is 2.

Cov
A positive definite matrix indicating the covariance matrix when computing the Mahalanobis distance. This parameter is only valid when Distance is 'mahalanobis'. Default is nancov ( X ).

\section*{Scale}

A vector \(S\) containing nonnegative values, with length equal to the number of columns in \(X\). Each coordinate of \(X\) and each query point is scaled by the corresponding element of \(S\) when computing the standardized Euclidean distance. This argument is only valid when Distance is 'seuclidean'. Default is nanstd ( \(X\) ).

\section*{BucketSize}

The maximum number of data points in the leaf node of the \(k \mathrm{~d}\)-tree. This argument is only meaningful when using the \(k \mathrm{~d}\)-tree search method. Default is 50 .

\section*{Examples}

Find the 10 nearest neighbors in x to each point in y using first the 'minkowski' distance metric with a \(p\) value of 5 , and then using the 'chebychev' distance metric. Visually compare the results:
```

load fisheriris
x = meas(:,3:4);
y = [5 1.45;6 2;2.75 .75];
% Perform a knnsearch between x and the query points in y,
% using first Minkowski then Chebychev distance metrics.
[n,d]=knnsearch(x,y,'k',10,'distance','minkowski','p',5);
[ncb,dcb] = knnsearch(x,y,'k',10,...

```
```

    'distance','chebychev');
    % Visualize the results of the two different nearest
% neighbors searches.
% First plot the training data.
gscatter(x(:,1),x(:,2),species)
% Plot an X for the query points.
line(y(:,1),y(:,2),'marker','x','color','k',...
'markersize', 10,'linewidth',2,'linestyle','none')
% Use circles to denote the Minkowski nearest neighbors.
line(x(n,1),x(n,2),'color',[.5 .5 .5],'marker','o',...
'linestyle','none','markersize',10)
% Use pentagrams to denote the Chebychev nearest neighbors.
line(x(ncb,1),x(ncb,2),'color',[.5 .5 .5],'marker','p',...
'linestyle','none','markersize',10)
legend('setosa','versicolor','virginica','query point',...
'minkowski','chebychev')
set(legend,'location','best')

```


\section*{Algorithms}

\section*{References}

See Also

How To

For information on a specific search algorithm, see "Distance Metrics" on page 13-9.
[1] Friedman, J. H., Bentely, J., and Finkel, R. A. (1977) An Algorithm for Finding Best Matches in Logarithmic Expected Time, ACM Transactions on Mathematical Software 3, 209.
createns | KDTreeSearcher.knnsearch | ExhaustiveSearcher.knnsearch | rangesearch
- " \(k\)-Nearest Neighbor Search and Radius Search" on page 13-12

\section*{kruskalwallis}
Purpose Kruskal-Wallis test
\begin{tabular}{|c|c|}
\hline Syntax & \(\mathrm{p}=\) kruskalwallis(X) \\
\hline & \(\mathrm{p}=\) kruskalwallis(X, group) \\
\hline & \(\mathrm{p}=\) kruskalwallis(X,group,displayopt) \\
\hline & [ p ,table] = kruskalwallis(...) \\
\hline & [p,table,stats] = kruskalwallis( \\
\hline
\end{tabular}

\section*{Description}
p = kruskalwallis(X) performs a Kruskal-Wallis test to compare samples from two or more groups. Each column of the \(m\)-by- \(n\) matrix X represents an independent sample containing \(m\) mutually independent observations. The function compares the medians of the samples in X , and returns the \(p\) value for the null hypothesis that all samples are drawn from the same population (or equivalently, from different populations with the same distribution). Note that the Kruskal-Wallis test is a nonparametric version of the classical one-way ANOVA, and an extension of the Wilcoxon rank sum test to more than two groups.

If the \(p\) value is near zero, this casts doubt on the null hypothesis and suggests that at least one sample median is significantly different from the others. The choice of a critical \(p\) value to determine whether the result is judged statistically significant is left to the researcher. It is common to declare a result significant if the \(p\) value is less than 0.05 or 0.01.

The kruskalwallis function displays two figures. The first figure is a standard ANOVA table, calculated using the ranks of the data rather than their numeric values. Ranks are found by ordering the data from smallest to largest across all groups, and taking the numeric index of this ordering. The rank for a tied observation is equal to the average rank of all observations tied with it. For example, the following table shows the ranks for a small sample.
\begin{tabular}{l|l|l|l|l|l|l|l}
\hline X value & 1.4 & 2.7 & 1.6 & 1.6 & 3.3 & 0.9 & 1.1 \\
\hline Rank & 3 & 6 & 4.5 & 4.5 & 7 & 1 & 2 \\
\hline
\end{tabular}

The entries in the ANOVA table are the usual sums of squares, degrees of freedom, and other quantities calculated on the ranks. The usual \(F\) statistic is replaced by a chi-square statistic. The \(p\) value measures the significance of the chi-square statistic.
The second figure displays box plots of each column of \(X\) (not the ranks of \(X\) ).
\(p\) = kruskalwallis(X, group) uses the values in group (a character array or cell array) as labels for the box plot of the samples in \(X\), when \(X\) is a matrix. Each row of group contains the label for the data in the corresponding column of \(X\), so group must have length equal to the number of columns in \(X\).

When \(X\) is a vector, kruskalwallis performs a Kruskal-Wallis test on the samples contained in \(X\), as indexed by input group (a categorical variable, vector, character array, or cell array). Each element in group identifies the group (i.e., sample) to which the corresponding element in vector \(X\) belongs, so group must have the same length as \(X\). The labels contained in group are also used to annotate the box plot. (See "Grouped Data" on page 2-34.)
It is not necessary to label samples sequentially (1, 2, 3, ...). For example, if \(X\) contains measurements taken at three different temperatures, \(-27^{\circ}, 65^{\circ}\), and \(110^{\circ}\), you could use these numbers as the sample labels in group. If a row of group contains an empty cell or empty string, that row and the corresponding observation in X are disregarded. NaNs in either input are similarly ignored.
\(\mathrm{p}=\) kruskalwallis(X,group,displayopt) enables the table and box plot displays when displayopt is 'on' (default) and suppresses the displays when displayopt is 'off'.
[p,table] = kruskalwallis(...) returns the ANOVA table (including column and row labels) in cell array table.
[p,table,stats] = kruskalwallis(...) returns a stats structure that you can use to perform a follow-up multiple comparison test. The kruskalwallis test evaluates the hypothesis that all samples come from populations that have the same median, against the alternative

\section*{kruskalwallis}
that the medians are not all the same. Sometimes it is preferable to perform a test to determine which pairs are significantly different, and which are not. You can use the multcompare function to perform such tests by supplying the stats structure as input.

\section*{Assumptions}

The Kruskal-Wallis test makes the following assumptions about the data in \(X\) :
- All samples come from populations having the same continuous distribution, apart from possibly different locations due to group effects.
- All observations are mutually independent.

The classical one-way ANOVA test replaces the first assumption with the stronger assumption that the populations have normal distributions.

\section*{Examples}

This example compares the material strength study used with the anova1 function, to see if the nonparametric Kruskal-Wallis procedure leads to the same conclusion. The example studies the strength of beams made from three alloys:
```

strength = [l 82 86 79 83 84 85 86 87 74 82 ...
78 75 76 77 79 79 77 78 82 79];
alloy = {'st','st','st','st','st','st','st','st',...
'al1','al1','al1','al1','al1','al1',...
'al2','al2','al2','al2','al2','al2'};

```

This example uses both classical and Kruskal-Wallis ANOVA, omitting displays:
```

anova1(strength,alloy,'off')
ans =
1.5264e-004
kruskalwallis(strength,alloy,'off')

```
```

ans =
0.0018

```

Both tests find that the three alloys are significantly different, though the result is less significant according to the Kruskal-Wallis test. It is typical that when a data set has a reasonable fit to the normal distribution, the classical ANOVA test is more sensitive to differences between groups.

To understand when a nonparametric test may be more appropriate, let's see how the tests behave when the distribution is not normal. You can simulate this by replacing one of the values by an extreme value (an outlier).
```

strength (20) =120;
anova1(strength,alloy,'off')
ans =
0.2501
kruskalwallis(strength, alloy, 'off')
ans =
0.0060

```

Now the classical ANOVA test does not find a significant difference, but the nonparametric procedure does. This illustrates one of the properties of nonparametric procedures - they are often not severely affected by changes in a small portion of the data.

\section*{References}
[1] Gibbons, J. D. Nonparametric Statistical Inference. New York: Marcel Dekker, 1985.
[2] Hollander, M., and D. A. Wolfe. Nonparametric Statistical Methods. Hoboken, NJ: John Wiley \& Sons, Inc., 1999.

\author{
See Also
}

How To
anova1 | boxplot | friedman
- "Grouped Data" on page 2-34

\section*{kruskalwallis}
- multcompare
- ranksum

Purpose Kernel smoothing density estimate
Syntax \(\quad[f, x i]=\) ksdensity \((x)\)
f = ksdensity(x,xi)
ksdensity(...)
ksdensity (ax,...)
[f,xi,u] = ksdensity(...)
[...] = ksdensity(...,'Name', value)
Description
[ \(\mathrm{f}, \mathrm{xi} \mathrm{x}=\mathrm{ksdensity}(\mathrm{x})\) computes a probability density estimate of the sample in the vector \(x . f\) is the vector of density values evaluated at the points in xi. The estimate is based on a normal kernel function, using a window parameter (width) that is a function of the number of points in \(x\). The density is evaluated at 100 equally spaced points that cover the range of the data in x . ksdensity works best with continuously distributed samples.
\(f=k s d e n s i t y(x, x i)\) specifies the vector xi of values, where the density estimate is to be evaluated.
ksdensity (...) without output arguments produces a plot of the results.
ksdensity (ax, ...) plots into axes ax instead of gca.
[f,xi,u] = ksdensity(...) also returns the width of the kernel-smoothing window.
[...] = ksdensity(...,'Name', value) specifies one or more optional parameter name/value pairs to control the density estimation. Specify Name inside single quotes. Valid parameter strings and their possible values are as follows:

\section*{ksdensity}
\begin{tabular}{l|l}
\hline Name & Value \\
\hline censoring & \begin{tabular}{l} 
A logical vector of the same length as x, indicating \\
which entries are censoring times. Default is no \\
censoring.
\end{tabular} \\
\hline kernel & \begin{tabular}{l} 
The type of kernel smoother to use. Choose the value \\
as \\
- ' normal' (default) \\
- 'box' \\
- 'triangle' \\
- 'epanechnikov' \\
Alternatively, you can specify some other function, \\
as a function handle or as a string, e.g., @normpdf \\
or 'normpdf '. The function must take a single \\
argument that is an array of distances between data \\
values and places where the density is evaluated. It \\
must return an array of the same size containing \\
corresponding values of the kernel function.
\end{tabular} \\
\hline npoints & \begin{tabular}{l} 
The number of equally spaced points in xi. Default \\
is 100.
\end{tabular} \\
\hline support & \begin{tabular}{l} 
- 'unbounded ' allows the density to extend over the \\
whole real line (default). \\
- 'positive ' restricts the density to positive \\
values.
\end{tabular} \\
\hline - A two-element vector gives finite lower and upper \\
bounds for the support of the density.
\end{tabular}
\begin{tabular}{l|l}
\hline Name & Value \\
\hline width & \begin{tabular}{l} 
The bandwidth of the kernel-smoothing window. The \\
default is optimal for estimating normal densities, \\
but you may want to choose a smaller value to reveal \\
features such as multiple modes.
\end{tabular} \\
\hline function & The function type to estimate, chosen from among \\
- 'pdf' - density \\
- 'cdf' - cumulative probability \\
- 'icdf' - inverse cumulative probability \\
- 'survivor' — survivor \\
- 'cumhazard' - cumulative hazard functions
\end{tabular}

In place of the kernel functions listed above, you can specify another kernel function by using @ (such as @normpdf) or quotes (such as 'normpdf'). ksdensity calls the function with a single argument that is an array containing distances between data values in \(x\) and locations in xi where the density is evaluated. The function must return an array of the same size containing corresponding values of the kernel function. When the function parameter value is 'pdf', this kernel function returns density values, otherwise it returns cumulative probability values. Specifying a custom kernel when the function parameter value is 'icdf' returns an error.

If the support parameter is 'positive', ksdensity transforms x using a \(\log\) function, estimates the density of the transformed values, and transforms back to the original scale. If support is a vector [L U],
ksdensity uses the transformation \(\log ((X-L) /(U-X))\). The width parameter and \(u\) outputs are on the scale of the transformed values.

\section*{Examples}

Generate a mixture of two normal distributions and plots the estimated density:


Generate a mixture of two normal distributions, and plot the estimated cumulative distribution at a specified set of values:
```

x = [randn(30,1); 5+randn(30,1)];

```
```

xi = linspace(-10,15,201);
f = ksdensity(x,xi,'function','cdf');
plot(xi,f);

```


Generate a mixture of two normal distributions, and plot the estimated inverse cumulative distribution function at a specified set of values:
```

x = [randn(30,1); 5+randn(30,1)];
yi = linspace(.01,.99,99);
g = ksdensity(x,yi,'function','icdf');

```

\section*{ksdensity}
plot(yi,g);


References [1] Bowman, A. W., and A. Azzalini. Applied Smoothing Techniques for Data Analysis. New York: Oxford University Press, 1997.

Purpose
Syntax

Description

One-sample Kolmogorov-Smirnov test
```

h = kstest(x)
h = kstest(x,CDF)
h = kstest(x,CDF,alpha)
h = kstest(x,CDF,alpha,type)
[h,p,ksstat,cv] = kstest(...)

```
\(h=k s t e s t(x)\) performs a Kolmogorov-Smirnov test to compare the values in the data vector \(x\) to a standard normal distribution. The null hypothesis is that \(x\) has a standard normal distribution. The alternative hypothesis is that \(x\) does not have that distribution. The result h is 1 if the test rejects the null hypothesis at the \(5 \%\) significance level, 0 otherwise.

The test statistic is:
\[
\max (|F(x)-G(x)|)
\]
where \(F(x)\) is the empirical cdf and \(G(x)\) is the standard normal cdf. \(\mathrm{h}=\) kstest \((\mathrm{x}, \mathrm{CDF})\) compares the distribution of x to the hypothesized continuous distribution defined by CDF, which is either a two-column matrix or a ProbDist object of the ProbDistUnivParam class or ProbDistUnivKernel class. When CDF is a matrix, column 1 contains a set of possible \(x\) values, and column 2 contains the corresponding hypothesized cumulative distribution function values \(G(x)\). If possible, define CDF so that column 1 contains the values in \(x\). If there are values in x not found in column 1 of CDF, kstest approximates \(G(x)\) by interpolation. All values in \(x\) must lie in the interval between the smallest and largest values in the first column of CDF. If the second argument is empty ([ ]), kstest uses the standard normal distribution.
The Kolmogorov-Smirnov test requires that CDF be predetermined. It is not accurate if CDF is estimated from the data. To test x against a normal distribution without specifying the parameters, use lillietest instead.

\section*{kstest}
\(h=k s t e s t(x, C D F, a l p h a)\) specifies the significance level alpha for the test. The default is 0.05 .
\(h=k s t e s t(x, C D F\), alpha, type) specifies the type of test using one of the following values for the string type:
- 'unequal ' - Tests the alternative hypothesis that the population cdf is unequal to the specified CDF. This is the default.
- 'larger' - Tests the alternative hypothesis that the population cdf is larger than the specified CDF. The test statistic does not use the absolute value.
- 'smaller' - Tests the alternative hypothesis that the population cdf is smaller than the specified CDF. The test statistic does not use the absolute value.
[h, p,ksstat, cv] = kstest(...) also returns the \(p\) value p , the test statistic ksstat, and the cutoff value cv for determining if ksstat is significant.

\section*{Examples}

Generate evenly spaced numbers and perform a Kolmogorov-Smirnov test to see if they come from a standard normal distribution:
```

$x=-2: 1: 4$
x =
$\begin{array}{lllllll}-2 & -1 & 0 & 1 & 2 & 3 & 4\end{array}$
[h,p,k,c] = kstest(x, [],0.05,0)
$\mathrm{h}=$
0
p =
0.13632
$\mathrm{k}=$
0.41277
c $=$
0.48342

```

The test fails to reject the null hypothesis that the values come from a standard normal distribution. This illustrates the difficulty of testing normality in small samples. (The Lilliefors test, implemented by the Statistics Toolbox function lillietest, may be more appropriate.)
The following figure illustrates the test statistic:
```

xx = -3:.1:5;
F = cdfplot(x);
hold on
G = plot(xx,normcdf(xx),'r-');
set(F,'LineWidth',2)
set(G,'LineWidth',2)
legend([F G],...
'Empirical','Standard Normal',...
'Location', 'NW')

```


The test statistic k is the maximum difference between the curves.
Setting type to 'smaller' tests the alternative that the population cdf is smaller than the normal cdf:
```

[h,p,ksstat] = kstest(x,[],0.05,'smaller')
h =
0
p =
0.068181
k =
0.41277

```

The test statistic is the same as before, but the \(p\) value is smaller.

Setting type to 'larger' changes the test statistic:
```

[h,p,k] = kstest(x,[],0.05,'larger')
$\mathrm{h}=$
0
$p=$
0.77533
$\mathrm{k}=$
0.12706

```

\section*{References [1] Massey, F. J. "The Kolmogorov-Smirnov Test for Goodness of Fit." Journal of the American Statistical Association. Vol. 46, No. 253, 1951, pp. 68-78.}
[2] Miller, L. H. "Table of Percentage Points of Kolmogorov Statistics." Journal of the American Statistical Association. Vol. 51, No. 273, 1956, pp. 111-121.
[3] Marsaglia, G., W. Tsang, and J. Wang. "Evaluating Kolmogorov's Distribution." Journal of Statistical Software. Vol. 8, Issue 18, 2003.

How To •kstest2
- lillietest

Purpose Two-sample Kolmogorov-Smirnov test
Syntax \(\quad h=\operatorname{kstest} 2(x 1, x 2)\)
h = kstest2(x1, x2,alpha,type)
[h,p] = kstest2(...)
[h,p,ks2stat] = kstest2(...)

\section*{Description}
\(h=k s t e s t 2(x 1, x 2)\) performs a two-sample Kolmogorov-Smirnov test to compare the distributions of the values in the two data vectors \(x 1\) and \(x 2\). The null hypothesis is that \(x 1\) and \(x 2\) are from the same continuous distribution. The alternative hypothesis is that they are from different continuous distributions. The result h is 1 if the test rejects the null hypothesis at the \(5 \%\) significance level; 0 otherwise.

The test statistic is:
\[
\max (|F 1(x)-F 2(x)|)
\]
where \(F 1(x)\) is the proportion of x 1 values less than or equal to \(x\) and \(F 1(x)\) is the proportion of x 2 values less than or equal to \(x\).
\(h=k s t e s t 2(x 1, x 2\), alpha) specifies the significance level alpha for the test. The default is 0.05 .
\(\mathrm{h}=\mathrm{kstest2}(\mathrm{x} 1, \mathrm{x} 2\), alpha, type) specifies the type of test using one of the following values for the string type:
- 'unequal ' - Tests the alternative hypothesis that the population cdfs are unequal. This is the default.
- 'larger' - Tests the alternative hypothesis that the first population cdf is larger than the second population cdf. The test statistic does not use the absolute value.
- 'smaller' - Tests the alternative hypothesis that the first population cdf is smaller than the second population cdf. The test statistic does not use the absolute value.
[h, p] = kstest2(...) also returns the asymptotic \(p\) value p . The asymptotic \(p\) value becomes very accurate for large sample sizes, and is believed to be reasonably accurate for sample sizes \(n 1\) and \(n 2\) such that \((\mathrm{n} 1 * \mathrm{n} 2) /(\mathrm{n} 1+\mathrm{n} 2)>=4\).
[h, \(\mathrm{p}, \mathrm{ks} 2 \mathrm{stat}]=\) kstest2(...) also returns the \(p\) value p and the test statistic ks2stat.

\section*{Examples}

The following test compares the distributions of a small evenly-spaced sample and a larger normal sample:
```

x = -1:1:5
y = randn(20,1);
[h,p,k] = kstest2(x,y)
h =
0
p =
0.0774
k =
0.5214

```

The following figure illustrates the test statistic:
```

F1 = cdfplot(x);
hold on
F2 = cdfplot(y)
set(F1,'LineWidth',2,'Color','r')
set(F2,'LineWidth',2)
legend([F1 F2],'F1(x)','F2(x)','Location','NW')

```


The test statistic k is the maximum difference between the curves.

\section*{References}
[1] Massey, F. J. "The Kolmogorov-Smirnov Test for Goodness of Fit." Journal of the American Statistical Association. Vol. 46, No. 253, 1951, pp. 68-78.
[2] Miller, L. H. "Table of Percentage Points of Kolmogorov Statistics." Journal of the American Statistical Association. Vol. 51, No. 273, 1956, pp. 111-121.
[3] Marsaglia, G., W. Tsang, and J. Wang. "Evaluating Kolmogorov's Distribution." Journal of Statistical Software. Vol. 8, Issue 18, 2003.
[4] Stephens, M. A. "Use of the Kolmogorov-Smirnov, Cramer-Von Mises and Related Statistics Without Extensive Tables." Journal of the Royal Statistical Society. Series B, Vol. 32, No. 1, 1970, pp. 115-122.

\section*{How To}
- kstest
- lillietest

Purpose Kurtosis
Syntax \(\quad k=\) kurtosis \((X)\)
k = kurtosis(X,flag)
k = kurtosis(X,flag,dim)

Description

Algorithms
\(k=\) kurtosis \((X)\) returns the sample kurtosis of \(X\). For vectors, kurtosis ( \(x\) ) is the kurtosis of the elements in the vector \(x\). For matrices kurtosis ( X ) returns the sample kurtosis for each column of \(X\). For N -dimensional arrays, kurtosis operates along the first nonsingleton dimension of \(X\).
\(\mathrm{k}=\) kurtosis \((\mathrm{X}, \mathrm{flag})\) specifies whether to correct for bias (flag is 0 ) or not ( flag is 1 , the default). When \(X\) represents a sample from a population, the kurtosis of \(X\) is biased, that is, it will tend to differ from the population kurtosis by a systematic amount that depends on the size of the sample. You can set flag to 0 to correct for this systematic bias.
\(\mathrm{k}=\) kurtosis( \(\mathrm{X}, \mathrm{flag}, \mathrm{dim})\) takes the kurtosis along dimension dim of \(X\).
kurtosis treats NaNs as missing values and removes them.
Kurtosis is a measure of how outlier-prone a distribution is. The kurtosis of the normal distribution is 3. Distributions that are more outlier-prone than the normal distribution have kurtosis greater than 3 ; distributions that are less outlier-prone have kurtosis less than 3.

The kurtosis of a distribution is defined as
\[
k=\frac{E(x-\mu)^{4}}{\sigma^{4}}
\]
where \(\mu\) is the mean of \(x, \sigma\) is the standard deviation of \(x\), and \(E(t)\) represents the expected value of the quantity \(t\). kurtosis computes a sample version of this population value.

Note Some definitions of kurtosis subtract 3 from the computed value, so that the normal distribution has kurtosis of 0 . The kurtosis function does not use this convention.

When you set flag to 1 , the following equation applies:
\[
k_{1}=\frac{\frac{1}{n} \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{4}}{\left(\frac{1}{n} \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2}\right)^{2}}
\]

When you set flag to 0 , the following equation applies:
\[
k_{0}=\frac{n-1}{(n-2)(n-3)}\left((n+1) k_{1}-3(n-1)\right)+3
\]

This bias-corrected formula requires that X contain at least four elements.

\section*{Examples}
```

$X=\operatorname{randn}\left(\left[\begin{array}{ll}5 & 4\end{array}\right]\right)$
X =
$\begin{array}{llll}1.1650 & 1.6961 & -1.4462 & -0.3600\end{array}$
$\begin{array}{llll}0.6268 & 0.0591 & -0.7012 & -0.1356\end{array}$
$\begin{array}{llll}0.0751 & 1.7971 & 1.2460 & -1.3493\end{array}$
$\begin{array}{llll}0.3516 & 0.2641 & -0.6390 & -1.2704\end{array}$
$\begin{array}{llll}-0.6965 & 0.8717 & 0.5774 & 0.9846\end{array}$
k = kurtosis(X)
$\mathrm{k}=$
$2.1658 \quad 1.2967 \quad 1.6378 \quad 1.9589$

```

\section*{See Also}
mean | moment | skewness | std | var

\section*{categorical.labels property}
Purpose Text labels for levels

Description Text labels for levels. Access labels with getlabels.
\begin{tabular}{|c|c|}
\hline Purpose & Regularized least-squares regression using lasso or elastic net algorithms \\
\hline Syntax & ```
B = lasso(X,Y)
[B,FitInfo] = lasso(X,Y)
[B,FitInfo] = lasso(X,Y,Name,Value)
``` \\
\hline Description & \begin{tabular}{l}
\(B=\) lasso( \(X, Y\) ) returns fitted least-squares regression coefficients for a set of regularization coefficients Lambda. \\
[B,FitInfo] = lasso \((X, Y)\) returns a structure containing information about the fits. \\
[B,FitInfo] = lasso(X,Y,Name, Value) fits regularized regressions with additional options specified by one or more Name, Value pair arguments.
\end{tabular} \\
\hline Input Arguments & Numeric matrix with \(n\) rows and \(p\) columns. Each row represents one observation, and each column represents one predictor (variable). \\
\hline
\end{tabular}
Y

Numeric vector of length \(n\), where \(n\) is the number of rows of \(X\). \(Y(i)\) is the response to row \(i\) of \(X\).

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

\section*{Alpha}

Scalar value from 0 to 1 (excluding 0 ) representing the weight of lasso \(\left(L^{1}\right)\) versus ridge \(\left(L^{2}\right)\) optimization. Alpha \(=1\) represents lasso regression, Alpha close to 0 approaches ridge regression, and
other values represent elastic net optimization. See "Definitions" on page 20-902.

Default: 1
CV
Method lasso uses to estimate mean squared error:
- K, a positive integer - lasso uses K-fold cross validation.
- cvp, a cvpartition object - lasso uses the cross-validation method expressed in cvp. You cannot use a 'leaveout' partition with lasso.
- 'resubstitution' - lasso uses \(X\) and \(Y\) to fit the model and to estimate the mean squared error, without cross validation.

Default: 'resubstitution'
DFmax
Maximum number of nonzero coefficients in the model. lasso returns results only for Lambda values that satisfy this criterion.

Default: Inf

\section*{Lambda}

Vector of nonnegative Lambda values. See "Definitions" on page 20-902.
- If you do not supply Lambda, lasso calculates the largest value of Lambda that gives a nonnull model. In this case, LambdaRatio gives the ratio of the smallest to the largest value of the sequence, and NumLambda gives the length of the vector.
- If you supply Lambda, lasso ignores LambdaRatio and NumLambda.

Default: Geometric sequence of NumLambda values, the largest just sufficient to produce \(B=0\)

\section*{LambdaRatio}

Positive scalar, the ratio of the smallest to the largest Lambda value when you do not set Lambda.

If you set LambdaRatio = 0, lasso generates a default sequence of Lambda values, and replaces the smallest one with 0 .

Default: 1e-4

\section*{MCReps}

Positive integer, the number of Monte Carlo repetitions for cross validation.
- If CV is 'resubstitution' or a cvpartition of type 'resubstitution', MCReps must be 1.
- If CV is a cvpartition of type 'holdout', MCReps must be greater than 1.

Default: 1

\section*{NumLambda}

Positive integer, the number of Lambda values lasso uses when you do not set Lambda.

Default: 100

\section*{Options}

Structure that specifies whether to cross validate in parallel, and specifies the random stream or streams. Create the Options structure with statset. Option fields:
- UseParallel - Set to 'always' to compute in parallel. Default is 'never'.
- UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'. Default is 'never'.
- Streams - A RandStream object or cell array consisting of one such object. If you do not specify Streams, lasso uses the default stream.

For details on using parallel computing, see Chapter 17, "Parallel Statistics".

\section*{PredictorNames}

Cell array of strings representing names of the predictor variables, in the order in which they appear in \(X\).

\section*{Default: \{\}}

RelTol
Convergence threshold for the coordinate descent algorithm (see Friedman, Tibshirani, and Hastie [3]). The algorithm terminates when successive estimates of the coefficient vector differ in the \(L^{2}\) norm by a relative amount less than RelTol.

Default: 1e-4

\section*{Standardize}

Boolean value specifying whether lasso scales \(X\) and \(Y\) before fitting the models. lasso always centers the data before fitting.

Default: true
```

Weights

```

Observation weights, a nonnegative vector of length \(n\), where n is the number of rows of \(X\). lasso scales Weights to sum to 1 .

Default: \(1 / \mathrm{n}\) * ones ( \(\mathrm{n}, 1\) )

\section*{Output \\ Arguments}

B
Fitted coefficients, a \(p\)-by-L matrix, where \(p\) is the number of predictors (columns) in \(X\), and \(L\) is the number of Lambda values.

\section*{FitInfo}

Structure containing information about the model fits.
\begin{tabular}{l|l}
\hline Field in Fitinfo & Description \\
\hline Intercept & \begin{tabular}{l} 
Intercept term \(\beta_{0}\) for each linear model, a \\
1-by-L vector
\end{tabular} \\
\hline Lambda & \begin{tabular}{l} 
Lambda parameters in ascending order, a \\
1-by-L vector
\end{tabular} \\
\hline Alpha & Value of Alpha parameter, a scalar \\
\hline DF & \begin{tabular}{l} 
Number of nonzero coefficients in B for each \\
value of Lambda, a 1-by-L vector
\end{tabular} \\
\hline MSE & Mean squared error (MSE), a 1-by-L vector \\
\hline
\end{tabular}

If you set the CV name-value pair to cross validate, the FitInfo structure contains additional fields.
\begin{tabular}{l|l}
\hline Field in Fitlnfo & Description \\
\hline SE & \begin{tabular}{l} 
The standard error of MSE for each Lambda, \\
as calculated during cross validation, a \\
1-by-L vector
\end{tabular} \\
\hline LambdaMinMSE & \begin{tabular}{l} 
The Lambda value with minimum MSE, a \\
scalar
\end{tabular} \\
\hline
\end{tabular}
\begin{tabular}{l|l}
\hline Field in FitInfo & Description \\
\hline IndexMinMSE & \begin{tabular}{l} 
The largest Lambda such that MSE is within \\
one standard error of the minimum, a scalar
\end{tabular} \\
\hline IndexMinMSE & \begin{tabular}{l} 
The index of Lambda with value \\
LambdaMinMSE, a scalar
\end{tabular} \\
\hline Index1SE & \begin{tabular}{l} 
The index of Lambda with value Lambda1SE, \\
a scalar
\end{tabular} \\
\hline
\end{tabular}

\section*{Definitions}

\section*{Lasso}

For a given value of \(\lambda\), a nonnegative parameter, lasso solves the problem
\[
\min _{\beta_{0}, \beta}\left(\frac{1}{2 N} \sum_{i=1}^{N}\left(y_{i}-\beta_{0}-x_{i}^{T} \beta\right)+\lambda \sum_{j=1}^{p}\left|\beta_{j}\right|\right),
\]
where
- \(N\) is the number of observations.
- \(y_{i}\) is the response at observation \(i\).
- \(x_{i}\) is data, a vector of \(p\) values at observation \(i\).
- \(\lambda\) is a nonnegative regularization parameter corresponding to one value of Lambda.
- The parameters \(\beta_{0}\) and \(\beta\) are scalar and \(p\)-vector respectively.

As \(\lambda\) increases, the number of nonzero components of \(\beta\) decreases.
The lasso problem involves the \(L^{1}\) norm of \(\beta\), as contrasted with the elastic net algorithm.

\section*{Elastic Net}

For an \(a\) strictly between 0 and 1, and a nonnegative \(\lambda\), elastic net solves the problem
\[
\min _{\beta_{0}, \beta}\left(\frac{1}{2 N} \sum_{i=1}^{N}\left(y_{i}-\beta_{0}-x_{i}^{T} \beta\right)+\lambda P_{\alpha}(\beta)\right),
\]
where
\[
P_{\alpha}(\beta)=\frac{(1-\alpha)}{2}\|\beta\|_{2}^{2}+\alpha\|\beta\|_{1}=\sum_{j=1}^{p}\left(\frac{(1-\alpha)}{2} \beta_{j}^{2}+\alpha\left|\beta_{j}\right|\right) .
\]

Elastic net is the same as lasso when \(a=1\). As \(a\) shrinks toward 0 , elastic net approaches ridge regression. For other values of \(a\), the penalty term \(P_{a}(\beta)\) interpolates between the \(L^{1}\) norm of \(\beta\) and the squared \(L^{2}\) norm of \(\beta\).

\section*{Examples}

Construct a data set with redundant predictors, and identify those predictors using cross-validated lasso:
```

X = randn(100,5);
r = [0;2;0;-3;0]; % only two nonzero coefficients
Y = X*r + randn(100,1)*.1; % small added noise
[b fitinfo] = lasso(X,Y,'CV',10);
lam = fitinfo.Index1SE; % find index of suggested lambda
b(:,lam)
ans =
0
1.9642
0
-2.9636
0

```
lasso identifies and removes the redundant predictors.

\section*{References}
[1] Tibshirani, R. Regression shrinkage and selection via the lasso. Journal of the Royal Statistical Society, Series B, Vol 58, No. 1, pp. 267-288, 1996.
[2] Zou, H. and T. Hastie. Regularization and variable selection via the elastic net. Journal of the Royal Statistical Society, Series B, Vol. 67, No. 2, pp. 301-320, 2005.
[3] Friedman, J., R. Tibshirani, and T. Hastie. Regularization paths for generalized linear models via coordinate descent. Journal of Statistical Software, Vol 33, No. 1, 2010. http://www.jstatsoft.org/v33/i01
[4] Hastie, T., R. Tibshirani, and J. Friedman. The Elements of Statistical Learning, 2nd edition. Springer, New York, 2008.

\section*{See Also}
lassoPlot | ridge
How To
- "Lasso and Elastic Net" on page 9-32

\section*{Purpose Trace plot of lasso fit}

Syntax \(\quad a x=\) lassoPlot \((B)\)
ax = lassoPlot(B,FitInfo)
ax = lassoPlot(B,FitInfo,Name,Value)
[ax,figh] = lassoPlot(B,...)

Description

\section*{Input Arguments}
ax = lassoPlot (B) creates a trace plot of the values in B against the \(L^{1}\) norm of \(B\). ax is a handle to the plot axis.
ax = lassoPlot (B,FitInfo) creates a plot with type depending on the data type of FitInfo and the value, if any, of the plotType name-value pair.
ax = lassoPlot(B,FitInfo, Name, Value) creates a plot with additional options specified by one or more Name, Value pair arguments.
[ax,figh] = lassoPlot ( \(B, \ldots\) ) returns a handle to the figure window.

B
Coefficients of a sequence of regression fits, as returned from the lasso function. \(B\) is a \(p\)-by-NLambda matrix, where \(p\) is the number of predictors, and each column of B is a set of coefficients lasso calculates using one Lambda penalty value.

FitInfo
Information controlling the plot:
- FitInfo is a structure, especially as returned from lasso lassoPlot creates a plot based on the PlotType name-value pair.
- FitInfo is a vector - lassoPlot forms the \(x\)-axis of the plot from the values in FitInfo. The length of FitInfo must equal the number of columns of \(B\).

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{Parent}

Axis in which to draw the plot.
Default: New plot
PlotType
Choose the plot type when you give a FitInfo vector or structure:
\begin{tabular}{l|l|l}
\hline Fitnfo Type & PlotType & Plot \\
\hline \begin{tabular}{l} 
Vector or \\
Structure
\end{tabular} & 'L1' & \begin{tabular}{l} 
plotLasso creates the \(x\)-axis from the \(L^{1}\) norm of the \\
coefficients in B. The \(x\)-axis at the top of the plot contains \\
the degrees of freedom (df), meaning the number of nonzero \\
coefficients of B.
\end{tabular} \\
\hline Structure & 'Lambda' & \begin{tabular}{l} 
plotLasso creates the x-axis from the Lambda field of \\
FitInfo. The x-axis at the top of the plot contains the \\
degrees of freedom (df), meaning the number of nonzero \\
coefficients of B.
\end{tabular} \\
\hline \begin{tabular}{l} 
Cross-Validated \\
Structure
\end{tabular} & 'CV' & \begin{tabular}{l} 
- For each Lambda, plots an estimate of the mean squared \\
prediction error on new data for the model fitted by lasso \\
with that value of Lambda.
\end{tabular} \\
- \begin{tabular}{l} 
Plots error bars for the estimates. \\
- \begin{tabular}{l} 
Plots the value of Lambda with minimum cross-validated \\
MSE. \\
- \begin{tabular}{l} 
Plots the greatest Lambda that is within one standard \\
error of minimum MSE (so makes the sparsest model \\
within that region).
\end{tabular} \\
\hline
\end{tabular} \\
\hline
\end{tabular} \\
\hline
\end{tabular}

Default: 'L1'

\section*{PredictorNames}

Cell array of strings to label each coefficient of \(B\). If the length of PredictorNames is less than the number of rows of \(B\), the remaining labels are padded with default values.
lassoPlot uses the predictor names in FitInfo only if:
- You created FitInfo with a call to lasso that included a PredictorNames name-value pair.
- You call lassoPlot without a PredictorNames name-value pair.
- You include FitInfo in your lassoPlot call.

Default: \{'B1','B2',...\}

\section*{XScale}
- 'linear' for linear x-axis
- ' \(\log\) ' for logarithmic scaled x -axis

Default: 'linear', except 'log' for the 'CV' plot type

\section*{Output Arguments}

Handle to the axis of the plot (see "Setting Axis Parameters").

\section*{figh}

Handle to the figure window (see "Graphics Windows - the Figure").

\section*{Examples}

Fit a regularized model of the acetylene data with lasso, and plot the fits with the default plot type:
load acetylene

\section*{lassoPlot}
```

X = [x1 x2 x3];
D = x2fx(X,'interaction');
D(:,1) = []; % No constant term
B = lasso(D,y);
lassoPlot(B);

```

Trace Plot of coefficients fit by Lasso.


Fit a regularized model of the acetylene data with lasso, and plot the fits with the Lambda plot type and logarithmic scaling:
```

load acetylene
X = [x1 x2 x3];
D = x2fx(X,'interaction');
D(:,1) = []; % No constant term

```
```

[B FitInfo] = lasso(D,y);
lassoPlot(B,FitInfo,'PlotType','Lambda','XScale','log');

```

Trace Plot of coefficients fit by Lasso


Fit a regularized model of the acetylene data with lasso and cross validation, and plot the cross-validated fits:
```

load acetylene
X = [x1 x2 x3];
D = x2fx(X,'interaction');
D(:,1) = []; % No constant term
[B FitInfo] = lasso(D,y,'CV',10);

```

\section*{lassoPlot}
lassoPlot(B,FitInfo,'PlotType','CV');

Cross-validated MSE of Lasso fit


\section*{See Also \\ lasso}
How To . "Lasso and Elastic Net" on page 9-32
\begin{tabular}{ll} 
Purpose & Less than or equal relation for handles \\
Syntax & h1 <= h2 \\
Description & \begin{tabular}{l} 
Handles are equal if they are handles for the same object. All \\
comparisons use a number associated with each handle object. Nothing \\
can be assumed about the result of a handle comparison except that the \\
repeated comparison of two handles in the same MATLAB session will \\
yield the same result. The order of handle values is purely arbitrary \\
and has no connection to the state of the handle objects being compared.
\end{tabular} \\
h1 <= h2 performs element-wise comparisons between handle arrays \\
h1 and h2. h1 and h2 must be of the same dimensions unless one is a \\
scalar. The result is a logical array of the same dimensions, where each \\
element is an element-wise <= result.
\end{tabular}

\section*{qrandset.Leap property}

\section*{Purpose Interval between points}

Description

Examples
Experiment with different leap values:
```

% No leaping produces the standard Halton sequence.
P = haltonset(5);
P(1:5,:)
% Set a leap of 1. The point set now includes every other
% point from the sequence.
P.Leap = 1;
P(1:5,:)

```

See Also net | qrandset | Skip | subsref | haltonset
Purpose Length of dataset array
Syntax \(\mathrm{n}=\) length \((\mathrm{A})\)
Description \(\mathrm{n}=\) length \((\mathrm{A})\) returns the number of observations in the dataset A . length is equivalent to size ( \(\mathrm{A}, 1\) ).
See Also ..... size

\section*{qrandset.length}
Purpose Length of point set
Syntax length(p)
Description length \((p)\) returns the number of points in the point set \(p\). It is equivalent to size ( \(p, 1\) ).
See Also qrandset | size

\section*{Purpose Length of categorical array}

\section*{Syntax \(\quad n=\) length \((A)\)}

Description \(\quad n=\) length \((A)\) returns the size of the longest dimension of the categorical array \(A\) when \(A\) is not empty. If \(A\) is a vector, this is the same as its length. length is equivalent to \(\max (\operatorname{size}(x))\) for nonempty arrays, and 0 for empty arrays.

See Also isempty | isscalar | size

Purpose Element counts by level
Syntax \(\quad \begin{aligned} C & =\text { levelcounts }(A) \\ C & =\text { levelcounts }(A, \operatorname{dim})\end{aligned}\)
Description
\(C=\) levelcounts(A) for a categorical vector \(A\) counts the number of elements in \(A\) equal to each of the possible levels in \(A\). The output is a vector \(C\) containing those counts, and has as many elements as \(A\) has levels. For matrix A, C is a matrix of column counts. For \(N\)-dimensional arrays, levelcounts operates along the first nonsingleton dimension. C = levelcounts(A,dim) operates along the dimension dim.

\section*{Examples}

Count the number of patients in each age group in the data in hospital.mat:
```

load hospital
edges = 0:10:100;
labels = strcat(num2str((0:10:90)','%d'),{'s'});
disp(labels')
'Os' '10s' '20s' '30s' '40s' '50s' '60s' '70s' '80s' '90s'
AgeGroup = ordinal(hospital.Age,labels,[],edges);
I = islevel(labels,AgeGroup);
disp(I')
0
c = levelcounts(AgeGroup);
disp(c')
0
AgeGroup = droplevels(AgeGroup);
I = islevel(labels,AgeGroup);
disp(I')
0
c = levelcounts(AgeGroup);
disp(c')
154142 2

```

\author{
See Also islevel | ismember | summary
}

\section*{leverage}

\section*{Purpose Leverage}
Syntax \(\quad\)\begin{tabular}{ll}
\(h\) & \(=\) leverage \((\) data \()\) \\
& \(h=\) leverage \((\) data, model \()\)
\end{tabular}

Description

Algorithms
\([Q, R]=\operatorname{qr}(x 2 f x(d a t a, ' m o d e l ')) ;\)
leverage = (sum(Q'.*Q'))'

\section*{Examples \\ One rule of thumb is to compare the leverage to \(2 p / n\) where \(n\) is the number of observations and \(p\) is the number of parameters in the model. For the Hald data set this value is 0.7692 . \\ ```
load hald \\ h = max(leverage(ingredients,'linear')) \\ h = \\ 0.7004
```}
\(\mathrm{h}=\) leverage (data) finds the leverage of each row (point) in the matrix data for a linear additive regression model.
\(\mathrm{h}=\) leverage(data, model) finds the leverage on a regression, using a specified model type, where model can be one of these strings:
- 'linear' - includes constant and linear terms
- 'interaction' - includes constant, linear, and cross product terms
- 'quadratic' - includes interactions and squared terms
- 'purequadratic ' - includes constant, linear, and squared terms

Leverage is a measure of the influence of a given observation on a regression due to its location in the space of the inputs.

Since \(0.7004<0.7692\), there are no high leverage points using this rule.

\author{
References [1] Goodall, C. R. "Computation Using the QR Decomposition." Handbook in Statistics. Vol. 9, Amsterdam: Elsevier/North-Holland, 1993.
}

How To • regstats

Purpose Latin hypercube sample
```

Syntax }\quadx=lhsdesign(n,p
X = lhsdesign(...,'smooth','off')
X = lhsdesign(...,'criterion',criterion)
X = lhsdesign(...,'iterations',k)

```

\section*{Description}

See Also haltonset | sobolset | lhsnorm | unifrnd
\(X=\operatorname{lnsdesign}(n, p)\) generates a latin hypercube sample \(X\) containing \(n\) values on each of \(p\) variables. For each column, the \(n\) values are randomly distributed with one from each interval \((0,1 / n),(1 / n, 2 / n)\), \(\ldots,(1-1 / n, 1)\), and they are randomly permuted.
X = lhsdesign(...,'smooth', 'off') produces points at the midpoints of the above intervals: \(0.5 / \mathrm{n}, 1.5 / \mathrm{n}, \ldots, 1-0.5 / \mathrm{n}\). The default is 'on'.
\(\mathrm{X}=\) lhsdesign(...,'criterion', criterion) iteratively generates latin hypercube samples to find the best one according to the criterion criterion, which can be one of the following strings.
\(X=\) lhsdesign(...,'iterations', \(k\) ) iterates up to \(k\) times in an attempt to improve the design according to the specified criterion. The default is \(\mathrm{k}=5\).
\begin{tabular}{l|l}
\hline Criterion & Description \\
\hline 'none' & No iteration \\
\hline 'maximin' & Maximize minimum distance between points \\
\hline 'correlation' & Reduce correlation \\
\hline
\end{tabular}

\section*{Purpose Latin hypercube sample from normal distribution}

Syntax
\(X=\) lhsnorm(mu,sigma, \(n\) )
X = lhsnorm(mu,sigma, \(n, f l a g)\)
[X,Z] = lhsnorm(...)

References

See Also
\(X=\operatorname{lnsnorm}(m u\), sigma, \(n)\) generates a latin hypercube sample \(X\) of size n from the multivariate normal distribution with mean vector mu and covariance matrix sigma. X is similar to a random sample from the multivariate normal distribution, but the marginal distribution of each column is adjusted so that its sample marginal distribution is close to its theoretical normal distribution.
\(X=\) lhsnorm(mu,sigma, \(n, f l a g)\) controls the amount of smoothing in the sample. If flag is 'off', each column has points equally spaced on the probability scale. In other words, each column is a permutation of the values \(G(0.5 / \mathrm{n}), G(1.5 / \mathrm{n}), \ldots, G(1-0.5 / \mathrm{n})\) where \(G\) is the inverse normal cumulative distribution for that column's marginal distribution. If flag is 'on' (the default), each column has points uniformly distributed on the probability scale. For example, in place of \(0.5 / \mathrm{n}\) you use a value having a uniform distribution on the interval ( \(0 / n, 1 / n\) ).
\([\mathrm{X}, \mathrm{Z}]=\operatorname{lnsnorm}(\ldots)\) also returns Z , the original multivariate normal sample before the marginals are adjusted to obtain \(X\).
[1] Stein, M. "Large sample properties of simulations using latin hypercube sampling." Technometrics. Vol. 29, No. 2, 1987, pp. 143-151. Correction, Vol. 32, p. 367.

\author{
lhsdesign | mvnrnd
}

\section*{lillietest}

\section*{Purpose Lilliefors test}

Syntax
```

h = lillietest(x)
h = lillietest(x,alpha)
h = lillietest(x,alpha,distr)
[h,p] = lillietest(...)
[h,p,kstat] = lillietest(...)
[h,p,kstat,critval] = lillietest(...)
[h,p,...] = lillietest(x,alpha,distr,mctol)

```

\section*{Description}
h = lillietest(x) performs a Lilliefors test of the default null hypothesis that the sample in vector \(x\) comes from a distribution in the normal family, against the alternative that it does not come from a normal distribution. The test returns the logical value \(h=1\) if it rejects the null hypothesis at the \(5 \%\) significance level, and \(h=0\) if it cannot. The test treats NaN values in x as missing values, and ignores them.
The Lilliefors test is a 2 -sided goodness-of-fit test suitable when a fully-specified null distribution is unknown and its parameters must be estimated. This is in contrast to the one-sample Kolmogorov-Smirnov test (see kstest), which requires that the null distribution be completely specified. The Lilliefors test statistic is the same as for the Kolmogorov-Smirnov test:
\[
K S=\max _{x}|S C D F(x)-C D F(x)|
\]
where \(S C D F\) is the empirical cdf estimated from the sample and \(C D F\) is the normal cdf with mean and standard deviation equal to the mean and standard deviation of the sample.
lillietest uses a table of critical values computed using Monte Carlo simulation for sample sizes less than 1000 and significance levels between 0.001 and 0.50 . The table is larger and more accurate than the table introduced by Lilliefors. Critical values for a test are computed by interpolating into the table, using an analytic approximation when extrapolating for larger sample sizes.
\(\mathrm{h}=\) lillietest(x, alpha) performs the test at significance level alpha. alpha is a scalar in the range [0.001, 0.50]. To perform the test at a significance level outside of this range, use the mctol input argument.
\(\mathrm{h}=\) lillietest(x, alpha, distr) performs the test of the null hypothesis that \(x\) came from the location-scale family of distributions specified by distr. Acceptable values for distr are 'norm' (normal, the default), 'exp' (exponential), and 'ev' (extreme value). The Lilliefors test can not be used when the null hypothesis is not a location-scale family of distributions.
[h, p] = lillietest(...) returns the \(p\) value p , computed using inverse interpolation into the table of critical values. Small values of p cast doubt on the validity of the null hypothesis. lillietest warns when \(p\) is not found within the tabulated range of [0.001, 0.50], and returns either the smallest or largest tabulated value. In this case, you can use the motol input argument to compute a more accurate \(p\) value.
[h,p,kstat] = lillietest(...) returns the test statistic kstat.
[h,p,kstat,critval] = lillietest(...) returns the critical value critval for the test. When kstat > critval, the null hypothesis is rejected at significance level alpha
[h,p,...] = lillietest(x, alpha, distr, mctol) computes a Monte Carlo approximation for \(p\) directly, rather than interpolating into the table of pre-computed values. This is useful when alpha or \(p\) lie outside the range of the table. lillietest chooses the number of Monte Carlo replications, mcreps, large enough to make the Monte Carlo standard error for \(p\), \(\operatorname{sqrt}\left(p^{*}(1-p) / m c r e p s\right)\), less than mctol.

\section*{Examples}

Use lillietest to determine if car mileage, in miles per gallon (MPG), follows a normal distribution across different makes of cars:
```

load carbig.mat
[h,p] = lillietest(MPG)
Warning: P is less than the smallest tabulated value, returning 0.001.
h =
1

```

\section*{lillietest}
```

p =
1.0000e-003

```

This is clear evidence for rejecting the null hypothesis of normality, but the \(p\) value returned is just the smallest value in the table of pre-computed values. To find a more accurate \(p\) value for the test, run a Monte Carlo approximation using the mctol input argument:
```

[h,p] = lillietest(MPG,0.05,'norm',1e-4)
h =
1
p =
8.3333e-006

```

\section*{References}

\section*{See Also}
jbtest | kstest | kstest2 | cdfplot


If H is omitted, c must have \(k\) elements and it specifies the null hypothesis values for the entire parameter vector.

Note The following functions return outputs suitable for use as the COVB input argument to linhyptest: nlinfit, coxphfit, glmfit, mnrfit, regstats, robustfit. nlinfit returns COVB directly; the other functions return COVB in stats.covb.
\([p, t, r]=\) linhyptest(...) also returns the test statistic \(t\) and the rank \(r\) of the hypothesis matrix \(H\). If dfe is Inf or is not given, \(t * r\) is a chi-square statistic with \(r\) degrees of freedom. If dfe is specified as a finite value, t is an \(F\) statistic with r and dfe degrees of freedom.

\section*{linhyptest}
linhyptest performs a test based on an asymptotic normal distribution for the parameter estimates. It can be used after any estimation procedure for which the parameter covariances are available, such as regstats or glmfit. For linear regression, the \(p\)-values are exact. For other procedures, the \(p\)-values are approximate, and may be less accurate than other procedures such as those based on a likelihood ratio.

\section*{Examples Fit a multiple linear model to the data in hald.mat:}
```

load hald
stats = regstats(heat,ingredients,'linear');
beta = stats.beta
beta =

```
    62.4054
        1.5511
        0.5102
        0.1019
        -0.1441

Perform an \(F\)-test that the last two coefficients are both 0 :
```

SIGMA = stats.covb;
dfe = stats.fstat.dfe;
H = [0 0 0 1 0;0 0 0 0 1];
c = [0;0];
[p,F] = linhyptest(beta,SIGMA,c,H,dfe)
p =
0.4668
F =
0.8391

```
See Also regstats | glmfit | robustfit | mnrfit | nlinfit | coxphfit

Purpose Agglomerative hierarchical cluster tree
Syntax \(\quad Z=\) linkage \((X)\)
Z = linkage(X, method)
Z = linkage( X, method,metric)
Z = linkage(X, method, pdist_inputs)
Z = linkage(X,method, metric,'savememory', value)
\(Z=\) linkage \((Y)\)
Z = linkage( Y , method)

\section*{Description}

Tips

Z = linkage (X) returns a matrix \(Z\) that encodes a tree of hierarchical clusters of the rows of the real matrix \(X\).

Z = linkage( X , method) creates the tree using the specified method, where method describes how to measure the distance between clusters.

Z = linkage(X,method,metric) performs clustering using the distance measure metric to compute distances between the rows of \(X\).

Z = linkage(X,method,pdist_inputs) passes parameters to the pdist function, which is the function that computes the distance between rows of \(X\).

Z = linkage(X,method,metric,'savememory', value) uses a memory-saving algorithm when value is 'true', and uses the standard algorithm when value is 'false'.
\(Z=\) linkage \((Y)\) uses a vector representation \(Y\) of a distance matrix. \(Y\) can be a distance matrix as computed by pdist, or a more general dissimilarity matrix conforming to the output format of pdist.

Z = linkage(Y,method) creates the tree using the specified method, where method describes how to measure the distance between clusters.
- Computing linkage \((\mathrm{Y})\) can be slow when Y is a vector representation of the distance matrix. For the 'centroid', 'median', and 'ward' methods, linkage checks whether Y is a Euclidean distance. Avoid this time-consuming check by passing in \(X\) instead of \(Y\).

\section*{linkage}
- The centroid and median methods can produce a cluster tree that is not monotonic. This occurs when the distance from the union of two clusters, \(r\) and \(s\), to a third cluster is less than the distance between \(r\) and \(s\). In this case, in a dendrogram drawn with the default orientation, the path from a leaf to the root node takes some downward steps. To avoid this, use another method. The following image shows a nonmonotonic cluster tree.


In this case, cluster 1 and cluster 3 are joined into a new cluster, while the distance between this new cluster and cluster 2 is less than the distance between cluster 1 and cluster 3. This leads to a nonmonotonic tree.
- You can provide the output \(Z\) to other functions including dendrogram to display the tree, cluster to assign points to clusters, inconsistent to compute inconsistent measures, and cophenet to compute the cophenetic correlation coefficient.

\section*{Input Arguments \\ X}

Matrix with two or more rows. The rows represent observations, the columns represent categories or dimensions.

\footnotetext{
method
}

Algorithm for computing distance between clusters.
\begin{tabular}{l|l}
\hline Method & Description \\
\hline 'average' & Unweighted average distance (UPGMA) \\
\hline 'centroid' & \begin{tabular}{l} 
Centroid distance (UPGMC), appropriate for \\
Euclidean distances only
\end{tabular} \\
\hline 'complete' & Furthest distance \\
\hline 'median' & \begin{tabular}{l} 
Weighted center of mass distance (WPGMC), \\
appropriate for Euclidean distances only
\end{tabular} \\
\hline 'single' & Shortest distance \\
\hline 'ward' & \begin{tabular}{l} 
Inner squared distance (minimum variance \\
algorithm), appropriate for Euclidean distances \\
only
\end{tabular} \\
\hline 'weighted ' & Weighted average distance (WPGMA) \\
\hline
\end{tabular}

Default: 'single'

\section*{metric}

Any distance metric that the pdist function accepts.
\begin{tabular}{l|l}
\hline Metric & Description \\
\hline 'euclidean' & Euclidean distance (default). \\
\hline 'seuclidean' & \begin{tabular}{l} 
Standardized Euclidean distance. Each \\
coordinate difference between rows in X \\
is scaled by dividing by the corresponding \\
element of the standard deviation \\
S=nanstd (X). To specify another value for S, \\
use D=pdist (X, 'seuclidean' S\().\)
\end{tabular} \\
\hline 'cityblock' & City block metric. \\
\hline
\end{tabular}

\section*{linkage}
\begin{tabular}{l|l}
\hline Metric & Description \\
\hline 'minkowski' & \begin{tabular}{l} 
Minkowski distance. The default exponent \\
is 2. To specify a different exponent, use D \\
= pdist (X, 'minkowski' , P), where P is a
\end{tabular} \\
\hline 'chebychev' & \begin{tabular}{l} 
Chebychev distance (maximum coordinate \\
difference).
\end{tabular} \\
\hline 'mahalanobis' & \begin{tabular}{l} 
Mahalanobis distance, using the \\
sample covariance of X as computed \\
by nancov. To compute the distance \\
with a different covariance, use D \\
pdist (X, 'manalanobis ' , C), where the \\
matrix C is symmetric and positive definite.
\end{tabular} \\
\hline 'cosine' & \begin{tabular}{l} 
One minus the cosine of the included angle \\
between points (treated as vectors).
\end{tabular} \\
\hline 'correlation' & \begin{tabular}{l} 
One minus the sample correlation between \\
points (treated as sequences of values).
\end{tabular} \\
\hline 'spearman' & \begin{tabular}{l} 
One minus the sample Spearman's rank \\
correlation between observations (treated as \\
sequences of values).
\end{tabular} \\
\hline 'hamming' & \begin{tabular}{l} 
Hamming distance, which is the percentage \\
of coordinates that differ.
\end{tabular} \\
\hline 'jaccard' & \begin{tabular}{l} 
One minus the Jaccard coefficient, which is \\
the percentage of nonzero coordinates that \\
differ.
\end{tabular} \\
\hline custom distance & \begin{tabular}{l} 
A distance function specified using @: \\
D = pdist (X, @distfun) \\
function
\end{tabular} \\
A distance function must be of form \\
d2 = distfun (XI , XJ)
\end{tabular}

\section*{Metric \(\quad\) Description}
\(\mid\) an \(m 2\)-by- \(n\) matrix XJ , corresponding to multiple rows of \(X\). distfun must accept a matrix \(X J\) with an arbitrary number of rows. distfun must return an \(m 2\)-by- 1 vector of distances d2, whose \(k\) th element is the distance between XI and \(X J(k,:)\).

Default: 'euclidean'
pdist_inputs
A cell array of parameters accepted by the pdist function. For example, to set the metric to minkowski and use an exponent of 5, set pdist_inputs to \{'minkowski', 5\}.

\section*{savememory}

A string, either 'on' or 'off'. When applicable, the 'on ' setting causes linkage to construct clusters without computing the distance matrix. savememory is applicable when:
- linkage is 'centroid', 'median', or 'ward'
- distance is 'euclidean' (default)

When savememory is 'on', linkage run time is proportional to the number of dimensions (number of columns of \(X\) ). When savememory is 'off', linkage memory requirement is proportional to \(\mathrm{N}^{2}\), where N is the number of observations. So choosing the best (least-time) setting for savememory depends on the problem dimensions, number of observations, and available memory. The default savememory setting is a rough approximation of an optimal setting.

Default: ' on ' when X has 20 columns or fewer, or the computer does not have enough memory to store the distance matrix; otherwise 'off'

\section*{linkage}

A vector of distances with the same format as the output of the pdist function:
- A row vector of length \(m(m-1) / 2\), corresponding to pairs of observations in a matrix \(X\) with \(m\) rows
- Distances arranged in the order \((2,1),(3,1), \ldots,(m, 1),(3,2), \ldots\), ( \(m, 2\) ), ..., ( \(m, m-1\) ))

Y can be a more general dissimilarity matrix conforming to the output format of pdist.

\section*{Output z \\ Arguments}

\section*{Definitions Linkages}

The following notation is used to describe the linkages used by the various methods:
- Cluster r is formed from clusters p and q .
- \(\mathrm{n}_{r}\) is the number of objects in cluster \(r\).
- \(\mathrm{x}_{r i}\) is the ith object in cluster r .
- Single linkage, also called nearest neighbor, uses the smallest distance between objects in the two clusters:
\[
d(r, s)=\min \left(\operatorname{dist}\left(x_{r i}, x_{s j}\right)\right), i \in\left(i, \ldots, n_{r}\right), j \in\left(1, \ldots, n_{s}\right)
\]
- Complete linkage, also called furthest neighbor, uses the largest distance between objects in the two clusters:
\[
d(r, s)=\max \left(\operatorname{dist}\left(x_{r i}, x_{s j}\right)\right), i \in\left(1, \ldots, n_{r}\right), j \in\left(1, \ldots, n_{s}\right)
\]
- Average linkage uses the average distance between all pairs of objects in any two clusters:
\[
d(r, s)=\frac{1}{n_{r} n_{s}} \sum_{i=1}^{n_{r}} \sum_{j=1}^{n_{s}} \operatorname{dist}\left(x_{r i}, x_{s j}\right)
\]
- Centroid linkage uses the Euclidean distance between the centroids of the two clusters:
\[
d(r, s)=\left\|\bar{x}_{r}-\bar{x}_{s}\right\|_{2}
\]
where
\[
\bar{x}_{r}=\frac{1}{n_{r}} \sum_{i=1}^{n_{r}} x_{r i}
\]
- Median linkage uses the Euclidean distance between weighted centroids of the two clusters,
\[
d(r, s)=\left\|\tilde{x}_{r}-\tilde{x}_{s}\right\|_{2}
\]

\section*{linkage}
where \(\tilde{x}_{r}\) and \(\tilde{x}_{s}\) are weighted centroids for the clusters \(r\) and \(s\). If cluster \(r\) was created by combining clusters \(p\) and \(q, \tilde{x}_{r}\) is defined recursively as
\[
\tilde{x}_{r}=\frac{1}{2}\left(\tilde{x}_{p}+\tilde{x}_{q}\right)
\]
- Ward's linkage uses the incremental sum of squares; that is, the increase in the total within-cluster sum of squares as a result of joining two clusters. The within-cluster sum of squares is defined as the sum of the squares of the distances between all objects in the cluster and the centroid of the cluster. The sum of squares measure is equivalent to the following distance measure \(d(r, s)\), which is the formula linkage uses:
\[
d(r, s)=\sqrt{\frac{2 n_{r} n_{s}}{\left(n_{r}+n_{s}\right)}}\left\|\bar{x}_{r}-\bar{x}_{s}\right\|_{2},
\]
where
- \(\left\|\|_{2}\right.\) is Euclidean distance
- \(\bar{x}_{r}\) and \(\bar{x}_{s}\) are the centroids of clusters \(r\) and \(s\)
- \(n_{r}\) and \(n_{s}\) are the number of elements in clusters \(r\) and \(s\)

In some references the Ward linkage does not use the factor of 2 multiplying \(n_{r} n_{s}\). The linkage function uses this factor so the distance between two singleton clusters is the same as the Euclidean distance.
- Weighted average linkage uses a recursive definition for the distance between two clusters. If cluster \(r\) was created by combining clusters \(p\) and \(q\), the distance between \(r\) and another cluster \(s\) is defined as the average of the distance between \(p\) and \(s\) and the distance between \(q\) and \(s\) :
\[
d(r, s)=\frac{(d(p, s)+d(q, s))}{2}
\]

\section*{Examples}

Compute four clusters of the Fisher iris data using Ward linkage and ignoring species information, and see how the cluster assignments correspond to the three species.
```

load fisheriris
Z = linkage(meas,'ward','euclidean');
c = cluster(Z,'maxclust',4);
crosstab(c,species)
firstfive = Z(1:5,:) % first 5 rows of Z
dendrogram(Z)
ans =
0 25 1
0 24 14
0}103
50 0 0
firstfive =

| 102.0000 | 143.0000 | 0 |
| ---: | ---: | ---: |
| 8.0000 | 40.0000 | 0.1000 |
| 1.0000 | 18.0000 | 0.1000 |
| 10.0000 | 35.0000 | 0.1000 |
| 129.0000 | 133.0000 | 0.1000 |

```

\section*{linkage}


Create a hierarchical cluster tree for a data with 20000 observations using Ward's linkage. If you set savememory to 'off', you can get an out-of-memory error if your machine doesn't have enough memory to hold the distance matrix. Cluster the data into four groups and plot the result.
```

X = rand(20000,3);
Z = linkage(X,'ward','euclidean','savememory','on');
c = cluster(Z,'maxclust',4);
scatter3(X(:,1),X(:,2),X(:,3),10,c)

```


\footnotetext{
See Also
cluster | clusterdata | cophenet | dendrogram | inconsistent | kmeans | pdist | silhouette | squareform

\author{
How To \\ - Chapter 11, "Cluster Analysis"
}
}

\section*{Purpose Lognormal cumulative distribution function}
```

Syntax $\quad P=\operatorname{logncdf}(X$, mu, sigma $)$
[P,PLO,PUP] = logncdf(X,mu,sigma,pcov,alpha)

```

\section*{Description}
\(P=\operatorname{logncdf}(X, m u\), sigma \()\) returns values at \(X\) of the lognormal cdf with distribution parameters mu and sigma. mu and sigma are the mean and standard deviation, respectively, of the associated normal distribution. X, mu, and sigma can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input for X , mu, or sigma is expanded to a constant array with the same dimensions as the other inputs.
[P, PLO, PUP] = logncdf(X,mu,sigma, pcov,alpha) returns confidence bounds for \(P\) when the input parameters mu and sigma are estimates. pcov is the covariance matrix of the estimated parameters. alpha specifies \(100(1-\) alpha \() \%\) confidence bounds. The default value of alpha is 0.05 . PLO and PUP are arrays of the same size as P containing the lower and upper confidence bounds.
logncdf computes confidence bounds for P using a normal approximation to the distribution of the estimate
\[
\frac{X-\hat{\mu}}{\hat{\sigma}}
\]
and then transforming those bounds to the scale of the output P. The computed bounds give approximately the desired confidence level when you estimate mu, sigma, and pcov from large samples, but in smaller samples other methods of computing the confidence bounds might be more accurate.

The lognormal cdf is
\[
p=F(x \mid \mu, \sigma)=\frac{1}{\sigma \sqrt{2 \pi}} \int_{0}^{x} \frac{e^{\frac{-(\ln (t)-\mu)^{2}}{2 \sigma^{2}}}}{t} d t
\]

Examples

\section*{References}

See Also
How To
```

    x = (0:0.2:10);
    y = logncdf(x,0,1);
    plot(x,y); grid;
    xlabel('x'); ylabel('p');
    ```

[1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions. 2nd ed., Hoboken, NJ: John Wiley \& Sons, Inc., 1993, pp. 102-105.
cdf | lognpdf | logninv | lognstat | lognfit | lognlike | lognrnd - "Lognormal Distribution" on page B-51

Purpose Lognormal parameter estimates
Syntax parmhat = lognfit(data)
[parmhat, parmci] = lognfit(data)
[parmhat, parmci] = lognfit(data,alpha)
[...] = lognfit(data, alpha, censoring)
[...] = lognfit(data, alpha, censoring,freq)
[...] = lognfit(data,alpha, censoring,freq,options)

\section*{Description}
parmhat = lognfit(data) returns a vector of maximum likelihood estimates parmhat (1) = mu and parmhat (2) = sigma of parameters for a lognormal distribution fitting data. mu and sigma are the mean and standard deviation, respectively, of the associated normal distribution.
[parmhat, parmci] = lognfit(data) returns \(95 \%\) confidence intervals for the parameter estimates mu and sigma in the 2 -by- 2 matrix parmci. The first column of the matrix contains the lower and upper confidence bounds for parameter mu, and the second column contains the confidence bounds for parameter sigma.
[parmhat, parmci] = lognfit(data,alpha) returns 100(1-alpha) \% confidence intervals for the parameter estimates, where alpha is a value in the range ( \(\left.\begin{array}{ll}0 & 1\end{array}\right)\) specifying the width of the confidence intervals. By default, alpha is 0.05 , which corresponds to \(95 \%\) confidence intervals.
[...] = lognfit(data, alpha, censoring) accepts a Boolean vector censoring, of the same size as data, which is 1 for observations that are right-censored and 0 for observations that are observed exactly.
[...] = lognfit(data, alpha, censoring,freq) accepts a frequency vector, freq, of the same size as data. Typically, freq contains integer frequencies for the corresponding elements in data, but can contain any nonnegative values. Pass in [] for alpha, censoring, or freq to use their default values.
[...] = lognfit(data, alpha, censoring,freq,options) accepts a structure, options, that specifies control parameters for the iterative algorithm the function uses to compute maximum likelihood estimates when there is censoring. The lognormal fit function accepts an options
structure which can be created using the function statset. Enter statset('lognfit') to see the names and default values of the parameters that lognfit accepts in the options structure. See the reference page for statset for more information about these options.
```

Examples This example generates }100\mathrm{ independent samples of lognormally distributed data with $\mu=0$ and $\sigma=3$. parmhat estimates $\mu$ and $\sigma$ and parmci gives $99 \%$ confidence intervals around parmhat. Notice that parmci contains the true values of $\mu$ and $\sigma$.
data = lognrnd(0,3,100,1);
[parmhat,parmci] = lognfit(data,0.01)
parmhat =
-0.2480 2.8902
parmci =
-1.0071 2.4393
0.5111 3.5262

```

See Also
How To
mle | lognlike | lognpdf | logncdf | logninv | lognstat | lognrnd
- "Lognormal Distribution" on page B-51

\section*{Purpose Lognormal inverse cumulative distribution function}
```

Syntax $\quad X=\operatorname{logninv}(P$, mu, sigma $)$
[X,XLO,XUP] = logninv(P,mu,sigma,pcov,alpha)

```
\(X=\) logninv( \(P\), mu, sigma) returns values at \(P\) of the inverse lognormal cdf with distribution parameters mu and sigma. mu and sigma are the mean and standard deviation, respectively, of the associated normal distribution. mu and sigma can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of X . A scalar input for P , mu, or sigma is expanded to a constant array with the same dimensions as the other inputs.
[X,XLO, XUP] = logninv(P,mu,sigma, pcov,alpha) returns confidence bounds for \(X\) when the input parameters mu and sigma are estimates. pcov is the covariance matrix of the estimated parameters. alpha specifies \(100(1-a l p h a) \%\) confidence bounds. The default value of alpha is 0.05 . XLO and XUP are arrays of the same size as \(X\) containing the lower and upper confidence bounds.
logninv computes confidence bounds for P using a normal approximation to the distribution of the estimate
\[
\hat{\mu}+\hat{\sigma} q
\]
where \(q\) is the Pth quantile from a normal distribution with mean 0 and standard deviation 1 . The computed bounds give approximately the desired confidence level when you estimate mu, sigma, and pcov from large samples, but in smaller samples other methods of computing the confidence bounds might be more accurate.

The lognormal inverse function is defined in terms of the lognormal cdf as
\[
x=F^{-1}(p \mid \mu, \sigma)=\{x: F(x \mid \mu, \sigma)=p\}
\]
where
\[
p=F(x \mid \mu, \sigma)=\frac{1}{\sigma \sqrt{2 \pi}} \int_{0}^{x} \frac{e^{\frac{-(\ln (t)-\mu)^{2}}{2 \sigma^{2}}}}{t} d t
\]

\section*{Examples}
```

p = (0.005:0.01:0.995);
crit = logninv(p,1,0.5);
plot(p,crit)
xlabel('Probability'); ylabel('Critical Value'); grid

```


References

How To

\author{
See Also
}
[1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions. Hoboken, NJ: Wiley-Interscience, 2000. pp. 102-105.
icdf | logncdf | lognpdf | lognstat | lognfit | lognlike | lognrnd
- "Lognormal Distribution" on page B-51
\begin{tabular}{ll} 
Purpose & Lognormal negative log-likelihood \\
Syntax & \(n \log L=\) lognlike (params, data) \\
& {\([n \log L\), avar \(]=\) lognlike (params, data) } \\
& {\([\ldots]=\) lognlike (params, data, censoring) } \\
& {\([\ldots]=\) lognlike (params, data, censoring, freq) }
\end{tabular}

\section*{Description}

See Also
How To
nlogL = lognlike(params, data) returns the negative log-likelihood of data for the lognormal distribution with parameters params. params (1) is the mean of the associated normal distribution, mu, and params (2) is the standard deviation of the associated normal distribution, sigma. The values of mu and sigma are scalars, and the output nlogL is a scalar.
[nlogL, avar] = lognlike(params, data) returns the inverse of Fisher's information matrix. If the input parameter value in params is the maximum likelihood estimate, avar is its asymptotic variance. avar is based on the observed Fisher's information, not the expected information.
[...] = lognlike(params, data, censoring) accepts a Boolean vector, censoring, of the same size as data, which is 1 for observations that are right-censored and 0 for observations that are observed exactly.
[...] = lognlike(params, data, censoring,freq) accepts a frequency vector, freq, of the same size as data. The vector freq typically contains integer frequencies for the corresponding elements in data, but can contain any nonnegative values. Pass in [] for censoring to use its default value.
lognfit | lognpdf | logncdf | logninv | lognstat | lognrnd
- "Lognormal Distribution" on page B-51

\section*{Purpose}

Lognormal probability density function

\section*{Syntax}

Description
\(Y=\) lognpdf( \(X, m u\), sigma \()\)
\(Y=\) lognpdf( \(X, \mathrm{mu}\), sigma \()\) returns values at \(X\) of the lognormal pdf with distribution parameters mu and sigma. mu and sigma are the mean and standard deviation, respectively, of the associated normal distribution. X , mu, and sigma can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of Y . A scalar input for X , mu, or sigma is expanded to a constant array with the same dimensions as the other inputs.

The lognormal pdf is
\[
y=f(x \mid \mu, \sigma)=\frac{1}{x \sigma \sqrt{2 \pi}} e^{\frac{-(\ln x-\mu)^{2}}{2 \sigma^{2}}}
\]

The normal and lognormal distributions are closely related. If \(X\) is distributed lognormally with parameters \(\mu\) and \(\sigma\), then \(\log (X)\) is distributed normally with mean \(\mu\) and standard deviation \(\sigma\).

The mean \(m\) and variance \(v\) of a lognormal random variable are functions of \(\mu\) and \(\sigma\) that can be calculated with the lognstat function. They are:
\[
\begin{aligned}
& m=\exp \left(\mu+\sigma^{2} / 2\right) \\
& v=\exp \left(2 \mu+\sigma^{2}\right)\left(\exp \left(\sigma^{2}\right)-1\right)
\end{aligned}
\]

A lognormal distribution with mean \(m\) and variance \(v\) has parameters
\[
\begin{aligned}
& \mu=\log \left(m^{2} / \sqrt{v+m^{2}}\right) \\
& \sigma=\sqrt{\log \left(v / m^{2}+1\right)}
\end{aligned}
\]

\section*{lognpdf}

The lognormal distribution is applicable when the quantity of interest must be positive, since \(\log (X)\) exists only when \(X\) is positive.

\section*{Examples}
\[
\begin{aligned}
& x=(0: 0.02: 10) ; \\
& y=\operatorname{lognpdf}(x, 0,1) ; \\
& \operatorname{plot}(x, y) ; \text { grid; } \\
& \text { xlabel('x'); ylabel('p') }
\end{aligned}
\]


\section*{References \\ [1] Mood, A. M., F. A. Graybill, and D. C. Boes. Introduction to the Theory of Statistics. 3rd ed., New York: McGraw-Hill, 1974. pp. 540-541.}

\footnotetext{
See Also
How To
pdf | logncdf | logninv | lognstat | lognfit | lognlike | lognrnd
- "Lognormal Distribution" on page B-51
}

\section*{Purpose}

Lognormal random numbers
Syntax
R = lognrnd(mu,sigma)
R = lognrnd(mu,sigma,m,n,...)
\(R\) = lognrnd(mu,sigma, \([m, n, \ldots])\)

\section*{Description}
\(R=\) lognrnd(mu, sigma) returns an array of random numbers generated from the lognormal distribution with parameters mu and sigma. mu and sigma are the mean and standard deviation, respectively, of the associated normal distribution. mu and sigma can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of R. A scalar input for mu or sigma is expanded to a constant array with the same dimensions as the other input.
\(R=\operatorname{lognrnd}(m u\), sigma, \(m, n, \ldots) \quad\) or \(R=\)
lognrnd(mu, sigma, \([m, n, \ldots]\) ) generates an \(m\)-by-n-by-... array. The mu, sigma parameters can each be scalars or arrays of the same size as R.

The normal and lognormal distributions are closely related. If \(X\) is distributed lognormally with parameters \(\mu\) and \(\sigma\), then \(\log (X)\) is distributed normally with mean \(\mu\) and standard deviation \(\sigma\).

The mean \(m\) and variance \(v\) of a lognormal random variable are functions of \(\mu\) and \(\sigma\) that can be calculated with the lognstat function. They are:
\[
\begin{aligned}
& m=\exp \left(\mu+\sigma^{2} / 2\right) \\
& v=\exp \left(2 \mu+\sigma^{2}\right)\left(\exp \left(\sigma^{2}\right)-1\right)
\end{aligned}
\]

A lognormal distribution with mean \(m\) and variance \(v\) has parameters
\[
\begin{aligned}
& \mu=\log \left(m^{2} / \sqrt{v+m^{2}}\right) \\
& \sigma=\sqrt{\log \left(v / m^{2}+1\right)}
\end{aligned}
\]

Examples Generate one million lognormally distributed random numbers with mean 1 and variance 2 :
```

$\mathrm{m}=1$;
$\mathrm{v}=2$;
$m u=\log \left(\left(m^{\wedge} 2\right) / s q r t\left(v+m^{\wedge} 2\right)\right) ;$
sigma $=\operatorname{sqrt}\left(\log \left(v /\left(m^{\wedge} 2\right)+1\right)\right)$;
[M, V]= lognstat(mu, sigma)
M =
1
$\mathrm{V}=$
2.0000
$X=\operatorname{lognrnd}(m u$, sigma, 1,1e6);
$M X=\operatorname{mean}(X)$
MX =
0.9974
$\mathrm{VX}=\operatorname{var}(\mathrm{X})$
$\mathrm{VX}=$
1.9776

```
References

[1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions.
 Hoboken, NJ: Wiley-Interscience, 2000. pp. 102-105.
See Also random | lognpdf | logncdf | logninv | lognstat | lognfit | lognlike | normrnd
How To - "Lognormal Distribution" on page B-51

\section*{Purpose}

Lognormal mean and variance

\section*{Syntax}
[M,V] = lognstat(mu,sigma)
\([M, V]=\) lognstat(mu, sigma) returns the mean of and variance of the lognormal distribution with parameters mu and sigma. mu and sigma are the mean and standard deviation, respectively, of the associated normal distribution. mu and sigma can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of \(M\) and \(V\). A scalar input for mu or sigma is expanded to a constant array with the same dimensions as the other input.

The normal and lognormal distributions are closely related. If \(X\) is distributed \(\operatorname{lognormally}\) with parameters \(\mu\) and \(\sigma\), then \(\log (X)\) is distributed normally with mean \(\mu\) and standard deviation \(\sigma\).

The mean \(m\) and variance \(v\) of a lognormal random variable are functions of \(\mu\) and \(\sigma\) that can be calculated with the lognstat function. They are:
\[
\begin{aligned}
& m=\exp \left(\mu+\sigma^{2} / 2\right) \\
& v=\exp \left(2 \mu+\sigma^{2}\right)\left(\exp \left(\sigma^{2}\right)-1\right)
\end{aligned}
\]

A lognormal distribution with mean \(m\) and variance \(v\) has parameters
\[
\begin{aligned}
& \mu=\log \left(m^{2} / \sqrt{v+m^{2}}\right) \\
& \sigma=\sqrt{\log \left(v / m^{2}+1\right)}
\end{aligned}
\]

\section*{Examples}

Generate one million lognormally distributed random numbers with mean 1 and variance 2:
\[
\begin{aligned}
& \mathrm{m}=1 ; \\
& \mathrm{v}=2 ;
\end{aligned}
\]

\section*{lognstat}
```

mu = log((m^2)/sqrt(v+m^2));
sigma = sqrt(log(v/(m^2)+1));
[M,V]= lognstat(mu,sigma)
M =
1
V =
2.0000
X = lognrnd(mu,sigma,1,1e6);
MX = mean(X)
MX =
0.9974
VX = var(X)
VX =
1.9776

```

References [1] Mood, A. M., F. A. Graybill, and D. C. Boes. Introduction to the Theory of Statistics. 3rd ed., New York: McGraw-Hill, 1974. pp. 540-541.

\section*{See Also lognpdf | logncdf | logninv | lognfit | lognlike | lognrnd} How To . "Lognormal Distribution" on page B-51

\section*{CompactClassificationDiscriminant.loss}
\begin{tabular}{|c|c|}
\hline Purpose & Classification error \\
\hline Syntax & \[
\begin{aligned}
L & =\operatorname{loss}(o b j, X, Y) \\
L & =\operatorname{loss}(o b j, X, Y, \text { Name, Value })
\end{aligned}
\] \\
\hline Description & \begin{tabular}{l}
L = loss(obj,X,Y) returns a scalar representing how well obj classifies the data in \(X\), when \(Y\) contains the true classifications. \\
\(\mathrm{L}=\) loss(obj, \(\mathrm{X}, \mathrm{Y}\), Name, Value) returns the loss with additional options specified by one or more Name, Value pair arguments.
\end{tabular} \\
\hline \multirow[t]{9}{*}{\begin{tabular}{l}
Input \\
Arguments
\end{tabular}} & \begin{tabular}{l}
obj \\
Discriminant analysis classifier of class ClassificationDiscriminant or CompactClassificationDiscriminant, typically constructed with ClassificationDiscriminant.fit.
\end{tabular} \\
\hline & X \\
\hline & Matrix where each row represents an observation, and each column represents a predictor. The number of columns in X must equal the number of predictors in obj. \\
\hline & Y \\
\hline & Class labels, with the same data type as exists in obj. The number of elements of \(Y\) must equal the number of rows of \(X\). \\
\hline & Name-Value Pair Arguments \\
\hline & Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN. \\
\hline & lossfun \\
\hline & Function handle or string representing a loss function. Built-in loss functions: \\
\hline
\end{tabular}

\section*{CompactClassificationDiscriminant.loss}
- 'binodeviance' - See "Loss Functions" on page 20-952.
- 'exponential ' - See "Loss Functions" on page 20-952.
- 'mincost' - Smallest misclassification cost as given by the obj. Cost matrix.

You can write your own loss function using the syntax described in "Loss Functions" on page 20-952.

Default: 'mincost'
weights
Numeric vector of length \(N\), where \(N\) is the number of rows of X . weights are nonnegative. When you supply weights, loss computes weighted classification error.

Default: ones ( \(\mathrm{N}, 1\) )

\section*{Output} Arguments

\section*{Definitions}

\section*{Classification Error}

The default classification error is the fraction of data \(X\) that obj misclassifies, where \(Y\) represents the true classifications.

Weighted classification error is the sum of weight \(i\) times the Boolean value that is 1 when obj misclassifies the \(i\) th row of \(X\), divided by the sum of the weights.

\section*{Loss Functions}

The built-in loss functions are:

\section*{CompactClassificationDiscriminant.loss}
- 'binodeviance' - For binary classification, assume the classes \(y_{n}\) are -1 and 1 . With weight vector \(w\) normalized to have sum 1 , and predictions of row \(n\) of data \(X\) as \(f\left(X_{n}\right)\), the binomial deviance is
\[
\sum w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(X_{n}\right)\right)\right) .
\]
- 'exponential' - With the same definitions as for 'binodeviance', the exponential loss is
\[
\sum w_{n} \exp \left(-y_{n} f\left(X_{n}\right)\right)
\]
- 'mincost' - Smallest expected misclassification cost, with expectation taken over the posterior probability, and cost as given by the obj. Cost matrix

To write your own loss function, create a function file of the form
```

function loss = lossfun(C,S,W,COST)

```
- \(N\) is the number of rows of \(X\).
- K is the number of classes in obj, represented in obj.ClassNames.
- C is an N-by-K logical matrix, with one true per row for the true class. The index for each class is its position in obj.ClassNames.
- \(S\) is an \(N\)-by-K numeric matrix. \(S\) is a matrix of posterior probabilities for classes with one row per observation, similar to the posterior output from predict.
- \(W\) is a numeric vector with \(N\) elements, the observation weights. If you pass W , the elements are normalized to sum to the prior probabilities in the respective classes.
- COST is a K-by-K numeric matrix of misclassification costs. For example, you can use COST=ones (K) - eye (K), which means a cost of 0 for correct classification, and 1 for misclassification.
- The output loss should be a scalar.

\section*{CompactClassificationDiscriminant.loss}

Pass the function handle @lossfun as the value of the lossfun name-value pair.

\section*{Examples Compute the resubstituted classification error for the Fisher iris data: \\ load fisheriris \\ obj = ClassificationDiscriminant.fit(meas,species); L = loss(obj,meas,species) \\ L = \\ 0.0200}

See Also ClassificationDiscriminant | edge | margin | predict
How To . "Discriminant Analysis" on page 12-3

\section*{CompactClassificationEnsemble.loss}

\section*{Purpose \\ Classification error}

Syntax
\(\mathrm{L}=\) loss(ens, \(\mathrm{X}, \mathrm{Y})\)
L = loss(ens, \(\mathrm{X}, \mathrm{Y}\), Name, Value)

Description

\section*{Input Arguments}
\(L=\) loss(ens, \(X, Y\) ) returns the classification error for ensemble ens computed using matrix of predictors X and true class labels Y .
\(\mathrm{L}=\) loss(ens, \(\mathrm{X}, \mathrm{Y}\), Name, Value) computes classification error with additional options specified by one or more Name, Value pair arguments.
ens
Classification ensemble created with fitensemble, or a compact classification ensemble created with compact.

X
Matrix of data to classify. Each row of X represents one observation, and each column represents one predictor. \(X\) must have the same number of columns as the data used to train ens. \(X\) should have the same number of rows as the number of elements in \(Y\).

\section*{Y}

Classification of \(X\). \(Y\) should be of the same type as the classification used to train ens, and its number of elements should equal the number of rows of \(X\).

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{learners}

Indices of weak learners in the ensemble ranging from 1 to ens. NTrained. loss uses only these learners for calculating loss.

\section*{CompactClassificationEnsemble.loss}

Default: 1:NTrained

\section*{lossfun}

Function handle or string representing a loss function. Built-in loss functions:
- 'binodeviance ' - See "Loss Functions" on page 20-957
- 'classiferror' - Fraction of misclassified data
- 'exponential' - See "Loss Functions" on page 20-957

You can write your own loss function in the syntax described in "Loss Functions" on page 20-957.

Default: 'classiferror'
mode
String representing the meaning of the output L:
- 'ensemble' - L is a scalar value, the loss for the entire ensemble.
- 'individual' - L is a vector with one element per trained learner.
- 'cumulative' - \(L\) is a vector in which element \(J\) is obtained by using learners \(1: J\) from the input list of learners.

Default: 'ensemble'
UseObsForLearner
A logical matrix of size N -by-T, where:
- \(N\) is the number of rows of \(X\).
- \(T\) is the number of weak learners in ens.

\section*{CompactClassificationEnsemble.loss}
hen UseObsForLearner( \(i, j\) ) is true, learner \(j\) is used in predicting the class of row i of \(X\).

Default: \(\operatorname{true}(\mathrm{N}, \mathrm{T})\)

\section*{weights}

Vector of observation weights, with nonnegative entries. The length of weights must equal the number of rows in X .

Default: ones(size (X, 1), 1)

Output
Arguments

\section*{Definitions}

L
Loss, by default the fraction of misclassified data. L can be a vector, and can mean different things, depending on the name-value pair settings.

\section*{Classification Error}

The default classification error is the fraction of the data \(X\) that ens misclassifies, where \(Y\) are the true classifications.

Weighted classification error is the sum of weight \(i\) times the Boolean value that is 1 when tree misclassifies the \(i\) th row of x , divided by the sum of the weights.

\section*{Loss Functions}

The built-in loss functions are:
- 'binodeviance' - For binary classification, assume the classes \(y_{n}\) are - 1 and 1. With weight vector \(w\) normalized to have sum 1, and predictions of row \(n\) of data \(X\) as \(f\left(X_{n}\right)\), the binomial deviance is
\[
\sum w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(X_{n}\right)\right)\right)
\]
- 'classiferror' - Fraction of misclassified data, weighted by \(w\).

\section*{CompactClassificationEnsemble.loss}
- 'exponential' - With the same definitions as for 'binodeviance', the exponential loss is
\[
\sum w_{n} \exp \left(-y_{n} f\left(X_{n}\right)\right)
\]

To write your own loss function, create a function file of the form
```

function loss = lossfun(C,S,W,COST)

```
- \(N\) is the number of rows of ens. \(X\).
- K is the number of classes in ens, represented in ens.ClassNames.
- C is an N-by-K logical matrix, with one true per row for the true class. The index for each class is its position in tree.ClassNames.
- \(S\) is an \(N\)-by-K numeric matrix. \(S\) is a matrix of posterior probabilities for classes with one row per observation, similar to the score output from predict.
- W is a numeric vector with N elements, the observation weights.
- COST is a K-by-K numeric matrix of misclassification costs. The default 'classiferror' gives a cost of 0 for correct classification, and 1 for misclassification.
- The output loss should be a scalar.

Pass the function handle @lossfun as the value of the lossfun name-value pair.

\section*{Examples Create a compact classification ensemble for the ionosphere data, and} find the fraction of training data that the ensemble misclassifies:
```

load ionosphere
ada = fitensemble(X,Y,'AdaBoostM1',100,'tree');
adb = compact(ada);
L = loss(adb,X,Y)
L =

```

\title{
CompactClassificationEnsemble.loss
}
0.0085
\begin{tabular}{ll} 
See Also loss | edge | margin | predict \\
How To & - Chapter 13, "Nonparametric Supervised Learning"
\end{tabular}

\section*{CompactClassificationTree.loss}

\section*{Purpose Classification error}

Syntax \(\quad L=\) loss (tree \(, X, Y\) )
L = loss(tree, \(X, Y\), Name, Value)
L = loss(tree, X, Y,'subtrees',subtreevector)
[L,se] = loss(tree, \(X, Y\), 'subtrees',subtreevector)
[L,se,NLeaf] = loss(tree, \(\mathrm{X}, \mathrm{Y}\), 'subtrees', subtreevector)
[L,se,NLeaf,bestlevel] = loss(tree, X, Y,'subtrees',
subtreevector)
[L,...] =
loss(tree, \(\mathrm{X}, \mathrm{Y}\), 'subtrees', subtreevector, Name, Value)
Description \(L=\) loss(tree \(, X, Y\) ) returns a scalar representing how well tree classifies the data in \(X\), when \(Y\) contains the true classifications.
\(\mathrm{L}=\) loss(tree, \(\mathrm{X}, \mathrm{Y}\), Name, Value) returns the loss with additional options specified by one or more Name, Value pair arguments.

L = loss(tree, X, Y, 'subtrees', subtreevector) returns a vector of classification errors for the trees in the pruning sequence subtreevector.
[L,se] = loss(tree, \(\mathrm{X}, \mathrm{Y}\), 'subtrees',subtreevector) returns the vector of standard errors of the classification errors.

Note loss returns se and further outputs only when the lossfun name-value pair is the default 'classiferror'.
[L,se,NLeaf] = loss(tree, X, Y,'subtrees', subtreevector) returns the vector of numbers of leaf nodes in the trees of the pruning sequence.
[L,se,NLeaf,bestlevel] = loss(tree, \(\mathrm{X}, \mathrm{Y}\), 'subtrees', subtreevector) returns the best pruning level as defined in the treesize name-value pair. By default, bestlevel is the pruning level that gives loss within one standard deviation of minimal loss.

\section*{CompactClassificationTree.loss}

\section*{[L,...] = \\ loss(tree, X, Y, 'subtrees', subtreevector, Name, Value) returns loss statistics with additional options specified by one or more Name, Value pair arguments. \\ Input Arguments \\ tree \\ A classification tree or compact classification tree constructed by ClassificationTree.fit or compact. \\ X \\ Matrix of data to classify. Each row of \(X\) represents one observation, and each column represents one predictor. X must have the same number of columns as the data used to train tree. \(X\) should have the same number of rows as the number of elements in \(Y\). \\ Y \\ Classification of \(X\). \(Y\) should be of the same type as the classification used to train tree, and its number of elements should equal the number of rows of \(X\). \\ Name-Value Pair Arguments \\ Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN. \\ lossfun \\ Function handle or string representing a loss function. Built-in loss functions: \\ - 'binodeviance ' - See "Loss Functions" on page 20-963 \\ - 'classiferror' - Fraction of misclassified data \\ - 'exponential' - See "Loss Functions" on page 20-963}

\section*{CompactClassificationTree.loss}
"Loss Functions" on page 20-963.
Default: 'classiferror'
weights
A numeric vector of length \(N\), where \(N\) is the number of rows of X. weights are nonnegative. When you supply weights, loss computes weighted classification error.

Default: ones (N, 1)
Name, Value arguments associated with pruning subtrees:

\section*{subtrees}

A vector with integer values from 0 (full unpruned tree) to the maximal pruning level max(tree. PruneList). You can set subtrees to 'all', meaning the entire pruning sequence.

Default: 0

\section*{treesize}

One of the following strings:
- 'se' - loss returns the highest pruning level with loss within one standard deviation of the minimum ( \(L+s e\), where \(L\) and se relate to the smallest value in subtrees).
- 'min' - loss returns the element of subtrees with smallest loss, usually the smallest element of subtrees.

\section*{Output Arguments}

\section*{L}

Classification error, a vector the length of subtrees. The meaning of the error depends on the values in weights and lossfun; see "Classification Error" on page 20-963.

\section*{CompactClassificationTree.loss}

\section*{se}

Standard error of loss, a vector the length of subtrees.

\section*{NLeaf}

Number of leaves (terminal nodes) in the pruned subtrees, a vector the length of subtrees.
```

bestlevel

```

A scalar whose value depends on treesize:
- treesize \(=\) 'se' - loss returns the highest pruning level with loss within one standard deviation of the minimum ( \(\mathrm{L}+\mathrm{se}\), where \(L\) and se relate to the smallest value in subtrees).
- treesize = 'min' - loss returns the element of subtrees with smallest loss, usually the smallest element of subtrees.

\section*{Definitions Classification Error}

The default classification error is the fraction of data \(X\) that tree misclassifies, where \(Y\) represents the true classifications.

Weighted classification error is the sum of weight \(i\) times the Boolean value that is 1 when tree misclassifies the \(i\) th row of x , divided by the sum of the weights.

\section*{Loss Functions}

The built-in loss functions are:
- 'binodeviance' - For binary classification, assume the classes \(y_{n}\) are - 1 and 1. With weight vector \(w\) normalized to have sum 1, and predictions of row \(n\) of data \(X\) as \(f\left(X_{n}\right)\), the binomial deviance is
\[
\sum w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(X_{n}\right)\right)\right) .
\]
- 'classiferror' - Fraction of misclassified data, weighted by \(w\).

\section*{CompactClassificationTree.loss}
- 'exponential' - With the same definitions as for 'binodeviance', the exponential loss is
\[
\sum w_{n} \exp \left(-y_{n} f\left(X_{n}\right)\right)
\]

To write your own loss function, create a function file of the form
```

function loss = lossfun(C,S,W,COST)

```
- \(N\) is the number of rows of \(X\).
- \(K\) is the number of classes in tree, represented in tree.ClassNames.
- C is an N-by-K logical matrix, with one true per row for the true class. The index for each class is its position in tree.ClassNames.
- \(S\) is an \(N\)-by-K numeric matrix. \(S\) is a matrix of posterior probabilities for classes with one row per observation, similar to the posterior output from predict.
- \(W\) is a numeric vector with \(N\) elements, the observation weights. If you pass \(W\), the elements are normalized to sum to the prior probabilities in the respective classes.
- COST is a K-by-K numeric matrix of misclassification costs. For example, you can use COST=ones (K) - eye (K), which means a cost of 0 for correct classification, and 1 for misclassification.
- The output loss should be a scalar.

Pass the function handle @lossfun as the value of the lossfun name-value pair.

\section*{Examples}

Compute the resubstituted classification error for the ionosphere data:
```

load ionosphere
tree = ClassificationTree.fit(X,Y);
L = loss(tree,X,Y)
L =

```

\section*{CompactClassificationTree.Ioss}
0.0114

\section*{See Also margin | edge | predict}

How To • Chapter 13, "Nonparametric Supervised Learning"

\section*{CompactRegressionEnsemble.loss}
\[
\begin{array}{ll}
\text { Purpose } & \text { Regression error } \\
\text { Syntax } & L=\operatorname{loss}(\mathrm{ens}, X, Y) \\
& L=\operatorname{loss}(\mathrm{ens}, X, Y, \text { Name, Value })
\end{array}
\]

Description

Input
Arguments
\(\mathrm{L}=\) loss (ens, \(\mathrm{X}, \mathrm{Y}\) ) returns the mean squared error between the predictions of ens to the data in X , compared to the true responses Y .
\(\mathrm{L}=\) loss(ens, \(\mathrm{X}, \mathrm{Y}\), Name, Value) computes the error in prediction with additional options specified by one or more Name, Value pair arguments. ens

A regression ensemble created with fitensemble, or the compact method.

X
A matrix of predictor values. Each column of \(X\) represents one variable, and each row represents one observation.
NaN values in X are taken to be missing values. Observations with all missing values for X are not used in the calculation of loss.

\section*{Y}

A numeric column vector with the same number of rows as \(X\). Each entry in \(Y\) is the response to the data in the corresponding row of \(X\).

NaN values in \(Y\) are taken to be missing values. Observations with missing values for \(Y\) are not used in the calculation of loss.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
learners

\section*{CompactRegressionEnsemble.loss}

Indices of weak learners in the ensemble ranging from 1 to ens.NTrained. oobEdge uses only these learners for calculating loss.

Default: 1:NTrained

\section*{lossfun}

Function handle for loss function, or the string 'mse', meaning mean squared error. If you pass a function handle fun, loss calls it as
fun(Y,Yfit,W)
where Y, Yfit, and W are numeric vectors of the same length.
- Y is the observed response.
- Yfit is the predicted response.
- \(W\) is the observation weights.

The returned value fun(Y,Yfit,W) should be a scalar.
Default: 'mse'

\section*{mode}

String representing the meaning of the output L:
- 'ensemble' - L is a scalar value, the loss for the entire ensemble.
- 'individual' - L is a vector with one element per trained learner.
- 'cumulative' - L is a vector in which element J is obtained by using learners \(1: J\) from the input list of learners.

Default: 'ensemble'

\section*{CompactRegressionEnsemble.loss}

\section*{UseObsForLearner}

A logical matrix of size N -by-NTrained, where N is the number of observations in ens. \(X\), and NTrained is the number of weak learners. When UseObsForLearner ( \(I, J\) ) is true, predict uses learner \(J\) in predicting observation I.

Default: true(N,NTrained)

\section*{weights}

Numeric vector of observation weights with the same number of elements as Y. The formula for loss with weights is in "Weighted Mean Squared Error" on page 20-968.

Default: ones(size(Y))

\section*{Output Arguments}

\section*{Definitions}

\section*{Examples}

Weighted mean squared error of predictions. The formula for loss is in "Weighted Mean Squared Error" on page 20-968.

\section*{Weighted Mean Squared Error}

Let \(n\) be the number of rows of data, \(x_{j}\) be the \(j\) th row of data, \(y_{j}\) be the true response to \(x_{j}\), and let \(f\left(x_{j}\right)\) be the response prediction of ens to \(x_{j}\). Let \(w\) be the vector of weights (all one by default).
First the weights are divided by their sum so they add to one: \(w \rightarrow w / \Sigma w\). The mean squared error \(L\) is
\[
L=\sum_{j=1}^{n} w_{j}\left(f\left(x_{j}\right)-y_{j}\right)^{2}
\]

Find the loss of an ensemble predictor of the carsmall data to find MPG as a function of engine displacement, horsepower, and vehicle weight:
```

load carsmall

```

\section*{CompactRegressionEnsemble.loss}
```

X = [Displacement Horsepower Weight];
ens = fitensemble(X,MPG,'LSBoost',100,'Tree');
L = loss(ens,X,MPG)
L =
4.3904

```

\section*{See Also \\ predict | fitensemble}

How To • Chapter 13, "Nonparametric Supervised Learning"

\section*{CompactRegressionTree.loss}

Purpose Regression error
```

Syntax $\quad L=$ loss (tree $, X, Y$ )
[L,se] = loss(tree, X,Y)
[L,se,NLeaf] = loss(tree, $\mathrm{X}, \mathrm{Y}$ )
[L,se,NLeaf,bestlevel] = loss(tree, $\mathrm{X}, \mathrm{Y}$ )
L = loss(tree, X, Y, Name, Value)

```

\section*{Description}

\section*{Input Arguments}
\(\mathrm{L}=\) loss (tree, \(\mathrm{X}, \mathrm{Y})\) returns the mean squared error between the predictions of tree to the data in X , compared to the true responses Y .
[L,se] = loss(tree, \(\mathrm{X}, \mathrm{Y}\) ) returns the standard error of the loss.
[L,se,NLeaf] = loss(tree, \(X, Y\) ) returns the number of leaves (terminal nodes) in the tree.
[L,se,NLeaf, bestlevel] = loss(tree, X, Y) returns the optimal pruning level for tree.
\(\mathrm{L}=\) loss(tree, \(\mathrm{X}, \mathrm{Y}\), Name, Value) computes the error in prediction with additional options specified by one or more Name, Value pair arguments.
tree
Regression tree created with RegressionTree.fit, or the compact method.

X
A matrix of predictor values. Each column of \(X\) represents one variable, and each row represents one observation.

Y
A numeric column vector with the same number of rows as \(X\). Each entry in \(Y\) is the response to the data in the corresponding row of \(X\).

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must

\section*{CompactRegressionTree.loss}
appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{lossfun}

Function handle for loss, or the string 'mse' representing mean-squared error. If you pass a function handle fun, loss calls fun as:
fun(Y,Yfit,W)
- \(Y\) is the vector of true responses.
- Yfit is the vector of predicted responses.
- W is the observation weights. If you pass W, the elements are normalized to sum to 1 .

All the vectors have the same number of rows as Y .
Default: 'mse'

\section*{subtrees}

A vector with integer values from 0 (full unpruned tree) to the maximal pruning level max(tree. PruneList). You can set subtrees to 'all', meaning the entire pruning sequence.

Default: 0

\section*{treesize}

A string, either:
- 'se' - loss returns bestlevel that corresponds to the smallest tree whose mean squared error (MSE) is within one standard error of the minimum MSE.
- 'min' - loss returns bestlevel that corresponds to the minimal MSE tree.

\section*{CompactRegressionTree.loss}
weights
Numeric vector of observation weights with the same number of elements as Y .

Default: ones(size(Y))

\section*{Output Arguments}

\section*{Definitions}

\section*{Mean Squared Error}

The mean squared error \(m\) of the predictions \(f\left(X_{n}\right)\) with weight vector \(w\) is
\[
m=\frac{\sum w_{n}\left(f\left(X_{n}\right)-Y_{n}\right)^{2}}{\sum w_{n}}
\]

\section*{CompactRegressionTree.loss}
```

Examples Find the loss of a regression tree predictor of the carsmall data to find
MPG as a function of engine displacement, horsepower, and vehicle
weight:
load carsmall
X = [Displacement Horsepower Weight];
tree = RegressionTree.fit(X,MPG);
L = loss(tree,X,MPG)
L =
4.8952
Find the pruning level that gives the optimal level of loss for the carsmall data:

```
```

load carsmall

```
load carsmall
X = [Displacement Horsepower Weight];
X = [Displacement Horsepower Weight];
tree = RegressionTree.fit(X,MPG);
tree = RegressionTree.fit(X,MPG);
[L,se,NLeaf,bestlevel] = loss(tree,X,MPG,'Subtrees','all');
[L,se,NLeaf,bestlevel] = loss(tree,X,MPG,'Subtrees','all');
bestlevel
bestlevel
bestlevel =
bestlevel =
4
4
See Also predict
How To - Chapter 13, "Nonparametric Supervised Learning"
```

Purpose Lower Pareto tails parameters
Syntax params = lowerparams(obj)
Description params = lowerparams(obj) returns the 2-element vector params of shape and scale parameters, respectively, of the lower tail of the Pareto tails object obj. lowerparams does not return a location parameter.

Examples $\quad$ Fit Pareto tails to a $t$ distribution at cumulative probabilities 0.1 and 0.9:
t $=$ trnd ( $3,100,1$ );
obj = paretotails(t, 0.1,0.9);
lowerparams(obj)
ans =
-0.1901 1.1898
upperparams(obj)
ans =
$0.3646 \quad 0.5103$

## See Also

paretotails | upperparams

## Purpose Less than relation for handles

## Syntax <br> h1 < h2

Description $\quad \mathrm{h} 1<\mathrm{h} 2$ performs element-wise comparisons between handle arrays h 1 and h2. h 1 and h2 must be of the same dimensions unless one is a scalar. The result is a logical array of the same dimensions, where each element is an element-wise < result.

If one of h1 or h2 is scalar, scalar expansion is performed and the result will match the dimensions of the array that is not scalar.
$\mathrm{tf}=\mathrm{lt}(\mathrm{h} 1, \mathrm{~h} 2)$ stores the result in a logical array of the same dimensions.

## See Also

qrandstream | eq | ge | gt | le | ne

## Isline

Purpose Add least-squares line to scatter plot

## Syntax <br> lsline <br> h = lsline

Description
lsline superimposes a least-squares line on each scatter plot in the current axes. Scatter plots are produced by the MATLAB scatter and plot functions. Data points connected with solid, dashed, or dash-dot lines (LineStyle '-', '--', or '.-') are not considered to be scatter plots by lsline, and are ignored.
$\mathrm{h}=$ lsline returns a column vector of handles h to the least-squares lines.

## Examples

Use lsline together with scatter plots produced by scatter and various line styles of plot:

```
x = 1:10;
y1 = x + randn(1,10);
scatter(x,y1,25,'b','*')
hold on
y2 = 2*x + randn(1,10);
plot(x,y2,'mo')
y3 = 3*x + randn(1,10);
plot(x,y3,'rx:')
y4 = 4*x + randn(1,10);
plot(x,y4,'g+--')
lsline
```



See Also
scatter | plot | refline | refcurve | gline

## Purpose Mean or median absolute deviation

## Syntax <br> $y=\operatorname{mad}(X)$ <br> $Y=\operatorname{mad}(X, 1)$ <br> $Y=\operatorname{mad}(X, 0)$

Description

Examples
$y=\operatorname{mad}(X)$ returns the mean absolute deviation of the values in $X$. For vector input, $y$ is mean (abs $(X-m e a n(X)))$. For a matrix input, $y$ is a row vector containing the mean absolute deviation of each column of X. For $N$-dimensional arrays, mad operates along the first nonsingleton dimension of $X$.
$Y=\operatorname{mad}(X, 1)$ returns the median absolute deviation of the values in $X$. For vector input, y is median(abs(X-median(X))). For a matrix input, $y$ is a row vector containing the median absolute deviation of each column of X. For $N$-dimensional arrays, mad operates along the first nonsingleton dimension of $X$.
$Y=\operatorname{mad}(X, 0)$ is the same as $\operatorname{mad}(X)$, and returns the mean absolute deviation of the values in $X$.
$\operatorname{mad}(\mathrm{X}, \mathrm{flag}, \mathrm{dim})$ computes absolute deviations along the dimension dim of X . flag is 0 or 1 to indicate mean or median absolute deviation, respectively.
mad treats NaNs as missing values and removes them.
For normally distributed data, multiply mad by one of the following factors to obtain an estimate of the normal scale parameter $\sigma$ :

- sigma $=1.253^{*} \operatorname{mad}(X, 0)-$ For mean absolute deviation
- sigma $=1.4826 * \operatorname{mad}(X, 1)$ - For median absolute deviation

The following compares the robustness of different scale estimates for normally distributed data in the presence of outliers:

```
x = normrnd(0,1,1,50);
xo = [x 10]; % Add outlier
```

```
r1 = std(xo)/std(x)
r1 =
        1.7385
r2 = mad(xo,0)/mad(x,0)
r2 =
    1.2306
r3 = mad(xo,1)/mad(x,1)
r3 =
        1.0602
```


## References <br> [1] Mosteller, F., and J. Tukey. Data Analysis and Regression. Upper Saddle River, NJ: Addison-Wesley, 1977. <br> [2] Sachs, L. Applied Statistics: A Handbook of Techniques. New York: Springer-Verlag, 1984, p. 253.

See Also
std | range | iqr

Purpose Mahalanobis distance

## Syntax $\quad d=\operatorname{mahal}(Y, X)$

Description
$d=\operatorname{mahal}(\mathrm{Y}, \mathrm{X})$ computes the Mahalanobis distance (in squared units) of each observation in $Y$ from the reference sample in matrix $X$. If $Y$ is $n$-by- $m$, where $n$ is the number of observations and $m$ is the dimension of the data, d is $n$-by- $1 . \mathrm{X}$ and Y must have the same number of columns, but can have different numbers of rows. $X$ must have more rows than columns.

For observation I, the Mahalanobis distance is defined by $d(I)=$ $(Y(I,:)-m u) * i n v(S I G M A) *(Y(I,:)-m u))^{\prime}$, where mu and SIGMA are the sample mean and covariance of the data in $X$. mahal performs an equivalent, but more efficient, computation.

## Examples

Generate some correlated bivariate data in $X$ and compare the Mahalanobis and squared Euclidean distances of observations in Y:

```
X = mvnrnd([0;0],[1 .9;.9 1],100);
Y = [1 1;1 -1;-1 1;-1 -1];
d1 = mahal(Y,X) % Mahalanobis
d1 =
    1.3592
    21.1013
    23.8086
        1.4727
d2 = sum((Y-repmat(mean(X),4,1)).^2, 2) % Squared Euclidean
d2 =
    1.9310
    1.8821
    2.1228
    2.0739
scatter(X(:,1),X(:,2))
```



The observations in Y with equal coordinate values are much closer to $X$ in Mahalanobis distance than observations with opposite coordinate values, even though all observations are approximately equidistant from the mean of X in Euclidean distance. The Mahalanobis distance, by considering the covariance of the data and the scales of the different variables, is useful for detecting outliers in such cases.

See Also pdist | mahal

## CompactClassificationDiscriminant.mahal

Purpose Mahalanobis distance to class means

M = mahal (obj, X)
M = mahal(obj,X,Name,Value)

## Syntax <br> Description

Input
Arguments
$M=\operatorname{mahal}(\mathrm{obj}, \mathrm{X})$ returns the squared Mahalanobis distances from observations in $X$ to the class means in obj.
$M=$ mahal(obj, X, Name, Value) computes the squared Mahalanobis distance with additional options specified by one or more Name, Value pair arguments.
obj
Discriminant analysis classifier of class ClassificationDiscriminant or CompactClassificationDiscriminant, typically constructed with ClassificationDiscriminant.fit.

X
Numeric matrix of size $n$-by- $p$, where $p$ is the number of predictors in obj, and $n$ is any positive integer. mahal computes the Mahalanobis distances from the rows of $X$ to each of the $K$ means of the classes in obj.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## ClassLabels

Class labels consisting of $n$ elements of obj. Y , where n is the number of rows of $X$.

## CompactClassificationDiscriminant.mahal

## Output M

## Definitions

## Examples

## Mahalanobis Distance

The Mahalanobis distance $d(x, y)$ between $n$-dimensional points $x$ and $y$, with respect to a given $n$-by- $n$ covariance matrix $S$, is

$$
d(x, y)=\sqrt{(x-y)^{T} S^{-1}(x-y)} .
$$

Find the Mahalanobis distances from the mean of the Fisher iris data to the class means, using distinct covariance matrices for each class:

```
load fisheriris
obj = ClassificationDiscriminant.fit(meas,species,...
    'DiscrimType','quadratic');
mahadist = mahal(obj,mean(meas))
mahadist =
    220.0667 5.0254 30.5804
```


## See Also

CompactClassificationDiscriminant | mahal | gmdistribution
How To . "Discriminant Analysis" on page 12-3

## gmdistribution.mahal

Purpose Mahalanobis distance to component means

## Syntax $\quad D=\operatorname{mahal}(o b j, x)$

Description $\quad D=$ mahal $(o b j, X)$ computes the Mahalanobis distance (in squared units) of each observation in $X$ to the mean of each of the $k$ components of the Gaussian mixture distribution defined by obj. obj is an object created by gmdistribution or fit. X is an $n$-by- $d$ matrix, where $n$ is the number of observations and $d$ is the dimension of the data. $D$ is $n$-by- $k$, with $\mathrm{D}(\mathrm{I}, \mathrm{J})$ the distance of observation I from the mean of component J.

## Examples

Generate data from a mixture of two bivariate Gaussian distributions using the mvnrnd function:

```
MU1 = [1 2];
SIGMA1 = [2 0; 0 .5];
MU2 = [-3 -5];
SIGMA2 = [1 0; 0 1];
X = [mvnrnd(MU1,SIGMA1,1000);mvnrnd(MU2,SIGMA2,1000)];
scatter(X(:,1),X(:,2),10,'.')
hold on
```



Fit a two-component Gaussian mixture model:

```
obj = gmdistribution.fit(X,2);
h = ezcontour(@(x,y)pdf(obj,[x y]),[-8 6],[-8 6]);
```


## gmdistribution.mahal



Compute the Mahalanobis distance of each point in $X$ to the mean of each component of obj:

```
D = mahal(obj,X);
delete(h)
scatter(X(:,1),X(:,2),10,D(:,1),'.')
hb = colorbar;
ylabel(hb,'Mahalanobis Distance to Component 1')
```



See Also gmdistribution | cluster | posterior | mahal

## maineffectsplot

Purpose Main effects plot for grouped data

```
Syntax maineffectsplot(Y,GROUP)
maineffectsplot(Y,GROUP,param1,val1,param2,val2,...)
[figh,AXESH] = maineffectsplot(...)
```


## Description

maineffectsplot (Y,GROUP) displays main effects plots for the group means of matrix $Y$ with groups defined by entries in the cell array GROUP. $Y$ is a numeric matrix or vector. If $Y$ is a matrix, the rows represent different observations and the columns represent replications of each observation. Each cell of GROUP must contain a grouping variable that can be a categorical variable, numeric vector, character matrix, or single-column cell array of strings. (See "Grouped Data" on page 2-34.) GROUP can also be a matrix whose columns represent different grouping variables. Each grouping variable must have the same number of rows as Y . The number of grouping variables must be greater than 1 .

The display has one subplot per grouping variable, with each subplot showing the group means of $Y$ as a function of one grouping variable.
maineffectsplot(Y,GROUP,param1,val1,param2,val2,...) specifies one or more of the following name/value pairs:

- 'varnames ' - Grouping variable names in a character matrix or a cell array of strings, one per grouping variable. Default names are 'X1', 'X2', ... .
- 'statistic' - String values that indicate whether the group mean or the group standard deviation should be plotted. Use 'mean' or 'std'. The default is 'mean'. If the value is 'std', Y is required to have multiple columns.
- 'parent' - A handle to the figure window for the plots. The default is the current figure window.
[figh,AXESH] = maineffectsplot(...) returns the handle figh to the figure window and an array of handles AXESH to the subplot axes.

Examples
Display main effects plots for car weight with two grouping variables, model year and number of cylinders:

```
load carsmall;
maineffectsplot(Weight, \{Model_Year, Cylinders\}, ...
'varnames',\{'Model Year','\# of Cylinders'\})
```




See Also interactionplot | multivarichart
How To

- "Grouped Data" on page 2-34


## ClassificationDiscriminant.make

## Purpose Construct discriminant analysis classifier from parameters

Syntax<br>Description

Tips

## Input <br> Arguments

cobj = ClassificationDiscriminant.make(Mu,Sigma)
cobj = ClassificationDiscriminant.make(Mu,Sigma,Name,Value)
cobj = ClassificationDiscriminant.make(Mu,Sigma) constructs a compact discriminant analysis classifier from the class means Mu and covariance matrix Sigma.
cobj = ClassificationDiscriminant.make(Mu,Sigma,Name,Value) constructs a compact classifier with additional options specified by one or more Name, Value pair arguments.

- You can change the discriminant type using dot addressing after constructing cobj:

```
cobj.DiscrimType = 'discrimType'
```

where discrimType is one of 'linear', 'quadratic', 'diagLinear', 'diagQuadratic', 'pseudoLinear', or 'pseudoQuadratic'. You can change between linear types or between quadratic types, but cannot change between a linear and a quadratic type.

- cobj is a linear classifier when Sigma is a matrix. cobj is a quadratic classifier when Sigma is a three-dimensional array.


## Mu

Matrix of class means of size K-by-p, where $K$ is the number of classes, and $p$ is the number of predictors. Each row of Mu represents the mean of the multivariate normal distribution of the corresponding class. The class indices are in the ClassNames attribute.

## Sigma

Within-class covariance matrix or matrices.

## ClassificationDiscriminant.make

- For a linear discriminant, Sigma is a symmetric, positive semidefinite matrix of size $p$-by- $p$, where $p$ is the number of predictors.
- For a quadratic discriminant, Sigma is an array of size p-by-p-by-K, where K is the number of classes. For each i, Sigma(:,:,i) is a symmetric, positive semidefinite matrix.


## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

## ClassNames

Array of class names. Use any data type for a grouping variable: a numeric vector, vector of categorical variables (nominal or ordinal), logical vector, character array, or cell array of strings.

ClassNames names the classes, as ordered in Mu.
Default: $1: K$, where $K$ is the number of classes (the number of rows of Mu)

## Cost

Square matrix, where $\operatorname{Cost}(i, j)$ is the cost of classifying a point into class $j$ if its true class is i. Alternatively, Cost can be a structure $S$ having two fields: S . ClassNames containing the group names as a variable of the same type as Y , and S.ClassificationCosts containing the cost matrix.

Default: $\operatorname{Cost}(i, j)=1$ if $i \sim=j$, and $\operatorname{Cost}(i, j)=0$ if $i=j$

## PredictorNames

Cell array of names for the predictor variables, in the order in which they appear in X .

## ClassificationDiscriminant.make

Default: \{'x1','x2',...\}
prior
Prior probabilities for each class. Specify as one of:

- 'uniform', a string meaning all class prior probabilities are equal
- A vector (one scalar value for each class)
- A structure S with two fields:
- S.ClassNames containing the class names as a variable of the same type as ClassNames
- S.ClassProbs containing a vector of corresponding probabilities

Default: 'uniform'

## ResponseName

Name of the response variable Y , a string.
Default: ' $Y$ '

## Output cobj <br> Discriminant analysis classifier, of class CompactClassificationDiscriminant.

## Examples Construct a compact linear discriminant analysis classifier from the means and covariances of the Fisher iris data:

```
load fisheriris
mu(1,:) = mean(meas(1:50,:));
mu(2,:) = mean(meas(51:100,:));
mu(3,:) = mean(meas(101:150,:));
```


## ClassificationDiscriminant.make

```
mm1 = repmat(mu(1,:),50,1);
mm2 = repmat(mu(2,:),50,1);
mm3 = repmat(mu(3,:),50,1);
cc = meas;
cc(1:50,:) = cc(1:50,:) - mm1;
cc(51:100,:) = cc(51:100,:) - mm2;
cc(101:150,:) = cc(101:150,:) - mm3;
sigstar = cc' * cc / 147; % unbiased estimator of sigma
cpct = ClassificationDiscriminant.make(mu,sigstar,...
    'ClassNames',{'setosa','versicolor','virginica'})
cpct =
classreg.learning.classif.CompactClassificationDiscriminant:
    PredictorNames: {'x1' 'x2' 'x3' 'x4'}
        ResponseName: 'Y'
            ClassNames: {'setosa' 'versicolor' 'virginica'}
        ScoreTransform: 'none'
            DiscrimType: 'linear'
            Mu: [3x4 double]
            Coeffs: [3x3 struct]
```

See Also ClassificationDiscriminant.fit | compact | CompactClassificationDiscriminant
How To . "Discriminant Analysis" on page 12-3

```
Purpose One-way multivariate analysis of variance
Syntax \(\quad d=\) manova1 \((X\), group \()\)
d = manova1(X,group,alpha)
[d,p] = manova1(...)
[d,p,stats] = manova1(...)
```

d = manova1(X,group) performs a one-way Multivariate Analysis of Variance (MANOVA) for comparing the multivariate means of the columns of X , grouped by group. X is an $m$-by- $n$ matrix of data values, and each row is a vector of measurements on $n$ variables for a single observation. group is a grouping variable defined as a categorical variable, vector, string array, or cell array of strings. Two observations are in the same group if they have the same value in the group array. (See "Grouped Data" on page 2-34.) The observations in each group represent a sample from a population.

The function returns $d$, an estimate of the dimension of the space containing the group means. manova1 tests the null hypothesis that the means of each group are the same $n$-dimensional multivariate vector, and that any difference observed in the sample X is due to random chance. If $d=0$, there is no evidence to reject that hypothesis. If $d=1$, then you can reject the null hypothesis at the $5 \%$ level, but you cannot reject the hypothesis that the multivariate means lie on the same line. Similarly, if $\mathrm{d}=2$ the multivariate means may lie on the same plane in $n$-dimensional space, but not on the same line.
d = manova1 ( X, group, alpha) gives control of the significance level, alpha. The return value $d$ will be the smallest dimension having $\mathrm{p}>$ alpha, where p is a $p$ value for testing whether the means lie in a space of that dimension.
$[\mathrm{d}, \mathrm{p}]=$ manova1(...) also returns a p , a vector of $p$-values for testing whether the means lie in a space of dimension 0,1 , and so on. The largest possible dimension is either the dimension of the space, or one less than the number of groups. There is one element of $p$ for each dimension up to, but not including, the largest.

If the $i$ th $p$ value is near zero, this casts doubt on the hypothesis that the group means lie on a space of $i-1$ dimensions. The choice of a critical $p$ value to determine whether the result is judged statistically significant is left to the researcher and is specified by the value of the input argument alpha. It is common to declare a result significant if the $p$ value is less than 0.05 or 0.01 .
[d,p,stats] = manova1(...) also returns stats, a structure containing additional MANOVA results. The structure contains the following fields.

| Field | Contents |
| :--- | :--- |
| W | Within-groups sum of squares and cross-products <br> matrix |
| B | Between-groups sum of squares and cross-products <br> matrix |
| T | Total sum of squares and cross-products matrix |
| dfW | Degrees of freedom for W |
| dfB | Degrees of freedom for B |
| dfT | Degrees of freedom for T |
| lambda | Vector of values of Wilk's lambda test statistic for <br> testing whether the means have dimension 0, 1, etc. |
| chisq | Transformation of lambda to an approximate <br> chi-square distribution |
| chisqdf | Degrees of freedom for chisq |
| eigenval | Eigenvalues of $W^{-1} B$ |
| eigenvec | Eigenvectors of $W^{-1} B ;$ these are the coefficients for <br> the canonical variables C, and they are scaled so the <br> within-group variance of the canonical variables is 1 |


| Field | Contents |
| :--- | :--- |
| canon | Canonical variables C, equal to XC*eigenvec, where XC <br> is X with columns centered by subtracting their means |
| mdist | A vector of Mahalanobis distances from each point <br> to the mean of its group |
| gmdist | A matrix of Mahalanobis distances between each pair <br> of group means |

The canonical variables C are linear combinations of the original variables, chosen to maximize the separation between groups. Specifically, C(:,1) is the linear combination of the X columns that has the maximum separation between groups. This means that among all possible linear combinations, it is the one with the most significant $F$ statistic in a one-way analysis of variance. $C(:, 2)$ has the maximum separation subject to it being orthogonal to $C(:, 1)$, and so on.
You may find it useful to use the outputs from manova1 along with other functions to supplement your analysis. For example, you may want to start with a grouped scatter plot matrix of the original variables using gplotmatrix. You can use gscatter to visualize the group separation using the first two canonical variables. You can use manovacluster to graph a dendrogram showing the clusters among the group means.

## Assumptions

The MANOVA test makes the following assumptions about the data in $X$ :

- The populations for each group are normally distributed.
- The variance-covariance matrix is the same for each population.
- All observations are mutually independent.


## Examples

you can use manova1 to determine whether there are differences in the averages of four car characteristics, among groups defined by the country where the cars were made.

```
load carbig
[d,p] = manova1([MPG Acceleration Weight Displacement],...
                                    Origin)
d =
    3
p =
        0
    0.0000
    0.0075
    0.1934
```

There are four dimensions in the input matrix, so the group means must lie in a four-dimensional space. manova1 shows that you cannot reject the hypothesis that the means lie in a 3-D subspace.

## References <br> [1] Krzanowski, W. J. Principles of Multivariate Analysis: A User's Perspective. New York: Oxford University Press, 1988.

See Also<br>anova1 | canoncorr | gscatter | gplotmatrix<br>How To<br>- "Grouped Data" on page 2-34<br>- manovacluster

Purpose Dendrogram of group mean clusters following MANOVA

Syntax $\quad$| manovacluster(stats) |  |
| :--- | :--- |
|  | manovacluster(stats, method) |
|  | $H=$ manovacluster(stats, method) |

## Description

manovacluster(stats) generates a dendrogram plot of the group means after a multivariate analysis of variance (MANOVA). stats is the output stats structure from manova1. The clusters are computed by applying the single linkage method to the matrix of Mahalanobis distances between group means.

See dendrogram for more information on the graphical output from this function. The dendrogram is most useful when the number of groups is large.
manovacluster(stats, method) uses the specified method in place of single linkage. method can be any of the following character strings that identify ways to create the cluster hierarchy. (See linkage for additional information.)

| Method | Description |
| :--- | :--- |
| 'single' | Shortest distance (default) |
| 'complete' | Largest distance |
| 'average' | Average distance |
| 'centroid' | Centroid distance |
| 'ward' | Incremental sum of squares |

H = manovacluster(stats, method) returns a vector of handles to the lines in the figure.

Examples $\quad \begin{aligned} & \text { Let's analyze the larger car data set to determine which countries } \\ & \text { produce cars with the most similar characteristics. }\end{aligned}$

```
load carbig
```



See Also
cluster | dendrogram | linkage | manova1

## CompactClassificationDiscriminant.margin

Purpose Classification margins
Syntax $\quad m=\operatorname{margin}(o b j, x, Y)$

Description

Input Arguments

## Output <br> Arguments <br> m

## Definitions

obj

X

Y
$m=$ margin (obj, $X, Y$ ) returns the classification margins for the matrix of predictors $X$ and class labels Y. For the definition, see "Definitions" on page 20-1000.

Discriminant analysis classifier of class ClassificationDiscriminant or CompactClassificationDiscriminant, typically constructed with ClassificationDiscriminant.fit.

Matrix where each row represents an observation, and each column represents a predictor. The number of columns in X must equal the number of predictors in obj.

Class labels, with the same data type as exists in obj. The number of elements of $Y$ must equal the number of rows of $X$.

Numeric column vector of length size( $X, 1$ ). Each entry in $m$ represents the margin for the corresponding rows of $X$ and (true class) Y , computed using obj.

## Margin

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix $X$. A high value of margin indicates a more reliable prediction than a low value.

## CompactClassificationDiscriminant.margin

## Score (discriminant analysis)

For discriminant analysis, the score of a classification is the posterior probability of the classification. For the definition of posterior probability in discriminant analysis, see "Posterior Probability" on page 12-7.

## Examples Compute the classification margin for the Fisher iris data, trained on

 its first two columns of data, and view the last 10 entries:```
load fisheriris
X = meas(:,1:2);
obj = ClassificationDiscriminant.fit(X,species);
M = margin(obj,X,species);
M(end-10:end)
ans =
    0.6551
        0.4838
        0.6551
        -0.5127
        0.5659
        0.4611
        0.4949
        0.1024
        0.2787
        -0.1439
        -0.4444
```

The classifier trained on all the data is better:

```
obj = ClassificationDiscriminant.fit(meas,species);
M = margin(obj,meas,species);
M(end-10:end)
ans =
    0.9983
    1.0000
```


## CompactClassificationDiscriminant.margin

0.9991
0.9978
1.0000
1.0000
0.9999
0.9882
0.9937
1.0000
0.9649
See Also
ClassificationDiscriminant | edge | loss | predict
How To

- "Discriminant Analysis" on page 12-3


## CompactClassificationEnsemble.margin

## Purpose Classification margins

Syntax $\quad$| $M$ | $=\operatorname{margin}(e n s, X, Y)$ |
| ---: | :--- |
| $M$ | $=\operatorname{margin}(e n s, X, Y$, Name, Value $)$ |

Description

## Input Arguments

$M=$ margin(ens, $X, Y$ ) returns the classification margin for the predictions of ens on data $X$, when the true classifications are Y .
$M=\operatorname{margin}(e n s, X, Y$, Name, Value) calculates margin with additional options specified by one or more Name, Value pair arguments.
ens
Classification ensemble created with fitensemble, or a compact classification ensemble created with compact.

X
Matrix of data to classify. Each row of X represents one observation, and each column represents one predictor. $X$ must have the same number of columns as the data used to train ens. $X$ should have the same number of rows as the number of elements in $Y$.

## Y

Classification of $X$. $Y$ should be of the same type as the classification used to train ens, and its number of elements should equal the number of rows of $X$.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
learners

## CompactClassificationEnsemble.margin

Indices of weak learners in the ensemble ranging from 1 to ens.NTrained. oobEdge uses only these learners for calculating loss.

Default: 1:NTrained

```
UseObsForLearner
```

A logical matrix of size N -by-T, where:

- $N$ is the number of rows of $X$.
- T is the number of weak learners in ens.

When UseObsForLearner ( $\mathrm{i}, \mathrm{j}$ ) is true, learner j is used in predicting the class of row i of $X$.

Default: $\operatorname{true}(\mathrm{N}, \mathrm{T})$

## Output

 Arguments
## Definitions

M
A numeric column vector with the same number of rows as $X$. Each row of $M$ gives the classification margin for that row of $X$.

## Margin

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix $X$.

## Score (ensemble)

For ensembles, a classification score represents the confidence of a classification into a class. The higher the score, the higher the confidence.

Different ensemble algorithms have different definitions for their scores. Furthermore, the range of scores depends on ensemble type. For example:

## CompactClassificationEnsemble.margin

- AdaBoostM1 scores range from $-\infty$ to $\infty$.
- Bag scores range from 0 to 1.

```
Examples
Find the margin for classifying an average flower from the Fisheriris data as 'versicolor':
```

```
load fisheriris % X = meas, Y = species
```

load fisheriris % X = meas, Y = species
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
flower = mean(meas);
flower = mean(meas);
predict(ens,flower)
predict(ens,flower)
ans =
ans =
'versicolor'
'versicolor'
margin(ens,mean(meas),'versicolor')
margin(ens,mean(meas),'versicolor')
ans =
ans =
3.2140

```
    3.2140
```

See Also

predict | edge | loss
How To - Chapter 13, "Nonparametric Supervised Learning"

## CompactClassificationTree.margin

## Purpose Classification margins

$$
\text { Syntax } \quad m=\operatorname{margin}(\text { tree }, X, Y)
$$

Description

## Input <br> Arguments

Output
Arguments

## Definitions

tree
A classification tree created by ClassificationTree.fit, or a compact classification tree created by compact.

X

A matrix where each row represents an observation, and each column represents a predictor. The number of columns in X must equal the number of predictors in tree.

Y
Class labels, with the same data type as exists in tree.
m
A numeric column vector of length size ( $X, 1$ ). Each entry in $m$ represents the margin for the corresponding rows of $X$ and (true class) Y , computed using tree.

## Margin

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix $X$.

## Score (tree)

For trees, the score of a classification of a leaf node is the posterior probability of the classification at that node. The posterior probability of the classification at a node is the number of training sequences that

## CompactClassificationTree.margin

lead to that node with the classification, divided by the number of training sequences that lead to that node.

For example, consider classifying a predictor $X$ as true when $X<0.15$ or $X>0.95$, and $X$ is false otherwise.

1
Generate 100 random points and classify them:

```
rng(0,'twister') % for reproducibility
X = rand(100,1);
Y = (abs(X - .55) > .4);
tree = ClassificationTree.fit(X,Y);
view(tree,'mode','graph')
```



2
Prune the tree:

```
tree1 = prune(tree,'level',1);
view(tree1,'mode','graph')
```


## CompactClassificationTree.margin



The pruned tree correctly classifies observations that are less than 0.15 as true. It also correctly classifies observations from .15 to .94 as false. However, it incorrectly classifies observations that are greater than . 94 as false. Therefore, the score for observations that are greater than .15 should be about $.05 / .85=.06$ for true, and about $.8 / .85=.94$ for false.

3
Compute the prediction scores for the first 10 rows of X :

```
[~,score] = predict(tree1,X(1:10));
[score X(1:10,:)]
ans =
\begin{tabular}{rrr}
0.9059 & 0.0941 & 0.8147 \\
0.9059 & 0.0941 & 0.9058 \\
0 & 1.0000 & 0.1270 \\
0.9059 & 0.0941 & 0.9134 \\
0.9059 & 0.0941 & 0.6324 \\
0 & 1.0000 & 0.0975 \\
0.9059 & 0.0941 & 0.2785
\end{tabular}
```


## CompactClassificationTree.margin

| 0.9059 | 0.0941 | 0.5469 |
| :--- | :--- | :--- |
| 0.9059 | 0.0941 | 0.9575 |
| 0.9059 | 0.0941 | 0.9649 |

Indeed, every value of $X$ (the rightmost column) that is less than 0.15 has associated scores (the left and center columns) of 0 and 1 , while the other values of $X$ have associated scores of 0.91 and 0.09 . The difference (score 0.09 instead of the expected .06) is due to a statistical fluctuation: there are 8 observations in X in the range (. 95,1 ) instead of the expected 5 observations.

## Examples

Compute the classification margin for the Fisher iris data, trained on its first two columns of data, and view the last 10 entries:

```
load fisheriris
X = meas(:,1:2);
tree = ClassificationTree.fit(X,species);
M = margin(tree,X,species);
M(end-10:end)
ans =
    0.1111
        0 . 1 1 1 1
        0.1111
        -0.2857
        0.6364
        0.6364
        0.1111
        0.7500
        1.0000
        0.6364
        0.2000
```

The classification tree trained on all the data is better:

```
tree = ClassificationTree.fit(meas,species);
M = margin(tree,meas,species);
M(end-10:end)
```


## CompactClassificationTree.margin

$$
\text { ans } \begin{aligned}
& = \\
& 0.9565 \\
& 0.9565 \\
& 0.9565 \\
& 0.9565 \\
& 0.9565 \\
& 0.9565 \\
& 0.9565 \\
& 0.9565 \\
& 0.9565 \\
& 0.9565 \\
& 0.9565
\end{aligned}
$$

See Also predict | loss | edge
How To - Chapter 13, "Nonparametric Supervised Learning"

## CompactTreeBagger.margin

| Purpose | Classification margin |
| :---: | :---: |
| Syntax | ```mar = margin(B,X,Y) mar = margin(B,X,Y,'param1',val1,'param2',val2,...)``` |
| Description | mar $=$ margin $(B, X, Y)$ computes the classification margins for predictors $X$ given true response $Y$. The $Y$ can be either a numeric vector, character matrix, cell array of strings, categorical vector or logical vector. mar is a numeric array of size Nobs-by-NTrees, where Nobs is the number of rows of $X$ and $Y$, and NTrees is the number of trees in the ensemble B. For observation I and tree $J$, mar $(I, J)$ is the difference between the score for the true class and the largest score for other classes. This method is available for classification ensembles only. <br> mar = margin(B,X,Y,'param1',val1,'param2',val2, $\ldots$ ) specifies optional parameter name/value pairs: |
|  | 'mode' String indicating how the method computes errors. If set to 'cumulative' (default), margin computes cumulative errors and mar is an Nobs-by-NTrees matrix, where the first column gives error from trees(1), second column gives error fromtrees (1:2) etc, up to trees(1:NTrees). If set to 'individual', mar is a Nobs-by-NTrees matrix, where each element is an error from each tree in the ensemble. If set to 'ensemble', mar a single column of length Nobs showing the cumulative margins for the entire ensemble. |
|  | 'trees ' Vector of indices indicating what trees to include in this calculation. By default, this argument is set to 'all' and the method uses all trees. If 'trees' is a numeric vector, the method returns a vector of length NTrees for 'cumulative' and 'individual' modes, where NTrees is the number of elements in the input vector, and a scalar for 'ensemble' mode. For example, in the 'cumulative' mode, the first element |

## CompactTreeBagger.margin

gives error from trees(1), the second element gives error from trees(1:2) etc.
'treeweights' Vector of tree weights. This vector must have the same length as the 'trees' vector. The method uses these weights to combine output from the specified trees by taking a weighted average instead of the simple non-weighted majority vote. You cannot use this argument in the 'individual' mode.
'useifort' Logical matrix of size Nobs-by-NTrees indicating which trees should be used to make predictions for each observation. By default the method uses all trees for all observations.

## See Also

TreeBagger.margin

## TreeBagger.margin



## TreeBagger.margin

gives error from trees(1), the second element gives error from trees(1:2) etc.
'treeweights ' Vector of tree weights. This vector must have the same length as the 'trees' vector. The method uses these weights to combine output from the specified trees by taking a weighted average instead of the simple non-weighted majority vote. You cannot use this argument in the 'individual' mode.
'useifort' Logical matrix of size Nobs-by-NTrees indicating which trees should be used to make predictions for each observation. By default the method uses all trees for all observations.

## See Also

CompactTreeBagger.margin

## Purpose Nonclassical multidimensional scaling

Syntax $\quad Y=\operatorname{mdscale}(D, p)$
[Y,stress] = mdscale(D, p)
[Y,stress,disparities] = mdscale(D, p)
[...] = mdscale(D, p,'Name', value)
Description
$Y=$ mdscale ( $D, p$ ) performs nonmetric multidimensional scaling on the $n$-by- $n$ dissimilarity matrix D , and returns Y , a configuration of $n$ points (rows) in $p$ dimensions (columns). The Euclidean distances between points in Y approximate a monotonic transformation of the corresponding dissimilarities in D. By default, mdscale uses Kruskal's normalized stress1 criterion.

You can specify D as either a full $n$-by- $n$ matrix, or in upper triangle form such as is output by pdist. A full dissimilarity matrix must be real and symmetric, and have zeros along the diagonal and non-negative elements everywhere else. A dissimilarity matrix in upper triangle form must have real, non-negative entries. mdscale treats NaNs in D as missing values, and ignores those elements. Inf is not accepted.

You can also specify D as a full similarity matrix, with ones along the diagonal and all other elements less than one. mdscale transforms a similarity matrix to a dissimilarity matrix in such a way that distances between the points returned in Y approximate sqrt(1-D). To use a different transformation, transform the similarities prior to calling mdscale.
[ $\mathrm{Y}, \mathrm{stress}$ ] = mdscale ( $\mathrm{D}, \mathrm{p}$ ) returns the minimized stress, i.e., the stress evaluated at $Y$.
[Y,stress,disparities] = mdscale(D,p) returns the disparities, that is, the monotonic transformation of the dissimilarities $D$.
[...] = mdscale(D, p, 'Name', value) specifies one or more optional parameter name/value pairs that control further details of mdscale. Specify Name in single quotes. Available parameters are

## mdscale

- Criterion- The goodness-of-fit criterion to minimize. This also determines the type of scaling, either non-metric or metric, that mdscale performs. Choices for non-metric scaling are:
- 'stress' - Stress normalized by the sum of squares of the inter-point distances, also known as stress1. This is the default.
- 'sstress' - Squared stress, normalized with the sum of 4th powers of the inter-point distances.

Choices for metric scaling are:

- 'metricstress' - Stress, normalized with the sum of squares of the dissimilarities.
- 'metricsstress' - Squared stress, normalized with the sum of 4th powers of the dissimilarities.
- 'sammon' - Sammon's nonlinear mapping criterion. Off-diagonal dissimilarities must be strictly positive with this criterion.
- 'strain' - A criterion equivalent to that used in classical multidimensional scaling.
- Weights - A matrix or vector the same size as D, containing nonnegative dissimilarity weights. You can use these to weight the contribution of the corresponding elements of $D$ in computing and minimizing stress. Elements of D corresponding to zero weights are effectively ignored.

Note When you specify weights as a full matrix, its diagonal elements are ignored and have no effect, since the corresponding diagonal elements of $D$ do not enter into the stress calculation.

- Start - Method used to choose the initial configuration of points for Y. The choices are
- 'cmdscale' - Use the classical multidimensional scaling solution. This is the default. 'cmdscale' is not valid when there are zero weights.
- 'random' - Choose locations randomly from an appropriately scaled p-dimensional normal distribution with uncorrelated coordinates.
- An $n$-by-p matrix of initial locations, where n is the size of the matrix $D$ and $p$ is the number of columns of the output matrix $Y$. In this case, you can pass in [] for $p$ and mdscale infers $p$ from the second dimension of the matrix. You can also supply a 3-D array, implying a value for 'Replicates' from the array's third dimension.
- Replicates - Number of times to repeat the scaling, each with a new initial configuration. The default is 1 .
- Options - Options for the iterative algorithm used to minimize the fitting criterion. Pass in an options structure created by statset. For example,

```
opts = statset(param1,val1,param2,val2, ...);
[...] = mdscale(...,'Options',opts)
```

The choices of statset parameters are

- 'Display ' - Level of display output. The choices are 'off' (the default), 'iter', and 'final'.
- 'MaxIter' - Maximum number of iterations allowed. The default is 200 .
- 'TolFun' - Termination tolerance for the stress criterion and its gradient. The default is 1e-4.
- 'TolX' - Termination tolerance for the configuration location step size. The default is $1 \mathrm{e}-4$.


## Examples

load cereal.mat
X = [Calories Protein Fat Sodium Fiber ...

## mdscale

```
Carbo Sugars Shelf Potass Vitamins];
% Take a subset from a single manufacturer.
X = X(strcmp('K',cellstr(Mfg)),:);
% Create a dissimilarity matrix.
dissimilarities = pdist(X);
% Use non-metric scaling to recreate the data in 2D,
% and make a Shepard plot of the results.
[Y,stress,disparities] = mdscale(dissimilarities,2);
distances = pdist(Y);
[dum,ord] = sortrows([disparities(:) dissimilarities(:)]);
plot(dissimilarities,distances,'bo', ...
dissimilarities(ord),disparities(ord),'r.-');
xlabel('Dissimilarities'); ylabel('Distances/Disparities')
legend({'Distances' 'Disparities'},'Location','NW');
```


\% Do metric scaling on the same dissimilarities. figure
[Y,stress] = ...
mdscale(dissimilarities,2, 'criterion','metricsstress');
distances = pdist(Y);
plot(dissimilarities,distances,'bo', ...
[0 max(dissimilarities)],[0 max(dissimilarities)],'r.-');
xlabel('Dissimilarities'); ylabel('Distances')


See Also<br>cmdscale | pdist | statset

## CompactTreeBagger.mdsProx

| Purpose | Multidimensional scaling of proximity matrix |
| :---: | :---: |
| Syntax | $\begin{aligned} {[S C, E I G E N] } & =\operatorname{mdsProx}(B, X) \\ {[S C, E I G E N] } & =\operatorname{mdsProx}(B, X, ' p a r a m 1 ', \text { val1,'param2',val2, } . .) \end{aligned}$ |
| Description | [SC,EIGEN] = mdsProx (B, X) applies classical multidimensional scaling to the proximity matrix computed for the data in the matrix $X$, and returns scaled coordinates SC and eigenvalues EIGEN of the scaling transformation. The method applies multidimensional scaling to the matrix of distances defined as 1-prox, where prox is the proximity matrix returned by the proximity method. |
|  | You can supply the proximity matrix directly by using the 'data' parameter. <br> [SC,EIGEN] = mdsProx (B, X,'param1', val1,'param2', val2, ...) specifies optional parameter name/value pairs: |
|  | 'data' Flag indicating how the method treats the X input argument. If set to 'predictors' (default), mdsProx assumes $X$ to be a matrix of predictors and used for computation of the proximity matrix. If set to 'proximity ', the method treats X as a proximity matrix returned by the proximity method. |
|  | 'colors' If you supply this argument, mdsProx makes overlaid scatter plots of two scaled coordinates using specified colors for different classes. You must supply the colors as a string with one character for each color. If there are more classes in the data than characters in the supplied string, mdsProx only plots the first C classes, where C is the length of the string. For regression or if you do not provide the vector of true class labels, the method uses the first color for all observations in $X$. |

## CompactTreeBagger.mdsProx

'labels' Vector of true class labels for a classification ensemble. True class labels can be either a numeric vector, character matrix, or cell array of strings. If supplied, this vector must have as many elements as there are observations (rows) in X. This argument has no effect unless you also supply the 'colors' argument.
'mdscoords' Indices of the two scaled coordinates to plot. By default, mdsProx makes a scatter plot of the first and second scaled coordinates which correspond to the two largest eigenvalues. You can specify any other two or three indices not exceeding the dimensionality of the scaled data. This argument has no effect unless you also supply the 'colors' argument.

See Also
cmdscale | TreeBagger.mdsProx | proximity

## TreeBagger.mdsProx

| Purpose | Multidimensional scaling of proximity matrix |
| :---: | :---: |
| Syntax | ```[S,E] = mdsProx(B) [S,E] = mdsProx(B,'param1',val1,'param2',val2,...)``` |
| Description | [S,E] = mdsProx(B) returns scaled coordinates, S, and eigenvalues, E , for the proximity matrix in the ensemble B . An earlier call to fillProximities(B) must create the proximity matrix. |
|  |  optional parameter name/value pairs: |
|  | keep ' Array of indices of observations in the training data to use for multidimensional scaling. By default, this argument is set to 'all'. If you provide numeric or logical indices, the method uses only the subset of the training data specified by these indices to compute the scaled coordinates and eigenvalues. |
|  | 'colors ' If you supply this argument, mdsProx makes overlaid scatter plots of two scaled coordinates using specified colors for different classes. You must supply the colors as a string with one character for each color. If there are more classes in the data than characters in the supplied string, mdsProx only plots the first C classes, where $C$ is the length of the string. For regression or if you do not provide the vector of true class labels, the method uses the first color for all observations in $X$. |
|  | 'mdscoords' Indices of the two scaled coordinates to plot. By default, mdsProx makes a scatter plot of the first and second scaled coordinates which correspond to the two largest eigenvalues. You can specify any other two or three indices not exceeding the dimensionality of the scaled data. This argument has no effect unless you also supply the 'colors' argument. |

See Also

## ProbDistUnivParam.mean

## Purpose Return mean of ProbDistUnivParam object

## Syntax $\quad M=\operatorname{mean}(P D)$

Description
$M=$ mean $(P D)$ returns $M$, the mean of the ProbDistUnivParam object $P D$.

Input
$P D$
An object of the class ProbDistUnivParam.
Arguments

Output $\quad M$
Arguments
The mean of the ProbDistUnivParam object $P D$.

See Also mean

## CompactTreeBagger.meanMargin

| Purpose | Mean classification margin |
| :---: | :---: |
| Syntax | ```mar = meanMargin(B,X,Y) mar = meanMargin(B,X,Y,'param1',val1,'param2',val2,\ldots)``` |
| Description | mar $=$ meanMargin $(B, X, Y)$ computes average classification margins for predictors $X$ given true response $Y$. The $Y$ can be either a numeric vector, character matrix, cell array of strings, categorical vector or logical vector. meanMargin averages the margins over all observations (rows) in $X$ for each tree. mar is a matrix of size 1-by-NTrees, where NTrees is the number of trees in the ensemble $B$. This method is available for classification ensembles only. <br> mar $=$ meanMargin( $B, X, Y$, 'param1', val1,'param2', val2, $\ldots$ ) specifies optional parameter name/value pairs: <br> 'mode' <br> String indicating how meanMargin computes errors. If set to 'cumulative' (default), is a vector of length NTrees where the first element gives mean margin from trees(1), second column gives mean margins from trees(1:2) etc, up to trees(1:NTrees). If set to 'individual', mar is a vector of length NTrees, where each element is a mean margin from each tree in the ensemble. If set to 'ensemble', mar is a scalar showing the cumulative mean margin for the entire ensemble. |
|  | 'trees ' Vector of indices indicating what trees to include in this calculation. By default, this argument is set to 'all' and the method uses all trees. If 'trees' is a numeric vector, the method returns a vector of length NTrees for 'cumulative' and 'individual' modes, where NTrees is the number of elements in the input vector, and a scalar for 'ensemble' mode. For example, in the 'cumulative' mode, the first element |

## CompactTreeBagger.meanMargin

> gives mean margin from trees(1), the second element gives mean margin from trees(1:2) etc.
> 'treeweights ' Vector of tree weights. This vector must have the same length as the 'trees' vector. meanMargin uses these weights to combine output from the specified trees by taking a weighted average instead of the simple nonweighted majority vote. You cannot use this argument in the 'individual' mode.

See Also
TreeBagger.meanMargin

## TreeBagger.meanMargin

$\left.\begin{array}{ll}\text { Purpose } & \text { Mean classification margin } \\ \text { Syntax } & \begin{array}{l}\text { mar }=\text { meanMargin }(B, X, Y) \\ \text { mar }=\text { meanMargin }(B, X, Y, ~ ' p a r a m 1 ' ~\end{array} \text {, val1, ' param2 ' , val2, } \ldots \text { ) }\end{array}\right\}$

## TreeBagger.meanMargin

gives mean margin from trees(1), the second element gives mean margin from trees (1:2) etc.
'treeweights ' Vector of tree weights. This vector must have the same length as the 'trees' vector. meanMargin uses these weights to combine output from the specified trees by taking a weighted average instead of the simple nonweighted majority vote. You cannot use this argument in the 'individual' mode.

See Also
CompactTreeBagger.meanMargin

| Purpose | Mean predictive measure of association for surrogate splits in decision tree |
| :---: | :---: |
| Syntax | ```MA = meansurrvarassoc(T) MA = meansurrvarassoc(T,N)``` |
| Description | MA $=$ meansurrvarassoc ( $T$ ) returns a $p$-by- $p$ matrix, MA, with predictive measures of association for $p$ predictors. Element MA $(i, j)$ is the predictive measure of association averaged over surrogate splits on predictor $j$ for which predictor $i$ is the optimal split predictor. This average is computed by summing positive values of the predictive measure of association over optimal splits on predictor $i$ and surrogate splits on predictor $j$ and dividing by the total number of optimal splits on predictor $i$, including splits for which the predictive measure of association between predictors $i$ and $j$ is negative. <br> MA = meansurrvarassoc ( $T, N$ ) takes an array $N$ of node numbers and returns the predictive measure of association averaged over the specified nodes. |
| See Also | classregtree \| surrcutvar | surrvarassoc | surrcutcategories | surrcuttype | surrcutpoint | surrcutflip |

## CompactClassificationTree.meanSurrVarAssoc

| Purpose | Mean predictive measure of association for surrogate splits in decision tree |
| :---: | :---: |
| Syntax | ma = meanSurrVarAssoc (tree) <br> ma $=$ meanSurrVarAssoc(tree,N) |
| Description | $\mathrm{ma}=$ meanSurrVarAssoc (tree) returns a matrix of predictive measures of association for the predictors in tree. <br> $\mathrm{ma}=$ meanSurrVarAssoc(tree,N) returns a matrix of predictive measures of association averaged over the nodes in vector N . |
| Input Arguments | tree <br> A classification tree constructed with ClassificationTree.fit, or a compact regression tree constructed with compact. |
|  | N <br> Vector of node numbers in tree. |
| Output Arguments | - ma $=$ meanSurrVarAssoc(tree) returns a P-by-P matrix, where $P$ is the number of predictors in tree. $m a(i, j)$ is the predictive measure of association between the optimal split on variable i and a surrogate split on variable j. See "Predictive Measure of Association" on page 20-1030. <br> - ma $=$ meanSurrVarAssoc (tree, $N$ ) returns a P-by-P representing the predictive measure of association between variables averaged over nodes in the vector $N$. $N$ contains node numbers from 1 to max (tree. NumNodes). |
| Definitions | Predictive Measure of Association |
|  | The predictive measure of association between the optimal split on variable $i$ and a surrogate split on variable $j$ is: |

## CompactClassificationTree.meanSurrVarAssoc

$$
\lambda_{i, j}=\frac{\min \left(P_{L}, P_{R}\right)-\left(1-P_{L_{i} L_{j}}-P_{R_{i} R_{j}}\right)}{\min \left(P_{L}, P_{R}\right)}
$$

Here

- $P_{L}$ and $P_{R}$ are the node probabilities for the optimal split of node $i$ into Left and Right nodes respectively.
- $P_{L_{i} L_{j}}$ is the probability that both (optimal) node $i$ and (surrogate) node $j$ send an observation to the Left.
- $P_{R_{j} R_{j}}$ is the probability that both (optimal) node $i$ and (surrogate) node $j$ send an observation to the Right.

Clearly, $\lambda_{i, j}$ lies from $-\infty$ to 1 . Variable $j$ is a worthwhile surrogate split for variable $i$ if $\lambda_{i, j}>0$.
Element ma( $\mathrm{i}, \mathrm{j})$ is the predictive measure of association averaged over surrogate splits on predictor $j$ for which predictor $i$ is the optimal split predictor. This average is computed by summing positive values of the predictive measure of association over optimal splits on predictor i and surrogate splits on predictor $j$ and dividing by the total number of optimal splits on predictor $i$, including splits for which the predictive measure of association between predictors $i$ and $j$ is negative.

## Examples

Find the mean predictive measure of association between the variables in the Fisher iris data:

```
load fisheriris
obj = ClassificationTree.fit(meas,species,'surrogate','on');
msva = meanSurrVarAssoc(obj)
msva =
\begin{tabular}{rrrr}
1.0000 & 0 & 0 & 0 \\
0 & 1.0000 & 0 & 0 \\
0.4633 & 0.2500 & 1.0000 & 0.5000
\end{tabular}
```


## CompactClassificationTree.meanSurrVarAssoc

$$
\begin{array}{llll}
0.2065 & 0.1413 & 0.4022 & 1.0000
\end{array}
$$

Find the mean predictive measure of association averaged over the odd-numbered nodes in obj:

```
N = 1:2:obj.NumNodes;
msva = meanSurrVarAssoc(obj,N)
msva =
\begin{tabular}{rrrr}
1.0000 & 0 & 0 & 0 \\
0 & 1.0000 & 0 & 0 \\
0.7600 & 0.5000 & 1.0000 & 1.0000 \\
0.4130 & 0.2826 & 0.8043 & 1.0000
\end{tabular}
```

See Also ClassificationTree
How To - Chapter 13, "Nonparametric Supervised Learning"

## CompactRegressionTree.meanSurrVarAssoc

## Purpose

Syntax

Description

Input
Arguments

## Output ma <br> Arguments

 tree
## tree

N

Mean predictive measure of association for surrogate splits in decision
ma = meanSurrVarAssoc(tree) ma = meanSurrVarAssoc(tree,N)
ma $=$ meanSurrVarAssoc(tree) returns a matrix of predictive measures of association for the predictors in tree.
$\mathrm{ma}=$ meanSurrVarAssoc (tree, N ) returns a matrix of predictive measures of association averaged over the nodes in vector $N$.

A regression tree constructed with RegressionTree.fit, or a compact regression tree constructed with compact.

Vector of node numbers in tree.

- ma = meanSurrVarAssoc(tree) returns a P-by-P matrix, where $P$ is the number of predictors in tree. ma( $i, j)$ is the predictive measure of association between the optimal split on variable i and a surrogate split on variable j. See "Predictive Measure of Association" on page 20-1033.
- ma $=$ meanSurrVarAssoc (tree, $N$ ) returns a P-by-P representing the predictive measure of association between variables averaged over nodes in the vector $N . N$ contains node numbers from 1 to max (tree. NumNodes).


## Definitions

## Predictive Measure of Association

The predictive measure of association between the optimal split on variable $i$ and a surrogate split on variable $j$ is:

## CompactRegressionTree.meanSurrVarAssoc

$$
\lambda_{i, j}=\frac{\min \left(P_{L}, P_{R}\right)-\left(1-P_{L_{i} L_{j}}-P_{R_{i} R_{j}}\right)}{\min \left(P_{L}, P_{R}\right)} .
$$

Here

- $P_{L}$ and $P_{R}$ are the node probabilities for the optimal split of node $i$ into Left and Right nodes respectively.
- $P_{L_{i} L_{j}}$ is the probability that both (optimal) node $i$ and (surrogate) node $j$ send an observation to the Left.
- $P_{R_{i} R_{j}}$ is the probability that both (optimal) node $i$ and (surrogate) node $j$ send an observation to the Right.

Clearly, $\lambda_{i, j}$ lies from $-\infty$ to 1 . Variable $j$ is a worthwhile surrogate split for variable $i$ if $\lambda_{i, j}>0$.
Element ma( $\mathrm{i}, \mathrm{j})$ is the predictive measure of association averaged over surrogate splits on predictor $j$ for which predictor $i$ is the optimal split predictor. This average is computed by summing positive values of the predictive measure of association over optimal splits on predictor i and surrogate splits on predictor $j$ and dividing by the total number of optimal splits on predictor $i$, including splits for which the predictive measure of association between predictors $i$ and $j$ is negative.

## Examples

Find the mean predictive measure of association between the Displacement, Horsepower, and Weight variables in the carsmall data:

```
load carsmall
X = [Displacement Horsepower Weight];
tree = RegressionTree.fit(X,MPG,'surrogate','on');
ma = meanSurrVarAssoc(tree)
ma =
\begin{tabular}{lll}
1.0000 & 0.2708 & 0.3854 \\
0.4764 & 1.0000 & 0.4568
\end{tabular}
```


## CompactRegressionTree.meanSurrVarAssoc

## $0.3472 \quad 0.2326 \quad 1.0000$

Find the mean predictive measure of association averaged over the odd-numbered nodes in tree:

```
N = 1:2:tree.NumNodes;
ma = meanSurrVarAssoc(tree,N)
ma =
            1.0000 0.2500 0.3750
            0.5910 1.0000 0.5861
            0.5000 0.2361 1.0000
```

See Also prune | RegressionTree
How To • Chapter 13, "Nonparametric Supervised Learning"

## ProbDistUnivKernel.median

Purpose Return median of ProbDistUnivKernel object
Syntax ..... $M=\operatorname{median}(P D)$
Description $M=$ median (PD) returns $M$, the median of the ProbDistUnivKernel object $P D$.
Input ..... PD
An object of the class ProbDistUnivKernel.
Output ..... M
Arguments
See Also
median

## ProbDistUnivParam.median

## Purpose Return median of ProbDistUnivParam object

Syntax
$M=$ median (PD)

Description $\quad \begin{aligned} & M=\text { median }(P D) \text { returns } M \text {, the median of the ProbDistUnivParam } \\ & \text { object } P D .\end{aligned}$

## Input <br> PD <br> An object of the class ProbDistUnivParam.

## Output <br> Arguments <br> M

See Also median

The median of the ProbDistUnivParam object $P D$.

## TreeBagger.MergeLeaves property

Purpose Flag to merge leaves that do not improve risk

# Description <br> The MergeLeaves property is true if decision trees have their leaves with the same parent merged for splits that do not decrease the total risk, and false otherwise. The default value is false. 

See Also<br>classregtree

## Purpose Merge levels

```
Syntax
B = mergelevels(A,oldlevels,newlevel)
B = mergelevels(A,oldlevels)
```

Description $\quad B=$ mergelevels( $A$, oldlevels, newlevel) merges two or more levels of the categorical array A into a single new level. oldlevels is a cell array of strings or a 2-D character matrix that specifies the levels to be merged. Any elements of A that have levels in oldlevels are assigned the new level in the corresponding elements of $B$. newlevel is a character string that specifies the label for the new level. For ordinal arrays, the levels of A specified by oldlevels must be consecutive, and mergelevels inserts the new level to preserve the order of the levels.
$B=$ mergelevels(A,oldlevels) merges two or more levels of A. For nominal arrays, mergelevels uses the first label in oldlevels as the label for the new level. For ordinal arrays, mergelevels uses the label corresponding to the lowest level in oldlevels as the label for the new level.

## Examples Example 1

For nominal data:

```
load fisheriris
species = nominal(species);
species = mergelevels(species,...
    {'setosa','virginica'},'parent');
species = setlabels(species,'hybrid','versicolor');
getlabels(species)
ans =
    'hybrid' 'parent'
```


## ordinal.mergelevels

## Example 2

For ordinal data:

```
A = ordinal([1 2 3 2 1],{'lo','med','hi'})
A =
    lo med hi med lo
A = mergelevels(A,{'lo','med'},'bad')
A =
    bad bad hi bad bad
```


## TreeBagger.Method property

Purpose Method used by trees (classification or regression)
Description The Method property is 'classification' for classification ensembles and 'regression' for regression ensembles.

Purpose
Metropolis-Hastings sample
Syntax

```
smpl = mhsample(start,nsamples,'pdf',pdf,'proppdf',proppdf,
    'proprnd',proprnd)
smpl = mhsample(...,'symmetric',sym)
smpl = mhsample(...,'burnin',K)
smpl = mhsample(...,'thin',m)
smpl = mhsample(...,'nchain',n)
[smpl,accept] = mhsample(...)
```


## Description

smpl = mhsample(start, nsamples,'pdf',pdf,'proppdf',proppdf, 'proprnd', proprnd) draws nsamples random samples from a target stationary distribution pdf using the Metropolis-Hastings algorithm.
start is a row vector containing the start value of the Markov Chain, nsamples is an integer specifying the number of samples to be generated, and pdf, proppdf, and proprnd are function handles created using @. proppdf defines the proposal distribution density, and proprnd defines the random number generator for the proposal distribution. pdf and proprnd take one argument as an input with the same type and size as start. proppdf takes two arguments as inputs with the same type and size as start.
smpl is a column vector or matrix containing the samples. If the log density function is preferred, 'pdf' and 'proppdf' can be replaced with 'logpdf' and 'logproppdf'. The density functions used in Metropolis-Hastings algorithm are not necessarily normalized.
The proposal distribution $q(x, y)$ gives the probability density for choosing $x$ as the next point when $y$ is the current point. It is sometimes written as $q(x \mid y)$.

If the proppdf or logproppdf satisfies $q(x, y)=q(y, x)$, that is, the proposal distribution is symmetric, mhsample implements Random Walk Metropolis-Hastings sampling. If the proppdf or logproppdf satisfies $q(x, y)=q(x)$, that is, the proposal distribution is independent of current values, mhsample implements Independent Metropolis-Hastings sampling.
smpl = mhsample(...,'symmetric',sym) draws nsamples random samples from a target stationary distribution pdf using the Metropolis-Hastings algorithm. sym is a logical value that indicates whether the proposal distribution is symmetric. The default value is false, which corresponds to the asymmetric proposal distribution. If sym is true, for example, the proposal distribution is symmetric, proppdf and logproppdf are optional.
smpl = mhsample(...,'burnin',K) generates a Markov chain with values between the starting point and the $\mathrm{k}^{\text {th }}$ point omitted in the generated sequence. Values beyond the $k^{\text {th }}$ point are kept. $k$ is a nonnegative integer with default value of 0 .
smpl = mhsample(...,'thin',m) generates a Markov chain with $m-1$ out of $m$ values omitted in the generated sequence. $m$ is a positive integer with default value of 1 .
smpl = mhsample(..., 'nchain', n) generates $n$ Markov chains using the Metropolis-Hastings algorithm. $n$ is a positive integer with a default value of 1 . smpl is a matrix containing the samples. The last dimension contains the indices for individual chains.
[smpl, accept] = mhsample(...) also returns accept, the acceptance rate of the proposed distribution. accept is a scalar if a single chain is generated and is a vector if multiple chains are generated.

## Examples

Estimate the second order moment of a Gamma distribution using the Independent Metropolis-Hastings sampling.

```
alpha = 2.43;
beta = 1;
pdf = @(x)gampdf(x,alpha,beta); %target distribution
proppdf = @(x,y)gampdf(x,floor(alpha),floor(alpha)/alpha);
proprnd = @(x)sum(...
    exprnd(floor(alpha)/alpha,floor(alpha),1));
nsamples = 5000;
smpl = mhsample(1,nsamples,'pdf',pdf,'proprnd',proprnd,...
    'proppdf',proppdf);
xxhat = cumsum(smpl.^2)./(1:nsamples)';
```



Generate random samples from $N(0,1)$ using the Random Walk Metropolis-Hastings sampling.

```
delta = .5;
pdf = @(x) normpdf(x);
proppdf = @(x,y) unifpdf(y-x,-delta,delta);
proprnd = @(x) x + rand*2*delta - delta;
nsamples = 15000;
x = mhsample(1,nsamples,'pdf',pdf,'proprnd',proprnd,'symmetric',1);
histfit(x,50)
h = get(gca,'Children');
set(h(2),'FaceColor',[.8 . 8 1])
```



See Also
slicesample | rand

## TreeBagger.MinLeaf property

Purpose Minimum number of observations per tree leaf
$\begin{array}{ll}\text { Description } & \text { The MinLeaf property specifies the minimum number of observations } \\ \text { per tree leaf. The default values are } 1 \text { for classification and } 5 \text { for } \\ \text { regression. For classregtree training, the 'minparent' value is set to } \\ & 2 * \text { MinLeaf. }\end{array}$
See Also
classregtree

## Purpose Maximum likelihood estimates

Syntax $\quad$ phat $=$ mle (data $)$
[phat,pci] = mle(data)
[...] = mle(data,'distribution',dist)
[...] = mle(data,...,name1,val1,name2,val2,...)
[...] = mle(data,'pdf',pdf,'cdf',cdf,'start',start,...)
[...] = mle(data,'logpdf',logpdf,'logsf',logsf,'start',start, ...)
[...] = mle(data,'nloglf',nloglf,'start',start,...)
Description
phat $=$ mle(data) returns maximum likelihood estimates (MLEs) for the parameters of a normal distribution, computed using the sample data in the vector data.
[phat, pci] = mle(data) returns MLEs and $95 \%$ confidence intervals for the parameters.
[...] = mle(data,'distribution',dist) computes parameter estimates for the distribution specified by dist. Acceptable strings for dist, and the returned parameters, are:

| name | Distribution | Parameter 1 | Parameter 2 | Parameter <br> $\mathbf{3}$ |
| :--- | :--- | :--- | :--- | :--- |
| 'bernoulli' | "Bernoulli <br> Distribution" on <br> page B-3 | p: probability of <br> success for each <br> trial | - | - |
| 'beta' or 'Beta' | "Beta <br> Distribution" on <br> page B-4 | a | b | - |
| 'bino' or <br> 'Binomial' | "Binomial <br> Distribution" on <br> page B-7 | n: number of trials | p: probability of <br> success for each <br> trial | - |


| name | Distribution | Parameter 1 | Parameter 2 | Parameter $3$ |
| :---: | :---: | :---: | :---: | :---: |
| 'birnbaumsaunder | s"Birnbaum-Saunder $\beta$ Distribution" on page B-10 |  | $\gamma$ | - |
| 'Discrete Uniform' or 'unid' | "Uniform Distribution (Discrete)" on page B-101 | N : maximum observable value | - | - |
| 'exp' or 'Exponential | "Exponential Distribution" on page B-16 | $\mu$ : mean | - | - |
| 'ev' or 'Extreme Value' | "Extreme Value Distribution" on page B-19 | $\mu$ : location parameter | $\sigma$ : scale parameter | - |
| 'gam' or 'Gamma' | "Gamma Distribution" on page B-27 | a: shape parameter | b: scale parameter | - |
| 'gev' or 'Generalized Extreme Value | "Generalized Extreme Value Distribution" on page B-32 | k: shape parameter | $\sigma$ : scale parameter | $\mu$ : location parameter |
| 'gp' or 'Generalized Pareto' | "Generalized <br> Pareto <br> Distribution" on page B-37 | k: tail index (shape) parameter | $\sigma$ : scale parameter | - |
| $\begin{aligned} & \hline \text { 'geo' or } \\ & \text { 'Geometric' } \end{aligned}$ | "Geometric Distribution" on page B-41 | p : probability parameter | - | - |


| name | Distribution | Parameter 1 | Parameter 2 | Parameter 3 |
| :---: | :---: | :---: | :---: | :---: |
| 'inversegaussian | '"Inverse <br> Gaussian Distribution" on page B-45 | $\mu$ | $\lambda$ | - |
| 'logistic' | "Logistic Distribution" on page B-49 | $\mu$ | $\sigma$ | - |
| 'loglogistic' | "Loglogistic Distribution" on page B-50 | $\mu$ | $\sigma$ | - |
| 'logn' or 'Lognormal' | "Lognormal Distribution" on page B-51 | $\mu$ | $\sigma$ | - |
| 'nakagami' | "Nakagami Distribution" on page B-70 | $\mu$ | $\omega$ | - |
| 'nbin' or 'Negative Binomial' | "Negative <br> Binomial <br> Distribution" on page B-72 | $r$ : number of successes | p: probability of success in a single trial | - |
| 'norm' or 'Normal' | "Normal Distribution" on page B-83 | $\mu$ : mean | $\sigma$ : standard deviation | - |
| 'poiss' or <br> 'Poisson' | "Poisson Distribution" on page B-89 | $\lambda:$ mean | - | - |
| 'rayl' or 'Rayleigh' | "Rayleigh Distribution" on page B-91 | b: scale parameter | - | - |

$\left.\begin{array}{l|l|l|l|l}\hline \text { name } & \text { Distribution } & \text { Parameter 1 } & \text { Parameter 2 } & \begin{array}{l}\text { Parameter } \\ \mathbf{3}\end{array} \\ \hline \text { 'rician' } & \begin{array}{l}\text { "Rician } \\ \text { Distribution" on } \\ \text { page B-93 }\end{array} & \begin{array}{l}\mathrm{s}: \text { noncentrality } \\ \text { parameter }\end{array} & \begin{array}{l}\sigma: \text { scale } \\ \text { parameter }\end{array} & - \\ \hline \text { 'tlocationscale } & \begin{array}{l}\text { "t Location-Scale } \\ \text { Distribution" on } \\ \text { page B-97 }\end{array} & \begin{array}{l}\mu: \text { location } \\ \text { parameter }\end{array} & \begin{array}{l}\text { }\end{array} \\ \hline \text { parameter scale }\end{array} \quad \begin{array}{l}v: \text { shape } \\ \text { parameter }\end{array}\right]$
[...] = mle(data,...,name1,val1,name2,val2,...) specifies optional argument name/value pairs chosen from the following list.

| Name | Value |
| :--- | :--- |
| 'censoring' | A Boolean vector of the same size as data, <br> containing ones when the corresponding <br> elements of data are right-censored <br> observations and zeros when the corresponding <br> elements are exact observations. The default <br> is that all observations are observed exactly. <br> Censoring is not supported for all distributions. |
| 'frequency ' | A vector of the same size as data, containing <br> nonnegative integer frequencies for the <br> corresponding elements in data. The default is <br> one observation per element of data. |


| Name | Value |
| :--- | :--- |
| 'alpha' | A value between 0 and 1 specifying a confidence <br> level of 100(1-alpha)\% for pci. The default is <br> 0.05 for $95 \%$ confidence. |
| 'ntrials ' | A scalar, or a vector of the same size as data, <br> containing the total number of trials for the <br> corresponding element of data. Applies only to <br> the binomial distribution. |
| 'options ' | A structure created by a call to statset, <br> containing numerical options for the fitting <br> algorithm. Not applicable to all distributions. |

mle can also fit custom distributions that you define using distribution functions, in one of three ways.
[...] = mle(data,'pdf',pdf,'cdf',cdf,'start',start,....) returns MLEs for the parameters of the distribution defined by the probability density and cumulative distribution functions pdf and cdf. pdf and cdf are function handles created using the @ sign. They accept as inputs a vector data and one or more individual distribution parameters, and return vectors of probability density values and cumulative probability values, respectively. If the 'censoring' name/value pair is not present, you can omit the 'cdf' name/value pair. mle computes the estimates by numerically maximizing the distribution's log-likelihood, and start is a vector containing initial values for the parameters.
[...] = mle(data,'logpdf',logpdf,'logsf',logsf,'start',start,...) returns MLEs for the parameters of the distribution defined by the log probability density and log survival functions logpdf and logsf. logpdf and logsf are function handles created using the @ sign. They accept as inputs a vector data and one or more individual distribution parameters, and return vectors of logged probability density values and logged survival function values, respectively. This form is sometimes more robust to the choice of starting point than using pdf and cdf
functions. If the 'censoring' name/value pair is not present, you can omit the 'logsf' name/value pair. start is a vector containing initial values for the distribution's parameters.
[...] = mle(data,'nloglf',nloglf,'start',start,...) returns MLEs for the parameters of the distribution whose negative log-likelihood is given by nloglf. nloglf is a function handle, specified using the @ sign, that accepts the four input arguments:

- params - a vector of distribution parameter values
- data - a vector of data
- cens - a Boolean vector of censoring values
- freq-a vector of integer data frequencies
nloglf must accept all four arguments even if you do not supply the 'censoring' or 'frequency' name/value pairs (see above). However, nloglf can safely ignore its cens and freq arguments in that case. nloglf returns a scalar negative log-likelihood value and, optionally, a negative log-likelihood gradient vector (see the 'GradObj' statset parameter below). start is a vector containing initial values for the distribution's parameters.
pdf, cdf, logpdf, logsf, or nloglf can also be cell arrays whose first element is a function handle as defined above, and whose remaining elements are additional arguments to the function. mle places these arguments at the end of the argument list in the function call.

The following optional argument name/value pairs are valid only when 'pdf' and 'cdf', 'logpdf' and 'logcdf', or 'nloglf' are given:

- 'lowerbound ' - A vector the same size as start containing lower bounds for the distribution parameters. The default is - Inf.
- 'upperbound ' - A vector the same size as start containing upper bounds for the distribution parameters. The default is Inf.
- 'optimfun' - A string, either 'fminsearch' or 'fmincon', naming the optimization function to be used in maximizing the likelihood.

The default is 'fminsearch'. You can only specify 'fmincon' if Optimization Toolbox software is available.

When fitting a custom distribution, use the 'options' parameter to control details of the maximum likelihood optimization. See statset('mlecustom') for parameter names and default values. mle interprets the following statset parameters for custom distribution fitting as follows:

| Parameter | Value |
| :--- | :--- |
| 'GradObj' | 'on' or ' off ', indicating whether or not fmincon <br> can expect the function provided with the ' nloglf' <br> name/value pair to return the gradient vector of <br> the negative log-likelihood as a second output. The <br> default is 'off'. Ignored when using fminsearch. |
| 'DerivStep' | The relative difference used in finite difference <br> derivative approximations when using fmincon, and <br> 'GradObj ' is ' off '. 'DerivStep' can be a scalar, or <br> the same size as 'start'. The default is eps^(1/3). <br> Ignored when using fminsearch. |
| 'FunValCheck ' | 'on' or 'off', indicating whether or not mle should <br> check the values returned by the custom distribution <br> functions for validity. The default is 'on '. A poor <br> choice of starting point can sometimes cause these <br> functions to return NaNs, infinite values, or out of <br> range values if they are written without suitable <br> error-checking. |
| 'TolBnd' | An offset for upper and lower bounds when using <br> fmincon. mle treats upper and lower bounds as <br> strict inequalities (i.e., open bounds). With fmincon, <br> this is approximated by creating closed bounds <br> inset from the specified upper and lower bounds by <br> TolBnd. The default is 1e-6. |

Examples The following returns an MLE and a $95 \%$ confidence interval for the success probability of a binomial distribution with 20 trials:

```
data = binornd(20,0.75,100,1); % Simulated data, p = 0.75
[phat,pci] = mle(data,'distribution','binomial',...
    'alpha',.05,'ntrials',20)
phat =
    0.7370
pci =
    0.7171
    0.7562
```

See Also
betafit | binofit | evfit | expfit | gamfit | gevfit | gpfit | lognfit | nbinfit | normfit | mlecov | poissfit | raylfit | statset | unifit | wblfit
Purpose Asymptotic covariance of maximum likelihood estimators
Syntax

ACOV = mlecov(params,data,...)
ACOV = mlecov(params,data,'pdf',pdf,'cdf',cdf)
ACOV = mlecov(params,data,'logpdf',logpdf,'logsf',logsf)
ACOV = mlecov(params,data,'nloglf', nloglf)
[...] = mlecov(params,data,...,param1,val1,param2,val2,...)
ACOV = mlecov(params,data,...) returns an approximation to the asymptotic covariance matrix of the maximum likelihood estimators of the parameters for a specified distribution. The following paragraphs describe how to specify the distribution. mlecov computes a finite difference approximation to the Hessian of the log-likelihood at the maximum likelihood estimates params, given the observed data, and returns the negative inverse of that Hessian. ACOV is a $p$-by- $p$ matrix, where $p$ is the number of elements in params.
You must specify a distribution after the input argument data, as follows.
ACOV = mlecov(params,data,'pdf',pdf,'cdf',cdf) enables you to define a distribution by its probability density and cumulative distribution functions, pdf and cdf, respectively. pdf and cdf are function handles that you create using the @ sign. They accept a vector of data and one or more individual distribution parameters as inputs and return vectors of probability density function values and cumulative distribution values, respectively. If the 'censoring' name/value pair (see below) is not present, you can omit the 'cdf' name/value pair.
ACOV = mlecov(params,data,'logpdf',logpdf,'logsf',logsf) enables you to define a distribution by its $\log$ probability density and log survival functions, logpdf and logsf, respectively. logpdf and logsf are function handles that you create using the @ sign. They accept as inputs a vector of data and one or more individual distribution parameters, and return vectors of logged probability density values and logged survival function values, respectively. If the 'censoring' name/value pair (see below) is not present, you can omit the 'logsf' name/value pair.

ACOV = mlecov(params,data,'nloglf',nloglf) enables you to define a distribution by its log-likelihood function. nloglf is a function handle, specified using the @ sign, that accepts the following four input arguments:

- params - Vector of distribution parameter values
- data - Vector of data
- cens - Boolean vector of censoring values
- freq - Vector of integer data frequencies
nloglf must accept all four arguments even if you do not supply the 'censoring' or 'frequency' name/value pairs (see below). However, nloglf can safely ignore its cens and freq arguments in that case. nloglf returns a scalar negative log-likelihood value and, optionally, the negative log-likelihood gradient vector (see the 'gradient' name/value pair below).
pdf, cdf, logpdf, logsf, and nloglf can also be cell arrays whose first element is a function handle, as defined above, and whose remaining elements are additional arguments to the function. The mle function places these arguments at the end of the argument list in the function call.
[...] =
mlecov(params,data,...,param1,val1,param2,val2,...) specifies optional parameter name/value pairs chosen from the following table.

| Parameter | Value |
| :--- | :--- |
| 'censoring ' | Boolean vector of the same size as data, containing <br> 1's when the corresponding elements of data are <br> right-censored observations and 0's when the <br> corresponding elements are exact observations. The <br> default is that all observations are observed exactly. <br> Censoring is not supported for all distributions. |
| 'frequency' | A vector of the same size as data containing <br> nonnegative frequencies for the corresponding <br> elements in data. The default is one observation per <br> element of data. |
| 'options' | A structure opts containing numerical options for <br> the finite difference Hessian calculation. You create <br> opts by calling statset. The applicable statset <br> parameters are: |
| - 'GradObj ' - ' on' or ' off ', indicating whether <br> or not the function provided with the 'nloglf ' <br> name/value pair can return the gradient vector of <br> the negative log-likelihood as its second output. The <br> default is 'off'. |  |
| - 'DerivStep ' - Relative step size used in finite |  |
| difference for Hessian calculations. Can be a |  |
| scalar, or the same size as params. The default is |  |
| eps^(1/4). A smaller value might be appropriate if |  |
| 'GradObj ' is 'on'. |  |

Examples
Create the following function:

```
function logpdf = betalogpdf(x,a,b)
logpdf = (a-1)*log(x)+(b-1)*log(1-x)-betaln(a,b);
```

Fit a beta distribution to some simulated data, and compute the approximate covariance matrix of the parameter estimates:

```
x = betarnd(1.23,3.45,25,1);
phat = mle(x,'dist','beta')
acov = mlecov(phat,x,'logpdf',@betalogpdf)
```

See Also mle

## Purpose

Multinomial probability density function

## Syntax

Description
$Y=\operatorname{mnpdf}(X, \operatorname{PROB})$ returns the pdf for the multinomial distribution with probabilities PROB, evaluated at each row of $X . X$ and PROB are $m$-by- $k$ matrices or 1 -by- $k$ vectors, where $k$ is the number of multinomial bins or categories. Each row of PROB must sum to one, and the sample sizes for each observation (rows of $X$ ) are given by the row sums sum ( $X, 2$ ). $Y$ is an $m$-by- $k$ matrix, and mnpdf computes each row of $Y$ using the corresponding rows of the inputs, or replicates them if needed.

## Examples

```
% Compute the distribution
p = [1/2 1/3 1/6]; % Outcome probabilities
n = 10; % Sample size
x1 = 0:n;
x2 = 0:n;
[X1,X2] = meshgrid(x1,x2);
X3 = n-(X1+X2);
Y = mnpdf([X1(:),X2(:),X3(:)],repmat(p,(n+1)^2,1));
% Plot the distribution
Y = reshape(Y, n+1,n+1);
bar3(Y)
set(gca,'XTickLabel',0:n)
set(gca,'YTickLabel',0:n)
xlabel('x_1')
ylabel('x_2')
zlabel('Probability Mass')
title('Trinomial Distribution')
```



Note that the visualization does not show $x_{3}$, which is determined by the constraint $x_{1}+x_{2}+x_{3}=n$.

See Also
mnrnd

- "Multinomial Distribution" on page B-54


## Purpose <br> Syntax <br> Description

Multinomial logistic regression
$B=\operatorname{mnrfit}(X, Y)$
B = mnrfit(X,Y,param1,val1,param2,val2,...)
[B,dev] = mnrfit(...)
[B,dev,stats] = mnrfit(...)
$B=\operatorname{mnrfit}(X, Y)$ returns a matrix $B$ of coefficient estimates for a multinomial logistic regression of the responses in $Y$ on the predictors in X . X is an $n$-by- $p$ matrix of $p$ predictors at each of $n$ observations. $Y$ is an $n$-by- $k$ matrix, where $Y(i, j)$ is the number of outcomes of the multinomial category j for the predictor combinations given by $X(i,:)$. The sample sizes for each observation are given by the row sums sum $(Y, 2)$.

Alternatively, Y can be an $n$-by- 1 column vector of scalar integers from 1 to $k$ indicating the value of the response for each observation, and all sample sizes are taken to be 1 .

The result B is a $(p+1)$-by- $(k-1)$ matrix of estimates, where each column corresponds to the estimated intercept term and predictor coefficients, one for each of the first $k-1$ multinomial categories. The estimates for the $k^{\text {th }}$ category are taken to be zero.

Note mnrfit automatically includes a constant term in all models. Do not enter a column of 1 s directly into X .
mnrfit treats NaNs in either X or Y as missing values, and ignores them. B = mnrfit(X,Y,param1,val1,param2,val2,...) allows you to specify optional parameter name/value pairs to control the model fit. Parameters are:

- 'model' - The type of model to fit; one of the text strings 'nominal' (the default), 'ordinal', or 'hierarchical'
- 'interactions' - Determines whether the model includes an interaction between the multinomial categories and the coefficients. Specify as 'off' to fit a model with a common set of coefficients for the predictor variables, across all multinomial categories. This is often described as parallel regression. Specify as 'on' to fit a model with different coefficients across categories. In all cases, the model has different intercepts across categories. Thus, B is a vector containing $k-1+p$ coefficient estimates when 'interaction' is 'off', and a ( $p+1$ )-by- $(k-1)$ matrix when it is 'on'. The default is 'off' for ordinal models, and 'on' for nominal and hierarchical models.
- 'link' - The link function to use for ordinal and hierarchical models. The link function defines the relationship $g\left(\mu_{i j}\right)=x_{i} b_{j}$ between the mean response for the $i^{\text {th }}$ observation in the $j^{\text {th }}$ category, $\mu_{i j}$, and the linear combination of predictors $x_{i} b_{j}$. Specify the link parameter value as one of the text strings 'logit'(the default), 'probit', 'comploglog', or 'loglog'. You may not specify the 'link' parameter for nominal models; these always use a multivariate logistic link.
- 'estdisp' - Specify as 'on ' to estimate a dispersion parameter for the multinomial distribution in computing standard errors, or 'off' (the default) to use the theoretical dispersion value of 1.
[B, dev] = mnrfit(...) returns the deviance of the fit dev.
[B,dev,stats] = mnrfit(...) returns a structure stats that contains the following fields:
- dfe - Degrees of freedom for error
- s-Theoretical or estimated dispersion parameter
- sfit - Estimated dispersion parameter
- se - Standard errors of coefficient estimates B
- coeffcorr - Estimated correlation matrix for B
- covb - Estimated covariance matrix for B
- $\mathrm{t}-t$ statistics for B
- $\mathrm{p}-p$-values for B
- resid - Residuals
- residp - Pearson residuals
- residd - Deviance residuals

Fit multinomial logistic regression models to data with one predictor variable and three categories in the response variable:

```
x = [-3 -2 -1 0-1 0 1 2 3]';
Y = [1 11 13; 2 9 14; 6 14 5; 5 10 10; 5 14 6; 7 13 5;...
    8 11 6];
bar(x,Y,'stacked'); ylim([0 25]);
```


\% response categories:
$x x=\operatorname{linspace}(-4,4)^{\prime}$;
pHatNom = mnrval(betaHatNom, xx,'model','nominal', ...
'interactions', 'on');
line (xx, cumsum(25*pHatNom, 2), 'LineWidth' , 2);


Fit a "parallel" ordinal model for the cumulative response category probabilities, with a common slope on the single predictor variable, $x$, across all categories:

```
% The first two elements of betaHatOrd are the
% intercept terms for the first two response categories.
% The last element of betaHatOrd is the common slope.
betaHatOrd = mnrfit(x,Y,'model','ordinal',...
    'interactions','off')
% Compute the predicted cumulative probabilities for the
% first two response categories. The cumulative
% probability for the third category is always 1.
pHatOrd = mnrval(betaHatOrd,xx,'type','cumulative',...
'model','ordinal','interactions','off');
bar(x,cumsum(Y,2),'grouped'); ylim([0 25]);
line(xx,25*pHatOrd,'LineWidth',2);
```



## References

[1] McCullagh, P., and J. A. Nelder. Generalized Linear Models. New York: Chapman \& Hall, 1990.

See Also mnrval | glmfit | glmval | regress | regstats
How To . "Multinomial Distribution" on page B-54

## Purpose Multinomial random numbers

```
Syntax r m mnrnd(n,p)
R = mnrnd(n,p,m)
R = mnrnd(N,P)
```


## Description

## Examples

$r=m n r n d(n, p)$ returns random values $r$ from the multinomial distribution with parameters n and p . n is a positive integer specifying the number of trials (sample size) for each multinomial outcome. $p$ is a 1 -by- $k$ vector of multinomial probabilities, where $k$ is the number of multinomial bins or categories. p must sum to one. (If p does not sum to one, $r$ consists entirely of NaN values.) $r$ is a $1-$ by- $k$ vector, containing counts for each of the $k$ multinomial bins.
$R=m n r n d(n, p, m)$ returns $m$ random vectors from the multinomial distribution with parameters $n$ and $p$. R is a $m$-by- $k$ matrix, where $k$ is the number of multinomial bins or categories. Each row of R corresponds to one multinomial outcome.
$R=m n r n d(N, P)$ generates outcomes from different multinomial distributions. P is a $m$-by- $k$ matrix, where $k$ is the number of multinomial bins or categories and each of the $m$ rows contains a different set of multinomial probabilities. Each row of P must sum to one. (If any row of $P$ does not sum to one, the corresponding row of $R$ consists entirely of NaN values.) N is a $m$-by- 1 vector of positive integers or a single positive integer (replicated by mnrnd to a $m$-by- 1 vector). R is a m-by- $k$ matrix. Each row of $R$ is generated using the corresponding rows of N and P .

Generate 2 random vectors with the same probabilities:

```
n = 1e3;
p = [0.2,0.3,0.5];
R = mnrnd(n,p,2)
R =
    215 282 503
    194 303 503
```

Generate 2 random vectors with different probabilities:

```
n = 1e3;
P = [0.2, 0.3, 0.5; ...
    0.3, 0.4, 0.3;];
R = mnrnd(n,P)
R =
    186 290 524
    290 389 321
```

See Also mnpdf

How To . "Multinomial Distribution" on page B-54

## Purpose Multinomial logistic regression values

```
Syntax \(\quad\) PHAT \(=\operatorname{mnrval}(B, X)\)
YHAT = mnrval(B,X,ssize)
[...,DLO,DHI] = mnrval(B,X,...,stats)
[...] = mnrval(...,param1,val1,param2,val2,...)
```


## Description

PHAT $=m n r v a l(B, X)$ computes predicted probabilities for the multinomial logistic regression model with predictors X. B contains intercept and coefficient estimates as returned by the mnrfit function. X is an $n$-by- $p$ matrix of $p$ predictors at each of $n$ observations. PHAT is an $n$-by- $k$ matrix of predicted probabilities for each multinomial category.

Note mnrval automatically includes a constant term in all models. Do not enter a column of 1 s directly into X .

YHAT $=$ mnrval( $B, X$, ssize) computes predicted category counts for sample sizes ssize. ssize is an $n$-by- 1 column vector of positive integers.
[..., DLO, DHI] = mnrval(B, X,...,stats) also computes $95 \%$ confidence bounds on the predicted probabilities PHAT or counts YHAT. stats is the structure returned by the mnrfit function. DLO and DHI define a lower confidence bound of PHAT or YHAT minus DLO and an upper confidence bound of PHAT or YHAT plus DHI. Confidence bounds are nonsimultaneous and they apply to the fitted curve, not to new observations.
[...] = mnrval(..., param1, val1, param2, val2,...) allows you to specify optional parameter name/value pairs to control the predicted values. These parameters must be set to the corresponding values used with the mnrfit function to compute B. Parameters are:

- 'model' - The type of model that was fit by mnrfit; one of the text strings 'nominal' (the default), 'ordinal', or 'hierarchical'.
- 'interactions' - Determines whether the model fit by mnrfit included an interaction between the multinomial categories and the coefficients. The default is 'off' for ordinal models, and 'on' for nominal and hierarchical models.
- 'link' - The link function that was used by mnrfit for ordinal and hierarchical models. Specify the link parameter value as one of the text strings 'logit'(the default), 'probit', 'comploglog', or 'loglog'. You may not specify the 'link' parameter for nominal models; these always use a multivariate logistic link.
- 'type' - Set to 'category' (the default) to return predictions and confidence bounds for the probabilities (or counts) of the $k$ multinomial categories. Set to 'cumulative' to return predictions and confidence bounds for the cumulative probabilities (or counts) of the first $k-1$ multinomial categories, as an $n$-by- $(k-1)$ matrix. The predicted cumulative probability for the $k$ th category is 1 . Set to 'conditional' to return predictions and confidence bounds in terms of the first $k-1$ conditional category probabilities, i.e., the probability for category $j$, given an outcome in category $j$ or higher. When 'type' is 'conditional', and you supply the sample size argument ssize, the predicted counts at each row of $X$ are conditioned on the corresponding element of ssize, across all categories.
- 'confidence' - The confidence level for the confidence bounds; a value between 0 and 1 . The default is 0.95 .


## Examples

Fit multinomial logistic regression models to data with one predictor variable and three categories in the response variable:

```
x = [\begin{array}{llllllll}{-3}&{-2}&{-1}&{0}&{1}&{2}&{3}\end{array}]}\mp@subsup{]}{}{\prime}
Y = [1 11 13; 2 9 14; 6 14 5; 5 10 10; 5 14 6; 7 13 5;...
    8 11 6];
bar(x,Y,'stacked');
ylim([0 25]);
```


\% Now fit a nominal model for the individual response \% category probabilities, with separate slopes on the
\% single predictor variable, x, for each
\% category:
\% The first row of betaHatNom contains the intercept terms
\% for the first two response categories. The second row \% contains the slopes.
betaHatNom = mnrfit(x,y,'model','nominal',...
'interactions', 'on')
\% Compute the predicted probabilities for the three
\% response categories:

```
xx = linspace(-4,4)';
pHatNom = mnrval(betaHatNom,xx,'model','nominal',...
    'interactions','on');
line(xx,cumsum(25*pHatNom,2),'LineWidth' ,2);
```



Fit a "parallel" ordinal model for the cumulative response category probabilities, with a common slope on the single predictor variable, $x$, across all categories:

```
% The first two elements of betaHatOrd are the
% intercept terms for the first two response categories.
% The last element of betaHatOrd is the common slope.
betaHatOrd = mnrfit(x,Y,'model','ordinal',...
    'interactions','off')
% Compute the predicted cumulative probabilities for the
% first two response categories. The cumulative
% probability for the third category is always 1.
pHatOrd = mnrval(betaHatOrd,xx,'type','cumulative',...
    'model','ordinal','interactions','off');
bar(x,cumsum(Y,2),'grouped');
ylim([0 25]);
line(xx,25*pHatOrd,'LineWidth',2);
```



## References

[1] McCullagh, P., and J. A. Nelder. Generalized Linear Models. New York: Chapman \& Hall, 1990.

See Also mnrfit | glmfit | glmval<br>How To . "Multinomial Distribution" on page B-54

## Purpose Central moments

## Syntax <br> m = moment(X,order) <br> moment (X,order,dim)

## Description

$m=$ moment ( $X$, order) returns the central sample moment of $X$ specified by the positive integer order. For vectors, moment ( $x$, order) returns the central moment of the specified order for the elements of $x$. For matrices, moment (X, order) returns central moment of the specified order for each column. For N-dimensional arrays, moment operates along the first nonsingleton dimension of X .
moment ( X , order, dim) takes the moment along dimension dim of X .

Tips

## Examples

Note that the central first moment is zero, and the second central moment is the variance computed using a divisor of $n$ rather than $n-$ 1 , where $n$ is the length of the vector x or the number of rows in the matrix $X$.

The central moment of order $k$ of a distribution is defined as

$$
m_{k}=E(x-\mu)^{k}
$$

where $E(x)$ is the expected value of $x$.

```
X = randn([6 5])
X =
        1.1650 0.0591 1.2460-1.2704 -0.0562
        0.6268 1.7971 -0.6390
        0.0751 0.2641 0.5774 -0.0449 0.3967
        0.3516 0.8717 -0.3600 -0.7989 0.7562
        -0.6965 -1.4462 -0.1356-0.7652 0.4005
        1.6961-0.7012 -1.3493 0.8617 -1.3414
m = moment(X,3)
m =
    -0.0282 0.0571 0.1253 0.1460 -0.4486
```

See Also
kurtosis | mean | skewness | std | var

## gmdistribution.Mu property

Purpose Input matrix of means MU
Description Input matrix of means mu.
Purpose Multiple comparison test
Syntax $\quad c=$ multcompare (stats)
c = multcompare(stats,param1,val1,param2,val2,...)
[c,m] = multcompare(...)
[c,m,h] = multcompare(...)
[c,m,h,gnames] = multcompare(...)

## Description

c = multcompare(stats) performs a multiple comparison test using the information in the stats structure, and returns a matrix c of pairwise comparison results. It also displays an interactive graph of the estimates with comparison intervals around them. See "Examples" on page 20-1084.
In a one-way analysis of variance, you compare the means of several groups to test the hypothesis that they are all the same, against the general alternative that they are not all the same. Sometimes this alternative may be too general. You may need information about which pairs of means are significantly different, and which are not. A test that can provide such information is called a multiple comparison procedure.
When you perform a simple t-test of one group mean against another, you specify a significance level that determines the cutoff value of the t statistic. For example, you can specify the value alpha $=0.05$ to insure that when there is no real difference, you will incorrectly find a significant difference no more than $5 \%$ of the time. When there are many group means, there are also many pairs to compare. If you applied an ordinary t-test in this situation, the alpha value would apply to each comparison, so the chance of incorrectly finding a significant difference would increase with the number of comparisons. Multiple comparison procedures are designed to provide an upper bound on the probability that any comparison will be incorrectly found significant.
The output c contains the results of the test in the form of a five-column matrix. Each row of the matrix represents one test, and there is one row for each pair of groups. The entries in the row indicate the means being compared, the estimated difference in means, and a confidence interval for the difference.

For example, suppose one row contains the following entries.

$$
\begin{array}{lllll}
2.0000 & 5.0000 & 1.9442 & 8.2206 & 14.4971
\end{array}
$$

These numbers indicate that the mean of group 2 minus the mean of group 5 is estimated to be 8.2206 , and a $95 \%$ confidence interval for the true difference of the means is [1.9442, 14.4971].

In this example the confidence interval does not contain 0.0 , so the difference is significant at the 0.05 level. If the confidence interval did contain 0.0 , the difference would not be significant at the 0.05 level.

The multcompare function also displays a graph with each group mean represented by a symbol and an interval around the symbol. Two means are significantly different if their intervals are disjoint, and are not significantly different if their intervals overlap. You can use the mouse to select any group, and the graph will highlight any other groups that are significantly different from it.
c = multcompare(stats,param1,val1,param2,val2,...) specifies one or more of the parameter name/value pairs described in the following table.

| Parameter | Values |
| :--- | :--- |
| ' alpha' | Scalar between 0 and 1 that determines the <br> confidence levels of the intervals in the matrix <br> c and in the figure (default is 0.05). The <br> confidence level is 100(1-alpha)\%. |
| 'display' | Either 'on ' (the default) to display a graph <br> of the estimates with comparison intervals <br> around them, or 'off ' to omit the graph. See <br> "Examples" on page 20-1084. |
| 'ctype' | Specifies the type of critical value to use for the <br> multiple comparison. "Values of ctype" on page <br> 20-1082 describes the allowed values for ctype. |


| Parameter | Values |
| :--- | :--- |
| 'dimension' | A vector specifying the dimension or dimensions <br> over which the population marginal means <br> are to be calculated. Use only if you create <br> stats with the function anovan. The default <br> is 1 to compute over the first dimension. See <br> "Dimension Parameter" on page 20-1084 for <br> more information. |
| 'estimate' | Specifies the estimate to be compared. The <br> allowable values of estimate depend on the <br> function that was the source of the stats <br> structure, as described in "Values of estimate" <br> on page 20-1083 |

[c,m] = multcompare(...) returns an additional matrix m. The first column of $m$ contains the estimated values of the means (or whatever statistics are being compared) for each group, and the second column contains their standard errors.
[ $\mathrm{c}, \mathrm{m}, \mathrm{h}]=$ multcompare (...) returns a handle h to the comparison graph. Note that the title of this graph contains instructions for interacting with the graph, and the $x$-axis label contains information about which means are significantly different from the selected mean. If you plan to use this graph for presentation, you may want to omit the title and the $x$-axis label. You can remove them using interactive features of the graph window, or you can use the following commands.

```
title('')
xlabel('')
```

[ $\mathrm{c}, \mathrm{m}, \mathrm{h}$, gnames ] = multcompare(...) returns gnames, a cell array with one row for each group, containing the names of the groups.

The intervals shown in the graph are computed so that to a very close approximation, two estimates being compared are significantly different if their intervals are disjoint, and are not significantly different if their intervals overlap. (This is exact for multiple comparison of means from
anova1, if all means are based on the same sample size.) You can click on any estimate to see which means are significantly different from it.

## Values of ctype

The following table describes the allowed values for the parameter ctype.

| Value | Description |
| :--- | :--- |
| 'hsd ' or <br> 'tukey-kramer' | Use Tukey's honestly significant difference <br> criterion. This is the default, and it is based on the <br> Studentized range distribution. It is optimal for <br> balanced one-way ANOVA and similar procedures <br> with equal sample sizes. It has been proven <br> to be conservative for one-way ANOVA with <br> different sample sizes. According to the unproven <br> Tukey-Kramer conjecture, it is also accurate for <br> problems where the quantities being compared <br> are correlated, as in analysis of covariance with <br> unbalanced covariate values. |
| 'lsd' | Use Tukey's least significant difference procedure. <br> This procedure is a simple t-test. It is reasonable <br> if the preliminary test (say, the one-way ANOVA <br> $F$ statistic) shows a significant difference. If it is <br> used unconditionally, it provides no protection <br> against multiple comparisons. |
| 'bonferroni' | Use critical values from the t distribution, after a <br> Bonferroni adjustment to compensate for multiple <br> comparisons. This procedure is conservative, but <br> usually less so than the Scheffé procedure. |


| Value | Description |
| :--- | :--- |
| 'dunn-sidak' | Use critical values from the t distribution, after <br> an adjustment for multiple comparisons that was <br> proposed by Dunn and proved accurate by Sidák. <br> This procedure is similar to, but less conservative <br> than, the Bonferroni procedure. |
| 'scheffe' | Use critical values from Scheffés S procedure, <br> derived from the F distribution. This procedure <br> provides a simultaneous confidence level for <br> comparisons of all linear combinations of the <br> means, and it is conservative for comparisons of <br> simple differences of pairs. |

## Values of estimate

The allowable values of the parameter 'estimate' depend on the function that was the source of the stats structure, according to the following table.

| Source | Values |
| :--- | :--- |
| 'anova1' | Ignored. Always compare the group means. |
| 'anova2' | Either 'column' (the default) or 'row ' to <br> compare column or row means. |
| 'anovan' | Ignored. Always compare the population <br> marginal means as specified by the dim <br> argument. |
| 'aoctool' | Either 'slope ', 'intercept', or 'pmm' to <br> compare slopes, intercepts, or population <br> marginal means. If the analysis of covariance <br> model did not include separate slopes, then <br> 'slope' is not allowed. If it did not include <br> separate intercepts, then no comparisons are <br> possible. |


| Source | Values |
| :--- | :--- |
| 'friedman ' | Ignored. Always compare average column ranks. |
| 'kruskalwallis' | Ignored. Always compare average group ranks. |

## Dimension Parameter

The dimension parameter is a vector specifying the dimension or dimensions over which the population marginal means are to be calculated. For example, if dim =1, the estimates that are compared are the means for each value of the first grouping variable, adjusted by removing effects of the other grouping variables as if the design were balanced. If dim = [ $\left.\begin{array}{ll}1 & 3\end{array}\right]$, population marginal means are computed for each combination of the first and third grouping variables, removing effects of the second grouping variable. If you fit a singular model, some cell means may not be estimable and any population marginal means that depend on those cell means will have the value NaN.
Population marginal means are described by Milliken and Johnson (1992) and by Searle, Speed, and Milliken (1980). The idea behind population marginal means is to remove any effect of an unbalanced design by fixing the values of the factors specified by dim, and averaging out the effects of other factors as if each factor combination occurred the same number of times. The definition of population marginal means does not depend on the number of observations at each factor combination. For designed experiments where the number of observations at each factor combination has no meaning, population marginal means can be easier to interpret than simple means ignoring other factors. For surveys and other studies where the number of observations at each combination does have meaning, population marginal means may be harder to interpret.

## Examples

## Example 1

The following example performs a 1 -way analysis of variance (ANOVA) and displays group means with their names.

```
load carsmall
```

```
[p,t,st] = anova1(MPG,Origin,'off');
[c,m,h,nms] = multcompare(st,'display','off');
[nms num2cell(m)]
ans =
\begin{tabular}{lll} 
'USA' & {\([21.1328]\)} & {\([0.8814]\)} \\
'Japan' & {\([31.8000]\)} & {\([1.8206]\)} \\
'Germany' & {\([28.4444]\)} & {\([2.3504]\)} \\
'France' & {\([23.6667]\)} & {\([4.0711]\)} \\
'Sweden' \(^{2}\) & {\([22.5000]\)} & {\([4.9860]\)} \\
'Italy' & {\([28.0000]\)} & {\([7.0513]\)}
\end{tabular}
```

multcompare also displays the following graph of the estimates with comparison intervals around them.


You can click the graphs of each country to compare its mean to those of other countries.

## Example 2

The following continues the example described in the anova1 reference page, which is related to testing the material strength in structural beams. From the anova1 output you found significant evidence that the three types of beams are not equivalent in strength. Now you can determine where those differences lie. First you create the data arrays and you perform one-way ANOVA.

```
strength = [llllllllll
    78 75 76 77 79 79 77 78 82 79];
alloy = {'st','st','st','st','st','st','st','st',...
    'al1','al1','al1','al1','al1','al1',...
    'al2','al2','al2','al2','al2','al2'};
[p,a,s] = anova1(strength,alloy);
```

Among the outputs is a structure that you can use as input to multcompare.

```
[c,m,h,nms] = multcompare(s);
[nms num2cell(c)]
ans =
\begin{tabular}{lllllll} 
'st' & {\([1]\)} & {\([2]\)} & {\([3.6064]\)} & {\([~ 7]\)} & {\([10.3936]\)} \\
'al1' & {\([1]\)} & {\([3]\)} & {\([1.6064]\)} & {\([5]\)} & {\([8.3936]\)} \\
'al2' & {\([2]\)} & {\([3]\)} & {\([-5.6280]\)} & {\([-2]\)} & {\([1.6280]\)}
\end{tabular}
```



The third row of the output matrix shows that the differences in strength between the two alloys is not significant. A $95 \%$ confidence interval for the difference is $[-5.6,1.6]$, so you cannot reject the hypothesis that the true difference is zero.

The first two rows show that both comparisons involving the first group (steel) have confidence intervals that do not include zero. In other words, those differences are significant. The graph shows the same information.

## References

[1] Hochberg, Y., and A. C. Tamhane. Multiple Comparison Procedures. Hoboken, NJ: John Wiley \& Sons, 1987.
[2] Milliken, G. A., and D. E. Johnson. Analysis of Messy Data, Volume 1: Designed Experiments. Boca Raton, FL: Chapman \& Hall/CRC Press, 1992.
[3] Searle, S. R., F. M. Speed, and G. A. Milliken. "Population marginal means in the linear model: an alternative to least-squares means." American Statistician. 1980, pp. 216-221.

See Also anova1| anova2 | anovan | aoctool | friedman | kruskalwallis

Purpose<br>Syntax<br>Description

Multivari chart for grouped data

```
multivarichart(y,GROUP)
multivarichart(Y)
multivarichart(...,param1,val1,param2,val2,...)
[charthandle,AXESH] = multivarichart(...)
```

multivarichart ( $y$, GROUP) displays the multivari chart for the vector y grouped by entries in the cell array GROUP. Each cell of GROUP must contain a grouping variable that can be a categorical variable, numeric vector, character matrix, or single-column cell array of strings. (See "Grouped Data" on page 2-34.) GROUP can also be a matrix whose columns represent different grouping variables. Each grouping variable must have the same number of elements as y . The number of grouping variables must be 2,3 , or 4 .

Each subplot of the plot matrix contains a multivari chart for the first and second grouping variables. The $x$-axis in each subplot indicates values of the first grouping variable. The legend at the bottom of the figure window indicates values of the second grouping variable. The subplot at position $(i, j)$ is the multivari chart for the subset of $y$ at the $i$ th level of the third grouping variable and the $j$ th level of the fourth grouping variable. If the third or fourth grouping variable is absent, it is considered to have only one level.
multivarichart $(\mathrm{Y})$ displays the multivari chart for a matrix Y . The data in different columns represent changes in one factor. The data in different rows represent changes in another factor.
multivarichart(..., param1, val1,param2,val2,...) specifies one or more of the following name/value pairs:

- 'varnames ' - Grouping variable names in a character matrix or a cell array of strings, one per grouping variable. Default names are 'X1', 'X2', ... .
- 'plotorder' - A string with the value 'sorted' or a vector containing a permutation of the integers from 1 to the number of grouping variables.


## multivarichart

If 'plotorder' is a string with value 'sorted', the grouping variables are rearranged in descending order according to the number of levels in each variable.
If 'plotorder' is a vector, it indicates the order in which each grouping variable should be plotted. For example, [2, 3, 1, 4] indicates that the second grouping variable should be used as the $x$-axis of each subplot, the third grouping variable should be used as the legend, the first grouping variable should be used as the columns of the plot, and the fourth grouping variable should be used as the rows of the plot.
[charthandle,AXESH] = multivarichart(...) returns a handle charthandle to the figure window and a matrix AXESH of handles to the subplot axes.

## Examples Display a multivari chart for data with two grouping variables:

```
y = randn(100,1); % response
group = [ceil(3*rand(100,1)) ceil(2*rand(100,1))];
multivarichart(y,group)
```



Display a multivari chart for data with four grouping variables:

```
y = randn(1000,1); % response
group = {ceil(2*rand(1000,1)),ceil(3*rand(1000,1)), ...
    ceil(2*rand(1000,1)),ceil(3*rand(1000,1))};
multivarichart(y,group)
```







$$
\circ \quad \times 2=1 \quad \square \quad \times 2=2 \quad \diamond \quad \times 2=3
$$

See Also
maineffectsplot | interactionplot
How To

- "Grouped Data" on page 2-34

```
Purpose Multivariate normal cumulative distribution function
Syntax \(\quad y=\operatorname{mvncdf}(x)\)
\(y=m v n c d f(X, m u, S I G M A)\)
\(y=m v n c d f(x l, x u, m u, S I G M A)\)
[y,err] = mvncdf(...)
[...] = mvncdf(...,options)
```

$y=m v n c d f(X)$ returns the cumulative probability of the multivariate normal distribution with zero mean and identity covariance matrix, evaluated at each row of X . Rows of the $n$-by- $d$ matrix X correspond to observations or points, and columns correspond to variables or coordinates. y is an $n$-by- 1 vector.
$y=m v n c d f(X$, mu, SIGMA $)$ returns the cumulative probability of the multivariate normal distribution with mean mu and covariance SIGMA, evaluated at each row of X. mu is a 1 -by- $d$ vector, and SIGMA is a $d$-by- $d$ symmetric, positive definite matrix. mu can also be a scalar value, which mvncdf replicates to match the size of $X$. If the covariance matrix is diagonal, containing variances along the diagonal and zero covariances off the diagonal, SIGMA may also be specified as a 1-by-d vector containing just the diagonal. Pass in the empty matrix [] for mu to use as its default value when you want to only specify SIGMA.

The multivariate normal cumulative probability at $X$ is defined as the probability that a random vector V , distributed as multivariate normal, will fall within the semi-infinite rectangle with upper limits defined by X , for example, $\operatorname{Pr}\{\mathrm{V}(1) \leq \mathrm{X}(1), \mathrm{V}(2) \leq \mathrm{X}(2), \ldots, \mathrm{V}(d) \leq \mathrm{X}(d)\}$.
$y=m v n c d f(x l, x u, m u, S I G M A)$ returns the multivariate normal cumulative probability evaluated over the rectangle with lower and upper limits defined by $x l$ and $x u$, respectively.
[y,err] = mvncdf(...) returns an estimate of the error in y. For bivariate and trivariate distributions, mvncdf uses adaptive quadrature on a transformation of the $t$ density, based on methods developed by Drezner and Wesolowsky and by Genz, as described in the references. The default absolute error tolerance for these cases is $1 \mathrm{e}-8$. For four or more dimensions, mvncdf uses a quasi-Monte Carlo integration
algorithm based on methods developed by Genz and Bretz, as described in the references. The default absolute error tolerance for these cases is $1 \mathrm{e}-4$.
[...] = mvncdf(..., options) specifies control parameters for the numerical integration used to compute $y$. This argument can be created by a call to statset. Choices of statset parameters:

- 'TolFun' - Maximum absolute error tolerance. Default is 1e-8 when $d<4$, or $1 \mathrm{e}-4$ when $d \geq 4$.
- 'MaxFunEvals' - Maximum number of integrand evaluations allowed when $d \geq 4$. Default is 1e7. 'MaxFunEvals' is ignored when $d<4$.
- 'Display' - Level of display output. Choices are 'off' (the default), 'iter', and 'final'. 'Display' is ignored when $d<4$.


## Examples

```
mu = [1 -1]; SIGMA = [.9 .4; .4 .3];
[X1,X2] = meshgrid(linspace(-1,3,25)',linspace(-3,1,25)');
X = [X1(:) X2(:)];
p = mvncdf(X,mu,SIGMA);
surf(X1,X2,reshape(p,25,25));
```



## References

[1] Drezner, Z. "Computation of the Trivariate Normal Integral." Mathematics of Computation. Vol. 63, 1994, pp. 289-294.
[2] Drezner, Z., and G. O. Wesolowsky. "On the Computation of the Bivariate Normal Integral." Journal of Statistical Computation and Simulation. Vol. 35, 1989, pp. 101-107.
[3] Genz, A. "Numerical Computation of Rectangular Bivariate and Trivariate Normal and t Probabilities." Statistics and Computing. Vol. 14, No. 3, 2004, pp. 251-260.
[4] Genz, A., and F. Bretz. "Numerical Computation of Multivariate t Probabilities with Application to Power Calculation of Multiple

Contrasts." Journal of Statistical Computation and Simulation. Vol. 63, 1999, pp. 361-378.
[5] Genz, A., and F. Bretz. "Comparison of Methods for the Computation of Multivariate $t$ Probabilities." Journal of Computational and Graphical Statistics. Vol. 11, No. 4, 2002, pp. 950-971.

## See Also mvnpdf | mvnrnd

How To . "Multivariate Normal Distribution" on page B-58

## Purpose Multivariate normal probability density function

Syntax $\quad$| $y$ | $=\operatorname{mvnpdf}(X)$ |
| ---: | :--- |
| $y$ | $=\operatorname{mvnpdf}(X, M U)$ |
| $y$ | $=\operatorname{mvnpdf}(X, M U, S I G M A)$ |

## Description

$y=m v n p d f(X)$ returns the $n$-by- 1 vector $y$, containing the probability density of the multivariate normal distribution with zero mean and identity covariance matrix, evaluated at each row of the $n$-by- $d$ matrix X . Rows of X correspond to observations and columns correspond to variables or coordinates.
$\mathrm{y}=\mathrm{mvnpdf}(\mathrm{X}, \mathrm{MU})$ returns the density of the multivariate normal distribution with mean mu and identity covariance matrix, evaluated at each row of X . MU is a 1 -by- $d$ vector, or an $n$-by- $d$ matrix. If MU is a matrix, the density is evaluated for each row of $X$ with the corresponding row of MU. MU can also be a scalar value, which mvnpdf replicates to match the size of X .
$y=m v n p d f(X, M U, S I G M A)$ returns the density of the multivariate normal distribution with mean MU and covariance SIGMA, evaluated at each row of X. SIGMA is a $d$-by- $d$ matrix, or a $d$-by- $d$-by- $n$ array, in which case the density is evaluated for each row of $X$ with the corresponding page of SIGMA, i.e., mvnpdf computes y(i) using X(i,:) and $\operatorname{SIGMA}(:,:, i)$. If the covariance matrix is diagonal, containing variances along the diagonal and zero covariances off the diagonal, SIGMA may also be specified as a 1-by- $d$ vector or a 1-by- $d$-by- $n$ array, containing just the diagonal. Specify [] for MU to use its default value when you want to specify only SIGMA.

If X is a 1 -by- $d$ vector, mvnpdf replicates it to match the leading dimension of mu or the trailing dimension of SIGMA.

## Examples

```
mu = [1 -1];
SIGMA = [.9 .4; .4 .3];
X = mvnrnd(mu,SIGMA,10);
p = mvnpdf(X,mu,SIGMA);
```

See Also mvncdf | mvnrnd | normpdf
How To . "Multivariate Normal Distribution" on page B-58

```
Purpose Multivariate linear regression
Syntax beta \(=\) mvregress \((X, Y)\)
[beta,SIGMA] = mvregress(X,Y)
[beta,SIGMA,RESID] = mvregress(X,Y)
[beta,SIGMA,RESID,COVB] = mvregress(...)
[beta,SIGMA,RESID,COVB,objective] = mvregress(...)
[...] = mvregress(X,Y,param1,val1,param2,val2,...)
Description
beta \(=\) mvregress \((X, Y)\) returns a vector beta of coefficient estimates for a multivariate regression of the \(d\)-dimensional responses in \(Y\) on the predictors in X . If \(d=1, \mathrm{X}\) can be an \(n\)-by- \(p\) matrix of \(p\) predictors at each of \(n\) observations. If \(d \geq 1, \mathrm{X}\) can be a cell array of length \(n\), with each cell containing a \(d\)-by- \(p\) design matrix for one multivariate observation. If all observations have the same \(d\)-by- \(p\) design matrix, X can be a single cell. Y is \(n\)-by- \(d\). beta is \(p\)-by- 1 .
```

Note To include a constant term in a model, a matrix $X$ should contain a column of 1 s . If $X$ is a cell array $X$, each cell should contain a matrix with a column of 1 s .
mvregress treats NaNs in $X$ or $Y$ as missing values. Missing values in $X$ are ignored. Missing values in $Y$ are handled according to the value of the 'algorithm' parameter described below.
[beta,SIGMA] = mvregress $(\mathrm{X}, \mathrm{Y})$ also returns a $d$-by- $d$ matrix SIGMA for the estimated covariance of $Y$.
[beta,SIGMA, RESID] = mvregress( $\mathrm{X}, \mathrm{Y}$ ) also returns an $n$-by- $d$ matrix RESID of residuals.

The RESID values corresponding to missing values in $Y$ are the differences between the conditionally imputed values for $Y$ and the fitted values. The SIGMA estimate is not the sample covariance matrix of the RESID matrix.
[beta,SIGMA,RESID,COVB] = mvregress(...) also returns a matrix COVB for the estimated covariance of the coefficients. By default, or if the 'varformat ' parameter is 'beta' (see below), COVB is the estimated covariance matrix of beta. If the 'varformat' parameter is 'full', COVB is the combined estimated covariance matrix for beta and SIGMA.
[beta,SIGMA,RESID,COVB,objective] = mvregress(...) also returns the value of the objective function, or log likelihood, objective, after the last iteration.
[...] = mvregress(X,Y,param1,val1,param2,val2,...) specifies additional parameter name/value pairs chosen from the following:

- 'algorithm' - Either 'ecm' to compute the maximum likelihood estimates via the ECM algorithm, 'cwls' to perform least squares (optionally conditionally weighted by an input covariance matrix), or 'mvn' to omit observations with missing data and compute the ordinary multivariate normal estimates. The default is 'mvn' for complete data, 'ecm' for missing data when the sample size is sufficient to estimate all parameters, and 'cwls' otherwise.
- 'covaro' - A d-by- $d$ matrix to be used as the initial estimate for SIGMA. The default is the identity matrix. For the 'cwls ' algorithm, this matrix is usually a diagonal matrix used as a weighting at each iteration. The 'cwls' algorithm uses the initial value of SIGMA at each iteration, without changing it.
- 'covtype' - Either 'full', to allow a full covariance matrix, or 'diagonal', to restrict the covariance matrix to be diagonal. The default is 'full'.
- 'maxiter' - Maximum number of iterations. The default is 100.
- 'outputfen' - An output function called with three arguments:

1. A vector of current parameter estimates.
2. A structure with fields 'Covar' for the current value of the covariance matrix, 'iteration' for the current iteration number, and 'fval' for the current value of the objective function.
3. A text string that is 'init' when called during initialization, 'iter' when called after an iteration, and 'done' when called after completion.

The function should return logical true if the iterations should stop, or logical false if they should continue.

- 'param0' - A vector of $p$ elements to be used as the initial estimate for beta. Default is a zero vector. Not used for the 'mvn' algorithm.
- 'tolbeta' - Convergence tolerance for beta. The default is sqrt(eps). Iterations continue until the tolbeta and tolobj conditions are met. The test for convergence at iteration k is

```
norm(beta(k)-beta(k-1))<sqrt(p)*tolbeta*(1+norm(beta(k)))
```

where $p=$ length(beta).

- 'tolobj' - Convergence tolerance for changes in the objective function. The default is $\mathrm{eps}^{\wedge}(3 / 4)$. The test is
abs(obj(k)-obj(k-1)) < tolobj*(1+abs(obj(k)))
where obj is the objective function. If both tolobj and tolbeta are 0 , the function performs maxiter iterations with no convergence test.
- 'beta0' - A vector of $p$ elements to be used as the initial estimate for beta. Default is a zero vector. Not used for the 'mvn' algorithm.
- 'covaro' - A $d$-by- $d$ matrix to be used as the initial estimate for SIGMA. Default is the identity matrix. For the 'cwls' algorithm, this matrix is usually a diagonal matrix and it is not changed during the iterations, so the input value is used as the weighting matrix at each iteration.
- 'outputfon' - An output function.
- 'varformat' - Either 'beta' to compute COVB for beta only (default), or 'full' to compute COVB for both beta and SIGMA.
- 'vartype' - Either 'hessian' to compute COVB using the Hessian or observed information (default), or 'fisher' to compute COVB using
the complete-data Fisher or expected information. The 'hessian' method takes into account the increased uncertainties due to missing data, while the 'fisher' method does not.


## Examples

Predict regional flu estimates based on Google ${ }^{\mathrm{TM}}$ queries using the national CDC estimates as a predictor:

```
load flu
% response: regional queries
y = double(flu(:,2:end-1));
% predictor: national CDC estimates
x = flu.WtdILI;
[nobs,nregions] = size(y);
% Create and fit model with separate intercepts but
% common slope
X = cell(nobs,1);
for j=1:nobs
    X{j} = [eye(nregions), repmat(x(j),nregions,1)];
end
[beta,sig,resid,vars,loglik] = mvregress(X,y);
% Plot raw data with fitted lines
B = [beta(1:nregions)';repmat(beta(end),1,nregions)]
axes1 = axes('Position',[0.13 0.5838 0.6191 0.3412]);
xx = linspace(.5,3.5)';
h = plot(x,y,'x', xx, [ones(size(xx)),xx]*B,'-');
for j=1:nregions;
        set(h(nregions+j),'color',get(h(j),'color'));
end
regions = flu.Properties.VarNames;
legend1 = legend(regions{2:end-1});
set(legend1,'Position', [0.7733 0.1967 0.2161 0.6667]);
% Create and fit model with separate intercepts and slopes
```

```
for j=1:nobs
    X{j} = [eye(nregions), x(j)*eye(nregions)];
end
[beta,sig,resid,vars,loglik2] = mvregress(X,y);
% Plot raw data with fitted lines
B = [beta(1:nregions)';beta(nregions+1:end)']
axes2 = axes('Parent',gcf,'Position',...
    [0.13 0.11 0.6191 0.3412]);
h = plot(x,y,'x', xx, [ones(size(xx)),xx]*B,' -');
for j=1:nregions;
    set(h(nregions+j),'color',get(h(j),'color'));
end
% Likelihood ratio test for significant difference
chisq = 2*(loglik2-loglik)
p = 1-chi2cdf(chisq, nregions-1)
chisq =
    96.4556
p =
0
```



## References

[3] Sexton, Joe, and A. R. Swensen. "ECM Algorithms that Converge at the Rate of EM." Biometrika. Vol. 87, No. 3, 2000, pp. 651-662.
[4] Dempster, A. P., N. M. Laird, and D. B. Rubin. "Maximum Likelihood from Incomplete Data via the EM Algorithm." Journal of the Royal Statistical Society. Series B, Vol. 39, No. 1, 1977, pp. 1-37.

See Also mvregresslike | manova1<br>How To . "Multivariate Normal Distribution" on page B-58

Purpose Negative log-likelihood for multivariate regression
Syntax $\quad \begin{array}{ll}\text { nlogL }=\text { mvregresslike }(X, Y, b, S I G M A, ~ a l g) ~ \\ {[n l o g L, C O V B]=\text { mvregresslike }(\ldots)} \\ & {[\text { nlogL }, \text { COVB }]=\text { mvregresslike }(\ldots, \text { type, format })}\end{array}$

## Description

nlogL $=$ mvregresslike $(X, Y, b$, SIGMA,$a l g)$ computes the negative log-likelihood nlogL for a multivariate regression of the $d$-dimensional multivariate observations in the $n$-by- $d$ matrix Y on the predictor variables in the matrix or cell array X , evaluated for the $p$-by- 1 column vector b of coefficient estimates and the $d$-by- $d$ matrix SIGMA specifying the covariance of a row of Y . If $d=1, \mathrm{X}$ can be an $n$-by- $p$ design matrix of predictor variables. For any value of $d, \mathrm{X}$ can also be a cell array of length $n$, with each cell containing a $d$-by- $p$ design matrix for one multivariate observation. If all observations have the same $d$-by-p design matrix, $X$ can be a single cell.
NaN values in $X$ or $Y$ are taken as missing. Observations with missing values in $X$ are ignored. Treatment of missing values in $Y$ depends on the algorithm specified by alg.
alg should match the algorithm used by mvregress to obtain the coefficient estimates $b$, and must be one of the following:

- 'ecm' - ECM algorithm
- 'cwls' - Least squares conditionally weighted by SIGMA
- 'mvn' - Multivariate normal estimates computed after omitting rows with any missing values in $Y$
[nlogL,COVB] = mvregresslike(...) also returns an estimated covariance matrix COVB of the parameter estimates $b$.
[nlogL,COVB] = mvregresslike(...,type,format) specifies the type and format of COVB.
type is either:
- 'hessian' - To use the Hessian or observed information. This method takes into account the increased uncertainties due to missing data. This is the default.
- 'fisher' - To use the Fisher or expected information. This method uses the complete data expected information, and does not include uncertainty due to missing data.
format is either:
- 'beta' - To compute COVB for $b$ only. This is the default.
- 'full' - To compute COVB for both $b$ and SIGMA.


## See Also mvregress | manova1

How To . "Multivariate Normal Distribution" on page B-58

## Purpose Multivariate normal random numbers

```
Syntax \(\quad R=\operatorname{mvnrnd}(M U, S I G M A)\)
\(r=\) mvnrnd(MU,SIGMA, cases)
```

$R=$ mvnrnd(MU,SIGMA) returns an $n$-by- $d$ matrix $R$ of random vectors chosen from the multivariate normal distribution with mean MU, and covariance SIGMA. MU is an $n$-by- $d$ matrix, and mvnrnd generates each row of R using the corresponding row of mu. SIGMA is a $d$-by- $d$ symmetric positive semi-definite matrix, or a $d$-by- $d$-by-n array. If SIGMA is an array, mvnrnd generates each row of R using the corresponding page of SIGMA, i.e., mvnrnd computes R(i,:) using MU(i,:) and SIGMA(:,:,i). If the covariance matrix is diagonal, containing variances along the diagonal and zero covariances off the diagonal, SIGMA may also be specified as a 1 -by- $d$ vector or a 1-by- $d$-by- $n$ array, containing just the diagonal. If $M U$ is a 1 -by- $d$ vector, mvnrnd replicates it to match the trailing dimension of SIGMA.
$r=$ mvnrnd(MU,SIGMA, cases) returns a cases-by-d matrix R of random vectors chosen from the multivariate normal distribution with a common 1-by-d mean vector MU, and a common d-by-d covariance matrix SIGMA.

## Examples

```
mu = [2 3];
SIGMA = [1 1.5; 1.5 3];
r = mvnrnd(mu,SIGMA,100);
plot(r(:,1),r(:,2),'+')
```



See Also
mvnpdf | mvncdf | normrnd
How To . "Multivariate Normal Distribution" on page B-58

```
Purpose Multivariate \(t\) cumulative distribution function
Syntax \(\quad y=\operatorname{mvtcdf}(X, C, D F)\)
\(y=\operatorname{mvtcdf}(x l, x u, C, D F)\)
[y,err] = mvtcdf(...)
[...] = mvntdf(...,options)
```


## Description

$y=m v t c d f(X, C, D F)$ returns the cumulative probability of the multivariate $t$ distribution with correlation parameters C and degrees of freedom DF, evaluated at each row of X. Rows of the $n$-by- $d$ matrix $X$ correspond to observations or points, and columns correspond to variables or coordinates. y is an n -by- 1 vector.

C is a symmetric, positive definite, $d$-by- $d$ matrix, typically a correlation matrix. If its diagonal elements are not 1, mvtcdf scales $C$ to correlation form. mvtcdf does not rescale X. DF is a scalar, or a vector with $n$ elements.

The multivariate $t$ cumulative probability at X is defined as the probability that a random vector T , distributed as multivariate $t$, will fall within the semi-infinite rectangle with upper limits defined by $x$, i.e., $\operatorname{Pr}\{T(1) \leq X(1), T(2) \leq X(2), \ldots T(d) \leq X(d)\}$.
$\mathrm{y}=\operatorname{mvtcdf}(\mathrm{xl}, \mathrm{xu}, \mathrm{C}, \mathrm{DF})$ returns the multivariate $t$ cumulative probability evaluated over the rectangle with lower and upper limits defined by $x l$ and $x u$, respectively.
[y,err] $=\operatorname{mvtcdf}(\ldots)$ returns an estimate of the error in $y$. For bivariate and trivariate distributions, mvtcdf uses adaptive quadrature on a transformation of the $t$ density, based on methods developed by Genz, as described in the references. The default absolute error tolerance for these cases is $1 e-8$. For four or more dimensions, mvtcdf uses a quasi-Monte Carlo integration algorithm based on methods developed by Genz and Bretz, as described in the references. The default absolute error tolerance for these cases is 1e-4.
[...] = mvntdf(..., options) specifies control parameters for the numerical integration used to compute $y$. This argument can be created by a call to statset. Choices of statset parameters are:

- 'TolFun' - Maximum absolute error tolerance. Default is 1e-8 when $d<4$, or $1 \mathrm{e}-4$ when $d \geq 4$.
- 'MaxFunEvals' - Maximum number of integrand evaluations allowed when $d \geq 4$. Default is 1 e 7 . 'MaxFunEvals' is ignored when $d<4$.
- 'Display' - Level of display output. Choices are 'off' (the default), 'iter', and 'final'. 'Display' is ignored when $d<4$.


## Examples

$C=[1.4 ; .41] ; d f=2$;
[X1,X2] = meshgrid(linspace(-2,2,25)',linspace(-2,2,25)');
X = [X1(:) X2(:)];
$\mathrm{p}=\mathrm{mvtcdf}(\mathrm{X}, \mathrm{C}, \mathrm{df})$;
$\operatorname{surf}(X 1, X 2$, reshape $(p, 25,25))$;


## References

[1] Genz, A. "Numerical Computation of Rectangular Bivariate and Trivariate Normal and t Probabilities." Statistics and Computing. Vol. 14, No. 3, 2004, pp. 251-260.
[2] Genz, A., and F. Bretz. "Numerical Computation of Multivariate t Probabilities with Application to Power Calculation of Multiple Contrasts." Journal of Statistical Computation and Simulation. Vol. 63, 1999, pp. 361-378.
[3] Genz, A., and F. Bretz. "Comparison of Methods for the Computation of Multivariate t Probabilities." Journal of Computational and Graphical Statistics. Vol. 11, No. 4, 2002, pp. 950-971.
$\begin{array}{ll}\text { See Also } & \text { mvtpdf | mvtrnd } \\ \text { How To } & \text { - "Multivariate t Distribution" on page B-64 }\end{array}$

Purpose Multivariate $t$ probability density function

## Syntax $\quad y=m v t p d f(x, C, d f)$

Description
$y=m v t p d f(X, C, d f)$ returns the probability density of the multivariate $t$ distribution with correlation parameters C and degrees of freedom df , evaluated at each row of $X$. Rows of the $n$-by- $d$ matrix X correspond to observations or points, and columns correspond to variables or coordinates. C is a symmetric, positive definite, $d$-by- $d$ matrix, typically a correlation matrix. If its diagonal elements are not 1 , mvtpdf scales $C$ to correlation form. mvtcdf does not rescale X . df is a scalar, or a vector with $n$ elements. y is an $n$-by- 1 vector.

## Examples Visualize a multivariate $t$ distribution:

```
[X1,X2] = meshgrid(linspace(-2,2,25)',linspace(-2,2,25)');
X = [X1(:) X2(:)];
C = [1 .4; .4 1];
df = 2;
p = mvtpdf(X,C,df);
surf(X1,X2,reshape(p,25,25))
```



See Also
mvtcdf | mvtrnd
How To

- "Multivariate t Distribution" on page B-64


## Purpose Multivariate $t$ random numbers

```
Syntax \(\quad R=\operatorname{mvtrnd}(C, d f\), cases \()\)
R = mvtrnd(C,df)
```


## Description

$R=m v t r n d(C, d f, c a s e s)$ returns a matrix of random numbers chosen from the multivariate $t$ distribution, where C is a correlation matrix. $d f$ is the degrees of freedom and is either a scalar or is a vector with cases elements. If $p$ is the number of columns in C , then the output $R$ has cases rows and $p$ columns.

Let $t$ represent a row of $R$. Then the distribution of $t$ is that of a vector having a multivariate normal distribution with mean 0 , variance 1 , and covariance matrix C , divided by an independent chi-square random value having df degrees of freedom. The rows of $R$ are independent.
$C$ must be a square, symmetric and positive definite matrix. If its diagonal elements are not all 1 (that is, if C is a covariance matrix rather than a correlation matrix), mvtrnd rescales C to transform it to a correlation matrix before generating the random numbers.
$R=m v t r n d(C, d f)$ returns a single random number from the multivariate $t$ distribution.

## Examples

$$
\begin{aligned}
& \text { SIGMA }=[10.8 ; 0.81] ; \\
& \text { R }=\text { mvtrnd (SIGMA, } 3,100) ; \\
& \operatorname{plot}\left(R(:, 1), R(:, 2),+^{\prime}\right)
\end{aligned}
$$



See Also
How To - "Multivariate $t$ Distribution" on page B-64

## cvpartition.N property

$\begin{array}{ll}\text { Purpose } & \begin{array}{l}\text { Number of observations (including observations with missing group } \\ \text { values) }\end{array} \\ \text { Description } & \begin{array}{l}\text { Number of observations (including observations with missing group } \\ \text { values). }\end{array}\end{array}$

## Purpose Naive Bayes classifier

Description A NaiveBayes object defines a Naive Bayes classifier. A Naive Bayes classifier assigns a new observation to the most probable class, assuming the features are conditionally independent given the class value.

## Construction

NaiveBayes
Create NaiveBayes object

## Methods

disp
display
fit
posterior
predict
subsasgn
subsref

Display NaiveBayes classifier object
Display NaiveBayes classifier object
Create Naive Bayes classifier object by fitting training data

Compute posterior probability of each class for test data

Predict class label for test data
Subscripted reference for NaiveBayes object
Subscripted reference for NaiveBayes object

## Properties

CIsNonEmpty
CLevels
CNames
CPrior

Flag for non-empty classes
Class levels
Class names
Class priors

## NaiveBayes

| Dist | Distribution names |
| :--- | :--- |
| NClasses | Number of classes |
| NDims | Number of dimensions |
| Params | Parameter estimates |

Copy
Semantics

## Examples Predict the class label using the Naive Bayes classifier

```
load fisheriris
```

Use the default Gaussian distribution and a confusion matrix:

```
01 = NaiveBayes.fit(meas,species);
C1 = 01.predict(meas);
cMat1 = confusionmat(species,C1)
```

This returns:

```
cMat1 =
    50 0 0
    0 47 3
    0}
```

Use the Gaussian distribution for features 1 and 3 and use the kernel density estimation for features 2 and 4 :

```
02 = NaiveBayes.fit(meas,species,'dist',...
{'normal','kernel','normal','kernel'});
C2 = 02.predict(meas);
cMat2 = confusionmat(species,C2)
```

This returns:

## NaiveBayes

```
cMat2 =
    50 0 0
    0 47 3
    0 3 47
```


## References [1] Mitchell, T. (1997) Machine Learning, McGraw Hill.

[2] Vangelis M., Ion A., and Geogios P. Spam Filtering with Naive Bayes - Which Naive Bayes? (2006) Third Conference on Email and Anti-Spam.
[3] George H. John and Pat Langley. Estimating continuous distributions in bayesian classifiers (1995) the Eleventh Conference on Uncertainty in Artificial Intelligence.

How To . "Naive Bayes Classification" on page 12-29

- "Grouped Data" on page 2-34


## NaiveBayes

$$
\begin{array}{ll}
\text { Purpose } & \text { Create NaiveBayes object } \\
\text { Description } & \begin{array}{l}
\text { You cannot create a NaiveBayes classifier by calling the constructor. } \\
\text { Use NaiveBayes.fit to create a NaiveBayes classifier by fitting the } \\
\text { object to training data. }
\end{array}
\end{array}
$$

See Also fit

## Purpose Covariance ignoring NaN values

Syntax
$Y=n a n c o v(X)$
$Y=n a n c o v(X 1, X 2)$
$Y=\operatorname{nancov}(. . ., 1)$
Y = nancov(...,' pairwise')

Description

## Examples

$Y=$ nancov $(X)$ is the covariance cov of $X$, computed after removing observations with NaN values.

For vectors $x$, $\operatorname{nancov}(x)$ is the sample variance of the remaining elements, once NaN values are removed. For matrices X, nancov (X) is the sample covariance of the remaining observations, once observations (rows) containing any NaN values are removed.
$Y=\operatorname{nancov}(X 1, X 2)$, where $X 1$ and $X 2$ are matrices with the same number of elements, is equivalent to $\operatorname{nancov}(X)$, where $X=[X 1(:)$ X2(:)].
nancov removes the mean from each variable (column for matrix $X$ ) before calculating Y. If $n$ is the number of remaining observations after removing observations with NaN values, nancov normalizes Y by either $n-1$ or $n$, depending on whether $n>1$ or $n=1$, respectively. To specify normalization by $n$, use $Y=\operatorname{nancov}(\ldots, 1)$.
$Y=$ nancov(...,'pairwise') computes $Y(i, j)$ using rows with no NaN values in columns i or $j$. The result $Y$ may not be a positive definite matrix.

Generate random data for two variables (columns) with random missing values:

```
X = rand(10,2);
p = randperm(numel(X));
X(p(1:5)) = NaN
X =
    0.8147 0.1576
    NaN NaN
    0.1270 0.9572
```

| 0.9134 | NaN |
| :--- | ---: |
| 0.6324 | NaN |
| 0.0975 | 0.1419 |
| 0.2785 | 0.4218 |
| 0.5469 | 0.9157 |
| 0.9575 | 0.7922 |
| 0.9649 | NaN |

Establish a correlation between a third variable and the other two variables:

| $\mathrm{X}(:, 3)=\operatorname{sum}(X, 2)$ |  |  |
| :--- | ---: | ---: |
| $X=$ |  |  |
| 0.8147 | 0.1576 | 0.9723 |
| NaN | NaN | NaN |
| 0.1270 | 0.9572 | 1.0842 |
| 0.9134 | NaN | NaN |
| 0.6324 | NaN | NaN |
| 0.0975 | 0.1419 | 0.2394 |
| 0.2785 | 0.4218 | 0.7003 |
| 0.5469 | 0.9157 | 1.4626 |
| 0.9575 | 0.7922 | 1.7497 |
| 0.9649 | NaN | NaN |

Compute the covariance matrix for the three variables after removing observations (rows) with NaN values:

```
Y = nancov(X)
Y =
\begin{tabular}{lll}
0.1311 & 0.0096 & 0.1407 \\
0.0096 & 0.1388 & 0.1483 \\
0.1407 & 0.1483 & 0.2890
\end{tabular}
```

See Also
NaN | cov | var | nanvar

```
Purpose Maximum ignoring NaN values
Syntax \(\quad y=\operatorname{nanmax}(x)\)
\(Y=\operatorname{nanmax}(X 1, X 2)\)
\(y=n a n m a x(X,[], d i m)\)
[y,indices] = nanmax(...)
```

Description $\quad y=n a n m a x(X)$ is the maximum max of $X$, computed after removing NaN values.

For vectors x , $\operatorname{nanmax}(\mathrm{x})$ is the maximum of the remaining elements, once NaN values are removed. For matrices X , nanmax $(\mathrm{X})$ is a row vector of column maxima, once NaN values are removed. For multidimensional arrays $X$, nanmax operates along the first nonsingleton dimension.
$Y=\operatorname{nanmax}(\mathrm{X} 1, \mathrm{X} 2)$ returns an array Y the same size as X 1 and X 2 with $Y(i, j)=$ nanmax (X1 $(i, j), X 2(i, j))$. Scalar inputs are expanded to an array of the same size as the other input.
$\mathrm{y}=\operatorname{nanmax}(\mathrm{X},[\mathrm{l}, \operatorname{dim})$ operates along the dimension $\operatorname{dim}$ of X .
[ $y$, indices] = nanmax(...) also returns the row indices of the maximum values for each column in the vector indices.

## Examples

Find column maxima and their indices for data with missing values:

```
X = magic(3);
X([1 6:9]) = repmat(NaN,1,5)
X =
    NaN 1 NaN
        3 5 NaN
        4 ~ N a N ~ N a N
[y,indices] = nanmax(X)
y =
    4 NaN
indices =
    3 2 1
```

See Also
NaN | max | nanmin

## Purpose Mean ignoring NaN values

```
Syntax
y = nanmean(X)
y = nanmean(X,dim)
```

$y=$ nanmean $(X)$ is the mean of $X$, computed after removing NaN values.
For vectors $x$, nanmean $(x)$ is the mean of the remaining elements, once NaN values are removed. For matrices $X$, nanmean $(X)$ is a row vector of column means, once NaN values are removed. For multidimensional arrays X , nanmean operates along the first nonsingleton dimension. $\mathrm{y}=$ nanmean( $\mathrm{X}, \mathrm{dim}$ ) takes the mean along dimension dim of X .

Note If X contains a vector of all NaN values along some dimension, the vector is empty once the NaN values are removed, so the sum of the remaining elements is 0 . Since the mean involves division by 0 , its value is NaN . The output NaN is not a mean of NaN values.

## Examples <br> Find column means for data with missing values:

```
X = magic(3);
X([1 6:9]) = repmat(NaN,1,5)
X =
    NaN 1 NaN
        3 5 NaN
        4 ~ N a N ~ N a N
y = nanmean(X)
y =
    3.5000 3.0000 NaN
```

See Also NaN | mean | nanmedian

## Purpose Median ignoring NaN values

```
Syntax y = nanmedian(X)
y = nanmedian(X,dim)
```


## Description

$\mathrm{y}=$ nanmedian( X$)$ is the median of X , computed after removing NaN values.

For vectors $x$, nanmedian ( $x$ ) is the median of the remaining elements, once NaN values are removed. For matrices $X$, nanmedian $(X)$ is a row vector of column medians, once NaN values are removed. For multidimensional arrays X , nanmedian operates along the first nonsingleton dimension.
$\mathrm{y}=$ nanmedian(X, dim) takes the mean along dimension dim of X .

## Examples Find column medians for data with missing values:

```
X = magic(3);
X([1 6:9]) = repmat(NaN,1,5)
X =
    NaN 1 NaN
        3 5 NaN
        4 NaN NaN
y = nanmedian(X)
y =
    3.5000 3.0000 NaN
```

See Also NaN | median | nanmean

Purpose Minimum ignoring NaN values

```
Syntax
\(y=n a n m i n(X)\)
\(Y=\operatorname{nanmin}(X 1, X 2)\)
y = nanmin(X,[],dim)
[y,indices] = nanmin(...)
```


## Description

## Examples

Find column minima and their indices for data with missing values:
$y=\operatorname{nanmin}(X)$ is the minimum min of $X$, computed after removing NaN values.

For vectors $x$, nanmin $(x)$ is the minimum of the remaining elements, once NaN values are removed. For matrices X, nanmin (X) is a row vector of column minima, once NaN values are removed. For multidimensional arrays $X$, nanmin operates along the first nonsingleton dimension.
$Y=$ nanmin $(X 1, X 2)$ returns an array $Y$ the same size as $X 1$ and $X 2$ with $Y(i, j)=$ nanmin $(X 1(i, j), X 2(i, j))$. Scalar inputs are expanded to an array of the same size as the other input.
$\mathrm{y}=$ nanmin(X, [], dim) operates along the dimension dim of X .
[y,indices] = nanmin(...) also returns the row indices of the minimum values for each column in the vector indices.

```
X = magic(3);
```

X = magic(3);
X([1 6:9]) = repmat(NaN,1,5)
X([1 6:9]) = repmat(NaN,1,5)
X =
X =
NaN 1 NaN
NaN 1 NaN
3 5 NaN
3 5 NaN
NaN NaN
NaN NaN
[y,indices] = nanmin(X)
[y,indices] = nanmin(X)
y =
y =
3 1 NaN
3 1 NaN
indices =
indices =
2 1 1

```
    2 1 1
```


## See Also

NaN | min | nanmax

## Purpose Standard deviation ignoring NaN values

```
Syntax
y = nanstd(X)
y = nanstd(X,1)
y = nanstd(X,flag,dim)
```


## Description

Examples
Find column standard deviations for data with missing values:

```
X = magic(3);
X([1 6:9]) = repmat(NaN,1,5)
X =
    NaN 1 NaN
        3 NaN
        4 ~ N a N ~ N a N
y = nanstd(X)
y =
    0.7071 2.8284 NaN
```

See Also NaN \| std | nanvar | nanmean

## Purpose Sum ignoring NaN values

Syntax
y = nansum( X )
y = nansum(X,dim)
$y=$ nansum $(X)$ is the sum of $X$, computed after removing NaN values.
For vectors $x$, nansum ( $x$ ) is the sum of the remaining elements, once NaN values are removed. For matrices $X$, nansum (X) is a row vector of column sums, once NaN values are removed. For multidimensional arrays X , nansum operates along the first nonsingleton dimension. $\mathrm{y}=$ nansum( $\mathrm{X}, \mathrm{dim}$ ) takes the sum along dimension $\operatorname{dim}$ of X .

Note If X contains a vector of all NaN values along some dimension, the vector is empty once the NaN values are removed, so the sum of the remaining elements is 0 . The output 0 is not a sum of NaN values.

## Examples Find column sums for data with missing values:

```
X = magic(3);
X([1 6:9]) = repmat(NaN,1,5)
X =
    NaN 1 NaN
        3 NaN
        NaN NaN
y = nansum(X)
y =
    7 6 0
```


## See Also <br> NaN | sum

## Purpose Variance, ignoring NaN values

```
Syntax
y = nanvar(X)
y = nanvar(X,1)
y = nanvar(X,w)
y = nanvar(X,w,dim)
```

Description

Examples
$\mathrm{y}=$ nanvar(X) is the variance var of X , computed after removing NaN values.

For vectors $x$, nanvar $(x)$ is the sample variance of the remaining elements, once NaN values are removed. For matrices X, nanvar (X) is a row vector of column sample variances, once NaN values are removed. For multidimensional arrays X, nanvar operates along the first nonsingleton dimension.
nancov removes the mean from each variable (column for matrix $X$ ) before calculating Y. If $n$ is the number of remaining observations after removing observations with NaN values, nanvar normalizes y by either $n-1$ or $n$, depending on whether $n>1$ or $n=1$, respectively. To specify normalization by $n$, use $\mathrm{y}=$ nanvar $(\mathrm{X}, 1)$.
$y=$ nanvar $(X, w)$ computes the variance using the weight vector $w$. The length of $w$ must equal the length of the dimension over which nanvar operates, and its elements must be nonnegative. Elements of X corresponding to NaN values of $w$ are ignored.
$\mathrm{y}=$ nanvar( $\mathrm{X}, \mathrm{w}, \mathrm{dim})$ takes the variance along the dimension $\operatorname{dim}$ of X . Set $w$ to [ ] to use the default normalization by $n-1$.

Find column standard deviations for data with missing values:

```
X = magic(3);
X([1 6:9]) = repmat(NaN,1,5)
X =
    NaN 1 NaN
        3 5 NaN
        NaN NaN
y = nanvar(X)
```


# $\mathrm{y}=$ <br> 0.50008 .0000 NaN 

See Also NaN | var | nanstd | nanmean

## Purpose Negative binomial cumulative distribution function

$$
\text { Syntax } \quad Y=\operatorname{nbincdf}(X, R, P)
$$

Description
$Y=\operatorname{nbincdf}(X, R, P)$ computes the negative binomial cdf at each of the values in $X$ using the corresponding number of successes, $R$ and probability of success in a single trial, P. X, R, and $P$ can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of Y . A scalar input for $\mathrm{X}, \mathrm{R}$, or P is expanded to a constant array with the same dimensions as the other inputs.

The negative binomial cdf is

$$
y=F(x \mid r, p)=\sum_{i=0}^{x}\binom{r+i-1}{i} p^{r} q^{i} I_{(0,1, \ldots)}(i)
$$

The simplest motivation for the negative binomial is the case of successive random trials, each having a constant probability P of success. The number of extra trials you must perform in order to observe a given number R of successes has a negative binomial distribution. However, consistent with a more general interpretation of the negative binomial, nbincdf allows $R$ to be any positive value, including nonintegers. When $R$ is noninteger, the binomial coefficient in the definition of the cdf is replaced by the equivalent expression

$$
\frac{\Gamma(r+i)}{\Gamma(r) \Gamma(i+1)}
$$

## Examples

```
x = (0:15);
p = nbincdf(x,3,0.5);
stairs(x,p)
```



See Also cdf | nbinpdf | nbininv | nbinstat | nbinfit | nbinrnd
How To . "Negative Binomial Distribution" on page B-72


Note The variance of a negative binomial distribution is greater than its mean. If the sample variance of the data in data is less than its sample mean, nbinfit cannot compute MLEs. You should use the poissfit function instead.

See Also
nbincdf | nbininv | nbinpdf | nbinrnd | nbinstat | mle | statset
How To • "Negative Binomial Distribution" on page B-72

## Purpose Negative binomial inverse cumulative distribution function

$$
\text { Syntax } \quad X=\operatorname{nbininv}(Y, R, P)
$$

Description

Examples
icdf | nbincdf | nbinpdf | nbinstat | nbinfit | nbinrnd
How To . "Negative Binomial Distribution" on page B-72
Note that you have to flip at least 10 times to get 10 heads. That is why the second term on the right side of the equals sign is a 10 .
How many times would you need to flip a fair coin to have a $99 \%$ probability of having observed 10 heads?
$X=$ nbininv (Y, R, P) returns the inverse of the negative binomial cdf with corresponding number of successes, $R$ and probability of success in a single trial, P. Since the binomial distribution is discrete, nbininv returns the least integer $X$ such that the negative binomial cdf evaluated at $X$ equals or exceeds Y. Y, R, and $P$ can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of X . A scalar input for Y , R , or P is expanded to a constant array with the same dimensions as the other inputs.
The simplest motivation for the negative binomial is the case of successive random trials, each having a constant probability P of success. The number of extra trials you must perform in order to observe a given number R of successes has a negative binomial distribution. However, consistent with a more general interpretation of the negative binomial, nbininv allows $R$ to be any positive value, including nonintegers.

```
flips = nbininv(0.99,10,0.5) + 10
```

flips = nbininv(0.99,10,0.5) + 10
flips =
flips =
33

```
    33
```


## Purpose Negative binomial probability density function

$$
\text { Syntax } \quad Y=\operatorname{nbinpdf}(X, R, P)
$$

Description
$Y=$ nbinpdf( $X, R, P$ ) returns the negative binomial pdf at each of the values in $X$ using the corresponding number of successes, $R$ and probability of success in a single trial, P. X, R, and P can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of Y . A scalar input for $\mathrm{X}, \mathrm{R}$, or P is expanded to a constant array with the same dimensions as the other inputs. Note that the density function is zero unless the values in X are integers.

The negative binomial pdf is

$$
y=f(x \mid r, p)=\binom{r+x-1}{x} p^{r} q^{x} I_{(0,1, \ldots)}(x)
$$

The simplest motivation for the negative binomial is the case of successive random trials, each having a constant probability $P$ of success. The number of extra trials you must perform in order to observe a given number $R$ of successes has a negative binomial distribution. However, consistent with a more general interpretation of the negative binomial, nbinpdf allows R to be any positive value, including nonintegers. When $R$ is noninteger, the binomial coefficient in the definition of the pdf is replaced by the equivalent expression

$$
\frac{\Gamma(r+x)}{\Gamma(r) \Gamma(x+1)}
$$

## Examples

```
x = (0:10);
y = nbinpdf(x,3,0.5);
plot(x,y,'+')
set(gca,'Xlim',[-0.5,10.5])
```



See Also
pdf | nbincdf | nbininv | nbinstat | nbinfit | nbinrnd
How To

- "Negative Binomial Distribution" on page B-72


## Purpose <br> Syntax <br> Description

Negative binomial random numbers

RND $=$ nbinrnd $(R, P)$
RND $=$ nbinrnd( $R, P, m, n, \ldots$ )
RND $=$ nbinrnd( $R, P,[m, n, \ldots])$

## Examples

See Also random | nbinpdf | nbincdf | nbininv | nbinstat | nbinfit
How To
RND $=$ nbinrnd ( $R, P$ ) is a matrix of random numbers chosen from a negative binomial distribution with corresponding number of successes, $R$ and probability of success in a single trial, P. R and P can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of RND. A scalar input for R or P is expanded to a constant array with the same dimensions as the other input.

RND $=n b i n r n d(R, P, m, n, \ldots)$ or RND $=n b i n r n d(R, P,[m, n, \ldots])$ generates an m-by-n-by-... array. The R, P parameters can each be scalars or arrays of the same size as R.

The simplest motivation for the negative binomial is the case of successive random trials, each having a constant probability P of success. The number of extra trials you must perform in order to observe a given number R of successes has a negative binomial distribution. However, consistent with a more general interpretation of the negative binomial, nbinrnd allows R to be any positive value, including nonintegers.

Suppose you want to simulate a process that has a defect probability of 0.01. How many units might Quality Assurance inspect before finding three defective items?

```
r = nbinrnd(3,0.01,1,6)+3
r =
\begin{tabular}{llllll}
496 & 142 & 420 & 396 & 851 & 178
\end{tabular}
```

- "Negative Binomial Distribution" on page B-72

Purpose Negative binomial mean and variance
Syntax $\quad[M, V]=$ nbinstat $(R, P)$
Description
$[M, V]=$ nbinstat $(R, P)$ returns the mean of and variance for the negative binomial distribution with corresponding number of successes, $R$ and probability of success in a single trial, P. R and P can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of $M$ and $V$. A scalar input for $R$ or $P$ is expanded to a constant array with the same dimensions as the other input.

The mean of the negative binomial distribution with parameters $r$ and $p$ is $r q / p$, where $q=1-p$. The variance is $r q / p^{2}$.

The simplest motivation for the negative binomial is the case of successive random trials, each having a constant probability $P$ of success. The number of extra trials you must perform in order to observe a given number R of successes has a negative binomial distribution. However, consistent with a more general interpretation of the negative binomial, nbinstat allows $R$ to be any positive value, including nonintegers.

## Examples

```
p = 0.1:0.2:0.9;
r = 1:5;
[R,P] = meshgrid(r,p);
[M,V] = nbinstat(R,P)
M =
\begin{tabular}{lllll}
9.0000 & 18.0000 & 27.0000 & 36.0000 & 45.0000 \\
2.3333 & 4.6667 & 7.0000 & 9.3333 & 11.6667 \\
1.0000 & 2.0000 & 3.0000 & 4.0000 & 5.0000 \\
0.4286 & 0.8571 & 1.2857 & 1.7143 & 2.1429 \\
0.1111 & 0.2222 & 0.3333 & 0.4444 & 0.5556
\end{tabular}
V =
        90.0000 180.0000 270.0000 360.0000 450.0000
        7.7778
        2.0000 4.0000 6.0000 8.0000 10.0000
```

| 0.6122 | 1.2245 | 1.8367 | 2.4490 | 3.0612 |
| :--- | :--- | :--- | :--- | :--- |
| 0.1235 | 0.2469 | 0.3704 | 0.4938 | 0.6173 |

See Also nbinpdf | nbincdf | nbininv | nbinfit | nbinrnd
How To . "Negative Binomial Distribution" on page B-72

Purpose Noncentral $F$ cumulative distribution function

$$
\text { Syntax } \quad P=\operatorname{ncfcdf}(X, N U 1, N U 2, D E L T A)
$$

Description $\quad \mathrm{P}=\operatorname{ncfcdf}(\mathrm{X}, \mathrm{NU} 1, \mathrm{NU} 2, \mathrm{DELTA})$ computes the noncentral $F$ cdf at each of the values in $X$ using the corresponding numerator degrees of freedom in NU1, denominator degrees of freedom in NU2, and positive noncentrality parameters in DELTA. NU1, NU2, and DELTA can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of P. A scalar input for X, NU1, NU2, or DELTA is expanded to a constant array with the same dimensions as the other inputs.

The noncentral $F \mathrm{cdf}$ is

$$
F\left(x \mid v_{1}, v_{2}, \delta\right)=\sum_{j=0}^{\infty}\left(\frac{\left(\frac{1}{2} \delta\right)^{j}}{j!} e^{\frac{-\delta}{2}}\right) Y\left(\left.\frac{v_{1} \cdot x}{v_{2}+v_{1} \cdot x} \right\rvert\, \frac{v_{1}}{2}+j, \frac{v_{2}}{2}\right)
$$

where $I(x \mid a, b)$ is the incomplete beta function with parameters $a$ and $b$.

## Examples

Compare the noncentral $F$ cdf with $\delta=10$ to the $F$ cdf with the same number of numerator and denominator degrees of freedom (5 and 20 respectively).

```
x = (0.01:0.1:10.01)';
p1 = ncfcdf(x,5,20,10);
p = fcdf(x,5,20);
plot(x,p,'-',x,p1,'-')
```



## References <br> [1] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 189-200.

See Also cdf | ncfpdf | ncfinv | ncfstat | ncfrnd
How To . "Noncentral F Distribution" on page B-78

## Purpose $\quad$ Noncentral $F$ inverse cumulative distribution function

$$
\text { Syntax } \quad X=\operatorname{ncfinv}(P, N U 1, N U 2, D E L T A)
$$

Description

Examples
$\mathrm{X}=$ ncfinv( $\mathrm{P}, \mathrm{NU} 1, \mathrm{NU} 2, \mathrm{DELTA})$ returns the inverse of the noncentral $F$ cdf with numerator degrees of freedom NU1, denominator degrees of freedom NU2, and positive noncentrality parameter DELTA for the corresponding probabilities in P. P, NU1, NU2, and DELTA can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of X . A scalar input for $\mathrm{P}, \mathrm{NU1}, \mathrm{NU}$, or DELTA is expanded to a constant array with the same dimensions as the other inputs.

One hypothesis test for comparing two sample variances is to take their ratio and compare it to an $F$ distribution. If the numerator and denominator degrees of freedom are 5 and 20 respectively, then you reject the hypothesis that the first variance is equal to the second variance if their ratio is less than that computed below.

```
critical = finv(0.95,5,20)
critical =
    2.7109
```

Suppose the truth is that the first variance is twice as big as the second variance. How likely is it that you would detect this difference?

```
prob = 1 - ncfcdf(critical,5,20,2)
prob =
    0.1297
```

If the true ratio of variances is 2 , what is the typical (median) value you would expect for the $F$ statistic?

```
ncfinv(0.5,5,20,2)
ans =
    1.2786
```

References

[1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions.
Hoboken, NJ: Wiley-Interscience, 2000.
[2] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 189-200.
See Also icdf | ncfcdf | ncfpdf | ncfstat | ncfrnd
How To . "Noncentral F Distribution" on page B-78

## ncfpdf

Purpose Noncentral $F$ probability density function

```
Syntax \(\quad Y=\operatorname{ncfpdf}(X, N U 1, N U 2, D E L T A)\)
```

Description

## Examples

## References

See Also

Compare the noncentral $F$ pdf with $\delta=10$ to the F pdf with the same number of numerator and denominator degrees of freedom (5 and 20 respectively).

```
    x = (0.01:0.1:10.01)';
    p1 = ncfpdf(x,5,20,10);
    p = fpdf(x,5,20);
    plot(x,p,'-',x,p1,'-')
```


[1] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 189-200.
pdf | ncfcdf | ncfinv | ncfstat | ncfrnd

How To . "Noncentral F Distribution" on page B-78

## Purpose $\quad$ Noncentral $F$ random numbers

```
Syntax R = ncfrnd(NU1,NU2,DELTA)
R = ncfrnd(NU1,NU2,DELTA,m,n,...)
R = ncfrnd(NU1,NU2,DELTA,[m,n,...])
```


## Description

## Examples

## References

$R=n c f r n d(N U 1, N U 2, D E L T A)$ returns a matrix of random numbers chosen from the noncentral $F$ distribution with corresponding numerator degrees of freedom in NU1, denominator degrees of freedom in NU2, and positive noncentrality parameters in DELTA. NU1, NU2, and DELTA can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of R. A scalar input for NU1, NU2, or DELTA is expanded to a constant matrix with the same dimensions as the other inputs.

R = ncfrnd(NU1,NU2,DELTA, m,n,...) or R = ncfrnd(NU1,NU2, DELTA, $[m, n, \ldots]$ ) generates an m-by-n-by-... array. The NU1, NU2, DELTA parameters can each be scalars or arrays of the same size as R.

Compute six random numbers from a noncentral $F$ distribution with 10 numerator degrees of freedom, 100 denominator degrees of freedom and a noncentrality parameter, $\delta$, of 4.0 . Compare this to the $F$ distribution with the same degrees of freedom.

```
r = ncfrnd(10,100,4,1,6)
r =
    2.5995 0.8824 0.8220}101.4485 1.4415 1.4864
r1 = frnd(10,100,1,6)
r1 =
    0.9826 0.5911 1.0967 0.9681 2.0096 0.6598
```

[1] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 189-200.

```
See Also random | ncfpdf | ncfcdf | ncfinv | ncfstat
How To . "Noncentral F Distribution" on page B-78
```

Purpose Noncentral $F$ mean and variance

## Syntax $\quad[M, V]=\operatorname{ncfstat}(N U 1, N U 2, D E L T A)$

Description
[M, V] = ncfstat(NU1, NU2, DELTA) returns the mean of and variance for the noncentral $F$ pdf with corresponding numerator degrees of freedom in NU1, denominator degrees of freedom in NU2, and positive noncentrality parameters in DELTA. NU1, NU2, and DELTA can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of M and V . A scalar input for NU1, NU2, or DELTA is expanded to a constant array with the same dimensions as the other input.

The mean of the noncentral F distribution with parameters $v_{1}, v_{2}$, and $\delta$ is

$$
\frac{v_{2}\left(\delta+v_{1}\right)}{v_{1}\left(v_{2}-2\right)}
$$

where $\mathrm{v}_{2}>2$.
The variance is

$$
2\left(\frac{v_{2}}{v_{1}}\right)^{2}\left[\frac{\left(\delta+v_{1}\right)^{2}+\left(2 \delta+v_{1}\right)\left(v_{2}-2\right)}{\left(v_{2}-2\right)^{2}\left(v_{2}-4\right)}\right]
$$

where $v_{2}>4$.

## Examples

References

```
[m,v]= ncfstat(10, 100, 4)
m \(=\)
    1.4286
\(\mathrm{v}=\)
    0.4252
```

[1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions. 2nd ed., Hoboken, NJ: John Wiley \& Sons, Inc., 1993, pp. 73-74.
[2] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 189-200.

See Also ncfpdf | ncfcdf | ncfinv | ncfrnd<br>How To . "Noncentral F Distribution" on page B-78

## NaiveBayes.NClasses property

## Purpose Number of classes

## Description

The NClasses property specifies the number of classes in the grouping variable used to create the Naive Bayes classifier.

## gmdistribution.NComponents property

Purpose Number $k$ of mixture components
Description The number $k$ of mixture components.

## Purpose Noncentral $t$ cumulative distribution function

$$
\text { Syntax } \quad P=\operatorname{nctcdf}(X, N U, D E L T A)
$$

Description

Examples
$\mathrm{P}=\operatorname{nctcdf}(\mathrm{X}, \mathrm{NU}, \mathrm{DELTA})$ computes the noncentral $t$ cdf at each of the values in X using the corresponding degrees of freedom in NU and noncentrality parameters in DELTA. X, NU, and DELTA can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of P. A scalar input for X, NU, or DELTA is expanded to a constant array with the same dimensions as the other inputs.

Compare the noncentral $t \mathrm{cdf}$ with DELTA $=1$ to the $t \mathrm{cdf}$ with the same number of degrees of freedom (10).

```
x = (-5:0.1:5)';
p1 = nctcdf(x,10,1);
p = tcdf(x,10);
plot(x,p,'-',x,p1,'-')
```



References

See Also cdf | nctpdf | nctinv | nctstat | nctrnd

How To . "Noncentral t Distribution" on page B-80

## Purpose Noncentral $t$ inverse cumulative distribution function

$$
\text { Syntax } \quad X=\operatorname{nctinv}(P, N U, D E L T A)
$$

Description $\quad X=\operatorname{nctinv}(P, N U, D E L T A)$ returns the inverse of the noncentral $t$ cdf with NU degrees of freedom and noncentrality parameter DELTA for the corresponding probabilities in P. P, NU, and DELTA can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of X . A scalar input for P , NU, or DELTA is expanded to a constant array with the same dimensions as the other inputs.

## Examples

## References

## See Also

How To

```
x = nctinv([0.1 0.2],10,1)
x =
    -0.2914 0.1618
```

[1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions. 2nd ed., Hoboken, NJ: John Wiley \& Sons, Inc., 1993, pp. 147-148.
[2] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 201-219.

```
icdf | nctcdf | nctpdf | nctstat | nctrnd
```

- "Noncentral t Distribution" on page B-80


## Purpose Noncentral $t$ probability density function

## Syntax $\quad Y=\operatorname{nctpdf}(X, V, D E L T A)$

Description $\quad Y=\operatorname{nctpdf}(X, V, D E L T A)$ computes the noncentral $t$ pdf at each of the values in $X$ using the corresponding degrees of freedom in $V$ and noncentrality parameters in DELTA. Vector or matrix inputs for $X, V$, and DELTA must have the same size, which is also the size of $Y$. A scalar input for $\mathrm{X}, \mathrm{V}$, or DELTA is expanded to a constant matrix with the same dimensions as the other inputs.

Examples Compare the noncentral $t$ pdf with DELTA $=1$ to the $t \mathrm{pdf}$ with the same number of degrees of freedom (10):

```
x = (-5:0.1:5)';
nct = nctpdf(x,10,1);
t = tpdf(x,10);
plot(x,nct,'b-','LineWidth', 2)
hold on
plot(x,t,'g--','LineWidth',2)
legend('nct','t')
```



[^13]See Also pdf | nctcdf | nctinv | nctstat | nctrnd
How To . "Noncentral t Distribution" on page B-80

| Purpose | Noncentral $t$ random numbers |
| :---: | :---: |
| Syntax | $\begin{aligned} & \mathrm{R}=\mathrm{nctrnd}(\mathrm{~V}, \mathrm{DELTA}) \\ & \mathrm{R}=\mathrm{nctrnd}(\mathrm{~V}, \mathrm{DELTA}, \mathrm{~m}, \mathrm{n}, \ldots) \\ & \mathrm{R}=\mathrm{nctrnd}(\mathrm{~V}, \mathrm{DELTA},[\mathrm{~m}, \mathrm{n}, \ldots]) \end{aligned}$ |
| Description | $R=$ nctrnd $(V, D E L T A)$ returns a matrix of random numbers chosen from the noncentral T distribution using the corresponding degrees of freedom in $V$ and noncentrality parameters in DELTA. $V$ and DELTA can be vectors, matrices, or multidimensional arrays. A scalar input for $V$ or DELTA is expanded to a constant array with the same dimensions as the other input. <br> $R=n c t r n d(V, D E L T A, m, n, \ldots)$ or $R=n c t r n d(V, D E L T A,[m, n, \ldots])$ generates an m-by-n-by-... array. The V, DELTA parameters can each be scalars or arrays of the same size as R. |
| Examples | $\begin{aligned} & \text { nctrnd }(10,1,5,1) \\ & \text { ans }= \\ & 1.6576 \\ & 1.0617 \\ & 1.4491 \\ & 0.2930 \\ & 3.6297 \end{aligned}$ |
| References | [1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions. 2nd ed., Hoboken, NJ: John Wiley \& Sons, Inc., 1993, pp. 147-148. <br> [2] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 201-219. |
| See Also | random \| nctpdf | nctcdf | nctinv | nctstat |
| How To | - "Noncentral t Distribution" on page B-80 |

## Purpose Noncentral $t$ mean and variance

## Syntax $\quad[M, V]=$ nctstat $(N U, D E L T A)$

## Description

$[M, V]=n c t s t a t(N U, D E L T A)$ returns the mean of and variance for the noncentral $t$ pdf with NU degrees of freedom and noncentrality parameter DELTA. NU and DELTA can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of M and V . A scalar input for NU or DELTA is expanded to a constant array with the same dimensions as the other input.

The mean of the noncentral t distribution with parameters $v$ and $\delta$ is

$$
\frac{\delta(v / 2)^{1 / 2} \Gamma((v-1) / 2)}{\Gamma(v / 2)}
$$

where $v>1$.
The variance is

$$
\frac{v}{(v-2)}\left(1+\delta^{2}\right)-\frac{v}{2} \delta^{2}\left[\frac{\Gamma((v-1) / 2)}{\Gamma(v / 2)}\right]^{2}
$$

where $v>2$.

## Examples

$$
\begin{aligned}
& {[m, v]=\operatorname{nctstat}(10,1)} \\
& m= \\
& 1.0837 \\
& v= \\
& 1.3255
\end{aligned}
$$

## References

[1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions. 2nd ed., Hoboken, NJ: John Wiley \& Sons, Inc., 1993, pp. 147-148.
[2] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 201-219.

| See Also | nctpdf \| nctcdf | nctinv | nctrnd |
| :--- | :--- |
| How To | - "Noncentral t Distribution" on page B- 80 |

Purpose Noncentral chi-square cumulative distribution function

$$
\text { Syntax } \quad P=\operatorname{ncx} 2 c d f(X, V, D E L T A)
$$

Description $\quad P=n c x 2 c d f(X, V, D E L T A)$ computes the noncentral chi-square cdf at each of the values in $X$ using the corresponding degrees of freedom in $V$ and positive noncentrality parameters in DELTA. $\mathrm{X}, \mathrm{V}$, and DELTA can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of P. A scalar input for X, V, or DELTA is expanded to a constant array with the same dimensions as the other inputs.

Some texts refer to this distribution as the generalized Rayleigh, Rayleigh-Rice, or Rice distribution.

The noncentral chi-square cdf is

$$
F(x \mid v, \delta)=\sum_{j=0}^{\infty}\left(\frac{\left(\frac{1}{2} \delta\right)^{j}}{j!} e^{\frac{-\delta}{2}}\right) \operatorname{Pr}\left[\chi_{v+2 j}^{2} \leq x\right]
$$

## Examples Compare the noncentral chi-square cdf with DELTA $=2$ to the

 chi-square cdf with the same number of degrees of freedom (4):```
x = (0:0.1:10)';
ncx2 = ncx2cdf(x,4,2);
chi2 = chi2cdf(x,4);
plot(x,ncx2,'b-','LineWidth',2)
hold on
plot(x,chi2,'g--','LineWidth',2)
legend('ncx2','chi2','Location','NW')
```



[^14]See Also
cdf | ncx2pdf | ncx2inv | ncx2stat | ncx2rnd
How To . "Noncentral Chi-Square Distribution" on page B-76

## ncx2inv

## Purpose Noncentral chi-square inverse cumulative distribution function

$$
\text { Syntax } \quad x=\operatorname{ncx} 2 \operatorname{inv}(P, v, D E L T A)
$$

Description

Algorithms
Examples

## References

See Also
How To - "Noncentral Chi-Square Distribution" on page B-76

## Purpose Noncentral chi-square probability density function

$$
\text { Syntax } \quad Y=\operatorname{ncx} 2 p d f(X, V, D E L T A)
$$

Description $\quad Y=n c x 2 p d f(X, V, D E L T A)$ computes the noncentral chi-square pdf at each of the values in X using the corresponding degrees of freedom in V and positive noncentrality parameters in DELTA. Vector or matrix inputs for $\mathrm{X}, \mathrm{V}$, and DELTA must have the same size, which is also the size of Y . A scalar input for $\mathrm{X}, \mathrm{V}$, or DELTA is expanded to a constant array with the same dimensions as the other inputs.

Some texts refer to this distribution as the generalized Rayleigh, Rayleigh-Rice, or Rice distribution.

## Examples

Compare the noncentral chi-square pdf with DELTA $=2$ to the chi-square pdf with the same number of degrees of freedom (4):

```
x = (0:0.1:10)';
ncx2 = ncx2pdf(x,4,2);
chi2 = chi2pdf(x,4);
plot(x,ncx2,'b-','LineWidth', 2)
hold on
plot(x,chi2,'g--','LineWidth',2)
legend('ncx2','chi2')
```


References $\quad$ [1] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous
Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc.,
1970, pp. 130-148.

See Also pdf | ncx2cdf | ncx2inv | ncx2stat | ncx2rnd
How To

- "Noncentral Chi-Square Distribution" on page B-76

| Purpose | Noncentral chi-square random numbers |
| :---: | :---: |
| Syntax | $\begin{aligned} & R=n c x 2 r n d(V, D E L T A) \\ & R=n c x 2 r n d(V, D E L T A, m, n, \ldots) \\ & R=n c x 2 r n d(V, D E L T A,[m, n, \ldots]) \end{aligned}$ |
| Description | $R=n c x 2 r n d(V, D E L T A)$ returns a matrix of random numbers chosen from the noncentral chi-square distribution using the corresponding degrees of freedom in $V$ and positive noncentrality parameters in DELTA. V and DELTA can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of R. A scalar input for $V$ or DELTA is expanded to a constant array with the same dimensions as the other input. |
|  | $R=n c x 2 r n d(V, D E L T A, m, n, \ldots) \quad$ or $R=$ ncx2rnd(V, DELTA, $[m, n, \ldots])$ generates an $m$-by-n-by-... array. The $V$, DELTA parameters can each be scalars or arrays of the same size as R. |
| Examples | $\begin{aligned} & \text { ncx2rnd }(4,2,6,3) \\ & \text { ans }= \end{aligned}$ |
|  | $6.8552 \quad 5.965011 .2961$ |
|  | $5.2631 \quad 4.2640 \quad 5.9495$ |
|  | $9.1939 \quad 6.71623 .8315$ |
|  | $10.3100 \quad 4.4828 \quad 7.1653$ |
|  | 2.11421 .98264 .6400 |
|  | $3.8852 \quad 5.39990 .9282$ |
| References | [1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions. 2nd ed., Hoboken, NJ: John Wiley \& Sons, Inc., 1993, pp. 50-52. |
|  | [2] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 130-148. |
| See Also | random \| ncx2pdf | ncx2cdf | ncx2inv | ncx2stat |
| How To | - "Noncentral Chi-Square Distribution" on page B-76 |

## Purpose Noncentral chi-square mean and variance

$$
\text { Syntax } \quad[M, V]=\text { ncx2stat }(N U, D E L T A)
$$

Description

Examples

References

See Also
How To

- "Noncentral Chi-Square Distribution" on page B-76

Purpose Number of dimensions of categorical array
Syntax $\quad n=\operatorname{ndims}(A)$
Description $\quad n=$ ndims $(A)$ returns the number of dimensions in the categorical array $A$. The number of dimensions in an array is always greater than or equal to 2 . Trailing singleton dimensions are ignored. Put simply, ndims(A) is length(size(A)).

See Also size

Purpose Dimension $d$ of multivariate Gaussian distributions

Description The dimension $d$ of the multivariate Gaussian distributions.
Purpose Number of dimensions of dataset array
Syntax $\mathrm{n}=$ ndims( A$)$
Description $\mathrm{n}=$ ndims $(\mathrm{A})$ returns the number of dimensions in the dataset $A$. Thenumber of dimensions in an array is always 2 .
See Also ..... size

## qrandset.ndims

Purpose Number of dimensions in matrix

## Syntax $\quad n=\operatorname{ndims}(p)$

Description $n=$ ndims $(p)$ returns the number of dimensions in the matrix that is created by the $\operatorname{syntax} p(:,:)$. Since this is always a 2 -D matrix, $n$ is always equal to 2 .

See Also qrandset | size

## Purpose Number of dimensions

Description The NDims property specifies the number of dimensions, which is equal to the number of features in the training data used to create the Naive Bayes classifier.

## qrandstream.ne

Purpose Not equal relation for handles

## Syntax <br> h1 ~= h2

Description Handles are equal if they are handles for the same object and are unequal otherwise.
h1 ~= h2 performs element-wise comparisons between handle arrays h 1 and h2. h1 and h2 must be of the same dimensions unless one is a scalar. The result is a logical array of the same dimensions, where each element is an element-wise $\sim=$ result.

If one of h 1 or h 2 is scalar, scalar expansion is performed and the result will match the dimensions of the array that is not scalar.
$\mathrm{tf}=\mathrm{ne}(\mathrm{h} 1, \mathrm{~h} 2)$ stores the result in a logical array of the same dimensions.

See Also qrandstream | eq | ge | gt | le | lt
Purpose

Nearest neighbor search object

## Construction

NeighborSearcher is an abstract class. You cannot create instances of this class directly. You can construct an object in a subclass, such as KDTreeSearcher or ExhaustiveSearcher, either by calling the subclass constructors or by using the createns function.

## Properties

## x

A matrix used to create the object.

## Distance

A string specifying a built-in distance metric (applies to both ExhaustiveSearcher and KDTreeSearcher) or a function handle (only applies to ExhaustiveSearcher) that you provide when you create the object. This property is the default distance metric used when you call the knnsearch method to find nearest neighbors for future query points.

## DistParameter

Specifies the additional parameter for the chosen distance metric. The value is:

- If 'Distance' is 'minkowski': A positive scalar indicating the exponent of the Minkowski distance. (Applies for both ExhaustiveSearcher and KDTreeSearcher.)
- If 'Distance' is 'mahalanobis': A positive definite matrix representing the covariance matrix used for computing the Mahalanobis distance. (Only applies for ExhaustiveSearcher.)


## NeighborSearcher

- If 'Distance' is 'seuclidean': A vector representing the scale value of the data when computing the 'seuclidean' distance. (Only applies for ExhaustiveSearcher.)
- Otherwise: Empty.

See Also createns | KDTreeSearcher | ExhaustiveSearcher

## Purpose

Generate quasi-random point set

## Syntax <br> $X=\operatorname{net}(p, n)$

$X=\operatorname{net}(p, n)$ returns the first $n$ points $X$ from the point set $p$ of the qrandset class. X is n -by- $d$, where $d$ is the dimension of the point set.

Objects p of the @qrandset class encapsulate properties of a specified quasi-random sequence. Values of the point set are not generated and stored in memory until $p$ is accessed using net or parenthesis indexing.

## Examples

Use haltonset to generate a 3-D Halton point set, skip the first 1000 values, and then retain every 101st point:

```
p = haltonset(3,'Skip',1e3,'Leap',1e2)
p =
    Halton point set in 3 dimensions (8.918019e+013 points)
    Properties:
                            Skip : }100
            Leap : 100
    ScrambleMethod : none
```

Use scramble to apply reverse-radix scrambling:

```
p = scramble(p,'RR2')
p =
    Halton point set in 3 dimensions (8.918019e+013 points)
    Properties:
            Skip : 1000
            Leap : 100
    ScrambleMethod : RR2
```

Use net to generate the first four points:

```
x0 = net(p,4)
X0 =
    0.0928 0.6950 0.0029
    0.6958 0.2958 0.8269
```


## qrandset.net

| 0.3013 | 0.6497 | 0.4141 |
| :--- | :--- | :--- |
| 0.9087 | 0.7883 | 0.2166 |

Use parenthesis indexing to generate every third point, up to the 11th point:

```
X = p(1:3:11,:)
X =
        0.0928 0.6950 0.0029
        0.9087 0.7883 0.2166
        0.3843 0.9840 0.9878
        0.6831 0.7357 0.7923
```

See Also
haltonset | sobolset | qrandstream

## Purpose <br> Syntax <br> Description

Tips
Nonlinear regression

```
beta = nlinfit(X,y,fun,beta0)
```

beta = nlinfit(X,y,fun,beta0)
[beta,r] = nlinfit(X,y,fun,betaO)
[beta,r] = nlinfit(X,y,fun,betaO)
[beta,r,J] = nlinfit(X,y,fun,beta0)
[beta,r,J] = nlinfit(X,y,fun,beta0)
[beta,r,J,COVB] = nlinfit(X,y,fun,betaO)
[beta,r,J,COVB] = nlinfit(X,y,fun,betaO)
[beta,r,J,COVB,mse] = nlinfit(X,y,fun,betaO)
[beta,r,J,COVB,mse] = nlinfit(X,y,fun,betaO)
[beta,...] = nlinfit(X,y,fun,beta0,options)

```
[beta,...] = nlinfit(X,y,fun,beta0,options)
```

beta $=$ nlinfit( $X, y, f u n$, betaO) returns a vector beta of coefficient estimates for a nonlinear regression of the responses in $y$ on the predictors in X using the model specified by fun.
[beta, $r$ ] = nlinfit( $X, y, f u n, b e t a 0)$ returns the residuals of the fit.
[beta, $\mathrm{r}, \mathrm{J}]=$ nlinfit( $\mathrm{X}, \mathrm{y}, \mathrm{fun}$, beta0 $)$ returns the Jacobian of fun.
[beta,r,J,COVB] = nlinfit(X,y,fun,betaO) returns an estimated covariance matrix of the fitted coefficients.
[beta,r,J,COVB,mse] = nlinfit(X,y,fun,betaO) returns an estimated variance of the error term.
[beta,...] = nlinfit(X,y,fun,beta0,options) specifies control parameters for the nlinfit algorithm.

- To produce error estimates on predictions, use nlpredci with its optional outputs r, J, COVB, or mse.
- To produce error estimates on the estimated coefficients beta, use nlparci with its optional outputs r, J, COVB, or mse.
- If you use the robust fitting option (see options in "Input Arguments" on page 20-1180), you must use COVB and may need mse as input to nlpredci or nlparci to ensure that the confidence intervals take the robust fit properly into account.


## nlinfit

## Input Arguments

Typically, X is a design matrix of predictor (independent variable) values, with one row for each value in $Y$ and one column for each coefficient. However, $X$ can be any array that fun accepts.
y
Vector of response (dependent variable) values.
nlinfit treats NaNs in $y$ as missing data and ignores the corresponding rows.
fun
Function specified using @ that accepts two arguments:

- A coefficient vector, say b
- The array X

For example, fun = @modelfun, where modelfun.m is a file such as:

```
function yhat = modelfun(b,X)
```

fun( $\mathrm{b}, \mathrm{X}$ ) returns a vector of fitted response values, matching Y as closely as possible in the least-squares sense.
nlinfit treats NaNs in fun(b, X) as missing data and ignores the corresponding rows.
beta0
Vector containing initial coefficient values. The number of elements in beta0 equals the number of elements in the beta output. A poor value of beta0 can lead to a poor solution beta, meaning one with large residual error.

Structure you can create using statset. Applicable statset parameters:

- 'DerivStep' - Relative difference used in finite difference gradient calculation. May be a scalar, or a vector the same size as beta0. The default is eps^(1/3).
- 'Display ' - Level of display output during estimation. The choices are
- 'off' (default) - Displays no information
- 'final' - Displays information after the final iteration
- 'iter' - Displays information at each iteration
- 'FunValCheck' - Check for invalid values, such as NaN or Inf, from the objective function. Values are 'off' or 'on' (the default).
- 'MaxIter' - Maximum number of iterations allowed. The default is 100 .
- 'Robust' - Invoke robust fitting option. Values are 'off' (the default) or 'on'.
- 'TolFun' - Termination tolerance on the residual sum of squares. The default is $1 \mathrm{e}-8$.
- 'TolX' - Termination tolerance on the estimated coefficients beta. The default is $1 \mathrm{e}-8$.
- 'Tune' - The tuning constant used in robust fitting to normalize the residuals before applying the weight function. The value is a positive scalar, with the default value dependent on the weight function. This parameter is required if the weight function is specified as a function handle.
- 'WgtFun' - A weight function for robust fitting. Valid only when 'Robust' is 'on'. It can be 'bisquare' (the default), 'andrews', 'cauchy', 'fair', 'huber', 'logistic', 'talwar', or 'welsch'. It can also be a function handle that


## nlinfit

accepts a normalized residual as input and returns the robust weights as output.

## Output Arguments

## beta

Vector of coefficients. beta minimizes the sum of the squares of $y$ - fun(beta, $x$ ).
$r$
Residual vector y - fun(beta, X).
J
Estimated Jacobian of fun with respect to beta.

## COVB

Estimated covariance matrix of the coefficients beta.
mse
Mean squared error norm( $r)^{\wedge} 2 /(n-p)$, where $n$ is the number of elements in $r$, and $p$ is the number of parameters (the length of beta).

## Examples Suppose you have data, and want to fit a model of the form

$$
y_{i}=a_{1}+a_{2} \exp \left(-a_{3} x_{i}\right)+\varepsilon_{i} .
$$

Here the $a_{i}$ are the parameters you want to estimate, $x_{i}$ are the data points, the $y_{i}$ are the responses, and the $\varepsilon_{i}$ are noise terms.

1 Write a function handle that represents the model:

```
mdl = @(a,x)(a(1) + a(2)*exp(-a(3)*x));
```

2 Generate synthetic data with parameters $a=[1 ; 3 ; 2]$, with the $x$ data points distributed exponentially with parameter 2 , and normally distributed noise with standard deviation 0.1:

```
rng(9845,'twister') % for reproducibility
a = [1;3;2];
x = exprnd(2,100,1);
epsn = normrnd(0,0.1,100,1);
y = mdl(a,x) + epsn;
```

3 Fit the model to data starting from the arbitrary guess a0 = [2;2;2]:

```
aO = [2;2;2];
[ahat,r,J,cov,mse] = nlinfit(x,y,mdl,aO);
ahat
ahat =
    1.0153
    3.0229
    2.1070
```

4 Examine the fit:

```
xrange = min(x):.01:max(x);
hold on
scatter(x,y)
plot(xrange,mdl(ahat,xrange),'r')
hold off
```



5 Check whether [1;3;2] is in a $95 \%$ confidence interval using nlparci:

```
ci = nlparci(ahat,r,'Jacobian',J)
ci =
    0.9869 1.0438
    2.9401 3.1058
    1.9963 2.2177
```


## Algorithms

nlinfit uses the Levenberg-Marquardt algorithm [1] for nonlinear least squares to compute non-robust fits.

For robust fits, nlinfit uses an algorithm [2] that iteratively refits a weighted nonlinear regression, where the weights at each iteration are based on each observation's residual from the previous iteration. These weights serve to downweight points that are outliers so that their
influence on the fit is decreased. Iterations continue until the weights converge. The computation of weights is identical to the iterative reweighting scheme used by robustfit for fitting a robust linear model, and in particular, the possible weighting functions and their tuning parameters are the same. The nonlinear regression at each iteration is a weighted version of the Levenberg-Marquardt algorithm that nlinfit uses for non-robust fits.

## References [1] Seber, G. A. F., and C. J. Wild. Nonlinear Regression. Hoboken, NJ: Wiley-Interscience, 2003.

[2] DuMouchel, W. H., and F. L. O'Brien. "Integrating a Robust Option into a Multiple Regression Computing Environment." Computer Science and Statistics: Proceedings of the 21st Symposium on the Interface. Alexandria, VA: American Statistical Association, 1989.
[3] Holland, P. W., and R. E. Welsch. "Robust Regression Using Iteratively Reweighted Least-Squares." Communications in Statistics: Theory and Methods, A6, 1977, pp. 813-827.

## Alternatives <br> nlintool is a graphical user interface to nlinfit. <br> See Also <br> nlparci | nlpredci | nlintool <br> Tutorials . "Nonlinear Regression" on page 9-72

Purpose Interactive nonlinear regression
$\begin{array}{ll}\text { Syntax } & \text { nlintool }(X, y, f u n, \text { beta0 }) \\ & \text { nlintool }(X, y, f u n, \text { beta0, alpha }) \\ & n l i n t o o l(X, y, f u n, b e t a 0, \text { alpha, 'xname' ' 'yname' })\end{array}$

## Description

## Examples

nlintool ( $X, y$, fun, beta0) is a graphical user interface to the nlinfit function, and uses the same input arguments. The interface displays plots of the fitted response against each predictor, with the other predictors held fixed. The fixed values are in the text boxes below each predictor axis. Change the fixed values by typing in a new value or by dragging the vertical lines in the plots to new positions. When you change the value of a predictor, all plots update to display the model at the new point in predictor space. Dashed red curves show $95 \%$ simultaneous confidence bands for the function.
nlintool (X,y,fun, beta0, alpha) shows 100(1-alpha) \% confidence bands. These are simultaneous confidence bounds for the function value. Using the Bounds menu you can switch between simultaneous and non-simultaneous bounds, and between bounds on the function and bounds for predicting a new observation.
nlintool(X,y,fun, beta0, alpha, 'xname', 'yname') labels the plots using the string matrix 'xname' for the predictors and the string 'yname' for the response.

The data in reaction.mat are partial pressures of three chemical reactants and the corresponding reaction rates. The function hougen implements the nonlinear Hougen-Watson model for reaction rates. The following fits the model to the data:

```
load reaction
nlintool(reactants,rate,@hougen, beta,0.01,xn,yn)
```



See Also nlinfit | polytool | rstool

## nlmefit

## Purpose Nonlinear mixed-effects estimation

```
Syntax beta = nlmefit(X,y,group,V,fun,beta0)
[beta,PSI] = nlmefit(X,y,group,V,fun,betaO)
[beta,PSI,stats] = nlmefit(X,y,group,V,fun,betaO)
[beta,PSI,stats,B] = nlmefit(X,y,group,V,fun,beta0)
[beta,PSI,stats,B] = nlmefit(X,y,group,V,fun,beta0,'Name',
    value)
```


## Description

beta $=$ nlmefit( $X, y, g r o u p, V, f u n, b e t a 0)$ fits a nonlinear mixed-effects regression model and returns estimates of the fixed effects in beta. By default, nlmefit fits a model in which each parameter is the sum of a fixed and a random effect, and the random effects are uncorrelated (their covariance matrix is diagonal).

X is an $n$-by- $h$ matrix of $n$ observations on $h$ predictors.
y is an $n$-by- 1 vector of responses.
group is a grouping variable indicating $m$ groups in the observations. group is a categorical variable, a numeric vector, a character matrix with rows for group names, or a cell array of strings. For more information on grouping variables, see "Grouped Data" on page 2-34.

V is an $m$-by- $g$ matrix or cell array of $g$ group-specific predictors. These are predictors that take the same value for all observations in a group. The rows of $V$ are assigned to groups using grp2idx, according to the order specified by grp2idx (group). Use a cell array for $V$ if group predictors vary in size across groups. Use [] for V if there are no group-specific predictors.
fun is a handle to a function that accepts predictor values and model parameters and returns fitted values. fun has the form
yfit = modelfun(PHI,XFUN,VFUN)

The arguments are:

- PHI — A 1-by- $p$ vector of model parameters.
- XFUN - A $k$-by- $h$ array of predictors, where:
- $k=1$ if XFUN is a single row of X.
- $k=n_{i}$ if XFUN contains the rows of X for a single group of size $n_{i}$.
- $k=n$ if XFUN contains all rows of X .
- VFUN - Group-specific predictors given by one of:
- A 1-by-g vector corresponding to a single group and a single row of V .
- An $n$-by-g array, where the $j$ th row is $\mathrm{V}(\mathrm{I},:)$ if the $j$ th observation is in group $I$.

If $V$ is empty, nlmefit calls modelfun with only two inputs.

- yfit - A $k$-by-1 vector of fitted values

When either PHI or VFUN contains a single row, it corresponds to all rows in the other two input arguments.

Note If modelfun can compute yfit for more than one vector of model parameters per call, use the 'Vectorization' parameter (described later) for improved performance.
beta0 is a $q$-by- 1 vector with initial estimates for $q$ fixed effects. By default, $q$ is the number of model parameters $p$.
nlmefit fits the model by maximizing an approximation to the marginal likelihood with random effects integrated out, assuming that:

- Random effects are multivariate normally distributed and independent between groups.
- Observation errors are independent, identically normally distributed, and independent of the random effects.
[beta, PSI] = nlmefit(X, y, group, $\mathrm{V}, \mathrm{fun}$, beta0 $)$ also returns PSI, an $r$-by- $r$ estimated covariance matrix for the random effects. By default, $r$ is equal to the number of model parameters $p$.
[beta, PSI, stats] = nlmefit (X,y,group, V,fun, beta0) also returns stats, a structure with fields:
- dfe - The error degrees of freedom for the model
- logl - The maximized log-likelihood for the fitted model
- rmse - The square root of the estimated error variance (computed on the log scale for the exponential error model)
- errorparam - The estimated parameters of the error variance model
- aic - The Akaike information criterion, calculated as aic $=-2 * \operatorname{logl}$ +2 * numParam, where numParam is the number of fitting parameters, including the degree of freedom for covariance matrix of the random effects, the number of fixed effects and the number of parameters of the error model, and logl is a field in the stats structure
- bic - The Bayesian information criterion, calculated as bic $=$ $-2 * \log l+\log (M)$ * numParam
- $M$ is the number of groups.
- numparam and logl are defined as in aic.

Note that some literature suggests that the computation of bic should be, bic $=-2^{*} \log l+\log (N)$ * numParam, where $N$ is the number of observations.

- covb - The estimated covariance matrix of the parameter estimates
- sebeta - The standard errors for beta
- ires - The population residuals (y-y_population), where y_population is the individual predicted values
- pres - The population residuals ( $\mathrm{y}-\mathrm{y}$ _population), where y_population is the population predicted values
- iwres - The individual weighted residuals
- pwres - The population weighted residuals
- cwres - The conditional weighted residuals
[beta,PSI, stats, $B]=$ nlmefit( $X, y$, group, V, fun, beta0 $)$ also returns B , an $r$-by- $m$ matrix of estimated random effects for the $m$ groups. By default, $r$ is equal to the number of model parameters $p$.
[beta,PSI,stats,B] =
nlmefit(X,y,group,V,fun, beta0, 'Name', value) specifies one or more optional parameter name/value pairs. Specify Name inside single quotes.

Use the following parameters to fit a model different from the default. (The default model is obtained by setting both FEConstDesign and REConstDesign to eye (p), or by setting both FEParamsSelect and REParamsSelect to 1:p.) Use at most one parameter with an 'FE' prefix and one parameter with an 'RE' prefix. The nlmefit function requires you to specify at least one fixed effect and one random effect.

| Parameter | Value |
| :--- | :--- |
| FEParamsSelect | A vector specifying which elements of the <br> parameter vector PHI include a fixed effect, <br> given as a numeric vector of indices from <br> 1 to $p$ or as a 1-by- $p$ logical vector. If $q$ is <br> the specified number of elements, then the <br> model includes $q$ fixed effects. |
| FEConstDesign | A $p$-by- $q$ design matrix ADESIGN, where <br> ADESIGN*beta are the fixed components of <br> the $p$ elements of PHI. |
| FEGroupDesign | A $p$-by- $q$-by- $m$ array specifying a different <br> $p-$ by- $q$ fixed-effects design matrix for each <br> of the $m$ groups. |
| FEObsDesign | A $p$-by- $q$-by- $n$ array specifying a different <br> $p$-by- $q$ fixed-effects design matrix for each <br> of the $n$ observations. |


| Parameter | Value |
| :--- | :--- |
| REParamsSelect | A vector specifying which elements of the <br> parameter vector PHI include a random <br> effect, given as a numeric vector of indices <br> from 1 to $p$ or as a 1-by- $p$ logical vector. The <br> model includes $r$ random effects, where $r$ <br> is the specified number of elements. |
| REConstDesign | A $p$-by- $r$ design matrix BDESIGN, where <br> BDESIGN*B are the random components of <br> the $p$ elements of PHI. |
| REGroupDesign | A $p$-by- $r$-by- $m$ array specifying a different <br> $p-$ by- random-effects design matrix for <br> each of $m$ groups. |
| REObsDesign | A $p$-by- $r$-by- $n$ array specifying a different <br> $p-$ by- $r$ random-effects design matrix for <br> each of $n$ observations. |

Use the following parameters to control the iterative algorithm for maximizing the likelihood:

| Parameter | Value |
| :--- | :--- |
| RefineBeta0 | Determines whether nlmefit makes an <br> initial refinement of beta0 by first fitting <br> modelfun without random effects and <br> replacing beta0 with beta. Choices are <br> 'on' and 'off' ' The default value is 'on ' |
| ErrorModel | A string specifying the form of the error <br> term. Default is 'constant '. Each model <br> defines the error using a standard normal <br> (Gaussian) variable $e$, the function value <br> $f$, and one or two parameters $a$ and $b$. <br> Choices are: |
| - 'constant ': $y=f+a^{*} e$ |  |


| Parameter | Value |
| :---: | :---: |
|  | - 'proportional': $y=f+b^{*} f^{*} e$ <br> - 'combined': $y=f+\left(a+b^{*} f\right) * e$ <br> - 'exponential': $y=f^{*} \exp \left(a^{*} e\right)$, or equivalently $\log (y)=\log (f)+a^{*} e$ <br> If this parameter is given, the output stats.errorparam field has the value <br> - $a$ for 'constant ' and 'exponential' <br> - $\quad b$ for 'proportional' <br> - $[a b]$ for 'combined' |
| ApproximationType | The method used to approximate the likelihood of the model. Choices are: <br> - 'LME' - Use the likelihood for the linear mixed-effects model at the current conditional estimates of beta and B. This is the default. <br> - 'RELME ' - Use the restricted likelihood for the linear mixed-effects model at the current conditional estimates of beta and $B$. <br> - 'FO' - First-order Laplacian approximation without random effects. <br> - 'FOCE' - First-order Laplacian approximation at the conditional estimates of B. |

## nlmefit

| Parameter | Value |
| :--- | :--- |
| Vectorization | Indicates acceptable sizes for the PHI, <br> XFUN, and VFUN input arguments to <br> modelfun. Choices are: |
|  | - SinglePhi' - modelfun can only <br> accept a single set of model parameters <br> at a time, so PHI must be a single <br> row vector in each call. nlmefit calls <br> modelfun in a loop, if necessary, with |
| a single PHI vector and with XFUN |  |
| containing rows for a single observation |  |
| or group at a time. VFUN may be a single |  |
| row that applies to all rows of XFUN, or a |  |
| matrix with rows corresponding to rows |  |
| in XFUN. This is the default. |  |
| - 'SingleGroup ' - modelfun can only |  |
| accept inputs corresponding to a single |  |
| group in the data, so XFUN must contain |  |
| rows of X from a single group in each |  |
| call. Depending on the model, PHI is |  |
| a single row that applies to the entire |  |
| group or a matrix with one row for each |  |
| observation. VFUN is a single row. |  |


| Parameter | Value |
| :--- | :--- |
| CovParameterization | Specifies the parameterization used <br> internally for the scaled covariance matrix. <br> Choices are 'chol' for the Cholesky <br> factorization or 'logm' the matrix <br> logarithm. The default is 'logm'. |
| CovPattern | Specifies an $r$-by- $r$ logical or numeric <br> matrix P that defines the pattern of the <br> random-effects covariance matrix PSI. <br> nlmefit estimates the variances along <br> the diagonal of PSI and the covariances <br> specified by nonzeroes in the off-diagonal <br> elements of P. Covariances corresponding <br> to zero off-diagonal elements in P are <br> constrained to be zero. If P does not specify <br> a row-column permutation of a block <br> diagonal matrix, nlmefit adds nonzero <br> elements to P as needed. The default <br> value of P is eye (r), corresponding to <br> uncorrelated random effects. |
|  | Alternatively, P may be a 1-by- $r$ vector <br> containing values in 1:r, with equal values <br> specifying groups of random effects. In this <br> case, nlmefit estimates covariances only <br> within groups, and constrains covariances <br> across groups to be zero. |

## nlmefit

| Parameter | Value |
| :---: | :---: |
| ParamTransform | A vector of P values specifying a transformation function $f()$ for each of the $P$ parameters: $X B=$ ADESIGN*BETA + BDESIGN*B PHI $=f(\mathrm{XB})$ Each element of the vector must be one of the following integer codes specifying the transformation for the corresponding value of PHI : <br> - $0: \mathrm{PHI}=\mathrm{XB}$ (default for all parameters) <br> - $1: \log (\mathrm{PHI})=X B$ <br> - 2: $\operatorname{probit}(\mathrm{PHI})=X B$ <br> - $3: \operatorname{logit}(\mathrm{PHI})=X B$ |
| Options | A structure of the form returned by statset. nlmefit uses the following statset parameters: <br> - 'DerivStep' - Relative difference used in finite difference gradient calculation. May be a scalar, or a vector whose length is the number of model parameters $p$. The default is eps^(1/3). <br> - 'Display ' - Level of iterative display during estimation. Choices are: <br> - 'off' (default) - Displays no information <br> - 'final' - Displays information after the final iteration <br> - 'iter' - Displays information at each iteration <br> - 'FunValCheck' - Check for invalid values, such as NaN or Inf, from |


| Parameter | Value |
| :--- | :--- |
|  | modelfun. Choices are 'on' and 'off '. <br> The default is 'on '. |
|  | - MaxIter' - Maximum number of <br> iterations allowed. The default is 200. <br> - 'OutputFcn' — Function handle <br> specified using @, a cell array with <br> function handles or an empty array <br> (default). The solver calls all output <br> functions after each iteration. |
| - 'TolFun' - Termination tolerance on |  |
| the log-likelihood function. The default |  |
| is 1e-4. |  |

Examples
Display data on the growth of five orange trees:

```
CIRC = [30 58 87 115 120 142 145;
    33 69 111 156 172 203 203;
    30 51 75 108 115 139 140;
    32 62 112 167 179 209 214;
    30 49 81 125 142 174 177];
time = [118 484 664 1004 1231 1372 1582];
h = plot(time,CIRC','o','LineWidth',2);
```


## nlmefit

```
xlabel('Time (days)')
ylabel('Circumference (mm)')
title('{\bf Orange Tree Growth}')
legend([repmat('Tree ',5,1),num2str((1:5)')],...
    Location','NW')
grid on
hold on
```



Use an anonymous function to specify a logistic growth model:

```
model=@(PHI,t)(PHI(:,1))./(1+exp(-(t-PHI(:,2))./PHI(:,3)));
```

Fit the model using nlmefit with default settings (that is, assuming each parameter is the sum of a fixed and a random effect, with no correlation among the random effects):

```
TIME = repmat(time,5,1);
NUMS = repmat((1:5)',size(time));
beta0 = [100 100 100];
[beta1,PSI1,stats1] = nlmefit(TIME(:),CIRC(:),NUMS(:),...
    [],model,beta0)
beta1 =
    191.3189
    723.7608
    346.2517
PSI1 =
    962.1534 0
    0 0.0000 0
    0 297.9881
stats1 =
            dfe: 28
            logl: -131.5457
                mse: 59.7882
            rmse: 7.9016
        errorparam: 7.7323
            aic: 277.0913
            bic: 274.3574
            covb: [3x3 double]
            sebeta: [15.2249 33.1579 26.8235]
            ires: [35x1 double]
            pres: [35\times1 double]
            iwres: [35\times1 double]
            pwres: [35\times1 double]
            cwres: [35\times1 double]
```


## nlmefit

The negligible variance of the second random effect, PSI1 $(2,2)$, suggests that it can be removed to simplify the model:

```
[beta2,PSI2,stats2,b2] = nlmefit(TIME(:),CIRC(:),...
    NUMS(:),[],model,beta0,'REParamsSelect',[1 3])
beta2 =
    191.3194
    723.7628
    346.2548
PSI2 =
    962.2114 0
            0 298.3989
stats2 =
                dfe: 29
            logl: -131.5456
                mse: 59.7851
            rmse: 7.7640
            errorparam: 7.7321
                aic: 275.0913
                bic: 272.7479
            covb: [3\times3 double]
            sebeta: [15.2252 33.1572 26.8246]
                    ires: [35x1 double]
                    pres: [35x1 double]
            iwres: [35x1 double]
            pwres: [35x1 double]
            cwres: [35x1 double]
b2 =
    -28.5250 31.6063 -36.5070 39.0735 -5.6479
    10.0097 -0.7638 6.0117 -9.4685 -5.7892
```

The log-likelihood logl is unaffected, and both the Akaike and Bayesian information criteria (aic and bic) are reduced, supporting the decision to drop the second random effect from the model.
Use the estimated fixed effects in beta2 and the estimated random effects for each tree in b2 to plot the model through the data:

```
PHI = repmat(beta2,1,5) + ... % Fixed effects
    [b2(1,:);zeros(1,5);b2(2,:)]; % Random effects
colors = get(h,'Color');
tplot = 0:0.1:1600;
for I = 1:5
    fitted_model=@(t)(PHI(1,I))./(1+exp(-(t-PHI(2,I))./ ...
PHI(3,I)));
    plot(tplot,fitted_model(tplot),'Color',}\operatorname{colors{I}, ...
        'LineWidth',2)
end
```


## nlmefit



## References

[1] Lindstrom, M. J., and D. M. Bates. "Nonlinear mixed-effects models for repeated measures data." Biometrics. Vol. 46, 1990, pp. 673-687.
[2] Davidian, M., and D. M. Giltinan. Nonlinear Models for Repeated Measurements Data. New York: Chapman \& Hall, 1995.
[3] Pinheiro, J. C., and D. M. Bates. "Approximations to the log-likelihood function in the nonlinear mixed-effects model." Journal of Computational and Graphical Statistics. Vol. 4, 1995, pp. 12-35.
[4] Demidenko, E. Mixed Models: Theory and Applications. Hoboken, NJ: John Wiley \& Sons, Inc., 2004.

## See Also

How To
nlinfit | nlpredci | nlmefitsa

- "Grouped Data" on page 2-34


## nlmefitsa

Description

## Input

Arguments

Purpose Fit nonlinear mixed effects model with stochastic EM algorithm

```
Syntax
```

[BETA,PSI,STATS,B] = nlmefitsa(X,Y,GROUP,V,MODELFUN,BETAO)

```
[BETA,PSI,STATS,B] = nlmefitsa(X,Y,GROUP,V,MODELFUN,BETAO)
[BETA,PSI,STATS,B] = nlmefitsa(X,Y,GROUP,V,MODELFUN,BETAO,
[BETA,PSI,STATS,B] = nlmefitsa(X,Y,GROUP,V,MODELFUN,BETAO,
    'Name',Value)
```

```
    'Name',Value)
```

```
\([B E T A, P S I, S T A T S, B]=\) nlmefitsa(X,Y,GROUP, V, MODELFUN,BETAO) fits a nonlinear mixed-effects regression model and returns estimates of the fixed effects in BETA. By default, nlmefitsa fits a model where each model parameter is the sum of a corresponding fixed and random effect, and the covariance matrix of the random effects is diagonal, i.e., uncorrelated random effects.

The BETA, PSI, and other values this function returns are the result of a random (Monte Carlo) simulation designed to converge to the maximum likelihood estimates of the parameters. Because the results are random, it is advisable to examine the plot of simulation to results to be sure that the simulation has converged. It may also be helpful to run the function multiple times, using multiple starting values, or use the 'Replicates' parameter to perform multiple simulations.
[BETA,PSI,STATS,B] = nlmefitsa(X, Y, GROUP, V, MODELFUN, BETAO, 'Name', Value) accepts one or more comma-separated parameter name/value pairs. Specify Name inside single quotes.

\section*{Definitions:}

In the following list of arguments, the following variable definitions apply:
- \(n\) - number of observations
- \(h\) - number of predictor variables
- \(m\) - number of groups
- \(g\) - number of group-specific predictor variables
- \(p\) - number of parameters
- \(f\) - number of fixed effects

X
An \(n\)-by- \(h\) matrix of \(n\) observations on \(h\) predictor variables.

\section*{Y}

An \(n\)-by- 1 vector of responses.
GROUP
A grouping variable indicating to which of \(m\) groups each observation belongs. GROUP can be a categorical variable, a numeric vector, a character matrix with rows for group names, or a cell array of strings.

\section*{V}

An \(m\)-by- \(g\) matrix of \(g\) group-specific predictor variables for each of the \(m\) groups in the data. These are predictor values that take on the same value for all observations in a group. Rows of V are ordered according to GRP2IDX (GROUP). Use an \(m\)-by- \(g\) cell array for \(V\) if any of the group-specific predictor values vary in size across groups. Specify [ ] for V if there are no group predictors.

\section*{MODELFUN}

A handle to a function that accepts predictor values and model parameters, and returns fitted values. MODELFUN has the form YFIT = MODELFUN(PHI,XFUN,VFUN) with input arguments
- PHI - A 1-by- \(p\) vector of model parameters.
- XFUN - An \(l\)-by- \(h\) array of predictor variables where
- \(l\) is 1 if XFUN is a single row of \(X\)
- \(l\) is \(n_{i}\) if XFUN contains the rows of X for a single group of size \(n_{i}\)
- \(l\) is \(n\) if XFUN contains all rows of X .
- VFUN - Either

\section*{nlmefitsa}
- A 1-by-g vector of group-specific predictors for a single group, corresponding to a single row of V
- An \(n\)-by- \(g\) matrix, where the \(k\)-th row of VFUN is \(\mathrm{V}(i\), :) if the \(k\)-th observation is in group \(i\).
If \(V\) is empty, nlmefitsa calls MODELFUN with only two inputs.
MODELFUN returns an \(l\)-by- 1 vector of fitted values YFIT. When either PHI or VFUN contains a single row, that one row corresponds to all rows in the other two input arguments. For improved performance, use the 'Vectorization' parameter name/value pair (described below) if MODELFUN can compute YFIT for more than one vector of model parameters in one call.

\section*{BETAO}

An \(f\)-by- 1 vector with initial estimates for the \(f\) fixed effects. By default, \(f\) is equal to the number of model parameters \(p\). BETAO can also be an \(f\)-by-REPS matrix, and the estimation is repeated REPS times using each column of BETAO as a set of starting values.

\section*{Name-Value Pair Arguments}

By default, nlmefitsa fits a model where each model parameter is the sum of a corresponding fixed and random effect. Use the following parameter name/value pairs to fit a model with a different number of or dependence on fixed or random effects. Use at most one parameter name with an 'FE' prefix and one parameter name with an 'RE' prefix. Note that some choices change the way nlmefitsa calls MODELFUN, as described further below.
```

FEParamsSelect

```

A vector specifying which elements of the model parameter vector PHI include a fixed effect, as a numeric vector with elements in \(1: p\), or as a 1-by-p logical vector. The model will include \(f\) fixed effects, where \(f\) is the specified number of elements.

\footnotetext{
FEConstDesign
}

A \(p\)-by- \(f\) design matrix ADESIGN, where ADESIGN*BETA are the fixed components of the \(p\) elements of PHI.

\section*{FEGroupDesign}

A \(p\)-by- \(f\)-by- \(m\) array specifying a different \(p\)-by- \(f\) fixed effects design matrix for each of the \(m\) groups.

\section*{REParamsSelect}

A vector specifying which elements of the model parameter vector PHI include a random effect, as a numeric vector with elements in \(1: p\), or as a 1-by- \(p\) logical vector. The model will include \(r\) random effects, where \(r\) is the specified number of elements.

\section*{REConstDesign}

A \(p\)-by- \(r\) design matrix BDESIGN, where BDESIGN*B are the random components of the \(p\) elements of PHI. This matrix must consist of 0 s and 1 s , with at most one 1 per row.

The default model is equivalent to setting both FEConstDesign and REConstDesign to eye ( \(p\) ), or to setting both FEParamsSelect and REParamsSelect to \(1: p\).

Additional optional parameter name/value pairs control the iterative algorithm used to maximize the likelihood:

\section*{CovPattern}

Specifies an \(r\)-by- \(r\) logical or numeric matrix PAT that defines the pattern of the random effects covariance matrix PSI. nlmefitsa computes estimates for the variances along the diagonal of PSI as well as covariances that correspond to non-zeroes in the off-diagonal of PAT. nlmefitsa constrains the remaining covariances, i.e., those corresponding to off-diagonal zeroes in PAT, to be zero. PAT must be a row-column permutation of a block diagonal matrix, and nlmefitsa adds non-zero elements to PAT as needed to produce such a pattern. The default value of PAT is eye ( \(r\) ), corresponding to uncorrelated random effects.

\section*{nlmefitsa}

Alternatively, specify PAT as a 1-by- \(r\) vector containing values in \(1: r\). In this case, elements of PAT with equal values define groups of random effects, nlmefitsa estimates covariances only within groups, and constrains covariances across groups to be zero.

\section*{Covo}

Initial value for the covariance matrix PSI. Must be an \(r\)-by- \(r\) positive definite matrix. If empty, the default value depends on the values of BETAO.

\section*{ComputeStdErrors}
true to compute standard errors for the coefficient estimates and store them in the output STATS structure, or false (default) to omit this computation.

\section*{ErrorModel}

A string specifying the form of the error term. Default is 'constant'. Each model defines the error using a standard normal (Gaussian) variable \(e\), the function value \(f\), and one or two parameters \(a\) and \(b\). Choices are
- 'constant' \(-y=f+a^{*} e\)
- 'proportional' - \(y=f+b^{*} f^{*} e\)
- 'combined' \(-y=f+\left(a+b^{*} f\right)^{*} e\)
- 'exponential' - \(y=f^{*} \exp \left(a^{*} e\right)\), or equivalently \(\log (y)=\log (f)\) \(+a^{*} e\)

If this parameter is given, the output STATS.rmse field has the value
- \(a\) for 'constant' and 'exponential'
- \(b\) for 'proportional'
- \([a b]\) for 'combined'

\section*{ErrorParameters}

A scalar or two-element vector specifying starting values for parameters of the error model. This specifies the \(a, b\), or \([a b\) ] values depending on the ErrorModel parameter.

\section*{LogLikMethod}

Specifies the method for approximating the log likelihood. Choices are:
- 'is' - Importance sampling
- 'gq' - Gaussian quadrature
- 'lin' - Linearization
- 'none ' - Omit the log likelihood approximation (default)

\section*{NBurnIn}

Number of initial burn-in iterations during which the parameter estimates are not recomputed. Default is 5 .

NChains
Number \(c\) of "chains" simulated. Default is 1 . Setting \(c>1\) causes \(c\) simulated coefficient vectors to be computed for each group during each iteration. Default depends on the data, and is chosen to provide about 100 groups across all chains.

\section*{NIterations}

Number of iterations. This can be a scalar or a three-element vector. Controls how many iterations are performed for each of three phases of the algorithm:

1 simulated annealing
2 full step size
3 reduced step size

\section*{nlmefitsa}

Default is [150 150 100]. A scalar is distributed across the three phases in the same proportions as the default.

\section*{NMCMCIterations}

Number of Markov Chain Monte Carlo (MCMC) iterations. This can be a scalar or a three-element vector. Controls how many of three different types of MCMC updates are performed during each phase of the main iteration:

1 full multivariate update
2 single coordinate update
3 multiple coordinate update
Default is [2 2 2]. A scalar value is treated as a three-element vector with all elements equal to the scalar.

\section*{OptimFun}

Either 'fminsearch' or 'fminunc', specifying the optimization function to be used during the estimation process. Default is 'fminsearch'. Use of 'fminunc' requires Optimization Toolbox.

\section*{Options}

A structure created by a call to statset. nlmefitsa uses the following statset parameters:
- 'DerivStep' - Relative difference used in finite difference gradient calculation. May be a scalar, or a vector whose length is the number of model parameters \(p\). The default is eps^(1/3).
- Display - Level of display during estimation.
- 'off' (default) - Displays no information
- 'final' - Displays information after the final iteration of the estimation algorithm
- 'iter' - Displays information at each iteration
- FunValCheck
- 'on' (sdefault) — Check for invalid values (such as NaN or Inf) from MODELFUN
- 'off' - Skip this check
- OutputFcn - Function handle specified using @, a cell array with function handles or an empty array. nlmefitsa calls all output functions after each iteration. See nlmefitoutputfon.m (the default output function for nlmefitsa) for an example of an output function.

\section*{ParamTransform}

A vector of \(p\) values specifying a transformation function f() for each of the \(p\) parameters:
```

XB = ADESIGN*BETA + BDESIGN*B
PHI = f(XB)

```

Each element of the vector must be one of the following integer codes specifying the transformation for the corresponding value of PHI:
- 0: PHI = XB (default for all parameters)
- 1: \(\log (\mathrm{PHI})=\mathrm{XB}\)
- 2: probit(PHI) = XB
- 3: logit(PHI) = XB

\section*{Replicates}

Number REPS of estimations to perform starting from the starting values in the vector BETAO. If BETAO is a matrix, REPS must match the number of columns in BETAO. Default is the number of columns in BETAO.
```

Vectorization

```

\section*{nlmefitsa}

Determines the possible sizes of the PHI, XFUN, and VFUN input arguments to MODELFUN. Possible values are:
- 'SinglePhi' - MODELFUN is a function (such as an ODE solver) that can only compute YFIT for a single set of model parameters at a time, i.e., PHI must be a single row vector in each call. nlmefitsa calls MODELFUN in a loop if necessary using a single PHI vector and with XFUN containing rows for a single observation or group at a time. VFUN may be a single row that applies to all rows of XFUN, or a matrix with rows corresponding to rows in XFUN.
- 'SingleGroup' - MODELFUN can only accept inputs corresponding to a single group in the data, i.e., XFUN must contain rows of X from a single group in each call. Depending on the model, PHI is a single row that applies to the entire group, or a matrix with one row for each observation. VFUN is a single row.
- 'Full' - MODELFUN can accept inputs for multiple parameter vectors and multiple groups in the data. Either PHI or VFUN may be a single row that applies to all rows of XFUN, or a matrix with rows corresponding to rows in XFUN. Using this option can improve performance by reducing the number of calls to MODELFUN, but may require MODELFUN to perform singleton expansion on PHI or V.

The default for 'Vectorization' is 'SinglePhi'. In all cases, if V is empty, nlmefitsa calls MODELFUN with only two inputs.

\section*{Output \\ BETA \\ Arguments \\ Estimates of the fixed effects}

An \(r\)-by- \(r\) estimated covariance matrix for the random effects. By default, \(r\) is equal to the number of model parameters \(p\).

A structure with the following fields:
- logl - The maximized log-likelihood for the fitted model; empty if the LogLikMethod parameter has its default value of 'none'
- rmse - The square root of the estimated error variance (computed on the log scale for the exponential error model)
- errorparam - The estimated parameters of the error variance model
- aic - The Akaike information criterion (empty if logl is empty), calculated as aic \(=-2 * \operatorname{logl}+2 *\) numParam, where
- logl is the maximized log-likelihood.
- numParam is the number of fitting parameters, including the degree of freedom for covariance matrix of the random effects, the number of fixed effects and the number of parameters of the error model.
- bic - The Bayesian information criterion (empty if logl is empty), calculated as bic \(=-2 * \log 1+\log (M) *\) numParam
- \(M\) is the number of groups.
- logl and numParam are defined as in aic.

Note that some literature suggests that the computation of bic should be, bic \(=-2^{*} \operatorname{logl}+\log (\mathrm{N})\) * numParam, where N is the number of observations. To adjust the value of the output you can redefine bic as follows: bic = bic - numel(unique(group)) + numel(Y)
- sebeta - The standard errors for BETA (empty if the ComputeStdErrors parameter has its default value of false)
- covb - The estimated covariance of the parameter estimates (empty if ComputeStdErrors is false)
- dfe - The error degrees of freedom

\section*{nlmefitsa}
- pres - The population residuals (y-y_population), where y_population is the population predicted values
- ires - The population residuals (y-y_population), where y_population is the individual predicted values
- pwres - The population weighted residuals
- cwres - The conditional weighted residuals
- iwres - The individual weighted residuals

\section*{Examples}

Fit a model to data on concentrations of the drug indomethacin in the bloodstream of six subjects over eight hours:
```

load indomethacin
model = @(phi,t)(phi(:,1).*exp(-phi(:,2).*t)+phi(:,3).*exp(-phi(:,4).*t));
phiO = [$$
\begin{array}{llll}{1}&{1}&{1}&{1}\end{array}
$$];
% log transform for 2nd and 4th parameters
xform = [0 1 0 1];
[beta,PSI,stats,br] = nlmefitsa(time,concentration,...
subject,[],model,phiO,'ParamTransform',xform)

```

\% Plot the data along with an overall "population" fit clf
```

phi = [beta(1), exp(beta(2)), beta(3), exp(beta(4))];

```
h = gscatter(time, concentration, subject);
xlabel('Time (hours)')
ylabel('Concentration (mcg/ml)')
title('\{\bf Indomethacin Elimination\}')
\(\mathrm{xx}=\operatorname{linspace}(0,8)\);
line(xx,model(phi, \(\left.x x), ' l i n e w i d t h ', 2, ' c o l o r^{\prime}, ' k^{\prime}\right)\)

\section*{nlmefitsa}
```

% Plot individual curves based on random effect estimates
for j=1:6
phir = [beta(1)+br(1,j), exp(beta(2)+br(2,j)), ...
beta(3)+br(3,j), exp(beta(4)+br(4,j))];
line(xx,model(phir,xx),'color',get(h(j),'color'))
end

```

Indomethacin Elimination


Algorithms
In order to estimate the parameters of a nonlinear mixed effects model, we would like to choose the parameter values that maximize a
likelihood function. These values are called the maximum likelihood estimates. The likelihood function can be written in the form
\[
p\left(y \mid \beta, \sigma^{2}, \Sigma\right)=\int p\left(y \mid \beta, b, \sigma^{2}\right) p(b \mid \Sigma) d b
\]
where
- \(y\) is the response data
- \(B\) is the vector of population coefficients
- \(\sigma^{2}\) is the residual variance
- \(\sum\) is the covariance matrix for the random effects
- \(b\) is the set of unobserved random effects

Each \(p()\) function on the right-hand-side is a normal (Gaussian) likelihood function that may depend on covariates.
Since the integral does not have a closed form, it is difficult to find parameters that maximize it. Delyon, Lavielle, and Moulines [1] proposed to find the maximum likelihood estimates using an Expectation-Maximization (EM) algorithm in which the E step is replaced by a stochastic procedure. They called their algorithm SAEM, for Stochastic Approximation EM. They demonstrated that this algorithm has desirable theoretical properties, including convergence under practical conditions and convergence to a local maximum of the likelihood function. Their proposal involves three steps:

1 Simulation: Generate simulated values of the random effects \(b\) from the posterior density \(p(b \mid \Sigma)\) given the current parameter estimates.

2 Stochastic approximation: Update the expected value of the log likelihood function by taking its value from the previous step, and moving part way toward the average value of the log likelihood calculated from the simulated random effects.

\section*{nlmefitsa}

3 Maximization step: Choose new parameter estimates to maximize the log likelihood function given the simulated values of the random effects.

\author{
References \\ [1] Delyon, B., M. Lavielle, and E. Moulines, Convergence of a stochastic approximation version of the EM algorithm, Annals of Statistics, 27, 94-128, 1999.
}
[2] Mentré, France, and Marc Lavielle, Stochastic EM algorithms in population PKPD analyses, 2008.

See Also nlinfit | nlmefit

\section*{gmdistribution.NlogL property}

\author{
Purpose Negative of log-likelihood \\ Description The negative of the log-likelihood of the data. \\ Note This property applies only to gmdistribution objects constructed with fit.
}

\section*{ProbDistParametric.NLogL property}
\begin{tabular}{ll} 
Purpose & \begin{tabular}{l} 
Read-only value specifying negative log likelihood for input data to \\
ProbDistParametric object
\end{tabular} \\
Description & \begin{tabular}{l} 
NLogL is a read-only property of the ProbDistParametric class. NLogL \\
is a value specifying the negative log likelihood for input data used to fit \\
a distribution represented by a ProbDistParametric object.
\end{tabular} \\
Values & \begin{tabular}{l} 
The value is a numeric scalar for a distribution fit to input data, that \\
is, a distribution created using the fitdist function. This property is \\
empty for distributions created without fitting to data, that is, by using \\
the ProbDistUnivParam. ProbDistUnivParam constructor. Use this \\
information to view and compare the negative log likelihood for input \\
data supplied to create distributions.
\end{tabular}
\end{tabular}

\section*{ProbDistUnivKernel.NLogL property}
\begin{tabular}{ll} 
Purpose & \begin{tabular}{l} 
Read-only value specifying negative log likelihood for input data to \\
ProbDistUnivKernel object
\end{tabular} \\
Description \(\quad\)\begin{tabular}{l} 
NLogL is a read-only property of the ProbDistUnivKernel class. NLogL \\
is a value specifying the negative log likelihood for input data used to fit \\
a distribution represented by a ProbDistUnivKernel object.
\end{tabular} \\
Values & \begin{tabular}{l} 
The value is a numeric scalar for a distribution fit to input data, that is, \\
a distribution created using the fitdist function. Use this information \\
to view and compare the negative log likelihood for input data used to \\
create distributions.
\end{tabular}
\end{tabular}

Purpose Nonlinear regression parameter confidence intervals
```

Syntax ci = nlparci(beta,resid,'covar',sigma)
ci = nlparci(beta,resid,'jacobian',J)
ci = nlparci(...,'alpha',alpha)

```

\section*{Description}

\section*{Examples}
ci = nlparci(beta, resid, 'covar', sigma) returns the \(95 \%\) confidence intervals ci for the nonlinear least squares parameter estimates beta. Before calling nlparci, use nlinfit to fit a nonlinear regression model and get the coefficient estimates beta, residuals resid, and estimated coefficient covariance matrix sigma.
ci = nlparci(beta, resid,'jacobian', J) is an alternative syntax that also computes \(95 \%\) confidence intervals. \(J\) is the Jacobian computed by nlinfit. If the 'robust' option is used with nlinfit, use the 'covar' input rather than the 'jacobian' input so that the required sigma parameter takes the robust fitting into account.
ci = nlparci(...,'alpha', alpha) returns 100(1-alpha)\% confidence intervals.
nlparci treats \(N a N s\) in resid or \(J\) as missing values, and ignores the corresponding observations.

The confidence interval calculation is valid for systems where the length of resid exceeds the length of beta and \(J\) has full column rank. When \(J\) is ill-conditioned, confidence intervals may be inaccurate.

\section*{Fit to exponential decay}

Suppose you have data, and want to fit a model of the form
\[
y_{i}=a_{1}+a_{2} \exp \left(-a_{3} x_{i}\right)+\varepsilon_{i} .
\]

Here the \(a_{i}\) are the parameters you want to estimate, \(x_{i}\) are the data points, the \(y_{i}\) are the responses, and the \(\varepsilon_{i}\) are noise terms.

1 Write a function handle that represents the model:
\[
m d l=a(a, x)(a(1)+a(2) * \exp (-a(3) * x)) ;
\]

2 Generate synthetic data with parameters \(a=[1 ; 3 ; 2]\), with the \(x\) data points distributed exponentially with parameter 2 , and normally distributed noise with standard deviation 0.1:
```

rng(9845,'twister') % for reproducibility
a = [1;3;2];
x = exprnd(2,100,1);
epsn = normrnd(0,0.1,100,1);
y = mdl(a,x) + epsn;

```

3 Fit the model to data starting from the arbitrary guess a0 \(=[2 ; 2 ; 2]\) :
```

a0 = [2;2;2];
[ahat,r,J,cov,mse] = nlinfit(x,y,mdl,aO);
ahat
ahat =
1.0153
3.0229
2.1070

```

4 Check whether [ \(1 ; 3 ; 2\) ] is in a \(95 \%\) confidence interval using the Jacobian argument in nlparci:
```

ci = nlparci(ahat,r,'Jacobian',J)
ci =
0.9869 1.0438
2.9401 3.1058
1.9963 2.2177

```

5 You can obtain the same result using the covariance argument:
```

ci = nlparci(ahat,r,'covar',cov)
ci =
0.9869 1.0438

```
\(\begin{array}{ll}2.9401 & 3.1058 \\ 1.9963 & 2.2177\end{array}\)

See Also nlinfit | nlpredci

\section*{Purpose}

Syntax

Nonlinear regression prediction confidence intervals
```

[ypred,delta] = nlpredci(modelfun,x,beta,resid,'covar',sigma)
[ypred,delta] = nlpredci(modelfun,x,beta,resid,'jacobian',J)
[...] = nlpredci(...,param1,val1,param2,val2,...)

```
[ypred,delta] =
nlpredci(modelfun, x,beta, resid,'covar',sigma) returns
predictions, ypred, and \(95 \%\) confidence interval half-widths, delta, for
the nonlinear regression model defined by modelfun, at input values \(x\).
modelfun is a function handle, specified using @, that accepts two
arguments-a coefficient vector and the array \(x\)-and returns a vector
of fitted y values. Before calling nlpredci, use nlinfit to fit modelfun
by nonlinear least squares and get estimated coefficient values beta,
residuals resid, and estimated coefficient covariance matrix sigma.
[ypred,delta] =
nlpredci(modelfun, \(x\), beta, resid,'jacobian', J) is an alternative
syntax that also computes \(95 \%\) confidence intervals. \(J\) is the Jacobian
computed by nlinfit. If the 'robust' option is used with nlinfit, use
the 'covar' input rather than the 'jacobian' input so that the
required sigma parameter takes the robust fitting into account.
[...] = nlpredci(...,param1,val1,param2,val2,...) accepts
optional parameter name/value pairs.
\begin{tabular}{l|l}
\hline Parameter & Value \\
\hline 'alpha' & \begin{tabular}{l} 
A value between 0 and 1 that specifies the confidence \\
level as 100(1-alpha)\%. Default is 0.05.
\end{tabular} \\
\hline 'mse ' & \begin{tabular}{l} 
The mean squared error returned by nlinfit. This is \\
required to predict new observations (see 'predopt ') if \\
the robust option is used with nlinfit; otherwise, the
\end{tabular} \\
\begin{tabular}{l} 
'mse' is computed from the residuals and does not take \\
the robust fitting into account.
\end{tabular} \\
\hline
\end{tabular}

\section*{nlpredci}
\begin{tabular}{l|l}
\hline Parameter & Value \\
\hline 'predopt' & \begin{tabular}{l} 
Either ' curve ' (the default) to compute confidence \\
intervals for the estimated curve (function value) at \\
x, or 'observation' for prediction intervals for a \\
new observation at x. If 'observation 'is specified \\
after using a robust option with nlinfit, the 'mse' \\
parameter must be supplied to specify the robust \\
estimate of the mean squared error.
\end{tabular} \\
\hline 'simopt' & \begin{tabular}{l} 
Either ' on ' for simultaneous bounds, or 'off' (the \\
default) for nonsimultaneous bounds.
\end{tabular} \\
\hline
\end{tabular}
nlpredci treats NaNs in resid or \(J\) as missing values, and ignores the corresponding observations.

The confidence interval calculation is valid for systems where the length of resid exceeds the length of beta and \(J\) has full column rank at beta. When \(J\) is ill-conditioned, predictions and confidence intervals may be inaccurate.

\section*{Examples}

\section*{Fit to exponential decay}

Suppose you have data, and want to fit a model of the form
\[
y_{i}=a_{1}+a_{2} \exp \left(-a_{3} x_{i}\right)+\varepsilon_{i} .
\]

Here the \(a_{i}\) are the parameters you want to estimate, \(x_{i}\) are the data points, the \(y_{i}\) are the responses, and the \(\varepsilon_{i}\) are noise terms.

1 Write a function handle that represents the model:
\[
m d l=@(a, x)(a(1)+a(2) * \exp (-a(3) * x)) ;
\]

2 Generate synthetic data with parameters \(a=[1 ; 3 ; 2]\), with the \(x\) data points distributed exponentially with parameter 2 , and normally distributed noise with standard deviation 0.1:
```

rng(9845,'twister') % for reproducibility

```
```

a = [1;3;2];
x = exprnd(2,100,1);
epsn = normrnd(0,0.1,100,1);
y = mdl(a,x) + epsn;

```

3 Fit the model to data starting from the arbitrary guess a0 = [2;2;2]:
```

a0 = [2;2;2];
[ahat,r,J,cov,mse] = nlinfit(x,y,mdl,a0);
ahat
ahat =
1.0153
3.0229
2.1070

```

4 Find the predicted response yhat together with a \(95 \%\) confidence interval about yhat, for the mean of \(x\) :
```

[ypred dlt] = nlpredci(mdl,mean(x),ahat,r,'Covar',cov)
ypred =
1.0669
dlt =
0.0231

```

The \(95 \%\) confidence interval about ypred is:
```

[ypred-dlt,ypred+dlt]
ans =
1.0438 1.0900

```

\section*{See Also}
nlinfit | nlparci

Purpose Nonnegative matrix factorization
```

Syntax
[W, H] $=\operatorname{nnmf}(A, k)$
[W,H] = nnmf(A,k,param1,val1,param2,val2,...)
[W,H,D] = nnmf(...)

```

\section*{Description}
\([\mathrm{W}, \mathrm{H}]=\operatorname{nnmf}(\mathrm{A}, \mathrm{k})\) factors the nonnegative \(n\)-by- \(m\) matrix A into nonnegative factors W ( \(n\)-by-k) and H (k-by- \(m\) ). The factorization is not exact; \(\mathrm{W} * \mathrm{H}\) is a lower-rank approximation to A . The factors W and H are chosen to minimize the root-mean-squared residual D between \(A\) and W*H:
```

D = sqrt(norm(A-W*H,'fro')/(N*M))

```

The factorization uses an iterative method starting with random initial values for \(W\) and \(H\). Because the root-mean-squared residual D may have local minima, repeated factorizations may yield different W and H . Sometimes the algorithm converges to a solution of lower rank than \(k\), which may indicate that the result is not optimal.

W and H are normalized so that the rows of H have unit length. The columns of \(W\) are ordered by decreasing length.
\([\mathrm{W}, \mathrm{H}]=\mathrm{nnmf}(\mathrm{A}, \mathrm{k}\), param1, val1, param2, val2, ...) specifies optional parameter name/value pairs from the following table.
\begin{tabular}{l|l}
\hline Parameter & Value \\
\hline 'algorithm' & \begin{tabular}{l} 
Either 'als ' (the default) to use an alternating \\
least-squares algorithm, or 'mult ' to use a \\
multiplicative update algorithm.
\end{tabular} \\
\begin{tabular}{l} 
In general, the 'als ' algorithm converges faster \\
and more consistently. The 'mult ' algorithm is \\
more sensitive to initial values, which makes it a
\end{tabular} \\
\hline
\end{tabular}
\begin{tabular}{l|l}
\hline Parameter & Value \\
\hline
\end{tabular}
\begin{tabular}{|c|c|}
\hline & good choice when using 'replicates' to find \(W\) and H from multiple random starting values. \\
\hline 'wo ' & An \(n\)-by- \(k\) matrix to be used as the initial value for \(W\). \\
\hline 'h0' & A \(k\)-by- \(m\) matrix to be used as the initial value for H . \\
\hline 'options' & \begin{tabular}{l}
An options structure as created by the statset function. nnmf uses the following fields of the options structure: \\
- Display - Level of display. Choices: \\
- 'off' (default) - No display \\
- 'final' - Display final result \\
- 'iter' - Iterative display of intermediate results \\
- MaxIter - Maximum number of iterations. Default is 100. Unlike in optimization settings, reaching MaxIter iterations is treated as convergence. \\
- TolFun - Termination tolerance on change in size of the residual. Default is \(1 \mathrm{e}-4\). \\
- TolX - Termination tolerance on relative change in the elements of \(W\) and \(H\). Default is 1e-4. \\
- UseParallel - Set to 'always' to compute in parallel. Default is 'never'. \\
- UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set
\end{tabular} \\
\hline
\end{tabular}
\begin{tabular}{|c|c|}
\hline Parameter & Value \\
\hline & \begin{tabular}{l}
Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'. \\
- Streams - A RandStream object or cell array of such objects. If you do not specify Streams, nnmf uses the default stream or streams. If you choose to specify Streams, use a single object except in the case \\
- You have an open MATLAB pool \\
- UseParallel is 'always' \\
- UseSubstreams is 'never' \\
In that case, use a cell array the same size as the MATLAB pool. \\
For more information on using parallel computing, see Chapter 17, "Parallel Statistics".
\end{tabular} \\
\hline 'replicates' & The number of times to repeat the factorization, using new random starting values for \(W\) and H , except at the first replication if 'w0' and 'h0' are given. This is most beneficial with the 'mult' algorithm. The default is 1 . \\
\hline
\end{tabular}
\([W, H, D]=\operatorname{nnmf}(. .\).\() also returns D\), the root mean square residual.

\section*{Examples Example 1}

Compute a nonnegative rank-two approximation of the measurements of the four variables in Fisher's iris data:
```

load fisheriris
[W,H] = nnmf(meas,2);
H
H =

| 0.6852 | 0.2719 | 0.6357 | 0.2288 |
| :--- | :--- | :--- | :--- |
| 0.8011 | 0.5740 | 0.1694 | 0.0087 |

```

The first and third variables in meas (sepal length and petal length, with coefficients 0.6852 and 0.6357 , respectively) provide relatively strong weights to the first column of W . The first and second variables in meas (sepal length and sepal width, with coefficients 0.8011 and 0.5740 ) provide relatively strong weights to the second column of \(W\).

Create a biplot of the data and the variables in meas in the column space of W :
```

    biplot(H','scores',W,'varlabels',{'sl','sw','pl','pw'});
    axis([0 1.1 0 1.1])
    xlabel('Column 1')
    ylabel('Column 2')
    ```


\section*{Example 2}

Starting from a random array \(X\) with rank 20 , try a few iterations at several replicates using the multiplicative algorithm:
```

X = rand(100,20)*rand(20,50);
opt = statset('MaxIter',5,'Display','final');
[WO,HO] = nnmf(X,5,'replicates',10,...
'options',opt,...
'algorithm','mult');
rep iteration rms resid |delta x|
1 5 0.560887 0.0245182
2 5 0.66418 0.0364471
3 5 0.609125 0.0358355
4 5 0.608894 0.0415491
5 5 0.619291 0.0455135
6 5 0.621549 0.0299965
7 5 0.640549 0.0438758
8 5 0.673015 0.0366856
9 5 0.606835 0.0318931
10 5 0.633526 0.0319591
Final root mean square residual = 0.560887

```

Continue with more iterations from the best of these results using alternating least squares:
```

opt = statset('Maxiter',1000,'Display','final');
[W,H] = nnmf(X,5,'wO',WO,'hO',HO,...
'options',opt,...
algorithm','als');
rep iteration rms resid |delta x|
1 80 0.256914 9.78625e-005
Final root mean square residual = 0.256914
[1] Berry, M. W., et al. "Algorithms and Applications for Approximate Nonnegative Matrix Factorization." Computational Statistics and Data Analysis. Vol. 52, No. 1, 2007, pp. 155-173.

```

\section*{References}

See Also princomp | factoran | statset

\section*{classregtree.nodeclass}

Purpose Class values of nodes of classification tree
```

Syntax
NAME=nodeclass(T)
NAME=nodeclass(T,J)
[ NAME, ID]=nodeclass(. . .)

```

\section*{Description}

See Also

NAME=nodeclass ( \(T\) ) returns an \(n\)-element cell array with the names of the most probable classes in each node of the tree T, where \(n\) is the number of nodes in the tree. Every element of this array is a string equal to one of the class names returned by classname( \(T\) ). For regression trees, nodeclass returns an empty cell array.

NAME=nodeclass \((T, J)\) takes an array \(J\) of node numbers and returns the class names for the specified nodes.
[NAME, ID]=nodeclass(...) also returns a numeric array with the class index for each node. The class index is determined by the order of classes classname returns.
classregtree | classname | numnodes

\section*{Purpose}

Return vector of node errors
Syntax
e = nodeerr(t)
e = nodeerr(t,nodes)

Description

\section*{Examples}
e = nodeerr( t ) returns an \(n\)-element vector e of the errors of the nodes in the tree \(t\), where \(n\) is the number of nodes. For a regression tree, the error e(i) for node \(i\) is the variance of the observations assigned to node i. For a classification tree, e(i) is the misclassification probability for node i.
e = nodeerr(t, nodes) takes a vector nodes of node numbers and returns the errors for the specified nodes.

The error e is the so-called resubstitution error computed by applying the tree to the same data used to create the tree. This error is likely to under estimate the error you would find if you applied the tree to new data. The test function provides options to compute the error (or cost) using cross-validation or a test sample.

Create a classification tree for Fisher's iris data:
```

load fisheriris;
t = classregtree(meas,species,...
'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica

```

\section*{classregtree.nodeerr}
view(t)
\begin{tabular}{|c|c|c|c|c|c|c|c|}
\hline Click to display: Identity & \(\checkmark\) & Magnification: & 100\% & & Pruning level: & 0 of 4 & - \\
\hline
\end{tabular}

\[
\begin{aligned}
& e= \text { nodeer }(t) \\
& e= 0.6667 \\
& 0 \\
& 0.5000 \\
& 0.0926 \\
& 0.0217 \\
& 0.0208
\end{aligned}
\]
0.3333

0
0

\author{
References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984. \\ See Also classregtree | numnodes | test
}

\section*{classregtree.nodemean}

Purpose Mean values of nodes of regression tree
Syntax
NM = nodemean(T)
NM = nodemean( \(\mathrm{T}, \mathrm{J}\) )

Description \(\quad N M=\) nodemean \((T)\) returns an \(n\)-element numeric array with mean values in each node of the tree \(T\), where \(n\) is the number of nodes in the tree. Every element of this array is computed by averaging true Y values over all observations in the node. For classification trees, nodemean returns an empty numeric array.
\(N M=\) nodemean \((T, J)\) takes an array \(J\) of node numbers and returns the mean values for the specified nodes.

\section*{See Also \\ classregtree | numnodes}

\section*{classregtree.nodeprob}

\section*{Purpose Node probabilities}
```

Syntax
p = nodeprob(t)
p = nodeprob(t,nodes)

```
\(p=\) nodeprob( \(t\) ) returns an \(n\)-element vector \(p\) of the probabilities of the nodes in the tree \(t\), where \(n\) is the number of nodes. The probability of a node is computed as the proportion of observations from the original data that satisfy the conditions for the node. For a classification tree, this proportion is adjusted for any prior probabilities assigned to each class.
\(p\) = nodeprob(t, nodes) takes a vector nodes of node numbers and returns the probabilities for the specified nodes.

\section*{Examples Create a classification tree for Fisher's iris data:}
```

load fisheriris;
t = classregtree(meas,species,...
'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
class = virginica
view(t)

```

\section*{classregtree.nodeprob}
\begin{tabular}{|c|c|c|c|c|c|c|c|c|}
\hline & & & & & & & & - \\
\hline Click to display: & Identity & \(\checkmark\) & Magnification: & 100\% & \(\checkmark\) & Pruning level: & 0 of 4 & \\
\hline
\end{tabular}

\[
\begin{aligned}
& p=\text { nodeprob(t) } \\
& p= \\
& 1.0000 \\
& 0.3333 \\
& 0.6667 \\
& 0.3600 \\
& 0.3067 \\
& 0.3200 \\
& 0.0400 \\
& 0.3133
\end{aligned}
\]

\author{
References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.
}

\author{
See Also \\ classregtree | nodesize | numnodes
}

\section*{classregtree.nodesize}

\section*{Purpose Return node size}
```

Syntax
sizes = nodesize(t)
sizes = nodesize(t,nodes)

```
sizes \(=\) nodesize (t) returns an \(n\)-element vector sizes of the sizes of the nodes in the tree \(t\), where \(n\) is the number of nodes. The size of a node is defined as the number of observations from the data used to create the tree that satisfy the conditions for the node.
sizes \(=\) nodesize(t, nodes) takes a vector nodes of node numbers and returns the sizes for the specified nodes.

\section*{Examples \\ Create a classification tree for Fisher's iris data:}
```

load fisheriris;
t = classregtree(meas,species,...
'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
view(t)

```
\begin{tabular}{|c|c|c|c|c|c|c|c|}
\hline Click to display: Identity & \(\checkmark\) & Magnification: & 100\% & \(\checkmark\) & Pruning level: & 0 of 4 & - \\
\hline
\end{tabular}

```

sizes = nodesize(t)
sizes =
150
50
100
54
46
48
6
4 7

```

\section*{classregtree.nodesize}

\title{
References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.
}

See Also classregtree | numnodes

\section*{Purpose Notify listeners of event}
```

Syntax notify(h,'eventname')
notify(h,'eventname',data)

```

Description notify (h, 'eventname') notifies listeners added to the event named eventname on handle object array \(h\) that the event is taking place. \(h\) is the array of handles to objects triggering the event, and eventname must be a string.
notify (h, 'eventname',data) provides a way of encapsulating information about an event which can then be accessed by each registered listener. data must belong to the event. eventdata class.

\author{
See Also
}
addlistener | event.EventData | events | qrandstream

\section*{Superclasses \\ categorical}

Purpose Arrays for nominal categorical data
Description Nominal arrays are used to store discrete values that are not numeric and that do not have an ordering. A nominal array provides efficient storage and convenient manipulation of such data, while also maintaining meaningful labels for the values. Nominal arrays are often used as grouping variables.

You can subscript, concatenate, reshape, etc. nominal arrays much like ordinary numeric arrays. You can test equality between elements of two nominal arrays, or between a nominal array and a single string representing a nominal value.

\section*{Construction}

Use the nominal constructor to create a nominal array from a numeric, logical, or character array, or from a cell array of strings.
nominal
Construct nominal categorical array

Methods
Each nominal array carries along a list of possible values that it can store, known as its levels. The list is created when you create a nominal array, and you can access it using the getlevels method, or modify it using the addlevels, mergelevels, or droplevels methods. Assignment to the array will also add new levels automatically if the values assigned are not already levels of the array.

You can change the order of the list of levels for a nominal array using the reorderlevels method, however, that order has no significance for the values in the array. The order is used only for display purposes, or when you convert the nominal array to numeric values using methods such as double or subsindex, or compare two arrays using isequal. If you need to work with values that have a mathematical ordering, you should use an ordinal array instead.

\section*{Inherited Methods}

Methods in the following table are inherited from categorical.
\begin{tabular}{ll} 
addlevels & Add levels to categorical array \\
cat \\
cellstr & \begin{tabular}{l} 
Concatenate categorical arrays \\
Convert categorical array to cell \\
array of strings
\end{tabular} \\
char & \begin{tabular}{l} 
Convert categorical array to \\
character array
\end{tabular} \\
circshift & Shift categorical array circularly \\
ctranspose & Transpose categorical matrix \\
disp & \begin{tabular}{l} 
Display categorical array \\
display
\end{tabular} \\
double & \begin{tabular}{l} 
Convert categorical array to \\
double array
\end{tabular} \\
droplevels & \begin{tabular}{l} 
Drop levels \\
end
\end{tabular} \\
\begin{tabular}{l} 
Last index in indexing expression \\
for categorical array
\end{tabular} \\
flipdim & \begin{tabular}{l} 
Flip categorical array along \\
specified dimension
\end{tabular} \\
fliplr & \begin{tabular}{l} 
Flip categorical matrix in \\
left/right direction
\end{tabular} \\
flipud & \begin{tabular}{l} 
Flip categorical matrix in \\
up/down direction
\end{tabular} \\
getlabels & \begin{tabular}{l} 
Access categorical array labels
\end{tabular} \\
getlevels & \begin{tabular}{l} 
Get categorical array levels
\end{tabular} \\
hist & Plot histogram of categorical data
\end{tabular}
\begin{tabular}{ll} 
horzcat & \begin{tabular}{l} 
Horizontal concatenation for \\
categorical arrays \\
Convert categorical array to \\
signed 16-bit integer array
\end{tabular} \\
int16 & \begin{tabular}{l} 
Convert categorical array to \\
signed 32-bit integer array \\
Convert categorical array to \\
signed 64-bit integer array
\end{tabular} \\
int64 & \begin{tabular}{l} 
Convert categorical array to \\
signed 8-bit integer array
\end{tabular} \\
int8 & \begin{tabular}{l} 
Set intersection for categorical \\
arrays
\end{tabular} \\
intersect & \begin{tabular}{l} 
Inverse permute dimensions of \\
categorical array
\end{tabular} \\
ipermute & \begin{tabular}{l} 
True for empty categorical array
\end{tabular} \\
isempty & \begin{tabular}{l} 
True if categorical arrays are \\
equal
\end{tabular} \\
isequal & \begin{tabular}{l} 
Test for levels
\end{tabular} \\
islevel & \begin{tabular}{l} 
True for elements of categorical \\
array in set
\end{tabular} \\
ismember & \begin{tabular}{l} 
True if categorical array is scalar
\end{tabular} \\
isscalar & Test for undefined elements
\end{tabular}
\begin{tabular}{ll} 
permute & \begin{tabular}{l} 
Permute dimensions of \\
categorical array
\end{tabular} \\
reorderlevels & \begin{tabular}{l} 
Reorder levels
\end{tabular} \\
repmat & \begin{tabular}{l} 
Replicate and tile categorical \\
array
\end{tabular} \\
reshape & \begin{tabular}{l} 
Resize categorical array \\
Rotate categorical matrix 90 \\
degrees
\end{tabular} \\
rot90 & \begin{tabular}{l} 
Set difference for categorical \\
arrays
\end{tabular} \\
setdiff & \begin{tabular}{l} 
Label levels \\
Set exclusive-or for categorical \\
arrays
\end{tabular} \\
setlabels & \begin{tabular}{l} 
Shift dimensions of categorical \\
array
\end{tabular} \\
setxor & \begin{tabular}{l} 
Convert categorical array to \\
single array
\end{tabular} \\
shiftdim & \begin{tabular}{l} 
Size of categorical array
\end{tabular} \\
single & \begin{tabular}{l} 
Squeeze singleton dimensions \\
from categorical array
\end{tabular} \\
size & \begin{tabular}{l} 
Subscripted assignment for \\
categorical array
\end{tabular} \\
squeeze & \begin{tabular}{l} 
Subscript index for categorical \\
array
\end{tabular} \\
subsasgn & \begin{tabular}{l} 
Subscripted reference for \\
categorical array
\end{tabular} \\
subsindex & \begin{tabular}{l} 
Summary statistics for categorical \\
array
\end{tabular} \\
subsref & Product of categorical arrays
\end{tabular}
\begin{tabular}{ll} 
transpose & \begin{tabular}{l} 
Transpose categorical matrix \\
uint16
\end{tabular} \\
uint32 & \begin{tabular}{l} 
unsigned 16-bit integers to \\
Convert categorical array to \\
unsigned 32-bit integers
\end{tabular} \\
uint64 & \begin{tabular}{l} 
Convert categorical array to \\
unsigned 64-bit integers
\end{tabular} \\
uint8 & \begin{tabular}{l} 
Convert categorical array to \\
unsigned 8-bit integers
\end{tabular} \\
union & \begin{tabular}{l} 
Set union for categorical arrays
\end{tabular} \\
unique & \begin{tabular}{l} 
Unique values in categorical \\
array
\end{tabular} \\
vertcat & \begin{tabular}{l} 
Vertical concatenation for \\
categorical arrays
\end{tabular} \\
\hline
\end{tabular}

\section*{Properties Inherited Properties}

Properties in the following table are inherited from categorical.
\begin{tabular}{ll} 
labels & Text labels for levels \\
undeflabel & Text label for undefined levels
\end{tabular}

\section*{Copy \\ Semantics}

\section*{Examples}

Value. To learn how this affects your use of the class, see Comparing Handle and Value Classes in the MATLAB Object-Oriented Programming documentation.

Create a nominal array from string data in a cell array:
```

colors = nominal({'r' 'b' 'g';'g' 'r' 'b';'b' 'r' 'g'},...
{'blue' 'green' 'red'})
% Find elements meeting a criterion
colors == 'red'

```
```

ismember(colors,{'red' 'blue'})
% Compare two nominal arrays
colors2 = fliplr(colors)
colors == colors2

```

See Also
histc | ordinal
```

Purpose Construct nominal categorical array
Syntax $\quad B=\operatorname{nominal}(A)$
B = nominal(A,labels)
$B=$ nominal(A,labels,levels)
$B=$ nominal(A,labels, [],edges)

```
\(B=\) nominal \((A)\) creates a nominal array \(B\) from the array \(A\). A can be numeric, logical, character, categorical, or a cell array of strings. nominal creates the levels of \(B\) from the sorted unique values in \(A\), and creates default labels for them.
\(B=\) nominal(A, labels) labels the levels in B using the character array or cell array of strings labels. nominal assigns labels to levels in B in order according to the sorted unique values in A.
\(B=\) nominal(A, labels, levels) creates a nominal array with possible levels defined by levels. levels is a vector whose values can be compared to those in A using the equality operator. nominal assigns labels to each level from the corresponding elements of labels. If A contains any values not present in levels, the levels of the corresponding elements of \(B\) are undefined.
\(B=\) nominal(A, labels, [], edges) creates a nominal array by binning the numeric array A with bin edges given by the numeric vector edges. The uppermost bin includes values equal to the right-most edge. nominal assigns labels to each level in B from the corresponding elements of labels. edges must have one more element than labels.

By default, an element of \(B\) is undefined if the corresponding element of \(A\) is NaN (when \(A\) is numeric), an empty string (when \(A\) is character), or undefined (when A is categorical). nominal treats such elements as "undefined" or "missing" and does not include entries for them among the possible levels for B. To create an explicit level for such elements instead of treating them as undefined, you must use the levels input, and include NaN , the empty string, or an undefined element.

You may include duplicate labels in labels in order to merge multiple values in A into a single level in \(B\).

\section*{Examples Create a nominal array from Fisher's iris data:}
```

load fisheriris
species = nominal(species);
summary(species)
setosa versicolor virginica
50 50 50

```

Create a nominal array from characters, and provide explicit labels:
```

colors1 = nominal({'r' 'b' 'g'; 'g' 'r' 'b'; 'b' 'r' 'g'},...
{'blue' 'green' 'red'})

```

Create a nominal array from characters, and provide both explicit labels and an explicit order for display:
```

colors2 = nominal({'r' 'b' 'g'; 'g' 'r' 'b'; 'b' 'r' 'g'}, ...
{'red' 'green' 'blue'},{'r' 'g' 'b'})

```

Create a nominal array from integer data, merging odd and even values into only two nominal levels. Provide explicit labels:
```

toss = nominal(randi([1 4],5,2),{'odd' 'even' 'odd' 'even'},1:4)

```

1 Load patient data from the CSV file hospital. dat and store the information in a dataset array with observation names given by the first column in the data (patient identification):
```

patients = dataset('file','hospital.dat',...
'delimiter',',',...
'ReadObsNames',true);

```

2 Make the \(\{0,1\}\)-valued variable smoke nominal, and change the labels to 'No' and 'Yes':
```

patients.smoke = nominal(patients.smoke,{'No','Yes'});

```

3 Add new levels to smoke as placeholders for more detailed histories of smokers:
```

patients.smoke = addlevels(patients.smoke,...
{'0-5 Years','5-10 Years','LongTerm'});

```

4 Assuming the nonsmokers have never smoked, relabel the 'No ' level: patients.smoke = setlabels(patients.smoke,'Never','No');

5 Drop the undifferentiated 'Yes' level from smoke:
```

patients.smoke = droplevels(patients.smoke,'Yes');
Warning: OLDLEVELS contains categorical levels that
were present in A, caused some array elements to have
undefined levels.

```

Note that smokers now have an undefined level.
6 Set each smoker to one of the new levels, by observation name:
```

patients.smoke('YPL-320') = '5-10 Years';

```

\section*{See Also histc | ordinal}

\section*{Purpose Normal cumulative distribution function}

\section*{Syntax \\ Description}
\(P=\operatorname{normcdf}(X, m u\), sigma \()\)
[P,PLO,PUP] = normcdf(X,mu,sigma,pcov,alpha)
\(P=\operatorname{normcdf}(X, m u\), sigma \()\) computes the normal cdf at each of the values in X using the corresponding mean mu and standard deviation sigma. X , mu, and sigma can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The parameters in sigma must be positive.
[P,PLO,PUP] = normcdf(X,mu,sigma,pcov,alpha) produces confidence bounds for \(P\) when the input parameters mu and sigma are estimates. pcov is the covariance matrix of the estimated parameters. alpha specifies \(100(1-a l p h a) \%\) confidence bounds. The default value of alpha is 0.05 . PLO and PUP are arrays of the same size as P containing the lower and upper confidence bounds.

The function normdf computes confidence bounds for P using a normal approximation to the distribution of the estimate
\[
\frac{X-\hat{\mu}}{\hat{\sigma}}
\]
and then transforming those bounds to the scale of the output \(P\). The computed bounds give approximately the desired confidence level when you estimate mu, sigma, and pcov from large samples, but in smaller samples other methods of computing the confidence bounds might be more accurate.

The normal cdf is
\[
p=F(x \mid \mu, \sigma)=\frac{1}{\sigma \sqrt{2 \pi}} \int_{-\infty}^{x} e^{\frac{-(t-\mu)^{2}}{2 \sigma^{2}}} d t
\]

The result, \(p\), is the probability that a single observation from a normal distribution with parameters \(\mu\) and \(\sigma\) will fall in the interval \((-\infty x]\).

The standard normal distribution has \(\mu=0\) and \(\sigma=1\).

\section*{Examples \\ ```
p = normcdf([[-1 1]); \\ p(2)-p(1) \\ ans = \\ 0.6827
```}

See Also
How To
cdf | normpdf | norminv | normstat | normfit | normlike | normrnd
More generally, about \(68 \%\) of the observations from a normal distribution fall within one standard deviation, \(\sigma\), of the mean, \(\mu\).
- "Normal Distribution" on page B-83
Purpose Normal parameter estimates
Syntax

[muhat,sigmahat] = normfit(data)
[muhat,sigmahat,muci,sigmaci] = normfit(data)
[muhat,sigmahat,muci,sigmaci] = normfit(data,alpha)
[...] = normfit(data,alpha, censoring)
[...] = normfit(data, alpha, censoring,freq)
[...] = normfit(data,alpha, censoring,freq,options)

\section*{Description}
[muhat, sigmahat] = normfit(data) returns estimates of the mean, \(\mu\), and standard deviation, \(\sigma\), of the normal distribution given the data in data.
[muhat,sigmahat,muci,sigmaci] = normfit(data) returns 95\% confidence intervals for the parameter estimates on the mean and standard deviation in the arrays muci and sigmaci, respectively. The first row of muci contains the lower bounds of the confidence intervals for \(\mu\) the second row contains the upper bounds. The first row of sigmaci contains the lower bounds of the confidence intervals for \(\sigma\), and the second row contains the upper bounds.
[muhat,sigmahat,muci,sigmaci] = normfit(data,alpha) returns 100(1-alpha) \% confidence intervals for the parameter estimates, where alpha is a value in the range [0 1] specifying the width of the confidence intervals. By default, alpha is 0.05 , which corresponds to \(95 \%\) confidence intervals.
[...] = normfit(data,alpha, censoring) accepts a Boolean vector, censoring, of the same size as data, which is 1 for observations that are right-censored and 0 for observations that are observed exactly. data must be a vector in order to pass in the argument censoring.
[...] = normfit(data, alpha, censoring,freq) accepts a frequency vector, freq, of the same size as data. Typically, freq contains integer frequencies for the corresponding elements in data, but can contain any nonnegative values. Pass in [] for alpha, censoring, or freq to use their default values.
[...] = normfit(data,alpha,censoring,freq,options) accepts a structure, options, that specifies control parameters for the iterative algorithm the function uses to compute maximum likelihood estimates when there is censoring. The normal fit function accepts an options structure which you can create using the function statset. Enter statset('normfit') to see the names and default values of the parameters that normfit accepts in the options structure. See the reference page for statset for more information about these options.

\section*{Examples}

In this example the data is a two-column random normal matrix. Both columns have \(\mu=10\) and \(\sigma=2\). Note that the confidence intervals below contain the "true values."
```

data = normrnd(10,2,100,2);
[mu,sigma,muci,sigmaci] = normfit(data)
mu =
10.1455 10.0527
sigma =
1.9072 2.1256
muci =
9.7652 9.6288
10.5258 10.4766
sigmaci =
1.6745 1.8663
2.2155 2.4693

```
See Also mle | normlike | normpdf | normcdf | norminv | normstat | normrnd
How To . "Normal Distribution" on page B-83

\section*{Purpose Normal inverse cumulative distribution function}

Syntax \(\quad X=\operatorname{norminv}(P\), mu, sigma \()\)
[X,XLO,XUP] = norminv(P,mu,sigma, pcov,alpha)

\section*{Description}
\(X=\) norminv( \(P\), mu, sigma) computes the inverse of the normal cdf using the corresponding mean mu and standard deviation sigma at the corresponding probabilities in P. P, mu, and sigma can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The parameters in sigma must be positive, and the values in \(P\) must lie in the interval [01].
[X,XLO,XUP] = norminv(P,mu,sigma,pcov,alpha) produces confidence bounds for \(X\) when the input parameters mu and sigma are estimates. pcov is the covariance matrix of the estimated parameters. alpha specifies \(100(1-\) alpha \() \%\) confidence bounds. The default value of alpha is 0.05 . XLO and XUP are arrays of the same size as X containing the lower and upper confidence bounds.

The function norminv computes confidence bounds for P using a normal approximation to the distribution of the estimate
\[
\hat{\mu}+\hat{\sigma} q
\]
where \(q\) is the Pth quantile from a normal distribution with mean 0 and standard deviation 1 . The computed bounds give approximately the desired confidence level when you estimate mu, sigma, and pcov from large samples, but in smaller samples other methods of computing the confidence bounds may be more accurate.

The normal inverse function is defined in terms of the normal cdf as
\[
x=F^{-1}(p \mid \mu, \sigma)=\{x: F(x \mid \mu, \sigma)=p\}
\]
where
\[
p=F(x \mid \mu, \sigma)=\frac{1}{\sigma \sqrt{2 \pi}} \int_{-\infty}^{x} e^{\frac{-(t-\mu)^{2}}{2 \sigma^{2}}} d t
\]

The result, \(x\), is the solution of the integral equation above where you supply the desired probability, \(p\).

\section*{Examples} normal distribution.
```

x = norminv([0.025 0.975],0,1)
x =
-1.9600 1.9600

``` shortest.
```

xl = norminv([0.01 0.96],0,1)
xl =
-2.3263 1.7507

``` than x .

\author{
\section*{See Also} \\ \section*{How To \\ \\ - "Normal Distribution" on page B-83} \\ icdf | normcdf | normpdf | normstat | normfit | normlike | normrnd
}

Find an interval that contains \(95 \%\) of the values from a standard

Note that the interval \(x\) is not the only such interval, but it is the

The interval xl also contains \(95 \%\) of the probability, but it is longer
Purpose Normal negative log-likelihood
Syntax

nlogL = normlike(params,data)
[nlogL,AVAR] = normlike(params,data)
[...] = normlike(param,data,censoring)
[...] = normlike(param,data,censoring,freq)
nlogL = normlike(params, data) returns the negative of the normal log-likelihood function. params(1) is the mean, mu, and params (2) is the standard deviation, sigma.
[nlogL,AVAR] = normlike(params,data) also returns the inverse of Fisher's information matrix, AVAR. If the input parameter values in params are the maximum likelihood estimates, the diagonal elements of AVAR are their asymptotic variances. AVAR is based on the observed Fisher's information, not the expected information.
[...] = normlike(param,data, censoring) accepts a Boolean vector, censoring, of the same size as data, which is 1 for observations that are right-censored and 0 for observations that are observed exactly.
[...] = normlike(param,data,censoring,freq) accepts a frequency vector, freq, of the same size as data. The vector freq typically contains integer frequencies for the corresponding elements in data, but can contain any nonnegative values. Pass in [] for censoring to use its default value.
normlike is a utility function for maximum likelihood estimation.

\footnotetext{
See Also normfit | normpdf | normcdf | norminv | normstat | normrnd
How To - "Normal Distribution" on page B-83
}

Purpose Normal probability density function
```

Syntax }\quadY=\operatorname{normpdf(X,mu,sigma)
Y = normpdf(X)
Y = normpdf(X,mu)

```

\section*{Description}

\section*{Examples}
```

mu = [0:0.1:2];
[y i] = max(normpdf(1.5,mu,1));
MLE $=$ mu(i)
MLE =

```

\subsection*{1.5000}

\author{
See Also \\ pdf | normcdf | norminv | normstat | normfit | normlike | normrnd | mvnpdf \\ How To \\ - "Normal Distribution" on page B-83
}

Purpose Normal probability plot

\section*{Syntax \(\quad h=\operatorname{normplot}(X)\)}

Description \(\quad h=\) normplot \((X)\) displays a normal probability plot of the data in \(X\). For matrix \(X\), normplot displays a line for each column of \(X\). \(h\) is a handle to the plotted lines.
The plot has the sample data displayed with the plot symbol ' + '. Superimposed on the plot is a line joining the first and third quartiles of each column of \(X\) (a robust linear fit of the sample order statistics.) This line is extrapolated out to the ends of the sample to help evaluate the linearity of the data.

The purpose of a normal probability plot is to graphically assess whether the data in \(X\) could come from a normal distribution. If the data are normal the plot will be linear. Other distribution types will introduce curvature in the plot. normplot uses midpoint probability plotting positions. Use probplot when the data included censored observations.

Examples Generate a normal sample and a normal probability plot of the data.
```

x = normrnd(10,1,25,1);
normplot(x)

```

\(\begin{array}{ll}\text { See Also } \quad \text { cdfplot | wblplot | probplot | hist | normfit | norminv | normpdf } \\ & \text { | normspec | normstat | normcdf | normrnd | normlike }\end{array}\)
How To . "Normal Distribution" on page B-83

\section*{Purpose Normal random numbers}
Syntax \(\quad\)\begin{tabular}{rl}
\(R\) & \(=\operatorname{normrnd}(m u\), sigma \()\) \\
\(R\) & \(=\operatorname{normrnd}(m u, \operatorname{sigma}, m, n, \ldots)\) \\
\(R\) & \(=\operatorname{normrnd}(m u, \operatorname{sigma},[m, n, \ldots])\)
\end{tabular}

\section*{Description}

\section*{Examples}

See Also \(\begin{gathered}\text { random | normpdf | normodf | norminv | normstat | normfit | } \\ \text { normlike | mvnrnd | lognrnd }\end{gathered}\)
How To . "Normal Distribution" on page B-83
\(R=\) normrnd(mu, sigma) generates random numbers from the normal distribution with mean parameter mu and standard deviation parameter sigma. mu and sigma can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of R. A scalar input for mu or sigma is expanded to a constant array with the same dimensions as the other input.
\(R=\) normrnd(mu,sigma, \(m, n, \ldots\) ) or \(R=\)
normrnd(mu, sigma, \([m, n, \ldots]\) ) generates an \(m\)-by-n-by-... array. The mu, sigma parameters can each be scalars or arrays of the same size as R.
```

n1 = normrnd(1:6,1./(1:6))
n1 =
2.1650
n2 = normrnd(0,1,[1 5])
n2 =
0.0591 1.7971 0.2641 0.8717 -1.4462
n3 = normrnd([1 2 3;4 5 6],0.1,2,3)
n3 =
0.9299 1.9361 2.9640
4.1246 5.0577 5.9864

```
        - "Normal Distribution" on page B-83
```

Purpose Normal density plot between specifications
Syntax normspec(specs)
normspec(specs,mu,sigma)
normspec(specs,mu,sigma,region)
p = normspec(...)
[p,h] = normspec(...)

```

\section*{Description}

Examples
normspec (specs) plots the standard normal density, shading the portion inside the specification limits given by the two-element vector specs. Set specs(1) to - Inf if there is no lower limit; set specs (2) to Inf if there is no upper limit.
normspec (specs, mu, sigma) shades the portion inside the specification limits of a normal density with parameters mu and sigma. The defaults are \(\mathrm{mu}=0\) and sigma \(=1\).
normspec(specs,mu,sigma,region) shades the region either 'inside' or 'outside' the specification limits. The default is 'inside'.
\(\mathrm{p}=\) normspec (...) also returns the probability, p , of the shaded area.
\([p, h]=\) normspec (...) also returns a handle \(h\) to the line objects.

A production process fills cans of paint. The average amount of paint in any can is 1 gallon, but variability in the process produces a standard deviation of 2 ounces ( \(2 / 128\) gallons). What is the probability that cans will be filled under specification by 3 or more ounces?
```

p = normspec([1-3/128,Inf],1,2/128,'outside')
p =
0.0668

```


\footnotetext{
See Also
capaplot | histfit
How To . "Normal Distribution" on page B-83
}
Purpose Normal mean and variance
Syntax [M,V] = normstat(mu,sigma)
Description [M, V] = normstat(mu, sigma) returns the mean of and variancearray with the same dimensions as the other input. the variance is \(\sigma^{2}\).

\section*{Examples}
```

n = 1:5;
[m,v] = normstat(n'*n,n'*n)
m =
1
2 4 4 6 8 10
3
4
5
v =
1
4
9% 36 81 144 225
16 64 144 256 400
25100 225400 625

```for the normal distribution using the corresponding mean mu andstandard deviation sigma. mu and sigma can be vectors, matrices, ormultidimensional arrays that all have the same size, which is also thesize of M and V . A scalar input for mu or sigma is expanded to a constant

The mean of the normal distribution with parameters \(\mu\) and \(\sigma\) is \(\mu\), and

\section*{See Also}
normpdf | normcdf | norminv | normfit | normlike | normrnd

\footnotetext{
How To
- "Normal Distribution" on page B-83
}

\section*{piecewisedistribution.nsegments}
Purpose Number of segments
Syntax n = nsegments(obj)
Description \(\mathrm{n}=\mathrm{nsegments}(\mathrm{obj})\) returns the number of segments n in thepiecewise distribution object obj.
Examples Fit Pareto tails to a \(t\) distribution at cumulative probabilities 0.1 and 0.9:
```

t = trnd(3,100,1);
obj = paretotails(t,0.1,0.9);
n = nsegments(obj)
n =
3

```
See Also
paretotails | boundary | segment
Purpose Number of decision trees in ensemble
Description The NTrees property is a scalar equal to the number of decision trees in the ensemble.
See Also ..... Trees

\section*{ProbDistParametric.NumParams property}
\begin{tabular}{ll} 
Purpose & \begin{tabular}{l} 
Read-only value specifying number of parameters of ProbDistParametric \\
object
\end{tabular} \\
Description & \begin{tabular}{l} 
NumParams is a read-only property of the ProbDistParametric class. \\
NumParams is a value specifying the number of parameters of a \\
distribution represented by a ProbDistParametric object.
\end{tabular} \\
Values & \begin{tabular}{l} 
This value is an integer that counts both the specified parameters and \\
parameters that are fit to the data. Use this information to view and \\
compare the number of parameters supplied to create distributions.
\end{tabular}
\end{tabular}
Purpose Number of elements in dataset array
Syntax \(\quad\)\begin{tabular}{rl}
\(n\) & \(=\operatorname{numel}(A)\) \\
\(n\) & \(=\operatorname{numel}(A\), varargin \()\)
\end{tabular}

Description \(\quad n=\) numel \((A)\) returns 1 . To find the number of elements, \(n\), in the dataset array \(A\), use prod(size(A)) or numel(A, ':', ':').
\(\mathrm{n}=\) numel(A, varargin) returns the number of subscripted elements, n , in \(\mathrm{A}(\) index1, index2, ..., indexn), where varargin is a cell array whose elements are index1, index2, ... indexn.

See Also length | size

Purpose Number of elements in categorical array
\[
\begin{array}{ll}
\text { Syntax } & n=\text { numel }(A) \\
& n=n u m e l(A, \text { varargin }) \\
\text { Description } \quad & n=\text { numel }(A) \text { returns the number of elements in the categorical array } \\
& \text { A. } \\
& n=\text { numel }(A, \text { varargin) returns the number of subscripted elements, } \\
& n, \text { in } A(\text { index } 1, \text { index2, } \ldots \text { indexN }) \text {, where varargin is a cell } \\
& \text { array whose elements are index1, index2, } \ldots \text { indexN. }
\end{array}
\]
See Also ..... size

\section*{classregtree.numnodes}

\section*{Purpose Number of nodes}

\section*{Syntax \(\quad n=\) numnodes \((t)\)}

Description \(n=\) numnodes \((t)\) returns the number of nodes \(n\) in the tree \(t\).

\section*{Examples Create a classification tree for Fisher's iris data:}
```

load fisheriris;
t = classregtree(meas,species,...
'names',{'SL' 'SW' 'PL' 'PW'})
t=
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
view(t)

```

\section*{classregtree.numnodes}
\begin{tabular}{|l|l|l|l|l|l|}
\hline Click to display: & Identity & \(\square\) & Magnification: \(100 \%\) & \(\rightarrow\) & Pruning level: \\
\hline 0 of 4 & \\
\hline
\end{tabular}

```

n = numnodes(t)
n =
9

```

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

\section*{See Also \\ classregtree}

\section*{cvpartition.NumTestSets property}

\section*{Purpose Number of test sets}

Description Value is the number of folds in partitions of type 'kfold' and 'leaveout'.

Value is 1 in partitions of type 'holdout' and 'resubstitution'.

\section*{TreeBagger.NVarToSample property}

Purpose Number of variables for random feature selection
Description The NVarToSample property specifies the number of predictor or feature variables to select at random for each decision split. By default, it is set to the square root of the total number of variables for classification and one third of the total number of variables for regression. Setting this argument to any valid value except 'all' invokes Breiman's "random forest" algorithm.

\author{
See Also \\ classregtree
}

Purpose

Description

Cell array of nonempty, distinct strings giving names of observations in data set

A cell array of nonempty, distinct strings giving the names of the observations in the data set. This property may be empty, but if not empty, the number of strings must equal the number of observations.

\section*{ClassificationBaggedEnsemble.oobEdge}

Purpose Out-of-bag classification edge
Syntax \(\quad \begin{aligned} \text { edge } & =\text { oobEdge }(\text { ens }) \\ \text { edge } & =\text { oobEdge (ens, Name, Value })\end{aligned}\)

Description

Input
Arguments
edge \(=\) oobEdge (ens) returns out-of-bag classification edge for ens.
edge \(=\) oobEdge(ens, Name, Value) computes classification edge with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
ens
A classification bagged ensemble, constructed with fitensemble.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
learners
Indices of weak learners in the ensemble ranging from 1 to ens.NTrained. oobEdge uses only these learners for calculating loss.

Default: 1:NTrained
mode
String representing the meaning of the output L:
- 'ensemble' - L is a scalar value, the loss for the entire ensemble.
- 'individual' - L is a vector with one element per trained learner.

\section*{ClassificationBaggedEnsemble.oobEdge}
- 'cumulative' - L is a vector in which element \(J\) is obtained by using learners \(1: J\) from the input list of learners.

Default: 'ensemble'

\section*{Output \\ edge}

Arguments
Classification edge, a weighted average of the classification margin.

\section*{Definitions}

\section*{Edge}

The edge is the weighted mean value of the classification margin. The weights are the class probabilities in ens. Prior.

\section*{Margin}

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix ens. \(x\).

\section*{Out of Bag}

Bagging, which stands for "bootstrap aggregation", is a type of ensemble learning. To bag a weak learner such as a decision tree on a dataset, fitensemble generates many bootstrap replicas of the dataset and grows decision trees on these replicas. fitensemble obtains each bootstrap replica by randomly selecting \(N\) observations out of \(N\) with replacement, where \(N\) is the dataset size. To find the predicted response of a trained ensemble, predict take an average over predictions from individual trees.

Drawing N out of N observations with replacement omits on average \(37 \%\) ( \(1 / e\) ) of observations for each decision tree. These are "out-of-bag" observations. For each observation, oobLoss estimates the out-of-bag prediction by averaging over predictions from all trees in the ensemble for which this observation is out of bag. It then compares the computed prediction against the true response for this observation. It calculates the out-of-bag error by comparing the out-of-bag predicted responses

\section*{ClassificationBaggedEnsemble.oobEdge}
against the true responses for all observations used for training. This out-of-bag average is an unbiased estimator of the true ensemble error.
```

Examples Find the out-of-bag edge for a bagged ensemble from the Fisher iris data:
load fisheriris
ens = fitensemble(meas,species,'Bag',100,...
'Tree','type','classification');
edge = oobEdge(ens)
edge =
0.8730
See Also oobMargin | oobPredict | oobLoss
How To

- Chapter 13, "Nonparametric Supervised Learning"

```

\section*{TreeBagger.oobError}
\begin{tabular}{|c|c|}
\hline Purpose & Out-of-bag error \\
\hline Syntax & ```
err = oobError(B)
err = oobError(B,'param1',val1,'param2',val2,...)
``` \\
\hline Description & \begin{tabular}{l}
err \(=\) oobError(B) computes the misclassification probability (for classification trees) or mean squared error (for regression trees) for out-of-bag observations in the training data, using the trained bagger \(B\). err is a vector of length NTrees, where NTrees is the number of trees in the ensemble. \\
err = oobError(B,'param1',val1,'param2', val2,...) specifies optional parameter name/value pairs: \\
'mode ' \\
String indicating how oobError computes errors. If set to 'cumulative' (default), the method computes cumulative errors and err is a vector of length NTrees, where the first element gives error from trees (1), second element gives error from trees(1:2) etc, up to trees ( \(1:\) NTrees). If set to 'individual', err is a vector of length NTrees, where each element is an error from each tree in the ensemble. If set to 'ensemble', err is a scalar showing the cumulative error for the entire ensemble.
\end{tabular} \\
\hline & \begin{tabular}{l}
'trees' \\
Vector of indices indicating what trees to include in this calculation. By default, this argument is set to 'all' and the method uses all trees. If 'trees' is a numeric vector, the method returns a vector of length NTrees for 'cumulative' and 'individual' modes, where NTrees is the number of elements in the input vector, and a scalar for 'ensemble' mode. For example, in the 'cumulative' mode, the
\end{tabular} \\
\hline
\end{tabular}

\section*{TreeBagger.oobError}
first element gives error from trees(1), the second element gives error from trees (1:2) etc.
'treeweights' Vector of tree weights. This vector must have the same length as the 'trees' vector. oobError uses these weights to combine output from the specified trees by taking a weighted average instead of the simple nonweighted majority vote. You cannot use this argument in the 'individual' mode.

See Also
CompactTreeBagger.error

\section*{TreeBagger.OOBIndices property}

\author{
Purpose Indicator matrix for out-of-bag observations \\ Description The 00BIndices property is a logical array of size Nobs-by-NTrees where Nobs is the number of observations in the training data and NTrees is the number of trees in the ensemble. The ( \(I, J\) ) element is true if observation I is out-of-bag for tree \(J\) and false otherwise. In other words, a true value means observation I was not selected for the training data used to grow tree J .
}

\author{
See Also \\ classregtree
}

\section*{TreeBagger.OOBInstanceWeight property}

Purpose Count of out-of-bag trees for each observation
Description The 00BInstanceWeight property is a numeric array of size Nobs-by-1 containing the number of trees used for computing out-of-bag response for each observation. Nobs is the number of observations in the training data used to create the ensemble.

\section*{ClassificationBaggedEnsemble.oobLoss}
\begin{tabular}{ll} 
Purpose & Out-of-bag classification error \\
Syntax & L \(=\) oobloss(ens) \\
& L \(=\) oobloss(ens, Name, Value)
\end{tabular}

Description \(L=\) oobloss (ens) returns the classification error for ens computed for out-of-bag data.

L = oobloss(ens, Name, Value) computes error with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1,Value1, ,NameN,ValueN.

\section*{Input \\ Arguments}
ens
A classification bagged ensemble, constructed with fitensemble.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

\section*{learners}

Indices of weak learners in the ensemble ranging from 1 to NTrained. oobLoss uses only these learners for calculating loss.

Default: 1:NTrained

\section*{lossfun}

Function handle or string representing a loss function. Built-in loss functions:
- 'binodeviance ' - See "Loss Functions" on page 20-1289
- 'classiferror' - Fraction of misclassified data

\section*{ClassificationBaggedEnsemble.oobLoss}
- 'exponential' - See "Loss Functions" on page 20-1289

You can write your own loss function in the syntax described in "Loss Functions" on page 20-1289.

Default: 'classiferror'

\section*{mode}

String representing the meaning of the output L:
- 'ensemble' - L is a scalar value, the loss for the entire ensemble.
- 'individual' - L is a vector with one element per trained learner.
- 'cumulative' - \(L\) is a vector in which element \(J\) is obtained by using learners \(1: \mathrm{J}\) from the input list of learners.

Default: 'ensemble'

\section*{Output}

Arguments

\section*{Definitions}

\section*{Out of Bag}

Bagging, which stands for "bootstrap aggregation", is a type of ensemble learning. To bag a weak learner such as a decision tree on a dataset, fitensemble generates many bootstrap replicas of the dataset and grows decision trees on these replicas. fitensemble obtains each bootstrap replica by randomly selecting N observations out of N with replacement, where \(N\) is the dataset size. To find the predicted response of a trained ensemble, predict take an average over predictions from individual trees.

\section*{ClassificationBaggedEnsemble.oobLoss}

Drawing N out of N observations with replacement omits on average \(37 \%\) (1/e) of observations for each decision tree. These are "out-of-bag" observations. For each observation, oobLoss estimates the out-of-bag prediction by averaging over predictions from all trees in the ensemble for which this observation is out of bag. It then compares the computed prediction against the true response for this observation. It calculates the out-of-bag error by comparing the out-of-bag predicted responses against the true responses for all observations used for training. This out-of-bag average is an unbiased estimator of the true ensemble error.

\section*{Loss Functions}

The built-in loss functions are:
- 'binodeviance' - For binary classification, assume the classes \(y_{n}\) are -1 and 1 . With weight vector \(w\) normalized to have sum 1 , and predictions of row \(n\) of data \(X\) as \(f\left(X_{n}\right)\), the binomial deviance is
\[
\sum w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(X_{n}\right)\right)\right)
\]
- 'classiferror' - Fraction of misclassified data, weighted by \(w\).
- 'exponential' - With the same definitions as for 'binodeviance', the exponential loss is
\[
\sum w_{n} \exp \left(-y_{n} f\left(X_{n}\right)\right)
\]

To write your own loss function, create a function file of the form
```

function loss = lossfun(C,S,W,COST)

```
- \(N\) is the number of rows of \(X\).
- K is the number of classes in tree, represented in tree.ClassNames.
- C is an N-by-K logical matrix, with one true per row for the true class. The index for each class is its position in tree.ClassNames.

\section*{ClassificationBaggedEnsemble.oobLoss}
- \(S\) is an N -by-K numeric matrix. S is a matrix of posterior probabilities for classes with one row per observation, similar to the posterior output from predict.
- \(W\) is a numeric vector with \(N\) elements, the observation weights.
- COST is a K-by-K numeric matrix of misclassification costs. The default 'classiferror' cost function uses a cost of 0 for correct classification, and 1 for misclassification. In other words, 'classiferror' uses COST=ones(K)-eye(K).
- The output loss should be a scalar.

Pass the function handle @lossfun as the value of the lossfun name-value pair.

\section*{Examples Find the out-of-bag error for a bagged ensemble from the Fisher iris data:}
```

load fisheriris
ens = fitensemble(meas,species,'Bag',100,...
'Tree','type','classification');
L = oobLoss(ens)
L =
0.0467

```

\section*{See Also}
loss | oobEdge | oobMargin | oobPredict

\section*{How To}
- Chapter 13, "Nonparametric Supervised Learning"

\section*{RegressionBaggedEnsemble.oobLoss}

\section*{Purpose}

Syntax

Description

Input
Arguments

Out-of-bag regression error

L = oobLoss(ens)
L = oobLoss(ens,Name,Value)
\(L=\) oobLoss(ens) returns the mean squared error for ens computed for out-of-bag data.

L = oobLoss(ens,Name, Value) computes error with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.
ens
A regression bagged ensemble, constructed with fitensemble.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

\section*{learners}

Indices of weak learners in the ensemble ranging from 1 to NTrained. oobLoss uses only these learners for calculating loss.

Default: 1:NTrained

\section*{lossfun}

Function handle for loss function, or the string 'mse', meaning mean squared error. If you pass a function handle fun, oobLoss calls it as
FUN(Y,Yfit,w)

\section*{RegressionBaggedEnsemble.oobLoss}
e Y, Yfit, and \(W\) are numeric vectors of the same length. \(Y\) is the observed response, Yfit is the predicted response, and \(W\) is the observation weights.

Default: 'mse'

\section*{mode}

String representing the meaning of the output L:
- 'ensemble' - L is a scalar value, the loss for the entire ensemble.
- 'individual' - L is a vector with one element per trained learner.
- 'cumulative' - \(L\) is a vector in which element \(J\) is obtained by using learners \(1: J\) from the input list of learners.

Default: 'ensemble'

\section*{Output}

Arguments

\section*{Definitions}

L
Mean squared error of the out-of-bag observations, a scalar. L can be a vector, or can represent a different quantity, depending on the name-value settings.

\section*{Out of Bag}

Bagging, which stands for "bootstrap aggregation", is a type of ensemble learning. To bag a weak learner such as a decision tree on a dataset, fitensemble generates many bootstrap replicas of the dataset and grows decision trees on these replicas. fitensemble obtains each bootstrap replica by randomly selecting N observations out of N with replacement, where \(N\) is the dataset size. To find the predicted response of a trained ensemble, predict take an average over predictions from individual trees.

Drawing N out of N observations with replacement omits on average \(37 \%\) (1/e) of observations for each decision tree. These are "out-of-bag"

\section*{RegressionBaggedEnsemble.oobLoss}

Examples Compute the out-of-bag error for the carsmall data:
```

load carsmall
X = [Displacement Horsepower Weight];
ens = fitensemble(X,MPG,'bag',100,'Tree',...
'type','regression');
L = oobLoss(ens)
L =
17.0665

```

\section*{See Also \\ oobPredict | loss}

How To - Chapter 13, "Nonparametric Supervised Learning"

\section*{ClassificationBaggedEnsemble.oobMargin}

Purpose Out-of-bag classification margins

Syntax

Description

\section*{Input \\ Arguments}

\section*{Output Arguments}

\section*{Definitions}
margin = oobMargin(ens)
margin \(=\) oobMargin(ens, Name, Value)
margin \(=\) oobMargin(ens) returns out-of-bag classification margins.
margin \(=\) oobMargin(ens, Name, Value) calculates margins with additional options specified by one or more Name, Value pair arguments.
ens
A classification bagged ensemble, constructed with fitensemble.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (''). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{learners}

Indices of weak learners in the ensemble ranging from 1 to ens.NTrained. oobEdge uses only these learners for calculating loss.

Default: 1:NTrained
margin
A numeric column vector of length size(ens. \(\mathrm{X}, 1\) ).

\section*{Out of Bag}

Bagging, which stands for "bootstrap aggregation", is a type of ensemble learning. To bag a weak learner such as a decision tree on a dataset, fitensemble generates many bootstrap replicas of the dataset and grows decision trees on these replicas. fitensemble obtains each bootstrap replica by randomly selecting N observations out of N with

\section*{ClassificationBaggedEnsemble.oobMargin}
replacement, where \(N\) is the dataset size. To find the predicted response of a trained ensemble, predict take an average over predictions from individual trees.

Drawing N out of N observations with replacement omits on average \(37 \%(1 / e)\) of observations for each decision tree. These are "out-of-bag" observations. For each observation, oobLoss estimates the out-of-bag prediction by averaging over predictions from all trees in the ensemble for which this observation is out of bag. It then compares the computed prediction against the true response for this observation. It calculates the out-of-bag error by comparing the out-of-bag predicted responses against the true responses for all observations used for training. This out-of-bag average is an unbiased estimator of the true ensemble error.

\section*{Margin}

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix ens. x .

Examples \(\begin{aligned} & \text { Find the out-of-bag margin for a bagged ensemble from the Fisher iris } \\ & \text { data: Find how many elements of margin are equal to } 1 .\end{aligned}\)
```

load fisheriris
ens = fitensemble(meas,species,'Bag',100,···
'Tree','type','classification');
margin = oobMargin(ens);
sum(margin == 1)
ans =
108

```

See Also oobPredict | oobLoss | oobEdge | margin
How To - Chapter 13, "Nonparametric Supervised Learning"

\section*{TreeBagger.oobMargin}
\begin{tabular}{|c|c|}
\hline Purpose & Out-of-bag margins \\
\hline Syntax & ```
mar = oobMargin(B)
mar = oobMargin(B,'param1',val1,'param2',val2,...)
``` \\
\hline Description & \begin{tabular}{l}
mar \(=\) oobMargin(B) computes an Nobs-by-NTrees matrix of classification margins for out-of-bag observations in the training data, using the trained bagger B. \\
mar = oobMargin(B,'param1', val1,'param2', val2,...) specifies optional parameter name/value pairs:
\end{tabular} \\
\hline & String indicating how oobMargin computes errors. If set to 'cumulative' (default), the method computes cumulative margins and mar is an Nobs-by-NTrees matrix, where the first column gives margins from trees (1), second column gives margins from trees (1:2) etc, up to trees(1:NTrees). If set to 'individual', mar is an Nobs-by-NTrees matrix, where each column gives margins from each tree in the ensemble. If set to 'ensemble', mar is a single column of length Nobs showing the cumulative margins for the entire ensemble. \\
\hline & 'trees ' Vector of indices indicating what trees to include in this calculation. By default, this argument is set to 'all' and the method uses all trees. If 'trees' is a numeric vector, the method returns an Nobs-by-NTrees matrix for 'cumulative' and 'individual' modes, where NTrees is the number of elements in the input vector, and a single column for 'ensemble' mode. For example, in the 'cumulative' mode, the first column gives \\
\hline
\end{tabular}

\section*{TreeBagger.oobMargin}
\[
\begin{array}{ll} 
& \begin{array}{l}
\text { margins from trees(1), the second column gives } \\
\text { margins from trees(1:2) etc. }
\end{array} \\
\text { 'treeweights' } & \begin{array}{l}
\text { Vector of tree weights. This vector must have the } \\
\text { same length as the 'trees' vector. oobMargin uses } \\
\text { these weights to combine output from the specified } \\
\text { trees by taking a weighted average instead of the } \\
\text { simple nonweighted majority vote. You cannot use } \\
\text { this argument in the 'individual ' mode. }
\end{array}
\end{array}
\]

See Also
CompactTreeBagger.margin

\section*{TreeBagger.oobMeanMargin}
\begin{tabular}{|c|c|}
\hline Purpose & Out-of-bag mean margins \\
\hline Syntax & ```
mar = oobMeanMargin(B)
mar = oobMeanMargin(B,'param1',val1,'param2',val2,...)
``` \\
\hline \multirow[t]{3}{*}{Description} & \begin{tabular}{l}
mar = oobMeanMargin(B) computes average classification margins for out-of-bag observations in the training data, using the trained bagger B. oobMeanMargin averages the margins over all out-of-bag observations. mar is a row-vector of length NTrees, where NTrees is the number of trees in the ensemble. \\
mar = oobMeanMargin(B,'param1',val1,'param2', val2,...) specifies optional parameter name/value pairs: \\
'mode' \\
String indicating how oobMargin computes errors. If set to 'cumulative' (default), is a vector of length NTrees where the first element gives mean margin from trees(1), second column gives mean margins from trees(1:2) etc, up to trees(1:NTrees). If set to 'individual', mar is a vector of length NTrees, where each element is a mean margin from each tree in the ensemble. If set to 'ensemble', mar is a scalar showing the cumulative mean margin for the entire ensemble.
\end{tabular} \\
\hline & 'trees ' Vector of indices indicating what trees to include in this calculation. By default, this argument is set to 'all' and the method uses all trees. If 'trees' is a numeric vector, the method returns a vector of length NTrees for 'cumulative' and 'individual' modes, where NTrees is the number of elements in the input vector, and a scalar for 'ensemble' mode. For example, in the 'cumulative' mode, the first element gives mean margin from trees(1), the second element gives mean margin from trees(1:2) etc. \\
\hline & 'treeweights ' Vector of tree weights. This vector must have the same length as the 'trees' vector. oobMeanMargin uses these weights to combine output from the specified trees by taking a weighted average instead of the simple nonweighted majority vote. You cannot use this argument in the 'individual' mode. \\
\hline
\end{tabular}

\section*{TreeBagger.oobMeanMargin}

See Also
CompactTreeBagger.meanMargin

\section*{TreeBagger.OOBPermutedVarCountRaiseMargin property}

Purpose Variable importance for raising margin
Description The 00BPermutedVarCountRaiseMargin property is a numeric array of size 1-by-Nvars containing a measure of variable importance for each predictor. For any variable, the measure is the difference between the number of raised margins and the number of lowered margins if the values of that variable are permuted across the out-of-bag observations. This measure is computed for every tree, then averaged over the entire ensemble and divided by the standard deviation over the entire ensemble. This property is empty for regression trees.

\section*{TreeBagger.OOBPermutedVarDeltaError property}

Purpose Variable importance for prediction error
Description The OOBPermutedVarDeltaError property is a numeric array of size 1-by-Nvars containing a measure of importance for each predictor variable (feature). For any variable, the measure is the increase in prediction error if the values of that variable are permuted across the out-of-bag observations. This measure is computed for every tree, then averaged over the entire ensemble and divided by the standard deviation over the entire ensemble.

\section*{TreeBagger.OOBPermutedVarDeltaMeanMargin property}

Purpose Variable importance for classification margin
Description
The OOBPermutedVarDeltaMeanMargin property is a numeric array of size 1-by-Nvars containing a measure of importance for each predictor variable (feature). For any variable, the measure is the decrease in the classification margin if the values of that variable are permuted across the out-of-bag observations. This measure is computed for every tree, then averaged over the entire ensemble and divided by the standard deviation over the entire ensemble. This property is empty for regression trees.

\section*{ClassificationBaggedEnsemble.oobPredict}

\section*{Purpose}

Predict out-of-bag response of ensemble

\section*{Syntax}

Description

Input Arguments

\section*{Arguments}

Output
label
Classification labels of the same data type as the training data Y . There are \(N\) elements or rows, where \(N\) is the number of training observations. The label is the class with the highest score. In case of a tie, the label is earliest in ens.ClassNames.
score

\section*{ClassificationBaggedEnsemble.oobPredict}

An N-by-K numeric matrix for N observations and K classes. A high score indicates that an observation is likely to come from this class. Scores are in the range 0 to 1 .

\section*{Definitions}

\section*{Out of Bag}

Bagging, which stands for "bootstrap aggregation", is a type of ensemble learning. To bag a weak learner such as a decision tree on a dataset, fitensemble generates many bootstrap replicas of the dataset and grows decision trees on these replicas. fitensemble obtains each bootstrap replica by randomly selecting N observations out of N with replacement, where \(N\) is the dataset size. To find the predicted response of a trained ensemble, predict take an average over predictions from individual trees.

Drawing N out of N observations with replacement omits on average \(37 \%\) ( \(1 / e\) ) of observations for each decision tree. These are "out-of-bag" observations. For each observation, oobLoss estimates the out-of-bag prediction by averaging over predictions from all trees in the ensemble for which this observation is out of bag. It then compares the computed prediction against the true response for this observation. It calculates the out-of-bag error by comparing the out-of-bag predicted responses against the true responses for all observations used for training. This out-of-bag average is an unbiased estimator of the true ensemble error.

\section*{Score (ensemble)}

For ensembles, a classification score represents the confidence of a classification into a class. The higher the score, the higher the confidence.

Different ensemble algorithms have different definitions for their scores. Furthermore, the range of scores depends on ensemble type. For example:
- AdaBoostM1 scores range from \(-\infty\) to \(\infty\).
- Bag scores range from 0 to 1 .

\section*{ClassificationBaggedEnsemble.oobPredict}
\begin{tabular}{|c|c|}
\hline Examples & Find the out-of-bag predictions and scores for the Fisher iris data. Find the scores in the range \((0.2,0.8)\); these are the scores where there is notable uncertainty in the resulting classifications.
```

load fisheriris
ens = fitensemble(meas,species,'Bag',100,...
'Tree','type','classification');
[label score] = oobPredict(ens);
unsure = ( (score > .2) \& (score < .8));
sum(sum(unsure)) % How many uncertain predictions?
ans =
1 6

``` \\
\hline See Also & oobMargin | oobPredict | oobLoss | oobEdge | predict \\
\hline How To & - Chapter 13, "Nonparametric Supervised Learning" \\
\hline
\end{tabular}

\section*{RegressionBaggedEnsemble.oobPredict}

Purpose Predict out-of-bag response of ensemble
Syntax \(\quad\) Yfit \(=\) oobPredict(ens)
Yfit = oobPredict(ens,Name, Value)

Description

Input
Arguments

\section*{Output Yfit}

Arguments
ens

Yfit \(=\) oobPredict(ens) returns the predicted responses for the out-of-bag data in ens.
Yfit \(=\) oobPredict(ens, Name, Value) predicts responses with additional options specified by one or more Name, Value pair arguments.

A regression bagged ensemble, constructed with fitensemble.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{learners}

Indices of weak learners in the ensemble ranging from 1 to NTrained. oobLoss uses only these learners for calculating loss.

Default: 1:NTrained

A vector of predicted responses for out-of-bag data. Yfit has size(ens.X,1) elements.

You can find the indices of out-of-bag observations for weak learner \(L\) with the command
```

~ens.UseObsForLearner(:,L)

```

\section*{RegressionBaggedEnsemble.oobPredict}

\section*{Definitions Out of Bag}

Bagging, which stands for "bootstrap aggregation", is a type of ensemble learning. To bag a weak learner such as a decision tree on a dataset, fitensemble generates many bootstrap replicas of the dataset and grows decision trees on these replicas. fitensemble obtains each bootstrap replica by randomly selecting \(N\) observations out of \(N\) with replacement, where \(N\) is the dataset size. To find the predicted response of a trained ensemble, predict take an average over predictions from individual trees.

Drawing N out of N observations with replacement omits on average \(37 \%\) ( \(1 / e\) ) of observations for each decision tree. These are "out-of-bag" observations. For each observation, oobLoss estimates the out-of-bag prediction by averaging over predictions from all trees in the ensemble for which this observation is out of bag. It then compares the computed prediction against the true response for this observation. It calculates the out-of-bag error by comparing the out-of-bag predicted responses against the true responses for all observations used for training. This out-of-bag average is an unbiased estimator of the true ensemble error.

\section*{Examples}

Compute out-of-bag predictions for the carsmall data. Look at the first three terms of the fit:
```

load carsmall
X = [Displacement Horsepower Weight];
ens = fitensemble(X,MPG,'bag',100,'Tree',...
'type','regression');
Yfit = oobPredict(ens);
Yfit(1:3) % first three terms
ans =
15.7964
14.7162
14.8062

```

See Also oobLoss | predict

\section*{RegressionBaggedEnsemble.oobPredict}

\author{
How To \\ - Chapter 13, "Nonparametric Supervised Learning"
}
\begin{tabular}{|c|c|}
\hline Purpose & Ensemble predictions for out-of-bag observations \\
\hline Syntax & \[
\begin{aligned}
& Y=\text { oobPredict }(B) \\
& Y=\text { oobPredict(B,'param1', val1,' param2', val2, } \ldots \text { ) }
\end{aligned}
\] \\
\hline \multirow[t]{4}{*}{Description} & \(Y=\) oobPredict( \(B\) ) computes predicted responses using the trained bagger B for out-of-bag observations in the training data. The output has one prediction for each observation in the training data. The returned \(Y\) is a cell array of strings for classification and a numeric array for regression. \\
\hline & \(Y=\) oobPredict (B,'param1', val1,'param2', val2, ...) specifies optional parameter name/value pairs: \\
\hline & trees ' Array of tree indices to use for computation of responses. Default is 'all'. \\
\hline & 'treeweights' Array of NTrees weights for weighting votes from the specified trees. \\
\hline See Also & CompactTreeBagger.predict | OOBIndices \\
\hline
\end{tabular}

\section*{Superclasses categorical}

Purpose Arrays for ordinal categorical data
Description Ordinal arrays are used to store discrete values that have an ordering but are not numeric. An ordinal array provides efficient storage and convenient manipulation of such data, while also maintaining meaningful labels for the values. Ordinal arrays are often used as grouping variables.

Like a numerical array, an ordinal array can have any size or dimension. You can subscript, concatenate, reshape, sort, etc. ordinal arrays, much like ordinary numeric arrays. You can make comparisons between elements of two ordinal arrays, or between an ordinal array and a single string representing a ordinal value.

\section*{Construction}

Use the ordinal constructor to create an ordinal array from a numeric, logical, or character array, or from a cell array of strings.
\[
\begin{array}{ll}
\text { ordinal } & \begin{array}{l}
\text { Construct ordinal categorical } \\
\text { array }
\end{array}
\end{array}
\]

\section*{Methods}

Each ordinal array carries along a list of possible values that it can store, known as its levels. The list is created when you create an ordinal array, and you can access it using the getlevels method, or modify it using the addlevels, mergelevels, or droplevels methods. Assignment to the array will also add new levels automatically if the values assigned are not already levels of the array. The ordering on values stored in an ordinal array is defined by the order of the list of levels. You can change that order using the reorderlevels method.

The following table lists operations available for ordinal arrays.
\begin{tabular}{ll} 
ismember & Test for membership \\
mergelevels & Merge levels
\end{tabular}
sort
sortrows

Sort elements of ordinal array
Sort rows

\section*{Inherited Methods}

Methods in the following table are inherited from categorical.
\begin{tabular}{ll} 
addlevels & Add levels to categorical array \\
cat \\
cellstr & \begin{tabular}{l} 
Concatenate categorical arrays \\
Convert categorical array to cell \\
array of strings
\end{tabular} \\
char & \begin{tabular}{l} 
Convert categorical array to \\
character array
\end{tabular} \\
circshift & Shift categorical array circularly \\
ctranspose & \begin{tabular}{l} 
Transpose categorical matrix
\end{tabular} \\
disp & \begin{tabular}{l} 
Display categorical array
\end{tabular} \\
display & \begin{tabular}{l} 
Display categorical array \\
Convert categorical array to \\
double array
\end{tabular} \\
double & \begin{tabular}{l} 
Drop levels
\end{tabular} \\
droplevels & \begin{tabular}{l} 
Last index in indexing expression \\
for categorical array
\end{tabular} \\
end & \begin{tabular}{l} 
Flip categorical array along \\
specified dimension
\end{tabular} \\
flipdim & \begin{tabular}{l} 
Flip categorical matrix in \\
left/right direction
\end{tabular} \\
fliplr & \begin{tabular}{l} 
Flip categorical matrix in \\
up/down direction
\end{tabular} \\
flipud & \begin{tabular}{l} 
Access categorical array labels \\
Get categorical array levels
\end{tabular} \\
getlabels & getlevels
\end{tabular}
\begin{tabular}{ll} 
hist & Plot histogram of categorical data \\
horzcat & \begin{tabular}{l} 
Horizontal concatenation for \\
categorical arrays \\
Convert categorical array to \\
signed 16-bit integer array
\end{tabular} \\
int16 & \begin{tabular}{l} 
Convert categorical array to \\
signed 32-bit integer array \\
Convert categorical array to \\
signed 64-bit integer array
\end{tabular} \\
int32 & \begin{tabular}{l} 
Convert categorical array to \\
signed 8-bit integer array
\end{tabular} \\
int64 & \begin{tabular}{l} 
Set intersection for categorical \\
arrays
\end{tabular} \\
int8 & \begin{tabular}{l} 
Inverse permute dimensions of \\
categorical array
\end{tabular} \\
intersect & \begin{tabular}{l} 
True for empty categorical array
\end{tabular} \\
ipermute & \begin{tabular}{l} 
True if categorical arrays are \\
equal
\end{tabular} \\
isempty & \begin{tabular}{l} 
Test for levels \\
isequal
\end{tabular} \\
True for elements of categorical \\
islevel & \begin{tabular}{l} 
array in set
\end{tabular} \\
ismember & True if categorical array is scalar \\
isscalar & Test for undefined elements
\end{tabular}
\(\left.\begin{array}{ll}\text { numel } & \begin{array}{l}\text { Number of elements in categorical } \\ \text { array }\end{array} \\ \text { permute } & \begin{array}{l}\text { Permute dimensions of } \\ \text { categorical array }\end{array} \\ \text { reorderlevels } & \begin{array}{l}\text { Reorder levels } \\ \text { Replicate and tile categorical } \\ \text { array }\end{array} \\ \text { repmat } & \begin{array}{l}\text { Resize categorical array } \\ \text { Rotate categorical matrix 90 } \\ \text { degrees }\end{array} \\ \text { reshape } & \begin{array}{l}\text { Set difference for categorical } \\ \text { arrays }\end{array} \\ \text { rot90 } & \begin{array}{l}\text { Label levels }\end{array} \\ \text { setdiff } & \begin{array}{l}\text { Set exclusive-or for categorical } \\ \text { arrays }\end{array} \\ \text { setlabels } & \begin{array}{l}\text { Shift dimensions of categorical } \\ \text { array }\end{array} \\ \text { setxor } & \begin{array}{l}\text { Convert categorical array to } \\ \text { single array }\end{array} \\ \text { shiftdim } & \begin{array}{l}\text { Size of categorical array } \\ \text { Squeeze singleton dimensions }\end{array} \\ \text { single } & \begin{array}{l}\text { Srom categorical array } \\ \text { fubscripted assignment for }\end{array} \\ \text { size } & \begin{array}{l}\text { Subegorical array } \\ \text { cater } \\ \text { squeeze }\end{array} \\ \text { subscript index for categorical } \\ \text { array } \\ \text { subscripted reference for } \\ \text { categorical array }\end{array}\right\}\)

\section*{ordinal}
\begin{tabular}{ll} 
summary & \begin{tabular}{l} 
Summary statistics for categorical \\
array
\end{tabular} \\
times & \begin{tabular}{l} 
Product of categorical arrays
\end{tabular} \\
transpose & \begin{tabular}{l} 
Transpose categorical matrix \\
Convert categorical array to \\
unsigned 16-bit integers \\
Convert categorical array to \\
unsigned 32-bit integers
\end{tabular} \\
uint32 & \begin{tabular}{l} 
Convert categorical array to \\
unsigned 64-bit integers
\end{tabular} \\
uint64 & \begin{tabular}{l} 
Convert categorical array to \\
unsigned 8-bit integers
\end{tabular} \\
uint8 & \begin{tabular}{l} 
Set union for categorical arrays
\end{tabular} \\
union & \begin{tabular}{l} 
Unique values in categorical \\
array
\end{tabular} \\
unique & \begin{tabular}{l} 
Vertical concatenation for \\
categorical arrays
\end{tabular} \\
vertcat &
\end{tabular}

\section*{Properties Inherited Properties}

Properties in the following table are inherited from categorical.
\begin{tabular}{ll} 
labels & Text labels for levels \\
undeflabel & Text label for undefined levels
\end{tabular}

Copy
Semantics

Examples

Value. To learn how this affects your use of the class, see Comparing Handle and Value Classes in the MATLAB Object-Oriented Programming documentation.

Create an ordinal array from integer data:
```

quality = ordinal([1 2 3; 3 2 1; 2 1 3],{'low' 'medium' 'high'})

```
```

% Find elements meeting a criterion
quality >= 'medium'
ismember(quality,{'low' 'high'})
% Compare two ordinal arrays
quality2 = fliplr(quality)
quality == quality2

```

\author{
References [1] Johnson, N. L., S. Kotz, and A. W. Kemp, Univariate Discrete Distributions, 2nd edition, Wiley, 1992, pp. 124-130.
}

See Also histc | nominal
\begin{tabular}{ll} 
Purpose & Construct ordinal categorical array \\
Syntax & \(B=\operatorname{ordinal}(A)\) \\
& \(B=\operatorname{ordinal}(A\), labels \()\) \\
& \(B=\operatorname{ordinal}(A\), labels, levels \()\) \\
& \(B=\operatorname{ordinal}(A\), labels, [], edges \()\)
\end{tabular}

\section*{Description}
\(B=\) ordinal(A) creates an ordinal array \(B\) from the array \(A\). A can be numeric, logical, character, categorical, or a cell array of strings. ordinal creates the levels of \(B\) from the sorted unique values in \(A\), and creates default labels for them.
\(B=\) ordinal(A, labels) labels the levels in B using the character array or cell array of strings labels. ordinal assigns labels to levels in B in order according to the sorted unique values in A .
\(B=\) ordinal(A, labels, levels) creates an ordinal array with possible levels defined by levels. levels is a vector whose values can be compared to those in A using the equality operator. ordinal assigns labels to each level from the corresponding elements of labels. If A contains any values not present in levels, the levels of the corresponding elements of B are undefined. Use [] for labels to allow ordinal to create default labels.
\(B=\) ordinal(A, labels, [], edges) creates an ordinal array by binning the numeric array \(A\), with bin edges given by the numeric vector edges. The uppermost bin includes values equal to the right-most edge. ordinal assigns labels to each level in B from the corresponding elements of labels. edges must have one more element than labels.

By default, an element of \(B\) is undefined if the corresponding element of A is NaN (when A is numeric), an empty string (when A is character), or undefined (when A is categorical). ordinal treats such elements as "undefined" or "missing" and does not include entries for them among the possible levels for B. To create an explicit level for such elements instead of treating them as undefined, you must use the levels input, and include NaN, the empty string, or an undefined element.

You may include duplicate labels in labels in order to merge multiple values in A into a single level in B.

\section*{Examples}

Create an ordinal array from integer data, and provide explicit labels:
```

quality1 = ordinal([1 2 3; 3 2 1; 2 1 3],...
{'low' 'medium' 'high'})

```

Create an ordinal array from integer data, and provide both explicit labels and an explicit order:
```

quality2 = ordinal([1 2 3; 3 2 1; 2 1 3],...
{'high' 'medium' 'low'},[3 2 1])

```

Create an ordinal array by binning floating point values:
```

size = ordinal(rand(5,2),{'small' 'medium' 'large'},...
[],[0 1/3 2/3 1])

```

Create an ordinal array from the measurements in Fisher's iris data, ignoring decimal lengths:
```

load fisheriris
m = floor(min(meas(:)));
M = floor(max(meas(:)));
labels = num2str((m:M)');
edges = m:M+1;
cms = ordinal(meas,labels,[],edges)
meas(1:5,:)
ans =

| 5.1000 | 3.5000 | 1.4000 | 0.2000 |
| :--- | :--- | :--- | :--- |
| 4.9000 | 3.0000 | 1.4000 | 0.2000 |

```
\begin{tabular}{rrrr}
4.7000 & 3.2000 & 1.3000 & 0.2000 \\
4.6000 & 3.1000 & 1.5000 & 0.2000 \\
5.0000 & 3.6000 & 1.4000 & 0.2000 \\
cms \((1.5,:)\) & & & \\
ans = & & & \\
5 & 3 & 1 & 0 \\
4 & 3 & 1 & 0 \\
4 & 3 & 1 & 0 \\
4 & 3 & 1 & 0
\end{tabular}

Create an age group ordinal array from the data in hospital.mat:
load hospital
edges = 0:10:100;
labels = strcat(num2str((0:10:90)', '\%d'), \{'s'\});
AgeGroup = ordinal(hospital.Age,labels,[],edges);
hospital.Age(1:5)
ans =
38
43
38
40
49

AgeGroup (1:5)
ans =
30s
40s
30s
40s
40s
See Also
histc | nominal

\section*{CompactTreeBagger.outlierMeasure}
\begin{tabular}{|c|c|}
\hline Purpose & Outlier measure for data \\
\hline Syntax & ```
out = outlierMeasure(B,X)
out = outlierMeasure(B,X,'param1',val1,'param2',val2,...)
``` \\
\hline \multirow[t]{4}{*}{Description} & out = outlierMeasure \((B, X)\) computes outlier measures for predictors \(X\) using trees in the ensemble B. The method computes the outlier measure for a given observation by taking an inverse of the average squared proximity between this observation and other observations. outlierMeasure then normalizes these outlier measures by subtracting the median of their distribution, taking the absolute value of this difference, and dividing by the median absolute deviation. A high value of the outlier measure indicates that this observation is an outlier. \\
\hline & \begin{tabular}{l}
You can supply the proximity matrix directly by using the 'data' parameter. \\
out = outlierMeasure (B, X,'param1', val1,'param2', val2, ...) specifies optional parameter name/value pairs:
\end{tabular} \\
\hline & \begin{tabular}{l}
'data' \\
Flag indicating how to treat the X input argument. If set to 'predictors' (default), the method assumes \(X\) is a matrix of predictors and uses it for computation of the proximity matrix. If set to 'proximity', the method treats \(X\) as a proximity matrix returned by the proximity method. If you do not supply the proximity matrix, outlierMeasure computes it internally. If you use the proximity method to compute a proximity matrix, supplying it as input to outlierMeasure reduces computing time.
\end{tabular} \\
\hline & \begin{tabular}{l}
'labels' \\
Vector of true class labels. True class labels can be either a numeric vector, character matrix, or cell array of strings. When you supply this parameter, the method performs the outlier calculation for any observations using only other observations from the same class. This parameter must specify one label for each observation (row) in \(X\).
\end{tabular} \\
\hline
\end{tabular}

\section*{CompactTreeBagger.outlierMeasure}

See Also proximity

Purpose Measure for determining outliers
Description \(\quad \begin{aligned} & \text { The OutlierMeasure property is a numeric array of size Nobs-by-1, } \\ & \text { where Nobs is the number of observations in the training data, } \\ & \text { containing outlier measures for each observation. }\end{aligned}\)
See Also CompactTreeBagger.outlierMeasure

\section*{parallelcoords}

Purpose
Parallel coordinates plot
Syntax
```

parallelcoords(X)
parallelcoords(X,...,'Standardize','on')
parallelcoords(X,...,'Standardize','PCA')
parallelcoords(X,...,'Standardize','PCAStd')
parallelcoords(X,...,'Quantile',alpha)
parallelcoords(X,...,'Group',group)
parallelcoords(X,...,'Labels',labels)
parallelcoords(X,···,PropertyName,PropertyValue,...)
h = parallelcoords(X,...)
parallelcoords(axes,...)

```

\section*{Description}
parallelcoords (X) creates a parallel coordinates plot of the multivariate data in the \(n\)-by- \(p\) matrix X . Rows of X correspond to observations, columns to variables. A parallel coordinates plot is a tool for visualizing high dimensional data, where each observation is represented by the sequence of its coordinate values plotted against their coordinate indices. parallelcoords treats NaNs in X as missing values and does not plot those coordinate values.
parallelcoords(X, ..,'Standardize', 'on') scales each column of X to have mean 0 and standard deviation 1 before making the plot.
parallelcoords(X, .., 'Standardize', 'PCA') creates a parallel coordinates plot from the principal component scores of \(X\), in order of decreasing eigenvalues. parallelcoords removes rows of \(X\) containing missing values (NaNs) for principal components analysis (PCA) standardization.
parallelcoords(X, ...,'Standardize', 'PCAStd') creates a parallel coordinates plot using the standardized principal component scores.
parallelcoords(X,..., 'Quantile', alpha) plots only the median and the alpha and 1 -alpha quantiles of \(f(t)\) at each value of \(t\). This is useful if \(X\) contains many observations.
parallelcoords(X,..., 'Group', group) plots the data in different groups with different colors. Groups are defined by group, a numeric
array containing a group index for each observation. (See "Grouped Data" on page 2-34.) group can also be a categorical variable, character matrix, or cell array of strings, containing a group name for each observation.
parallelcoords(X, ..., 'Labels', labels) labels the coordinate tick marks along the horizontal axis using labels, a character array or cell array of strings.
parallelcoords(X,...,PropertyName,PropertyValue,...) sets properties to the specified property values for all line graphics objects created by parallelcoords.
\(h=\) parallelcoords \((X, \ldots)\) returns a column vector of handles to the line objects created by parallelcoords, one handle per row of \(X\). If you use the 'Quantile' input argument, \(h\) contains one handle for each of the three lines objects created. If you use both the 'Quantile' and the 'Group ' input arguments, h contains three handles for each group.
parallelcoords(axes,...) plots into the axes with handle axes.

\section*{Examples}
```

% Make a grouped plot of the raw data
load fisheriris
labels = {'Sepal Length','Sepal Width',...
'Petal Length','Petal Width'};
parallelcoords(meas,'group',species,'labels',labels);
% Plot only the median and quartiles of each group
parallelcoords(meas,'group',species,'labels',labels,...
'quantile',.25);

```

\section*{See Also \\ andrewsplot | glyphplot}

How To
- "Grouped Data" on page 2-34

\section*{ProbDistUnivParam.paramci}
\begin{tabular}{ll} 
Purpose & Return parameter confidence intervals of ProbDistUnivParam object \\
Syntax & \begin{tabular}{l} 
CI = paramci(PD) \\
CI = paramci(PD, Alpha)
\end{tabular} \\
Description & \begin{tabular}{l} 
CI = paramci(PD) returns CI, a 2-by-N array containing \(95 \%\) \\
confidence intervals for the parameters of the ProbDistUnivParam \\
object PD. \(N\) is the number of parameters in the distribution. \\
When you create PD by specifying parameters (such as using the
\end{tabular} \\
\begin{tabular}{l} 
ProbDistUnivParam. ProbDistUnivParam constructor or using the \\
fitdist function and specifying a 'binomial' or 'generalized \\
pareto' distribution) rather than by fitting to data, the confidence \\
intervals have a width of 0 because the parameters are viewed as \\
estimates of an unknown parameter.
\end{tabular} \\
\begin{tabular}{l} 
CI = paramci(PD, Alpha) returns 100*(1 - Alpha) \(\%\) confidence \\
intervals. Default Alpha is 0.05, which specifies \(95 \%\) confidence \\
intervals.
\end{tabular}
\end{tabular}

Note If you create PD with a distribution that does not support confidence intervals, then CI contains NaN values.

\section*{Input \\ Arguments}

\section*{PD}

Alpha

An object of the class ProbDistUnivParam.
A value between 0 and 1 that specifies a confidence interval. Default is 0.05 , which specifies \(95 \%\) confidence intervals.

A 2 -by-N array containing 100*(1 - Alpha) \% confidence intervals for the parameters of the ProbDistUnivParam object PD. \(N\) is the number of parameters in the distribution.

See Also fitdist

\section*{ProbDistParametric.ParamCov property}
\begin{tabular}{ll} 
Purpose & \begin{tabular}{l} 
Read-only covariance matrix of parameter estimates of \\
ProbDistParametric object
\end{tabular} \\
Description \(\quad\)\begin{tabular}{l} 
ParamCov is a read-only property of the ProbDistParametric class. \\
ParamCov is a covariance matrix containing the parameter estimates of \\
a distribution represented by a ProbDistParametric object. ParamCov \\
has a size of NumParams-by-NumParams.
\end{tabular} \\
Values & \begin{tabular}{l} 
This covariance matrix includes estimates for both the specified \\
parameters and parameters that are fit to the data. For specified \\
parameters, the covariance is 0, indicating the parameter is known \\
exactly. Use this information to view and compare the descriptions of \\
parameters supplied to create distributions.
\end{tabular}
\end{tabular}

\section*{ProbDistParametric.ParamDescription property}
\begin{tabular}{ll} 
Purpose & \begin{tabular}{l} 
Read-only cell array specifying descriptions of parameters of \\
ProbDistParametric object
\end{tabular} \\
Description \(\quad\)\begin{tabular}{l} 
ParamDescription is a read-only property of the ProbDistParametric \\
class. ParamDescription is a cell array of strings specifying the \\
descriptions or meanings of the parameters of a distribution represented \\
by a ProbDistParametric object. ParamDescription has a length of \\
NumParams.
\end{tabular} \\
Values & \begin{tabular}{l} 
This cell array includes a brief description of the meaning of both \\
the specified parameters and parameters that are fit to the data. \\
The description is the same as the parameter name when no further \\
description information is available. Use this information to view and \\
compare the descriptions of parameters used to create distributions.
\end{tabular}
\end{tabular}

\section*{ProbDistParametric.ParamlsFixed property}
\begin{tabular}{ll} 
Purpose & \begin{tabular}{l} 
Read-only logical array specifying fixed parameters of \\
ProbDistParametric object
\end{tabular} \\
Description & \begin{tabular}{l} 
ParamIsFixed is a read-only property of the ProbDistParametric class. \\
ParamIsFixed is a logical array specifying the fixed parameters of a \\
distribution represented by a ProbDistParametric object. ParamIsFixed \\
has a length of NumParams.
\end{tabular} \\
Values & \begin{tabular}{l} 
This array specifies a 1 (true) for fixed parameters, and a 0 (false) \\
for parameters that are estimated from the input data. Use this \\
information to view and compare the fixed parameters used to create \\
distributions.
\end{tabular}
\end{tabular}

\section*{ProbDistParametric.ParamNames property}
\begin{tabular}{ll} 
Purpose & \begin{tabular}{l} 
Read-only cell array specifying names of parameters of \\
ProbDistParametric object
\end{tabular} \\
Description & \begin{tabular}{l} 
ParamNames is a read-only property of the ProbDistParametric class. \\
ParamNames is a cell array of strings specifying the names of the \\
parameters of a distribution represented by a ProbDistParametric \\
object. ParamNames has a length of NumParams.
\end{tabular} \\
Values & \begin{tabular}{l} 
This cell array includes the names of both the specified parameters and \\
parameters that are fit to the data. Use this information to view and \\
compare the names of parameters used to create distributions.
\end{tabular}
\end{tabular}

\section*{NaiveBayes.Params property}

\section*{Purpose Parameter estimates}

\section*{Description}

The Params property is an NClasses-by-NDims cell array containing the parameter estimates, excluding the class priors. Params(i,j) contains the parameter estimates for the \(j\) th feature in the ith class. Params ( \(i, j\) ) is an empty cell if the ith class is empty.

The entry in Params ( \(i, j\) ) depends on the distribution type used for the \(j\) th feature, as follows:
\begin{tabular}{|c|c|}
\hline 'normal' & A vector of length two. The first element is the mean, and the second element is standard deviation \\
\hline 'kernel' & A ProbDistUnivKernel object \\
\hline 'mvmn' & A vector containing the probability for each possible value of the \(j\) th feature in the ith class. The order of the probabilities is decided by the sorted order of all the unique values of the \(j\) th feature. \\
\hline 'mn' & A scalar representing the probability the \(j\) th token appearing in the ith class, Prob(token j | class i). It is estimated as (1 + the number of occurrence of token \(J\) in class I)/(NDims + the total number of token occurrence in class I). \\
\hline
\end{tabular}

\section*{ProbDistParametric.Params property}
\begin{tabular}{ll} 
Purpose & \begin{tabular}{l} 
Read-only array specifying values of parameters of ProbDistParametric \\
object
\end{tabular} \\
Description \(\quad\)\begin{tabular}{l} 
Params is a read-only property of the ProbDistParametric class. \\
Params is an array of values specifying the values of the parameters of \\
a distribution represented by a ProbDistParametric object. Params \\
has a length of NumParams.
\end{tabular} \\
Values & \begin{tabular}{l} 
This array includes the values of both the specified parameters and \\
parameters that are fit to the data. Use this information to view and \\
compare the values of parameters used to create distributions.
\end{tabular}
\end{tabular}

\section*{classregtree.parent}

Purpose Parent node
Syntax
p = parent(t)
p = parent(t, nodes)

Description
\(p=\) parent ( \(t\) ) returns an \(n\)-element vector \(p\) containing the number of the parent node for each node in the tree \(t\), where \(n\) is the number of nodes. The parent of the root node is 0 .
\(p\) = parent(t, nodes) takes a vector nodes of node numbers and returns the parent nodes for the specified nodes.

Examples Create a classification tree for Fisher's iris data:
```

load fisheriris;
t = classregtree(meas,species,...
'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 else node 3
2 class = setosa
3 if PW<1.75 then node 4 else node 5
4 if PL<4.95 then node 6 else node 7
5 class = virginica
6 if PW<1.65 then node 8 else node 9
7 class = virginica
8 class = versicolor
9 class = virginica
view(t)

```
\begin{tabular}{|c|c|c|c|c|c|c|c|c|}
\hline & & & & & & & & - \\
\hline Click to display: & Identity & \(\checkmark\) & Magnification: & 100\% & \(\checkmark\) & Pruning level: & 0 of 4 & \\
\hline
\end{tabular}

\[
\begin{aligned}
& p=\text { parent }(t) \\
& p=
\end{aligned}
\]
\[
0
\]
\[
1
\]
\[
1
\]
\[
3
\]
\[
3
\]
\[
4
\]
\[
4
\]
\[
6
\]

\section*{classregtree.parent}

\section*{6}

\section*{References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.}

\author{
See Also \\ classregtree | children | numnodes
}

\section*{Purpose Pareto chart}

\section*{Syntax pareto(y,names) \\ [h,ax] = pareto(...)}

Description

Examples
pareto ( \(y\), names) displays a Pareto chart where the values in the vector y are drawn as bars in descending order. Each bar is labeled with the associated value in the string matrix or cell array, names. pareto (y) labels each bar with the index of the corresponding element in y .

The line above the bars shows the cumulative percentage.
[ \(\mathrm{h}, \mathrm{ax}\) ] = pareto(...) returns a combination of patch and line object handles to the two axes created in ax.

Create a Pareto chart from data measuring the number of manufactured parts rejected for various types of defects.
```

defects = {'pits';'cracks';'holes';'dents'};
quantity = [5 3 19 25];
pareto(quantity,defects)

```


See Also bar | hist

\title{
Superclasses piecewisedistribution
}

Purpose Empirical distributions with Pareto tails

\section*{Construction}
paretotails
Construct Pareto tails object

Methods
lowerparams
Lower Pareto tails parameters
upperparams
Upper Pareto tails parameters

\section*{Inherited Methods}

Methods in the following table are inherited from piecewisedistribution.
\begin{tabular}{|c|c|}
\hline boundary & Piecewise distribution boundaries \\
\hline cdf & Cumulative distribution function for piecewise distribution \\
\hline disp & Display piecewisedistribution object \\
\hline display & Display piecewisedistribution object \\
\hline icdf & Inverse cumulative distribution function for piecewise distribution \\
\hline nsegments & Number of segments \\
\hline pdf & Probability density function for piecewise distribution \\
\hline random & Random numbers from piecewise distribution \\
\hline segment & Segments containing values \\
\hline
\end{tabular}

Properties

Copy
Semantics

How To
- "Generalized Pareto Distribution" on page B-37

Objects of the paretotails class have no properties accessible by dot indexing, get methods, or set methods. To obtain information about a paretotails object, use the appropriate method.

Value. To learn how this affects your use of the class, see Comparing Handle and Value Classes in the MATLAB Object-Oriented Programming documentation.

\section*{Purpose \\ Syntax \\ Description}

Construct Pareto tails object
obj = paretotails(x,pl,pu)
obj = paretotails(x,pl,pu,cdffun)
obj = paretotails( \(\mathrm{x}, \mathrm{pl}, \mathrm{pu}\) ) creates an object obj defining a distribution consisting of the empirical distribution of \(x\) in the center and Pareto distributions in the tails. \(x\) is a real-valued vector of data values whose extreme observations are fit to generalized Pareto distributions (GPDs). pl and pu identify the lower- and upper-tail cumulative probabilities such that \(100 * \mathrm{pl}\) and 100*(1-pu) percent of the observations in \(x\) are, respectively, fit to a GPD by maximum likelihood. If pl is 0 , or if there are not at least two distinct observations in the lower tail, then no lower Pareto tail is fit. If pu is 1 , or if there are not at least two distinct observations in the upper tail, then no upper Pareto tail is fit.
obj \(=\) paretotails(x, pl, pu,cdffun) uses cdffun to estimate the cdf of \(x\) between the lower and upper tail probabilities. cdffun may be any of the following:
- 'ecdf' - Uses an interpolated empirical cdf, with data values as the midpoints in the vertical steps in the empirical cdf, and computed by linear interpolation between data values. This is the default.
- 'kernel' - Uses a kernel-smoothing estimate of the cdf.
- @fun - Uses a handle to a function of the form [p,xi] = fun (x) that accepts the input data vector \(x\) and returns a vector \(p\) of cdf values and a vector xi of evaluation points. Values in xi must be sorted and distinct but need not equal the values in \(x\).
cdffun is used to compute the quantiles corresponding to pl and pu by inverse interpolation, and to define the fitted distribution between these quantiles.
The output object obj is a Pareto tails object with methods to evaluate the cdf, inverse cdf, and other functions of the fitted distribution. These methods are well-suited to copula and other Monte Carlo simulations.

The pdf method in the tails is the GPD density, but in the center it is computed as the slope of the interpolated cdf.

The paretotails class is a subclass of the piecewisedistribution class, and many of its methods are derived from that class.

\section*{Examples}

Fit Pareto tails to a \(t\) distribution at cumulative probabilities 0.1 and 0.9:
```

t = trnd(3,100,1);
obj = paretotails(t,0.1,0.9);
[p,q] = boundary(obj);
x = linspace(-5,5);
plot(x,cdf(obj,x),'b-','LineWidth',2)
hold on
plot(x,tcdf(x,3),'r:','LineWidth',2)
plot(q,p,'bo','LineWidth',2,'MarkerSize',5)
legend('Pareto Tails Object','t Distribution',...
'Location','NW')

```


See Also
cdf | ecdf | gpfit | icdf | ksdensity

Purpose
Linear or rank partial correlation coefficients
Syntax
RHO = partialcorr (X)
RHO = partialcorr (X,Z)
RHO = partialcorr(X,Y,Z)
[RHO, PVAL] = partialcorr(...)
[...] = partialcorr(...,param1,val1,param2,val2,...)

\section*{Description}

RHO = partialcorr \((X)\) returns the sample linear partial correlation
coefficients between pairs of variables in X , controlling for the remaining variables in X . X is an \(n\)-by- \(p\) matrix, with rows corresponding to observations, and columns corresponding to variables. RHO is a symmetric \(p\)-by-p matrix, where the \((i, j)\)-th entry is the sample linear partial correlation between the \(i\)-th and \(j\)-th columns in \(X\).
RHO = partialcorr ( \(X, Z\) ) returns the sample linear partial correlation coefficients between pairs of variables in \(X\) controlling for the variables in Z . X is an \(n\)-by \(-p\) matrix, and Z is an \(n\)-by- \(q\) matrix with rows corresponding to observations, and columns corresponding to variables. The output, RHO, is a symmetric \(p\)-by- \(p\) matrix.

RHO = partialcorr (X,Y,Z) returns the sample linear partial correlation coefficients between pairs of variables between \(X\) and Y , controlling for the variables in Z . X is an \(n\)-by- \(p_{1}\) matrix, Y an \(n\)-by- \(p_{2}\) matrix, and \(Z\) is an \(n\)-by- \(q\) matrix, with rows corresponding to observations, and columns corresponding to variables. RHO is a \(p_{1}\)-by- \(p_{2}\) matrix, where the \((i, j)\) th entry is the sample linear partial correlation between the \(i\) th column in X and the \(j\) th column in Y .

If the covariance matrix of \([X, Z]\) is
\[
S=\left(\begin{array}{cc}
S_{11} & S_{12} \\
S_{12}^{T} & S_{22}
\end{array}\right)
\]
then the partial correlation matrix of X , controlling for Z , can be defined formally as a normalized version of the covariance matrix \(S_{x y}=S_{11}\) \(-\left(S_{12} S_{22}{ }^{-1} S_{12}{ }^{T}\right)\)
[RHO,PVAL] = partialcorr(...) also returns PVAL, a matrix of \(p\)-values for testing the hypothesis of no partial correlation against the alternative that there is a nonzero partial correlation. Each element of PVAL is the \(p\) value for the corresponding element of RHO. If PVAL ( \(\mathrm{I}, \mathrm{J})\) is small, say less than 0.05 , then the partial correlation, \(\mathrm{RHO}(\mathrm{I}, \mathrm{J})\), is significantly different from zero.
[...] = partialcorr(...,param1,val1,param2,val2,...) specifies additional parameters and their values. Valid parameter/value pars are listed in the following table.
\begin{tabular}{|c|c|}
\hline Parameter & Values \\
\hline 'type' & \begin{tabular}{l}
- 'Pearson' - To compute Pearson (linear) partial correlations. This is the default. \\
- 'Spearman ' - To compute Spearman (rank) partial correlations.
\end{tabular} \\
\hline 'rows ' & \begin{tabular}{l}
- 'all' - To use all rows regardless of missing ( NaN ) values. This is the default. \\
- 'complete ' - To use only rows with no missing values. \\
- 'pairwise' - To compute RHO (I , J) using rows with no missing values in column I or \(J\).
\end{tabular} \\
\hline \begin{tabular}{l}
'tail' \\
The alternative hypothesis against which to compute \(p\)-values for testing the hypothesis of no partial correlation.
\end{tabular} & \begin{tabular}{l}
- 'both' (the default) - the correlation is not zero. \\
- 'right ' - the correlation is greater than zero. \\
- 'left' - the correlation is less than zero.
\end{tabular} \\
\hline
\end{tabular}

\section*{partialcorr}

A 'pairwise' value for the rows parameter can produce a RHO that is not positive definite. A 'complete' value always produces a positive definite RHO, but when data is missing, the estimates will be based on fewer observations, in general.
partialcorr computes \(p\)-values for linear and rank partial correlations using a Student's \(t\) distribution for a transformation of the correlation. This is exact for linear partial correlation when \(X\) and \(Z\) are normal, but is a large-sample approximation otherwise.

\section*{See Also}
corr | tiedrank | corrcoef

\section*{Purpose \\ Principal component analysis on covariance matrix}

\section*{Syntax}

COEFF = pcacov(V)
[COEFF,latent] = pcacov(V)
[COEFF,latent,explained] = pcacov(V)

\section*{Description}

Examples
COEFF \(=\operatorname{pcacov}(\mathrm{V})\) performs principal components analysis on the p-by-p covariance matrix \(V\) and returns the principal component coefficients, also known as loadings. COEFF is a p-by-p matrix, with columns are in order of decreasing component variance.
pcacov does not standardize \(V\) to have unit variances. To perform principal components analysis on standardized variables, use the data matrix, use princomp. principal component variances, that is, the eigenvalues of V . containing the percentage of the total variance explained by each principal component. each column containing coefficients for one principal component. The correlation matrix \(R=V . /\left(S D^{*} S D^{\prime}\right)\), where \(S D=\operatorname{sqrt}(\operatorname{diag}(V))\), in place of \(V\). To perform principal components analysis directly on the
[COEFF, latent] = pcacov(V) returns latent, a vector containing the
[COEFF,latent, explained] = pcacov(V) returns explained, a vector
```

load hald
covx = cov(ingredients);
[COEFF,latent,explained] = pcacov(covx)
COEFF =
0.0678 -0.6460 0.5673-0.5062
0.6785-0.0200 -0.5440-0.4933
-0.0290 0.7553 0.4036-0.5156
-0.7309 -0.1085 -0.4684 -0.4844
latent =
517.7969
67.4964
12.4054

```
0.2372
explained =
86.5974
11.2882
2.0747
0.0397

References [1] Jackson, J. E. A User's Guide to Principal Components. Hoboken, NJ: John Wiley and Sons, 1991.
[2] Jolliffe, I. T. Principal Component Analysis. 2nd ed., New York: Springer-Verlag, 2002.
[3] Krzanowski, W. J. Principles of Multivariate Analysis: A User's Perspective. New York: Oxford University Press, 1988.
[4] Seber, G. A. F., Multivariate Observations, Wiley, 1984.
See Also barttest | biplot | factoran | pcares | princomp | rotatefactors

\section*{Purpose}

Residuals from principal component analysis

\section*{Syntax}
```

residuals = pcares(X,ndim)
[residuals,reconstructed] = pcares(X,ndim)

```
residuals \(=\) pcares \((X\), ndim) returns the residuals obtained by retaining ndim principal components of the n-by-p matrix \(X\). Rows of \(X\) correspond to observations, columns to variables. ndim is a scalar and must be less than or equal to p. residuals is a matrix of the same size as \(X\). Use the data matrix, not the covariance matrix, with this function.
pcares does not normalize the columns of \(X\). To perform the principal components analysis based on standardized variables, that is, based on correlations, use pcares(zscore(X), ndim). You can perform principal components analysis directly on a covariance or correlation matrix, but without constructing residuals, by using pcacov.
[residuals, reconstructed] \(=\) pcares(X, ndim) returns the reconstructed observations; that is, the approximation to \(X\) obtained by retaining its first ndim principal components.

\section*{Examples}

This example shows the drop in the residuals from the first row of the Hald data as the number of component dimensions increases from one to three.
```

load hald
r1 = pcares(ingredients,1);
r2 = pcares(ingredients,2);
r3 = pcares(ingredients,3);
r11 = r1(1,:)
r11 =
2.0350
r21 = r2(1,:)
r21 =
-2.4037 2.6930 -1.6482 2.3425

```
```

$r 31=r 3(1,:)$
r31 =
$0.2008 \quad 0.1957 \quad 0.2045 \quad 0.1921$

```
References

[1] Jackson, J. E., A User's Guide to Principal Components, John Wiley
 and Sons, 1991.
[2] Jolliffe, I. T., Principal Component Analysis, 2nd Edition, Springer, 2002.
[3] Krzanowski, W. J. Principles of Multivariate Analysis: A User's Perspective. New York: Oxford University Press, 1988.
[4] Seber, G. A. F. Multivariate Observations. Hoboken, NJ: John Wiley \& Sons, Inc., 1984.

See Also factoran | pcacov | princomp

\section*{gmdistribution.PComponents property}
\(\begin{array}{ll}\text { Purpose } & \text { Input vector of mixing proportions } \\ \text { Description } & \text { Optional input vector of mixing proportions } p \text {, or its default value. }\end{array}\)

Purpose Probability density functions
Syntax \(\quad \begin{aligned} Y & =\operatorname{pdf}(\text { name }, X, A) \\ Y & =\operatorname{pdf}(\text { name }, X, A, B) \\ & Y\end{aligned}\)

\section*{Description}
\(Y=\operatorname{pdf}(\) name \(, X, A)\) computes the probability density function for the one-parameter family of distributions specified by name. Parameter values for the distribution are given in A. Densities are evaluated at the values in X and returned in Y .

If \(X\) and \(A\) are arrays, they must be the same size. If \(X\) is a scalar, it is expanded to a constant matrix the same size as \(A\). If \(A\) is a scalar, it is expanded to a constant matrix the same size as \(X\).
\(Y\) is the common size of \(X\) and \(A\) after any necessary scalar expansion.
\(Y=p d f(\) name \(, X, A, B)\) computes the probability density function for two-parameter families of distributions, where parameter values are given in \(A\) and \(B\).

If \(X, A\), and \(B\) are arrays, they must be the same size. If \(X\) is a scalar, it is expanded to a constant matrix the same size as A and B. If either A or B are scalars, they are expanded to constant matrices the same size as \(X\).
\(Y\) is the common size of \(X, A\), and \(B\) after any necessary scalar expansion.
\(Y=p d f(\) name \(, X, A, B, C)\) computes the probability density function for three-parameter families of distributions, where parameter values are given in \(A, B\), and \(C\).

If \(X, A, B\), and \(C\) are arrays, they must be the same size. If \(X\) is a scalar, it is expanded to a constant matrix the same size as \(A, B\), and \(C\). If any of A, B or C are scalars, they are expanded to constant matrices the same size as \(X\).
\(Y\) is the common size of \(X, A, B\) and \(C\) after any necessary scalar expansion.
Acceptable strings for name are:
\begin{tabular}{l|l|l|l|l}
\hline name & Distribution & \begin{tabular}{l} 
Input \\
Parameter A
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter B
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter \\
C
\end{tabular} \\
\hline 'beta' or 'Beta' & \begin{tabular}{l} 
"Beta \\
Distribution" \\
on page B-4
\end{tabular} & a & b & - \\
\hline \begin{tabular}{l} 
'bino' or \\
'Binomial'
\end{tabular} & \begin{tabular}{l} 
"Binomial \\
Distribution" \\
on page B-7
\end{tabular} & \begin{tabular}{l}
\(\mathrm{n}:\) number of \\
trials
\end{tabular} & \begin{tabular}{l}
\(\mathrm{p}:\) probability of \\
success for each \\
trial
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'chi2' or \\
'Chisquare'
\end{tabular} & \begin{tabular}{l} 
"Chi-Square \\
Distribution" on \\
page B-12
\end{tabular} & \begin{tabular}{l}
\(v:\) degrees of \\
freedom
\end{tabular} & - & - \\
\hline \begin{tabular}{l} 
'exp' or \\
'Exponential'
\end{tabular} & \begin{tabular}{l} 
"Exponential \\
Distribution" on \\
page B-16
\end{tabular} & \begin{tabular}{l}
\(\mu:\) mean
\end{tabular} & - & - \\
\hline \begin{tabular}{l} 
'ev' or 'Extreme \\
Value'
\end{tabular} & \begin{tabular}{l} 
"Extreme Value \\
Distribution" on \\
page B-19
\end{tabular} & \begin{tabular}{l}
\(\mu:\) location \\
parameter
\end{tabular} & \begin{tabular}{l}
\(\sigma:\) scale \\
parameter
\end{tabular} & - \\
\hline 'f' or 'F' & \begin{tabular}{l} 
"F Distribution" \\
on page B-25
\end{tabular} & \begin{tabular}{l}
\(v 1:\) numerator \\
degrees of \\
freedom
\end{tabular} & \begin{tabular}{l}
\(v 2:\) denominator \\
degrees of \\
freedom
\end{tabular} & - \\
\hline 'gam' or 'Gamma' & \begin{tabular}{l} 
"Gamma \\
Distribution" \\
on page B-27
\end{tabular} & \begin{tabular}{l} 
a: shape \\
parameter
\end{tabular} & \begin{tabular}{l} 
b: scale \\
parameter
\end{tabular} & - \\
\hline \begin{tabular}{l} 
"gev' or \\
'Generalized \\
'Generalized \\
Extreme Value' \\
Distribution" on \\
page B-32
\end{tabular} & \begin{tabular}{l} 
k: shape \\
parameter
\end{tabular} & \begin{tabular}{l}
\(\sigma:\) scale \\
parameter
\end{tabular} & \begin{tabular}{l}
\(\mu:\) location \\
parameter
\end{tabular} \\
\hline 'gp' or \\
'Generalized & \begin{tabular}{l} 
"Generalized \\
Pareto \\
Distribution" \\
on page B-37
\end{tabular} & \begin{tabular}{l} 
k: tail \\
index (shape) \\
parameter
\end{tabular} & \begin{tabular}{l}
\(\sigma:\) scale \\
parameter
\end{tabular} & \begin{tabular}{l}
\(\mu:\) threshold \\
(location) \\
parameter
\end{tabular} \\
\hline
\end{tabular}
\begin{tabular}{l|l|l|l|l}
\hline name & Distribution & \begin{tabular}{l} 
Input \\
Parameter A
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter B
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter \\
C
\end{tabular} \\
\hline \begin{tabular}{l} 
'geo' or \\
'Geometric'
\end{tabular} & \begin{tabular}{l} 
"Geometric \\
Distribution" \\
on page B-41
\end{tabular} & \begin{tabular}{l} 
p: probability \\
parameter
\end{tabular} & - & - \\
\hline \begin{tabular}{l} 
'hyge' or \\
'Hypergeometric '
\end{tabular} & \begin{tabular}{l} 
"Hypergeometric \\
Distribution" on \\
page B-43
\end{tabular} & \begin{tabular}{l} 
M: size of the \\
population
\end{tabular} & \begin{tabular}{l} 
K: number of \\
items with \\
the desired \\
characteristic in \\
the population
\end{tabular} & \begin{tabular}{l} 
n: number \\
of samples \\
drawn
\end{tabular} \\
\hline \begin{tabular}{l} 
'logn' or \\
'Lognormal'
\end{tabular} & \begin{tabular}{l} 
"Lognormal \\
Distribution" \\
on page B-51
\end{tabular} & \(\mu\) & \(\sigma\) & - \\
\hline \begin{tabular}{l} 
'nbin' or \\
'Negative \\
Binomial'
\end{tabular} & \begin{tabular}{l} 
"Negative \\
Binomial \\
Distribution" \\
on page B-72
\end{tabular} & \begin{tabular}{l} 
r: number of \\
successes
\end{tabular} & \begin{tabular}{l} 
p: probability \\
of success in a \\
single trial
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'ncf' or \\
'Noncentral F'
\end{tabular} & \begin{tabular}{l} 
"Noncentral F \\
Distribution" on \\
page B-78
\end{tabular} & \begin{tabular}{l}
\(v 1:\) numerator \\
degrees of \\
freedom
\end{tabular} & \begin{tabular}{l}
\(v 2:\) denominator \\
degrees of \\
freedom
\end{tabular} & \begin{tabular}{l}
\(\delta:\) \\
noncentrality \\
parameter
\end{tabular} \\
\hline \begin{tabular}{l} 
'nct' or \\
'Noncentral t'
\end{tabular} & \begin{tabular}{l} 
"Noncentral t \\
Distribution" on \\
page B-80
\end{tabular} & \begin{tabular}{l}
\(v:\) degrees of \\
freedom
\end{tabular} & \begin{tabular}{l}
\(\delta:\) noncentrality \\
parameter
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'ncx2' or \\
'Noncentral \\
Chi-square'
\end{tabular} & \begin{tabular}{l} 
"Noncentral \\
Chi-Square \\
Distribution" \\
on page B-76
\end{tabular} & \begin{tabular}{l}
\(v:\) degrees of \\
freedom
\end{tabular} & \begin{tabular}{l}
\(\delta:\) noncentrality \\
parameter
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'norm' or \\
'Normal'
\end{tabular} & \begin{tabular}{l} 
"Normal \\
Distribution" \\
on page B-83
\end{tabular} & \(\mu:\) mean & \begin{tabular}{l}
\(\sigma:\) standard \\
deviation
\end{tabular} & - \\
\hline
\end{tabular}
\begin{tabular}{l|l|l|l|l}
\hline name & Distribution & \begin{tabular}{l} 
Input \\
Parameter A
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter B
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter \\
C
\end{tabular} \\
\hline \begin{tabular}{l} 
'poiss' or \\
'Poisson'
\end{tabular} & \begin{tabular}{l} 
"Poisson \\
Distribution" \\
on page B-89
\end{tabular} & \(\lambda:\) mean & - & - \\
\hline \begin{tabular}{l} 
'rayl' or \\
'Rayleigh'
\end{tabular} & \begin{tabular}{l} 
"Rayleigh \\
Distribution" \\
on page B-91
\end{tabular} & \begin{tabular}{l} 
b: scale \\
parameter
\end{tabular} & - & - \\
\hline 't' or 'T' & \begin{tabular}{l} 
"Student's t \\
Distribution" \\
on page B-95
\end{tabular} & \begin{tabular}{l}
\(v:\) degrees of \\
freedom
\end{tabular} & - & - \\
\hline \begin{tabular}{l} 
'unif' or \\
'Uniform'
\end{tabular} & \begin{tabular}{l} 
"Uniform \\
Distribution \\
(Continuous)" \\
on page B-99
\end{tabular} & \begin{tabular}{l} 
a: lower \\
endpoint \\
(minimum)
\end{tabular} & \begin{tabular}{l} 
b: upper \\
endpoint \\
(maximum)
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'unid' or \\
'Discrete \\
Uniform'
\end{tabular} & \begin{tabular}{l} 
"Uniform \\
Distribution \\
(Discrete)" on \\
page B-101
\end{tabular} & \begin{tabular}{l} 
N: maximum \\
observable value
\end{tabular} & - & - \\
\hline \begin{tabular}{l} 
'wbl' or \\
'Weibull'
\end{tabular} & \begin{tabular}{l} 
"Weibull \\
Distribution" \\
on page B-103
\end{tabular} & \begin{tabular}{l} 
a: scale \\
parameter
\end{tabular} & \begin{tabular}{l} 
b: shape \\
parameter
\end{tabular} & - \\
\hline
\end{tabular}

Examples
Compute the pdf of the normal distribution with mean 0 and standard deviation 1 at inputs \(-2,-1,0,1,2\) :
```

p1 = pdf('Normal',-2:2,0,1)
p1 =

```


The order of the parameters is the same as for normpdf.

Compute the pdfs of Poisson distributions with rate parameters \(0,1, \ldots\), 4 at inputs \(1,2, \ldots, 5\), respectively:
```

p2 = pdf('Poisson',0:4,1:5)
p2 =
0.3679

```

The order of the parameters is the same as for poisspdf.

\section*{See Also}
cdf | icdf | mle | random

\section*{gmdistribution.pdf}

Purpose
Probability density function for Gaussian mixture distribution

\section*{Syntax \\ y = pdf(obj, X)}
\(y=\operatorname{pdf}(o b j, X)\) returns a vector \(y\) of length \(n\) containing the values of the probability density function (pdf) for the gmdistribution object obj, evaluated at the \(n\)-by- \(d\) data matrix \(X\), where \(n\) is the number of observations and \(d\) is the dimension of the data. obj is an object created by gmdistribution or fit. \(\mathrm{y}(\mathrm{I})\) is the pdf of observation I.

Examples Create a gmdistribution object defining a two-component mixture of bivariate Gaussian distributions:
```

MU = [1 2;-3 -5];
SIGMA = cat(3,[2 0;0 .5],[1 0;0 1]);
p = ones(1,2)/2;
obj = gmdistribution(MU,SIGMA,P);
ezsurf(@(x,y)pdf(obj,[x y]),[-10 10],[-10 10])

```

\section*{gmdistribution.pdf}


See Also
gmdistribution | fit | cdf | mvnpdf

\section*{Purpose}

Probability density function for piecewise distribution

\section*{Syntax \\ P = pdf(obj, X\()\)}
\(P=\operatorname{pdf}(o b j, X)\) returns an array \(P\) of values of the probability density function for the piecewise distribution object obj, evaluated at the values in the array \(X\).

Note For a Pareto tails object, the pdf is computed using the generalized Pareto distribution in the tails. In the center, the pdf is computed using the slopes of the cdf, which are interpolated between a set of discrete values. Therefore the pdf in the center is piecewise constant. It is noisy for a cdffun specified in paretotails via the 'ecdf' option, and somewhat smoother for the 'kernel' option, but generally not a good estimate of the underlying density of the original data.

\section*{Examples}

Fit Pareto tails to a \(t\) distribution at cumulative probabilities 0.1 and 0.9:
```

t = trnd(3,100,1);
obj = paretotails(t,0.1,0.9);
[p,q] = boundary(obj)
p =
0.1000
0.9000
q =
-1.7766
1.8432
pdf(obj,q)
ans =
0.2367
0.1960

```

See Also paretotails | cdf
\(\left.\begin{array}{ll}\text { Purpose } & \text { Return probability density function (PDF) for ProbDist object } \\ \text { Syntax } & Y=\operatorname{pdf}(P D, X) \\ \text { Description } & \begin{array}{l}Y=\operatorname{pdf}(P D, X) \text { returns } Y, \text { an array containing the probability density } \\ \text { function (PDF) for the ProbDist object } P D, \text { evaluated at values in } X .\end{array} \\ \begin{array}{l}\text { Input } \\ \text { Arguments }\end{array} & P D \\ \begin{array}{l}\text { Output } \\ \text { Arguments }\end{array} & Y\end{array} \begin{array}{l}\text { An object of the class ProbDistUnivParam or } \\ \text { ProbDistUnivKernel. } \\ \text { A numeric array of values where you want to } \\ \text { evaluate the PDF. }\end{array}\right\}\)

Purpose Pairwise distance between pairs of objects
Syntax \(\quad \begin{aligned} D & =\operatorname{pdist}(X) \\ & D=\operatorname{pdist}(X, \text { distance })\end{aligned}\)
Description
\(\mathrm{D}=\) pdist \((\mathrm{X})\) computes the Euclidean distance between pairs of objects in \(m\)-by- \(n\) data matrix \(X\). Rows of \(X\) correspond to observations, and columns correspond to variables. D is a row vector of length \(m(m-1) / 2\), corresponding to pairs of observations in X. The distances are arranged in the order \((2,1),(3,1), \ldots,(m, 1),(3,2), \ldots,(m, 2), \ldots\), ( \(m, m-1\) )). D is commonly used as a dissimilarity matrix in clustering or multidimensional scaling.

To save space and computation time, D is formatted as a vector. However, you can convert this vector into a square matrix using the squareform function so that element \(i, j\) in the matrix, where \(i<j\), corresponds to the distance between objects \(i\) and \(j\) in the original data set.
\(\mathrm{D}=\) pdist(X, distance) computes the distance between objects in the data matrix, X , using the method specified by distance, which can be any of the following character strings.
\begin{tabular}{l|l}
\hline Metric & Description \\
\hline 'euclidean' & Euclidean distance (default). \\
\hline 'seuclidean' & \begin{tabular}{l} 
Standardized Euclidean distance. Each \\
coordinate difference between rows in X is \\
scaled by dividing by the corresponding \\
element of the standard deviation \\
S=nanstd (X). To specify another value for \\
S, use D=pdist (X, 'seuclidean ' , S).
\end{tabular} \\
\hline 'cityblock' & \begin{tabular}{l} 
City block metric.
\end{tabular} \\
\hline 'minkowski' & \begin{tabular}{l} 
Minkowski distance. The default exponent is \\
2. To specify a different exponent, use D \(=\) \\
pdist (X, 'minkowski' , P), where P is a scalar
\end{tabular} \\
\hline
\end{tabular} positive value of the exponent.
\(\left.\begin{array}{l|l}\hline \text { Metric } & \text { Description } \\ \hline \text { 'chebychev' } & \begin{array}{l}\text { Chebychev distance (maximum coordinate } \\ \text { difference). }\end{array} \\ \hline \text { 'mahalanobis' } & \begin{array}{l}\text { Mahalanobis distance, using the sample } \\ \text { covariance of X as computed by nancov. To } \\ \text { compute the distance with a different covariance, } \\ \text { use D = pdist (X, 'mahalanobis ' , C), where } \\ \text { the matrix C is symmetric and positive definite. }\end{array} \\ \hline \text { 'cosine' } & \begin{array}{l}\text { One minus the cosine of the included angle } \\ \text { between points (treated as vectors). }\end{array} \\ \hline \text { 'correlation' } & \begin{array}{l}\text { One minus the sample correlation between } \\ \text { points (treated as sequences of values). }\end{array} \\ \hline \text { 'spearman' } & \begin{array}{l}\text { One minus the sample Spearman's rank } \\ \text { correlation between observations (treated as } \\ \text { sequences of values). }\end{array} \\ \hline \text { 'hamming' } & \begin{array}{l}\text { Hamming distance, which is the percentage of } \\ \text { coordinates that differ. }\end{array} \\ \hline \text { 'jaccard' } & \begin{array}{l}\text { One minus the Jaccard coefficient, which is the } \\ \text { percentage of nonzero coordinates that differ. }\end{array} \\ \hline \begin{array}{l}\text { custom distance } \\ \text { function }\end{array} & \begin{array}{l}\text { A distance function specified using @: } \\ \mathrm{D}=\text { pdist (X, @distfun) } \\ \text { A distance function must be of form }\end{array} \\ \hline \text { d2 = distfun (XI, XJ) }\end{array}\right\}\)

\section*{pdist}

The output D is arranged in the order of \(((2,1),(3,1), \ldots\), ( \(m, 1\) ),(3,2),...( \(m, 2\) ),..... \((m, m-1)\) ), i.e. the lower left triangle of the full \(m\)-by- \(m\) distance matrix in column order. To get the distance between the \(i\) th and \(j\) th observations \((i<j)\), either use the formula \(D((i-1) *(m-i / 2)+j-i)\), or use the helper function \(Z=\) squareform \((D)\), which returns an \(m\)-by- \(m\) square symmetric matrix, with the ( \(i, j\) ) entry equal to distance between observation \(i\) and observation \(j\).

\section*{Metrics}

Given an \(m\)-by- \(n\) data matrix \(X\), which is treated as \(m\) (1-by- \(n\) ) row vectors \(\mathrm{x}_{1}, \mathrm{x}_{2}, \ldots, \mathrm{x}_{m}\), the various distances between the vector \(\mathrm{x}_{s}\) and \(\mathrm{x}_{t}\) are defined as follows:
- Euclidean distance
\[
d_{s t}^{2}=\left(x_{s}-x_{t}\right)\left(x_{s}-x_{t}\right)^{\prime}
\]

Notice that the Euclidean distance is a special case of the Minkowski metric, where \(\mathrm{p}=2\).
- Standardized Euclidean distance
\[
d_{s t}^{2}=\left(x_{s}-x_{t}\right) V^{-1}\left(x_{s}-x_{t}\right)^{\prime}
\]
where V is the \(n\)-by- \(n\) diagonal matrix whose \(j\) th diagonal element is \(S(j)^{2}\), where \(S\) is the vector of standard deviations.
- Mahalanobis distance
\[
d_{s t}^{2}=\left(x_{s}-x_{t}\right) C^{-1}\left(x_{s}-x_{t}\right)^{\prime}
\]
where \(C\) is the covariance matrix.
- City block metric
\[
d_{s t}=\sum_{j=1}^{n}\left|x_{s j}-x_{t j}\right|
\]

Notice that the city block distance is a special case of the Minkowski metric, where \(\mathrm{p}=1\).
- Minkowski metric
\[
d_{s t}=\sqrt[p]{\sum_{j=1}^{n}\left|x_{s j}-x_{t j}\right|^{p}}
\]

Notice that for the special case of \(p=1\), the Minkowski metric gives the city block metric, for the special case of \(p=2\), the Minkowski metric gives the Euclidean distance, and for the special case of \(p=\infty\), the Minkowski metric gives the Chebychev distance.
- Chebychev distance
\[
d_{s t}=\max _{j}\left\{\left|x_{s j}-x_{t j}\right|\right\}
\]

Notice that the Chebychev distance is a special case of the Minkowski metric, where \(p=\infty\).
- Cosine distance
\[
d_{s t}=1-\frac{x_{s} x_{t}^{\prime}}{\sqrt{\left(x_{s} x_{s}^{\prime}\right)\left(x_{t} x_{t}^{\prime}\right)}}
\]
- Correlation distance
\[
d_{s t}=1-\frac{\left(x_{s}-\bar{x}_{s}\right)\left(x_{t}-\bar{x}_{t}\right)^{\prime}}{\sqrt{\left(x_{s}-\bar{x}_{s}\right)\left(x_{s}-\bar{x}_{s}\right)^{\prime}} \sqrt{\left(x_{t}-\bar{x}_{t}\right)\left(x_{t}-\bar{x}_{t}\right)^{\prime}}}
\]
where
\(\bar{x}_{s}=\frac{1}{n} \sum_{j} x_{s j}\) and \(\bar{x}_{t}=\frac{1}{n} \sum_{j} x_{t j}\)
- Hamming distance
\[
d_{s t}=\left(\#\left(x_{s j} \neq x_{t j}\right) / n\right)
\]
- Jaccard distance
\[
d_{s t}=\frac{\#\left[\left(x_{s j} \neq x_{t j}\right) \cap\left(\left(x_{s j} \neq 0\right) \cup\left(x_{t j} \neq 0\right)\right)\right]}{\#\left[\left(x_{s j} \neq 0\right) \cup\left(x_{t j} \neq 0\right)\right]}
\]
- Spearman distance
\[
d_{s t}=1-\frac{\left(r_{s}-\bar{r}_{s}\right)\left(r_{t}-\bar{r}_{t}\right)^{\prime}}{\sqrt{\left(r_{s}-\bar{r}_{s}\right)\left(r_{s}-\bar{r}_{s}\right)^{\prime}} \sqrt{\left(r_{t}-\bar{r}_{t}\right)\left(r_{t}-\bar{r}_{t}\right)^{\prime}}}
\]
where
- \(r_{s j}\) is the rank of \(x_{s j}\) taken over \(x_{1 j}, x_{2 j}, \ldots x_{m j}\), as computed by tiedrank
- \(r_{s}\) and \(r_{t}\) are the coordinate-wise rank vectors of \(x_{s}\) and \(x_{t}\), i.e., \(r_{s}=\left(r_{s 1}, r_{s 2}, \ldots r_{s n}\right)\)
- \(\quad \bar{r}_{s}=\frac{1}{n} \sum_{j} r_{s j}=\frac{(n+1)}{2}\)
- \(\bar{r}_{t}=\frac{1}{n} \sum_{j} r_{t j}=\frac{(n+1)}{2}\)

\section*{Examples}

Generate random data and find the unweighted Euclidean distance and then find the weighted distance using two different methods:
```

% Compute the ordinary Euclidean distance.
X = randn(100, 5);
D = pdist(X,'euclidean'); % euclidean distance

```
```

% Compute the Euclidean distance with each coordinate
% difference scaled by the standard deviation.
Dstd = pdist(X,'seuclidean');
% Use a function handle to compute a distance that weights
% each coordinate contribution differently.
Wgts = [.1 .3 .3 .2 .1]; % coordinate weights
weuc = @(XI,XJ,W)(sqrt(bsxfun(@minus,XI,XJ).^2 * W'));
Dwgt = pdist(X, @(Xi,Xj) weuc(Xi,Xj,Wgts));

```

See Also
cluster | clusterdata | cmdscale | cophenet | dendrogram | inconsistent | linkage | pdist2 | silhouette | squareform

Purpose
Pairwise distance between two sets of observations
Syntax
```

D = pdist2(X,Y)
D = pdist2(X,Y,distance)
D = pdist2(X,Y,'minkowski',P)
D = pdist2(X,Y,'mahalanobis',C)
D = pdist2(X,Y,distance,'Smallest',K)
D = pdist2(X,Y,distance,'Largest',K)
[D,I] = pdist2(X,Y,distance,'Smallest',K)
[D,I] = pdist2(X,Y,distance,'Largest',K)

```

\section*{Description}

D = pdist2 \((X, Y)\) returns a matrix \(D\) containing the Euclidean distances between each pair of observations in the \(m x\)-by- \(n\) data matrix X and \(m y\)-by- \(n\) data matrix Y . Rows of X and Y correspond to observations, columns correspond to variables. D is an \(m x\)-by- \(m y\) matrix, with the \((i, j)\) entry equal to distance between observation \(i\) in X and observation \(j\) in Y . The \((i, j)\) entry will be NaN if observation \(i\) in X or observation \(j\) in Y contain NaNs.
\(\mathrm{D}=\) pdist2(X,Y,distance) computes D using distance. Choices are:
\begin{tabular}{|c|c|}
\hline Metric & Description \\
\hline 'euclidean' & Euclidean distance (default). \\
\hline 'seuclidean' & Standardized Euclidean distance. Each coordinate difference between rows in \(X\) and \(Y\) is scaled by dividing by the corresponding element of the standard deviation computed from X , \(S=\) nanstd \((X)\). To specify another value for \(S\), use \(D=\) PDIST2 ( \(\mathrm{X}, \mathrm{Y}\), ' seuclidean ' S ). \\
\hline 'cityblock' & City block metric. \\
\hline 'minkowski' & Minkowski distance. The default exponent is 2. To compute the distance with a different exponent, use D = pdist2(X,Y,'minkowski', P), where the exponent \(P\) is a scalar positive value. \\
\hline
\end{tabular}
\begin{tabular}{l|l}
\hline Metric & Description \\
\hline 'chebychev' & \begin{tabular}{l} 
Chebychev distance (maximum coordinate \\
difference).
\end{tabular} \\
\hline 'mahalanobis ' & \begin{tabular}{l} 
Mahalanobis distance, using the sample covariance \\
of X as computed by nancov. To compute the \\
distance with a different covariance, use D \(=\) \\
pdist2 (X, Y, ' mahalanobis ' , C) where the matrix C \\
is symmetric and positive definite.
\end{tabular} \\
\hline 'cosine' & \begin{tabular}{l} 
One minus the cosine of the included angle between \\
points (treated as vectors).
\end{tabular} \\
\hline 'correlation ' & \begin{tabular}{l} 
One minus the sample correlation between points \\
(treated as sequences of values).
\end{tabular} \\
\hline 'spearman' & \begin{tabular}{l} 
One minus the sample Spearman's rank correlation \\
between observations, treated as sequences of \\
values.
\end{tabular} \\
\hline 'hamming' & \begin{tabular}{l} 
Hamming distance, the percentage of coordinates \\
that differ.
\end{tabular} \\
\hline 'jaccard' & \begin{tabular}{l} 
One minus the Jaccard coefficient, the percentage of \\
nonzero coordinates that differ.
\end{tabular} \\
\hline function & \begin{tabular}{l} 
A distance function specified using @: \\
D = pdist2(X,Y, @distfun). \\
A distance function must be of the form
\end{tabular} \\
\hline function D2 = distfun (ZI, ZJ)
\end{tabular}
\begin{tabular}{|l|l} 
Metric & Description
\end{tabular}

If your data is not sparse, generally it is faster to use a built-in distance than to use a function handle.
\(\mathrm{D}=\) pdist2(X,Y, distance, 'Smallest', K) returns a K-by-my matrix \(D\) containing the \(K\) smallest pairwise distances to observations in \(X\) for each observation in Y. pdist2 sorts the distances in each column of \(D\) in ascending order. \(\mathrm{D}=\) pdist2( \(\mathrm{X}, \mathrm{Y}\), distance, 'Largest' K ) returns the \(K\) largest pairwise distances sorted in descending order. If \(K\) is greater than \(m x\), pdist2 returns an \(m x\)-by- \(m y\) distance matrix. For each observation in Y , pdist2 finds the K smallest or largest distances by computing and comparing the distance values to all the observations in X .
[ \(\mathrm{D}, \mathrm{I}]=\) pdist2(X,Y,distance, 'Smallest', K) returns a K -by-my matrix I containing indices of the observations in X corresponding to the K smallest pairwise distances in D . \([\mathrm{D}, \mathrm{I}]=\) pdist2(X,Y,distance, 'Largest', K) returns indices corresponding to the K largest pairwise distances.

\section*{Metrics}

Given an \(m x\)-by- \(n\) data matrix X , which is treated as \(m x\) (1-by- \(n\) ) row vectors \(\mathrm{x}_{1}, \mathrm{x}_{2}, \ldots, \mathrm{x}_{m \mathrm{x}}\), and my-by- \(n\) data matrix Y , which is treated as \(m y\) (1-by- \(n\) ) row vectors \(\mathrm{y}_{1}, \mathrm{y}_{2}, \ldots, \mathrm{y}_{m \mathrm{y}}\), the various distances between the vector \(\mathrm{x}_{s}\) and \(\mathrm{y}_{t}\) are defined as follows:
- Euclidean distance
\[
d_{s t}^{2}=\left(x_{s}-y_{t}\right)\left(x_{s}-y_{t}\right)^{\prime}
\]

Notice that the Euclidean distance is a special case of the Minkowski metric, where \(\mathrm{p}=2\).
- Standardized Euclidean distance
\[
d_{s t}^{2}=\left(x_{s}-y_{t}\right) V^{-1}\left(x_{s}-y_{t}\right)^{\prime}
\]
where V is the \(n\)-by- \(n\) diagonal matrix whose \(j\) th diagonal element is \(S(j)^{2}\), where \(S\) is the vector of standard deviations.
- Mahalanobis distance
\[
d_{s t}^{2}=\left(x_{s}-y_{t}\right) C^{-1}\left(x_{s}-y_{t}\right)^{\prime}
\]
where C is the covariance matrix.
- City block metric
\[
d_{s t}=\sum_{j=1}^{n}\left|x_{s j}-y_{t j}\right|
\]

Notice that the city block distance is a special case of the Minkowski metric, where \(\mathrm{p}=1\).
- Minkowski metric
\[
d_{s t}=\sqrt[p]{\sum_{j=1}^{n}\left|x_{s j}-y_{t j}\right|^{p}}
\]

Notice that for the special case of \(p=1\), the Minkowski metric gives the City Block metric, for the special case of \(p=2\), the Minkowski metric gives the Euclidean distance, and for the special case of \(p=\infty\), the Minkowski metric gives the Chebychev distance.
- Chebychev distance
\[
d_{s t}=\max _{j}\left\{\left|x_{s j}-y_{t j}\right|\right\}
\]

Notice that the Chebychev distance is a special case of the Minkowski metric, where \(p=\infty\).
- Cosine distance
\[
d_{s t}=\left(1-\frac{x_{s} y_{t}^{\prime}}{\sqrt{\left(x_{s} x_{s}^{\prime}\right)\left(y_{t} y_{t}^{\prime}\right)}}\right)
\]
- Correlation distance
\[
d_{s t}=1-\frac{\left(x_{s}-\bar{x}_{s}\right)\left(y_{t}-\bar{y}_{t}\right)^{\prime}}{\sqrt{\left(x_{s}-\bar{x}_{s}\right)\left(x_{s}-\bar{x}_{s}\right)^{\prime}} \sqrt{\left(y_{t}-\bar{y}_{t}\right)\left(y_{t}-\bar{y}_{t}\right)^{\prime}}}
\]
where
\[
\begin{gathered}
\bar{x}_{s}=\frac{1}{n} \sum_{j} x_{s j} \text { and } \\
\bar{y}_{t}=\frac{1}{n} \sum_{j} y_{t j}
\end{gathered}
\]
- Hamming distance
\[
d_{s t}=\left(\#\left(x_{s j} \neq y_{t j}\right) / n\right)
\]
- Jaccard distance
\[
d_{s t}=\frac{\#\left[\left(x_{s j} \neq y_{t j}\right) \cap\left(\left(x_{s j} \neq 0\right) \cup\left(y_{t j} \neq 0\right)\right)\right]}{\#\left[\left(x_{s j} \neq 0\right) \cup\left(y_{t j} \neq 0\right)\right]}
\]
- Spearman distance
\[
d_{s t}=1-\frac{\left(r_{s}-\bar{r}_{s}\right)\left(r_{t}-\bar{r}_{t}\right)^{\prime}}{\sqrt{\left(r_{s}-\bar{r}_{s}\right)\left(r_{s}-\bar{r}_{s}\right)^{\prime}} \sqrt{\left(r_{t}-\bar{r}_{t}\right)\left(r_{t}-\bar{r}_{t}\right)^{\prime}}}
\]
where
- \(r_{s j}\) is the rank of \(x_{s j}\) taken over \(x_{1 j}, x_{2 j}, \ldots x_{m x, j}\), as computed by tiedrank
- \(r_{t j}\) is the rank of \(y_{t j}\) taken over \(y_{1 j}, y_{2 j}, \ldots y_{m y, j}\), as computed by tiedrank
- \(r_{s}\) and \(r_{t}\) are the coordinate-wise rank vectors of \(x_{s}\) and \(y_{t}\), i.e. \(r_{s}=\) \(\left(r_{s 1}, r_{s 2}, \ldots r_{s n}\right)\) and \(r_{t}=\left(r_{t 1}, r_{t 2}, \ldots r_{t n}\right)\)
- \(\quad \bar{r}_{s}=\frac{1}{n} \sum_{j} r_{s j}=\frac{(n+1)}{2}\)
- \(\bar{r}_{t}=\frac{1}{n} \sum_{j} r_{t j}=\frac{(n+1)}{2}\)

\section*{Examples}

Generate random data and find the unweighted Euclidean distance, then find the weighted distance using two different methods:
```

% Compute the ordinary Euclidean distance
X = randn(100, 5);
Y = randn(25, 5);
D = pdist2(X,Y,'euclidean'); % euclidean distance
% Compute the Euclidean distance with each coordinate
% difference scaled by the standard deviation
Dstd = pdist2(X,Y,'seuclidean');
% Use a function handle to compute a distance that weights
% each coordinate contribution differently.
Wgts = [.1 .3 .3 .2 .1];
weuc = @(XI,XJ,W)(sqrt(bsxfun(@minus,XI,XJ).^2 * W'));
Dwgt = pdist2(X,Y, @(Xi,Xj) weuc(Xi,Xj,Wgts));

```

See Also
pdist | createns | knnsearch | KDTreeSearcher | ExhaustiveSearcher

Purpose Pearson system random numbers
```

Syntax $\quad r=$ pearsrnd(mu, sigma, skew, kurt, m, n)
$r=$ pearsrnd(mu,sigma, skew, kurt)
$r=$ pearsrnd(mu,sigma,skew,kurt,m,n,...)
$r$ = pearsrnd(mu, sigma, skew, kurt, [m, n, ...])
[r,type] = pearsrnd(...)
[r,type,coefs] = pearsrnd(...)

```

\section*{Description}
\(r=\) pearsrnd(mu, sigma, skew, kurt, \(m, n\) ) returns an m-by-n matrix of random numbers drawn from the distribution in the Pearson system with mean mu, standard deviation sigma, skewness skew, and kurtosis kurt. The parameters mu, sigma, skew, and kurt must be scalars.

Note Because \(r\) is a random sample, its sample moments, especially the skewness and kurtosis, typically differ somewhat from the specified distribution moments.
pearsrnd uses the definition of kurtosis for which a normal distribution has a kurtosis of 3 . Some definitions of kurtosis subtract 3 , so that a normal distribution has a kurtosis of 0 . The pearsrnd function does not use this convention.

Some combinations of moments are not valid; in particular, the kurtosis must be greater than the square of the skewness plus 1. The kurtosis of the normal distribution is defined to be 3 .
\(r=\) pearsrnd(mu, sigma, skew, kurt) returns a scalar value.
\(r=\) pearsrnd(mu,sigma, skew,kurt, m, n, ...) or \(r=\) pearsrnd(mu, sigma, skew, kurt, [m,n,...]) returns an m-by-n-by-... array.
[ \(r\), type] \(=\) pearsrnd (...) returns the type of the specified distribution within the Pearson system. type is a scalar integer from 0
to 7 . Set m and n to 0 to identify the distribution type without generating any random values.

The seven distribution types in the Pearson system correspond to the following distributions:
- 0 - Normal distribution
- 1 - Four-parameter beta distribution
- 2 - Symmetric four-parameter beta distribution
- 3 - Three-parameter gamma distribution
- 4 - Not related to any standard distribution. The density is proportional to:
\[
\left(1+((x-a) / b)^{2}\right)^{-c} \exp (-d \arctan ((x-a) / b)) .
\]
- 5 - Inverse gamma location-scale distribution
- \(6-F\) location-scale distribution
- 7 - Student's \(t\) location-scale distribution
[r,type,coefs] = pearsrnd(...) returns the coefficients coefs of the quadratic polynomial that defines the distribution via the differential equation
\[
\frac{d}{d x} \log (p(x))=\frac{-(a+x)}{c(0)+c(1) x+c(2) x^{2}} .
\]

\section*{Examples}

Generate random values from the standard normal distribution:
```

r = pearsrnd(0,1,0,3,100,1); % Equivalent to randn(100,1)

```

Determine the distribution type:
```

[r,type] = pearsrnd(0,1,1,4,0,0);
r =

```
        []
type
        1
References [1] Johnson, N.L., S. Kotz, and N. Balakrishnan (1994) Continuous
        Univariate Distributions, Volume 1, Wiley-Interscience, Pg 15, Eqn
        12.33 .
See Also random | johnsrnd

\section*{Purpose}

\section*{Syntax}
```

```
[X,Y] = perfcurve(labels,scores,posclass)
```

```
[X,Y] = perfcurve(labels,scores,posclass)
[X,Y] = perfcurve(labels,scores,posclass,'Name',value)
[X,Y] = perfcurve(labels,scores,posclass,'Name',value)
[X,Y,T,AUC,OPTROCPT,SUBY,SUBYNAMES] = perfcurve(labels,scores,
[X,Y,T,AUC,OPTROCPT,SUBY,SUBYNAMES] = perfcurve(labels,scores,
    posclass)
    posclass)
[X,Y,T,AUC] = perfcurve(labels,scores,posclass)
```

```
[X,Y,T,AUC] = perfcurve(labels,scores,posclass)
```

```

\section*{Description}

Compute Receiver Operating Characteristic (ROC) curve or other performance curve for classifier output
[ \(\mathrm{X}, \mathrm{Y}\) ] = perfcurve(labels, scores, posclass) computes a ROC curve for a vector of classifier predictions scores given true class labels, labels. labels can be a numeric vector, logical vector, character matrix, cell array of strings or categorical vector. scores is a numeric vector of scores returned by a classifier for some data. posclass is the positive class label (scalar), either numeric (for numeric labels), logical (for logical labels), or char. The returned values \(X\) and \(Y\) are coordinates for the performance curve and can be visualized with plot ( \(\mathrm{X}, \mathrm{Y}\) ). For more information on labels, scores, and posclass, see "Input Arguments" on page 20-1377 . For more information on X and Y, see "Output Arguments" on page 20-1381.
[ \(\mathrm{X}, \mathrm{Y}\) ] = perfcurve(labels,scores, posclass,' \({ }^{2}\). specifies one or more optional parameter name/value pairs, with Name in single quotes. See "Input Arguments" on page 20-1377 for a list of inputs, parameter name/value pairs, and respective explanations.

See "Grouped Data" on page 2-34 for more information on grouping variables.
[X,Y,T,AUC,OPTROCPT,SUBY,SUBYNAMES] = perfcurve(labels, scores, posclass) returns:
- An array of thresholds on classifier scores for the computed values of \(X\) and \(Y(T)\).
- The area under curve (AUC) for the computed values of \(X\) and \(Y\).
- The optimal operating point of the ROC curve (OPTROCPT).
- An array of \(Y\) values for negative subclasses (SUBY).
- A cell array of negative class names (SUBYNAMES).

For more information on each output, see "Output Arguments" on page 20-1381.
[ \(\mathrm{X}, \mathrm{Y}, \mathrm{T}, \mathrm{AUC}]=\) perfcurve(labels,scores, posclass) also returns pointwise confidence bounds for the computed values \(X, Y, T\), and AUC if you supply cell arrays for labels and scores or set NBoot (see "Input Arguments" on page 20-1377) to a positive integer. To compute the confidence bounds, perfcurve uses either vertical averaging (VA) or threshold averaging (TA). The returned values \(Y\) are an \(m\)-by- 3 array in which the 1st element in every row gives the mean value, the 2nd element gives the lower bound and the 3rd element gives the upper bound. The returned AUC is a row-vector with 3 elements following the same convention. For VA, the returned values T are an \(m\)-by- 3 array and X is a column-vector. For TA, the returned values X are an \(m\)-by- 3 matrix and \(T\) is a column-vector.
perfcurve computes confidence bounds using either cross-validation or bootstrap. If you supply cell arrays for labels and scores, perfcurve uses cross-validation and treats elements in the cell arrays as cross-validation folds. labels can be a cell array of numeric vectors, logical vectors, character matrices, cell arrays of strings or categorical vectors. All elements in labels must have the same type. scores is a cell array of numeric vectors. The cell arrays for labels and scores must have the same number of elements, and the number of labels in cell \(k\) must be equal to the number of scores in cell \(k\) for any \(k\) in the range from 1 to the number of elements in scores.

If you set NBoot to a positive integer, perfcurve generates nooot bootstrap replicas to compute pointwise confidence bounds. You cannot supply cell arrays for labels and scores and set NBoot to a positive integer at the same time.
perfcurve returns pointwise confidence bounds. It does not return a simultaneous confidence band for the entire curve.

If you use 'XCrit' or 'YCrit' options described below to set the criterion for \(X\) or \(Y\) to an anonymous function, perfcurve can only compute confidence bounds by bootstrap.

\section*{Input}

Arguments labels
scores
posclass
labels can be a numeric vector, logical vector, character matrix, cell array of strings or categorical vector.
scores is a numeric vector of scores returned by a classifier for some data. This vector must have as many elements as labels does.
posclass is the positive class label (scalar),
xCrit

\section*{Name-Value Pair Arguments}

\section*{Value and Description}

List of negative classes. Can be either a numeric array or an array of chars or a cell array of strings. By default, negClass is set to 'all' and all classes found in the input array of labels that are not the positive class are considered negative. If negClass is a subset of the classes found in the input array of labels, instances with labels that do not belong to either positive or negative classes are discarded. either numeric (for numeric labels) or char. The specified positive class must be in the array of input labels.

Criterion to compute for X . This criterion must be a monotone function of the positive class score. perfcurve supports the following criteria:
- TP - Number of true positive instances.
- FN - Number of false negative instances.
- FP - Number of false positive instances.

\section*{Name}

\section*{Value and Description}
- TN - Number of true negative instances.
- TP+FP - Sum of TP and FP.
- RPP - Rate of positive predictions. RPP \(=(T P+F P) /(T P+F N+F P+T N)\)
- RNP - Rate of negative predictions. \(R N P=(T N+F N) /(T P+F N+F P+T N)\)
- accu - Accuracy. accu \(=(T P+T N) /(T P+F N+F P+T N)\)
- TPR, sens, reca - True positive rate, sensitivity, recall. TPR, sens, reca \(=\) TP/(TP+FN)
- FNR, miss - False negative rate, miss. FNR, miss=FN/ (TP+FN)
- FPR, fall - False positive rate, fallout. FPR,fall=FP/(TN+FP)
- TNR, spec - True negative rate, specificity. TNR, spec \(=T N /(T N+F P)\)
- PPV, prec - Positive predictive value, precision. PPV, prec=TP/(TP+FP)
- NPV - Negative predictive value. NPV=TN/(TN+FN)
- ecost - Expected cost. ecost \(=(T P * \operatorname{COST}(P \mid P)+F N * \operatorname{COST}(N \mid P)+F P *\) \(\operatorname{COST}(P \mid N)+T N * \operatorname{COST}(N \mid N)) /(T P+F N+F P+T N)\)

In addition, you can define an arbitrary criterion by supplying an anonymous function of three arguments, (C, scale, cost), where \(C\) is a 2 -by- 2 confusion matrix, scale is a 2 -by- 1 array of class scales, and cost is a 2 -by- 2 misclassification cost matrix.

Caution Some of these criteria return NaN values at one of the two special thresholds, 'reject all' and 'accept all'.

\section*{Name}
yCrit

XVals

TVals

UseNearest

\section*{Value and Description}

Criterion to compute for Y. perfcurve supports the same criteria as for \(X\). This criterion does not have to be a monotone function of the positive class score.
Values for the X criterion. The default value for XVals is 'all' and perfcurve computes \(X\) and \(Y\) values for all scores. If the value for xVals is not 'all', it must be a numeric array. In this case, perfcurve computes \(X\) and \(Y\) only for the specified \(X V a l s\).

Thresholds for the positive class score. By default, TVals is unset and perfcurve computes \(X\), Y , and T values for all scores. You can set TVals to either 'all' or a numeric array. If TVals is set to 'all' or unset and XVals is unset, perfcurve returns X, Y , and T values for all scores and computes pointwise confidence bounds for \(Y\) and \(X\) using threshold averaging. If TVals is set to a numeric array, perfcurve returns \(X, Y\), and \(T\) values for the specified thresholds and computes pointwise confidence bounds for \(Y\) and \(X\) at these thresholds using threshold averaging. You cannot set XVals and TVals at the same time.
'on' to use nearest values found in the data instead of the specified numeric XVals or TVals and 'off' otherwise. If you specify numeric XVals and set UseNearest to 'on', perfcurve returns nearest unique values \(X\) found in the data, as well as corresponding values of \(Y\) and \(T\). If you specify numeric XVals and set UseNearest to 'off', perfcurve returns these XVals sorted. By default this parameter is set to 'on'. If you compute confidence bounds by cross-validation or bootstrap, this parameter is always 'off'.

Name
ProcessNaN

Prior

Cost

Alpha

Weights

\section*{Value and Description}

Specifies how perfcurve processes NaN scores. The default value is 'ignore' and perfcurve removes observations with NaN scores from the data. If you set the parameter to 'addtofalse ', perfcurve adds instances with NaN scores to false classification counts in the respective class. That is, perfcurve always counts instances from the positive class as false negative (FN), and always counts instances from the negative class as false positive (FP).
Either string or array with two elements. It represents prior probabilities for the positive and negative class, respectively. Default is 'empirical', that is, perfcurve derives prior probabilities from class frequencies. If set to 'uniform', perfcurve sets all prior probabilities equal.

A 2-by-2 matrix of misclassification costs [C(P|P) C(N|P); \(C(P \mid N) C(N \mid N)]\) where \(C(I \mid J)\) is the cost of misclassifying class \(J\) as class I. By default set to [0 0.5; 0.5 0].
A numeric value between 0 and 1 . perfcurve returns 100*(1-Alpha) percent pointwise confidence bounds for X, Y, T and AUC. By default set to 0.05 for \(95 \%\) confidence interval.
A numeric vector of non-negative observation weights. This vector must have as many elements as scores or labels do. If you supply cell arrays for scores and labels and you need to supply weights, you must supply them as a cell array too. In this case, every element in weights must be a numeric vector with as many elements as the corresponding element in scores: numel (weights \(\{1\}\) ) \(==\) numel ( scores \(\{1\}\) ) etc. To compute \(X\), Y and T or to compute confidence bounds by cross-validation, perfcurve uses these observation weights instead of observation counts. To compute confidence bounds by bootstrap, perfcurve samples \(N\) out of \(N\) with replacement using these weights as multinomial sampling probabilities.

\section*{Name}

NBoot

BootType

BootArg

\section*{Value and Description}

Number of bootstrap replicas for computation of confidence bounds. Must be a positive integer. By default this parameter is set to zero, and bootstrap confidence bounds are not computed. If you supply cell arrays for labels and scores, this parameter must be set to zero because perfcurve cannot use both cross-validation and bootstrap to compute confidence bounds.
Confidence interval type bootci uses to compute confidence bounds. You can specify any type supported by bootci. By default set to 'bca'.
Optional input arguments for bootci used to compute confidence bounds. You can specify all arguments bootci supports. Empty by default.

\section*{Output \(x\) \\ Arguments}
\(Y \quad y\)-coordinates for the performance curve. By default, Y is true positive rate, TPR, (equivalently, recall, or sensitivity). To change this output, use the 'yCrit' input. For accepted criterion, see 'xCrit' in "Input Arguments" on page 20-1377 for more information.

T

AUC The area under curve (AUC) for the computed values of \(X\) and Y. If you set \(x\) Vals to 'all' (the default), perfcurve computes AUC using the returned \(X\) and \(Y\) values. If \(x\) Vals is a numeric array, perfcurve computes AUC using \(X\) and \(Y\) values found from all distinct scores in the interval specified by the smallest and largest elements of xVals. More precisely, perfcurve finds \(X\) values for all distinct thresholds as if XVals were set to 'all', then uses a subset of these (with corresponding \(Y\) values) between min(xVals) and max (xVals) to compute AUC. The function uses trapezoidal approximation to estimate the area. If the first or last value of \(X\) or \(Y\) are NaNs , perfcurve removes them to allow calculation of AUC. This takes care of criteria that produce NaNs for the special 'reject all' or 'accept all' thresholds, for example, positive predictive value (PPV) or negative predictive value (NPV).

OPTROCPT The optimal operating point of the ROC curve as an array of size 1-by- 2 with FPR and TPR values for the optimal ROC operating point. perfcurve computes optrocpt only for the standard ROC curve and sets to NaNs otherwise. To obtain the optimal operating point for the ROC curve, perfcurve first finds the slope, S , using
 of class \(J\) to class \(I\), and \(P=T P+F N\) and \(N=T N+F P\) are the total instance counts in the positive and negative class, respectively. perfcurve then finds the optimal operating point by moving the straight line with slope \(S\) from the upper left corner of the ROC plot (FPR=0, TPR=1) down and to the right until it intersects the ROC curve.
SUBY An array of \(Y\) values for negative subclasses. If you only specify one negative class, SUBY is identical to Y . Otherwise SUBY is a matrix of size \(M-b y-K\), where \(M\) is the number of returned values for \(X\) and \(Y\), and \(K\) is the number of negative classes. perfcurve computes \(Y\) values by summing counts over all negative classes. SUBY gives values of the \(Y\) criterion for each negative class separately. For each negative class, perfcurve places a new column in SUBY and fills it with \(Y\) values for TN and FP counted just for this class.
SUBYNAMES A cell array of negative class names. If you provide an input array, negClass, of negative class names, perfcurve copies it into SUBYNAMES. If you do not provide negClass, perfcurve extracts SUBYNAMES from input labels. The order of SUBYNAMES is the same as the order of columns in \(\operatorname{SUBY}\), that is \(\operatorname{SUBY}(:, 1)\) is for negative class SUBYNAMES \(\{1\}\) etc.

Examples Plot the ROC curve for classification by logistic regression:
```

load fisheriris
x = meas(51:end,1:2);
% iris data, 2 classes and 2 features
y = (1:100)'>50;
% versicolor=0, virginica=1
b = glmfit(x,y,'binomial');
% logistic regression
p = glmval(b,x,'logit');
% fit probabilities for scores
[X,Y,T,AUC] = perfcurve(species(51:end,:),p,'virginica');
plot(X,Y)
xlabel('False positive rate'); ylabel('True positive rate')
title('ROC for classification by logistic regression')

```


Obtain errors on TPR by vertical averaging
```

[X,Y] = perfcurve(species(51:end,:),p,'virginica',...
'nboot',1000,'xvals','all');
% plot errors

```
```

errorbar(X,Y(:,1),Y(:,2)-Y(:,1),Y(:,3)-Y(:,1));

```

\section*{References}

See Also
How To
[1] T. Fawcett, ROC Graphs: Notes and Practical Considerations for Researchers, 2004.
[2] M. Zweig and G. Campbell, Receiver-Operating Characteristic (ROC) Plots: A Fundamental Evaluation Tool in Clinical Medicine, Clin. Chem. 39/4, 561-577, 1993.
[3] J. Davis and M. Goadrich, The relationship between precision-recall and ROC curves, in Proceedings of ICML '06, 233-240, 2006.
[4] C. Moskowitz and M. Pepe, Quantifying and comparing the predictive accuracy of continuous prognostic factors for binary outcomes, Biostatistics 5, 113-127, 2004.
[5] Y. Huang, M. Pepe and Z. Feng, Evaluating the Predictiveness of a Continuous Marker, U. Washington Biostatistics Paper Series, 282, 2006.
[6] W. Briggs and R. Zaretzki, The Skill Plot: A Graphical Technique for Evaluating Continuous Diagnostic Tests, Biometrics 63, 250-261, 2008.
[7] http://www2.cs.uregina.ca/~hamilton/courses/831/notes/lift_chart/lift_chart.htm http://www.dmreview.com/news/5329-1.html.
[8] R. Bettinger, Cost-Sensitive Classifier Selection Using the ROC Convex Hull Method, SAS Institute.
[9] http://www.stata.com/statalist/archive/2003-02/msg00060.html
bootci | glmfit | mnrfit | classify | NaiveBayes | classregtree
- "Grouped Data" on page 2-34
- "Performance Curves" on page 12-32
- "Plotting a Classification Performance Curve" on page 13-115
Purpose Enumeration of permutations
Syntax \(P=\operatorname{perms}(v)\)
Description \(P=\operatorname{perms}(v)\), where \(v\) is a row vector of length \(n\), creates a matrixwhose rows consist of all possible permutations of the \(n\) elements of \(v\).The matrix \(P\) contains \(n!\) rows and \(n\) columns.perms is only practical when \(n\) is less than about 11 (for \(n=11\), theoutput takes over 3 gigabytes).
Examples
perms([2 4 6])
ans \(=\)
\(6 \quad 4\) ..... 2
\(6 \quad 2\) ..... 4
46 ..... 2
42 ..... 6
246
264
See Also
combnk

\section*{Purpose Permute dimensions of categorical array}
Syntax
B = permute(A,order)

Description
\(B=\) permute ( \(A\), order) rearranges the dimensions of the categorical array A so that they are in the order specified by the vector order. The array produced has the same values as A but the order of the subscripts needed to access any particular element are rearranged as specified by order. The elements of order must be a rearrangement of the numbers from 1 to \(n\).

See Also circshift | ipermute

\section*{piecewisedistribution}

Purpose Piecewise-defined distributions

Construction

Methods
boundary
cdf
disp
display
icdf
nsegments
pdf
random
segment

Piecewise distribution boundaries
Cumulative distribution function for piecewise distribution
Display piecewisedistribution object

Display piecewisedistribution object

Inverse cumulative distribution function for piecewise distribution

Number of segments
Probability density function for piecewise distribution
Random numbers from piecewise distribution
Segments containing values

\section*{Properties}

Copy
Semantics

Objects of the piecewisedistribution class have no properties accessible by dot indexing, get methods, or set methods. To obtain information about a piecewisedistribution object, use the appropriate method.

Value. To learn how this affects your use of the class, see Comparing Handle and Value Classes in the MATLAB Object-Oriented Programming documentation.

\title{
Purpose Create piecewise distribution object
}

Description \(\begin{aligned} & \text { piecewisedistribution is an abstract class, and you cannot create } \\ & \text { instances of it directly. You can create paretotails objects that are } \\ & \text { derived from this class. }\end{aligned}\)

\author{
See Also \\ paretotails
}

Purpose Partial least-squares regression
Syntax
```

[XL,YL] = plsregress(X,Y,ncomp)
[XL,YL,XS] = plsregress(X,Y,ncomp)
[XL,YL,XS,YS] = plsregress(X,Y, ncomp)
[XL,YL,XS,YS,BETA] = PLSREGRESS(X,Y,ncomp,...)
[XL,YL,XS,YS,BETA,PCTVAR] = plsregress(X,Y, ncomp)
[XL,YL,XS,YS,BETA, PCTVAR,MSE] = plsregress(X,Y,ncomp)
[XL,YL,XS,YS,BETA,PCTVAR,MSE] = plsregress(...,param1,val1,
param2,val2,...)
[XL,YL,XS,YS,BETA, PCTVAR,MSE,stats] = PLSREGRESS(X,Y, ncomp,
...)

```

\section*{Description}
\([\mathrm{XL}, \mathrm{YL}]=\) plsregress \((\mathrm{X}, \mathrm{Y}, \mathrm{ncomp})\) computes a partial least-squares (PLS) regression of \(Y\) on \(X\), using ncomp PLS components, and returns the predictor and response loadings in XL and YL, respectively. X is an \(n\)-by- \(p\) matrix of predictor variables, with rows corresponding to observations and columns to variables. Y is an \(n\)-by- \(m\) response matrix. XL is a \(p\)-by-ncomp matrix of predictor loadings, where each row contains coefficients that define a linear combination of PLS components that approximate the original predictor variables. YL is an \(m\)-by-ncomp matrix of response loadings, where each row contains coefficients that define a linear combination of PLS components that approximate the original response variables.
[ \(\mathrm{XL}, \mathrm{YL}, \mathrm{XS}\) ] = plsregress ( \(\mathrm{X}, \mathrm{Y}, \mathrm{ncomp}\) ) returns the predictor scores XS , that is, the PLS components that are linear combinations of the variables in X. XS is an \(n\)-by-ncomp orthonormal matrix with rows corresponding to observations and columns to components.
[XL, YL, XS, YS] = plsregress(X,Y, ncomp) returns the response scores YS, that is, the linear combinations of the responses with which the PLS components XS have maximum covariance. YS is an \(n\)-by-ncomp matrix with rows corresponding to observations and columns to components. YS is neither orthogonal nor normalized.
plsregress uses the SIMPLS algorithm, first centering \(X\) and \(Y\) by subtracting off column means to get centered variables XO and YO.

However, it does not rescale the columns. To perform PLS with standardized variables, use zscore to normalize \(X\) and \(Y\).
If ncomp is omitted, its default value is \(\min (\operatorname{size}(X, 1)-1, \operatorname{size}(X, 2))\).
The relationships between the scores, loadings, and centered variables XO and YO are:
\(X L=(X S \backslash X O)^{\prime}=X O^{\prime *} X S\),
YL = (XS \(\backslash Y O)^{\prime}=O^{\prime}{ }^{\prime * X S}\),
\(X L\) and \(Y L\) are the coefficients from regressing \(X O\) and \(Y O\) on \(X S\), and \(X S^{*} \mathrm{XL}^{\prime}\) and \(\mathrm{XS} * \mathrm{YL}\) ' are the PLS approximations to XO and YO .
plsregress initially computes YS as:
YS = YO*YL = YO*YO'*XS,
By convention, however, plsregress then orthogonalizes each column of YS with respect to preceding columns of XS, so that XS' *YS is lower triangular.
[XL, YL, XS , YS , BETA] = PLSREGRESS (X, Y, ncomp, ...) returns the PLS regression coefficients BETA. BETA is a ( \(p+1\) )-by- \(m\) matrix, containing intercept terms in the first row:
```

Y = [ones(n,1),X]*BETA + RESIDUALS,
YO = XO*BETA(2:end,:) + RESIDUALS.

```
[XL,YL,XS, YS, BETA, PCTVAR] = plsregress(X,Y,ncomp) returns a 2-by-ncomp matrix PCTVAR containing the percentage of variance explained by the model. The first row of PCTVAR contains the percentage of variance explained in \(X\) by each PLS component, and the second row contains the percentage of variance explained in Y .
[ \(\mathrm{XL}, \mathrm{YL}, \mathrm{XS}, \mathrm{YS}, \mathrm{BETA}, \mathrm{PCTVAR}, \mathrm{MSE}\) ] = plsregress ( \(\mathrm{X}, \mathrm{Y}, \mathrm{ncomp}\) ) returns a 2 -by-(ncomp+1) matrix MSE containing estimated mean-squared errors for PLS models with 0 : ncomp components. The first row of MSE contains mean-squared errors for the predictor variables in \(X\), and the second row contains mean-squared errors for the response variable(s) in \(Y\).
[XL, YL, XS , YS , BETA, PCTVAR,MSE] = plsregress(..., param1, val1, param2, val2,...) specifies optional parameter name/value pairs from the following table to control the calculation of MSE.
\begin{tabular}{l|l}
\hline Parameter & Value \\
\hline 'cv' & \begin{tabular}{l} 
The method used to compute MSE. \\
- When the value is a positive integer k, plsregress uses k-fold \\
cross-validation.
\end{tabular} \\
& \begin{tabular}{l} 
- When the value is an object of the cvpartition class, other forms of \\
cross-validation can be specified. \\
- When the value is 'resubstitution', plsregress uses X and Y both \\
to fit the model and to estimate the mean-squared errors, without \\
cross-validation.
\end{tabular} \\
\hline 'mcreps' & \begin{tabular}{l} 
Ahe default is 'resubstitution '.
\end{tabular} \\
\hline cross-validation. The default value is 1. The value must be 1 if the value of \\
'cv' is 'resubstitution'.
\end{tabular}

\section*{Parameter Value}

For more information on using parallel computing, see Chapter 17, "Parallel Statistics".
[XL, YL, XS , YS , BETA, PCTVAR, MSE, stats] =
PLSREGRESS ( \(\mathrm{X}, \mathrm{Y}, \mathrm{ncomp}, \ldots\) ) returns a structure stats with the following fields:
- W - A \(p\)-by-ncomp matrix of PLS weights so that XS \(=\mathrm{XO}\) *W.
- T2 - The \(T^{2}\) statistic for each point in XS.
- Xresiduals - The predictor residuals, that is, \(\mathrm{XO}-\mathrm{XS}\) *XL' .
- Yresiduals - The response residuals, that is, YO-XS*YL'.

Examples Load data on near infrared (NIR) spectral intensities of 60 samples of gasoline at 401 wavelengths, and their octane ratings:
```

load spectra
X = NIR;
y = octane;

```

Perform PLS regression with ten components:
```

[XL,yl,XS,YS,beta,PCTVAR] = plsregress(X,y,10);

```

Plot the percent of variance explained in the response variable as a function of the number of components:
```

plot(1:10,cumsum(100*PCTVAR(2,:)),'-bo');
xlabel('Number of PLS components');
ylabel('Percent Variance Explained in y');

```


Compute the fitted response and display the residuals:
```

yfit = [ones(size(X,1),1) X]*beta;
residuals = y-yfit;
stem(residuals)
xlabel('Observation');
ylabel('Residual');

```


\section*{References}

See Also
[1] de Jong, S. "SIMPLS: An Alternative Approach to Partial Least Squares Regression." Chemometrics and Intelligent Laboratory Systems. Vol. 18, 1993, pp. 251-263.
[2] Rosipal, R., and N. Kramer. "Overview and Recent Advances in Partial Least Squares." Subspace, Latent Structure and Feature Selection: Statistical and Optimization Perspectives Workshop (SLSFS 2005), Revised Selected Papers (Lecture Notes in Computer Science 3940). Berlin, Germany: Springer-Verlag, 2006, pp. 34-51.
```

regress | sequentialfs

```

\section*{sobolset.PointOrder property}

Purpose Point generation method
Description The PointOrder property contains a string that specifies the order in which the Sobol sequence points are produced. The property value must be one of 'standard' or 'graycode'. When set to 'standard' the points produced match the original Sobol sequence implementation. When set to 'graycode', the sequence is generated using an implementation that uses the Gray code of the index instead of the index itself.

Purpose
Point set from which stream is drawn
Description
The PointSet property contains a copy of the point set from which the stream is providing points. The point set is specified during construction of a quasi-random stream and cannot subsequently be altered.

\section*{Examples}
```

Q = qrandstream('sobol', 5, 'Skip', 8);
% Create a new stream based on the same sequence as that in Q
Q2 = qrandstream(Q.PointSet);
u1 = qrand (Q, 10)
u2 = qrand(Q2, 10) % contains exactly the same values as u1

```

Purpose Poisson cumulative distribution function
Syntax \(\quad P=\operatorname{poisscdf}(X, l a m b d a)\)
Description
\(P=\) poisscdf( \(X\), lambda) computes the Poisson cdf at each of the values in \(X\) using the corresponding mean parameters in lambda. \(X\) and lambda can be vectors, matrices, or multidimensional arrays that have the same size. A scalar input is expanded to a constant array with the same dimensions as the other input. The parameters in lambda must be positive.

The Poisson cdf is
\[
p=F(x \mid \lambda)=e^{-\lambda} \sum_{i=0}^{\text {floor }(x)} \frac{\lambda^{i}}{i!}
\]

\section*{Examples}

For example, consider a Quality Assurance department that performs random tests of individual hard disks. Their policy is to shut down the manufacturing process if an inspector finds more than four bad sectors on a disk. What is the probability of shutting down the process if the mean number of bad sectors ( \(\lambda\) ) is two?
```

probability = 1-poisscdf(4,2)
probability =
0.0527

```

About \(5 \%\) of the time, a normally functioning manufacturing process produces more than four flaws on a hard disk.

Suppose the average number of flaws ( \(\lambda\) ) increases to four. What is the probability of finding fewer than five flaws on a hard drive?
```

probability = poisscdf(4,4)
probability =
0.6288

```

This means that this faulty manufacturing process continues to operate after this first inspection almost \(63 \%\) of the time.

See Also cdf | poisspdf | poissinv | poisstat | poissfit | poissrnd
How To . "Poisson Distribution" on page B-89

Purpose Poisson parameter estimates
```

Syntax lambdahat = poissfit(data)
[lambdahat,lambdaci] = poissfit(data)
[lambdahat,lambdaci] = poissfit(data,alpha)

```

\section*{Description}
lambdahat \(=\) poissfit(data) returns the maximum likelihood estimate (MLE) of the parameter of the Poisson distribution, \(\lambda\), given the data data.
[lambdahat,lambdaci] = poissfit(data) also gives \(95 \%\) confidence intervals in lamdaci.
[lambdahat,lambdaci] = poissfit(data,alpha) gives 100(1-alpha)\% confidence intervals. For example alpha \(=0.001\) yields \(99.9 \%\) confidence intervals.
The sample mean is the MLE of \(\lambda\).
\[
\hat{\lambda}=\frac{1}{n} \sum_{i=1}^{n} x_{i}
\]

\section*{Examples}
```

r = poissrnd(5,10,2);
[l,lci] = poissfit(r)
l =
7.4000 6.3000
lci =
5.8000 4.8000
9.1000 7.9000

```
See Also mle | poisspdf | poisscdf | poissinv | poisstat | poissrnd
How To . "Poisson Distribution" on page B-89
Purpose Poisson inverse cumulative distribution function
Syntax X = poissinv(P,lambda)
Description \(X=\) poissinv( \(P\), lambda) returns the smallest value \(X\) such that thePoisson cdf evaluated at \(X\) equals or exceeds \(P\), using mean parametersin lambda. P and lambda can be vectors, matrices, or multidimensionalarrays that all have the same size. A scalar input is expanded to aconstant array with the same dimensions as the other input.
Examples If the average number of defects \((\lambda)\) is two, what is the 95 th percentile of the number of defects?
```

poissinv(0.95,2)
ans =

```
    5
What is the median number of defects?
```

median_defects = poissinv(0.50,2)
median_defects =
2

```
See Also icdf | poisscdf | poisspdf | poisstat | poissfit | poissrnd
How To - "Poisson Distribution" on page B-89

\section*{poisspdf}

Purpose Poisson probability density function
Syntax \(\quad Y=\operatorname{poisspdf}(X\), lambda \()\)
Description
\(Y=\) poisspdf( \(X\), lambda) computes the Poisson pdf at each of the values in \(X\) using mean parameters in lambda. \(X\) and lambda can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other input. The parameters in lambda must all be positive.

The Poisson pdf is
\[
y=f(x \mid \lambda)=\frac{\lambda^{x}}{x!} e^{-\lambda} I_{(0,1, \ldots)}(x)
\]
where \(x\) can be any nonnegative integer. The density function is zero unless \(x\) is an integer.

\section*{Examples}

A computer hard disk manufacturer has observed that flaws occur randomly in the manufacturing process at the average rate of two flaws in a 4 GB hard disk and has found this rate to be acceptable. What is the probability that a disk will be manufactured with no defects?

In this problem, \(\lambda=2\) and \(x=0\).
```

p = poisspdf(0,2)
p =
0.1353

```

See Also pdf | poisscdf | poissinv | poisstat | poissfit | poissrnd How To . "Poisson Distribution" on page B-89

\section*{Purpose Poisson random numbers}
```

Syntax R = poissrnd(lambda)
R = poissrnd(lambda,m,n,...)
R = poissrnd(lambda,[m,n,...])

```

Description

Examples
Generate a random sample of 10 pseudo-observations from a Poisson distribution with \(\lambda=2\).
```

lambda = 2;
random_sample1 = poissrnd(lambda,1,10)
random_sample1 =
1
random_sample2 = poissrnd(lambda,[1 10])
random_sample2 =
1
random_sample3 = poissrnd(lambda(ones(1,10)))
random_sample3 =
3

```
See Also random | poisspdf | poisscdf | poissinv | poisstat | poissfit
How To
- "Poisson Distribution" on page B-89

\section*{Purpose Poisson mean and variance}

\author{
Syntax \(\quad M=\operatorname{poisstat}(\) lambda) \\ [M,V] = poisstat(lambda)
}

Description \(\quad M=\) poisstat (lambda) returns the mean of the Poisson distribution using mean parameters in lambda. The size of \(M\) is the size of lambda.
\([M, V]=\) poisstat (lambda) also returns the variance \(V\) of the Poisson distribution.

For the Poisson distribution with parameter \(\lambda\), both the mean and variance are equal to \(\lambda\).

\section*{Examples \\ Find the mean and variance for the Poisson distribution with \(\lambda=2\).}
```

[m,v] = poisstat([1 2; 3 4])
m =
12
3 4
v =
1 2
3 4

```

See Also poisspdf | poisscdf | poissinv | poissfit | poissrnd
How To . "Poisson Distribution" on page B-89

\section*{Purpose}

Polynomial confidence intervals

\section*{Syntax}

Y = polyconf( \(\mathrm{p}, \mathrm{X}\) )
[Y,DELTA] = polyconf(p,X,S)
[Y,DELTA] = polyconf(p,X,S,param1,val1,param2,val2,...)

\section*{Description}
\(Y=\) polyconf( \(p, X)\) evaluates the polynomial \(p\) at the values in \(X . p\) is
a vector of coefficients in descending powers.
[ \(\mathrm{Y}, \mathrm{DELTA}\) ] \(=\) polyconf( \(\mathrm{p}, \mathrm{X}, \mathrm{S})\) takes outputs p and S from polyfit and generates \(95 \%\) prediction intervals \(Y\) - DELTA for new observations at the values in \(X\).
[Y,DELTA] = polyconf(p,X,S,param1,val1,param2,val2,...) specifies optional parameter name/value pairs chosen from the following list.
\begin{tabular}{l|l}
\hline Parameter & Value \\
\hline 'alpha' & \begin{tabular}{l} 
A value between 0 and 1 specifying a confidence level \\
of 100*(1-alpha)\%. The default is 0.05.
\end{tabular} \\
\hline 'mu' & \begin{tabular}{l} 
A two-element vector containing centering and \\
scaling parameters. With this option, polyconf uses \\
(X-mu(1))/mu(2) in place of X.
\end{tabular} \\
\hline 'predopt' & \begin{tabular}{l} 
Either 'observation' (the default) to compute \\
prediction intervals for new observations at the \\
values in X, or 'curve' to compute confidence \\
intervals for the fit evaluated at the values in X. See \\
below.
\end{tabular} \\
\hline 'simopt' & \begin{tabular}{l} 
Either 'off ' (the default) for nonsimultaneous \\
bounds, or 'on' for simultaneous bounds. See below.
\end{tabular} \\
\hline
\end{tabular}

The 'predopt' and 'simopt ' parameters can be understood in terms of the following functions:
- \(p(x)\) - the unknown mean function estimated by the fit
- \(l(x)\) - the lower confidence bound
- \(u(x)\) - the upper confidence bound

Suppose you make a new observation \(y_{n+1}\) at \(x_{n+1}\), so that
\[
y_{n+1}\left(x_{n+1}\right)=p\left(x_{n+1}\right)+\varepsilon_{n+1}
\]

By default, the interval \(\left[l_{n+1}\left(x_{n+1}\right), u_{n+1}\left(x_{n+1}\right)\right]\) is a \(95 \%\) confidence bound on \(y_{n+1}\left(x_{n+1}\right)\).
The following combinations of the 'predopt' and 'simopt' parameters allow you to specify other bounds.
\begin{tabular}{l|l|l}
\hline 'simopt & 'predopt' & Bounded Quantity \\
\hline 'off' & 'observation' & \(y_{n+1}\left(x_{n+1}\right)\) (default) \\
\hline 'off' & 'curve' & \(p\left(x_{n+1}\right)\) \\
\hline 'on' & 'observation' & \(y_{n+1}(x)\), for all \(x\) \\
\hline 'on' & 'curve' & \(p(x)\), for all \(x\) \\
\hline
\end{tabular}

In general, 'observation' intervals are wider than 'curve' intervals, because of the additional uncertainty of predicting a new response value (the curve plus random errors). Likewise, simultaneous intervals are wider than nonsimultaneous intervals, because of the additional uncertainty of bounding values for all predictors \(x\).


\section*{Examples}

This example uses code from the documentation example function polydemo, and calls the documentation example function polystr to convert the coefficient vector \(p\) into a string for the polynomial expression displayed in the figure title. It combines the functions polyfit, polyval, roots, and polyconf to produce a formatted display of data with a polynomial fit.

Note Statistics Toolbox documentation example files are located in the \help\toolbox\stats \(\backslash\) examples subdirectory of your MATLAB root folder (matlabroot). This subdirectory is not on the MATLAB path at installation. To use the files in this subdirectory, either add the subdirectory to the MATLAB path (addpath) or make the subdirectory your current working folder (cd).

Display simulated data with a quadratic trend, a fitted quadratic polynomial, and \(95 \%\) prediction intervals for new observations:
```

xdata = -5:5;
ydata = x.^2 - 5*x - 3 + 5*randn(size(x));
degree = 2; % Degree of the fit
alpha = 0.05; % Significance level
% Compute the fit and return the structure used by
% POLYCONF.
[p,S] = polyfit(xdata,ydata,degree);
% Compute the real roots and determine the extent of the
% data.
r = roots(p)'; % Roots as a row vector.
real_r = r(imag(r) == 0); % Real roots.
% Assure that the data are row vectors.
xdata = reshape(xdata,1,length(xdata));
ydata = reshape(ydata,1,length(ydata));
% Extent of the data.
mx = min([real_r,xdata]);
Mx = max([real_r,xdata]);
my = min([ydata,0]);
My = max([ydata,0]);

```
```

% Scale factors for plotting.
sx = 0.05*(Mx-mx);
sy = 0.05*(My-my);
% Plot the data, the fit, and the roots.
hdata = plot(xdata,ydata,'md','MarkerSize',5,...
'LineWidth',2);
hold on
xfit = mx-sx:0.01:Mx+sx;
yfit = polyval(p,xfit);
hfit = plot(xfit,yfit,'b-','LineWidth',2);
hroots = plot(real_r,zeros(size(real_r)),...
'bo','MarkerSize',5,...
'LineWidth',2,...
'MarkerFaceColor','b');
grid on
plot(xfit,zeros(size(xfit)),'k-','LineWidth',2)
axis([mx-sx Mx+sx my-sy My+sy])
% Add prediction intervals to the plot.
[Y,DELTA] = polyconf(p,xfit,S,'alpha',alpha);
hconf = plot(xfit,Y+DELTA,'b--');
plot(xfit,Y-DELTA,'b--')
% Display the polynomial fit and the real roots.
approx_p = round(100*p)/100; % Round for display.
htitle = title(['{\bf Fit: }',...
texlabel(polystr(approx_p))]);
set(htitle,'Color','b')
approx_real_r = round(100*real_r)/100; % Round for display.
hxlabel = xlabel(['{\bf Real Roots: }',...
num2str(approx_real_r)]);
set(hxlabel,'Color','b')
% Add a legend.
legend([hdata,hfit,hroots,hconf],...
'Data','Fit','Real Roots of Fit',...

```
'95\% Prediction Intervals')
Fit: \(0.81 \mathrm{x}^{2}-4.51 \mathrm{x}-1.19\)


See Also
polyfit | polyval | polytool
Purpose Interactive polynomial fitting
Syntax polytool
polytool(x,y)
polytool( \(x, y, n\) )
polytool(x,y,n,alpha)
polytool( \(x, y, n\), alpha, xname, yname)
h = polytool(...)

\section*{Description}

See Also
polytool
polytool ( \(x, y\) ) fits a line to the vectors \(x\) and \(y\) and displays an interactive plot of the result in a graphical interface. You can use the interface to explore the effects of changing the parameters of the fit and to export fit results to the workspace.
polytool ( \(x, y, n\) ) initially fits a polynomial of degree \(n\). The default is 1 , which produces a linear fit.
polytool(x,y,n,alpha) initially plots 100(1-alpha)\% confidence intervals on the predicted values. The default is 0.05 which results in \(95 \%\) confidence intervals.
polytool( \(x, y, n\), alpha, xname, yname) labels the \(x\) and \(y\) values on the graphical interface using the strings xname and yname. Specify \(n\) and alpha as [] to use their default values.
\(h=\) polytool(...) outputs a vector of handles, \(h\), to the line objects in the plot. The handles are returned in the degree: data, fit, lower bounds, upper bounds.

Purpose Posterior probabilities of components
\(\begin{array}{ll}\text { Syntax } & P=\text { posterior }(o b j, X) \\ & {[P, n l o g l]=\operatorname{posterior}(o b j, X)}\end{array}\)
Description \(\quad P=\) posterior (obj, X) returns the posterior probabilities of each of the \(k\) components in the Gaussian mixture distribution defined by obj for each observation in the data matrix X . X is \(n\)-by- \(d\), where \(n\) is the number of observations and \(d\) is the dimension of the data. obj is an object created by gmdistribution or fit. P is \(n\)-by- \(k\), with \(\mathrm{P}(\mathrm{I}, \mathrm{J})\) the probability of component \(J\) given observation I.
posterior treats NaN values as missing data. Rows of X with NaN values are excluded from the computation.
[P, nlogl] = posterior(obj, X) also returns nlogl, the negative log-likelihood of the data.

\section*{Examples}

Generate data from a mixture of two bivariate Gaussian distributions using the mvnrnd function:
```

MU1 = [1 2];
SIGMA1 = [2 0; 0 .5];
MU2 = [-3 -5];
SIGMA2 = [1 0; 0 1];
X = [mvnrnd(MU1,SIGMA1,1000);mvnrnd(MU2,SIGMA2,1000)];
scatter(X(:,1),X(:,2),10,'.')
hold on

```


Fit a two-component Gaussian mixture model:
```

obj = gmdistribution.fit(X,2);
h = ezcontour(@(x,y)pdf(obj,[x y]),[-8 6],[-8 6]);

```

\section*{gmdistribution.posterior}


Compute posterior probabilities of the components:
```

P = posterior(obj,X);
delete(h)
scatter(X(:,1),X(:,2),10,P(:,1),'.')
hb = colorbar;
ylabel(hb,'Component 1 Probability')

```


See Also
gmdistribution | fit | cluster | mahal

\section*{NaiveBayes.posterior}

\section*{Purpose Compute posterior probability of each class for test data}

\author{
Syntax \\ \section*{Description}
}
post \(=\) posterior(nb,test)
[post,cpre] = posterior(nb,test)
[post,cpre,logp] = posterior(nb,test)
[...] = posterior(..., 'HandleMissing', val)
post = posterior(nb,test) returns the posterior probability of the observations in test according to the NaiveBayes object nb. test is a N -by-nb.ndims matrix, where N is the number of observations in the test data. Rows of test correspond to points, columns of test correspond to features. post is a N -by-nb.nclasses matrix containing the posterior probability of each observation for each class. post (i, \(j\) ) is the posterior probability of point I belonging to class j . Classes are ordered the same as nb.clevels, i.e., column \(j\) of post corresponds to the jth class in nb.clevels. The posterior probabilities corresponding to any empty classes are NaN.
[post, cpre] = posterior(nb,test) returns cpre, an N-by-1 vector, containing the class to which each row of test has been assigned. cpre has the same type as nb.CLevels.
[post, cpre, logp] = posterior(nb,test) returns logp, an N-by-1 vector containing estimates of the log of the probability density function (PDF). logp (i) is the log of the PDF of point i. The PDF value of point \(i\) is the sum of \(\operatorname{Prob}(\) point \(\mathrm{I} \mid\) class J\()\) * \(\operatorname{Pr}\{\) class J\(\}\) taken over all classes.
[...] = posterior(..., 'HandleMissing', val) specifies how posterior treats NaN (missing values). val can be one of the following:
\begin{tabular}{ll} 
'off' \\
(default) & \begin{tabular}{l} 
Observations with NaN in any of the columns are not \\
classified into any class. The corresponding rows in \\
post and logp are NaN. The corresponding rows in \\
cpre are NaN (if obj.clevels is numeric or logical), \\
empty strings (if obj.clevels is char or cell array of \\
strings) or (if obj.clevels is categorical).
\end{tabular} \\
'on' & \begin{tabular}{l} 
For observations having NaN in some (but not all) \\
columns, post and cpre are computed using the \\
columns with non-NaN values. Corresponding logp \\
values are NaN.
\end{tabular}
\end{tabular}

See Also NaiveBayes | fit | predict

\section*{prctile}

Purpose Calculate percentile values
Syntax \(\left.\quad \begin{array}{rl}Y & =\operatorname{prctile}(X, p) \\ & Y\end{array}\right) \operatorname{prctile}(X, p, \operatorname{dim})\)
Description
\(Y=\operatorname{prctile}(X, p)\) returns percentiles of the values in \(X . p\) is a scalar or a vector of percent values. When \(X\) is a vector, \(Y\) is the same size as \(p\) and \(Y(i)\) contains the \(p(i)\) th percentile. When \(X\) is a matrix, the \(i\) th row of \(Y\) contains the \(p(i)\) th percentiles of each column of \(X\). For N -dimensional arrays, prctile operates along the first nonsingleton dimension of \(X\).
\(Y=p r c t i l e(X, p, d i m)\) calculates percentiles along dimension dim.
The dim'th dimension of \(Y\) has length length ( \(p\) ).
Percentiles are specified using percentages, from 0 to 100 . For an \(n\)-element vector X, prctile computes percentiles as follows:

1 The sorted values in X are taken to be the \(100(0.5 / n), 100(1.5 / n), \ldots\), \(100([n-0.5] / n)\) percentiles.

2 Linear interpolation is used to compute percentiles for percent values between 100(0.5/n) and 100 ([n-0.5]/n).

3 The minimum or maximum values in X are assigned to percentiles for percent values outside that range.
prctile treats NaNs as missing values and removes them.

\section*{Examples}
```

x = (1:5)'*(1:5)
x =
1 2 3 3 4 5
2
3
4
5
y = prctile(x,[25 50 75])

```
\begin{tabular}{lrrrr}
\(\mathrm{y}=\) & & & \\
1.7500 & 3.5000 & 5.2500 & 7.0000 & 8.7500 \\
3.0000 & 6.0000 & 9.0000 & 12.0000 & 15.0000 \\
4.2500 & 8.5000 & 12.7500 & 17.0000 & 21.2500
\end{tabular}

\section*{CompactClassificationDiscriminant.predict}

Purpose Predict classification
Syntax label = predict (obj, x)
[label,score] = predict(obj,X)
[label,score,cost] = predict(obj,X)

Description

Input
Arguments

\section*{Output Arguments}
label \(=\) predict \((\mathrm{obj}, \mathrm{X})\) returns a vector of predicted class labels for a matrix X , based on obj, a trained full or compact classifier.
[label,score] = predict(obj,X) returns a matrix of scores (posterior probabilities).
[label, score, cost] = predict(obj, X) returns a matrix of costs; label is the vector or minimal costs for each row of cost.
obj
Discriminant analysis classifier of class ClassificationDiscriminant or CompactClassificationDiscriminant, typically constructed with ClassificationDiscriminant.fit.

X
Matrix where each row represents an observation, and each column represents a predictor. The number of columns in X must equal the number of predictors in obj.
label
Vector of class labels of the same type as the response data used in training obj. Each entry of labels corresponds to a predicted class label for the corresponding row of X; see "Predicted Class Label" on page 20-1422.

\section*{score}

Numeric matrix of size N -by-K, where N is the number of observations (rows) in \(X\), and \(K\) is the number of classes (in

\section*{CompactClassificationDiscriminant.predict}
obj.ClassNames). score( \(i, j\) ) is the posterior probability that row \(i\) of \(X\) is of class \(j\); see "Posterior Probability" on page 20-1421. cost

Matrix of expected costs of size N-by-K. cost (i,j) is the cost of classifying row i of X as class j. See "Cost" on page 20-1422.

\section*{Definitions}

\section*{Posterior Probability}

The posterior probability that a point \(z\) belongs to class \(j\) is the product of the prior probability and the multivariate normal density. The density function of the multivariate normal with mean \(\mu_{j}\) and covariance \(\Sigma_{j}\) at a point \(z\) is
\[
P(x \mid k)=\frac{1}{\left(2 \pi\left|\Sigma_{k}\right|\right)^{1 / 2}} \exp \left(-\frac{1}{2}\left(x-\mu_{k}\right)^{T} \Sigma_{k}^{-1}\left(x-\mu_{k}\right)\right),
\]
where \(\left|\Sigma_{k}\right|\) is the determinant of \(\Sigma_{k}\), and \(\Sigma_{k}^{-1}\) is the inverse matrix.
Let \(P(k)\) represent the prior probability of class \(k\). Then the posterior probability that an observation \(x\) is of class \(k\) is
\[
\hat{P}(k \mid x)=\frac{P(x \mid k) P(k)}{P(x)}
\]
where \(P(x)\) is a normalization constant, the sum over \(k\) of \(P(x \mid k) P(k)\).

\section*{Prior Probability}

The prior probability is one of three choices:
- 'uniform' - The prior probability of class k is one over the total number of classes.
- 'empirical' - The prior probability of class \(k\) is the number of training samples of class \(k\) divided by the total number of training samples.

\section*{CompactClassificationDiscriminant.predict}
- Custom - The prior probability of class k is the kth element of the prior vector. See ClassificationDiscriminant.fit.

After creating a classifier obj, you can set the prior by dot addressing:
```

obj.Prior = v;

```
where \(v\) is a vector of positive elements representing the frequency with which each element occurs. You do not need to retrain the classifier when you set a new prior.

\section*{Cost}

The matrix of expected costs per observation is defined in "Cost" on page 12-8.

\section*{Predicted Class Label}
predict classifies so as to minimize the expected classification cost:
\[
\hat{y}=\underset{y=1, \ldots, K}{\arg \min } \sum_{k=1}^{K} \hat{P}(k \mid x) C(y \mid k),
\]
where
- \(\hat{y}\) is the predicted classification.
- \(K\) is the number of classes.
- \(\hat{P}(k \mid x)\) is the posterior probability of class \(k\) for observation \(x\).
- \(C(y \mid k)\) is the cost of classifying an observation as \(y\) when its true class is \(k\).

\author{
Examples Examine predictions for a few rows in the Fisher iris data: \\ load fisheriris \\ obj = ClassificationDiscriminant.fit(meas,species); X = meas(99:102,:); \% take four rows
}

\section*{CompactClassificationDiscriminant.predict}
```

[label score cost] = predict(obj,X)
label =
'versicolor'
'versicolor'
'virginica'
'virginica'
score =

| 0.0000 | 1.0000 | 0.0000 |
| :--- | :--- | :--- |
| 0.0000 | 0.9999 | 0.0001 |
| 0.0000 | 0.0000 | 1.0000 |
| 0.0000 | 0.0011 | 0.9989 |

cost =
1.0000 0.0000 1.0000
1.0000 0.0001 0.9999
1.0000 1.0000 0.0000
1.0000 0.9989 0.0011

```

\section*{See Also ClassificationDiscriminant | ClassificationDiscriminant.fit | edge | loss | margin}

How To . "Discriminant Analysis" on page 12-3

\section*{CompactClassificationEnsemble.predict}

Purpose Predict classification
Syntax labels = predict(ens, \(x\) )
[labels,score] = predict(ens,X)
[labels,...] = predict(ens,X,Name,Value)

Description

Input
Arguments
labels \(=\) predict (ens, \(X\) ) returns a vector of predicted class labels for a matrix \(X\), based on ens, a trained full or compact classification ensemble.
[labels, score] = predict(ens, X) also returns scores for all classes.
[labels,...] = predict(ens,X,Name,Value) predicts classifications with additional options specified by one or more Name, Value pair arguments.
ens
A classification ensemble created by fitensemble, or a compact classification ensemble created by compact.

X
A matrix where each row represents an observation, and each column represents a predictor. The number of columns in X must equal the number of predictors in ens.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{learners}

Indices of weak learners predict uses for computation of responses, a numeric vector.

Default: 1:T, where \(T\) is the number of weak learners in ens

\section*{CompactClassificationEnsemble.predict}

\section*{UseObsForLearner}

A logical matrix of size N -by-T, where:
- \(N\) is the number of rows of \(X\).
- T is the number of weak learners in ens.

When UseObsForLearner( \(\mathrm{i}, \mathrm{j})\) is true, learner j is used in predicting the class of row i of \(X\).

Default: \(\operatorname{true}(\mathrm{N}, \mathrm{T})\)

Output
Arguments
labels
Vector of classification labels. labels has the same data type as the labels used in training ens.

\section*{score}

A matrix with one row per observation and one column per class. For each observation and each class, the score generated by each tree is the probability of this observation originating from this class computed as the fraction of observations of this class in a tree leaf. predict averages these scores over all trees in the ensemble.

\section*{Definitions}

\section*{Score (ensemble)}

For ensembles, a classification score represents the confidence of a classification into a class. The higher the score, the higher the confidence.

Different ensemble algorithms have different definitions for their scores. Furthermore, the range of scores depends on ensemble type. For example:
- AdaBoostM1 scores range from \(-\infty\) to \(\infty\).
- Bag scores range from 0 to 1 .

\section*{CompactClassificationEnsemble.predict}
```

Examples Train a boosting ensemble for the ionosphere data, and predict the
classification of the mean of the data:
load ionosphere;
ada = fitensemble(X,Y,'AdaBoostM1',100,'tree');
Xbar = mean(X);
[ypredict score] = predict(ada,Xbar)
ypredict =
'g'
score =
-2.9460 2.9460
See Also margin | edge | loss
How To

- Chapter 13, "Nonparametric Supervised Learning"

```

\section*{CompactClassificationTree.predict}

Purpose
Predict classification
Syntax
label = predict(tree, X)
[label, score] = predict(tree, X)
[label, score, node] = predict(tree, X)
[label,score, node, cnum] = predict(tree, X)
[label,...] = predict(tree, X, Name, Value)
Description

\section*{Input Arguments}
label \(=\) predict(tree, \(X\) ) returns a vector of predicted class labels for a matrix \(X\), based on tree, a trained full or compact classification tree.
[label, score] = predict(tree, X) returns a matrix of scores, indicating the likelihood that a label comes from a particular class.
[label, score, node] = predict(tree, X) returns a vector of predicted node numbers for the classification, based on tree.
[label, score, node, cnum] = predict(tree, \(X\) ) returns a vector of predicted class number for the classification, based on tree.
[label,...] = predict(tree, X,Name, Value) returns labels with additional options specified by one or more Name, Value pair arguments.
tree
A classification tree created by ClassificationTree.fit, or a compact classification tree created by compact.

X
A matrix where each row represents an observation, and each column represents a predictor. The number of columns in \(X\) must equal the number of predictors in tree.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{CompactClassificationTree.predict}

\section*{subtrees}

Numeric vector of pruning levels, with 0 representing the full, unpruned tree. To use the subtrees name-value pair, tree must include a pruning sequence as created by the ClassificationTree.fit or prune methods. If subtrees has T elements, and X has N rows, then labels is an N -by- T matrix. The ith column of labels contains the fitted values produced by the subtrees (I) subtree. Similarly, score is an N-by-K-by-T array, and node and cnum are N -by-T matrices. subtrees must be sorted in ascending order. (To compute fitted values for a tree that is not part of the optimal pruning sequence, first use prune to prune the tree.)

Default: 0

\section*{Output Arguments}

\section*{label}

Vector of class labels of the same type as the response data used in training tree. Each entry of labels corresponds to a predicted class label for the corresponding row of \(X\).

\section*{score}

Numeric matrix of size N -by-K, where N is the number of observations (rows) in \(X\), and \(K\) is the number of classes (in tree.ClassNames). score ( \(\mathrm{i}, \mathrm{j}\) ) is the posterior probability that row \(i\) of \(X\) is of class \(j\).
node
Numeric vector of node numbers for the predicted classes. Each entry corresponds to the predicted node in tree for the corresponding row of X .
cnum
Numeric vector of class numbers corresponding to the predicted labels. Each entry of cnum corresponds to a predicted class number for the corresponding row of \(X\).

\section*{CompactClassificationTree.predict}

\section*{Definitions Predictive Measure of Association}

The predictive measure of association between the optimal split on variable \(i\) and a surrogate split on variable \(j\) is:
\[
\lambda_{i, j}=\frac{\min \left(P_{L}, P_{R}\right)-\left(1-P_{L_{i} L_{j}}-P_{R_{i} R_{j}}\right)}{\min \left(P_{L}, P_{R}\right)} .
\]

Here
- \(P_{L}\) and \(P_{R}\) are the node probabilities for the optimal split of node \(i\) into Left and Right nodes respectively.
- \(P_{L_{i} L_{i}}\) is the probability that both (optimal) node \(i\) and (surrogate) node \(j\) send an observation to the Left.
- \(P_{R_{i} R_{j}}\) is the probability that both (optimal) node \(i\) and (surrogate) node \(j\) send an observation to the Right.

Clearly, \(\lambda_{i, j}\) lies from \(-\infty\) to 1 . Variable \(j\) is a worthwhile surrogate split for variable \(i\) if \(\lambda_{i, j}>0\).

\section*{Examples Examine predictions for a few rows in the Fisher iris data:}
```

load fisheriris
tree = ClassificationTree.fit(meas,species);
X = meas(99:102,:); % take four rows
[label score node cnum] = predict(tree,X)
label =
'versicolor'
'versicolor'
'virginica'
'virginica'
score =

```

\section*{CompactClassificationTree.predict}
```

                0 1.0000 0
    0 1.0000 0
0 0.0217 0.9783
0 0.0217 0.9783
node =
8
8
5
5
cnum =
2
2
3
3

```
```

Examine predictions from pruned trees for the Fisher iris model:

```
```

load fisheriris

```
load fisheriris
tree = ClassificationTree.fit(meas,species);
tree = ClassificationTree.fit(meas,species);
X = meas(99:102,:); % taking four rows
X = meas(99:102,:); % taking four rows
[label score node cnum] = predict(tree,X,'subtrees',[2 3 4])
[label score node cnum] = predict(tree,X,'subtrees',[2 3 4])
label =
label =
    'versicolor' 'versicolor' 'setosa'
    'versicolor' 'versicolor' 'setosa'
    'versicolor' 'versicolor' 'setosa'
    'versicolor' 'versicolor' 'setosa'
    'virginica' 'versicolor' 'setosa'
    'virginica' 'versicolor' 'setosa'
    'virginica' 'versicolor' 'setosa'
    'virginica' 'versicolor' 'setosa'
score(:,:,1) =
score(:,:,1) =
        0 0.9074 0.0926
        0 0.9074 0.0926
        0 0.9074 0.0926
        0 0.9074 0.0926
        0 0.0217 0.9783
        0 0.0217 0.9783
        0 0.0217 0.9783
```

        0 0.0217 0.9783
    ```

\section*{CompactClassificationTree.predict}
\(\left.\begin{array}{rrr}\text { score }(:,:, 2)= & \\ 0 & 0.5000 & 0.5000 \\ 0 & 0.5000 & 0.5000 \\ 0 & 0.5000 & 0.5000 \\ 0 & 0.5000 & 0.5000 \\ \text { score }(:,:, 3) & = & \\ 0.3333 & 0.3333 & 0.3333 \\ 0.3333 & 0.3333 & 0.3333 \\ 0.3333 & 0.3333 & 0.3333 \\ 0.3333 & 0.3333 & 0.3333 \\ & & \\ \text { node } & & \\ 4 & 3 & 1\end{array}\right]\)

\section*{Algorithms}
predict generates predictions by following the branches of tree until it reaches a leaf node or a missing value. If predict reaches a leaf node, it returns the classification of that node.

If predict reaches a node with a missing value for a predictor, its behavior depends on the setting of the Surrogate name-value pair when ClassificationTree.fit constructs tree.
- \(\operatorname{Surrogate}=\) 'off ' \((\) default \()-\) predict returns the label with the largest number of training samples that reach the node.
- Surrogate \(=\) ' on' - predict uses the best surrogate split at the node. If all surrogate split variables with positive predictive measure of association are missing, predict returns the label with the largest

\section*{CompactClassificationTree.predict}
number of training samples that reach the node. For a definition, see "Predictive Measure of Association" on page 20-1429.

\author{
See Also ClassificationTree.fit | compact | prune | loss | edge | margin \\ How To • Chapter 13, "Nonparametric Supervised Learning"
}

\section*{CompactRegressionEnsemble.predict}

\section*{Purpose Predict response of ensemble}
```

Syntax
Yfit = predict(ens,Xdata)
Yfit = predict(ens,Xdata,Name,Value)

```

Description \(\quad\) Yfit \(=\) predict (ens, Xdata) returns predicted responses to the data in Xdata, based on the ens regression ensemble model.

Yfit = predict(ens,Xdata, Name, Value) predicts with additional options specified by one or more Name, Value pair arguments.

\section*{Input Arguments}

Regression ensemble created by fitensemble, or by the compact method.

\section*{Xdata}

Numeric array with the same number of columns as the array used for creating ens. Each row of Xdata corresponds to one data point, and each column corresponds to one predictor.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

\section*{learners}

Indices of weak learners in the ensemble ranging from 1 to ens.NTrained. oobEdge uses only these learners for calculating loss.

Default: 1:NTrained
UseObsForLearner

\section*{CompactRegressionEnsemble.predict}
ical matrix of size N -by-NTrained, where N is the number of observations in ens. \(X\), and NTrained is the number of weak learners. When UseObsForLearner \((I, J)\) is true, predict uses learner \(J\) in predicting observation I.

Default: true (N,NTrained)

\section*{Output Arguments}
Examples

Find the predicted mileage for a four-cylinder car, with 200 cubic inch
 engine displacement, 150 horsepower, weighing 3000 lbs , based on the
 carsmall data:
```

load carsmall
X = [Cylinders Displacement Horsepower Weight];
rens = fitensemble(X,MPG,'LSBoost',100,'Tree');
Mileage = predict(rens,[4 200 150 3000])
Mileage =
20.4982

```
See Also ..... loss | fitensemble
How To - Chapter 13, "Nonparametric Supervised Learning"

\section*{CompactRegressionTree.predict}

Purpose
Predict response of regression tree
Syntax
Yfit = predict(tree,Xdata)
[Yfit,node] = predict(tree,Xdata)
[Yfit, node] = predict(tree,Xdata, Name, Value)

\section*{Description}

Input
Arguments
Yfit \(=\) predict(tree, Xdata) returns predicted responses to the data in Xdata, based on the tree regression tree.
[Yfit, node] = predict(tree,Xdata) returns the predicted node numbers of tree in response to Xdata.
[Yfit, node] = predict(tree,Xdata,Name, Value) predicts response with additional options specified by one or more Name, Value pair arguments.
tree
Regression tree created by RegressionTree.fit, or by the compact method.

Xdata
Numeric array with the same number of columns as the array used for creating tree. Each row of Xdata corresponds to one data point, and each column corresponds to one predictor.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{subtrees}

Numeric vector of pruning levels, with 0 representing the full, unpruned tree. To use the subtrees name-value pair, tree must include a pruning sequence as created by the RegressionTree.fit or prune methods. If subtrees has \(T\)

\section*{CompactRegressionTree.predict}
nts, and \(X\) has \(N\) rows, then Yfit is an N -by- \(T\) matrix. The ith column of Yfit contains the fitted values produced by the subtrees (I) subtree. Similarly, node is an N-by-T matrix. subtrees must be sorted in ascending order. (To compute fitted values for a tree that is not part of the optimal pruning sequence, first use prune to prune the tree.)

Default: 0

\section*{Output Yfit Arguments}

A numeric column vector with the same number of rows as Xdata. Each row of Yfit gives the predicted response to the corresponding row of Xdata, based on the tree regression model. node

Numeric vector of node numbers for the predictions. Each entry corresponds to the predicted leaf node in tree for the corresponding row of Xdata.

\section*{Examples Find the predicted mileage for a car with 200 cubic inch engine displacement, 150 horsepower, weighing 3000 lbs , based on the carsmall data:}
```

load carsmall
X = [Displacement Horsepower Weight];
tree = RegressionTree.fit(X,MPG);
Mileage = predict(tree,[200 150 3000])
Mileage =
21.9375

```

\section*{See Also RegressionTree.fit | compact | loss}

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\section*{CompactTreeBagger.predict}
\begin{tabular}{|c|c|}
\hline Purpose & Predict response \\
\hline \multirow[t]{5}{*}{Syntax} & YFIT \(=\) predict (B, X \\
\hline & [YFIT, stdevs] = predict (B,X) \\
\hline & [YFIT, scores] = predict (B,X) \\
\hline & [YFIT, scores, stdevs] = predict (B,X) \\
\hline & Y = predict (B, X, 'param1', val1, 'param2', val2,...) \\
\hline
\end{tabular}

Description
YFIT \(=\) predict \((B, X)\) computes the predicted response of the trained ensemble B for predictors X. By default, predict takes a democratic (nonweighted) average vote from all trees in the ensemble. In X, rows represent observations and columns represent variables. YFIT is a cell array of strings for classification and a numeric array for regression.

For regression, [YFIT, stdevs] = predict (B,X) also returns standard deviations of the computed responses over the ensemble of the grown trees.

For classification, [YFIT, scores] = predict (B,X) returns scores for all classes. scores is a matrix with one row per observation and one column per class. For each observation and each class, the score generated by each tree is the probability of this observation originating from this class computed as the fraction of observations of this class in a tree leaf. predict averages these scores over all trees in the ensemble.
[YFIT, scores, stdevs] = predict (B,X) also returns standard deviations of the computed scores for classification. stdevs is a matrix with one row per observation and one column per class, with standard deviations taken over the ensemble of the grown trees.

Y = predict( \(B, X\), 'param1', val1, 'param2', val2, ...) specifies optional parameter name/value pairs:

\section*{CompactTreeBagger.predict}
\begin{tabular}{ll} 
'trees' & \begin{tabular}{l} 
Array of tree indices to use for computation of \\
responses. Default is 'all'.
\end{tabular} \\
'treeweights' & \begin{tabular}{l} 
Array of NTrees weights for weighting votes from \\
the specified trees.
\end{tabular} \\
'useifort' & \begin{tabular}{l} 
Logical matrix of size Nobs-by-NTrees indicating \\
which trees to use to make predictions for each \\
observation. By default all trees are used for all \\
observations.
\end{tabular}
\end{tabular}

\section*{See Also}
classregtree.eval | TreeBagger.predict

\section*{NaiveBayes.predict}
Purpose Predict class label for test data
Syntax

cpre = predict(nb,test)

cpre = predict(...,'HandleMissing', val)

\section*{Description}
cpre \(=\) predict (nb,test) classifies each row of data in test into one of the classes according to the NaiveBayes classifier nb, and returns the predicted class level cpre. test is an \(N\)-by-nb. ndims matrix, where \(N\) is the number of observations in the test data. Rows of test correspond to points, columns of test correspond to features. cpre is an N-by-1 vector of the same type as nb. CLevels, and it indicates the class to which each row of test has been assigned.
cpre = predict(...,'HandleMissing', val) specifies how predict treats NaN (missing values). val can be one of the following:
\begin{tabular}{ll} 
'off' \\
(default) & \begin{tabular}{l} 
Observations with NaN in any of the columns are \\
not classified into any class. The corresponding \\
rows in post and logp are NaN. The corresponding \\
rows in cpre are NaN (if obj.clevels is numeric or
\end{tabular} \\
logical), empty strings (if obj.clevels is char or cell \\
array of strings) or <undefined> (if obj.clevels is \\
categorical).
\end{tabular}

\footnotetext{
See Also NaiveBayes | fit | posterior
}

\section*{TreeBagger.predict}

Purpose Predict response
```

Syntax $\quad Y=\operatorname{predict}(B, X)$
[Y,stdevs] = predict(B,X)
[Y,scores] = predict(B, X)
[Y,scores,stdevs] = predict(B,X)
$Y=\operatorname{predict}(B, X, ' p a r a m 1 ', v a l 1, ' p a r a m 2 ', v a l 2, \ldots)$

```

\section*{Description}
\(Y=p r e d i c t(B, X)\) computes predicted response of the trained ensemble B for data \(X\). The output has one prediction for each row of \(X\). The returned \(Y\) is a cell array of strings for classification and a numeric array for regression.

For regression, \([\mathrm{Y}\), stdevs] \(=\) predict \((B, X)\) also returns standard deviations of the computed responses over the ensemble of the grown trees.

For classification, [ Y, scores] \(=\) predict \((\mathrm{B}, \mathrm{X})\) returns scores for all classes. scores is a matrix with one row per observation and one column per class. For each observation and each class, the score generated by each tree is the probability of this observation originating from this class computed as the fraction of observations of this class in a tree leaf. predict averages these scores over all trees in the ensemble.
[ Y, scores, stdevs] \(=\) predict \((\mathrm{B}, \mathrm{X})\) also returns standard deviations of the computed scores for classification. stdevs is a matrix with one row per observation and one column per class, with standard deviations taken over the ensemble of the grown trees.

Y = predict(B,X,'param1',val1,'param2', val2,...) specifies optional parameter name/value pairs:
\begin{tabular}{ll} 
'trees' & \begin{tabular}{l} 
Array of tree indices to use for computation of \\
responses. Default is 'all'.
\end{tabular} \\
'treeweights' & \begin{tabular}{l} 
Array of NTrees weights for weighting votes from \\
the specified trees.
\end{tabular} \\
'useifort' & \begin{tabular}{l} 
Logical matrix of size Nobs-by-NTrees indicating \\
which trees to use to make predictions for each \\
observation. By default all trees are used for all \\
observations.
\end{tabular}
\end{tabular}

\section*{See Also}

CompactTreeBagger.predict

\section*{CompactClassificationEnsemble.predictorImportance}
\begin{tabular}{ll} 
Purpose & Estimates of predictor importance \\
Syntax & \begin{tabular}{l} 
imp \(=\) predictorImportance(ens) \\
{\([i m p, m a]=\) predictorImportance (ens) }
\end{tabular} \\
&
\end{tabular}

Description

\section*{Input}

Arguments

\section*{Output \\ imp}

Arguments
ens
imp = predictorImportance(ens) computes estimates of predictor importance for ens by summing these estimates over all weak learners in the ensemble. imp has one element for each input predictor in the data used to train this ensemble. A high value indicates that this predictor is important for ens.
[imp,ma] = predictorImportance(ens) returns a P-by-P matrix with predictive measures of association for \(P\) predictors, when the learners in ens contain surrogate splits. See "Definitions" on page 20-1443.

A classification ensemble created by fitensemble, or by the compact method.

A row vector with the same number of elements as the number of predictors (columns) in ens.X. The entries are the estimates of predictor importance, with 0 representing the smallest possible importance.

A P-by-P matrix of predictive measures of association for \(P\) predictors. Element ma \((I, J)\) is the predictive measure of association averaged over surrogate splits on predictor \(J\) for which predictor I is the optimal split predictor. predictor Importance averages this predictive measure of association over all trees in the ensemble.

\section*{CompactClassificationEnsemble.predictorImportance}

\section*{Definitions}

\section*{Predictor Importance}
predictorImportance computes estimates of predictor importance for ens by summing changes in the risk due to splits on every predictor and dividing the sum by the number of tree nodes. If ens is grown without surrogate splits, this sum is taken over best splits found at each branch node. If ens is grown with surrogate splits, this sum is taken over all splits at each branch node including surrogate splits. imp has one element for each input predictor in the data used to train ens. Predictor importance associated with this split is computed as the difference between the risk for the parent node and the total risk for the two children.

\section*{Impurity and Node Error}

ClassificationTree splits nodes based on either impurity or node error. Impurity means one of several things, depending on your choice of the SplitCriterion name-value pair:
- Gini's Diversity Index (gdi) - The Gini index of a node is
\[
1-\sum_{i} p^{2}(i)
\]
where the sum is over the classes \(i\) at the node, and \(p(i)\) is the observed fraction of classes with class \(i\) that reach the node. A node with just one class (a pure node) has Gini index 0; otherwise the Gini index is positive. So the Gini index is a measure of node impurity.
- Deviance ('deviance') — With \(p(i)\) defined as for the Gini index, the deviance of a node is
\[
-\sum_{i} p(i) \log p(i)
\]

A pure node has deviance 0 ; otherwise, the deviance is positive.
- Twoing rule ('twoing') — Twoing is not a purity measure of a node, but is a different measure for deciding how to split a node. Let \(L(i)\) denote the fraction of members of class \(i\) in the left child node after a

\section*{CompactClassificationEnsemble.predictorImportance}
split, and \(R(i)\) denote the fraction of members of class \(i\) in the right child node after a split. Choose the split criterion to maximize
\[
P(L) P(R)\left(\sum_{i}|L(i)-R(i)|\right)^{2}
\]
where \(P(L)\) and \(P(R)\) are the fractions of observations that split to the left and right respectively. If the expression is large, the split made each child node purer. Similarly, if the expression is small, the split made each child node similar to each other, and hence similar to the parent node, and so the split did not increase node purity.
- Node error - The node error is the fraction of misclassified classes at a node. If \(j\) is the class with largest number of training samples at a node, the node error is
\[
1-p(j)
\]

\section*{Predictive Measure of Association}

The predictive measure of association between the optimal split on variable \(i\) and a surrogate split on variable \(j\) is:
\[
\lambda_{i, j}=\frac{\min \left(P_{L}, P_{R}\right)-\left(1-P_{L_{i} L_{j}}-P_{R_{i} R_{j}}\right)}{\min \left(P_{L}, P_{R}\right)} .
\]

\section*{Here}
- \(P_{L}\) and \(P_{R}\) are the node probabilities for the optimal split of node \(i\) into Left and Right nodes respectively.
- \(P_{L_{i} L_{j}}\) is the probability that both (optimal) node \(i\) and (surrogate) node \(j\) send an observation to the Left.
- \(P_{R_{i} R_{j}}\) is the probability that both (optimal) node \(i\) and (surrogate) node \(j\) send an observation to the Right.

\section*{CompactClassificationEnsemble.predictorImportance}

Clearly, \(\lambda_{i, j}\) lies from \(-\infty\) to 1 . Variable \(j\) is a worthwhile surrogate split for variable \(i\) if \(\lambda_{i, j}>0\).
Element ma( \(\mathrm{i}, \mathrm{j})\) is the predictive measure of association averaged over surrogate splits on predictor \(j\) for which predictor \(i\) is the optimal split predictor. This average is computed by summing positive values of the predictive measure of association over optimal splits on predictor i and surrogate splits on predictor \(j\) and dividing by the total number of optimal splits on predictor \(i\), including splits for which the predictive measure of association between predictors \(i\) and \(j\) is negative.

\section*{Examples}

Estimate the predictor importance for all variables in the Fisher iris data:
```

load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
imp = predictorImportance(ens)
imp =
0.0001 0.0005 0.0384 0.0146

```

The first two predictors are not very important in ens.

Estimate the predictor importance for all variables in the Fisher iris data for an ensemble where the trees contain surrogate splits:
```

load fisheriris
surrtree = ClassificationTree.template('Surrogate','on');
ens2 = fitensemble(meas,species,'AdaBoostM2',100,surrtree);
[imp2,ma] = predictorImportance(ens2)
imp2 =
0.0224 0.0142 0.0525 0.0508
ma =
1.0000 0 0.0001 0.0001
0.0115 1.0000 0.0023 0.0054

```

\section*{CompactClassificationEnsemble.predictorImportance}
\begin{tabular}{llll}
0.2810 & 0.1747 & 1.0000 & 0.5372 \\
0.0789 & 0.0463 & 0.2339 & 1.0000
\end{tabular}

The first two predictors show much more importance than in the previous example.

\section*{See Also \\ predictorImportance}

How To • Chapter 13, "Nonparametric Supervised Learning"

\section*{CompactClassificationTree.predictorImportance}

Purpose
Syntax imp = predictorImportance(tree)
Description

\section*{Input Arguments}

\section*{Output imp \\ Arguments}

\section*{Definitions}

Estimates of predictor importance

\section*{tree} the compact method. importance.

\section*{Predictor Importance}
imp = predictorImportance(tree) computes estimates of predictor importance for tree by summing changes in the risk due to splits on every predictor and dividing the sum by the number of tree nodes.

A classification tree created by ClassificationTree.fit, or by

A row vector with the same number of elements as the number of predictors (columns) in tree. X. The entries are the estimates of predictor importance, with 0 representing the smallest possible
predictor Importance computes estimates of predictor importance for tree by summing changes in the risk due to splits on every predictor and dividing the sum by the number of tree nodes. If tree is grown without surrogate splits, this sum is taken over best splits found at each branch node. If tree is grown with surrogate splits, this sum is taken over all splits at each branch node including surrogate splits. imp has one element for each input predictor in the data used to train tree. Predictor importance associated with this split is computed as the difference between the risk for the parent node and the total risk for the two children.

Estimates of predictor importance do not depend on the order of predictors if you use surrogate splits, but do depend on the order if you do not use surrogate splits.
If you use surrogate splits, predictorImportance computes estimates before the tree is reduced by pruning or merging leaves. If you do not

\section*{CompactClassificationTree.predictorImportance}
use surrogate splits, predictorImportance computes estimates after the tree is reduced by pruning or merging leaves. Therefore, reducing the tree by pruning affects the predictor importance for a tree grown without surrogate splits, and does not affect the predictor importance for a tree grown with surrogate splits.

\section*{Impurity and Node Error}

ClassificationTree splits nodes based on either impurity or node error. Impurity means one of several things, depending on your choice of the SplitCriterion name-value pair:
- Gini's Diversity Index (gdi) - The Gini index of a node is
\[
1-\sum_{i} p^{2}(i),
\]
where the sum is over the classes \(i\) at the node, and \(p(i)\) is the observed fraction of classes with class \(i\) that reach the node. A node with just one class (a pure node) has Gini index 0; otherwise the Gini index is positive. So the Gini index is a measure of node impurity.
- Deviance ('deviance') — With \(p(i)\) defined as for the Gini index, the deviance of a node is
\[
-\sum_{i} p(i) \log p(i)
\]

A pure node has deviance 0 ; otherwise, the deviance is positive.
- Twoing rule ('twoing') - Twoing is not a purity measure of a node, but is a different measure for deciding how to split a node. Let \(L(i)\) denote the fraction of members of class \(i\) in the left child node after a split, and \(R(i)\) denote the fraction of members of class \(i\) in the right child node after a split. Choose the split criterion to maximize
\[
P(L) P(R)\left(\sum_{i}|L(i)-R(i)|\right)^{2}
\]

\section*{CompactClassificationTree.predictorImportance}
where \(P(L)\) and \(P(R)\) are the fractions of observations that split to the left and right respectively. If the expression is large, the split made each child node purer. Similarly, if the expression is small, the split made each child node similar to each other, and hence similar to the parent node, and so the split did not increase node purity.
- Node error - The node error is the fraction of misclassified classes at a node. If \(j\) is the class with largest number of training samples at a node, the node error is
\[
1-p(j)
\]

\section*{Examples}

Estimate the predictor importance for all variables in the Fisher iris data:
```

load fisheriris
tree = ClassificationTree.fit(meas,species);
imp = predictorImportance(tree)
imp =

```
\[
\begin{array}{llll}
0 & 0 & 0.0403 & 0.0303
\end{array}
\]

The first two elements of imp are zero. Therefore, the first two predictors do not enter into tree calculations for classifying irises.

Estimate the predictor importance for all variables in the Fisher iris data for a tree grown with surrogate splits:
```

tree2 = ClassificationTree.fit(meas,species,...
'Surrogate','on');
imp2 = predictorImportance(tree2)
imp2 =
0.0287 0.0136 0.0560 0.0556

```

In this case, all predictors have some importance. As you expect by comparing to the first example, the first two predictors are less important than the final two.

\section*{CompactClassificationTree.predictorImportance}

Estimates of predictor importance do not depend on the order of predictors if you use surrogate splits, but do depend on the order if you do not use surrogate splits. For example, permute the order of the data columns in the previous example:
```

load fisheriris
meas3 = meas(:,[4 1 3 2]);
tree3 = ClassificationTree.fit(meas3,species);
imp3 = predictorImportance(tree2)
imp3 =
0.0674 0 0.0033 0

```

The estimates of predictor importance are not a permutation of imp from the first example.

Estimate the predictor importance using surrogate splits.
```

tree4 = ClassificationTree.fit(meas3,species,...
'Surrogate','on');
imp4 = predictorImportance(tree4)
imp4 =
0.0556 0.0287 0.0560 0.0136

```
imp4 is a permutation of imp2, demonstrating that estimates of predictor importance do not depend on the order of predictors with surrogate splits.

\section*{See Also}
predictorImportance

\section*{How To \\ - Chapter 13, "Nonparametric Supervised Learning"}

\section*{CompactRegressionEnsemble.predictorImportance}
\begin{tabular}{ll} 
Purpose & Estimates of predictor importance \\
Syntax & \begin{tabular}{l} 
imp \(=\) predictorImportance(ens) \\
{\([i m p, m a]=\) predictorImportance(ens) }
\end{tabular} \\
&
\end{tabular}

Description

\section*{Input \\ Arguments}

\section*{Output \\ Arguments}

\section*{ens}
imp \(=\) predictorImportance(ens) computes estimates of predictor importance for ens by summing these estimates over all weak learners in the ensemble. imp has one element for each input predictor in the data used to train this ensemble. A high value indicates that this predictor is important for ens.
[imp,ma] = predictorImportance(ens) returns a P-by-P matrix with predictive measures of association for P predictors.

A regression ensemble created by fitensemble, or by the compact method.

A row vector with the same number of elements as the number of predictors (columns) in ens.X. The entries are the estimates of predictor importance, with 0 representing the smallest possible importance.
ma
A P-by-P matrix of predictive measures of association for \(P\) predictors. Element ma \((I, J)\) is the predictive measure of association averaged over surrogate splits on predictor \(J\) for which predictor I is the optimal split predictor. predictor Importance averages this predictive measure of association over all trees in the ensemble.

\section*{Definitions}

\section*{Predictor Importance}
predictorImportance computes estimates of predictor importance for tree by summing changes in the mean squared error (MSE) due to

\section*{CompactRegressionEnsemble.predictorImportance}
splits on every predictor and dividing the sum by the number of tree nodes. If the tree is grown without surrogate splits, this sum is taken over best splits found at each branch node. If the tree is grown with surrogate splits, this sum is taken over all splits at each branch node including surrogate splits. imp has one element for each input predictor in the data used to train this tree. At each node, MSE is estimated as node error weighted by the node probability. Variable importance associated with this split is computed as the difference between MSE for the parent node and the total MSE for the two children.

\section*{Predictive Measure of Association}

The predictive measure of association between the optimal split on variable \(i\) and a surrogate split on variable \(j\) is:
\[
\lambda_{i, j}=\frac{\min \left(P_{L}, P_{R}\right)-\left(1-P_{L_{i} L_{j}}-P_{R_{i} R_{j}}\right)}{\min \left(P_{L}, P_{R}\right)}
\]

Here
- \(P_{L}\) and \(P_{R}\) are the node probabilities for the optimal split of node \(i\) into Left and Right nodes respectively.
- \(P_{L_{i} L_{j}}\) is the probability that both (optimal) node \(i\) and (surrogate) node \(j\) send an observation to the Left.
- \(P_{R_{i} R_{j}}\) is the probability that both (optimal) node \(i\) and (surrogate) node \(j\) send an observation to the Right.

Clearly, \(\lambda_{i, j}\) lies from \(-\infty\) to 1 . Variable \(j\) is a worthwhile surrogate split for variable \(i\) if \(\lambda_{i, j}>0\).
Element ma(i,j) is the predictive measure of association averaged over surrogate splits on predictor \(j\) for which predictor \(i\) is the optimal split predictor. This average is computed by summing positive values of the predictive measure of association over optimal splits on predictor i and surrogate splits on predictor j and dividing by the total number of

\section*{CompactRegressionEnsemble.predictorImportance}

\section*{Examples}
optimal splits on predictor i, including splits for which the predictive measure of association between predictors \(i\) and \(j\) is negative.

Estimate the predictor importance for all numeric variables in the carsmall data:
```

load carsmall
X = [Acceleration Cylinders Displacement ...
Horsepower Model_Year Weight];
ens = fitensemble(X,MPG,'LSBoost',100,'Tree');
imp = predictorImportance(ens)
imp =
0.0082 0

```

The weight (last predictor) has the most impact on mileage (MPG). The second predictor has importance 0 ; this means the number of cylinders has no impact on predictions made with ens.

Estimate the predictor importance for all variables in the carsmall data for an ensemble where the trees contain surrogate splits:
```

load carsmall
surrtree = RegressionTree.template('Surrogate','on');
X = [Acceleration Cylinders Displacement ...
Horsepower Model_Year Weight];
ens2 = fitensemble(X,MPG,'LSBoost',100,surrtree);
[imp2,ma] = predictorImportance(ens2)
imp2 =
0.0725 0.1342 0.1425 0.1397 0.1380
ma =

| 1.0000 | 0.0414 | 0.0607 | 0.0782 | 0.0102 | 0.0322 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 0 | 1.0000 | 0 | 0 | 0 | 0 |
| 0.0441 | 0.0704 | 1.0000 | 0.0883 | 0.0175 | 0.0913 |
| 0.0944 | 0.1166 | 0.1400 | 1.0000 | 0.0390 | 0.1308 |

```

\section*{CompactRegressionEnsemble.predictorImportance}
\begin{tabular}{llllll}
0.0121 & 0.0139 & 0.0127 & 0.0127 & 1.0000 & 0.0113 \\
0.0818 & 0.1317 & 0.2072 & 0.1878 & 0.0340 & 1.0000
\end{tabular}

While weight (last predictor) still has the most impact on mileage (MPG), this estimate has the second predictor (number of cylinders) is essentially tied for third most important predictor.

\section*{See Also}
predictorImportance

\section*{How To}
- Chapter 13, "Nonparametric Supervised Learning"

\section*{CompactRegressionTree.predictorImportance}

Purpose
Syntax imp = predictorImportance(tree)
Estimates of predictor importance

\section*{Input Arguments}

\section*{Output \\ imp} Arguments

\section*{Definitions}
tree
imp \(=\) predictorImportance(tree) computes estimates of predictor importance for tree by summing changes in the mean squared error due to splits on every predictor and dividing the sum by the number of tree nodes.

A regression tree created by RegressionTree.fit, or by the compact method.

A row vector with the same number of elements as the number of predictors (columns) in tree. X . The entries are the estimates of predictor importance, with 0 representing the smallest possible importance.

\section*{Predictor Importance}
predictor Importance computes estimates of predictor importance for tree by summing changes in the mean squared error (MSE) due to splits on every predictor and dividing the sum by the number of tree nodes. If the tree is grown without surrogate splits, this sum is taken over best splits found at each branch node. If the tree is grown with surrogate splits, this sum is taken over all splits at each branch node including surrogate splits. imp has one element for each input predictor in the data used to train this tree. At each node, MSE is estimated as node error weighted by the node probability. Variable importance associated with this split is computed as the difference between MSE for the parent node and the total MSE for the two children.
Estimates of predictor importance do not depend on the order of predictors if you use surrogate splits, but do depend on the order if you do not use surrogate splits.

\section*{CompactRegressionTree.predictorImportance}

If you use surrogate splits, predictorImportance computes estimates before the tree is reduced by pruning or merging leaves. If you do not use surrogate splits, predictorImportance computes estimates after the tree is reduced by pruning or merging leaves. Therefore, reducing the tree by pruning affects the predictor importance for a tree grown without surrogate splits, and does not affect the predictor importance for a tree grown with surrogate splits.

\section*{Examples}

Find predictor importance for the carsmall data. Use just the numeric predictors:
```

load carsmall
X = [Acceleration Cylinders Displacement ...
Horsepower Model_Year Weight];
tree = RegressionTree.fit(X,MPG);
imp = predictorImportance(tree)
imp =
0.0315 0 0.1082 0.0686 0.1629 1.2924

```

The weight (last predictor) has the most impact on mileage (MPG). The second predictor has importance 0 ; this means the number of cylinders has no impact on predictions made with tree.

Estimate the predictor importance for all variables in the carsmall data for a tree grown with surrogate splits:
```

load carsmall
X = [Acceleration Cylinders Displacement ...
Horsepower Model_Year Weight];
tree2 = RegressionTree.fit(X,MPG,...
'Surrogate','on');
imp2 = predictorImportance(tree2)
imp2 =
0.5287 1.1977 1.2400 0.7059 1.0677 1.4106

```

\section*{CompactRegressionTree.predictorImportance}
While weight (last predictor) still has the most impact on mileage (MPG), this estimate has the second predictor (number of cylinders) as the third most important predictor.

\author{
See Also \\ predictorImportance \\ How To • Chapter 13, "Nonparametric Supervised Learning"
}

Purpose
Principal component analysis (PCA) on data

\section*{Syntax}
[COEFF,SCORE] = princomp(X)
[COEFF, SCORE, latent] = princomp(X)
[COEFF, SCORE, latent, tsquare] = princomp(X)
[...] \(=\) princomp( \(X\), 'econ')

COEFF = princomp (X) performs principal components analysis (PCA) on the \(n\)-by- \(p\) data matrix X , and returns the principal component coefficients, also known as loadings. Rows of X correspond to observations, columns to variables. COEFF is a \(p\)-by- \(p\) matrix, each column containing coefficients for one principal component. The columns are in order of decreasing component variance.
princomp centers \(X\) by subtracting off column means, but does not rescale the columns of \(X\). To perform principal components analysis with standardized variables, that is, based on correlations, use princomp (zscore(X)). To perform principal components analysis directly on a covariance or correlation matrix, use pcacov.
[COEFF,SCORE] = princomp(X) returns SCORE, the principal component scores; that is, the representation of \(X\) in the principal component space. Rows of SCORE correspond to observations, columns to components.
[COEFF, SCORE, latent] = princomp(X) returns latent, a vector containing the eigenvalues of the covariance matrix of \(X\).
[COEFF,SCORE, latent,tsquare] = princomp(X) returns tsquare, which contains Hotelling's \(\mathrm{T}^{2}\) statistic for each data point.

The scores are the data formed by transforming the original data into the space of the principal components. The values of the vector latent are the variance of the columns of SCORE. Hotelling's \(T^{2}\) is a measure of the multivariate distance of each observation from the center of the data set.

When \(n<=p, \operatorname{SCORE}(:, n: p)\) and latent \((n: p)\) are necessarily zero, and the columns of \(\operatorname{COEFF}(:, n: p)\) define directions that are orthogonal to X .
[...] = princomp(X,'econ') returns only the elements of latent that are not necessarily zero, and the corresponding columns of COEFF and SCORE, that is, when \(n<=p\), only the first \(n-1\). This can be significantly faster when \(p\) is much larger than \(n\).

\section*{Examples}

Compute principal components for the ingredients data in the Hald data set, and the variance accounted for by each component.
```

load hald;
[pc,score,latent,tsquare] = princomp(ingredients);
pc,latent
pc =
0.0678 -0.6460 0.5673-0.5062
0.6785-0.0200 -0.5440-0.4933
-0.0290 0.7553 0.4036 -0.5156
-0.7309 -0.1085 -0.4684 -0.4844
latent =
517.7969
67.4964
12.4054
0.2372

```

The following command and plot show that two components account for \(98 \%\) of the variance:
```

cumsum(latent)./sum(latent)
ans =
0.86597
0.97886
0.9996
1
biplot(pc(:,1:2),'Scores',score(:,1:2),'VarLabels',...
{'X1' 'X2' 'X3' 'X4'})

```


For a more detailed example and explanation of this analysis method, see "Principal Component Analysis (PCA)" on page 10-31.

\section*{References}
[1] Jackson, J. E., A User's Guide to Principal Components, John Wiley and Sons, 1991, p. 592.
[2] Jolliffe, I. T., Principal Component Analysis, 2nd edition, Springer, 2002.
[3] Krzanowski, W. J. Principles of Multivariate Analysis: A User's Perspective. New York: Oxford University Press, 1988.
[4] Seber, G. A. F., Multivariate Observations, Wiley, 1984.

\section*{See Also \\ barttest | biplot | canoncorr | factoran | pcacov | pcares | rotatefactors}

How To
- "Principal Component Analysis (PCA)" on page 10-31

\section*{TreeBagger.Prior property}

\section*{Purpose Prior class probabilities}

Description The Prior property is a vector with prior probabilities for classes. This property is empty for ensembles of regression trees.

\author{
See Also \\ classregtree
}

\section*{Purpose \\ Description}

Object representing probability distribution

\section*{Construction}

Methods

\section*{Properties}

InputData

Support

Return cumulative distribution function (CDF) for ProbDist object

Return probability density function (PDF) for ProbDist object

Generate random number drawn from ProbDist object

Read-only string containing probability distribution name of ProbDist object

Read-only structure containing information about input data to ProbDist object

Read-only structure containing information about support of ProbDist object

\section*{Copy \\ Semantics}

Value. To learn how this affects your use of the class, see Copying Objects in the MATLAB Programming Fundamentals documentation.

\section*{ProbDist}
\(\begin{array}{ll}\text { See Also } \quad & \text { fitdist | ProbDistParametric | ProbDistKernel } \\ & \text { | ProbDistUnivParam | ProbDistUnivKernel } \\ & \text { | ProbDistUnivParam. ProbDistUnivParam | } \\ & \text { ProbDistUnivKernel.ProbDistUnivKernel }\end{array}\)

\section*{ProbDistKernel}

\section*{Superclasses ProbDist}

Purpose Object representing nonparametric probability distribution defined by kernel smoothing

Description

\section*{Construction}

\section*{Methods}
cdf
pdf
random

Return cumulative distribution function (CDF) for ProbDist object

Return probability density function (PDF) for ProbDist object

Generate random number drawn from ProbDist object

Note The above methods are inherited from the ProbDist class.

\section*{Properties}

BandWidth

DistName

Read-only value specifying bandwidth of kernel smoothing function for ProbDistKernel object

Read-only string containing probability distribution name of ProbDist object

\section*{ProbDistKernel}
\begin{tabular}{ll} 
InputData & \begin{tabular}{l} 
Read-only structure containing \\
information about input data to \\
ProbDist object
\end{tabular} \\
Kernel & \begin{tabular}{l} 
Read-only string specifying name \\
of kernel smoothing function for \\
ProbDistKernel object
\end{tabular} \\
Support & \begin{tabular}{l} 
Read-only structure containing \\
information about support of
\end{tabular} \\
& ProbDist object
\end{tabular}

Note Some of the above properties are inherited from the ProbDist class.

Copy
Semantics
See Also

Value. To learn how this affects your use of the class, see Copying Objects in the MATLAB Programming Fundamentals documentation.
fitdist | ProbDist | ProbDistUnivKernel | ProbDistUnivKernel. ProbDistUnivKernel

\section*{ProbDistParametric}

\section*{Superclasses ProbDist}

Purpose Object representing parametric probability distribution
Description ProbDistParametric is an abstract class defining the properties and methods of a parametric probability distribution.

\section*{Construction ProbDistParametric is an abstract class. You cannot create instances of this class directly. You can construct an object in its subclass, ProbDistUnivParam, either by calling the subclass constructor, ProbDistUnivParam. ProbDistUnivParam, or by using the fitdist function.}

\section*{Methods}
\begin{tabular}{ll} 
cdf & \begin{tabular}{l} 
Return cumulative distribution \\
function (CDF) for ProbDist object
\end{tabular} \\
pdf & \begin{tabular}{l} 
Return probability density \\
function (PDF) for ProbDist object
\end{tabular} \\
random & \begin{tabular}{l} 
Generate random number drawn \\
from ProbDist object
\end{tabular}
\end{tabular}

Note The above methods are inherited from the ProbDist class.

\section*{Properties}

DistName

InputData

Read-only string containing probability distribution name of ProbDist object

Read-only structure containing information about input data to ProbDist object

\section*{ProbDistParametric}
\begin{tabular}{ll} 
NLogL & \begin{tabular}{l} 
Read-only value specifying \\
negative log likelihood for input \\
data to ProbDistParametric \\
object
\end{tabular} \\
NumParams & \begin{tabular}{l} 
Read-only value specifying \\
number of parameters of \\
ProbDistParametric object
\end{tabular} \\
ParamCov & \begin{tabular}{l} 
Read-only covariance matrix \\
of parameter estimates of \\
ProbDistParametric object
\end{tabular} \\
ParamDescription & \begin{tabular}{l} 
Read-only cell array specifying \\
descriptions of parameters of \\
ProbDistParametric object
\end{tabular} \\
ParamIsFixed & \begin{tabular}{l} 
Read-only logical array \\
specifying fixed parameters \\
of ProbDistParametric object
\end{tabular} \\
ParamNames & \begin{tabular}{l} 
Read-only cell array specifying \\
names of parameters of
\end{tabular} \\
ProbDistParametric object
\end{tabular}

Note Some of the above properties are inherited from the ProbDist class.

Value. To learn how this affects your use of the class, see Copying Objects in the MATLAB Programming Fundamentals documentation.

See Also
fitdist | ProbDist | ProbDistUnivParam | ProbDistUnivParam.ProbDistUnivParam

\section*{ProbDistUnivKernel}

Superclasses ProbDistKernel
Purpose Object representing univariate kernel probability distribution
Description A ProbDistUnivKernel object represents a univariate nonparametric probability distribution defined by kernel smoothing. You create this object using the fitdist function to fit the distribution to data.

\section*{Construction}

\section*{fitdist}
cdf
icdf
iqr
median
pdf
random

Fit probability distribution to data

\section*{Methods}

Return cumulative distribution function (CDF) for ProbDist object

Return inverse cumulative distribution function (ICDF) for ProbDistUnivKernel object
Return interquartile range (IQR) for ProbDistUnivKernel object
Return median of ProbDistUnivKernel object
Return probability density function (PDF) for ProbDist object
Generate random number drawn from ProbDist object

Note Some of the above methods are inherited from the ProbDistKernel class.

\section*{ProbDistUnivKernel}

\section*{Properties}

BandWidth
\begin{tabular}{ll} 
DistName & \begin{tabular}{l} 
Read-only string containing \\
probability distribution name of \\
ProbDist object
\end{tabular} \\
InputData & \begin{tabular}{l} 
Read-only structure containing \\
information about input data to \\
ProbDist object
\end{tabular} \\
Kernel & \begin{tabular}{l} 
Read-only string specifying name \\
of kernel smoothing function for \\
ProbDistKernel object
\end{tabular} \\
NLogL & \begin{tabular}{l} 
Read-only value specifying \\
negative log likelihood for input \\
data to ProbDistUnivKernel \\
object
\end{tabular} \\
Support & \begin{tabular}{l} 
Read-only structure containing \\
information about support of \\
ProbDist object
\end{tabular}
\end{tabular}

Note Some of the above properties are inherited from the ProbDistKernel class.
\begin{tabular}{ll} 
Copy & Value. To learn how this affects your use of the class, see Copying \\
Semantics & \begin{tabular}{l} 
Objects in the MATLAB Programming Fundamentals documentation.
\end{tabular} \\
References & \begin{tabular}{l} 
[1] Bowman, A. W., and A. Azzalini. Applied Smoothing Techniques for \\
\\
\\
Data Analysis. New York: Oxford University Press, 1997.
\end{tabular}
\end{tabular}

\section*{ProbDistUnivKernel}

\author{
See Also \\ fitdist | ksdensity | ProbDist | ProbDistKernel | ProbDistUnivKernel. ProbDistUnivKernel
}

\section*{ProbDistUnivKernel}

\section*{Purpose Construct ProbDistUnivKernel object}

Syntax \(\quad \begin{aligned} P D & =\operatorname{ProbDistUnivKernel}(X) \\ P D & =\operatorname{ProbDistUnivKernel}(X,\end{aligned}\)
\(P D=\) ProbDistUnivKernel(X, param1, val1, param2, val2, ...)

\section*{Description}

Tip Although you can use this constructor function to create a ProbDistUnivKernel object, using the fitdist function is an easier way to create the ProbDistUnivKernel object.
\(P D=\) ProbDistUnivKernel \((X)\) creates \(P D\), a ProbDistUnivKernel object, which represents a nonparametric probability distribution, based on a normal kernel smoothing function.
\(P D=\) ProbDistUnivKernel(X, param1, val1, param2, val2, ...) specifies optional parameter name/value pairs, as described in the Parameter/Values table. Parameter and value names are case insensitive.

\section*{Input \\ \(x\) \\ Arguments}

A column vector of data.

Note Any NaN values in \(X\) are ignored by the fitting calculations.

\section*{ProbDistUnivKernel}
\begin{tabular}{l|l}
\hline Parameter & Values \\
\hline 'censoring & \begin{tabular}{l} 
A Boolean vector the same size as \(X\), containing 1s when \\
the corresponding elements in \(X\) are right-censored \\
observations and 0s when the corresponding elements \\
are exact observations. Default is a vector of 0s.
\end{tabular} \\
\hline \begin{tabular}{l} 
Note Any NaN values in this censoring vector are \\
ignored by the fitting calculations.
\end{tabular} \\
\hline frequency & \begin{tabular}{l} 
A vector the same size as \(X\), containing nonnegative \\
integers specifying the frequencies for the corresponding \\
elements in \(X\). Default is a vector of 1s.
\end{tabular} \\
\hline \begin{tabular}{l} 
Note Any NaN values in this frequency vector are \\
ignored by the fitting calculations.
\end{tabular} \\
\hline 'kernel' & \begin{tabular}{l} 
A string specifying the type of kernel smoother to use. \\
Choices are: \\
- 'normal' (default) \\
- 'box' \\
- 'triangle ' \\
- 'epanechnikov'
\end{tabular} \\
\hline
\end{tabular}

\section*{ProbDistUnivKernel}
\begin{tabular}{l|l}
\hline Parameter & Values \\
\hline 'support' & \begin{tabular}{l} 
Any of the following to specify the support: \\
- 'unbounded ' — Default. If the density can extend \\
over the whole real line.
\end{tabular} \\
- 'positive' — To restrict it to positive values. \\
- A two-element vector giving finite lower and upper \\
limits for the support of the density.
\end{tabular}

\section*{Output \\ Arguments \\ PD}

\section*{References}
[1] Bowman, A. W., and A. Azzalini. Applied Smoothing Techniques for Data Analysis. New York: Oxford University Press, 1997.

See Also fitdist \| ksdensity

\section*{ProbDistUnivParam}

Superclasses ProbDistParametric
Purpose Object representing univariate parametric probability distribution
Description A ProbDistUnivParam object represents a univariate parametric probability distribution. You create this object by using the constructor (ProbDistUnivParam. ProbDistUnivParam) and supplying parameter values, or by using the fitdist function to fit the distribution to data.

\author{
Construction \\ fitdist \\ ProbDistUnivParam
}

\section*{Methods}
cdf
icdf
iqr
mean
median
paramci
pdf

Fit probability distribution to data

Construct ProbDistUnivParam object

Return cumulative distribution function (CDF) for ProbDist object

Return inverse cumulative distribution function (ICDF) for ProbDistUnivParam object
Return interquartile range (IQR) for ProbDistUnivParam object
Return mean of ProbDistUnivParam object
Return median of ProbDistUnivParam object

Return parameter confidence intervals of ProbDistUnivParam object
Return probability density function (PDF) for ProbDist object
\begin{tabular}{ll} 
random & \begin{tabular}{l} 
Generate random number drawn \\
from ProbDist object
\end{tabular} \\
std & \begin{tabular}{l} 
Return standard deviation of \\
ProbDistUnivParam object
\end{tabular} \\
var & \begin{tabular}{l} 
Return variance of \\
ProbDistUnivParam object
\end{tabular}
\end{tabular}

Note Some of the above methods are inherited from the ProbDistParametric class.
\begin{tabular}{|c|c|c|}
\hline Properties & DistName & Read-only string containing probability distribution name of ProbDist object \\
\hline & InputData & Read-only structure containing information about input data to ProbDist object \\
\hline & NLogL & Read-only value specifying negative log likelihood for input data to ProbDistParametric object \\
\hline & NumParams & Read-only value specifying number of parameters of ProbDistParametric object \\
\hline & ParamCov & Read-only covariance matrix of parameter estimates of ProbDistParametric object \\
\hline & ParamDescription & Read-only cell array specifying descriptions of parameters of ProbDistParametric object \\
\hline
\end{tabular}

\section*{ProbDistUnivParam}
\begin{tabular}{|c|c|c|}
\hline & ParamIsFixed & Read-only logical array specifying fixed parameters of ProbDistParametric object \\
\hline & ParamNames & Read-only cell array specifying names of parameters of ProbDistParametric object \\
\hline & Params & Read-only array specifying values of parameters of ProbDistParametric object \\
\hline & Support & Read-only structure containing information about support of ProbDist object \\
\hline & Note The above class. & erited from the ProbDistParametric \\
\hline & Note Parameter create PD, a uni that represents properties that & roperties. For example, if you probability distribution object on, then PD.mu and PD.sigma are e mu and sigma parameters. \\
\hline Copy Semantics & Value. To learn Objects in the M & ur use of the class, see Copying ing Fundamentals documentation. \\
\hline References & [1] Johnson, N. Univariate Distr 1993. & Balakrishnan. Continuous oboken, NJ: Wiley-Interscience, \\
\hline
\end{tabular}
[2] Johnson, N. L., S. Kotz, and N. Balakrishnan. Continuous Univariate Distributions. Vol. 2, Hoboken, NJ: Wiley-Interscience, 1994.

\section*{See Also fitdist | ProbDist | ProbDistParametric | ProbDistUnivParam.ProbDistUnivParam}

How To - Appendix B, "Distribution Reference"

\section*{ProbDistUnivParam}

\section*{Purpose Construct ProbDistUnivParam object}
```

Syntax $\quad P D=$ ProbDistUnivParam(DistName, Params)

```

Description \(\quad P D=\) ProbDistUnivParam(DistName, Params) creates PD, a ProbDistUnivParam object, which represents a probability distribution. This distribution is defined by the parametric distribution specified by DistName, with parameters specified by the numeric vector Params.

\section*{Input DistName \\ Arguments}

A string specifying a distribution. Choices are:
- 'beta'
- 'binomial'
- 'birnbaumsaunders'
- 'exponential'
- 'extreme value' or ev'
- 'gamma'
- 'generalized extreme value' or 'gev'
- 'generalized pareto' or 'gp'
- 'inversegaussian'
- 'logistic'
- 'loglogistic'
- 'lognormal'
- 'nakagami'
- 'negative binomial' or 'nbin'
- 'normal'
- 'poisson'


For more information on these parametric distributions, see Appendix B, "Distribution Reference".

Params
Numeric vector of distribution parameters. The number and type of parameters depends on the distribution you specify with DistName. For information on parameters for each distribution type, see Appendix B, "Distribution Reference".

\section*{Output \\ \(P D\) \\ Arguments}

\section*{Examples}

An object in the ProbDistUnivParam class, which is derived from the ProbDist class. It represents a parametric probability distribution.

1 Create an object representing a normal distribution with a mean of 100 and a standard deviation of 10 .
pd = ProbDistUnivParam('normal',[100 10])
pd \(=\)
normal distribution
\[
\mathrm{mu}=100
\]
sigma \(=10\)

\section*{ProbDistUnivParam}

2 Generate a 4-by- 5 matrix of random values from this distribution.


\author{
References [1] Johnson, N. L., S. Kotz, and N. Balakrishnan. Continuous Univariate Distributions. Vol. 1, Hoboken, NJ: Wiley-Interscience, 1993.
}
[2] Johnson, N. L., S. Kotz, and N. Balakrishnan. Continuous Univariate Distributions. Vol. 2, Hoboken, NJ: Wiley-Interscience, 1994.

See Also fitdist
How To • Appendix B, "Distribution Reference"
\begin{tabular}{|c|c|}
\hline Purpose & Probability plots \\
\hline \multirow[t]{8}{*}{Syntax} & probplot(Y) \\
\hline & probplot (distribution, Y ) \\
\hline & probplot(Y, cens,freq) \\
\hline & probplot(ax, Y) \\
\hline & probplot(...,'noref') \\
\hline & probplot(ax, PD) \\
\hline & probplot(ax,fun, params) \\
\hline & \(\mathrm{h}=\operatorname{probplot}(\ldots)\) \\
\hline
\end{tabular}

\section*{Description}
probplot \((\mathrm{Y})\) produces a normal probability plot comparing the distribution of the data \(Y\) to the normal distribution. \(Y\) can be a single vector, or a matrix with a separate sample in each column. The plot includes a reference line useful for judging whether the data follow a normal distribution.
probplot uses midpoint probability plotting positions. The \(i^{\text {th }}\) sorted value from a sample of size \(N\) is plotted against the midpoint in the jump of the empirical CDF on the \(y\) axis. With uncensored data, that midpoint is \((i-0.5) / N\). With censored data (see below), the \(y\) value is more complicated to compute.
probplot(distribution, Y ) creates a probability plot for the distribution specified by distribution. Acceptable strings for distribution are:
- 'exponential' - Exponential probability plot (nonnegative values)
- 'extreme value' - Extreme value probability plot (all values)
- 'lognormal' - Lognormal probability plot (positive values)
- 'normal' - Normal probability plot (all values)
- 'rayleigh' - Rayleigh probability plot (positive values)
- 'weibull' - Weibull probability plot (positive values)

\section*{probplot}

The \(y\) axis scale is based on the selected distribution. The \(x\) axis has a log scale for the Weibull and lognormal distributions, and a linear scale for the others.

Not all distributions are appropriate for all data sets, and probplot will error when asked to create a plot with a data set that is inappropriate for a specified distribution. Appropriate data ranges for each distribution are given parenthetically in the list above.
probplot(Y,cens,freq) or probplot(distname, Y , cens,freq) requires a vector \(Y\). cens is a vector of the same size as \(Y\) and contains 1 for observations that are right-censored and 0 for observations that are observed exactly. freq is a vector of the same size as Y , containing integer frequencies for the corresponding elements in \(Y\).
probplot (ax, Y ) takes a handle ax to an existing probability plot, and adds additional lines for the samples in Y . ax is a handle for a set of axes.
probplot(...,'noref') omits the reference line.
probplot (ax, PD) takes a probability distribution object, PD, and adds a fitted line to the axes specified by ax to represent the probability distribution specified by PD. PD is a ProbDist object of the ProbDistUnivParam class or ProbDistUnivKernel class.
probplot (ax, fun, params) takes a function fun and a set of parameters, params, and adds fitted lines to the axes of an existing probability plot specified by ax. fun is a function handle to a cdf function, specified with @ (for example, @wblcdf). params is the set of parameters required to evaluate fun, and is specified as a cell array or vector. The function must accept a vector of \(X\) values as its first argument, then the optional parameters, and must return a vector of cdf values evaluated at \(X\).
\(h=\) probplot(...) returns handles to the plotted lines.

\section*{Examples}

\section*{Example 1}

The following plot assesses two samples, one from a Weibull distribution and one from a Rayleigh distribution, to see if they may have come from a Weibull population.
```

x1 = wblrnd(3,3,100,1);
x2 = raylrnd(3,100,1);
probplot('weibull',[x1 x2])
legend('Weibull Sample','Rayleigh Sample','Location','NW')

```

Probability plot for Weibull distribution


\section*{Example 2}

Consider the following data, with about \(20 \%\) outliers:
```

left_tail = - exprnd(1,10,1);
right_tail = exprnd(5,10,1);
center = randn(80,1);

```
```

data = [left_tail;center;right_tail];

```

Neither a normal distribution nor a \(t\) distribution fits the tails very well:
```

probplot(data);
p = mle(data,'dist','tlo');
t = @(data,mu,sig,df)cdf('tlocationscale',data,mu,sig,df);
h = probplot(gca,t,p);
set(h,'color','r','linestyle',' -')
title('{\bf Probability Plot}')
legend('Data','Normal','t','Location','NW')

```

Probability Plot


See Also normplot | ecdf | wblplot

\section*{Purpose Procrustes analysis}

\section*{Syntax \(\quad d=\operatorname{procrustes}(X, Y)\)}
[d,Z] = procrustes(X,Y)
[d,Z,transform] = procrustes( \(\mathrm{X}, \mathrm{Y}\) )
[...] = procrustes(...,'scaling',flag)
[...] = procrustes(...,'reflection',flag)

\section*{Description}
\(d=p r o c r u s t e s(X, Y)\) determines a linear transformation (translation, reflection, orthogonal rotation, and scaling) of the points in matrix \(Y\) to best conform them to the points in matrix \(X\). The goodness-of-fit criterion is the sum of squared errors. procrustes returns the minimized value of this dissimilarity measure in \(d\). \(d\) is standardized by a measure of the scale of \(X\), given by:
```

sum(sum((X-repmat(mean(X,1),size(X,1),1)).^2,1))

```

That is, the sum of squared elements of a centered version of \(X\). However, if \(X\) comprises repetitions of the same point, the sum of squared errors is not standardized.
\(X\) and \(Y\) must have the same number of points (rows), and procrustes matches \(Y(i)\) to \(X(i)\). Points in \(Y\) can have smaller dimension (number of columns) than those in \(X\). In this case, procrustes adds columns of zeros to Y as necessary.
\([d, Z]=\) procrustes \((X, Y)\) also returns the transformed \(Y\) values.
[d, \(Z\), transform] = procrustes \((X, Y)\) also returns the transformation that maps \(Y\) to \(Z\). transform is a structure array with fields:
- c - Translation component
- T-Orthogonal rotation and reflection component
- b-Scale component

That is:
c = transform.c;
```

T = transform.T;
b = transform.b;
Z = b*Y*T + c;

```
[...] = procrustes(...,'scaling',flag), when flag is false,
allows you to compute the transformation without a scale component (that is, with b equal to 1). The default flag is true.
[...] = procrustes(...,'reflection',flag), when flag is false, allows you to compute the transformation without a reflection component (that is, with \(\operatorname{det}(T)\) equal to 1 ). The default flag is 'best', which computes the best-fitting transformation, whether or not it includes a reflection component. A flag of true forces the transformation to be computed with a reflection component (that is, with \(\operatorname{det}(T)\) equal to -1 )

\section*{Examples}

This example creates some random points in two dimensions, then rotates, scales, translates, and adds some noise to those points. It uses procrustes to conform \(Y\) to \(X\), then plots the original \(X\) and \(Y\) with the transformed \(Y\).
```

n = 10;
X = normrnd(0,1,[n 2]);
S = [0.5 -sqrt(3)/2; sqrt(3)/2 0.5];
Y = normrnd(0.5*X*S+2,0.05,n,2);
[d,Z,tr] = procrustes(X,Y);
plot(X(:,1),X(:,2),'rx',...
Y(:,1),Y(:,2),'b.',...
Z(:,1),Z(:,2),'bx');

```


\section*{References}

See Also
[1] Kendall, David G. "A Survey of the Statistical Theory of Shape." Statistical Science. Vol. 4, No. 2, 1989, pp. 87-99.
[2] Bookstein, Fred L. Morphometric Tools for Landmark Data. Cambridge, UK: Cambridge University Press, 1991.
[3] Seber, G. A. F. Multivariate Observations. Hoboken, NJ: John Wiley \& Sons, Inc., 1984.

\author{
cmdscale | factoran
}

\section*{CompactTreeBagger.proximity}

\section*{Purpose Proximity matrix for data}

\section*{Syntax \\ prox = proximity (B, X)}

Description
prox \(=\) proximity \((B, X)\) computes a numeric matrix of size Nobs-by-Nobs of proximities for data \(X\), where Nobs is the number of observations (rows) in X. Proximity between any two observations in the input data is defined as a fraction of trees in the ensemble \(B\) for which these two observations land on the same leaf. This is a symmetric matrix with ones on the diagonal and off-diagonal elements ranging from 0 to 1 .

\section*{TreeBagger.Proximity property}

Purpose Proximity matrix for observations
Description
The Proximity property is a numeric matrix of size Nobs-by-Nobs, where Nobs is the number of observations in the training data, containing measures of the proximity between observations. For any two observations, their proximity is defined as the fraction of trees for which these observations land on the same leaf. This is a symmetric matrix with 1 s on the diagonal and off-diagonal elements ranging from 0 to 1 .

See Also CompactTreeBagger.proximity | classregtree.varimportance

\section*{Purpose}

Produce sequence of subtrees by pruning
Syntax

Description

Tips

Input
Arguments
tree1 = prune(tree)
tree1 \(=\) prune(tree, Name, Value) with its optimal pruning sequence filled in.
tree1 \(=\) prune(tree, Name, Value) creates a pruned tree with can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN. sequence first.

\section*{tree}
tree1 = prune(tree) creates a copy of the classification tree tree additional options specified by one Name, Value pair argument. You
- tree1 \(=\) prune (tree) returns the decision tree tree1 that is the full, unpruned tree, but with optimal pruning information added. This is useful only if you created tree by pruning another tree, or by using the ClassificationTree.fit method with pruning set 'off'. If you plan to prune a tree multiple times along the optimal pruning sequence, it is more efficient to create the optimal pruning

A classification tree created with ClassificationTree.fit.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
alpha
A numeric scalar. prune prunes tree to the specified value of the pruning cost.
level

\section*{ClassificationTree.prune}
eric scalar from 0 (no pruning) to the largest pruning level of this tree max (tree. PruneList). prune returns the tree pruned to this level.

\section*{nodes}

A numeric vector with elements from 1 to tree. NumNodes. Any tree branch nodes listed in nodes become leaf nodes in tree1, unless their parent nodes are also pruned.

\section*{Output Arguments \\ tree 1 \\ A classification tree.}

Examples Display a full tree for Fisher's iris data, as well as the next largest tree from the optimal pruning sequence:
```

load fisheriris;
varnames = {'SL' 'SW' 'PL' 'PW'};
t1 = ClassificationTree.fit(meas,species,...
'minparent',5,'predictornames',varnames);
view(t1,'mode','graph');

```

\section*{ClassificationTree.prune}

\[
\begin{aligned}
& \text { t2 }=\text { prune(t1,'level', } 1) \text {; } \\
& \text { view(t2,'mode','graph'); }
\end{aligned}
\]

\section*{ClassificationTree.prune}

See Also
ClassificationTree.fit
How To
- Chapter 13, "Nonparametric Supervised Learning"

\section*{Purpose Prune tree}
```

Syntax
t2 = prune(t1,'level',level)
t2 = prune(t1,'nodes',nodes)
t2 = prune(t1)

```

\section*{Description}

\section*{Examples}

Display the full tree for Fisher's iris data:
```

load fisheriris;
t1 = classregtree(meas,species,...
names',{'SL' 'SW' 'PL' 'PW'},...
splitmin',5)
t1 =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa

```

\section*{classregtree.prune}
```

3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
class = versicolor
9 class = virginica
view(t1)

```
\begin{tabular}{|c|c|c|c|c|c|c|c|}
\hline Click to display: Identity & \(\checkmark\) & Magnification: & 100\% & \(\checkmark\) & Pruning level: & 0 of 4 & - \\
\hline
\end{tabular}


Display the next largest tree from the optimal pruning sequence:
```

t2 = prune(t1,'level',1)
t2 =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 class = versicolor
7 class = virginica
view(t2)

```

\section*{classregtree.prune}
\begin{tabular}{|c|c|c|c|c|c|c|c|}
\hline Click to display: Identity & \(\checkmark\) & Magnification: & 100\% & \(\checkmark\) & Pruning level: & 0 of 3 & - \\
\hline
\end{tabular}


References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

See Also classregtree | test | view

\section*{Purpose}

Produce sequence of subtrees by pruning
Syntax

Description

Tips

Input
Arguments
tree1 = prune(tree)
tree1 \(=\) prune(tree, Name, Value) its optimal pruning sequence filled in.
tree1 = prune(tree, Name, Value) creates a pruned tree with can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN. is more efficient to create the optimal pruning sequence first.

\section*{tree}
tree1 \(=\) prune(tree) creates a copy of the regression tree tree with additional options specified by one Name, Value pair argument. You
- tree1 = prune(tree) returns the decision tree tree1 that is the full, unpruned tree, but with optimal pruning information added. This is useful only if you created tree by pruning another tree, or by using the RegressionTree.fit with pruning set 'off'. If you plan to prune a tree multiple times along the optimal pruning sequence, it

A regression tree created with RegressionTree.fit.

\section*{Name-Value Pair Arguments}

Optional comma-separated pair of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify only one name-value pair argument.

\section*{alpha}

A numeric scalar from 0 (no pruning) to 1 (prune to one node). Prunes to minimize the sum of (alpha times the number of leaf nodes) and a cost (mean squared error).
level

\section*{RegressionTree.prune} of this tree max (tree. PruneList). prune returns the tree pruned to this level.
```

nodes

```

A numeric vector with elements from 1 to tree. NumNodes. Any tree branch nodes listed in nodes become leaf nodes in tree1, unless their parent nodes are also pruned.

\section*{Output Arguments}

Examples
tree 1
A regression tree.
Display a full tree for the carsmall data, as well as the tree pruned to level 10:
```

load carsmall;

```
load carsmall;
varnames = {'Weight' 'Horsepower'};
varnames = {'Weight' 'Horsepower'};
t1 = RegressionTree.fit([Weight Horsepower],MPG,...
t1 = RegressionTree.fit([Weight Horsepower],MPG,...
    'predictornames',varnames)
    'predictornames',varnames)
view(t1,'mode','graph');
```

view(t1,'mode','graph');

```

t2 = prune(t1,'level',10);

\section*{RegressionTree.prune}
view(t2,'mode','graph');

See Also
RegressionTree.fit
How To
- Chapter 13, "Nonparametric Supervised Learning"

\section*{TreeBagger.Prune property}

Purpose Flag to prune trees
Description \(\quad \begin{aligned} & \text { The Prune property is true if decision trees are pruned and false if they } \\ & \text { are not. Pruning decision trees is not recommended for ensembles. The } \\ & \text { default value is false. }\end{aligned}\)
See Also classregtree.prune

Purpose Pruning levels for decision tree nodes
Syntax \(\quad P=\operatorname{prunelist}(T)\)
P = prunelist( \(\mathrm{T}, \mathrm{J}\) )
Description \(\quad P=\) prunelist \((T)\) returns an \(n\)-element numeric vector with the pruning levels in each node of the tree \(T\), where \(n\) is the number of nodes. When you call prune(T,'level',level), nodes with the pruning levels below level are pruned, and nodes with the pruning levels greater or equal to level are not pruned.
\(P=p r u n e l i s t(T, J)\) takes an array \(J\) of node numbers and returns the pruning levels for the specified nodes.

\author{
See Also
}
classregtree | numnodes

\section*{qrandstream.qrand}

\section*{Purpose Generate quasi-random points from stream}
Syntax
\(x=q r a n d(q)\)
\(x=\operatorname{arand}(q, n)\)

Description

\section*{Examples}
\(x=q r a n d(q)\) returns the next value \(x\) in the quasi-random number stream \(q\) of the qrandstream class. x is a 1-by- \(d\) vector, where \(d\) is the dimension of the stream. The command sets \(q\). State to the index in the underlying point set of the next value to be returned.
\(\mathrm{X}=\mathrm{qrand}(\mathrm{q}, \mathrm{n})\) returns the next n values X in an n -by- \(d\) matrix.
Objects q of the qrandstream class encapsulate properties of a specified quasi-random number stream. Values of the stream are not generated and stored in memory until \(q\) is accessed using qrand.

Use qrandstream to construct a 3-D Halton stream, based on a point set that skips the first 1000 values and then retains every 101st point:
```

q = qrandstream('halton',3,'Skip',1e3,'Leap',1e2)
q =
Halton quasi-random stream in 3 dimensions
Point set properties:
Skip : 1000
Leap : 100
ScrambleMethod : none
nextIdx = q.State
nextIdx =
1

```

Use qrand to generate two samples of size four:
```

X1 = qrand(q,4)
X1 =

| 0.0928 | 0.3475 | 0.0051 |
| :--- | :--- | :--- |
| 0.6958 | 0.2035 | 0.2371 |
| 0.3013 | 0.8496 | 0.4307 |

```
```

    0.9087 0.5629 0.6166
    nextIdx = q.State
nextIdx =
5
X2 = qrand(q,4)
X2 =
0.2446 0.0238 0.8102
0.5298 0.7540 0.0438
0.3843 0.5112 0.2758
0.8335 0.2245 0.4694
nextIdx = q.State
nextIdx =
9

```

Use reset to reset the stream, then generate another sample:
```

reset(q)
nextIdx = q.State
nextIdx =
1
X = qrand(q,4)
X =

| 0.0928 | 0.3475 | 0.0051 |
| :--- | :--- | :--- |
| 0.6958 | 0.2035 | 0.2371 |
| 0.3013 | 0.8496 | 0.4307 |
| 0.9087 | 0.5629 | 0.6166 |

```

See Also qrandstream | reset

\section*{qrandset}

\section*{Purpose Quasi-random point sets}

Description
qrandset is a base class that encapsulates a sequence of multidimensional quasi-random numbers. This base class is abstract and cannot be instantiated directly. Concrete subclasses include sobolset and haltonset.

\section*{Construction \\ qrandset}
disp
end
length
ndims
net
scramble
size
subsref

\section*{Properties}

Dimensions
Leap
ScrambleMethod

Abstract quasi-random point set class

Display qrandset object
Last index in indexing expression for point set

Length of point set
Number of dimensions in matrix
Generate quasi-random point set
Scramble quasi-random point set
Number of dimensions in matrix
Subscripted reference for qrandset

Number of dimensions
Interval between points
Settings that control scrambling
\begin{tabular}{ll} 
Skip & \begin{tabular}{l} 
Number of initial points to omit \\
from sequence
\end{tabular} \\
Type & \begin{tabular}{l} 
Name of sequence on which point \\
set \(P\) is based
\end{tabular}
\end{tabular}

Copy
Semantics

See Also haltonset | sobolset
How To Programming documentation.

Value. To learn how this affects your use of the class, see Comparing Handle and Value Classes in the MATLAB Object-Oriented
- "Quasi-Random Point Sets" on page 6-16

\section*{qrandset}

\section*{Purpose Abstract quasi-random point set class}
\(\begin{array}{ll}\text { Description } & \begin{array}{l}\text { qrandset is an abstract class, and you cannot create instances of it } \\ \text { directly. You must use haltonset or sobolset to create a qrandset } \\ \text { object. }\end{array}\end{array}\)
See Also haltonset | sobolset

\section*{Purpose \\ Quasi-random number streams}

\section*{Construction qrandstream}

Construct quasi-random number stream

\section*{Methods}
\begin{tabular}{|c|c|}
\hline addlistener & Add listener for event \\
\hline delete & Delete handle object \\
\hline disp & Display qrandstream object \\
\hline eq & Test handle equality \\
\hline findobj & Find objects matching specified conditions \\
\hline findprop & Find property of MATLAB handle object \\
\hline ge & Greater than or equal relation for handles \\
\hline gt & Greater than relation for handles \\
\hline isvalid & Test handle validity \\
\hline le & Less than or equal relation for handles \\
\hline lt & Less than relation for handles \\
\hline ne & Not equal relation for handles \\
\hline notify & Notify listeners of event \\
\hline qrand & Generate quasi-random points from stream \\
\hline rand & Generate quasi-random points from stream \\
\hline reset & Reset state \\
\hline
\end{tabular}

\section*{qrandstream}

\author{
Properties PointSet \\ State \\ Point set from which stream is drawn \\ Current state of the stream \\ \footnotetext{
Copy \\ Handle. To learn how this affects your use of the class, see \\ Semantics \\ Comparing Handle and Value Classes in the MATLAB Object-Oriented Programming documentation.
}
}

\section*{Purpose Construct quasi-random number stream}
```

Syntax $\quad q=$ qrandstream (type, $d$ )
q = qrandstream(type,d,prop1,val1,prop2,val2,...)
$q$ = qrandstream(p)

```

\section*{Description}

\section*{Examples}
\(\mathrm{q}=\mathrm{qrandstream}(\) type, d\()\) constructs a d-dimensional quasi-random number stream \(q\) of the qrandstream class, of type specified by the string type. type is either 'halton' or 'sobol', and q is based on a point set from either the haltonset class or sobolset class, respectively, with default property settings.
q = qrandstream(type,d,prop1,val1,prop2,val2,...) specifies property name/value pairs for the point set on which the stream is based. Applicable properties depend on type.
\(q=q r a n d s t r e a m(p)\) constructs a stream based on the specified point set \(p\). \(p\) must be a point set from either the haltonset class or sobolset class.

Construct a 3-D Halton stream, based on a point set that skips the first 1000 values and then retains every 101st point:
```

q = qrandstream('halton',3,'Skip',1e3,'Leap',1e2)
q =
Halton quasi-random stream in 3 dimensions
Point set properties:
Skip : 1000
Leap : 100
ScrambleMethod : none
nextIdx = q.State
nextIdx =
1

```

Use qrand to generate two samples of size four:
\[
\mathrm{X} 1=\operatorname{qrand}(\mathrm{q}, 4)
\]

\section*{qrandstream}
```

X1 =
0.0928 0.3475 0.0051
0.6958 0.2035 0.2371
0.3013 0.8496 0.4307
0.9087 0.5629 0.6166
nextIdx = q.State
nextIdx =
5
X2 = qrand(q,4)
x2 =
0.2446 0.0238 0.8102
0.5298 0.7540 0.0438
0.3843 0.5112 0.2758
0.8335 0.2245 0.4694
nextIdx = q.State
nextIdx =
9

```

Use reset to reset the stream, and then generate another sample:
```

reset(q)
nextIdx = q.State
nextIdx =
1
X = qrand(q,4)
X =

| 0.0928 | 0.3475 | 0.0051 |
| :--- | :--- | :--- |
| 0.6958 | 0.2035 | 0.2371 |
| 0.3013 | 0.8496 | 0.4307 |
| 0.9087 | 0.5629 | 0.6166 |

```

See Also
haltonset | qrand | reset | sobolset

\section*{Purpose}

Quantile-quantile plot
Syntax
```

qqplot(X)
qqplot(X,Y)
qqplot(X,PD)
qqplot(X,Y,pvec)
h = qqplot(X,Y,pvec)

```

\section*{Description}
qqplot ( \(X\) ) displays a quantile-quantile plot of the sample quantiles of \(X\) versus theoretical quantiles from a normal distribution. If the distribution of \(X\) is normal, the plot will be close to linear.
qqplot ( \(\mathrm{X}, \mathrm{Y}\) ) displays a quantile-quantile plot of two samples. If the samples do come from the same distribution, the plot will be linear.
qqplot ( \(X, P D\) ) makes an empirical quantile-quantile plot of the quantiles of the data in the vector \(X\) versus the quantiles of the distribution specified by PD, a ProbDist object of the ProbDistUnivParam class or ProbDistUnivKernel class.

For matrix \(X\) and \(Y\), qqplot displays a separate line for each pair of columns. The plotted quantiles are the quantiles of the smaller data set.

The plot has the sample data displayed with the plot symbol ' + '. Superimposed on the plot is a line joining the first and third quartiles of each distribution (this is a robust linear fit of the order statistics of the two samples). This line is extrapolated out to the ends of the sample to help evaluate the linearity of the data.

Use qqplot ( \(\mathrm{X}, \mathrm{Y}, \mathrm{pvec}\) ) to specify the quantiles in the vector pvec.
\(h=q q p l o t(X, Y, p v e c)\) returns handles to the lines in \(h\).

\section*{Examples}

The following example shows a quantile-quantile plot of two samples from Poisson distributions.
```

x = poissrnd(10,50,1);
y = poissrnd(5,100,1);
qqplot(x,y);

```


\footnotetext{
How To
- normplot
}

\section*{Purpose}

Quantiles

\section*{Syntax}
\(Y=\) quantile \((X, p)\)
\(Y=\) quantile \((X, p, d i m)\)
\(Y=\) quantile \((X, N, \ldots)\)

\section*{Examples}
```

y = quantile(x,.50); % the median of x
y = quantile(x,[.025 .25 .50 .75 .975]); % Summary of x
y = quantile(x,[.25 . 50 . 75]); % the quartiles of x
y = quantile(x,3); % another way to get quartiles

```

\section*{quantile}

See Also prctile | iqr | median
Purpose Generate quasi-random points from stream
Syntax rand
rand (q, n)
rand(q)
rand(q,m,n)
rand ( \(q,[m, n]\) )
rand \((q, m, n, p, \ldots)\)
\(\operatorname{rand}(q,[m, n, p, \ldots])\)

\section*{Description}

\section*{Examples}
Generate the first 256 points from a 5-D Sobol sequence:
```

q = qrandstream('sobol',5);
X = rand(q,256,5);

```
See Also qrandstream | qrand | rand

Purpose Gamma random numbers
\begin{tabular}{|c|c|}
\hline Syntax & \(Y=\) randg \\
\hline & \(Y=\operatorname{randg}(\mathrm{A})\) \\
\hline & \(Y=\operatorname{randg}(\mathrm{A}, \mathrm{m})\) \\
\hline & \(Y=\operatorname{randg}(A, m, n, p, \ldots)\) \\
\hline & \(Y=\operatorname{randg}(A,[m, n, p, \ldots])\) \\
\hline
\end{tabular}

\section*{Description}
\(Y\) = randg returns a scalar random value chosen from a gamma distribution with unit scale and shape.
\(Y=r a n d g(A)\) returns a matrix of random values chosen from gamma distributions with unit scale. Y is the same size as A , and randg generates each element of \(Y\) using a shape parameter equal to the corresponding element of A .
\(Y=\operatorname{randg}(A, m)\) returns an \(m\)-by- \(m\) matrix of random values chosen from gamma distributions with shape parameters A. A is either an \(m-b y-m\) matrix or a scalar. If \(A\) is a scalar, randg uses that single shape parameter value to generate all elements of \(Y\).
\(Y=r a n d g(A, m, n, p, \ldots)\) or \(Y=r a n d g(A,[m, n, p, \ldots])\) returns an m-by-n-by-p-by-... array of random values chosen from gamma distributions with shape parameters A. A is either an m-by-n-by-p-by-... array or a scalar.
randg produces pseudo-random numbers using the MATLAB functions rand and randn. The sequence of numbers generated is determined by the state of the default random number stream. Since MATLAB resets the state at start-up, the sequence of numbers randg generates will be the same in each session unless those states are changed.

To create reproducible output from randg, reset the state of the random number stream before calling randg. For example:
```

s = rng % Obtain the current state of the random stream
% call randg
rng(s) % Reset the stream to the previous state
% call randg again, obtain identical results

```

See the rng documentation for more information.
Calling randg changes the current states of rand, randn, and randi, and therefore alters the outputs of subsequent calls to those functions.

To generate gamma random numbers and specify both the scale and shape parameters, you should call gamrnd rather than calling randg directly.

\section*{References [1] Marsaglia, G., and W. W. Tsang. "A Simple Method for Generating Gamma Variables." ACM Transactions on Mathematical Software. Vol. 26, 2000, pp. 363-372.}

See Also gamrnd

\section*{Purpose Random numbers}
```

Syntax }\quadY== random(name,A
Y = random(name,A,B)
Y = random(name,A,B,C)
Y = random(name,A,m,n,...)
Y = random(name,A,[m,n,···.])
Y = random(name, A, B,m,n,···.)
Y = random(name,A,B,[m,n,...])
Y = random(name,A,B,C,m,n,...)
Y = random(name,A,B,C,[m,n,...])

```

\section*{Description}
\(\mathrm{Y}=\) random(name, A\()\) where name is the name of a distribution that takes a single parameter, returns random numbers \(Y\) from the one-parameter family of distributions specified by name. Parameter values for the distribution are given in A .
\(Y\) is the same size as \(A\).
\(Y=\) random(name, \(A, B\) ) returns random numbers \(Y\) from a two-parameter family of distributions. Parameter values for the distribution are given in \(A\) and \(B\).

If \(A\) and \(B\) are arrays, they must be the same size. If either \(A\) or \(B\) are scalars, they are expanded to constant matrices of the same size.
\(Y=\) random(name, \(A, B, C\) ) returns random numbers \(Y\) from a three-parameter family of distributions. Parameter values for the distribution are given in \(A, B\), and \(C\).

If \(A, B\), and \(C\) are arrays, they must be the same size. If any of \(A, B\), or \(C\) are scalars, they are expanded to constant matrices of the same size.
\(Y=r a n d o m(n a m e, A, m, n, \ldots)\) or \(Y=r a n d o m(n a m e, A,[m, n, \ldots])\) returns an m-by-n-by... matrix of random numbers.

Similarly, \(Y=\) random(name, \(A, B, m, n, \ldots\) ) or \(Y=\) random(name, A, B, [m, n, ...]) returns an m-by-n-by... matrix of random numbers for distributions that require two parameters. \(Y=\) random(name, \(\mathrm{A}, \mathrm{B}, \mathrm{C}, \mathrm{m}, \mathrm{n}, \ldots\) ) or \(\mathrm{Y}=\)
random(name , \(A, B, C,[m, n, \ldots])\) returns an \(m\)-by- \(n\)-by... matrix of random numbers for distributions that require three parameters.
If any of \(A, B\), or \(C\) are arrays, then the specified dimensions must match the common dimensions of \(A, B\), and \(C\) after any necessary scalar expansion.
The following table denotes the acceptable strings for name, as well as the parameters for that distribution:
\begin{tabular}{l|l|l|l|l}
\hline name & Distribution & \begin{tabular}{l} 
Input \\
Parameter A
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter B
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter \\
C
\end{tabular} \\
\hline 'beta' or 'Beta' & \begin{tabular}{l} 
"Beta \\
Distribution" \\
on page B-4
\end{tabular} & a & b & - \\
\hline \begin{tabular}{l} 
'bino' or \\
'Binomial'
\end{tabular} & \begin{tabular}{l} 
"Binomial \\
Distribution" \\
on page B-7
\end{tabular} & \begin{tabular}{l} 
n: number of \\
trials
\end{tabular} & \begin{tabular}{l} 
p: probability of \\
success for each \\
trial
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'chi2' or \\
'Chisquare'
\end{tabular} & \begin{tabular}{l} 
"Chi-Square \\
Distribution" \\
on page B-12
\end{tabular} & \begin{tabular}{l} 
v: degrees of \\
freedom
\end{tabular} & - & - \\
\hline \begin{tabular}{l} 
'exp' or \\
'Exponential'
\end{tabular} & \begin{tabular}{l} 
"Exponential \\
Distribution" on \\
page B-16
\end{tabular} & \(\mu:\) mean & - & - \\
\hline \begin{tabular}{l} 
'ev' or 'Extreme \\
Value'
\end{tabular} & \begin{tabular}{l} 
"Extreme Value \\
Distribution" on \\
page B-19
\end{tabular} & \begin{tabular}{l}
\(\mu:\) location \\
parameter
\end{tabular} & \begin{tabular}{l}
\end{tabular} \\
\hline parameter scale
\end{tabular}
\begin{tabular}{l|l|l|l|l}
\hline name & Distribution & \begin{tabular}{l} 
Input \\
Parameter A
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter B
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter \\
C
\end{tabular} \\
\hline \begin{tabular}{l} 
'gev' or \\
Extreme Value'
\end{tabular} & \begin{tabular}{l} 
"Generalized \\
Extreme Value \\
Distribution" on \\
page B-32
\end{tabular} & \begin{tabular}{l} 
k: shape \\
parameter
\end{tabular} & \begin{tabular}{l}
\(\sigma:\) scale \\
parameter
\end{tabular} & \begin{tabular}{l}
\(\mu:\) location \\
parameter
\end{tabular} \\
\hline \begin{tabular}{l} 
'gp' or \\
'Generalized \\
Pareto'
\end{tabular} & \begin{tabular}{l} 
"Generalized \\
Pareto \\
Distribution" \\
on page B-37
\end{tabular} & \begin{tabular}{l} 
k: tail \\
index (shape) \\
parameter
\end{tabular} & \begin{tabular}{l}
\(\sigma:\) scale \\
parameter
\end{tabular} & \begin{tabular}{l}
\(\mu:\) threshold \\
(location) \\
parameter
\end{tabular} \\
\hline \begin{tabular}{l} 
'geo' or \\
'Geometric'
\end{tabular} & \begin{tabular}{l} 
"Geometric \\
Distribution" \\
on page B-41
\end{tabular} & \begin{tabular}{l} 
p: probability \\
parameter
\end{tabular} & - & - \\
\hline \begin{tabular}{l} 
'hyge' or \\
'Hypergeometric'
\end{tabular} & \begin{tabular}{l} 
"Hypergeometric \\
Distribution" on \\
page B-43
\end{tabular} & \begin{tabular}{l} 
M: size of the \\
population
\end{tabular} & \begin{tabular}{l} 
K: number of \\
items with \\
the desired \\
characteristic in \\
the population
\end{tabular} & \begin{tabular}{l} 
n: number \\
of samples \\
drawn
\end{tabular} \\
\hline \begin{tabular}{l} 
'logn' or \\
'Lognormal'
\end{tabular} & \begin{tabular}{l} 
"Lognormal \\
Distribution" \\
on page B-51
\end{tabular} & \(\mu\) & \(\sigma\) & - \\
\hline \begin{tabular}{l} 
'nbin' or \\
'Negative \\
Binomial'
\end{tabular} & \begin{tabular}{l} 
"Negative \\
Binomial \\
Distribution" \\
on page B-72
\end{tabular} & \begin{tabular}{l} 
r: number of \\
successes
\end{tabular} & \begin{tabular}{l} 
p: probability \\
of success in a \\
single trial
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'ncf' or \\
'Noncentral F'
\end{tabular} & \begin{tabular}{l} 
"Noncentral F \\
Distribution" on \\
page B-78
\end{tabular} & \begin{tabular}{l}
\(v 1:\) numerator \\
degrees of \\
freedom
\end{tabular} & \begin{tabular}{l}
\(v 2:\) denominator \\
degrees of \\
freedom
\end{tabular} & \begin{tabular}{l}
\(\delta:\) \\
noncentrality \\
parameter
\end{tabular} \\
\hline \begin{tabular}{l} 
'nct' or \\
'Noncentral t'
\end{tabular} & \begin{tabular}{l} 
"Noncentral t \\
Distribution" on \\
page B-80
\end{tabular} & \begin{tabular}{l}
\(v:\) degrees of \\
freedom
\end{tabular} & \begin{tabular}{l}
\(\delta:\) noncentrality \\
parameter
\end{tabular} & - \\
\hline
\end{tabular}
\begin{tabular}{l|l|l|l|l}
\hline name & Distribution & \begin{tabular}{l} 
Input \\
Parameter A
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter B
\end{tabular} & \begin{tabular}{l} 
Input \\
Parameter \\
C
\end{tabular} \\
\hline \begin{tabular}{l} 
'ncx2' or \\
'Noncentral \\
Chi-square'
\end{tabular} & \begin{tabular}{l} 
"Noncentral \\
Chi-Square \\
Distribution" \\
on page B-76
\end{tabular} & \begin{tabular}{l}
\(v:\) degrees of \\
freedom
\end{tabular} & \begin{tabular}{l}
\(\delta:\) noncentrality \\
parameter
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'norm' or \\
'Normal'
\end{tabular} & \begin{tabular}{l} 
"Normal \\
Distribution" \\
on page B-83
\end{tabular} & \(\mu:\) mean & \begin{tabular}{l} 
o: standard \\
deviation
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'poiss' or \\
'Poisson'
\end{tabular} & \begin{tabular}{l} 
"Poisson \\
Distribution" \\
on page B-89
\end{tabular} & \(\lambda:\) mean & - & - \\
\hline \begin{tabular}{l} 
'rayl' or \\
'Rayleigh'
\end{tabular} & \begin{tabular}{l} 
"Rayleigh \\
Distribution" \\
on page B-91
\end{tabular} & \begin{tabular}{l} 
b: scale \\
parameter
\end{tabular} & - & - \\
\hline 't' or 'T' & \begin{tabular}{l} 
"Student's t \\
Distribution" \\
on page B-95
\end{tabular} & \begin{tabular}{l}
\(v:\) degrees of \\
freedom
\end{tabular} & - & - \\
\hline \begin{tabular}{l} 
'unif' or \\
'Uniform'
\end{tabular} & \begin{tabular}{l} 
"Uniform \\
Distribution \\
(Continuous)" \\
on page B-99
\end{tabular} & \begin{tabular}{l} 
a: lower \\
endpoint \\
(minimum)
\end{tabular} & \begin{tabular}{l} 
b: upper \\
endpoint \\
(maximum)
\end{tabular} & - \\
\hline \begin{tabular}{l} 
'unid' or \\
'Discrete \\
Uniform'
\end{tabular} & \begin{tabular}{l} 
"Uniform \\
Distribution \\
(Discrete" on \\
page B-101
\end{tabular} & \begin{tabular}{l} 
N: maximum \\
observable \\
value
\end{tabular} & - & - \\
\hline \begin{tabular}{l} 
'wbl' or \\
'Weibull'
\end{tabular} & \begin{tabular}{l} 
"Weibull \\
Distribution" \\
on page B-103
\end{tabular} & \begin{tabular}{l} 
a: scale \\
parameter
\end{tabular} & \begin{tabular}{l} 
b: shape \\
parameter
\end{tabular} & - \\
\hline
\end{tabular}

Examples Generate a 2-by-4 array of random values from the normal distribution with mean 0 and standard deviation 1 :
```

x1 = random('Normal', 0,1,2,4)
x1 =
1.1650 0.0751 -0.6965 0.0591
0.6268 0.3516 1.6961 1.7971

```

The order of the parameters is the same as for normrnd.
Generate a single random value from Poisson distributions with rate parameters \(1,2, \ldots, 6\), respectively:
```

x2 = random('Poisson',1:6,1,6)
x2 =
0

```

See Also cdf | pdf | icdf | mle

\section*{Purpose Random numbers from Gaussian mixture distribution}
Syntax \(\quad\)\begin{tabular}{ll}
\(y\) & \(=\operatorname{random}(o b j)\) \\
& \(Y=\operatorname{random}(o b j, n)\) \\
& {\([Y, i d x]=\operatorname{random}(o b j, n)\)}
\end{tabular}

Description \(\quad y=\) random(obj) generates a 1 -by- \(d\) vector \(y\) drawn at random from the \(d\)-dimensional Gaussian mixture distribution defined by obj. obj is an object created by gmdistribution or fit.
\(\mathrm{Y}=\) random (obj, n ) generates an n -by- \(d\) matrix Y of \(\mathrm{n} d\)-dimensional random samples.
[ \(\mathrm{Y}, \mathrm{idx}\) ] = random(obj, n ) also returns an n -by-1 vector idx, where \(\operatorname{idx}(\mathrm{I})\) is the index of the component used to generate \(Y(I,:)\).

\section*{Examples Create a gmdistribution object defining a two-component mixture of} bivariate Gaussian distributions:
```

MU = [1 2;-3 -5];
SIGMA = cat(3,[2 0;0 .5],[1 0;0 1]);
p = ones(1,2)/2;
obj = gmdistribution(MU,SIGMA,p);
ezcontour(@(x,y)pdf(obj,[x y]),[-10 10],[-10 10])
hold on

```

\section*{gmdistribution.random}


Generate 1000 random values:
```

Y = random(obj,1000);
scatter(Y(:,1),Y(:,2),10,'.')

```


\section*{See Also}
gmdistribution | fit | mvnrnd

\section*{piecewisedistribution.random}

Purpose Random numbers from piecewise distribution
```

Syntax r = random(obj)
R = random(obj,n)
R = random(obj,m,n)
R = random(obj,[m,n])
R = random(obj,m,n,p,···...
R = random(obj,[m,n,p,···..])

```

Description \(r=\) random \((\mathrm{obj})\) generates a pseudo-random number \(r\) drawn from the piecewise distribution object obj.
\(R=\operatorname{random}(o b j, n)\) generates an \(n\)-by- \(n\) matrix of pseudo-random numbers R.
\(R=\operatorname{random}(o b j, m, n)\) or \(R=\operatorname{random}(o b j,[m, n])\) generates an \(m\)-by- \(n\) matrix of pseudo-random numbers \(R\).
\(R=\operatorname{random}(o b j, m, n, p, \ldots)\) or \(R=\operatorname{random}(o b j,[m, n, p, \ldots])\) generates an \(m\)-by- \(n\)-by- \(p\)-by-... array of pseudo-random numbers R .

\section*{Examples \(\quad\) Fit Pareto tails to a \(t\) distribution at cumulative probabilities 0.1 and} 0.9:
```

t = trnd(3,100,1);
obj = paretotails(t,0.1,0.9);
r = random(obj)
r=
0.8285

```

\section*{See Also paretotails | cdf | icdf}

\section*{Purpose}

Generate random number drawn from ProbDist object
Syntax
\(Y=\operatorname{random}(P D)\)
\(Y=\operatorname{random}(P D, N)\)
\(Y=\operatorname{random}(P D, N, M, \ldots)\)
\(Y=\operatorname{random}(P D)\) generates a random number drawn from the distribution specified by \(P D\), a ProbDist object.
\(Y=\operatorname{random}(P D, N)\) generates an \(N\)-by- \(N\) array of random numbers drawn from the distribution specified by \(P D\), a ProbDist object.
\(Y=\operatorname{random}(P D, N, M, \ldots)\) generates an \(N\)-by- \(M\)-by ... array of random numbers drawn from the distribution specified by \(P D\), a ProbDist object.

\section*{Input \\ Arguments}

\section*{Output \\ Arguments}

See Also

An object of the class ProbDistUnivParam or ProbDistUnivKernel.

A positive integer.
A positive integer.

A random number drawn from the distribution specified by \(P D\).

\section*{randsample}

Purpose Random sample
```

Syntax y = randsample(n,k)
y = randsample(population,k)
y = randsample(n,k,replacement)
y = randsample(population, k,replacement)
y = randsample(n,k,true,w)
y = randsample(population,k,true,w)
y = randsample(s,...)

```

\section*{Description}

\section*{Examples}

Draw a single value from the integers 1 through 10:
```

n = 10;
x = randsample(n,1);

```

Draw a single value from the population 1 through \(n\), where \(n>1\) :
\[
y=\text { randsample(1:n,1); }
\]

Note If population is a numeric vector containing only nonnegative integer values, and population can have length 1, use
```

y = population(randsample(length(population),k))

```
instead of \(y=\) randsample(population, \(k\) ).

Generate a random sequence of the characters A, C, G, and T, with replacement, according to the specified probabilities.
\[
\mathrm{R}=\text { randsample('ACGT',48,true,[0.15 } 0.350 .350 .15])
\]

\section*{randtool}

Purpose Interactive random number generation

\section*{Syntax randtool}

Description randtool opens the Random Number Generation Tool.
The Random Number Generation Tool is a graphical user interface that generates random samples from specified probability distributions and displays the samples as histograms. Use the tool to explore the effects of changing parameters and sample size on the distributions.


\section*{randtool}

Start by selecting a distribution, then enter the desired sample size.
You can also
- Use the controls at the bottom of the window to set parameter values for the distribution and to change their upper and lower bounds.
- Draw another sample from the same distribution, with the same size and parameters.
- Export the current sample to your workspace. A dialog box enables you to provide a name for the sample.

\section*{See Also}
disttool | dfittool
Purpose Range of values
Syntax range (X) ..... y = range(X, dim)
Description
range \((X)\) returns the difference between the maximum and the minimum of a sample. For vectors, range \((x)\) is the range of the elements. For matrices, range ( X ) is a row vector containing the range of each column of X. For N-dimensional arrays, range operates along the first nonsingleton dimension of \(X\).
\(\mathrm{y}=\) range( \(\mathrm{X}, \mathrm{dim}\) ) operates along the dimension dim of X .
range treats NaNs as missing values and ignores them.
The range is an easily-calculated estimate of the spread of a sample. Outliers have an undue influence on this statistic, which makes it an unreliable estimator.
```

Examples The range of a large sample of standard normal random numbers is
approximately six. This is the motivation for the process capability
indices $C_{\mathrm{p}}$ and $C_{\mathrm{pk}}$ in statistical quality control applications.

```
```

rv = normrnd(0,1,1000,5);

```
rv = normrnd(0,1,1000,5);
near6 = range(rv)
near6 = range(rv)
near6 =
near6 =
    6.1451 6.4986 6.2909 5.8894 7.0002
```

    6.1451 6.4986 6.2909 5.8894 7.0002
    ```
See Also std \| iq \| mad

Purpose Find all neighbors within specified distance
```

Syntax idx = rangesearch(X,Y,r)
[idx,D]= rangesearch(X,Y,r)
[idx,D]= rangesearch(X,Y,r,Name,Value)

```

\section*{Description}

\section*{Tips}

Input
Arguments
\(i d x=\) rangesearch \((X, Y, r)\) finds all the \(X\) points that are within distance \(r\) of the \(Y\) points. Rows of \(X\) and \(Y\) correspond to observations, and columns correspond to variables.
[idx, D]= rangesearch ( \(\mathrm{X}, \mathrm{Y}, \mathrm{r}\) ) returns the distances between each row of \(Y\) and the rows of \(X\) that are \(r\) or less distant.
[idx, D]= rangesearch ( \(X, Y, r\), Name, Value) finds nearby points with
- For a fixed positive integer K, knnsearch finds K points in \(X\) that are nearest each \(Y\) point. In contrast, for a fixed positive real value \(r\), rangesearch finds all the \(X\) points that are within a distance \(r\) of each \(Y\) point.

X
additional options specified by one or more Name, Value pair arguments.
\(m x\)-by- \(n\) numeric matrix, where each row represents one \(n\)-dimensional point. The number of columns \(n\) must equal as the number of columns in \(Y\).

Y
\(m y\)-by- \(n\) numeric matrix, where each row represents one \(n\)-dimensional point. The number of columns \(n\) must equal as the number of columns in \(X\).
\(r\)

Search radius, a scalar. rangesearch finds all \(X\) points (rows) that are within distance \(r\) of each \(Y\) point. The meaning of distance depends on the Distance name-value pair.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

\section*{BucketSize}

Maximum number of data points in the leaf node of the \(k\) d-tree. This argument is only meaningful when \(k d\)-tree is used for finding nearest neighbors.

Default: 50

Cov
Positive definite matrix indicating the covariance matrix when computing the Mahalanobis distance. This argument is only valid when the Distance name-value pair is 'mahalanobis'.

Default: nancov (X)

\section*{Distance}

String or function handle specifying the distance metric.
\begin{tabular}{l|l}
\hline Value & Description \\
\hline 'euclidean' & Euclidean distance. \\
\hline 'seuclidean' & \begin{tabular}{l} 
Standardized Euclidean distance. Each \\
coordinate difference between X and a \\
query point is scaled, meaning divided by \\
a scale value S. The default value of S is \\
the standard deviation computed from X, \\
S=nanstd (X). To specify another value for \\
S, use the Scale name-value pair.
\end{tabular} \\
\hline
\end{tabular}
\begin{tabular}{l|l}
\hline Value & Description \\
\hline 'mahalanobis' & \begin{tabular}{l} 
Mahalanobis distance, computed using a \\
positive definite covariance matrix C. The \\
default value of C is the sample covariance \\
matrix of X, as computed by nancov (X). \\
To specify a different value for C, use the \\
'Cov' name-value pair.
\end{tabular} \\
\hline 'cityblock' & City block distance. \\
\hline 'minkowski' & \begin{tabular}{l} 
Minkowski distance. The default exponent \\
is 2. To specify a different exponent, use \\
the 'P' name-value pair.
\end{tabular} \\
\hline 'chebychev' & \begin{tabular}{l} 
Chebychev distance (maximum coordinate \\
difference).
\end{tabular} \\
\hline 'cosine' & \begin{tabular}{l} 
One minus the cosine of the included angle \\
between observations (treated as vectors).
\end{tabular} \\
\hline 'correlation' & \begin{tabular}{l} 
One minus the sample linear correlation \\
between observations (treated as \\
sequences of values).
\end{tabular} \\
\hline 'hamming' & \begin{tabular}{l} 
Hamming distance, percentage of \\
coordinates that differ.
\end{tabular} \\
\hline 'jaccard' & \begin{tabular}{l} 
One minus the Jaccard coefficient, the \\
percentage of nonzero coordinates that \\
differ.
\end{tabular} \\
\hline
\end{tabular}
\begin{tabular}{|c|c|}
\hline Value & Description \\
\hline 'spearman' & One minus the sample Spearman's rank correlation between observations (treated as sequences of values). \\
\hline @distfun & \begin{tabular}{l}
Distance function handle. distfun has the form \\
function D2 = DISTFUN(ZI,ZJ) \\
where \\
- ZI is a 1 -by-N vector containing one row of \(X\) or \(Y\). \\
- ZJ is an M 2 -by- N matrix containing multiple rows of X or Y . \\
- D2 is an M2-by-1 vector of distances, \(\mathrm{D} 2(\mathrm{k})\) is the distance between the observations ZI and ZJ (J,: ).
\end{tabular} \\
\hline
\end{tabular}

For definitions, see "Distance Metrics" on page 13-9.
Default: 'euclidean'

\section*{NSMethod}

Nearest neighbors search method.
\begin{tabular}{l|l}
\hline Value & Meaning \\
\hline 'kdtree' & \begin{tabular}{l} 
Creates and uses a \(k\) d-tree to find nearest \\
neighbors. 'kdtree ' is only valid when \\
the distance metric is one of:
\end{tabular} \\
& - 'chebyshev' \\
& - 'cityblock' \\
& - 'euclidean ' \\
& - 'minkowski'
\end{tabular}

Default: 'kdtree' when the number of columns of X is not greater than \(10, X\) is not sparse, and the distance metric is one of the valid 'kdtree' metrics. Otherwise, the default is 'exhaustive'.

Positive scalar indicating the exponent of Minkowski distance. This argument is only valid when the Distance name-value pair is 'minkowski'.

\section*{Default: 2}

\section*{Scale}

Vector \(S\) containing nonnegative values, with length equal to the number of columns in X. Each coordinate difference between X and a query point is scaled by the corresponding element of S . This argument is only valid when the Distance name-value pair is 'seuclidean'.

Default: nanstd(X)

\section*{Output \\ idx}

Arguments

\section*{Definitions}

\section*{Examples}
\(m y\)-by- 1 cell array, where \(m y\) is the number of rows in Y. idx \(\{\mathrm{I}\}\) contains the indices of points (rows) in X whose distances to \(Y(I,:)\) are not greater than \(r\). The entries in idx \(\{I\}\) are in ascending order of distance.

\section*{D}
\(m y\)-by- 1 cell array, where \(m y\) is the number of rows in Y . \(\mathrm{D}\{\mathrm{I}\}\) contains the distance values between \(\mathrm{Y}(\mathrm{I},:\) ) and the corresponding points in idx\{I\}.

\section*{Distance Metrics}

For definitions, see "Distance Metrics" on page 13-9.

Find the \(X\) points that are within a Euclidean distance 1.5 of each \(Y\) point. Both \(X\) and \(Y\) are samples of 5-D normally distributed variables.
```

rng('default') % for reproducibility
X = randn(100,5);
Y = randn(10,5);
[idx, dist] = rangesearch(X,Y,1.5)
idx =
[1x7 double]
[1x2 double]
[1\times11 double]
[1x2 double]
[1x12 double]
[1x9 double]
[[ 89]
[1x0 double]
[1x0 double]
[1x0 double]
dist =
[1x7 double]

```
```

[1x2 double]
[1\times11 double]
[1x2 double]
[1x12 double]
[1x9 double]
[ 1.1739]
[1x0 double]
[1x0 double]
[1x0 double]

```

In this case, the last three \(Y\) points are more than 1.5 distant from any \(X\) point. \(X(89,:)\) is 1.1739 distant from \(Y(7,:)\), and there is no other \(X\) point that is within distance 1.5 of \(Y(7,:)\). There are 12 points in \(X\) within distance 1.5 of \(Y(5,:)\).

\footnotetext{
Algorithms For an overview of the \(k\) d-tree algorithm, see " \(k\)-Nearest Neighbor Search Using a \(k\) d-Tree" on page 13-14.

The exhaustive search algorithm finds the distance of each point in \(X\) to each point in \(Y\).
Alternatives \begin{tabular}{l} 
rangesearch is the ExhaustiveSearcher method for distance search. \\
It is equivalent to the rangesearch function with the NSMethod \\
name-value pair set to 'exhaustive '. \\
rangesearch is the KDTreeSearcher method for distance search. It is \\
equivalent to the rangesearch function with the NSMethod name-value \\
pair set to 'kdtree'.
\end{tabular}
See Also createns | ExhaustiveSearcher | KDTreeSearcher | knnsearch | pdist2

How To . " \(k\)-Nearest Neighbor Search and Radius Search" on page 13-12
}

\section*{Purpose}

Syntax

Description

\section*{Tips}

Input
Arguments

Find all neighbors within specified distance using ExhaustiveSearcher object
idx = rangesearch(NS, \(Y, r\) )
[idx, D]= rangesearch(NS, Y, r)
[idx, D] = rangesearch(NS, Y, r, Name, Value)
idx \(=\) rangesearch(NS, \(Y, r\) ) finds all points in NS. \(X\) that are within distance \(r\) of the \(Y\) points. Rows of NS. \(X\) and \(Y\) correspond to observations, and columns correspond to variables.
[idx, D]= rangesearch(NS, Y, r) returns the distances between each row of \(Y\) and the rows of NS. \(X\) that are \(r\) or less distant.
[idx, D] = rangesearch(NS, \(Y\), \(r\), Name, Value) finds nearby points with additional options specified by one or more Name, Value pair arguments.
- For a fixed positive integer K, knnsearch finds the K points in NS.X that are nearest each \(Y\) point. In contrast, for a fixed positive real value \(r\), rangesearch finds all the points in NS. \(X\) that are within a distance \(r\) of each \(Y\) point.

NS
ExhaustiveSearcher object, constructed using ExhaustiveSearcher or createns.

Y
\(m y\)-by- \(n\) numeric matrix, where each row represents one \(n\)-dimensional point. The number of columns \(n\) must equal the number of columns in NS.X.

\section*{\(r\)}

Search radius, a scalar. rangesearch finds all NS.X points (rows) that are within distance \(r\) of each \(Y\) point. The meaning of distance depends on the Distance name-value pair.

\section*{ExhaustiveSearcher.rangesearch}

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

Cov
Positive definite matrix indicating the covariance matrix when computing the Mahalanobis distance. This argument is only valid when the Distance name-value pair is 'mahalanobis'.

Default: nancov (X)

\section*{Distance}

String or function handle specifying the distance metric.
\begin{tabular}{l|l}
\hline Value & Description \\
\hline 'euclidean' & Euclidean distance. \\
\hline 'seuclidean' & \begin{tabular}{l} 
Standardized Euclidean distance. Each \\
coordinate difference between X and a \\
query point is scaled, meaning divided by \\
a scale value S. The default value of S is \\
the standard deviation computed from X, \\
S=nanstd (X). To specify another value for \\
S, use the Scale name-value pair.
\end{tabular} \\
\hline 'mahalanobis' & \begin{tabular}{l} 
Mahalanobis distance, computed using a \\
positive definite covariance matrix C. The \\
default value of C is the sample covariance \\
matrix of X, as computed by nancov (X). \\
To specify a different value for C, use the \\
'Cov' name-value pair.
\end{tabular} \\
\hline 'cityblock' & City block distance. \\
\hline
\end{tabular}
\(\left.\left.\begin{array}{l|l}\hline \text { Value } & \text { Description } \\ \hline \text { 'minkowski' } & \begin{array}{l}\text { Minkowski distance. The default exponent } \\ \text { is 2. To specify a different exponent, use } \\ \text { the 'P' name-value pair. }\end{array} \\ \hline \text { 'chebychev' } & \begin{array}{l}\text { Chebychev distance (maximum coordinate } \\ \text { difference). }\end{array} \\ \hline \text { 'cosine' } & \begin{array}{l}\text { One minus the cosine of the included angle } \\ \text { between observations (treated as vectors). }\end{array} \\ \hline \text { 'correlation' } & \begin{array}{l}\text { One minus the sample linear correlation } \\ \text { between observations (treated as } \\ \text { sequences of values). }\end{array} \\ \hline \text { 'hamming' } & \begin{array}{l}\text { Hamming distance, percentage of } \\ \text { coordinates that differ. }\end{array} \\ \hline \text { 'jaccard' } & \begin{array}{l}\text { One minus the Jaccard coefficient, the } \\ \text { percentage of nonzero coordinates that } \\ \text { differ. }\end{array} \\ \hline \text { 'spearman' } & \begin{array}{l}\text { One minus the sample Spearman's rank } \\ \text { correlation between observations (treated } \\ \text { as sequences of values). }\end{array} \\ \hline \text { @distfun } & \begin{array}{l}\text { Distance function handle. distfun has } \\ \text { the form }\end{array} \\ \text { function D2 = DISTFUN(ZI , ZJ) }\end{array}\right\} \begin{array}{l}\text { I. } \\ \text { where } \\ \text { - ZI is a 1-by-N vector containing one row } \\ \text { of X or Y. } \\ \text { - ZJ is an M2-by-N matrix containing } \\ \text { multiple rows of X or Y. }\end{array}\right\}\)

\section*{ExhaustiveSearcher.rangesearch}
\begin{tabular}{l|l} 
Value & Description \\
\hline & - D2 is an M2-by-1 vector of distances, \\
& D2 (k) is the distance between the \\
observations ZI and \(\mathrm{ZJ}(J,:)\). \\
\hline
\end{tabular}

For definitions, see "Distance Metrics" on page 13-9.
Default: 'euclidean'
P
Positive scalar indicating the exponent of Minkowski distance. This argument is only valid when the Distance name-value pair is 'minkowski'.

\section*{Default: 2}

\section*{Scale}

Vector \(S\) containing nonnegative values, with length equal to the number of columns in X. Each coordinate difference between X and a query point is scaled by the corresponding element of \(S\). This argument is only valid when the Distance name-value pair is 'seuclidean'.

Default: nanstd (X)

\section*{Output idx \\ Arguments}
\(m y\)-by- 1 cell array, where \(m y\) is the number of rows in Y . idx \(\{\mathrm{I}\}\) contains the indices of points (rows) in NS. X whose distances to \(Y(I,:)\) are not greater than \(r\). The entries in idx \(\{I\}\) are in ascending order of distance.

\footnotetext{
D
}

\section*{ExhaustiveSearcher.rangesearch}
\(m y\)-by- 1 cell array, where \(m y\) is the number of rows in Y . \(D\{I\}\) contains the distance values between \(Y(I,:)\) and the corresponding points in idx\{I\}.

\section*{Definitions Distance Metrics}

For definitions, see "Distance Metrics" on page 13-9.
Examples Create \(X\) and \(Y\) as samples of 5-D normally distributed variables. Create an ExhaustiveSearcher object NS from X. Find the points in NS.X that are within a Euclidean distance 1.5 of each point in \(Y\).
```

rng('default') % for reproducibility
X = randn(100,5);
Y = randn(10,5);
NS = ExhaustiveSearcher(X);
[idx, dist] = rangesearch(NS,Y,1.5)
idx =
[1x7 double]
[1x2 double]
[1x11 double]
[1x2 double]
[1\times12 double]
[1x9 double]
[[ 89]
[1x0 double]
[1x0 double]
[1x0 double]
dist =
[1x7 double]
[1x2 double]
[1\times11 double]
[1x2 double]
[1x12 double]
[1x9 double]

```

\section*{ExhaustiveSearcher.rangesearch}
```

[ 1.1739]
[1x0 double]
[1x0 double]
[1x0 double]

```

In this case, the last three points in \(Y\) are more than 1.5 distant from any point in NS.X. NS. \(X(89,:)\) is 1.1739 distant from \(Y(7,:)\), and there is no other point in NS. \(X\) that is within distance 1.5 of \(Y(7,:)\). There are 12 points in NS. \(X\) within distance 1.5 of \(Y(5,:)\).

\section*{Algorithms The exhaustive search algorithm finds the distance of each point in \(X\) to each point in \(Y\).}

\title{
Alternatives rangesearch is the ExhaustiveSearcher method for distance search. It is equivalent to the rangesearch function with the NSMethod name-value pair set to 'exhaustive'. \\ rangesearch is the KDTreeSearcher method for distance search. It is equivalent to the rangesearch function with the NSMethod name-value pair set to 'kdtree'.
}

\author{
See Also \\ createns | ExhaustiveSearcher | knnsearch | pdist2 | rangesearch
}
\begin{tabular}{|c|c|}
\hline Purpose & Find all neighbors within specified distance using KDTreeSearcher object \\
\hline Syntax & ```
idx = rangesearch(NS,Y,r)
[idx,D]= rangesearch(NS,Y,r)
[idx,D]= rangesearch(NS,Y,r,Name,Value)
``` \\
\hline Description & \begin{tabular}{l}
idx = rangesearch(NS, \(Y, r\) ) finds all points in NS. \(X\) that are within distance \(r\) of the \(Y\) points. Rows of NS. \(X\) and \(Y\) correspond to observations, and columns correspond to variables. \\
[idx, D]= rangesearch(NS, \(Y, r\) ) returns the distances between each row of \(Y\) and the rows of NS. \(X\) that are \(r\) or less distant. \\
[idx, D]= rangesearch(NS, \(Y\), r,Name, Value) finds nearby points with additional options specified by one or more Name, Value pair arguments.
\end{tabular} \\
\hline Tips & - For a fixed positive integer K, knnsearch finds the K points in NS.X that are nearest each \(Y\) point. In contrast, for a fixed positive real value \(r\), rangesearch finds all the points in NS. \(X\) that are within a distance \(r\) of each \(Y\) point. \\
\hline Input Arguments & \begin{tabular}{l}
NS \\
KDTreeSearcher object, constructed using KDTreeSearcher or createns.
\end{tabular} \\
\hline & Y \\
\hline & \(m y\)-by- \(n\) numeric matrix, where each row represents one \(n\)-dimensional point. The number of columns \(n\) must equal the number of columns in NS.X. \\
\hline & \(r\) \\
\hline & Search radius, a scalar. rangesearch finds all NS.X points (rows) that are within distance \(r\) of each \(Y\) point. The meaning of distance depends on the Distance name-value pair. \\
\hline
\end{tabular}

\section*{KDTreeSearcher.rangesearch}

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{Distance}

String or function handle specifying the distance metric.
\begin{tabular}{l|l}
\hline Value & Description \\
\hline 'euclidean' & Euclidean distance. \\
\hline 'cityblock' & City block distance. \\
\hline 'minkowski' & \begin{tabular}{l} 
Minkowski distance. The default exponent \\
is 2. To specify a different exponent, use \\
the 'P' name-value pair.
\end{tabular} \\
\hline 'chebychev' & \begin{tabular}{l} 
Chebychev distance (maximum coordinate \\
difference).
\end{tabular} \\
\hline
\end{tabular}

For definitions, see "Distance Metrics" on page 13-9.
Default: NS.Distance

\section*{P}

Positive scalar indicating the exponent of Minkowski distance. This argument is only valid when the Distance name-value pair is 'minkowski'.

Default: 2

Output
Arguments
idx
\(m y\)-by- 1 cell array, where \(m y\) is the number of rows in Y . idx \(\{\mathrm{I}\}\) contains the indices of points (rows) in NS. X whose distances
to \(Y(I,:)\) are not greater than \(r\). The entries in idx\{I\} are in ascending order of distance.

\section*{D}
\(m y\)-by- 1 cell array, where \(m y\) is the number of rows in Y . \(\mathrm{D}\{\mathrm{I}\}\) contains the distance values between \(\mathrm{Y}(\mathrm{I},:()\) and the corresponding points in idx\{I\}.

\section*{Definitions \\ Distance Metrics}

For definitions, see "Distance Metrics" on page 13-9.
Examples Create \(X\) and \(Y\) as samples of 5-D normally distributed variables. Create a KDTreeSearcher object NS from X. Find the points in NS.X that are within a Euclidean distance 1.5 of each point in Y .
```

rng('default') % for reproducibility
X = randn(100,5);
Y = randn(10,5);
NS = KDTreeSearcher(X);
[idx, dist] = rangesearch(NS,Y,1.5)
idx =
[1x7 double]
[1x2 double]
[1\times11 double]
[1x2 double]
[1\times12 double]
[1x9 double]
[ 89]
[1x0 double]
[1x0 double]
[1x0 double]
dist =
[1x7 double]
[1x2 double]

```

\section*{KDTreeSearcher.rangesearch}
```

[1x11 double]
[1x2 double]
[1\times12 double]
[1x9 double]
[ 1.1739]
[1x0 double]
[1x0 double]
[1x0 double]

```

In this case, the last three points in \(Y\) are more than 1.5 distant from any point in NS.X. NS. \(\mathrm{X}(89,:)\) is 1.1739 distant from \(Y(7,:)\), and there is no other point in NS. \(X\) that is within distance 1.5 of \(Y(7,:)\). There are 12 points in NS. \(X\) within distance 1.5 of \(Y(5,:)\).

\section*{Algorithms \\ For an overview of the \(k\) d-tree algorithm, see " \(k\)-Nearest Neighbor Search Using a \(k\) d-Tree" on page 13-14.}

\footnotetext{
Alternatives rangesearch is the KDTreeSearcher method for distance search. It is equivalent to the rangesearch function with the NSMethod name-value pair set to 'kdtree'.
rangesearch is the ExhaustiveSearcher method for distance search. It is equivalent to the rangesearch function with the NSMethod name-value pair set to 'exhaustive'.

See Also createns | KDTreeSearcher | knnsearch | pdist2 | rangesearch
}

\section*{Purpose Wilcoxon rank sum test}

\section*{Syntax}
```

p = ranksum(x,y)
[p,h] = ranksum(x,y)
[p,h] = ranksum(x,y,'alpha',alpha)
[p,h] = ranksum(...,'method',method)
[p,h,stats] = ranksum(...)

```

\section*{Description}
\(\mathrm{p}=\) ranksum ( \(\mathrm{x}, \mathrm{y}\) ) performs a two-sided rank sum test of the null hypothesis that data in the vectors \(x\) and \(y\) are independent samples from identical continuous distributions with equal medians, against the alternative that they do not have equal medians. \(x\) and \(y\) can have different lengths. The \(p\) value of the test is returned in p . The test is equivalent to a Mann-Whitney \(U\)-test.
\([p, h]=\operatorname{ranksum}(x, y)\) returns the result of the test in \(h . h=1\) indicates a rejection of the null hypothesis at the \(5 \%\) significance level. \(\mathrm{h}=0\) indicates a failure to reject the null hypothesis at the \(5 \%\) significance level.
[p,h] = ranksum(x,y,'alpha',alpha) performs the test at the (100*alpha)\% significance level. The default, when unspecified, is alpha \(=0.05\).
\([\mathrm{p}, \mathrm{h}]=\) ranksum(...,'method', method) computes the \(p\) value using either an exact algorithm, when method is 'exact', or a normal approximation, when method is 'approximate'. Because the exact method can be very slow on large samples, the default is exact for small samples and approximate for large samples.
[ \(p, h\), stats] \(=\) ranksum(...) returns the structure stats with the following fields:
- ranksum - Value of the rank sum test statistic
- zval - Value of the \(z\)-statistic (computed only for large samples)
```

Examples Test the hypothesis of equal medians for two independent unequal-sized
samples. The sampling distributions are identical except for a shift
of 0.25 .

```
```

x = unifrnd(0,1,10,1);

```
x = unifrnd(0,1,10,1);
y = unifrnd(0.25,1.25,15,1);
y = unifrnd(0.25,1.25,15,1);
[p,h] = ranksum(x,y)
[p,h] = ranksum(x,y)
p =
p =
    0.0375
    0.0375
h =
h =
    1
```

    1
    ```

The test rejects the null hypothesis of equal medians at the default \(5 \%\) significance level.

References [1] Gibbons, J. D. Nonparametric Statistical Inference. New York: Marcel Dekker, 1985.
[2] Hollander, M., and D. A. Wolfe. Nonparametric Statistical Methods. Hoboken, NJ: John Wiley \& Sons, Inc., 1999.

See Also kruskalwallis | signrank | signtest | ttest2

\section*{Purpose}

Rayleigh cumulative distribution function

\section*{Syntax \\ \(P=\operatorname{raylcdf}(X, B)\)}
\(P=\) raylcdf \((X, B)\) computes the Rayleigh cdf at each of the values in \(X\) using the corresponding scale parameter, \(B . X\) and \(B\) can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input for \(X\) or \(B\) is expanded to a constant array with the same dimensions as the other input.

The Rayleigh cdf is
\[
y=F(x \mid b)=\int_{0}^{x} \frac{t}{b^{2}} e^{\left(\frac{-t^{2}}{2 b^{2}}\right)} d t
\]

\section*{Examples}

\section*{References}

See Also
How To
\[
\begin{aligned}
& x=0: 0.1: 3 ; \\
& p=\operatorname{raylcdf}(x, 1) ; \\
& \operatorname{plot}(x, p)
\end{aligned}
\]

[1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions. Hoboken, NJ: Wiley-Interscience, 2000. pp. 134-136.
```

cdf | raylpdf | raylinv | raylstat | raylfit | raylrnd

```
- "Rayleigh Distribution" on page B-91

Purpose Rayleigh parameter estimates
```

Syntax raylfit(data, alpha)
[phat,pci] = raylfit(data,alpha)

```

Description raylfit(data, alpha) returns the maximum likelihood estimates of the parameter of the Rayleigh distribution given the data in the vector data.
[phat, pci] = raylfit(data,alpha) returns the maximum likelihood estimate and 100(1-alpha)\% confidence interval given the data. The default value of the optional parameter alpha is 0.05 , corresponding to \(95 \%\) confidence intervals.

\section*{See Also \\ mle | raylpdf | raylcdf | raylinv | raylstat | raylrnd}

How To . "Rayleigh Distribution" on page B-91
Purpose Rayleigh inverse cumulative distribution function
Syntax X = raylinv(P,B)
Description \(X=\) raylinv( \(P, B\) ) returns the inverse of the Rayleigh cumulativedistribution function using the corresponding scale parameter, \(B\) atthe corresponding probabilities in P. P and B can be vectors, matrices,or multidimensional arrays that all have the same size. A scalar inputfor P or B is expanded to a constant array with the same dimensionsas the other input.
Examples

\(x=\operatorname{raylinv}(0.9,1)\)

\(\mathrm{x}=\)

    2.1460
See Also raylcdf | raylpdf | raylrnd | raylstat
How To ..... - icdf- "Rayleigh Distribution" on page B-91

Purpose Rayleigh probability density function

\section*{Syntax \(\quad Y=\operatorname{raylpdf}(X, B)\)}

Description
\(Y=\operatorname{raylpdf}(X, B)\) computes the Rayleigh pdf at each of the values in \(X\) using the corresponding scale parameter, \(B . X\) and \(B\) can be vectors, matrices, or multidimensional arrays that all have the same size, which is also the size of Y . A scalar input for X or B is expanded to a constant array with the same dimensions as the other input.

The Rayleigh pdf is
\[
y=f(x \mid b)=\frac{x}{b^{2}} e^{\left(\frac{-x^{2}}{2 b^{2}}\right)}
\]

\section*{Examples}


\section*{See Also \\ pdf | raylcdf | raylinv | raylstat | raylfit | raylrnd}

\footnotetext{
How To
- "Rayleigh Distribution" on page B-91
}

\section*{Purpose Rayleigh random numbers}

\section*{Syntax}
\(R=\) raylrnd(B)
\(R=\) raylrnd(B, v)
\(R=\operatorname{raylrnd}(B, m, n)\)

\section*{Examples}

See Also
How To
\(R=\) raylrnd \((B)\) returns a matrix of random numbers chosen from the Rayleigh distribution with scale parameter, B. B can be a vector, a matrix, or a multidimensional array. The size of \(R\) is the size of \(B\).
\(R=\) raylrnd \((B, v)\) returns a matrix of random numbers chosen from the Rayleigh distribution with parameter \(B\), where \(v\) is a row vector. If \(v\) is a 1 -by- 2 vector, \(R\) is a matrix with \(v(1)\) rows and \(v(2)\) columns. If \(v\) is 1 -by- \(\mathrm{n}, \mathrm{R}\) is an n -dimensional array.
\(R=\operatorname{raylrnd}(B, m, n)\) returns a matrix of random numbers chosen from the Rayleigh distribution with parameter \(B\), where scalars \(m\) and \(n\) are the row and column dimensions of R.
```

r = raylrnd(1:5)
r =
1.7986 0.8795 3.3473 8.9159 3.5182

```
random | raylpdf | raylcdf | raylinv | raylstat | raylfit
- "Rayleigh Distribution" on page B-91

\section*{raylstat}

Purpose Rayleigh mean and variance

\section*{Syntax \(\quad[M, V]=\operatorname{raylstat}(B)\)}

Description \([M, V]=\) raylstat \((B)\) returns the mean of and variance for the Rayleigh distribution with scale parameter B.

The mean of the Rayleigh distribution with parameter \(b\) is \(b \sqrt{\pi / 2}\) and the variance is
\[
\frac{4-\pi}{2} b^{2}
\]

\section*{Examples}
\[
\begin{aligned}
& {[\mathrm{mn}, \mathrm{v}]=\text { raylstat }(1)} \\
& \mathrm{mn}= \\
& 1.2533 \\
& \mathrm{v}= \\
& 0.4292
\end{aligned}
\]

See Also raylpdf | raylcdf | raylinv | raylfit | raylrnd
How To . "Rayleigh Distribution" on page B-91

Purpose
Residual case order plot

\section*{Syntax}

Description

Examples
```

rcoplot(r,rint)

```
rcoplot(r, rint) displays an errorbar plot of the confidence intervals on the residuals from a regression. The residuals appear in the plot in case order. Inputs \(r\) and rint are outputs from the regress function.

The following plots residuals and prediction intervals from a regression of a linearly additive model to the data in moore. mat:
```

    load moore
    X = [ones(size(moore,1),1) moore(:,1:5)];
    y = moore(:,6);
    alpha = 0.05;
    [betahat,Ibeta,res,Ires,stats] = regress(y,X,alpha);
    rcoplot(res,Ires)
    ```


The interval around the first residual, shown in red, does not contain zero. This indicates that the residual is larger than expected in \(95 \%\) of new observations, and suggests the data point is an outlier.

\section*{See Also \\ regress}

\section*{Purpose Add reference curve to plot}
```

Syntax refcurve(p)
refcurve
hcurve = refcurve(...)

```

\section*{Description}
refcurve ( \(p\) ) adds a polynomial reference curve with coefficients \(p\) to the current axes. If \(p\) is a vector with \(n+1\) elements, the curve is:
\[
y=p(1) * x^{\wedge} n+p(2) * x^{\wedge}(n-1)+\ldots+p(n) * x+p(n+1)
\]
refcurve with no input arguments adds a line along the \(x\) axis.
hcurve \(=\) refcurve(...) returns the handle hcurve to the curve.

\section*{Examples Example 1}

Plot data from a population with a polynomial trend and use refcurve to add both the population and fitted mean functions:
```

p = [1 -2 -1 0];
t = 0:0.1:3;
y = polyval(p,t) + 0.5*randn(size(t));
plot(t,y,'ro')
h = refcurve(p);
set(h,'Color','r')
q = polyfit(t,y,3);
refcurve(q)
legend('Data','Population Mean','Fitted Mean',...
'Location','NW')

```


\section*{Example 2}

Plot trajectories of a batted baseball, with and without air resistance.
Relevant physical constants are:
```

M = 0.145; % Mass (kg)
R = 0.0366; % Radius (m)
A = pi*R^2; %Area (m^2)
rho = 1.2; % Density of air (kg/m^3)
C = 0.5; % Drag coefficient

```
```

D = rho*C*A/2;
% Drag proportional to the square of the speed
g = 9.8; %Acceleration due to gravity (m/s^2)

```

First, simulate the trajectory with drag proportional to the square of the speed, assuming constant acceleration in each time interval:
```

dt = 1e-2; % Simulation time interval (s)
rO = [0 1]; % Initial position (m)
s0 = 50; % Initial speed (m/s)
alpha0 = 35; % Initial angle (deg)
v0=s0*[cosd(alpha0) sind(alphaO)]; % Initial velocity (m/s)
r = r0;
v = v0;
trajectory = r0;
while r(2) > 0
a = [0 -g]-(D/M)*norm(v)*v;
v = v + a*dt;
r = r + v*dt + (1/2)*a*(dt^2);
trajectory = [trajectory;r];
end

```

Second, use refcurve to add the drag-free parabolic trajectory (found analytically) to a plot of trajectory:
```

plot(trajectory(:,1),trajectory(:,2),'m','LineWidth',2)
xlim([0,250])
h = refcurve([-g/(2*v0(1)^2),...
(g*r0(1)/v0(1)^2)+(v0(2)/v0(1)),···
(-g*r0(1)^2/(2*v0(1)^2))-(v0(2)*r0(1)/v0(1))+r0(2)]);
set(h,'Color','c','LineWidth',2)
axis equal
ylim([0,50])
grid on
xlabel('Distance (m)')
ylabel('Height (m)')
title('{\bf Baseball Trajectories}')

```

\section*{refcurve}

> legend('With Drag','Without Drag')

Baseball Trajectories


See Also
refline | lsline | gline | polyfit

\section*{Purpose Add reference line to plot}
```

Syntax refline(m,b)
refline(coeffs)
refline
hline = refline(...)

```

Description refline ( \(m, b\) ) adds a reference line with slope \(m\) and intercept \(b\) to the current axes.
refline(coeffs), where coeffs is a two-element coefficient vector, adds the line
\[
y=\operatorname{coeffs}(1) * x+\operatorname{coeffs}(2)
\]
to the figure.
refline with no input arguments is equivalent to lsline.
hline = refline(...) returns the handle hline to the line.

\section*{Examples \\ Add a reference line at the mean of a data scatter and its least-squares} line:
```

x = 1:10;
y = x + randn(1,10);
scatter(x,y,25,'b','*')
lsline
mu = mean(y);
hline = refline([0 mu]);
set(hline,'Color','r')

```

refcurve | lsline | gline
Purpose Multiple linear regression
Syntax

b = regress ( \(\mathrm{y}, \mathrm{X}\) )
[b,bint] = regress(y, X)
[b,bint,r] = regress(y,X)
[b,bint,r,rint] = regress(y,X)
[b,bint, \(r\),rint,stats] \(=\) regress \((y, X)\)
\([\ldots]=\operatorname{regress}(y, x, a l p h a)\)

\section*{Description}
\(\mathrm{b}=\) regress \((\mathrm{y}, \mathrm{X})\) returns a \(p\)-by- 1 vector b of coefficient estimates for a multilinear regression of the responses in y on the predictors in \(\mathrm{X} . \mathrm{X}\) is an \(n\)-by- \(p\) matrix of \(p\) predictors at each of \(n\) observations. y is an \(n\)-by- 1 vector of observed responses.
regress treats \(N a N s\) in \(X\) or \(y\) as missing values, and ignores them.
If the columns of \(X\) are linearly dependent, regress obtains a basic solution by setting the maximum number of elements of \(b\) to zero.
[b, bint] \(=\) regress \((y, X)\) returns a \(p\)-by- 2 matrix bint of \(95 \%\) confidence intervals for the coefficient estimates. The first column of bint contains lower confidence bounds for each of the \(p\) coefficient estimates; the second column contains upper confidence bounds.
If the columns of \(X\) are linearly dependent, regress returns zeros in elements of bint corresponding to the zero elements of \(b\).
[ \(b\), bint, \(r\) ] \(=\) regress \((y, x)\) returns an \(n\)-by- 1 vector \(r\) of residuals.
[b, bint, \(r, r i n t]=r e g r e s s(y, X)\) returns an \(n\)-by- 2 matrix rint of intervals that can be used to diagnose outliers. If the interval rint (i,: ) for observation i does not contain zero, the corresponding residual is larger than expected in \(95 \%\) of new observations, suggesting an outlier.
In a linear model, observed values of y are random variables, and so are their residuals. Residuals have normal distributions with zero mean but with different variances at different values of the predictors. To put residuals on a comparable scale, they are "Studentized," that is, they are divided by an estimate of their standard deviation that is independent of their value. Studentized residuals have \(t\) distributions with known
degrees of freedom. The intervals returned in rint are shifts of the \(95 \%\) confidence intervals of these \(t\) distributions, centered at the residuals.
[b,bint,r,rint,stats] = regress ( \(y, X\) ) returns a 1-by-4 vector stats that contains, in order, the \(R^{2}\) statistic, the \(F\) statistic and its \(p\) value, and an estimate of the error variance.

Note When computing statistics, X should include a column of 1 s so that the model contains a constant term. The \(F\) statistic and its \(p\) value are computed under this assumption, and they are not correct for models without a constant. The \(R^{2}\) statistic can be negative for models without a constant, indicating that the model is not appropriate for the data.
[...] = regress(y, X, alpha) uses a 100*(1-alpha)\% confidence level to compute bint and rint.

\section*{Examples}

Load data on cars; identify weight and horsepower as predictors, mileage as the response:
```

load carsmall
x1 = Weight;
x2 = Horsepower; % Contains NaN data
y = MPG;

```

Compute regression coefficients for a linear model with an interaction term:
```

X = [ones(size(x1)) x1 x2 x1.*x2];
b = regress(y,X) % Removes NaN data
b =
60.7104
-0.0102
-0.1882
0.0000

```

Plot the data and the model:
```

scatter3(x1,x2,y,'filled')
hold on
x1fit = min(x1):100:max(x1);
x2fit = min(x2):10:max(x2);
[X1FIT,X2FIT] = meshgrid(x1fit,x2fit);
YFIT = b(1) + b(2)*X1FIT + b(3)*X2FIT + b(4)*X1FIT.*X2FIT;
mesh(X1FIT,X2FIT,YFIT)
xlabel('Weight')
ylabel('Horsepower')
zlabel('MPG')
view(50,10)

```


Weight
Horsepower

\title{
References
}

See Also regstats | mvregress | robustfit | stepwisefit | rcoplot

\section*{RegressionBaggedEnsemble}
Superclasses RegressionEnsemble
Purpose Regression ensemble grown by resampling
Description RegressionBaggedEnsemble combines a set of trained weak learnermodels and data on which these learners were trained. It can predictensemble response for new data by aggregating predictions from itsweak learners.
Construction
ens =
fitensemble(X,Y, 'bag', nlearn, learners, 'type', 'regression')creates a bagged regression ensemble. For more information on thesyntax, see the fitensemble function reference page.
Properties
CategoricalPredictors
List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of \(X\).

\section*{CombineWeights}
A string describing how the ensemble combines learner predictions.

\section*{FitInfo}
A numeric array of fit information. The FitInfoDescription property describes the content of this array.
FitInfoDescription
String describing the meaning of the FitInfo array.

\section*{FResample}
A numeric scalar between 0 and 1. FResample is the fraction of training data fitensemble resampled at random for every weak learner when constructing the ensemble.

\section*{LearnerNames}

\section*{RegressionBaggedEnsemble}

Cell array of strings with names of the weak learners in the ensemble. The name of each learner appears just once. For example, if you have an ensemble of 100 trees, LearnerNames is \{'Tree'\}.
Method
A string with the name of the algorithm fitensemble used for training the ensemble.

\section*{ModelParams}

Parameters used in training ens.

\section*{NObservations}

Numeric scalar containing the number of observations in the training data.

NTrained
Number of trained learners in the ensemble, a positive scalar.

\section*{PredictorNames}

A cell array of names for the predictor variables, in the order in which they appear in \(X\).

\section*{ReasonForTermination}

A string describing the reason fitensemble stopped adding weak learners to the ensemble.

\section*{Regularization}

A structure containing the result of the regularize method. Use Regularization with shrink to lower resubstitution error and shrink the ensemble.

\section*{Replace}

Boolean flag indicating if training data for weak learners in this ensemble were sampled with replacement. Replace is true for sampling with replacement, false otherwise.

ResponseName
A string with the name of the response variable \(Y\).

\section*{ResponseTransform}

Function handle for transforming scores, or string representing a built-in transformation function. ' none' means no transformation; equivalently, ' none' means @(x)x.

Add or change a ResponseTransform function by dot addressing:
```

ens.ResponseTransform = @function

```

Trained
The trained learners, a cell array of compact regression models.

\section*{TrainedWeights}

A numeric vector of weights the ensemble assigns to its learners. The ensemble computes predicted response by aggregating weighted predictions from its learners.

\section*{UseObsForLearner}

A logical matrix of size \(N\)-by-NTrained, where \(N\) is the number of rows (observations) in the training data X , and NTrained is the number of trained weak learners. UseObsForLearner (I , J) is true if observation I was used for training learner \(J\), and is false otherwise.
w
The scaled weights, a vector with length \(n\), the number of rows in \(x\). The sum of the elements of \(W\) is 1 .

X
The matrix of predictor values that trained the ensemble. Each column of \(X\) represents one variable, and each row represents one observation.

\section*{Y}

\section*{RegressionBaggedEnsemble}

The numeric column vector with the same number of rows as \(X\) that trained the ensemble. Each entry in \(Y\) is the response to the data in the corresponding row of X .

\section*{Methods}

\author{
oobLoss \\ oobPredict
}

\section*{Inherited Methods}
\begin{tabular}{ll} 
compact & \begin{tabular}{l} 
Create compact regression \\
ensemble
\end{tabular} \\
crossval & \begin{tabular}{l} 
Cross validate ensemble
\end{tabular} \\
crshrink & \begin{tabular}{l} 
Cross validate shrinking \\
(pruning) ensemble
\end{tabular} \\
regularize & \begin{tabular}{l} 
Find weights to minimize \\
resubstitution error plus penalty \\
term
\end{tabular} \\
resubLoss & \begin{tabular}{l} 
Regression error by resubstitution \\
Predict response of ensemble by \\
resubstitution
\end{tabular} \\
resubPredict & Resume training ensemble \\
shrink & Prune ensemble
\end{tabular}\(\quad\)\begin{tabular}{l} 
Regression error \\
loss \\
predict \\
predictorImportance
\end{tabular}

Out-of-bag regression error
Predict out-of-bag response of ensemble

Find weights to minimize resubstitution error plus penalty Regression error by resubstitution Predict response of ensemble by Resume training ensemble Prune ensemble Estimates of predictor importance
\(\begin{array}{ll}\text { Copy } & \text { Value. To learn how value classes affect copy operations, see Copying } \\ \text { Semantics } & \text { Objects in the MATLAB Programming Fundamentals documentation. }\end{array}\)

Examples
Create a bagged regression ensemble to predict the mileage of cars in the carsmall data set based on their engine displacement, horsepower, and weight:
```

load carsmall
X = [Displacement Horsepower Weight];
ens = fitensemble(X,MPG,'bag',100,'Tree',...
'type','regression')
ens =
classreg.learning.regr.RegressionBaggedEnsemble:
PredictorNames: {'x1' 'x2' 'x3'}
CategoricalPredictors: []
ResponseName: 'Y'
ResponseTransform: 'none'
NObservations: 94
NTrained: 100
Method: 'Bag'
LearnerNames: {'Tree'}
ReasonForTermination: [1x77 char]
FitInfo: []
FitInfoDescription: 'None'
Regularization: []
FResample: 1
Replace: 1
UseObsForLearner: [94x100 logical]

```

Predict the mileage of a car whose characteristics are the average of those of the first 10 cars:
```

car10 = mean(X(1:10,:));
predict(ens,car10)
ans =
14.6569

```

RegressionEnsemble | fitensemble

\section*{RegressionBaggedEnsemble}

\author{
How To \\ - Chapter 13, "Nonparametric Supervised Learning"
}

\section*{RegressionEnsemble}

\section*{Superclasses CompactRegressionEnsemble}

\section*{Purpose Ensemble regression}

Description RegressionEnsemble combines a set of trained weak learner models and data on which these learners were trained. It can predict ensemble response for new data by aggregating predictions from its weak learners.

\section*{Construction}
ens = fitensemble(X,Y,method, nlearn,learners) returns an ensemble model that can predict responses to data. The ensemble consists of models listed in learners. For more information on the syntax, see the fitensemble function reference page.
ens = fitensemble(X,Y,method, nlearn, learners,Name,Value) returns an ensemble model with additional options specified by one or more Name, Value pair arguments. For more information on the syntax, see the fitensemble function reference page.

\section*{Properties}

\section*{CategoricalPredictors}

List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of \(X\).

CombineWeights
A string describing how the ensemble combines learner predictions.

\section*{FitInfo}

A numeric array of fit information. The FitInfoDescription property describes the content of this array.

\section*{FitInfoDescription}

String describing the meaning of the FitInfo array.

\footnotetext{
LearnerNames
}

\section*{RegressionEnsemble}

Cell array of strings with names of the weak learners in the ensemble. The name of each learner appears just once. For example, if you have an ensemble of 100 trees, LearnerNames is \{'Tree'\}.
Method
A string with the name of the algorithm fitensemble used for training the ensemble.

\section*{ModelParams}

Parameters used in training ens.

\section*{NObservations}

Numeric scalar containing the number of observations in the training data.

NTrained
Number of trained learners in the ensemble, a positive scalar.

\section*{PredictorNames}

A cell array of names for the predictor variables, in the order in which they appear in \(X\).

\section*{ReasonForTermination}

A string describing the reason fitensemble stopped adding weak learners to the ensemble.

\section*{Regularization}

A structure containing the result of the regularize method. Use Regularization with shrink to lower resubstitution error and shrink the ensemble.

\section*{ResponseName}

A string with the name of the response variable Y .
ResponseTransform

Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none ' means @(x)x.

Add or change a ResponseTransform function by dot addressing:
```

ens.ResponseTransform = @function

```

\section*{Trained}

The trained learners, a cell array of compact regression models.

\section*{TrainedWeights}

A numeric vector of weights the ensemble assigns to its learners. The ensemble computes predicted response by aggregating weighted predictions from its learners.

The scaled weights, a vector with length n , the number of rows in \(X\). The sum of the elements of \(W\) is 1 .

X
The matrix of predictor values that trained the ensemble. Each column of \(X\) represents one variable, and each row represents one observation.

\section*{Y}

The numeric column vector with the same number of rows as \(X\) that trained the ensemble. Each entry in \(Y\) is the response to the data in the corresponding row of X .

\section*{Methods}
\begin{tabular}{ll} 
compact & \begin{tabular}{l} 
Create compact regression \\
ensemble
\end{tabular} \\
crossval & Cross validate ensemble \\
cvshrink & \begin{tabular}{l} 
Cross validate shrinking \\
(pruning) ensemble
\end{tabular}
\end{tabular}

\section*{RegressionEnsemble}
\begin{tabular}{ll} 
regularize & \begin{tabular}{l} 
Find weights to minimize \\
resubstitution error plus penalty \\
term
\end{tabular} \\
resubLoss & Regression error by resubstitution \\
resubPredict & \begin{tabular}{l} 
Predict response of ensemble by \\
resubstitution
\end{tabular} \\
resume & Resume training ensemble \\
shrink & Prune ensemble
\end{tabular}

\section*{Inherited Methods}
```

loss
predict
predictorImportance

```

Regression error
Predict response of ensemble
Estimates of predictor importance

Copy
Semantics

\section*{Examples}

Value. To learn how value classes affect copy operations, see Copying Objects in the MATLAB Programming Fundamentals documentation.

Create a boosted regression ensemble to predict the mileage of cars in the carsmall data set based on their weights and numbers of cylinders:
```

load carsmall
learner = RegressionTree.template('MinParent',20);
ens = fitensemble([Weight, Cylinders],MPG,...
'LSBoost',100,learner,'PredictorNames',{'W','C'},...
'categoricalpredictors',2)
ens =
classreg.learning.regr.RegressionEnsemble:
PredictorNames: {'W' 'C'}
CategoricalPredictors: 2
ResponseName: 'Response'
ResponseTransform: 'none'

```
```

NObservations: 94
NTrained: 100
Method: 'LSBoost'
LearnerNames: \{'Tree'\}
ReasonForTermination: [1x77 char]
FitInfo: [100x1 double]
FitInfoDescription: [2x83 char]
Regularization: []

```

Predict the mileage of 4,000-pound cars with 4,6 , and 8 cylinders:
```

mileage4K = predict(ens,[4000 4; 4000 6; 4000 8])
mileage4K =
20.0294
19.4206
15.5000

```
See Also ClassificationEnsemble | fitensemble |
How To
- Chapter 13, "Nonparametric Supervised Learning"

\section*{RegressionPartitionedEnsemble}
\(\begin{array}{ll}\text { Superclasses } & \text { RegressionPartitionedModel } \\ \text { Purpose } & \text { Cross-validated regression ensemble }\end{array}\)
Description RegressionPartitionedEnsemble is a set of regression ensembles trained on cross-validated folds. Estimate the quality of classification by cross validation using one or more "kfold" methods: kfoldfun, kfoldLoss, or kfoldPredict. Every "kfold" method uses models trained on in-fold observations to predict response for out-of-fold observations. For example, suppose you cross validate using five folds. In this case, every training fold contains roughly \(4 / 5\) of the data and every test fold contains roughly \(1 / 5\) of the data. The first model stored in Trained\{1\} was trained on \(X\) and \(Y\) with the first \(1 / 5\) excluded, the second model stored in Trained \(\{2\}\) was trained on \(X\) and \(Y\) with the second \(1 / 5\) excluded, and so on. When you call kfoldPredict, it computes predictions for the first \(1 / 5\) of the data using the first model, for the second \(1 / 5\) of data using the second model and so on. In short, response for every observation is computed by kfoldPredict using the model trained without this observation.

\section*{Construction}
cvens \(=\) crossval(ens) creates a cross-validated ensemble from ens, a regression ensemble. For syntax details, see the crossval method reference page.
cvens = fitensemble(X,Y,method, nlearn, learners, name, value) creates a cross-validated ensemble when name is one of 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. For syntax details, see the fitensemble function reference page.

\section*{Input Arguments}

\section*{ens}

A regression ensemble constructed with fitensemble.

\section*{RegressionPartitionedEnsemble}

\section*{Properties}

\section*{CategoricalPredictors}

List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of \(X\).

\section*{CrossValidatedModel}

Name of the cross-validated model, a string.

\section*{Kfold}

Number of folds used in a cross-validated tree, a positive integer.

\section*{ModelParams}

Object holding parameters of tree.

\section*{NObservations}

Numeric scalar containing the number of observations in the training data.

\section*{NTrainedPerFold}

Vector of Kfold elements. Each entry contains the number of trained learners in this cross-validation fold.

\section*{Partition}

The partition of class cvpartition used in creating the cross-validated ensemble.

\section*{PredictorNames}

A cell array of names for the predictor variables, in the order in which they appear in X .

\section*{ResponseName}

Name of the response variable Y , a string.

\section*{ResponseTransform}

Function handle for transforming scores, or string representing a built-in transformation function. 'none' means no transformation; equivalently, 'none' means @(x)x.

\section*{RegressionPartitionedEnsemble}

Add or change a ResponseTransform function by dot addressing:
```

ens.ResponseTransform = @function

```

\section*{Trainable}

Cell array of ensembles trained on cross-validation folds. Every ensemble is full, meaning it contains its training data and weights. Trained

Cell array of compact ensembles trained on cross-validation folds. w

The scaled weights, a vector with length n , the number of rows in \(X\).

X
A matrix of predictor values. Each column of X represents one variable, and each row represents one observation.

\section*{Y}

A numeric column vector with the same number of rows as \(X\). Each entry in Y is the response to the data in the corresponding row of X .

\section*{Methods}
\begin{tabular}{ll} 
kfoldLoss & \begin{tabular}{l} 
Cross-validation loss of \\
partitioned regression ensemble
\end{tabular} \\
resume & Resume training ensemble
\end{tabular}

\section*{Inherited Methods}
\begin{tabular}{ll} 
kfoldfun & Cross validate function \\
kfoldLoss & \begin{tabular}{l} 
Cross-validation loss of \\
partitioned regression model
\end{tabular} \\
kfoldPredict & \begin{tabular}{l} 
Predict response for observations \\
not used for training.
\end{tabular}
\end{tabular}

\section*{RegressionPartitionedEnsemble}
\begin{tabular}{|c|c|}
\hline Copy Semantics & Value. To learn how value classes affect copy operations, see Copying Objects in the MATLAB Programming Fundamentals documentation. \\
\hline \multirow[t]{13}{*}{Examples} & Construct a partitioned regression ensemble, and examine the cross-validation losses for the folds: \\
\hline & ```
load carsmall
XX = [Cylinders Displacement Horsepower Weight];
YY = MPG;
rens = fitensemble(XX,YY,'LSBoost',100,'Tree');
cvrens = crossval(rens);
L = kfoldLoss(cvrens,'mode','individual')
``` \\
\hline & L = \\
\hline & 42.4468 \\
\hline & 12.3158 \\
\hline & 65.9432 \\
\hline & 39.0019 \\
\hline & 30.5908 \\
\hline & 16.6225 \\
\hline & 17.3071 \\
\hline & 46.1769 \\
\hline & 8.0561 \\
\hline & 12.9689 \\
\hline \multirow[t]{2}{*}{See Also} & ClassificationPartitionedEnsemble | \\
\hline & RegressionPartitionedModel | RegressionEnsemble \\
\hline How To & - Chapter 13, "Nonparametric Supervised Learning" \\
\hline
\end{tabular}

\section*{RegressionPartitionedModel}

\section*{Purpose Cross-validated regression model}

Description

\section*{Construction}
cvmodel \(=\) crossval(tree) creates a cross-validated classification model from a regression tree. For syntax details, see the crossval method reference page.
cvmodel = RegressionTree.fit(X,Y, name, value) creates a cross-validated model when name is one of 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'. For syntax details, see the RegressionTree.fit function reference page.

\section*{Input Arguments}

\section*{tree}

A regression tree constructed with RegressionTree.fit.

\section*{Properties}

\section*{CategoricalPredictors}

RegressionPartitionedModel is a set of regression models trained on cross-validated folds. Estimate the quality of regression by cross validation using one or more "kfold" methods: kfoldPredict, kfoldLoss, and kfoldfun. Every "kfold" method uses models trained on in-fold observations to predict response for out-of-fold observations. For example, suppose you cross validate using five folds. In this case, every training fold contains roughly \(4 / 5\) of the data and every test fold contains roughly \(1 / 5\) of the data. The first model stored in Trained\{1\} was trained on \(X\) and \(Y\) with the first \(1 / 5\) excluded, the second model stored in Trained \(\{2\}\) was trained on \(X\) and \(Y\) with the second \(1 / 5\) excluded, and so on. When you call kfoldPredict, it computes predictions for the first \(1 / 5\) of the data using the first model, for the second \(1 / 5\) of data using the second model and so on. In short, response for every observation is computed by kfoldPredict using the model trained without this observation.

List of categorical predictors. CategoricalPredictors is a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of X .

\section*{RegressionPartitionedModel}

\section*{CrossValidatedModel}

Name of the cross-validated model, a string.

\section*{Kfold}

Number of folds used in a cross-validated tree, a positive integer.

\section*{ModelParams}

Object holding parameters of tree.

\section*{Partition}

The partition of class cupartition used in the cross-validated model.

\section*{PredictorNames}

A cell array of names for the predictor variables, in the order in which they appear in X .

\section*{ResponseName}

Name of the response variable \(Y\), a string.

\section*{ResponseTransform}

Function handle for transforming the raw response values (mean squared error). The function handle should accept a matrix of response values and return a matrix of the same size. The default string ' none' means @(x)x, or no transformation.

Add or change a ResponseTransform function by dot addressing:
```

ctree.ResponseTransform = @function

```

\section*{Trained}

The trained learners, a cell array of compact regression models.
W
The scaled weights, a vector with length n , the number of rows in \(X\).

\section*{RegressionPartitionedModel}

X
A matrix of predictor values. Each column of X represents one variable, and each row represents one observation.

\section*{Y}

A numeric column vector with the same number of rows as \(X\). Each entry in \(Y\) is the response to the data in the corresponding row of \(X\).

\section*{Methods}
\begin{tabular}{ll} 
kfoldfun & Cross validate function \\
kfoldLoss & \begin{tabular}{l} 
Cross-validation loss of \\
partitioned regression model
\end{tabular} \\
kfoldPredict & \begin{tabular}{l} 
Predict response for observations \\
not used for training.
\end{tabular}
\end{tabular}

Copy
Semantics
Examples
Evaluate the cross-validation error of the carsmall data using Horsepower and Weight as predictor variables for mileage (MPG):
load carsmall
X = [Horsepower Weight];
tree \(=\) RegressionTree.fit(X,MPG);
cvtree = crossval(tree);
L = kfoldLoss(cvtree)
L =
26.4414

See Also RegressionPartitionedEnsemble | ClassificationPartitionedModel

\section*{RegressionPartitionedModel}

How To • Chapter 13, "Nonparametric Supervised Learning"

\section*{RegressionTree}

\section*{Superclasses CompactRegressionTree}

\section*{Purpose Regression tree}

Description A decision tree with binary splits for regression. An object of class RegressionTree can predict responses for new data with the predict method. The object contains the data used for training, so can compute resubstitution predictions.

\section*{Construction}
tree \(=\) RegressionTree.fit \((X, Y)\) returns a regression tree based on the input variables (also known as predictors, features, or attributes) \(X\) and output (response) \(Y\). tree is a binary tree where each branching node is split based on the values of a column of \(X\).
tree = RegressionTree.fit(X,Y,Name, Value) fits a tree with additional options specified by one or more Name, Value pair arguments.

\section*{Input Arguments}

\section*{X}

A matrix of predictor values. Each column of X represents one variable, and each row represents one observation.

RegressionTree.fit considers NaN values in X as missing values. RegressionTree.fit does not use observations with all missing values for \(X\) the fit. RegressionTree.fit uses observations with some missing values for X to find splits on variables for which these observations have valid values.

Y
A numeric column vector with the same number of rows as \(X\). Each entry in \(Y\) is the response to the data in the corresponding row of \(X\).

RegressionTree.fit considers NaN values in \(Y\) to be missing values. RegressionTree.fit does not use observations with missing values for \(Y\) in the fit.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

\section*{CategoricalPredictors}

List of categorical predictors. Pass CategoricalPredictors as one of:
- A numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of \(X\).
- A logical vector of length \(p\), where a true entry means that the corresponding column of X is a categorical variable.
- 'all', meaning all predictors are categorical.
- A cell array of strings, where each element in the array is the name of a predictor variable. The names must match entries in the PredictorNames property.
- A character matrix, where each row of the matrix is a name of a predictor variable. Pad the names with extra blanks so each row of the character matrix has the same length.

\section*{Default: []}

\section*{crossval}

If 'on', grows a cross-validated decision tree with 10 folds. You can use 'kfold', 'holdout', 'leaveout', or 'cvpartition' parameters to override this cross-validation setting. You can only use one of these four parameters ('kfold', 'holdout', 'leaveout', or 'cvpartition') at a time when creating a cross-validated tree.

Alternatively, cross-validate tree later using the crossval method.

\section*{RegressionTree}

Default: 'off'

\section*{cvpartition}

A partition created with cvpartition to use in cross-validated tree. You can only use one of these four options at a time: 'kfold', 'holdout', 'leaveout' and 'cvpartition'.
holdout
Holdout validation tests the specified fraction of the data, and uses the rest of the data for training. Specify a numeric scalar from 0 to 1 . You can only use one of these four options at a time for creating a cross-validated tree: 'kfold', 'holdout', 'leaveout' and 'cvpartition'.
kfold
Number of folds to use in a cross-validated tree, a positive integer.
You can only use one of these four options at a time: 'kfold', 'holdout', 'leaveout', or 'cvpartition'.

Default: 10

\section*{leaveout}

Use leave-one-out cross validation by setting to 'on '. You can only use one of these four options at a time: 'kfold', 'holdout ', 'leaveout', or 'cvpartition'.

\section*{MergeLeaves}

When 'on', RegressionTree merges leaves that originate from the same parent node, and that give a sum of risk values greater or equal to the risk associated with the parent node. When 'off', RegressionTree does not merge leaves.

Default: 'on'
MinLeaf

Each leaf has at least MinLeaf observations per tree leaf. If you supply both MinParent and MinLeaf, RegressionTree uses the setting that gives larger leaves: MinParent=max(MinParent, 2*MinLeaf).

Default: 1

\section*{MinParent}

Each branch node in the tree has at least MinParent observations. If you supply both MinParent and MinLeaf, RegressionTree uses the setting that gives larger leaves: MinParent=max(MinParent,2*MinLeaf).

Default: 10

\section*{NVarToSample}

Number of predictors to select at random for each split. Can be a positive integer or 'all', which means use all available predictors.

Default: 'all'

\section*{PredictorNames}

A cell array of names for the predictor variables, in the order in which they appear in X .

Default: \{'x1','x2',...\}
Prune
When 'on', RegressionTree computes the full tree and the optimal sequence of pruned subtrees. When 'off' RegressionTree computes the full tree without pruning.

Default: 'on'

\section*{RegressionTree}

PruneCriterion
String with the pruning criterion, always 'error'.
Default: 'error'
ResponseName
Name of the response variable Y , a string.
Default: ' \(Y\) '
ResponseTransform
Function handle for transforming the raw response values (mean squared error). The function handle should accept a matrix of response values and return a matrix of the same size. The default string 'none' means @(x)x, or no transformation.

Add or change a ResponseTransform function by dot addressing:
```

tree.ResponseTransform = @function

```

Default: 'none'

\section*{SplitCriterion}

Criterion for choosing a split, always the string 'MSE', meaning mean squared error.

Default: 'MSE'
Weights
Vector of observation weights. The length of weights is the number of rows in \(X\).

Default: ones(size(X,1),1)

\section*{RegressionTree}

\section*{Properties}

\section*{CategoricalPredictors}

List of categorical predictors, a numeric vector with indices from 1 to \(p\), where \(p\) is the number of columns of \(X\).

\section*{CatSplit}

An \(n\)-by- 2 cell array, where n is the number of nodes in tree. Each row in CatSplit gives left and right values for a categorical split. For each branch node \(j\) based on a categorical predictor variable \(z\), the left child is chosen if \(z\) is in CatSplit ( \(j, 1\) ) and the right child is chosen if \(z\) is in CatSplit ( \(\mathrm{j}, 2\) ). The splits are in the same order as nodes of the tree. Nodes for these splits can be found by running cuttype and selecting 'categorical' cuts from top to bottom.

\section*{Children}

An \(n\)-by- 2 array containing the numbers of the child nodes for each node in tree, where \(n\) is the number of nodes. Leaf nodes have child node 0 .

\section*{CutCategories}

An \(n\)-by- 2 cell array of the categories used at branches in tree, where \(n\) is the number of nodes. For each branch node i based on a categorical predictor variable \(x\), the left child is chosen if \(x\) is among the categories listed in CutCategories \(\{\mathrm{i}, 1\}\), and the right child is chosen if \(x\) is among those listed in CutCategories \(\{i, 2\}\). Both columns of CutCategories are empty for branch nodes based on continuous predictors and for leaf nodes.

CutVar contains the cut points for 'continuous' cuts, and CutCategories contains the set of categories.

\section*{CutPoint}

An \(n\)-element vector of the values used as cut points in tree, where \(n\) is the number of nodes. For each branch node i based on a continuous predictor variable \(x\), the left child is chosen if \(x<\) CutPoint (i) and the right child is chosen if \(x>=\) CutPoint (i).

\section*{RegressionTree}

CutPoint is NaN for branch nodes based on categorical predictors and for leaf nodes.

\section*{CutType}

An \(n\)-element cell array indicating the type of cut at each node in tree, where \(n\) is the number of nodes. For each node i, CutType\{i\} is:
- 'continuous' - If the cut is defined in the form \(x<v\) for a variable x and cut point v .
- 'categorical' - If the cut is defined by whether a variable \(x\) takes a value in a set of categories.
- ' ' - If i is a leaf node.

CutVar contains the cut points for 'continuous ' cuts, and CutCategories contains the set of categories.

\section*{CutVar}

An \(n\)-element cell array of the names of the variables used for branching in each node in tree, where \(n\) is the number of nodes. These variables are sometimes known as cut variables. For leaf nodes, CutVar contains an empty string.

CutVar contains the cut points for 'continuous' cuts, and CutCategories contains the set of categories.

\section*{IsBranch}

An \(n\)-element logical vector ib that is true for each branch node and false for each leaf node of tree.

\section*{ModelParams}

Object holding parameters of tree.
NObservations

Number of observations in the training data, a numeric scalar. NObservations can be less than the number of rows of input data \(X\) when there are missing values in \(X\) or response \(Y\).

\section*{NodeErr}

An \(n\)-element vector e of the errors of the nodes in tree, where \(n\) is the number of nodes. \(\mathrm{e}(\mathrm{i})\) is the misclassification probability for node i.

\section*{NodeMean}

An \(n\)-element numeric array with mean values in each node of tree, where \(n\) is the number of nodes in the tree. Every element in NodeMean is the average of the true \(Y\) values over all observations in the node.

\section*{NodeProb}

An \(n\)-element vector \(p\) of the probabilities of the nodes in tree, where \(n\) is the number of nodes. The probability of a node is computed as the proportion of observations from the original data that satisfy the conditions for the node. This proportion is adjusted for any prior probabilities assigned to each class.

\section*{NodeSize}

An \(n\)-element vector sizes of the sizes of the nodes in tree, where \(n\) is the number of nodes. The size of a node is defined as the number of observations from the data used to create the tree that satisfy the conditions for the node.

\section*{NumNodes}

The number of nodes \(n\) in tree.

\section*{Parent}

An \(n\)-element vector \(p\) containing the number of the parent node for each node in tree, where \(n\) is the number of nodes. The parent of the root node is 0 .
```

PredictorNames

```

\section*{RegressionTree}

A cell array of names for the predictor variables, in the order in which they appear in \(X\).

\section*{PruneList}

An \(n\)-element numeric vector with the pruning levels in each node of tree, where \(n\) is the number of nodes.

\section*{ResponseName}

Name of the response variable Y , a string.

\section*{ResponseTransform}

Function handle for transforming the raw response values (mean squared error). The function handle should accept a matrix of response values and return a matrix of the same size. The default string 'none' means @(x)x, or no transformation.

Add or change a ResponseTransform function by dot addressing:
tree.ResponseTransform = @function
Risk
An \(n\)-element vector \(r\) of the risk of the nodes in tree, where \(n\) is the number of nodes. The risk \(r(i)\) for node \(i\) is the node error NodeErr(i) multiplied by the node probability NodeProb(i).

\section*{SurrCutCategories}

An \(n\)-element cell array of the categories used for surrogate splits in tree, where \(n\) is the number of nodes in tree. For each node k, SurrCutCategories \(\{\mathrm{k}\}\) is a cell array. The length of SurrCutCategories \(\{\mathrm{k}\}\) is equal to the number of surrogate predictors found at this node. Every element of SurrCutCategories \(\{\mathrm{k}\}\) is either an empty string for a continuous surrogate predictor, or is a two-element cell array with categories for a categorical surrogate predictor. The first element of this two-element cell array lists categories assigned to the left child by this surrogate split, and the second element of this two-element cell array lists categories assigned to the right child by this

\section*{RegressionTree}
surrogate split. The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutCategories contains an empty cell.

\section*{SurrCutFlip}

An \(n\)-element cell array of the numeric cut assignments used for surrogate splits in tree, where \(n\) is the number of nodes in tree. For each node k, SurrCutFlip \(\{\mathrm{k}\}\) is a numeric vector. The length of SurrCutFlip\{k\} is equal to the number of surrogate predictors found at this node. Every element of SurrCutFlip \(\{\mathrm{k}\}\) is either zero for a categorical surrogate predictor, or a numeric cut assignment for a continuous surrogate predictor. The numeric cut assignment can be either -1 or +1 . For every surrogate split with a numeric cut \(C\) based on a continuous predictor variable \(Z\), the left child is chosen if \(Z<C\) and the cut assignment for this surrogate split is +1 , or if \(Z \geq C\) and the cut assignment for this surrogate split is -1 . Similarly, the right child is chosen if \(Z \geq C\) and the cut assignment for this surrogate split is +1 , or if \(Z<C\) and the cut assignment for this surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutFlip contains an empty array.

\section*{SurrCutPoint}

An \(n\)-element cell array of the numeric values used for surrogate splits in tree, where \(n\) is the number of nodes in tree. For each node \(k\), SurrCutPoint \(\{k\}\) is a numeric vector. The length of SurrCutPoint \(\{k\}\) is equal to the number of surrogate predictors found at this node. Every element of SurrCutPoint \{k\} is either NaN for a categorical surrogate predictor, or a numeric cut for a continuous surrogate predictor. For every surrogate split with a numeric cut \(C\) based on a continuous predictor variable \(Z\), the left child is chosen if \(Z<C\) and SurrCutFlip for this surrogate split is -1 . Similarly, the right child is chosen if \(Z \geq C\) and SurrCutFlip for this surrogate split is +1 , or if \(Z<C\) and SurrCutFlip for this

\section*{RegressionTree}
surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables returned by SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutPoint contains an empty cell.

\section*{SurrCutType}

An \(n\)-element cell array indicating types of surrogate splits at each node in tree, where \(n\) is the number of nodes in tree. For each node \(k\), SurrCutType \(\{k\}\) is a cell array with the types of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The order of the surrogate split variables at each node is matched to the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutType contains an empty cell. A surrogate split type can be either 'continuous ' if the cut is defined in the form \(Z<V\) for a variable \(Z\) and cut point V or 'categorical' if the cut is defined by whether \(Z\) takes a value in a set of categories.

\section*{SurrCutVar}

An \(n\)-element cell array of the names of the variables used for surrogate splits in each node in tree, where \(n\) is the number of nodes in tree. Every element of SurrCutVar is a cell array with the names of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrCutVar contains an empty cell.

\section*{SurrVarAssoc}

An \(n\)-element cell array of the predictive measures of association for surrogate splits in tree, where \(n\) is the number of nodes in tree. For each node k, SurrVarAssoc \(\{k\}\) is a numeric vector. The
length of SurrVarAssoc \(\{k\}\) is equal to the number of surrogate predictors found at this node. Every element of SurrVarAssoc \{k\} gives the predictive measure of association between the optimal split and this surrogate split. The order of the surrogate split variables at each node is the order of variables in SurrCutVar. The optimal-split variable at this node does not appear. For nonbranch (leaf) nodes, SurrVarAssoc contains an empty cell.

The scaled weights, a vector with length n , the number of rows in \(X\).

A matrix of predictor values. Each column of \(X\) represents one variable, and each row represents one observation.

\section*{Y}

A numeric column vector with the same number of rows as \(X\). Each entry in \(Y\) is the response to the data in the corresponding row of \(X\).

\section*{Methods}
\begin{tabular}{ll} 
compact & Compact regression tree \\
crossval & Cross-validated decision tree \\
cvloss & \begin{tabular}{l} 
Regression error by cross \\
validation
\end{tabular} \\
fit & \begin{tabular}{l} 
Binary decision tree for regression \\
Produce sequence of subtrees by \\
pruning
\end{tabular} \\
resubLoss & \begin{tabular}{l} 
Regression error by resubstitution
\end{tabular} \\
resubPredict & \begin{tabular}{l} 
Predict resubstitution response of \\
tree
\end{tabular} \\
template & Create regression template
\end{tabular}

\section*{RegressionTree}

Inherited Methods
\begin{tabular}{ll} 
loss & Regression error \\
meanSurrVarAssoc & \begin{tabular}{l} 
Mean predictive measure of \\
association for surrogate splits in \\
decision tree
\end{tabular} \\
predict & \begin{tabular}{l} 
Predict response of regression \\
tree
\end{tabular} \\
predictorImportance & \begin{tabular}{l} 
Estimates of predictor importance \\
view
\end{tabular} \\
& View tree
\end{tabular}

Value. To learn how value classes affect copy operations, see Copying Objects in the MATLAB Programming Fundamentals documentation.

\section*{Examples Load the data in carsmall.mat, and make a regression tree to predict} the mileage of cars based on their weights and numbers of cylinders:
```

load carsmall
tree = RegressionTree.fit([Weight, Cylinders],MPG,...
'categoricalpredictors',2,'MinParent',20,...
'PredictorNames',{'W','C'})
tree =
RegressionTree:
PredictorNames: {'W' 'C'}
CategoricalPredictors: 2
ResponseName: 'Response'
ResponseTransform: 'none'
X: [94x2 double]
Y: [94x1 double]
W: [94x1 double]
ModelParams: [1x1 classreg.learning.modelparams.TreeParams]

```

\section*{RegressionTree}

Predict the mileage of 4,000 -pound cars with 4,6 , and 8 cylinders:
```

mileage4K = predict(tree,[4000 4; 4000 6; 4000 8])
mileage4K =
19.2778
19.2778
14.3889

```

\section*{See Also ClassificationTree | RegressionEnsemble | RegressionTree.fit | CompactRegressionTree | predict}

Tutorials . "Classification Trees and Regression Trees" on page 13-27
```

Purpose Regression diagnostics
Syntax regstats(y,x,model)
stats = regstats(...)
stats = regstats(y,X,model,whichstats)

```

\section*{Description}
regstats ( \(y, x\), model) performs a multilinear regression of the responses in y on the predictors in X . X is an \(n\)-by- \(p\) matrix of \(p\) predictors at each of \(n\) observations. y is an \(n\)-by- 1 vector of observed responses.

Note By default, regstats adds a first column of 1s to X , corresponding to a constant term in the model. Do not enter a column of 1s directly into X .

The optional input model controls the regression model. By default, regstats uses a linear additive model with a constant term. model can be any one of the following strings:
- 'linear' - Constant and linear terms (the default)
- 'interaction ' - Constant, linear, and interaction terms
- 'quadratic ' - Constant, linear, interaction, and squared terms
- 'purequadratic' - Constant, linear, and squared terms

Alternatively, model can be a matrix of model terms accepted by the x2fx function. See x2fx for a description of this matrix and for a description of the order in which terms appear. You can use this matrix to specify other models including ones without a constant term.

With this syntax, the function displays a graphical user interface (GUI) with a list of diagnostic statistics, as shown in the following figure.


When you select check boxes corresponding to the statistics you want to compute and click OK, regstats returns the selected statistics to the MATLAB workspace. The names of the workspace variables are displayed on the right-hand side of the interface. You can change the name of the workspace variable to any valid MATLAB variable name.
stats \(=\) regstats(...) creates the structure stats, whose fields contain all of the diagnostic statistics for the regression. This syntax does not open the GUI. The fields of stats are listed in the following table.
\begin{tabular}{l|l}
\hline Field & Description \\
\hline\(Q\) & \(Q\) from the \(Q R\) decomposition of the design matrix \\
\hline R & \(R\) from the \(Q R\) decomposition of the design matrix \\
\hline beta & Regression coefficients \\
\hline covb & Covariance of regression coefficients \\
\hline yhat & Fitted values of the response data \\
\hline\(r\) & Residuals \\
\hline mse & Mean squared error \\
\hline rsquare & \(R^{2}\) statistic \\
\hline adjrsquare & Adjusted \(R^{2}\) statistic \\
\hline leverage & Leverage \\
\hline hatmat & Hat matrix \\
\hline s2_i & Delete-1 variance \\
\hline beta_i & Delete-1 coefficients \\
\hline standres & Standardized residuals \\
\hline studres & Studentized residuals \\
\hline dfbetas & Scaled change in regression coefficients \\
\hline dffit & Change in fitted values \\
\hline
\end{tabular}
\begin{tabular}{l|l}
\hline Field & Description \\
\hline dffits & Scaled change in fitted values \\
\hline covratio & Change in covariance \\
\hline cookd & Cook's distance \\
\hline tstat & \(t\) statistics and \(p\)-values for coefficients \\
\hline fstat & \(F\) statistic and \(p\)-value \\
\hline dwstat & Durbin-Watson statistic and \(p\)-value \\
\hline
\end{tabular}

Note that the fields names of stats correspond to the names of the variables returned to the MATLAB workspace when you use the GUI. For example, stats.beta corresponds to the variable beta that is returned when you select Coefficients in the GUI and click OK.
stats \(=\) regstats ( \(\mathrm{y}, \mathrm{X}\), model, whichstats) returns only the statistics that you specify in whichstats. whichstats can be a single string such as 'leverage' or a cell array of strings such as \{'leverage' 'standres' 'studres'\}. Set whichstats to 'all' to return all of the statistics.

Note The \(F\) statistic is computed under the assumption that the model contains a constant term. It is not correct for models without a constant. The \(R^{2}\) statistic can be negative for models without a constant, which indicates that the model is not appropriate for the data.

\section*{Examples}

Open the regstats GUI using data from hald.mat:
```

load hald
regstats(heat,ingredients,'linear');

```

Select Fitted Values and Residuals in the GUI:


Click OK to export the fitted values and residuals to the MATLAB workspace in variables named yhat and \(r\), respectively.

You can create the same variables using the stats output, without opening the GUI:
```

whichstats = {'yhat','r'};
stats = regstats(heat,ingredients,'linear',whichstats);
yhat = stats.yhat;
r = stats.r;

```

\section*{References}
[1] Belsley, D. A., E. Kuh, and R. E. Welsch. Regression Diagnostics. Hoboken, NJ: John Wiley \& Sons, Inc., 1980.
[2] Chatterjee, S., and A. S. Hadi. "Influential Observations, High Leverage Points, and Outliers in Linear Regression." Statistical Science. Vol. 1, 1986, pp. 379-416.
[3] Cook, R. D., and S. Weisberg. Residuals and Influence in Regression. New York: Chapman \& Hall/CRC Press, 1983.
[4] Goodall, C. R. "Computation Using the QR Decomposition." Handbook in Statistics. Vol. 9, Amsterdam: Elsevier/North-Holland, 1993.

\footnotetext{
See Also x2fx | regress | stepwise | leverage
}

\section*{RegressionEnsemble.regularize}

\section*{Purpose}

Find weights to minimize resubstitution error plus penalty term
```

ens1 = regularize(ens)

```
ens1 = regularize(ens,Name, Value)
ens1 = regularize(ens) finds optimal weights for learners in ens by lasso regularization. regularize returns a regression ensemble identical to ens, but with a populated Regularization property.
ens1 = regularize(ens,Name,Value) computes optimal weights with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{Input \\ Arguments}
ens
A regression ensemble, created by fitensemble.

\section*{Name-Value Pair Arguments}

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
lambda
Vector of nonnegative regularization parameter values for lasso. For the default setting of lambda, regularize calculates the smallest value lambda_max for which all optimal weights for learners are 0 . The default value of lambda is a vector including 0 and nine exponentially-spaced numbers from lambda_max/1000 to lambda_max.

Default: [0
logspace \(\left(\log 10\left(\operatorname{lambda\_ max/1000)}, \log 10(1\right.\right.\) ambda_max), 9\(\left.)\right]\)
npass

\section*{RegressionEnsemble.regularize}

Maximal number of passes for lasso optimization, a positive integer.

Default: 10

\section*{reltol}

Relative tolerance on the regularized loss for lasso, a numeric positive scalar.

Default: 1e-3

\section*{verbose}

Verbosity level, either 0 or 1 . When set to 1, regularize displays more information as it runs.

Default: 0

\section*{Output} Arguments

A regression ensemble. Usually you set ens1 to the same name as ens.

\section*{Definitions}

\section*{Lasso}

The lasso algorithm finds an optimal set of learner weights \(a_{t}\) that minimize
\[
\sum_{n=1}^{N} w_{n} g\left(\left(\sum_{t=1}^{T} \alpha_{t} h_{t}\left(x_{n}\right)\right), y_{n}\right)+\lambda \sum_{t=1}^{T}\left|\alpha_{t}\right| .
\]

Here
- \(\lambda \geq 0\) is a parameter you provide, called the lasso parameter.
- \(h_{t}\) is a weak learner in the ensemble trained on \(N\) observations with predictors \(x_{n}\), responses \(y_{n}\), and weights \(w_{n}\).

\section*{RegressionEnsemble.regularize}
- \(g(f, y)=(f-y)^{2}\) is the squared error.

\section*{Examples}

Regularize an ensemble of bagged trees:
```

X = rand(2000,20);
Y = repmat(-1,2000,1);
Y(sum(X(:,1:5),2)>2.5) = 1;
bag = fitensemble(X,Y,'Bag',300,'Tree',...
'type','regression');
bag = regularize(bag,'lambda',[0.001 0.1],'verbose',1);

```
regularize reports on its progress.
To see the resulting regularization structure:
```

bag.Regularization
ans =
Method: 'Lasso'
TrainedWeights: [300x2 double]
Lambda: [1.0000e-003 0.1000]
ResubstitutionMSE: [0.0616 0.0812]
CombineWeights: @classreg.learning.combiner.WeightedSum

```

See how many learners in the regularized ensemble have positive weights (so would be included in a shrunken ensemble):
```

sum(bag.Regularization.TrainedWeights > 0)
ans =
116 91

```

To shrink the ensemble using the weights from Lambda \(=0.1\) :
```

cmp = shrink(bag,'weightcolumn',2)
cmp =

```
classreg.learning.regr.CompactRegressionEnsemble:
    PredictorNames: \{1x20 cell\}

\section*{RegressionEnsemble.regularize}

\section*{CategoricalPredictors: [] \\ ResponseName: ' \(Y\) ' \\ ResponseTransform: 'none' \\ NTrained: 91}

There are 91 members in the regularized ensemble, which is less than \(1 / 3\) of the original 300.

\section*{See Also}
shrink | cvshrink

\section*{How To}
- Chapter 13, "Nonparametric Supervised Learning"

\section*{gmdistribution.RegV property}

\author{
Purpose Value of 'Regularize' parameter \\ Description The value of the parameter 'Regularize'. \\ Note This property applies only to gmdistribution objects constructed with fit.
}
```

Purpose Importance of attributes (predictors) using ReliefF algorithm
Syntax [RANKED, WEIGHT] = relieff $(X, Y, K)$
[RANKED, WEIGHT] = relieff(X,Y,K,'PARAM1', val1,'PARAM2', val2,
...)

```

Name-Value Pair Arguments
[RANKED, WEIGHT] = relieff \((X, Y, K)\) computes ranks and weights of attributes (predictors) for input data matrix \(X\) and response vector \(Y\) using the ReliefF algorithm for classification or RReliefF for regression with K nearest neighbors. For classification, relieff uses K nearest neighbors per class. RANKED are indices of columns in \(X\) ordered by attribute importance, meaning RANKED (1) is the index of the most important predictor. WEIGHT are attribute weights ranging from -1 to 1 with large positive weights assigned to important attributes.

If \(Y\) is numeric, relieff by default performs RReliefF analysis for regression. If \(Y\) is categorical, logical, a character array, or a cell array of strings, relieff by default performs ReliefF analysis for classification.

Attribute ranks and weights computed by relieff usually depend on K. If you set K to 1 , the estimates computed by relieff can be unreliable for noisy data. If you set K to a value comparable with the number of observations (rows) in X, relieff can fail to find important attributes. You can start with \(K=10\) and investigate the stability and reliability of relieff ranks and weights for various values of \(K\).
[RANKED,WEIGHT] =
relieff(X,Y,K,'PARAM1',val1,'PARAM2',val2,...) specifies optional parameter name/value pairs.

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
method
```

Either 'regression' (default if $Y$ is numeric) or 'classification' (default if Y is not numeric).

```

\section*{prior}
```

Prior probabilities for each class, specified as a string ('empirical' or 'uniform') or as a vector (one value for each distinct group name) or as a structure $S$ with two fields:

```
- S.group containing the group names as a categorical variable, character array, or cell array of strings
- S.prob containing a vector of corresponding probabilities

If the input value is 'empirical' (default), class probabilities are determined from class frequencies in Y . If the input value is 'uniform', all class probabilities are set equal.

\section*{updates}

Number of observations to select at random for computing the weight of every attribute. By default all observations are used.
```

categoricalx

```
'on' or 'off', 'off' by default. If 'on', treat all predictors in \(X\) as categorical. If ' off', treat all predictors in \(X\) as numerical. You cannot mix numerical and categorical predictors.

\section*{sigma}

Distance scaling factor. For observation \(i\), influence on the attribute weight from its nearest neighbor \(j\) is multiplied by \(\exp \left((-\operatorname{rank}(i, j) / \text { sigma })^{\wedge} 2\right)\), where \(\operatorname{rank}(i, j)\) is the position of \(j\) in the list of nearest neighbors of \(i\) sorted by distance in the ascending order. Default is Inf (all nearest neighbors have the same influence) for classification and 50 for regression.

\section*{Examples Identify important predictors in the ionosphere dataset:}
```

load ionosphere;

```
```

[ranked,weights] = relieff(X,Y,10);
bar(weights(ranked));
xlabel('Predictor rank');
ylabel('Predictor importance weight');

```


Examine the Fisher iris data to find the important predictors:
```

load fisheriris
[ranked,weight] = relieff(meas,species,10)
ranked =

```
```

    4 3 1 2
    weight =
0.1399
0.1226
0.3590
0.3754

```

The fourth predictor is the most important, and the second predictor is the least important.

\author{
References [1] Kononenko, I., Simec, E., \& Robnik-Sikonja, M. (1997). Overcoming the myopia of inductive learning algorithms with RELIEFF. Retrieved from CiteSeerX: http://citeseerx.ist.psu.edu/viewdoc/summary?doi=10.1.1.56.4740 \\ [2] Robnik-Sikonja, M., \& Kononenko, I. (1997). An adaptation of Relief for attribute estimation in regression. Retrieved from CiteSeerX: http://citeseerx.ist.psu.edu/viewdoc/summary?doi=10.1.1.34.8381 \\ [3] Robnik-Sikonja, M., \& Kononenko, I. (2003). Theoretical and empirical analysis of ReliefF and RReliefF. Machine Learning , 53, 23-69. \\ \section*{See Also knnsearch | pdist2}
}

Purpose Reorder levels
Syntax \(\quad B=\) reorderlevels \((A\), newlevels \()\)
Description
\(B=\) reorderlevels(A, newlevels) reorders the levels of the categorical array A. newlevels is a cell array of strings or a \(2-\mathrm{D}\) character matrix that specifies the new order. newlevels must be a reordering of getlabels(A).

The order of the levels of an ordinal array has significance for relational operators, minimum and maximum, and for sorting.

\section*{Examples}

Reorder hockey standings:
```

standings = ordinal(1:3,{'Leafs','Canadiens','Bruins'});
getlabels(standings)
ans =
'Leafs' 'Canadiens' 'Bruins'
standings = reorderlevels(standings,...
{'Canadiens','Leafs','Bruins'});
getlabels(standings)
ans =
'Canadiens' 'Leafs' 'Bruins'

```
See Also addlevels | droplevels | getlabels | islevel | mergelevels

\section*{Purpose Repartition data for cross-validation}

\section*{Syntax \\ cnew = repartition(c)}
cnew = repartition(c) constructs an object cnew of the cvpartition class defining a random partition of the same type as \(c\), where \(c\) is also an object of the cvpartition class.

Repartitioning is useful for Monte-Carlo repetitions of cross-validation analyses. repartition is called by crossval when the 'mcreps' parameter is specified.

\section*{Examples}

Partition and repartition 100 observations for 3 -fold cross-validation:
```

c = cvpartition(100,'kfold',3)
c =
K-fold cross validation partition
N: 100
NumTestSets: 3
TrainSize: 67 66 67
TestSize: 33 34 33
cnew = repartition(c)
cnew =
K-fold cross validation partition
N: 100
NumTestSets: 3
TrainSize: 67 66 67
TestSize: 33 34 33

```

Check for equality of the test data in the first fold:
```

isequal(test(c,1),test(cnew,1))
ans =
0

```
See Also cvpartition

\section*{dataset.replacedata}
\begin{tabular}{ll} 
Purpose & Replace dataset variables \\
Syntax & \(B=\operatorname{replacedata}(A, X)\) \\
& \(B=\operatorname{replacedata}(A, X\), vars \()\) \\
& \(B=\) replacedata \((A, f u n)\) \\
& \(B=\operatorname{replacedata}(A, f u n\), vars \()\)
\end{tabular}

Description

\section*{Examples}

\section*{dataset.replacedata}
```

X = double(data);
X = zscore(X);
data = replacedata(data,X)
% Equivalently, provide a handle to ZSCORE.
data = replacedata(data,@zscore)
% Use ZSCORE to normalize each observation in a dataset
% array separately by creating an anonymous function.
data = replacedata(data,@(x) zscore(x,[],2))

```

See Also dataset

Purpose Replicate and tile categorical array
```

Syntax
$B=\operatorname{repmat}(A, m, n)$
B = repmat(A,[m n p ...])

```

Description \(\quad B=\operatorname{repmat}(A, m, n)\) creates a large array \(B\) consisting of an \(m-b y-n\) tiling of copies of the categorical array \(A\). The size of \(B\) is \([\operatorname{size}(A, 1) * M\) \(\operatorname{size}(A, 2) * N, \operatorname{size}(A, 3), \ldots]\). repmat \((A, n)\) creates an \(n\)-by-n tiling.
\(B=\operatorname{repmat}(A,[m \cap p \ldots])\) tiles the categorical array \(A\) to produce a multidimensional array \(B\) composed of copies of \(A\). The size of \(B\) is [size(A,1)*M, size(A,2)*N, size(A,3)*P, ...].

\section*{See Also ndims | size}

\section*{Purpose Reset state}

\section*{Syntax \(\quad\) reset (q)}

Description
reset (q) resets the state of the quasi-random number stream \(q\) of the qrandstream class back to its initial state, 1. Subsequent points drawn from the stream will be the same as those drawn from a new stream. The command is equivalent to \(q\). State \(=1\).

\section*{Examples}

Use qrandstream to construct a 3-D Halton stream, based on a point set that skips the first 1000 values and then retains every 101st point:
```

q = qrandstream('halton',3,'Skip',1e3,'Leap',1e2)
q =
Halton quasi-random stream in 3 dimensions
Point set properties:
Skip : }100
Leap : 100
ScrambleMethod : none
nextIdx = q.State
nextIdx =
1

```

Use qrand to generate two samples of size four:
```

X1 = qrand(q,4)
X1 =
0.0928 0.3475 0.0051
0.6958 0.2035 0.2371
0.3013 0.8496 0.4307
0.9087 0.5629 0.6166
nextIdx = q.State
nextIdx =
5
X2 = qrand(q,4)

```

\section*{qrandstream.reset}
```

X2 =
0.2446 0.0238 0.8102
0.5298 0.7540 0.0438
0.3843 0.5112 0.2758
0.8335 0.2245 0.4694
nextIdx = q.State
nextIdx =
9

```

Use reset to reset the stream, then generate another sample:
```

reset(q)
nextIdx = q.State
nextIdx =
1
X = qrand(q,4)
X =
0.0928 0.3475 0.0051
0.6958 0.2035 0.2371
0.3013 0.8496 0.4307
0.9087 0.5629 0.6166

```
See Also
qrandstream | qrand
\begin{tabular}{ll} 
Purpose & Resize categorical array \\
Syntax & \(B=\operatorname{reshape}(A, M, N)\) \\
& \(B=\operatorname{reshape}(A, m, n, p, \ldots)\) \\
& reshape \((A,[m, n, \ldots])\) \\
& \(B=\operatorname{reshape}(A, \ldots,[], \ldots)\)
\end{tabular}

Description
\(B=r e s h a p e(A, M, N)\) returns an m-by-n categorical matrix whose elements are taken columnwise from the categorical array A. An error results if A does not have \(m * n\) elements.
\(B=r e s h a p e(A, m, n, p, \ldots)\) or reshape(A,[m n p ...]) returns an array with the same elements as \(A\) but reshaped to have the size \(m-b y-n-b y-p-b y-. . . m^{*} n * p * \ldots\) must be the same as numel \((A)\).
\(B=r e s h a p e(A, \ldots,[], \ldots)\) calculates the length of the dimension represented by [], such that the product of the dimensions equals numel (A). numel (A) must be evenly divisible by the product of the known dimensions. You can use only one occurrence of [].

In general, reshape ( A, siz) returns an array with the same elements as A but reshaped to the size siz. prod(siz) must be the same as numel(A).

\section*{See Also}

\section*{ClassificationDiscriminant.resubEdge}

Purpose Classification edge by resubstitution
Syntax edge = resubEdge \((\) obj \()\)
Description edge \(=\) resubEdge \((o b j)\) returns the classification edge obtained by obj on its training data.

Input obj
Arguments

Output
Arguments

\section*{Definitions}
edge
Classification edge obtained by resubstituting the training data into the calculation of edge.

\section*{Edge}

The edge is the weighted mean value of the classification margin. The weights are the class probabilities in obj. Prior. If you supply weights in the weights name-value pair, those weights are normalized to sum to the prior probabilities in the respective classes, and are then used to compute the weighted average.

\section*{Margin}

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix obj.x. A high value of margin indicates a more reliable prediction than a low value.

\section*{Score (discriminant analysis)}

For discriminant analysis, the score of a classification is the posterior probability of the classification. For the definition of posterior probability in discriminant analysis, see "Posterior Probability" on page 12-7.

\title{
ClassificationDiscriminant.resubEdge
}

\author{
Examples Estimate the quality of a discriminant analysis classifier for the Fisher iris data by resubstitution: \\ ```
load fisheriris \\ obj = ClassificationDiscriminant.fit(meas,species); \\ redge = resubEdge(obj) \\ redge = \\ 0.9454
``` \\ See Also \\ ClassificationDiscriminant | edge | resubMargin \\ How To \\ - "Discriminant Analysis" on page 12-3
}

\section*{ClassificationEnsemble.resubEdge}
Purpose Classification edge by resubstitution
Syntax

edge = resubEdge(ens)

edge = resubEdge(ens,Name,Value)
Descriptionedge \(=\) resubEdge(ens) returns the classification edge obtained byens on its training data.
edge = resubEdge(ens,Name, Value) calculates edge with additionaloptions specified by one or more Name, Value pair arguments. Youcan specify several name-value pair arguments in any order asName1, Value1, ,NameN, ValueN.
Input ..... ens
Arguments
A classification ensemble created with fitensemble.

\section*{Name-Value Pair Arguments}
Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

\section*{learners}
Indices of weak learners in the ensemble ranging from 1 to NTrained. resubEdge uses only these learners for calculating edge.
Default: 1:NTrained
mode
String representing the meaning of the output edge:
- 'ensemble' - edge is a scalar value, the loss for the entire ensemble.

\section*{ClassificationEnsemble.resubEdge}
- 'individual' - edge is a vector with one element per trained learner.
- 'cumulative' - edge is a vector in which element \(J\) is obtained by using learners \(1: J\) from the input list of learners.

Default: 'ensemble'

\section*{Output Arguments}

\section*{Definitions}

\section*{Edge}

The edge is the weighted mean value of the classification margin. The weights are the class probabilities in ens. Prior.

\section*{Margin}

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix ens. x .

\section*{Score (ensemble)}

For ensembles, a classification score represents the confidence of a classification into a class. The higher the score, the higher the confidence.

Different ensemble algorithms have different definitions for their scores. Furthermore, the range of scores depends on ensemble type. For example:
- AdaBoostM1 scores range from \(-\infty\) to \(\infty\).
- Bag scores range from 0 to 1.

\section*{ClassificationEnsemble.resubEdge}
```

Examples Find the resubstitution edge for an ensemble that classifies the Fisher iris data:

```
```

load fisheriris

```
load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
edge = resubEdge(ens)
edge = resubEdge(ens)
edge =
edge =
    3.2486
```

    3.2486
    ```
See Also resubMargin | resubLoss | resubPredict | resubEdge
How To • Chapter 13, "Nonparametric Supervised Learning"

\section*{ClassificationTree.resubEdge}

\section*{Purpose Classification edge by resubstitution}

Syntax edge = resubEdge (tree)
Description edge \(=\) resubEdge(tree) returns the classification edge obtained by tree on its training data.

\section*{Input \\ Arguments}
tree
A classification tree created by ClassificationTree.fit.
Output
Arguments
edge
Classification edge obtained by resubstituting the training data into the calculation of edge.

\section*{Definitions}

\section*{Edge}

The edge is the weighted mean value of the classification margin. The weights are the class probabilities in tree. Prior.

\section*{Margin}

The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix \(X\).

\section*{Score (tree)}

For trees, the score of a classification of a leaf node is the posterior probability of the classification at that node. The posterior probability of the classification at a node is the number of training sequences that lead to that node with the classification, divided by the number of training sequences that lead to that node.

For example, consider classifying a predictor X as true when \(\mathrm{X}<0.15\) or \(X>0.95\), and \(X\) is false otherwise.

\section*{ClassificationTree.resubEdge}

Generate 100 random points and classify them:
```

rng(0,'twister') % for reproducibility
X = rand(100,1);
Y = (abs(X - . 55) > .4);
tree = ClassificationTree.fit(X,Y);
view(tree,'mode','graph')

```
\begin{tabular}{|c|c|c|c|c|c|c|c|c|}
\hline Click to display: & Identity & \(\checkmark\) & Magnification: & 100\% & \(\checkmark\) & Pruning level: & 0 of 2 & ㅂ \\
\hline
\end{tabular}


2
Prune the tree:
```

tree1 = prune(tree,'level',1);
view(tree1,'mode','graph')

```

\section*{ClassificationTree.resubEdge}


The pruned tree correctly classifies observations that are less than 0.15 as true. It also correctly classifies observations from .15 to .94 as false. However, it incorrectly classifies observations that are greater than . 94 as false. Therefore, the score for observations that are greater than . 15 should be about \(.05 / .85=.06\) for true, and about \(.8 / .85=.94\) for false.

3
Compute the prediction scores for the first 10 rows of X :
```

[~,score] = predict(tree1,X(1:10));
[score X(1:10,:)]
ans =

| 0.9059 | 0.0941 | 0.8147 |
| ---: | ---: | ---: |
| 0.9059 | 0.0941 | 0.9058 |
| 0 | 1.0000 | 0.1270 |
| 0.9059 | 0.0941 | 0.9134 |
| 0.9059 | 0.0941 | 0.6324 |
| 0 | 1.0000 | 0.0975 |
| 0.9059 | 0.0941 | 0.2785 |

```

\section*{ClassificationTree.resubEdge}
\begin{tabular}{lll}
0.9059 & 0.0941 & 0.5469 \\
0.9059 & 0.0941 & 0.9575 \\
0.9059 & 0.0941 & 0.9649
\end{tabular}

Indeed, every value of \(X\) (the rightmost column) that is less than 0.15 has associated scores (the left and center columns) of 0 and 1 , while the other values of \(X\) have associated scores of 0.91 and 0.09 . The difference (score 0.09 instead of the expected .06) is due to a statistical fluctuation: there are 8 observations in \(X\) in the range (.95,1) instead of the expected 5 observations.
```

Examples Estimate the quality of a classification tree for the Fisher iris data by
resubstitution:

```
```

load fisheriris

```
load fisheriris
tree = ClassificationTree.fit(meas,species);
tree = ClassificationTree.fit(meas,species);
redge = resubEdge(tree)
redge = resubEdge(tree)
redge =
redge =
    0.9384
    0.9384
See Also edge | resubMargin | resubLoss | resubPredict
How To • Chapter 13, "Nonparametric Supervised Learning"
```


## ClassificationDiscriminant.resubLoss

## Purpose

Classification error by resubstitution
Syntax

Input
Arguments

L = resubLoss(obj)
L = resubLoss(obj,Name, Value)

Description create obj.

## obj

$\mathrm{L}=$ resubLoss(obj) returns the resubstitution loss, meaning the loss computed for the data that ClassificationDiscriminant.fit used to

L = resubLoss(obj,Name, Value) returns loss statistics with additional options specified by one or more Name, Value pair arguments.

Discriminant analysis classifier, produced using ClassificationDiscriminant.fit.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## lossfun

Function handle or string representing a loss function. Built-in loss functions are:

- 'binodeviance' - See "Loss Functions" on page 20-1640.
- 'exponential' - See "Loss Functions" on page 20-1640.
- 'mincost' - Smallest misclassification cost as given by the obj. Cost matrix.

You can write your own loss function using the syntax described in "Loss Functions" on page 20-1640.

Default: 'mincost'

## ClassificationDiscriminant.resubLoss

## weights

Numeric vector of length $N$, where $N$ is the number of rows of obj. X . weights are nonnegative. When you supply weights, resubLoss computes weighted classification error.

Default: ones ( $\mathrm{N}, 1$ )

## Output Arguments

## Definitions

L
Classification error, a scalar. The meaning of the error depends on the values in weights and lossfun. See "Classification Error" on page 20-1640.

## Classification Error

The default classification error is the fraction of the training data $X$ that obj misclassifies.

Weighted classification error is the sum of weight $i$ times the Boolean value that is 1 when obj misclassifies the $i$ th row of $X$, divided by the sum of the weights.

## Loss Functions

The built-in loss functions are:

- 'binodeviance' - For binary classification, assume the classes $y_{n}$ are -1 and 1 . With weight vector $w$ normalized to have sum 1 , and predictions of row $n$ of data $X$ as $f\left(X_{n}\right)$, the binomial deviance is

$$
\sum w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(X_{n}\right)\right)\right)
$$

- 'exponential' - With the same definitions as for 'binodeviance', the exponential loss is

$$
\sum w_{n} \exp \left(-y_{n} f\left(X_{n}\right)\right)
$$

## ClassificationDiscriminant.resubLoss

- 'mincost' - Smallest expected misclassification cost, with expectation taken over the posterior probability, and cost as given by the obj. Cost matrix.

To write your own loss function, create a function file of the form

```
function loss = lossfun(C,S,W,COST)
```

- $N$ is the number of rows of obj. X .
- K is the number of classes in obj, represented in obj.ClassNames.
- C is an N-by-K logical matrix, with one true per row for the true class. The index for each class is its position in obj.ClassNames.
- $S$ is an N-by-K numeric matrix. $S$ is a matrix of posterior probabilities for classes with one row per observation, similar to the posterior output from predict.
- $W$ is a numeric vector with $N$ elements, the observation weights.
- COST is a K-by-K numeric matrix of misclassification costs. The default 'classiferror' gives a cost of 0 for correct classification, and 1 for misclassification.
- The output loss should be a scalar.

Pass the function handle @lossfun as the value of the lossfun name-value pair.

## Examples Compute the resubstituted classification error for the Fisher iris data:

```
load fisheriris
obj = ClassificationDiscriminant.fit(meas,species);
L = resubLoss(obj)
L =
    0.0200
```

See Also
ClassificationDiscriminant | loss

## ClassificationDiscriminant.resubLoss

How To<br>- "Discriminant Analysis" on page 12-3

## ClassificationEnsemble.resubLoss

## Purpose

Classification error by resubstitution
Syntax
L = resubLoss(ens)
L = resubLoss(ens,Name,Value)

Description

Input
Arguments
$\mathrm{L}=$ resubLoss(ens) returns the resubstitution loss, meaning the loss computed for the data that fitensemble used to create ens.
L = resubLoss(ens, Name, Value) calculates loss with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN,ValueN.
ens
A classification ensemble created with fitensemble.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

## learners

Indices of weak learners in the ensemble ranging from 1 to NTrained. resubLoss uses only these learners for calculating loss.

Default: 1:NTrained

## lossfun

Function handle or string representing a loss function. Built-in loss functions:

- 'binodeviance ' - See "Loss Functions" on page 20-1644
- 'classiferror' - Fraction of misclassified data


## ClassificationEnsemble.resubLoss

- 'exponential' - See "Loss Functions" on page 20-1644

You can write your own loss function in the syntax described in "Loss Functions" on page 20-1644.

Default: 'classiferror'

## mode

String representing the meaning of the output L:

- 'ensemble' - L is a scalar value, the loss for the entire ensemble.
- 'individual' - L is a vector with one element per trained learner.
- 'cumulative' - $L$ is a vector in which element $J$ is obtained by using learners $1: J$ from the input list of learners.

Default: 'ensemble'

## Output

Arguments

## Definitions

L
Loss, by default the fraction of misclassified data. L can be a vector, and can mean different things, depending on the name-value pair settings.

## Classification Error

The default classification error is the fraction of the training data $X$ that ens misclassifies.

Weighted classification error is the sum of weight $i$ times the Boolean value that is 1 when tree misclassifies the $i$ th row of X , divided by the sum of the weights.

## Loss Functions

The built-in loss functions are:

## ClassificationEnsemble.resubLoss

- 'binodeviance' - For binary classification, assume the classes $y_{n}$ are -1 and 1 . With weight vector $w$ normalized to have sum 1 , and predictions of row $n$ of data $X$ as $f\left(X_{n}\right)$, the binomial deviance is

$$
\sum w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(X_{n}\right)\right)\right)
$$

- 'classiferror' - Fraction of misclassified data, weighted by $w$.
- 'exponential' - With the same definitions as for 'binodeviance', the exponential loss is

$$
\sum w_{n} \exp \left(-y_{n} f\left(X_{n}\right)\right)
$$

To write your own loss function, create a function file of the form

```
function loss = lossfun(C,S,W,COST)
```

- $N$ is the number of rows of ens. $X$.
- K is the number of classes in ens, represented in ens.ClassNames.
- C is an N-by-K logical matrix, with one true per row for the true class. The index for each class is its position in tree.ClassNames.
- $S$ is an N-by-K numeric matrix. $S$ is a matrix of posterior probabilities for classes with one row per observation, similar to the score output from predict.
- W is a numeric vector with $N$ elements, the observation weights.
- COST is a K-by-K numeric matrix of misclassification costs. The default 'classiferror' gives a cost of 0 for correct classification, and 1 for misclassification.
- The output loss should be a scalar.

Pass the function handle @lossfun as the value of the lossfun name-value pair.

## ClassificationEnsemble.resubLoss

```
Examples Compute the resubstitution loss for a classification ensemble for the Fisher iris data:
```

```
load fisheriris
```

load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
loss = resubLoss(ens)
loss = resubLoss(ens)
loss =
loss =
0.0333

```
    0.0333
```

See Also resubEdge | resubMargin | resubPredict | resubLoss
How To • Chapter 13, "Nonparametric Supervised Learning"

## ClassificationTree.resubLoss

## Purpose

Classification error by resubstitution
Syntax

## Description

$\mathrm{L}=$ resubLoss(tree)
L = resubLoss(tree, Name, Value)
L = resubLoss(tree,'subtrees',subtreevector)
[L,se] = resubLoss(tree,'subtrees',subtreevector)
subtreevector)
[L, ...] =
[L,se,NLeaf] = resubLoss(tree, 'subtrees', subtreevector)
[L,se,NLeaf,bestlevel] = resubLoss(tree,'subtrees',
resubLoss(tree, 'subtrees', subtreevector, Name, Value)
$\mathrm{L}=$ resubLoss(tree) returns the resubstitution loss, meaning the loss computed for the data that ClassificationTree.fit used to create tree.

L = resubLoss(tree, Name, Value) returns the loss with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

L = resubLoss(tree,'subtrees', subtreevector) returns a vector of classification errors for the trees in the pruning sequence subtreevector.
[L,se] = resubLoss(tree,'subtrees', subtreevector) returns the vector of standard errors of the classification errors.
[L,se,NLeaf] = resubLoss(tree,'subtrees',subtreevector) returns the vector of numbers of leaf nodes in the trees of the pruning sequence.
[L,se,NLeaf,bestlevel] =
resubLoss(tree, 'subtrees', subtreevector) returns the best pruning level as defined in the treesize name-value pair. By default, bestlevel is the pruning level that gives loss within one standard deviation of minimal loss.

```
[L,...] =
resubLoss(tree,'subtrees',subtreevector,Name,Value) returns
```


## ClassificationTree.resubLoss

loss statistics with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

## Input <br> Arguments <br> tree

A classification tree constructed by ClassificationTree.fit.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## lossfun

Function handle or string representing a loss function. Built-in loss functions:

- 'binodeviance ' - See "Loss Functions" on page 20-1650
- 'classiferror' - Fraction of misclassified data
- 'exponential' - See "Loss Functions" on page 20-1650

You can write your own loss function in the syntax described in "Loss Functions" on page 20-1650.

Default: 'classiferror'
weights
A numeric vector of length $N$, where $N$ is the number of rows of tree. X. weights are nonnegative. When you supply weights, loss computes weighted classification error.

Default: ones ( $\mathrm{N}, 1$ )
Name, Value arguments associated with pruning subtrees:

## ClassificationTree.resubLoss

## subtrees

A vector with integer values from 0 (full unpruned tree) to the maximal pruning level max(tree. PruneList).

Default: 0

## treesize

One of the following strings:

- 'se' - loss returns the highest pruning level with loss within one standard deviation of the minimum ( $L+s e$, where $L$ and se relate to the smallest value in subtrees).
- 'min' - loss returns the element of subtrees with smallest loss, usually the smallest element of subtrees.


## Output Arguments

## L

Classification error, a vector the length of subtrees. The meaning of the error depends on the values in weights and lossfun; see "Classification Error" on page 20-1650.

Standard error of loss, a vector the length of subtrees.

## NLeaf

Number of leaves (terminal nodes) in the pruned subtrees, a vector the length of subtrees.

## bestlevel

A scalar whose value depends on treesize:

- treesize $=$ 'se' - loss returns the highest pruning level with loss within one standard deviation of the minimum ( $\mathrm{L}+\mathrm{se}$, where $L$ and se relate to the smallest value in subtrees).


## ClassificationTree.resubLoss

- treesize $=$ 'min' - loss returns the element of subtrees with smallest loss, usually the smallest element of subtrees.


## Definitions

## Classification Error

The default classification error is the fraction of the training data $X$ that tree misclassifies.

Weighted classification error is the sum of weight $i$ times the Boolean value that is 1 when tree misclassifies the $i$ th row of $X$, divided by the sum of the weights.

## Loss Functions

The built-in loss functions are:

- 'binodeviance' - For binary classification, assume the classes $y_{n}$ are -1 and 1 . With weight vector $w$ normalized to have sum 1, and predictions of row $n$ of data $X$ as $f\left(X_{n}\right)$, the binomial deviance is

$$
\sum w_{n} \log \left(1+\exp \left(-2 y_{n} f\left(X_{n}\right)\right)\right)
$$

- 'classiferror' - Fraction of misclassified data, weighted by $w$.
- 'exponential' - With the same definitions as for 'binodeviance', the exponential loss is

$$
\sum w_{n} \exp \left(-y_{n} f\left(X_{n}\right)\right)
$$

To write your own loss function, create a function file of the form

$$
\text { function loss }=\text { lossfun(C, } \mathrm{S}, \mathrm{~W}, \mathrm{COST})
$$

- $N$ is the number of rows of tree. $X$.
- K is the number of classes in tree, represented in tree.ClassNames.
- C is an N-by-K logical matrix, with one true per row for the true class. The index for each class is its position in tree.ClassNames.


## ClassificationTree.resubLoss

- $S$ is an $N$-by-K numeric matrix. $S$ is a matrix of posterior probabilities for classes with one row per observation, similar to the posterior output from predict.
- $W$ is a numeric vector with $N$ elements, the observation weights.
- COST is a K-by-K numeric matrix of misclassification costs. The default 'classiferror' gives a cost of 0 for correct classification, and 1 for misclassification.
- The output loss should be a scalar.

Pass the function handle @lossfun as the value of the lossfun name-value pair.

```
Examples
    Compute the resubstituted classification error for the ionosphere data:
load ionosphere
tree \(=\) ClassificationTree.fit(X,Y);
L = resubLoss(tree)
L =
    0.0114
```

See Also loss | resubEdge | resubMargin | resubPredict
How To - Chapter 13, "Nonparametric Supervised Learning"

## RegressionEnsemble.resubLoss

Purpose Regression error by resubstitution
Syntax L = resubLoss(ens)
L = resubLoss(ens,Name,Value)
Description$\mathrm{L}=$ resubLoss(ens) returns the resubstitution loss, meaning themean squared error computed for the data that fitensemble used tocreate ens.L = resubLoss(ens,Name, Value) calculates loss with additionaloptions specified by one or more Name, Value pair arguments. Youcan specify several name-value pair arguments in any order asName1, Value1, ,NameN, ValueN.
Input ..... ens
ArgumentsA regression ensemble created with fitensemble.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## learners

Indices of weak learners in the ensemble ranging from 1 to NTrained. resubLoss uses only these learners for calculating loss.
Default: 1:NTrained

## lossfun

Function handle for loss function, or the string 'mse', meaning mean squared error. If you pass a function handle fun, resubLoss calls it as
FUN(Y,Yfit,w)
where $\mathrm{Y}, \mathrm{Yfit}$, and W are numeric vectors of the same length. Y is the observed response, Yfit is the predicted response, and $W$ is the observation weights.

Default: 'mse'

## mode

String representing the meaning of the output L:

- 'ensemble' - L is a scalar value, the loss for the entire ensemble.
- 'individual' - L is a vector with one element per trained learner.
- 'cumulative' - $L$ is a vector in which element $J$ is obtained by using learners $1: J$ from the input list of learners.

Default: 'ensemble'

## Output

Arguments

## Examples Find the resubstitution predictions of mileage from the carsmall data based on horsepower and weight, and look at their mean square difference from the training data.

```
load carsmall
X = [Horsepower Weight];
ens = fitensemble(X,MPG,'LSBoost',100,'Tree');
MSE = resubLoss(ens)
MSE =
    6.4336
```

See Also
resubPredict | resubLoss | loss

## RegressionEnsemble.resubLoss

How To<br>- Chapter 13, "Nonparametric Supervised Learning"

## RegressionTree.resubLoss

## Purpose

Regression error by resubstitution
Syntax

## Description

L = resubLoss(tree)
L = resubLoss(tree, Name, Value)
L = resubLoss(tree,'subtrees',subtreevector) subtreevector)
[ $\mathrm{L}, \ldots$. $]=$
[L,se] = resubLoss(tree,'subtrees',subtreevector)
[L,se,NLeaf] = resubLoss(tree,'subtrees',subtreevector)
[L,se,NLeaf,bestlevel] = resubLoss(tree,'subtrees',
resubLoss(tree,'subtrees', subtreevector, Name, Value)
$\mathrm{L}=$ resubLoss(tree) returns the resubstitution loss, meaning the loss computed for the data that RegressionTree.fit used to create tree.

L = resubLoss(tree, Name, Value) returns the loss with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

L = resubLoss(tree,'subtrees', subtreevector) returns a vector of mean squared errors for the trees in the pruning sequence subtreevector.
[L,se] = resubLoss(tree, 'subtrees', subtreevector) returns the vector of standard errors of the classification errors.
[L,se,NLeaf] = resubLoss(tree,'subtrees',subtreevector) returns the vector of numbers of leaf nodes in the trees of the pruning sequence.
[L,se,NLeaf,bestlevel] =
resubLoss(tree, 'subtrees', subtreevector) returns the best pruning level as defined in the treesize name-value pair. By default, bestlevel is the pruning level that gives loss within one standard deviation of minimal loss.
[L, ...] =
resubLoss(tree, 'subtrees', subtreevector, Name, Value) returns loss statistics with additional options specified by one or more

## RegressionTree.resubLoss

Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

Input
Arguments
tree
A regression tree constructed by RegressionTree.fit.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
lossfun
Function handle, or the string 'mse' meaning mean squared error.

You can write your own loss function in the syntax described in "Loss Functions" on page 20-1657.

Default: 'mse'
weights
A numeric vector of length $N$, where $N$ is the number of rows of tree. X. weights are nonnegative. When you supply weights, loss computes weighted classification error.

Default: ones (N, 1)
Name, Value arguments associated with pruning subtrees:
subtrees
A vector with integer values from 0 (full unpruned tree) to the maximal pruning level max(tree. PruneList).

Default: 0

## RegressionTree.resubLoss

## treesize

One of the following strings:

- 'se' - loss returns the highest pruning level with loss within one standard deviation of the minimum ( $L+s e$, where $L$ and se relate to the smallest value in subtrees).
- 'min' - loss returns the element of subtrees with smallest loss, usually the smallest element of subtrees.


## Output <br> Arguments

## Definitions

## L

Mean squared error, a vector the length of subtrees. The meaning of the error depends on the values in weights and lossfun.

## se

Standard error of loss, a vector the length of subtrees.
NLeaf
Number of leaves (terminal nodes) in the pruned subtrees, a vector the length of subtrees.

## bestlevel

A scalar whose value depends on treesize:

- treesize = 'se' - loss returns the highest pruning level with loss within one standard deviation of the minimum ( $L+$ se, where $L$ and se relate to the smallest value in subtrees).
- treesize = 'min' - loss returns the element of subtrees with smallest loss, usually the smallest element of subtrees.


## Loss Functions

The built-in loss function is 'mse', meaning mean squared error.
To write your own loss function, create a function file of the form

```
function loss = lossfun(Y,Yfit,W)
```


## RegressionTree.resubLoss

- $N$ is the number of rows of tree. $X$.
- $Y$ is an $N$-element vector representing the observed response.
- Yfit is an N -element vector representing the predicted responses.
- $W$ is an $N$-element vector representing the observation weights.
- The output loss should be a scalar.

Pass the function handle @lossfun as the value of the lossfun name-value pair.

## Examples

Find the mean square error of a model of the carsmall data:

```
load carsmall
X = [Displacement Horsepower Weight];
tree = RegressionTree.fit(X,MPG);
resubLoss(tree)
ans =
    4.8952
```

See Also resubPredict | loss
How To

- Chapter 13, "Nonparametric Supervised Learning"


## ClassificationDiscriminant.resubMargin

| Purpose | Classification margins by resubstitution |
| :---: | :---: |
| Syntax | M = resubMargin(obj) |
| Description | M = resubMargin(obj) returns resubstitution classification margins for obj. |
| Input Arguments | obj <br> Discriminant analysis classifier, produced using ClassificationDiscriminant.fit. |
| Output Arguments | M <br> Numeric column-vector of length size(obj. $\mathrm{X}, 1$ ) containing the classification margins. |
| Definitions | Margin |
|  | The classification margin is the difference between the classification score for the true class and maximal classification score for the false classes. Margin is a column vector with the same number of rows as in the matrix obj. x . A high value of margin indicates a more reliable prediction than a low value. |
|  | Score (discriminant analysis) |
|  | For discriminant analysis, the score of a classification is the posterior probability of the classification. For the definition of posterior probability in discriminant analysis, see "Posterior Probability" on page 12-7. |
| Examples | Find the margins for a classification tree for the Fisher iris data by resubstitution. Examine several entries: |
|  | ```load fisheriris obj = ClassificationDiscriminant.fit(meas,species); M = resubMargin(obj); M(1:25:end)``` |

## ClassificationDiscriminant.resubMargin

$$
\begin{aligned}
& \text { ans }= \\
& 1.0000 \\
& 1.0000 \\
& 0.9998 \\
& 0.9998 \\
& 1.0000 \\
& 0.9946
\end{aligned}
$$

## See Also ClassificationDiscriminant | margin

How To . "Discriminant Analysis" on page 12-3

## ClassificationEnsemble.resubMargin

## Purpose Classification margins by resubstitution

```
Syntax
margin = resubMargin(ens)
margin = resubMargin(ens,Name,Value)
```

Description

## Input Arguments

## Definitions

margin
A numeric column-vector of length size(ens. $\mathrm{X}, 1$ ) containing the classification margins.

## Margin

The classification margin is the difference between the classification score for the true class and maximal classification score for the false

## ClassificationEnsemble.resubMargin

classes. Margin is a column vector with the same number of rows as in the matrix ens. x .

## Score (ensemble)

For ensembles, a classification score represents the confidence of a classification into a class. The higher the score, the higher the confidence.

Different ensemble algorithms have different definitions for their scores. Furthermore, the range of scores depends on ensemble type. For example:

- AdaBoostM1 scores range from $-\infty$ to $\infty$.
- Bag scores range from 0 to 1.

```
Examples Find the resubstitution margins for an ensemble that classifies the Fisher iris data:
```

```
load fisheriris
```

load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
margin = resubMargin(ens);
margin = resubMargin(ens);
[min(margin) mean(margin) max(margin)]
[min(margin) mean(margin) max(margin)]
ans =
ans =
-0.5674 3.2486 4.6245

```
    -0.5674 3.2486 4.6245
```

See Also resubEdge | resubLoss | resubPredict | resubMargin
How To • Chapter 13, "Nonparametric Supervised Learning"

## ClassificationTree.resubMargin

Purpose Classification margins by resubstitution
Syntax M = resubMargin(tree)
Description M = resubMargin(tree) returns resubstitution classification margins for tree.
Input treeArgumentsA classification tree created by ClassificationTree.fit.
Output ..... M
Arguments
A numeric column-vector of length size(tree. $X, 1$ ) containing the classification margins.

## Definitions

## Margin

Classification margin is the difference between classification score for the true class and maximal classification score for the false classes. A high value of margin indicates a more reliable prediction than a low value.

## Score (tree)

For trees, the score of a classification of a leaf node is the posterior probability of the classification at that node. The posterior probability of the classification at a node is the number of training sequences that lead to that node with the classification, divided by the number of training sequences that lead to that node.
For example, consider classifying a predictor X as true when $\mathrm{X}<0.15$ or $X>0.95$, and $X$ is false otherwise.

## 1

Generate 100 random points and classify them:

```
rng(0,'twister') % for reproducibility
X = rand(100,1);
```


## ClassificationTree.resubMargin

```
Y = (abs(X - . 55) > .4);
tree = ClassificationTree.fit(X,Y);
view(tree,'mode','graph')
```

| Click to display: | Identity | Magnification: $100 \%$ | Pruning level: 0 of 2 |
| :--- | :--- | :--- | :--- | :--- |



2
Prune the tree:

```
tree1 = prune(tree,'level',1);
view(tree1,'mode','graph')
```


## ClassificationTree.resubMargin



The pruned tree correctly classifies observations that are less than 0.15 as true. It also correctly classifies observations from .15 to .94 as false. However, it incorrectly classifies observations that are greater than . 94 as false. Therefore, the score for observations that are greater than . 15 should be about $.05 / .85=.06$ for true, and about $.8 / .85=.94$ for false.

## 3

Compute the prediction scores for the first 10 rows of X :

```
[~,score] = predict(tree1,X(1:10));
[score X(1:10,:)]
ans =
\begin{tabular}{rrr}
0.9059 & 0.0941 & 0.8147 \\
0.9059 & 0.0941 & 0.9058 \\
0 & 1.0000 & 0.1270 \\
0.9059 & 0.0941 & 0.9134 \\
0.9059 & 0.0941 & 0.6324 \\
0 & 1.0000 & 0.0975 \\
0.9059 & 0.0941 & 0.2785
\end{tabular}
```


## ClassificationTree.resubMargin

| 0.9059 | 0.0941 | 0.5469 |
| :--- | :--- | :--- |
| 0.9059 | 0.0941 | 0.9575 |
| 0.9059 | 0.0941 | 0.9649 |

Indeed, every value of $X$ (the rightmost column) that is less than 0.15 has associated scores (the left and center columns) of 0 and 1 , while the other values of $X$ have associated scores of 0.91 and 0.09 . The difference (score 0.09 instead of the expected .06) is due to a statistical fluctuation: there are 8 observations in $X$ in the range (.95,1) instead of the expected 5 observations.

## Examples

Find the margins for a classification tree for the Fisher iris data by resubstitution. Examine several entries:

```
load fisheriris
tree = ClassificationTree.fit(meas,species);
M = resubMargin(tree);
M(1:25:end)
ans =
    1.0000
    1.0000
    1.0000
    1.0000
    0.9565
    0.9565
```

See Also margin | resubLoss | resubPredict | resubEdge
How To - Chapter 13, "Nonparametric Supervised Learning"

## ClassificationDiscriminant.resubPredict



## ClassificationDiscriminant.resubPredict

Definitions Posterior Probabilityposterior ( $i, k$ ) is the posterior probability of class $k$ for observationi. For the mathematical definition, see "Posterior Probability" on page12-7.
Examples Find the total number of misclassifications of the Fisher iris data for a discriminant analysis classifier:
load fisheriris

obj = ClassificationDiscriminant.fit(meas,species);

Ypredict $=$ resubPredict(obj); \% the predictions

Ysame = strcmp(Ypredict,species); \% true when ==

sum( $\sim$ Ysame) \% how many are different?

ans =
3
See Also ClassificationDiscriminant | predict
How To - "Discriminant Analysis" on page 12-3

## ClassificationEnsemble.resubPredict

Purpose
Predict ensemble response by resubstitution
Syntax
label = resubPredict(ens)
[label,score] = resubPredict(ens)
[label, score] = resubPredict(ens, Name, Value)

Description

Input
Arguments

## Output Arguments

label $=$ resubPredict(ens) returns the labels ens predicts for the data ens. $x$. label is the predictions of ens on the data that fitensemble used to create ens.
[label,score] = resubPredict(ens) also returns scores for all classes.
[label, score] = resubPredict(ens,Name, Value) finds resubstitution predictions with additional options specified by one or more Name, Value pair arguments.
ens
A classification ensemble created with fitensemble.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## learners

Indices of weak learners in the ensemble ranging from 1 to NTrained. oobLoss uses only these learners for calculating loss.

Default: 1:NTrained
label
The response ens predicts for the training data. label is the same data type as the training response data ens. Y , and has the same number of entries as the number of rows in ens.X.

## ClassificationEnsemble.resubPredict

score
An $N$-by- K matrix, where N is the number of rows in ens. X , and K is the number of classes in ens. High score value indicates that an observation likely comes from this class.

## Definitions

Examples
Find the total number of misclassifications of the Fisher iris data for a classification ensemble:

```
load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,'Tree');
Ypredict = resubPredict(ens); % the predictions
Ysame = strcmp(Ypredict,species); % true when ==
sum(~Ysame) % how many are different?
ans =
    5
```


## See Also resubEdge | resubMargin | resubLoss | resubPredict

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## ClassificationTree.resubPredict

Purpose
Syntax
Description

Input Arguments
Description

Predict resubstitution response of tree

```
label = resubPredict(tree)
[label,posterior] = resubPredict(tree)
[label,posterior,node] = resubPredict(tree)
[label,posterior,node,cnum] = resubPredict(tree)
[label,...] = resubPredict(tree,Name,Value)
```

label $=$ resubPredict(tree) returns the labels tree predicts for the data tree. X. label is the predictions of tree on the data that ClassificationTree.fit used to create tree.
[label, posterior] = resubPredict(tree) returns the posterior class probabilities for the predictions.
[label, posterior, node] = resubPredict(tree) returns the node numbers of tree for the resubstituted data.
[label, posterior, node, cnum] = resubPredict(tree) returns the predicted class numbers for the predictions.
[label,...] = resubPredict(tree,Name,Value) returns resubstitution predictions with additional options specified by one or more Name, Value pair arguments.

## tree

A classification tree constructed by ClassificationTree.fit.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
subtrees

## ClassificationTree.resubPredict

A vector with integer values from 0 (full unpruned tree) to the maximal pruning level max (tree. PruneList). subtrees must be in ascending order.

Default: 0

## Output label <br> Arguments

The response tree predicts for the training data. label is the same data type as the training response data tree. Y.

If the subtrees name-value argument contains $m>1$ entries, label has $m$ columns, each of which represents the predictions of the corresponding subtree. Otherwise, label is a vector.

## posterior

Matrix or array of posterior probabilities for classes tree predicts.
If the subtrees name-value argument is a scalar or is missing, posterior is an $n$-by-k matrix, where $n$ is the number of rows in the training data tree. $X$, and $k$ is the number of classes.
If subtrees contains $m>1$ entries, posterior is an $n$-by- $k$-by- $m$ array, where the matrix for each $m$ gives posterior probabilities for the corresponding subtree.
node
The node numbers of tree where each data row resolves.
If the subtrees name-value argument is a scalar or is missing, node is a numeric column vector with n rows, the same number of rows as tree. X .

If subtrees contains $m>1$ entries, node is a $n$-by-m matrix. Each column represents the node predictions of the corresponding subtree.
cnum
The class numbers that tree predicts for the resubstituted data.

## ClassificationTree.resubPredict

If the subtrees name-value argument is a scalar or is missing, cnum is a numeric column vector with $n$ rows, the same number of rows as tree. X .

If subtrees contains $m>1$ entries, cnum is a $n$-by-m matrix. Each column represents the class predictions of the corresponding subtree.

## Definitions Posterior Probability

The posterior probability of the classification at a node is the number of training sequences that lead to that node with this classification, divided by the number of training sequences that lead to that node.
For example, consider classifying a predictor $X$ as true when $X<0.15$ or $X>0.95$, and $X$ is false otherwise.

1
Generate 100 random points and classify them:

```
rng(0,'twister') % for reproducibility
X = rand(100,1);
Y = (abs(X - .55) > .4);
tree = ClassificationTree.fit(X,Y);
view(tree,'mode','graph')
```


## ClassificationTree.resubPredict



2
Prune the tree:

```
tree1 = prune(tree,'level',1);
view(tree1,'mode','graph')
```


## ClassificationTree.resubPredict



The pruned tree correctly classifies observations that are less than 0.15 as true. It also correctly classifies observations between .15 and .94 as false. However, it incorrectly classifies observations that are greater than .94 as false. Therefore the score for observations that are greater than .15 should be about $.05 / .85=.06$ for true, and about $.8 / .85=.94$ for false.

3
Compute the prediction scores for the first 10 rows of X :

```
[~,score] = predict(tree1,X(1:10));
[score X(1:10,:)]
ans =
\begin{tabular}{rrr}
0.9405 & 0.0595 & 0.6555 \\
0.9405 & 0.0595 & 0.1712 \\
0.9405 & 0.0595 & 0.7060 \\
0 & 1.0000 & 0.0318 \\
0.9405 & 0.0595 & 0.2769 \\
0 & 1.0000 & 0.0462
\end{tabular}
```


## ClassificationTree.resubPredict

| 0 | 1.0000 | 0.0971 |
| ---: | ---: | ---: |
| 0.9405 | 0.0595 | 0.8235 |
| 0.9405 | 0.0595 | 0.6948 |
| 0.9405 | 0.0595 | 0.3171 |

Indeed, every value of $X$ (the rightmost column) that is less than 0.15 has associated scores (the left and center columns) of 0 and 1 , while the other values of $X$ have associated scores of 0.94 and 0.06 .

## Examples

Find the total number of misclassifications of the Fisher iris data for a classification tree:

```
load fisheriris
tree = ClassificationTree.fit(meas,species);
Ypredict = resubPredict(tree); % the predictions
Ysame = strcmp(Ypredict,species); % true when ==
sum(~Ysame) % how many are different?
ans =
```

    3
    See Also resubEdge | resubMargin | resubLoss | predict
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## RegressionEnsemble.resubPredict

## Purpose

Predict response of ensemble by resubstitution
Syntax
Yfit = resubPredict(ens)
Yfit = resubPredict(ens,Name,Value)

Input Arguments

## Examples

Yfit
A vector of predicted responses to the training data, with ens. $X$ elements.

Find the resubstitution predictions of mileage from the carsmall data based on horsepower and weight, and look at their mean square difference from the training data.

## RegressionEnsemble.resubPredict

```
load carsmall
X = [Horsepower Weight];
ens = fitensemble(X,MPG,'LSBoost',100,'Tree');
Yfit = resubPredict(ens);
MSE = mean((Yfit - ens.Y).^2)
MSE =
    6.4336
```

This is the same as the result of resubLoss:

```
resubLoss(ens)
ans =
    6.4336
```

See Also
resubLoss | resubPredict | predict
How To

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## RegressionTree.resubPredict

| Purpose | Predict resubstitution response of tree |
| :--- | :--- |
| Syntax | Yfit $=$ resubPredict (tree) |
|  | $[$ Yfit, node $]=$ resubPredict $($ tree $)$ |
|  | $[$ Yfit, node $]=$ resubPredict $($ tree, Name, Value) |

## Description

## Input <br> Arguments

Yfit = resubPredict(tree) returns the responses tree predicts for the data tree.X. Yfit is the predictions of tree on the data that RegressionTree.fit used to create tree.
[Yfit, node] = resubPredict(tree) returns the node numbers of tree for the resubstituted data.
[Yfit, node] = resubPredict(tree,Name,Value) predicts with additional options specified by one or more Name, Value pair arguments.

## tree

A regression tree constructed by RegressionTree.fit.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## subtrees

A vector with integer values from 0 (full unpruned tree) to the maximal pruning level max(tree. PruneList). subtrees must be in ascending order.

Default: 0
Output
Arguments

Yfit
The response tree predicts for the training data.

## RegressionTree.resubPredict

If the subtrees name-value argument is a scalar or is missing, label is the same data type as the training response data tree. Y.
If subtrees contains $m>1$ entries, label has $m$ columns, each of which represents the predictions of the corresponding subtree. node

The tree node numbers where tree sends each data row.
If the subtrees name-value argument is a scalar or is missing, node is a numeric column vector with n rows, the same number of rows as tree. X .

If subtrees contains $m>1$ entries, node is a $n$-by-m matrix. Each column represents the node predictions of the corresponding subtree.

## Examples Find the mean square error of a model of the carsmall data:

```
load carsmall
X = [Displacement Horsepower Weight];
tree = RegressionTree.fit(X,MPG);
Yfit = resubPredict(tree);
mean((Yfit - tree.Y).^2)
ans =
    4.8952
```

You can get the same answer using resubLoss:

```
resubLoss(tree)
ans =
    4.8952
```

See Also resubloss | predict
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## ClassificationEnsemble.resume

| Purpose | Resume training ensemble |
| :---: | :---: |
| Syntax | ```ens1 = resume(ens,nlearn) ens1 = resume(ens,nlearn,Name,Value)``` |
| Description | ens1 = resume(ens, nlearn) trains ens for nlearn more cycles. resume uses the same training options fitensemble used to create ens. <br> ens1 = resume(ens, nlearn, Name, Value) trains ens with additional options specified by one or more Name, Value pair arguments. |
| Input Arguments | ens ${ }^{\text {A classification ensemble, created with fitensemble. }}$ |
|  | nlearn <br> A positive integer, the number of cycles for additional training of ens. |
|  | Name-Value Pair Arguments |
|  | Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN. <br> nprint |
|  | Printout frequency, a positive integer scalar or 'off' (no printouts). Returns to the command line the number of weak learners trained so far. Useful when you train ensembles with many learners on large data sets. |
|  | Default: 'off' |
| Output Arguments | ens1 <br> The classification ensemble ens, augmented with additional training. |

## ClassificationEnsemble.resume

```
Examples Train a classification ensemble for 10 cycles. Examine the resubstitution
error. Then train for 10 more cycles and examine the new resubstitution
error.
load ionosphere
ens = fitensemble(X,Y,'GentleBoost',10,'Tree');
\(\mathrm{L}=\) resubLoss(ens)
L =
        0.0484
ens1 = resume(ens,10);
L = resubLoss(ens1)
\(\mathrm{L}=\)
    0.0256
```

The new ensemble has much less resubstitution error than the original.

## See Also <br> fitensemble

$$
\text { How To } \quad \text { Chapter 13, "Nonparametric Supervised Learning" }
$$

## Purpose

Resume training learners on cross-validation folds
Syntax

Description

## Input Arguments

ens1 = resume(ens,nlearn)
ens1 = resume(ens, nlearn, Name, Value)
ens1 = resume(ens, nlearn) trains ens in every fold for nlearn more cycles. resume uses the same training options fitensemble used to create ens.
ens1 = resume(ens, nlearn, Name, Value) trains ens with additional options specified by one or more Name, Value pair arguments.
ens
A cross-validated classification ensemble. ens is the result of
either:

- The fitensemble function with a cross-validation name-value pair. The names are 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
- The crossval method applied to a classification ensemble.


## nlearn

A positive integer, the number of cycles for additional training of ens.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.
nprint
Printout frequency, a positive integer scalar or 'off' (no printouts). Returns to the command line the number of weak

## ClassificationPartitionedEnsemble.resume

learners trained so far. Useful when you train ensembles with many learners on large data sets.

Default: 'off'

```
Output
Arguments
```


## Examples

```
Train a partitioned classification ensemble for 10 cycles. Examine the error. Then train for 10 more cycles and examine the new error.
```

```
load ionosphere
```

load ionosphere
cvens = fitensemble(X,Y,'GentleBoost',10,'Tree',...
cvens = fitensemble(X,Y,'GentleBoost',10,'Tree',...
'crossval','on');
'crossval','on');
L = kfoldLoss(cvens)
L = kfoldLoss(cvens)
L =
L =
0.0883
0.0883
cvens = resume(cvens,10);
cvens = resume(cvens,10);
L = kfoldLoss(cvens)
L = kfoldLoss(cvens)
L =
L =
0.0769

```
    0.0769
```

The ensemble has less cross-validation error after training for ten more cycles.

See Also kfoldPredict | kfoldEdge | kfoldMargin | kfoldLoss
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## RegressionEnsemble.resume

Purpose Resume training ensemble
Syntax

ens1 = resume(ens,nlearn)
ens1 = resume(ens, nlearn, Name, Value)
Description
Input Arguments
OutputArguments

## ens 1

The regression ensemble ens, augmented with additional training.

## RegressionEnsemble.resume

Examples Train a regression ensemble for 50 cycles. Examine the resubstitution error. Then train for 50 more cycles and examine the new resubstitution error.

```
load carsmall
X = [Displacement Horsepower Weight];
ens = fitensemble(X,MPG,'LSBoost',50,'Tree');
    L = resubLoss(ens)
    L =
            6.2681
ens = resume(ens,50);
L = resubLoss(ens)
L =
    4.3904
```

The new ensemble has much less resubstitution error than the original.

## See Also

fitensemble

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## RegressionPartitionedEnsemble.resume

## Purpose

Resume training ensemble
Syntax

Description

## Input

Arguments

```
ens1 = resume(ens,nlearn)
```

ens1 = resume(ens, nlearn, Name, Value) create ens.

## ens

ens1 = resume(ens, nlearn) trains ens in every fold for nlearn more cycles. resume uses the same training options fitensemble used to
ens1 = resume(ens, nlearn, Name, Value) trains ens with additional options specified by one or more Name, Value pair arguments.

A cross-validated regression ensemble. ens is the result of either:

- The fitensemble function with a cross-validation name-value pair. The names are 'crossval', 'kfold', 'holdout', 'leaveout', or 'cvpartition'.
- The crossval method applied to a regression ensemble.
nlearn
A positive integer, the number of cycles for additional training of ens.


## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.
nprint
Printout frequency, a positive integer scalar or 'off' (no printouts). Returns to the command line the number of weak learners trained so far. Useful when you train ensembles with many learners on large data sets.

## RegressionPartitionedEnsemble.resume

Default: 'off'

## Output <br> ens1 <br> Arguments

The cross-validated regression ensemble ens, augmented with additional training.

## Examples <br> Train a regression ensemble for 50 cycles, and cross validate it. Examine the cross-validation error. Then train for 50 more cycles and examine the new cross-validation error.

```
load carsmall
X = [Displacement Horsepower Weight];
ens = fitensemble(X,MPG,'LSBoost',50,'Tree');
cvens = crossval(ens);
L = kfoldLoss(cvens)
L =
    25.6573
cvens = resume(cvens,50);
L = kfoldLoss(cvens)
L =
    26.7563
```

The additional training did not improve the cross-validation error.
See Also ..... fitensemble | kfoldLoss
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## Purpose <br> Ridge regression

Syntax

Description
b = ridge( $\mathrm{y}, \mathrm{x}, \mathrm{k}$ )
$\mathrm{b}=$ ridge $(\mathrm{y}, \mathrm{X}, \mathrm{k}$, scaled)
$b=r i d g e(y, x, k)$ returns a vector $b$ of coefficient estimates for $a$ multilinear ridge regression of the responses in $y$ on the predictors in $X$. X is an $n$-by- $p$ matrix of $p$ predictors at each of $n$ observations. y is an $n$-by- 1 vector of observed responses. k is a vector of ridge parameters. If k has $m$ elements, b is $p$-by- $m$. By default, b is computed after centering and scaling the predictors to have mean 0 and standard deviation 1. The model does not include a constant term, and $X$ should not contain a column of 1 s .
$b=r i d g e(y, x, k, s c a l e d)$ uses the $\{0,1\}$-valued flag scaled to determine if the coefficient estimates in $b$ are restored to the scale of the original data. ridge $(y, x, k, 0)$ performs this additional transformation. In this case, b contains $p+1$ coefficients for each value of $k$, with the first row corresponding to a constant term in the model. ridge ( $\mathrm{y}, \mathrm{X}, \mathrm{k}, 1$ ) is the same as ridge $(y, X, k)$. In this case, $b$ contains $p$ coefficients, without a coefficient for a constant term.

The relationship between $b 0=r i d g e(y, x, k, 0)$ and $b 1=$ ridge ( $y, x, k, 1$ ) is given by

```
m = mean(X);
s = std(X,0,1)';
b1_scaled = b1./s;
b0 = [mean(y)-m*b1_scaled; b1_scaled]
```

This can be seen by replacing the $x_{i}(i=1, \ldots, n)$ in the multilinear model $y=b_{0}{ }^{0}+b_{1}{ }^{0} x_{1}+\ldots+b_{n}{ }^{0} x_{n}$ with the $z$-scores $z_{i}=\left(x_{i}-\mu_{i}\right) / \sigma_{i}$, and replacing $y$ with $y-\mu_{y}$.
In general, b 1 is more useful for producing plots in which the coefficients are to be displayed on the same scale, such as a ridge trace (a plot of the regression coefficients as a function of the ridge parameter). b0 is more useful for making predictions.

Coefficient estimates for multiple linear regression models rely on the independence of the model terms. When terms are correlated and the columns of the design matrix $X$ have an approximate linear dependence, the matrix $\left(X^{T} X\right)^{-1}$ becomes close to singular. As a result, the least-squares estimate

$$
\hat{\beta}=\left(X^{T} X\right)^{-1} X^{T} y
$$

becomes highly sensitive to random errors in the observed response $y$, producing a large variance. This situation of multicollinearity can arise, for example, when data are collected without an experimental design.

Ridge regression addresses the problem by estimating regression coefficients using

$$
\hat{\beta}=\left(X^{T} X+k I\right)^{-1} X^{T} y
$$

where $k$ is the ridge parameter and $I$ is the identity matrix. Small positive values of $k$ improve the conditioning of the problem and reduce the variance of the estimates. While biased, the reduced variance of ridge estimates often result in a smaller mean square error when compared to least-squares estimates.

## Examples

Load the data in acetylene.mat, with observations of the predictor variables x 1 , x 2 , x 3 , and the response variable y :

```
load acetylene
```

Plot the predictor variables against each other:

```
subplot(1,3,1)
plot(x1,x2,'.')
xlabel('x1'); ylabel('x2'); grid on; axis square
subplot(1,3,2)
plot(x1,x3,'.')
xlabel('x1'); ylabel('x3'); grid on; axis square
```



Note the correlation between $\times 1$ and the other two predictor variables.
Use ridge and x2fx to compute coefficient estimates for a multilinear model with interaction terms, for a range of ridge parameters:

```
X = [x1 x2 x3];
D = x2fx(X,'interaction');
D(:,1) = []; % No constant term
k = 0:1e-5:5e-3;
b = ridge(y,D,k);
```

Plot the ridge trace:

```
figure
plot(k,b,'LineWidth',2)
ylim([-100 100])
grid on
xlabel('Ridge Parameter')
ylabel('Standardized Coefficient')
title('{\bf Ridge Trace}')
legend('x1','x2','x3','x1x2','x1x3','x2x3')
```



The estimates stabilize to the right of the plot. Note that the coefficient of the $\times 2 \times 3$ interaction term changes sign at a value of the ridge parameter $\approx 5 \times 10^{-4}$.

[^15]
## classregtree.risk

## Purpose Node risks

```
Syntax
\(r=r i s k(t)\)
\(r=r i s k(t, n o d e s)\)
```

Description
$r=r i s k(t)$ returns an $n$-element vector $r$ of the risk of the nodes in the tree $t$, where $n$ is the number of nodes. The risk $r(i)$ for node $i$ is the node error $e(i)$ (computed by nodeerr) weighted by the node probability $\mathrm{p}(\mathrm{i})$ (computed by nodeprob).
$r=r i s k(t, n o d e s)$ takes a vector nodes of node numbers and returns the risk values for the specified nodes.

## Examples <br> Create a classification tree for Fisher's iris data:

```
load fisheriris;
t = classregtree(meas,species,...
    'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
view(t)
```

| Click to display: Identity | $\checkmark$ | Magnification: | 100\% | $\checkmark$ | Pruning level: | 0 of 4 | - |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |



$$
\begin{aligned}
& e=\text { nodeerr }(t) ; \\
& p= \text { nodeprob }(t) ; \\
& r= \text { risk }(t) ; \\
& r= \\
& r= 0.6667 \\
& 0.3333 \\
& 0.0333
\end{aligned}
$$

## classregtree.risk

$$
\begin{array}{r}
0.0067 \\
0.0067 \\
0.0133 \\
0 \\
0 \\
\text { e. *p } \\
\text { ans }= \\
0.6667 \\
0 \\
0.3333 \\
0.0333 \\
0.0067 \\
0.0067 \\
0.0133 \\
0 \\
0
\end{array}
$$

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.
See Also classregtree | nodeerr | nodeprob

## Purpose Interactive robust regression

## Syntax <br> robustdemo <br> robustdemo(x,y)

## Description

robustdemo shows the difference between ordinary least squares and robust regression for data with a single predictor. With no input arguments, robustdemo displays a scatter plot of a sample of roughly linear data with one outlier. The bottom of the figure displays equations of lines fitted to the data using ordinary least squares and robust methods, together with estimates of the root mean squared errors.

Use the right mouse button to click on a point and view its least-squares leverage and robust weight.

Use the left mouse button to click-and-drag a point. The displays will update.
robustdemo ( $x, y$ ) uses $x$ and $y$ data vectors you supply, in place of the sample data supplied with the function.

## Examples The following steps show you how to use robustdemo.

1 Start the demo. To begin using robustdemo with the built-in data, simply type the function name:
robustdemo

## robustdemo

Use left mouse button to select and drag points
Use right mouse button to query point properties


Least squares:
$Y=-0.188327+1.10351^{*} X$
Robust:
$Y=-1.77278+1.50415^{*} X$
RMS error $=2.21375$
RMS error $=1.43663$

The resulting figure shows a scatter plot with two fitted lines. The red line is the fit using ordinary least-squares regression. The green line is the fit using robust regression. At the bottom of the figure are the equations for the fitted lines, together with the estimated root mean squared errors for each fit.

2 View leverages and robust weights. Right-click on any data point to see its least-squares leverage and robust weight:

## Use left mouse button to select and drag points <br> Use right mouse button to query point properties



Least squares:
Robust:

$$
\begin{aligned}
& Y=-0.188327+1.10351^{*} X \\
& Y=-1.77278+1.50415^{*} X
\end{aligned}
$$

$$
\text { RMS error }=2.21375
$$

$$
\text { RMS error }=1.43663
$$

In the built-in data, the right-most point has a relatively high leverage of 0.35 . The point exerts a large influence on the least-squares fit, but its small robust weight shows that it is effectively excluded from the robust fit.

3 See how changes in the data affect the fits. With the left mouse button, click and hold on any data point and drag it to a new location. When you release the mouse button, the displays update:

## robustdemo

Use left mouse button to select and drag points
Use right mouse button to query point properties


Least squares:
Robust:

$$
\begin{aligned}
& Y=-1.0661+1.33785^{*} X \\
& Y=-1.18916+1.36459^{*} X
\end{aligned}
$$

RMS error $=1.21477$
RMS error $=1.27697$

Bringing the right-most data point closer to the least-squares line makes the two fitted lines nearly identical. The adjusted right-most data point has significant weight in the robust fit.

## See Also

robustfit | leverage

## Purpose <br> Robust regression

Syntax
b = robustfit (X, y$)$
b = robustfit( $\mathrm{X}, \mathrm{y}$, wfun,tune)
b = robustfit(X,y,wfun,tune, const)
[b,stats] = robustfit(...)
Description
$\mathrm{b}=$ robustfit $(\mathrm{X}, \mathrm{y})$ returns a $p$-by- 1 vector b of coefficient estimates for a robust multilinear regression of the responses in y on the predictors in X . X is an $n$-by- $p$ matrix of $p$ predictors at each of $n$ observations. y is an $n$-by- 1 vector of observed responses. By default, the algorithm uses iteratively reweighted least squares with a bisquare weighting function.

Note By default, robustfit adds a first column of 1 s to $X$, corresponding to a constant term in the model. Do not enter a column of 1 s directly into $X$. You can change the default behavior of robustfit using the input const, below.
robustfit treats NaNs in X or y as missing values, and removes them.
$b=$ robustfit( $X, y, w f u n$, tune) specifies a weighting function wfun. tune is a tuning constant that is divided into the residual vector before computing weights.

The weighting function wfun can be any one of the following strings:

| Weight Function | Equation | Default <br> Tuning Constant |
| :---: | :---: | :---: |
| 'andrews ' | w = (abs(r)<pi) .* sin r$)$. / r | 1.339 |
| 'bisquare' <br> (default) | $\begin{aligned} & w=(\operatorname{abs}(r)<1) \cdot *(1- \\ & r . \wedge 2) .^{\wedge} \end{aligned}$ | 4.685 |
| 'cauchy ' | $w=1 . /(1+r . \wedge 2)$ | 2.385 |

## robustfit

| Weight <br> Function | Equation | Default <br> Tuning <br> Constant |
| :--- | :--- | :--- |
| 'fair' | $w=1 . /(1+\operatorname{abs}(r))$ | 1.400 |
| 'huber' | $w=1 . / \max (1$, abs $(r))$ | 1.345 |
| 'logistic' | $w=\tanh (r) . / \mathrm{r}$ | 1.205 |
| 'ols' | Ordinary least squares (no <br> weighting function) | None |
| 'talwar' | $w=1 \quad$ * abs (r)<1) | 2.795 |
| 'welsch' | $w=\exp (-(r . \wedge 2))$ | 2.985 |

If tune is unspecified, the default value in the table is used. Default tuning constants give coefficient estimates that are approximately $95 \%$ as statistically efficient as the ordinary least-squares estimates, provided the response has a normal distribution with no outliers. Decreasing the tuning constant increases the downweight assigned to large residuals; increasing the tuning constant decreases the downweight assigned to large residuals.

The value $r$ in the weight functions is

```
r = resid/(tune*s*sqrt(1-h))
```

where resid is the vector of residuals from the previous iteration, $h$ is the vector of leverage values from a least-squares fit, and $s$ is an estimate of the standard deviation of the error term given by

```
s = MAD/0.6745
```

Here MAD is the median absolute deviation of the residuals from their median. The constant 0.6745 makes the estimate unbiased for the normal distribution. If there are $p$ columns in X , the smallest $p$ absolute deviations are excluded when computing the median.

You can write your own weight function. The function must take a vector of scaled residuals as input and produce a vector of weights as
output. In this case, wfun is specified using a function handle @ (as in @myfun), and the input tune is required.
b = robustfit ( $\mathrm{X}, \mathrm{y}$, wfun, tune, const) controls whether or not the model will include a constant term. const is 'on' to include the constant term (the default), or 'off' to omit it. When const is 'on', robustfit adds a first column of 1 s to $X$. When const is 'off', robustfit does not alter X .
[b,stats] = robustfit(...) returns the structure stats, whose fields contain diagnostic statistics from the regression. The fields of stats are:

- ols_s - Sigma estimate (RMSE) from ordinary least squares
- robust_s - Robust estimate of sigma
- mad_s - Estimate of sigma computed using the median absolute deviation of the residuals from their median; used for scaling residuals during iterative fitting
- s - Final estimate of sigma, the larger of robust_s and a weighted average of ols_s and robust_s
- resid - Residual
- rstud - Studentized residual (see regress for more information)
- se - Standard error of coefficient estimates
- covb - Estimated covariance matrix for coefficient estimates
- coeffcorr - Estimated correlation of coefficient estimates
- $t$ - Ratio of $b$ to se
- $\mathrm{p}-p$-values for t
- $w$ - Vector of weights for robust fit
- $\mathrm{R}-R$ factor in $Q R$ decomposition of x
- dfe - Degrees of freedom for error
- $h$ - Vector of leverage values for least-squares fit


## robustfit

The robustfit function estimates the variance-covariance matrix of the coefficient estimates using inv ( $\mathrm{X}^{\prime *} \mathrm{X}$ )*stats. $\mathrm{s}^{\wedge} 2$. Standard errors and correlations are derived from this estimate.

## Examples

Generate data with the trend $y=10-2 * x$, then change one value to simulate an outlier:

```
x = (1:10)';
y = 10 - 2*x + randn(10,1);
y(10) = 0;
```

Use both ordinary least squares and robust regression to estimate a straight-line fit:

```
bls = regress(y,[ones(10,1) x])
bls =
    7.2481
    -1.3208
brob = robustfit(x,y)
brob =
    9.1063
-1.8231
```

A scatter plot of the data together with the fits shows that the robust fit is less influenced by the outlier than the least-squares fit:

```
scatter(x,y,'filled'); grid on; hold on
plot(x,bls(1)+bls(2)*x,'r','LineWidth',2);
plot(x,brob(1)+brob(2)*x,'g','LineWidth',2)
legend('Data','Ordinary Least Squares','Robust Regression')
```



## References

[1] DuMouchel, W. H., and F. L. O’Brien. "Integrating a Robust Option into a Multiple Regression Computing Environment." Computer Science and Statistics: Proceedings of the 21st Symposium on the Interface. Alexandria, VA: American Statistical Association, 1989.
[2] Holland, P. W., and R. E. Welsch. "Robust Regression Using Iteratively Reweighted Least-Squares." Communications in Statistics: Theory and Methods, A6, 1977, pp. 813-827.
[3] Huber, P. J. Robust Statistics. Hoboken, NJ: John Wiley \& Sons, Inc., 1981.

## robustfit

[4] Street, J. O., R. J. Carroll, and D. Ruppert. "A Note on Computing Robust Regression Estimates via Iteratively Reweighted Least Squares." The American Statistician. Vol. 42, 1988, pp. 152-154.

## See Also regress | robustdemo

## Purpose Rotate categorical matrix 90 degrees

Syntax
$B=\operatorname{rot} 90(A)$
$B=\operatorname{rot} 90(A, k)$

Description $\quad B=\operatorname{rot} 90(A)$ returns the 90 degree counterclockwise rotation of the 2-D categorical matrix $A$.
$B=\operatorname{rot} 90(A, k)$ returns the $k * 90$ degree rotation $\mathrm{of} A, k=$ $+-1,+-2, \ldots$

See Also flipdim | fliplr | flipud

## rotatefactors

Purpose Rotate factor loadings
Syntax
Description
$B=$ rotatefactors(A)
B = rotatefactors(A,'Method','orthomax','Coeff',gamma)
B = rotatefactors(A,'Method','procrustes','Target',target)
B = rotatefactors(A,'Method','pattern','Target',target)
B = rotatefactors(A,'Method','promax')
$[B, T]=$ rotatefactors( $\mathrm{A}, \ldots$ )
$\mathrm{B}=$ rotatefactors(A) rotates the $d$-by $-m$ loadings matrix A to maximize the varimax criterion, and returns the result in $B$. Rows of A and B correspond to variables and columns correspond to factors, for example, the $(i, j)$ th element of A is the coefficient for the $i$ th variable on the $j$ th factor. The matrix A usually contains principal component coefficients created with princomp or pcacov, or factor loadings estimated with factoran.

B = rotatefactors(A,'Method','orthomax','Coeff',gamma) rotates A to maximize the orthomax criterion with the coefficient gamma, i.e., $B$ is the orthogonal rotation of $A$ that maximizes

```
sum(D*sum(B.^4,1) - GAMMA*sum(B.^2,1).^2)
```

The default value of 1 for gamma corresponds to varimax rotation. Other possibilities include gamma $=0, m / 2$, and $d(m-1) /(d+m-2)$, corresponding to quartimax, equamax, and parsimax. You can also supply the strings 'varimax', 'quartimax', 'equamax', or 'parsimax' for the 'method' parameter and omit the 'Coeff' parameter.

If 'Method' is 'orthomax', 'varimax', 'quartimax', 'equamax', or 'parsimax', then additional parameters are

- 'Normalize ' - Flag indicating whether the loadings matrix should be row-normalized for rotation. If ' on' (the default), rows of A are normalized prior to rotation to have unit Euclidean norm, and unnormalized after rotation. If 'off', the raw loadings are rotated and returned.
- 'Reltol' - Relative convergence tolerance in the iterative algorithm used to find $T$. The default is sqrt(eps).
- 'Maxit' - Iteration limit in the iterative algorithm used to find $T$. The default is 250 .

B = rotatefactors(A,'Method','procrustes','Target',target) performs an oblique procrustes rotation of A to the $d$-by- $m$ target loadings matrix target.

B = rotatefactors(A,'Method','pattern','Target',target) performs an oblique rotation of the loadings matrix A to the $d$-by-m target pattern matrix target, and returns the result in $B$. target defines the "restricted" elements of B, i.e., elements of B corresponding to zero elements of target are constrained to have small magnitude, while elements of B corresponding to nonzero elements of target are allowed to take on any magnitude.

If 'Method' is 'procrustes' or 'pattern', an additional parameter is 'Type', the type of rotation. If 'Type' is 'orthogonal', the rotation is orthogonal, and the factors remain uncorrelated. If 'Type' is 'oblique' (the default), the rotation is oblique, and the rotated factors might be correlated.

When 'Method' is 'pattern', there are restrictions on target. If A has $m$ columns, then for orthogonal rotation, the $j$ th column of target must contain at least $m-j$ zeros. For oblique rotation, each column of target must contain at least $m-1$ zeros.

B = rotatefactors(A,'Method','promax') rotates A to maximize the promax criterion, equivalent to an oblique Procrustes rotation with a target created by an orthomax rotation. Use the four orthomax parameters to control the orthomax rotation used internally by promax.

An additional parameter for 'promax' is 'Power', the exponent for creating promax target matrix. 'Power' must be 1 or greater. The default is 4.
$[B, T]=$ rotatefactors $(A, \ldots)$ returns the rotation matrix $T$ used to create $B$, that is, $B=A * T$. $\operatorname{inv}\left(T^{\prime} * T\right)$ is the correlation matrix of the
rotated factors. For orthogonal rotation, this is the identity matrix, while for oblique rotation, it has unit diagonal elements but nonzero off-diagonal elements.

```
Examples
Promax rotation:
\% first three factors.
LFA = factoran(X,3,'Rotate','none');
\([L 3, T]=\) rotatefactors(LFA(:,1:3),...
    'method','promax',...
    'power',2);
\% Pattern rotation:
\% first three factors.
Tgt \(=\left[\begin{array}{llllllllll}1 & 1 & 1 & 1 & 1 & 0 & 1 & 0 & 1 & 1 ;\end{array} .\right.\).
    \(0001110000 ; \ldots\)
    \(1001011110]\) ';
[L4,T] = rotatefactors(LFA(:,1:3),...
    'method','pattern',...
    'target',Tgt);
inv(T'*T) \% Correlation matrix of the rotated factors
References
[1] Harman, H. H. Modern Factor Analysis. 3rd ed. Chicago: University of Chicago Press, 1976.
```

[2] Lawley, D. N., and A. E. Maxwell. Factor Analysis as a Statistical Method. 2nd ed. New York: American Elsevier Publishing, 1971.

See Also<br>biplot | factoran | princomp | pcacov | procrustes

## Purpose Row exchange

Syntax

```
dRE = rowexch(nfactors,nruns)
[dRE,X] = rowexch(nfactors,nruns)
[dRE,X] = rowexch(nfactors,nruns,model)
[dRE,X] = rowexch(...,param1,val1,param2,val2,...)
```


## Description

dRE = rowexch(nfactors, nruns) uses a row-exchange algorithm to generate a $D$-optimal design dRE with nruns runs (the rows of dRE) for a linear additive model with nfactors factors (the columns of dRE). The model includes a constant term.
[dRE, X] = rowexch(nfactors, nruns) also returns the associated design matrix $X$, whose columns are the model terms evaluated at each treatment (row) of dRE.
[dRE,X] = rowexch(nfactors, nruns,model) uses the linear regression model specified in model. model is one of the following strings:

- 'linear' - Constant and linear terms. This is the default.
- 'interaction' - Constant, linear, and interaction terms
- 'quadratic ' - Constant, linear, interaction, and squared terms
- 'purequadratic' - Constant, linear, and squared terms

The order of the columns of X for a full quadratic model with $n$ terms is:
1 The constant term
2 The linear terms in order $1,2, \ldots, n$
3 The interaction terms in order (1, 2), (1, 3), $\ldots,(1, n),(2,3), \ldots,(n-1, n)$
4 The squared terms in order $1,2, \ldots, n$
Other models use a subset of these terms, in the same order.

Alternatively, model can be a matrix specifying polynomial terms of arbitrary order. In this case, model should have one column for each factor and one row for each term in the model. The entries in any row of model are powers for the factors in the columns. For example, if a model has factors $\mathrm{X} 1, \mathrm{X} 2$, and X 3 , then a row [ $\left.\begin{array}{lll}0 & 1 & 2\end{array}\right]$ in model specifies the term (X1.^0).*(X2.^1).*(X3.^2). A row of all zeros in model specifies a constant term, which can be omitted.
[dRE,X] = rowexch(...,param1,val1,param2,val2,...) specifies additional parameter/value pairs for the design. Valid parameters and their values are listed in the following table.

| Parameter | Value |
| :--- | :--- |
| 'bounds ' | Lower and upper bounds for each factor, specified as <br> a 2-by-nfactors matrix. Alternatively, this value <br> can be a cell array containing nfactors elements, <br> each element specifying the vector of allowable <br> values for the corresponding factor. |
| 'categorical' | Indices of categorical predictors. |
| 'display' | Either ' on ' or 'off ' to control display of the <br> iteration counter. The default is 'on'. |
| 'excludefun' | Handle to a function that excludes undesirable <br> runs. If the function is $f$, it must support the syntax <br> $b=f(S)$, where $S$ is a matrix of treatments with <br> nfactors columns and $b$ is a vector of Boolean <br> values with the same number of rows as $S . b(i)$ is <br> true if the $i$ th row $S$ should be excluded. |
| 'init' | Initial design as an nruns-by-nfactors matrix. The <br> default is a randomly selected set of points. |
| 'levels' | Vector of number of levels for each factor. |
| 'maxiter' | Maximum number of iterations. The default is 10. |


| Parameter | Value |
| :--- | :--- |
| options | A structure that specifies whether to run in parallel, <br> and specifies the random stream or streams. Create <br> the options structure with statset. Option fields: |
| - UseParallel - Set to ' always' to compute in <br> parallel. Default is 'never'. <br> - UseSubstreams — Set to 'always' to compute <br> in parallel in a reproducible fashion. Default is <br> 'never'. To compute reproducibly, set Streams <br> to a type allowing substreams: 'mlfg6331_64' <br> or 'mrg32k3a'. <br> - Streams - A RandStream object or cell array <br> of such objects. If you do not specify Streams, <br> rowexch uses the default stream or streams. If <br> you choose to specify Streams, use a single object <br> except in the case <br> - You have an open MATLAB pool |  |
| - UseParallel is 'always' |  |
| - UseSubstreams is 'never' |  |

Both cordexch and rowexch use iterative search algorithms. They operate by incrementally changing an initial design matrix $X$ to increase $D=\left|X^{T} X\right|$ at each step. In both algorithms, there is randomness
built into the selection of the initial design and into the choice of the incremental changes. As a result, both algorithms may return locally, but not globally, $D$-optimal designs. Run each algorithm multiple times and select the best result for your final design. Both functions have a 'tries ' parameter that automates this repetition and comparison.

At each step, the row-exchange algorithm exchanges an entire row of $X$ with a row from a design matrix $C$ evaluated at a candidate set of feasible treatments. The rowexch function automatically generates a $C$ appropriate for a specified model, operating in two steps by calling the candgen and candexch functions in sequence. Provide your own $C$ by calling candexch directly. In either case, if $C$ is large, its static presence in memory can affect computation.

## Examples

Suppose you want a design to estimate the parameters in the following three-factor, seven-term interaction model:

$$
y=\beta_{0}+\beta_{1} x_{1}+\beta_{2} x_{2}+\beta_{3} x_{3}+\beta_{12} x_{1} x_{2}+\beta_{13} x_{1} x_{3}+\beta_{23} x_{2} x_{3}+\varepsilon
$$

Use rowexch to generate a $D$-optimal design with seven runs:

```
nfactors = 3;
nruns = 7;
[dRE,X] = rowexch(nfactors,nruns,'interaction','tries',10)
dRE =
\begin{tabular}{rrrrrr}
-1 & -1 & 1 & & & \\
1 & -1 & 1 & & & \\
1 & -1 & -1 & & & \\
1 & 1 & 1 & & & \\
-1 & -1 & -1 & & & \\
-1 & 1 & -1 & & & \\
\hline-1 & 1 & 1 & & & -1 \\
\hline & & & & 1 & -1 \\
1 & -1 & -1 & 1 & 1 & -1 \\
1 & 1 & -1 & 1 & -1 & 1 \\
1 & 1 & -1 & -1 & -1 & -1 \\
1 & 1 & 1 & 1 & 1 & 1
\end{tabular}
```

| 1 | -1 | 1 | -1 | -1 | 1 | -1 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | -1 | 1 | 1 | -1 | -1 | 1 |

Columns of the design matrix X are the model terms evaluated at each row of the design dRE. The terms appear in order from left to right: constant term, linear terms (1, 2, 3), interaction terms (12, 13, 23). Use $X$ to fit the model, as described in "Linear Regression" on page 9-3, to response data measured at the design points in dRE.

## Purpose Interactive response surface demonstration

## Syntax rsmdemo

Description
rsmdemo opens a group of three graphical user interfaces for interactively investigating response surface methodology (RSM), nonlinear fitting, and the design of experiments.

The interfaces allow you to collect and model data from a simulated chemical reaction. Experimental predictors are concentrations of three reactants (hydrogen, $n$-Pentane, and isopentane) and the response is the reaction rate. The reaction rate is simulated by a Hougen-Watson model (Bates and Watts, [2], pp. 271-272):

$$
\text { rate }=\frac{\beta_{1} x_{2}-x_{3} / \beta_{5}}{1+\beta_{2} x_{1}+\beta_{3} x_{2}+\beta_{4} x_{3}}
$$

where rate is the reaction rate, $x_{1}, x_{2}$, and $x_{3}$ are the concentrations of hydrogen, $n$-pentane, and isopentane, respectively, and $\beta_{1}, \beta_{2}, \ldots, \beta_{5}$ are fixed parameters. Random errors are used to perturb the reaction rate for each combination of reactants.

Collect data using one of two methods:
1 Manually set reactant concentrations in the Reaction Simulator interface by editing the text boxes or by adjusting the associated sliders.


When you click Run, the concentrations and simulated reaction rate are recorded on the Trial and Error Data interface.


You are allowed up to 13 independent experimental runs for data collection.

2 Use a designed experiment to set reactant concentrations in the Experimental Data interface by clicking the Do Experiment button.


A 13 -run $D$-optimal design for a full quadratic model is generated by the cordexch function, and the concentrations and simulated reaction rates are recorded on the same interface.


Once data is collected, scatter plots of reaction rates vs. individual predictors are generated by selecting one of the following from the Plot pop-up menu below the recorded data:

- Hydrogen vs. Rate
- n-Pentane vs. Rate
- Isopentane vs. Rate

Fit a response surface model to the data by clicking the Analyze button below the trial-and-error data or the Response Surface button below the experimental data. Both buttons load the data into the Response Surface Tool rstool. By default, trial-and-error data is fit with a linear additive model and experimental data is fit with a full quadratic model, but the models can be adjusted in the Response Surface Tool.

For experimental data, you have the additional option of fitting a Hougen-Watson model. Click the Nonlinear Model button to load the data and the model in hougen into the Nonlinear Fitting Tool nlintool.

See Also hougen | cordexch | rstool | nlintool

Purpose Interactive response surface modeling

| Syntax | rstool |
| :---: | :---: |
|  | rstool(X, Y,model) |
|  | rstool(x,y,model, alpha) |
|  | rstool(x,y,model, alpha, xname, yname) |

## Description

rstool opens a graphical user interface for interactively investigating one-dimensional contours of multidimensional response surface models.


By default, the interface opens with the data from hald.mat and a fitted response surface with constant, linear, and interaction terms.

A sequence of plots is displayed, each showing a contour of the response surface against a single predictor, with all other predictors held fixed.
rstool plots a $95 \%$ simultaneous confidence band for the fitted response surface as two red curves. Predictor values are displayed in the text boxes on the horizontal axis and are marked by vertical dashed blue lines in the plots. Predictor values are changed by editing the text boxes or by dragging the dashed blue lines. When you change the value of a predictor, all plots update to show the new point in predictor space.

The pop-up menu at the lower left of the interface allows you to choose among the following models:

- Linear - Constant and linear terms (the default)
- Pure Quadratic - Constant, linear, and squared terms
- Interactions - Constant, linear, and interaction terms
- Full Quadratic - Constant, linear, interaction, and squared terms

Click Export to open the following dialog box:


The dialog allows you to save information about the fit to MATLAB workspace variables with valid names.
rstool(X,Y,model) opens the interface with the predictor data in $X$, the response data in $Y$, and the fitted model model. Distinct predictor variables should appear in different columns of $X$. $Y$ can be a vector, corresponding to a single response, or a matrix, with columns corresponding to multiple responses. Y must have as many elements (or rows, if it is a matrix) as $X$ has rows.

The optional input model can be any one of the following strings:

- 'linear' - Constant and linear terms (the default)
- 'purequadratic' - Constant, linear, and squared terms
- 'interaction' - Constant, linear, and interaction terms
- 'quadratic' - Constant, linear, interaction, and squared terms

To specify a polynomial model of arbitrary order, or a model without a constant term, use a matrix for model as described in x2fx.
rstool(x,y,model, alpha) uses 100(1-alpha)\% global confidence intervals for new observations in the plots.
rstool( $x, y$, model, alpha, xname, yname) labels the axes using the strings in xname and yname. To label each subplot differently, xname and yname can be cell arrays of strings.

## Examples

The following uses rstool to visualize a quadratic response surface model of the 3-D chemical reaction data in reaction.mat:

```
load reaction
alpha = 0.01; % Significance level
rstool(reactants,rate,'quadratic',alpha, xn,yn)
```



The rstool interface is used by rsmdemo to visualize the results of simulated experiments with data like that in reaction.mat. As described in "Response Surface Designs" on page 15-9, rsmdemo uses a response surface model to generate simulated data at combinations of predictors specified by either the user or by a designed experiment.

See Also $\quad x 2 f x$ | rsmdemo | nlintool

## Purpose Run test for randomness

```
Syntax
h = runstest(x)
h = runstest(x, v)
h = runstest(x,'ud')
h = runstest(...,param1,val1,param2,val2,...)
[h, p] = runstest(...)
[h,p,stats] = runstest(...)
```


## Description

$\mathrm{h}=$ runstest(x) performs a runs test on the sequence of observations in the vector $x$. This is a test of the null hypothesis that the values in $x$ come in random order, against the alternative that they do not. The test is based on the number of runs of consecutive values above or below the mean of $x$. Too few runs indicate a tendency for high and low values to cluster. Too many runs indicate a tendency for high and low values to alternate. The test returns the logical value $h=1$ if it rejects the null hypothesis at the $5 \%$ significance level, and $\mathrm{h}=0$ if it cannot. The test treats NaN values in x as missing values, and ignores them.
runstest uses a test statistic which is approximately normally distributed when the null hypothesis is true. It is the difference between the number of runs and its mean, divided by its standard deviation.
$\mathrm{h}=$ runstest $(\mathrm{x}, \mathrm{v})$ performs the test using runs above or below the value $v$. Values exactly equal to $v$ are discarded. The default value of $v$ is the mean of $x$.
$h=r u n s t e s t\left(x,{ }^{\prime} u d '\right)$ performs a test for the number of runs up or down. This also tests the hypothesis that the values in $x$ come in random order. Too few runs indicate a trend. Too many runs indicate an oscillation. Values exactly equal to the preceding value are discarded.
h = runstest(...,param1,val1,param2,val2,...) specifies additional parameters and their values. Valid parameter/value pairs are the following:

- 'alpha' - A scalar giving the significance level of the test
- 'method ' - Either 'exact' to compute the $p$ value using an exact algorithm, or 'approximate' to use a normal approximation. The default is 'exact', except for runs up/down when the length of is 51 or more. The 'exact' method is not available for runs up/down when the length of $x$ is over 50 .
- 'tail' - Performs the test against one of the following alternative hypotheses:
- 'both' - two-tailed test (sequence is not random)
- 'right' - right-tailed test (like values separate for runs above/below, direction alternates for runs up/down)
- 'left' - left-tailed test (like values cluster for runs above/below, values trend for runs up/down)
$[\mathrm{h}, \mathrm{p}]=$ runstest $(\ldots)$ returns the $p$ value of the test. The output p is computed from either the test statistic or the exact distribution of the number of runs, depending on the value of the 'method' parameter.
[h,p,stats] = runstest(...) returns a structure stats with the following fields:
- nruns - The number of runs
- n 1 - The number of values above v
- no - The number of values below v
- z - The test statistic


## Examples

```
\(x=\operatorname{randn}(40,1)\);
    [h,p] = runstest(x,median(x))
    h =
            0
    p \(=\)
        0.6286
```

signrank | signtest

## TreeBagger.SampleWithReplacement property

Purpose Flag to sample with replacement
$\begin{array}{ll}\text { Description } & \begin{array}{l}\text { The SampleWithReplacement property is a logical flag specifying if } \\ \text { data are sampled for each decision tree with replacement. True if }\end{array} \\ & \begin{array}{l}\text { TreeBagger samples data with replacement and false otherwise. True } \\ \text { by default. }\end{array}\end{array}$

## Purpose <br> Syntax <br> Description

Sample size and power of test

```
n = sampsizepwr(testtype,p0,p1)
n = sampsizepwr(testtype,p0,p1,power)
power = sampsizepwr(testtype,p0,p1,[],n)
p1 = sampsizepwr(testtype,p0,[],power,n)
[...] = sampsizepwr(...,n,param1,val1,param2,val2,...)
```

$\mathrm{n}=$ sampsizepwr(testtype, p0,p1) returns the sample size n required
for a two-sided test of the specified type to have a power (probability of rejecting the null hypothesis when the alternative hypothesis is true) of 0.90 when the significance level (probability of rejecting the null hypothesis when the null hypothesis is true) is 0.05 . p0 specifies parameter values under the null hypothesis. p1 specifies the single parameter value being tested under the alternative hypothesis.

The following values are available for testtype:

- 'z' - z-test for normally distributed data with known standard deviation. pO is a two-element vector [mu0 sigma0] of the mean and standard deviation, respectively, under the null hypothesis. p1 is the value of the mean under the alternative hypothesis.
- ' t ' - $t$-test for normally distributed data with unknown standard deviation. p0 is a two-element vector [mu0 sigma0] of the mean and standard deviation, respectively, under the null hypothesis. p1 is the value of the mean under the alternative hypothesis.
- 'var' - Chi-square test of variance for normally distributed data. p 0 is the variance under the null hypothesis. p 1 is the variance under the alternative hypothesis.
- ' p ' - Test of the $p$ parameter (success probability) for a binomial distribution. p 0 is the value of $p$ under the null hypothesis. p 1 is the value of $p$ under the alternative hypothesis.

The ' $p$ ' test is a discrete test for which increasing the sample size does not always increase the power. For $n$ values larger than 200,
there may be values smaller than the returned n value that also produce the desired size and power.
$\mathrm{n}=$ sampsizepwr(testtype, $\mathrm{p} 0, \mathrm{p} 1$, power) returns the sample size n such that the power is power for the parameter value p 1 .
power = sampsizepwr(testtype, p0, p1, [],n) returns the power achieved for a sample size of $n$ when the true parameter value is $p 1$.
p1 = sampsizepwr(testtype, p0,[], power,n) returns the parameter value detectable with the specified sample size n and power power.

When computing p 1 for the ' p ' test, if no alternative can be rejected for a given null hypothesis and significance level, the function displays a warning message and returns NaN .
[...] = sampsizepwr(...,n,param1,val1,param2,val2,...) specifies one or more of the following name/value pairs:

- 'alpha' - Significance level of the test (default 0.05)
- 'tail' - The type of test is one of the following:
- 'both' - Two-sided test for an alternative not equal to po
- 'right ' - One-sided test for an alternative larger than po
- 'left' - One-sided test for an alternative smaller than po


## Examples Compute the mean closest to 100 that can be determined to be

 significantly different from 100 using a $t$-test with a sample size of 60 and a power of 0.8.```
mu1 = sampsizepwr('t',[100 10],[],0.8,60)
mu1 =
    103.6770
```

Compute the sample size $n$ required to distinguish $p=0.26$ from $p=0.6$ with a binomial test. The result is approximate, so make a plot to see if any smaller $n$ values also have the required power of 0.5 .

```
    napprox = sampsizepwr('p',0.2,0.26,0.6)
    Warning: Values N>200 are approximate. Plotting the power as a function
    of N may reveal lower N values that have the required power.
    napprox =
    244
    nn = 1:250;
    pwr = sampsizepwr('p',0.2,0.26,[],nn);
    nexact = min(nn(pwr>=0.6))
    nexact =
    213
```

    plot(nn, pwr,'b-',[napprox nexact],pwr([napprox nexact]),'ro');
    grid on


## sampsizepwr

See Also vartest | ttest | ztest | binocdf

## Purpose <br> Syntax <br> Description

## Examples

Scatter plot with marginal histograms
scatterhist(x,y)
h = scatterhist(...)
scatterhist(...,'param1', val1,'param2', val2,...)
scatterhist ( $\mathrm{x}, \mathrm{y}$ ) creates a 2-D scatterplot of the data in the vectors x and $y$, and puts a univariate histogram on the horizontal and vertical axes of the plot. $x$ and $y$ must be the same length.

The function is useful for viewing properties of random samples produced by functions such as copularnd, mvnrnd, lhsdesign.
$h=$ scatterhist (...) returns a vector of three axes handles for the scatterplot, the histogram along the horizontal axis, and the histogram along the vertical axis, respectively.
scatterhist(...,'param1',val1,'param2',val2,...) specifies additional parameter name/values pairs to control how the plot is made. Valid parameters are the following:

- 'NBins ' - A scalar or a two-element vector specifying the number of bins for the $X$ and $Y$ histograms. The default is to compute the number of bins using Scott's rule based on the sample standard deviation.
- 'Location' - A string controlling the location of the marginal histograms within the figure. 'SouthWest' (the default) plots the histograms below and to the left of the scatterplot, 'SouthEast ' plots them below and to the right, 'NorthEast' above and to the right, and 'NorthWest' above and to the left.
- 'Direction' - A string controlling the directation of the marginal histograms in the figure. 'in' (the default) plots the histograms with bars directed in towards the scatterplot, 'out' plots the histograms with bars directed out away from the scatterplot.


## Example 1

Independent normal and lognormal random samples:

```
x = randn(1000,1);
y = exp(.5*randn(1000,1));
scatterhist(x,y)
```



## Example 2

Marginal uniform samples that are not independent:

```
u = copularnd('Gaussian',.8,1000);
scatterhist(u(:,1),u(:,2),'Direction','out')
```



## Example 3

Mixed discrete and continuous data:

```
load('carsmall');
scatterhist(Weight,Cylinders,'NBins',[10 3],'Direction','out')
```



See Also
scatter | hist

Purpose
Scramble quasi-random point set
Syntax
ps = scramble(p,type)
ps $=$ scramble( $p$, 'clear')
ps = scramble(p)

## Examples

$p s=\operatorname{scramble}(p$, type $)$ returns a scrambled copy $p s$ of the point set $p$ of the qrandset class, created using the scramble type specified in the string type. Point sets from different subclasses of qrandset support different scramble types, as indicated in the following table.

| Subclass | Scramble Types |
| :--- | :--- |
| haltonset | 'RR2' — A permutation of the radical inverse <br> coefficients derived by applying a reverse-radix <br> operation to all of the possible coefficient values. <br> The scramble is described in [1]. |
| sobolset | 'MatousekAffineOwen' - A random linear <br> scramble combined with a random digital shift. <br> The scramble is described in [2]. |

$p s=\operatorname{scramble}(p$, 'clear') removes all scramble settings from $p$ and returns the result in ps .
$p s=s c r a m b l e(p)$ removes all scramble settings from $p$ and then adds them back in the order they were originally applied. This typically results in a different point set because of the randomness of the scrambling algorithms.

Use haltonset to generate a 3-D Halton point set, skip the first 1000 values, and then retain every 101st point:

```
p = haltonset(3,'Skip',1e3,'Leap',1e2)
p=
    Halton point set in 3 dimensions (8.918019e+013 points)
    Properties:
            Skip : }100
```


## qrandset.scramble

Leap : 100
ScrambleMethod : none
Use scramble to apply reverse-radix scrambling:

```
p = scramble(p,'RR2')
p =
    Halton point set in 3 dimensions (8.918019e+013 points)
    Properties:
        Skip : 1000
        Leap : 100
    ScrambleMethod : RR2
```

Use net to generate the first four points:

```
XO = net(p,4)
XO =
    0.0928 0.6950 0.0029
    0.6958 0.2958 0.8269
    0.3013 0.6497 0.4141
    0.9087 0.7883 0.2166
```

Use parenthesis indexing to generate every third point, up to the 11th point:

```
\(x=p(1: 3: 11,:)\)
\(x=\)
\begin{tabular}{lll}
0.0928 & 0.6950 & 0.0029 \\
0.9087 & 0.7883 & 0.2166 \\
0.3843 & 0.9840 & 0.9878 \\
0.6831 & 0.7357 & 0.7923
\end{tabular}
```


## References

[1] Kocis, L., and W. J. Whiten. "Computational Investigations of Low-Discrepancy Sequences." ACM Transactions on Mathematical Software. Vol. 23, No. 2, 1997, pp. 266-294.
[2] Matousek, J. "On the L2-Discrepancy for Anchored Boxes." Journal of Complexity. Vol. 14, No. 4, 1998, pp. 527-556.

See Also haltonset | sobolset

## qrandset.ScrambleMethod property

Purpose
Description

## Examples

See Also

Settings that control scrambling
The ScrambleMethod property contains a structure that defines which scrambles to apply to the sequence. The structure consists of two fields:

- Type: A string containing the name of the scramble.
- Options: A cell array of parameter values for the scramble.

Different point sets support different scramble types as outlined in the help for each point set class. An error occurs if you set an invalid scramble type for a given point set.

The ScrambleMethod property also accepts an empty matrix as a value. This will clear all scrambling and set the property to contain a (0x0) structure.

The scramble method provides an alternative, easier way to set scrambles.

Apply a random linear scramble combined with a random digital shift to a sobolset point set class:

```
P = sobolset(5);
P = scramble(P, 'MatousekAffineOwen');
P.ScrambleMethod
```

sobolset | scramble
Purpose Segments containing values
Syntax S = segment(obj, X, P)
Description $S=\operatorname{segment}(o b j, X, P)$ returns an array $S$ of integers indicating whichsegment of the piecewise distribution object obj contains each valueof $X$ or, alternatively, $P$. One of $X$ and $P$ must be empty ([ ]). If $X$ isnonempty, $S$ is determined by comparing $X$ with the quantile boundaryvalues defined for obj. If $P$ is nonempty, $S$ is determined by comparing$P$ with the probability boundary values.
Examples Fit Pareto tails to a $t$ distribution at cumulative probabilities 0.1 and 0.9:

```
t = trnd(3,100,1);
obj = paretotails(t,0.1,0.9);
pvals = 0:0.2:1;
s = segment(obj,[],pvals)
s =
    1 2 2 2 2 2 0
```

See Also paretotails | boundary | nsegments

Purpose
Syntax

## Description

Sequential feature selection

```
inmodel = sequentialfs(fun, X,y)
```

inmodel = sequentialfs(fun, X,y)
inmodel = sequentialfs(fun, X,Y,Z,···.)
inmodel = sequentialfs(fun, X,Y,Z,···.)
[inmodel,history] = sequentialfs(fun,X,...)
[inmodel,history] = sequentialfs(fun,X,...)
[] = sequentialfs(...,param1,val1,param2,val2,...)

```
[] = sequentialfs(...,param1,val1,param2,val2,...)
```

inmodel = sequentialfs(fun, $X, y$ ) selects a subset of features from the data matrix $X$ that best predict the data in $y$ by sequentially selecting features until there is no improvement in prediction. Rows of $X$ correspond to observations; columns correspond to variables or features. $y$ is a column vector of response values or class labels for each observation in $X$. $X$ and $y$ must have the same number of rows. fun is a function handle to a function that defines the criterion used to select features and to determine when to stop. The output inmodel is a logical vector indicating which features are finally chosen.
Starting from an empty feature set, sequentialfs creates candidate feature subsets by sequentially adding each of the features not yet selected. For each candidate feature subset, sequentialfs performs 10 -fold cross-validation by repeatedly calling fun with different training subsets of $X$ and $y$, XTRAIN and ytrain, and test subsets of $X$ and $y$, XTEST and ytest, as follows:

```
criterion = fun(XTRAIN,ytrain,XTEST,ytest)
```

XTRAIN and ytrain contain the same subset of rows of $X$ and $Y$, while XTEST and ytest contain the complementary subset of rows. XTRAIN and XTEST contain the data taken from the columns of $X$ that correspond to the current candidate feature set.

Each time it is called, fun must return a scalar value criterion. Typically, fun uses XTRAIN and ytrain to train or fit a model, then predicts values for XTEST using that model, and finally returns some measure of distance, or loss, of those predicted values from ytest. In the cross-validation calculation for a given candidate feature set, sequentialfs sums the values returned by fun and divides that sum
by the total number of test observations. It then uses that mean value to evaluate each candidate feature subset.

Typical loss measures include sum of squared errors for regression models (sequentialfs computes the mean-squared error in this case), and the number of misclassified observations for classification models (sequentialfs computes the misclassification rate in this case).

Note sequentialfs divides the sum of the values returned by fun across all test sets by the total number of test observations. Accordingly, fun should not divide its output value by the number of test observations.

After computing the mean criterion values for each candidate feature subset, sequentialfs chooses the candidate feature subset that minimizes the mean criterion value. This process continues until adding more features does not decrease the criterion.
inmodel = sequentialfs(fun, $X, Y, Z, \ldots$ ) allows any number of input variables $X, Y, Z, \ldots$. sequentialfs chooses features (columns) only from $X$, but otherwise imposes no interpretation on $X, Y, Z, \ldots$. All data inputs, whether column vectors or matrices, must have the same number of rows. sequentialfs calls fun with training and test subsets of $X, Y, Z, \ldots$ as follows:

```
criterion \(=\) fun(XTRAIN, YTRAIN,ZTRAIN, ...,
    XTEST, YTEST, ZTEST, . . .)
```

sequentialfs creates XTRAIN, YTRAIN, ZTRAIN, ... , XTEST, YTEST, ZTEST, ... by selecting subsets of the rows of $X, Y, Z, \ldots$. fun must return a scalar value criterion, but may compute that value in any way. Elements of the logical vector inmodel correspond to columns of $X$ and indicate which features are finally chosen.
[inmodel,history] = sequentialfs(fun, X, ...) returns information on which feature is chosen at each step. history is a scalar structure with the following fields:

- Crit - A vector containing the criterion values computed at each step.
- In - A logical matrix in which row i indicates the features selected at step i.
[] = sequentialfs(...,param1,val1,param2,val2,...) specifies optional parameter name/value pairs from the following table.

| Parameter | Value |
| :--- | :--- |
| 'cv' | The validation method used to compute the <br> criterion for each candidate feature subset. |
|  | - When the value is a positive integer $k$, <br> sequentialfs uses k-fold cross-validation <br> without stratification. |
|  | - When the value is an object of the cvpartition <br> class, other forms of cross-validation can be <br> specified. |
|  | - When the value is 'resubstitution ', the <br> original data are passed to fun as both the <br> training and test data to compute the criterion. |
|  | When the value is 'none ', sequentialfs <br> calls fun as criterion $=$ fun $(X, Y, Z, \ldots)$, <br> without separating test and training sets. |
|  |  |

The default value is 10 , that is, 10 -fold cross-validation without stratification.

So-called wrapper methods use a function fun that implements a learning algorithm. These methods usually apply cross-validation to select features. So-called filter methods use a function

| Parameter | Value |
| :--- | :--- |
|  | fun that measures characteristics of the data <br> (such as correlation) to select features. |
| 'mcreps' | A positive integer indicating the number of <br> Monte-Carlo repetitions for cross-validation. The <br> default value is 1. The value must be 1 if the <br> value of 'cv' is 'resubstitution' or 'none '. |
| 'direction' | The direction of the sequential search. The <br> default is 'forward '. A value of 'backward ' <br> specifies an initial candidate set including all <br> features and an algorithm that removes features <br> sequentially until the criterion increases. |
| 'keepin' | A logical vector or a vector of column numbers <br> specifying features that must be included. The <br> default is empty. |
| 'keepout' | A logical vector or a vector of column numbers <br> specifying features that must be excluded. The <br> default is empty. |
| 'nfeatures' | The number of features at which sequentialfs <br> should stop. inmodel includes exactly this <br> many features. The default value is empty, <br> indicating that sequentialfs should stop when <br> a local minimum of the criterion is found. A <br> nonempty value overrides values of 'MaxIter' <br> and 'TolFun' in 'options '. |
| 'nullmodel' | A logical value, indicating whether or not the null <br> model (containing no features from X) should be <br> included in feature selection and in the history <br> output. The default is false. |


| Parameter | Value |
| :---: | :---: |
| 'options' | Options structure for the iterative sequential search algorithm, as created by statset. sequentialfs uses the following statset parameters: <br> - Display - Amount of information displayed by the algorithm. The default is 'off'. <br> - MaxIter - Maximum number of iterations allowed. The default is Inf. <br> - TolFun - Termination tolerance for the objective function value. The default is $1 \mathrm{e}-6$ if 'direction' is 'forward'; 0 if 'direction' is 'backward'. <br> - TolTypeFun - Use absolute or relative objective function tolerances. The default is 'rel'. <br> - UseParallel - Set to 'always' to compute in parallel. Default is 'never'. <br> - UseSubstreams - Set to 'always' to compute in parallel in a reproducible fashion. Default is 'never'. To compute reproducibly, set Streams to a type allowing substreams: 'mlfg6331_64' or 'mrg32k3a'. <br> - Streams - A RandStream object or cell array consisting of one such object. If you do not specify Streams, sequentialfs uses the default stream. <br> For more information on using parallel computing, see Chapter 17, "Parallel Statistics". |

Examples
Perform sequential feature selection for classification of noisy features:

```
load fisheriris;
X = randn(150,10);
X(:,[11 3 5 7 ])= meas;
y = species;
c = cvpartition(y,'k',10);
opts = statset('display','iter');
fun = @(XT,yT,Xt,yt)...
        (sum(~strcmp(yt,classify(Xt,XT,yT,'quadratic'))));
[fs,history] = sequentialfs(fun,X,y,'cv',c,'options',opts)
Start forward sequential feature selection:
Initial columns included: none
Columns that can not be included: none
Step 1, added column 7, criterion value 0.04
Step 2, added column 5, criterion value 0.0266667
Final columns included: 5 7
fs =
            0
history =
            In: [2x10 logical]
            Crit: [0.0400 0.0267]
history.In
ans =
\begin{tabular}{llllllllll}
0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 1 & 0 & 1 & 0 & 0 & 0
\end{tabular}
```

See Also
Tutorials
How To
crossval | cvpartition | stepwisefit | statset

- "Example: Sequential Feature Selection" on page 10-24
- "Sequential Feature Selection" on page 10-23

Purpose Set and display properties

```
Syntax
set (A)
set(A,PropertyName)
A = set(A,PropertyName,PropertyValue,...)
B \(=\operatorname{set}(A\), PropertyName, value)
```


## Description

$\operatorname{set}(A)$ displays all properties of the dataset array $A$ and their possible values.
set (A,PropertyName) displays possible values for the property specified by the string PropertyName.

A $=\operatorname{set}(\mathrm{A}$, PropertyName, PropertyValue,...) sets property name/value pairs.
$B=\operatorname{set}(A$, PropertyName, value) returns a dataset array $B$ that is a copy of A, but with the property 'PropertyName' set to the value value.

Note Using set(A, 'PropertyName', value) without assigning to a variable does not modify A's properties. Use A = set (A, 'PropertyName', value) to modify A.

## Examples

Create a dataset array from Fisher's iris data and add a description:

```
load fisheriris
NumObs = size(meas,1);
NameObs = strcat({'Obs'},num2str((1:NumObs)','%-d'));
iris = dataset({nominal(species),'species'},...
    {meas,'SL','SW','PL','PW'},...
    'ObsNames',NameObs);
iris = set(iris,'Description','Fisher''s Iris Data');
get(iris)
    Description: 'Fisher's Iris Data'
    Units: {}
    DimNames: {'Observations' 'Variables'}
```

UserData: []
ObsNames: \{150x1 cell\}
VarNames: \{'species' 'SL' 'SW' 'PL' 'PW'\}
See Also get | summary

## CompactTreeBagger.SetDefaultYfit

Purpose Set default value for predict
Syntax $\quad B=\operatorname{SetDefaultYfit(B,Yfit)~}$
Description $B=$ SetDefaultYfit( $B, Y f i t)$ sets the default prediction for ensemble $B$ to Yfit. The default prediction must be a character variable for classification or a numeric scalar for regression. This setting controls what predicted value CompactTreeBagger returns when no prediction is possible, for example when the predict method needs to predict for an observation which has only false values in the matrix supplied through 'useifort' argument.

See Also<br>predict | TreeBagger.DefaultYfit

## Purpose Set difference for categorical arrays

Syntax
C = setdiff(A,B)
[C,I] = setdiff(A,B)

Description $\quad C=\operatorname{setdiff}(A, B)$ when $A$ and $B$ are categorical arrays returns a categorical vector $C$ containing the values in $A$ that are not in $B$. The result $C$ is sorted. The set of categorical levels for $C$ is the sorted union of the sets of levels of the inputs, as determined by their labels.
[C,I] = setdiff(A,B) also returns index vectors I such that $C=$ A(I).

See Also<br>intersect | ismember | setxor | union | unique

Purpose Label levels
Syntax $\quad A=\operatorname{set} \operatorname{labels}(A$, labels $)$
A = setlabels(A,labels,levels)
Description $\quad A=\operatorname{setlabels}(A, l a b e l s)$ labels the levels in the categorical array $A$ using the cell array of strings or 2-D character matrix labels. Labels are assigned in the order given in labels.
A = setlabels(A, labels,levels) labels only the levels specified in the cell array of strings or 2-D character matrix levels.

## Examples <br> Example 1

Relabel the species in Fisher's iris data using new categories:

```
load fisheriris
species = nominal(species);
species = mergelevels(...
    species,{'setosa','virginica'},'parent');
species = setlabels(species,'hybrid','versicolor');
getlabels(species)
ans =
    'hybrid' 'parent'
```


## Example 2

1 Load patient data from the CSV file hospital. dat and store the information in a dataset array with observation names given by the first column in the data (patient identification):

```
patients = dataset('file','hospital.dat',...
    'delimiter',',',...
    'ReadObsNames',true);
```

2 Make the $\{0,1\}$-valued variable smoke nominal, and change the labels to 'No' and 'Yes':

```
patients.smoke = nominal(patients.smoke,{'No','Yes'});
```

3 Add new levels to smoke as placeholders for more detailed histories of smokers:

```
patients.smoke = addlevels(patients.smoke,...
    {'0-5 Years','5-10 Years','LongTerm'});
```

4 Assuming the nonsmokers have never smoked, relabel the 'No ' level:

```
patients.smoke = setlabels(patients.smoke,'Never','No');
```

5 Drop the undifferentiated 'Yes' level from smoke:

```
patients.smoke = droplevels(patients.smoke,'Yes');
Warning: OLDLEVELS contains categorical levels that
were present in A, caused some array elements to have
undefined levels.
```

Note that smokers now have an undefined level.
6 Set each smoker to one of the new levels, by observation name:

```
patients.smoke('YPL-320') = '5-10 Years';
```

See Also
getlabels
Purpose Set exclusive-or for categorical arrays

```
Syntax
C = setxor(A,B)
[C,IA,IB] = setxor(A,B)
```

Description
$C=\operatorname{set} x \operatorname{cor}(A, B)$ when $A$ and $B$ are categorical arrays returns a categorical vector $C$ containing the values not in the intersection of $A$ and B. The result C is sorted. The set of categorical levels for C is the sorted union of the sets of levels of the inputs, as determined by their labels.
$[C, I A, I B]=\operatorname{setxor}(A, B)$ also returns index vectors IA and IB such that $C$ is a sorted combination of the elements $A(I A)$ and $B(I B)$.

See Also<br>intersect | ismember | setdiff | union | unique true if all covariance matrices are restricted to be the same

Description Logical true if all the covariance matrices are restricted to be the same (pooled estimate); logical false otherwise.

## categorical.shiftdim

Purpose Shift dimensions of categorical array

```
Syntax
B = shiftdim(A, n)
[B,nshifts] = shiftdim(A)
```

Description
$B=\operatorname{shiftdim}(A, n)$ shifts the dimensions of the categorical array $A$ by $N$. When $n$ is positive, shiftdim shifts the dimensions to the left and wraps the $n$ leading dimensions to the end. When $n$ is negative, shiftdim shifts the dimensions to the right and pads with singletons.
[B,nshifts] = shiftdim(A) returns the array B with the same number of elements as $A$ but with any leading singleton dimensions removed. nshifts returns the number of dimensions that are removed. If A is a scalar, shiftdim has no effect.

## See Also

circshift | reshape | squeeze

## RegressionEnsemble.shrink

Purpose Prune ensemble
Syntax

cmp = shrink(ens)
cmp = shrink(ens,Name,Value)
Description
InputArguments
$\mathrm{cmp}=\operatorname{shrink}(\mathrm{ens})$ returns a compact shrunken version of ens, a regularized ensemble. cmp retains only learners with weights above a threshold.
cmp $=$ shrink(ens, Name, Value) returns an ensemble with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.
ens
A regression ensemble created with fitensemble.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## lambda

Vector of nonnegative regularization parameter values for lasso. If ens. Regularization is nonempty (populate it with regularize), shrink regularizes ens using lambda. If ens contains a Regularization structure, you cannot pass lambda.

Default: []

## threshold

Lower cutoff on weights for weak learners, a numeric nonnegative scalar. shrink creates cmp from those learners with weights above threshold.

## RegressionEnsemble.shrink

Default: 0
weightcolumn
Column index of ens.Regularization.TrainedWeights, a positive integer. shrink creates cmp with learner weights from this column.

Default: 1

## Output

 Arguments
## Examples

Shrink a 300 -member bagged regression ensemble using 0.1 for the parameter lambda, and view the number of members of the resulting ensemble:

```
X = rand(2000,20);
Y = repmat(-1,2000,1);
Y(sum(X(:, 1:5),2)>2.5) = 1;
bag = fitensemble(X,Y,'Bag',300,'Tree','type','regression');
cmp = shrink(bag,'lambda',0.1);
cmp.NTrained
ans =
83
```

See Also regularize | cvshrink | predict
How To • Chapter 13, "Nonparametric Supervised Learning"

## gmdistribution.Sigma property

| Purpose | Input array of covariances |
| :--- | :--- |
| Description | Input array of covariances SIGMA. |

Purpose Wilcoxon signed rank test

```
Syntax
\(\mathrm{p}=\) signrank(x)
\(p=\operatorname{signrank}(x, m)\)
\(\mathrm{p}=\) signrank( \(\mathrm{x}, \mathrm{y})\)
[p,h] = signrank(...)
[ \(p, h\) ] = signrank(...,'alpha',alpha)
[p,h] = signrank(...,'method',method)
[p,h,stats] = signrank(...)
```


## Description

$\mathrm{p}=$ signrank( x ) performs a two-sided signed rank test of the null hypothesis that data in the vector $x$ comes from a continuous, symmetric distribution with zero median, against the alternative that the distribution does not have zero median. The $p$ value of the test is returned in $p$.
$p=$ signrank $(x, m)$ performs a two-sided signed rank test of the null hypothesis that data in the vectors $x$ and $y$ are independent samples from a continuous, symmetric distribution with median m , against the alternative that the distribution does not have median $m$. $m$ must be a scalar.
$\mathrm{p}=$ signrank $(\mathrm{x}, \mathrm{y})$ performs a paired, two-sided signed rank test of the null hypothesis that data in the vector $x-y$ come from a continuous, symmetric distribution with zero median, against the alternative that the distribution does not have zero median. $x$ and $y$ must have equal lengths. Note that a hypothesis of zero median for $\mathrm{x}-\mathrm{y}$ is not equivalent to a hypothesis of equal median for $x$ and $y$.
$[p, h]=$ signrank(...) returns the result of the test in $h . h=1$ indicates a rejection of the null hypothesis at the $5 \%$ significance level. $\mathrm{h}=0$ indicates a failure to reject the null hypothesis at the $5 \%$ significance level.
$[p, h]=\operatorname{signrank}(\ldots$, 'alpha', alpha) performs the test at the (100*alpha)\% significance level. The default, when unspecified, is alpha $=0.05$.
[ $\mathrm{p}, \mathrm{h}$ ] = signrank(...,'method',method) computes the $p$ value using either an exact algorithm, when method is 'exact', or a normal approximation, when method is 'approximate'. The default, when unspecified, is the exact method for small samples and the approximate method for large samples.
[ $\mathrm{p}, \mathrm{h}$, stats] = signrank(...) returns the structure stats with the following fields:

- signedrank - Value of the signed rank test statistic
- zval - Value of the $z$-statistic (computed only for large samples)


## Examples

## References

See Also

Test the hypothesis of zero median for the difference between two paired samples.

```
before = lognrnd(2,.25,10,1);
after = before+trnd(2,10,1);
[p,h] = signrank(before,after)
p =
    0.5566
h =
    0
```

The sampling distribution of the difference between before and after is symmetric with zero median. At the default $5 \%$ significance level, the test fails to reject to the null hypothesis of zero median in the difference.
[1] Gibbons, J. D. Nonparametric Statistical Inference. New York: Marcel Dekker, 1985.
[2] Hollander, M., and D. A. Wolfe. Nonparametric Statistical Methods. Hoboken, NJ: John Wiley \& Sons, Inc., 1999.

```
ranksum | ttest | ztest
```


## signtest

## Purpose Sign test

Syntax

```
p = signtest(x)
p = signtest(x,m)
p = signtest(x,y)
[p,h] = signtest(...)
[p,h] = signtest(...,'alpha',alpha)
[p,h] = signtest(...,'method',method)
[p,h,stats] = signtest(...)
```


## Description

$\mathrm{p}=$ signtest( x ) performs a two-sided sign test of the null hypothesis that data in the vector $x$ come from a continuous distribution with zero median, against the alternative that the distribution does not have zero median. The $p$ value of the test is returned in p
$\mathrm{p}=$ signtest ( $\mathrm{x}, \mathrm{m}$ ) performs a two-sided sign test of the null hypothesis that data in the vector x come from a continuous distribution with median m , against the alternative that the distribution does not have median $\mathrm{m} . \mathrm{m}$ must be a scalar.
$p=$ signtest $(x, y)$ performs a paired, two-sided sign test of the null hypothesis that data in the vector $x-y$ come from a continuous distribution with zero median, against the alternative that the distribution does not have zero median. $x$ and $y$ must be the same length. Note that a hypothesis of zero median for $x-y$ is not equivalent to a hypothesis of equal median for x and y .
$[p, h]=\operatorname{signtest}(\ldots)$ returns the result of the test in $h . h=1$ indicates a rejection of the null hypothesis at the $5 \%$ significance level. $\mathrm{h}=0$ indicates a failure to reject the null hypothesis at the $5 \%$ significance level.
[ $p, h$ ] = signtest(...,'alpha', alpha) performs the test at the (100*alpha)\% significance level. The default, when unspecified, is alpha $=0.05$.
$[p, h]=$ signtest(...,'method', method) computes the $p$ value using either an exact algorithm, when method is 'exact', or a normal approximation, when method is 'approximate'. The default, when
unspecified, is the exact method for small samples and the approximate method for large samples.
[ $p, h$, stats] $=$ signtest(...) returns the structure stats with the following fields:

- sign - Value of the sign test statistic
- zval - Value of the $z$-statistic (computed only for large samples)


## Examples

## References

See Also

Test the hypothesis of zero median for the difference between two paired samples.

```
before = lognrnd(2,.25,10,1);
after = before + (lognrnd(0,.5,10,1) - 1);
[p,h] = signtest(before,after)
p =
    0.3438
h =
    0
```

The sampling distribution of the difference between before and after is symmetric with zero median. At the default $5 \%$ significance level, the test fails to reject to the null hypothesis of zero median in the difference.
[1] Gibbons, J. D. Nonparametric Statistical Inference. New York: Marcel Dekker, 1985.
[2] Hollander, M., and D. A. Wolfe. Nonparametric Statistical Methods. Hoboken, NJ: John Wiley \& Sons, Inc., 1999.

ranksum | signrank | ttest | ztest

Purpose Silhouette plot

Syntax
silhouette(X,clust) $s=$ silhouette(X,clust) [s,h] = silhouette(X,clust) [...] $=$ silhouette(X,clust,metric) $[\ldots]=$ silhouette(X,clust,distfun, $\mathrm{p} 1, \mathrm{p} 2, \ldots)$

## Description

silhouette ( X, clust) plots cluster silhouettes for the $n$-by-p data matrix $X$, with clusters defined by clust. Rows of $X$ correspond to points, columns correspond to coordinates. clust can be a categorical variable, numeric vector, character matrix, or cell array of strings containing a cluster name for each point. (See "Grouped Data" on page $2-34$.) silhouette treats NaNs or empty strings in clust as missing values, and ignores the corresponding rows of $X$. By default, silhouette uses the squared Euclidean distance between points in $X$.
$s=$ silhouette (X, clust) returns the silhouette values in the $n$-by- 1 vector s, but does not plot the cluster silhouettes.
[s,h] = silhouette(X, clust) plots the silhouettes, and returns the silhouette values in the $n$-by- 1 vector s , and the figure handle in h .
[...] = silhouette(X, clust,metric) plots the silhouettes using the inter-point distance function specified in metric. Choices for metric are given in the following table.

| Metric | Description |
| :--- | :--- |
| 'Euclidean' | Euclidean distance |
| 'sqEuclidean' | Squared Euclidean distance (default) |
| 'cityblock' | Sum of absolute differences |
| 'cosine' | One minus the cosine of the included angle <br> between points (treated as vectors) |
| 'correlation' | One minus the sample correlation between points <br> (treated as sequences of values) |


| Metric | Description |
| :--- | :--- |
| 'Hamming' | Percentage of coordinates that differ |
| 'Jaccard' | Percentage of nonzero coordinates that differ |
| Vector | A numeric distance matrix in upper triangular <br> vector form, such as is created by pdist. X is <br> not used in this case, and can safely be set to [ ]. |

For more information on each metric, see "Distance Metrics" on page 13-9.
[...] = silhouette(X,clust,distfun, p1,p2,...) accepts a function handle distfun to a metric of the form

$$
d=\operatorname{distfun}(X 0, x, p 1, p 2, \ldots)
$$

where XO is a 1-by-p point, X is an n -by-p matrix of points, and $\mathrm{p} 1, \mathrm{p} 2, \ldots$ are optional additional arguments. The function distfun returns an n -by- 1 vector d of distances between X0 and each point (row) in X. The arguments $\mathrm{p} 1, \mathrm{p} 2, \ldots$ are passed directly to the function distfun.

Tips

Examples

The silhouette value for each point is a measure of how similar that point is to points in its own cluster compared to points in other clusters, and ranges from -1 to +1 . It is defined as

$$
S(i)=(\min (b(i,:), 2)-a(i)) . / \max (a(i), \min (b(i,:)))
$$

where $a(i)$ is the average distance from the ith point to the other points in its cluster, and $b(i, k)$ is the average distance from the $i$ th point to points in another cluster k .

```
X = [randn(10,2)+ones(10,2);
randn(10,2)-ones(10,2)];
cidx = kmeans(X,2,'distance','sqeuclid');
s = silhouette(X,cidx,'sqeuclid');
```


## silhouette

| References | [1] Kaufman L., and P. J. Rousseeuw. Finding Groups in Data: An <br> Introduction to Cluster Analysis. Hoboken, NJ: John Wiley \& Sons, <br> Inc., 1990. |
| :--- | :--- |
| See Also | dendrogram \| kmeans | linkage | pdist |
| How To | - "Grouped Data" on page 2-34 |

Purpose Convert categorical array to single array

## Syntax <br> $B=$ single $(A)$

Description $\quad B=$ single $(A)$ converts the categorical array $A$ to a single array. Each element of $B$ contains the internal categorical level code for the corresponding element of $A$.

See Also double

## dataset.single

Purpose Convert dataset variables to single array
Syntax
$B=$ single(A)
B = single(A,vars)

Description
$B=$ single $(A)$ returns the contents of the dataset $A$, converted to one single array. The classes of the variables in the dataset must support the conversion.
$B=$ single(A,vars) returns the contents of the dataset variables specified by vars. vars is a positive integer, a vector of positive integers, a variable name, a cell array containing one or more variable names, or a logical vector.

## See Also

dataset | double | replacedata

## Purpose Size of categorical array

Syntax

```
d = size(A)
[m,n] = size(A)
[m1,m2,m3,...,mn] = size(A)
m = size(A,dim)
```

Description

$n>\operatorname{ndims}(A) \quad$| size returns ones in the "extra" variables, |
| :--- |
| i.e., outputs ndims $(A)+1$ through $n$. |


| mn contains the product of the sizes of |
| :--- |
| dimensions $n$ through ndims $(A)$. |

$m=\operatorname{size}(A, \operatorname{dim})$ returns the length of the dimension specified by the

| scalar dim. For example, size $(A, 1)$ returns the number of rows. If |
| :--- |
| dim $>\operatorname{ndims}(A), m$ will be 1. |

See Also length | ndims | numel
Purpose Size of dataset array

Syntax $\quad$|  | $D=\operatorname{SIZE}(A)$ |
| :--- | :--- |
|  | $[\operatorname{NOBS}, \operatorname{NVARS}]=\operatorname{SIZE}(A)$ |
|  | $[M 1, M 2, M 3, \ldots, \operatorname{MN}]=\operatorname{SIZE}(A)$ |
|  | $M=\operatorname{size}(A, \operatorname{dim})$ |

## Description

$D=\operatorname{SIZE}(A)$ returns the two-element row vector $D=[$ NOBS, NVARS] containing the number of observations and number of variables in the dataset A. A dataset array always has two dimensions.
[NOBS,NVARS] = SIZE(A) returns the numbers of observations and variables in the dataset A as separate output variables.
[M1 , M2, M3,$\ldots$, MN] $=\operatorname{SIZE}(A)$, for $N>2$, returns M1 = NOBS, M2 = NVARS, and M3,..,MN = 1 .
$M=\operatorname{size}(A, d i m)$ returns the length of the dimension specified by the scalar dim:

- $M=\operatorname{size}(A, 1)$ returns NOBS
- $M=\operatorname{size}(A, 2)$ returns NVARS
- $M=\operatorname{size}(A, k)$ returns 1 for $k>2$


## See Also length | ndims | numel

## Purpose Number of dimensions in matrix

## Syntax $\quad d=\operatorname{size}(p)$ <br> [m,n] = size(p) <br> m = size(p,dim)

## Description

## Examples The commands

```
    P = sobolset(12);
d = size(P)
return
\[
d=\left[\begin{array}{ll}
9.0072 e+015 & 12
\end{array}\right]
\]
```

The command

$$
[m, n]=\operatorname{size}(P)
$$

returns

$$
\begin{aligned}
& \mathrm{m}=9.0072 \mathrm{e}+015 \\
& \mathrm{n}=12
\end{aligned}
$$

The command

## qrandset.size

$$
\mathrm{m} 2=\operatorname{size}(\mathrm{P}, 2)
$$

returns

$$
m 2=12
$$

See Also length | ndims | qrandset

## Purpose

Slice sampler

Syntax<br>\section*{Description}

rnd = slicesample(initial, nsamples, 'pdf',pdf)
rnd = slicesample(initial,nsamples,'logpdf',logpdf)
[rnd,neval] = slicesample(initial,...)
[rnd, neval] = slicesample(initial,...,Name,Value)

Tips

Input Arguments
rnd = slicesample(initial, nsamples,'pdf',pdf) generates nsamples random samples using the slice sampling method (see "Algorithms" on page 20-1777). pdf gives the target probability density function (pdf). initial is a row vector or scalar containing the initial value of the random sample sequences.
rnd = slicesample(initial, nsamples,'logpdf',logpdf) generates samples using the logarithm of the pdf.
[rnd,neval] = slicesample(initial,...) returns the average number of function evaluations that occurred in the slice sampling.
[rnd,neval] = slicesample(initial,...,Name,Value) generates random samples with additional options specified by one or more Name, Value pair arguments.

- There are no definitive suggestions for choosing appropriate values for burnin, thin, or width. Choose starting values of burnin and thin, and increase them, if necessary, to give the requisite independence and marginal distributions. See Neal [1] for details of the effect of adjusting width.
initial
Initial point, a scalar or row vector. Set initial so pdf(initial) is a strictly positive scalar. length(initial) is the number of dimensions of each sample.
nsamples
Positive integer, the number of samples that slicesample generates.


## slicesample

pdf
Handle to a function that generates the probability density function, specified with @. pdf can be unnormalized, meaning it need not integrate to 1 .

## logpdf

Handle to a function that generates the logarithm of the probability density function, specified with @. logpdf can be the logarithm of an unnormalized pdf.

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

## burnin

Nonnegative integer, the number of samples to generate and discard before generating the samples to return. The slice sampling algorithm is a Markov chain whose stationary distribution is proportional to that of the pdf argument. Set burnin to a high enough value that you believe the Markov chain approximately reaches stationarity after burnin samples.

Default: 0
thin
Positive integer, where slicesample discards every thin - 1 samples and returns the next. The slice sampling algorithm is a Markov chain, so the samples are serially correlated. To reduce the serial correlation, choose a larger value of thin.

Default: 1
width

Width of the interval around the current sample, a scalar or vector of positive values. slicesample begins with this interval and searches for an appropriate region containing the points of pdf that evaluate to a large enough value.

- If width is a scalar and the samples have multiple dimensions, slicesample uses width for each dimension.
- If width is a vector, it should have the same length as initial.

Default: 10

## Output <br> rnd

Arguments
nsamples-by-length(initial) matrix, where each row is one sample.
neval
Scalar, the mean number of function evaluations per sample. neval includes the burnin and thin evaluations, not just the evaluations of samples returned in rnd. Therefore the total number of function evaluations is

```
neval*(nsamples*thin + burnin).
```


## Examples

Generate random samples from a multimodal density using slicesample.

1 Define a function proportional to a multimodal density:

```
f = @(x) exp(-x.^2/2).*(1 + (sin(3*x)).^2).*...
    (1 + (cos(5*x).^2));
area = quad(f,-5,5);
```

2 Generate 2000 samples from the density, using a burn-in period of 1000, and keeping one in five samples:
$N=2000 ;$

## slicesample

```
x = slicesample(1,N,'pdf',f,'thin',5,'burnin',1000);
```

3 Plot a histogram of the sample:

```
[binheight,bincenter] = hist(x,50);
h = bar(bincenter,binheight,'hist');
set(h,'facecolor',[0.8 .8 1]);
```

4 Scale the density to have the same area as the histogram, and superimpose it on the histogram:

```
hold on
xd = get(gca,'XLim');
xgrid = linspace(xd(1),xd(2),1000);
binwidth = (bincenter(2)-bincenter(1));
y = (N*binwidth/area) * f(xgrid);
plot(xgrid,y,'r','LineWidth',2)
hold off
```



## Algorithms

At each point in the sequence of random samples, slicesample selects the next point by "slicing" the density to form a neighborhood around the previous point where the density is above some value. Consequently, the sample points are not independent. Nearby points in the sequence tend to be closer together than they would be from a sample of independent values. For many purposes, the entire set of points can be used as a sample from the target distribution. However, when this type of serial correlation is a problem, the burnin and thin parameters can help reduce that correlation.
slicesample uses the slice sampling algorithm of Neal [1]. For numerical stability, it converts a pdf function into a logpdf function. The algorithm to resize the support region for each level, called "stepping-out" and "stepping-in," was suggested by Neal.

References [1] Neal, Radford M. Slice Sampling. Ann. Stat. Vol. 31, No. 3, pp. 705-767, 2003. Available at Project Euclid.

## See Also

mhsample | rand | randsample

## Purpose

Skewness

## Syntax

$y=$ skewness (X)
$y=$ skewness $(X, f l a g)$

## Algorithms

$y=$ skewness $(X)$ returns the sample skewness of $X$. For vectors, skewness $(x)$ is the skewness of the elements of $x$. For matrices, skewness $(X)$ is a row vector containing the sample skewness of each column. For N-dimensional arrays, skewness operates along the first nonsingleton dimension of $X$.
$y=s k e w n e s s(X, f l a g)$ specifies whether to correct for bias (flag = 0 ) or not (flag = 1, the default). When X represents a sample from a population, the skewness of $X$ is biased; that is, it will tend to differ from the population skewness by a systematic amount that depends on the size of the sample. You can set flag $=0$ to correct for this systematic bias.
skewness (X,flag, dim) takes the skewness along dimension dim of X .
skewness treats NaNs as missing values and removes them.
Skewness is a measure of the asymmetry of the data around the sample mean. If skewness is negative, the data are spread out more to the left of the mean than to the right. If skewness is positive, the data are spread out more to the right. The skewness of the normal distribution (or any perfectly symmetric distribution) is zero.

The skewness of a distribution is defined as

$$
s=\frac{E(x-\mu)^{3}}{\sigma^{3}}
$$

where $\mu$ is the mean of $x, \sigma$ is the standard deviation of $x$, and $E(t)$ represents the expected value of the quantity $t$. skewness computes a sample version of this population value.

When you set flag to 1 , the following equation applies:

$$
s_{1}=\frac{\frac{1}{n} \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{3}}{\left(\sqrt{\frac{1}{n} \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2}}\right)^{3}}
$$

When you set flag to 0 , the following equation applies:

$$
s_{0}=\frac{\sqrt{n(n-1)}}{n-2} s_{1}
$$

This bias-corrected formula requires that X contain at least three elements.

```
Examples
X = randn([5 4])
X =
        1.1650 1.6961 -1.4462 -0.3600
        0.6268 0.0591-0.7012 -0.1356
        0.0751 1.7971 1.2460-1.3493
        0.3516 0.2641 -0.6390-1.2704
        -0.6965 0.8717 0.5774 0.9846
y = skewness(X)
y =
    -0.2933 0.0482 0.2735 0.4641
```

See Also kurtosis | mean | moment | std | var

## qrandset.Skip property

| Purpose | Number of initial points to omit from sequence |
| :---: | :---: |
| Description | The Skip property of a point set contains a positive integer which specifies the number of initial points in the sequence to omit from the point set. The default Skip value is 0 . |
|  | Initial points of a sequence sometimes exhibit undesirable properties, for example the first point is often $(0,0,0, \ldots)$ and this may "unbalance" the sequence since its counterpart, ( $1,1,1, \ldots$ ), never appears. Another common reason is that initial points often exhibit correlations among different dimensions which disappear later in the sequence. |
| Examples | Examine the difference between skipping and not skipping points: |
|  | \% No skipping produces the standard Sobol sequence. $P=\operatorname{sobolset}(5) ;$ $P(1: 3,:)$ |
|  | ```% Skip the first point of the sequence. The point set now % starts at the second point of the basic Sobol sequence. P.Skip = 1; P(1:3,:)``` |
| See Also | Leap \| net | qrandset | subsref |

Superclasses qrandset
Purpose Sobol quasi-random point sets
Description sobolset is a quasi-random point set class that produces points from the Sobol sequence. The Sobol sequence is a base-2 digital sequence that fills space in a highly uniform manner.

## Construction

sobolset
Construct Sobol quasi-random point set

## Methods

## Inherited Methods

Methods in the following table are inherited from qrandset.

| disp | Display qrandset object |
| :--- | :--- |
| end | Last index in indexing expression <br> for point set |
| length | Length of point set |
| ndims | Number of dimensions in matrix |
| net | Generate quasi-random point set |
| scramble | Scramble quasi-random point set |
| size | Number of dimensions in matrix |
| subsref | Subscripted reference for <br> qrandset |

## Properties

PointOrder
Point generation method

## Inherited Properties

Properties in the following table are inherited from qrandset.

|  | Dimensions | Number of dimensions |
| :---: | :---: | :---: |
|  | Leap | Interval between points |
|  | ScrambleMethod | Settings that control scrambling |
|  | Skip | Number of initial points to omit from sequence |
|  | Type | Name of sequence on which point set $P$ is based |
| Copy <br> Semantics | Value. To learn how this affects your use of the class, see Comparing Handle and Value Classes in the MATLAB Object-Oriented Programming documentation. |  |
| References | [1] Bratley, P., and B. L. Fox, "ALGORITHM 659 Implementing Sobol's Quasirandom Sequence Generator," ACM Transactions on Mathematical Software, Vol. 14, No. 1, pp. 88-100, 1988. |  |
|  | [2] Joe, S., and F. Y. Kuo, "Remark on Algorithm 659: Implementing Sobol's Quasirandom Sequence Generator," ACM Transactions on Mathematical Software, Vol. 29, No. 1, pp. 49-57, 2003. |  |
|  | [3] Hong, H. S., and F. J. Hickernell, "ALGORITHM 823: Implementing Scrambled Digital Sequences," ACM Transactions on Mathematical Software, Vol. 29, No. 2, pp. 95-109, 2003. |  |
|  | [4] Matousek, J., "On the L2-discrepancy for anchored boxes," Journal of Complexity, Vol. 14, pp. 527-556, 1998. |  |
| See Also | haltonset |  |
| How To | - "Quasi-Random | ge 6-16 |

## Purpose Construct Sobol quasi-random point set

```
Syntax
p = sobolset(d)
p = sobolset(d,prop1,val1,prop2,val2,...)
```

Description

## Examples

Generate a 3-D Sobol point set, skip the first 1000 values, and then retain every 101st point:

```
p = sobolset(3,'Skip',1e3,'Leap',1e2)
p =
    Sobol point set in 3 dimensions (8.918019e+013 points)
    Properties:
        Skip : 1000
        Leap : 100
    ScrambleMethod : none
        PointOrder : standard
```

Use scramble to apply a random linear scramble combined with a random digital shift:

```
p = scramble(p,'MatousekAffineOwen')
p =
    Sobol point set in 3 dimensions (8.918019e+013 points)
    Properties:
    Skip : 1000
```

```
Leap : 100
ScrambleMethod : MatousekAffineOwen
PointOrder : standard
```

Use net to generate the first four points:

```
X0 = net(p,4)
XO =
    0.7601 0.5919 0.9529
    0.1795 0.0856 0.0491
    0.5488 0.0785 0.8483
    0.3882 0.8771 0.8755
```

Use parenthesis indexing to generate every third point, up to the 11th point:

```
X = p(1:3:11,:)
X =
    0.7601 0.5919 0.9529
    0.3882 0.8771 0.8755
    0.6905 0.4951 0.8464
    0.1955 0.5679 0.3192
```


## References

[1] Bratley, P., and B. L. Fox. "Algorithm 659 Implementing Sobol's Quasirandom Sequence Generator." ACM Transactions on Mathematical Software. Vol. 14, No. 1, 1988, pp. 88-100.
[2] Joe, S., and F. Y. Kuo. "Remark on Algorithm 659: Implementing Sobol's Quasirandom Sequence Generator." ACM Transactions on Mathematical Software. Vol. 29, No. 1, 2003, pp. 49-57.
[3] Hong, H. S., and F. J. Hickernell. "Algorithm 823: Implementing Scrambled Digital Sequences." ACM Transactions on Mathematical Software. Vol. 29, No. 2, 2003, pp. 95-109.
[4] Matousek, J. "On the L2-Discrepancy for Anchored Boxes." Journal of Complexity. Vol. 14, No. 4, 1998, pp. 527-556.

See Also haltonset | net | scramble

## Purpose Sort elements of ordinal array

Syntax

```
B = sort(A)
B = sort(A,dim)
B = sort(A,dim,mode)
[B,I] = sort(A,...)
```

Description
$B=\operatorname{sort}(A)$, when $A$ is an ordinal vector, sorts the elements of $A$ in ascending order. For ordinal matrices, sort(A) sorts each column of A in ascending order. For $N$-D ordinal arrays, sort (A) sorts the along the first nonsingleton dimension of A. B is an ordinal array with the same levels as A.
$B=\operatorname{sort}(A, d i m)$ sorts $A$ along dimension dim.
$B=\operatorname{sort}(A, \operatorname{dim}$, mode $)$ sorts $A$ in the order specified by mode. mode is ' ascend ' for ascending order, or 'descend ' for descending order.
$[B, I]=\operatorname{sort}(A, \ldots)$ also returns an index matrix I. If $A$ is a vector, then $B=A(I)$. If $A$ is an $m$-by- $n$ matrix and $\operatorname{dim}$ is 1 , then $B(:, j)=$ $A(I(:, j), j)$ for $j=1: n$.

Elements with undefined levels are sorted to the end.

## Examples <br> Sort the columns of an ordinal array in ascending order:

```
A = ordinal([6 2 5; 2 4 1; 3 2 4],...
    {'lo','med','hi'},[],[0 2 4 6])
A =
    hi med hi
    med hi lo
    med med hi
B = sort(A)
B =
    med med lo
    med med hi
    hi hi hi
```


## ordinal.sort

See Also sortrows

## Purpose Sort rows of dataset array

```
Syntax B = sortrows(A)
B = sortrows(A,vars)
B = sortrows(A,'obsnames')
B = sortrows(A,vars,mode)
[B,idx] = sortrows(A)
```

Description

## Examples

$B=$ sortrows(A) returns a copy of the dataset array $A$, with the observations sorted in ascending order by all of the variables in A. The observations in B are sorted first by the first variable, next by the second variable, and so on. The variables in A must be scalar valued (i.e., column vectors) and be from a class for which a sort method exists.
$B=$ sortrows(A, vars) sorts the observations in $A$ by the variables specified by vars. vars is a positive integer, a vector of positive integers, variable names, a cell array containing one or more variable names, or a logical vector.

B = sortrows(A,'obsnames') sorts the observations in A by observation name.

B = sortrows(A, vars,mode) sorts in the direction specified by mode. mode is 'ascend' (the default) or 'descend'. Use [] for vars to sort using all variables.
$[B, i d x]=$ sortrows $(A)$ also returns an index vector $i d x$ such that B = A(idx,:).

Sort the data in hospital.mat by age and then by last name:

| $l$ |  |  |  |
| :--- | :--- | :--- | :--- |
| load hospital |  |  |  |
| hospital(1:5,1:3) |  |  |  |
| ans = |  |  |  |
|  | LastName | Sex | Age |
| YPL-320 | 'SMITH' | Male | 38 |
| GLI-532 | 'JOHNSON' | Male | 43 |
| PNI-258 | 'WILLIAMS' | Female | 38 |
| MIJ-579 | 'JONES' | Female | 40 |

## dataset.sortrows

| XLK-030 | 'BROWN ' | Female | 49 |
| :---: | :---: | :---: | :---: |
| ```hospital = sortrows(hospital,{'Age','LastName'}); hospital(1:5,1:3) ans =``` |  |  |  |
|  | LastName | Sex | Age |
| REV-997 | 'ALEXANDER ' | Male | 25 |
| FZR-250 | 'HALL' | Male | 25 |
| LIM-480 | 'HILL' | Female | 25 |
| XUE-826 | 'JACKSON ' | Male | 25 |
| SCQ-914 | 'JAMES' | Male | 25 |

See Also sortrows

## Purpose

Sort rows
Syntax
$B=$ sortrows $(A)$
$B=$ sortrows $(A, c o l)$
[B,I] = sortrows(A)
$[B, I]=$ sortrows $(A, c o l)$
Description
$B=$ sortrows $(A)$ sorts the rows of the 2-D ordinal matrix $A$ in ascending order, as a group. $B$ is an ordinal array with the same levels as A.
$B=$ sortrows(A, col) sorts A based on the columns specified in the vector col. If an element of col is positive, the corresponding column in A is sorted in ascending order; if an element of col is negative, the corresponding column in $A$ is sorted in descending order.
$[B, I]=\operatorname{sortrows}(A)$ and $[B, I]=\operatorname{sortrows}(A, C o l)$ also returns an index matrix $I$ such that $B=A(I,:)$.
Elements with undefined levels are sorted to the end.

## Examples

Sort the rows of an ordinal array in ascending order for the first column, and then in descending order for the second column:

```
A = ordinal([6 2 5; 2 4 1; 3 2 4],...
    {'lo','med','hi'},[],[0 2 4 6])
A =
    hi med hi
    med hi lo
    med med hi
B = sortrows(A,[1 -2])
B =
        med hi lo
        med med hi
        hi med hi
```

    See Also sort | sortrows
    
## Purpose Format distance matrix

```
Syntax z = squareform(y)
y = squareform(Z)
Z = squareform(y,'tovector')
Y = squareform(Z,'tomatrix')
```


## Description

$Z=\operatorname{squareform(y),~where~} y$ is a vector as created by the pdist function, converts y into a square, symmetric format $Z$, in which $Z(i, j)$ denotes the distance between the ith and $j$ th objects in the original data.
$y=$ squareform( $Z$ ), where $Z$ is a square, symmetric matrix with zeros along the diagonal, creates a vector $y$ containing the $Z$ elements below the diagonal. $y$ has the same format as the output from the pdist function.

Z = squareform(y,'tovector') forces squareform to treat y as a vector.
$Y=$ squareform(Z,'tomatrix') forces squareform to treat $Z$ as a matrix.

The last two formats are useful if the input has a single element, so that it is ambiguous whether the input is a vector or square matrix.

## Examples

```
\(y=1: 6\)
y \(=\)
    \(\begin{array}{llllll}1 & 2 & 3 & 4 & 5 & 6\end{array}\)
\(X=[0123 ; 1045 ; 2406 ; 3560]\)
\(x=\)
    \(0 \quad 1 \quad 2 \quad 3\)
    \(1 \quad 0 \quad 4 \quad 5\)
    2406
    \(3 \quad 5 \quad 6 \quad 0\)
```

Then squareform(y) $=X$ and squareform $(X)=y$.

See Also pdist

Purpose Squeeze singleton dimensions from categorical array

## Syntax <br> $B=$ squeeze (A)

Description $B=$ squeeze $(A)$ returns an array $B$ with the same elements as the categorical array $A$ but with all the singleton dimensions removed. A singleton is a dimension such that size (A, dim) $=1.2$-D arrays are unaffected by squeeze so that row vectors remain rows.

See Also shiftdim

## Purpose <br> Syntax <br> Description

Stack data from multiple variables into single variable
tall = stack(wide,datavars)
[tall,iwide] = stack(wide,datavars)
tall = stack(wide,datavars,Parameter,value)
tall = stack(wide, datavars) converts a wide-format dataset array into a tall-format array, by stacking multiple variables in wide into a single variable in tall. In general, tall contains fewer variables but more observations than wide.
datavars specifies a group of $m$ data variables in wide. stack creates a single data variable in tall by interleaving their values, and if wide has $n$ observations, then tall has m-by-n observations. In other words, stack takes the $m$ data values from each observation in wide and stacks them up to create $m$ observations in tall. datavars is a positive integer, a vector of positive integers, a variable name, a cell array containing one or more variable names, or a logical vector. stack also creates a grouping variable in tall to indicate which of the $m$ data variables in wide each observation in tall corresponds to.
stack assigns values for the "per-variable properties (e.g., Units and VarDescription) for the new data variable in tall from the corresponding property values for the first variable listed in datavars.
stack copies the remaining variables from wide to tall without stacking, by replicating each of their values $m$ times. These variables are typically grouping variables. Because their values are constant across each group of $m$ observations in tall, they identify which observation in wide an observation in tall came from.
[tall,iwide] = stack(wide,datavars) returns an index vector iwide indicating the correspondence between observations in tall and those in wide. stack creates tall( $\mathrm{j},:$ ) using wide(iwide( j ), datavarss).

For more information on grouping variables, see "Grouping Variables" on page 2-34.

## Input Arguments

## Examples

tall = stack(wide,datavars, Parameter, value) uses the following parameter name/value pairs to control how stack converts variables in wide to variables in tall:

| 'ConstVars' | Variables in wide to copy to tall <br> without stacking. ConstVars <br> is a positive integer, a vector <br> of positive integers, a variable <br> name, a cell array containing <br> one or more variable names, or a <br> logical vector. The default is all <br> variables in wide not specified in <br> datavars. |
| :--- | :--- |
| 'NewDataVarName' | A name for the data variable to be <br> created in tall. The default is a <br> concatenation of the names of the <br> m variables that are stacked up. |
| 'IndVarName' | A name for the grouping variable <br> to create in tall to indicate the <br> source of each value in the new <br> data variable. The default is <br> based on the 'NewDataVarName' <br> parameter. |

You can also specify multiple groups of data variables in wide, each of which becomes a variable in tall. All groups must contain the same number of variables. Use a cell array to contain multiple parameter values for datavars, and a cell array of strings to contain multiple 'NewDataVarName'.

Convert a wide format data set to tall format, and then back to a different wide format:
load flu
\% FLU has a 'Date' variable, and 10 variables for estimated
\% influenza rates (in 9 different regions, estimated from
\% Google searches, plus a nationwide extimate from the \% CDC). Combine those 10 variables into a "tall" array that

\% has a single data variable, 'FluRate', and an indicator

\% variable, 'Region', that says which region each estimate

\% is from.

[flu2,iflu] = stack(flu, 2:11, 'NewDataVarName','FluRate', ...

    'IndVarName','Region')
    \% The second observation in FLU is for $10 / 16 / 2005$. Find the
\% observations in FLU2 that correspond to that date.
flu(2,:)
flu2(iflu==2,:)
\% Use the 'Date' variable from that tall array to split
\% 'FluRate' into 52 separate variables, each containing the
\% estimated influenza rates for each unique date. The new
\% "wide" array has one observation for each region. In
\% effect, this is the original array FLU "on its side".
dateNames = cellstr(datestr(flu.Date,'mmm_DD_YYYY'));
[flu3,iflu2] = unstack(flu2, 'FluRate', 'Date', ...
'NewDataVarNames ', dateNames)
\% Since observations in FLU3 represent regions, IFLU2
\% indicates the first occurrence in FLU2 of each region.
flu2(iflu2,:)
See Also
dataset.unstack | dataset.join | dataset.grpstats
How To . "Grouping Variables" on page 2-34

## qrandstream.State property

Purpose Current state of the stream
Description The State property of a quasi-random stream contains the index into the associated point set of the next point to draw in the stream. Getting and resetting the State property allows you to return a stream to a previous state. The initial value of State is 1 .

Examples

```
Q = qrandstream('sobol', 5);
s = Q.State;
u1 = qrand(Q, 10)
Q.State = s;
u2 = qrand(Q, 10) % contains exactly the same values as u1
```

See Also qrand
Purpose Access values in statistics options structure
Syntax

val = statget(options, param)
val = statget(options,param,default)

Description

## Input <br> Arguments

val $=$ statget (options, param) returns the value of the parameter specified by the string param in the statistics options structure options. If the parameter is undefined in options, statget returns []. You need to type only enough leading characters to define the parameter name uniquely. statget ignores case for parameter names. For available options, see Inputs.
val = statget(options, param, default) returns default if the specified parameter is undefined in the optimization options structure options.

DerivStep
Relative difference used in finite difference derivative calculations. A positive scalar, or a vector of positive scalars the same size as the vector of parameters estimated by the Statistics Toolbox function using the options structure.

Display
Amount of information displayed by the algorithm.

- 'off' - Displays no information.
- 'final' - Displays the final output.
- 'iter' - Displays iterative output to the command window for some functions; otherwise displays the final output.


## FunValCheck

Check for invalid values, such as NaN or Inf, from the objective function.

- 'off'
- 'on'

GradObj
Flags whether the objective function returns a gradient vector as a second output.

- 'off'
- 'on'

Jacobian
Flags whether the objective function returns a Jacobian as a second output.

- 'off'
- 'on'

MaxFunEvals
Maximum number of objective function evaluations allowed. Positive integer.

MaxIter
Maximum number of iterations allowed. Positive integer.

## OutputFen

The solver calls all output functions after each iteration.

- Function handle specified using @
- a cell array with function handles
- an empty array (default)

Robust
Invoke robust fitting option.

- 'off'
- 'on'


## Streams

A single instance of the RandStream class, or a cell array of RandStream instances. The Streams option is accepted by some functions to govern what stream(s) to use in generating random numbers within the function. If 'UseSubstreams' is 'always', the Streams value must be a scalar, or must be empty. If 'UseParallel' is 'always' and 'UseSubstreams' is 'never', then the Streams argument must either be empty, or its length must match the number of processors used in the computation: equal to the matlabpool size if a matlabpool is open, a scalar otherwise. For more information on parallel computing, see Chapter 17, "Parallel Statistics".

TolBnd
Parameter bound tolerance. Positive scalar.
TolFun
Termination tolerance for the objective function value. Positive scalar.

## TolTypeFun

Use TolFun for absolute or relative objective function tolerances.

- 'abs'
- 'rel'


## TolTypeX

Use TolX for absolute or relative parameter tolerances.

- 'abs'
- 'rel'

TolX
Termination tolerance for the parameters. Positive scalar.

## Tune

The tuning constant used in robust fitting to normalize the residuals before applying the weight function. The default value depends upon the weight function. This parameter is necessary if you specify the weight function as a function handle. Positive scalar.

UseParallel
Flag indicating whether eligible functions should use capabilities of the Parallel Computing Toolbox (PCT), if the capabilities are available. That is, if the PCT is installed, and a PCT matlabpool is in effect. Valid values are 'never' (the default), for serial computation, and 'always', for parallel computation. For more information on parallel computing, see Chapter 17, "Parallel Statistics".

## UseSubstreams

Flag indicating whether the random number generator in eligible functions should use Substream property of the RandStream class. 'never' (default) or 'always'. 'always', high level iterations within the function will set the Substream property to the value of the iteration. This behavior helps to generate reproducible random number streams in parallel and/or serial mode computation. For more information on parallel computing, see Chapter 17, "Parallel Statistics".

## WgtFun

A weight function for robust fitting. Valid only when Robust is ' on '. Can also be a function handle that accepts a normalized residual as input and returns the robust weights as output.

- 'bisquare'
- 'andrews'
- 'cauchy'
- 'fair'
- 'huber'
- 'logistic'
- 'talwar'
- 'welsch'


## Examples

This statement returns the value of the Display statistics options parameter from the structure called my_options.

```
val = statget(my_options,'Display')
```

Return the value of the Display statistics options parameter from the structure called my_options (as in the previous example). If the Display parameter is undefined, statget returns the value 'final'.

```
optnew = statget(my_options,'Display','final');
```

See Also statset

## Purpose <br> Create statistics options structure

Syntax

Description

```
statset
statset(statfun)
options = statset(...)
options = statset(fieldname1,val1,fieldname2,val2,...)
options = statset(oldopts,fieldname1,val1,fieldname2,val2,
    ...)
options = statset(oldopts,newopts)
```

statset with no input arguments and no output arguments displays all fields of a statistics options structure and their possible values.
statset(statfun) displays fields and default values used by the Statistics Toolbox function statfun. Specify statfun using a string name or a function handle.
options = statset(...) creates a statistics options structure options. With no input arguments, all fields of the options structure are an empty array ([ ]). With a specified statfun, function-specific fields are default values and the remaining fields are [ ]. Function-specific fields set to [ ] indicate that the function is to use its default value for that parameter. For available options, see Inputs.
options = statset(fieldname1,val1,fieldname2,val2,...) creates an options structure in which the named fields have the specified values. Any unspecified values are []. Use strings for field names. For fields that are string-valued, you must input the complete string for the value. If you provide an invalid string for a value, statset uses the default.
options =
statset(oldopts,fieldname1,val1,fieldname2, val2,...)
creates a copy of oldopts with the named parameters changed to the specified values.
options = statset(oldopts, newopts) combines an existing options structure, oldopts, with a new options structure, newopts. Any
parameters in newopts with nonempty values overwrite corresponding parameters in oldopts.

## Input <br> Arguments

## DerivStep

Relative difference used in finite difference derivative calculations. A positive scalar, or a vector of positive scalars the same size as the vector of parameters estimated by the Statistics Toolbox function using the options structure.

Display
Amount of information displayed by the algorithm.

- 'off' - Displays no information.
- 'final' - Displays the final output.
- 'iter' - Displays iterative output to the command window for some functions; otherwise displays the final output.


## FunValCheck

Check for invalid values, such as NaN or Inf, from the objective function.

- 'off'
- 'on'

GradObj
Flags whether the objective function returns a gradient vector as a second output.

- 'off'
- 'on'

Jacobian

Flags whether the objective function returns a Jacobian as a second output.

- 'off'
- 'on'


## MaxFunEvals

Maximum number of objective function evaluations allowed. Positive integer.

MaxIter
Maximum number of iterations allowed. Positive integer.

## OutputFen

The solver calls all output functions after each iteration.

- Function handle specified using @
- a cell array with function handles
- an empty array (default)


## Robust

Invoke robust fitting option.

- 'off'
- 'on'


## Streams

A single instance of the RandStream class, or a cell array of RandStream instances. The Streams option is accepted by some functions to govern what stream(s) to use in generating random numbers within the function. If 'UseSubstreams' is 'always', the Streams value must be a scalar, or must be empty. If 'UseParallel' is 'always' and 'UseSubstreams' is 'never',
then the Streams argument must either be empty, or its length must match the number of processors used in the computation: equal to the matlabpool size if a matlabpool is open, a scalar otherwise. For more information on parallel computing, see Chapter 17, "Parallel Statistics".

TolBnd
Parameter bound tolerance. Positive scalar. TolFun

Termination tolerance for the objective function value. Positive scalar.

## TolTypeFun

Use TolFun for absolute or relative objective function tolerances.

- 'abs'
- 'rel'

TolTypeX
Use TolX for absolute or relative parameter tolerances.

- 'abs'
- 'rel'

TolX
Termination tolerance for the parameters. Positive scalar.

## Tune

The tuning constant used in robust fitting to normalize the residuals before applying the weight function. The default value depends upon the weight function. This parameter is necessary if you specify the weight function as a function handle. Positive scalar.

## UseParallel

Flag indicating whether eligible functions should use capabilities of the Parallel Computing Toolbox (PCT), if the capabilities are available. That is, if the PCT is installed, and a PCT matlabpool is in effect. Valid values are 'never' (the default), for serial computation, and 'always', for parallel computation. For more information on parallel computing, see Chapter 17, "Parallel Statistics".

## UseSubstreams

Flag indicating whether the random number generator in eligible functions should use Substream property of the RandStream class. 'never' (default) or 'always'. 'always', high level iterations within the function will set the Substream property to the value of the iteration. This behavior helps to generate reproducible random number streams in parallel and/or serial mode computation. For more information on parallel computing, see Chapter 17, "Parallel Statistics".

## WgtFun

A weight function for robust fitting. Valid only when Robust is 'on'. Can also be a function handle that accepts a normalized residual as input and returns the robust weights as output.

- 'bisquare'
- 'andrews'
- 'cauchy'
- 'fair'
- 'huber'
- 'logistic'
- 'talwar'
- 'welsch'

Examples

```
statset('evfit')
ans =
    Display: 'off'
    MaxFunEvals: []
            MaxIter: []
            TolBnd: []
            TolFun: []
            TolX: 1.0000e-006
            GradObj: []
            DerivStep: []
    FunValCheck: []
            Robust: []
            WgtFun: []
            Tune: []
```

The only parameters that evfit uses are Display and TolX. To create an options structure with the value of TolX set to $1 \mathrm{e}-8$, enter:

```
options = statset('TolX',1e-8)
% Pass options to evfit:
mu = 1;
sigma = 1;
data = evrnd(mu,sigma,1,100);
paramhat = evfit(data,[],[],[],options)
```

See Also statget

## ProbDistUnivParam.std

Purpose Return standard deviation of ProbDistUnivParam object
Syntax $S=\operatorname{std}(P D)$
Description $S=\operatorname{std}(P D)$ returns $S$, the standard deviation of the ProbDistUnivParam object PD.
Input ..... $P D$
An object of the class ProbDistUnivParam.
Arguments
Output
Arguments ..... $s$
See Also ..... std

## Purpose Interactive stepwise regression

```
Syntax stepwise
stepwise(X,y)
stepwise(X,y,inmodel,penter,premove)
```


## Description

stepwise uses the sample data in hald.mat to display a graphical user interface for performing stepwise regression of the response values in heat on the predictive terms in ingredients.

## stepwise



The upper left of the interface displays estimates of the coefficients for all potential terms, with horizontal bars indicating $90 \%$ (colored) and $95 \%$ (grey) confidence intervals. The red color indicates that, initially, the terms are not in the model. Values displayed in the table are those that would result if the terms were added to the model.

The middle portion of the interface displays summary statistics for the entire model. These statistics are updated with each step.

The lower portion of the interface, Model History, displays the RMSE for the model. The plot tracks the RMSE from step to step, so you can compare the optimality of different models. Hover over the blue dots in the history to see which terms were in the model at a particular step. Click on a blue dot in the history to open a copy of the interface initialized with the terms in the model at that step.

Initial models, as well as entrance/exit tolerances for the $p$-values of $F$-statistics, are specified using additional input arguments to stepwise. Defaults are an initial model with no terms, an entrance tolerance of 0.05 , and an exit tolerance of 0.10 .

To center and scale the input data (compute $z$-scores) to improve conditioning of the underlying least-squares problem, select Scale Inputs from the Stepwise menu.

You proceed through a stepwise regression in one of two ways:
1 Click Next Step to select the recommended next step. The recommended next step either adds the most significant term or removes the least significant term. When the regression reaches a local minimum of RMSE, the recommended next step is "Move no terms." You can perform all of the recommended steps at once by clicking All Steps.

2 Click a line in the plot or in the table to toggle the state of the corresponding term. Clicking a red line, corresponding to a term not currently in the model, adds the term to the model and changes the line to blue. Clicking a blue line, corresponding to a term currently in the model, removes the term from the model and changes the line to red.

To call addedvarplot and produce an added variable plot from the stepwise interface, select Added Variable Plot from the Stepwise menu. A list of terms is displayed. Select the term you want to add, and then click OK.

Click Export to display a dialog box that allows you to select information from the interface to save to the MATLAB workspace.

Check the information you want to export and, optionally, change the names of the workspace variables to be created. Click OK to export the information.
stepwise ( $\mathrm{X}, \mathrm{y}$ ) displays the interface using the $p$ predictive terms in the $n$-by- $p$ matrix X and the response values in the $n$-by- 1 vector y . Distinct predictive terms should appear in different columns of $X$.

Note stepwise automatically includes a constant term in all models. Do not enter a column of 1 s directly into X .
stepwise treats NaN values in either X or y as missing values, and ignores them.
stepwise( $\mathrm{X}, \mathrm{y}$, inmodel, penter, premove) additionally specifies the initial model (inmodel) and the entrance (penter) and exit (premove) tolerances for the $p$-values of $F$-statistics. inmodel is either a logical vector with length equal to the number of columns of $X$, or a vector of indices, with values ranging from 1 to the number of columns in $X$. The value of penter must be less than or equal to the value of premove.

## Algorithms

Stepwise regression is a systematic method for adding and removing terms from a multilinear model based on their statistical significance in a regression. The method begins with an initial model and then compares the explanatory power of incrementally larger and smaller models. At each step, the $p$ value of an $F$-statistic is computed to test models with and without a potential term. If a term is not currently in the model, the null hypothesis is that the term would have a zero coefficient if added to the model. If there is sufficient evidence to reject the null hypothesis, the term is added to the model. Conversely, if a term is currently in the model, the null hypothesis is that the term has a zero coefficient. If there is insufficient evidence to reject the null hypothesis, the term is removed from the model. The method proceeds as follows:

1 Fit the initial model.

2 If any terms not in the model have $p$-values less than an entrance tolerance (that is, if it is unlikely that they would have zero coefficient if added to the model), add the one with the smallest $p$ value and repeat this step; otherwise, go to step 3.

3 If any terms in the model have $p$-values greater than an exit tolerance (that is, if it is unlikely that the hypothesis of a zero coefficient can be rejected), remove the one with the largest $p$ value and go to step 2 ; otherwise, end.

Depending on the terms included in the initial model and the order in which terms are moved in and out, the method may build different models from the same set of potential terms. The method terminates when no single step improves the model. There is no guarantee, however, that a different initial model or a different sequence of steps will not lead to a better fit. In this sense, stepwise models are locally optimal, but may not be globally optimal.

## See Also

addedvarplot | regress | stepwisefit

## Purpose Stepwise regression

```
Syntax
b = stepwisefit(X,y)
[b,se,pval,inmodel,stats,nextstep,history] = stepwisefit(...)
[...] = stepwisefit(X,y,param1,val1,param2,val2,...)
```


## Description

b = stepwisefit ( $X, y$ ) uses a stepwise method to perform a multilinear regression of the response values in the $n$-by- 1 vector y on the $p$ predictive terms in the $n$-by- $p$ matrix X. Distinct predictive terms should appear in different columns of $\mathrm{X} . \mathrm{b}$ is a $p$-by- 1 vector of estimated coefficients for all of the terms in $X$. The value in $b$ for a term not included in the final model is the coefficient estimate that would result from adding the term to the model.

Note stepwisefit automatically includes a constant term in all models. Do not enter a column of 1 s directly into $X$.
stepwisefit treats NaN values in either X or y as missing values, and ignores them.
[b,se,pval,inmodel,stats, nextstep, history] = stepwisefit(...) returns the following additional information:

- se - A vector of standard errors for b
- pval - A vector of $p$-values for testing whether elements of b are 0
- inmodel - A logical vector, with length equal to the number of columns in $X$, specifying which terms are in the final model
- stats - A structure of additional statistics with the following fields. All statistics pertain to the final model except where noted.
- source - The string 'stepwisefit'
- dfe - Degrees of freedom for error
- df0 - Degrees of freedom for the regression
- SStotal - Total sum of squares of the response
- SSresid - Sum of squares of the residuals
- fstat - $F$-statistic for testing the final model vs. no model (mean only)
- pval - $p$ value of the $F$-statistic
- rmse - Root mean square error
- xr - Residuals for predictors not in the final model, after removing the part of them explained by predictors in the model
- yr - Residuals for the response using predictors in the final model
- B - Coefficients for terms in final model, with values for a term not in the model set to the value that would be obtained by adding that term to the model
- SE - Standard errors for coefficient estimates
- TSTAT - $t$ statistics for coefficient estimates
- PVAL — $p$-values for coefficient estimates
- intercept - Estimated intercept
- wasnan - Indicates which rows in the data contained NaN values
- nextstep - The recommended next step-either the index of the next term to move in or out of the model, or 0 if no further steps are recommended
- history - A structure containing information on steps taken, with the following fields:
- rmse - Root mean square errors for the model at each step
- df0 - Degrees of freedom for the regression at each step
- in - Logical array indicating which predictors are in the model at each step


## stepwisefit

[...] = stepwisefit(X,y,param1,val1,param2,val2,...) specifies one or more of the name/value pairs described in the following table.

| Parameter | Value |
| :--- | :--- |
| 'inmodel' | A logical vector specifying terms to include in the <br> initial fit. The default is to specify no terms. |
| 'penter' | The maximum $p$ value for a term to be added. The <br> default is 0.05. |
| 'premove' | The minimum $p$ value for a term to be removed. The <br> default is the maximum of the value of 'penter' and <br> 0.10. |
| 'display' | 'on' displays information about each step in the <br> command window. This is the default. <br> 'off' omits the display. |
| 'maxiter' | The maximum number of steps in the regression. The <br> default is Inf. |
| 'keep' | A logical vector specifying terms to keep in their initial <br> state. The default is to specify no terms. |
| 'scale' | 'on' centers and scales each column of $X$ (computes <br> $z$-scores) before fitting. <br> 'off' does not scale the terms. This is the default. |

## Algorithms

Stepwise regression is a systematic method for adding and removing terms from a multilinear model based on their statistical significance in a regression. The method begins with an initial model and then compares the explanatory power of incrementally larger and smaller models. At each step, the $p$ value of an $F$-statistic is computed to test models with and without a potential term. If a term is not currently in the model, the null hypothesis is that the term would have a zero coefficient if added to the model. If there is sufficient evidence to reject
the null hypothesis, the term is added to the model. Conversely, if a term is currently in the model, the null hypothesis is that the term has a zero coefficient. If there is insufficient evidence to reject the null hypothesis, the term is removed from the model. The method proceeds as follows:

1 Fit the initial model.
2 If any terms not in the model have $p$-values less than an entrance tolerance (that is, if it is unlikely that they would have zero coefficient if added to the model), add the one with the smallest $p$ value and repeat this step; otherwise, go to step 3.

3 If any terms in the model have $p$-values greater than an exit tolerance (that is, if it is unlikely that the hypothesis of a zero coefficient can be rejected), remove the one with the largest $p$ value and go to step 2 ; otherwise, end.

Depending on the terms included in the initial model and the order in which terms are moved in and out, the method may build different models from the same set of potential terms. The method terminates when no single step improves the model. There is no guarantee, however, that a different initial model or a different sequence of steps will not lead to a better fit. In this sense, stepwise models are locally optimal, but may not be globally optimal.

## Examples

Load the data in hald.mat, which contains observations of the heat of reaction of various cement mixtures:

| load hald <br> whos <br> Name |  |  |  |
| :--- | :--- | ---: | :--- |
|  | Size | Bytes | Class Attributes |
| Description | $22 \times 58$ | 2552 | char |
| hald | $13 \times 5$ | 520 | double |
| heat | $13 \times 1$ | 104 | double |
| ingredients | $13 \times 4$ | 416 | double |

The response (heat) depends on the quantities of the four predictors (the columns of ingredients).
Use stepwisefit to carry out the stepwise regression algorithm, beginning with no terms in the model and using entrance/exit tolerances of $0.05 / 0.10$ on the $p$-values:

```
stepwisefit(ingredients,heat,...
    'penter',0.05,'premove',0.10);
Initial columns included: none
Step 1, added column 4, p=0.000576232
Step 2, added column 1, p=1.10528e-006
Final columns included: 1 4
    'Coeff' 'Std.Err.' 'Status' 'P'
    [ 1.4400] [ 0.1384] 'In' [1.1053e-006]
    [ 0.4161] [ 0.1856] 'Out' [ 0.0517]
    [-0.4100] [ 0.1992] 'Out' [ 0.0697]
    [-0.6140] [ 0.0486] 'In' [1.8149e-007]
```

stepwisefit automatically includes an intercept term in the model, so you do not add it explicitly to ingredients as you would for regress. For terms not in the model, coefficient estimates and their standard errors are those that result if the term is added.

The inmodel parameter is used to specify terms in an initial model:

```
initialModel = ...
    [false true false false]; % Force in 2nd term
stepwisefit(ingredients,heat,...
    'inmodel',initialModel,...
    'penter',.05,'premove',0.10);
Initial columns included: 2
Step 1, added column 1, p=2.69221e-007
Final columns included: 1 2
    'Coeff' 'Std.Err.' 'Status' 'P'
    [ 1.4683] [ 0.1213] 'In' [2.6922e-007]
    [ 0.6623] [ 0.0459] 'In' [5.0290e-008]
    [ 0.2500] [ 0.1847] 'Out' [ 0.2089]
```

$$
\left[\begin{array}{ccc}
-0.2365] & 0.1733]
\end{array}\right] \text { 'Out' }\left[\begin{array}{l}
\text { [2054] }
\end{array}\right.
$$

The preceding two models, built from different initial models, use different subsets of the predictive terms. Terms 2 and 4 , swapped in the two models, are highly correlated:

```
term2 = ingredients(:,2);
term4 = ingredients(:,4);
R = corrcoef(term2,term4)
R =
    1.0000 -0.9730
    -0.9730 1.0000
```

To compare the models, use the stats output of stepwisefit:

```
[betahat1,se1,pval1,inmodel1,stats1] = ...
    stepwisefit(ingredients,heat,...
    'penter',.05,'premove',0.10,...
    'display','off');
[betahat2,se2,pval2,inmodel2,stats2] = ...
    stepwisefit(ingredients,heat,...
                                    'inmodel',initialModel,...
                                    'penter',.05,'premove',0.10,...
                                    'display','off');
RMSE1 = stats1.rmse
RMSE1 =
    2.7343
RMSE2 = stats2.rmse
RMSE2 =
    2.4063
```

The second model has a lower Root Mean Square Error (RMSE).

## References

See Also
[1] Draper, N. R., and H. Smith. Applied Regression Analysis. Hoboken, NJ: Wiley-Interscience, 1998. pp. 307-312.

Purpose Subscripted assignment for categorical array

## Syntax <br> $A=\operatorname{subsasgn}(A, S, B)$

Description
$A=\operatorname{subsasgn}(A, S, B)$ is called for the $\operatorname{syntax} A(i)=B . S$ is a structure array with the fields:

| type | String containing ' ( ) ' specifying <br> the subscript type. Only <br> parenthesis subscripting is <br> allowed. |
| :--- | :--- |
| subs | Cell array or string containing <br> the actual subscripts. |

## See Also <br> categorical | subsref

| Purpose | Subscripted reference for classregtree object |
| :--- | :--- |
| Syntax |  |
| Description | Subscript assignment is not allowed for a classregtree object. |
| See Also | classregtree |

Purpose
Description

Subscripted assignment to dataset array
$A=\operatorname{subsasgn}(A, S, B)$ is called for the $\operatorname{syntax} A(i, j)=B, A\{i, j\}=B$, or $A$. var $=B$ when $A$ is a dataset array. $S$ is a structure array with the fields:

| type | String containing ' ( ) ', ' $\}$ ', or |
| :--- | :--- |
| subs | '. specifying the subscript type. |
|  | Cell array or string containing |
|  | the actual subscripts. |

$A(i, j)=B$ assigns the contents of the dataset array $B$ to a subset of the observations and variables in the dataset array A. i and j are one of the following types:

- positive integers
- vectors of positive integers
- observation/variable names
- cell arrays containing one or more observation/variable names
- logical vectors

The assignment does not use observation names, variable names, or any other properties of $B$ to modify properties of $A$; however properties of $A$ are extended with default values if the assignment expands the number of observations or variables in A. Elements of B are assigned into A by position, not by matching names.
$A\{i, j\}=B$ assigns the value $B$ into an element of the dataset array $A$. $i$ and $J$ are positive integers, or logical vectors. Cell indexing cannot assign into multiple dataset elements, that is, the subscripts i and $j$ must each refer to only a single observation or variable. B is cast to the type of the target variable if necessary. If the dataset element already exists, A\{i, j\} may also be followed by further subscripting as supported by the variable.

For dataset variables that are cell arrays, assignments such as $\mathrm{A}\{1$, 'CellVar' $\}=\mathrm{B}$ assign into the contents of the target dataset element in the same way that \{\}-indexing of an ordinary cell array does.

For dataset variables that are n-D arrays, i.e., each observation is a matrix or array, an assignment such as A\{1,'ArrayVar'\} = B assigns into the second and following dimensions of the target dataset element, i.e., the assignment adds a leading singleton dimension to $B$ to account for the observation dimension of the dataset variable.
A.var $=B$ or $A .($ varname $)=B$ assigns B to a dataset variable. var is a variable name literal, or varname is a character variable containing a variable name. If the dataset variable already exists, the assignment completely replaces that variable. To assign into an element of the variable, A. var or A. (varname) may be followed by further subscripting as supported by the variable. In particular, A.var(obsnames,...) = $B$ and A.var\{obsnames, ...\} = B (when supported by var) provide assignment into a dataset variable using observation names.
A.properties.propertyname $=P$ assigns to a dataset property. propertyname is one of the following:

- 'ObsNames'
- 'VarNames'
- 'Description'
- 'Units'
- 'DimNames'
- 'UserData'
- 'VarDescription'

To assign into an element of the property, A. properties.propertyname may also be followed by further subscripting as supported by the property.

You cannot assign multiple values into dataset variables or properties using assignments such as [A.CellVar\{1:2\}] = B,

## dataset.subsasgn

[A.StructVar(1:2).field] = B, or [A.Properties.ObsNames\{1:2\}] $=$ B. Use multiple assignments of the form A.CellVar $\{1\}=$ B instead.
Similarly, if a dataset variable is a cell array with multiple columns or is an n-D cell array, then the contents of that variable for a single observation consists of multiple cells, and you cannot assign to all of them using the syntax A\{1,'CellVar' $\}=B$. Use multiple assignments of the form [A.CellVar\{1,1\}] = B instead.

## See Also dataset | set | subsref

| Purpose | Subscripted reference for Gaussian mixture distribution object |
| :--- | :--- |
| Description | Subscript assignment is not allowed for gmdistribution objects. |
| See Also | gmdistribution |

## NaiveBayes.subsasgn

## Purpose Subscripted reference for NaiveBayes object

Description Subscript assignment is not allowed for a NaiveBayes object.

Purpose Subscript index for categorical array
Syntax I $=$ subsindex $(A)$
Description
I = subsindex (A) is called for the syntax ${ }^{\prime} X(A)^{\prime}$ ' when $A$ is a categorical array and $X$ is one of the built-in types (most commonly 'double'). subsindex returns the internal categorical level codes of A converted to zero-based integer indices. subsindex is invoked separately on all the subscripts in an expression such as $X(A, B)$.

```
Examples
load fisheriris
a = ordinal(species,[],unique(species));
colmeans = grpstats(meas,a,@mean);
residuals = meas - colmeans(a,:);
```

See Also

categorical | double

## classregtree.subsref

Purpose Subscripted reference for classregtree object

## Syntax <br> $B=\operatorname{subsref}(T, S)$

Description
$B=$ subsref $(T, S)$ is called for the syntax $T(X)$ when $T$ is a classregtree object. $S$ is a structure array with the fields:
type String containing '()', '\{\}', or '.' specifying the subscript type.
subs Cell array or string containing the actual subscripts.
[...]=T(...) invokes the eval method for the tree T.

## See Also <br> classregtree | eval

Purpose Subscripted reference for categorical array
Syntax $\quad A=\operatorname{subsref}(A, S, B)$
Description
$A=\operatorname{subsref}(A, S, B)$ is called for the $\operatorname{syntax} A(I)=B$. S is a structure array with the fields:

| type | String containing ' ()' specifying <br> the subscript type. Only <br> parenthesis subscripting is <br> allowed. |
| :--- | :--- |
| subs | Cell array or string containing <br> the actual subscripts. |

[^16]
## dataset.subsref

## Purpose Subscripted reference for dataset array

## Syntax <br> $B=\operatorname{subsref}(A, S)$

Description
$B=\operatorname{subsref}(A, S)$ is called for the syntax $A(i, j), A\{i, j\}$, or $A . v a r$ when $A$ is a dataset array. $S$ is a structure array with the fields:

```
type String containing '()', '{}', or '.' specifying the subscript type.
subs Cell array or string containing the actual subscripts.
```

$B=A(i, j)$ returns a dataset array that contains a subset of the observations and variables in the dataset array A. i and $j$ are one of the following types:

- positive integers
- vectors of positive integers
- observation/variable names
- cell arrays containing one or more observation/variable names
- logical vectors

B contains the same property values as A, subsetted for observations or variables where appropriate.
$B=A\{i, j\}$ returns an element of a dataset variable. $i$ and $j$ are positive integers, or logical vectors. Cell indexing cannot return multiple dataset elements, that is, the subscripts $i$ and $j$ must each refer to only a single observation or variable. A\{i,j\} may also be followed by further subscripting as supported by the variable.

For dataset variables that are cell arrays, expressions such as A\{1, 'CellVar' \} return the contents of the referenced dataset element in the same way that \{ $\}$-indexing on an ordinary cell array does. If the dataset variable is a single column of cells, the contents of a single cell
is returned. If the dataset variable has multiple columns or is n-D, multiple outputs containing the contents of multiple cells are returned.

For dataset variables that are n-D arrays, i.e., each observation is a matrix or an array, expressions such as A\{1, 'ArrayVar'\} return A. ArrayVar $(1,:, \ldots)$ with the leading singleton dimension squeezed out.

B = A.var or A. (varname) returns a dataset variable. var is a variable name literal, or varname is a character variable containing a variable name. A.var or A. (varname) may also be followed by further subscripting as supported by the variable. In particular, A.var(obsnames,...) and A.var\{obsnames,...\} (when supported by var) provide subscripting into a dataset variable using observation names.
$P=A . P r o p e r t i e s . p r o p e r t y n a m e ~ r e t u r n s ~ a ~ d a t a s e t ~ p r o p e r t y . ~$ propertyname is one of the following:

- 'ObsNames'
- 'VarNames'
- 'Description'
- 'Units'
- 'DimNames'
- 'UserData'
- 'VarDescription'
A.properties.propertyname may also be followed by further subscripting as supported by the property.


## Limitations

Subscripting expressions such as A.CellVar\{1:2\}, A.StructVar(1:2).field, or A.Properties.ObsNames $\{1: 2\}$ are valid, but result in subsref returning multiple outputs in the form of a comma-separated list. If you explicitly assign to output
arguments on the left-hand side of an assignment, for example, [cellval1,cellval2] = A.CellVar\{1:2\}, those variables will receive the corresponding values. However, if there are no output arguments, only the first output in the comma-separated list is returned.
Similarly, if a dataset variable is a cell array with multiple columns or is an $n$-D cell array, then subscripting expressions such as A\{1, 'CellVar'\} result in subsref returning the contents of multiple cells. You should explicitly assign to output arguments on the left-hand side of an assignment, for example, [cellval1, cellval2] = A\{1, 'CellVar'\}.

## See Also

dataset | set | subsasgn

Purpose
Subscripted reference for Gaussian mixture distribution object

## Syntax

$B=\operatorname{subsref}(T, S)$
Description
$B=\operatorname{subsref}(T, S)$ is called for the syntax $T(X)$ when $T$ is a gmdistribution object. $S$ is a structure array with the following fields:
$\begin{array}{ll}\text { type } & \begin{array}{l}\text { String containing ' ( })^{\prime} \text {, ' }\} \text { ', or '.' specifying the } \\ \text { subscript type. }\end{array} \\ \text { subs } & \text { Cell array or string containing the actual subscripts. }\end{array}$

See Also gmdistribution

## NaiveBayes.subsref

## Purpose Subscripted reference for NaiveBayes object

## Syntax <br> b = subsref(nb,s)

Description
$\mathrm{b}=\operatorname{subsref}(\mathrm{nb}, \mathrm{s})$ is called for the syntax $\mathrm{nb}(\mathrm{s})$ when nb is a NaiveBayes object. $S$ is a structure array with the fields:
type
subs
string containing '()', '\}, or '.' specifying the subscript type.

Cell array or string containing the actual subscripts.

## Purpose Subscripted reference for qrandset

Syntax
$x=p(i, j)$
$x=\operatorname{subsref}(p, s)$

Description
$x=p(i, j)$ returns a matrix that contains a subset of the points from the point set $p$. The indices in i select points from the set and the indices in $j$ select columns from those points. i and $j$ are vector of positive integers or logical vectors. A colon used as a subscript, as in $p(i,:)$, indicates the entire row (or column).
$x=\operatorname{subsref}(p, s)$ is called for the syntax $p(i), p\{i\}$, or $p . i . s$ is a structure array with the fields:
type string containing '()', '\}', or '.' specifying the subscript type.
subs Cell array or string containing the actual subscripts.

## Examples

| Command | Returns |
| :--- | :--- |
| $p=\operatorname{sobolset}(5) ;$ | The fifth point |
| $x=p(1: 10,:)$ | All columns of the first 10 points |
| $x=p(e n d, 1)$ | The first column of the last point |
| $x=p([1,4,5],:)$ | Points 1,4, and 5 |

See Also qrandset

Purpose Summary statistics for categorical array
Syntax summary ( $A$ )
C = summary (A)
[C,labels] = summary(A)

## Description

summary (A) displays the number of elements in the categorical array A equal to each of the possible levels in A. If A contains any undefined elements, the output also includes the number of undefined elements.
$C=$ summary (A) returns counts of the number of elements in the categorical array A equal to each of the possible levels in A. If A is a matrix or $N$-dimensional array, C is a matrix or array with rows corresponding to the levels of A. If A contains any undefined elements, C contains one more row than the number of levels of $A$, with the number of undefined elements in $c$ (end) or $c$ (end,: ).
[C,labels] = summary (A) also returns the list of categorical level labels corresponding to the counts in C .

Examples Count the number of patients in each age group in the data in hospital.mat:

```
load hospital
edges = 0:10:100;
labels = strcat(num2str((0:10:90)','%d'),{'s'});
AgeGroup = ordinal(hospital.Age,labels,[],edges);
[c,labels] = summary(AgeGroup);
Table = dataset({labels,'AgeGroup'},{c,'Count'});
Table(3:6,:)
ans =
    AgeGroup Count
    '20s' }1
    '30s' 41
    '40s' 42
    '50s' 2
```

See Also islevel | ismember | levelcounts

## Purpose

Print summary of dataset array

| Syntax | summary $(A)$ |
| :--- | :--- |
| $s=\operatorname{summary}(A)$ |  |

summary (A) prints a summary of a dataset array and the variables that it contains.
s = summary (A) returns a scalar structure $s$ that contains a summary of the dataset A and the variables that A contains. For more information on the fields in s, see Outputs.

Summary information depends on the type of the variables in the data set:

- For numerical variables, summary computes a five-number summary of the data, giving the minimum, the first quartile, the median, the third quartile, and the maximum.
- For logical variables, summary counts the number of trues and falses in the data.
- For categorical variables, summary counts the number of data at each level.


## Output <br> Arguments

The following list describes the fields in the structure s:

- Description - A character array containing the dataset description.
- Variables - A structure array with one element for each dataset variable in A. Each element has the following fields:
- Name - A character string containing the name of the variable.
- Description - A character string containing the variable's description.
- Units - A character string containing the variable's units.
- Size - A numeric vector containing the size of the variable.
- Class - A character string containing the class of the variable.
- Data - A scalar structure containing the following fields.

For numeric variables:

- Probabilities - A numeric vector containing the probabilities [0.0 . 25.50 .751 .0 ] and NaN (if any are present in the corresponding dataset variable).
- Quantiles - A numeric vector containing the values that correspond to 'Probabilities' for the corresponding dataset variable, and a count of NaNs (if any are present).
For logical variables:
- Values - The logical vector [true false].
- Counts - A numeric vector of counts for each logical value.

For categorical variables:

- Levels - A cell array containing the labels for each level of the corresponding dataset variable.
- Counts - A numeric vector of counts for each level.
'Data' is empty if variable is not numeric, categorical, or logical. If a dataset variable has more than one column, then the corresponding 'Quantiles' or 'Counts' field is a matrix or an array.


## Examples

Summarize Fisher's iris data:

```
load fisheriris
species = nominal(species);
data = dataset(species,meas);
summary(data)
species: [150x1 nominal]
    setosa versicolor virginica
        50 50 50
```

| meas: [150x4 double] |  |  |  |  |
| :--- | :---: | ---: | ---: | ---: |
| min | 4.3000 | 2 | 1 | 0.1000 |
| 1st Q | 5.1000 | 2.8000 | 1.6000 | 0.3000 |
| median | 5.8000 | 3 | 4.3500 | 1.3000 |
| 3rd Q | 6.4000 | 3.3000 | 5.1000 | 1.8000 |
| max | 7.9000 | 4.4000 | 6.9000 | 2.5000 |

Summarize the data in hospital.mat:

```
load hospital
summary(hospital)
```

Dataset array created from the data file hospital.dat.
The first column of the file ("id") is used for observation names. Other columns ("sex" and "smoke") have been converted from their original coded values into categorical and logical variables. Two sets of columns ("sys" and "dia", "trial1" through "trial4") have been combined into single variables with multivariate observations. Column headers have been replaced with more descriptive variable names. Units have been added where appropriate.

```
LastName: [100x1 cell string]
```

Sex: [100x1 nominal]
Female Male
$53 \quad 47$

Age: [100x1 double, Units = Yrs]
min 1 st $Q$ median 3 rd $Q$ max

| 25 | 32 | 39 | 44 | 50 |
| :--- | :--- | :--- | :--- | :--- |


| Weight: $[100 \times 1$ | double, Units $=$ Lbs] |  |  |  |
| :---: | :---: | :---: | :--- | :--- |
| min | 1 st $Q$ | median | 3 rd $Q$ | max |
| 111 | 130.5000 | 142.5000 | 180.5000 | 202 |

```
Smoker: [100x1 logical]
    true false
            34 66
BloodPressure: [100x2 double, Units = mm Hg]
Systolic/Diastolic
    min }1096
    1st Q 117.5000 77.5000
    median 122 81.5000
    3rd Q 127.5000 89
    max 138 99
    Trials: [100x1 cell, Units = Counts]
    From zero to four measurement trials performed
```


## See Also

get | set | grpstats

## ProbDist.Support property

Purpose Read-only structure containing information about support of ProbDist object

## Description

Support is a read-only property of the ProbDist class. Support is a structure containing information about the support of a ProbDist object. It includes the following fields:

- range
- closedbound
- iscontinuous


## Values

The values for the three fields in the structure are:

- range - A two-element vector [L, U], such that all of the probability is contained from $L$ to $U$.
- closedbound - A two-element logical vector indicating whether the corresponding range endpoint is included. Possible values for each endpoint are 1 (true) or 0 (false).
- iscontinuous - A logical value indicates if the distribution takes values on the entire interval from $L$ to $U$ (true), or if it takes only integer values within this range (false). Possible values are 1 (true) or 0 (false).

Use this information to view and compare information about the support of distributions.

## Purpose <br> Interactive contour plot

Syntax
surfht (Z)
surfht ( $x, y, Z$ )
surfht ( $Z$ ) is an interactive contour plot of the matrix $Z$ treating the values in $Z$ as height above the plane. The $x$-values are the column indices of $Z$ while the $y$-values are the row indices of $Z$.
surfht $(x, y, Z)$ where $x$ and $y$ are vectors specify the $x$ and $y$-axes on the contour plot. The length of $x$ must match the number of columns in $Z$, and the length of $y$ must match the number of rows in $Z$.

There are vertical and horizontal reference lines on the plot whose intersection defines the current $x$ value and $y$ value. You can drag these dotted white reference lines and watch the interpolated $z$ value (at the top of the plot) update simultaneously. Alternatively, you can get a specific interpolated $z$ value by typing the $x$ value and $y$ value into editable text fields on the $x$-axis and $y$-axis respectively.

## classregtree.surrcutcategories

Purpose Categories used for surrogate splits in decision tree
Syntax
C = surrcutcategories(T)
C = surrcutcategories(T,J)

Description
$C=$ surrcutcategories $(T)$ returns an $n$-element cell array $C$ of the categories used for surrogate splits in the decision tree T, where $n$ is the number of nodes in the tree. For each node $K, C\{K\}$ is a cell array. The length of $\mathrm{C}\{\mathrm{K}\}$ is equal to the number of surrogate predictors found at this node. Every element of C\{K\} is either an empty string for a continuous surrogate predictor or a two-element cell array with categories for a categorical surrogate predictor. The first element of this two-element cell array lists categories assigned to the left child by this surrogate split and the second element of this two-element cell array lists categories assigned to the right child by this surrogate split. The order of the surrogate split variables at each node is matched to the order of variables returned by surrcutvar. The optimal-split variable at this node is not included. For non-branch (leaf) nodes, C contains an empty cell.
$C=$ surrcutcategories $(T, J)$ takes an array $J$ of node numbers and returns the categories for the specified nodes.

See Also classregtree | surrcutvar | cutcategories | surrcuttype | surrcutpoint

## classregtree.surrcutflip

## Purpose

Numeric cutpoint assignments used for surrogate splits in decision tree

Syntax

Description
$V=$ surrcutflip(T)
V = surrcutflip(T,J)
$V=$ surrcutflip $(T)$ returns an $n$-element cell array $V$ of the numeric cut assignments used for surrogate splits in the decision tree $T$, where $n$ is the number of nodes in the tree. For each node $\mathrm{K}, \mathrm{V}\{\mathrm{K}\}$ is a numeric vector. The length of $\vee\{K\}$ is equal to the number of surrogate predictors found at this node. Every element of $\vee\{K\}$ is either zero for a categorical surrogate predictor or a numeric cut assignment for a continuous surrogate predictor. The numeric cut assignment can be either -1 or +1 . For every surrogate split with a numeric cut $C$ based on a continuous predictor variable $Z$, the left child is chosen if $Z<C$ and the cut assignment for this surrogate split is +1 , or if $Z>=C$ and the cut assignment for this surrogate split is -1 . Similarly, the right child is chosen if $Z>=C$ and the cut assignment for this surrogate split is +1 , or if $Z<C$ and the cut assignment for this surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables returned by surrcutvar. The optimal-split variable at this node is not included. For non-branch (leaf) nodes, V contains an empty array.
$V=$ surrcutflip $(T, J)$ takes an array $J$ of node numbers and returns the cutpoint assignments for the specified nodes.

See Also<br>classregtree | surrcutvar | surrcutcategories | surrcuttype | surrcutpoint | cutpoint

## classregtree.surrcutpoint

Purpose Cutpoints used for surrogate splits in decision tree
Syntax $\quad \begin{aligned} & V=\operatorname{surrcutpoint}(T) \\ & V=\operatorname{surrcutpoint}(T, J)\end{aligned}$
Description
$\mathrm{V}=$ surrcutpoint $(\mathrm{T})$ returns an $n$-element cell array V of the numeric values used for surrogate splits in the decision tree T , where $n$ is the number of nodes in the tree. For each node K, V $\{K\}$ is a numeric vector. The length of $V\{K\}$ is equal to the number of surrogate predictors found at this node. Every element of $V\{K\}$ is either either NaN for a categorical surrogate predictor or a numeric cut for a continuous surrogate predictor. For every surrogate split with a numeric cut $C$ based on a continuous predictor variable $Z$, the left child is chosen if $Z<C$ and surrcutflip for this surrogate split is -1 . Similarly, the right child is chosen if $Z>=C$ and surrcutflip for this surrogate split is +1 , or if $Z<C$ and surrcutflip for this surrogate split is -1 . The order of the surrogate split variables at each node is matched to the order of variables returned by surrcutvar. The optimal-split variable at this node is not included. For non-branch (leaf) nodes, V contains an empty cell.
$V=$ surrcutpoint $(T, J)$ takes an array $J$ of node numbers and returns the cutpoint assignments for the specified nodes.

See Also classregtree | surrcutvar | surrcutcategories | surrcuttype | surrcutflip | cutpoint

## Purpose

Types of surrogate splits used at branches in decision tree

## Syntax

C = surrcuttype( T )
C = surrcuttype(T,J)
$C=$ surrcuttype ( $T$ ) returns an $n$-element cell array $C$ indicating types of surrogate splits at each node in the tree T , where $n$ is the number of nodes in the tree. For each node K, C $\{K\}$ is a cell array with the types of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The order of the surrogate split variables at each node is matched to the order of variables returned by surrcutvar. The optimal-split variable at this node is not included. For non-branch (leaf) nodes, C contains an empty cell. A surrogate split type can be either 'continuous' if the cut is defined in the form $Z<V$ for a variable $Z$ and cutpoint $V$ or 'categorical' if the cut is defined by whether $Z$ takes a value in a set of categories.
$C=$ surrcuttype $(T, J)$ takes an array $J$ of node numbers and returns the cut types for the specified nodes.

[^17]
## classregtree.surrcutvar

Purpose Variables used for surrogate splits in decision tree

| Syntax | $V=\operatorname{surrcutvar}(T)$ |
| :--- | :--- |
|  | $V=\operatorname{surrcutvar}(T, J)$ |
|  | $[V, N U M]=\operatorname{surrcutvar}(\ldots)$ |

Description
$\mathrm{V}=$ surrcutvar $(\mathrm{T})$ returns an $n$-element cell array V of the names of the variables used for surrogate splits in each node of the tree T, where $n$ is the number of nodes in the tree. Every element of V is a cell array with the names of the surrogate split variables at this node. The variables are sorted by the predictive measure of association with the optimal predictor in the descending order, and only variables with the positive predictive measure are included. The optimal-split variable at this node is not included. For non-branch (leaf) nodes, V contains an empty cell.
$V=$ surrcutvar $(T, J)$ takes an array $J$ of node numbers and returns the cut types for the specified nodes.
[V,NUM] = surrcutvar(...) also returns a cell array NUM with indices for each variable.

See Also classregtree | numnodes | children | cutvar

Purpose
Predictive measure of association for surrogate splits in decision tree
Syntax
$A=$ surrvarassoc ( $T$ )
$A=\operatorname{sur} r v a r a s s o c(T, J)$
$\mathrm{A}=\operatorname{sur} r$ varassoc $(\mathrm{T})$ returns an $n$-element cell array A of the predictive measures of association for surrogate splits in the decision tree T, where $n$ is the number of nodes in the tree. For each node K, $A\{K\}$ is a numeric vector. The length of $A\{K\}$ is equal to the number of surrogate predictors found at this node. Every element of A\{K\} gives the predictive measure of association between the optimal split and this surrogate split. The order of the surrogate split variables at each node is matched to the order of variables returned by surrcutvar. The optimal-split variable at this node is not included. For non-branch (leaf) nodes, V contains an empty cell.
$A=\operatorname{surrvarassoc}(T, J)$ takes an array $J$ of node numbers and returns the predictive measure of association for the specified nodes.

## See Also

classregtree | surrcutcategories | surrcuttype | surrcutflip | surrcutpoint | surrcutvar
Purpose Frequency table
Syntax TABLE = tabulate(x) tabulate(x)
Description TABLE $=$ tabulate $(x)$ creates a frequency table of data in vector $x$.Information in TABLE is arranged as follows:

- 1st column - The unique values of $x$
- 2nd column - The number of instances of each value
- 3rd column - The percentage of each value
If $x$ is a numeric array, TABLE is a numeric matrix. If the elements of $x$ are nonnegative integers, TABLE includes 0 counts for integers between 1 and $\max (x)$ that do not appear in $x$.
If $x$ is a categorical variable, character array, or cell array of strings, TABLE is a cell array. (See "Grouped Data" on page 2-34.)
tabulate ( $x$ ) with no output arguments displays the table in the command window.


## Examples

| Value | Count | Percent |
| :---: | :---: | :---: |
| 1 | 1 | 16.67\% |
| 2 | 1 | 16.67\% |
| 3 | 1 | 16.67\% |
| 4 | 3 | 50.00\% |

See Also pareto
How To - "Grouped Data" on page 2-34

## Purpose Read tabular data from file

## Syntax <br> Description

[data,varnames,casenames] = tblread [data,varnames,casenames] = tblread(filename)
[data, varnames,casenames] = tblread(filename,delimiter)
[data, varnames, casenames] = tblread displays the File Open dialog box for interactive selection of a tabular data file. The file format has variable names in the first row, case names in the first column and data starting in the $(2,2)$ position. Outputs are:

- data - Numeric matrix with a value for each variable-case pair
- varnames - String matrix containing the variable names in the first row of the file
- casenames - String matrix containing the names of each case in the first column of the file
[data,varnames,casenames] = tblread(filename) allows command line specification of the name of a file in the current folder, or the complete path name of any file, using the string filename.
[data, varnames, casenames] = tblread(filename,delimiter) reads from the file using delimiter as the delimiting character. Accepted values for delimiter are:
- ' ' or 'space'
- '\t' or 'tab'
- ', ' or 'comma'
- ';' or 'semi'
- '|' or 'bar'

The default value of delimiter is 'space'.

## tblread

```
Examples
    [data,varnames,casenames] = tblread('sat.dat')
    data =
        470 530
        520480
    varnames =
    Male
    Female
    casenames =
    Verbal
    Quantitative
See Also
    tblwrite | tdfread | caseread
```

Purpose Write tabular data to file

```
Syntax
tblwrite(data, varnames, casenames)
tblwrite(data, varnames, casenames,filename)
tblwrite(data, varnames, casenames,filename,delimiter)
```


## Description

## Examples

Continuing the example from tblread:

```
tblwrite(data,varnames,casenames,'sattest.dat')
```

```
type sattest.dat
                            Male Female
Verbal 470530
Quantitative 520480
```

See Also
casewrite | tblread

## Purpose

Student's $t$ cumulative distribution function

## Syntax

$P=\operatorname{tcdf}(X, V)$
$P=\operatorname{tcdf}(X, V)$ computes Student's $t c d f$ at each of the values in $X$ using the corresponding degrees of freedom in $V$. $X$ and $V$ can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs.

The $t \mathrm{cdf}$ is

$$
p=F(x \mid v)=\int_{-\infty}^{x} \frac{\Gamma\left(\frac{v+1}{2}\right)}{\Gamma\left(\frac{v}{2}\right)} \frac{1}{\sqrt{v \pi}} \frac{1}{\left(1+\frac{t^{2}}{v}\right)^{\frac{v+1}{2}}} d t
$$

The result, $p$, is the probability that a single observation from the $t$ distribution with $v$ degrees of freedom will fall in the interval $[-\infty, x)$.

## Examples

```
mu = 1; % Population mean
sigma = 2; % Population standard deviation
n = 100; % Sample size
x = normrnd(mu,sigma,n,1); % Random sample from population
xbar = mean(x); % Sample mean
s = std(x); % Sample standard deviation
t = (xbar-mu)/(s/sqrt(n)) % t-statistic
t =
    0.2489
p = 1-tcdf(t,n-1) % Probability of larger t-statistic
p =
    0.4020
```

This probability is the same as the $p$ value returned by a $t$-test of the null hypothesis that the sample comes from a normal population with mean $\mu$ :

```
[h,ptest] = ttest(x,mu,0.05,'right')
h =
        0
ptest =
    0.4020
```

See Also cdf | tpdf | tinv | tstat | trnd
How To . "Student's t Distribution" on page B-95

## Purpose Read tab-delimited file

## Syntax

```
tdfread
tdfread(filename)
tdfread(filename,delimiter)
s = tdfread(filename,...)
```

Description
tdfread displays the File Open dialog box for interactive selection of a data file, then reads data from the file. The file should have variable names separated by tabs in the first row, and data values separated by tabs in the remaining rows. tdfread creates variables in the workspace, one for each column of the file. The variable names are taken from the first row of the file. If a column of the file contains only numeric data in the second and following rows, tdfread creates a double variable. Otherwise, tdfread creates a char variable. After all values are imported, tdfread displays information about the imported values using the format of the tdfread command.
tdfread(filename) allows command line specification of the name of a file in the current folder, or the complete path name of any file, using the string filename.
tdfread(filename, delimiter) indicates that the character specified by delimiter separates columns in the file. Accepted values for delimiter are:

- ' ' or 'space'
- '\t' or 'tab'
- ', ' or 'comma
- ';' or 'semi'
- '|' or 'bar'

The default delimiter is 'tab '.
s = tdfread(filename, ...) returns a scalar structure s whose fields each contain a variable.

## tdfread

Examples The following displays the contents of the file sat2.dat:

```
type sat2.dat
Test,Gender, Score
Verbal,Male,470
Verbal,Female,530
Quantitative,Male,520
Quantitative,Female,480
```

The following creates the variables Gender, Score, and Test from the file sat2.dat and displays the contents of the MATLAB workspace:

```
tdfread('sat2.dat',',')
```

Name Size Bytes Class Attributes

| Gender | $4 \times 6$ | 48 | char |
| :--- | :--- | :--- | :--- |
| Score | $4 \times 1$ | 32 | double |
| Test | $4 \times 12$ | 96 | char |

See Also tblread | caseread

## ClassificationTree.template

| Purpose | Create classification template |
| :---: | :---: |
| Syntax | t = ClassificationTree.template |
|  | $\mathrm{t}=$ ClassificationTree.template (Name, Value) |
| Description | $t=$ ClassificationTree.template returns a learner template suitable to use in the fitensemble function. |
|  | $\mathrm{t}=$ ClassificationTree.template (Name, Value) creates a template with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN. |
| Input Arguments | Name-Value Pair Arguments |
|  | Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN. |
|  | MergeLeaves |
|  | String that specifies whether to merge leaves after the tree is grown. Values are 'on' or 'off'. |
|  | When 'on', ClassificationTree merges leaves that originate from the same parent node, and that give a sum of risk values greater or equal to the risk associated with the parent node. When 'off', ClassificationTree does not merge leaves. |
|  | Default: 'off' |
|  | MinLeaf |
|  | Each leaf has at least MinLeaf observations per tree leaf. If you supply both MinParent and MinLeaf, ClassificationTree uses the setting that gives larger leaves: MinParent=max (MinParent, 2*MinLeaf). |

## ClassificationTree.template

Default: Half the number of training observations for boosting, 1 for bagging

MinParent
Each branch node in the tree has at least MinParent observations. If you supply both MinParent and MinLeaf, ClassificationTree uses the setting that gives larger leaves: MinParent=max(MinParent,2*MinLeaf).

Default: Number of training observations for boosting, 2 for bagging

## NVarToSample

Number of predictors to select at random for each split. Can be a positive integer or 'all', which means use all available predictors.

Default: 'all' for boosting, square root of number of predictors for bagging

## Prune

When 'on', ClassificationTree grows the classification tree and computes the optimal sequence of pruned subtrees. When 'off' ClassificationTree grows the tree without pruning.

Default: 'off'
PruneCriterion
String with the pruning criterion, either 'error' or 'impurity'.
Default: 'error'

SplitCriterion

## ClassificationTree.template

Criterion for choosing a split. One of 'gdi' (Gini's diversity index), 'twoing' for the twoing rule, or 'deviance' for maximum deviance reduction (also known as cross entropy).
Default: 'gdi'

## Surrogate

String describing whether to find surrogate decision splits at each branch node. This setting improves the accuracy of predictions for data with missing values. The setting also enables you to compute measures of predictive association between predictors. This setting can use much time and memory. Values are 'on' or 'off'.
Default: 'off'

## Output Arguments

## Examples

t
Classification tree template suitable to use in the fitensemble function. In an ensemble, $t$ specifies how to grow the classification trees.

Create a classification template with surrogate splits, and train an ensemble for the Fisher iris model with the template:

```
t = ClassificationTree.template('surrogate','on');
load fisheriris
ens = fitensemble(meas,species,'AdaBoostM2',100,t);
```

See Also
ClassificationTree | ClassificationTree.fit | fitensemble
How To - Chapter 13, "Nonparametric Supervised Learning"

## RegressionTree.template

Purpose Create regression template
Syntax $\quad \begin{aligned} \mathrm{t} & =\text { RegressionTree.template } \\ \mathrm{t} & =\text { RegressionTree.template (Name, Value) }\end{aligned}$
Description
$\mathrm{t}=$ RegressionTree. template returns a learner template suitable to use in the fitensemble function.
$\mathrm{t}=$ RegressionTree.template (Name, Value) creates a template with additional options specified by one or more Name, Value pair arguments. You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN.

## Input Arguments

## Name-Value Pair Arguments

Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' '). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN.

MergeLeaves
String that specifies whether to merge leaves after the tree is grown. Values are 'on' or 'off'.

When 'on', RegressionTree merges leaves that originate from the same parent node, and that give a sum of risk values greater or equal to the risk associated with the parent node. When 'off', RegressionTree does not merge leaves.

Default: 'off'
MinLeaf
Each leaf has at least MinLeaf observations per tree leaf. If you supply both MinParent and MinLeaf, RegressionTree uses the setting that gives larger leaves: MinParent=max(MinParent,2*MinLeaf).

Default: Half the number of training observations for boosting, 5 for bagging

## MinParent

Each branch node in the tree has at least MinParent observations. If you supply both MinParent and MinLeaf, RegressionTree uses the setting that gives larger leaves: MinParent=max (MinParent, 2*MinLeaf).

Default: Number of training observations for boosting, 10 for bagging

## NVarToSample

Number of predictors to select at random for each split. Can be a positive integer or 'all', which means use all available predictors.

Default: 'all' for boosting, one third of the number of predictors for bagging

## Prune

When 'on', RegressionTree grows the regression tree and computes the optimal sequence of pruned subtrees. When 'off' RegressionTree grows the tree without pruning.

Default: 'off'

## Surrogate

String describing whether to find surrogate decision splits at each branch node. This setting improves the accuracy of predictions for data with missing values. The setting also enables you to compute measures of predictive association between predictors. This setting can use much time and memory. Values are 'on' or 'off'.

Default: 'off'

## RegressionTree.template

Output tArguments
Regression tree template suitable to use in the fitensemble function. In an ensemble, $t$ specifies how to grow the regression trees.
Examples $\quad \begin{aligned} & \text { Create a regression template with surrogate splits, and train an } \\ & \text { ensemble for the carsmall data with the template: }\end{aligned}$ ensemble for the carsmall data with the template:

```
t = RegressionTree.template('surrogate','on');
load carsmall
X = [Acceleration Displacement Horsepower Weight];
ens = fitensemble(X,MPG,'LSBoost',100,t);
```

See Also
RegressionTree | RegressionTree.fit | fitensemble
How To • Chapter 13, "Nonparametric Supervised Learning"

Purpose<br>Syntax<br>Description

Error rate
cost = test(t,'resubstitution')
cost = test(t,'test', X,y)
cost = test(t,'crossvalidate',X,y)
[cost,secost,ntnodes,bestlevel] = test(...)
[...] = test(...,param1,val1,param2,val2,...)
cost = test(t,'resubstitution') computes the cost of the tree t using a resubstitution method. t is a decision tree as created by classregtree. The cost of the tree is the sum over all terminal nodes of the estimated probability of a node times the cost of a node. If $t$ is a classification tree, the cost of a node is the sum of the misclassification costs of the observations in that node. If $t$ is a regression tree, the cost of a node is the average squared error over the observations in that node. cost is a vector of cost values for each subtree in the optimal pruning sequence for $t$. The resubstitution cost is based on the same sample that was used to create the original tree, so it under estimates the likely cost of applying the tree to new data.
cost $=$ test( t, 'test', $\mathrm{X}, \mathrm{y}$ ) uses the matrix of predictors X and the response vector y as a test sample, applies the decision tree $t$ to that sample, and returns a vector cost of cost values computed for the test sample. $X$ and $y$ should not be the same as the learning sample, that is, the sample that was used to fit the tree $t$.
cost $=$ test( $\mathrm{t}, \mathrm{C}$ crossvalidate', $\mathrm{X}, \mathrm{y}$ ) uses 10 -fold cross-validation to compute the cost vector. X and y should be the learning sample, that is, the sample that was used to fit the tree $t$. The function partitions the sample into 10 subsamples, chosen randomly but with roughly equal size. For classification trees, the subsamples also have roughly the same class proportions. For each subsample, test fits a tree to the remaining data and uses it to predict the subsample. It pools the information from all subsamples to compute the cost for the whole sample.
[cost, secost, ntnodes, bestlevel] = test(...) also returns the vector secost containing the standard error of each cost value, the vector ntnodes containing the number of terminal nodes for each

## classregtree.test

subtree, and the scalar bestlevel containing the estimated best level of pruning. A bestlevel of 0 means no pruning. The best level is the one that produces the smallest tree that is within one standard error of the minimum-cost subtree.
[...] = test(...,param1,val1,param2,val2,...) specifies optional parameter name/value pairs for methods other than 'resubstitution', chosen from the following:

- 'weights' - Observation weights.
- 'nsamples' - The number of cross-validation samples (default is 10).
- 'treesize' - Either 'se' (default) to choose the smallest tree whose cost is within one standard error of the minimum cost, or 'min' to choose the minimal cost tree.


## Examples

Find the best tree for Fisher's iris data using cross-validation. Start with a large tree:

```
load fisheriris;
t = classregtree(meas,species,...
    'names',{'SL' 'SW' 'PL' 'PW'},...
    'splitmin',5)
t =
Decision tree for classification
    if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
    class = setosa
    if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
    if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
    class = virginica
    if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
    if PW<1.55 then node 10 elseif PW>=1.55 then node 11 else virginica
    class = versicolor
    class = virginica
    class = virginica
```

| Click to display: | Identity | $\rightarrow$ Magnification: | $100 \%$ | $\rightarrow$ | Pruning level: 0 of $\mathbf{4}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |



Find the minimum-cost tree:

```
[c,s,n,best] = test(t,'cross',meas,species);
tmin = prune(t,'level',best)
tmin =
Decision tree for classification
```

```
if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 class = versicolor
5 class = virginica
view(tmin)
```

| Click to display: | Identity | $\square$ | Magnification: | $100 \%$ | $\square$ | Pruning level: |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 0 of 2 |  |  |  |  |  |  |



Plot the smallest tree within one standard error of the minimum cost tree:

```
[mincost,minloc] = min(c);
plot(n,c,'b-o',...
    n(best+1),c(best+1),'bs',...
    n,(mincost+s(minloc))*ones(size(n)),'k--')
xlabel('Tree size (number of terminal nodes)')
ylabel('Cost')
```



The solid line shows the estimated cost for each tree size, the dashed line marks one standard error above the minimum, and the square marks the smallest tree under the dashed line.

## classregtree.test

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.<br>See Also classregtree | eval | view | prune

## Purpose Test indices for cross-validation

## Syntax <br> idx $=$ test (c) <br> idx = test(c,i)

Description
$i d x=$ test (c) returns the logical vector idx of test indices for an object c of the cvpartition class of type 'holdout' or 'resubstitution'.
If c.Type is 'holdout', idx specifies the observations in the test set.
If c.Type is 'resubstitution', idx specifies all observations.
idx $=$ test $(c, i)$ returns the logical vector idx of test indices for repetition $i$ of an object $c$ of the cvpartition class of type 'kfold' or 'leaveout'.

If c.Type is 'kfold', idx specifies the observations in the test set in fold i.

If c. Type is 'leaveout', idx specifies the observation left out at repetition i.

Examples Identify the test indices in the first fold of a partition of 10 observations for 3 -fold cross-validation:

```
c = cvpartition(10,'kfold',3)
c =
K-fold cross validation partition
            N: 10
        NumTestSets: 3
        TrainSize: 7 6 7
            TestSize: 3 4 3
test(c,1)
ans =
        1
        1
        0
        0
```


## cvpartition.test

0
0
0
0
1
0
See Also
cvpartition | training

## Purpose Size of each test set

$\begin{array}{ll}\text { Description } \quad & \text { Value is a vector in partitions of type 'kfold' and 'leaveout'. } \\ & \text { Value is a scalar in partitions of type 'holdout' and } \\ \text { 'resubstitution'. }\end{array}$

Purpose Rank adjusted for ties

Syntax $\quad$| $[R$, TIEADJ $]$ | $=$ tiedrank $(X)$ |
| ---: | :--- |
| $[R, T I E A D J]$ | $=$ tiedrank $(X, 1)$ |
| $[R, T I E A D J]$ | $=$ tiedrank $(X, 0,1)$ |

## Description

## Examples

## See Also ansaribradley | corr | partialcorr | ranksum | signrank

Purpose Product of categorical arrays
Syntax $\quad C=\operatorname{times}(A, B)$
Description $\quad C=$ times $(A, B)$ returns a categorical array each of whose elements has the level formed from the concatenation of the levels of the corresponding elements of $A$ and $B$. The set of levels of $C$ is the cartesian product of the sets of levels of A and of B. The syntax A .* B calls C $=$ times(A,B).

See Also categorical

Purpose
Student's $t$ inverse cumulative distribution function

## Syntax <br> $X=\operatorname{tinv}(P, V)$

$X=\operatorname{tinv}(P, V)$ computes the inverse of Student's $t$ cdf using the degrees of freedom in $V$ for the corresponding probabilities in $P$. $P$ and $V$ can be vectors, matrices, or multidimensional arrays that are the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The values in P must lie on the interval [01].

The t inverse function in terms of the $t$ cdf is

$$
x=F^{-1}(p \mid v)=\{x: F(x \mid v)=p\}
$$

where

$$
p=F(x \mid v)=\int_{-\infty}^{x} \frac{\Gamma\left(\frac{v+1}{2}\right)}{\Gamma\left(\frac{v}{2}\right)} \frac{1}{\sqrt{v \pi}} \frac{1}{\left(1+\frac{t^{2}}{v}\right)^{\frac{v+1}{2}}} d t
$$

The result, $x$, is the solution of the cdf integral with parameter $v$, where you supply the desired probability $p$.

Examples What is the 99th percentile of the $t$ distribution for one to six degrees of freedom?

```
percentile = tinv(0.99,1:6)
percentile =
    31.8205 6.9646 4.5407
```

See Also icdf | tcdf | tpdf | trnd | tstat
How To

- "Student's t Distribution" on page B-95


## Purpose

Student's $t$ probability density function

## Syntax

$Y=\operatorname{tpdf}(X, V)$
$Y=\operatorname{tpdf}(X, V)$ computes Student's $t$ pdf at each of the values in $X$ using the corresponding degrees of freedom in $V$. $X$ and $V$ can be vectors, matrices, or multidimensional arrays that have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs.
Student's $t \mathrm{pdf}$ is

$$
y=f(x \mid v)=\frac{\Gamma\left(\frac{v+1}{2}\right)}{\Gamma\left(\frac{v}{2}\right)} \frac{1}{\sqrt{v \pi}} \frac{1}{\left(1+\frac{x^{2}}{v}\right)^{\frac{v+1}{2}}}
$$

## Examples

See Also
pdf | tcdf | tinv \| tstat | trnd

How To . "Student's t Distribution" on page B-95

## Purpose Training indices for cross-validation

```
Syntax
idx = training(c)
idx = training(c,i)
```

Identify the training indices in the first fold of a partition of 10 observations for 3 -fold cross-validation:

```
c = cvpartition(10,'kfold',3)
c =
K-fold cross validation partition
            N: 10
        NumTestSets: 3
        TrainSize: 7 6 7
            TestSize: 3 4 3
training(c,1)
ans =
        0
            0
```

1
1
1
1
1
1
0
1

## See Also <br> cvpartition | test

Purpose Size of each training set
Description
Value is a vector in partitions of type 'kfold' and 'leaveout'.
Value is a scalar in partitions of type 'holdout' and 'resubstitution'.

## See Also <br> type

## categorical.transpose

Purpose Transpose categorical matrix

## Syntax <br> $B=t r a n s p o s e(A)$

Description $B=$ transpose ( $A$ ) returns the transpose of the 2-D categorical matrix A. ctranspose is identical to transpose for categorical arrays. The syntax A.' calls transpose.

See Also ctranspose | permute

Purpose Cell array of arguments for classregtree
Description The TreeArgs property is a cell array of arguments for the classregtree constructor. TreeBagger uses these arguments in growing new trees for the ensemble.

## TreeBagger

## Purpose Bootstrap aggregation for ensemble of decision trees

TreeBagger bags an ensemble of decision trees for either classification or regression. Bagging stands for bootstrap aggregation. Every tree in the ensemble is grown on an independently drawn bootstrap replica of input data. Observations not included in this replica are "out of bag" for this tree. To compute prediction of an ensemble of trees for unseen data, TreeBagger takes an average of predictions from individual trees. To estimate the prediction error of the bagged ensemble, you can compute predictions for each tree on its out-of-bag observations, average these predictions over the entire ensemble for each observation and then compare the predicted out-of-bag response with the true value at this observation.

TreeBagger relies on the classregtree functionality for growing individual trees. In particular, classregtree accepts the number of features selected at random for each decision split as an optional input argument.

The Compact property contains another class, CompactTreeBagger, with sufficient information to make predictions using new data. This information includes the tree ensemble, variable names, and class names (for classification). CompactTreeBagger requires less memory than TreeBagger, but only TreeBagger has methods for growing more trees for the ensemble. Once you grow an ensemble of trees using TreeBagger and no longer need access to the training data, you can opt to work with the compact version of the trained ensemble from then on.

## Construction TreeBagger

Create ensemble of bagged decision trees

## Methods

append
compact

Append new trees to ensemble
Compact ensemble of decision trees

| error | Error (misclassification <br> probability or MSE) |
| :--- | :--- |
| fillProximities | Proximity matrix for training <br> data |
| growTrees | Train additional trees and add to <br> ensemble <br> Classification margin |
| margin | Multidimensional scaling of <br> proximity matrix |
| mdsProx | Mean classification margin |
| meanMargin | Out-of-bag error |
| oobError | Out-of-bag margins |
| oobMargin | Out-of-bag mean margins |
| oobMeanMargin | Ensemble predictions for |
| oobPredict | out-of-bag observations |
| predict | Predict response |

## Properties

## ClassNames

A cell array containing the class names for the response variable Y . This property is empty for regression trees.

Compute00BPrediction
A logical flag specifying whether out-of-bag predictions for training observations should be computed. The default is false.

If this flag is true, the following properties are available:

- OOBIndices
- OOBInstanceWeight

If this flag is true, the following methods can be called:

## TreeBagger

- oobError
- oobMargin
- oobMeanMargin

See also oobError, OOBIndices, OOBInstanceWeight, oobMargin, oobMeanMargin.

Cost
A matrix with misclassification costs. This property is empty for ensembles of regression trees.

See also classregtree.

## DeltaCritDecisionSplit

A numeric array of size 1-by-Nvars of changes in the split criterion summed over splits on each variable, averaged across the entire ensemble of grown trees.

See also classregtree.varimportance.
NVarSplit
A numeric array of size 1-by-Nvars, where every element gives a number of splits on this predictor summed over all trees.

## OOBPermutedVarDeltaError

A numeric array of size 1-by-Nvars containing a measure of importance for each predictor variable (feature). For any variable, the measure is the increase in prediction error if the values of that variable are permuted across the out-of-bag observations. This measure is computed for every tree, then averaged over the entire ensemble and divided by the standard deviation over the entire ensemble.

## OOBPermutedVarDeltaMeanMargin

A numeric array of size 1-by-Nvars containing a measure of importance for each predictor variable (feature). For any variable, the measure is the decrease in the classification margin if the
values of that variable are permuted across the out-of-bag observations. This measure is computed for every tree, then averaged over the entire ensemble and divided by the standard deviation over the entire ensemble. This property is empty for regression trees.

## OutlierMeasure

A numeric array of size Nobs-by-1, where Nobs is the number of observations in the training data, containing outlier measures for each observation.

See also CompactTreeBagger.OutlierMeasure.
Prior
A vector with prior probabilities for classes. This property is empty for ensembles of regression trees.

See also classregtree.
Proximity
A numeric matrix of size Nobs-by-Nobs, where Nobs is the number of observations in the training data, containing measures of the proximity between observations. For any two observations, their proximity is defined as the fraction of trees for which these observations land on the same leaf. This is a symmetric matrix with 1 s on the diagonal and off-diagonal elements ranging from 0 to 1 .

See also CompactTreeBagger.proximity, classregtree.varimportance.

Prune
The Prune property is true if decision trees are pruned and false if they are not. Pruning decision trees is not recommended for ensembles. The default value is false.

See also classregtree. prune.
SampleWithReplacement

## TreeBagger

A logical flag specifying if data are sampled for each decision tree with replacement. True if TreeBagger samples data with replacement and false otherwise. True by default.

## Trees

A cell array of size NTrees-by- 1 containing the trees in the ensemble.

See also NTrees.
VarAssoc
A matrix of size Nuars-by-Nuars with predictive measures of variable association, averaged across the entire ensemble of grown trees. If you grew the ensemble setting 'surrogate' to ' on ', this matrix for each tree is filled with predictive measures of association averaged over the surrogate splits. If you grew the ensemble setting 'surrogate' to 'off' (default), VarAssoc is diagonal.

## VarNames

A cell array containing the names of the predictor variables (features). TreeBagger takes these names from the optional 'names ' parameter. The default names are 'x1', 'x2', etc.

X
A numeric matrix of size Nobs-by-Nvars, where Nobs is the number of observations (rows) and Nuars is the number of variables (columns) in the training data. This matrix contains the predictor (or feature) values.

## Y

An array of true class labels for classification, or response values for regression. Y can be a numeric column vector, a character matrix, or a cell array of strings.

| Copy | Value. To learn how this affects your use of the class, see Comparing <br> Semantics |
| :--- | :--- |
| Handle and Value Classes in the MATLAB Object-Oriented <br> Programming documentation. |  |
| How To | - "Ensemble Methods" on page 13-51 |
|  | - "Classification Trees and Regression Trees" on page 13-27 |
|  | - Grouped Data |

## TreeBagger

| Purpose | Create ensemble of bagged decision trees |
| :---: | :---: |
| Syntax | ```B = TreeBagger(ntrees,X,Y) B = TreeBagger(ntrees,X,Y,'param1',val1,'param2',val2,...)``` |
| Description | $B=$ TreeBagger (ntrees, $X, Y$ ) creates an ensemble $B$ of ntrees decision trees for predicting response $Y$ as a function of predictors $X$. By default TreeBagger builds an ensemble of classification trees. The function can build an ensemble of regression trees by setting the optional input argument 'method' to 'regression'. |
|  | $X$ is a numeric matrix of training data. Each row represents an observation and each column represents a predictor or feature. Y is an array of true class labels for classification or numeric function values for regression. True class labels can be a numeric vector, character matrix, vector cell array of strings or categorical vector. TreeBagger converts labels to a cell array of strings for classification. |
|  | For more information on grouping variables, see "Grouped Data" on page 2-34. |
|  | $B=$ TreeBagger(ntrees, $\mathrm{X}, \mathrm{Y}$, 'param1', val1, 'param2', val2, $\ldots$ ) specifies optional parameter name/value pairs: |
|  | 'FBoot ' Fraction of input data to sample with replacement <br> from the input data for growing each new tree. <br> Default value is 1. |
|  | 'OOBPred' 'on' to store info on what observations are out of bag for each tree. This info can be used by oobPredict to compute the predicted class probabilities for each tree in the ensemble. Default is 'off'. |
|  | 'OOBVarImp' 'on' to store out-of-bag estimates of feature importance in the ensemble. Default is 'off'. Specifying 'on' also sets the 'OOBPred' value to 'on'. |


| 'Method ' | Either 'classification' or 'regr Regression requires a numeric Y . |
| :---: | :---: |
| 'NVarToSample' | Number of variables to select at random for each decision split. Default is the square root of the number of variables for classification and one third of the number of variables for regression. Valid values are 'all' or a positive integer. Setting this argument to any valid value but 'all' invokes Breiman's 'random forest' algorithm. |
| 'NPrint' | Number of training cycles (grown trees) after which TreeBagger displays a diagnostic message showing training progress. Default is no diagnostic messages. |
| 'MinLeaf' | Minimum number of observations per tree leaf. Default is 1 for classification and 5 for regression. |
| 'Options ' | A structure that specifies options that govern the computation when growing the ensemble of decision trees. One option requests that the computation of decision trees on multiple bootstrap replicates uses multiple processors, if the Parallel Computing Toolbox is available. Two options specify the random number streams to use in selecting bootstrap replicates. You can create this argument with a call to statset. You can retrieve values of the individual fields with a call to statget. Applicable statset parameters are: <br> - 'UseParallel' - If 'always' and if a matlabpool of the Parallel Computing Toolbox is open, compute decision trees drawn on separate boostrap replicates in parallel. If the Parallel Computing Toolbox is not installed, or a matlabpool is not open, computation occurs in serial mode. Default is 'never', or serial computation. |

## TreeBagger

- 'UseSubstreams' - If 'always' select each bootstrap replicate using a separate Substream of the random number generator (aka Stream). This option is available only with RandStream types that support Substreams: 'mlfg6331_64' or 'mrg32k3a'. Default is 'never', do not use a different Substream to compute each bootstrap replicate.
- Streams - A RandStream object or cell array of such objects. If you do not specify Streams, TreeBagger uses the default stream or streams. If you choose to specify Streams, use a single object except in the case
- You have an open MATLAB pool
- UseParallel is 'always'
- UseSubstreams is 'never'

In that case, use a cell array the same size as the MATLAB pool.

For more information on using parallel computing, see Chapter 17, "Parallel Statistics".
In addition to the optional arguments above, this method accepts all optional classregtree arguments with the exception of 'minparent'. Refer to the documentation for classregtree for more detail.

## Examples

```
load fisheriris
b = TreeBagger(50,meas,species,'00BPred','on')
plot(oobError(b))
xlabel('number of grown trees')
ylabel('out-of-bag classification error')
```

returns
b =

```
        Ensemble with 50 bagged decision trees:
            Training X:
            [150\times4]
            Training Y: [150\times1]
                                    Method: classification
                                    Nvars: 4
            NVarToSample: 2
            MinLeaf: 1
            FBoot: 1
                SampleWithReplacement: 1
            Compute00BPrediction: 1
            Compute00BVarImp: 0
                    Proximity: []
                            Prune: 0
                            MergeLeaves: 0
            TreeArgs:
            ClassNames:'setosa' 'versicolor' 'virginica'
```


classregtree | CompactTreeBagger
How To

- "Ensemble Methods" on page 13-51


## TreeBagger

- "Grouped Data" on page 2-34


## Purpose

Plot tree

Note treedisp will be removed in a future release. Use classregtree.view instead.

```
Syntax
treedisp(t)
treedisp(t,param1,val1,param2,val2,...)
```


## Description

Note This function is superseded by the view method of the classregtree class and is maintained only for backwards compatibility. It accepts objects $t$ created with the classregtree constructor.
treedisp ( $t$ ) takes as input a decision tree $t$ as computed by the treefit function, and displays it in a figure window. Each branch in the tree is labeled with its decision rule, and each terminal node is labeled with the predicted value for that node.

For each branch node, the left child node corresponds to the points that satisfy the condition, and the right child node corresponds to the points that do not satisfy the condition.

The Click to display pop-up menu at the top of the figure enables you to display more information about each node, as described in the following table.

| Menu Choice | Displays |
| :--- | :--- |
| Identity | The node number, whether the node is a branch <br> or a leaf, and the rule that governs the node |
| Variable ranges | The range of each of the predictor variables for <br> that node |
| Node statistics | Descriptive statistics for the observations falling <br> into this node |

## treedisp

After you select the type of information you want, click any node to display the information for that node.

The Pruning level button displays the number of levels that have been cut from the tree and the number of levels in the unpruned tree. For example, 1 of 6 indicates that the unpruned tree has six levels, and that one level has been cut from the tree. Use the spin button to change the pruning level.
treedisp(t,param1,val1,param2,val2,...) specifies optional parameter name-value pairs, listed in the following table.

| Parameter | Value |
| :--- | :--- |
| 'names ' | A cell array of names for the predictor variables, <br> in the order in which they appear in the $X$ matrix <br> from which the tree was created (see treefit) |
| 'prunelevel' | Initial pruning level to display |

## Examples

Create and graph classification tree for Fisher's iris data. The names in this example are abbreviations for the column contents (sepal length, sepal width, petal length, and petal width).

```
load fisheriris;
t = treefit(meas,species);
treedisp(t,'names',{'SL' 'SW' 'PL' 'PW'});
```

| Click to display: | Identity | $\square$ | Magnification: | $100 \%$ | $\square$ |
| :--- | :--- | :--- | :--- | :--- | :--- |



References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

See Also treefit | treeprune | treetest

Purpose Fit tree

Note treefit will be removed in a future release. Use classregtree instead.

```
Syntax \(\quad t=\operatorname{treefit}(x, y)\)
t = treefit(X,y,param1,val1,param2,val2,...)
```


## Description

Note This function is superseded by the classregtree constructor of the classregtree class and is maintained only for backwards compatibility. It returns objects $t$ in the classregtree class.
$t=\operatorname{treefit}(X, y)$ creates a decision tree $t$ for predicting response $y$ as a function of predictors $X$. $X$ is an $n$-by-m matrix of predictor values. $y$ is either a vector of $n$ response values (for regression), or a character array or cell array of strings containing $n$ class names (for classification). Either way, t is a binary tree where each non-terminal node is split based on the values of a column of $X$.
$\mathrm{t}=$ treefit $(\mathrm{X}, \mathrm{y}$, param1, val1, param2, val2, ...) specifies optional parameter name-value pairs. Valid parameter strings are:

The following table lists parameters available for all trees.

| Parameter | Value |
| :--- | :--- |
| 'catidx' | Vector of indices of the columns of X . treefit <br> treats these columns as unordered categorical <br> values. |
| 'method' | Either 'classification' (default if y is text) or <br> 'regression' (default if y is numeric). |


| Parameter | Value |
| :--- | :--- |
| 'splitmin' | A number n such that impure nodes must have n <br> or more observations to be split (default 10). |
| 'prune' | 'on' (default) to compute the full tree and a <br> sequence of pruned subtrees, or 'off' for the full <br> tree without pruning. |

The following table lists parameters available for classification trees only.

| Parameter | Value |
| :--- | :--- |
| 'cost' | p-by-p matrix C, where $p$ is the number of distinct <br> response values or class names in the input y. <br> $C(i, j)$ is the cost of classifying a point into class <br> $j$ if its true class is i. (The default has C(i,j)=1 <br> if i~=j, and C(i,j)=0 if i=j.) C can also be a <br> structure S with two fields: S.group containing <br> the group names (see "Grouped Data" on page <br> $2-34)$, and S.cost containing a matrix of cost <br> values. |
| 'splitcriterion ' | Criterion for choosing a split: either 'gdi' <br> (default) for Gini's diversity index, 'twoing' for <br> the twoing rule, or 'deviance' for maximum <br> deviance reduction. |
| 'priorprob' | Prior probabilities for each class, specified as a <br> vector (one value for each distinct group name) <br> or as a structure S with two fields: S.group <br> containing the group names, and S.prob <br> containing a vector of corresponding probabilities. |

Examples Create a classification tree for Fisher's iris data:

```
load fisheriris;
t = treefit(meas,species);
```


## treefit

```
treedisp(t,'names',{'SL' 'SW' 'PL' 'PW'});
```

| Click to display: | Identity | $\square$ | Magnification: | $100 \%$ | $\square$ | Pruning level: |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 0 of 4 |  |  |  |  |  |  |



References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification
See Also
treedisp | treetest
How To

- "Grouped Data" on page 2-34


## Purpose

Prune tree

Note treeprune will be removed in a future release. Use classregtree. prune instead.

## Syntax

```
t2 = treeprune(t1,'level',level)
t2 = treeprune(t1,'nodes',nodes)
t2 = treeprune(t1)
```


## Description

Note This function is superseded by the prune method of the classregtree class and is maintained only for backwards compatibility. It accepts objects $t 1$ created with the classregtree constructor and returns objects t 2 in the classregtree class.
t2 = treeprune(t1,'level',level) takes a decision tree t1 as created by the treefit function, and a pruning level, and returns the decision tree t2 pruned to that level. Setting level to 0 means no pruning. Trees are pruned based on an optimal pruning scheme that first prunes branches giving less improvement in error cost.
t2 = treeprune(t1, 'nodes', nodes) prunes the nodes listed in the nodes vector from the tree. Any t1 branch nodes listed in nodes become leaf nodes in t2, unless their parent nodes are also pruned. The treedisp function can display the node numbers for any node you select.
t2 $=$ treeprune( t 1 ) returns the decision tree t2 that is the same as t 1 , but with the optimal pruning information added. This is useful only if you created t1 by pruning another tree, or by using the treefit function with pruning set 'off'. If you plan to prune a tree multiple times, it is more efficient to create the optimal pruning sequence first.

Pruning is the process of reducing a tree by turning some branch nodes into leaf nodes, and removing the leaf nodes under the original branch.

Examples Display the full tree for Fisher's iris data, as well as the next largest tree from the optimal pruning sequence:

```
load fisheriris;
t1 = treefit(meas,species,'splitmin',5);
treedisp(t1,'names',{'SL' 'SW' 'PL' 'PW'});
```

| Click to display: Identity | $\checkmark$ | Magnification: | 100\% | $\checkmark$ | Pruning level: | 0 of 4 | - |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |





References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

See Also treefit | treetest | treedisp

## TreeBagger.Trees property

Purpose Decision trees in ensemble
Description The Trees property is a cell array of size NTrees-by- 1 containing the trees in the ensemble.
See Also ..... NTrees

## Purpose

Error rate
Syntax

```
cost = treetest(t,'resubstitution')
cost = treetest(t,'test',X,y)
cost = treetest(t,'crossvalidate',X,y)
[cost,secost,ntnodes,bestlevel] = treetest(...)
[...] = treetest(...,param1,val1,param2,val2,...)
```

Note treetest will be removed in a future release. Use classregtree.test instead.

## Description

Note This function is superseded by the test method of the classregtree class class and is maintained only for backwards compatibility. It accepts objects $t$ created with the classregtree constructor.
cost $=$ treetest( $t$, 'resubstitution') computes the cost of the tree $t$ using a resubstitution method. $t$ is a decision tree as created by the treefit function. The cost of the tree is the sum over all terminal nodes of the estimated probability of that node times the node's cost. If $t$ is a classification tree, the cost of a node is the sum of the misclassification costs of the observations in that node. If $t$ is a regression tree, the cost of a node is the average squared error over the observations in that node. cost is a vector of cost values for each subtree in the optimal pruning sequence for $t$. The resubstitution cost is based on the same sample that was used to create the original tree, so it underestimates the likely cost of applying the tree to new data.
cost $=$ treetest( $t$, 'test', $\mathrm{X}, \mathrm{y}$ ) uses the predictor matrix X and response $y$ as a test sample, applies the decision tree $t$ to that sample, and returns a vector cost of cost values computed for the test sample. X and y should not be the same as the learning sample, which is the sample that was used to fit the tree $t$.
cost $=$ treetest(t, 'crossvalidate', $\mathrm{X}, \mathrm{y}$ ) uses 10 -fold cross-validation to compute the cost vector. $X$ and $y$ should be the learning sample, which is the sample that was used to fit the tree $t$. The function partitions the sample into 10 subsamples, chosen randomly but with roughly equal size. For classification trees, the subsamples also have roughly the same class proportions. For each subsample, treetest fits a tree to the remaining data and uses it to predict the subsample. It pools the information from all subsamples to compute the cost for the whole sample.
[cost, secost, ntnodes,bestlevel] = treetest(...) also returns the vector secost containing the standard error of each cost value, the vector ntnodes containing number of terminal nodes for each subtree, and the scalar bestlevel containing the estimated best level of pruning. bestlevel $=0$ means no pruning, i.e., the full unpruned tree. The best level is the one that produces the smallest tree that is within one standard error of the minimum-cost subtree.
[...] = treetest(...,param1,val1,param2,val2,...) specifies optional parameter name-value pairs chosen from the following table.

| Parameter | Value |
| :--- | :--- |
| 'nsamples ' | The number of cross-validations samples (default <br> ic -10) |
| 'treesize' | Either 'se' (default) to choose the smallest tree <br> whose cost is within one standard error of the <br> minimum cost, or 'min' to choose the minimal cost <br> tree. |

## Examples

Find the best tree for Fisher's iris data using cross-validation. The solid line shows the estimated cost for each tree size, the dashed line marks one standard error above the minimum, and the square marks the smallest tree under the dashed line.

```
% Start with a large tree.
load fisheriris;
t = treefit(meas,species','splitmin',5);
```

```
% Find the minimum-cost tree.
[c,s,n,best] = treetest(t,'cross',meas,species);
tmin = treeprune(t,'level',best);
% Plot smallest tree within 1 std of minimum cost tree.
[mincost,minloc] = min(c);
plot(n,c,'b-o',...
    n(best+1),c(best+1),'bs',...
    n,(mincost+s(minloc))*ones(size(n)),'k--');
xlabel('Tree size (number of terminal nodes)')
ylabel('Cost')
```



References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.<br>See Also treefit | treedisp

## Purpose

Predicted responses

Note treeval will be removed in a future release. Use classregtree.eval instead.

Syntax<br>\section*{Description}

```
yfit = treeval(t,x)
yfit = treeval(t,X,subtrees)
[yfit,node] = treeval(...)
[yfit,node,cname] = treeval(...)
```

Note This function is superseded by the eval method of the classregtree class and is maintained only for backwards compatibility. It accepts objects $t$ created with the classregtree constructor.
yfit $=$ treeval $(\mathrm{t}, \mathrm{X})$ takes a classification or regression tree t as produced by the treefit function and a matrix $X$ of predictor values, and produces a vector yfit of predicted response values. For a regression tree, yfit(i) is the fitted response value for a point having the predictor values $\mathrm{X}(\mathrm{i},:)$. For a classification tree, yfit(i) is the class number into which the tree would assign the point with data X(i,:). To convert the number into a class name, use the third output argument, cname (described below).
yfit $=$ treeval( $\mathrm{t}, \mathrm{X}$, subtrees) takes an additional vector subtrees of pruning levels, with 0 representing the full, unpruned tree. T must include a pruning sequence as created by the treefit or prunetree function. If subtree has $k$ elements and X has $n$ rows, the output yfit is an $n$-by- $k$ matrix, with the $j$ th column containing the fitted values produced by the subtrees (j) subtree. subtrees must be sorted in ascending order.
[yfit, node] = treeval(...) also returns an array node of the same size as yfit containing the node number assigned to each row of $X$. The treedisp function can display the node numbers for any node you select.
[yfit, node, cname] = treeval(...) is valid only for classification trees. It returns a cell array cname containing the predicted class names.

## Examples Find the predicted classifications for Fisher's iris data: <br> load fisheriris; <br> t = treefit(meas,species); \% Create decision tree <br> sfit = treeval(t,meas); \% Find assigned class numbers <br> sfit = t.classname(sfit); \% Get class names <br> mean(strcmp(sfit,species)) \% Proportion in correct class <br> ans = <br> 0.9800

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

## See Also treefit | treeprune | treetest

| Purpose | Mean excluding outliers |
| :--- | :--- |
| Syntax | $m=\operatorname{trimmean}(X$, percent $)$ |
|  | trimmean $(X$, percent, dim $)$ |
|  | $m=\operatorname{trimmean}(X$, percent, flag $)$ |
|  | $m=\operatorname{trimmean}(x$, percent,flag, dim $)$ |

Description

Tips
'floor' Round $k$ down to the next smaller integer.
'weight ' If $k=i+f$ where $i$ is the integer part and $f$ is the fraction, compute a weighted mean with weight ( $1-f$ ) for the $(i+1)$ th and ( $n-i$ )th values, and full weight for the values between them.
$m=$ trimmean(x, percent,flag,dim) takes the trimmed mean along dimension dim of $x$.
$m=$ trimmean(X, percent) calculates the trimmed mean of the values in $X$. For a vector input, $m$ is the mean of $X$, excluding the highest and lowest k data values, where $\mathrm{k}=\mathrm{n}$ * (percent/100)/2 and where n is the number of values in $X$. For a matrix input, $m$ is a row vector containing the trimmed mean of each column of $X$. For $n$-D arrays, trimmean operates along the first non-singleton dimension. percent is a scalar between 0 and 100 .
trimmean (X, percent, dim) takes the trimmed mean along dimension dim of $X$.
$m=$ trimmean(X, percent,flag) controls how to trim when $k$ is not an integer. flag can be chosen from the following:

| 'round ' | Round $k$ to the nearest integer (round to a smaller <br> integer if $k$ is a half integer). This is the default. |
| :--- | :--- |
| 'floor' | Round $k$ down to the next smaller integer. |
| 'weight ' | If $k=i+f$ where $i$ is the integer part and $f$ is the <br> fraction, compute a weighted mean with weight <br> $(1-f)$ for the $(i+1)$ th and $(n-i)$ th values, and |
| full weight for the values between them. |  |

,
The trimmed mean is a robust estimate of the location of a sample. If there are outliers in the data, the trimmed mean is a more representative estimate of the center of the body of the data than the mean. However, if the data is all from the same probability distribution,

## trimmean

then the trimmed mean is less efficient than the sample mean as an estimator of the location of the data.

## Examples

## Example 1

This example shows a Monte Carlo simulation of the efficiency of the $10 \%$ trimmed mean relative to the sample mean for normal data.

```
x = normrnd(0,1,100,100);
m = mean(x);
trim = trimmean(x,10);
sm = std(m);
strim = std(trim);
efficiency = (sm/strim).^2
efficiency =
    0.9702
```


## Example 2

Generate random data from the $t$ distribution, which tends to have outliers:

```
rng('default') % to reproduce the plot exactly
x = trnd(1,40,1);
probplot(x)
```



Though the distribution is symmetric around zero, there are several outliers which will affect the mean. The trimmed mean is much closer to zero, which is much more representative of the data:

```
mean(x)
ans =
2.7991
trimmean(x,25)
ans =
```


## trimmean

0.8797

See Also mean | median | geomean | harmmean

## Purpose <br> Syntax <br> Description

Student's $t$ random numbers
$\mathrm{R}=\operatorname{trnd}(\mathrm{V})$
$R=\operatorname{trnd}(V, m, n, \ldots)$
$R=\operatorname{trnd}(V,[m, n, \ldots])$
$R=\operatorname{trnd}(\mathrm{V})$ generates random numbers from Student's $t$ distribution with $V$ degrees of freedom. V can be a vector, a matrix, or a multidimensional array. The size of $R$ is the size of $V$.
$R=\operatorname{trnd}(V, m, n, \ldots)$ or $R=\operatorname{trnd}(V,[m, n, \ldots])$ generates an $m$-by- $n$-by-... array. The $V$ parameter can be a scalar or an array of the same size as R.

## Examples

```
noisy = trnd(ones(1,6))
noisy =
    19.7250
numbers = trnd(1:6,[1 6])
numbers =
    -1.9500 -0.9611 -0.9038
numbers = trnd(3,2,6)
numbers =
    -0.3177 -0.0812 -0.6627 0.1905 -1.5585 -0.0433
    0.2536}00.5502 0.8646 0.8060 -0.5216 0.0891
```

See Also random | tpdf | tcdf | tinv | tstat
How To . "Student's t Distribution" on page B-95

Purpose
Student's $t$ mean and variance

## Syntax <br> [M,V] = tstat(NU)

$[M, V]=\operatorname{tstat}(N U)$ returns the mean of and variance for Student's $t$ distribution using the degrees of freedom in NU. M and $V$ are the same size as NU.

The mean of the Student's $t$ distribution with parameter $v$ is zero for values of $v$ greater than 1 . If $v$ is one, the mean does not exist. The variance for values of $v$ greater than 2 is $v /(v-2)$.

## Examples

Find the mean of and variance for 1 to 30 degrees of freedom.


Note that the variance does not exist for one and two degrees of freedom.

## See Also

## Purpose <br> One-sample and paired-sample $t$-test

## Syntax

```
h = ttest(x)
h = ttest(x,m)
h = ttest(x,y)
h = ttest(...,alpha)
h = ttest(...,alpha,tail)
h = ttest(...,alpha,tail,dim)
[h,p] = ttest(...)
[h,p,ci] = ttest(...)
[h,p,ci,stats] = ttest(...)
```


## Description

$h=$ ttest $(x)$ performs a $t$-test of the null hypothesis that data in the vector x are a random sample from a normal distribution with mean 0 and unknown variance, against the alternative that the mean is not 0 . The result of the test is returned in $\mathrm{h} . \mathrm{h}=1$ indicates a rejection of the null hypothesis at the $5 \%$ significance level. $\mathrm{h}=0$ indicates a failure to reject the null hypothesis at the $5 \%$ significance level.
x can also be a matrix or an $N$-dimensional array. For matrices, ttest performs separate $t$-tests along each column of x and returns a vector of results. For $N$-dimensional arrays, ttest works along the first non-singleton dimension of $x$.

The test treats NaN values as missing data, and ignores them. $\mathrm{h}=\mathrm{ttest}(\mathrm{x}, \mathrm{m})$ performs a $t$-test of the null hypothesis that data in the vector x are a random sample from a normal distribution with mean $m$ and unknown variance, against the alternative that the mean is not $m$.
$\mathrm{h}=\mathrm{ttest}(\mathrm{x}, \mathrm{y})$ performs a paired $t$-test of the null hypothesis that data in the difference $x-y$ are a random sample from a normal distribution with mean 0 and unknown variance, against the alternative that the mean is not $0 . x$ and $y$ must be vectors of the same length, or arrays of the same size.
$\mathrm{h}=\mathrm{ttest}(. .$. , alpha) performs the test at the (100*alpha)\% significance level. The default, when unspecified, is alpha $=0.05$.
h = ttest(..., alpha, tail) performs the test against the alternative specified by the string tail. There are three options for tail:

- 'both' - Mean is not 0 (or m) (two-tailed test). This is the default, when tail is unspecified.
- 'right ' — Mean is greater than 0 (or m) (right-tail test)
- 'left' - Mean is less than 0 (or m) (left-tail test)
tail must be a single string, even when x is a matrix or an $N$-dimensional array.
$h=$ ttest(..., alpha, tail, dim) works along dimension dim of $x$, or of $x-y$ for a paired test. Use [ ] to pass in default values for $m$, alpha, or tail.
$[\mathrm{h}, \mathrm{p}]=\operatorname{ttest}(\ldots)$ returns the $p$ value of the test. The $p$ value is the probability, under the null hypothesis, of observing a value as extreme or more extreme of the test statistic

$$
t=\frac{\bar{x}-\mu}{s / \sqrt{n}}
$$

where $\bar{x}$ is the sample mean, $\mu=0$ (or m ) is the hypothesized population mean, $s$ is the sample standard deviation, and $n$ is the sample size. Under the null hypothesis, the test statistic will have Student's $t$ distribution with $n-1$ degrees of freedom.
[h,p,ci] = ttest(...) returns a 100*(1 - alpha)\% confidence interval on the population mean, or on the difference of population means for a paired test.
[h,p,ci,stats] = ttest(...) returns the structure stats with the following fields:

- tstat - Value of the test statistic
- df - Degrees of freedom of the test
- sd - Sample standard deviation


## Examples

Simulate a random sample of size 100 from a normal distribution with mean 0.1:

$$
x=\operatorname{normrnd}(0.1,1,1,100) ;
$$

Test the null hypothesis that the sample comes from a normal distribution with mean 0 :

```
[h,p,ci] = ttest(x,0)
\(\mathrm{h}=\)
    0
\(p=\)
    0.8323
ci \(=\)
    \(-0.1650 \quad 0.2045\)
```

The test fails to reject the null hypothesis at the default $\alpha=0.05$ significance level. Under the null hypothesis, the probability of observing a value as extreme or more extreme of the test statistic, as indicated by the $p$ value, is much greater than $\alpha$. The $95 \%$ confidence interval on the mean contains 0 .

Simulate a larger random sample of size 1000 from the same distribution:

```
y = normrnd(0.1,1,1,1000);
```

Test again if the sample comes from a normal distribution with mean 0 :

```
[h,p,ci] = ttest(y,0)
h =
    1
p =
    0.0160
ci =
    0.0142 0.1379
```

This time the test rejects the null hypothesis at the default $\alpha=0.05$ significance level. The $p$ value has fallen below $\alpha=0.05$ and the $95 \%$ confidence interval on the mean does not contain 0 .

Because the $p$ value of the sample y is greater than 0.01 , the test will fail to reject the null hypothesis when the significance level is lowered to $\alpha=0.01$ :

```
[h,p,ci] = ttest(y, 0, 0.01)
\(h=\)
        0
\(p=\)
        0.0160
ci \(=\)
    \(-0.0053 \quad 0.1574\)
```

Notice that at the lowered significance level the $99 \%$ confidence interval on the mean widens to contain 0 .

This example will produce slightly different results each time it is run, because of the random sampling.

## See Also

ttest2 | ztest

## Purpose <br> Syntax <br> Description

Two-sample $t$-test

```
\(h=t t e s t 2(x, y)\)
\(h=t t e s t 2(x, y, a l p h a)\)
\(h=t t e s t 2(x, y, a l p h a, t a i l)\)
h = ttest2(x,y,alpha,tail,vartype)
\(h=\) ttest2(x,y,alpha,tail, vartype,dim)
[h, p] = ttest2(...)
[h, p, ci] = ttest2(...)
[h,p,ci,stats] = ttest2(...)
```

$\mathrm{h}=\mathrm{ttest2}(\mathrm{x}, \mathrm{y})$ performs a $t$-test of the null hypothesis that data in the vectors $x$ and $y$ are independent random samples from normal distributions with equal means and equal but unknown variances, against the alternative that the means are not equal. The result of the test is returned in $\mathrm{h} . \mathrm{h}=1$ indicates a rejection of the null hypothesis at the $5 \%$ significance level. $\mathrm{h}=0$ indicates a failure to reject the null hypothesis at the $5 \%$ significance level. $x$ and $y$ need not be vectors of the same length.
x and y can also be matrices or $N$-dimensional arrays. Matrices x and y must have the same number of columns, in which case ttest2 performs separate $t$-tests along each column and returns a vector of results. $N$-dimensional arrays x and y must have the same size along all but the first non-singleton dimension, in which case ttest2 works along the first non-singleton dimension.

The test treats NaN values as missing data, and ignores them.
$h=t t e s t 2(x, y, a l p h a)$ performs the test at the (100*alpha)\% significance level. The default, when unspecified, is alpha $=0.05$.
$h=t t e s t 2(x, y, a l p h a, t a i l)$ performs the test against the alternative specified by the string tail. There are three options for tail:

- 'both' - Means are not equal (two-tailed test). This is the default, when tail is unspecified.
- 'right ' - Mean of $x$ is greater than mean of $y$ (right-tail test)
- 'left' - Mean of $x$ is less than mean of $y$ (left-tail test)
tail must be a single string, even when x is a matrix or an $N$-dimensional array.
h = ttest2(x,y,alpha,tail,vartype) performs the test under the assumption of equal or unequal population variances, as specified by the string vartype. There are two options for vartype:
- 'equal' - Assumes equal variances. This is the default, when vartype is unspecified.
- 'unequal' - Does not assume equal variances. This is the Behrens-Fisher problem.
vartype must be a single string, even when x is a matrix or an $N$-dimensional array.

If vartype is 'equal', the test computes a pooled sample standard deviation using

$$
s=\sqrt{\frac{(n-1) s_{x}^{2}+(m-1) s_{y}^{2}}{n+m-2}}
$$

where $s_{x}$ and $s_{y}$ are the sample standard deviations of x and y , respectively, and $n$ and $m$ are the sample sizes of $x$ and $y$, respectively.
h = ttest2(x,y,alpha,tail, vartype,dim) works along dimension dim of $x$ and $y$. Use [ ] to pass in default values for alpha, tail, or vartype.
$[\mathrm{h}, \mathrm{p}]=\mathrm{ttest2}(\ldots)$ returns the $p$ value of the test. The $p$ value is the probability, under the null hypothesis, of observing a value as extreme or more extreme of the test statistic

$$
t=\frac{\bar{x}-\bar{y}}{\sqrt{\frac{s_{x}^{2}}{n}+\frac{s_{y}^{2}}{m}}}
$$

where $\bar{x}$ and $\bar{y}$ are the sample means, $s_{x}$ and $s_{y}$ are the sample standard deviations (replaced by the pooled standard deviation $s$ in the default case where vartype is 'equal'), and $n$ and $m$ are the sample sizes.

In the default case where vartype is 'equal', the test statistic, under the null hypothesis, has Student's $t$ distribution with $n+m-2$ degrees of freedom.

In the case where vartype is 'unequal', the test statistic, under the null hypothesis, has an approximate Student's $t$ distribution with a number of degrees of freedom given by Satterthwaite's approximation.
[h,p,ci] = ttest2(...) returns a 100*(1 - alpha)\% confidence interval on the difference of population means.
[h,p,ci,stats] = ttest2(...) returns structure stats with the following fields:

- tstat - Value of the test statistic
- df - Degrees of freedom of the test
- sd - Pooled sample standard deviation (in the default case where vartype is 'equal') or a vector with the sample standard deviations (in the case where vartype is 'unequal').


## Examples

Simulate random samples of size 1000 from normal distributions with means 0 and 0.1 , respectively, and standard deviations 1 and 2, respectively:

```
x = normrnd(0,1,1,1000);
y = normrnd(0.1,2,1,1000);
```

Test the null hypothesis that the samples come from populations with equal means, against the alternative that the means are unequal. Perform the test assuming unequal variances:

```
[h, p, ci] = ttest2(x,y,[],[],'unequal')
\(\mathrm{h}=\)
        1
\(p=\)
    0.0102
ci \(=\)
    \(-0.3227-0.0435\)
```

The test rejects the null hypothesis at the default $\alpha=0.05$ significance level. Under the null hypothesis, the probability of observing a value as extreme or more extreme of the test statistic, as indicated by the $p$ value, is less than $\alpha$. The $95 \%$ confidence interval on the mean of the difference does not contain 0 .

This example will produce slightly different results each time it is run, because of the random sampling.

## See Also

ttest | ztest

## Purpose Tree type

## Syntax ttype $=$ type $(t)$

Description ttype $=$ type( $t$ ) returns the type of the tree $t$. ttype is 'regression ' for regression trees and 'classification' for classification trees.

Examples Create a classification tree for Fisher's iris data:

```
load fisheriris;
t = classregtree(meas,species,...
    'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
class = versicolor
9 class = virginica
```

view(t)

| Click to display: Identity | $\checkmark$ | Magnification: | 100\% |  | Pruning level: | 0 of 4 | - |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |



```
ttype = type(t)
ttype =
classification
```

References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

## See Also

classregtree
Purpose Type of partition
Description The type of validation partition. It is 'kfold', 'holdout', 'leaveout', or 'resubstitution'.
See Also ..... trainsize

## qrandset.Type property

Purpose Name of sequence on which point set $P$ is based
Description P.Type returns a string that contains the name of the sequence on which the point set $P$ is based, for example 'Sobol'. You cannot change the Type property for a point set.

## Purpose Convert categorical array to unsigned 8-bit integers

Syntax
B = uint8(A)

Description $\quad B=$ uint8 $(A)$ converts the categorical array $A$ to unsigned 8-bit integers. Each element of B contains the internal categorical level code for the corresponding element of A. Undefined elements of A are assigned the value 0 in $B$. If A contains more than intmax ('uint8') levels, the internal codes will saturate to intmax ('uint8')when cast to int8.

See Also double | int8

## categorical.uint 16

Purpose Convert categorical array to unsigned 16-bit integers

## Syntax <br> $B=$ uint16(A)

Description $\quad B=$ uint16(A) converts the categorical array A to unsigned 16-bit integers. Each element of $B$ contains the internal categorical level code for the corresponding element of $A$.

Undefined elements of A are assigned the value 0 in B.
See Also double | int16

# Purpose Convert categorical array to unsigned 32-bit integers 

## Syntax <br> $B=$ uint32(A)

Description $\quad B=$ uint32 $(A)$ converts the categorical array A to unsigned 32-bit integers. Each element of B contains the internal categorical level code for the corresponding element of $A$.

Undefined elements of A are assigned the value 0 in B.
See Also double | int32

Purpose Convert categorical array to unsigned 64-bit integers

## Syntax <br> $B=$ uint64(A)

Description
$B=$ uint64 (A) converts the categorical array $A$ to unsigned 64-bit integers. Each element of B contains the internal categorical level code for the corresponding element of $A$.

Undefined elements of A are assigned the value 0 in B.
See Also
double | int64

Purpose Text label for undefined levels
$\begin{array}{ll}\text { Description } & \begin{array}{l}\text { Text label for undefined levels. Constant property with value } \\ \text { '<undefined> ' . }\end{array}\end{array}$

## dataset.unique

Purpose Unique observations in dataset array
Syntax
$B=$ unique $(A)$
$B=$ unique(A,vars)
[ $\mathrm{B}, \mathrm{i}, \mathrm{j}]=$ unique( A$)$
[...] = unique(A,vars,'first')

## Description

$B=$ unique $(A)$ returns a copy of the dataset $A$ that contains only the sorted unique observations. A must contain only variables whose class has a unique method, including:

- numeric
- character
- logical
- categorical
- cell arrays of strings

For a variable with multiple columns, its class's unique method must support the 'rows ' flag.
$B=$ unique(A, vars) returns a dataset that contains only one observation for each unique combination of values for the variables in A specified in vars. vars is a positive integer, a vector of positive integers, a variable name, a cell array containing one or more variable names, or a logical vector. B includes all variables from A. The values in B for the variables not specified in vars are taken from the last occurrence among observations in A with each unique combination of values for the variables specified in vars.
$[B, i, j]=$ unique $(A)$ also returns index vectors $i$ and $j$ such that $B=$ $A(i,:)$ and $A=B(j,:)$.
[...] = unique(A,vars,'first') returns the vector ito index the first occurrence of each unique observation in $A$. unique (A, vars, 'last'), the default, returns the vector ito index
the last occurrence. Specify vars as [ ] to use the default value of all variables.

See Also<br>dataset | set | subsasgn

## dataset.Units property

Purpose Units of variables in data set
Description A cell array of strings giving the units of the variables in the data set. This property may be empty, but if not empty, the number of strings must equal the number of variables. Any individual string may be empty for a variable that does not have units defined. The default is an empty cell array.

## Purpose

Discrete uniform cumulative distribution function

## Syntax <br> $P=u n i d c d f(X, N)$

$P=$ unidcdf $(X, N)$ computes the discrete uniform cdf at each of the values in $X$ using the corresponding maximum observable value in $N . X$ and $N$ can be vectors, matrices, or multidimensional arrays that have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The maximum observable values in $N$ must be positive integers.

The discrete uniform cdf is

$$
p=F(x \mid N)=\frac{\text { floor }(x)}{N} I_{(1, \ldots, N)}(x)
$$

The result, $p$, is the probability that a single observation from the discrete uniform distribution with maximum $N$ will be a positive integer less than or equal to $x$. The values $x$ do not need to be integers.

Examples What is the probability of drawing a number 20 or less from a hat with the numbers from 1 to 50 inside?

```
probability = unidcdf(20,50)
probability =
    0.4000
```

See Also cdf | unidpdf | unidinv | unidstat | unidrnd | mle
How To . "Uniform Distribution (Discrete)" on page B-101

Purpose Discrete uniform inverse cumulative distribution function
Syntax $\quad X=\operatorname{unidinv}(P, N)$
Description
$X=$ unidinv ( $P, N$ ) returns the smallest positive integer $X$ such that the discrete uniform cdf evaluated at $X$ is equal to or exceeds $P$. You can think of $P$ as the probability of drawing a number as large as $X$ out of a hat with the numbers 1 through $N$ inside.

P and N can be vectors, matrices, or multidimensional arrays that have the same size, which is also the size of $X$. A scalar input for $N$ or $P$ is expanded to a constant array with the same dimensions as the other input. The values in $P$ must lie on the interval [01] and the values in $N$ must be positive integers.

## Examples <br> ```x = unidinv(0.7,20)``` <br> x = <br> 14 <br> $y=u n i d i n v(0.7+e p s, 20)$ <br> $y=$ <br> 15

A small change in the first parameter produces a large jump in output. The cdf and its inverse are both step functions. The example shows what happens at a step.

## See Also icdf | unidcdf | unidpdf | unidstat | unidrnd <br> How To . "Uniform Distribution (Discrete)" on page B-101

## Purpose

Discrete uniform probability density function

## Syntax

Description
$Y=\operatorname{unidpdf}(X, N)$
$Y=u n i d p d f(X, N)$ computes the discrete uniform pdf at each of the values in $X$ using the corresponding maximum observable value in $N . X$ and $N$ can be vectors, matrices, or multidimensional arrays that have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The parameters in $N$ must be positive integers.

The discrete uniform pdf is

$$
y=f(x \mid N)=\frac{1}{N} I_{(1, \ldots, N)}(x)
$$

You can think of $y$ as the probability of observing any one number between 1 and $n$.

## Examples For fixed n , the uniform discrete pdf is a constant.

```
y = unidpdf(1:6,10)
y =
    0.1000}00.1000 0.1000 0.1000 0.1000 0.1000 
```

Now fix x , and vary n .

```
likelihood = unidpdf(5,4:9)
likelihood =
\begin{tabular}{llllll}
0 & 0.2000 & 0.1667 & 0.1429 & 0.1250 & 0.1111
\end{tabular}
```

See Also
pdf | unidcdf | unidinv | unidstat | unidrnd
How To . "Uniform Distribution (Discrete)" on page B-101

## unidrnd

Purpose Discrete uniform random numbers

Syntax $\quad$| $R$ | $=\operatorname{unidrnd}(N)$ |
| ---: | :--- |
| $R$ | $=\operatorname{unidrnd}(N, m, n, \ldots)$ |
| $R$ | $=\operatorname{unidrnd}(N,[m, n, \ldots])$ |

Description

## Examples In the Massachusetts lottery, a player chooses a four-digit number. Generate random numbers for Monday through Saturday.

See Also random | unidpdf | unidcdf | unidinv | unidstat
How To . "Uniform Distribution (Discrete)" on page B-101
$R=$ unidrnd( $N$ ) generates random numbers for the discrete uniform distribution with maximum $N$. The parameters in $N$ must be positive integers. $N$ can be a vector, a matrix, or a multidimensional array. The size of $R$ is the size of $N$. The discrete uniform distribution arises from experiments equivalent to drawing a number from one to $N$ out of a hat.
$R=$ unidrnd( $N, m, n, \ldots$ ) or $R=$ unidrnd( $N,[m, n, \ldots])$ generates an m-by-n-by-... array. The N parameter can be a scalar or an array of the same size as R . Ger

```
numbers = unidrnd(10000,1,6)-1
```

numbers = unidrnd(10000,1,6)-1
numbers =
numbers =
4564}10185 8214 4447 6154 7919

```
    4564}10185 8214 4447 6154 7919 
```


## Purpose

Discrete uniform mean and variance

## Syntax

Description
$[M, V]=$ unidstat $(N)$ returns the mean and variance of the discrete uniform distribution with minimum value 1 and maximum value $N$.

The mean of the discrete uniform distribution with parameter $N$ is $(N+1) / 2$. The variance is $\left(N^{2}-1\right) / 12$.

```
[m,v] = unidstat(1:6)
m =
    1.0000 1.5000 2.0000 2.5000 3.0000 3.5000
v =
    0
```

See Also

unidpdf | unidcdf | unidinv | unidrnd
How To . "Uniform Distribution (Discrete)" on page B-101

## unifcdf

Purpose Continuous uniform cumulative distribution function

$$
\text { Syntax } \quad P=\operatorname{unifcdf}(X, A, B)
$$

Description

## Examples

## See Also

How To
$P=$ unifcdf $(X, A, B)$ computes the uniform cdf at each of the values in X using the corresponding lower endpoint (minimum), A and upper endpoint (maximum), B. X, A, and B can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant matrix with the same dimensions as the other inputs.

The uniform cdf is

$$
p=F(x \mid a, b)=\frac{x-a}{b-a} I_{[a, b]}(x)
$$

The standard uniform distribution has $\mathrm{A}=0$ and $\mathrm{B}=1$.

What is the probability that an observation from a standard uniform distribution will be less than 0.75 ?

```
probability = unifcdf(0.75)
probability =
    0.7500
```

What is the probability that an observation from a uniform distribution with $\mathrm{a}=-1$ and $\mathrm{b}=1$ will be less than 0.75 ?

```
probability = unifcdf(0.75,-1,1)
probability =
    0.8750
```

cdf | unifpdf | unifinv | unifstat | unifit | unifrnd

- "Uniform Distribution (Continuous)" on page B-99


## Purpose

Continuous uniform inverse cumulative distribution function

## Syntax

X = unifinv( $\mathrm{P}, \mathrm{A}, \mathrm{B}$ )
$X=$ unifinv $(P, A, B)$ computes the inverse of the uniform cdf with parameters $A$ and $B$ (the minimum and maximum values, respectively) at the corresponding probabilities in P. P, A, and B can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs.

The inverse of the uniform cdf is

$$
x=F^{-1}(p \mid a, b)=a+p(a-b) I_{[0,1]}(p)
$$

The standard uniform distribution has $\mathrm{A}=0$ and $\mathrm{B}=1$.
Examples
What is the median of the standard uniform distribution?

```
median_value = unifinv(0.5)
median_value =
    0.5000
```

What is the 99 th percentile of the uniform distribution between -1 and 1 ?

```
percentile = unifinv(0.99,-1,1)
percentile =
    0.9800
```

See Also icdf | unifcdf | unifpdf | unifstat | unifit | unifrnd

How To . "Uniform Distribution (Continuous)" on page B-99

| Purpose | Continuous uniform parameter estimates |
| :--- | :--- |
| Synfax | $[$ ahat, bhat $]=$ unifit(data) |
|  | $[$ ahat, bhat, ACI, BCI $]=$ unifit(data) |
|  | $[$ ahat, bhat, ACI, BCI $]=$ unifit(data, alpha) |

## Description

## Examples

See Also mle | unifpdf | unifcdf | unifinv | unifstat | unifrnd
How To . "Uniform Distribution (Continuous)" on page B-99

## Purpose

Continuous uniform probability density function

## Syntax

Description
$Y=\operatorname{unifpdf}(X, A, B)$
$Y=$ unifpdf( $X, A, B$ ) computes the continuous uniform pdf at each of the values in $X$ using the corresponding lower endpoint (minimum), $A$ and upper endpoint (maximum), B. $X, A$, and $B$ can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array with the same dimensions as the other inputs. The parameters in B must be greater than those in A.

The continuous uniform distribution pdf is

$$
y=f(x \mid a, b)=\frac{1}{b-a} I_{[a, b]}(x)
$$

The standard uniform distribution has $\mathrm{A}=0$ and $\mathrm{B}=1$.
For fixed a and b, the uniform pdf is constant.

```
x = 0.1:0.1:0.6;
y = unifpdf(x)
y =
    1
```

What if x is not between a and b ?

$$
\begin{aligned}
& y=\operatorname{unifpdf}(-1,0,1) \\
& y=
\end{aligned}
$$

0

## See Also <br> pdf | unifcdf | unifinv | unifstat | unifit | unifrnd

How To . "Uniform Distribution (Continuous)" on page B-99

## Purpose Continuous uniform random numbers

Syntax $\quad$| $R$ | $=\operatorname{unifrnd}(A, B)$ |
| ---: | :--- |
| $R$ | $=\operatorname{unifrnd}(A, B, m, n, \ldots)$ |
| $R$ | $=\operatorname{uniffnd}(A, B,[m, n, \ldots])$ |

## Description

## Examples

$R=$ unifrnd $(A, B)$ returns an array $R$ of random numbers generated from the continuous uniform distributions with lower and upper endpoints specified by A and B, respectively. If A and B are arrays, $R(i, j)$ is generated from the distribution specified by the corresponding elements of $A$ and $B$. If either $A$ or $B$ is a scalar, it is expanded to the size of the other input.
$R=$ unifrnd( $A, B, m, n, \ldots)$ or $R=u n i f r n d(A, B,[m, n, \ldots])$ returns an m-by-n-by-... array. If $A$ and $B$ are scalars, all elements of $R$ are generated from the same distribution. If either $A$ or $B$ is an array, they must be m-by-n-by-... .

Generate one random number each from the continuous uniform distributions on the intervals $(0,1),(0,2), \ldots,(0,5)$ :

```
a = 0; b = 1:5;
r1 = unifrnd(a,b)
r1 =
    0.8147 1.8116 0.3810 3.6535 3.1618
```

Generate five random numbers each from the same distributions:

```
B = repmat(b,5,1);
R = unifrnd(a,B)
R =
\begin{tabular}{lllll}
0.0975 & 0.3152 & 0.4257 & 2.6230 & 3.7887 \\
0.2785 & 1.9412 & 1.2653 & 0.1428 & 3.7157 \\
0.5469 & 1.9143 & 2.7472 & 3.3965 & 1.9611 \\
0.9575 & 0.9708 & 2.3766 & 3.7360 & 3.2774 \\
0.9649 & 1.6006 & 2.8785 & 2.7149 & 0.8559
\end{tabular}
```

Generate five random numbers from the continuous uniform distribution on $(0,2)$ :

```
r2 = unifrnd(a,b(2),1,5)
r2 =
    1.4121 0.0637 0.5538 0.0923 0.1943
```

See Also
rand | random | unifpdf | unifcdf | unifinv | unifstat | unifit
How To . "Uniform Distribution (Continuous)" on page B-99

Purpose Continuous uniform mean and variance

## Syntax <br> [ $\mathrm{M}, \mathrm{V}$ ] = unifstat (A, B)

$[M, V]=$ unifstat $(A, B)$ returns the mean of and variance for the continuous uniform distribution using the corresponding lower endpoint (minimum), A and upper endpoint (maximum), B. Vector or matrix inputs for $A$ and $B$ must have the same size, which is also the size of $M$ and V. A scalar input for A or B is expanded to a constant matrix with the same dimensions as the other input.

The mean of the continuous uniform distribution with parameters $a$ and $b$ is $(a+b) / 2$, and the variance is $(a-b)^{2} / 12$.

## Examples

See Also
unifpdf | unifcdf | unifinv | unifit | unifrnd
How To

- "Uniform Distribution (Continuous)" on page B-99

```
a = 1:6;
b = 2.*a;
[m,v] = unifstat(a,b)
m =
    1.5000}3.0000 4.5000 6.0000 7.5000 9.0000
v =
    0.0833 0.3333 0.7500 1.3333 2.0833 3.0000
```

nifpdf | unifcdf | unifinv | unifit | unifrnd

## Purpose Set union for categorical arrays

Syntax
$C=$ union $(A, B)$
[C,IA,IB] = union(A,B)

Description
$C=$ union $(A, B)$ when $A$ and $B$ are categorical arrays returns a categorical vector $C$ containing the combined values from $A$ and $B$ but with no repetitions. The result C is sorted. The set of categorical levels for C is the sorted union of the sets of levels of the inputs, as determined by their labels.
[C,IA, IB] = union(A,B) also returns index vectors IA and IB such that $C$ is a sorted combination of the elements $A(I A)$ and $B(I B)$.

See Also<br>intersect | ismember | setdiff | setxor | unique

Purpose Unique values in categorical array

| Syntax | $B=\operatorname{unique}(A)$ |
| :--- | :--- |
|  | $[B, I, J]=\operatorname{unique}(A)$ |
|  | $[B, I, J]=\operatorname{unique}\left(A\right.$, ' first' $\left.^{\prime}\right)$ |

## Description

$B=$ unique ( $A$ ) returns a categorical array containing the unique elements of A, sorted by the order of A's levels.
$[B, I, J]=$ unique $(A)$ also returns index vectors $I$ and $J$ such that $B=A(I)$ and $A=B(J)$.
$[B, I, J]=$ unique( $A$, 'first') returns the vector $I$ to index the first occurrence of each unique value in $A$. unique ( $A$, 'last'), the default, returns the vector I to index the last occurrence.

## See Also

intersect | ismember | setdiff | setxor | union

| Purpose | Unstack data from single variable into multiple variables |
| :--- | :--- |
| Syntax | wide $=$ unstack(tall, datavar, indvar) <br> [wide, itall] = unstack(tall, datavar, indvar) <br> wide $=$ unstack(tall, datavar, indvar, 'Parameter', value) |
|  | wide = unstack(tall, datavar, indvar) converts a dataset array |
| tall to an equivalent dataset array wide that is in wide format, by |  |
| unstacking a single variable in tall into multiple variables in wide. In |  |
| general wide contains more variables, but fewer observations, than |  |
| tall. |  |
| datavar specifies the data variable in tall to unstack. indvar specifies |  |
| an indicator variable in tall that determines which variable in wide |  |
| each value in datavar is unstacked into. unstack treats the remaining |  |
| variables in tall as grouping variables. Each unique combination |  |
| of their values defines a group of observations in tall that will be |  |
| unstacked into a single observation in wide. |  |

Input Arguments
wide = unstack(tall,datavar,indvar,'Parameter', value) uses the following parameter name/value pairs to control how unstack converts variables in tall to variables in wide:

| 'GroupVars ' | Grouping variables in tall that define groups <br> of observations. groupvars is a positive <br> integer, a vector of positive integers, a variable <br> name, a cell array containing one or more <br> variable names, or a logical vector. The default <br> is all variables in tall not listed in datavar <br> or indvar. |
| :--- | :--- |
| 'NewDataVarNames' | A cell array of strings containing names for the <br> data variables unstack should create in wide. <br> Default is the group names of the grouping <br> variable specified in indvar. |
| 'AggregationFun' | A function handle that accepts a subset of <br> values from datavar and returns a single value. <br> stack applies this function to observations from <br> the same group that have the same value of <br> indvar. The function must aggregate the data <br> values into a single value, and in such cases <br> it is not possible to recover tall from wide <br> using stack. The default is @sum for numeric <br> data variables. For non-numeric variables, <br> there is no default, and you must specify <br> 'AggregationFun' if multiple observations <br> in the same group have the same values of <br> indvar. |
| 'ConstVars' | Variables in tall to copy to wide without <br> unstacking. The values for these variables <br> in wide are taken from the first observation <br> in each group in tall, so these variables <br> should typically be constant within each group. <br> ConstVars is a positive integer, a vector of <br> positive integers, a variable name, a cell array |


|  | containing one or more variable names, or a <br> logical vector. The default is no variables. |
| :--- | :--- |

You can also specify more than one data variable in tall, each of which becomes a set of $m$ variables in wide. In this case, specify datavar as a vector of positive integers, a cell array containing variable names, or a logical vector. You may specify only one variable with indvar. The names of each set of data variables in wide are the name of the corresponding data variable in tall concatenated with the names specified in 'NewDataVarNames'. The function specified in 'AggregationFun' must return a value with a single row.

## Examples

Convert a "wide format" data set to "tall format", and then back to a different "wide format":
load flu

```
% FLU has a 'Date' variable, and 10 variables for estimated
% influenza rates (in 9 different regions, estimated from
% Google searches, plus a nationwide extimate from the
% CDC). Combine those 10 variables into a "tall" array that
% has a single data variable, 'FluRate', and an indicator
% variable, 'Region', that says which region each estimate
% is from.
[flu2,iflu] = stack(flu, 2:11, 'NewDataVarName','FluRate', ...
    'IndVarName','Region')
% The second observation in FLU is for 10/16/2005. Find the
% observations in FLU2 that correspond to that date.
flu(2,:)
flu2(iflu==2,:)
% Use the 'Date' variable from that tall array to split
% 'FluRate' into 52 separate variables, each containing the
% estimated influenza rates for each unique date. The new
% "wide" array has one observation for each region. In
% effect, this is the original array FLU "on its side".
dateNames = cellstr(datestr(flu.Date,'mmm_DD_YYYY'));
```

```
[flu3,iflu2] = unstack(flu2, 'FluRate', 'Date', ...
            'NewDataVarNames ',dateNames)
% Since observations in FLU3 represent regions, IFLU2
% indicates the first occurrence in FLU2 of each region.
flu2(iflu2,:)
```


## See Also

dataset.stack| dataset.join | dataset.grpstats

How To<br>- "Grouping Variables" on page 2-34

Purpose Upper Pareto tails parameters
Syntax params = upperparams(obj)
Description params = upperparams(obj) returns the 2-element vector params ofshape and scale parameters, respectively, of the upper tail of the Paretotails object obj. upperparams does not return a location parameter.
Examples Fit Pareto tails to a $t$ distribution at cumulative probabilities 0.1 and 0.9:

```
t = trnd(3,100,1);
obj = paretotails(t,0.1,0.9);
    lowerparams(obj)
    ans =
        -0.1901 1.1898
        upperparams(obj)
        ans =
        0.3646 0.5103
```

See Also
paretotails | lowerparams

## dataset.UserData property

Purpose Variable containing additional information associated with data set
Description Any variable containing additional information to be associated with the data set. The default is an empty array.

## ProbDistUnivParam.var

| Purpose | Return variance of ProbDistUnivParam object |
| :---: | :---: |
| Syntax | $V=\operatorname{var}(P D)$ |
| Description | $V=\operatorname{var}(P D)$ returns $V$, the variance of the ProbDistUnivParam object $P D$. |
| Input Arguments | $P D \quad$ An object of the class ProbDistUnivParam. |
| Output Arguments | The variance of the ProbDistUnivParam object $P D$. |
| See Also | var |

## dataset.VarDescription property

Purpose Cell array of strings giving descriptions of variables in data set
Description
A cell array of strings giving the descriptions of the variables in the data set. This property may be empty, but if not empty, the number of strings must equal the number of variables. Any individual string may be empty for a variable that does not have a description defined. The default is an empty cell array.

| Purpose | Compute embedded estimates of input feature importance |
| :--- | :--- |
| Syntax | imp = varimportance ( $t$ ) |
| Description | imp = varimportance $(t)$ computes estimates of input feature <br> importance for tree $t$ by summing changes in the risk due to splits on <br> every feature. The returned vector imp has one element for each input <br> variable in the data used to train this tree. At each node, the risk is <br> estimated as node impurity if impurity was used to split nodes and <br> node error otherwise. This risk is weighted by the node probability. |
| Variable importance associated with this split is computed as the <br> difference between the risk for the parent node and the total risk for <br> the two children. |  |

See Also ..... risk

## dataset.VarNames property

Purpose Cell array giving names of variables in data set
Description
A cell array of nonempty, distinct strings giving the names of the variables in the data set. The number of strings must equal the number of variables. The default is the cell array of string names for the variables used to create the data set.

Purpose Variable names
Description The VarNames property is a cell array containing the names of the predictor variables (features). TreeBagger takes these names from the optional 'names' parameter. The default names are 'x1', 'x2', etc.

## Purpose Chi-square variance test

```
Syntax \(\quad H=\operatorname{vartest}(X, V)\)
H = vartest(X,V,alpha)
H = vartest(X,V,alpha,tail)
[H,P] = vartest(...)
[ \(\mathrm{H}, \mathrm{P}, \mathrm{CI}\) ] = vartest(...)
[H,P,CI,STATS] = vartest(...)
[...] = vartest(X,V,alpha,tail,dim)
```


## Description

$\mathrm{H}=$ vartest $(\mathrm{X}, \mathrm{V})$ performs a chi-square test of the hypothesis that the data in the vector $X$ comes from a normal distribution with variance $V$, against the alternative that $X$ comes from a normal distribution with a different variance. The result is $\mathrm{H}=0$ if the null hypothesis (variance is V ) cannot be rejected at the $5 \%$ significance level, or $\mathrm{H}=1$ if the null hypothesis can be rejected at the $5 \%$ level.
$X$ may also be a matrix or an n-dimensional array. For matrices, vartest performs separate tests along each column of $X$, and returns a row vector of results. For n-dimensional arrays, vartest works along the first nonsingleton dimension of X . V must be a scalar.
$\mathrm{H}=$ vartest( $\mathrm{X}, \mathrm{V}$, alpha) performs the test at the significance level (100*alpha) \%. alpha has a default value of 0.05 and must be a scalar.

H = vartest(X,V,alpha, tail) performs the test against the alternative hypothesis specified by tail, where tail is a single string from the following choices:

- 'both' - Variance is not V (two-tailed test). This is the default.
- 'right' - Variance is greater than V (right-tailed test).
- 'left' - Variance is less than V (left-tailed test).
$[\mathrm{H}, \mathrm{P}]=$ vartest (...) returns the $p$ value, i.e., the probability of observing the given result, or one more extreme, by chance if the null hypothesis is true. Small values of P cast doubt on the validity of the null hypothesis.
[ $\mathrm{H}, \mathrm{P}, \mathrm{CI}]=$ vartest(...) returns a 100*(1-alpha) \% confidence interval for the true variance.
[H,P,CI,STATS] = vartest(...) returns the structure STATS with the following fields:
- 'chisqstat' - Value of the test statistic
- 'df' - Degrees of freedom of the test
[...] = vartest(X,V,alpha,tail,dim) works along dimension dim of $X$. Pass in [] for alpha or tail to use their default values.


## Examples

Determine whether the standard deviation is significantly different from 7?

```
load carsmall
```

[h,p,ci] = vartest(MPG,7^2)
See Also
ttest | ztest | vartest2

Purpose
Two-sample $F$-test for equal variances
Syntax
H = vartest2 (X,Y)
H = vartest2(X,Y,alpha)
H = vartest2(X,Y,alpha,tail)
[ $\mathrm{H}, \mathrm{P}$ ] = vartest2(...)
[ $\mathrm{H}, \mathrm{P}, \mathrm{CI}$ ] $=$ vartest2(...)
[H,P,CI,STATS] = vartest2(...)
[...] = vartest2(X,Y,alpha,tail,dim)

## Description

H = vartest2(X,Y) performs an $F$ test of the hypothesis that two
independent samples, in the vectors $X$ and $Y$, come from normal distributions with the same variance, against the alternative that they come from normal distributions with different variances. The result is H $=0$ if the null hypothesis (variances are equal) cannot be rejected at the $5 \%$ significance level, or $\mathrm{H}=1$ if the null hypothesis can be rejected at the $5 \%$ level. $X$ and $Y$ can have different lengths. $X$ and $Y$ can also be matrices or n-dimensional arrays.

For matrices, vartest2 performs separate tests along each column, and returns a vector of results. $X$ and $Y$ must have the same number of columns. For n-dimensional arrays, vartest2 works along the first nonsingleton dimension. $X$ and $Y$ must have the same size along all the remaining dimensions.

H = vartest2( $\mathrm{X}, \mathrm{Y}$, alpha) performs the test at the significance level (100*alpha) \%. alpha must be a scalar.

H = vartest2(X,Y,alpha,tail) performs the test against the alternative hypothesis specified by tail, where tail is one of the following single strings:

- 'both' - Variance is not Y (two-tailed test). This is the default.
- 'right ' - Variance is greater than $Y$ (right-tailed test).
- 'left' - Variance is less than $Y$ (left-tailed test).
[ $\mathrm{H}, \mathrm{P}]=$ vartest2(...) returns the $p$ value, i.e., the probability of observing the given result, or one more extreme, by chance if the null hypothesis is true. Small values of P cast doubt on the validity of the null hypothesis.
[H,P,CI] = vartest2(...) returns a 100*(1-alpha)\% confidence interval for the true variance ratio $\operatorname{var}(X) / \operatorname{var}(Y)$.
[H,P,CI,STATS] = vartest2(...) returns a structure with the following fields:
- 'fstat ${ }^{\prime}$ - Value of the test statistic
- 'df1' - Numerator degrees of freedom of the test
- 'df2' - Denominator degrees of freedom of the test
[...] = vartest2(X,Y,alpha,tail,dim) works along dimension dim of $X$. To pass in the default values for alpha or tail use [ ].


## Examples

See Also

Is the variance significantly different for two model years, and what is a confidence interval for the ratio of these variances?

## load carsmall

```
[H,P,CI] =
vartest2(MPG(Model_Year==82),MPG(Model_Year==76))
```

```
ansaribradley | vartest | vartestn | ttest2
```

Purpose
Bartlett multiple-sample test for equal variances

## Syntax

vartestn (X)
vartestn(X, group)
p = vartestn(...)
[p,STATS] = vartestn(...)
[...] = vartestn(...,displayopt)
[...] = vartestn(....,testtype)

## Description

vartestn (X) performs Bartlett's test for equal variances for the columns of the matrix $X$. This is a test of the null hypothesis that the columns of $X$ come from normal distributions with the same variance, against the alternative that they come from normal distributions with different variances. The result is a display of a box plot of the groups, and a summary table of statistics.
vartestn ( $X$, group) requires a vector $X$, and a group argument that is a categorical variable, vector, string array, or cell array of strings with one row for each element of $X$. The $X$ values corresponding to the same value of group are placed in the same group. (See "Grouped Data" on page 2-34.) The function tests for equal variances across groups.
vartestn treats NaNs as missing values and ignores them.
$\mathrm{p}=$ vartestn (...) returns the $p$ value, i.e., the probability of observing the given result, or one more extreme, by chance if the null hypothesis of equal variances is true. Small values of $p$ cast doubt on the validity of the null hypothesis.
[p,STATS] = vartestn(...) returns a structure with the following fields:

- 'chistat ' - Value of the test statistic
- 'df' - Degrees of freedom of the test
[...] = vartestn(...,displayopt) determines if a box plot and table are displayed. displayopt may be 'on' (the default) or 'off'.
[...] = vartestn(...,testtype) sets the test type. When testtype is 'robust', vartestn performs Levene's test in place of Bartlett's test, which is a useful alternative when the sample distributions are not normal, and especially when they are prone to outliers. For this test the STATS output structure has a field named 'fstat ' containing the test statistic, and 'df1' and 'df2' containing its numerator and denominator degrees of freedom. When testtype is 'classical' vartestn performs Bartlett's test.


## Examples

Does the variance of mileage measurements differ significantly from one model year to another?

```
load carsmall
p = vartestn(MPG,Model_Year)
p =
    0.8327
```


## Group Summary Table

| Group | Count | Mean | Std Dev | - |
| :---: | :---: | :---: | :---: | :---: |
| 70 | 29 | 17.6897 | 5.33923 |  |
| 76 | 34 | 21.5735 | 5.8893 |  |
| 82 | 31 | 31.7097 | 5.39255 |  |
| Pooled | 94 | 23.7181 | 5.562 |  |
| Bartlett's statistic | 0.36619 |  |  |  |
| Degrees of freedom | 2 |  |  |  |
| p-value | 0.83269 |  |  | $\checkmark$ |



See Also vartest | vartest2 | anova1
How To . "Grouped Data" on page 2-34

## Purpose Vertical concatenation for categorical arrays

Syntax
$C=\operatorname{vertcat}(\operatorname{dim}, A, B, \ldots)$
C = vertcat (A,B)

Description $\quad C=$ vertcat (dim, $A, B, \ldots$ ) vertically concatenates the categorical arrays $A, B, \ldots$. For matrices, all inputs must have the same number of rows. For n-D arrays, all inputs must have the same sizes except in the second dimension. The set of categorical levels for C is the sorted union of the sets of levels of the inputs, as determined by their labels. $C=\operatorname{vertcat}(A, B)$ is called for the syntax $[A B]$.

See Also<br>cat | horzcat

## dataset.vertcat

Purpose Vertical concatenation for dataset arrays

```
Syntax
ds = vertcat(ds1, ds2, ...)
```

Description ds = vertcat(ds1, ds2, ...) vertically concatenates the dataset arrays ds1, ds2, ... . Observation names, when present, must be unique across datasets. vertcat fills in default observation names for the output when some of the inputs have names and some do not.

Variable names for all dataset arrays must be identical except for order. vertcat concatenates by matching variable names. vertcat assigns values for the "per-variable" properties (e.g., Units and VarDescription) in ds from the corresponding property values in ds1.

## See Also <br> cat | horzcat

## Purpose <br> Plot tree

## Syntax

view(t)
view(t,param1,val1,param2,val2,...)
view( $t$ ) displays the decision tree $t$ as computed by classregtree in a figure window. Each branch in the tree is labeled with its decision rule, and each terminal node is labeled with the predicted value for that node. Click any node to get more information about it. The information displayed is specified by the Click to display pop-up menu at the top of the figure.
view(t,param1,val1,param2,val2,...) specifies optional parameter name/value pairs:

- 'names ' - A cell array of names for the predictor variables, in the order in which they appear in the matrix $X$ from which the tree was created. (See classregtree.)
- 'prunelevel' - Initial pruning level to display.

For each branch node, the left child node corresponds to the points that satisfy the condition, and the right child node corresponds to the points that do not satisfy the condition.

## Examples <br> Create a classification tree for Fisher's iris data:

```
load fisheriris;
t = classregtree(meas,species,...
    'names',{'SL' 'SW' 'PL' 'PW'})
t =
Decision tree for classification
1 if PL<2.45 then node 2 elseif PL>=2.45 then node 3 else setosa
2 class = setosa
3 if PW<1.75 then node 4 elseif PW>=1.75 then node 5 else versicolor
4 if PL<4.95 then node 6 elseif PL>=4.95 then node 7 else versicolor
5 class = virginica
```


## classregtree.view

```
6 if PW<1.65 then node 8 elseif PW>=1.65 then node 9 else versicolor
7 class = virginica
8 class = versicolor
9 class = virginica
view(t)
```

| Click to display: | Identity | $\square$ | Magnification: $100 \%$ | $\rightarrow$ | Pruning level: 0 of $\mathbf{4}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |



References [1] Breiman, L., J. Friedman, R. Olshen, and C. Stone. Classification and Regression Trees. Boca Raton, FL: CRC Press, 1984.

See Also classregtree | eval | prune | test

## CompactClassificationTree.view

| Purpose | View tree |
| :---: | :---: |
| Syntax | $\begin{aligned} & \text { view(tree) } \\ & \text { view(tree, Name, Value) } \end{aligned}$ |
| Description | view(tree) returns a text description of tree, a decision tree. <br> view(tree, Name, Value) describes tree with additional options specified by one or more Name, Value pair arguments. |
| Input Arguments | tree <br> A classification tree or compact classification tree created by ClassificationTree.fit or compact. |
|  | Name-Value Pair Arguments |
|  | Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, ,NameN, ValueN. <br> mode |
|  | String describing the display of tree, either 'graph ' or 'text'. 'graph ' opens a GUI displaying tree, and containing controls for querying the tree. 'text' sends output to the Command Window describing tree. |
|  | Default: 'text' |
| Examples | View the classification tree for Fisher's iris model in both textual and graphical displays: |
|  | load fisheriris tree $=$ ClassificationTree.fit(meas,species); |
|  | view(tree) |
|  | Decision tree for classification |

## CompactClassificationTree.view

```
1 if x3<2.45 then node 2 elseif x3>=2.45 then node 3 else setosa
class = setosa
if x4<1.75 then node 4 elseif x4>=1.75 then node 5 else versicolor
if x3<4.95 then node 6 elseif x3>=4.95 then node 7 else versicolor
class = virginica
if x4<1.65 then node 8 elseif x4>=1.65 then node 9 else versicolor
class = virginica
class = versicolor
class = virginica
view(tree,'mode','graph')
```

| - Classification tree viewer |  |  |  | 回 | $x$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| File Tools Desktop | Tree | Window | Help |  | v |
|  |  |  |  |  |  |
| Click to display: Identity |  | $\checkmark$ | Magnification: | 100\% | $\checkmark$ |



## CompactClassificationTree.view

See Also ClassificationTree<br>How To • Chapter 13, "Nonparametric Supervised Learning"

## CompactRegressionTree.view

| Purpose | View tree |
| :---: | :---: |
| Syntax | view(tree) <br> view(tree, Name, Value) |
| Description | view(tree) returns a text description of tree, a decision tree. <br> view(tree, Name, Value) describes tree with additional options specified by one or more Name, Value pair arguments. |
| Input Arguments | tree <br> A regression tree or compact regression tree created by RegressionTree.fit or compact. |
|  | Name-Value Pair Arguments |
|  | Optional comma-separated pairs of Name, Value arguments, where Name is the argument name and Value is the corresponding value. Name must appear inside single quotes (' ' ). You can specify several name-value pair arguments in any order as Name1, Value1, , NameN, ValueN. <br> mode |
|  | String describing the display of tree, either 'graph' or 'text'. 'graph ' opens a GUI displaying tree, and containing controls for querying the tree. 'text' sends output to the Command Window describing tree. |
|  | Default: 'text' |
| Examples | View a regression tree for the carsmall data in both textual and graphical displays: |

Examples View a regression tree for the carsmall data in both textual and graphical displays:

```
load carsmall
tree = RegressionTree.fit([Weight, Cylinders],MPG,...
    categoricalpredictors',2,'MinParent',20,...
    PredictorNames',{'W','C'});
view(tree)
```


## CompactRegressionTree.view

Decision tree for regression
1 if $W<3085.5$ then node 2 elseif $W>=3085.5$ then node 3 else 23.7181
2 if $W<2371$ then node 4 elseif $W>=2371$ then node 5 else 28.7931
3 if $C=8$ then node 6 elseif $C$ in $\{46\}$ then node 7 else 15.5417
4 if $W<2162$ then node 8 elseif $W>=2162$ then node 9 else 32.0741
5 if $C=6$ then node 10 elseif $C=4$ then node 11 else 25.9355
6 if $W<4381$ then node 12 elseif $W>=4381$ then node 13 else 14.2963
7 fit $=19.2778$
8 fit $=33.3056$
9 fit $=29.6111$
10 fit $=23.25$
11 if $W<2827.5$ then node 14 elseif $W>=2827.5$ then node 15 else 27.2143
12 if $W<3533.5$ then node 16 elseif $W>=3533.5$ then node 17 else 14.8696
13 fit $=11$
14 fit $=27.6389$
15 fit $=24.6667$
16 fit $=16.6$
17 fit $=14.3889$


See Also

RegressionTree

## CompactRegressionTree.view

How To<br>- Chapter 13, "Nonparametric Supervised Learning"

## Purpose Weibull cumulative distribution function

```
Syntax P = wblcdf(X,A,B)
[P,PLO,PUP] = wblcdf(X,A,B,PCOV,alpha)
```


## Description $\quad P=\operatorname{wblcdf}(X, A, B)$ computes the $c d f$ of the Weibull distribution with

 scale parameter $A$ and shape parameter $B$, at each of the values in $X$. $X$, $A$, and $B$ can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array of the same size as the other inputs. The default values for A and B are both 1. The parameters A and B must be positive.[P,PLO, PUP] = wblcdf(X,A,B,PCOV,alpha) returns confidence bounds for $P$ when the input parameters $A$ and $B$ are estimates. PCOV is the 2 -by- 2 covariance matrix of the estimated parameters. alpha has a default value of 0.05 , and specifies $100(1-\mathrm{alpha}) \%$ confidence bounds. PLO and PUP are arrays of the same size as P containing the lower and upper confidence bounds.

The function wblcdf computes confidence bounds for P using a normal approximation to the distribution of the estimate

$$
\hat{b}(\log x-\log \hat{a})
$$

and then transforms those bounds to the scale of the output P. The computed bounds give approximately the desired confidence level when you estimate mu, sigma, and PCOV from large samples, but in smaller samples other methods of computing the confidence bounds might be more accurate.

The Weibull cdf is

$$
p=F(x \mid a, b)=\int_{0}^{x} b a^{-b} t^{b-1} e^{-\left(\frac{t}{a}\right)^{b}} d t=1-e^{-\left(\frac{x}{a}\right)^{b}} I_{(0, \infty)}(x)
$$

## Examples

What is the probability that a value from a Weibull distribution with parameters $\mathrm{a}=0.15$ and $\mathrm{b}=0.8$ is less than 0.5 ?

```
probability = wblcdf(0.5, 0.15, 0.8)
probability =
    0.9272
```

How sensitive is this result to small changes in the parameters?

```
[A, B] = meshgrid(0.1:0.05:0.2,0.2:0.05:0.3);
probability = wblcdf(0.5, A, B)
probability =
    0.7484 0.7198
    0.7758}00.7411 0.715
    0.8022 0.7619 0.7319
```

See Also
cdf | wblpdf | wblinv | wblstat | wblfit | wbllike | wblrnd
How To . "Weibull Distribution" on page B-103

Purpose Weibull parameter estimates

```
Syntax parmhat = wblfit(data)
[parmhat,parmci] = wblfit(data)
[parmhat,parmci] = wblfit(data,alpha)
[...] = wblfit(data,alpha,censoring)
[...] = wblfit(data,alpha,censoring,freq)
[...] = wblfit(...,options)
```


## Description

parmhat = wblfit(data) returns the maximum likelihood estimates, parmhat, of the parameters of the Weibull distribution given the values in the vector data, which must be positive. parmhat is a two-element row vector: parmhat(1) estimates the Weibull parameter $a$, and parmhat (2) estimates the Weibull parameter $b$, in the pdf

$$
y=f(x \mid a, b)=b a^{-b} x^{b-1} e^{-\left(\frac{x}{a}\right)^{b}} I_{(0, \infty)}(x)
$$

[parmhat, parmci] = wblfit(data) returns $95 \%$ confidence intervals for the estimates of $a$ and $b$ in the 2 -by- 2 matrix parmci. The first row contains the lower bounds of the confidence intervals for the parameters, and the second row contains the upper bounds of the confidence intervals.
[[parmhat,parmci] = wblfit(data,alpha) returns 100(1-alpha)\% confidence intervals for the parameter estimates.
[...] = wblfit(data, alpha, censoring) accepts a Boolean vector, censoring, of the same size as data, which is 1 for observations that are right-censored and 0 for observations that are observed exactly.
[...] = wblfit(data, alpha, censoring,freq) accepts a frequency vector, freq, of the same size as data. The vector freq typically contains integer frequencies for the corresponding elements in data, but can contain any non-negative values. Pass in [] for alpha, censoring, or freq to use their default values.
[...] = wblfit(...,options) accepts a structure, options, that specifies control parameters for the iterative algorithm the function uses to compute maximum likelihood estimates. The Weibull fit function accepts an options structure that can be created using the function statset. Enter statset ('wblfit') to see the names and default values of the parameters that lognfit accepts in the options structure. See the reference page for statset for more information about these options.

## Examples

```
data = wblrnd(0.5,0.8,100,1);
[parmhat, parmci] = wblfit(data)
parmhat =
    0.5861 0.8567
parmci =
    0.4606 0.7360
    0.7459 0.9973
```

mle | wbllike | wblpdf | wblcdf | wblinv | wblstat | wblrnd
How To . "Weibull Distribution" on page B-103

## Purpose Weibull inverse cumulative distribution function

| Syntax | $X=\operatorname{wblinv}(P, A, B)$ |
| :--- | :--- |
|  | $[X, X L O, X U P]=\operatorname{wblinv}(P, A, B, P C O V, a l p h a)$ |

$X=$ wblinv $(P, A, B)$ returns the inverse cumulative distribution function (cdf) for a Weibull distribution with scale parameter $A$ and shape parameter B, evaluated at the values in P. P, A, and B can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array of the same size as the other inputs. The default values for $A$ and $B$ are both 1 .
[X,XLO,XUP] = wblinv(P,A,B,PCOV,alpha) returns confidence bounds for $X$ when the input parameters $A$ and $B$ are estimates. PCOV is a 2 -by- 2 matrix containing the covariance matrix of the estimated parameters. alpha has a default value of 0.05 , and specifies 100(1-alpha) $\%$ confidence bounds. XLO and XUP are arrays of the same size as X containing the lower and upper confidence bounds.

The function wblinv computes confidence bounds for X using a normal approximation to the distribution of the estimate

$$
\log a+\frac{\log q}{b}
$$

where $q$ is the Pth quantile from a Weibull distribution with scale and shape parameters both equal to 1 . The computed bounds give approximately the desired confidence level when you estimate mu, sigma, and PCOV from large samples, but in smaller samples other methods of computing the confidence bounds might be more accurate.

The inverse of the Weibull cdf is

$$
x=F^{-1}(p \mid a, b)=-a[\ln (1-p)]^{1 / b} I_{[0,1]}(p)
$$

## Examples

The lifetimes (in hours) of a batch of light bulbs has a Weibull distribution with parameters $\mathrm{a}=200$ and $\mathrm{b}=6$.

Find the median lifetime of the bulbs:

```
life = wblinv(0.5, 200, 6)
life =
    188.1486
```

Generate 100 random values from this distribution, and estimate the 90th percentile (with confidence bounds) from the random sample

```
x = wblrnd(200,6,100,1);
p = wblfit(x)
[nlogl,pcov] = wbllike(p,x)
[q90,q90lo,q90up] = wblinv(0.9,p(1),p(2),pcov)
p =
    204.8918 6.3920
nlogl =
    496.8915
pcov =
    11.3392 0.5233
            0.5233 0.2573
q90 =
    233.4489
q901o =
```

```
226.0092
q90up =
241.1335
```


## See Also <br> icdf | wblcdf | wblpdf | wblstat | wblfit | wbllike | wblrnd

How To . "Weibull Distribution" on page B-103

## Purpose Weibull negative log-likelihood

Syntax nlogL = wbllike(params, data)
[logL,AVAR] = wbllike(params,data)
[...] = wbllike(params,data,censoring)
[...] = wbllike(params,data,censoring,freq)

Description

## Examples

This example continues the example from wblfit.

$$
r=\text { wblrnd }(0.5,0.8,100,1)
$$

```
[logL, AVAR] = wbllike(wblfit(r),r)
logL =
    47.3349
AVAR =
    0.0048 0.0014
    0.0014 0.0040
```

| References | [1] Patel, J. K., C. H. Kapadia, and D. B. Owen. Handbook of Statistical |
| :--- | :--- |
| Distributions. New York: Marcel Dekker, 1976. |  |

## Purpose

Weibull probability density function

## Syntax

Description
$Y=\operatorname{wblpdf}(X, A, B)$
$Y=$ wblpdf( $X, A, B)$ computes the Weibull pdf at each of the values in $X$ using the corresponding scale parameter, $A$ and shape parameter, $B$. $X, A$, and $B$ can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array of the same size as the other inputs. The parameters in A and B must be positive.

The Weibull pdf is

$$
y=f(x \mid a, b)=b a^{-b} x^{b-1} e^{-\left(\frac{x}{a}\right)^{b}} I_{(0, \infty)}(x)
$$

Some references refer to the Weibull distribution with a single parameter. This corresponds to wblpdf with $A=1$.

The exponential distribution is a special case of the Weibull distribution.

```
lambda = 1:6;
y = wblpdf(0.1:0.1:0.6,lambda,1)
y =
    0.9048
y1 = exppdf(0.1:0.1:0.6,lambda)
y1 =
    0.9048}00.4524 0.3016 0.2262 0.1810 0.1508 
```

[1] Devroye, L. Non-Uniform Random Variate Generation. New York: Springer-Verlag, 1986.

See Also pdf | wblcdf

## How To

- wblfit
- wblinv
- wbllike
- wblplot
- wblrnd
- wblstat
- "Weibull Distribution" on page B-103


## Purpose

Syntax

Description

## Examples

Weibull probability plot
wblplot (X)
$h=w b l p l o t(X)$
wblplot ( X ) displays a Weibull probability plot of the data in X . If X is a matrix, wblplot displays a plot for each column.
$\mathrm{h}=$ wblplot (X) returns handles to the plotted lines.
The purpose of a Weibull probability plot is to graphically assess whether the data in X could come from a Weibull distribution. If the data are Weibull the plot will be linear. Other distribution types might introduce curvature in the plot. wblplot uses midpoint probability plotting positions. Use probplot when the data included censored observations.

```
r = wblrnd(1.2,1.5,50,1);
wblplot(r)
```


## wblplot



See Also
probplot | normplot | wblcdf
How To

- wblfit
- wblinv
- wbllike
- wblpdf
- wblrnd
- wblstat
- "Weibull Distribution" on page B-103

Purpose Weibull random numbers

Syntax $\quad$| $R$ | $=\operatorname{wblrnd}(A, B)$ |
| ---: | :--- |
| $R$ | $=\operatorname{wblrnd}(A, B, m, n, \ldots)$ |
| $R$ | $=\operatorname{wblrnd}(A, B,[m, n, \ldots])$ |

Description

## Examples

References

See Also
How To
$R=$ wblrnd $(A, B)$ generates random numbers for the Weibull distribution with scale parameter, $A$ and shape parameter, $B$. The input arguments A and B can be either scalars or matrices. A and B, can be vectors, matrices, or multidimensional arrays that all have the same size. A scalar input is expanded to a constant array of the same size as the other input.
$R=$ wblrnd $(A, B, m, n, \ldots)$ or $R=$ wblrnd $(A, B,[m, n, \ldots])$ generates an m-by-n-by-... array. The A, B parameters can each be scalars or arrays of the same size as R .

Devroye [1] refers to the Weibull distribution with a single parameter; this is wblrnd with $\mathrm{A}=1$.

```
n1 = wblrnd(0.5:0.5:2,0.5:0.5:2)
n1 =
    0.0178
n2 = wblrnd(1/2,1/2,[1 6])
n2 =
    0.0046 1.7214 2.2108 0.0367 0.0531
```

[1] Devroye, L. Non-Uniform Random Variate Generation. New York: Springer-Verlag, 1986.
random | wblpdf | wblcdf | wblinv | wblstat | wblfit | wbllike

- "Weibull Distribution" on page B-103


## Purpose Weibull mean and variance

## Syntax <br> $[M, V]=$ wblstat $(A, B)$

Description

## Examples

```
[m,v] = wblstat(1:4,1:4)
m =
    1.0000}101.7725 2.6789 3.6256 
v =
    1.0000}00.8584 0.9480 1.0346 
wblstat(0.5,0.7)
ans =
    0.6329
```

See Also
How To
wblpdf | wblcdf | wblinv | wblfit | wbllike | wblrnd

- "Weibull Distribution" on page B-103

Purpose Wishart random numbers

| Syntax | W $=$ wishrnd $($ Sigma, df $)$ |
| :--- | :--- |
|  | W $=$ wishrnd $($ Sigma, df, D $)$ |
|  | $[W, D]=$ wishrnd $($ Sigma,$d f)$ |

Description
W = wishrnd(Sigma, df ) generates a random matrix $W$ having the Wishart distribution with covariance matrix Sigma and with df degrees of freedom. The inverse of $w$ has the Inverse Wishart distribution with parameters Tau $=$ inv(Sigma) and df degrees of freedom.

W = wishrnd(Sigma, df, D) expects D to be the Cholesky factor of Sigma. If you call wishrnd multiple times using the same value of Sigma, it's more efficient to supply D instead of computing it each time.
[W, D] = wishrnd(Sigma,df) returns D so you can provide it as input in future calls to wishrnd.

This function defines the parameter Sigma so that the mean of the output matrix is Sigma*df

See Also iwishrnd
How To . "Wishart Distribution" on page B-106

Purpose $\quad \mathrm{X}$ data used to create ensemble
Description The $X$ property is a numeric matrix of size Nobs-by-Nvars, where Nobs is the number of observations (rows) and Nvars is the number of variables (columns) in the training data. This matrix contains the predictor (or feature) values.

## Purpose <br> Create dataset array from data stored in SAS XPORT format file

## Syntax <br> Description

data = xptread
data = xptread(filename)
[data,missing] = xptread(filename)
xptread(...,'ReadObsNames',true)

Examples
Read in a SAS XPORT format dataset:

See Also
dataset | dataset.export
data $=$ xptread displays a dialog box for selecting a file, then reads data from the file into a dataset array. The file must be in the SAS XPORT format.
data $=$ xptread(filename) retrieves data from a SAS XPORT format file filename. The XPORT format allows for 28 missing data types, represented in the file by an upper case letter, '. ' or '_'. xptread converts All missing data to NaN values in data. However, if you need the specific missing types then you can recover this information by specifying a second output.
[data,missing] = xptread(filename) returns a nominal array, missing, of the same size as data containing the missing data type information from the xport format file. The entries are undefined for values that are not present and are one of '.', '_', 'A',...,'Z' for missing values.
xptread(..., 'ReadObsNames', true) treats the first variable in the file as observation names. The default value is false.
xptread only supports single data sets per file. xptread does not support compressed files.

```
data \(=\) xptread('sample.xpt')
```

```
data \(=\) xptread('sample.xpt')
```


## Purpose <br> Convert predictor matrix to design matrix

## Syntax <br> Description

D = x2fx(X,model)
D = x2fx(X,model,categ)
D = x2fx(X,model, categ,catlevels)
$D=x 2 f x(X$, model $)$ converts a matrix of predictors $X$ to a design matrix D for regression analysis. Distinct predictor variables should appear in different columns of X .

The optional input model controls the regression model. By default, x2fx returns the design matrix for a linear additive model with a constant term. model is one of the following strings:

- 'linear' - Constant and linear terms. This is the default.
- 'interaction' - Constant, linear, and interaction terms
- 'quadratic' - Constant, linear, interaction, and squared terms
- 'purequadratic' - Constant, linear, and squared terms

If $X$ has $n$ columns, the order of the columns of $D$ for a full quadratic model is:

1 The constant term
2 The linear terms (the columns of X , in order $1,2, \ldots, n$ )
3 The interaction terms (pairwise products of the columns of X, in order $(1,2),(1,3), \ldots,(1, n),(2,3), \ldots,(n-1, n))$

4 The squared terms (in order $1,2, \ldots, n$ )
Other models use a subset of these terms, in the same order.
Alternatively, model can be a matrix specifying polynomial terms of arbitrary order. In this case, model should have one column for each column in X and one row for each term in the model. The entries in any row of model are powers for the corresponding columns of X. For
example, if $X$ has columns $X 1, X 2$, and $X 3$, then a row [ $\left.\begin{array}{lll}0 & 1 & 2\end{array}\right]$ in model specifies the term (X1.^0).*(X2.^1).*(X3.^2). A row of all zeros in model specifies a constant term, which can be omitted.
$D=x 2 f x(X$, model, categ $)$ treats columns with numbers listed in the vector categ as categorical variables. Terms involving categorical variables produce dummy variable columns in D. Dummy variables are computed under the assumption that possible categorical levels are completely enumerated by the unique values that appear in the corresponding column of $X$.

D = x2fx(X,model, categ, catlevels) accepts a vector catlevels the same length as categ, specifying the number of levels in each categorical variable. In this case, values in the corresponding column of X must be integers in the range from 1 to the specified number of levels. Not all of the levels need to appear in X.

## Examples

## Example 1

The following converts 2 predictors X 1 and X 2 (the columns of X ) into a design matrix for a full quadratic model with terms constant, $\mathrm{X} 1, \mathrm{X} 2$, X1.*X2, X1.^2, and X2. ${ }^{\wedge} 2$.

X $=\left[\begin{array}{ll}1 & 10\end{array}\right.$
220
310
420
515
6 15];
$D=x 2 f x\left(X\right.$, 'quadratic $\left.{ }^{\prime}\right)$
$D=$

| 1 | 1 | 10 | 10 | 1 | 100 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 2 | 20 | 40 | 4 | 400 |
| 1 | 3 | 10 | 30 | 9 | 100 |
| 1 | 4 | 20 | 80 | 16 | 400 |
| 1 | 5 | 15 | 75 | 25 | 225 |
| 1 | 6 | 15 | 90 | 36 | 225 |

## Example 2

The following converts 2 predictors X1 and X2 (the columns of X) into a design matrix for a quadratic model with terms constant, $\mathrm{X} 1, \mathrm{X} 2$, X1.*X2, and X1. ${ }^{\wedge} 2$.

$$
X=\left[\begin{array}{ll}
1 & 10
\end{array}\right.
$$

220
310
420
515
6 15];
model = [0 0
10
01
11
$20] ;$
D $=x 2 f x(X$, model $)$
D =

| 1 | 1 | 10 | 10 | 1 |
| ---: | ---: | ---: | ---: | ---: |
| 1 | 2 | 20 | 40 | 4 |
| 1 | 3 | 10 | 30 | 9 |
| 1 | 4 | 20 | 80 | 16 |
| 1 | 5 | 15 | 75 | 25 |
| 1 | 6 | 15 | 90 | 36 |

See Also
regstats | rstool | candexch | candgen | cordexch | rowexch

## TreeBagger.Y property

Purpose Y data used to create ensemble
Description The $Y$ property is an array of true class labels for classification, or response values for regression. Y can be a numeric column vector, a character matrix, or a cell array of strings.

## Purpose Standardized $z$-scores

## Syntax <br> Z = zscore(X)

[Z,mu,sigma] = zscore(X)
[...] = zscore(X,1)
[...] = zscore(X,flag,dim)

## Description

## Examples

$Z=z \operatorname{core}(X)$ returns a centered, scaled version of $X$, the same size as $X$. For vector input $x$, output is the vector of $z$-scores $z=$ (x mean(x))./std(x). For matrix input $X, z$-scores are computed using the mean and standard deviation along each column of X. For higher-dimensional arrays, $z$-scores are computed using the mean and standard deviation along the first non-singleton dimension.

The columns of $Z$ have mean zero and standard deviation one (unless a column of $X$ is constant, in which case that column of $Z$ is constant at 0 ). $z$-scores are used to put data on the same scale before further analysis.
[Z,mu,sigma] = zscore(X) also returns mean(X) in mu and std $(X)$ in sigma.
[...] = zscore(X,1) normalizes $X$ using std $(X, 1)$, that is, by computing the standard deviation(s) using $n$ rather than $n-1$, where $n$ is the length of the dimension along which zscore works. zscore $(X, 0)$ is the same as zscore (X).
[...] = zscore(X,flag,dim) standardizes X by working along the dimension dim of $X$. Set flag to 0 to use the default normalization by $n-1$; set flag to 1 to use $n$.

Compare the predictors in the Moore data on original and standardized scales:

```
load moore
predictors = moore(:,1:5);
subplot(2,1,1),plot(predictors)
subplot(2,1,2),plot(zscore(predictors))
```



## See Also

mean | std

## Purpose <br> Syntax <br> Description

$z$-test

```
h = ztest(x,m,sigma)
h = ztest(...,alpha)
h = ztest(...,alpha,tail)
h = ztest(...,alpha,tail,dim)
[h,p] = ztest(...)
[h,p,ci] = ztest(...)
[h,p,ci,zval] = ztest(...)
```

$\mathrm{h}=$ ztest( $\mathrm{x}, \mathrm{m}$, sigma) performs a $z$-test of the null hypothesis that data in the vector $x$ are a random sample from a normal distribution with mean m and standard deviation sigma, against the alternative that the mean is not m . The result of the test is returned in $\mathrm{h} . \mathrm{h}=$ 1 indicates a rejection of the null hypothesis at the $5 \%$ significance level. $\mathrm{h}=0$ indicates a failure to reject the null hypothesis at the $5 \%$ significance level.
x can also be a matrix or an $N$-dimensional array. For matrices, ztest performs separate $z$-tests along each column of x and returns a vector of results. For $N$-dimensional arrays, ztest works along the first non-singleton dimension of $x$.

The test treats NaN values as missing data, and ignores them. h = ztest(..., alpha) performs the test at the (100*alpha)\% significance level. The default, when unspecified, is alpha $=0.05$.
h = ztest(..., alpha, tail) performs the test against the alternative specified by the string tail. There are three options for tail:

- 'both' - Mean is not m (two-tailed test). This is the default, when tail is unspecified.
- 'right' - Mean is greater than m (right-tail test)
- 'left' - Mean is less than m (left-tail test)
tail must be a single string, even when x is a matrix or an $N$-dimensional array.
h = ztest(..., alpha,tail, dim) works along dimension dim of $x$. Use [] to pass in default values for alpha or tail.
$[\mathrm{h}, \mathrm{p}]=\mathrm{ztest}(\ldots)$ returns the $p$ value of the test. The $p$ value is the probability, under the null hypothesis, of observing a value as extreme or more extreme of the test statistic

$$
z=\frac{\bar{x}-\mu}{\sigma / \sqrt{n}}
$$

where $\bar{x}$ is the sample mean, $\mu=\mathrm{m}$ is the hypothesized population mean, $\sigma$ is the population standard deviation, and $n$ is the sample size. Under the null hypothesis, the test statistic will have a standard normal distribution, $N(0,1)$.
[h,p,ci] = ztest(...) returns a 100*(1 - alpha)\% confidence interval on the population mean.
[h,p,ci,zval] = ztest(...) returns the value of the test statistic.

## Examples

Simulate a random sample of size 100 from a normal distribution with mean 0.1 and standard deviation 1 :

```
x = normrnd(0.1,1,1,100);
```

Test the null hypothesis that the sample comes from a standard normal distribution:

```
[h,p,ci] = ztest(x,0,1)
\(h=\)
    0
p \(=\)
    0.1391
ci \(=\)
    \(-0.0481 \quad 0.3439\)
```

The test fails to reject the null hypothesis at the default $\alpha=0.05$ significance level. Under the null hypothesis, the probability of observing a value as extreme or more extreme of the test statistic, as
indicated by the $p$ value, is greater than $a$. The $95 \%$ confidence interval on the mean contains 0 .

Simulate a larger random sample of size 1000 from the same distribution:

```
y = normrnd(0.1,1,1,1000);
```

Test again if the sample comes from a normal distribution with mean 0 :

```
[h,p,ci] = ztest(y,0,1)
h =
    1
\(p=\)
    5.5160e-005
ci \(=\)
    \(0.0655 \quad 0.1895\)
```

This time the test rejects the null hypothesis at the default $a=0.05$ significance level. The $p$ value has fallen below $\alpha=0.05$ and the $95 \%$ confidence interval on the mean does not contain 0 .

Because the $p$ value of the sample y is less than 0.01 , the test will still reject the null hypothesis when the significance level is lowered to $a$ $=0.01$ :

```
[h,p,ci] = ztest(y,0,1,0.01)
h =
    1
p \(=\)
    5.5160e-005
ci \(=\)
    \(0.0461 \quad 0.2090\)
```

This example will produce slightly different results each time it is run, because of the random sampling.

See Also ttest | ttest2

Data Sets

Statistics Toolbox software includes the sample data sets in the following table.

To load a data set into the MATLAB workspace, type:
load filename
where filename is one of the files listed in the table.
Data sets contain individual data variables, description variables with references, and dataset arrays encapsulating the data set and its description, as appropriate.

| File | Description of Data Set |
| :--- | :--- |
| acetylene.mat | Chemical reaction data with correlated predictors |
| arrhythmia.mat | Cardiac arrhythmia data from the UCI machine <br> learning repository |
| carbig.mat | Measurements of cars, 1970-1982 |
| carsmall.mat | Subset of carbig.mat. Measurements of cars, 1970, <br> 1976,1982 |
| cereal.mat | Breakfast cereal ingredients |
| cities.mat | Quality of life ratings for U.S. metropolitan areas |
| discrim.mat | A version of cities.mat used for discriminant <br> analysis |
| examgrades.mat | Exam grades on a scale of 0-100 |
| fisheriris.mat | Fisher's 1936 iris data |
| flu.mat | Google Flu Trends estimated ILI (influenza-like <br> illness) percentage for various regions of the US, <br> and CDC weighted ILI percentage based on sentinel <br> provider reports |
| gas.mat | Gasoline prices around the state of Massachusetts <br> in 1993 |
| hald.mat | Heat of cement vs. mix of ingredients |
| hogg.mat | Bacteria counts in different shipments of milk |


| File | Description of Data Set |
| :--- | :--- |
| hospital.mat | Simulated hospital data |
| imports-85.mat | 1985 Auto Imports Database from the UCI <br> repository |
| ionosphere.mat | Ionosphere dataset from the UCI machine learning <br> repository |
| kmeansdata.mat | Four-dimensional clustered data |
| lawdata.mat | Grade point average and LSAT scores from 15 law <br> schools |
| mileage.mat | Mileage data for three car models from two factories |
| moore.mat | Biochemical oxygen demand on five predictors |
| morse.mat | Recognition of Morse code distinctions by non-coders |
| ovariancancer.mat | Grouped observations on 4000 predictors |
| parts.mat | Dimensional run-out on 36 circular parts |
| polydata.mat | Sample data for polynomial fitting |
| popcorn.mat | Popcorn yield by popper type and brand |
| reaction.mat | Reaction kinetics for Hougen-Watson model |
| sat.dat | Scholastic Aptitude Test averages by gender and <br> test (table) |
| sat2.dat | Scholastic Aptitude Test averages by gender and <br> test (csv) |
| spectra.mat | NIR spectra and octane numbers of 60 gasoline <br> samples |
| stockreturns.mat | Simulated stock returns |

## Distribution Reference

- "Bernoulli Distribution" on page B-3
- "Beta Distribution" on page B-4
- "Binomial Distribution" on page B-7
- "Birnbaum-Saunders Distribution" on page B-10
- "Chi-Square Distribution" on page B-12
- "Copulas" on page B-14
- "Custom Distributions" on page B-15
- "Exponential Distribution" on page B-16
- "Extreme Value Distribution" on page B-19
- "F Distribution" on page B-25
- "Gamma Distribution" on page B-27
- "Gaussian Distribution" on page B-30
- "Gaussian Mixture Distributions" on page B-31
- "Generalized Extreme Value Distribution" on page B-32
- "Generalized Pareto Distribution" on page B-37
- "Geometric Distribution" on page B-41
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- "Uniform Distribution (Discrete)" on page B-101
- "Weibull Distribution" on page B-103
- "Wishart Distribution" on page B-106


## Bernoulli Distribution

## Definition of the Bernoulli Distribution

The Bernoulli distribution is a special case of the binomial distribution, with $n=1$.

## See Also

"Discrete Distributions" on page 5-7

## Beta Distribution

In this section...<br>"Definition" on page B-4<br>"Background" on page B-4<br>"Parameters" on page B-5<br>"Example" on page B-6<br>"See Also" on page B-6

## Definition

The beta pdf is

$$
y=f(x \mid a, b)=\frac{1}{B(a, b)} x^{a-1}(1-x)^{b-1} I_{(0,1)}(x)
$$

where $B(\cdot)$ is the Beta function. The indicator function $I_{(0,1)}(x)$ ensures that only values of $x$ in the range ( 01 ) have nonzero probability.

## Background

The beta distribution describes a family of curves that are unique in that they are nonzero only on the interval (01). A more general version of the function assigns parameters to the endpoints of the interval.

The beta cdf is the same as the incomplete beta function.
The beta distribution has a functional relationship with the $t$ distribution. If $Y$ is an observation from Student's $t$ distribution with $v$ degrees of freedom, then the following transformation generates $X$, which is beta distributed.

$$
X=\frac{1}{2}+\frac{1}{2} \frac{Y}{\sqrt{v+Y^{2}}}
$$

If $Y \sim t(v)$, then $X \square \beta\left(\frac{v}{2}, \frac{v}{2}\right)$
This relationship is used to compute values of the $t \mathrm{cdf}$ and inverse function as well as generating $t$ distributed random numbers.

## Parameters

Suppose you are collecting data that has hard lower and upper bounds of zero and one respectively. Parameter estimation is the process of determining the parameters of the beta distribution that fit this data best in some sense.

One popular criterion of goodness is to maximize the likelihood function. The likelihood has the same form as the beta pdf. But for the pdf, the parameters are known constants and the variable is $x$. The likelihood function reverses the roles of the variables. Here, the sample values (the $x$ 's) are already observed. So they are the fixed constants. The variables are the unknown parameters. Maximum likelihood estimation (MLE) involves calculating the values of the parameters that give the highest likelihood given the particular set of data.

The function betafit returns the MLEs and confidence intervals for the parameters of the beta distribution. Here is an example using random numbers from the beta distribution with $a=5$ and $b=0.2$.

```
r = betarnd(5,0.2,100,1);
[phat, pci] = betafit(r)
phat =
    4.5330 0.2301
pci =
    2.8051 0.1771
    6.2610 0.2832
```

The MLE for parameter $a$ is 4.5330, compared to the true value of 5 . The $95 \%$ confidence interval for $a$ goes from 2.8051 to 6.2610 , which includes the true value.

Similarly the MLE for parameter $b$ is 0.2301 , compared to the true value of 0.2 . The $95 \%$ confidence interval for $b$ goes from 0.1771 to 0.2832 , which
also includes the true value. In this made-up example you know the "true value." In experimentation you do not.

## Example

The shape of the beta distribution is quite variable depending on the values of the parameters, as illustrated by the plot below.


The constant pdf (the flat line) shows that the standard uniform distribution is a special case of the beta distribution.

## See Also

"Continuous Distributions (Data)" on page 5-4

## Binomial Distribution

## In this section...

"Definition" on page B-7
"Background" on page B-7
"Parameters" on page B-8
"Example" on page B-9
"See Also" on page B-9

## Definition

The binomial pdf is

$$
f(k \mid n, p)=\binom{n}{k} p^{k}(1-p)^{n-k}
$$

where $k$ is the number of successes in $n$ trials of a Bernoulli process with probability of success $p$.

The binomial distribution is discrete, defined for integers $k=0,1,2, \ldots n$, where it is nonzero.

## Background

The binomial distribution models the total number of successes in repeated trials from an infinite population under the following conditions:

- Only two outcomes are possible on each of $n$ trials.
- The probability of success for each trial is constant.
- All trials are independent of each other.

The binomial distribution is a generalization of the Bernoulli distribution; it generalizes to the multinomial distribution.

## Parameters

Suppose you are collecting data from a widget manufacturing process, and you record the number of widgets within specification in each batch of 100 . You might be interested in the probability that an individual widget is within specification. Parameter estimation is the process of determining the parameter, $p$, of the binomial distribution that fits this data best in some sense.

One popular criterion of goodness is to maximize the likelihood function. The likelihood has the same form as the binomial pdf above. But for the pdf, the parameters ( $n$ and $p$ ) are known constants and the variable is $x$. The likelihood function reverses the roles of the variables. Here, the sample values (the $x$ 's) are already observed. So they are the fixed constants. The variables are the unknown parameters. MLE involves calculating the value of $p$ that give the highest likelihood given the particular set of data.

The function binofit returns the MLEs and confidence intervals for the parameters of the binomial distribution. Here is an example using random numbers from the binomial distribution with $n=100$ and $p=0.9$.

```
r = binornd(100,0.9)
r =
    88
[phat, pci] = binofit(r,100)
phat =
    0.8800
pci =
    0.7998
    0.9364
```

The MLE for parameter $p$ is 0.8800 , compared to the true value of 0.9 . The $95 \%$ confidence interval for $p$ goes from 0.7998 to 0.9364 , which includes the true value. In this made-up example you know the "true value" of $p$. In experimentation you do not.

## Example

The following commands generate a plot of the binomial pdf for $n=10$ and $p=1 / 2$.

$$
\begin{aligned}
& x=0: 10 ; \\
& y=\operatorname{binopdf}(x, 10,0.5) ; \\
& \operatorname{plot}\left(x, y, '^{\prime}\right)
\end{aligned}
$$



## See Also

"Discrete Distributions" on page 5-7

## Birnbaum-Saunders Distribution

## In this section...

"Definition" on page B-10
"Background" on page B-10
"Parameters" on page B-11
"See Also" on page B-11

## Definition

The Birnbaum-Saunders distribution has the density function

$$
\frac{1}{\sqrt{2 \pi}} \exp \left\{-\frac{(\sqrt{x / \beta}-\sqrt{\beta / x})^{2}}{2 \gamma^{2}}\right\}\left(\frac{(\sqrt{x / \beta}+\sqrt{\beta / x})}{2 \gamma x}\right)
$$

with scale parameter $B>0$ and shape parameter $\gamma>0$, for $x>0$.
If $x$ has a Birnbaum-Saunders distribution with parameters $B$ and $\gamma$, then

$$
\frac{(\sqrt{x / \beta}-\sqrt{\beta / x})}{\gamma}
$$

has a standard normal distribution.

## Background

The Birnbaum-Saunders distribution was originally proposed as a lifetime model for materials subject to cyclic patterns of stress and strain, where the ultimate failure of the material comes from the growth of a prominent flaw. In materials science, Miner's Rule suggests that the damage occurring after $n$ cycles, at a stress level with an expected lifetime of $N$ cycles, is proportional
to $n / N$. Whenever Miner's Rule applies, the Birnbaum-Saunders model is a reasonable choice for a lifetime distribution model.

## Parameters

See mle, dfittool.

## See Also

"Continuous Distributions (Data)" on page 5-4

## Chi-Square Distribution

## In this section...

"Definition" on page B-12
"Background" on page B-12
"Example" on page B-13
"See Also" on page B-13

## Definition

The $x^{2} p d f$ is

$$
y=f(x \mid v)=\frac{x^{(v-2) / 2} e^{-x / 2}}{2^{\frac{v}{2}} \Gamma(v / 2)}
$$

where $\Gamma(\cdot)$ is the Gamma function, and $v$ is the degrees of freedom.

## Background

The $x^{2}$ distribution is a special case of the gamma distribution where $b=2$ in the equation for gamma distribution below.

$$
y=f(x \mid a, b)=\frac{1}{b^{a} \Gamma(a)} x^{a-1} e^{\frac{x}{b}}
$$

The $x^{2}$ distribution gets special attention because of its importance in normal sampling theory. If a set of $n$ observations is normally distributed with variance $\sigma^{2}$, and $s^{2}$ is the sample standard deviation, then

$$
\frac{(n-1) s^{2}}{\sigma^{2}} \square \chi^{2}(n-1)
$$

This relationship is used to calculate confidence intervals for the estimate of the normal parameter $\sigma^{2}$ in the function normfit.

## Example

The $x^{2}$ distribution is skewed to the right especially for few degrees of freedom $(v)$. The plot shows the $x^{2}$ distribution with four degrees of freedom.

```
    x = 0:0.2:15;
    y = chi2pdf(x,4);
plot(x,y)
```



## See Also

"Continuous Distributions (Statistics)" on page 5-6

## Copulas

See "Copulas" on page 5-107.

## Custom Distributions

User-defined custom distributions, created using files and function handles, are supported by the Statistics Toolbox functions pdf, cdf, icdf, and mle, and the Statistics Toolbox GUI dfittool.

## Exponential Distribution

## In this section...

"Definition" on page B-16
"Background" on page B-16
"Parameters" on page B-16
"Example" on page B-17
"See Also" on page B-18

## Definition

The exponential pdf is

$$
y=f(x \mid \mu)=\frac{1}{\mu} e^{\frac{-x}{\mu}}
$$

## Background

Like the chi-square distribution, the exponential distribution is a special case of the gamma distribution (obtained by setting $a=1$ )

$$
y=f(x \mid a, b)=\frac{1}{b^{a} \Gamma(a)} x^{a-1} e^{\frac{x}{b}}
$$

where $\Gamma(\cdot)$ is the Gamma function.
The exponential distribution is special because of its utility in modeling events that occur randomly over time. The main application area is in studies of lifetimes.

## Parameters

Suppose you are stress testing light bulbs and collecting data on their lifetimes. You assume that these lifetimes follow an exponential distribution.

You want to know how long you can expect the average light bulb to last. Parameter estimation is the process of determining the parameters of the exponential distribution that fit this data best in some sense.

One popular criterion of goodness is to maximize the likelihood function. The likelihood has the same form as the exponential pdf above. But for the pdf, the parameters are known constants and the variable is $x$. The likelihood function reverses the roles of the variables. Here, the sample values (the $x$ 's) are already observed. So they are the fixed constants. The variables are the unknown parameters. MLE involves calculating the values of the parameters that give the highest likelihood given the particular set of data.

The function expfit returns the MLEs and confidence intervals for the parameters of the exponential distribution. Here is an example using random numbers from the exponential distribution with $\mu=700$.

```
lifetimes = exprnd(700,100,1);
[muhat, muci] = expfit(lifetimes)
muhat =
```

    672.8207
    muci =
547.4338
810.9437

The MLE for parameter $\mu$ is 672 , compared to the true value of 700 . The $95 \%$ confidence interval for $\mu$ goes from 547 to 811 , which includes the true value.

In the life tests you do not know the true value of $\mu$ so it is nice to have a confidence interval on the parameter to give a range of likely values.

## Example

For exponentially distributed lifetimes, the probability that an item will survive an extra unit of time is independent of the current age of the item. The example shows a specific case of this special property.

$$
1=10: 10: 60 ;
$$

```
lpd = l+0.1;
deltap = (expcdf(lpd,50)-\operatorname{expcdf(1,50))./(1-\operatorname{expcdf(1,50))}}\mathbf{(1,50}(1)
deltap =
    0.0020 0.0020 0.0020 0.0020 0.0020 0.0020
```

The following commands generate a plot of the exponential pdf with its parameter (and mean), $\mu$, set to 2 .

```
x = 0:0.1:10;
y = exppdf(x,2);
plot(x,y)
```



## See Also

"Continuous Distributions (Data)" on page 5-4

## Extreme Value Distribution

## In this section...

"Definition" on page B-19
"Background" on page B-19
"Parameters" on page B-21
"Example" on page B-22
"See Also" on page B-24

## Definition

The probability density function for the extreme value distribution with location parameter $\mu$ and scale parameter $\sigma$ is

$$
y=f(x \mid \mu, \sigma)=\sigma^{-1} \exp \left(\frac{x-\mu}{\sigma}\right) \exp \left(-\exp \left(\frac{x-\mu}{\sigma}\right)\right)
$$

If $T$ has a Weibull distribution with parameters $a$ and $b$, then $\log T$ has an extreme value distribution with parameters $\mu=\log a$ and $\sigma=1 / b$.

## Background

Extreme value distributions are often used to model the smallest or largest value among a large set of independent, identically distributed random values representing measurements or observations. The extreme value distribution is appropriate for modeling the smallest value from a distribution whose tails decay exponentially fast, for example, the normal distribution. It can also model the largest value from a distribution, such as the normal or exponential distributions, by using the negative of the original values.

For example, the following fits an extreme value distribution to minimum values taken over 1000 sets of 500 observations from a normal distribution:

```
xMinima = min(randn(1000,500), [], 2);
paramEstsMinima = evfit(xMinima);
y = linspace(-5,-1.5,1001);
```

```
    hist(xMinima,-4.75:.25:-1.75);
    p = evpdf(y,paramEstsMinima(1),paramEstsMinima(2));
line(y,.25*length(xMinima)*p,'color','r')
```



The following fits an extreme value distribution to the maximum values in each set of observations:

```
xMaxima = max(randn(1000,500), [], 2);
paramEstsMaxima = evfit(-xMaxima);
y = linspace(1.5,5,1001);
hist(xMaxima,1.75:.25:4.75);
p = evpdf(-y,paramEstsMaxima(1),paramEstsMaxima(2));
line(y,.25*length(xMaxima)*p,'color','r')
```



Although the extreme value distribution is most often used as a model for extreme values, you can also use it as a model for other types of continuous data. For example, extreme value distributions are closely related to the Weibull distribution. If $T$ has a Weibull distribution, then $\log (T)$ has a type 1 extreme value distribution.

## Parameters

The function evfit returns the maximum likelihood estimates (MLEs) and confidence intervals for the parameters of the extreme value distribution. The following example shows how to fit some sample data using evfit, including estimates of the mean and variance from the fitted distribution.

Suppose you want to model the size of the smallest washer in each batch of 1000 from a manufacturing process. If you believe that the sizes are
independent within and between each batch, you can fit an extreme value distribution to measurements of the minimum diameter from a series of eight experimental batches. The following code returns the MLEs of the distribution parameters as parmhat and the confidence intervals as the columns of parmci.

```
x = [19.774 20.141 19.44 20.511 21.377 19.003 19.66 18.83];
[parmhat, parmci] = evfit(x)
parmhat =
    20.2506 0.8223
parmci =
    19.644 0.49861
    20.857 1.3562
```

You can find mean and variance of the extreme value distribution with these parameters using the function evstat.

```
[meanfit, varfit] = evstat(parmhat(1),parmhat(2))
meanfit =
    19.776
varfit =
    1.1123
```


## Example

The following code generates a plot of the pdf for the extreme value distribution.

```
t = [-5:.01:2];
y = evpdf(t);
plot(t,y)
```



The extreme value distribution is skewed to the left, and its general shape remains the same for all parameter values. The location parameter, mu, shifts the distribution along the real line, and the scale parameter, sigma, expands or contracts the distribution. This example plots the probability function for different combinations of mu and sigma.

```
x = -15:.01:5;
plot(x,evpdf(x,2,1),'-', ...
    x,evpdf(x,0,2),':', ...
    x,evpdf(x,-2,4),' -.');
legend({'mu = 2, sigma = 1', ...
    'mu = 0, sigma = 2', ...
    'mu = -2, sigma = 4'}, ...
    'Location','NW')
xlabel('x')
ylabel('f(x|mu,sigma)')
```



## See Also

"Continuous Distributions (Data)" on page 5-4

## F Distribution

## In this section...

"Definition" on page B-25
"Background" on page B-25
"Example" on page B-26
"See Also" on page B-26

## Definition

The pdf for the $F$ distribution is

$$
y=f\left(x \mid v_{1}, v_{2}\right)=\frac{\Gamma\left[\frac{\left(v_{1}+v_{2}\right)}{2}\right]}{\Gamma\left(\frac{v_{1}}{2}\right) \Gamma\left(\frac{v_{2}}{2}\right)}\left(\frac{v_{1}}{v_{2}}\right)^{\frac{v_{1}}{2}} \frac{x^{\frac{v_{1}-2}{2}}}{\left[1+\left(\frac{v_{1}}{v_{2}}\right) x\right]^{\frac{v_{1}+v_{2}}{2}}}
$$

where $\Gamma(\cdot)$ is the Gamma function.

## Background

The $F$ distribution has a natural relationship with the chi-square distribution. If $X_{1}$ and $X_{2}$ are both chi-square with $v_{1}$ and $v_{2}$ degrees of freedom respectively, then the statistic $F$ below is $F$-distributed.

$$
F\left(v_{1}, v_{2}\right)=\frac{\chi_{1} / v_{1}}{\chi_{2} / v_{2}}
$$

The two parameters, $v_{1}$ and $v_{2}$, are the numerator and denominator degrees of freedom. That is, $v_{1}$ and $v_{2}$ are the number of independent pieces of information used to calculate $x_{1}$ and $x_{2}$, respectively.

## Example

The most common application of the $F$ distribution is in standard tests of hypotheses in analysis of variance and regression.

The plot shows that the $F$ distribution exists on the positive real numbers and is skewed to the right.

$$
\begin{aligned}
& x=0: 0.01: 10 ; \\
& y=\operatorname{fpdf}(x, 5,3) ; \\
& \operatorname{plot}(x, y)
\end{aligned}
$$



## See Also

"Continuous Distributions (Statistics)" on page 5-6

## Gamma Distribution

## In this section...

"Definition" on page B-27
"Background" on page B-27
"Parameters" on page B-28
"Example" on page B-29
"See Also" on page B-29

## Definition

The gamma pdf is

$$
y=f(x \mid a, b)=\frac{1}{b^{a} \Gamma(a)} x^{a-1} e^{\frac{-x}{b}}
$$

where $\Gamma(\cdot)$ is the Gamma function.

## Background

The gamma distribution models sums of exponentially distributed random variables.

The gamma distribution family is based on two parameters. The chi-square and exponential distributions, which are children of the gamma distribution, are one-parameter distributions that fix one of the two gamma parameters.

The gamma distribution has the following relationship with the incomplete Gamma function.

$$
f(x \mid a, b)=\text { gammainc }\left(\frac{x}{b}, a\right)
$$

When $a$ is large, the gamma distribution closely approximates a normal distribution with the advantage that the gamma distribution has density only for positive real numbers.

## Parameters

Suppose you are stress testing computer memory chips and collecting data on their lifetimes. You assume that these lifetimes follow a gamma distribution. You want to know how long you can expect the average computer memory chip to last. Parameter estimation is the process of determining the parameters of the gamma distribution that fit this data best in some sense.

One popular criterion of goodness is to maximize the likelihood function. The likelihood has the same form as the gamma pdf above. But for the pdf, the parameters are known constants and the variable is $x$. The likelihood function reverses the roles of the variables. Here, the sample values (the $x$ 's) are already observed. So they are the fixed constants. The variables are the unknown parameters. MLE involves calculating the values of the parameters that give the highest likelihood given the particular set of data.

The function gamfit returns the MLEs and confidence intervals for the parameters of the gamma distribution. Here is an example using random numbers from the gamma distribution with $a=10$ and $b=5$.

```
lifetimes = gamrnd(10,5,100,1);
[phat, pci] = gamfit(lifetimes)
phat =
    10.9821 4.7258
pci =
    7.4001 3.1543
    14.5640 6.2974
```

Note phat (1) $=\hat{a}$ and phat (2) $=\hat{b}$. The MLE for parameter $a$ is 10.98 , compared to the true value of 10 . The $95 \%$ confidence interval for $a$ goes from 7.4 to 14.6 , which includes the true value.

Similarly the MLE for parameter $b$ is 4.7, compared to the true value of 5 . The $95 \%$ confidence interval for $b$ goes from 3.2 to 6.3 , which also includes the true value.

In the life tests you do not know the true value of $a$ and $b$ so it is nice to have a confidence interval on the parameters to give a range of likely values.

## Example

In the example the gamma pdf is plotted with the solid line. The normal pdf has a dashed line type.

```
    x = gaminv((0.005:0.01:0.995),100,10);
    y = gampdf(x,100,10);
    y1 = normpdf(x,1000,100);
    plot(x,y,'-',x,y1,'-.')
```



## See Also

"Continuous Distributions (Data)" on page 5-4

## Gaussian Distribution

See "Normal Distribution" on page B-83.

## Gaussian Mixture Distributions

See the discussion of the gmdistribution class in the "Gaussian Mixture Models" on page 5-99 section of "Probability Distributions Used for Multivariate Modeling" on page 5-99 and the "Gaussian Mixture Models" on page 11-28 section of Chapter 11, "Cluster Analysis".

## Generalized Extreme Value Distribution

## In this section...

"Definition" on page B-32
"Background" on page B-32
"Parameters" on page B-33
"Example" on page B-34
"See Also" on page B-36

## Definition

The probability density function for the generalized extreme value distribution with location parameter $\mu$, scale parameter $\sigma$, and shape parameter $k \neq 0$ is

$$
y=f(x \mid \mathrm{k}, \mu, \sigma)=\left(\frac{1}{\sigma}\right) \exp \left(-\left(1+\mathrm{k} \frac{(x-\mu)}{\sigma}\right)^{-\frac{1}{\mathrm{k}}}\right)\left(1+\mathrm{k} \frac{(x-\mu)}{\sigma}\right)^{-1-\frac{1}{\mathrm{k}}}
$$

for

$$
1+k \frac{(x-\mu)}{\sigma}>0
$$

k > 0 corresponds to the Type II case, while k < 0 corresponds to the Type III case. For $\mathrm{k}=0$, corresponding to the Type I case, the density is

$$
y=f(x \mid 0, \mu, \sigma)=\left(\frac{1}{\sigma}\right) \exp \left(-\exp \left(-\frac{(x-\mu)}{\sigma}\right)-\frac{(x-\mu)}{\sigma}\right)
$$

## Background

Like the extreme value distribution, the generalized extreme value distribution is often used to model the smallest or largest value among a large set of independent, identically distributed random values representing measurements or observations. For example, you might have batches of 1000
washers from a manufacturing process. If you record the size of the largest washer in each batch, the data are known as block maxima (or minima if you record the smallest). You can use the generalized extreme value distribution as a model for those block maxima.

The generalized extreme value combines three simpler distributions into a single form, allowing a continuous range of possible shapes that includes all three of the simpler distributions. You can use any one of those distributions to model a particular dataset of block maxima. The generalized extreme value distribution allows you to "let the data decide" which distribution is appropriate.

The three cases covered by the generalized extreme value distribution are often referred to as the Types I, II, and III. Each type corresponds to the limiting distribution of block maxima from a different class of underlying distributions. Distributions whose tails decrease exponentially, such as the normal, lead to the Type I. Distributions whose tails decrease as a polynomial, such as Student's $t$, lead to the Type II. Distributions whose tails are finite, such as the beta, lead to the Type III.

Types I, II, and III are sometimes also referred to as the Gumbel, Frechet, and Weibull types, though this terminology can be slightly confusing. The Type I (Gumbel) and Type III (Weibull) cases actually correspond to the mirror images of the usual Gumbel and Weibull distributions, for example, as computed by the functions evcdf and evfit, or wblcdf and wblfit, respectively. Finally, the Type II (Frechet) case is equivalent to taking the reciprocal of values from a standard Weibull distribution.

## Parameters

If you generate 250 blocks of 1000 random values drawn from Student's $t$ distribution with 5 degrees of freedom, and take their maxima, you can fit a generalized extreme value distribution to those maxima.

```
blocksize = 1000;
nblocks = 250;
t = trnd(5,blocksize,nblocks);
x = max(t); % 250 column maxima
paramEsts = gevfit(x)
paramEsts =
```

$$
0.2438 \quad 1.1760 \quad 5.8045
$$

Notice that the shape parameter estimate (the first element) is positive, which is what you would expect based on block maxima from a Student's $t$ distribution.

```
hist(x,2:20);
set(get(gca,'child'),'FaceColor',[.8 . 8 1])
xgrid = linspace(2,20,1000);
line(xgrid,nblocks*...
    gevpdf(xgrid,paramEsts(1), paramEsts(2), paramEsts(3)));
```



## Example

The following code generates examples of probability density functions for the three basic forms of the generalized extreme value distribution.


Notice that for $\mathrm{k}>0$, the distribution has zero probability density for x such that

$$
x<-\frac{\sigma}{\mathrm{k}}+\mu
$$

For $\mathrm{k}<0$, the distribution has zero probability density for

$$
x>-\frac{\sigma}{\mathrm{k}}+\mu
$$

For $\mathrm{k}=0$, there is no upper or lower bound.

## See Also

"Continuous Distributions (Data)" on page 5-4

## Generalized Pareto Distribution

## In this section...

"Definition" on page B-37
"Background" on page B-37
"Parameters" on page B-38
"Example" on page B-39
"See Also" on page B-40

## Definition

The probability density function for the generalized Pareto distribution with shape parameter $k \neq 0$, scale parameter $\sigma$, and threshold parameter $\theta$, is

$$
y=f(x \mid k, \sigma, \theta)=\left(\frac{1}{\sigma}\right)\left(1+k \frac{(x-\theta)}{\sigma}\right)^{-1-\frac{1}{k}}
$$

for $\theta<x$, when $k>0$, or for $\theta<x<-\sigma / k$ when $k<0$.
For $k=0$, the density is

$$
y=f(x \mid 0, \sigma, \theta)=\left(\frac{1}{\sigma}\right) \mathrm{e}^{-\frac{(x-\theta)}{\sigma}}
$$

for $\theta<x$.
If $k=0$ and $\theta=0$, the generalized Pareto distribution is equivalent to the exponential distribution. If $k>0$ and $\theta=\sigma / k$, the generalized Pareto distribution is equivalent to the Pareto distribution.

## Background

Like the exponential distribution, the generalized Pareto distribution is often used to model the tails of another distribution. For example, you might have washers from a manufacturing process. If random influences in the process lead to differences in the sizes of the washers, a standard probability
distribution, such as the normal, could be used to model those sizes. However, while the normal distribution might be a good model near its mode, it might not be a good fit to real data in the tails and a more complex model might be needed to describe the full range of the data. On the other hand, only recording the sizes of washers larger (or smaller) than a certain threshold means you can fit a separate model to those tail data, which are known as exceedences. You can use the generalized Pareto distribution in this way, to provide a good fit to extremes of complicated data.

The generalized Pareto distribution allows a continuous range of possible shapes that includes both the exponential and Pareto distributions as special cases. You can use either of those distributions to model a particular dataset of exceedences. The generalized Pareto distribution allows you to "let the data decide" which distribution is appropriate.

The generalized Pareto distribution has three basic forms, each corresponding to a limiting distribution of exceedence data from a different class of underlying distributions.

- Distributions whose tails decrease exponentially, such as the normal, lead to a generalized Pareto shape parameter of zero.
- Distributions whose tails decrease as a polynomial, such as Student's $t$, lead to a positive shape parameter.
- Distributions whose tails are finite, such as the beta, lead to a negative shape parameter.

The generalized Pareto distribution is used in the tails of distribution fit objects of the paretotails class.

## Parameters

If you generate a large number of random values from a Student's $t$ distribution with 5 degrees of freedom, and then discard everything less than 2, you can fit a generalized Pareto distribution to those exceedences.

```
t = trnd(5,5000,1);
y = t(t > 2) - 2;
paramEsts = gpfit(y)
paramEsts =
```

$$
0.1267 \quad 0.8134
$$

Notice that the shape parameter estimate (the first element) is positive, which is what you would expect based on exceedences from a Student's $t$ distribution.

```
hist(y+2,2.25:.5:11.75);
set(get(gca,'child'),'FaceColor',[.8 . 8 1])
xgrid = linspace(2,12,1000);
line(xgrid,.5*length(y)*...
    gppdf(xgrid,paramEsts(1),paramEsts(2),2));
```



## Example

The following code generates examples of the probability density functions for the three basic forms of the generalized Pareto distribution.

```
x = linspace(0,10,1000);
```



Notice that for $\mathrm{k}<0$, the distribution has zero probability density for $x>-\frac{\sigma}{\mathrm{k}}$, while for $k \geq 0$, there is no upper bound.

## See Also

"Continuous Distributions (Data)" on page 5-4

## Geometric Distribution

In this section...<br>"Definition" on page B-41<br>"Background" on page B-41<br>"Example" on page B-41<br>"See Also" on page B-42

## Definition

The geometric pdf is

$$
y=f(x \mid p)=p q^{x} I_{(0,1, \ldots)}(x)
$$

where $q=1-p$. The geometric distribution is a special case of the negative binomial distribution, with $r=1$.

## Background

The geometric distribution is discrete, existing only on the nonnegative integers. It is useful for modeling the runs of consecutive successes (or failures) in repeated independent trials of a system.

The geometric distribution models the number of successes before one failure in an independent succession of tests where each test results in success or failure.

## Example

Suppose the probability of a five-year-old battery failing in cold weather is 0.03. What is the probability of starting 25 consecutive days during a long cold snap?

```
1 - geocdf(25,0.03)
```

ans $=$

$$
0.4530
$$

The plot shows the cdf for this scenario.


## See Also

"Discrete Distributions" on page 5-7

## Hypergeometric Distribution

## In this section...

"Definition" on page B-43
"Background" on page B-43
"Example" on page B-44
"See Also" on page B-44

## Definition

The hypergeometric pdf is

$$
y=f(x \mid M, K, n)=\frac{\binom{K}{x}\binom{M-K}{n-x}}{\binom{M}{n}}
$$

## Background

The hypergeometric distribution models the total number of successes in a fixed-size sample drawn without replacement from a finite population.

The distribution is discrete, existing only for nonnegative integers less than the number of samples or the number of possible successes, whichever is greater. The hypergeometric distribution differs from the binomial only in that the population is finite and the sampling from the population is without replacement.

The hypergeometric distribution has three parameters that have direct physical interpretations.

- $M$ is the size of the population.
- $K$ is the number of items with the desired characteristic in the population.
- $n$ is the number of samples drawn.

Sampling "without replacement" means that once a particular sample is chosen, it is removed from the relevant population for all subsequent selections.

## Example

The plot shows the cdf of an experiment taking 20 samples from a group of 1000 where there are 50 items of the desired type.

```
    x = 0:10;
    y = hygecdf(x,1000,50,20);
    stairs(x,y)
```



## See Also

"Discrete Distributions" on page 5-7

## Inverse Gaussian Distribution

## In this section...

"Definition" on page B- 45
"Background" on page B-45
"Parameters" on page B-45
"See Also" on page B-45

## Definition

The inverse Gaussian distribution has the density function

$$
\sqrt{\frac{\lambda}{2 \pi x^{3}}} \exp \left\{-\frac{\lambda}{2 \mu^{2} x}(x-\mu)^{2}\right\}
$$

## Background

Also known as the Wald distribution, the inverse Gaussian is used to model nonnegative positively skewed data. The distribution originated in the theory of Brownian motion, but has been used to model diverse phenomena. Inverse Gaussian distributions have many similarities to standard Gaussian (normal) distributions, which lead to applications in inferential statistics.

## Parameters

See mle, dfittool.

## See Also

"Continuous Distributions (Data)" on page 5-4

## Inverse Wishart Distribution

## Definition

The probability density function of the $d$-dimensional Inverse Wishart distribution is given by

$$
y=f(\mathrm{X}, \Sigma, v)=\frac{|T|^{(v / 2)} e^{\left(-\frac{1}{2} \operatorname{trace}\left(T X^{-1}\right)\right)}}{2^{(v d) / 2} \pi^{(d(d-1)) / 4}|X|^{(v+d+1) / 2} \Gamma(v / 2) \ldots \Gamma(v-(d-1)) / 2}
$$

where $X$ and $T$ are $d$-by- $d$ symmetric positive definite matrices, and $v$ is a scalar greater than or equal to $d$. While it is possible to define the Inverse Wishart for singular $T$, the density cannot be written as above.

If a random matrix has a Wishart distribution with parameters $T^{-1}$ and $v$, then the inverse of that random matrix has an inverse Wishart distribution with parameters $T$ and $v$. The mean of the distribution is given by

$$
\frac{1}{v-d-1} T
$$

where $d$ is the number of rows and columns in $T$.
Only random matrix generation is supported for the inverse Wishart, including both singular and nonsingular $T$.

## Background

The inverse Wishart distribution is based on the Wishart distribution. In Bayesian statistics it is used as the conjugate prior for the covariance matrix of a multivariate normal distribution.

## Example

Notice that the sampling variability is quite large when the degrees of freedom is small.

```
Tau = [1 .5; .5 2];
df = 10; S1 = iwishrnd(Tau,df)*(df-2-1)
```

```
S1 =
            1.7959 0.64107
            0.64107 1.5496
df = 1000; S2 = iwishrnd(Tau,df)*(df-2-1)
S2 =
            0.9842 0.50158
            0.50158 2.1682
```


## See Also

"Wishart Distribution" on page B-106, "Multivariate Distributions" on page 5-8

## Johnson System

See "Pearson and Johnson Systems" on page 6-25.

## Logistic Distribution

In this section...<br>"Definition" on page B-49<br>"Background" on page B-49<br>"Parameters" on page B-49<br>"See Also" on page B-49

## Definition

The logistic distribution has the density function

$$
\frac{e^{\frac{x-\mu}{\sigma}}}{\sigma\left(1+e^{\frac{x-\mu}{\sigma}}\right)^{2}}
$$

with location parameter $\mu$ and scale parameter $\sigma>0$, for all real $x$.

## Background

The logistic distribution originated with Verhulst's work on demography in the early 1800s. The distribution has been used for various growth models, and is used in logistic regression. It has longer tails and a higher kurtosis than the normal distribution.

## Parameters

See mle, dfittool.

## See Also

"Continuous Distributions (Data)" on page 5-4

## Loglogistic Distribution

In this section...<br>"Definition" on page B-50<br>"Parameters" on page B-50<br>"See Also" on page B-50

## Definition

The variable $x$ has a loglogistic distribution with location parameter $\mu$ and scale parameter $\sigma>0$ if $\ln x$ has a logistic distribution with parameters $\mu$ and $\sigma$. The relationship is similar to that between the lognormal and normal distribution.

## Parameters

See mle, dfittool.

## See Also

"Continuous Distributions (Data)" on page 5-4

## Lognormal Distribution

## In this section...

"Definition" on page B-51
"Background" on page B-51
"Example" on page B-52
"See Also" on page B-53

## Definition

The lognormal pdf is

$$
y=f(x \mid \mu, \sigma)=\frac{1}{x \sigma \sqrt{2 \pi}} e^{\frac{-(\ln x-\mu)^{2}}{2 \sigma^{2}}}
$$

## Background

The normal and lognormal distributions are closely related. If $X$ is distributed $\operatorname{lognormally}$ with parameters $\mu$ and $\sigma$, then $\log (X)$ is distributed normally with mean $\mu$ and standard deviation $\sigma$.

The mean $m$ and variance $v$ of a lognormal random variable are functions of $\mu$ and $\sigma$ that can be calculated with the lognstat function. They are:

$$
\begin{aligned}
& m=\exp \left(\mu+\sigma^{2} / 2\right) \\
& v=\exp \left(2 \mu+\sigma^{2}\right)\left(\exp \left(\sigma^{2}\right)-1\right)
\end{aligned}
$$

A lognormal distribution with mean $m$ and variance $v$ has parameters

$$
\begin{aligned}
& \mu=\log \left(m^{2} / \sqrt{v+m^{2}}\right) \\
& \sigma=\sqrt{\log \left(v / m^{2}+1\right)}
\end{aligned}
$$

The lognormal distribution is applicable when the quantity of interest must be positive, since $\log (X)$ exists only when $X$ is positive.

## Example

Suppose the income of a family of four in the United States follows a lognormal distribution with $\mu=\log (20,000)$ and $\sigma^{2}=1.0$. Plot the income density.

```
    x = (10:1000:125010)';
    y = lognpdf(x,log(20000),1.0);
    plot(x,y)
    set(gca,'xtick',[0 30000 60000 90000 120000])
    set(gca,'xticklabel',{'0','$30,000','$60,000',...
        '$90,000', '$120,000'})
```



## See Also

"Continuous Distributions (Data)" on page 5-4

## Multinomial Distribution

## In this section...

"Definition" on page B-54
"Background" on page B-54
"Example" on page B-54

## Definition

The multinomial pdf is

$$
f(x \mid n, p)=\frac{n!}{x_{1}!\cdots x_{k}!} p_{1}^{x_{1}} \cdots p_{k}^{x_{k}}
$$

where $x=\left(x_{1}, \ldots, x_{k}\right)$ gives the number of each of $k$ outcomes in $n$ trials of a process with fixed probabilities $p=\left(p_{1}, \ldots, p_{k}\right)$ of individual outcomes in any one trial. The vector $x$ has non-negative integer components that sum to $n$. The vector $p$ has non-negative integer components that sum to 1 .

## Background

The multinomial distribution is a generalization of the binomial distribution. The binomial distribution gives the probability of the number of "successes" and "failures" in $n$ independent trials of a two-outcome process. The probability of "success" and "failure" in any one trial is given by the fixed probabilities $p$ and $q=1-p$. The multinomial distribution gives the probability of each combination of outcomes in $n$ independent trials of a $k$-outcome process. The probability of each outcome in any one trial is given by the fixed probabilities $p_{1}, \ldots, p_{\mathrm{k}}$.

The expected value of outcome $i$ is $n p_{i}$. The variance of outcome $i$ is $n p_{i}\left(1-p_{i}\right)$. The covariance of outcomes $i$ and $j$ is $-n p_{i} p_{j}$ for distinct $i$ and $j$.

## Example

The following uses mnpdf to produce a visualization of a trinomial distribution:

```
    % Compute the distribution
    p = [1/2 1/3 1/6]; % Outcome probabilities
    n = 10; % Sample size
    x1 = 0:n;
    x2 = 0:n;
    [X1,X2] = meshgrid(x1,x2);
    X3 = n-(X1+X2);
    Y = mnpdf([X1(:),X2(:),X3(:)],repmat(p,(n+1)^2,1));
    % Plot the distribution
    Y = reshape(Y,n+1,n+1);
    bar3(Y)
    set(gca,'XTickLabel',0:n)
    set(gca,'YTickLabel',0:n)
    xlabel('x_1')
    ylabel('x_2')
    zlabel('Probability Mass')
```

Trinomial Distribution


Note that the visualization does not show $x_{3}$, which is determined by the constraint $x_{1}+x_{2}+x_{3}=n$.

# Multivariate Gaussian Distribution 

See "Multivariate Normal Distribution" on page B-58.

## Multivariate Normal Distribution

## In this section...

"Definition" on page B-58
"Background" on page B-58
"Example" on page B-59
"See Also" on page B-63

## Definition

The probability density function of the $d$-dimensional multivariate normal distribution is given by

$$
y=f(x, \mu, \Sigma)=\frac{1}{\sqrt{|\Sigma|(2 \pi)^{d}}} e^{-\frac{1}{2}(x-\mu) \Sigma^{-1}(x-\mu)^{\prime}}
$$

where $x$ and $\mu$ are 1-by- $d$ vectors and $\Sigma$ is a $d$-by- $d$ symmetric positive definite matrix. While it is possible to define the multivariate normal for singular $\Sigma$, the density cannot be written as above. Only random vector generation is supported for the singular case. Note that while most textbooks define the multivariate normal with $x$ and $\mu$ oriented as column vectors, for the purposes of data analysis software, it is more convenient to orient them as row vectors, and Statistics Toolbox software uses that orientation.

## Background

The multivariate normal distribution is a generalization of the univariate normal to two or more variables. It is a distribution for random vectors of correlated variables, each element of which has a univariate normal distribution. In the simplest case, there is no correlation among variables, and elements of the vectors are independent univariate normal random variables.

The multivariate normal distribution is parameterized with a mean vector, $\mu$, and a covariance matrix, $\Sigma$. These are analogous to the mean $\mu$ and variance $\sigma^{2}$ parameters of a univariate normal distribution. The diagonal elements of $\Sigma$
contain the variances for each variable, while the off-diagonal elements of $\Sigma$ contain the covariances between variables.

The multivariate normal distribution is often used as a model for multivariate data, primarily because it is one of the few multivariate distributions that is tractable to work with.

## Example

This example shows the probability density function (pdf) and cumulative distribution function (cdf) for a bivariate normal distribution with unequal standard deviations. You can use the multivariate normal distribution in a higher number of dimensions as well, although visualization is not easy.

```
mu = [0 0];
Sigma = [.25 .3; .3 1];
x1 = -3:.2:3; x2 = -3:.2:3;
[X1,X2] = meshgrid(x1,x2);
F = mvnpdf([X1(:) X2(:)],mu,Sigma);
F = reshape(F,length(x2),length(x1));
surf(x1,x2,F);
caxis([min(F(:))-.5*range(F(:)),max(F(:))]);
axis([-3 3 -3 3 0 .4])
xlabel('x1'); ylabel('x2'); zlabel('Probability Density');
```



```
F = mvncdf([X1(:) X2(:)],mu,Sigma);
F = reshape(F,length(x2),length(x1));
surf(x1,x2,F);
caxis([min(F(:))-.5*range(F(:)),max(F(:))]);
axis([-3 3 -3 3 0 1])
xlabel('x1'); ylabel('x2'); zlabel('Cumulative Probability');
```



Since the bivariate normal distribution is defined on the plane, you can also compute cumulative probabilities over rectangular regions. For example, this contour plot illustrates the computation that follows, of the probability contained within the unit square.

```
contour(x1, x2,F,[.0001 .001 .01 .05:.1:.95 .99 .999 .9999]);
xlabel('x'); ylabel('y');
line([0 0 1 1 0],[1 0 0 1 1],'linestyle','--','color','k');
```



Computing a multivariate cumulative probability requires significantly more work than computing a univariate probability. By default, the mvncdf function computes values to less than full machine precision, and returns an estimate of the error as an optional second output:

```
[F,err] = mvncdf([0 0],[1 1],mu,Sigma)
F =
    0.20974
err =
    1e-008
```


## See Also

"Multivariate Distributions" on page 5-8

## Multivariate $\dagger$ Distribution

In this section...<br>"Definition" on page B-64<br>"Background" on page B-64<br>"Example" on page B-65<br>"See Also" on page B-69

## Definition

The probability density function of the $d$-dimensional multivariate Student's $t$ distribution is given by

$$
f(x, \Sigma, v)=\frac{1}{|\Sigma|^{1 / 2}} \frac{1}{\sqrt{(v \pi)^{d}}} \frac{\Gamma((v+d) / 2)}{\Gamma(v / 2)}\left(1+\frac{x^{\prime} \Sigma^{-1} x}{v}\right)^{-(v+d) / 2}
$$

where $x$ is a 1 -by- $d$ vector, $\Sigma$ is a $d$-by- $d$ symmetric, positive definite matrix, and $v$ is a positive scalar. While it is possible to define the multivariate Student's $t$ for singular $\Sigma$, the density cannot be written as above. For the singular case, only random number generation is supported. Note that while most textbooks define the multivariate Student's $t$ with $x$ oriented as a column vector, for the purposes of data analysis software, it is more convenient to orient $x$ as a row vector, and Statistics Toolbox software uses that orientation.

## Background

The multivariate Student's $t$ distribution is a generalization of the univariate Student's $t$ to two or more variables. It is a distribution for random vectors of correlated variables, each element of which has a univariate Student's $t$ distribution. In the same way as the univariate Student's $t$ distribution can be constructed by dividing a standard univariate normal random variable by the square root of a univariate chi-square random variable, the multivariate Student's $t$ distribution can be constructed by dividing a multivariate normal random vector having zero mean and unit variances by a univariate chi-square random variable.

The multivariate Student's $t$ distribution is parameterized with a correlation matrix, $\Sigma$, and a positive scalar degrees of freedom parameter, $v . v$ is analogous to the degrees of freedom parameter of a univariate Student's $t$ distribution. The off-diagonal elements of $\Sigma$ contain the correlations between variables. Note that when $\Sigma$ is the identity matrix, variables are uncorrelated; however, they are not independent.

The multivariate Student's $t$ distribution is often used as a substitute for the multivariate normal distribution in situations where it is known that the marginal distributions of the individual variables have fatter tails than the normal.

## Example

This example shows the probability density function (pdf) and cumulative distribution function (cdf) for a bivariate Student's $t$ distribution. You can use the multivariate Student's $t$ distribution in a higher number of dimensions as well, although visualization is not easy.

```
Rho = [1 .6; .6 1];
nu = 5;
x1 = -3:.2:3; x2 = -3:.2:3;
[X1,X2] = meshgrid(x1,x2);
F = mvtpdf([X1(:) X2(:)],Rho,nu);
F = reshape(F,length(x2),length(x1));
surf(x1,x2,F);
caxis([min(F(:))-.5*range(F(:)),max(F(:))]);
axis([-3 3 -3 3 0 .2])
xlabel('x1'); ylabel('x2'); zlabel('Probability Density');
```




Since the bivariate Student's $t$ distribution is defined on the plane, you can also compute cumulative probabilities over rectangular regions. For example, this contour plot illustrates the computation that follows, of the probability contained within the unit square.

```
contour(x1,x2,F,[.0001 .001 .01 .05:.1:.95 .99 .999 .9999]);
xlabel('x'); ylabel('y');
line([0 0 1 1 0],[1 0 0 1 1],'linestyle','--','color','k');
```



Computing a multivariate cumulative probability requires significantly more work than computing a univariate probability. By default, the mvtcdf function computes values to less than full machine precision, and returns an estimate of the error as an optional second output:

```
[F,err] = mvtcdf([0 0],[1 1],Rho,nu)
F =
    0.14013
err =
    1e-008
```


## See Also

"Multivariate Distributions" on page 5-8

## Nakagami Distribution

## In this section...

"Definition" on page B-70
"Background" on page B-70
"Parameters" on page B-70
"See Also" on page B-71

## Definition

The Nakagami distribution has the density function

$$
2\left(\frac{\mu}{\omega}\right)^{\mu} \frac{1}{\Gamma(\mu)} x^{(2 \mu-1)} e^{\frac{-\mu}{\omega} x^{2}}
$$

with shape parameter $\mu$ and scale parameter $\omega>0$, for $x>0$. If $x$ has a Nakagami distribution with parameters $\mu$ and $\omega$, then $x^{2}$ has a gamma distribution with shape parameter $\mu$ and scale parameter $\omega / \mu$.

## Background

In communications theory, Nakagami distributions, Rician distributions, and Rayleigh distributions are used to model scattered signals that reach a receiver by multiple paths. Depending on the density of the scatter, the signal will display different fading characteristics. Rayleigh and Nakagami distributions are used to model dense scatters, while Rician distributions model fading with a stronger line-of-sight. Nakagami distributions can be reduced to Rayleigh distributions, but give more control over the extent of the fading.

## Parameters

See mle, dfittool.

## See Also

"Continuous Distributions (Data)" on page 5-4

## Negative Binomial Distribution

## In this section...

"Definition" on page B-72
"Background" on page B-72
"Parameters" on page B-73
"Example" on page B-75
"See Also" on page B-75

## Definition

When the $r$ parameter is an integer, the negative binomial pdf is

$$
y=f(x \mid r, p)=\binom{r+x-1}{x} p^{r} q^{x} I_{(0,1, \ldots)}(x)
$$

where $q=1-p$. When $r$ is not an integer, the binomial coefficient in the definition of the pdf is replaced by the equivalent expression

$$
\frac{\Gamma(r+x)}{\Gamma(r) \Gamma(x+1)}
$$

## Background

In its simplest form (when $r$ is an integer), the negative binomial distribution models the number of failures $x$ before a specified number of successes is reached in a series of independent, identical trials. Its parameters are the probability of success in a single trial, $p$, and the number of successes, $r$. A special case of the negative binomial distribution, when $r=1$, is the geometric distribution, which models the number of failures before the first success.

More generally, $r$ can take on non-integer values. This form of the negative binomial distribution has no interpretation in terms of repeated trials, but, like the Poisson distribution, it is useful in modeling count data. The negative binomial distribution is more general than the Poisson distribution because it has a variance that is greater than its mean, making it suitable for count data
that do not meet the assumptions of the Poisson distribution. In the limit, as $r$ increases to infinity, the negative binomial distribution approaches the Poisson distribution.

## Parameters

Suppose you are collecting data on the number of auto accidents on a busy highway, and would like to be able to model the number of accidents per day. Because these are count data, and because there are a very large number of cars and a small probability of an accident for any specific car, you might think to use the Poisson distribution. However, the probability of having an accident is likely to vary from day to day as the weather and amount of traffic change, and so the assumptions needed for the Poisson distribution are not met. In particular, the variance of this type of count data sometimes exceeds the mean by a large amount. The data below exhibit this effect: most days have few or no accidents, and a few days have a large number.

```
accident = [llllllllllllllllll
mean(accident)
ans =
    8.0667
var(accident)
ans =
    79.352
```

The negative binomial distribution is more general than the Poisson, and is often suitable for count data when the Poisson is not. The function nbinfit returns the maximum likelihood estimates (MLEs) and confidence intervals for the parameters of the negative binomial distribution. Here are the results from fitting the accident data:

```
[phat,pci] = nbinfit(accident)
phat =
    1.0060 0.1109
pci =
    0.2152 0.0171
    1.7968 0.2046
```

It is difficult to give a physical interpretation in this case to the individual parameters. However, the estimated parameters can be used in a model for the number of daily accidents. For example, a plot of the estimated cumulative probability function shows that while there is an estimated $10 \%$ chance of no accidents on a given day, there is also about a $10 \%$ chance that there will be 20 or more accidents.

```
plot(0:50,nbincdf(0:50,phat(1),phat(2)),'.-');
xlabel('Accidents per Day')
ylabel('Cumulative Probability')
```



## Example

The negative binomial distribution can take on a variety of shapes ranging from very skewed to nearly symmetric. This example plots the probability function for different values of $r$, the desired number of successes: .1, 1, 3, 6 .

```
x = 0:10;
plot(x,nbinpdf(x,.1,.5),'s-', ...
    x,nbinpdf(x,1,.5),'o-', ...
    x,nbinpdf(x,3,.5),'d-', ...
    x,nbinpdf(x,6,.5),'^-');
legend({'r = .1' 'r = 1' 'r = 3' 'r = 6'})
xlabel('x')
ylabel('f(x|r,p)')
```



## See Also

"Discrete Distributions" on page 5-7

## Noncentral Chi-Square Distribution

## In this section...

"Definition" on page B-76
"Background" on page B-76
"Example" on page B-77

## Definition

There are many equivalent formulas for the noncentral chi-square distribution function. One formulation uses a modified Bessel function of the first kind. Another uses the generalized Laguerre polynomials. The cumulative distribution function is computed using a weighted sum of $x^{2}$ probabilities with the weights equal to the probabilities of a Poisson distribution. The Poisson parameter is one-half of the noncentrality parameter of the noncentral chi-square

$$
F(x \mid v, \delta)=\sum_{j=0}^{\infty}\left(\frac{\left(\frac{1}{2} \delta\right)^{j}}{j!} e^{\frac{-\delta}{2}}\right) \operatorname{Pr}\left[\chi_{v+2 j}^{2} \leq x\right]
$$

where $\delta$ is the noncentrality parameter.

## Background

The $x^{2}$ distribution is actually a simple special case of the noncentral chi-square distribution. One way to generate random numbers with a $x^{2}$ distribution (with $v$ degrees of freedom) is to sum the squares of $v$ standard normal random numbers (mean equal to zero.)

What if the normally distributed quantities have a mean other than zero? The sum of squares of these numbers yields the noncentral chi-square distribution. The noncentral chi-square distribution requires two parameters: the degrees of freedom and the noncentrality parameter. The noncentrality parameter is the sum of the squared means of the normally distributed quantities.

The noncentral chi-square has scientific application in thermodynamics and signal processing. The literature in these areas may refer to it as the "Rician Distribution" on page B-93 or generalized "Rayleigh Distribution" on page B-91.

## Example

The following commands generate a plot of the noncentral chi-square pdf.


## Noncentral F Distribution

## In this section...

"Definition" on page B-78
"Background" on page B-78
"Example" on page B-79
"See Also" on page B-79

## Definition

Similar to the noncentral $x^{2}$ distribution, the toolbox calculates noncentral $F$ distribution probabilities as a weighted sum of incomplete beta functions using Poisson probabilities as the weights.

$$
F\left(x \mid v_{1}, v_{2}, \delta\right)=\sum_{j=0}^{\infty}\left(\frac{\left(\frac{1}{2} \delta\right)^{j}}{j!} e^{\frac{-\delta}{2}}\right) z\left(\left.\frac{v_{1} \cdot x}{v_{2}+v_{1} \cdot x} \right\rvert\, \frac{v_{1}}{2}+j, \frac{v_{2}}{2}\right)
$$

$I(x \mid a, b)$ is the incomplete beta function with parameters $a$ and $b$, and $\delta$ is the noncentrality parameter.

## Background

As with the $X^{2}$ distribution, the $F$ distribution is a special case of the noncentral $F$ distribution. The $F$ distribution is the result of taking the ratio of $x^{2}$ random variables each divided by its degrees of freedom.

If the numerator of the ratio is a noncentral chi-square random variable divided by its degrees of freedom, the resulting distribution is the noncentral $F$ distribution.

The main application of the noncentral $F$ distribution is to calculate the power of a hypothesis test relative to a particular alternative.

## Example

The following commands generate a plot of the noncentral $F \mathrm{pdf}$.

$$
\begin{aligned}
& x=(0.01: 0.1: 10.01)^{\prime} ; \\
& p 1=\operatorname{ncfpdf}(x, 5,20,10) \\
& p=\operatorname{fpdf}(x, 5,20) \\
& p l o t\left(x, p, '^{\prime}, x, p 1, \prime^{\prime}\right)
\end{aligned}
$$



## See Also

"Continuous Distributions (Statistics)" on page 5-6

## Noncentral t Distribution

## In this section...

"Definition" on page B-80
"Background" on page B-80
"Example" on page B-81
"See Also" on page B-81

## Definition

The most general representation of the noncentral $t$ distribution is quite complicated. Johnson and Kotz [60] give a formula for the probability that a noncentral $t$ variate falls in the range $[-u, u]$.

$$
P(-u<x<u \mid v, \delta)=\sum_{j=0}^{\infty}\left(\frac{\left(\frac{1}{2} \delta^{2}\right)^{j}}{j!} e^{\frac{-\delta^{2}}{2}}\right) I\left(\left.\frac{u^{2}}{v+u^{2}} \right\rvert\, \frac{1}{2}+j, \frac{v}{2}\right)
$$

$I(x \mid v, \delta)$ is the incomplete beta function with parameters $v$ and $\delta . \delta$ is the noncentrality parameter, and $v$ is the number of degrees of freedom.

## Background

The noncentral $t$ distribution is a generalization of Student's $t$ distribution.
Student's $t$ distribution with $n-1$ degrees of freedom models the $t$-statistic

$$
t=\frac{\bar{x}-\mu}{s / \sqrt{n}}
$$

where $\bar{x}$ is the sample mean and $s$ is the sample standard deviation of a random sample of size $n$ from a normal population with mean $\mu$. If the population mean is actually $\mu_{0}$, then the $t$-statistic has a noncentral $t$ distribution with noncentrality parameter

$$
\delta=\frac{\mu_{0}-\mu}{\sigma / \sqrt{n}}
$$

The noncentrality parameter is the normalized difference between $\mu_{0}$ and $\mu$.
The noncentral $t$ distribution gives the probability that a $t$ test will correctly reject a false null hypothesis of mean $\mu$ when the population mean is actually $\mu_{0}$; that is, it gives the power of the $t$ test. The power increases as the difference $\mu_{0}-\mu$ increases, and also as the sample size $n$ increases.

## Example

The following commands generate a plot of the noncentral $t \mathrm{pdf}$.

```
    x = (-5:0.1:5)';
    p1 = nctcdf(x,10,1);
    p = tcdf(x,10);
    plot(x, p,'-',x,p1, ' -')
```



## See Also

"Continuous Distributions (Statistics)" on page 5-6

## Nonparametric Distributions

See the discussion of ksdensity in "Estimating PDFs without Parameters" on page 5-55.

## Normal Distribution

## In this section...

"Definition" on page B-83
"Background" on page B-83
"Parameters" on page B-84
"Example" on page B-85
"See Also" on page B-85

## Definition

The normal pdf is

$$
y=f(x \mid \mu, \sigma)=\frac{1}{\sigma \sqrt{2 \pi}} e^{\frac{-(x-\mu)^{2}}{2 \sigma^{2}}}
$$

## Background

The normal distribution is a two-parameter family of curves. The first parameter, $\mu$, is the mean. The second, $\sigma$, is the standard deviation. The standard normal distribution (written $\Phi(x)$ ) sets $\mu$ to 0 and $\sigma$ to 1 .
$\Phi(x)$ is functionally related to the error function, erf.

$$
\operatorname{erf}(x)=2 \Phi(x \sqrt{2})-1
$$

The first use of the normal distribution was as a continuous approximation to the binomial.

The usual justification for using the normal distribution for modeling is the Central Limit Theorem, which states (roughly) that the sum of independent samples from any distribution with finite mean and variance converges to the normal distribution as the sample size goes to infinity.

## Parameters

To use statistical parameters such as mean and standard deviation reliably, you need to have a good estimator for them. The maximum likelihood estimates (MLEs) provide one such estimator. However, an MLE might be biased, which means that its expected value of the parameter might not equal the parameter being estimated. For example, an MLE is biased for estimating the variance of a normal distribution. An unbiased estimator that is commonly used to estimate the parameters of the normal distribution is the minimum variance unbiased estimator (MVUE). The MVUE has the minimum variance of all unbiased estimators of a parameter.

The MVUEs of parameters $\mu$ and $\sigma^{2}$ for the normal distribution are the sample mean and variance. The sample mean is also the MLE for $\mu$. The following are two common formulas for the variance.

$$
\begin{align*}
& s^{2}=\frac{1}{n} \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2}  \tag{B-1}\\
& s^{2}=\frac{1}{n-1} \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2} \tag{B-2}
\end{align*}
$$

where

$$
\bar{x}=\sum_{i=1}^{n} \frac{x_{i}}{n}
$$

Equation 1 is the maximum likelihood estimator for $\sigma^{2}$, and equation 2 is the MVUE.

As an example, suppose you want to estimate the mean, $\mu$, and the variance, $\sigma^{2}$, of the heights of all fourth grade children in the United States. The function normfit returns the MVUE for $\mu$, the square root of the MVUE for $\sigma^{2}$, and confidence intervals for $\mu$ and $\sigma^{2}$. Here is a playful example modeling the heights in inches of a randomly chosen fourth grade class.

```
height = normrnd(50,2,30,1); % Simulate heights.
[mu,s,muci,sci] = normfit(height)
```

```
mu =
    50.2025
S =
    1.7946
muci =
    49.5210
    50.8841
sci =
    1.4292
    2.4125
```

Note that $\mathrm{s}^{\wedge} 2$ is the MVUE of the variance.

```
s^2
ans =
    3.2206
```


## Example

The plot shows the bell curve of the standard normal pdf, with $\mu=0$ and $\sigma=1$.


## See Also

"Continuous Distributions (Data)" on page 5-4

## Pareto Distribution

See "Generalized Pareto Distribution" on page B-37.

## Pearson System

See "Pearson and Johnson Systems" on page 6-25.

## Piecewise Distributions

See the discussion of the @piecewisedistribution class in "Fitting Piecewise Distributions" on page 5-72.

## Poisson Distribution

## In this section...

"Definition" on page B-89
"Background" on page B-89
"Parameters" on page B-90
"Example" on page B-90
"See Also" on page B-90

## Definition

The Poisson pdf is

$$
y=f(x \mid \lambda)=\frac{\lambda^{x}}{x!} e^{-\lambda} I_{(0,1, \ldots)}(x)
$$

## Background

The Poisson distribution is appropriate for applications that involve counting the number of times a random event occurs in a given amount of time, distance, area, etc. Sample applications that involve Poisson distributions include the number of Geiger counter clicks per second, the number of people walking into a store in an hour, and the number of flaws per 1000 feet of video tape.

The Poisson distribution is a one-parameter discrete distribution that takes nonnegative integer values. The parameter, $\lambda$, is both the mean and the variance of the distribution. Thus, as the size of the numbers in a particular sample of Poisson random numbers gets larger, so does the variability of the numbers.

The Poisson distribution is the limiting case of a binomial distribution where $N$ approaches infinity and $p$ goes to zero while $N p=\lambda$.

The Poisson and exponential distributions are related. If the number of counts follows the Poisson distribution, then the interval between individual counts follows the exponential distribution.

## Parameters

The MLE and the MVUE of the Poisson parameter, $\lambda$, is the sample mean. The sum of independent Poisson random variables is also Poisson distributed with the parameter equal to the sum of the individual parameters. This is used to calculate confidence intervals $\lambda$. As $\lambda$ gets large the Poisson distribution can be approximated by a normal distribution with $\mu=\lambda$ and $\sigma^{2}$ $=\lambda$. This approximation is used to calculate confidence intervals for values of $\lambda$ greater than 100 .

## Example

The plot shows the probability for each nonnegative integer when $\lambda=5$.


## See Also

"Discrete Distributions" on page 5-7

## Rayleigh Distribution

## In this section...

"Definition" on page B-91
"Background" on page B-91
"Parameters" on page B-92
"Example" on page B-92
"See Also" on page B-92

## Definition

The Rayleigh pdf is

$$
y=f(x \mid b)=\frac{x}{b^{2}} e^{\left(\frac{-x^{2}}{2 b^{2}}\right)}
$$

## Background

The Rayleigh distribution is a special case of the Weibull distribution. If $A$ and $B$ are the parameters of the Weibull distribution, then the Rayleigh distribution with parameter $b$ is equivalent to the Weibull distribution with parameters $A=\sqrt{2} b$ and $B=2$.

If the component velocities of a particle in the $x$ and $y$ directions are two independent normal random variables with zero means and equal variances, then the distance the particle travels per unit time is distributed Rayleigh.

In communications theory, Nakagami distributions, Rician distributions, and Rayleigh distributions are used to model scattered signals that reach a receiver by multiple paths. Depending on the density of the scatter, the signal will display different fading characteristics. Rayleigh and Nakagami distributions are used to model dense scatters, while Rician distributions model fading with a stronger line-of-sight. Nakagami distributions can be reduced to Rayleigh distributions, but give more control over the extent of the fading.

## Parameters

The raylfit function returns the MLE of the Rayleigh parameter. This estimate is

$$
b=\sqrt{\frac{1}{2 n} \sum_{i=1}^{n} x_{i}^{2}}
$$

## Example

The following commands generate a plot of the Rayleigh pdf.

$$
\begin{aligned}
& x=[0: 0.01: 2] ; \\
& p=\operatorname{raylpdf}(x, 0.5) ; \\
& \operatorname{plot}(x, p)
\end{aligned}
$$



## See Also

"Continuous Distributions (Data)" on page 5-4

## Rician Distribution

## In this section...

"Definition" on page B-93
"Background" on page B-93
"Parameters" on page B-93
"See Also" on page B-94

## Definition

The Rician distribution has the density function

$$
I_{0}\left(\frac{x s}{\sigma^{2}}\right) \frac{x}{\sigma^{2}} e^{-\left(\frac{x^{2}+s^{2}}{2 \sigma^{2}}\right)}
$$

with noncentrality parameter $s \geq 0$ and scale parameter $\sigma>0$, for $x>0 . I_{0}$ is the zero-order modified Bessel function of the first kind. If $x$ has a Rician distribution with parameters $s$ and $\sigma$, then $(x / \sigma)^{2}$ has a noncentral chi-square distribution with two degrees of freedom and noncentrality parameter $(s / \sigma)^{2}$.

## Background

In communications theory, Nakagami distributions, Rician distributions, and Rayleigh distributions are used to model scattered signals that reach a receiver by multiple paths. Depending on the density of the scatter, the signal will display different fading characteristics. Rayleigh and Nakagami distributions are used to model dense scatters, while Rician distributions model fading with a stronger line-of-sight. Nakagami distributions can be reduced to Rayleigh distributions, but give more control over the extent of the fading.

## Parameters

See mle, dfittool.

## See Also

"Continuous Distributions (Data)" on page 5-4

## Student's † Distribution

## In this section...

"Definition" on page B-95
"Background" on page B-95
"Example" on page B-96
"See Also" on page B-96

## Definition

Student's $t \mathrm{pdf}$ is

$$
y=f(x \mid v)=\frac{\Gamma\left(\frac{v+1}{2}\right)}{\Gamma\left(\frac{v}{2}\right)} \frac{1}{\sqrt{v \pi}} \frac{1}{\left(1+\frac{x^{2}}{v}\right)^{\frac{v+1}{2}}}
$$

where $\Gamma(\cdot)$ is the Gamma function.

## Background

The $t$ distribution is a family of curves depending on a single parameter $v$ (the degrees of freedom). As $v$ goes to infinity, the $t$ distribution approaches the standard normal distribution.
W. S. Gossett discovered the distribution through his work at the Guinness brewery. At the time, Guinness did not allow its staff to publish, so Gossett used the pseudonym "Student."

If $x$ is a random sample of size $n$ from a normal distribution with mean $\mu$, then the statistic

$$
t=\frac{\bar{x}-\mu}{s / \sqrt{n}}
$$

where $\bar{x}$ is the sample mean and $s$ is the sample standard deviation, has Student's $t$ distribution with $n-1$ degrees of freedom.

## Example

The plot compares the $t$ distribution with $v=5$ (solid line) to the shorter tailed, standard normal distribution (dashed line).

```
    x = -5:0.1:5;
    y = tpdf(x,5);
    z = normpdf(x,0,1);
    plot(x,y,'-',x,z,'-.')
```



## See Also

"Continuous Distributions (Statistics)" on page 5-6

## t Location-Scale Distribution

## In this section...

"Definition" on page B-97
"Background" on page B-97
"Parameters" on page B-97
"See Also" on page B-98

## Definition

The $t$ location-scale distribution has the density function

$$
\frac{\Gamma\left(\frac{v+1}{2}\right)}{\sigma \sqrt{v \pi} \Gamma\left(\frac{v}{2}\right)}\left[\frac{v+\left(\frac{x-\mu}{\sigma}\right)^{2}}{v}\right]^{-\left(\frac{v+1}{2}\right)}
$$

with location parameter $\mu$, scale parameter $\sigma>0$, and shape parameter $v>0$. If $x$ has a $t$ location-scale distribution, with parameters $\mu, \sigma$, and $v$, then

$$
\frac{x-\mu}{\sigma}
$$

has a Student's $t$ distribution with $v$ degrees of freedom.

## Background

The $t$ location-scale distribution is useful for modeling data distributions with heavier tails (more prone to outliers) than the normal distribution. It approaches the normal distribution as $v$ approaches infinity, and smaller values of $v$ yield heavier tails.

## Parameters

See mle, dfittool.

## See Also

"Continuous Distributions (Statistics)" on page 5-6

## Uniform Distribution (Continuous)

## In this section...

"Definition" on page B-99
"Background" on page B-99
"Parameters" on page B-99
"Example" on page B-99
"See Also" on page B-100

## Definition

The uniform cdf is

$$
p=F(x \mid a, b)=\frac{x-a}{b-a} I_{[a, b]}(x)
$$

## Background

The uniform distribution (also called rectangular) has a constant pdf between its two parameters $a$ (the minimum) and $b$ (the maximum). The standard uniform distribution ( $a=0$ and $b=1$ ) is a special case of the beta distribution, obtained by setting both of its parameters to 1 .

The uniform distribution is appropriate for representing the distribution of round-off errors in values tabulated to a particular number of decimal places.

## Parameters

The sample minimum and maximum are the MLEs of $a$ and $b$ respectively.

## Example

The example illustrates the inversion method for generating normal random numbers using rand and norminv. Note that the MATLAB function, randn, does not use inversion since it is not efficient for this case.

$$
u=\operatorname{rand}(1000,1)
$$



## See Also

"Continuous Distributions (Data)" on page 5-4

## Uniform Distribution (Discrete)

## In this section...

"Definition" on page B-101
"Background" on page B-101
"Example" on page B-101
"See Also" on page B-102

## Definition

The discrete uniform pdf is

$$
y=f(x \mid N)=\frac{1}{N} I_{(1, \ldots, N)}(x)
$$

## Background

The discrete uniform distribution is a simple distribution that puts equal weight on the integers from one to $N$.

## Example

As for all discrete distributions, the cdf is a step function. The plot shows the discrete uniform cdf for $N=10$.

```
x = 0:10;
y = unidcdf(x,10);
stairs(x,y)
set(gca,'Xlim',[0 11])
```



Pick a random sample of 10 from a list of 553 items:

```
numbers = unidrnd(553,1,10)
numbers =
293
```


## See Also

"Discrete Distributions" on page 5-7

## Weibull Distribution

## In this section...

"Definition" on page B-103
"Background" on page B-103
"Parameters" on page B-104
"Example" on page B-104
"See Also" on page B-105

## Definition

The Weibull pdf is

$$
y=f(x \mid a, b)=b a^{-b} x^{b-1} e^{-\left(\frac{x}{a}\right)^{b}} I_{(0, \infty)}(x)
$$

## Background

Waloddi Weibull offered the distribution that bears his name as an appropriate analytical tool for modeling the breaking strength of materials. Current usage also includes reliability and lifetime modeling. The Weibull distribution is more flexible than the exponential for these purposes.

To see why, consider the hazard rate function (instantaneous failure rate). If $f(t)$ and $F(t)$ are the pdf and cdf of a distribution, then the hazard rate is

$$
h(t)=\frac{f(t)}{1-F(t)}
$$

Substituting the pdf and cdf of the exponential distribution for $f(t)$ and $F(t)$ above yields a constant. The example below shows that the hazard rate for the Weibull distribution can vary.

## Parameters

Suppose you want to model the tensile strength of a thin filament using the Weibull distribution. The function wblfit gives maximum likelihood estimates and confidence intervals for the Weibull parameters.

```
strength = wblrnd(0.5,2,100,1); % Simulated strengths.
[p,ci] = wblfit(strength)
p =
0.4715 1.9811
ci =
    0.4248 1.7067
    0.5233 2.2996
```

The default $95 \%$ confidence interval for each parameter contains the true value.

## Example

The exponential distribution has a constant hazard function, which is not generally the case for the Weibull distribution.

The plot shows the hazard functions for exponential (dashed line) and Weibull (solid line) distributions having the same mean life. The Weibull hazard rate here increases with age (a reasonable assumption).

```
t = 0:0.1:4.5;
h1 = exppdf(t,0.6267) ./ (1-expcdf(t,0.6267));
h2 = wblpdf(t,2,2) ./ (1-wblcdf(t,2,2));
plot(t,h1,'--',t,h2,'-')
```



## See Also

"Continuous Distributions (Data)" on page 5-4

## Wishart Distribution

## In this section...

"Definition" on page B-106
"Background" on page B-106
"Example" on page B-107
"See Also" on page B-107

## Definition

The probability density function of the $d$-dimensional Wishart distribution is given by

$$
\mathrm{y}=\mathrm{f}(\mathrm{X}, \Sigma, v)=\frac{|\mathrm{X}|^{((\mathrm{v}-\mathrm{d}-1) / 2)} e^{\left(-\frac{1}{2} \operatorname{trace}\left(\Sigma^{-1} \mathrm{X}\right)\right)}}{2^{(\mathrm{vd}) / 2} \pi^{(\mathrm{d}(\mathrm{~d}-1)) / 4}|\Sigma|^{\mathrm{v} / 2} \Gamma(\mathrm{v} / 2) \ldots \Gamma(\mathrm{v}-(\mathrm{d}-1)) / 2}
$$

where $X$ and $\Sigma$ are $d$-by- $d$ symmetric positive definite matrices, and $v$ is a scalar greater than $d-1$. While it is possible to define the Wishart for singular $\Sigma$, the density cannot be written as above.

Only random matrix generation is supported for the Wishart distribution, including both singular and nonsingular $\Sigma$.

## Background

The Wishart distribution is a generalization of the univariate chi-square distribution to two or more variables. It is a distribution for symmetric positive semidefinite matrices, typically covariance matrices, the diagonal elements of which are each chi-square random variables. In the same way as the chi-square distribution can be constructed by summing the squares of independent, identically distributed, zero-mean univariate normal random variables, the Wishart distribution can be constructed by summing the inner products of independent, identically distributed, zero-mean multivariate normal random vectors.

The Wishart distribution is parameterized with a symmetric, positive semidefinite matrix, $\Sigma$, and a positive scalar degrees of freedom parameter, $v$. $v$ is analogous to the degrees of freedom parameter of a univariate chi-square distribution, and $\Sigma \mathrm{v}$ is the mean of the distribution.

The Wishart distribution is often used as a model for the distribution of the sample covariance matrix for multivariate normal random data, after scaling by the sample size.

## Example

If $x$ is a bivariate normal random vector with mean zero and covariance matrix

$$
\Sigma=\left(\begin{array}{ll}
1 & .5 \\
.5 & 2
\end{array}\right)
$$

then you can use the Wishart distribution to generate a sample covariance matrix without explicitly generating $x$ itself. Notice how the sampling variability is quite large when the degrees of freedom is small.

```
Sigma = [1 .5; .5 2];
df = 10; S1 = wishrnd(Sigma,df)/df
S1 =
    1.7959 0.64107
    0.64107 1.5496
df = 1000; S2 = wishrnd(Sigma,df)/df
S2 =
        0.9842 0.50158
        0.50158 2.1682
```


## See Also

"Inverse Wishart Distribution" on page B-46, "Multivariate Distributions" on page 5-8

## $c$

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[^0]:    See Also
    delete | dynamicprops | event.listener | events | meta.property | notify | qrandstream | reset

[^1]:    See Also
    cdf | binopdf | binoinv | binostat | binofit | binornd

[^2]:    See Also anova1 | axes_props | kruskalwallis | multcompare
    How To . "Grouped Data" on page 2-34

[^3]:    CutPoint

[^4]:    CutCategories

[^5]:    See Also

[^6]:    learners

[^7]:    How To

    - "Grouped Data" on page 2-34

[^8]:    References [1] Dobson, A. J. An Introduction to Generalized Linear Models. New York: Chapman \& Hall, 1990.
    [2] McCullagh, P., and J. A. Nelder. Generalized Linear Models. New York: Chapman \& Hall, 1990.
    [3] Collett, D. Modeling Binary Data. New York: Chapman \& Hall, 2002.

    See Also glmfit

[^9]:    random | hygepdf | hygecdf | hygeinv | hygestat

[^10]:    Examples Display interaction plots for data with four 3-level factors named ' $A$ ', 'B','C', and 'D':

[^11]:    ib = isbranch(t)
    ib $=$
    1
    0
    1
    1
    0
    1
    0
    0

[^12]:    load carsmall

[^13]:    References [1] Evans, M., N. Hastings, and B. Peacock. Statistical Distributions. 2nd ed., Hoboken, NJ: John Wiley \& Sons, Inc., 1993, pp. 147-148.
    [2] Johnson, N., and S. Kotz. Distributions in Statistics: Continuous Univariate Distributions-2. Hoboken, NJ: John Wiley \& Sons, Inc., 1970, pp. 201-219.

[^14]:    References

[^15]:    References

    See Also regress | stepwise
    [1] Hoerl, A. E., and R. W. Kennard. "Ridge Regression: Biased Estimation for Nonorthogonal Problems." Technometrics. Vol. 12, No. 1, 1970, pp. 55-67.
    [2] Hoerl, A. E., and R. W. Kennard. "Ridge Regression: Applications to Nonorthogonal Problems." Technometrics. Vol. 12, No. 1, 1970, pp. 69-82.
    [3] Marquardt, D.W. "Generalized Inverses, Ridge Regression, Biased Linear Estimation, and Nonlinear Estimation." Technometrics. Vol. 12, No. 3, 1970, pp. 591-612.
    [4] Marquardt, D. W., and R.D. Snee. "Ridge Regression in Practice." The American Statistician. Vol. 29, No. 1, 1975, pp. 3-20.

[^16]:    See Also
    categorical | subsasgn

[^17]:    See Also
    classregtree | numnodes | cuttype | surrcutvar

